## SHORT COMMUNICATION

# A NOTE ON BILINEAR TIME SERIES MODELS 

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Conditions for the existence of a stationary solution for certain forms of bilinear difference equations are derived.

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In [1] certain non-linear 'superdiagonal' time series models are discussed, of the form

$$
\begin{equation*}
\sum_{0}^{p} \alpha_{j} x(t-j)+\sum_{j=0}^{Q} \beta_{j} \varepsilon(t-j)+\sum_{1=k<l=2}^{Q P} \gamma_{k l} \varepsilon(t-k) x(t-l)=0 . \tag{1}
\end{equation*}
$$

Here we assume

$$
\begin{equation*}
\alpha_{0}=\beta_{0}=1, \quad \sum \alpha_{i} z^{i} \neq 0, \quad|z| \leqslant 1 \tag{2}
\end{equation*}
$$

and that the $\varepsilon(t)$ are ormally and independently distributed witio zero mean and variance $\sigma^{2}$. In [1] the question of the existence of a stationary solution to (1), in terms of the $\varepsilon(s), s \leqslant t$, is discussed but not completely. In [2] the somewhat mos: difficult question is discussed as to whether, given a stationary solution, the $\varepsilon(t)$ are the (non-linear) innovations, i.e., whether $\varepsilon(t)=x(t)-\mathscr{E}\left\{x(t) \mid \mathscr{F}_{t-1}\right\}$ wherc $\mathscr{F}_{t}$ is the $\sigma$-algebra determined by $x(s), s \leqslant t$. Thus it has to be determined whe aer the $\varepsilon(t)$ are measurable, $\mathscr{F}_{i}$. In [1, p. 41] it is said that a necessary and suficient condition for the existence of a second order stationary solution to (1) is, when $p=q=0$, that

$$
\sigma^{2} \sum_{k<l} \gamma_{k l}^{2} z^{\prime}=1
$$

should have all solutions outside of the closed unit disc. This condition is derived from the equation

$$
\mathscr{E}\left\{x(t)^{2}\right\}=\sigma^{2} \sum \sum \gamma_{k l}^{2} \mathscr{E}\left\{x(t-l)^{2}\right\}+\sigma^{2}
$$

which is, however, incorrect. In any case the argument would be incomplete. However it is fairly easy to obtain a necessary and sufficient condition in this superdiagonal case. If $x(t)$ is to be a stationary solution of the required form then successive substitutions using (1) show that this solution must be

$$
\begin{equation*}
x(t)=\sum_{r=1}^{\infty} \Sigma_{r} \mathrm{~d}_{r}\left(j_{1}, \ldots, j_{r}\right) \varepsilon\left(t-j_{1}\right) \cdots \varepsilon\left(t-j_{r}\right) \tag{3}
\end{equation*}
$$

where $\Sigma_{r}$ is over $0<j_{1}<\cdots<j_{r}<\infty$ except for $r=1$ when $0 \leqslant j_{1}<\infty$. Indeed at any stage in this process of substitution there will be a single term, $x(t-m)$ let us say, in any product, any other factors being a constant or an $\varepsilon(t-j)$. Moreover $m$ will be larger than any such lag, $j$, for an $\varepsilon(t-j)$. The form of (3) now follows.

Because of the condition $j_{1}<j_{2}<\cdots<j_{r}$ the individual terms in (3) are easily seen to be orthogonal. Moreover if $r \neq s$ it also follows that

$$
\begin{align*}
& \mathscr{E}\left[\Sigma_{r}\left\{\mathrm{~d}_{r}\left(j_{1}, \ldots, j_{r}\right) \varepsilon\left(t-j_{1}\right) \cdots \varepsilon\left(t-j_{r}\right)\right\}\right. \\
& \left.\quad \times \Sigma_{s}\left\{\mathrm{~d}_{s}\left(k_{1}, \ldots, k_{s}\right) \varepsilon\left(\tau-k_{1}\right) \cdots \varepsilon\left(\tau-k_{s}\right)\right\}\right]=0 \tag{4}
\end{align*}
$$

for any $t, \tau$.
We consider the necessary and sufficient condition for mean square convergence, i.e.,

$$
\begin{equation*}
\sum_{1}^{\infty} \sigma^{2 r} \Sigma_{r} \mathbf{d}_{r}\left(j_{1}, \ldots, j_{r}\right)^{2}<\infty \tag{5}
\end{equation*}
$$

Put

$$
\begin{aligned}
& \alpha(z)=\sum \alpha_{i} z^{j}, \quad \beta(z)=\sum \beta_{i} z^{j}, \quad \gamma\left(z_{1}, z_{2}\right)=\sum \sum \gamma_{k,} z_{1}^{k} z_{2}^{l}, \\
& \delta_{r}\left(z_{1}, \ldots, z_{r}\right)=\Sigma_{r} \mathrm{~d}_{r}\left(j_{1}, \ldots, j_{r}\right) z_{1}^{i_{1}} \cdots z_{r}^{j_{r} .}
\end{aligned}
$$

Now substituting in (1) we see that

$$
\delta_{1}\left(z_{1}\right)=-\beta\left(z_{1}\right) / \alpha\left(z_{1}\right)
$$

and that

$$
\alpha\left(z_{1} z_{2} \cdots z_{r}\right) \delta_{r}\left(z_{1}, \ldots, z_{r}\right)+\gamma\left(z_{1}, z_{2} \cdots z_{r}\right) \delta_{r-1}\left(z_{2}, \ldots, z_{r}\right)=0
$$

so that we obtain, in general,

$$
\begin{align*}
& \delta_{r}\left(z_{1}, \ldots, z_{r}\right)= \\
& \quad=(-)^{r} \frac{\gamma\left(z_{1}, z_{2} z_{3} \cdots z_{r}\right) \gamma\left(z_{2}, z_{3} \cdots z_{r}\right) \cdots \gamma\left(z_{r-1}, z_{r}\right) \beta\left(z_{r}\right)}{\alpha\left(z_{1} z_{2} \cdots z_{r}\right) \alpha\left(z_{2} z_{3} \cdots z_{r}\right) \cdots \alpha\left(z_{r-1} z_{r}\right) \alpha\left(z_{r}\right)} . \tag{6}
\end{align*}
$$

It is evident that this gives the unique solution of (1) in the form (3). Now using (6) the left side of (5) is

$$
\begin{equation*}
\sum_{r=1}^{\infty}\left(\frac{\sigma^{2}}{2 \pi}\right)^{r} \int_{-\pi}^{\pi} \cdots \int\left|\delta_{r}\left(\mathrm{e}^{\mathrm{i} \omega_{1}}, \mathrm{e}^{\mathrm{i} \omega_{2}}, \ldots, \mathrm{e}^{\mathrm{i} \omega_{r}}\right)\right|^{2} \prod_{1}^{r} \mathrm{~d} \omega_{j} \tag{7}
\end{equation*}
$$

It follows easily that a simple sufficient condition for (5) is

$$
\begin{equation*}
\frac{\sigma^{2}}{2 \pi} \int_{-\pi}^{\pi}\left|\frac{\gamma\left(\mathrm{e}^{\mathrm{i} \omega}, \mathrm{e}^{\mathrm{i} \phi}\right)}{\alpha\left(\mathrm{e}^{\mathrm{i}(\omega+\phi)}\right)}\right|^{2} \mathrm{~d} \omega \leqslant c<1, \quad \phi \in[-\pi, \pi] \tag{8}
\end{equation*}
$$

since then (7) is dominated by a geometrically convergent series. The expression (6) may also be used to evaluate the spectral density of a stationary, mean square convergent, solution to (1). Because of (4) it is

$$
\sum_{r=0}^{\infty}\left(\frac{\sigma^{2}}{2 \pi}\right)^{r} \int_{\Sigma \omega_{l}=\omega}\left|\delta_{r}\left(\mathrm{e}^{\mathrm{i} \omega_{1}}, \ldots, \mathrm{e}^{\mathrm{i} \omega_{r}}\right)\right|^{2} \Pi \mathrm{~d} \omega_{j} .
$$

Fut $\xi(r)=\Sigma_{r} \mathrm{~d}_{r}\left(j_{1}, \ldots, j_{r}\right) \varepsilon\left(t-j_{1}\right) \cdots \varepsilon\left(t-j_{r}\right)$. Then under (8) the series $\sum \xi(r)$ converges almost surely, sin:e $\mathscr{E}\{|\xi(r)|\} \leqslant\left[\mathscr{E}\left\{\xi(r)^{2}\right\}\right]^{1 / 2}$ and thus $\sum \mathscr{E}\{|\xi(r)|\}$ is dominated by a geometrically convergent series. Also

$$
\begin{equation*}
\sum_{1=j_{1}<\cdots<j_{r}=r}^{N} \mathrm{~d}_{r}\left(j_{1}, \ldots, j_{r}\right) \varepsilon\left(t-j_{1}\right) \cdots \varepsilon\left(t-j_{r}\right) \tag{9}
\end{equation*}
$$

converges almost surely as can be seen by rearranging (9), for example for $r=2$, as

$$
\sum_{j_{2}=2}^{N} \varepsilon\left(t-j_{2}\right) \sum_{j_{1}=1}^{i_{2}-1} \mathrm{~d}_{2}\left(j_{1}, j_{2}\right) \varepsilon\left(t-j_{1}\right) .
$$

This shows that (9) is, for each fixed $t$, a square integrable martingale, with bounded mean square, with respect to a sequence of $\sigma$-algebras $G_{N}$ where $G_{N}$ is generated by $\varepsilon(t-1), \ldots, \varepsilon(t-N)$.

If the assumption that the model is superdiagonal is eliminated, when the last term in (1) becomes

$$
\sum_{k=0}^{O} \sum_{l=1}^{P} \gamma_{k \mid} \varepsilon(t-k) x(t-l),
$$

then (6) still holds but, because $0<j_{1}<\cdots<j_{r}$ in (3) no longer holds, then (7) is invalid. Thus it seems difficult in general to use (6) and to study convergence in general. However in certain special cases, not superdiagonal, a treatment is given in [3].

The discussion of the question as to whether $\varepsilon(t)$ is measurable $\mathscr{F}_{t}$ seems intrinsically difficult. One might consider cases for which (7) is finite and express $\varepsilon(t)$ in terms of the $x(t-j)$ by a formula of the form of (3) but with the place; of $x(t)$ and $\varepsilon(t)$ reversed. Of course $j_{1}<j_{2}<\cdots<j_{r}$ will not hold. Then the coefficients
are generated by formulae of the form of

$$
(-1)^{r} \frac{\gamma\left(z_{2}, z_{3}, \ldots, z_{r}, z_{1}\right) \cdots \gamma\left(z_{r}, z_{r-1}\right) \alpha\left(z_{r}\right)}{\beta\left(z_{1} z_{2} \cdots z_{r}\right) \cdots \beta\left(z_{r-1} z_{r}\right) \beta\left(z_{r}\right)} .
$$

However the lack of orthogonality makes this difficult to use. Again some special cases have been dealt with in [3].

## References

[1] C.W.J. Granger and A.P. Andersen, An Introduction to Bilinear Time Series Models (Vandenhoeck and Ruprecht, Göttingen, 1978).
[2] C.W.J. Grariger and A.P. Andersen, On the invertibility of time series models, Stochastic Process. Appl. 8 (19"8) 87-92.
[3] T.D. Phan and L.T. Tran, Quelques résultats sur les modèles bilinéaires de séries chronologique, C.R. Acad. Sci. Paris 290, Série A (1980) 335-338.

