

A treelet transform analysis to relate nutrient patterns to the risk of hormonal receptor-defined breast cancer in the European Prospective Investigation into Cancer and Nutrition (EPIC)

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Abstract

Objective: Pattern analysis has emerged as a tool to depict the role of multiple nutrients/foods in relation to health outcomes. The present study aimed at extracting nutrient patterns with respect to breast cancer (BC) aetiology.

Design: Nutrient patterns were derived with treelet transform (TT) and related to BC risk. TT was applied to twenty-three log-transformed nutrient densities from dietary questionnaires. Hazard ratios (HR) and 95% confidence intervals computed using Cox proportional hazards models quantified the association between quintiles of nutrient pattern scores and risk of overall BC, and by hormonal receptor and menopausal status. Principal component analysis was applied for comparison.

Setting: The European Prospective Investigation into Cancer and Nutrition (EPIC).

Subjects: Women (n 334 850) from the EPIC study.

Results: The first TT component (TC1) highlighted a pattern rich in nutrients found in animal foods loading on cholesterol, protein, retinol, vitamins B₁₂ and D, while the second TT component (TC2) reflected a diet rich in β -carotene, riboflavin, thiamin, vitamins C and B₆, fibre, Fe, Ca, K, Mg, P and folate. While TC1 was not associated with BC risk, TC2 was inversely associated with BC risk overall (HR_{Q5 v. Q1} = 0.89, 95% CI 0.83, 0.95, $P_{\text{trend}} < 0.01$) and showed a significantly lower risk in oestrogen receptor-positive (HR_{Q5 v. Q1} = 0.89, 95% CI 0.81, 0.98, $P_{\text{trend}} = 0.02$) and progesterone receptor-positive tumours (HR_{Q5 v. Q1} = 0.87, 95% CI 0.77, 0.98, $P_{\text{trend}} < 0.01$).

Conclusions: TT produces readily interpretable sparse components explaining similar amounts of variation as principal component analysis. Our results suggest that participants with a nutrient pattern high in micronutrients found in vegetables, fruits and cereals had a lower risk of BC.

Keywords

Nutrient patterns
Treelet transform
Breast cancer

European Prospective Investigation
into Cancer and Nutrition
Principal component analysis

Breast cancer (BC) remains the highest incident cancer affecting women worldwide, with almost 1 670 000 cases registered in 2012. It is a major public health concern with mortality from BC accounting for over 522 000 deaths in 2012, including almost 198 000 deaths in Western countries and about 324 000 in less developed regions⁽¹⁾. Established BC risk factors include age, genetic mutations, ethnicity, height, reproductive history, breast-feeding, hormone therapy and diabetes^(2–6). Besides these, a number of modifiable lifestyle factors are associated with BC such as smoking^(7,8), body fat and obesity^(9–11), physical inactivity^(10,12,13), alcohol consumption^(14–16) and diet^(5,17,18). Diet has been suggested to account for up to 25–40% of preventable causes of cancers; in particular, 50% of BC deaths are linked to diet, although the consensus around this estimate is not unanimous^(12,19,20). Standard approaches customarily evaluate the risk of BC associated with one or a group of dietary items, i.e. food(s) or nutrient(s). Nevertheless, associations between diet and disease might be missed when one parses the effect of a limited list of dietary constituents. Although this simplified approach of examining a single food or nutrient at a time has led to important results on the role of an individual dietary component in BC aetiology, such as fibre from vegetables, alcohol, tea consumption, folate and other micronutrients^(12,14,18,20–23), research might benefit from a more comprehensive approach by exploring BC aetiology in terms of an integrated ensemble of dietary characteristics.

To capture the complexity of individuals' dietary habits, dietary pattern analysis has emerged as a complementary holistic methodology focusing on sets of dietary variables and addressing their inherent interrelations⁽²⁴⁾. This approach is justified as components of dietary exposure are not independent^(25,26) and because it allows to account for complex relationships between nutrients in biological pathways⁽²⁵⁾. In addition, BC is a multifactorial disease^(2–18), the aetiology of which possibly depends on more than a restricted list of dietary items.

Recent investigations carried out in Western populations^(27–32) have consistently identified two main dietary patterns: the prudent/healthy and the Western/unhealthy^(29,33). While diet is related to cultural background, common nutrients are present in different combinations of foods; hence looking into diet–disease associations on the nutrient scale could lead to the identification of specific nutritional profiles relevant to BC aetiology.

In the present study, nutrient patterns within the European Investigation into Cancer and Nutrition (EPIC) were related to BC risk. Nutrient patterns were obtained by applying the treelet transform (TT) that has recently been introduced into nutritional epidemiology^(34–36) and the well-known principal component analysis (PCA) was used for the sake of comparison⁽³⁷⁾. TT yields sparse components and reveals the intrinsic structure of the data, thus simplifying interpretability. Aspects related to the application of TT to dietary data in the context of a multi-centre study are described and discussed. The association between nutrient

patterns and BC was evaluated using all BC cases and by taking into account the heterogeneity of BC subtypes by integrating information on menopausal and hormone receptor status.

Materials and methods

Study population and exclusion criteria

EPIC is a large prospective cohort of 521 330 healthy men and women designed to evaluate the relationships between dietary habits, nutrition, lifestyle factors and the incidence of cancer. The EPIC cohort includes participants from twenty-three centres in France, Germany, Denmark, Sweden, Norway, Greece, Italy, the Netherlands, Spain and the UK. In most centres, participants were recruited from the general population, the exceptions being France (women were enrolled from a national health insurance scheme covering teachers in the French education system employees), Italy (Turin and Ragusa: blood donors; Florence: screening programme participants), Spain (blood donors) and the Netherlands (Utrecht: women participating in BC screening). In Norway, only women from the general population were recruited and in the UK, one-half of the cohort (the Oxford sub-cohort) consisted of 'health-conscious' individuals from England, Wales, Scotland and Northern Ireland. The design of the study and its rationale along with the recruitment process have been described elsewhere⁽³⁸⁾.

Among the 521 330 EPIC participants, men were first removed (n 153 427). Women with prevalent cancers at any site at baseline (other than non-melanoma skin cancer; n 19 853) or lost to follow-up (n 2892) were excluded, as were women who did not complete any dietary questionnaire (n 3315) and those who did not complete a lifestyle questionnaire (n 26). To avoid including extreme values, participants in the top and bottom 1% of the distribution of the ratio of reported total energy intake to energy requirement (n 6753) were excluded. After exclusion of non-first BC cases (n 2) the cohort included 335 062 women upon whom the dietary patterns were derived. An additional number of women (n 212) with missing information on BC status were excluded, which left 334 850 women retained for the statistical analyses.

Cancer assessment

Incident BC cases were identified through population cancer registries (Denmark, Italy, Netherlands, Norway, Spain, Sweden and UK) or through active follow-up (France, Germany, Naples and Greece), as detailed in Ferrari *et al.*⁽²¹⁾. Information on oestrogen receptor (ER) and progesterone receptor (PR) statuses was provided by each centre on the basis of pathology reports.

Dietary assessment

Long-term usual dietary intake was assessed at baseline using country-specific and validated dietary questionnaires

(self-administered FFQ, semi-quantitative or interviewer-performed)^(38–40). In the validation studies, the dietary questionnaires were compared with a reference method which was in most centres 24 h dietary recalls, except in Sweden and the UK, where food records were used. Generally, the correlation coefficients were between 0.40 and 0.70 for all nutrients examined which was considered satisfactory⁽⁴¹⁾. Individual intakes of twenty-three nutrients and total energy were estimated using a common food composition database, the EPIC Nutrient Database (ENDB), which was compiled from national food composition databases of the ten countries represented in EPIC following standardized procedures^(42,43).

Lifestyle questionnaires

Information on sociodemographic characteristics, including education, and lifestyle habits such as levels of physical activity, tobacco smoking, as well as consumption of alcohol and drinking habits, were collected using lifestyle questionnaires. In addition, anthropometric measures and past medical information were gathered at recruitment⁽³⁸⁾.

Nutrient pattern assessment

EPIC-wide nutrient patterns were derived among female participants in EPIC using TT in the main analysis and PCA in the sensitivity analysis. The sample covariance matrix of twenty-three log-transformed nutrient densities, computed using alcohol-free energy intake⁽⁴⁴⁾, was consistently used. The use of the sample covariance matrix allows variability to be informative in the pattern discovery phase. The distribution of nutrient consumption tends to be log-normal and may not be best described by the mean and variance on the original scale. Moreover micro- and macronutrients are expressed on different scales (micrograms, milligrams or grams). The nutrient densities were log-transformed to remove scale dependence and render their variance (or covariance) independent of the unit of measure. In line with previous work^(28,45,46), alcohol intake was not included and was considered as a lifestyle factor. Total fat was divided into MUFA, PUFA and SFA, and total carbohydrates were broken down into starch and sugar. The micro- and macronutrients studied were Ca, β -carotene, cholesterol, MUFA, PUFA, SFA, Fe, fibre, K, Mg, P, protein, retinol, riboflavin, starch, sugar, thiamin, vitamins B₆, B₁₂, C, D, E and folate. The list of nutrients as well as the approach described for their handling is consistent with the nutrient patterns initiative within EPIC described by Moskal *et al.*⁽⁴⁵⁾.

Pattern extraction

The TT method used for pattern extraction is described in detail by Gorst-Rasmussen and co-workers^(35,47). Briefly, TT is a dimension reduction technique aimed at converting a set of observations of possibly correlated variables into orthogonal components. TT scores, corresponding to

the projection of data onto components, generally have a small degree of correlation, unlike PCA scores that are always uncorrelated. The number of retained components was based on the percentage of explained variance, scree plots and interpretability. The nutrient patterns were defined after the inspection of factor loadings, i.e. eigenvectors, expressing the contribution of nutrients to a given component. Score variables were determined for each component of TT and reflected adherence to a given type of diet/nutrient profile. TT combines the quantitative pattern extraction capabilities of PCA with interpretational advantages of hierarchical clustering of variables. In TT, the two variables displaying the highest correlation (or covariance) are identified, and a PCA is performed on them. The two variables are then replaced with the score of their first PCA component and a merge is indicated in

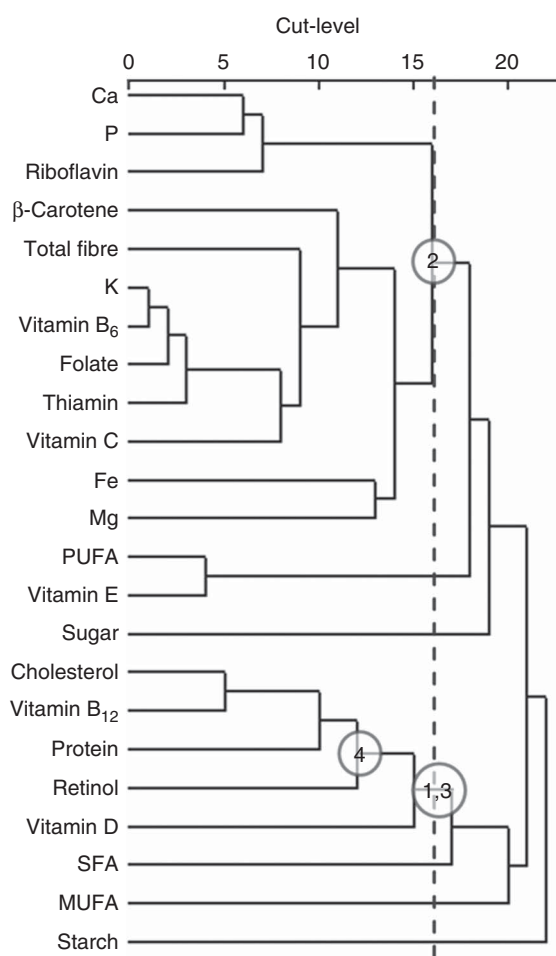


Fig. 1 Cluster tree produced by the treelet transform algorithm applied to twenty-three log-transformed nutrient densities for 335 062 women in the European Prospective Investigation into Cancer and Nutrition (EPIC). The dashed line indicates the chosen cut-level (16) to extract components. The highest-variance factors, i.e. treelet components at this level of the tree, are indicated with numbered circles. The nutrients related to these nodes have non-zero loadings on the given component. Components 1 and 3 share the same node but the variable loadings differ

the cluster tree. This operation is re-iterated until all variables have joined the cluster tree. In this way, TT produces a hierarchical grouping of variables which may reveal intrinsic characteristics of data structure. An important feature of TT is that it introduces sparsity into factors, making many factors loadings exactly equal to zero, potentially simplifying the interpretation. Alongside the cluster tree dendrogram produced by TT (as exemplified in Fig. 1), TT yields a coordinate system for the data at each level of the cluster tree. Selecting a cluster tree level (cut-level) for the TT cluster tree amounts to choosing the level of detail desired in the dimension reduction of data. More variation can be explained at the cost of factor sparsity when the cluster tree is cut near its 'root'. If the data have p variables, there are $p - 1$ possible cut-levels. After deciding on the number of components to retain, we performed a tenfold cross-validation to identify the optimal cut-level, i.e. the point at which increasing the cut-level does not substantially increase the variation of the retained patterns. We also performed a sensitivity analysis to assess the effect of different cut-levels^(35,48).

Consistently, a PCA was also applied for the sake of comparison⁽³⁷⁾. This technique yields orthogonal components that are invariant to the number of subsequent components retained. PCA identifies the best linear combination of the variables accounting for the most variance observed in the original data, producing components with uncorrelated scores. Results of TT analysis were compared with findings obtained with the more classic PCA method. To make the comparison easier, and because TT returns sparse vectors, only nutrients with absolute loadings greater than 0.2 were retained to identify a given pattern in PCA.

Patterns and breast cancer risk

The associations between nutrient patterns and risk of BC were investigated by using Cox proportional hazards regression models to estimate hazard ratios (HR) and 95% confidence intervals. Breslow's method was adopted for handling time ties⁽⁴⁹⁾. The time at entry was the age at recruitment and the time of exit was the age at cancer diagnosis, death, loss or end of follow-up, whichever happened first. Models were stratified by centre, to control for differences in questionnaire designs, follow-up procedures and other centre-specific effects, as well as for age at recruitment (1-year categories)⁽⁵⁰⁾. Analyses were performed by considering the TT (and principal component (PC)) scores in quintiles to appreciate potential departure from linearity. Statistical analyses were adjusted for baseline menopausal status (premenopausal and perimenopausal (reference) or postmenopausal and women who underwent an ovariectomy), baseline alcohol intake (never drinkers (reference), former drinkers, drinkers only at recruitment, lifetime drinkers, unknown), height (continuous), BMI (below (reference) or above 25 kg/m²), schooling level (none, primary (reference), technical/

professional/secondary, longer education, unknown/unspecified), age at first full-term pregnancy (nulliparous (reference), ≤ 21 years, 21–30 years, > 30 years, unknown or missing), age at menarche (≤ 12 years (reference), 12–14 years, > 14 years, missing), age at menopause (≤ 50 years (reference), > 50 years, premenopausal or missing), use of hormone replacement therapy (never (reference), ever, unknown), level of physical activity (categorical, metabolic equivalents of task (MET)/h: inactive (reference), moderately inactive, moderately active, active, unknown) and alcohol-free energy (continuous). Use of oral contraceptive pills (never (reference), ever or unknown) and smoking status (never smokers (reference), ex-smokers, current smokers, unknown) were evaluated but not retained in the final models, due to limiting confounding exerted by these variables.

The overall significance of a score variable in categories was evaluated using the likelihood ratio test statistics (P_{LRT}) with $df = 4$. Additionally, P values for trend (P_{trend}) were computed by modelling a score variable with quintile-specific medians as continuous. The association between nutrient patterns and BC risk was evaluated in pre- and postmenopausal women and according to BC hormonal receptor status (ER/PR status). Interaction between menopausal status and pattern scores was explored. In addition, tests of heterogeneity of associations according to receptor status were performed using the data-augmentation method⁽⁵¹⁾ by comparing the difference in the log likelihood between a model with receptor status-specific variable and a model with a single HR estimate for the two categories of receptor status to a χ^2 distribution with $df = 1$ ($P_{heterogeneity}$).

Departure from linearity was explored with restricted cubic splines⁽⁵²⁾, using five knots corresponding to the 1st and 99th percentiles and medians of the centred scores of quintiles 1, 3 and 5. Spline plots were produced by taking the median of the first quintile as reference. Departures from linearity were assessed via an evaluation of the joint significance of variables other than the linear one included in the model using Wald's test on $df = 3$. Associations

between all of the PC and BC were investigated in a consistent way.

Statistical tests were two-sided, the per-test significance level was set to $\alpha = 0.05$. All analyses were performed using the SAS statistical software package version 9.3; the 'tt' package in the STATA statistical software package release 12 was used to perform TT.

Results

A total of 11 576 BC cases were recorded in 11.5 years of median follow-up time and 3 670 439 person-years. Based on the information obtained at baseline, 2827 cases were premenopausal, 5872 were postmenopausal, 2548 were perimenopausal and 328 cases had a bilateral ovariectomy. Among incident cases, information on hormone receptor status for ER and PR was available only in 62% and 52% of total cancer cases, respectively, and was distributed as follows: 81% ER⁺ and 19% ER⁻ tumours and 63% PR⁺ and 37% PR⁻ tumours. Descriptive information of the study sample by EPIC country is available in Table 1.

Identification of nutrient patterns

Inspection of factor loadings allowed an initial identification of four nutrient patterns with TT, explaining 62% of total nutrient intake variability within individuals. After a tenfold cross-validation along with a sensitivity analysis strategy and after evaluating the interpretability of each pattern, we chose to cut the cluster tree at level 16. Loadings of components 1 and 2 are shown in Table 2. TT yielded a dendrogram shown in Fig. 1, with numbered nodes indicating the four highest-variance factors, where factors 1 and 2 were identified as the first two components after setting the cut-level to 16 indicated by the dashed line. This dendrogram reveals the correlation structure of the log-transformed nutrient densities. The first treelet component (TC1) loaded on vitamin D, vitamin B₁₂, cholesterol, protein and retinol, suggesting a diet rich in animal products. The second treelet component (TC2)

Table 1 Numbers of women and breast cancer (BC) cases (first tumours only) in the European Prospective Investigation into Cancer and Nutrition (EPIC) cohort by country

Country	No. of women	Person-years	No. of BC cases	Follow-up time (years)*	Age at enrolment (years)*
France	67 356	699 216	3187	11.8	51.5
Italy	30 498	341 417	1047	11.7	50.9
Spain	24 846	299 575	495	12.6	47.7
UK general population	17 145	200 812	719	12.3	55.6
UK health-conscious	35 368	385 353	761	11.3	41.5
Netherlands	26 839	315 554	916	12.2	52.7
Greece	15 224	148 594	198	10.7	53.6
Germany	27 390	272 011	834	10.9	48.4
Sweden	26 339	349 110	1095	13.9	50.6
Denmark	28 693	316 601	1340	11.6	56.3
Norway	35 152	342 195	984	10.1	48.0
Total	334 850	3 670 439	11 576	11.5	51.0

*Median is given for follow-up time and age at enrolment.

Table 2 Loadings of the first two components from treelet transform (TT; cut-level 16)

Variable*	TT 16 loadings	
	TC1	TC2
Ca		0.153
β -Carotene		0.721
Cholesterol	0.294	
MUFA		
PUFA		
SFA		
Fe		0.109
Fibre		0.183
K		0.157
Mg		0.144
P		0.074
Protein	0.086	
Retinol	0.679	
Riboflavin		0.141
Starch		
Sugar		
Thiamin		0.217
Vitamin B ₆		0.185
Vitamin B ₁₂	0.421	
Vitamin C		0.452
Vitamin D	0.517	
Vitamin E		
Folate		0.235
Explained variance	26 %	21 %

TC1, treelet component 1; TC2, treelet component 2.

*Log-transformed nutrient variables.

presented high positive loadings on β -carotene, thiamin, fibre, vitamin C and folate, and singled out some nutrients with mild loadings (<0.2), i.e. Fe, Ca, K, Mg and P (Table 2). TC2 may evoke a diet rich in vegetables, fruits and cereals. While the third treelet component (TC3) was largely driven by vitamin D, the fourth treelet component (TC4) was less straightforward to characterize, as displayed in the online supplementary material, Supplemental Table 1. Distributions of known risk factors for BC by quintiles of TT scores for the first two components are displayed in Table 3.

PC loadings are displayed in the online supplementary material, Supplemental Table 2. PCA produced patterns similar to TT with respect to the amount of variability explained and the nutrients contributing to the definition of each component: with PC1 displaying high loadings for cholesterol, retinol, vitamin B₁₂ and vitamin D and negative loadings for vitamin C and β -carotene; and PC2 suggesting a micronutrient-dense pattern rich in fruits, vegetables, plant foods and dairy. The first two components (in TT and PCA) explained the most variability and were the most informative with respect to capturing meaningful nutrient patterns, and thus were further related to BC risk in disease models.

Nutrient patterns and breast cancer risk

Scores of nutrient patterns were related to BC risk. TC1 showed no statistically significant association with BC risk with $HR_{TC1\ Q5\ v.\ Q1} = 1.05$ (95 % CI 0.98, 1.13, $P_{trend} = 0.36$,

$P_{LRT} = 0.39$), while TC2 was significantly associated with BC risk with $HR_{TC2\ Q5\ v.\ Q1} = 0.89$ (95 % CI 0.83, 0.95, $P_{trend} < 0.001$, $P_{LRT} = 0.02$), as shown in Table 4. The relationship between TT scores and BC risk was modelled through restricted cubic splines (RCS) and is presented in Fig. 2. Overall, there was a significant progressive decrease in BC risk for the second component. TC2 scores showed a linear decrease in BC risk ($RCS_{TC2\ P_{trend}} = 0.02$). However, no departure from linearity was observed ($P_{Wald\ non-linearity} = 0.94$ and 0.77, respectively, in TC1 and TC2; Fig. 2). Analyses of interaction between TC (or PC) scores and menopausal status were not statistically significant (results not shown).

Hormonal receptor status

In ER⁻ tumours, no significant association with BC risk was observed for TC1 and TC2 scores (Table 4). For ER⁺ tumours there was a decrease in BC risk in the fourth and fifth quintiles of TC2 scores with $HR_{Q4\ v.\ Q1} = 0.90$ (95 % CI 0.83, 0.99) and $HR_{Q5\ v.\ Q1} = 0.89$ (95 % CI 0.81, 0.98, $P_{trend} = 0.02$; Table 4). Regarding PR⁻ tumours (see online supplementary material, Supplemental Table 3), the second component TC2 showed a decreased BC risk with $HR_{Q5\ v.\ Q1} = 0.84$ (95 % CI 0.72, 0.98). For PR⁺ tumours, TC2 was linked with a decreased BC risk in participants in the fifth quintile with $HR_{Q5\ v.\ Q1} = 0.87$ (95 % CI 0.77, 0.98). No significant association was seen for ER⁻/PR⁻ tumours (Table 5). TC2 was linked with a decreased BC risk trend in ER⁺/PR⁺ tumours with $HR_{Q5\ v.\ Q1} = 0.86$ (0.76, 0.98, $P_{trend} < 0.01$; Table 5). Tests of heterogeneity yielded no significant results.

PCA derived components displayed a significant increase in BC risk for PC1 in participants in the highest quintile and a decreasing trend of BC risk for PC2, as shown in the online supplementary material, Supplemental Table 4 and Supplemental Fig. 1. Results of associations of PC with tumours by hormone receptor status are displayed in the online supplementary material, Supplemental Tables 4 and 5.

Discussion

In the present study, the role of nutrient patterns in the aetiology of BC was explored through the use of TT, a multivariate method recently introduced to the landscape of nutritional epidemiology⁽³⁴⁻³⁶⁾. The association was evaluated in the context of the EPIC study, characterized by large variability of dietary habits and by a large number of incident cancer cases across participating centres⁽³⁸⁾.

In recent years, dietary pattern analysis has emerged as a promising technique, complementary to methods focusing on individual foods or food components, to investigate the relationships between diet and risk of disease⁽²⁵⁾. A systematic review and meta-analysis on dietary patterns in BC aetiology⁽³³⁾ selected eighteen

Table 3 Lifestyle and dietary baseline characteristics* according to the lowest, middle and highest quintiles of treelet transform (cut-level 16) scores for the first and second components among 334 850 women in the European Prospective Investigation into Cancer and Nutrition (EPIC)

	TC1						TC2					
	Q1		Q3		Q5		Q1		Q3		Q5	
	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD
No. of women	66 988		66 977		66 955		66 961		66 969		66 970	
Age (years)	50.2	11.8	50.8	9.5	52.0	8.1	49.6	9.3	51.1	9.5	52.2	10.9
Weight (kg)	63.0	11.6	64.8	11.8	65.0	11.9	64.0	11.9	64.0	11.7	63.8	11.5
Height (cm)	160.1	7.1	162.6	6.5	163.0	6.5	162.0	6.9	162.5	6.7	162.0	6.5
Non-alcohol energy (kJ/d)	7565	2280	7573	2171	7368	2121	8309	2406	7623	2138	6820	1929
Non-alcohol energy (kcal/d)	1808	545	1810	519	1761	507	1986	575	1822	511	1630	461
	%		%		%		%		%		%	
BMI class												
Below 25 kg/m ²	57		59		57		58		58		59	
Above 25 kg/m ²	43		41		43		42		42		41	
Schooling level												
None	11		3		2		5		5		4	
Primary	25		22		26		33		23		17	
Technical/professional/secondary	35		47		50		44		46		44	
Longer education	25		23		19		16		23		28	
Unspecified/unknown	4		5		3		2		3		8	
Use of hormone replacement therapy												
Never	82		68		60		71		68		69	
Ever	16		25		31		20		25		27	
Unknown	2		7		9		9		7		4	
Age at first term pregnancy												
Nulliparous	21		13		11		13		14		19	
≤21 years	16		18		24		20		18		17	
21–30 years	52		56		54		54		56		52	
>30 years	9		9		7		8		8		8	
Unknown	3		5		4		5		4		5	
Age at menarche												
≤12 years	38		35		33		33		35		39	
12–14	46		46		47		46		47		45	
>14 years	15		15		17		16		16		14	
Unknown	1		4		4		5		3		3	
Age at menopause												
≤50 years	19		16		18		17		17		18	
>50 years	19		18		19		16		18		19	
Unknown	63		66		63		67		65		62	
Menopausal status												
Pre and peri	55		55		49		60		53		49	
Post and ovariectomy	45		45		51		40		47		51	
Alcohol drinkers												
Never	16		6		4		8		8		9	
Former	6		3		2		4		3		4	
Only at recruitment	17		11		8		6		11		19	
Lifetime	51		56		46		44		54		57	
Unknown	10		22		40		38		24		11	
Physical activity												
Inactive	31		20		16		25		20		21	
Moderately inactive	33		33		28		30		31		33	
Moderately active	21		23		18		18		22		24	
Active	13		15		12		12		14		17	
Unknown	2		10		25		15		13		5	

TC1, treelet component 1; TC2, treelet component 2; Q1, quintile 1; Q3, quintile 3; Q5, quintile 5.

*Means and standard deviations are presented for continuous variables, and frequencies are presented for categorical variables.

relevant studies from case–control and cohort studies that used combinations of foods and micronutrients to identify dietary patterns^(17,27,53–66). Two *a posteriori* defined patterns emerged consistently: the Western/unhealthy (in seventeen studies) and the prudent/healthy (eighteen

studies)⁽³³⁾. In the aforementioned meta-analysis⁽³³⁾, the prudent/healthy dietary pattern, rich in intakes of vegetables, leafy vegetables, legumes and fish, was associated to decreased BC risk (relative risk comparing top *v.* bottom categories = 0.89, 95 % CI 0.82, 0.99), while the Western/

Table 4 Hazard ratios (HR) and 95 % confidence intervals for breast cancer (BC) by quintiles of pattern scores (first and second components of treelet transform, cut-level 16) for overall, oestrogen receptor-positive (ER⁺) and oestrogen receptor-negative (ER⁻) tumours in 334 850 women in the European Prospective Investigation into Cancer and Nutrition (EPIC)

Model*	TC1						TC2					
	Person-years	No. of BC cases	HR	95 % CI	$P_{LRT}†$	$P_{trend}‡$	Person-years	No. of BC cases	HR	95 % CI	$P_{LRT}†$	$P_{trend}‡$
Overall												
Q1	730 785	1784	1.00	Ref.	0.39	0.36	747 690	2317	1.00	Ref.	0.02	<0.001
Q2	738 136	2342	1.06	0.99, 1.13			736 718	2307	0.95	0.89, 1.00		
Q3	735 683	2376	1.04	0.97, 1.11			729 544	2365	0.95	0.89, 1.01		
Q4	737 533	2513	1.06	0.99, 1.14			725 903	2350	0.94	0.88, 1.00		
Q5	728 303	2561	1.05	0.98, 1.13			730 584	2237	0.89	0.83, 0.95		
ER ⁺												
Q1	725 634	885	1.00	Ref.	0.55	0.47	740 268	1133	1.00	Ref.	0.13	0.02
Q2	731 571	1214	1.07	0.98, 1.17			729 915	1140	0.92	0.84, 1.00		
Q3	728 782	1212	1.06	0.97, 1.16			722 467	1192	0.92	0.84, 1.00		
Q4	729 703	1247	1.08	0.98, 1.19			719 201	1193	0.90	0.83, 0.99		
Q5	720 422	1272	1.05	0.95, 1.16			724 261	1172	0.89	0.81, 0.98		
ER ⁻												
Q1	721 118	227	1.00	Ref.	0.94	0.43	734 469	287	1.00	Ref.	0.25	0.06
Q2	725 180	302	1.03	0.86, 1.23			724 168	318	1.06	0.90, 1.24		
Q3	722 496	301	0.99	0.82, 1.18			716 332	288	0.93	0.78, 1.10		
Q4	723 410	316	1.01	0.83, 1.22			713 221	288	0.93	0.78, 1.12		
Q5	714 166	292	0.95	0.78, 1.16			718 180	257	0.87	0.71, 1.05		
$P_{heterogeneity}§$						0.70						0.12

TC1, treelet component 1; TC2, treelet component 2; Q1, quintile 1; Q2, quintile 2; Q3, quintile 3; Q4, quintile 4; Q5, quintile 5; Ref., reference category.

*Models were stratified by study centre and age in 1-year categories and adjusted for baseline menopausal status (premenopausal and perimenopausal (reference) or postmenopausal and women who underwent an ovariectomy), baseline alcohol intake (never drinkers (reference), former drinkers, drinkers only at recruitment, lifetime drinkers, unknown), height (continuous), BMI (below (reference) or above 25 kg/m²), schooling level (none, primary (reference), technical/professional/secondary, longer education, unknown/unspecified), age at first full-term pregnancy (nulliparous (reference), ≤21 years, 21–30 years, >30 years, unknown or missing), age at menarche (≤12 years (reference), 12–14 years, >14 years, missing), age at menopause (≤50 years (reference), >50 years, premenopause or missing), use of hormone replacement therapy (never (reference), ever, unknown), level of physical activity (inactive (reference), moderately inactive, moderately active, active, unknown) and alcohol-free energy (continuous).

† P_{LRT} , P values for the likelihood ratio test (LRT) that was used to evaluate the overall significance of a score variable in quintile categories compared with a χ^2 distribution with $df=4$.

‡ P_{trend} , P values obtained by modelling score variables with quintile-specific medians as continuous variables.

§ $P_{heterogeneity}$, P values for BC risks across ER status with $df=1$ obtained using a data augmentation method.

unhealthy pattern, characterized by intakes of high-fat dairy products, red meat, processed meats and French fries, was not associated with BC risk. A recent study of the California Teachers Cohort identified a plant-based pattern, which was related to a reduction of BC risk⁽⁶⁷⁾. In parallel, increasing evidence is accumulating that adherence to the *a priori* defined Mediterranean pattern is associated with a decreased BC risk^(68–70), although results from these studies are not totally consistent, particularly for premenopausal women^(70,71).

The dimension reduction techniques used herein were applied to nutrient densities. Nutrients are present in different combinations of foods, are less country-specific and are directly involved in biological reactions⁽⁷²⁾. By exploring macro- and micronutrients, the present study aimed to provide an exhaustive representation of individuals' diet. Log-transformation was used to address scaling issues that can arise because macro- and micronutrients are expressed in different units. In this way, the variance and the components' decomposition are invariant to the unit of measure. Dietary normalization was achieved using equal energy, i.e. by dividing nutrient intakes by energy intake, minus energy from alcohol intake⁽⁴⁴⁾. Most nutrients are associated with total energy because

either they contribute to total energy directly or because people with higher energy values tend to display larger intakes of specific nutrients^(44,73).

The first two patterns were retained as they were the most interpretable and depicted realistic nutrient patterns that could ultimately be linked with disease risk. The first pattern identified a diet characterized by animal products as opposed to a vegetarian diet, and was associated with a non-significant increase of 5 % in BC risk (TT). TC1 was quite comparable to a Western pattern. Two recent reviews on dietary patterns and BC^(74,75) showed that diets rich in high-fat foods and processed meats were associated with an increased BC risk, although the findings described in both reviews have not been conclusive in this respect with most results reporting a positive association between Western-like dietary pattern and BC being not statistically significant^(74,75). In our study, the micronutrient-dense pattern characterized by a diet rich in vitamins and minerals, akin to a prudent pattern, was associated with an 11 % reduction in BC risk (TT), in line with previous findings^(33,74,75). The protective effect may come from the anti-carcinogenic properties of nutrients such as β -carotene, vitamins C and E, that may exert an antioxidant effect on oestrogen metabolism and reduce cell proliferation⁽⁷⁵⁾. The TT components were

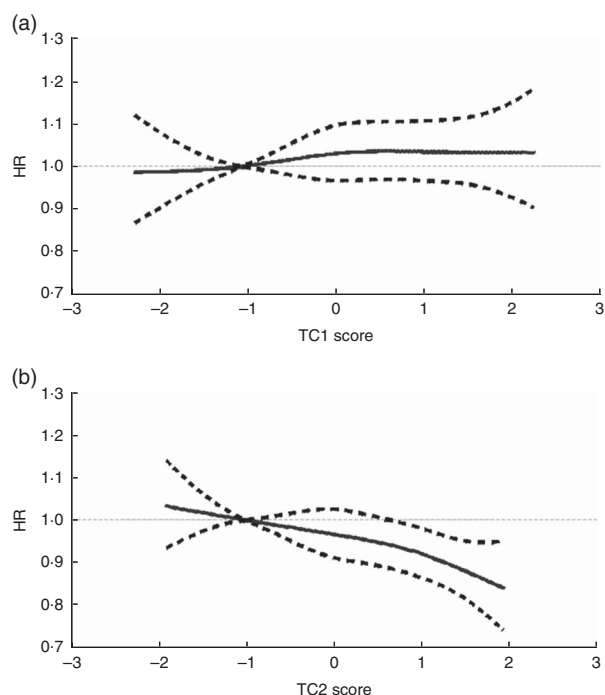


Fig. 2 Relationship between nutrient patterns derived from treelet transform and breast cancer risk (—, hazard ratio (HR); ---, associated 95% CI), obtained by using restrictive cubic splines with values of 1st and 99th percentiles and medians of quintiles 1, 3 and 5 used as knots, among 334 850 women in the European Prospective Investigation into Cancer and Nutrition (EPIC): (a) first treelet component (TC1), $P_{\text{non-linearity}}=0.94$, $P_{\text{trend}}=0.88$; (b) second treelet component (TC2), $P_{\text{non-linearity}}=0.77$, $P_{\text{trend}}=0.02$. Models were stratified by study centre and age in 1-year categories and adjusted for baseline menopausal status (premenopausal and perimenopausal (reference) or postmenopausal and women who underwent an ovariectomy), baseline alcohol intake (never drinkers (reference), former drinkers, drinkers only at recruitment, lifetime drinkers, unknown), height (continuous), BMI (below (reference) or above 25 kg/m²), schooling level (none, primary (reference), technical/professional/secondary, longer education, unknown/unspecified), age at first full-term pregnancy (nulliparous (reference), ≤ 21 years, 21–30 years, >30 years, unknown or missing), age at menarche (≤ 12 years (reference), 12–14 years, >14 years, missing), age at menopause (≤ 50 years (reference), >50 years, pre-menopause or missing), use of hormone replacement therapy (never (reference), ever, unknown), level of physical activity (inactive (reference), moderately inactive, moderately active, active, unknown) and alcohol-free energy (continuous). P_{trend} was obtained by evaluating the joint significance of variables other than the linear one in the model by using Wald's test with $df=3$

highly correlated with those of PCA ($\rho_{\text{TC1,PC1}}=0.91$, $\rho_{\text{TC2,PC2}}=0.86$). TT and PCA provided overall consistent findings in terms of pattern identification and amount of total variability explained. Further analyses were conducted by menopausal status at cohort enrolment, showing no differential association in pre- and postmenopausal women. Analyses carried out by hormonal receptor status showed that the second TT nutrient pattern was related to a significant decrease in BC risk for ER⁺, PR⁺, PR⁻ and ER⁺/PR⁺

tumours. These results are complementary to previous literature findings on dietary patterns and hormonal defined risk of BC^(58,67,70,75). Indeed, Fung *et al.* found that a prudent dietary pattern was linked with decreased ER⁻ risk (relative risk = 0.62, 95% CI 0.45, 0.91)⁽⁷⁶⁾. ER⁻/PR⁻ tumour risk was reduced in postmenopausal women among participants in the highest quintiles of a plant-based pattern and an *a priori* defined Mediterranean diet by 34% and 20%, respectively^(67,70). Results from the Pooling Project of Prospective Studies of Diet and Cancer found a protective association between total fruit or fruit and vegetable consumption in ER⁻ tumours but not in ER⁺ tumours or overall BC risk⁽⁷⁷⁾.

Whereas a large portion of the scientific literature on dietary patterns has used factor analysis or principal component factor analysis⁽⁷⁴⁾, the current paper promotes the use of TT. While PCA produces patterns that are eigenvectors of a covariance/correlation matrix of starting variables, TT is a multivariate technique that yields components by aggregating variables according to covariance/correlation⁽⁷⁸⁾, while at the same time exploring the clustering structure of variables, combining features of PCA with those of cluster analysis. Eventually, TT produces a cluster tree revealing the hierarchical grouping structure of variables. The dendrogram allows a visual inspection of the way different nutrients cluster, possibly easing interpretability of patterns. In addition, loadings are sparse, i.e. some of them are equal to zero as they do not pertain to the clustering node of the component so that a limited number of variables contributes to each treelet component.

In line with other clustering techniques⁽⁷⁹⁾, TT users are confronted with subjective decisions to select the appropriate cut-level for the cluster tree. Information on the grouping structure of variables that have joined (or not) the tree are specific to each level of the TT tree. By choosing a cut-level, the user decides on how much information to extract and the degree of sparsity of the components. If the tree is cut near the 'root', all nutrient variables join the tree. The information would be comparable to PCA output, i.e. all variables would contribute to treelet components. If the tree is cut closer to the 'leaves', i.e. when the cut-level is lower, loadings are sparse as many are equal to zero, possibly making the interpretation easier. By contrast, this may lead to components that do not capture dietary complexity and are therefore not informative. As pointed out by Meinhäuser and Bühlmann, the use of TT leads to a trade-off between amount of variability explained and sparsity. The objective is to 'make the results as sparse as possible but not any sparser'⁽⁴⁸⁾. To identify an optimal cut-level, cross-validation can be used. Once the cut-level is chosen, the loadings computed are invariant to the number of components to be retained; hence keeping n components is an *a priori* parameter to be specified in the cross-validation step.

The present study relied on dietary questionnaires to assess nutrient intakes, which are prone to measurement errors and may lack information on some relevant nutrients. Questionnaires were country-specific, potentially

Table 5 Hazard ratios (HR) and 95 % confidence intervals for breast cancer (BC) by quintiles of pattern scores (first and second components of treelet transform, cut-level 16) for oestrogen receptor-positive + progesterone receptor-positive (ER⁺/PR⁺) and oestrogen receptor-negative + progesterone receptor-negative (ER⁻/PR⁻) tumours in 334 850 women in the European Prospective Investigation into Cancer and Nutrition (EPIC)

Model*	TC1						TC2					
	Person-years	No. of BC cases	HR	95 % CI	$P_{LRT}†$	$P_{trend}‡$	Person-years	No. of BC cases	HR	95 % CI	$P_{LRT}†$	$P_{trend}‡$
ER ⁺ /PR ⁺												
Q1	723 508	568	1.00	Ref.	0.16	0.26	737 812	753	1.00	Ref.	0.15	<0.01
Q2	728 884	811	1.15	1.03, 1.29			727 617	777	0.95	0.86, 1.05		
Q3	725 948	750	1.10	0.98, 1.23			719 931	777	0.94	0.84, 1.04		
Q4	726 667	751	1.11	0.98, 1.25			716 303	720	0.89	0.79, 0.99		
Q5	717 569	773	1.11	0.98, 1.26			720 914	626	0.86	0.76, 0.98		
ER ⁻ /PR ⁻												
Q1	720 830	172	1.00	Ref.	0.60	0.31	734 117	218	1.00	Ref.	0.26	0.08
Q2	724 871	235	1.09	0.89, 1.33			723 844	241	1.05	0.87, 1.26		
Q3	722 003	207	0.93	0.75, 1.15			715 963	207	0.88	0.72, 1.08		
Q4	722 988	222	0.98	0.79, 1.23			712 804	210	0.93	0.76, 1.14		
Q5	713 798	214	0.97	0.77, 1.22			717 762	174	0.85	0.68, 1.06		
$P_{heterogeneity}§$						0.19						0.27

TC1, treelet component 1; TC2, treelet component 2; Q1, quintile 1; Q2, quintile 2; Q3, quintile 3; Q4, quintile 4; Q5, quintile 5; Ref., reference category.

*Models were stratified by study centre and age in 1-year categories and adjusted for baseline menopausal status (premenopausal and perimenopausal (reference) or postmenopausal and women who underwent an ovariectomy), baseline alcohol intake (never drinkers (reference), former drinkers, drinkers only at recruitment, lifetime drinkers, unknown), height (continuous), BMI (below (reference) or above 25 kg/m²), schooling level (none, primary (reference), technical/professional/secondary, longer education, unknown/unspecified), age at first full-term pregnancy (nulliparous (reference), ≤21 years, 21–30 years, >30 years, unknown or missing), age at menarche (≤12 years (reference), 12–14 years, >14 years, missing), age at menopause (≤50 years (reference), >50 years, premenopause or missing), use of hormone replacement therapy (never (reference), ever, unknown), level of physical activity (inactive (reference), moderately inactive, moderately active, active, unknown) and alcohol-free energy (continuous).

† P_{LRT} , P values for the likelihood ratio test (LRT) that was used to evaluate the overall significance of a score variable in quintile categories compared with a χ^2 distribution with $df=4$.

‡ P_{trend} , P values obtained by modelling score variables with quintile-specific medians as continuous variables.

§ $P_{heterogeneity}$, P values for BC risks across ER/PR status with $df=1$ obtained using a data augmentation method.

introducing systematic between-country differences in nutrient assessment. However, in the EPIC study, harmonized composition tables across European countries were used to translate food into nutrient intakes⁽⁴²⁾, thus sizeably improving the comparability of nutrient intakes.

One key element in pattern literature is reproducibility of patterns across populations. With twenty-three centres from ten countries, EPIC accounts for a wide heterogeneity in diet^(80,81). Previous findings in Moskal *et al.*'s study⁽⁴⁵⁾ on the EPIC data showed that more than 75 % of the variance that would be captured by centre-specific PC was captured by PC from overall PCA. This evidence suggested that overall PCA combining data from all EPIC centres allows capturing a good proportion of the variance explained by each EPIC centre. This motivated the choice of applying pattern decomposition on the overall data.

Conclusion

The current study presented results of a nutrient pattern analysis in an international setting using a new tool, TT, and subsequently related the patterns to risk of developing BC. TT is a complementary method to PCA in nutritional epidemiology as it produces readily interpretable sparse components. In the EPIC study, nutrient patterns characterized by a diet rich in macronutrients of animal origin, such as

cholesterol or SFA, were associated with a non-significant increase in BC risk while a diet rich in vitamins, minerals and β -carotene, indicating a more plant-based diet, was associated with a significant decreased BC risk. This decrease was also significant for ER⁺, PR⁺, PR⁻ and ER⁺/PR⁺ tumours.

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Supplementary material

To view supplementary material for this article, please visit <http://dx.doi.org/10.1017/S1368980015000294>

References

- Bray F, Ren JS, Masuyer E *et al.* (2013) Estimates of global cancer prevalence for 27 sites in the adult population in 2008. *Int J Cancer* **132**, 1133–1145.
- Key TJ, Verkasalo PK & Banks E (2001) Epidemiology of breast cancer. *Lancet Oncol* **2**, 133–140.
- Collaborative Group on Hormonal Factors in Breast Cancer (2001) Familial breast cancer: collaborative reanalysis of individual data from 52 epidemiological studies including 58 209 women with breast cancer and 101 986 women without the disease. *Lancet* **358**, 1389–1399.
- Green J, Cairns BJ, Casabonne D *et al.* (2011) Height and cancer incidence in the Million Women Study: prospective cohort, and meta-analysis of prospective studies of height and total cancer risk. *Lancet Oncol* **12**, 785–794.
- Chlebowski R (2007) Lifestyle change including dietary fat reduction and breast cancer outcome. *J Nutr* **137**, 1 Suppl., 233S–235S.
- Anothaisintawee T, Wiratkapun C, Lerdstitthichai P *et al.* (2013) Risk factors of breast cancer: a systematic review and meta-analysis. *Asia Pac J Public Health* **25**, 368–387.
- McKenzie F, Ellison-Loschmann L, Jeffreys M *et al.* (2013) Cigarette smoking and risk of breast cancer in a New Zealand multi-ethnic case-control study. *PLoS One* **8**, e63132.
- Terry PD & Goodman M (2006) Is the association between cigarette smoking and breast cancer modified by genotype? A review of epidemiologic studies and meta-analysis. *Cancer Epidemiol Biomarkers Prev* **15**, 602–611.
- Rohan TE, Heo M, Choi L *et al.* (2013) Body fat and breast cancer risk in postmenopausal women: a longitudinal study. *J Cancer Epidemiol* **2013**, 754815.
- McCullough LE, Eng SM, Bradshaw PT *et al.* (2012) Fat or fit: the joint effects of physical activity, weight gain, and body size on breast cancer risk. *Cancer* **118**, 4860–4568.
- Amadou A, Hainaut P & Romieu I (2013) Role of obesity in the risk of breast cancer: lessons from anthropometry. *J Oncol* **2013**, 906495.
- World Cancer Research Fund/American Institute for Cancer Research (2010) Continuous Update Project Report. Food, Nutrition, Physical Activity, and the Prevention of Breast Cancer. http://www.dietandcancerreport.org/cancer_resource_center/downloads/cu/Breast-Cancer-2010-Report.pdf
- Monninkhof EM, Elias SG, Vlems FA *et al.* (2007) Physical activity and breast cancer: a systematic review. *Epidemiology* **18**, 137–157.
- Fagherazzi G, Vilier A, Boutron-Ruault M-C *et al.* (2014) Alcohol consumption and breast cancer risk subtypes in the E3N-EPIC cohort. *Eur J Cancer Prev* (Epublication ahead of print version).
- Tjønneland A, Christensen J, Olsen A *et al.* (2007) Alcohol intake and breast cancer risk: the European Prospective Investigation into Cancer and Nutrition (EPIC). *Cancer Causes Control* **18**, 361–373.
- Zhang SM, Lee I-M, Manson JE *et al.* (2007) Alcohol consumption and breast cancer risk in the Women's Health Study. *Am J Epidemiol* **165**, 667–676.
- Cui X, Dai Q, Tseng M *et al.* (2007) Dietary patterns and breast cancer risk in the Shanghai breast cancer study. *Cancer Epidemiol Biomarkers Prev* **16**, 1443–1448.
- Levi F, Pasche C, Lucchini F *et al.* (2001) Dietary intake of selected nutrients and breast-cancer risk. *Int J Cancer* **91**, 260–263.
- Doll R (1992) The lessons of life: keynote address to the nutrition and cancer conference. *Cancer Res* **52**, 7 Suppl., 2024S–2029S.
- Anand P, Kunnumakkara AB, Kunnumakara AB *et al.* (2008) Cancer is a preventable disease that requires major lifestyle changes. *Pharm Res* **25**, 2097–2116.
- Ferrari P, Rinaldi S, Jenab M *et al.* (2013) Dietary fiber intake and risk of hormonal receptor-defined breast cancer in the European Prospective Investigation into Cancer and Nutrition study. *Am J Clin Nutr* **97**, 344–353.
- Wu AH, Yu MC, Tseng C-C *et al.* (2003) Green tea and risk of breast cancer in Asian Americans. *Int J Cancer* **106**, 574–579.
- Shrubsole MJ, Jin F, Dai Q *et al.* (2001) Dietary folate intake and breast cancer risk: results from the Shanghai Breast Cancer Study. *Cancer Res* **61**, 7136–7141.

24. Jacques PF & Tucker KL (2001) Are dietary patterns useful for understanding the role of diet in chronic diseases? *Am J Clin Nutr* **73**, 1–2.
25. Hu FB (2002) Dietary pattern analysis: a new direction in nutritional epidemiology. *Curr Opin Lipidol* **13**, 3–9.
26. Jacobs DR & Steffen LM (2003) Nutrients, foods, and dietary patterns as exposures in research: a framework for food synergy. *Am J Clin Nutr* **78**, 3 Suppl., 508S–513S.
27. De Stefani E, Deneo-Pellegrini H, Boffetta P *et al.* (2009) Dietary patterns and risk of cancer: a factor analysis in Uruguay. *Int J Cancer* **124**, 1391–1397.
28. Edefonti V, Bravi F, Garavello W *et al.* (2010) Nutrient-based dietary patterns and laryngeal cancer: evidence from an exploratory factor analysis. *Cancer Epidemiol Biomarkers Prev* **19**, 18–27.
29. Nkondjock A, Krewski D, Johnson KC *et al.* (2005) Dietary patterns and risk of pancreatic cancer. *Int J Cancer* **114**, 817–823.
30. Schulze MB, Hoffmann K, Kroke A *et al.* (2001) Dietary patterns and their association with food and nutrient intake in the European Prospective Investigation into Cancer and Nutrition (EPIC)-Potsdam study. *Br J Nutr* **85**, 363–373.
31. Vrieling A, Buck K, Seibold P *et al.* (2013) Dietary patterns and survival in German postmenopausal breast cancer survivors. *Br J Cancer* **108**, 188–192.
32. De Stefani E, Boffetta P, Ronco AL *et al.* (2008) Nutrient patterns and risk of lung cancer: a factor analysis in Uruguayan men. *Lung Cancer* **61**, 283–291.
33. Brennan SF, Cantwell MM, Cardwell CR *et al.* (2010) Dietary patterns and breast cancer risk: a systematic review and meta-analysis. *Am J Clin Nutr* **91**, 1294–1302.
34. Dahm CC, Gorst-Rasmussen A, Crowe FL *et al.* (2012) Fatty acid patterns and risk of prostate cancer in a case-control study nested within the European Prospective Investigation into Cancer and Nutrition. *Am J Clin Nutr* **96**, 1354–1361.
35. Gorst-Rasmussen A, Dahm CC, Dethlefsen C *et al.* (2011) Exploring dietary patterns by using the treelet transform. *Am J Epidemiol* **173**, 1097–1104.
36. Schoenaker DAJM, Dobson AJ, Soedamah-Muthu SS *et al.* (2013) Factor analysis is more appropriate to identify overall dietary patterns associated with diabetes when compared with Treelet transform analysis. *J Nutr* **143**, 392–398.
37. Jolliffe IT (2002) *Principal Component Analysis*, 2nd ed. New York: Springer-Verlag.
38. Riboli E, Hunt KJ, Slimani N *et al.* (2002) European Prospective Investigation into Cancer and Nutrition (EPIC): study populations and data collection. *Public Health Nutr* **5**, 1113–1124.
39. Riboli E & Kaaks R (1997) The EPIC project: rationale and study design. European Prospective Investigation into Cancer and Nutrition. *Int J Epidemiol* **26**, Suppl. 1, S6–S14.
40. Kaaks R, Slimani N & Riboli E (1997) Pilot phase studies on the accuracy of dietary intake measurements in the EPIC project: overall evaluation of results. European Prospective Investigation into Cancer and Nutrition. *Int J Epidemiol* **26**, Suppl. 1, S26–S36.
41. Margetts B & Pietinen P (1997) European Prospective Investigation into Cancer and Nutrition: validity studies on dietary assessment methods. *Int J Epidemiol* **26**, Suppl. 1, S1–S5.
42. Slimani N, Deharveng G, Unwin I *et al.* (2007) The EPIC nutrient database project (ENDB): a first attempt to standardize nutrient databases across the 10 European countries participating in the EPIC study. *Eur J Clin Nutr* **61**, 1037–1056.
43. Bouckaert KP, Slimani N, Nicolas G *et al.* (2011) Critical evaluation of folate data in European and international databases: recommendations for standardization in international nutritional studies. *Mol Nutr Food Res* **55**, 166–180.
44. Willett WC, Howe GR & Kushi LH (1997) Adjustment for total energy intake in epidemiologic studies. *Am J Clin Nutr* **65**, 4 Suppl., 1220S–1228S.
45. Moskal A, Pisa P, Ferrari P *et al.* (2014) Nutrient patterns and their food sources in an international study setting: report from the EPIC study. *PLoS One* **9**, e98647.
46. Imamura F, Lichtenstein AH, Dallal GE *et al.* (2009) Confounding by dietary patterns of the inverse association between alcohol consumption and type 2 diabetes risk. *Am J Epidemiol* **170**, 37–45.
47. Gorst-Rasmussen A (2011) tt: treelet transform with Stata. *Stata J* **12**, 130–146.
48. Meinshausen N & Bühlmann P (2008) Discussion of: treelets – an adaptive multi-scale basis for sparse unordered data. *Ann Appl Stat* **2**, 478–481.
49. Thiébaud ACM & Bénichou J (2004) Choice of time-scale in Cox's model analysis of epidemiologic cohort data: a simulation study. *Stat Med* **23**, 3803–3820.
50. Ferrari P, Day NE, Boshuizen HC *et al.* (2008) The evaluation of the diet/disease relation in the EPIC study: considerations for the calibration and the disease models. *Int J Epidemiol* **37**, 368–378.
51. Lunn M & McNeil D (1995) Applying Cox regression to competing risks. *Biometrics* **51**, 524–532.
52. Heinzl H & Kaider A (1997) Gaining more flexibility in Cox proportional hazards regression models with cubic spline functions. *Comput Methods Programs Biomed* **54**, 201–208.
53. Männistö S, Dixon LB, Balder HF *et al.* (2005) Dietary patterns and breast cancer risk: results from three cohort studies in the DIETSCAN project. *Cancer Causes Control* **16**, 725–733.
54. Agurs-Collins T, Rosenberg L, Makambi K *et al.* (2009) Dietary patterns and breast cancer risk in women participating in the Black Women's Health Study. *Am J Clin Nutr* **90**, 621–628.
55. Terry P, Suzuki R & Hu FB (2001) A prospective study of major dietary patterns and the risk of breast cancer. *Cancer Epidemiol Biomarkers Prev* **10**, 1281–1285.
56. Sieri S, Krogh V, Pala V *et al.* (2004) Dietary patterns and risk of breast cancer in the ORDET Cohort. *Cancer Epidemiol Biomarkers Prev* **13**, 567–572.
57. Adebamowo CA, Hu FB, Cho E *et al.* (2005) Dietary patterns and the risk of breast cancer. *Ann Epidemiol* **15**, 789–795.
58. Fung TT, Hu FB, Holmes MD *et al.* (2005) Dietary patterns and the risk of postmenopausal breast cancer. *Int J Cancer* **116**, 116–121.
59. Nkondjock A & Ghadirian P (2005) Associated nutritional risk of breast and colon cancers: a population-based case-control study in Montreal, Canada. *Cancer Lett* **223**, 85–91.
60. Velie EM, Schairer C, Flood A *et al.* (2005) Empirically derived dietary patterns and risk of postmenopausal breast cancer in a large prospective cohort study. *Am J Clin Nutr* **82**, 1308–1319.
61. Hirose K, Matsuo K, Iwata H *et al.* (2007) Dietary patterns and the risk of breast cancer in Japanese women. *Cancer Sci* **98**, 1431–1438.
62. Murtaugh MA, Sweeney C, Giuliano AR *et al.* (2008) Diet patterns and breast cancer risk in Hispanic and non-Hispanic white women: the Four-Corners Breast Cancer Study. *Am J Clin Nutr* **87**, 978–984.
63. Wu AH, Yu MC, Tseng C *et al.* (2009) Dietary patterns and breast cancer risk in Asian American women. *Am J Clin Nutr* **89**, 1145–1154.
64. Cottet V, Touvier M, Fournier A *et al.* (2009) Postmenopausal breast cancer risk and dietary patterns in the E3N-EPIC prospective cohort study. *Am J Epidemiol* **170**, 1257–1267.
65. Ronco AL, de Stefani E, Aune D *et al.* (2010) Nutrient patterns and risk of breast cancer in Uruguay. *Asian Pac J Cancer Prev* **11**, 519–524.

66. Edefonti V, Decarli A, La Vecchia C *et al.* (2008) Nutrient dietary patterns and the risk of breast and ovarian cancers. *Int J Cancer* **122**, 609–613.
67. Link LB, Canchola AJ, Bernstein L *et al.* (2013) Dietary patterns and breast cancer risk in the California Teachers Study cohort. *Am J Clin Nutr* **98**, 1524–1532.
68. Trichopoulou A, Bamia C, Lagiou P *et al.* (2010) Conformity to traditional Mediterranean diet and breast cancer risk in the Greek EPIC (European Prospective Investigation into Cancer and Nutrition) cohort. *Am J Clin Nutr* **92**, 620–625.
69. Demetriou CA, Hadjisavvas A, Loizidou MA *et al.* (2012) The Mediterranean dietary pattern and breast cancer risk in Greek-Cypriot women: a case–control study. *BMC Cancer* **12**, 113.
70. Buckland G, Travier N, Cottet V *et al.* (2013) Adherence to the Mediterranean diet and risk of breast cancer in the European prospective investigation into cancer and nutrition cohort study. *Int J Cancer* **132**, 2918–2927.
71. Couto E, Sandin S, Lo M *et al.* (2013) Mediterranean dietary pattern and risk of breast cancer. *PLoS One* **8**, e55374.
72. Edefonti V, Hashibe M, Ambrogi F *et al.* (2012) Nutrient-based dietary patterns and the risk of head and neck cancer: a pooled analysis in the International Head and Neck Cancer Epidemiology consortium. *Ann Oncol* **23**, 1869–1880.
73. Freedman LS, Hartman AM, Kipnis V *et al.* (1997) Comments on: adjustment for total energy intake in epidemiologic studies. *Am J Clin Nutr* **65**, 1229–1231.
74. Edefonti V, Randi G, La Vecchia C *et al.* (2009) Dietary patterns and breast cancer: a review with focus on methodological issues. *Nutr Rev* **67**, 297–314.
75. Albuquerque RCR, Baltar VT & Marchioni DML (2014) Breast cancer and dietary patterns: a systematic review. *Nutr Rev* **72**, 1–17.
76. Fung TT, Rimm EB, Spiegelman D *et al.* (2001) Association between dietary patterns and plasma biomarkers of obesity and cardiovascular disease risk. *Am J Clin Nutr* **73**, 61–67.
77. Jung S, Spiegelman D, Baglietto L *et al.* (2013) Fruit and vegetable intake and risk of breast cancer by hormone receptor status. *J Natl Cancer Inst* **105**, 219–236.
78. Gorst-Rasmussen A, Dahm CC, Dethlefsen C *et al.* (2011) Gorst-Rasmussen *et al.* respond to 'Dietary pattern analysis'. *Am J Epidemiol* **173**, 1109–1110.
79. Krzanowski WJ (2000) *Principles of Multivariate Analysis: A User's Perspective*, 2nd ed. New York: Oxford University Press Inc.
80. Freisling H, Fahey MT, Moskal A *et al.* (2010) Region-specific nutrient intake patterns exhibit a geographical gradient within and between European countries. *J Nutr* **140**, 1280–1286.
81. Slimani N & Margetts B (2009) Nutrient intakes and patterns in the EPIC cohorts from 10 European countries. *Eur J Clin Nutr* **63**, Suppl. 4, S1–S274.