

Three Hundred Million Points Suffice

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Communicated by the Managing Editors

Received February 3, 1987

There is a graph G with 300,000,000 vertices and no clique on four points, such that if its edges are two colored these must be a monochromatic triangle. © 1988 Academic Press, Inc.

HISTORY AND SUMMARY

In the late 1960s Paul Erdős asked what graphs G , other than K_6 , had the property that $G \rightarrow (K_3)$. We use the Rado arrow notation: $G \rightarrow (H)$ is the statement that if the edges of G are two colored there exists a monochromatic H and, more generally, $G \rightarrow (H)_r$ is the statement that if the edges of G are r -colored there exists a monochromatic H . In particular, Erdős asked if there is a graph G satisfying

$$\begin{aligned} G \rightarrow (K_3) \\ \omega(G) = 3. \end{aligned} \quad (*)$$

A proof of the existence of such a G was first given by Jon Folkman [2]. This supremely ingenious proof had two drawbacks. First, the graph G given was extremely large. Second, the proof did not generalize to give for all r a graph G with $\omega(G) = 3$ and $G \rightarrow (K_3)_r$. At the Combinatorial Conference in Kesthely, Hungary 1973 this problem was given to the Czechoslovakian mathematician Jarik Nešetřil and his young student Vojtech Rödl. They [4] found a completely different argument that for all r graphs G exist with $\omega(G) = 3$ and $G \rightarrow (K_3)_r$. Those of us at that meeting (see [5] for an anecdotal account) recall the sense of excitement accompanying that discovery and I feel it played a critical role in the development of modern Ramsey Theory. The graphs given by the Nešetřil–Rödl methods were still extremely large and Erdős offered a reward for the discovery of a G satisfying (*) having less than 10^{10} vertices. Here we claim this reward.

The method used has been known for several years to Szemerédi, Nešetřil, Rödl, Frankl, and others. Frankl and Rödl [3] calculated that a graph G satisfying (*) with roughly 7×10^{11} vertices exists. Our note may be considered a case study in the application of asymptotic methods to give precise bounds. The method is extremely case specific. It does not give, for example, graphs G of moderate size satisfying $\omega(G)=3$ and $G \rightarrow (K_3)_3$. This remains an intriguing open problem.

1. THE METHOD

Let $G = G(n, p)$ be the random graph on n vertices with edge probability p . For each K_4 in G randomly select an edge. Delete these edges from G , giving G^* . We show that for appropriate n, p (*) is satisfied by G^* with positive probability. It shall be convenient to write $p = cn^{-1/2}$. In the end we will minimize n by taking c roughly 6, and n roughly $3E8$. Set

$$U = \{(x, xyz) : xyz \text{ is a triangle of } G\}$$

$$U^* = \{(x, xyz) : xyz \text{ is a triangle of } G^*\}.$$

Note. xy, xyz shall denote the sets $\{x, y\}, \{x, y, z\}$ throughout. For each vertex x set

$$N(x) = \{y : xy \in G\}$$

and

$A(x) =$ maximum over all partitions $N(x) = T \cup B$ of the number of edges $yz \in G$ with $y \in T$ and $z \in B$.

THEOREM. *If*

$$\sum_x A(x) < \frac{2}{3}|U^*| \tag{**}$$

then G^ satisfies (*).*

Proof. Clearly G^* has no K_4 ; suppose there is a coloring χ with no monochromatic triangle. We count pairs (x, xyz) such that xyz is a triangle of G^* and $\chi(xy) \neq \chi(xz)$. For each triangle xyz the coloring is essentially unique (two red edges and a blue edge or vice versa) and there are two choices of x so that (x, xyz) is counted so the number of pairs is precisely $\frac{2}{3}|U^*|$. (The unique nature of two colorings of K_3 is unusual and does not seem to generalize to the case of more colors.) For each x let $B(x) =$

$\{y \in N(x) : \chi(xy) = \text{blue}\}$, $T(x) = N(x) - B(x)$. Then the number of (x, xyz) counted is precisely the number of edges $yz \in G^*$ with $y \in T(x)$, $z \in B(x)$. Replacing G^* by the larger G can only increase this number, and replacing the partition $T(x), B(x)$ by the optimal partition T, B can only increase this number so that the number of (x, xyz) is at most $A(x)$ and the total number of such pairs is at most $\sum A(x)$ which would contradict (**). ■

We shall show for appropriate n, p that (**) holds with positive probability.

2. THE CALCULATION IGNORING VARIANCE

Let

T = number of triangles in G

Q = number of K_4 in G

R = number of (xy, uv, a) with x, y, u, v, a distinct, $ax, ay \in G$, $xyuv$ forming a K_4 in G , xy selected from $xyuv$ to be removed from F^* .

Clearly $|U| = 3T$. Also $|U - U^*| \leq 2Q + R$. For suppose $(a, axy) \in U - U^*$. Then xy was in a K_4 of G and was deleted and $ax, ay \in G$. If the K_4 does not contain a it is counted in R ; those (a, axy) , where the K_4 contains a are at most $2Q$ in number, since each K_4 $abxy$ chooses one edge xy and contributes axy, bxy to $U - U^*$. Together,

$$|U^*| > 3T - 2Q - R.$$

We find expectations

$$E(T) = \binom{n}{3} p^3 \sim (c^3/6)n^{3/2} \tag{1}$$

$$E(Q) = \binom{n}{4} p^6 \sim (c^6/24)n \tag{2}$$

$$E(R) = 30 \binom{n}{5} p^8/6 \sim (c^8/24)n \tag{3}$$

so that

$$E(|U^*|) > \frac{1}{2}c^3n^{3/2} - (c^6/12 + c^8/24)n. \tag{4}$$

In the next section we examine variances and show that $|U^*|$ is “very often” “very close” to its expectation.

Now we examine $A(x)$. Set

$$d = d(x) = |N(x)|$$

$$e = e(x) = \text{number of edges of } G \text{ in } N(x).$$

Conditioning on values $d, e, N(x)$ becomes a random graph H with d vertices and e edges.

For a partition $N(x) = T \cup B$ let X_T be the number of edges of H from T to B . Assume $|T| = |B| = d/2$, that being the extreme case. Then X_T has basically binomial distribution $B(e, \frac{1}{2})$ as e edges are selected and each has probability $\frac{1}{2}$ of “crossing.” Employing the basic Chernoff bound

$$\Pr[X_T > \frac{1}{2}e + \frac{1}{2}e^{1/2}s] < \exp(-s^2/2). \quad (5)$$

We set $s = (2d \ln 2)^{1/2}(1.01)$ so that this probability is $\ll 2^{-d}$. But $A(x) = \max X_T$, over 2^d possible T , so

$$\Pr[A(x) > \frac{1}{2}e + \frac{1}{2}e^{1/2}s] \ll 2^d 2^{-d} \ll 1. \quad (6)$$

That is, “almost always,” all

$$A(x) < \frac{1}{2}e(x) + d(x)^{1/2} e(x)^{1/2} (\frac{1}{2} \ln 2)^{1/2} (1.01). \quad (7)$$

Now $\sum_x e(x) = 3T \sim \frac{1}{2}c^3 n^{3/2}$, all $d(x) \sim np$, all $e(x) \sim \frac{1}{2}c^3 n^{1/2}$ so

$$\sum |A(x)| < c^3 n^{3/2}/4 + n(np)^{1/2} (c^3 n^{1/2}/2)^{1/2} (\ln 2/2)^{1/2}. \quad (8)$$

Combining (4), (8), (**) holds if

$$\begin{aligned} & c^3 n^{3/2}/4 + n(cn^{1/2})^{1/2} (c^3 n^{1/2}/2)^{1/2} (\ln 2/2)^{1/2} \\ & < c^3 n^{3/2}/3 - [c^6/18 + c^8/36]n; \end{aligned} \quad (9)$$

i.e., if

$$\left[\frac{c^6}{18} \left(1 + \frac{c^2}{2} \right) / \left(\frac{c^3}{12} - \frac{c^2(\ln 2)^{1/2}}{2} \right) \right]^2 < n, \quad (10)$$

where the LHS must have positive denominator. We take $c \sim 6$ to minimize this inequality so that $n \sim 2.7 \times 10^8$. We allow ourselves a little room and set $c = 6$, $n = 3 \times 10^8$ in the next section. We know that (**) holds “almost always”—i.e., with probability approaching unity as n approaches infinity—but our object is to show that with these particular values (**) holds with positive probability.

3. THE CALCULATION

Set $c = 6$, $n = 3E8$, $p = cn^{-1/2}$. We find (to three significant decimals)

$$E(E) = 1.87E14 \quad \text{Var}(T) < 5E16 \quad (11)$$

$$E(Q) = 5.83E11 \quad \text{Var}(Q) < E12 \quad (12)$$

$$E(R) = 2.10E13 \quad \text{Var}(R) < 6E16. \quad (13)$$

The variance calculations are cumbersome though elementary exercises. We employ Chebyshev's inequality in the form

$$\Pr[|X - E(X)| > tE(X)] < t^{-2} \text{Var}(X)/E(X)^2. \quad (14)$$

Taking $t = 10^{-3}$ with $X = T, Q$, and R above we find

$$\Pr[1.88E14 > T > 1.86E14] > 0.999 \quad (15)$$

$$\Pr[Q < 5.84E11] > 0.999 \quad (16)$$

$$\Pr[R < 2.11E13] > 0.999. \quad (17)$$

Let $\text{BAD}(x)$ be the event, setting $e = e(x)$, $d = d(x)$ given by

$$\text{BAD}(x): A(x) > \frac{1}{2}e/(d-1) + e^{1/2} d^{1/2} (\frac{1}{2} \ln 2)^{1/2} (1.01) \quad (18)$$

and let BAD be the disjunction of the events $\text{BAD}(x)$ over all vertices x . We show

$$\Pr[\text{BAD}] < 0.01, \quad (19)$$

for which it suffices to show

$$\Pr[\text{BAD}(x)] < 3E - 10. \quad (20)$$

The degree $d(x)$ has distribution $B(n-1, p)$ which has mean $(n-1)p = 1.04E5$ and variance $(n-1)p(1-p) = 1.04E5$. We use the Chernoff bounds (see, e.g., [6; or 1, sect. I.3])

$$\Pr[B(m, p) < mp - a] < \exp[-a^2/2pm] \quad (a > 0) \quad (21)$$

$$\Pr[B(m, p) > mp + a] < \exp[-a^2/2pm + a^3/2(pm)^2] \quad (a > 0). \quad (22)$$

First, quite roughly, take $a = E4$ and note

$$\Pr[d(x) < 0.9E5] < \exp[-10^8/2p(n-1)] < 10^{-100}. \quad (23)$$

To show (20) it suffices to show

$$\Pr[\text{BAD}(x) | d(x) = d, e(x) = e] < 3 \times 10^{-10} - 10^{-100} \quad (24)$$

for every d, e with $d \geq 0.9E5$. Conditioning on d, e we may consider $N(x)$ as a random graph $H = (V(H), E(H))$ with d vertices and e edges. For each $S \subset V(H)$ let Y_S be the number of $yz \in E(H)$ with $y \in S, z \notin S$. Let $\text{HYP}[N, M, r]$ denote the hypergeometric distribution of the number of red balls from an urn of M red and $(N - M)$ nonred balls selected in r trials without replacement. Letting $|S| = s, Y_S$ has precisely the distribution $\text{HYP}[\binom{d}{2}, s(d - s), e]$. Set

$$b = \frac{1}{2}e(d/(d - 1)) + e^{1/2} d^{1/2} (\frac{1}{2} \ln 2)^{1/2} (1.01), \tag{25}$$

for convenience. Clearly $\Pr[Y_S > b]$ is maximized when $s(d - s)$ is maximized, i.e., at $s = \lfloor d/2 \rfloor$. Setting

$$q' = \lfloor d/2 \rfloor (d - \lfloor d/2 \rfloor) / \binom{d}{2}, \tag{26}$$

for convenience,

$$\Pr[Y_S > b] \leq \Pr \left[\text{HYP} \left[\binom{d}{2}, q' \binom{d}{2}, e \right] > b \right] \tag{27}$$

W. Uhlmann [7] has made a systematic comparison between $\text{HYP}[N, Nq, r]$ and the corresponding binomial $B(r, q)$ —the distribution given by electing balls with replacement. For our values,

$$\begin{aligned} \Pr \left[\text{HYP} \left[\binom{d}{2}, q' \binom{d}{2}, e \right] > b \right] &\leq \Pr[B(e, q') > b] \\ &\leq \Pr[B(e, q) > b], \end{aligned} \tag{28}$$

setting $q = \frac{1}{2}(d/(d - 1))$, a convenient upper bound on q' . We use the bound (again see, e.g., [6 or 1])

$$\Pr[B(e, q) > eq + a] < \exp(-2a^2/e) \quad (a > 0), \tag{29}$$

valid for all e, q . Then

$$\Pr[Y_S > b] < \exp[-2(1.01)^2 d(\ln 2)/2] < 2^{-d(1.02)}. \tag{30}$$

Hence

$$\begin{aligned} \Pr[\text{BAD}(x) | d(x) = d, e(x) = e] &< \sum \Pr[Y_S] < 2^{d^2 - d(1.02)} \\ &= 2^{-0.02d} < 2^{-1800}, \end{aligned} \tag{31}$$

giving (24) with “plenty of room.”

Application of (21), (22) with precise values give

$$\Pr[d(x) > 1.06E5] < 0.2/n \quad (32)$$

$$\Pr[d(x) < 1.01E5] < 0.1/n, \quad (33)$$

so that, with room to spare,

$$\Pr[1.01E5 \leq d(x) \leq 1.06E5 \text{ for all } x] > 0.7. \quad (34)$$

Combining (15)–(17), (19), (34) we have that, with probability at least 0.65, the pair G, G^* satisfy

$$\begin{aligned} 1.86E14 &< T < 1.88E14 \\ Q &< 5.84E11 \\ R &< 2.11E13 \\ A(x) &< b, \quad \text{all } x \\ 101000 &\leq d(x) \leq 106000, \quad \text{all } x. \end{aligned} \quad (35)$$

Let G, G^* be a specific graph pair satisfying the above. Then

$$\sum A(x) = \frac{1}{2}(1.00001) \sum e(x) + (1.01)(\frac{1}{2} \ln 2)^{1/2} \sum e(x)^{1/2} d(x)^{1/2}. \quad (36)$$

We note $\sum e(x) = 3T$ and bound

$$\begin{aligned} \sum e(x)^{1/2} d(x)^{1/2} &\leq (106000)^{1/2} \sum e(x)^{1/2} \\ &\leq (106000)^{1/2} (3Tn)^{1/2} \end{aligned} \quad (37)$$

as, in general' $y_1^{1/2} + \dots + y_n^{1/2} \leq (y_1 + \dots + y_n)^{1/2} n^{1/2}$. Plugging in values

$$\sum A(x) < 2.83E14. \quad (38)$$

On the other side,

$$2|U^*|/3 \geq 2T - (2/3)(2Q + R) > 3.57E14, \quad (39)$$

so that, indeed, the conditions of the theorem hold and $G^* \rightarrow (K_3)$.

There was plenty of room in our variance arguments. But even if all variances were zero without further argumentation we could not improve on the value $c = 6.0157$ and a graph G with 266, 930, 400 vertices.

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