

Orthogonal Polynomials on the Negative Multinomial Distribution

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Orthogonal polynomials on the multivariate negative binomial distribution,

$$(1 + \theta)^{-\alpha-x} \left\{ \prod_{j=0}^p \theta_j^{x_j} / x_j! \right\} \Gamma(\alpha + x) / \Gamma(\alpha),$$

where $\alpha > 0$, $\theta_j > 0$, $x = \Sigma x_j$, $\theta = \Sigma \theta_j$, $x_0, x_1, \dots, x_p = 0, 1, \dots$ are constructed and their properties studied.

INTRODUCTION

One generalization of the negative binomial to higher dimensions is given by the distribution

$$(1 + \theta)^{-\alpha-x} \left(\prod_{j=0}^p \theta_j^{x_j} / x_j! \right) \Gamma(\alpha + x) / \Gamma(\alpha), \quad (1)$$

where $\alpha > 0$, $\theta_j > 0$, $x = \sum_0^p x_j$, $\theta = \sum_0^p \theta_j$, $x_0, x_1, \dots, x_p = 0, 1, \dots$.

The aim of this paper is to generalize the Meixner polynomials on the negative binomial distribution to a set of orthogonal polynomials on the distribution (1). If \mathbf{k} is any vector the notation $k = \sum_i k_i$ will be used, and at times the parameters of a function will be suppressed to avoid cumbersome notation.

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GENERATING FUNCTION

Define a set of functions $\{m_{\mathbf{k}}(\mathbf{X})\}$ by

$$G(\mathbf{X}, \mathbf{w}) = \sum_{\mathbf{k} \geq 0} m_{\mathbf{k}}(\mathbf{X}) w_0^{k_0} \cdots w_p^{k_p} / (k_0! \cdots k_p!), \tag{2}$$

$$= (1 - w_0)^{-X-\alpha} \prod_{i=0}^p \left(1 - w_0(1 + \theta)/\theta + \sum_{j=1}^p v_i^{(j)} w_j \right)^{X_i},$$

where $\{v^{(j)}; j = 0, 1, \dots, p\}$ is a complete set of orthogonal functions on $\{\theta_i\}$ with $v^{(0)} \equiv 1$ and the summation is taken over $k_j = 0, 1, \dots, j = 0, 1, \dots, p$. The domain D of \mathbf{w} is given by $\sum_0^p a_i^2 w_i^2 < 1$, where $a_0^2 = (1 + \theta)/\theta$,

$$a_j^2 = \sum_{i=0}^p v_i^{(j)2} \theta_i.$$

Let $b_i(\mathbf{w}) = 1 - w_0(1 + \theta)/\theta + \sum_{j=1}^p v_i^{(j)} w_j$. Then

$$EG(\mathbf{X}, \mathbf{w}) G(\mathbf{X}, \mathbf{v}) = EE \left\{ [(1 - w_0)(1 - v_0)]^{-\alpha-X} \prod_{i=0}^p [b_i(\mathbf{w}) b_i(\mathbf{v})]^{X_i} \mid X \right\}$$

$$= E[(1 - w_0)(1 - v_0)]^{-\alpha-X} \left[\sum_{i=0}^p b_i(\mathbf{w}) b_i(\mathbf{v}) \theta_i / \theta \right]^X$$

$$= \left[(1 + \theta)(1 - w_0)(1 - v_0) - \sum_{i=0}^p \theta_i b_i(\mathbf{w}) b_i(\mathbf{v}) \right]^{-\alpha}$$

$$= \left(1 - \sum_{j=0}^p a_j^2 w_j v_j \right)^{-\alpha}.$$

Thus $Em_{\mathbf{h}}(\mathbf{X})m_{\mathbf{k}}(\mathbf{X}) = \delta_{\mathbf{h}\mathbf{k}} a_0^{2k_0} \cdots a_p^{2k_p} k_0! \cdots k_p! \Gamma(\alpha + k) / \Gamma(\alpha)$,

$$\delta_{\mathbf{h}\mathbf{k}} = \begin{cases} 0 & \text{if } \mathbf{h} \neq \mathbf{k}, \\ 1 & \text{if } \mathbf{h} = \mathbf{k}, \end{cases}$$

and $M_{\mathbf{k}}(\mathbf{X}) = \{\Gamma(\alpha) / (\Gamma(\alpha + k) k_0! \cdots k_p!)\}^{1/2} a_0^{-k_0} \cdots a_p^{-k_p} m_{\mathbf{k}}(\mathbf{X})$ are the orthonormal functions.

TRANSFORM

Denote $Ef(\mathbf{X})s_0^{X_0} \cdots s_p^{X_p}$ by $f^*(\mathbf{s})$ for any function f for which $E|f(\mathbf{X})| < \infty$. It is straightforward to show

$$G^*(\mathbf{s}, \mathbf{w}) = \left\{ (1 + \theta)(1 - w_0) - \sum_{i=0}^p \theta_i s_i \left(1 - w_0(1 + \theta)/\theta + \sum_{j=1}^p v_i^{(j)} w_j \right) \right\}^{-\alpha},$$

$$= \left\{ (1 + \theta - S_0) - (1 + \theta)(\theta - S_0) w_0/\theta - \sum_{j=1}^p S_j w_j \right\}^{-\alpha},$$

where $S_j = \sum_{i=0}^p s_i \theta_i v_i^{(j)}$, $j = 0, 1, \dots, p$. Thus

$$m_{\mathbf{k}}^*(\mathbf{s}) = a_0^{2k_0} (1 + \theta - S_0)^{-\alpha - k} (\theta - S_0)^{k_0} \left(\prod_{j=1}^p S_j^{k_j} \right) \Gamma(\alpha + k) / \Gamma(\alpha).$$

COMPLETENESS

Suppose $Ef(\mathbf{X})M_{\mathbf{k}}(\mathbf{X}) = 0$ for every index vector \mathbf{k} , then this is equivalent to

$$Ef(\mathbf{X}) G(\mathbf{X}, \mathbf{w}) = 0 \quad \text{for all } \mathbf{w} \in D.$$

This means that

$$\sum_{\mathbf{x} \geq 0} f(\mathbf{x}) \Gamma(\alpha + x) \prod_{i=0}^p (\phi_i^{x_i} / x_i!) = 0,$$

where $\phi_i = \theta_i \{ 1 - w_0(1 + \theta)/\theta + \sum_{j=1}^p v_i^{(j)} w_j \} / (1 - w_0)(1 + \theta)$, for $i = 0, 1, \dots, p$, which is a one-to-one and onto transformation of \mathbf{w} on account of the complete orthogonal functions $\{v^{(j)}\}$ on $\{\theta_i\}$. Hence $Ef(\mathbf{X}) M_{\mathbf{k}}(\mathbf{X}) = 0$ for every $\mathbf{k} \geq 0$ is equivalent to $f(\mathbf{x}) \equiv 0$ for all \mathbf{x} .

Let \mathbf{e}_h be the vector (δ_{jh}) , $h = 0, 1, \dots, p$. It will be shown that $\{m_{k\mathbf{e}_0+r\mathbf{e}_j}(\mathbf{X}); k = 0, 1, \dots, r = 0, 1, \dots\}$ is a complete orthogonal polynomial set in X_0 and X alone if and only if $v_1^{(j)} = v_2^{(j)} = \cdots = v_p^{(j)}$ and $v_0^{(j)} \neq v_1^{(j)}$.

To prove this, write $X_0 = Y_0$, $X = Y$, $X_i = Y_i$ for $i = 2, 3, \dots, p$ in (2). Then, for any w_0 and w_j ,

$$G(\mathbf{Y}, w_0, w_j) = \sum_{k,r} m_{k\mathbf{e}_0+r\mathbf{e}_j}(\mathbf{X}) w_0^k w_j^r / k! r!$$

$$= (1 - w_0)^{-Y-\alpha} (1 - w_0(1 + \theta)/\theta + v_0^{(j)} w_j)^{Y_0}$$

$$\cdot (1 - w_0(1 + \theta)/\theta + v_1^{(j)} w_j)^{Y - Y_0}$$

$$\cdot \prod_{i=2}^p \left[1 + \frac{(v_i^{(j)} - v_1^{(j)}) w_j}{1 - w_0(1 + \theta)/\theta + v_1^{(j)} w_j} \right]^{Y_i}.$$

If $v_1^{(j)} = \dots = v_p^{(j)}$ and $v_1^{(j)} \neq v_0^{(j)}$, then the result is obvious and conversely if $G(\mathbf{Y}, w_0, w_j)$ is a function of Y and Y_0 for any Y_2, \dots, Y_p, w_0 and w_j ; then this implies

$$\prod_{i=2}^p [1 + (v_i^{(j)} - v_1^{(j)}) w_j / \{1 + w_0(1 + \theta)/\theta + v_1^{(j)} w_j\}]^{Y_i} = 1,$$

which gives $v_i^{(j)} = v_1^{(j)}$ for $i = 2, 3, \dots, p$ and $v_0^{(j)} \neq v_1^{(j)}$, because otherwise, it will be a function of Y only.

POLYNOMIAL STRUCTURE

Define the degree of a multidimensional polynomial

$$\sum_{\mathbf{a}} q(\mathbf{a}) z_0^{a_0} \dots z_p^{a_p}$$

to be the degree of the polynomial $\sum_{\mathbf{a}} q(\mathbf{a}) u^{\mathbf{a}}$.

If $\chi_j = \sum_{i=0}^p v_i^{(j)} X_i, j = 0, 1, \dots, p$, then any polynomial of degree δ in \mathbf{X} is a polynomial of degree δ in χ , and vice versa.

$$\begin{aligned} E_{\chi_0^{\delta_0} \dots \chi_p^{\delta_p}} G(\mathbf{X}, \mathbf{w}) &= E \sum_{\beta \leq \delta} a_{\beta} \mathbf{X}^{(\beta)} G(\mathbf{X}, \mathbf{w}) \\ &= \sum_{\beta \leq \delta} a_{\beta} (\alpha + \beta - 1)^{(\beta)} \prod_{i=0}^p \left[\theta_i \left(1 - w_0(1 + \theta)/\theta + \sum_{j=1}^p v_i^{(j)} w_j \right) \right]^{\beta_i} \\ &= (\alpha + \delta - 1)^{(\delta)} \prod_{l=0}^p \left\{ \sum_{i=0}^p v_i^{(l)} \theta_i \left(1 - w_0(1 + \theta)/\theta + \sum_{j=1}^p v_i^{(j)} w_j \right) \right\}^{\delta_l} + H(\mathbf{w}) \\ &= (\alpha + \delta - 1)^{(\delta)} \theta^{\delta_0} (1 - w_0(1 + \theta)/\theta)^{\delta_0} \prod_{l=1}^p a_l^{2^{\delta_l}} w_l^{\delta_l} + H(\mathbf{w}), \end{aligned}$$

where $\{\alpha_{\beta}\}$ is a set of constants and $H(\mathbf{w})$ is a polynomial in \mathbf{w} of degree at most $\delta - 1$. Thus $E_{\chi_0^{\delta_0} \dots \chi_p^{\delta_p}} M_{\mathbf{k}}(\mathbf{X}) = 0$ if either $k > \delta$, or $\mathbf{k} \neq \delta$ with $k = \delta$, which implies that

$$M_{\mathbf{k}}(\mathbf{X}) = c_{\mathbf{k}} \chi_0^{k_0} \dots \chi_p^{k_p} + Q(\mathbf{X}),$$

where $c_{\mathbf{k}}$ is a nonzero constant and $Q(\mathbf{X})$ is a polynomial of degree at most $k - 1$. The above representation means that $\{M_{\mathbf{k}}(\mathbf{X})\}$ can be constructed by applying the

Gram-Schmidt orthogonalization process to the sequence

$$1, \chi_0^{a_0} \chi_1^{a_1} \cdots \chi_p^{a_p}, \chi_0^{b_0} \chi_1^{b_1} \cdots \chi_p^{b_p}, \dots,$$

where each of the products is distinct,

$$\sum a_i \leq \sum b_i \leq \dots \text{ and the ordering when } \sum a_i = \sum b_i \text{ is arbitrary.}$$

Clearly, the $M_k(\mathbf{X})$ are polynomials of maximum degree k and if $P_h(\mathbf{X})$ is any polynomial of degree $h < k$, then $EM_k(\mathbf{X}) P_h(\mathbf{X}) = 0$ and so any other set of orthogonal polynomials $\{N_h(\mathbf{X})\}$ can be written as

$$N_h(\mathbf{X}) = \sum_{k=h} (EN_h(\mathbf{X}) M_k(\mathbf{X})) M_k(\mathbf{X}).$$

RUNGE-TYPE IDENTITY

Suppose \mathbf{X}_1 and \mathbf{X}_2 are independent random vectors having the distribution (1) with parameters α_1 and α_2 , then

$$G(\mathbf{X}_1 + \mathbf{X}_2, \mathbf{w}; \alpha_1 + \alpha_2) = G(\mathbf{X}_1, \mathbf{w}; \alpha_1) G(\mathbf{X}_2, \mathbf{w}; \alpha_2).$$

Thus

$$m_k(\mathbf{X}_1 + \mathbf{X}_2; \alpha_1 + \alpha_2) = \sum_{h_1+h_2=k} \binom{k_0}{h_0} \cdots \binom{k_p}{h_p} m_{h_1}(\mathbf{X}_1; \alpha_1) m_{h_2}(\mathbf{X}_2; \alpha_2).$$

A Runge-type identity holds for a class of polynomials and the reader is referred to Eagleson [1] for background material.

BILINEAR SUMS

Define

$$g_h(\mathbf{x}, \mathbf{y}) = \sum_{k=h} M_k(\mathbf{x}) M_k(\mathbf{y}).$$

The transform of $g_h(\mathbf{x}, \mathbf{y})$ is

$$\begin{aligned} g_h^*(\mathbf{s}, \mathbf{t}) &= [\Gamma(\alpha + h)/\Gamma(\alpha) h!](1 + \theta - S_0)^{-\alpha-h} (1 + \theta - T_0)^{-\alpha-h} \\ &\quad \cdot \left\{ (\theta - S_0)(\theta - T_0)(1 + \theta)/\theta + \sum_{i=1}^p S_i T_i / a_i^2 \right\}^h, \\ &= [\Gamma(\alpha + h)/\Gamma(\alpha) h!](1 + \theta - S_0)^{-\alpha} (1 + \theta - T_0)^{-\alpha} \\ &\quad \cdot \left\{ 1 - \left[(1 + \theta) - \sum_{i=0}^p \theta_i s_i t_i \right] / (1 + \theta - S_0)(1 + \theta - T_0) \right\}^h, \end{aligned}$$

where

$$S_j = \sum_{i=0}^p \theta_i s_i v_i^{(j)}, \quad T_j = \sum_{i=0}^p \theta_i t_i v_i^{(j)}, \quad j = 0, 1, 2, \dots, p.$$

Thus $g_h(\mathbf{x}, \mathbf{y})$ is invariant under any choice of $\{v^{(j)}\}$.

If $\{N_k(\mathbf{X})\}$ is any other orthonormal polynomial set

$$g_h(\mathbf{X}, \mathbf{Y}) = \sum_{k=h} N_k(\mathbf{X}) N_k(\mathbf{Y}),$$

and if $R(\mathbf{x})$ is a polynomial of degree n , then

$$R(\mathbf{x}) = \sum_{h \leq n} E g_h(\mathbf{x}, \mathbf{X}) R(\mathbf{X}).$$

POISSON LIMIT

Suppose $\theta_i \rightarrow 0, \alpha \rightarrow \infty$ while $\alpha\theta_i \rightarrow \mu_i, 0 < \mu_i < \infty, i = 0, 1, \dots, p$. It is easily shown that the distribution of \mathbf{X} converges to that of $p + 1$ independent Poisson random variables with means $\mu_i, i = 0, 1, \dots, p$. Further suppose $\alpha^{-1}v_i^{(j)} \rightarrow u_i^{(j)}, i = 0, 1, \dots, p, j = 1, 2, \dots, p$ and $\{u^{(j)}\}$ is complete on $\{\mu_i\}$; then

$$G(\mathbf{X}, \alpha^{-1}\mathbf{w}) \rightarrow e^{w_0} \prod_{i=0}^p \left(1 - w_0/\mu_i + \sum_{j=1}^p u_i^{(j)} w_j \right)^{X_i},$$

which generates a complete set of orthogonal functions on the limit distribution.

NORMAL LIMIT

If $\alpha \rightarrow \infty$, while θ remains fixed, the distribution of $\mathbf{Y} = (\mathbf{X} - \alpha\theta)/\alpha^{1/2}$ converges to that of $p + 1$ normal random variables with variances

$$\theta_0(1 + \theta_0), \dots, \theta_p(1 + \theta_p)$$

and covariances $\theta_0\theta_1, \dots, \theta_{p-1}\theta_p$.

$$G(\alpha^{1/2}\mathbf{Y} + \alpha\theta, \alpha^{-1/2}\mathbf{w}) \rightarrow \exp \left(\sum_{j=0}^p \eta_j w_j - \frac{1}{2} \sum_{j=0}^p a_j^2 w_j^2 \right),$$

where $\eta_0 = -Y/\theta, \eta_j = -\sum_{i=0}^p Y_i v_i^{(j)}, j = 1, 2, \dots, p$.

The limiting generating function generates the product set of Hermite-Chebysheff polynomials on the independent normal variables $\eta_0, \eta_1, \dots, \eta_p$.

GAMMA LIMIT

Suppose $\theta_i \rightarrow \infty$, $\theta_i/\theta \rightarrow \nu_i$, $i = 0, 1, \dots, p$ and $\theta \nu^{(j)} \rightarrow 0$, $j = 1, 2, \dots, p$ while α remains fixed.

The distribution of \mathbf{X}/θ converges to that of a random vector \mathbf{Y} ; where $Y_i = \nu_i Z$ and Z has a gamma (α) distribution.

$$G(\theta \mathbf{Y}, \mathbf{w}) \rightarrow (1 - w_0)^{-\alpha} \exp\{-\nu Z w_0 / (1 - w_0)\},$$

which generates the set of Laguerre polynomials on Z .

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REFERENCES

- [1] EAGLESON, G. K. (1964). Polynomial expansions of bivariate distributions. *Ann. Math. Statist.* 35 1208-1215.
- [2] GRIFFITHS, R. C. (1971). Orthogonal polynomials on the multinomial distribution. *Austral. J. Statist.* 13 27-34; *Corregenda* 14 270.