

A Note on Rational Approximation to $(1 - x)^\sigma$

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THEOREM 1. For $0 < \sigma \leq 1$, $0 \leq x \leq 1$ and $n \geq 0$, we have

$$0 \leq S_n^{-1}(x) - (1 - x)^\sigma \leq \frac{x^{n+1}}{S_n(x)} \leq \binom{\sigma + n}{n}^{-1}, \tag{1}$$

where $S_n(x) = \sum_{j=0}^n \binom{\sigma + j - 1}{j} x^j$. (Note that $(1 - x)^{-\sigma} = \sum_{j=0}^\infty \binom{\sigma + j - 1}{j} x^j$ for $0 < \sigma \leq 1$, $0 \leq x < 1$.)

Proof. For $0 < \sigma \leq 1$, $0 \leq x < 1$, $j \geq 0$, and $n \geq 0$, we obtain, by observing the fact that

$$\begin{aligned} \binom{n + \sigma + j}{n + 1 + j} &\leq \binom{\sigma + j - 1}{j}, \\ 0 \leq S_n^{-1}(x) - (1 - x)^\sigma &= \frac{\sum_{j=n+1}^\infty \binom{\sigma + j - 1}{j} x^j}{S_n(x)(1 - x)^{-\sigma}} \\ &= \frac{x^{n+1} \sum_{j=0}^\infty \binom{n + \sigma + j}{n + 1 + j} x^j}{S_n(x)(1 - x)^{-\sigma}} \\ &\leq \frac{x^{n+1} \sum_{j=0}^\infty \binom{\sigma + j - 1}{j} x^j}{S_n(x)(1 - x)^{-\sigma}} = \frac{x^{n+1}}{S_n(x)}. \end{aligned}$$

Let $0 < \sigma \leq 1$, $n \geq 0$. As

$$(n + 1) S_n(x) > x S_n'(x) \quad \text{for } x > 0,$$

$$\left(\frac{x^{n+1}}{S_n(x)} \right)' > 0 \quad \text{for } x > 0,$$

and hence $x^{n+1}/S_n(x)$ is increasing on $[0, \infty)$. Thus, for $0 \leq x < 1$,

$$0 \leq [S_n(x)]^{-1} - (1-x)^\sigma \leq \frac{x^{n+1}}{S_n(x)} < \frac{1}{S_n(1)}$$

$$= \left(\sum_{j=0}^n \binom{\sigma+j-1}{j} \right)^{-1} = \binom{\sigma+n}{n}^{-1}.$$

The last two equalities prove (1) also for $x = 1$. It follows that

$$\| (1-x)^\sigma - S_n^{-1}(x) \|_{L^\infty[0,1]} \leq \binom{\sigma+n}{n}^{-1}. \tag{2}$$

Remarks on Theorem 1. (i) For $\sigma = \frac{1}{2}$, (2) was obtained in [1].

(ii) For $0 < \sigma < 1$, (2) is sharper than the estimate given in [2], as for $n = 1, 2, \dots$,

$$\binom{\sigma+n}{n}^{-1} = \frac{\Gamma(\sigma+1)\Gamma(n+1)}{\Gamma(\sigma+n+1)} < \frac{n}{n+1} \frac{\Gamma(\sigma)\Gamma(n)}{\Gamma(\sigma+n)}.$$

THEOREM 2. Let $P(x)$ and $Q(x)$ be polynomials of degree at most n (≥ 1) having only real, non-negative coefficients ($Q(0) > 0$) and let $0 < \sigma < 1$. Then

$$\left\| (1-x)^\sigma - \frac{P(x)}{Q(x)} \right\|_{L^\infty[0,1]}$$

$$\geq (n+1)^{-\sigma} \left(1 + \left(\frac{n+1}{n} \right)^n \right)^{-1} > \frac{(n+1)^{-\sigma}}{1+e}. \tag{3}$$

Proof. Set

$$\left\| (1-x)^\sigma - \frac{P(x)}{Q(x)} \right\|_{L^\infty[0,1]} = \varepsilon.$$

Then

$$\varepsilon \geq \left(1 - \frac{n}{n+1} \right)^\sigma - \frac{P\left(\frac{n}{n+1}\right)}{Q\left(\frac{n}{n+1}\right)}$$

$$\geq \frac{1}{(n+1)^\sigma} + \left(0 - \frac{P(1)}{Q(1)} \right) \left(\frac{n+1}{n} \right)^n$$

$$\geq \frac{1}{(n+1)^\sigma} - \varepsilon \left(\frac{n+1}{n} \right)^n$$

Hence

$$\varepsilon \geq (n + 1)^{-\sigma} \left(1 + \left(\frac{n + 1}{n} \right)^n \right)^{-1}.$$

Remarks on Theorem 2. (i) The estimate (3) is sharper than these obtained in [2, 3]. The technique used here is different from the one used in [3].

(ii) Actually the inequality $P(1) \geq \max(0, P(n/(n + 1)))$ suffices for the proof of Theorem 2 rather than assuming $P(x)$ has only real, non-negative coefficients.

THEOREM 3. *Let $P(x)$ and $Q(x)$ be polynomials of degree at most n (≥ 2) and let $P(x)$ have only real, non-negative coefficients. Let $Q(x) \neq 0$ throughout $[0, 1]$ and let $0 < \sigma \leq 1$. Then*

$$\left\| (1 - x)^\sigma - \frac{P(x)}{Q(x)} \right\|_{L^\infty[0,1]} \geq (16n^{2\sigma})^{-1}.$$

We need the following well known

LEMMA [5, p. 68, Eq. (9)]. *If a real polynomial $Q_n(x)$ of degree $\leq n$ ($n \geq 0$) satisfies the inequality $|Q_n(x)| \leq L$ on $[a, b]$ ($-\infty < a < b < \infty$), then at each real x outside $[a, b]$ we have*

$$|Q_n(x)| \leq L \left| T_n \left(\frac{2x - a - b}{b - a} \right) \right|,$$

where T_n is the n th degree Chebychev polynomial of the first kind.

Proof of Theorem 3. Suppose

$$\left\| (1 - x)^\sigma - \frac{P(x)}{Q(x)} \right\|_{L^\infty[0,1]} < (16n^{2\sigma})^{-1}. \tag{4}$$

Then on $[0, 1 - n^{-2}]$,

$$\begin{aligned} \left| \frac{P(x)}{Q(x)} \right| &> (1 - x)^\sigma - \frac{1}{16n^{2\sigma}} \geq \frac{1}{n^{2\sigma}} - \frac{1}{16n^{2\sigma}} = \frac{15}{16n^{2\sigma}}, \\ |Q(x)| &< \frac{16n^{2\sigma}}{15} P(1). \end{aligned} \tag{5}$$

In (5) we have used the assumption that $P(x)$ has only real, non-negative coefficients. Now, by applying the lemma to (5) with $a = 0, b = 1 - n^{-2}$, we

get on $[0, 1]$, by familiar elementary properties of T_n , and by [4, p. 38, Problem 172],

$$\begin{aligned} \text{Max}_{[0,1]} |Q(x)| &\leq \frac{16n^{2\sigma}}{15} P(1) T_n\left(\frac{n^2+1}{n^2-1}\right) \\ &< \frac{16n^{2\sigma}}{15} P(1) \left(\frac{n^2+1}{n^2-1} + \sqrt{\left(\frac{n^2+1}{n^2-1}\right)^2 - 1}\right)^n \\ &= \frac{16n^{2\sigma}}{15} P(1) \left(\frac{n+1}{n-1}\right)^n \\ &\leq \frac{16n^{2\sigma}}{15} P(1) \cdot 9 = \frac{48}{5} n^{2\sigma} P(1). \end{aligned} \tag{6}$$

From (6) and (4) we obtain $\frac{5}{48} < n^{2\sigma} |P(1)/Q(1)| < \frac{1}{16}$, which is false.

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