

## AN IMPROVED ALGORITHM FOR TRANSITIVE CLOSURE ON ACYCLIC DIGRAPHS

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**Abstract.** In [6] Goralčíková and Koubek describe an algorithm for finding the transitive closure of an acyclic digraph  $G$  with worst-case runtime  $O(n \cdot e_{\text{red}})$ , where  $n$  is the number of nodes and  $e_{\text{red}}$  is the number of edges in the transitive reduction of  $G$ . We present an improvement on their algorithm which runs in worst-case time  $O(k \cdot e_{\text{red}})$  and space  $O(n \cdot k)$ , where  $k$  is the width of a chain decomposition. For the expected values in the  $G_{n,p}$  model of a random acyclic digraph with  $0 < p < 1$  we have

$$E(k) = O\left(\frac{\log(p \cdot n)}{p}\right), \quad E(e_{\text{red}}) = O(n \cdot \log n),$$

$$E(k \cdot e_{\text{red}}) = \begin{cases} O(n^2) & \text{for } \log^2 n / n \leq p < 1, \\ O(n^2 \cdot \log \log n) & \text{otherwise,} \end{cases}$$

where “log” means the natural logarithm.

### 1. Introduction

A directed graph  $G = (V, E)$  consists of a vertex set  $V = \{1, 2, 3, \dots, n\}$  and an edge set  $E \subseteq V \times V$ . Each element  $(v, w)$  of  $E$  is an edge and joins  $v$  to  $w$ . If  $G_1 = (V_1, E_1)$  and  $G_2 = (V_2, E_2)$  are directed graphs,  $G_1$  is a subgraph of  $G_2$  if  $V_1 \subseteq V_2$  and  $E_1 \subseteq E_2$ . The subgraph of  $G_2$  induced by the subset  $V_1$  of  $V_2$  is the graph  $G_1 = (V_1, E_1)$ , where  $E_1$  is the set of all elements of  $E_2$  which join pairs of elements of  $V_1$ . Unless we specify otherwise, any subgraph referred to in this paper is the subgraph induced by its vertex set. A path in a graph from vertex  $v_0$  to vertex  $v_s$  is a sequence of vertices  $v_0, v_1, \dots, v_s$  such that  $(v_{i-1}, v_i)$  is an edge for  $i \in \{1, 2, \dots, s\}$ ;  $s$  is the length of the path. The path is simple if all its vertices are pairwise distinct. A path  $v_0, \dots, v_s$  is a cycle if  $s \geq 1$  and  $v_0 = v_s$ , and a simple cycle if in addition  $v_1, \dots, v_{s-1}$  are pairwise distinct. A graph without cycles is acyclic. A topological sorting of a digraph  $G = (V, E)$  is a mapping  $\text{ord}: V \rightarrow \{1, 2, \dots, |V|\}$  such that for all edges  $(v, w) \in E$  we have  $\text{ord}(v) < \text{ord}(w)$ . The relation between an acyclic digraph  $G$  and a topological sorting is given in Theorem 1.1 (for a proof see [8, pp. 320–323]).

**Theorem 1.1.**  $G = (V, E)$  is acyclic if and only if it has a topological sorting. A topological sorting of an acyclic graph can be computed in linear time  $O(|V| + |E|)$ .

Throughout this paper we will assume that the acyclic digraph  $G = (V, E)$  is topologically sorted, i.e.,  $(i, j) \in E$  implies  $i < j$  and that the adjacency lists  $\text{out}(v) =$

$\{w \in V \mid (v, w) \in E\}$  are sorted in increasing order. This can be achieved in linear time  $O(|V| + |E|)$  (see [8, p. 323]). Now we need some more definitions. The node  $w$  is in the reflexive, transitive closure  $\text{out}^*(v)$  if and only if there is a path from  $v$  to  $w$  in  $G$ . The set  $\text{out}^{\text{red}}(v) = \{w \in \text{out}(v) \mid \text{there is no path of length at least 2 from } v \text{ to } w \text{ in } G\}$  is called the transitive reduction of the node  $v$ . Let  $E^* = \{(v, w) \mid w \in \text{out}^*(v)\}$ ,  $E_{\text{red}} = \{(v, w) \mid w \in \text{out}^{\text{red}}(v)\}$ ; then  $G^* = (V, E^*)$  ( $G_{\text{red}} = (V, E_{\text{red}})$ ) is called the transitive closure (transitive reduction) of  $G$ . Further we use the usual notation  $e = |E|$ ,  $e^* = |E^*|$ ,  $e_{\text{red}} = |E_{\text{red}}|$ ,  $n = |V|$ ,  $\gamma(v) = |\text{out}(v)|$ ,  $\gamma^*(v) = |\text{out}^*(v)|$ , and  $\gamma^{\text{red}}(v) = |\text{out}^{\text{red}}(v)|$ .

Let  $\text{out}(v) = \{w_1 < \dots < w_s\}$ ; then it is easy to see that we obtain

$$\text{out}^*(v) = \{v\} \cup \text{out}^*(w_1) \cup \text{out}^*(w_2) \cup \dots \cup \text{out}^*(w_s). \quad (1)$$

For a topologically sorted digraph  $G = (V, E)$  the observation that

$$\begin{aligned} w_i \notin \text{out}^{\text{red}}(v) &\Leftrightarrow w_i \in \text{out}^*(w_1) \cup \dots \cup \text{out}^*(w_{i-1}) \\ &\Leftrightarrow \text{out}^*(w_i) \subseteq \text{out}^*(w_1) \cup \dots \cup \text{out}^*(w_{i-1}) \end{aligned} \quad (2)$$

was made by Goralčíková and Koubek to show  $G^* = G_{\text{red}}^*$ . Through negation from (2) we get

$$\text{out}^{\text{red}}(v) = \{w_i \in \text{out}(v) \mid w_i \notin \text{out}^*(w_1) \cup \dots \cup \text{out}^*(w_{i-1})\} \quad (3)$$

and further

$$\bigcup_{w \in \text{out}(v)} \text{out}^*(w) = \bigcup_{z \in \text{out}^{\text{red}}(v)} \text{out}^*(z). \quad (4)$$

Here we give their algorithm.

**Algorithm A** (Goralčíková, Koubek [6])

*Input:*  $G = (V, E)$

*Output:*  $\text{out}^*(v)$  and  $\text{out}^{\text{red}}(v)$ ,  $\forall v \in V$

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(1)   for  $v \leftarrow n$  downto 1 (*  $V = \{1, \dots, n\}$  *)
(2)   do
(3)      $\text{out}^*(v) \leftarrow \{v\}$ ;
(4)      $\text{out}^{\text{red}}(v) \leftarrow \emptyset$ ;
(5)     for  $\forall w \in \text{out}(v)$  (* in increasing order *)
(6)     do
(7)       if  $w \notin \text{out}^*(v)$ 
(8)       then
(9)          $\text{out}^*(v) \leftarrow \text{out}^*(v) \cup \text{out}^*(w)$ ;
(10)         $\text{out}^{\text{red}}(v) \leftarrow \text{out}^{\text{red}}(v) \cup \{w\}$ ;
(11)      fi
(12)    od
(13)  od

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Inside loop (2)–(13) we use a bitvector for  $\text{out}^*(v)$ . So the test “ $w \in \text{out}^*(v)$ ” takes time  $O(1)$ . Outside loop (2)–(13) the set  $\text{out}^*(v)$  is kept as a linear list. This implies that the operation  $\text{out}^*(v) \cup \text{out}^*(w)$  has execution time  $O(|\text{out}^*(w)|) = O(n)$  and so we need total time  $O(n \cdot e_{\text{red}})$ . We will show later that the expected execution time is  $O(n^2 \cdot \log n)$  for this algorithm. In this paper we give a better method for computing  $\text{out}^*(v) \cup \text{out}^*(w)$ . Our improvement is based on an efficient data structure, the so-called chain decomposition.

## 2. The algorithm

**Definition 2.1.** Let  $G = (V, E)$  be an acyclic digraph. A partition  $Z_1, \dots, Z_k$  of  $V$  ( $Z_i \neq \emptyset$  for  $1 \leq i \leq k$  and  $Z_1 \cup \dots \cup Z_k = V$ ) is called a *chain decomposition* of  $G = (V, E)$  if and only if every  $Z_i, 1 \leq i \leq k$ , is a path in  $G$ . Because  $G$  is topologically sorted, we obtain for a path  $Z_i = \{v_1 < \dots < v_s\}$

$$(v_j, v_{j+1}) \in E \quad \forall 1 \leq j \leq k.$$

The integer  $k$  is called the *width* of the decomposition. A chain decomposition  $Z_1, \dots, Z_k$  induces the maps  $\text{id}$ ,  $\text{niv}$ , and  $\text{niv}_j$ . Let  $v \in V, A \subseteq V, i, j \in \{1, \dots, k\}$ ; then

$$\text{id}(v) = i \Leftrightarrow v \in Z_i,$$

$$\text{niv}_j(A) = \min(A \cap Z_j), \quad \text{niv}(A) = \{\text{niv}_j(A) \mid 1 \leq j \leq k\}.$$

In particular, we use the notation  $\text{niv}_j(v)$  for  $\text{niv}_j(\text{out}^*(v))$  and  $\text{niv}(v)$  for  $\text{niv}(\text{out}^*(v))$ . In this paper we have only  $A = \text{out}^*(w_1) \cup \dots \cup \text{out}^*(w_s)$  and therefore  $A \cap Z_j$  is a path in  $G$  with first node  $\text{niv}_j(A)$  (see Fig. 1).

Now we come to the question how we can use a chain decomposition to speed up the computation of the transitive closure? The critical observation is given in Theorem 2.2.

**Theorem 2.2.**  $\forall v \in V$ :

$$\text{out}^*(v) = \bigcup_{1 \leq j \leq k} \{w \in Z_j \mid w \geq \text{niv}_j(v)\}.$$

**Proof.** Let  $v \in V$  and let  $Z_1, \dots, Z_k$  be a chain decomposition. Then  $Z_1, \dots, Z_k$  are a partition of  $V$  and we have  $Z_1 \cup \dots \cup Z_k = V$ . Since  $\text{out}^*(v) \subseteq V$ , we get

$$\text{out}^*(v) = \bigcup_{1 \leq j \leq k} (\text{out}^*(v) \cap Z_j).$$

Let  $Z_j$  be a path  $v_1, \dots, v_s$  with  $v_1 < v_2 < \dots < v_s$ ,  $\text{niv}_j(v) = \min(\text{out}^*(v) \cap Z_j)$  and  $\text{out}^*(v) \cap Z_j \neq \emptyset$ . Now  $\text{niv}_j(v)$  is an element of  $\text{out}^*(v)$  and therefore the set

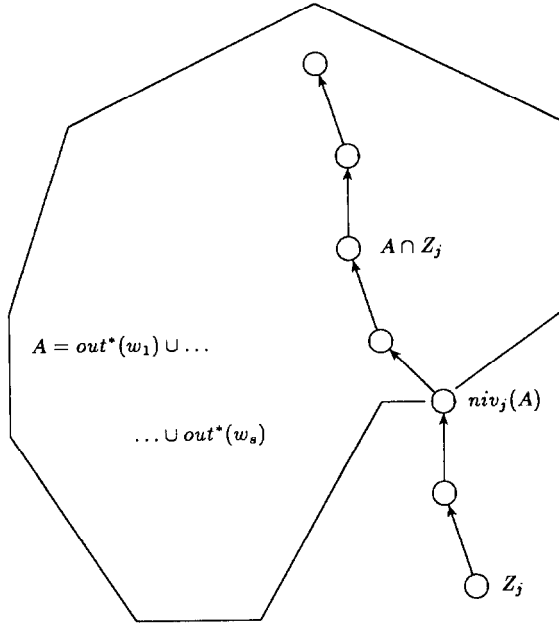


Fig. 1.

$(out^*(v) \cap Z_j)$  is a tail of  $Z_j$ , namely  $v_l, v_{l+1}, \dots, v_s$ , where  $v_l = niv_j(v)$ . Then we have

$$out^*(v) = \bigcup_{1 \leq j \leq k} (out^*(v) \cap Z_j) = \bigcup_{1 \leq j \leq k} \{w \in Z_j \mid w \geq niv_j(v)\}. \quad \square$$

This theorem shows that it is sufficient to compute  $niv_j(v)$  for all  $1 \leq j \leq k$  since the computation of the set  $\{w \in Z_j \mid w \geq niv_j(v)\}$  is trivial for a given chain  $Z_j$ . In the following we will use the convention  $\min(\emptyset) = \infty$ . Now we describe relations that we need for the efficient computation of  $niv(v)$ . Let

$$niv_j(w_1) = niv_j(out^*(w_1)), \dots, niv_j(w_s) = niv_j(out^*(w_s))$$

are given; then we infer

$$\begin{aligned} & niv_j(out^*(w_1) \cup \dots \cup out^*(w_s)) \\ &= \min((out^*(w_1) \cup \dots \cup out^*(w_s)) \cap Z_j) \\ &= \min((out^*(w_1) \cap Z_j) \cup \dots \cup (out^*(w_s) \cap Z_j)) \\ &= \min(\min(out^*(w_1) \cap Z_j), \dots, \min(out^*(w_s) \cap Z_j)) \\ &= \min(niv_j(out^*(w_1)), \dots, niv_j(out^*(w_s))) \\ &= \min(niv_j(w_1), \dots, niv_j(w_s)). \end{aligned} \tag{5}$$

Now we combine expression (5) with (1) and this leads us to Theorem 2.3.

**Theorem 2.3.**  $\forall v \in V, \forall j, 1 \leq j \leq k$ :

$$\text{niv}_j(v) = \begin{cases} \min(\{\text{niv}_j(w) \mid w \in \text{out}(v)\}) & \text{if } j \neq \text{id}(v), \\ v & \text{otherwise.} \end{cases}$$

**Proof.** Let  $\text{out}(v) = \{w_1, \dots, w_s\}$ ; then we get

$$\begin{aligned} \text{niv}_j(v) &= \text{niv}_j(\text{out}^*(v)) \\ &\stackrel{(1)}{=} \text{niv}_j(\{v\} \cup \underbrace{\text{out}^*(w_1) \cup \dots \cup \text{out}^*(w_s)}_{= \text{def } A}) \\ &= \min((\{v\} \cup A) \cap Z_j) \\ &= \min((\{v\} \cap Z_j) \cup (A \cap Z_j)) \\ &= \min(\min(\{v\} \cap Z_j), \underbrace{\min(A \cap Z_j)}_{= \text{niv}_j(A)}) \\ &\stackrel{(5)}{=} \min(\min(\{v\} \cap Z_j), \min(\text{niv}_j(w_1), \dots, \text{niv}_j(w_s))) \\ &= \begin{cases} \min(\text{niv}_j(w_1), \dots, \text{niv}_j(w_s)) & \text{if } v \notin Z_j, \\ v & \text{otherwise.} \quad \square \end{cases} \end{aligned}$$

With Theorem 2.3 it is clear how we compute  $\text{niv}(v) = \text{niv}(\text{out}^*(v))$ . In Algorithm A we replace the operation

$$\text{out}^*(v) \leftarrow \text{out}^*(v) \cup \text{out}^*(w)$$

by

$$\text{niv}(\text{out}^*(v)) \leftarrow \text{niv}(\text{out}^*(v) \cup \text{out}^*(w)).$$

With expression (5) this reduces the execution time from

$$O(|\text{out}^*(w)|) \stackrel{(L.3.5)}{=} O(n)$$

to

$$O(|\text{niv}(\text{out}^*(w))|) = O(k).$$

In general we find that  $k$  is very much smaller as  $|\text{out}^*(w)|$  (see Section 3). Now it remains the problem how we implement the test “ $w \in \text{out}^*(v)$ ” from Algorithm A. If we want to use our new data structure, we cannot realize set  $\text{out}^*(v)$  as a bitvector. But now we use an array of integers for set  $\text{niv}(\text{out}^*(v))$  and Lemma 2.4 shows this is sufficient.

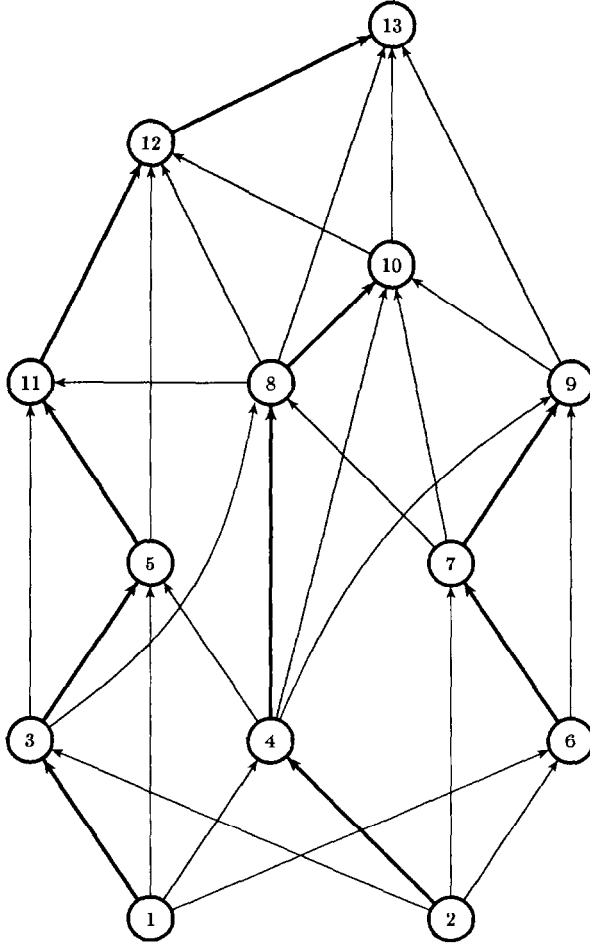
**Lemma 2.4.** Let  $v, w_1, \dots, w_s \in V$  and  $w \in Z_j$  ( $\Leftrightarrow \text{id}(w) = j$ ). Then there is

$$w \notin \text{out}^*(w_1) \cup \dots \cup \text{out}^*(w_s) \Leftrightarrow w < \min(\text{niv}_{\text{id}(w)}(w_1), \dots, \text{niv}_{\text{id}(w)}(w_s)).$$

**Proof.** With  $A = \text{out}^*(w_1) \cup \dots \cup \text{out}^*(w_s)$  we get

$$\begin{aligned}
 w \in A &\Leftrightarrow w \in A \cap Z_{\text{id}(w)} \\
 &\Leftrightarrow w \geq \min(A \cap Z_{\text{id}(w)}) \\
 &\Leftrightarrow w \geq \text{niv}_{\text{id}(w)}(A) \\
 &\stackrel{(5)}{\Leftrightarrow} w \geq \min(\text{niv}_{\text{id}(w)}(w_1), \dots, \text{niv}_{\text{id}(w)}(w_s)).
 \end{aligned}$$

Our claim is inferred through negation.  $\square$



$$Z_1 = \{ 1, 3, 5, 11, 12, 13 \}$$

$$Z_2 = \{ 2, 4, 8, 10 \}$$

$$Z_3 = \{ 6, 7, 9 \}$$

Fig. 2. An example for chain decomposition.

A chain decomposition of an acyclic graph is easily constructed in time  $O(n + e)$ . Now we give one particular algorithm. In a greedy manner we find a first path  $Z$ . We remove  $Z$  from  $G = (V, E)$  and restart the method (see Fig. 2).

Now have a look at Algorithm B. At line (4)  $V_i$  is realized as a bitvector  $S$ . Line (6) is implemented by sliding a pointer  $pt$  across bitvector  $S$ . All elements to the left of  $pt$  are not in  $V_i$ . The sets  $Z_i$  are kept as linear lists. Then loop (7)–(13) takes time  $O(\gamma(v))$  for fixed  $v$ . Hence, the total time (1)–(18) is  $O(n + e)$ . Note that the total time spent in line (6) is  $O(n)$  since  $pt$  is slid one across vector  $S$ .

**Algorithm B** (*computation of a chain decomposition*)

*Input:*  $G = (V, E)$

*Output:*  $Z_1, \dots, Z_k, id$

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(1)    $i \leftarrow 1$ 
(2)   for  $\forall v \in V$  do  $id[v] \leftarrow 0$  od
(3)    $V_i \leftarrow V$ 
(4)   while  $V_i \neq \emptyset$ 
(5)   do
(6)      $x \leftarrow \min(V_i)$ 
(7)      $Z \leftarrow \{x\}$ 
(8)     while  $\exists y \in V_i$  with  $(x, y) \in E$ 
(9)     do
(10)      let  $y$  be minimal with  $y \in V_i$  and  $(x, y) \in E$ 
(11)       $Z \leftarrow Z + \{y\}$ 
(12)       $x \leftarrow y$ 
(13)    od
(14)     $Z_i \leftarrow Z$ 
(15)     $V_{i+1} \leftarrow V_i - Z_i$ 
(16)    for  $\forall v \in Z_i$  do  $id[v] \leftarrow i$  od
(17)     $i \leftarrow i + 1$ 
(18)  od

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In the following when we speak about a chain decomposition we will mean the decomposition constructed by this algorithm and we define the width  $k = k(G)$  of the chain decomposition by

$$k \stackrel{\text{def}}{=} \max(\{s \in \mathbb{N} \mid Z_s \neq \emptyset\}).$$

Now we can compute  $niv(v)$  for all  $v \in V$ . In Algorithm C we use the linear list  $nivlist[v]$  for keeping  $niv(v)$ ;  $nivfield, id$  are arrays of integers. So we come to Algorithm C.

**Algorithm C** (*computation of  $niv(v)$* )

*Input:*  $G = (V, E), id: V \rightarrow \{1, \dots, k\}$

*Output:*  $\text{niv}(v)$ ,  $\text{out}^{\text{red}}(v)$ ,  $\forall v \in V$

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(1)   for  $s \leftarrow 1$  to  $k$  do  $\text{nivfield}[s] \leftarrow \infty$  od
(2)   for  $v \leftarrow n$  downto 1
(3)   do
(4)      $\text{out}^{\text{red}}[v] \leftarrow \emptyset$ ;  $\text{nivlist}[v] \leftarrow \emptyset$ ;
(5)     for  $\forall w \in \text{out}(v)$  (* in increasing order *)
(6)     do
(7)       if  $w < \text{nivfield}[\text{id}[w]]$ 
(8)       then
(9)          $\text{out}^{\text{red}}[v] \leftarrow \text{out}^{\text{red}}[v] \cup \{w\}$ 
(10)      for  $\forall p \in \text{nivlist}[w]$ 
(11)      do
(12)         $\text{nivfield}[\text{id}[p]] \leftarrow \min(\text{nivfield}[\text{id}[p]], p)$ 
(13)      od
(14)    fi
(15)  od
(16)   $\text{nivfield}[\text{id}[v]] \leftarrow v$ ;
(17)  for  $s \leftarrow 1$  to  $k$ 
(18)  do
(19)    if  $\text{nivfield}[s] \neq \infty$ 
(20)    then  $\text{nivlist}[v] \leftarrow \text{nivlist}[v] \cup \text{nivfield}[s]$  fi
(21)     $\text{nivfield}[s] \leftarrow \infty$ 
(22)  od
(23) od

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*Correctness* is shown by induction, starting with  $v = n$ . In particular we get, after every execution of loop (2)-(23),

$$\text{nivlist}[v] = \text{niv}(v) \quad \text{and} \quad \text{out}^{\text{red}}[v] = \text{out}^{\text{red}}(v).$$

( $v = n$ ): Since loop (5)-(15) is not executed, we obtain  $\text{nivlist}[n] = \{n\} = \text{niv}(n)$ .  
( $v < n$ ). Now the induction hypothesis is

$$\forall w, v < w \leq n: \quad \text{nivlist}[w] = \text{niv}(w).$$

There is, after line (13),

$$\forall p, p \in \text{nivlist}[w]: \quad \text{nivfield}[\text{id}[p]]_{\text{new}} = \min(\text{nivfield}[\text{id}[p]]_{\text{old}}, p).$$

With the induction hypothesis this is equivalent to

$$\forall j, 1 \leq j \leq k; \quad \text{nivfield}[j]_{\text{new}} = \min(\text{nivfield}[j]_{\text{old}}, \text{niv}_j(w)) \quad (*)$$

since

$$p = \text{niv}_{\text{id}(p)}(w) \quad \forall p \in \text{nivlist}[w] \stackrel{(\text{I.H.})}{=} \text{niv}(w).$$



We infer inductively from (\*) that before the execution of loop (6)–(15) it is valid that

$$\text{nivfield}[j] = \min(\{\text{niv}_j(z) \mid z \in \text{out}^{\text{red}}(v) \wedge z < w\}). \quad (**)$$

Consequently, line (7) is executed if and only if

$$w < \text{nivfield}[\text{id}[w]] \stackrel{(**)}{\Leftrightarrow} w < \text{niv}_{\text{id}(w)}(v) \stackrel{(\text{L.2.4})}{\Leftrightarrow} w \in \text{out}^{\text{red}}(v).$$

This shows the correct construction of  $\text{out}^{\text{red}}$ . With induction on the number of executions of loop (6)–(15) we get from (\*\*) after line (16)  $\forall j, 1 \leq j \leq k$ :

$$\text{nivfield}[j] = \begin{cases} \min(\{\text{niv}_j(w) \mid w \in \text{out}^{\text{red}}(v)\}) & \text{if } j \neq \text{id}(v), \\ v & \text{otherwise} \end{cases}$$

and this leads with Theorem 2.3 to

$$\text{nivfield}[j] = \text{niv}_j(v) \quad \forall j, 1 \leq j \leq k.$$

This ends the correctness proof of Algorithm C.

*Running time:* Outside lines (6)–(15) the cost of the algorithm is clearly  $O(e + n \cdot k)$ . One execution of the loop (10)–(13) has cost  $O(k)$  and this loop is executed only for  $(v, w) \in E_{\text{red}}$ . Hence, for Algorithm C we have total cost

$$O(e + n \cdot k) + O(e_{\text{red}} \cdot k) = O(e + e_{\text{red}} \cdot k).$$

**Theorem 2.5.** *The improved algorithm computes the transitive closure of an acyclic digraph in time  $O(e^* + e_{\text{red}} \cdot k)$ .*

**Proof.** Running time of the decomposition algorithm is  $O(n + e)$ . The computation of  $\text{niv}(v)$ ,  $\forall v \in V$  needs time  $O(e + e_{\text{red}} \cdot k)$ . From a chain decomposition of  $Z_1, \dots, Z_k$  with given  $\text{id}(v)$ ,  $\text{niv}(v)$ ,  $\forall v \in V$ , it is now trivial to compute  $E^*$  in time  $O(e^*)$  (recall Theorem 2.2). Hence, we get a total running time

$$O(n + e) + O(e + e_{\text{red}} \cdot k) + O(e^*) = O(e^* + e_{\text{red}} \cdot k). \quad \square$$

### 3. Average case

For the average case analysis we use the  $G_{n,p}$  model of a random acyclic digraph with vertex set  $\{1, \dots, n\}$  in which the possible edges  $(i, j)$ ,  $1 \leq i < j \leq n$ , occur independently with probability  $p$ ,  $0 < p < 1$ . An introduction to the theory of random graphs was given by Erdős and Spencer in [3]. By this model the size of  $k$ ,  $e_{\text{red}}$ , or  $k \cdot e_{\text{red}}$  is a random variable. Our aspiration is to obtain good upper bounds for the expected values of these random variables, especially for the product  $k \cdot e_{\text{red}}$ . Note that the latter is a product of two dependent random variables and its analysis takes a lot of time. Therefore we give the main findings first and see the proofs and other results later in their logical order. We write  $\text{Pr}(A)$  for the probability of event  $A$

and further  $\Pr(A|B)$  for the probability of  $A$  on condition of event  $B$ . Let  $X$  be a random variable; then  $E(X)$  means the expected value of  $X$ . “log” stands for the natural logarithm.

$$\text{(Lemma 3.4)} \quad E(k) \leq \frac{\log(p \cdot n)}{p} + 1,$$

$$\text{(Corollary 3.8)} \quad E(e_{\text{red}}) \leq n \cdot (\log n + 2),$$

$$\text{(Lemma 3.12)} \quad E(k \cdot e_{\text{red}}) \leq 4 \cdot \frac{\log n}{p} \cdot E(e_{\text{red}}) + 1,$$

$$\text{(Lemma 3.13)} \quad E(k \cdot e_{\text{red}}) \leq (E(k) + 1) \cdot E(e).$$

**Theorem 3.1.** *By use of the  $G_{n,p}$  model of a random digraph our Algorithm C computes the transitive closure of an acyclic digraph in expected time:*

$$O(n^2) \quad \text{for } \frac{\log^2 n}{n} \leq p < 1 \quad \text{and}$$

$$O(n^2 \cdot \log \log n) \quad \text{for } 0 < p < \frac{\log^2 n}{n}.$$

**Proof.** For  $p \geq (\log^2 n)/n$  we use Lemma 3.12 and otherwise Lemma 3.13.

*Case  $p \geq (\log^2 n)/n$ :*

$$\begin{aligned} E(e_{\text{red}}) &\leq 4 \cdot \frac{\log n}{p} \cdot E(e_{\text{red}}) + 1 \\ &\leq 4 \cdot \frac{\log n}{(\log^2 n)/n} \cdot n \cdot (\log n + 2) + 1 \\ &\leq O\left(n^2 \cdot \frac{\log^2 n}{\log^2 n}\right) \leq O(n^2). \end{aligned}$$

*Case  $p \leq (\log^2 n)/n$ :*

$$\begin{aligned} E(k \cdot e_{\text{red}}) &\leq (E(k) + 1) \cdot E(e) \leq O\left(\frac{\log(p \cdot n)}{p} \cdot p \cdot n^2\right) \\ &\leq O\left(n^2 \cdot \left(\log\left(\frac{\log^2 n}{n} \cdot n\right)\right)\right) \leq O(n^2 \cdot (2 \cdot \log(\log n))) \\ &\leq O(n^2 \cdot \log \log n). \quad \square \end{aligned}$$

In the random graph  $G_{n,p}$  we describe the size of the chain decomposition  $k = k(G_{n,p}) = k_n$  and the size of the transitive closure of the first vertex  $\gamma^*(1) = \gamma^*(1)(G_{n,p}) = \gamma_n^*$  as a Markov chain with discrete time  $t = n$  [4]. First we consider the behaviour of  $k$  at the point of transition from  $n$  to  $n+1$ . If  $k_n = l$ ,  $1 \leq l \leq n$ , then

we get  $l \leq k_{n+1} \leq l+1$  since the width of the chain decomposition either increases by one or does not change. When it increases, i.e., when the vertex  $n+1$  is a new chain  $Z_{l+1} = \{n+1\}$ , then there is no chain  $Z_j$ ,  $1 \leq j \leq l$ , which can be extended to  $n+1$ . More formally,

$$k_{n+1} = l+1 \Leftrightarrow \forall j, 1 \leq j \leq l: (\max(Z_j), n+1) \notin E.$$

This leads to the transition probability

$$\begin{aligned} \Pr(k_{n+1} = l+1 | k_n = l) &= \Pr(\forall j, 1 \leq j \leq l: (\max(Z_j), n+1) \notin E) \\ &= (1-p)^l. \end{aligned} \tag{6}$$

Note that the  $l$  possible edges are independent. This implies

$$\Pr(k_{n+1} = l | k_n = l) = 1 - \Pr(k_{n+1} = l+1 | k_n = l) = 1 - (1-p)^l. \tag{7}$$

**Remark.** Let  $k_{n+1} = k_n$ ; then the current chain decomposition  $Z_1, \dots, Z_l$  can be extended to the new node  $n+1$ , and Algorithm B links the vertex  $n+1$  to the chain with lowest index. Let  $(\max(Z_i), n+1) \in E$  and

$$(\max(Z_1), n+1) \notin E, \dots, (\max(Z_{i-1}), n+1) \notin E;$$

Then Algorithm B connects  $n+1$  with chain  $Z_i$ . It is easy to see that this special choice has no influence on the future growth of the chain decomposition.

In the same way as  $k$  we treat  $\gamma_n^* = |\text{out}_n^*(1)|$ :

$$\begin{aligned} \Pr(\gamma_{n+1}^* = l+1 | \gamma_n^* = l) &= \Pr(\exists w \in \text{out}_n^*(1): (w, n+1) \in E) \\ &= 1 - \Pr(\forall w \in \text{out}_n^*(1): (w, n+1) \notin E) \\ &= 1 - (1-p)^l \end{aligned} \tag{8}$$

and

$$\Pr(\gamma_{n+1}^*(1) = l | \gamma_n^* = l) = 1 - \Pr(\gamma_{n+1}^* = l+1 | \gamma_n^* = l) = (1-p)^l. \tag{9}$$

With the additional notation  $k_{n,l} =_{\text{def}} \Pr(k_n = l)$  and  $\gamma_{n,l}^* =_{\text{def}} \Pr(\gamma_n^* = l)$  we find that  $k_{n,l}, \gamma_{n,l}^*$  satisfy the following recurrence.

**Lemma 3.2.**  $\forall n \in \mathbb{N}, \forall l, 1 \leq l \leq n,$

$$k_{1,1} = 1, \quad k_{n,l} = (1 - (1-p)^l) \cdot k_{n-1,l} + (1-p)^{l-1} \cdot k_{n-1,l-1} \tag{10}$$

and

$$\gamma_{1,1}^* = 1, \quad \gamma_{n,l}^* = (1-p)^l \cdot \gamma_{n-1,l}^* + (1 - (1-p)^{l-1}) \cdot \gamma_{n-1,l-1}^*. \tag{11}$$

**Proof.** Clear by the preceding discussion.  $\square$

This description shows that  $k_n$  ( $\gamma_n^*$  respectively) is a discrete-time, pure-birth process (see [4]). By a discrete-time, pure-birth process we understand a sequence of random variables  $X_t, t \in \mathbb{N}$ , assuming the states  $l = 1, 2, 3, \dots$  with corresponding probabilities  $P_{t,l}$  and a sequence of transition probabilities  $\lambda_l, 0 \leq \lambda_l \leq 1$  and  $l \in \mathbb{N}$ , so that

$$P_{1,1} = 1 \quad \text{and} \quad P_{t,l} = (1 - \lambda_l) \cdot P_{t-1,l} + \lambda_{l-1} \cdot P_{t-1,l-1},$$

i.e., the process starts at epoch 1 from state 1; direct transitions from a state  $l$  are only possible to  $l + 1$ ; these transitions have probability  $\lambda_l$ . To provide for an easy treatment we first give a very useful identity of this kind of birth process.

Let  $\varphi$  be a real function with

$$\varphi(l) = \sum_{j=1}^{l-1} \frac{1}{\lambda_j} \quad \forall l \in \mathbb{N}.$$

Then we state the following lemma.

**Lemma 3.3.**  $E(\varphi(X_t)) = t - 1, \forall t \in \mathbb{N}.$

**Proof.** We use induction on  $t$ .

$$(t = 1): \quad E(X_1) = \sum_{l=1}^1 \varphi(l) \cdot P_{1,1} = 0 \cdot 1 = 0.$$

$(t \geq 2)$ : Our induction hypothesis (I.H.) is  $E(\varphi(X_{t-1})) = t - 2$ . Then we have

$$\begin{aligned} E(\varphi(X_t)) &= \sum_{l=1}^t \varphi(l) \cdot P_{t,l} \\ &= \sum_{l=1}^t \varphi(l) \cdot ((1 - \lambda_l) \cdot P_{t-1,l} + \lambda_{l-1} \cdot P_{t-1,l-1}) \\ &= \underbrace{\sum_{l=1}^t \varphi(l) \cdot P_{t-1,l}}_{= E(\varphi(X_{t-1})) = t-2} + \sum_{l=1}^t \varphi(l) \cdot \lambda_{l-1} \cdot P_{t-1,l-1} - \sum_{l=1}^t \varphi(l) \cdot \lambda_l \cdot P_{t-1,l} \\ &= t - 2 + \underbrace{\sum_{l=1}^t \varphi(l) \cdot \lambda_{l-1} \cdot P_{t-1,l-1} - \sum_{l=1}^t \varphi(l) \cdot \lambda_l \cdot P_{t-1,l}}_{= \sum_{l=0}^{t-1} \varphi(l+1) \cdot \lambda_l \cdot P_{t-1,l}} \end{aligned}$$

(note that  $P_{t-1,0} = 0 = P_{t-1,t}$ )

$$= t - 2 + \sum_{l=1}^{t-1} \varphi(l+1) \cdot \lambda_l \cdot P_{t-1,l} - \sum_{l=1}^{t-1} \varphi(l) \cdot \lambda_l \cdot P_{t-1,l}$$

(since  $\varphi(l+1) = \varphi(l) + 1/\lambda_l$ )

$$\begin{aligned} &= t - 2 + \sum_{l=1}^{t-1} \underbrace{(\varphi(l) \cdot \lambda_l \cdot P_{t-1,l} - \varphi(l) \cdot \lambda_l \cdot P_{t-1,l})}_{=0} + \frac{1}{\lambda_l} \cdot P_{t-1,l} \\ &= t - 2 + \underbrace{\sum_{l=1}^{t-1} \frac{1}{\lambda_l} \cdot \lambda_l \cdot P_{t-1,l}}_{=1} \\ &= t - 2 + 1 = t - 1. \quad \square \end{aligned}$$

By use of Lemma 3.3 a simple deduction leads to Lemmas 3.4 and 3.5.

**Lemma 3.4**

$$E(k(G_{n,p})) \leq \frac{\log(p \cdot (n-1) + 1)}{-\log(1-p)} + 1.$$

**Proof.** With expression (6) there is

$$\lambda_l(k) = \Pr(k_n = l + 1 \mid k_{n-1} = l) = (1-p)^l.$$

By a simple application of the summation formula for geometric series (see Appendix A) we get

$$\varphi(l) = \sum_{j=1}^{l-1} \frac{1}{\lambda_j(k)} = \sum_{j=1}^{l-1} \frac{1}{(1-p)^j} = \frac{1}{p} \cdot \left( \left( \frac{1}{1-p} \right)^{l-1} - 1 \right) \quad \forall l, 1 \leq l \leq n.$$

Of course,  $\varphi(l)$  is an exponential function and so  $\varphi(l)$  is convex, i.e.,

$$\forall x_1, \dots, x_n \in \mathbb{R}: \quad \alpha_1, \dots, \alpha_n \in \mathbb{R}^+ \text{ with } \sum_{l=1}^n \alpha_l = 1.$$

We get

$$\varphi(\alpha_1 \cdot x_1 + \dots + \alpha_n \cdot x_n) \leq \alpha_1 \cdot \varphi(x_1) + \dots + \alpha_n \cdot \varphi(x_n).$$

This implies with  $\alpha_l = k_{n,l}$  and  $x_l = l, 1 \leq l \leq n$ , Jensen's inequality

$$\varphi(E(k_n)) \stackrel{\text{def}}{=} \varphi\left(\sum_{l=1}^n l \cdot k_{n,l}\right) \leq \sum_{l=1}^n \varphi(l) \cdot k_{n,l} \stackrel{\text{def}}{=} E(\varphi(k_n)).$$

Now we apply Lemma 3.3 and the inverse function  $\varphi^{-1}$  to  $\varphi$  given by

$$\varphi^{-1}(x) = \frac{\log(p \cdot x + 1)}{-\log(1-p)} + 1.$$

We find

$$\begin{aligned} \varphi(E(k_n)) &\leq E(\varphi(k_n)) \stackrel{(L.3.3)}{=} n-1 \\ \Leftrightarrow \varphi(E(k_n)) &\leq n-1 \\ \Leftrightarrow \varphi^{-1}(\varphi(E(k_n))) &\leq \varphi^{-1}(n-1) \\ \Leftrightarrow E(k_n) &\leq \frac{\log(p \cdot (n-1) + 1)}{-\log(1-p)} + 1. \quad \square \end{aligned}$$

**Lemma 3.5**

$$E(|\text{out}^*(1)(G_{n,p})|) \geq n+1 - \frac{|\log p|+1}{p}.$$

**Proof.** By term (7) we have  $\lambda_l = \lambda_l(\gamma^*) = 1 - (1-p)^l$  and, consequently, for  $l \in \mathbb{N}$ ,

$$\begin{aligned} \varphi(l) &= \sum_{j=1}^{l-1} \frac{1}{\lambda_j} = \frac{1}{p} + \sum_{j=2}^{l-1} \frac{1}{1 - (1-p)^j} \\ &\leq \frac{1}{p} + \sum_{j=2}^{l-1} \int_{j-1}^j \frac{dx}{1 - (1-p)^x} = \frac{1}{p} + \int_1^{l-1} \frac{dx}{1 - (1-p)^x} \end{aligned}$$

(see [1, p. 87])

$$\begin{aligned} &= \frac{1}{p} + (l-1) - \underbrace{\frac{\log(1 - (1-p)^{l-1})}{\log(1-p)}}_{\geq 0} - 1 + \frac{\log(1 - (1-p))}{\log(1-p)} \\ &\leq (l-2) + \frac{|\log p|+1}{p}. \end{aligned}$$

And, consequently,  $\varphi(l) \leq (l-2) + (|\log p|+1)/p =_{\text{def}} \phi(l)$ . This implies

$$E(\varphi(\gamma_n^*)) = \sum_{l=1}^n \varphi(l) \cdot \gamma_{n,l}^* \leq \sum_{l=1}^n \phi(l) \cdot \gamma_{n,l}^* = E(\phi(\gamma_n^*))$$

also by Lemma 3.3 and, by the fact that  $\phi$  is a straight line,

$$\begin{aligned} n-1 &= E(\varphi(\gamma_n^*)) \leq E(\phi(\gamma_n^*)) = \phi(E(\gamma_n^*)) = E(\gamma_n^*) + \frac{|\log p|+1}{p} - 2 \\ \Leftrightarrow E(\gamma_n^*) &\geq n+1 - \frac{|\log p|+1}{p}. \quad \square \end{aligned}$$

Using the notation  $\gamma_n^{\text{red}} = \gamma^{\text{red}}(1)(G_{n,p})$  for the size of the transitive reduction of the first vertex, we show that the lower bound for  $E(\gamma_n^*)$  gives an upper bound for the expected value of the reduction  $E(\gamma_n^{\text{red}})$ . This can be reached by comparing  $\Delta E(\gamma_n^*) = \Pr((1, n) \in E^*)$  with  $\Delta E(\gamma_n^{\text{red}}) = \Pr((1, n) \in E_{\text{red}})$ .

**Lemma 3.6**

$$E(\gamma_n^{\text{red}}) = \frac{p}{1-p} \cdot (n - E(\gamma_n^*)).$$

**Proof.** Let  $A$  be the event “ $(1, n) \in E^*$ ” with

$$\Pr((1, n) \in E^*) = \Pr(\exists w \in \text{out}_{n-1}^*(1) : (w, n) \in E).$$

Then, by splitting  $A$  on  $\gamma_{n-1}^*$ , we have

$$\begin{aligned} \Pr((1, n) \in E^*) &= \sum_{l=1}^{n-1} \underbrace{\Pr(A \mid \gamma_{n-1}^* = l)}_{=(1-(1-p)^l)} \cdot \underbrace{\Pr(\gamma_{n-1}^* = l)}_{=\gamma_{n-1,l}^*} \\ &= \sum_{l=1}^{n-1} (1 - (1-p)^l) \cdot \gamma_{n-1,l}^*. \end{aligned} \tag{12}$$

Let  $n \geq 2$ ; then we find

$$\begin{aligned} \Pr((1, n) \in E_{\text{red}}) &= \Pr((1, n) \in E) \cdot \Pr(\forall w \in \text{out}_{n-1}^*(1), w \neq 1 : (w, n) \notin E) \\ &= p \cdot \sum_{l=1}^{n-1} (1-p)^{l-1} \cdot \gamma_{n-1,l}^* \end{aligned}$$

(splitting by  $\gamma_{n-1}^*$ )

$$\begin{aligned} &= \frac{p}{1-p} \cdot \sum_{l=1}^{n-1} (1 - 1 + (1-p)^l) \cdot \gamma_{n-1,l}^* \\ &= \frac{p}{1-p} \cdot \left( 1 - \sum_{l=1}^{n-1} (1 - (1-p)^l) \cdot \gamma_{n-1,l}^* \right) \end{aligned}$$

Thus we arrive at

$$\Pr((1, n) \in E_{\text{red}}) \stackrel{(12)}{=} \frac{p}{1-p} \cdot (1 - \Pr((1, n) \in E^*)) \tag{13}$$

Now the proof is easily completed:

$$\begin{aligned} E(\gamma_n^{\text{red}}) &= \sum_{j=2}^n E((1, j) \in E_{\text{red}}) = \sum_{j=2}^n \Pr((1, j) \in E_{\text{red}}) \\ &\stackrel{(13)}{=} \sum_{j=2}^n \frac{p}{1-p} \cdot (1 - \Pr((1, j) \in E^*)) \\ &= \frac{p}{1-p} \cdot \left( n - 1 - \sum_{j=2}^n \Pr((1, j) \in E^*) \right) \\ &= \frac{p}{1-p} \cdot (n - E(\gamma_n^*)). \quad \square \end{aligned}$$

**Corollary 3.7.**  $E(e_{\text{red}}) \leq n \cdot (|\log p| + 2)$ .

**Proof**

$$E(e_{\text{red}}) = E(\gamma^{\text{red}}(1)) + \dots + E(\gamma^{\text{red}}(n)) \leq n \cdot E(\gamma^{\text{red}}(1)) = n \cdot E(\gamma_n^{\text{red}}).$$

By Lemma 3.6,

$$E(e^{\text{red}}) = n \cdot \frac{p}{1-p} \cdot (n - E(\gamma_n^*))$$

(by Lemma 5)

$$\begin{aligned} &\leq n \cdot \frac{p}{1-p} \cdot \left( n - \left( n + 1 - \frac{|\log p| + 1}{p} \right) \right) \\ &\leq n \cdot \left( \frac{|\log p|}{1-p} + 1 \right) \end{aligned}$$

(by Taylor)

$$\begin{aligned} &\leq n \cdot \left( \frac{(1-p) + \frac{1}{2}(1-p)^2 + \frac{1}{3}(1-p)^3 + \dots}{1-p} + 1 \right) \\ &\leq n \cdot (|\log p| + 2). \quad \square \end{aligned}$$

**Corollary 3.8.** *The expected running time of Algorithm B, according to Goralčíková and Koubek [6], is*

$$O(\min(n^2 \cdot (|\log p| + 2), p \cdot n^3)) = O(n^2 \cdot \log n).$$

**Proof.** We have, by the preceding discussion,

*Case*  $p > (\log n)/n$ :

$$E(n \cdot e_{\text{red}}) \leq n^2 \cdot (|\log p| + 2) \leq O(n^2 \cdot \log n).$$

*Case*  $p \leq (\log n)/n$ :

$$E(n \cdot e_{\text{red}}) \leq E(n \cdot e) \leq p \cdot n^3 \leq O(n^2 \cdot \log n). \quad \square$$

In the remainder of this section we develop upper bounds for  $E(k \cdot e_{\text{red}})$ . Recall that  $k \cdot e_{\text{red}}$  is a product of two dependent random variables. In Lemma 3.12 we will understand  $e_{\text{red}}$  as the dependent variable, i.e., we use the interpretation

$$E(k \cdot e_{\text{red}}) = E(k) \cdot E(e_{\text{red}} | k).$$

This analysis is prepared in the next three lemmas. First we determine the expected size of chain  $Z = Z(G_{n,p})$  constructed in Algorithm B. The repeated application of Lemma 3.9 shows that

$$E(|V_i|) = (1-p)^{i-1} \cdot n - h(\Pr(k(G_{n,p}) \geq i)).$$

By simple transformations we can see that function  $f(i) = \Pr(k(G_{n,p}) \geq i)$  decreases exponentially.



Now we analyse the construction of the chain decomposition (Algorithm B) more exactly. The algorithm iteratively constructs a chain  $Z = Z(G)$ , deletes  $Z$  from  $G$  and starts again with graph  $G' = G - Z$ . Retain the notations:

$$G_1 = (V_1, E) = G_{n,p}, \quad G_i = (V_i, E),$$

$$V_{i+1} = V_i - Z(G_i) = V_i - Z_i.$$

For the chain  $Z = Z(G = (V, E))$  there is

- (1)  $\min(V) \in Z$ , and
- (2)  $v \in Z \Rightarrow \min(\text{out}(v)) \in Z$ .

Now we are going to compute the expected size of  $Z(G_{n,p})$  in our random graph model.

**Lemma 3.9.**  $E(|Z(G_{n,p})|) = 1 + p \cdot (|V| - 1)$ .

**Proof.** Let  $T \subseteq V = \{1, \dots, n\}$  with  $T = \{1 = t_1 < t_2 < \dots < t_s\}$ . Then  $T = Z(G)$  if and only if

$$(1, h) \notin E, \forall 1 < h < t_2 \wedge (1, t_2) \in E \wedge (t_2, h) \notin E, \forall t_2 < h < t_3 \wedge \dots \\ \wedge (t_s, h) \notin E, \forall t_s < h \leq n.$$

Since all these events are independent, we get

$$\Pr(T = Z(G)) = p^{s-1} \cdot (1-p)^{(n-1)-(s-1)} = p^{|T|-1} \cdot (1-p)^{n-|T|}.$$

For  $T_1, T_2 \subseteq V$ , we obtain

$$T_1 = Z(G) \wedge T_2 = Z(G) \Rightarrow T_1 = T_2$$

such that we have  $\forall T_1, T_2 \subseteq V, T_1 \neq T_2$ :

$$\Pr(T_1 = Z \vee T_2 = Z) = \Pr(T_1 = Z) + \Pr(T_2 = Z)$$

and further

$$\Pr(|Z(G)| = l + 1) = \binom{n-1}{l} \cdot p^l \cdot (1-p)^{(n-1)-l}.$$

But this is a binomial distribution (see [4, Vol. 1, p. 146]), which implies

$$E(|Z(G)|) = 1 + p \cdot (n-1). \quad \square$$

Since removal of  $Z_1, \dots, Z_i$  turns a random graph into a random graph, we get, by Lemma 3.9,  $E(|Z_i|) = 1 + p \cdot (j-1)$  on the condition that  $|V_i| = j$ . This implies the following lemma.

**Lemma 3.10**

$$E(|V_i|) = (1-p)^{i-1} \cdot n - \sum_{j=1}^{i-1} \Pr(k(G) \geq i-j) \cdot (1-p)^j.$$

**Proof.** We first show by induction on  $i$  that

$$\begin{aligned}
 E(|V_i|) &= (1-p)^{i-1} \cdot n - \sum_{j=1}^{i-1} (1-p)^j + \sum_{j=1}^{i-1} \Pr(|V_j|=0) \cdot (1-p)^{i-j}. \\
 (i=1): E(|V_1|) &= |V| = n = p^0 \cdot n - 0. \\
 (i \geq 2): E(|V_{i+1}|) &= E(|V_i|) - E(|Z_i|) \\
 &= E(|V_i|) - \left( \sum_{j=0}^n E(|Z_i| \mid |V_i|=j) \cdot \Pr(|V_i|=j) \right) \\
 &\stackrel{(L.3.9)}{=} E(|V_i|) - \sum_{j=1}^n (1-p \cdot (j-1)) \cdot \Pr(|V_i|=j) \\
 &= E(|V_i|) - \left( \underbrace{\sum_{j=0}^n (1-p \cdot (j-1)) \cdot \Pr(|V_i|=j)}_{= 1-p + p \cdot \underbrace{\sum_{j=0}^n j \cdot \Pr(|V_i|=j)}_{= E(|V_i|)}} - (1-p) \cdot \Pr(|V_i|=0) \right) \\
 &= E(|V_i|) - (p \cdot E(|V_i|) + (1-p) - (1-p) \cdot \Pr(|V_i|=0)) \\
 &= (1-p) \cdot E(|V_i|) - (1-p) + (1-p) \cdot \Pr(|V_i|=0) \\
 &\stackrel{(I.H.)}{=} (1-p) \left( (1-p)^{i-1} \cdot n - \sum_{j=1}^{i-1} (1-p)^j + \sum_{j=1}^{i-1} \Pr(|V_j|=0) (1-p)^{i-j} \right) \\
 &\quad - (1-p) + (1-p) \cdot \Pr(|V_i|=0) \\
 &= (1-p)^i \cdot n - \sum_{j=1}^i (1-p)^j + \sum_{j=1}^i \Pr(|V_j|=0) \cdot (1-p)^{i+1-j}.
 \end{aligned}$$

And this ends the induction proof of our first statement. This expression for  $E(|V_i|)$  is equivalent to

$$\begin{aligned}
 E(|V_i|) &= (1-p)^{i-1} \cdot n - \sum_{j=1}^{i-1} (1-p)^j + \underbrace{\sum_{j=1}^{i-1} \Pr(|V_j|=0) \cdot (1-p)^{i-j}}_{= \sum_{j=1}^i \Pr(|V_{i-j}|=0) \cdot (1-p)^j} \\
 &= (1-p)^{i-1} \cdot n - \sum_{j=1}^{i-1} \underbrace{(1 - \Pr(|V_{i-j}|=0))}_{= \Pr(|V_{i-j}| \neq 0)} \cdot (1-p)^j \\
 &\quad = \Pr(k(G) \geq i-j) \\
 E(|V_i|) &= (1-p)^{i-1} \cdot n - \sum_{j=1}^{i-1} \Pr(k(G) \geq i-j) \cdot (1-p)^j. \quad \square
 \end{aligned}$$

This last lemma allows us to give an upper bound for  $\Pr(k(G_{n,p}) \geq i)$ .

**Lemma 3.11.**  $\forall 1 \leq i \leq n$ :

$$\Pr(k(G_{n,p}) \geq i) \leq (1-p)^{i-1} \cdot n.$$

**Proof.** Since  $E(|V_i|) \geq 0$ , we infer from Lemma 3.10

$$\sum_{j=1}^{i-1} \Pr(k(G) \geq i-j) \cdot (1-p)^j \leq (1-p)^{i-1} \cdot n.$$

Now, note that there is

$$\Pr(k(G) \geq i-1) \leq \Pr(k(G) \geq i-2) \leq \dots \leq \Pr(k(G) \geq 1) = 1.$$

This implies

$$\begin{aligned} \Pr(k(G) \geq i-1) \cdot \sum_{j=1}^{i-1} (1-p)^j &\leq (1-p)^{i-1} \cdot n \\ \Rightarrow \Pr(k(G) \geq i-1) \cdot (1-p) \cdot \left(\frac{1-(1-p)^i}{1-(1-p)}\right) &\leq (1-p)^{i-1} \cdot n \\ \Rightarrow \Pr(k(G) \geq i-1) &\leq (1-p)^{i-2} \cdot n. \quad \square \end{aligned}$$

Lemma 3.11 leads to the first upper bound for  $E(k \cdot e_{\text{red}})$ .

**Lemma 3.12.** Let  $k_0 =_{\text{def}} \lceil -(4 \cdot \log n) / \log(1-p) \rceil$ . Then

$$E(k \cdot e_{\text{red}}) \leq k_0 \cdot E(e_{\text{red}}) + 1.$$

**Proof.** Let  $k_0 = \lceil -(4 \cdot \log n) / \log(1-p) \rceil$ ; then we get by, Lemma 3.11,

$$\begin{aligned} \Pr(k(G) \geq k_0 + 1) &\leq (1-p)^{k_0} \cdot n \\ &\leq (1-p)^{- (4 \cdot \log n) / \log(1-p)} \cdot n \\ &\leq \frac{n}{\exp(4 \cdot \log n)} \leq \frac{1}{n^3}. \end{aligned} \tag{14}$$

This implies for our product  $k \cdot e_{\text{red}}$ :

$$\begin{aligned} E(k \cdot e_{\text{red}}) &= \sum_{l=1}^n l \cdot E(e_{\text{red}} | k(G) = l) \cdot \Pr(k(G) = l) \\ &= \sum_{l=1}^{k_0} l \cdot E(e_{\text{red}} | k(G) = l) \cdot \Pr(k(G) = l) \\ &\quad + \sum_{l=k_0+1}^n l \cdot E(e_{\text{red}} | k(G) = l) \cdot \Pr(k(G) = l) \\ &\leq k_0 \cdot \underbrace{\sum_{l=1}^{k_0} E(e_{\text{red}} | k(G) = l) \cdot \Pr(k(G) = l)}_{\leq E(e_{\text{red}})} + n^3 \cdot \underbrace{\sum_{l=k_0+1}^n \Pr(k(G) = l)}_{= \Pr(k(G) \geq k_0+1)} \end{aligned}$$

(by (14))

$$\leq k_0 \cdot E(e_{\text{red}}) + n^3 \cdot \frac{1}{n^3}. \quad \square$$

We need a better bound for small  $p$ 's.

**Lemma 3.13.**  $E(k \cdot e_{\text{red}}) \leq n \cdot E(|\text{out}(1)|) \cdot (E(k) + 1)$ .

**Proof.** First we need an upper bound for  $E(k(G) | \gamma(1) = l)$ .

We claim

$$E(k(G_{n,p}) | \gamma(1) = l) \leq 1 + E(k(G_{n,p})). \tag{15}$$

( $l = 0$ ): Then  $Z_1(G_{n,p}) = \{1\}$  and  $G - \{1\}$  is a random graph with  $n - 1$  vertices. So we get

$$E(k(G_{n,p}) | \gamma(1) = 0) = 1 + E(k(G_{n-1,p})) \leq 1 + E(k(G_{n,p})).$$

( $l \geq 1$ ): Then  $\gamma(1) \geq 1$  implies  $|Z_1(G_{n,p})| \geq 2$  and  $G - Z_1(G)$  is a random graph with  $\leq n - 2$  vertices. We have

$$E(k(G_{n,p}) | \gamma(1) = l) \leq 1 + E(k(G_{n-2,p})) \leq 1 + E(k(G_{n,p})).$$

This implies for the product

$$E(k \cdot e_{\text{red}}) \leq E(k \cdot e) = E(k \cdot \gamma(1)) + \dots + E(k \cdot \gamma(n)) \leq n \cdot E(k \cdot \gamma(1))$$

which implies

$$E(k \cdot e_{\text{red}}) \leq n \cdot \sum_{l=0}^{n-1} l \cdot E(k(G) | \gamma(1) = l) \cdot \Pr(\gamma(1) = l)$$

(by (15))

$$\leq n \cdot \sum_{l=0}^{n-1} l \cdot (E(k(G)) + 1) \cdot \Pr(\gamma(1) = l)$$

$$\leq n \cdot (E(k(G)) + 1) \cdot E(\gamma(1)). \quad \square$$

#### 4. Conclusion

We presented an improved algorithm for computing the transitive closure of an acyclic digraph with running time  $O(k \cdot e_{\text{red}})$ , where  $e_{\text{red}}$  is the number of edges in the transitive reduction and  $k$  is the width of the chain decomposition, a partition of  $V$  into distinct paths. To analyse the expected values of  $k$ ,  $e_{\text{red}}$ ,  $e^*$ ,  $k \cdot e_{\text{red}}$  we used the  $G_{n,p}$  model of a random graph. We found

$$\begin{aligned} \Pr(k(G_{n,p}) \geq i) &\leq (1-p)^{i-1} \cdot n \quad \forall 1 \leq i \leq n, \\ E(k) &= O\left(\frac{\log(p \cdot n)}{p}\right), \quad E(\gamma^*) = \Omega\left(n - \frac{\log p}{p}\right), \\ E(\gamma_n^{\text{red}}) &= \frac{p}{1-p} \cdot (n - E(\gamma_n^*)), \\ E(e_{\text{red}}) &= O(\min(n \cdot |\log p|, p \cdot n^2)) = O(n \cdot \log n), \end{aligned}$$

$$E(k \cdot e_{\text{red}}) = \begin{cases} O(n^2) & \text{for } \log^2 n / n \leq p < 1, \\ O(n^2 \cdot \log \log n) & \text{otherwise.} \end{cases}$$

Our data structure for representing the transitive closure, namely  $\text{niv}(v)$  for a vertex  $v$ , the map  $\text{id}(v)$  and the chain decomposition  $Z_1, \dots, Z_k$  only used space  $O(n \cdot k)$  in contrast to  $O(e^*)$  of previous methods. Nevertheless we can execute the test “ $w \in \text{out}^*(v)$ ” in time  $O(1)$ . Moreover, with this data structure it is easy to compute, e.g.,  $\text{out}^*(v) \cap \text{out}^*(v)$  in time  $O(k)$  and in general  $k$  is very much smaller than  $|\text{out}^*(w)| + |\text{out}^*(v)|$

**5. Further remarks**

We think that the following questions are interesting:

- $E(k \cdot e_{\text{red}}) \leq E(k) \cdot E(e_{\text{red}})$ ? Conjecture: Yes.
- Do simple limit theorems exist for the  $k$ -, or  $\gamma^*$ -distribution?

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**A. Appendix**

**Claim A.1**

$$\sum_{j=1}^{l-1} \frac{1}{(1-p)^j} = \frac{1}{p} \cdot \left( \left( \frac{1}{1-p} \right)^{l-1} - 1 \right) \quad \forall l \in \mathbb{N}.$$

**Proof**

$$\begin{aligned} \sum_{j=1}^{l-1} \frac{1}{(1-p)^j} &= \sum_{j=1}^{l-1} \left( \frac{1}{1-p} \right)^j = \left( \frac{1}{1-p} \right) \cdot \left( \sum_{j=0}^{l-2} \left( \frac{1}{1-p} \right)^j \right) \\ &= \left( \frac{1}{1-p} \right) \cdot \left( \frac{1 - (1-p)^{-(l-1)}}{1 - (1-p)^{-1}} \right) = \frac{1}{p} \cdot \left( \left( \frac{1}{1-p} \right)^{l-1} - 1 \right). \quad \square \end{aligned}$$

**Claim A.2**

$$\varphi(x) = \frac{1}{p} \cdot \left( \left( \frac{1}{1-p} \right)^{x-1} - 1 \right) \Rightarrow \varphi^{-1}(x) = \frac{\log(p \cdot x - 1)}{-\log(1-p)} + 1.$$

**Proof**

$$\begin{aligned}
\varphi(\varphi^{-1}(x)) &= \frac{1}{p} \cdot \left( \left( \frac{1}{1-p} \right)^{(\log(p \cdot x + 1) / (-\log(1-p))) + 1} - 1 \right) \\
&= \frac{1}{p} \cdot \left( \left( \frac{1}{1-p} \right)^{\log(p \cdot x + 1) / (\log 1 - \log(1-p))} - 1 \right) \\
&= \frac{1}{p} \cdot \left( \left( \left( \frac{1}{1-p} \right)^{1 / -\log(1-p)} \right)^{\log(p \cdot x + 1)} - 1 \right) \\
&= \frac{1}{p} \cdot (\exp(\log(p \cdot x + 1)) - 1) \\
&= \frac{1}{p} \cdot (p \cdot x + 1 - 1) = x. \quad \square
\end{aligned}$$

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