



# Limit theorem for the statistical solution of Burgers equation

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## Abstract

In this work we study limit theorems for the Hopf–Cole solution of the Burgers equation when the initial value is a functional of some Gaussian processes. We use the Gaussian chaos decomposition, and we get “Gaussian scenario” with new normalization factors. © 1999 Elsevier Science B.V. All rights reserved.

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## 1. Introduction

The one-dimensional Burgers equation without force has the form

$$\partial_t u + u \partial_x u = \frac{\nu}{2} \partial_{xx}^2 u, \quad x \in \mathbb{R}, t > 0. \tag{1}$$

Here  $\nu > 0$  is the viscosity and the initial value  $u(x, 0) = -d\xi(x)/dx$  is given, and satisfies  $\xi(x) = o(x^2)$  as  $|x| \rightarrow +\infty$ . If we introduce a potential function  $\psi$  defined as  $u = -\partial_x \psi$ , then the Hopf–Cole substitution  $\psi = \nu \ln \theta$  shows that  $\theta$  satisfies the heat equation  $\partial_t \theta = (\nu/2) \partial_{xx}^2 \theta$ . Using this fact one can write down for the solution  $u = u(x, t, \nu)$  the explicit expression

$$u(x, t, \nu) = \frac{I(x, t, \nu)}{J(x, t, \nu)}, \tag{2}$$

where

$$I(x, t, \nu) = \int_{-\infty}^{\infty} \frac{x-y}{t} g\left(\frac{x-y}{\sqrt{t\nu}}\right) \exp\left\{\frac{\xi(y)}{\nu}\right\} dy, \tag{3}$$

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and

$$J(x, t, v) = \int_{-\infty}^{\infty} g\left(\frac{x-y}{\sqrt{tv}}\right) \exp\left\{\frac{\zeta(y)}{v}\right\} dy. \tag{4}$$

The function  $g$  is the density of the standard Gaussian random variable  $\mathcal{N}(0, 1)$ .

Many authors have investigated the solution  $u(x, t, v)$  on different types of initial conditions which are stationary Gaussian processes. It is well-known (Albeverio et al., 1994; Bulinskii and Molchanov, 1991) that, if  $v=1$  and  $\zeta(x)$  is a stationary zero-mean Gaussian process, with covariance function  $B(x) = \mathbb{E}[\zeta(x)\zeta(0)]$  is assumed to satisfy

$$\int_{\mathbb{R}} |B(x)| dx < \infty, \tag{5}$$

then the solution  $u(x, t, 1)$  obeys the ‘‘Gaussian scenario’’, that is, as  $L \rightarrow \infty$ ,

$$L^{3/2}u(Lx, L^2t, 1) \Rightarrow X(x, t), \tag{6}$$

where  $X$  is a Gaussian process, and  $\Rightarrow$  stands for the weak convergence of the finite-dimensional distributions. In the case where

$$\begin{aligned} B(x) &= 1 - \frac{\lambda_2}{2}x^2 + \frac{\lambda_4}{4!}x^4 + o(x^4) \quad \text{as } x \rightarrow 0 \quad \text{and} \quad B(x) = o(1/\ln x) \\ &\text{as } x \rightarrow \infty, \end{aligned} \tag{7}$$

Molchanov et al. (1995) have proved that, as  $L \rightarrow +\infty$ ,

$$L\sqrt{2 \ln L}u(Lx, L^2\sqrt{2 \ln L}t, 1/L^2\sqrt{2 \ln L}) \Rightarrow \frac{x - y_{j(x,t)}}{t},$$

where  $y_{j(x,t)}$  is defined via a Poisson process  $(y_j, u_j)_{j \in \mathbb{Z}}$  on  $\mathbb{R}^2$  with the intensity  $e^{-u} dy du$ , by maximizing the difference

$$\max_j \left[ u_j - \frac{(x - y_j)^2}{2t} \right] = u_{j(x,t)} - \frac{(x - y_{j(x,t)})^2}{2t}. \tag{8}$$

On the other hand, Surgailis (1996) has proved that if  $B(0) := \lambda_0 > 1$ ,

$$B(x) = \lambda_0 - \frac{\lambda_2}{2}x^2 + o(x^2) \quad \text{as } x \rightarrow 0 \quad \text{and} \quad B(x) = o(1/\ln x) \quad \text{as } x \rightarrow \infty, \tag{9}$$

then

$$L\sqrt{2 \ln L}u(Lx, L^2\sqrt{2 \ln L}t, 1/\sqrt{2 \ln L}) \Rightarrow v(x, t),$$

where

$$\begin{aligned} v(x, t) &= \sum_{i \in \mathbb{Z}} t^{-1}(x - y_i) \exp \left\{ \sqrt{\lambda_0} u_i - \frac{(x - y_i)^2}{2t} \right\} \\ &\times \left\{ \sum_{i \in \mathbb{Z}} \exp \left\{ \sqrt{\lambda_0} u_i - \frac{(x - y_i)^2}{2t} \right\} \right\}^{-1}. \end{aligned}$$

See also Hu and Woyczynski (1995) for other cases. In our work we unify these results and we broaden this list. This is the aim of Section 2 Theorem 2.1. In Section 3 we study the case when  $\zeta(x) = \int_{\mathbb{R}} f(x, y) dW(y) - \frac{1}{2} \int_{\mathbb{R}} f^2(x, y) dy$ , where  $W$  is the Gaussian white noise, and we obtain the ‘‘Gaussian scenario’’.

## 2. General scaling limit behavior

We set, for  $L > 0$ , for positive function  $T(L)$ , and for functions  $A(L)$ ,  $b(L)$ ,

$$V(Ly, v) = \exp\{v^{-1}(\zeta(Ly) - b(L))\} - A(L).$$

The parameter  $L$  may depend on  $v$ . We suppose that for all  $(x, t) \in \mathbb{R} \times \mathbb{R}_+$ ,  $t \neq 0$ , there exists scaling  $\beta(x, t, L) > 0$ , such that for any  $\varepsilon \in \{0, 1\}$ , as  $L \rightarrow \infty$ ,

$$\begin{aligned} \beta(x, t, L) \int_{\mathbb{R}} V(Ly, v) g\left(\frac{L(y-x)}{\sqrt{T(L)tv}}\right) ((x-y)/t)^\varepsilon dy &\Rightarrow (Z^{x,t}, ((x-y)/t)^\varepsilon) \\ &:= \int_{\mathbb{R}} ((x-y)/t)^\varepsilon Z^{x,t}(dy), \end{aligned} \tag{10}$$

where  $Z^{x,t}$  is a generalized process.

The following theorem gives a convergence of finite-dimensional distributions of the rescaled solutions of (1).

**Theorem 2.1.** (i) *Suppose that  $v = 1$ , and  $T(L) = L^2$ . If the scaling  $\beta(x, t, L) := \beta(L)$  does not depend on  $(x, t)$ , then  $Z^{x,t}(dy) = g((y-x)/\sqrt{t})Z(dy)$ , where  $Z$  is a generalized process. Moreover, if for all  $\varphi$  belonging to the Schwartz space  $\mathcal{S}(\mathbb{R})$ ,*

$$\lim_{L \rightarrow \infty} \int_{\mathbb{R}} \varphi(y) \exp\{\zeta(Ly) - b(L)\} dy = a \int_{\mathbb{R}} \varphi(y) dy, \tag{11}$$

*then, as  $L \rightarrow \infty$ , the finite-dimensional distributions of the two parameter random fields,*

$$L\beta(L)u(Lx, L^2t, 1), \quad x \in \mathbb{R} \text{ and } t > 0,$$

*converge to the corresponding finite-dimensional distributions of the random field*

$$a^{-1} \int_{\mathbb{R}} \frac{x-y}{t} g\left(\frac{x-y}{\sqrt{t}}\right) Z(dy), \quad x \in \mathbb{R} \text{ and } t > 0.$$

(ii) *If  $A(L) = 0$ , then the finite-dimensional distributions of the two parameter random field,*

$$T(L)L^{-1}u(Lx, T(L)t, v), \quad x \in \mathbb{R} \text{ and } t > 0,$$

*converge to the corresponding finite-dimensional distributions of the random field*

$$\frac{\int_{\mathbb{R}} (x-y)/(t) Z^{x,t}(dy)}{\int_{\mathbb{R}} Z^{x,t}(dy)}, \quad x \in \mathbb{R} \text{ and } t > 0.$$

**Proof.** The proof of (i) is similar to the proof of Theorem 2.1 in Funaki et al. (1995), and it is based on the well-known following lemma.

**Lemma 2.1.** *If  $\{X_t, t > 0\}$  and  $\{Y_t, t > 0\}$  are families of random variables such that  $X_t \rightarrow X$  in law and  $Y_t \rightarrow c$  in probability where  $c$  is a real constant, then  $X_t + Y_t \rightarrow X + c, X_t Y_t \rightarrow cX, X_t/Y_t \rightarrow X/c$  (if  $c \neq 0$ ), in law as  $t \rightarrow +\infty$ .*

Now we come back to the proof of (ii). In this case  $V(Ly, v) = \exp\{v^{-1}(\zeta(Ly) - b(L))\}$ , and from (2), (3) and (4) we have

$$I(Lx, T(L)t, v) = \frac{L^2}{T(L)} \int_{\mathbb{R}} \frac{x-y}{t} V(Ly, v) g\left(\frac{L(y-x)}{\sqrt{T(L)tv}}\right) dy \exp(v^{-1}b(L)),$$

and

$$J(Lx, T(L)t, v) = L \int_{\mathbb{R}} V(Ly, v) g\left(\frac{L(y-x)}{\sqrt{T(L)tv}}\right) dy \exp(v^{-1}b(L)).$$

We have for all  $\alpha_i, \gamma_i, x_i \in \mathbb{R}, t_i > 0; 1 \leq i \leq n$ ,

$$\begin{aligned} & \left[ \sum_{i=1}^n \alpha_i \beta(x_i, t_i, L) T(L) L^{-2} I(Lx_i, T(L)t_i, v) \right. \\ & \quad \left. + \sum_{i=1}^n \gamma_i L^{-1} \beta(x_i, t_i, L) J(Lx_i, T(L)t_i, v) \right] \exp(-v^{-1}b(L)) \\ & = \int_{\mathbb{R}} \left\{ \sum_{i=1}^n \left( \alpha_i \frac{x_i - y}{t_i} + \gamma_i \right) \beta(x_i, t_i, L) V(Ly, v) g\left(\frac{L(y-x_i)}{\sqrt{T(L)t_i v}}\right) \right\} dy. \end{aligned}$$

From (10) the member of the left-hand side converges weakly, as  $L \rightarrow \infty$ , to

$$\begin{aligned} & \int_{\mathbb{R}} \left\{ \sum_{i=1}^n \left( \alpha_i \frac{x_i - y}{t_i} + \gamma_i \right) Z^{x_i, t_i}(dy) \right\} \\ & = \sum_{i=1}^n \alpha_i \int_{\mathbb{R}} \frac{x_i - y}{t_i} Z^{x_i, t_i}(dy) + \sum_{i=1}^n \gamma_i \int_{\mathbb{R}} Z^{x_i, t_i}(dy). \end{aligned}$$

It follows that

$$T(L)L^{-1} \frac{I(Lx, T(L)t, v)}{J(Lx, T(L)t, v)} = T(L)L^{-1} u(Lx, T(L)t, v) \Rightarrow \frac{\int_{\mathbb{R}} (x-y)/(t) Z^{x, t}(dy)}{\int_{\mathbb{R}} Z^{x, t}(dy)},$$

whence the claimed result.  $\square$

**Remark 2.1.** If  $A(L) = A$ , and  $\beta(L) \rightarrow \infty$ , as  $L \rightarrow \infty$ , then (11) is satisfied, with  $a = A$ .

**Corollary 2.1.** Let  $\zeta(x)$  be a stationary zero-mean Gaussian process.

(1) If  $\zeta(x)$  satisfies (5), then part (i) of Theorem 2.1 is satisfied with

$$b(L) = 0, \quad A(L) = 0, \quad \beta(L) = \sqrt{L}, \quad a = \exp\{E[\zeta(0)^2]/2\},$$

and  $Z(dy) = cW(dy)$ ; where  $W$  is the Gaussian white noise on  $\mathbb{R}$  and  $c = a(\int_{\mathbb{R}} (e^{B(z)} - 1) dz)^{1/2}$ .

(2) If  $\zeta(x)$  satisfies (9), then part (ii) of Theorem 2.1. is satisfied with

$$\begin{aligned} v &= 1/\sqrt{2 \ln L}, \quad T(L) = L^2 \sqrt{2 \ln L}, \\ b(L) &= \sqrt{\lambda_0}(\sqrt{2 \ln L} + \ln[(\lambda_2/\lambda_0)^{1/2}/2\pi])/\sqrt{2 \ln L} \end{aligned}$$

and

$$\beta(x, t, L) = (\lambda_2/2\pi\lambda_0)^{1/2}L\sqrt{2 \ln L},$$

$$Z^{x,t}(\mathrm{d}y) = g\left(\frac{x-y}{\sqrt{t}}\right) \sum_{i \in \mathbb{Z}} \pi^{-1/2} \exp(\sqrt{\lambda_0}u_i)\delta(y - y_i).$$

Here  $\sum_i \delta_{(y_i, u_i)}$  is a random Poisson measure on  $\mathbb{R}^2$  with intensity  $e^{-u} \mathrm{d}u \mathrm{d}y$ .

(3) If  $\zeta(x)$  satisfies (7) then part (ii) of Theorem 2.1 is satisfied with

$$v = \frac{1}{L^2\sqrt{2 \ln L}}, \quad T(L) = L^2\sqrt{2 \ln L}, \quad b(L) = \sqrt{2 \ln L} + \ln[(\lambda_2)^{1/2}/2\pi]/\sqrt{2 \ln L}$$

and the normalization  $\beta(x, t, L)$  is given by

$$1/\beta(x, t, L) = \int_{\Delta_{j(x,t)}(L)} \exp(v^{-1}\{\zeta(Ly) - b(L)\})g\left(\frac{L(x-y)}{\sqrt{t}}\right) \mathrm{d}y,$$

where  $\Delta_{j(x,t)}(L) = \{y \in \mathbb{R} : |y - y_{j(x,t)}| \leq \frac{1}{L\sqrt{2 \ln L}}\}$ , and  $Z^{x,t}(\mathrm{d}y) = \delta(y - y_{j(x,t)})$ . Here  $y_{j(x,t)}$  is given by (8).

The proofs of (1), (2), and (3) are due, respectively, to Albeverio et al. (1994), Surgailis (1996) and Molchanov et al. (1995).

### 3. The Gaussian scenario

In the sequel  $v = 1$ , and  $T(L) = L^2$ . The initial value is a functional of the one-dimensional Gaussian white noise  $W$  with intensity the Lebesgue measure. More precisely, we consider the setting where  $\zeta(x)$  is such that the random variables  $\exp\{\zeta(x)\}$  belong to  $L^2(\mathbb{P})$ , the Hilbert space of square integrable functions with respect to  $\mathbb{P}$ , the measure of the underlying probability space supporting  $W$ .

We are interested in the asymptotic behavior of the ratio  $u(Lx, L^2t, 1) := u(Lx, L^2t)$  as  $L \rightarrow +\infty$ . In view to obtain the normalization  $\beta(L)$ , the centering constants  $b(L)$ , and  $A(L)$ , we employ the Gaussian chaos decomposition Itô (1951). The nonlinear stochastic functional  $\exp\{\zeta(x)\}$  has the chaos decomposition

$$\exp\{\zeta(x)\} = \sum_{k=0}^{\infty} \frac{1}{k!} I_k(f_k(x)), \tag{12}$$

where for  $k \geq 1$ ,  $I_k(f_k(x)) = \int \cdots \int_{\mathbb{R}^k} f_k(x, y_1, \dots, y_k) \mathrm{d}W(y_1) \dots \mathrm{d}W(y_k)$  is the stochastic multiple Wiener integral, and for all  $x \in \mathbb{R}$ ,  $f_k(x) : y = (y_1, \dots, y_k) \rightarrow f_k(x, y)$  is a symmetric function belonging to  $L^2(\mathbb{R}^k, \mathrm{d}x^{\otimes k})$ . The latter space is endowed with the natural scalar product denoted  $\langle \cdot, \cdot \rangle$ , and with the norm  $\|\cdot\| = \langle \cdot, \cdot \rangle^{1/2}$ . If  $\zeta(x)$  is a Gaussian process then (12) is reduced to the Hermite expansion. The Gaussian chaos decomposition has a long history of application to the Burgers, and Navier-Stokes turbulence, both in the mathematical, and in the fluid dynamic communities, for references see e.g. Funaki et al. (1995).

We set for  $k \geq 0$ , and  $\varphi$  belonging to the Schwartz space  $\mathcal{S}(\mathbb{R})$ ,

$$v_k(\varphi, L) = \frac{1}{k!} \int_{\mathbb{R}} \varphi(y) I_k(f_k(Ly)) \mathrm{d}y.$$

From (12) we get

$$v(\varphi, L) := \int_{\mathbb{R}} \varphi(y) \exp\{\xi(Ly)\} dy = \sum_{k=0}^{\infty} v_k(\varphi, L).$$

From the well-known formula Itô (1951),

$$\mathbb{E} [I_k(f_k(x_1))I_j(f_j(x_2))] = k! \delta_k^j \langle f_k(x_1), f_k(x_2) \rangle,$$

we have

$$\mathbb{E} [v(\varphi, L)] = v_0(\varphi, L), \quad \text{and} \quad \text{Var}(v(\varphi, L)) = \sum_{k=1}^{+\infty} \text{Var}(v_k(\varphi, L)).$$

For  $k \geq 1$ , we set  $B_k(x_1, x_2) = \langle f_k(x_1), f_k(x_2) \rangle$ , and we get

$$\text{Var}(v_k(\varphi, L)) = \frac{1}{k!} \int \int_{\mathbb{R}^2} \varphi(y_1)\varphi(y_2)B_k(Ly_1, Ly_2) dy_1 dy_2.$$

**Theorem 3.1.** *Suppose that  $\mathbb{E}[\exp\{\xi(x)\}] := A$ , a constant which does not depend on  $x$ , and there exist a positive measurable function  $\beta(L) \rightarrow \infty$ , as  $L \rightarrow \infty$ , a measure  $\mu$  belonging to the space of tempered distributions  $\mathcal{S}'(\mathbb{R}^2)$  such that: for all  $G$  belonging to the Schwartz space  $\mathcal{S}(\mathbb{R}^2)$ , as  $L \rightarrow \infty$ ,*

$$[H_1] : \int \int_{\mathbb{R}^2} \beta(L)^2 B_1(Lx, Ly) G(x, y) dx dy \rightarrow \int \int_{\mathbb{R}^2} G(x, y) \mu(dx, dy),$$

$$[H_2] : \beta(L)^2 \sup_{k \geq 2} \int \int_{\mathbb{R}^2} |B_k(Lx, Ly) G(x, y)| dx dy \rightarrow 0.$$

(1) *Then we have for all  $\varphi \in \mathcal{S}(\mathbb{R})$ , as  $L \rightarrow \infty$ ,*

$$\beta(L) \int_{\mathbb{R}} [\exp\{\xi(Ly)\} - A] \varphi(y) dy = \beta(L)(v(\varphi, L) - v_0(\varphi, L)) \rightarrow \int_{\mathbb{R}} \varphi(y) Z(dy),$$

where  $Z$  is a Gaussian field whose covariance function is

$$\text{cov}((Z, \varphi_1), (Z, \varphi_2)) = \int \int_{\mathbb{R}^2} \varphi_1(y_1)\varphi_2(y_2)\mu(dy_1, dy_2).$$

(2) *It follows that*

$$L\beta(L)u(Lx, L^2t, 1) \Rightarrow A^{-1} \int_{\mathbb{R}} \frac{x-y}{t} g\left(\frac{x-y}{\sqrt{t}}\right) Z(dy).$$

**Proof.** The proof of part (1) is a consequence of Lemma 2.1. Indeed, thanks to the hypotheses  $[H_1]$  and  $[H_2]$ , we have, as  $L \rightarrow \infty$ , the convergence in law of  $\beta(L)v_1(\varphi, L)$  to the Gaussian process  $\int_{\mathbb{R}} \varphi(y)Z(dy)$ , and the convergence of  $\beta(L)(v(\varphi, L) - v_0(\varphi, L) - v_1(\varphi, L))$  to 0 in  $L^2(\mathbb{P})$ . Hence according to Lemma 2.1  $\beta(L)(v(\varphi, L) - v_0(\varphi, L))$  converges in law to  $\int_{\mathbb{R}} \varphi(y)Z(dy)$ .

The proof of part (2) is a consequence of part (i) of Theorem 2.1, and from Remark 2.1.

**Corollary 3.1.** *Suppose that*

$$\xi(x) = \int_{\mathbb{R}} f(x, y) dW(y) - \frac{\|f(x, \cdot)\|^2}{2},$$

where  $f$  is such that  $B_1(x, x) = \|f(x, \cdot)\|^2 \leq 1$  for all  $x$ . If the hypotheses  $[C_1]$  and  $[C_2]$  below are satisfied:

$[C_1]$  there exist a positive measurable function  $\beta(L) \rightarrow \infty$ , as  $L \rightarrow \infty$ , and two positive measures  $\mu_1, \mu_2 \in \mathcal{S}'(\mathbb{R}^2)$  such that for any  $G \in \mathcal{S}(\mathbb{R}^2)$ ,

$$\int_{\mathbb{R}^2} \beta(L)^2 B_1^+(Lx, Ly) G(x, y) \, dx \, dy \rightarrow \int_{\mathbb{R}^2} G(x, y) \, d\mu_1(x, y)$$

and

$$\int_{\mathbb{R}^2} \beta(L)^2 B_1^-(Lx, Ly) G(x, y) \, dx \, dy \rightarrow \int_{\mathbb{R}^2} G(x, y) \, d\mu_2(x, y),$$

where  $B_1^+, B_1^-$  are, respectively, the positive and negative parts of  $B_1$ .

$[C_2]$  For all  $\varepsilon > 0$ , and for all compact set  $K$ , subset of  $\mathbb{R}^2$ , as  $L \rightarrow \infty$ ,

$$\beta(L)^2 \lambda(\{(x, y) \in K, |B_1(Lx, Ly)| \geq \varepsilon\}) \rightarrow 0,$$

where  $\lambda$  is the Lebesgue measure.

Then, for all  $\varphi \in \mathcal{S}(\mathbb{R})$ ,

$$\beta(L) \int_{\mathbb{R}} [\exp(\xi(Ly)) - 1] \varphi(y) \, dy$$

converges in distribution to the random variable

$$\int_{\mathbb{R}} \varphi(y) Z(dy) := (Z, \varphi),$$

where  $Z(dy)$  is a Gaussian field whose covariance function is

$$\text{cov}((Z, \varphi_1), (Z, \varphi_2)) = \int \int_{\mathbb{R}^2} \varphi_1(y_1) \varphi_2(y_2) \, d\mu(y_1, y_2),$$

where  $\mu = \mu_1 - \mu_2$ .

**Proof.** For any  $x \in \mathbb{R}$ ,  $\exp(\xi(x)) = \sum_{k=0}^{\infty} (1/k!) I_k(f(x, \cdot)^{\otimes k})$ , hence the constant  $A$  of Theorem 3.1 is equal to 1. The condition  $[C_1]$  implies the condition  $[H_1]$  of Theorem 3.1. For  $k \geq 2$ ,  $|B_k(x, y)| = |B_1(x, y)|^k \leq B_2(x, y) \, dx \, dy - a.s.$  So, to get the condition  $[H_2]$  of Theorem 3.1 we have to prove that for all  $G \in \mathcal{S}(\mathbb{R}^2)$ , as  $L \rightarrow \infty$ ,

$$\int \int_{\mathbb{R}^2} \beta(L)^2 (B_1(Lx, Ly))^2 G(x, y) \, dx \, dy \rightarrow 0.$$

We have for  $\varepsilon > 0$ ,  $\int \int_{\mathbb{R}^2} \beta(L)^2 (B_1(Lx, Ly))^2 G(x, y) \, dx \, dy = I + J$ , where

$$I = \int \int_{\{|B_1(Lx, Ly)| < \varepsilon\}} \beta(L)^2 (B_1(Lx, Ly))^2 G(x, y) \, dx \, dy,$$

and

$$J = \int \int_{\{|B_1(Lx, Ly)| \geq \varepsilon\}} \beta(L)^2 (B_1(Lx, Ly))^2 G(x, y) \, dx \, dy.$$

It follows that there exist two constants  $c_1$  and  $c_2$  such that  $|I| \leq \varepsilon c_1$  (for  $L$  large enough), and

$$|J| \leq c_2 \int \int_{\{|B_1(Lx, Ly)| \geq \varepsilon\}} \beta(L)^2 |B_1(Lx, Ly) G(x, y)| \, dx \, dy.$$

For  $L$  sufficiently large there exist  $M > 0$  such that

$$\int \int_{\{x^2+y^2 \geq M^2\}} \beta(L)^2 |B_1(Lx, Ly)G(x, y)| \, dx \, dy \leq \varepsilon$$

and for  $K = \{(x, y) \in \mathbb{R}^2, x^2 + y^2 \leq M\}$ ,

$$\begin{aligned} & \int \int_{\{|B_1(Lx, Ly)| \geq \varepsilon\}} \beta(L)^2 |B_1(Lx, Ly)G(x, y)| \, dx \, dy \\ & \leq \varepsilon + \int \int_{K \cap \{|B_1(Lx, Ly)| \geq \varepsilon\}} \beta(L)^2 |B_1(Lx, Ly)G(x, y)| \, dx \, dy \\ & \leq \varepsilon + c\beta(L)^2 \lambda(\{(x, y) \in K, |B_1(Lx, Ly)| \geq \varepsilon\}), \end{aligned}$$

where  $c > 0$  is some constant. From that we have for  $L$  large enough,

$$I + J \leq \varepsilon c_1 + c_2[\varepsilon + c\beta(L)^2 \lambda(\{(x, y) \in K, |B_1(Lx, Ly)| \geq \varepsilon\})].$$

Now taking account of [C<sub>2</sub>] we get the desired result.

**Remark 3.1.** The case where the initial value  $u(x, 0)$ , is the Brownian motion, is out of the scope of Theorem 3.1, and Corollary 3.1.

In this part we study some situations when Corollary 3.1 is satisfied. A measurable function  $L : (0, +\infty) \rightarrow (0, +\infty)$  is slowly varying at  $+\infty$  if, for each  $t > 0$ , as  $x \rightarrow +\infty$ ,  $L(tx)/L(x) = 1$ . It is well known (see for example Appendix III in Galambos, 1978; ch. 0 in Bertoin, 1996) that  $L$  has the form

$$L(x) = \exp \left\{ u(x) + \int_1^x \frac{e(z)}{z} \, dz \right\}, \tag{13}$$

where  $u, e : (0, +\infty) \rightarrow \mathbb{R}$  are bounded measurable functions with  $u(x) \rightarrow u^* \in \mathbb{R}$ , and  $e(x) \rightarrow 0$ , as  $x \rightarrow +\infty$ .

**Corollary 3.2.** Let  $f \in L^2(\mathbb{R}, dx)$ , and  $\xi(x) = \int_{\mathbb{R}} f(x - y) \, dW(y) - \|f\|^2/2$ . The process  $\xi(x)$  is a stationary Gaussian process with covariance function  $B_1(x, y) = \int_{\mathbb{R}} f(x - z)f(y - z) \, dz := B(x - y)$ . Suppose that  $B(x - y) = (\Phi(|x - y|)/|x - y|^\alpha)$ , where  $0 < \alpha < 1$ , and  $\Phi$  is a slowly varying function such that  $0 < \Phi(|x|) \leq |x|^\alpha$  for all  $x$ . Then, for all  $\varphi \in \mathcal{S}(\mathbb{R})$ ,

$$\frac{L^{\alpha/2}}{\sqrt{\Phi(L)}} \int_{\mathbb{R}} [\exp(\xi(Ly)) - 1] \varphi(y) \, dy$$

converges in distribution to  $\int_{\mathbb{R}} \varphi(y)Z(dy) := (Z, \varphi)$ , where  $Z$  is a Gaussian field whose covariance function is

$$\text{cov}((Z, \varphi_1), (Z, \varphi_2)) = \int \int_{\mathbb{R}^2} \varphi_1(y_1)\varphi_2(y_2) \frac{dy_1 \, dy_2}{|y_1 - y_2|^\alpha}.$$

**Proof.** This result is well known, it has been considered by Albeverio et al. (1994), and Leonenko and Orsingher (1995). Corollary 3.1 gives a simplified proof. Indeed, in this case  $B_k(x, y) = ((\Phi(|x - y|)/|x - y|^\alpha))^k$ , and it is easy to see that  $|B_k(x, y)| \leq |B_2(x, y)|$  for all  $k \geq 2$ . Let us prove the condition [C<sub>1</sub>] with  $\beta(L) = \frac{L^{\alpha/2}}{\sqrt{\Phi(L)}}$  and  $\mu = (dx \, dy/|x - y|^\alpha)$ .



We have for all  $A > 0$ ,  $\int \int_{\mathbb{R}^2} \beta(L)^2 B(L|x - y|)G(x, y) dx dy = I + J$ , where

$$I = \int \int_{[L|x-y|>A]} \beta^2(L)B(L|x - y|)G(x, y) dx dy$$

and

$$J = \int \int_{[L|x-y|<A]} \beta^2(L)B(L|x - y|)G(x, y) dx dy.$$

Let us prove that  $I \rightarrow \int \int_{\mathbb{R}^2} |x - y|^{-\alpha}G(x, y) dx dy$  as  $L \rightarrow \infty$ . First we have

$$\beta^2(L)B(L|x - y|) = \Phi(L|x - y|)\Phi(L)^{-1}.$$

Using (13), we have for all  $\varepsilon > 0$ , and for  $L$  large enough,

$$\begin{aligned} |\Phi(L|x - y|)\Phi(L)^{-1}| &= \exp \left\{ u(L|x - y|) - u(L) + \int_L^{L|x-y|} \frac{e(z)}{z} dz \right\} \\ &\leq c\{|x - y|^\varepsilon + |x - y|^{-\varepsilon}\}, \end{aligned}$$

where  $c$  is some constant which does not depend on  $L$ . From the dominated convergence theorem we get

$$I \rightarrow \int \int_{\mathbb{R}^2} |x - y|^{-\alpha}G(x, y) dx dy \text{ as } L \rightarrow \infty.$$

Now let us prove that  $J \rightarrow 0$ , as  $L \rightarrow \infty$ . Since for any  $x \in \mathbb{R}$ ,  $|B(x)| \leq 1$ , and

$$J = \frac{L^\alpha}{\Phi(L)} \int \int_{[L|x-y|<A]} B(L|x - y|)G(x, y) dx dy,$$

we have the estimate

$$\begin{aligned} |J| &\leq \frac{L^\alpha}{\Phi(L)} \int \int_{[L|x-y|<A]} |G(x, y)| dx dy \\ &\leq \Phi^{-1}(L) \int \int_{[L|x-y|<A]} A^\alpha |x - y|^{-\alpha} |G(x, y)| dx dy. \end{aligned}$$

We have for  $\varepsilon > 0$ ,

$$\begin{aligned} \Phi^{-1}(L) \int \int_{[L|x-y|<A]} A^\alpha |x - y|^{-\alpha} |G(x, y)| dx dy \\ \leq L^{-\varepsilon} \Phi^{-1}(L) \int \int_{[L|x-y|<A]} A^{\alpha+\varepsilon} |x - y|^{-(\alpha+\varepsilon)} |G(x, y)| dx dy. \end{aligned}$$

But  $L^{-\varepsilon} \Phi^{-1}(L) \rightarrow 0$ , as  $L \rightarrow \infty$ , and for  $\alpha + \varepsilon < 1$ , the integral  $\int \int_{\mathbb{R}^2} A^{\alpha+\varepsilon} |x - y|^{-\alpha-\varepsilon} |G(x, y)| dx dy$  is bounded. It follows that  $J \rightarrow 0$ , as  $L \rightarrow \infty$ , whence the condition  $[C_1]$  is satisfied.

Now we prove the condition  $[C_2]$ . Since  $(\Phi(|L|)/L^\alpha) \rightarrow 0$ , as  $L \rightarrow \infty$ , then for  $L$  large enough, and for some constant  $A$ , which does not depend on  $L$ , we have

$$[B(L|x - y|) \geq \varepsilon] \subset [|x - y| \leq AL^{-1}].$$

It follows that for any compact set  $K$  and for some constant  $c$ ,

$$\begin{aligned} \beta(L)^2 \lambda(\{(x, y) \in K, B_1(Lx, Ly) \geq \varepsilon\}) &\leq \beta(L)^2 \lambda(\{(x, y) \in K, |x - y| \leq AL^{-1}\}) \\ &\leq c\beta(L)^2 L^{-1} \rightarrow 0 \end{aligned}$$

as  $L \rightarrow \infty$ , whence the desired result.

**Remarks.**

(1) Theorem 3.1 can be rephrased as a statement that the stochastic process  $\{\exp(\xi(y)), y \in \mathbb{R}\}$  has a large-scale Gaussian limit in the sense of Dobrushin (1980, p. 169) with normalization  $A, \beta(L)$ . It is well known (Dobrushin, 1979; Dobrushin and Major, 1979) that necessarily,  $\beta(L) = L^\kappa / \sqrt{\Phi(L)}$ , for some constant  $\kappa \in \mathbb{R}$ . The function  $\Phi(L)$  is slowly varying, as  $L \rightarrow \infty$ , and locally bounded. From that we conclude that if  $B(x, y) \rightarrow 0$ , as  $x, y$  go to infinity, and satisfies condition  $[C_1]$  of Corollary 3.1 with the normalization  $\beta(L) = L^\kappa / \sqrt{\Phi(L)}$  with  $0 < \kappa < \frac{1}{2}$ , then condition  $[C_2]$  holds.

(2) If instead of  $B(x) = (\Phi(|x|)/|x|^\alpha)$ ,  $\alpha \in ]0, 1[$ , we suppose only that  $B(x) \sim (\Phi(|x|)/|x|^\alpha)$ , as  $x \rightarrow \infty$ , and  $B(x) \leq 1$  for any  $x \in \mathbb{R}$ , we can show, essentially in the same way, that the conditions  $[C_1]$  and  $[C_2]$  still hold. Such a situation happens if the function  $f$  which defines the covariance function  $B$  satisfies the following conditions:

(A1)  $\int_{\mathbb{R}} f^2(x) dx < \infty$ .

(A2)  $|f(x)| \leq Cx^{H_0-3/2}\Phi(x)$  for almost all  $x > 0$ , and  $1 - 1/2m < H_0 < 1$  for some integer  $m \geq 1$ . Here  $\Phi$  is a slowly varying function at  $+\infty$ .

(A3)  $f(x) \sim x^{H_0-3/2}\Phi(x)$  as  $x \rightarrow \infty$ .

(A4) There exists a constant  $\gamma$  satisfying  $0 < \gamma < \min\{H_0 - (1 - 1/2m), 1 - H_0\}$ , such that

$$\int_{-\infty}^0 |f(u)f(xy + u)| du = o(x^{2H_0-2}\Phi(x))y^{2H_0-2-2\gamma},$$

as  $x \rightarrow \infty$ , uniformly in  $y \in (0, t]$ , for a given  $t > 0$ . See Taquq (1979, p. 57, Section 2) for more details.

**Corollary 3.3.** Let  $\xi(x) = \int_{\mathbb{R}} f(x, y) dW(y) - \|f(x, \cdot)\|^2/2$ .

(A) If  $f(x, y) = \Phi(|x|)|x|^{-\alpha}h(y)$ ,  $\Phi$  is slowly varying at  $\infty$ ;  $0 < \alpha < \frac{1}{2}$ , and  $\|h\| = 1$ , then, for all  $\varphi \in \mathcal{S}(\mathbb{R})$ , as  $L \rightarrow \infty$ ,

$$\frac{L^\alpha}{\Phi(L)} \int_{\mathbb{R}} [\exp(\xi(Ly)) - 1]\varphi(y) dy$$

converges in distribution to  $\int_{\mathbb{R}} \varphi(y)Z(dy) := (Z, \varphi)$ , where  $Z$  is a Gaussian field whose covariance function is

$$\text{cov}((Z, \varphi_1), (Z, \varphi_2)) = \int \int_{\mathbb{R}^2} \varphi_1(y_1)\varphi_2(y_2) \frac{dy_1 dy_2}{|y_1|^\alpha |y_2|^\alpha}.$$

(B) Suppose that  $f(x, \cdot) = f(|x|, \cdot) \geq 0, \forall x \in \mathbb{R}$ , non-increasing with respect to  $x$ , and for all  $x \neq 0$

$$\left( \int_{\mathbb{R}} f^2(L, z) dz \right)^{-1/2} f(Lx, \cdot)$$

converges in  $L^2(\mathbb{R})$  to some function  $g(x, \cdot)$ . We suppose also that the covariance function  $B_1(x, y) := \langle f(|x|, \cdot), f(|y|, \cdot) \rangle$  is slowly varying at infinity, namely, for all  $x \neq 0, y \neq 0$ ,

$$\lim_{L \rightarrow \infty} \frac{B_1(Lx, Ly)}{B_1(L, L)} = \int_{\mathbb{R}} g(x, z)g(y, z) dz := m(x, y)$$

and that  $m(Lx, Ly) = L^{-\alpha}m(x, y)$ , for some  $0 < \alpha < 1$ .

Then  $B_1(L, L)^{-1/2} \int_{\mathbb{R}} \varphi(y)[\exp(\zeta(Ly)) - 1] dy$  converges, as  $L \rightarrow \infty$ , to  $(Z, \varphi)$ , where  $Z$  is a Gaussian field with covariance function  $\text{cov}((Z, \varphi_1), (Z, \varphi_2)) = \int \int_{\mathbb{R}^2} \varphi_1(x)\varphi_2(y)m(x, y) dx dy$ .

**Proof.** We use Corollary 3.1 for the proof. In the assertion (A), we have  $B_k(x, y) = (\Phi(x)^k \Phi^k(y))/(|x|^{\alpha k} |y|^{\alpha k})$ , and it is easy to see, for all  $k \geq 2$ , that  $|B_k(x, y)| \leq |B_2(x, y)|$ .

Let us prove the condition [C<sub>1</sub>] of Corollary 3.1, with  $\beta(L) = (L^\alpha/\Phi(L))$ , and  $\mu = (dx dy/|x|^\alpha |y|^\alpha)$ . We split  $\int \int_{\mathbb{R}^2} \beta(L)^2 B_1(Lx, Ly)G(x, y) dx dy = I + J$ , where

$$I = \int \int_{[L(|x|+|y|) > A]} \beta^2(L)B_1(Lx, Ly)G(x, y) dx dy,$$

and  $J = \int \int_{[L(|x|+|y|) \leq A]} \beta^2(L)B_1(Lx, Ly)G(x, y) dx dy$ . Let us prove that  $I \rightarrow \int \int_{\mathbb{R}^2} |x|^{-\alpha} |y|^{-\alpha} G(x, y) dx dy$ , as  $L \rightarrow \infty$ . From (13), we have for all  $\varepsilon > 0$ , and for  $L$  large enough,

$$|\Phi(L|x|)\Phi(L)^{-1}| = \exp \left\{ u(L|x|) - u(L) + \int_L^{L|x|} \frac{e(z)}{z} dz \right\} \leq c\{|x|^\varepsilon + |x|^{-\varepsilon}\},$$

where  $c$  is constant which does not depend on  $L$ . From the dominated convergence theorem we get

$$I \rightarrow \int \int_{\mathbb{R}^2} |x|^{-\alpha} |y|^{-\alpha} G(x, y) dx dy \quad \text{as } L \rightarrow \infty.$$

Now let us prove that  $J \rightarrow 0$  as  $L \rightarrow \infty$ . Since for any  $(x, y) \in \mathbb{R}^2$ ,  $|B_1(x, y)| \leq 1$ , and

$$J = \frac{L^{2\alpha}}{\Phi^2(L)} \int \int_{[L(|x|+|y|) < A]} B_1(Lx, Ly)G(x, y) dx dy,$$

we have the estimate

$$\begin{aligned} |J| &\leq \frac{L^{2\alpha}}{\Phi^2(L)} \int \int_{[L(|x|+|y|) < A]} |G(x, y)| dx dy \\ &\leq \Phi^{-2}(L) \int \int_{[L(|x|+|y|) < A]} A^{2\alpha}(|x| + |y|)^{-2\alpha} |G(x, y)| dx dy. \end{aligned}$$

We have for  $\varepsilon > 0$ ,

$$\begin{aligned} &\Phi^{-2}(L) \int \int_{[L(|x|+|y|) < A]} A^{2\alpha}(|x| + |y|)^{-2\alpha} |G(x, y)| dx dy \\ &\leq L^{-\varepsilon} \Phi^{-2}(L) \int \int_{[L(|x|+|y|) < A]} A^{2\alpha+\varepsilon}(|x| + |y|)^{-(2\alpha+\varepsilon)} |G(x, y)| dx dy. \end{aligned}$$

If  $2\alpha + \varepsilon < 1$ , then the integral  $\int \int_{\mathbb{R}^2} A^{2\alpha+\varepsilon}(|x| + |y|)^{-(2\alpha+\varepsilon)} |G(x, y)| dx dy$  is bounded. From that, and from the fact that  $L^{-\varepsilon} \Phi^{-2}(L) \rightarrow 0$ , as  $L \rightarrow \infty$ , we have  $J \rightarrow 0$  as  $L \rightarrow \infty$ , whence the condition [C<sub>1</sub>] is satisfied.

Now, we prove the condition [C<sub>2</sub>]. Since  $B_1(x, y) \rightarrow 0$ , as  $|x| + |y| \rightarrow \infty$ , then for  $L$  large enough, and for some constant  $A$ , which does not depend on  $L$ , we have

$$[B_1(Lx, Ly) \geq \varepsilon] \subset [|x| + |y| \leq AL^{-1}].$$

It follows that for any compact set  $K$ , and for some constant  $c$ ,

$$\begin{aligned} \beta(L)^2 \lambda(\{(x, y) \in K, B_1(Lx, Ly) \geq \varepsilon\}) &\leq \beta(L)^2 \lambda(\{(x, y) \in K, |x| + |y| \leq AL^{-1}\}) \\ &\leq c\beta(L)^2 L^{-2} \rightarrow 0, \end{aligned}$$

as  $L \rightarrow \infty$ , which yields the desired result.  $\square$

Now we prove the assertion (B). To prove [C<sub>1</sub>] it is sufficient to take  $G > 0$ . Let  $a > 0, b > 0$ . We split  $B_1(L, L)^{-1} \int \int_{\mathbb{R}^2} B_1(Lx, Ly)G(x, y) dx dy = I + J$ , where

$$I = \int \int_{[|x| > a, |y| > b]} B_1(L, L)^{-1} B_1(Lx, Ly)G(x, y) dx dy$$

and

$$J = \int \int_{[|x| \leq a] \cup [ |y| \leq b]} B_1(L, L)^{-1} B_1(Lx, Ly)G(x, y) dx dy.$$

From the monotonicity of  $B$ , and the dominated convergence theorem we have

$$I \rightarrow \int \int_{[|x| > a, |y| > b]} G(x, y)m(x, y) dx dy \quad \text{as } L \rightarrow \infty.$$

Now,

$$G(x, y)B_1(L, L)^{-1} B_1(Lx, Ly)1_{[0 < |x| < |y|]} \leq G(x, y)B_1(L, L)^{-1} B_1(Lx, Lx)1_{[0 < |x| < |y|]}$$

and each of these functions converges as  $L \rightarrow \infty$ . Using the same proof as in Corollary 3.2 we show that the integral of the latter function converges. Combining the fact that  $\int \int_{[0,1]^2} m(x, y) dx dy < \infty$ , see Haan and Resnick (1979), with the following variant of Fatou’s lemma: if  $0 \leq h_n \leq g_n$  are real-valued functions on some measure space, and  $h_n \rightarrow h, g_n \rightarrow g$ , then  $\int h_n \rightarrow \int h$  provided  $\int g < \infty$ , we can show that

$$J \rightarrow \int \int_{[|x| \leq a] \cup [ |y| \leq b]} G(x, y)m(x, y) dx dy,$$

which yields [C<sub>1</sub>]. [C<sub>2</sub>] follows from the fact that  $B_1(x, y) \rightarrow 0$  as  $x, y \rightarrow \infty$  and from  $B_1(L, L) \sim L^{-\alpha}$ , as  $L \rightarrow \infty$ .

We finish this work by considering the case where  $\langle f(y_1, \cdot), f(y_2, \cdot) \rangle = 0$  for  $|y_1| + |y_2| \geq 1$ . This example is out of the scope of Theorem 3.1. The covariance function  $B_1$  is positive and vanishing outside the ball  $|x| + |y| \leq 1$ , then the condition [C<sub>1</sub>] holds with  $\beta(L) = L$  but [C<sub>2</sub>] does not hold. We want to prove the following theorem.

**Theorem 3.2.** *Under the latter hypothesis we have for all  $\varphi \in \mathcal{S}(\mathbb{R})$*

$$L \int_{\mathbb{R}} \varphi(y) \left[ \exp \left\{ \int f(Ly, u) dW(u) - \frac{B_1(Ly, Ly)}{2} \right\} - 1 \right] dy \rightarrow \mathcal{N}(0, \varphi^2(0)c),$$

where  $c = \int \int_{\mathbb{R}^2} (\exp\{B_1(x, y)\} - 1) dx dy$ , and

$$L^2 u(Lx, L^2 t) \rightarrow c \frac{xZ}{t},$$

where  $Z$  is a normal random variable.

**Proof.** We will use the technique of Theorem 1.1 in Hu and Woyczynski (1995). Let us prove that

$$L \int_{-\infty}^{\infty} \varphi(y) \left[ \exp \left\{ \int f(Ly, u) dW(u) - \frac{B_1(Ly, Ly)}{2} \right\} - 1 \right] dy$$

converges in distribution to  $\mathcal{N}(0, \varphi(0)^2 c)$ . Let

$$H(L) = L \int_{-\infty}^{-L} \varphi(y) \left[ \exp \left\{ \int f(Ly, u) dW(u) - \frac{B_1(Ly, Ly)}{2} \right\} - 1 \right] dy.$$

Then

$$\mathbb{E}[|H(L)|^2] \leq kL^2 \left( \int_{-\infty}^{-L} \varphi(y) dy \right)^2.$$

Using Chebyshev’s inequality and since  $\varphi \in \mathcal{S}(\mathbb{R})$  we get  $H(L) \rightarrow 0$  in probability as  $L \rightarrow \infty$ . Similarly, we can get as  $L \rightarrow \infty$  the convergence in probability to 0 of

$$L \int_L^{\infty} \varphi(y) \left[ \exp \left\{ \int f(Ly, u) dW(u) - \frac{B_1(Ly, Ly)}{2} \right\} - 1 \right] dy.$$

The remainder term is

$$\begin{aligned} L \int_{-L}^L \varphi(y) \left[ \exp \left\{ \int f(Ly, u) dW(u) - \frac{B_1(Ly, Ly)}{2} \right\} - 1 \right] dy \\ = I + \sum_{-L^2 \leq k \leq L^2} \eta_k(L), \end{aligned}$$

where

$$\eta_k(L) = L \int_{k/L}^{(k+1)/L} \varphi(y) \left[ \exp \left\{ \int f(Ly, u) dW(u) - \frac{B_1(Ly, Ly)}{2} \right\} - 1 \right] dy$$

and  $I$  converges in probability to 0. Since  $B(Ly_1, Ly_2) = \int_{\mathbb{R}} f(Ly_1, u) f(Ly_2, u) du = 0$  for  $y_1 \in [k/L, (k + 1)/L], y_2 \in [j/L, (j + 1)/L]$  with  $|k - j| > 2$ , then the sequence  $\eta_k(L), -L^2 \leq k \leq L^2$  is 2-dependent sequence. It is useful to recall the following Bulinskii (1987) result: let  $X_j(t), j \in U(t)$  be an  $m(t)$ -dependent field on a finite set  $U(t) \subset \mathbb{Z}$ , and let, for some  $s \in (2, 3]$  and all  $t > 0$ ,

$$\sup_{j \in U(t)} (\mathbb{E}[|X_j(t)|^s])^{1/s} = C_s(t) < \infty.$$

Then

$$\begin{aligned} \sup_{x \in \mathbb{R}} \left| \mathbb{P} \left( \delta^{-1}(t) \sum_{j \in U(t)} (X_j(t) - \mathbb{E}[X_j(t)]) \leq x \right) - F(x) \right| \\ \leq k_0 |U(t)| M_s^s(t) m^{s-1}(t) + M_s(t) m(t) + |U(t)|^{1/2} M_s^2(t) m^{2/s}(t), \end{aligned}$$

where  $\delta^2(t) = \text{Var} \sum_{j \in U(t)} X_j(t)$ ,  $k_0$  is some constant,  $|U(t)|$  is a number of points in  $U(t)$ ,  $M_s(t) = \delta^{-1}(t) C_s(t)$ , and  $F(x)$  is the distribution function of  $\mathcal{N}(0, 1)$ .

Now we return to our proof. For  $(\eta_k(L), -L^2 \leq k \leq L^2)$  we get  $|U(L)| = 2L^2, C_3(L) \leq k_1 L^{-1}, 0 < k_2 \leq \delta(L) \leq k_3, M_3(L) \leq k_4 L^{-1}$ , and then we have

$$k_0 |U(L)| M_3^3(L) 2^2 + M_3(L) 2 + |U(L)|^{1/2} M_3^2(L) 2^{2/3} \leq kL^{-1},$$

so that

$$L \int_{-\infty}^{\infty} \varphi(y) \left[ \exp \left\{ \int f(Ly, u) dW(u) - \frac{\|f(Ly)\|^2}{2} \right\} - 1 \right] dy \rightarrow \mathcal{N}(0, \varphi(0)^2 c).$$

Now part (i) of Theorem 2.1 combined with Remark 2.1 achieve the proof.

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