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Estimating influenza vaccine effectiveness using routine surveillance data among children aged 6–59 months for five consecutive influenza seasons



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SUMMARY

Objectives: We aimed to estimate the pooled vaccine effectiveness (VE) in children over five winters through data linkage of two existing surveillance systems.

Methods: Five test-negative case–control studies were conducted from November to February during the 2004/2005 to 2008/2009 seasons. Sentinel physicians from the Viral Surveillance Network enrolled children aged 6–59 months with influenza-like illness to collect throat swabs. Through linking with a nationwide vaccination registry, we measured the VE with a logistic regression model adjusting for age, gender, and week of symptom onset. Both fixed-effects and random-effects models were used in the meta-analysis.

Results: Four thousand four hundred and ninety-four subjects were included. The proportion of influenza test-positive subjects across the five seasons was 11.5% (132/1151), 7.2% (41/572), 23.9% (189/791), 6.6% (75/1135), and 11.2% (95/845), respectively. The pooled VE was 62% (95% confidence interval (CI) 48–83%) in both meta-analysis models. By age category, VE was 51% (95% CI 23–68%) for those aged 6–23 months and 75% (95% CI 60–84%) for those aged 24–59 months.

Conclusions: Influenza vaccination provided measurable protection against laboratory-confirmed influenza among children aged 6–59 months despite variations in the vaccine match during the 2004/2005 to 2008/2009 influenza seasons in Taiwan.

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1. Introduction

Influenza viruses cause annual epidemics and the occasional pandemic of acute respiratory disease, which pose a threat to the health of the population.¹ Vaccination is considered a priority in public health departments and is an effective way to prevent influenza-associated morbidity, mortality, and expense.² Since

1998, the Department of Health in Taiwan has gradually endorsed annual influenza vaccination campaigns to encourage susceptible subjects, including the elderly, healthcare workers, poultry workers, and young children, to receive free influenza immunization, based on the recommendations of the Advisory Committee on Immunization Practices in Taiwan.

The recommendation of universal influenza vaccination of young children was not popular in many countries initially, probably because of the absence of studies providing solid evidence of effectiveness in the targeted population.^{3,4} During the 2004/2005 influenza season, the Centers for Disease Control in Taiwan (Taiwan CDC) started to vaccinate groups of children aged 6–23 months; this program was extended to those aged 24–35

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months starting in 2008/2009. All vaccination target groups had the same opportunity to receive free influenza shots beginning in October each year. In addition to the recommended groups, all people could receive free influenza vaccination after December 1 each season in order to best utilize the influenza vaccine resources and to increase the vaccine coverage of the entire population.

It is important to determine the influenza vaccine effectiveness (VE) after the implementation of such a program. Previous studies have encouraged large studies to assess the impact of influenza vaccination on children in terms of specific outcome measurements.^{5,6} Furthermore, multiyear studies are preferred for estimating robust influenza VE over time through a meta-analysis methodology.^{7,8} The Taiwan CDC has successfully coordinated a laboratory-based surveillance network for influenza virus for all ages since 2000 and established the National Immunization Information System (NIIS) for children aged <6 years in 2003.⁹ By using the retrospective laboratory-confirmed influenza surveillance data and linking these to individual vaccination records, we were able to rapidly and efficiently demonstrate the influenza VE in children for the 2004/2005 to 2008/2009 seasons.

Previous reports have demonstrated influenza VE using routinely collected laboratory and/or surveillance data and directly pooling results from multiple years to provide the overall VE.^{10–12} In this study, we implemented a fixed-effects and a random-effects meta-analysis of case–control studies to estimate the pooled VE for children aged 6–59 months across the five consecutive influenza seasons, and considered the variation in antigenic match across seasons and epidemics year by year as the heterogeneity between studies. Such effectiveness studies of inactivated influenza vaccine among young children could assist public health sectors in reassessing the current national influenza vaccine match varies year to year.

2. Subjects and methods

2.1. Study population

Children aged 6–59 months with an influenza-like illness (ILI) during the November to February winter epidemics over five seasons from 2004/2005 to 2008/2009 were investigated. The Viral Surveillance Network required sentinel physicians to collect throat or nasal swabs among verbally consenting ILI patients regardless of the patient's influenza vaccination status and underlying medical conditions. ILI was defined as a body temperature \geq 38 °C plus one of the following four clinical manifestations: cough, sore throat, hoarseness and running nose, or headache and myalgia/fatigue. This study was initiated as a public health response and used routinely collected surveillance data and vaccination records to assess influenza VE. The Taiwan CDC determined these activities to be non-research and thus the study did not require review by an institutional review board.

2.2. Viral surveillance network and virological testing

The Viral Surveillance Network coordinated by the Taiwan CDC was started in October 2000; it comprises 10–13 collaborating laboratories (the number is affected by the annual budget) and aims to survey and isolate nationwide circulating viruses related to respiratory tract infections year-round.⁹ Clinical specimens obtained from nasal or throat swabs were collected by the sentinel physicians and sent to the local collaborating laboratories for virus identification using viral culture and/or reverse transcriptase PCR. Methods of virus isolation have been described previously.¹³ The Taiwan CDC collected and analyzed these results on a weekly basis

and posted this information on their website. The antigenic match between vaccine and circulating strains in each season was evaluated by hemagglutination inhibition (HAI) assay.

2.3. Determination of case and control subjects

Children whose specimens tested positive for laboratoryconfirmed influenza infection during the study periods were defined as case subjects. Control subjects were those with the same symptoms but who were negative for influenza. For cases and controls, information about age, sex, week of symptom onset, and personal identifiers were obtained from the reports submitted by the sentinel physicians.

2.4. Influenza vaccination status

Information on the influenza vaccination status of the subjects was obtained from the NIIS, which was established by the Taiwan CDC to collect vaccination records for children at a national level. Children were classified as vaccinated if they had received one or more vaccine doses in the current influenza season and it was administered \geq 14 days before the onset of ILI. Children were classified as unvaccinated for the given season if they were not vaccinated in that study season or if they had received the first vaccine dose within 14 days before respiratory tract infection. In this study, we did not define the status of partially vaccinated children because many influenza epidemic strains in Taiwan become the vaccine strains 2-3 years later, as shown by hemagglutination sequence comparisons.^{9,14} Therefore, children aged 6-59 months were considered immunized if they had received one or more vaccine doses in the current influenza season regardless of previous influenza immunization history.

2.5. Statistical analysis

We linked the NIIS and National Viral Surveillance System using the personal identifier. We used a logistic regression model to adjust for age, gender, and week of symptom onset, with the first week including November 1 and the last week including February 28 in the five different epidemic seasons.^{15,16} The adjusted odds ratio (aOR) was used to model the association between influenza vaccination and laboratory-confirmed influenza-related medical visits in each season. VE and 95% confidence intervals (CI) were estimated using the formula VE = $(1 - OR) \times 100\%$. Stratified VE estimates were calculated according to age (6-23 months or 24-59 months) and adjusted for gender and week of symptom onset. We used both a fixed-effects model with inverse variance method and a random-effects model with DerSimonian-Laird weighting method¹⁷ to run the synthesis results. A forest plot was used to display the estimated overall ORs and separate ORs in the five epidemic seasons according to the two age groups.¹⁸ We used the 'meta' package for the R system for statistical computing to implement the meta-analysis.^{19,20} Annual vaccination rates among control groups were examined using the Cochran-Armitage test for trend in SAS, version 9.3 (SAS Institute Inc., Cary, NC, USA).

3. Results

In Taiwan, five winter epidemics occurred between November 2004 and February 2009, which were dominated by influenza A H1N1 in 2005/2006, 2007/2008, and 2008/2009, influenza B in 2004/2005, and influenza B followed by influenza A H3N2 in 2006/2007 (Figure 1, Table 1). Information on influenza activity obtained from the Viral Surveillance Network demonstrated that positive rates of influenza isolates for 6–59-month-old children varied each winter; from high to low, these rates were 23.9% (189/791) in the



Figure 1. Nationwide laboratory-based influenza surveillance in Taiwan from July 2004 to June 2009 is illustrated. Five winter epidemics occurred between November 2004 and February 2009. Vaccine match: The antigenic match between recommended vaccines and circulating viruses in each season was evaluated using the hemagglutination inhibition (HAI) assay. The vaccine match (%) was calculated through the ratios of collected influenza A H1N1, influenza A H3N2, and influenza B viruses and the respective antigenic match of three tested type/subtype viruses.

2006/2007 season, 11.5% (132/1151) in the 2004/2005 season, 11.2% (95/845) in the 2008/2009 season, 7.2% (41/572) in the 2005/2006 season, and 6.6% (75/1135) in the 2007/2008 season (Table 2). Furthermore, matches between circulating and vaccine influenza strains were analyzed based on HAI assays, and a range of 12–93% of influenza virus isolates were antigenically similar to influenza vaccine strains during 2004/2005 to 2008/2009 (Figure 1).

VE was estimated in a total of 4494 children aged 6–59 months for whom laboratory results and vaccination status were available for the five winter epidemics. Table 2 shows the demographic distribution of the cases and controls in each season. Categorized according to epidemic, the sex distribution was similar among case and control subjects; however, the proportion of subjects who tested positive was significantly higher in children aged 24–59 months (15.2%) than in those aged 6–23 months (6.1%). From 2004/ 2005 through 2008/2009 seasons, annual vaccination rates among

Table 1

Antigenic characteristics of influenza viruses and vaccine strains from the 2004/2005 to 2008/2009 seasons

Season and type or subtype	Vaccine component	Circulating strains
2004/2005		
H1N1	A/New Caledonia/20/99-like	-
H3N2	A/Fujian/411/2002-like	A/California/7/2004(H3N2)-like
В	B/Shanghai/361/2002-like	B/Malaysia/2506/2004-like B/Shanghai/361/2002-like
2005/2006		
H1N1	A/New Caledonia/20/99-like	A/New Caledonia/20/99*
H3N2	A/California/7/2004-like	A/Wisconsin/67/2005-like
В	B/Shanghai/361/2002-like	-
2006/2007		
H1N1	A/New Caledonia/20/99	-
H3N2	A/Wisconsin/67/2005 or	A/Wisconsin/67/2005-like
	A/Hiroshima/52/2005	
В	B/Malaysia/2506/2004	B/Malaysia/2506/2004-like
2007/2008		
H1N1	A/Solomon Islands/3/2006	A/Brisbane/59/2007-like
H3N2	A/Wisconsin/67/2005 or A/Hiroshima/52/2005	A/Brisbane/10/2007-like
В	B/Malaysia/2506/2004	B/Florida/4/2006-like
2008/2009		
H1N1	A/Brisbane/59/2007	A/Brisbane/59/2007-like*
H3N2	A/Brisbane/10/2007	A/Brisbane/10/2007-like
В	B/Florida/4/2006-like	B/Florida/4/2006-like

Dominant types/subtypes during the given season.

control groups in children aged 6-23 months were 66.6% (213/320), 61.8% (139/225), 57.1% (129/226), 41.0% (236/575), and 47.1% (115/ 244), which were higher than the rates in those aged 24–59 months (10.7% (75/699), 21.2% (65/306), 18.6% (70/376), 19.4% (94/485), and 19.0% (96/506)) (test for trend, p < 0.0001 and p < 0.001 for the two age groups, respectively). Vaccine coverage rates among the control groups were close to those in the corresponding population nationwide. The national influenza vaccine coverage available for children aged 6-23 months was 68.9% (260 499/377 933), 66.6% (243 149/365 335), 60.3% (212 605/352 502), and 44.8% (150 675/ 335 972) for seasons 2004/2005 to 2007/2008, respectively, which were estimated through the NIIS database. The national vaccine coverage changed to 48.2% (255 565/530 561) for children aged 6-35 months in the 2008/2009 season when the influenza vaccination program was extended to those aged 24-35 months. Otherwise, influenza was detected in 532 (11.8%) enrollees. Approximately 11.1% (59/532) of those who tested positive and 31.1% (1232/3962) of those who tested negative had been vaccinated.

The pooled estimate of VE for children aged 6–59 months during 2004/2005 to 2008/2009 was 62% (95% CI 48–83%) using the meta-analysis method (Figure 2). By age category (6–23 months and 24–59 months), the VE estimates were 51% (95% CI 23–68%) and 75% (95% CI 60–84%) for those aged 6–23 months and 24–59 months, respectively. The I^2 value of 0% possibly indicates that statistical heterogeneity was not observed across the five winter epidemics, and across the age ranges of 6–23 months and 24–59 months (Figure 2). The VE estimates were higher among those aged 24–59 months than among those aged 6–23 months across the five seasons.

4. Discussion

In this study, we found clear evidence that the current public health policy to reduce laboratory-confirmed influenza among young children through immunization is effective. The linkage of routinely collected data is considered an efficient method for estimating influenza vaccine effectiveness more accurately.^{21,22} We successfully demonstrated an efficient way to evaluate the influenza VE for each winter epidemic through data linkage of two already established systems in the public health sector and used meta-analysis to estimate the pooled VE for children aged 6–59 months over consecutive seasons.

Meta-analysis could be appropriate when a group of studies is sufficiently homogeneous in terms of participants, interventions, and outcomes to provide a meaningful summary.²³ The reason we

Table 2

Characteristics of influenza-positive case subjects and influenza-negative control subjects according to influenza season, 2004/2005 through 2008/2009

$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Characteristics	Cases		Control	p-Value	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	2004/2005	(<i>n</i> =132)		(<i>n</i> =101		
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Gender					0.33
Female5440.9%44944.1%Unknown75.3%828.0%Age group<0.01	Male	71	53.8%	488	47.9%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Female	54	40.9%	449	44.1%	
Age group<0.01 $6-23$ months11 8.3% 320 31.4% $24-59$ months121 91.7% 699 68.6% Vaccinated12 9.1% 288 28.3% Unvaccinated120 90.9% 731 71.7% $2005/2006$ $(n=41)$ $(n=531)$ $(n=531)$ Gender0.85 0.85 0.85 Male21 51.2% 295 55.6% Female19 46.3% 226 42.6% Unknown1 2.4% 10 1.9% Age group<0.01	Unknown	7	5.3%	82	8.0%	
$\begin{array}{c ccccc} 6-23 \mbox{ months} & 11 & 8.3\% & 320 & 31.4\% \\ 24-59 \mbox{ months} & 121 & 91.7\% & 699 & 68.6\% \\ Vaccinated & 12 & 91.7\% & 288 & 28.3\% \\ Unvaccinated & 120 & 90.9\% & 731 & 71.7\% \\ \hline \\ \hline 2005/2006 & (n=41) & (n=531) \\ \hline \\ \hline \\ Gender & & & & & & & & & & & & & & & & & & &$	Age group					< 0.01
$\begin{array}{c cccc} 24-59 \mbox{ months} & 121 & 91.7\% & 699 & 68.6\% \\ Vaccination status & & <0.01 \\ Vaccinated & 12 & 9.1\% & 288 & 28.3\% \\ Unvaccinated & 120 & 90.9\% & 731 & 71.7\% \\ \hline \\ \hline 2005/2006 & (n=41) & (n=531) \\ \hline \\ Gender & & & & & & & & & & & & & & & & & & &$	6–23 months	11	8.3%	320	31.4%	
Vaccination status <0.01 Vaccinated 12 9.1% 288 28.3% Unvaccinated 120 90.9% 731 71.7% 2005/2006 $(n=41)$ $(n=531)$ 0.85 Male 21 51.2% 295 55.6% Female 19 46.3% 226 42.6% Unknown 1 2.4% 10 1.9% Age group <0.01	24–59 months	121	91.7%	699	68.6%	
Vaccinated 12 9.1% 288 28.3% Unvaccinated 120 90.9% 731 71.7% 2005/2006 $(n=41)$ $(n=531)$ Gender 0.85 Male 21 51.2% 295 55.6% Female 19 46.3% 226 42.6% Unknown 1 2.4% 10 1.9% Age group 6-23 months 4 9.8% 225 42.4% 24-59 months 37 90.2% 306 57.6% Vaccination status <	Vaccination status					< 0.01
Unvaccinated 120 90.9% 731 71.7% 2005/2006 $(n=41)$ $(n=531)$ 0.85 Male 21 51.2% 295 55.6% Female 19 46.3% 226 42.6% Unknown 1 2.4% 10 1.9% Age group 6-23 months 4 9.8% 225 42.4% 24-59 months 37 90.2% 306 57.6% 7.6% Vaccinated 3 7.3% 204 38.4% 0.01 Unvaccinated 38 92.7% 327 61.6% $2006/2007$ $(n=189)$ $(n=602)$ Gender 0 0.0% 7 1.2% Age $geonp$ <0.01 6-23 months 45 23.8% 226 37.5% $24-59$ <0.01 6-23 months 45 23.8% 226 37.5% <0.01 Vaccinated 25 13.2% 199	Vaccinated	12	9.1%	288	28.3%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Unvaccinated	120	90.9%	731	71.7%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	2005/2006	(n = 41))	(<i>n</i> = 53	1)	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Gender					0.85
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	Male	21	51.2%	295	55.6%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Female	19	46.3%	226	42.6%	
Age group<0.01 $6-23 \text{ months}$ 49.8%22542.4% $24-59 \text{ months}$ 3790.2%30657.6%Vaccinated37.3%20438.4%Unvaccinated3892.7%32761.6%2006/2007(n=189)(n=602)0.29Gender0.00%71.2%Age group0.00%71.2%Age group0.00%71.2%Age group0.00%71.2%Vaccinated2513.2%19933.1%Unvaccinated16486.8%40366.9%2007/20082007/2008(n=75)(n=1060)0.40Male4661.3%58054.7%Female2837.3%47344.6%Unknown11.3%70.7%Age group<	Unknown	1	2.4%	10	1.9%	
	Age group					<0.01
$\begin{array}{c ccccc} 24-59 \mbox{ months} & 37 & 90.2\% & 306 & 57.6\% \\ Vaccination status & Vaccinated & 3 & 7.3\% & 204 & 38.4\% \\ Unvaccinated & 38 & 92.7\% & 327 & 61.6\% \\ \hline 2006/2007 & (n=189) & (n=602) \\ \hline Gender & & & & & & & & & & & & & & & & & & &$	6–23 months	4	9.8%	225	42.4%	
Vaccination status <0.01	24-59 months	37	90.2%	306	57.6%	
Vaccinated37.3%20438.4%Unvaccinated3892.7%32761.6%2006/2007 $(n = 189)$ $(n = 602)$ Gender0.29Male10957.7%35559.0%Female8042.3%24039.9%Unknown00.0%71.2%Age group </td <td>Vaccination status</td> <td></td> <td></td> <td></td> <td></td> <td><0.01</td>	Vaccination status					<0.01
Unvaccinated 38 92.7% 327 61.6% 2006/2007 $(n = 189)$ $(n = 602)$ Gender 0.29 Male 109 57.7% 355 59.0% Female 80 42.3% 240 39.9% Unknown 0 0.0% 7 1.2% Age group <0.01	Vaccinated	3	7.3%	204	38.4%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Unvaccinated	38	92.7%	327	61.6%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	2006/2007	2006/2007 (<i>n</i> = 189)		(<i>n</i> = 60	2)	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Gender					0.29
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Male	109	57.7%	355	59.0%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Female	80	42.3%	240	39.9%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Unknown	0	0.0%	7	1.2%	
	Age group					< 0.01
$\begin{array}{c ccccc} 24-59 \text{ months} & 144 & 76.2\% & 376 & 62.5\% \\ \hline Vaccination status & & <& <0.01 \\ \hline Vaccinated & 25 & 13.2\% & 199 & 33.1\% \\ \hline Unvaccinated & 164 & 86.8\% & 403 & 66.9\% \\ \hline 2007/2008 & (n=75) & (n=1060) \\ \hline Gender & & & 0.40 \\ \hline Male & 46 & 61.3\% & 580 & 54.7\% \\ \hline Female & 28 & 37.3\% & 473 & 44.6\% \\ \hline Unknown & 1 & 1.3\% & 7 & 0.7\% \\ \hline Age group & & & <0.01 \\ \hline 6-23 \text{ months} & 28 & 37.3\% & 575 & 54.2\% \\ \hline 24-59 \text{ months} & 47 & 62.7\% & 485 & 45.8\% \\ \hline Vaccination status & & <<0.01 \\ \hline Vaccinated & 8 & 10.7\% & 330 & 31.1\% \\ \hline \end{array}$	6–23 months	45	23.8%	226	37.5%	
Vaccination status <0.01	24–59 months	144	76.2%	376	62.5%	
Vaccinated 25 13.2% 199 33.1% Unvaccinated 164 86.8% 403 66.9% 2007/2008 (n = 75) (n = 1060) Gender 0.40 Male 46 61.3% 580 54.7% Female 28 37.3% 473 44.6% Unknown 1 1.3% 7 0.7% Age group <0.01 $6-23$ months 28 37.3% 575 54.2% 24–59 months 47 62.7% 485 45.8% <0.01 Vaccination status <0.01 <0.01	Vaccination status	25	12.20/	100	22.1%	<0.01
Onvacchated 164 86.8% 403 66.9% 2007/2008 $(n=75)$ $(n=1060)$ Gender 0.40 Male 46 61.3% 580 54.7% Female 28 37.3% 473 44.6% Unknown 1 1.3% 7 0.7% Age group <0.01 $6-23$ months 28 37.3% 575 54.2% 24-59 months 47 62.7% 485 45.8% <0.01 Vaccination status <0.01 Vaccinated 8 10.7% 330 31.1%	Vaccinated	25	13.2%	199	33.1%	
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Unvaccinated	164	80.8%	403	66.9%	
Gender 0.40 Male 46 61.3% 580 54.7% Female 28 37.3% 473 44.6% Unknown 1 1.3% 7 0.7% Age group <0.01	2007/2008	(<i>n</i> =75)	(n = 100)	50)	
Male 46 61.3% 580 54.7% Female 28 37.3% 473 44.6% Unknown 1 1.3% 7 0.7% Age group 6-23 months 28 37.3% 575 54.2% 24–59 months 47 62.7% 485 45.8% Vaccination status <<0.01	Gender					0.40
Female 28 37.3% 47.3 44.6% Unknown 1 1.3% 7 0.7% Age group 6-23 months 28 37.3% 575 54.2% 24-59 months 47 62.7% 485 45.8% Vaccination status <<0.01	Male	46	61.3%	580	54.7%	
Unknown 1 1.3% 7 0.7% Age group <0.01	Female	28	37.3%	473	44.6%	
Age group <0.01 6-23 months 28 37.3% 575 54.2% 24-59 months 47 62.7% 485 45.8% Vaccination status <0.01	Unknown	I	1.3%	/	0.7%	-0.01
24-59 months 27 62.7% 485 45.8% Vaccination status <0.01	Age group	20	27.2%	575	51.2%	<0.01
Vaccination status 47 62.7% 465 43.8% Vaccination status <0.01	24_{50} months	20 47	57.5% 62.7%	185	J4.2% 45.8%	
Vaccinated 8 10.7% 330 31.1%	Vaccination status	47	02.7%	405	45.6%	<0.01
Vacchiated 0 10.7% 550 51.1%	Vaccinated	8	10.7%	330	31.1%	<0.01
Unvaccinated 67 89.3% 730 68.9%	Unvaccinated	67	89.3%	730	68.9%	
2008/2009 (n - 95) (n - 750)	2008/2009	(n - 05)	(n - 75)	1)	
	2000/2005	(11-55)	(11 - 750	5)	
Gender 0.74	Gender	50	- 4	207	50.000	0.74
Ividie 52 54.1% 397 52.9% Formula 42 45.2% 252 47.1%	Iviale Famala	52	54./%	397	52.9%	
Feilidie 43 40.3% 303 47.1%	Feilidie	43	45.5%	353	4/.1%	
UIIKIIUWII U U.U% U U.U%		U	0.0%	U	0.0%	-0.01
Age group <0.01 6-23 months 16 16.8% 244 22.5%	6-23 months	16	16.8%	2//	32.5%	<0.01
24-59 months 79 83.2% 506 67.5%	24_{59} months	79	83.7%	506	67.5%	
Vaccination status /0.01	Vaccination status	,5	03.2/0	500	07.3/0	< 0.01
Vaccinated 11 11.6% 211 28.1%	Vaccinated	11	11.6%	211	28.1%	20.01
Unvaccinated 84 88.4% 539 71.9%	Unvaccinated	84	88.4%	539	71.9%	

did not simply add up the subjects in each season to summarize the influenza VE across the five seasons of test-negative case-control studies was to consider heterogeneity such as variation in the antigenic match among years, influenza activity year by year, vaccine policy changes over time with expansion of the age groups targeted, and changes in the dominant circulating subtypes of influenza virus, etc. Although statistical heterogeneity between studies was not found in our study, to estimate VE across years using a meta-analysis methodology might be a feasible and applicable approach. Further stratified analysis of influenza type-specific effectiveness was not possible due to the small sample of laboratory-confirmed cases with virus type categories.

We estimated VE against laboratory-confirmed influenza infection using the test-negative case-control study design, which is less susceptible to bias due to misclassification of infection and to confounding by health-seeking behavior, relative to traditional case-control or cohort studies.¹⁵ Virological swab tests for influenza as part of routine influenza surveillance to estimate influenza VE more specifically in time and to compare VEs internationally is encouraged and practiced in many countries.²⁴ Recent studies have suggested that vaccine-induced protection against influenza may decline over time among young children and older adults in the test-negative case-control study design.^{25,26} Nunes et al. raised the issue of the best influenza-negative control group to use in the test-negative study design.²⁷ They observed that a VE difference existed when choosing a non-influenza virus control group and a pan-negative control group. Further studies should be conducted to clarify such important issues. Another challenge exists in establishing and maintaining the quality of the vaccination register for vaccination programs, especially for annual influenza vaccinations; such a register will provide information for resource allocation, comprehensive evaluation, and a timely response to all vaccine-preventable diseases.²⁸

Our findings demonstrated a significant pooled VE of 51% (95% CI 23-68%) in 6-23-month-old children over five seasons. Previous reports have shown that influenza vaccines are effective in healthy children against laboratory-confirmed influenza, serologically confirmed influenza, and clinical illness by systematic metaanalysis.^{29,30} A recent study reported by Yang et al. with a study design similar to ours found an influenza VE against medicallyattended influenza illness of 16% for those aged 6-35 months in the 2012/2013 season.³¹ Evidence for VE against outcomes other than laboratory-confirmed influenza infection such as preventing emergency department visits and hospitalizations for ILI in children has been evaluated, although more robust evidence is needed in the future.³² Nevertheless, few of these studies have provided adequate evidence of influenza VE among children younger than 2 years of age, who, without chronic or serious medical conditions, are still at increased risk of hospitalization during the influenza season.^{3,4,33–35} Findings regarding the efficacy and effectiveness of inactivated influenza vaccines in children younger than 2 years are inconsistent.²⁴ Although a smaller reduction in laboratory-confirmed influenza infection was shown in this age group compared to those older than 2 years of age, the clinical relevance and public health implications of routine influenza vaccination for both age groups has been confirmed.^{3,30,36}

Several factors may affect the efficacy and effectiveness of influenza vaccine, including (1) antigenic similarity between the circulating and vaccine types or strains of influenza virus; (2) specificity of the outcome measurement of VE; (3) yearly variability in influenza illness rates; (4) host characteristics (e.g., age and underlying medical conditions) in relation to immune responses; (5) vaccine coverage and herd immunity; and (6) relatively small sample sizes for influenza-positive cases within each stratum evaluated.^{7,30,33,37-41}

Annual vaccination rates in children aged 6–23 months appeared to decline over the study period. This might be explained by the severe acute respiratory syndrome (SARS) outbreaks in Taiwan in 2003 and the introduction of the influenza vaccination program in the 2004/2005 season for children aged 6–23 months. Both of these events were an incentive to parents to have their young children vaccinated with the influenza vaccine during the early part of the study.

The possibility of cross-protection by influenza vaccine with a suboptimal match has been debated.⁴² In years with a suboptimal

(A)								
Study	TE	seTE	Odds Ratio		OR	95%-CI	W(fixed)	W(random)
2004-2005 (6-59 mo)	-0.74	0.3593		C	.48	[0.24; 0.97]	20.4%	20.4%
2005-2006 (6-59 mo)	-1.23	0.6872		C	.29	[0.08; 1.12]	5.6%	5.6%
2006-2007 (6-59 mo)	-1.03	0.2696		C	0.36	[0.21; 0.60]	36.2%	36.2%
2007-2008 (6-59 mo) 2008-2009 (6-59 mo)	-1.18 -0.89	0.3898		0	.31	[0.14; 0.66] [0.20; 0.83]	20.5%	20.5%
Fixed effect model Random effects mode	 % tourse	warad-0		C C	.38 .38	[0.27; 0.52] [0.27; 0.52]	100% 	 100%
neterogeneny. Psquared~0.	70, Iau-sq	uareu-o,		_				
			0.1 0.5 1 2	10				
(B)								
Study	TE	seTE	Odds Ratio		OR	95%-CI	W(fixed)	W(random)
2004-2005 (6-23 mo)	-0.51	0.658		(0.60	[0.17; 2.19]	11.7%	11.7%
2005-2006 (6-23 mo)	-0.17	1.050	*	— 9	0.85	[0.11; 6.62]	4.6%	4.6%
2005-2007 (6-23 mo)	-0.88	0.340		l l	J.41	[0.21; 0.80]	43.1%	43.1%
2008-2009 (6-23 mo)	-0.26	0.535		().40).77	[0.10, 1.02]	17.6%	17.6%
Fixed effect model				(0.49	[0.32; 0.77]	100%	
Random effects mode Heterogeneity: Hsquared=0	el %. tau-so	nuared=0	→ p=0.8244	().49	[0.32; 0.77]		100%
neterogeneny. rsquareu-o	70, 188-50	400100-0	, p=0.0244	7				
			0.2 0.5 1 2	5				
(C) Study	TE	a TE	Odds Ratio		0 0	05% CI	W/fived)	W/rondom)
Study	16	Seic	Ouds Mado		UR	95%-01	w(lixed)	w(random)
2004-2005 (24-59 mo)	-0.77	0.422	+	0	.46	[0.20; 1.06]	30.8%	30.8%
2005-2006 (24-59 mo)	-2.38	1.055 -		0	.09	[0.01; 0.73]	4.9%	4.9%
2006-2007 (24-59 mo)	-1.63	0.397		0	.20	[0.09; 0.43]	34.8%	34.8%
2007-2008 (24-59 mo)	-1.66	0.733		0	.19	[0.05; 0.80]	10.2%	10.2%
2008-2009 (24-59 mo)	-1.45	0.533		C	.23	[0.08; 0.67]	19.3%	19.3%
Fixed effect model			4	0	.25	[0.16; 0.40]	100%	
Random effects mode Heterogeneity: I-squared=09	 %, <i>tau-</i> sq	uared=0	ρ=0.4697	0	.25	[0.16; 0.40]		100%

Figure 2. A fixed-effects model with inverse variance method and a random-effects model with DerSimonian–Laird weighting method were applied to illustrate the vaccine effectiveness by meta-analysis. Forest plots are used to display the estimated overall and separate adjusted odds ratios (aORs) and 95% confidence interval (CI) in the five epidemic seasons among children aged 6–59 months, stratified by age (6–23 months and 24–59 months). TE: log odds ratio; seTE: standard error of the log odds ratio; weight.

match, the VE has typically been lower or even not clearly demonstrated.^{40,41} However, we found that only 12% of circulating strains were similar to the vaccine strains in the 2007/2008 season, but the VE was not as low as expected. The lack of significant VE in 6–23-month-old children despite a high vaccine match in the 2005/2006 and 2008/2009 seasons might be related to the low ILI rates, which were associated with the relatively small sample size of children who tested positive for influenza.⁴⁰ The seasonal variations in terms of the proportions of case subjects in our data obtained from the Viral Surveillance Network may reflect the variable nature of influenza epidemics.³⁸ The effect measure modification of age might be a concern in the analysis of vaccine effectiveness; therefore it was appropriate to demonstrate both age strata and age adjustment in the statistical models to estimate vaccine effectiveness.⁴³

Children aged 6 months to 8 years who have never received seasonal influenza vaccines previously or who have received only one dose in their first year of vaccination should receive two doses of seasonal influenza vaccine to be categorized as fully immunized according to the recommendations of the US CDC.^{38,44} However, influenza epidemic strains in Taiwan often circulate earlier than

vaccine strains recommended by the World Health Organization.⁹ Therefore, in this study, we defined children aged 6–59 months who had received at least one dose of seasonal influenza vaccine as immunized regardless of their previous influenza vaccination history. This may not be appropriate for infants aged 6–23 months.⁴⁵ We speculate that the younger the children were, the less chance they had had to experience the circulating local strains in previous seasons for immune priming before vaccination, which could have led to an underestimation of VE among younger children if they were only partially immunized. Besides, high vaccine coverage (50–70%) in children could possibly have an impact on reducing influenza-related morbidity and mortality, not only among the vaccinated children, but also in other age groups.^{30,37} The challenge of increasing vaccine coverage still exists.

The findings of this study are subject to several limitations. First, children with underlying medical conditions are at even greater risk of an adverse outcome related to influenza than healthy children.^{33,39} Data linkage of the Viral Surveillance Network and the NIIS did not provide personal medical information, so we could not examine this issue. Second, only two

categories of vaccination status were defined in this study. To resolve this problem, a validation study to confirm the serology indicator of an HAI antibody titer >1:40 after one dose of trivalent influenza vaccine between two age strata might assist in confirming the immunization status in different geographic areas.⁴⁶ Third, there might have been selection bias with regard to the vaccination status of subjects who were enrolled for laboratory testing. It is possible that the testing of samples was biased as this was done by the clinicians who provided the influenza shot to the patient. However, in Taiwan, most patients can see the doctor of their choice, therefore, most clinicians would not know the patient's influenza vaccination status. If sampling was done on an informed basis such that the sentinel physician collected fewer swabs from vaccinated ILI patients, a potential biasing of estimates of effectiveness upwards might have occurred.⁴⁷ In the test-negative design for estimating influenza VE, effectiveness does not vary by health-seeking behavior with the assumption that the distribution of non-influenza causes of acute respiratory infection does not vary by influenza vaccination status.¹⁵ Therefore, cases and controls in our study probably had similar characteristics with regard to their willingness to seek medical care and their willingness to be swabbed, which could eliminate the uncontrolled confounding between cases and controls. The Viral Surveillance Network did not include all ILI patients (those willing to be swabbed and those not willing to be swabbed), thus it was not possible to determine the characteristics of the children who were swabbed and those who were not. The voluntary enrolment of swabbed participants in our study raises the question of potential selection bias, which might bias the VE estimation and limit the generalizability.³ Fourth, the impact of repeated seasonal influenza vaccination on vaccine effectiveness against influenza A and B virus has been debated, and further studies are needed to provide more clear evidence.^{48,49} As this was an observational study, residual confounding may still be present despite different statistical models. We plan to investigate this further in the future.⁵⁰

In conclusion, our study rapidly and efficiently determined VE across five influenza seasons using data linkage of immunization records and viral surveillance data at the national level. Because the annual burden of influenza illness and vaccine match could influence VE, we combined studies using a case–control design across consecutive influenza seasons using a novel method of meta-analysis. Influenza vaccination provided measurable protection against laboratory-confirmed influenza among children aged 6–23 months, 24–59 months, and the entire range of 6–59 months, despite variation in vaccine match during the 2004/2005 to 2008/2009 influenza seasons in Taiwan.

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