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Trans-ancestry genome-wide association study identifies 12 genetic loci influencing blood pressure and implicates a role for DNA methylation

A full list of authors and affiliations appears at the end of the article.

These authors contributed equally to this work.

Abstract

We carried out a trans-ancestry genome-wide association and replication study of blood pressure phenotypes among up to 320,251 individuals of East Asian, European and South Asian ancestry. We find genetic variants at 12 new loci to be associated with blood pressure ($P = 3.9 \times 10^{-11}$ to 5.0×10^{-21}). The sentinel blood pressure SNPs are enriched for association with DNA methylation at multiple nearby CpG sites, suggesting that, at some of the loci identified, DNA methylation may lie on the regulatory pathway linking sequence variation to blood pressure. The sentinel SNPs at the 12 new loci point to genes involved in vascular smooth muscle (*IGFBP3*, *KCNK3*, *PDE3A* and *PRDM6*) and renal (*ARHGAP24*, *OSRI*, *SLC22A7* and *TBX2*) function. The new and known genetic variants predict increased left ventricular mass, circulating levels of NT-proBNP, and cardiovascular and all-cause mortality ($P = 0.04$ to 8.6×10^{-6}). Our results provide new evidence for the role of DNA methylation in blood pressure regulation.

High blood pressure, which affects more than 1 billion people worldwide, is a major risk factor for myocardial infarction, stroke and chronic kidney disease. Approximately 9 million deaths each year are attributable to high blood pressure, including >50% of deaths from coronary heart disease and stroke^{1,2}. High blood pressure is more prevalent in people of East Asian and South Asian ancestry and is a major contributor to their increased risk of stroke and coronary heart disease^{3,4}. Genome-wide association studies (GWAS) have identified over 50 genetic loci influencing blood pressure in predominantly European populations^{5–16}. A role for epigenetic mechanisms in blood pressure regulation has also been suggested^{17–20}.

We carried out a GWAS in East Asians and South Asians, as well as Europeans, to seek both cosmopolitan and population-specific genetic effects for five blood pressure phenotypes: systolic blood pressure (SBP), diastolic blood pressure (DBP), pulse pressure,

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Correspondence should be addressed to J.C.C. (john.chambers@ic.ac.uk), P.E. (p.elliott@ic.ac.uk), P.v.d.H. (p.van.der.harst@umcg.nl), Jiang He (jhe@tulane.edu), N.K. (nokato@ri.ncgm.go.jp), J.S.K. (j.kooner@ic.ac.uk) or E.S.T. (e_shyong_tai@nuhs.edu.sg).

¹⁸³Present address: Institute of Medical Informatics and Statistics, Christian Albrechts University of Kiel, Kiel, Germany.

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Note: Any Supplementary Information and Source Data files are available in the online version of the paper.

COMPETING FINANCIAL INTERESTS

The authors declare no competing financial interests.

mean arterial pressure (MAP) and hypertension (Supplementary Fig. 1) (ref. 5). We then sought DNA coding and gene regulatory mechanisms, including DNA methylation and gene transcription, to help explain the relationships we observed between sequence variation and blood pressure.

RESULTS

Genome-wide association and replication testing

We used genome-wide association data from 99,994 individuals of East Asian ($n = 31,516$), European ($n = 35,352$) and South Asian ($n = 33,126$) ancestry. Characteristics of the participants and information on the genotyping arrays and imputation are summarized in Supplementary Tables 1–3. Phenotype-specific meta-analysis was carried out separately for East Asian, European and South Asian samples, followed by a meta-analysis across the three ancestral population groups.

The trans-ancestry genome-wide association results identified 4,077 variants with $P < 1 \times 10^{-4}$ against any blood pressure phenotype, distributed among 630 genetic loci. At each locus, we identified the sentinel SNP (the SNP with the lowest P value against any phenotype) and carried out combined analysis with phenotype-specific results from the International Consortium on Blood Pressure (ICBP) GWAS (maximum $n = 87,205$) (refs. 8,9). This analysis identified 19 previously unreported loci where the sentinel SNP had suggestive evidence for association with blood pressure ($P < 1 \times 10^{-7}$; Supplementary Table 4). We performed further testing of these 19 SNPs in additional samples of up to 133,052 individuals (48,268 East Asian, 68,456 European and 16,328 South Asian; Supplementary Table 5). Twelve of the 19 SNPs reached both $P < 0.05$ in replication testing and $P < 1 \times 10^{-9}$ in the combined analysis of data from across all stages (Table 1, Supplementary Figs. 2 and 3, and Supplementary Table 6). We set the threshold for genome-wide significance as $P = 1 \times 10^{-9}$ to provide a conservative Bonferroni correction for testing ~ 2.1 million SNPs against the 5 blood pressure phenotypes, in the 3 ancestry groups and overall.

Regional association plots for the 12 newly identified loci are shown in Figures 1–4 and Supplementary Figure 4; associations of the 12 sentinel SNPs with other blood pressure phenotypes are shown in Supplementary Figure 5 and Supplementary Table 7. There was little evidence for heterogeneity of effect between the ancestry groups in either the genome-wide association or replication data. We also replicated previously reported associations with blood pressure at 23 genetic loci at genome-wide significance; a further 17 loci were associated with blood pressure phenotypes at $P < 0.05$ (Supplementary Fig. 6 and Supplementary Table 8).

In population-specific analyses, we identified two further SNPs (rs9425586 in East Asians and rs13395018 in Europeans) that reached $P < 1 \times 10^{-7}$ against a blood pressure phenotype in their respective discovery meta-analyses. We carried out ancestry-specific testing in the East Asian and European replication samples. Neither SNP reached $P < 0.05$ in replication testing or $P < 1 \times 10^{-9}$ in combined analysis with the discovery data (Supplementary Table 6).

Candidate sequence variants and genes at new loci

Taking advantage of trans-ancestry differences in linkage disequilibrium (LD), we used MANTRA and varLD^{21,22} to narrow the 99% credible SNP sets and facilitate future efforts to identify the causal variants underlying blood pressure variability (Supplementary Figs. 7 and 8, and Supplementary Table 9).

Next, we searched for genetic variants at the newly identified blood pressure loci that might influence protein coding or gene transcription and that were in high LD ($r^2 > 0.8$) with sentinel blood pressure SNPs. We identified SNPs that were nonsynonymous ($n = 9$) or splicing variants ($n = 2$) and/or were present in regulatory regions (including transcription factor binding sites, promoter and enhancer regions, DNase I hypersensitivity sites, regulatory motifs and CpG islands; $n = 825$; Supplementary Table 10) (refs. 23,24).

Analysis of coding variation and gene regulatory signatures (Supplementary Tables 10 and 11) identified 20 genes as possible candidates underlying the associations with blood pressure at the newly identified loci (Table 1). Current knowledge on gene function for all 20 candidates is summarized in Supplementary Table 12.

Association of sentinel SNPs with DNA methylation

We investigated the relationships of the sentinel blood pressure SNPs with local DNA methylation (within 1 Mb of each SNP) in 1,904 South Asians with whole-genome methylation data available (peripheral blood; Illumina HumanMethylation450 BeadChip (450K) array; Supplementary Table 13). We found a ~2-fold enrichment for association between sequence variation and DNA methylation in comparison with expectations under the null hypothesis ($P = 0.01$; Supplementary Fig. 9). Twenty-eight of the 35 sentinel blood pressure SNPs were associated with one or more methylation markers at $P < 3.8 \times 10^{-6}$ ($P < 0.05$ after Bonferroni correction for the 13,275 SNP-CpG association tests; Supplementary Table 14); the 28 leading CpG sites (the CpG sites with the lowest P value for association with each sentinel blood pressure SNP) are summarized in Table 2. All 28 leading CpG sites showed replication in further testing among 4,780 European and South Asian samples ($P < 0.05$ and consistent direction of effect; Supplementary Table 15). Regional plots of DNA methylation are shown in Figures 1–4. There was little evidence for heterogeneity of effect of SNPs on methylation between Europeans and South Asians (Supplementary Fig. 10).

We found evidence of replication of the relationships of the sentinel blood pressure SNPs with methylation of their respective leading CpG sites in genomic DNA from cord blood ($P = 4.0 \times 10^{-4}$, binomial test for directionally consistent effects, $n = 237$ samples; Supplementary Table 16). The presence of these associations at an early stage of life, before substantial environmental exposure, lends support to the view that the sequence variants have a direct effect on DNA methylation and argues against reverse causation. We separately showed that association of sentinel SNPs with local DNA methylation is generalizable across multiple phenotypic traits and not unique to blood pressure phenotypes (Supplementary Fig. 11).

Sequence variation, DNA methylation and blood pressure

We used genetic association and the concept of Mendelian randomization to test whether DNA methylation might contribute, at least in part, to the relationship of the sentinel SNPs with blood pressure. For the 28 sentinel SNPs that were associated with methylation of *cis* CpG sites (Supplementary Table 15), we quantified the three-way relationships between the sentinel SNPs, their leading CpG sites and blood pressure among the 6,757 Europeans and South Asians with DNA methylation data available (Supplementary Table 17). Across all 28 loci, we found that the observed effects of SNPs on blood pressure were correlated with the effects predicted through association with methylation ($r = 0.52$; $P = 0.005$; Fig. 5). Of the 14 sentinel SNPs with the highest predicted genetic effects (above the median for the distribution), 13 were directionally consistent ($P = 1.2 \times 10^{-4}$, sign test), with a close correlation between the observed and predicted effects ($r = 0.72$; $P = 0.004$). Our results support the view that DNA methylation may be involved in the regulatory pathway linking DNA sequence variants to blood pressure.

Fine mapping the association of SNPs and DNA methylation

The 450K methylation array assays ~2% of the estimated ~30 million CpG sites in the human genome. To further evaluate the relationship between the sentinel blood pressure SNPs and DNA methylation at the 19p13.3 locus near *AMH*, we used next-generation sequencing to fine map DNA methylation at all CpG sites within 1 kb on either side of the leading 450K CpG site in 168 samples. We successfully quantified DNA methylation at 34 CpG sites, of which only 2 are assayed by the 450K array (Supplementary Fig. 12). The sentinel blood pressure SNP at the *AMH* locus (rs740406) had a directionally consistent effect on methylation at 29 of the 34 CpG sites assayed ($P = 4 \times 10^{-5}$, sign test; Supplementary Fig. 13), consistent with published data suggesting that clusters of adjacent CpG sites are co-regulated^{25,26}. Of the 34 CpG sites assayed, we found that 28 had a positive relationship with blood pressure ($P = 2 \times 10^{-4}$, sign test), and 10 were associated with blood pressure at $P < 0.05$ ($P = 5 \times 10^{-7}$ for enrichment; Supplementary Fig. 13).

Cross-tissue patterns of DNA methylation

DNA methylation can show tissue-specific patterns that contribute to differences in transcriptional regulation and cellular differentiation²⁷. We investigated the cross-tissue patterns of DNA methylation at the 26 leading CpG sites associated with the sentinel blood pressure SNPs in the present study. Using data from 7 tissue samples (including muscle, liver, and subcutaneous and visceral fat), we showed that DNA methylation in blood at the 26 CpG sites was closely correlated with methylation in a wide range of tissues (Pearson correlation coefficient, 0.61–0.97; $P = 1.2 \times 10^{-4}$ to 1.3×10^{-47} ; Supplementary Figs. 14 and 15). Our findings support the view that, for the CpG sites examined, methylation levels in blood provide a surrogate for patterns of methylation in other tissues.

Clinical relevance of our findings

We tested whether the genetic variants singly or in aggregate contribute to risk of clinical phenotypes associated with high blood pressure. In single-variant tests, we found that the 35 (known and new) sentinel SNPs were enriched for variants associated with adiposity, type 2

diabetes, coronary heart disease and kidney function in published GWAS ($P = 2.5 \times 10^{-3}$ to 1.8×10^{-11} ; Supplementary Table 18). We further showed that weighted genetic risk scores comprising known and new variants predicted increased left ventricular mass by electrocardiographic criteria, circulating levels of NT-proBNP (a marker of heart function), clinical coronary heart disease, and cardiovascular and all-cause mortality ($P = 0.04$ to 8.6×10^{-6} ; Supplementary Table 19). Our findings provide evidence that the genetic loci associated with blood pressure contribute to cardiovascular outcomes.

DISCUSSION

Our genome-wide association and replication study in 320,251 people identifies 12 new genetic loci influencing blood pressure phenotypes in 3 ancestry groups. Among the genetic loci and candidate genes identified, several have been implicated in other cardiovascular and metabolic phenotypes through genome-wide association. *IGFBP3*, *KCNK3*, *PDE3A* and *PRDM6* have a role in vascular smooth muscle cell biology. *PDE3A* is a phosphodiesterase involved in cyclic GMP (cGMP) metabolism, vascular smooth muscle contraction and cardiovascular function²⁸. Pharmacological inhibitors of *PDE3A* lower blood pressure²⁹. *KCNK3* is a potassium channel involved in the regulation of vascular tone; mutations in *KCNK3* are associated with pulmonary hypertension³⁰. *PRDM6* acts as an epigenetic regulator of vascular smooth muscle cell phenotypic plasticity by suppressing differentiation and maintaining proliferative potential. Genetic variants near *PRDM6* are associated with intracranial aneurysm³¹. *IGFBP3* modulates the actions of insulin-like growth factors (IGFs), circulating hormones that influence vascular smooth muscle cell function. Serum levels of *IGFBP3* are associated with cardiovascular disease³². We also note several candidate genes that are involved in renal function, a determinant of blood pressure. *ARHGAP24* influences podocyte formation³³, *OSR1* encodes a transcription factor that influences renal mass and function³⁴, and *SLC22A7* encodes a key renal solute transporter³⁵; genetic variants at *TBX2* are determinants of renal function and chronic kidney disease³⁶.

The mechanisms underlying the associations between common genetic variants and blood pressure are incompletely understood. The majority of the loci identified do not contain common or low-frequency coding variants to account for the association between the sentinel SNP and blood pressure. Using both the 450K methylation array and fine mapping through targeted bisulfite sequencing, we show that SNPs influencing blood pressure are associated with methylation at multiple local CpG sites and that DNA methylation is associated with blood pressure. Using genetic association and the concept of Mendelian randomization, we further show that the observed effect of SNPs on blood pressure is closely correlated with the effect predicted through association with methylation. The effects of genetic variation on methylation can be demonstrated in the newborn, in the absence of substantial adverse environmental exposures, further supporting a causal relationship. Our results suggest that DNA methylation may be involved in the regulatory pathway linking common genetic variants with blood pressure at some of the loci identified, consistent with findings from experimental models of hypertension³⁷. We note an effect of genome-wide associated sentinel SNPs on DNA methylation for traits in addition to blood pressure, suggesting that DNA methylation might have a wider role in linking common genetic variation to multiple phenotypes.

URLs

Sequenom EpiDesigner BETA, <http://www.epidesigner.com/>.

ONLINE METHODS

Populations and phenotypes

Details of the participating cohorts are summarized in Supplementary Table 1 and in the Supplementary Note. Phenotype definitions were based on the published literature⁶. SBP, DBP, pulse pressure and MAP were continuous variables measured in millimeters of mercury. SBP and DBP were directly measured in millimeters of mercury, and pulse pressure and MAP were calculated by $SBP - DBP$ and $(2 \times DBP + SBP)/3$, respectively. SNP associations for SBP, DBP, pulse pressure and MAP were tested by linear regression with age and sex using an additive genetic model. For individuals being treated with blood pressure-lowering medication, the following adjustments to the blood pressure values were made before performing the regression analysis: SBP (+15), DBP (+10), pulse pressure (+5) and MAP (+11.667). For hypertension, logistic regression with sex as a covariate was applied, with cases and controls defined as follows: cases: (i) SBP ≥ 160 mm Hg or DBP ≥ 100 mm Hg or on antihypertensive treatment and (ii) age of onset ≤ 65 years; controls: (i) SBP < 130 mm Hg and DBP < 85 mm Hg and not on antihypertensive treatment and (ii) age ≥ 50 years). Data and sample collection by the cohorts participating in the study was approved by respective research ethics committees, and all research participants gave written consent to take part.

Genome-wide association

Genome-wide association was analyzed in a total of 99,994 subjects, of whom 31,516 were of East Asian ancestry, 35,352 were of European ancestry and 33,126 were of South Asian ancestry. Imputation was carried out using haplotypes from HapMap Phase 2. Details of genotyping arrays and imputation are summarized in Supplementary Table 2. Quality control checks included a check of the distribution of effect sizes across phenotypes and comparison of allele frequencies against those expected from HapMap populations. There were between 2,127,883 (SBP) and 2,166,286 (hypertension) SNPs for analysis after quality control. Genomic control inflation factors ranged from 1.01 to 1.09 in the ancestry-specific meta-analyses and from 1.05 to 1.12 in global analyses (Supplementary Table 3).

Genome-wide significance was inferred at $P < 1 \times 10^{-9}$. This conservative choice fully corrects for the ~ 10 million SNP-phenotype combinations tested, in 3 ancestry groups and overall, and makes no adjustment for the potential correlations between the SNPs or phenotypes tested. We adopted this strategy to ensure that the results reported are robust and to reduce the risk of spurious findings in our multi-stage trans-ancestry GWAS.

Associations of SNPs with phenotype were tested in each cohort separately in single-marker tests, using regression analysis and an additive genetic model. Principal components and other study-specific factors were included as covariates to account for population substructure as described in Supplementary Table 2. Test statistics from each cohort were then corrected for their respective genomic control inflation factor to adjust for residual

population substructure; the genomic control inflation factors are summarized in Supplementary Table 3. We then performed inverse variance meta-analysis of the results from the individual cohorts; meta-analysis was carried out among East Asian, European and South Asian populations separately. SNPs with information score <0.5 and minor allele frequency (MAF) $<1\%$ (weighted average across the cohorts) as well as sample size $<50\%$ of the maximum n for the phenotype were removed. We also removed SNPs showing heterogeneity of effect ($P_{\text{het}} < 1 \times 10^{-8}$) within any one of the three ancestry groups.

Finally, we carried out inverse variance meta-analyses of the results from the three ancestry groups. There was little evidence for inflation of test statistics at SNPs not known to be associated with blood pressure phenotypes, and genomic control was not applied to the final meta-analysis results.

Identification of candidate SNPs

We identified all common genetic variants that were in LD with one or more of the sentinel SNPs at $r^2 > 0.8$. LD was calculated using pooled haplotypes for (i) European and East Asian samples in the 1000 Genomes Project data set (March 2012 release) and (ii) 168 South Asians with whole-genome sequence data. We annotated the sentinel SNPs and their proxies for regulatory regions (promoter and enhancer histone marks, DNase I hypersensitivity, protein binding and regulatory motifs) with HaploRegv2 (Broad Institute)²⁴. VEP (Variant Effect Predictor) was used for the identification of transcription factor binding sites and nonsynonymous and splicing variants²³. EpiExplorer and the UCSC Genome Browser were used to annotate CpG islands³⁸.

Identification of candidate genes

We considered the nearest gene and any other gene located within 10 kb of the sentinel SNP to be candidates for mediating the association with the blood pressure phenotype, along with any gene containing a SNP predicted to be nonsynonymous or affecting a splice site. We also examined the associations of the sentinel SNPs and their proxies with eQTL data from Zeller *et al.*, consisting of data from circulating monocytes in 1,490 unrelated individuals³⁹. SNPs were tested for association with the expression of nearby genes (within 1 Mb of the sentinel SNP; $P < 1 \times 10^{-5}$). Finally, for significant SNP-methylation associations, the gene nearest the leading CpG site was also included as a candidate.

Association between sentinel SNPs, DNA methylation and phenotype

The associations of the 36 sentinel blood pressure SNPs with DNA methylation were first examined among 1,904 South Asian individuals from the LOLIPOP cohort. Bisulfite conversion of genomic DNA was performed using the EZ DNA methylation kit according to the manufacturer's instructions (Zymo Research). Methylation of genomic DNA was quantified using the Illumina HumanMethylation450 array according to the manufacturer's instructions. To facilitate the comparison of effects between CpG sites, methylation levels were z -transformed for all analyses; the scale for methylation is thus 'standard deviations'. Whole-genome genotyping was carried out using the Illumina 317, 610 or OmniExome microarray, with genomic DNA and according to the manufacturer's instructions. SNPs and samples with low call rates ($<98\%$) were excluded, as were SNPs with departure from

Hardy-Weinberg equilibrium ($P < 1 \times 10^{-6}$). We used IMPUTE2 to predict (impute) unmeasured genotypes, using phased haplotypes from the whole-genome sequencing of 168 Indian Asians as a reference panel.

The association of the sentinel blood pressure SNPs with *cis* DNA methylation (within 1 Mb) was tested by linear regression and an additive genetic model. We used an analytic strategy validated to reduce batch and other technical confounding effects in quantification of DNA methylation and adjusted for the white blood cell composition of blood^{40–42}. We inferred statistical significance at $P < 3.8 \times 10^{-6}$ (Bonferroni correction for 13,275 SNP–CpG marker tests). We identified the leading CpG site (having the lowest P value for association with the sentinel SNP) at each blood pressure locus. We then carried out replication testing of the leading SNP–CpG associations among independent samples of South Asians (LOLIPOP, $n = 1,373$) and Europeans (LOLIPOP, $n = 166$; LifeLines Deep, $n = 752$; RS-BIOS, $n = 762$; KORA, $n = 1,727$; Supplementary Table 13).

Next, we quantified the relationship of the 28 leading CpG sites with blood pressure (Supplementary Tables 15 and 17). We then calculated the predicted effect of each SNP on blood pressure as the product of the regression coefficients between (i) the SNP and methylation ($n = 6,684$) and (ii) methylation and blood pressure ($n = 6,757$). We used linear regression and sign tests to compare the predicted effect of a SNP on blood pressure via methylation with the directly observed effect of this SNP on blood pressure in genome-wide association (Fig. 5).

Association of methylation with gene expression

The relationship between methylation and the expression of nearest genes was investigated in samples from LOLIPOP ($n = 1,082$; 907 South Asians and 175 Europeans) and the EnviroGenoMarkers project, a nested case-control study of incident breast cancer and B cell leukemia ($n = 638$ Europeans)^{43,44}.

LOLIPOP—Details of the LOLIPOP cohort and methylation analysis have been provided above. Gene expression analysis was performed with the Illumina HumanHT-12 v4 BeadChip according to the manufacturer's protocol. Background correction using negative controls was performed, and data were subsequently quantile normalized and \log_2 transformed. Linear models were fitted with \log_2 -transformed gene expression as the response variable and quantile normalized with β values (methylation), age, sex, the top 24 control probe principal components from methylation measurement and technical covariates related to the measurement of expression, including RNA integrity number (RIN), RNA extraction batch, RNA conversion batch, scanning batch, array and array position. Analyses were conducted separately in South Asians and Europeans, followed by inverse variance-weighted meta-analysis. Calculations were performed using R, version 3.0.1.

EnviroGenoMarkers—Methylation and gene expression were quantified in the baseline blood samples collected 1–17 years before disease onset. Transcriptomic profiles were obtained using the Agilent 4x44K Whole Human Genome microarray and subjected to extensive quality control procedures⁴⁵. DNA methylation profiles were obtained using the Illumina Infinium HumanMethylation450 BeadChip according to the manufacturer's

protocol. Bisulfite conversion was carried out using the Zymo EZ DNA Methylation kit. Probes that had missing values in more than 20% of the samples were excluded. We used linear regression to determine the association between methylation and gene transcription.

Enrichment of reported sentinel SNPs for association with DNA methylation

SNPs reported to be associated with phenotype were retrieved from the National Human Genome Research Institute (NHGRI) GWAS catalog. We considered studies with a sample size greater than 1,000 and retained SNPs with association $P < 5 \times 10^{-8}$. For simplicity, we removed data for Crohn's disease and ulcerative colitis (both represented by inflammatory bowel disease) and obesity (represented by body mass index (BMI)). To account for biases due to LD, SNPs were pruned for each trait on the basis of a 1-Mb flanking window (by consecutively selecting the SNP with the lowest P value and removing any variant within 1 Mb). Traits were then ranked by the number of significant associations, and the top 20 traits were tested for enrichment with methylation quantitative trait locus (methQTL) SNPs. For this purpose, we derived 1 million sets of matched background SNPs for each trait. These background SNPs were chosen randomly but had properties matched to the associated SNPs (MAF $\pm 2\%$, distance to gene ± 10 kb, CpGs in *cis* ± 200 kb). The proportion of *cis* methQTLs among the associated SNPs was then compared to the proportion among each of the 1 million sets of background SNPs, thereby deriving an empirical P value.

Cross-tissue methylation

Publicly available data (GSE48472) were downloaded from the Gene Expression Omnibus (GEO)⁴⁶. Briefly, the data set consisted of 41 samples from blood, liver, muscle, pancreas, subcutaneous fat, omentum and spleen analyzed on the 450K methylation array. Data from the 28 CpG sites of interest were extracted and plotted using the heatmap.2 function in the gplots library with R. Mean methylation levels for each CpG site across all samples within each tissue type were used to test for pairwise correlation between tissue types.

Relationship of sentinel SNPs with methylation in cord blood

We tested the relationship of sentinel SNPs with methylation for the 28 SNP-CpG pairs of interest in cord blood to investigate whether reverse causation might account for the observed associations between SNPs and methylation. This analysis was conducted in the GUSTO (Growing Up in Singapore Toward Healthy Outcomes) study⁴⁷. Extracted DNA from cord blood ($n = 237$ samples) was genotyped using the Illumina OmniExpress + exome array, and DNA methylation profiling was performed using the Infinium HumanMethylation450 BeadChip. Data were processed as described⁴⁸. Both data sets have been described previously and are deposited in GEO under accessions GSE53816 and GSE54445 (ref. 49). Genotype data were imputed with reference to HapMap 2 East Asian populations. SNPs with MAF $< 1\%$ in GUSTO and CpGs that failed quality control were excluded from further analysis. Linear regression was used to quantify SNP-CpG associations, adjusting for sex.

Targeted resequencing for regional methylation

The 450K array assays <2% of the estimated ~30 million CpG sites in the human genome. To better describe the patterns of regional methylation, we carried out resequencing of the *AMH* locus in 168 samples. We used sequence capture and next-generation sequencing to assay 34 predicted CpG sites within 1 kb of the sentinel methylation marker at the *AMH* locus (chr. 19, 2,250,061–2,252,061). Primers were designed using Sequenom EpiDesigner BETA. Target DNA enrichment was carried out using the Fluidigm 48.48 Access Array IFC system, followed by PCR to attach sequence-specific adaptors and sample barcodes. Pooled sequencing was performed using the Illumina MiSeq platform (300-bp paired-end runs). We then used Burrows-Wheeler Aligner to map the directional, paired-end Illumina sequencing reads to the reference genome (hg19 build) and quantified methylation from the frequencies of converted and unconverted cytosine residues observed in reads mapped to each CpG site.

Fine mapping

To take advantage of any variation in LD structure between ancestry groups, we used MANTRA and varLD for further trans-ancestry fine mapping^{21,22}. MANTRA, a Bayesian approach, allows for heterogeneity in effect sizes between ancestry or ethnic groups, which arises as a result of underlying differences in LD patterns but with a shared underlying causal variant across diverse populations that cannot be accommodated in fixed-effects meta-analysis. At each locus, 99% credible SNP sets were also constructed, which can be interpreted in a similar way to confidence intervals in a frequentist statistical framework^{21,50}.

Genetic risk scores

We calculated weighted genetic risk scores for each of the 5 blood pressure phenotypes, using all 35 sentinel SNPs reaching genome-wide significance or the 12 sentinel SNPs from the newly identified genetic loci; this yielded 10 genetic risk scores per person. Each score was calculated as the sum of the effect allele counts weighted by β coefficients for association with the respective phenotype. To facilitate comparisons between genetic risk scores, each score was then standardized. We examined the relationships between genetic risk scores and phenotypes relevant to blood pressure in three cohorts—LOLIPOP, LifeLines and PREVEND—using regression analysis, including age and sex as covariates. Results were combined across cohorts by inverse variance meta-analysis where necessary. Where possible, we also used the *in silico* approach from T. Johnson for comparison⁸.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Authors

Norihiro Kato^{#1}, Marie Loh^{#2,3,4}, Fumihiko Takeuchi^{#1}, Niek Verweij^{#5}, Xu Wang^{#6}, Weihua Zhang^{#3,7}, Tanika N Kelly^{#8}, Danish Saleheen^{#9,10,11}, Benjamin Lehne^{#3}, Irene Mateo Leach^{#5}, Alexander W Drong¹², James Abbott¹³, Simone Wahl^{14,14,16}, Sian-Tsung Tan^{7,17}, William R Scott^{3,17}, Gianluca Campanella³, Marc Chadeau-

Hyam³, Uzma Afzal^{3,7}, Tarunveer S Ahluwalia^{18,19,20}, Marc Jan Bonder²¹, Peng Chen⁶, Abbas Dehghan²², Todd L Edwards²³, Tõnu Esko^{24,25,26,27}, Min Jin Go²⁸, Sarah E Harris^{29,30}, Jaana Hartiala^{31,32}, Silva Kasela²⁴, Anuradhani Kasturiratne³³, Chiea-Chuen Khor^{6,34,35,36}, Marcus E Kleber³⁷, Huaixing Li³⁸, Zuan Yu Mok³⁹, Masahiro Nakatochi⁴⁰, Nur Sabrina Sapari³⁹, Richa Saxena⁴¹, Alexandre F R Stewart^{42,43}, Lisette Stolk⁴⁴, Yasuharu Tabara⁴⁵, Ai Ling Teh⁴⁶, Ying Wu⁴⁷, Jer-Yuarn Wu^{48,49}, Yi Zhang^{50,51}, Imke Aits⁵², Alexessander Da Silva Couto Alves⁵³, Shikta Das⁵³, Rajkumar Dorajoo³⁴, Jemma C Hopewell⁵⁴, Yun Kyoung Kim²⁸, Robert W Koivula⁵⁵, Jian'an Luan⁵⁶, Leo-Pekka Lyytikäinen^{57,58}, Quang N Nguyen⁵⁹, Mark A Pereira⁶⁰, Iris Postmus^{61,62}, Olli T Raitakari^{63,64}, Molly Scannell Bryan⁶⁵, Robert A Scott⁵⁶, Rossella Sorice⁶⁶, Vinicius Tragante⁶⁷, Michela Traglia^{68,69}, Jon White⁷⁰, Ken Yamamoto⁷¹, Yonghong Zhang⁷², Linda S Adair⁷³, Alauddin Ahmed⁷⁴, Koichi Akiyama¹, Rasheed Asif⁹, Tin Aung⁷⁵, Inês Barroso^{76,77,78}, Andrew Bjornes⁴¹, Timothy R Braun⁷⁹, Hui Cai^{80,81}, Li-Ching Chang⁴⁸, Chien-Hsiun Chen^{48,49}, Ching-Yu Cheng^{6,35,75,82}, Yap-Seng Chong^{46,83}, Rory Collins⁵⁴, Regina Courtney^{80,81}, Gail Davies^{30,84}, Graciela Delgado³⁷, Loi D Do⁵⁹, Pieter A Doevendans⁶⁷, Ron T Gansevoort⁸⁵, Yu-Tang Gao⁸⁶, Tanja B Grammer³⁷, Niels Grarup¹⁸, Jagvir Grewal^{3,7}, Dongfeng Gu⁸⁷, Gurpreet S Wander⁸⁸, Anna-Liisa Hartikainen^{89,90,91}, Stanley L Hazen^{92,93}, Jing He⁹⁴, Chew-Kiat Heng⁹⁵, James E Hixson⁹⁶, Albert Hofman²², Chris Hsu⁹⁷, Wei Huang⁹⁸, Lise L N Husemoen⁹⁹, Joo-Yeon Hwang²⁸, Sahoko Ichihara¹⁰⁰, Michiya Igase¹⁰¹, Masato Isono¹, Johanne M Justesen¹⁸, Tomohiro Katsuya^{102,103}, Muhammad G Kibriya⁶⁵, Young Jin Kim²⁸, Miyako Kishimoto¹⁰⁴, Woon-Puay Koh^{6,105}, Katsuhiko Kohara¹⁰¹, Meena Kumari¹⁰⁶, Kenneth Kwek¹⁰⁷, Nanette R Lee^{108,109}, Jeannette Lee⁶, Jiemin Liao^{35,75}, Wolfgang Lieb⁵², David C M Liewald^{30,84}, Tatsuki Matsubara¹¹⁰, Yumi Matsushita¹⁰⁴, Thomas Meitinger^{111,112}, Evelin Mihailov²⁴, Lili Milani²⁴, Rebecca Mills⁷, Nina Mononen^{57,58}, Martina Müller-Nurasyid^{113,114,115}, Toru Nabika¹¹⁶, Eitaro Nakashima^{117,118}, Hong Kiat Ng³⁹, Kjell Nikus¹¹⁹, Teresa Nutile⁶⁶, Takayoshi Ohkubo¹²⁰, Keizo Ohnaka¹²¹, Sarah Parish⁵⁴, Lavinia Paternoster¹²², Hao Peng⁷², Annette Peters^{15,115}, Son T Pham⁵⁹, Mohitha J Pinidiyapathirage³³, Mahfuzar Rahman^{74,123}, Hiromi Rakugi¹⁰³, Olov Rolandsson¹²⁴, Michelle Ann Rozario³⁹, Daniela Ruggiero⁶⁶, Cinzia F Sala⁶⁸, Ralhan Sarju⁸⁸, Kazuro Shimokawa¹, Harold Snieder¹²⁵, Thomas Sparsø¹⁸, Wilko Spiering¹²⁶, John M Starr^{30,127}, David J Stott¹²⁸, Daniel O Stram⁹⁷, Takao Sugiyama¹²⁹, Silke Szymczak^{130,183}, W H Wilson Tang^{92,131}, Lin Tong⁶⁵, Stella Trompet¹³², Väinö Turjanmaa^{133,134}, Hirotugu Ueshima^{135,136}, André G Uitterlinden^{22,44}, Satoshi Umemura¹³⁷, Marja Vaarasmaki^{89,90,91}, Rob M van Dam⁶, Wiek H van Gilst⁵, Dirk J van Veldhuisen⁵, Jorma S Viikari^{138,139}, Melanie Waldenberger^{14,15}, Yiqin Wang³⁸, Aili Wang⁷², Rory Wilson^{14,15}, Tien-Yin Wong^{35,75}, Yong-Bing Xiang⁸⁶, Shuhei Yamaguchi¹⁴⁰, Xingwang Ye³⁸, Robin D Young¹⁰, Terri L Young^{141,142}, Jian-Min Yuan¹⁴³, Xueya Zhou^{144,145,146}, Folkert W Asselbergs^{67,147,148}, Marina Ciullo⁶⁶, Robert Clarke⁵⁴, Panos Deloukas^{149,150}, Andre Franke¹³⁰, Paul W Franks^{55,124,151}, Steve Franks¹⁵², Yechiel Friedlander¹⁵³, Myron D Gross¹⁵⁴, Zhirong Guo⁷², Torben Hansen¹⁸, Marjo-Riitta Jarvelin^{2,53,91,155,156}, Torben Jørgensen⁹⁹, J Wouter Jukema^{132,147,157}, Mika

kähönen^{133,134}, Hiroshi Kajio¹⁰⁴, Mika Kivimaki¹⁰⁶, Jong-Young Lee^{158,159}, Terho Lehtimäki^{57,58}, Allan Linneberg^{99,160,161}, Tetsuro Miki¹⁰¹, Oluf Pedersen¹⁸, Nilesh J Samani^{162,163}, Thorkild I A Sørensen^{18,122,164}, Ryoichi Takayanagi¹⁶⁵, Daniela Toniolo^{68,166}, BIOS-consortium¹⁶⁷, CARDIo GRAMplusCD¹⁶⁷, LifeLines Cohort Study¹⁶⁷, The InterAct Consortium¹⁶⁷, Habibul Ahsan⁶⁵, Hooman Allayee^{31,32}, Yuan-Tsong Chen⁴⁸, John Danesh^{10,168}, Ian J Deary^{30,84}, Oscar H Franco²², Lude Franke²¹, Bastiaan T Heijman¹⁶⁹, Joanna D Holbrook⁴⁶, Aaron Isaacs²², Bong-Jo Kim²⁸, Xu Lin³⁸, Jianjun Liu^{6,34}, Winfried März^{37,170,171}, Andres Metspalu²⁴, Karen L Mohlke⁴⁷, Dharambir K Sanghera⁷⁹, Xiao-Ou Shu^{80,81}, Joyce B J van Meurs⁴⁴, Eranga Vithana^{35,75,141}, Ananda R Wickremasinghe³³, Cisca Wijmenga²¹, Bruce H W Wolffenbuttel¹⁷², Mitsuhiro Yokota¹⁷³, Wei Zheng^{80,81}, Dingliang Zhu^{50,51}, Paolo Vineis³, Soterios A Kyrtopoulos¹⁷⁴, Jos C S Kleinjans¹⁷⁵, Mark I McCarthy^{12,176}, Richie Soong^{39,177}, Christian Gieger^{14,15}, James Scott¹⁷, Yik-Ying Teo^{#6,34,75,178,179,180}, Jiang He^{#8}, Paul Elliott^{#3}, E Shyong Tai^{#6,102,181}, Pim van der Harst^{#5,21,147}, Jaspal S Kooner^{#7,17,182}, and John C Chambers^{#3,7,182}

Affiliations

¹Department of Gene Diagnostics and Therapeutics, Research Institute, National Center for Global Health and Medicine, Tokyo, Japan. ²Institute of Health Sciences, University of Oulu, Oulu, Finland. ³Department of Epidemiology and Biostatistics, Imperial College London, London, UK. ⁴Translational Laboratory in Genetic Medicine, Agency for Science, Technology and Research (A*STAR), Singapore. ⁵Department of Cardiology, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands. ⁶Saw Swee Hock School of Public Health, National University of Singapore and National University Health System, Singapore. ⁷Ealing Hospital National Health Service (NHS) Trust, Middlesex, UK. ⁸Department of Epidemiology, Tulane University School of Public Health and Tropical Medicine, New Orleans, Louisiana, USA. ⁹Center for Non-Communicable Diseases, Karachi, Pakistan. ¹⁰Department of Public Health and Primary Care, University of Cambridge, Strangeways Research Laboratory, Cambridge, UK. ¹¹Cardiovascular Institute, Perelman School of Medicine at the University of Pennsylvania, Philadelphia, Pennsylvania, USA. ¹²Wellcome Trust Centre for Human Genetics, University of Oxford, Oxford, UK. ¹³Bioinformatics Support Service, Imperial College London, London, UK. ¹⁴Research Unit of Molecular Epidemiology, Helmholtz Zentrum München–German Research Center for Environmental Health, Neuherberg, Germany. ¹⁵Institute of Epidemiology II, Helmholtz Zentrum München–German Research Center for Environmental Health, Neuherberg, Germany. ¹⁶German Center for Diabetes Research (DZD), Neuherberg, Germany. ¹⁷National Heart and Lung Institute, Imperial College London, London, UK. ¹⁸Novo Nordisk Foundation Centre for Basic Metabolic Research, Section of Metabolic Genetics, Faculty of Health and Medical Sciences, University of Copenhagen, Copenhagen, Denmark. ¹⁹Copenhagen Prospective Studies on Asthma in Childhood (COSPAC), Herlev and Gentofte Hospital, University of Copenhagen, Copenhagen, Denmark. ²⁰Steno Diabetes Center, Gentofte, Denmark. ²¹Department of Genetics, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands.

²²Department of Epidemiology, Erasmus University Medical Center, Rotterdam, the Netherlands. ²³Vanderbilt Epidemiology Center, Center for Human Genetics Research, Division of Epidemiology, Department of Medicine, Vanderbilt University, Nashville, Tennessee, USA. ²⁴Estonian Genome Center, University of Tartu, Tartu, Estonia. ²⁵Division of Endocrinology, Children's Hospital Boston, Boston, Massachusetts, USA. ²⁶Department of Genetics, Harvard Medical School, Boston, Massachusetts, USA. ²⁷Program in Medical and Population Genetics, Broad Institute, Cambridge, Massachusetts, USA. ²⁸Center for Genome Science, National Institute of Health, Osong Health Technology Administration Complex, Chungcheongbuk-do, Republic of Korea. ²⁹Medical Genetics Section, University of Edinburgh Centre for Genomic and Experimental Medicine and Medical Research Council (MRC) Institute of Genetics and Molecular Medicine, Western General Hospital, Edinburgh, UK. ³⁰Centre for Cognitive Aging and Cognitive Epidemiology, University of Edinburgh, Edinburgh, UK. ³¹Department of Preventive Medicine, University of Southern California Keck School of Medicine, Los Angeles, California, USA. ³²Institute for Genetic Medicine, University of Southern California Keck School of Medicine, Los Angeles, California, USA. ³³Department of Public Health, Faculty of Medicine, University of Kelaniya, Ragama, Sri Lanka. ³⁴Genome Institute of Singapore, A*STAR, Singapore. ³⁵Department of Ophthalmology, National University of Singapore, Singapore. ³⁶Department of Paediatrics, National University of Singapore, Singapore. ³⁷Medical Clinic V, Mannheim Medical Faculty, University of Heidelberg, Mannheim, Germany. ³⁸Key Laboratory of Nutrition and Metabolism, Institute for Nutritional Sciences, Shanghai Institutes for Biological Sciences, Chinese Academy of Sciences, Shanghai, China. ³⁹Cancer Science Institute of Singapore, National University of Singapore, Singapore. ⁴⁰Center for Advanced Medicine and Clinical Research, Nagoya University Hospital, Nagoya, Japan. ⁴¹Center for Human Genetic Research, Massachusetts General Hospital, Harvard Medical School, Boston, Massachusetts, USA. ⁴²University of Ottawa Heart Institute, Cardiovascular Research Methods Centre, Ottawa, Ontario, Canada. ⁴³Ruddy Canadian Cardiovascular Genetics Centre, Ottawa, Ontario, Canada. ⁴⁴Department of Internal Medicine, Erasmus University Medical Center, Rotterdam, the Netherlands. ⁴⁵Center for Genomic Medicine, Kyoto University Graduate School of Medicine, Kyoto, Japan. ⁴⁶Singapore Institute for Clinical Sciences (SICS), A*STAR, Singapore. ⁴⁷Department of Genetics, University of North Carolina, Chapel Hill, North Carolina, USA. ⁴⁸Institute of Biomedical Sciences, Academia Sinica, Taipei, Taiwan. ⁴⁹School of Chinese Medicine, China Medical University, Taichung, Taiwan. ⁵⁰State Key Laboratory of Medical Genetics, Shanghai Ruijin Hospital, Shanghai Jiao Tong University School of Medicine, Shanghai, China. ⁵¹Shanghai Institute of Hypertension, Shanghai, China. ⁵²Institute of Epidemiology and Biobank popgen, Christian Albrechts University of Kiel, Kiel, Germany. ⁵³Department of Epidemiology and Biostatistics, MRC Health Protection Agency (PHE) Centre for Environment and Health, School of Public Health, Imperial College London, London, UK. ⁵⁴Clinical Trials Support Unit (CTSU), Nuffield Department of Population Health, University of Oxford, Oxford, UK. ⁵⁵Department of Clinical

Sciences, Genetic and Molecular Epidemiology Unit, Skåne University Hospital Malmö, Malmö, Sweden. ⁵⁶MRC Epidemiology Unit, Institute of Metabolic Science, University of Cambridge, Cambridge, UK. ⁵⁷Department of Clinical Chemistry, Fimlab Laboratories, Tampere, Finland. ⁵⁸Department of Clinical Chemistry, University of Tampere School of Medicine, Tampere, Finland. ⁵⁹Vietnam National Heart Institute, Bach Mai Hospital, Hanoi, Vietnam. ⁶⁰School of Public Health, University of Minnesota, Minneapolis, Minnesota, USA. ⁶¹Department of Gerontology and Geriatrics, Leiden University Medical Center, Leiden, the Netherlands. ⁶²Netherlands Consortium for Healthy Ageing, Leiden, the Netherlands. ⁶³Department of Clinical Physiology and Nuclear Medicine, Turku University Hospital, Turku, Finland. ⁶⁴Research Centre of Applied and Preventive Cardiovascular Medicine, University of Turku, Turku, Finland. ⁶⁵Department of Public Health Sciences, University of Chicago, Chicago, Illinois, USA. ⁶⁶Institute of Genetics and Biophysics A Buzzati-Traverso, CNR, Naples, Italy. ⁶⁷Division of Heart and Lungs, Department of Cardiology, University Medical Center Utrecht, Utrecht, the Netherlands. ⁶⁸Division of Genetics and Cell Biology, San Raffaele Scientific Institute, Milan, Italy. ⁶⁹Institute for Maternal and Child Health, Istituto di Ricovero e Cura a Carattere Scientifico (IRCCS) 'Burlo Garofolo', Trieste, Italy. ⁷⁰University College London Genetics Institute, Department of Genetics, Environment and Evolution, University College London, London, UK. ⁷¹Division of Genomics, Medical Institute of Bioregulation, Kyushu University, Fukuoka, Japan. ⁷²Department of Epidemiology, School of Public Health, Medical College of Soochow University, Suzhou, China. ⁷³Department of Nutrition, University of North Carolina, Chapel Hill, North Carolina, USA. ⁷⁴UChicago Research Bangladesh, Uttara, Dhaka, Bangladesh. ⁷⁵Singapore Eye Research Institute, Singapore National Eye Centre, Singapore. ⁷⁶Metabolic Disease Group, Wellcome Trust Sanger Institute, Cambridge, UK. ⁷⁷National Institute for Health Research (NIHR) Cambridge Biomedical Research Centre, Institute of Metabolic Science, Addenbrooke's Hospital, Cambridge, UK. ⁷⁸University of Cambridge Metabolic Research Laboratories, Institute of Metabolic Science, Addenbrooke's Hospital, Cambridge, UK. ⁷⁹Department of Pediatrics, University of Oklahoma Health Sciences Center, Oklahoma City, Oklahoma, USA. ⁸⁰Vanderbilt Epidemiology Center, Division of Epidemiology, Department of Medicine, Vanderbilt Genetics Institute, Vanderbilt University, Nashville, Tennessee, USA. ⁸¹Vanderbilt-Ingram Cancer Center, Division of Epidemiology, Department of Medicine, Vanderbilt University, Nashville, Tennessee, USA. ⁸²Centre for Quantitative Medicine, Office of Clinical Sciences, Duke–National University of Singapore Graduate Medical School, Singapore. ⁸³Yong Loo Lin School of Medicine, National University of Singapore, Singapore. ⁸⁴Department of Psychology, University of Edinburgh, Edinburgh, UK. ⁸⁵Department of Internal Medicine, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands. ⁸⁶Department of Epidemiology, Shanghai Cancer Institute, Shanghai, China. ⁸⁷Fu Wai Hospital and Cardiovascular Institute, Chinese Academy of Medical Sciences and Peking Union Medical College, Beijing, China. ⁸⁸Dayanand Medical College and Hospital Unit, Hero DMC Heart Institute,

Ludhiana, India. ⁸⁹Department of Obstetrics and Gynecology, Oulu University Hospital, Oulu, Finland. ⁹⁰Medical Research Center, University of Oulu, Oulu, Finland. ⁹¹Unit of Primary Care, Oulu University Hospital, Oulu, Finland. ⁹²Department of Cardiovascular Medicine, Heart and Vascular Institute, Cleveland Clinic, Cleveland, Ohio, USA. ⁹³Department of Cell Biology, Lerner Research Institute, Cleveland Clinic, Cleveland, Ohio, USA. ⁹⁴Department of Biostatistics, Vanderbilt University, Nashville, Tennessee, USA. ⁹⁵Department of Paediatrics, Yong Loo Lin School of Medicine, Singapore. ⁹⁶Human Genetics Center, University of Texas School of Public Health at Houston, Houston, Texas, USA. ⁹⁷University of Southern California Keck School of Medicine, Los Angeles, California, USA. ⁹⁸Department of Genetics, Chinese National Human Genomic Center, Shanghai, China. ⁹⁹Research Centre for Prevention and Health, Glostrup University Hospital, Glostrup, Denmark. ¹⁰⁰Graduate School of Regional Innovation Studies, Mie University, Tsu, Japan. ¹⁰¹Department of Geriatric Medicine, Ehime University Graduate School of Medicine, Toon, Ehime, Japan. ¹⁰²Department of Clinical Gene Therapy, Osaka University Graduate School of Medicine, Suita, Japan. ¹⁰³Department of Geriatric Medicine and Nephrology, Osaka University Graduate School of Medicine, Suita, Japan. ¹⁰⁴National Center for Global Health and Medicine, Toyama, Japan. ¹⁰⁵Duke–National University of Singapore Graduate Medical School, Singapore. ¹⁰⁶Department of Epidemiology and Public Health, University College London, London, UK. ¹⁰⁷K.K. Women’s and Children’s Hospital, Singapore. ¹⁰⁸University of San Carlos Office of Population Studies Foundation, University of San Carlos, Cebu City, Philippines. ¹⁰⁹Department of Anthropology, Sociology and History, University of San Carlos, Cebu City, Philippines. ¹¹⁰Department of Internal Medicine, Aichi-Gakuin University School of Dentistry, Nagoya, Japan. ¹¹¹Institute of Human Genetics, Helmholtz Zentrum München–German Research Center for Environmental Health, Neuherberg, Germany. ¹¹²Institute of Human Genetics, Technische Universität München, Munich, Germany. ¹¹³Institute of Genetic Epidemiology, Helmholtz Zentrum München–German Research Center for Environmental Health, Neuherberg, Germany. ¹¹⁴Department of Medicine I, Ludwig Maximilians University Munich, Munich, Germany. ¹¹⁵German Center for Cardiovascular Research (DZHK), partner site Munich Heart Alliance, Munich, Germany. ¹¹⁶Department of Functional Pathology, Shimane University Faculty of Medicine, Izumo, Japan. ¹¹⁷Division of Endocrinology and Diabetes, Department of Internal Medicine, Nagoya University Graduate School of Medicine, Nagoya, Japan. ¹¹⁸Department of Diabetes and Endocrinology, Chubu Rosai Hospital, Nagoya, Japan. ¹¹⁹Heart Centre, Department of Cardiology, Tampere University Hospital and University of Tampere School of Medicine, Tampere, Finland. ¹²⁰Department of Hygiene and Public Health, Teikyo University School of Medicine, Tokyo, Japan. ¹²¹Department of Geriatric Medicine, Graduate School of Medical Sciences, Kyushu University, Fukuoka, Japan. ¹²²MRC Integrative Epidemiology Unit, University of Bristol, Bristol, UK. ¹²³Research and Evaluation Division, Bangladesh Rehabilitation Assistance Committee (BRAC), Dhaka, Bangladesh. ¹²⁴Department of Public Health and Clinical Medicine, Section

for Family Medicine, Umeå Universitet, Umeå, Sweden. ¹²⁵Department of Epidemiology, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands. ¹²⁶Department of Vascular Medicine, University Medical Center Utrecht, Utrecht, the Netherlands. ¹²⁷Alzheimer Scotland Dementia Research Centre, University of Edinburgh, Edinburgh, UK. ¹²⁸Academic Section of Geriatric Medicine, Institute of Cardiovascular and Medical Sciences, Faculty of Medicine, University of Glasgow, Glasgow, UK. ¹²⁹Institute for Adult Diseases, Asahi Life Foundation, Tokyo, Japan. ¹³⁰Institute of Clinical Molecular Biology, Christian Albrechts University of Kiel, Kiel, Germany. ¹³¹Department of Cellular and Molecular Medicine, Lerner Research Institute, Cleveland Clinic, Cleveland, Ohio, USA. ¹³²Department of Cardiology, Leiden University Medical Center, Leiden, the Netherlands. ¹³³Department of Clinical Physiology, Tampere University Hospital, Tampere, Finland. ¹³⁴Department of Clinical Physiology, University of Tampere School of Medicine, Tampere, Finland. ¹³⁵Department of Health Science, Shiga University of Medical Science, Otsu, Japan. ¹³⁶Center for Epidemiologic Research in Asia, Shiga University of Medical Science, Otsu, Japan. ¹³⁷Department of Medical Science and Cardiorenal Medicine, Yokohama City University Graduate School of Medicine, Yokohama, Japan. ¹³⁸Division of Medicine, Turku University Hospital, Turku, Finland. ¹³⁹Department of Medicine, University of Turku, Turku, Finland. ¹⁴⁰Third Department of Internal Medicine, Shimane University Faculty of Medicine, Izumo, Japan. ¹⁴¹Neuroscience and Behavioural Disorders (NBD) Program, Duke–National University of Singapore Graduate Medical School, Singapore. ¹⁴²Duke Eye Center, Duke University Medical Center, Durham, North Carolina, USA. ¹⁴³Cancer Control and Population Sciences, University of Pittsburgh Cancer Institute, Pittsburgh, Pennsylvania, USA. ¹⁴⁴Bioinformatics Division, Tsinghua National Laboratory for Informatics Science and Technology (TNLIST), Ministry of Education Key Laboratory of Bioinformatics, Department of Automation, Tsinghua University, Beijing, China. ¹⁴⁵Center for Synthetic and Systems Biology, TNLIST, Ministry of Education Key Laboratory of Bioinformatics, Department of Automation, Tsinghua University, Beijing, China. ¹⁴⁶Department of Psychiatry, University of Hong Kong, Hong Kong. ¹⁴⁷Durrer Center for Cardiogenetic Research, Interuniversity Cardiology Institute of the Netherlands (ICIN)–Netherlands Heart Institute, Utrecht, the Netherlands. ¹⁴⁸Institute of Cardiovascular Science, Faculty of Population Health Sciences, University College London, London, UK. ¹⁴⁹William Harvey Research Institute, Barts and the London School of Medicine and Dentistry, Queen Mary University of London, London, UK. ¹⁵⁰King Abdulaziz University, Jeddah, Saudi Arabia. ¹⁵¹Department of Nutrition, Harvard School of Public Health, Boston, Massachusetts, USA. ¹⁵²Institute of Reproductive and Developmental Biology, Imperial College London, Hammersmith Hospital, London, UK. ¹⁵³Hebrew University, School of Public Health, Jerusalem, Israel. ¹⁵⁴School of Medicine, University of Minnesota, Minneapolis, Minnesota, USA. ¹⁵⁵Biocenter Oulu, University of Oulu, Oulu, Finland. ¹⁵⁶Center for Life Course Epidemiology, Faculty of Medicine, University of Oulu, Oulu, Finland. ¹⁵⁷ICIN, Utrecht, the Netherlands. ¹⁵⁸Ministry of Health and Welfare, Seoul, Republic of Korea. ¹⁵⁹THERAGEN ETEX

Bio Institute, Suwon, Republic of Korea. ¹⁶⁰Department of Clinical Experimental Research, Rigshospitalet, Copenhagen, Denmark. ¹⁶¹Department of Clinical Medicine, Faculty of Health and Medical Sciences, University of Copenhagen, Copenhagen, Denmark. ¹⁶²Department of Cardiovascular Sciences, University of Leicester, Glenfield Hospital, Leicester, UK. ¹⁶³NIHR Leicester Cardiovascular Biomedical Research Unit, Glenfield Hospital, Leicester, UK. ¹⁶⁴Institute of Preventive Medicine, Bispebjerg and Frederiksberg Hospital, Copenhagen, Denmark. ¹⁶⁵Department of Medicine and Bioregulatory Science, Graduate School of Medical Sciences, Kyushu University, Kyushu, Japan. ¹⁶⁶Institute of Molecular Genetics, National Research Council (CNR), Pavia, Italy. ¹⁶⁷A full list of members appears in the Supplementary Note. ¹⁶⁸Wellcome Trust Sanger Institute, Hinxton, UK. ¹⁶⁹Molecular Epidemiology, Leiden University Medical Center, Leiden, the Netherlands. ¹⁷⁰Clinical Institute of Medical and Chemical Laboratory Diagnostics, Medical University of Graz, Graz, Austria. ¹⁷¹Synlab Academy, Synlab Services, Mannheim, Germany. ¹⁷²Department of Endocrinology, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands. ¹⁷³Department of Genome Science, Aichi-Gakuin University School of Dentistry, Nagoya, Japan. ¹⁷⁴National Hellenic Research Foundation, Institute of Biology, Medicinal Chemistry and Biotechnology, Athens, Greece. ¹⁷⁵Department of Toxicogenomics, Maastricht University, Maastricht, the Netherlands. ¹⁷⁶Oxford Centre for Diabetes Endocrinology and Metabolism, University of Oxford, Oxford, UK. ¹⁷⁷Department of Pathology, National University of Singapore, Singapore. ¹⁷⁸National University of Singapore Graduate School for Integrative Science and Engineering, National University of Singapore, Singapore. ¹⁷⁹Life Sciences Institute, National University of Singapore, Singapore. ¹⁸⁰Department of Statistics and Applied Probability, National University of Singapore, Singapore. ¹⁸¹Department of Medicine, Yong Loo Lin School of Medicine, National University of Singapore, Singapore. ¹⁸²Imperial College Healthcare NHS Trust, London, UK.

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AUTHOR CONTRIBUTIONS

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G. Davies, I.J.D., J.M.S., S.E.H.; Ludwigshafen Risk and Cardiovascular Health Study: G. Delgado, M.E.K., T.B.G., W.M.; Northern Finland Birth Cohort 1986: A.D.S.C.A., A.-L.H., M.-R.J., M.V., S.D., S.F.; Nutrition and Health of Aging Population in China: H.L., X.L., X.Y., Y. Wang; Pakistan Risk Of Myocardial Infarction Study: D.S., J.D., R.A., R.D.Y.; POPGEN study: A.F., I.A., S.S., W.L.; Prevention of Renal and Vascular End-Stage Disease: I.M.L., N.V., R.T.G., W.H.v.G.; Prospective Study of Pravastatin in the Elderly at Risk: D.J.S., I.P., S.T., J.W.J.; Ragama Health Study: A.K., A.R.W., K.S., M.J.P.; Rotterdam Study: A.D., A.G.U., A.H., J.B.J.v.M., L.S., O.H.F.; Secondary Manifestations of Arterial Disease: F.W.A., P.A.D., V. Tragante, W.S.; Shanghai Men's and Women's Health Studies: H.C., Jing He, R. Courtney, T.L.E., W. Zheng, X.-O.S., Y.-B.X., Y.-T.G.; Shanghai-Ruijin Study: D.Z., W.H., X.Z., Yi Zhang; Singapore Chinese Eye Study: C.-C.K., C.-Y.C., J. Liu, T.-Y.W.; Singapore Chinese Health Study: C.H., D.O.S., M.A.P., M.D.G., C.-K.H., J.-M.Y., R.D., R.M.v.D., W.-P.K., Y.F.; Singapore Indian Eye Study: E.S.T., E.V., J. Liao, T.A.; Singapore Malay Eye Study: Y.-Y.T.; Singapore Prospective Study Program: J. Lee, P.C., T.L.Y., X.W.; Suzhou Metabolic Syndrome Study: A.W., H.P., Yonghong Zhang, Z.G.; Taiwan Super Control Study: C.-H.C., J.-Y.W., L.-C.C., Y.-T.C.; Tartu: S.K.; Whitehall II study: J.W., M. Kivimaki, M. Kumari; Young Finns Study: J.S.V., N.M., O.T.R., T.L.

Functional genomics and targeted resequencing: M.L., H.K.N., M.A.R., Z.Y.M., R. Soong, N.S.S. **Statistical analyses:** M.L., F.T., N.V., X.W., W. Zhang, B.L., I.M.L., N.K., J.C.C.

Steering and manuscript writing committee: N.K., M.L., F.T., T.N.K., Y.-Y.T., Jiang He, P.E., E.S.T., P.v.d.H., J.S.K., J.C.C.

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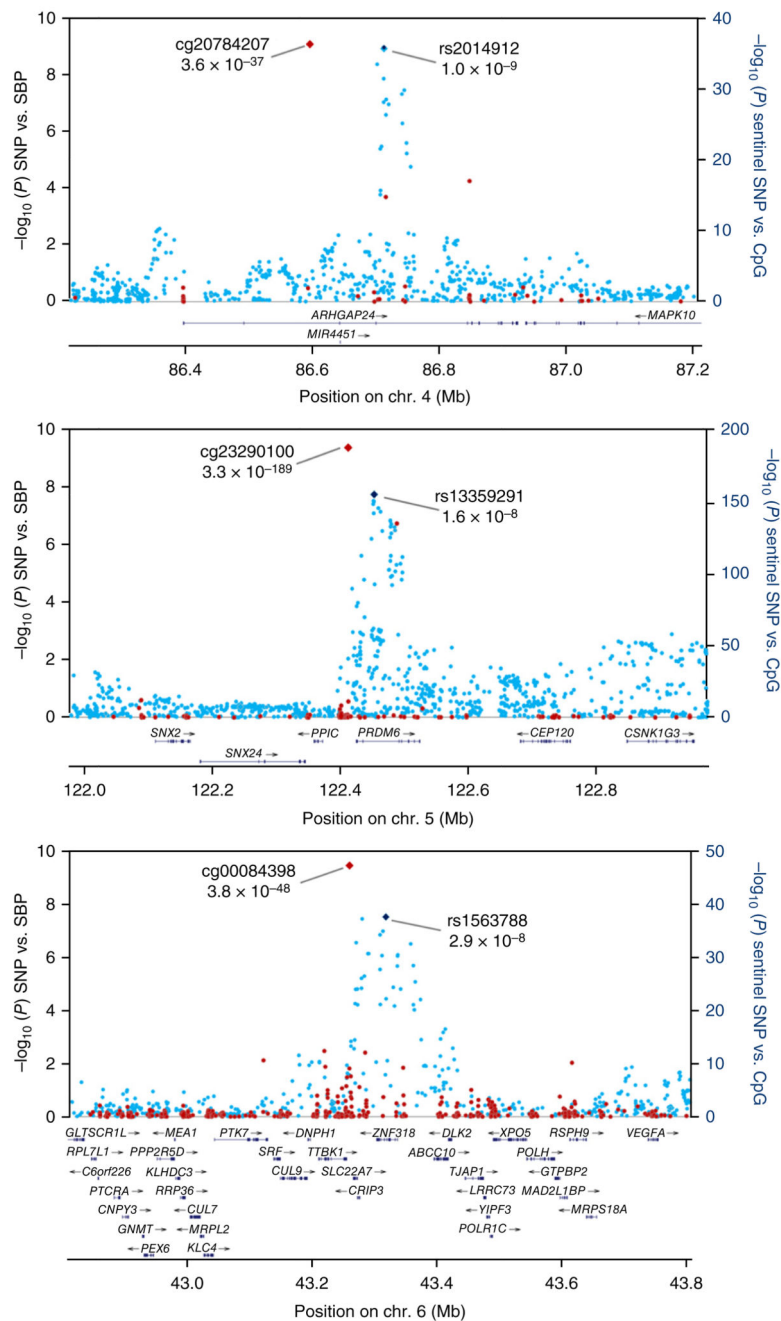


Figure 1.

Regional plots for the three newly identified loci associated with SBP. Associations of SNPs with SBP in the trans-ancestry GWAS (blue markers; $n = 99,994$) and of sentinel SNP with methylation at nearby CpG sites (red markers; $n = 2,664$) are shown. The identities of the sentinel SNP and most closely associated CpG site are provided; correlations between markers are shown in Supplementary Figure 4.

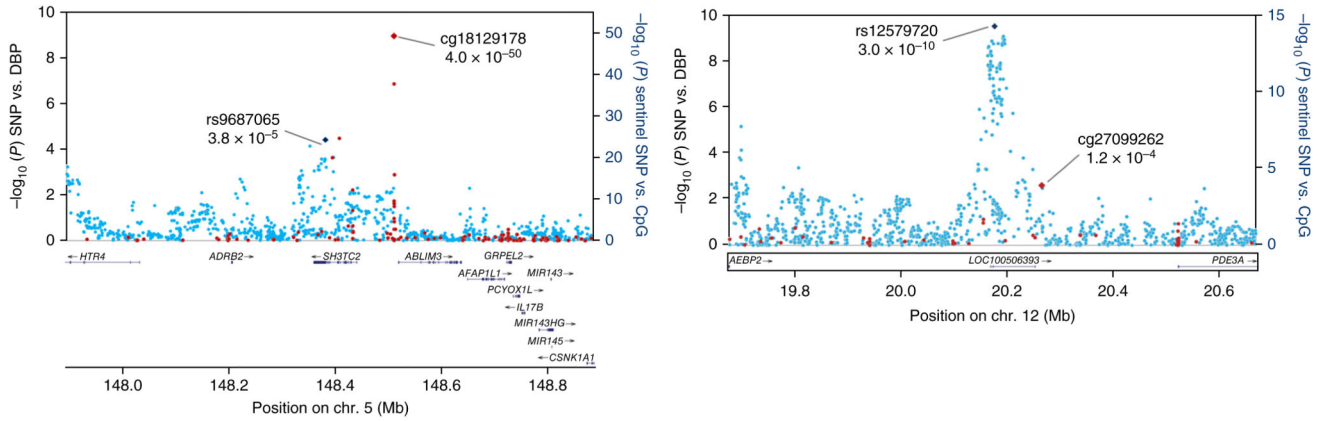


Figure 2.

Regional plots for the two newly identified loci associated with DBP. Associations of SNPs with DBP in the trans-ancestry GWAS (blue markers; $n = 99,994$) and of sentinel SNPs with methylation at nearby CpG sites (red markers; $n = 2,664$) are shown. The identities of the sentinel SNP and most closely associated CpG site are provided; correlations between markers are shown in Supplementary Figure 4.

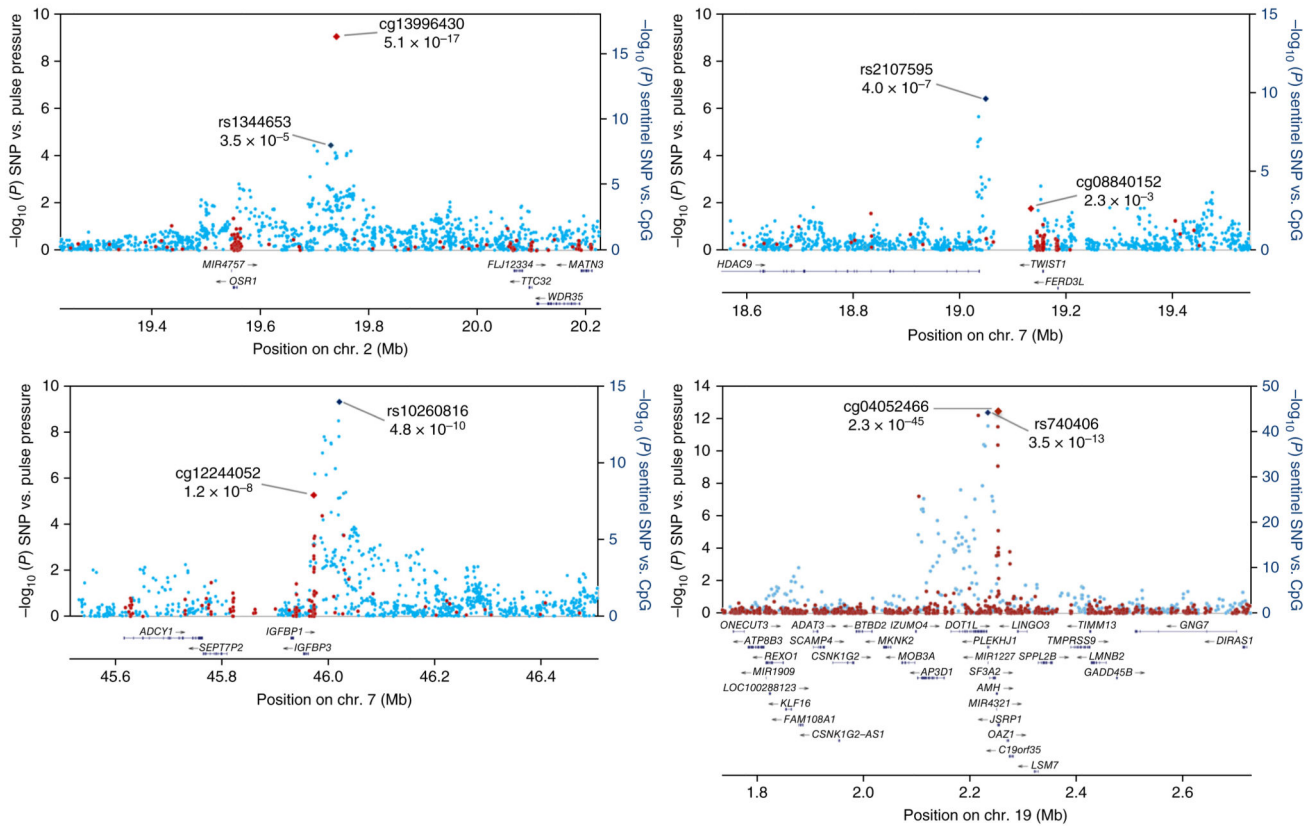


Figure 3.

Regional plots for the four newly identified loci associated with pulse pressure. Associations of SNPs with pulse pressure in the trans-ancestry GWAS (blue markers; $n = 99,994$) and of sentinel SNPs with methylation at nearby CpG sites (red markers; $n = 2,664$) are shown. The identities of the sentinel SNP and most closely associated CpG site are provided; correlations between SNPs and between methylation markers are shown in Supplementary Figure 4.

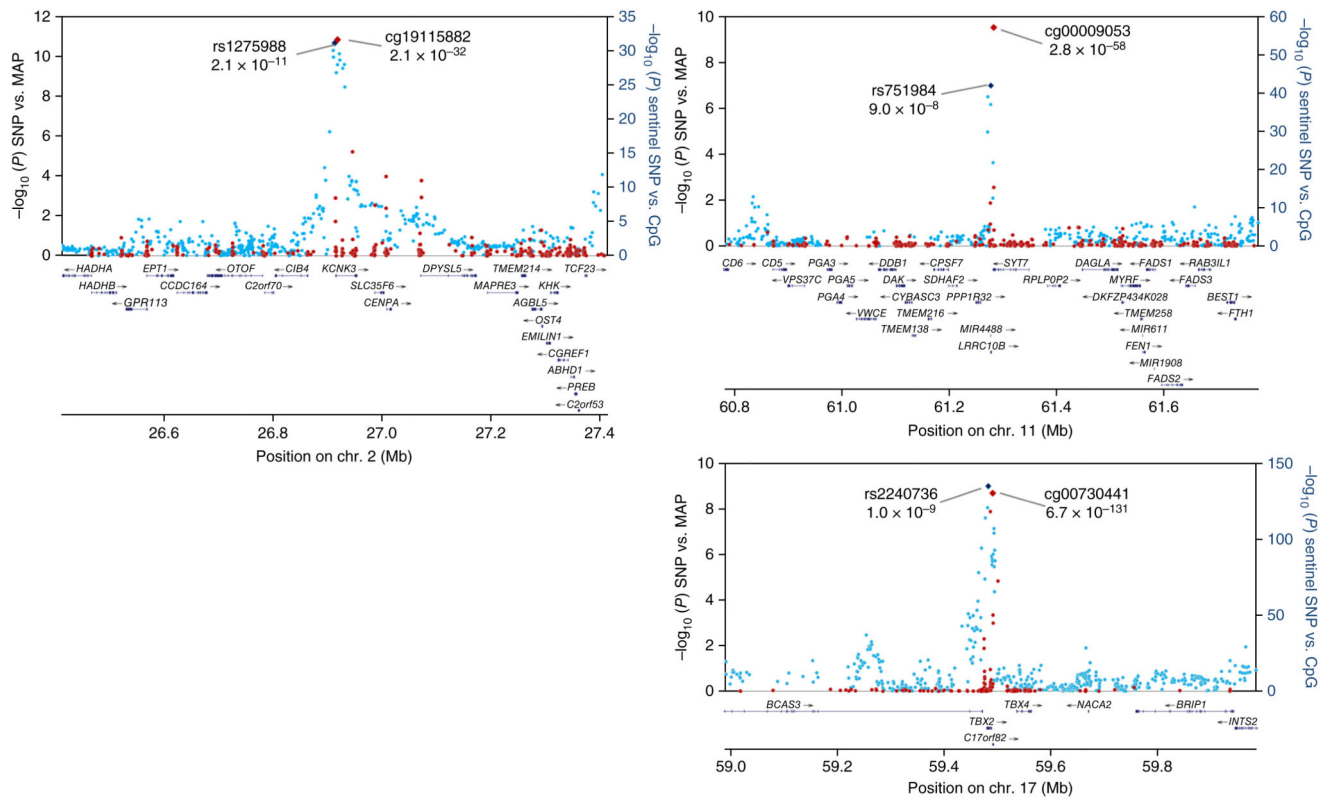


Figure 4.

Regional plots for the three newly identified loci associated with MAP. Associations of SNPs with MAP in the trans-ancestry GWAS (blue markers; $n = 99,994$) and of sentinel SNPs with methylation at nearby CpG sites (red markers; $n = 2,664$) are shown. The identities of the sentinel SNP and most closely associated CpG site are provided; correlations between markers are shown in Supplementary Figure 4.

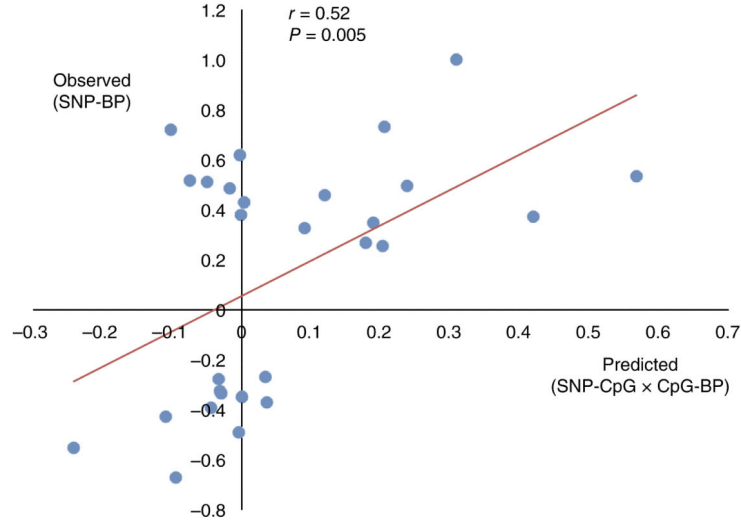


Figure 5.

DNA methylation as a potential mediator of the relationship between sentinel SNPs and blood pressure at the loci reaching genome-wide significance in our study. Results are shown for the 28 sentinel SNPs that are associated with methylation at $P < 0.05$ after Bonferroni correction for multiple tests. Predicted effects on blood pressure are based on the relationship of sentinel SNPs with methylation and the relationship of methylation with blood pressure (BP); observed effects represent the direct relationship between the sentinel SNPs and blood pressure (discovery phenotype). The P value is for the correlation of the observed versus predicted effects (solid line).

Table 1

Genetic loci newly identified to be associated with blood pressure

Sentinel SNP	Chr.	Position (bp)	Candidate gene	EA	AA	EAF	Phenotype	n	Effect (mm Hg)	P
rs1344653	2	19,730,845	<i>OSR1</i> ^{n,m}	A	G	0.54	PP	220,853	-0.27 (0.04)	7.8×10^{-12}
rs1275988	2	26,914,364	<i>KCNK3</i> ^{n,m}	T	C	0.50	MAP	236,311	-0.37 (0.04)	5.0×10^{-21}
rs2014912	4	86,715,670	<i>ARHGAP24</i> ^{n,m}	T	C	0.16	SBP	242,456	0.62 (0.08)	5.4×10^{-17}
rs13359291	5	122,476,457	<i>PRDM6</i> ^{n,m}	A	G	0.31	SBP	229,584	0.53 (0.07)	8.9×10^{-16}
rs9687065	5	148,391,140	<i>ABLIM3</i> ^m , <i>SH3TC2</i> ^{n,ns}	A	G	0.76	DBP	259,216	0.26 (0.04)	7.4×10^{-11}
rs1563788	6	43,308,363	<i>TTBK1</i> ^m , <i>SLC22A7</i> ^v , <i>ZNF318</i> ^{m,e}	T	C	0.31	SBP	220,757	0.51 (0.06)	2.2×10^{-16}
rs2107595	7	19,049,388	<i>HDAC9</i> ⁿ	A	G	0.24	PP	209,305	0.31 (0.05)	3.9×10^{-11}
rs10260816	7	46,010,100	<i>IGFBP3</i> ^{n,m,ns}	C	G	0.62	PP	207,070	0.32 (0.04)	1.5×10^{-14}
rs751984	11	61,278,246	<i>LRRCL10B</i> ⁿ , <i>SYT7</i> ^{n,m}	T	C	0.76	MAP	233,082	0.33 (0.05)	7.7×10^{-12}
rs12579720	12	20,173,764	<i>PDE3A</i> ⁿ	C	G	0.33	DBP	218,606	-0.32 (0.04)	2.2×10^{-16}
rs2240736	17	59,485,393	<i>C17orf82</i> ⁿ , <i>TBX2</i> ^{n,m,ns}	T	C	0.65	MAP	217,197	0.35 (0.04)	2.2×10^{-16}
rs740406	19	2,232,221	<i>AMPH</i> ^m , <i>DOTIL1</i> ⁿ , <i>PLEKHJ1</i> ⁿ , <i>SF3A2</i> ⁿ	A	G	0.85	PP	193,219	-0.55 (0.07)	3.1×10^{-15}

Candidate genes are annotated by the nature of the variant: e, expression quantitative trait locus (eQTL); n, nearby gene (± 10 kb); ns, nonsynonymous; sv, splicing variant; m, DNA methylation marker. Position is based on Build 37 of the reference genome. Effect is shown as unit change (mm Hg) in blood pressure (standard error, SE) per copy of the risk allele (SBP, DBP, PP (pulse pressure), MAP). SNPs rs751984, rs2240736 and rs740406 are near or in annotated microRNA genes. Chr., chromosome; EA, effect allele; AA, alternate allele; EAF, effect allele frequency; n, sample size.

Table 2

CpG sites associated in *cis* with the sentinel blood pressure SNPs

Sentinel SNP	Chr.	EA	Lead CpG	CpG position (bp)	SNP-CpG ^a		Nearest gene to CpG	Relation to gene (CpG)	CpG-eQTL ^b		
					SNP-CpG distance (bp)	P			Effect	P	
rs880315	1	T	cg02903756	10,750,680	46,186	-0.17	7.0×10^{-24}	CASZ1	Body	0.09	2.5×10^{-2}
rs12567136	1	T	cg05228408	11,865,352	18,379	0.6	2.8×10^{-248}	MTHFR	5' UTR	2.34	6.5×10^{-4}
rs1344653	2	A	cg13996430	19,741,587	-10,742	-0.12	7.0×10^{-14}	OSR1	Intergenic	0.20	2.4×10^{-1}
rs1275988	2	T	cg19115882	26,919,145	-4,781	-0.3	1.8×10^{-74}	KCNK3	Body	0.25	1.5×10^{-4}
rs7629767	3	T	cg02108620	42,002,230	41,279	0.57	2.1×10^{-741}	ULK4	5' UTR	-0.1	4.4×10^{-1}
rs13149993	4	A	cg05452645	81,117,647	40,898	-0.26	3.7×10^{-47}	PRDM8	5' UTR	0.03	5.8×10^{-1}
rs2014912	4	T	cg20784207	86,597,598	118,072	-0.27	9.7×10^{-51}	ARHGAP24	Body	-0.51	2.4×10^{-1}
rs7733331	5	T	cg24363955	32,788,467	40,379	-0.22	1.6×10^{-41}	NPR3	Upstream	0.09	5.9×10^{-1}
rs13359291	5	A	cg23290100	122,435,626	40,831	-0.88	6.8×10^{-372}	PRDM6	Body	-0.05	4.4×10^{-1}
rs9687065	5	A	cg18129178	148,520,854	-129,714	-0.45	2.0×10^{-138}	ABLLM3	TSS	-0.07	3.5×10^{-1}
rs11960210	5	T	cg22790839	157,883,933	-66,299	-0.28	3.1×10^{-65}	EBF1	Intergenic	-0.11	1.7×10^{-1}
rs1563788	6	T	cg00084398	43,249,983	58,380	-0.42	5.0×10^{-139}	TTBK1	Body	0.06	5.3×10^{-1}
rs17080102	6	C	cg02784464	151,121,916	-117,146	0.27	7.2×10^{-29}	PLEKHG1	Body	0	3.0×10^{-2}
rs10260816	7	C	cg12244052	45,961,469	48,631	-0.08	4.6×10^{-6}	IGFBP3	Upstream	0.59	7.6×10^{-15}
rs731141	10	A	cg10751070	96,143,568	-244,887	0.14	8.3×10^{-16}	TBC1D12	Intergenic	0.1	5.2×10^{-2}
rs11191375	10	T	cg07119830	104,412,306	52,351	0.97	$3. \times 10^{-746}$	TRIM8	Body	0.08	2.5×10^{-2}
rs2484294	10	A	cg05575054	115,804,968	-12,906	-0.26	2.7×10^{-49}	ADRB1	Body	-0.23	1.7×10^{-1}
rs751984	11	T	cg00009053	61,283,865	-5,619	0.46	1.2×10^{-167}	SYT7	3' UTR	0.1	5.1×10^{-1}
rs2055450	11	A	cg05925497	100,734,094	-183,677	0.19	1.2×10^{-30}	ARHGAP42	Body	-0.09	2.7×10^{-5}
rs10894192	11	A	cg03927812	130,271,903	-5,786	-0.41	5.1×10^{-136}	ADAMTS8	Intergenic	-0.07	4.3×10^{-1}
rs11105354	12	A	cg00757033	89,920,650	105,873	-0.76	9.6×10^{-462}	GALNT4	Intergenic	1.02	2.1×10^{-7}
rs3184504	12	T	cg10833066	111,807,467	96,904	-0.59	4.8×10^{-222}	FAM109A	Intergenic	-0.02	6.7×10^{-1}
rs1378942	15	A	cg02696790	75,250,997	-173,630	0.53	3.1×10^{-223}	RPP25	Intergenic	-0.23	1.7×10^{-1}
rs8032315	15	A	cg06330618	91,428,456	-10,159	0.45	3.0×10^{-493}	FES	Body	-3.19	1.3×10^{-7}
rs2301597	17	T	cg19407385	43,099,144	74,129	-0.72	6.0×10^{-1257}	DCAKD	Intergenic	0.74	7.8×10^{-6}
rs7405452	17	T	cg22053945	46,651,360	23,310	-0.72	4.0×10^{-358}	HOXB3	5' UTR	-0.07	3.3×10^{-1}

Sentinel SNP	Chr.	EA	Lead CpG	CpG position (bp)	SNP-CpG ^a		Nearest gene to CpG	Relation to gene (CpG)	CpG-eQTL ^b	
					SNP-CpG distance (bp)	P			Effect	P
rs2240736	17	T	cg00730441	59,483,863	1,530	1.4×10^{-330}	TBX2	Body	Effect	3.1×10^{-1}
rs740406	19	A	cg04052466	2,251,061	-18,840	3.7×10^{-71}	AMH	Body	Effect	1.5×10^{-2}

Results are shown for SNP-CpG associations reaching both $P < 3.8 \times 10^{-6}$ in discovery (Bonferroni correction for 13,275 SNP-CpG marker tests) and $P < 0.05$ with consistent direction of effect in replication testing (Supplementary Table 15). For each sentinel SNP, the lead CpG site is provided (lowest P value for association of the SNP with the CpG; $P_{\text{SNP-CpG}}$), along with the genomic context of the CpG site. The gene nearest to the CpG site is listed, as well as the P value for association between the CpG site and expression of the nearest gene ($P_{\text{CpG-eQTL}}$). Chr., chromosome; EA, effect allele; NA, not available.

^aThe P value shown is for combined analysis of discovery and replication data for SNP-CpG association.

^bStatistical significance inferred at $P < 1.8 \times 10^{-3}$ (Bonferroni correction for 26 CpG-eQTL tests).