Measuring economic integration:
the case of Asian economies[1]

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Introduction

Since the 1997 Asian financial crisis, both intra-Asia trade and Asian financial markets have experienced considerable growth. Anecdotal evidence indicates that the economic integration of the Asian economies has been steadily progressing. The degree of economic integration is of substantial interest to both academics and policymakers because of its implications for economic efficiency, risk-sharing and the feasibility of forming a currency union.

How integrated are the Asian economies? This is not an easy question to answer. Roughly speaking, economic integration refers to increased interactions and strengthened links between economies. Eatwell, Milgate and Newman (1987, p 43), for example, define economic integration as "a process and as a state of affairs. Considered as a process, it encompasses measures designed to eliminate discrimination between economic units that belong to different national states; viewed as a state of affairs, it represents the absence of various forms of discrimination between national economies". Translating economic concepts into real-world measures may not be straightforward. Assessing the extent of economic integration is no exception.

In the literature, a number of criteria have been developed to evaluate the degree of economic integration. The criteria can be broadly classified in two categories, namely quantity- and price-based measures. The quantity-based category includes measurements of openness and restrictiveness in trade and financial transactions, capital flows, output correlation, savings-investment correlation and consumption correlation.[4] A greater degree of openness (or a lesser degree of restrictiveness) is associated with greater economic integration. The price-based category consists of tests derived from price differentials in goods and financial markets. A greater degree of economic integration is implied by a smaller price differential. Variables including interest rates, price indices and asset prices have been used to assess integration. The use of macro variables such as output, saving, investment and consumption to assess integration is sometimes labelled the macroeconomic approach, while the microeconomic approach refers to the use of financial and goods prices.[5]

It is not an exaggeration to say that we have an embarrassment of riches. There is no consensus on which of these different measures is the most appropriate one to use. We

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[4] Sometimes, the regulatory and institutional measures are included.

[5] See Bayoumi (1997).

anticipate that the multitude of measures, with different implementation methods, will yield different inferences about the degree of integration. For instance, using different approaches, Yu, Fung and Tam (2007) and McCauley, Fung and Gadanecz (2002) offer different assessments of the integration of bond markets in Asia. Indeed, it is reasonable to ask which of the available measures should be used in assessing the degree of integration among the Asian economies.

Instead of arguing in favour of one measure over another, we propose an alternative framework. The economic intuition is that, in general, individual measures focus on different aspects and implications of economic integration, and, therefore, no one by itself gives a complete picture. Thus, it is useful to combine information from individual measures to form an overall assessment of the degree of integration.

The proposed framework is based on the premise that integration is driven by common factors that affect all economies, that some factors affect a group of economies with common characteristics and that there are also economy-specific, idiosyncratic factors. Suppose we have a measure of trade integration and a measure of financial integration. To combine information from these two measures, we assume there is an overall common factor driving both trade and financial integration. Further, some common and group factors are specific to trade, others to financial integration. Thus, a given economy's observed degree of integration is decomposed into several components – an overall common factor that drives both trade and financial integration, one common factor that drives trade (or financial) integration, one factor that drives a group of economies that share some common characteristics and an idiosyncratic component.

The common factors required for the analysis can be constructed using two approaches. One approach is to assume that the common factors are represented by a set of observed economic variables. With this approach, it is desirable to have a theory that relates integration to these variables. The same applies to the use of common elements of these economic variables as proxies of common factors. The second approach is to assume that the common factors are unobservable. We can extract the latent common factors directly from the measures of integration. This approach implicitly assumes that the observed measures of integration contain information on the common force that drives integration. Although the approach is atheoretical, it is quite intuitive and can be implemented easily. Indeed, the technical aspect is drawn mainly from factor models, which have been used to analyse various economic issues. In the current exercise, we will follow the latent common factor approach.

In the next section, we describe the basic econometric framework and its variants. The third section illustrates the practical relevance of the proposed framework. Specifically, the proposed framework is used to examine data on two measures of integration. Some concluding remarks are provided in the final section.

Econometric framework

To simplify the presentation, we first consider the case of one common and one group factor. Then we discuss the variants of the basic setup. The basic specification is given by

$$X_{ij,t} = \gamma_{ij}F_t + \nu_{ij,t}; \qquad\qquad i, j = 1, 2, \dots, \text{N and } i < j, \ t = 1, \dots, \text{T}, \qquad (1)$$

$$X_{ij,t} = \gamma_{ij}F_t + \delta_{ij}Q_{ij,t} + \nu_{ij,t}; \qquad i, j = 1, 2, \dots, \text{N and } i < j, \ t = 1, \dots, \text{T}, \qquad (2)$$

where $X_{ij,t}$ is a measure of integration between economies i and j at time t, F_t is the common factor that affects the level of integration among all the economies, $Q_{ij,t}$ is the group factor defined by some common characteristics of economies in the sample, $\nu_{ij,t}$ is the regression

error term that captures the idiosyncratic component of integration, *N* is the number of economies under consideration and *T* gives the time dimension of the sample.

To fix the idea, we can interpret $X_{ij,t}$ as the measure of trade integration between economies *i* and *j* at time *t*, F_t as a latent variable that summarises the effects of, say, common economic growth and institutional changes on trade and $Q_{ij,t}$ as a group variable that captures the trade effect of, say, the two economies sharing a similar culture.

In the literature, equation (1) is known as a factor model. The specification has been adapted in finance to investigate asset pricing, in macroeconomics to study business cycles and generate economic forecasts; see, for example, Chamberlain and Rothschild (1983), Forni and Reichlin (1998), Giannone, Reichlin and Small (2005) and Stock and Watson (1989, 2002a,b). In the current context, it is implicitly assumed that the effects of economic variables on the evolution of global trade integration can be represented by a few latent common factors. Alternatively, one can view F_t as the common component of $X_{ij,t}$ in the analysis. One advantage of the data-driven approach is that we do not have to commit to a specific theory on the determinants of global trade integration and the specific (dynamic) channels through which these determinants affect integration.

We deem equation (2), which includes the group factor, to be a relevant specification for data analysis. For instance, in the trade literature some attributes such as culture and participation in a trade agreement have implications for trade intensity. In the current exercise, we appeal to some observable economic characteristics to define the group factor.

The coefficient γ_{ij} pertaining to the common factor effect is allowed to vary across economies. We consider that cross-economy heterogeneity is a real phenomenon and, hence, that a homogeneous restriction on the global factor coefficients is undesirable. For the same reason, the coefficient δ_{ij} of the group effect is also economy-specific.

Two remarks are in order. First, the model can be easily modified to accommodate a case in which there is more than one measure of integration, as illustrated below. Further, the model can be extended to include more than one factor in F_t and $Q_{ij,t}$ and the lags of these factors.

Second, the principal component approach can be used to estimate the latent factor F_t. Forni et al (2000) and Stock and Watson (2002a,b), for example, show that under some regularity conditions and for large *N* and *T*, the principal component of $X_{ij,t}$ is a consistent estimator of the common factor that drives $X_{ij,t}$. By the same token, the latent factor $Q_{ij,t}$ can be estimated by the principal component derived from the subset of $X_{ij,t}$ determined by the common economic characteristic defining the group factor.

Now, suppose $Y_{ij,t}$ is a measure of financial integration. Its common-group-factors specification is given by

$$Y_{ij,t} = \gamma_{ij}G_t + \delta_{ij}R_{ij,t} + \varepsilon_{ij,t}, \tag{3}$$

where G_t, $R_{ij,t}$ and $\varepsilon_{ij,t}$ are the common, group and idiosyncratic components, respectively, of the integration measure $Y_{ij,t}$.

For the sake of argument, we assume that the two measures of integration, $X_{ij,t}$ and $Y_{ij,t}$, represent different aspects of integration and that individually neither gives a complete picture of the degree of integration of the two economies. An analysis that combines information from these two measures can be expressed as follows:

$$X_{ij,t} = \beta_{ij,x}W_t + \gamma_{ij,x}F_t + \delta_{ij,x}Q_{ij,t} + \nu_{ij,t} \tag{4}$$

and

$$Y_{ij,t} = \beta_{ij,y}W_t + \gamma_{ij,y}G_t + \delta_{ij,y}R_{ij,t} + \varepsilon_{ij,t} \tag{5}$$

The system (4) and (5) is a combination of (2) and (3) with an added variable, W_t, which represents the overall common factor that affects, in the current example, both trade and financial integration. The subscripts of ß indicate the effect of the overall common factor on trade and financial channels, respectively. Thus, the setup allows us to infer latent common factors that affect the overall (or, to be more precise in the current example, combined) level of integration, trade (financial) integration and group-specific trade (financial) integration.

We apologise for the imprecise use of language. The meaning of the "common" factor is situation-dependent. For instance, F_t is the common factor when only $X_{ij,t}$ is under consideration. When both $X_{ij,t}$ and $Y_{ij,t}$ are considered, W_t is the overall common factor and, strictly speaking, F_t becomes the trade integration-specific factor. Of course, when we change the sample of economies and the measures of integration, the interpretation of these latent common factors will be altered accordingly. Similarly, the meaning of group factor can be situation-specific. We will make the interpretations of these factors appropriate to the content of the discussion.

Empirical results

In the aftermath of the 1997 Asian financial crisis, there was an intense interest in assessing the integration of Asian economies, not only because of the contribution of integration to economic efficiency but also because integration is believed to promote policy coordination and to be capable of deterring future crises in the region. Further, the level of integration is usually deemed to be one of the preconditions for forming an economic or currency union. Indeed, in the post-crisis period, there has been a substantial increase in intraregional trade, and various initiatives, including the development of local bond markets, have been taken to foster integration. To shed some light on integration, we consider 14 economies in Asia: Australia, China, Hong Kong SAR (hereinafter referred to as Hong Kong), India, Indonesia, Japan, Korea, Malaysia, New Zealand, the Philippines, Singapore, Taiwan (China) (hereinafter referred to as Taiwan), Thailand and Vietnam.

It is quite common to discuss economic integration in terms of trade and financial integration. It has been found that both trade and financial integration increase over time and, typically, go hand in hand, at least in the postwar period.[6] Thus, in our exercise, we consider one measure each of trade and financial integration.

For simplicity, we retain $X_{ij,t}$ as our notation of the measure of trade integration. It is given by:

$$X_{ij,t} = (Ex_{ij,t} + Ex_{ji,t})/(GDP_{i,t} + GDP_{j,t}), \tag{6}$$

where $Ex_{ij,t}$ denotes the exports of economy i to economy j, $Ex_{ji,t}$ denotes the exports of economy j to economy i, and $GDP_{i,t}$ and $GDP_{j,t}$ are the output of economy i and economy j, respectively, at time t. The variable $X_{ij,t}$ is also known as the trade intensity between the two economies and is customarily scaled by 100 to make it a percentage of the sum of the two GDPs.

Figure 1 shows nine selected trade intensity series from our sample of 14 economies for the period January 1998 to December 2006. It is clear that China's trade with its partners grew significantly during the sample period.

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[6] See IMF (2002). Obstfeld and Taylor (2004) observe that the degree of international integration was greater, by some measures, at the end of the 1800s.

Figure 1

Selected trade intensity series

1998M1 to 2006M12
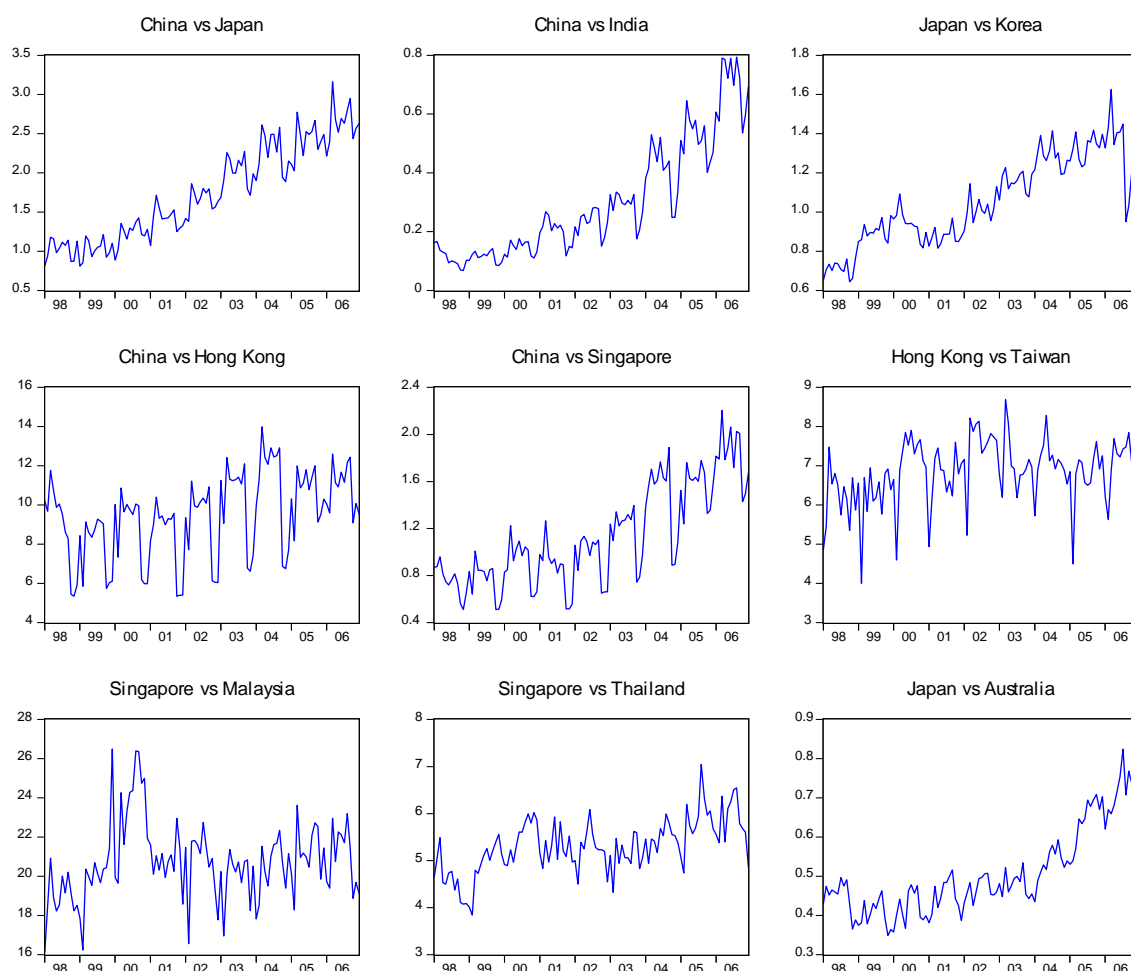
We use interest rate co-movement to assess the degree of financial integration. Specifically, our measure of financial integration is defined by $Y_{ij,t} = corr(IR_{i,t}, IR_{j,t})$, the correlation of interest rates of economies i and j over a moving window of 12 months.[7] Because of the lack of data, Vietnam is not included in the sample for financial integration analysis and the sample period is restricted to January 2000–December 2006. Figure 2 depicts nine selected interest rate correlation series.

As discussed in the previous section, the principal component approach is used to extract from the trade intensity series the common factors that drive the evolution of bilateral trade among the sample economies. Table 1 shows the five largest principal components, which explain 70% of the total variation. The largest principal component accounts for around 44% of the total variation. The presence of a strong common component suggests that trade among the 14 sample economies is driven by an influential common latent factor.

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[7] There are other measures of financial integration, such as interest rate parity conditions and financial openness. See, for example, Cheung, Chinn and Fujii (2007).

Figure 2

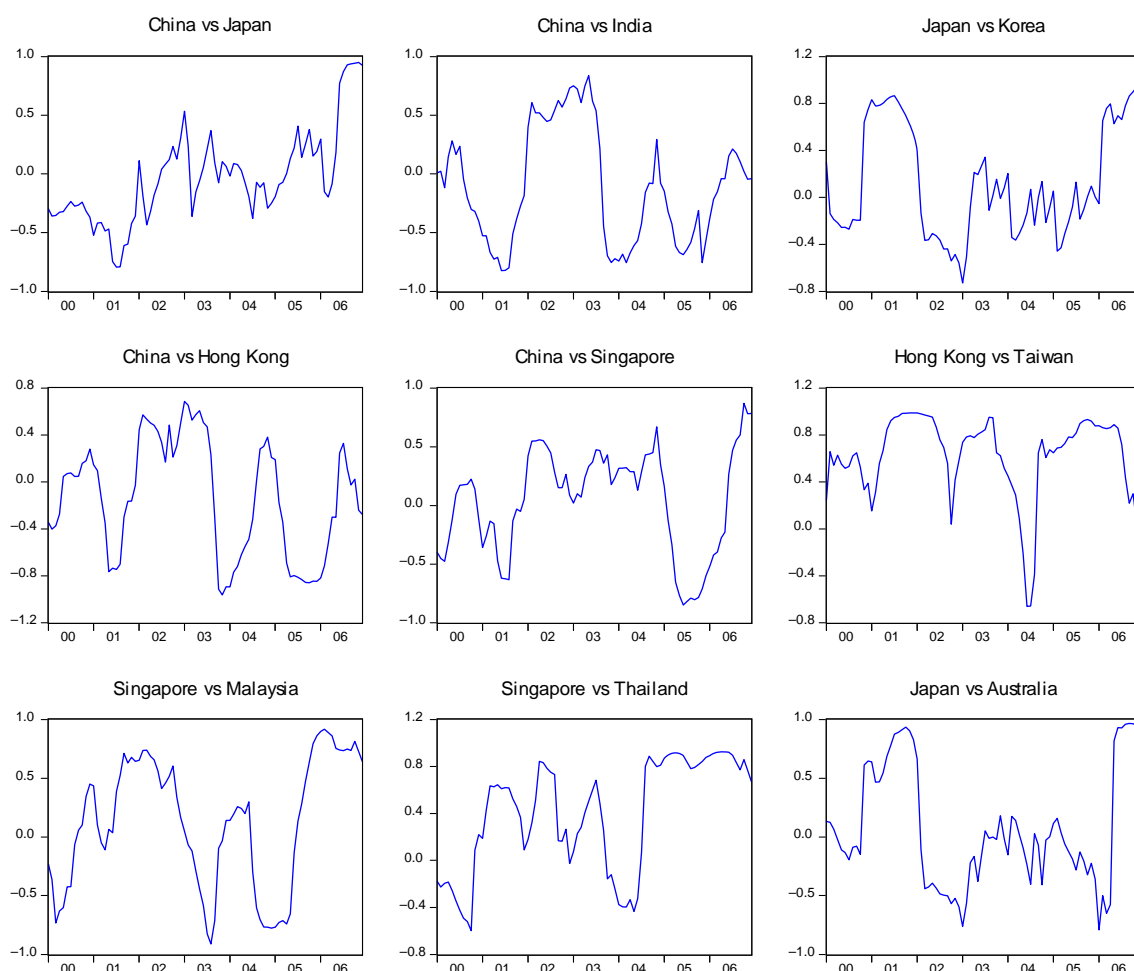**Selected interest rate correlation series**

2000M1 to 2006M12

Table 2 describes the five largest principal components derived from the interest rate correlation series. Unlike the trade intensity series, the interest rate correlation series do not display a dominant principal component. The largest principal component accounts for only 16% of the total variation, whereas each of the next three largest principal components accounts for more than 10% of the total variation. The evidence indicates that, compared with the trade intensity series, the interest rate correlation series have relatively weak common components. The result should not be too surprising because the interest rate is an instrument of the monetary policy pursued by these economies to manage diverse economic conditions.[8] Further, most of these economies do not have full capital account convertibility.

To investigate the relevance of the largest principal component, we estimate equation (1) and calculate the proportion of trade intensity variation explained by the common factor F_t. The results are presented in Table 3. The common factor plays a significant role in explaining the bilateral trade of China, Japan and India, the three largest economies in the region. The average of the explained variability for each economy is shown in the last row of the table.

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[8]   Hong Kong may be the only exception in the group, given its currency board arrangement, which pegs the Hong Kong dollar to the US dollar.

These averages range from 24% (New Zealand) to 77% (China), indicating a diverse common factor effect.

Table 1

Principal component analysis of trade intensity series

1998M1 to 2006M12

	First principal component	Second principal component	Third principal component	Fourth principal component	Fifth principal component
Eigenvalue	43.88	7.72	5.12	4.65	2.55
Cumulative value	43.88	51.59	56.71	61.36	63.91
Variance proportion	0.48	0.08	0.06	0.05	0.03
Cumulative proportion	0.48	0.57	0.62	0.67	0.70

Table 2

Principal component analysis of interest rate correlation series

2000M1 to 2006M12

	First principal component	Second principal component	Third principal component	Fourth principal component	Fifth principal component
Eigenvalue	15.72	11.00	8.87	7.52	6.18
Cumulative value	15.72	26.72	35.59	43.10	49.29
Variance proportion	0.20	0.14	0.11	0.10	0.08
Cumulative proportion	0.20	0.34	0.46	0.55	0.63

Table 4 reports the same results from the interest rate correlation series – each interest rate correlation series is regressed on the largest principal component. Contrary to the results in Table 4, the proportion of variability in interest rate correlation explained by the largest component is quite small. Specifically, average explained variability ranges from 2.19% (China) to 31% (Malaysia). The relatively low explanatory power reflects the absence of a dominating interest rate correlation principal component.

Next, we investigate the role of group factors. Table 5 reports the regression results of equation (2), with Chinese culture as the group factor. The Chinese culture group comprises China, Hong Kong, Singapore and Taiwan. We first estimate the principal components from these economies' trade intensity series. To capture the marginal Chinese culture effect, the Chinese culture principal component is regressed on the common factor F_t, and the resulting residuals, labelled U_t, are used to define the group factor $Q_{ij,t}$ in the regression exercise. Hong Kong and Singapore give the only insignificant estimate of the latent group factor. In general, the results indicate that the Chinese culture factor offers a significant marginal explanation of bilateral trade between these economies. Except in the case of Hong Kong and Taiwan, adjusted R^2 is quite high. Indeed, in three of the six cases adjusted R^2 equals nearly 90%.

Table 6 considers the ASEAN trade agreement group effect. The members of ASEAN included in our sample are Indonesia, Malaysia, the Philippines, Singapore and Thailand (ASEAN 5). The ASEAN trade agreement group factor is constructed using the procedure for estimating the Chinese culture group factor. Only two of the coefficient estimates in Table 6 are insignificant – that is, a portion of bilateral trade between these economies is attributable to the ASEAN trade agreement. A comparison of adjusted R^2 in Tables 5 and 6 suggests that the Chinese culture group factor has a stronger influence on bilateral trade.

Estimates of bilateral interest rate correlation (equation (3)) are given in Tables 7 and 8. As shown in these tables, the latent factors are insignificant more often in the bilateral interest rate correlation series than in the trade intensity series and are less able to explain bilateral interest rate correlation, as indicated by the adjusted R^2 estimates.

Next, we consider an overall common factor that affects both the trade intensity and the interest rate correlation series. The availability of interest rate data dictates the size of the combined dataset; specifically, Vietnam is not included because of the lack of interest rate data, and the sample period is limited to January 2000–December 2006. Table 9 describes the first five principal components. The first principal component accounts for 26% of the total variation in the dataset, and the second principal component explains another 13%.

Table 10 shows that the first principal component (ie W_t in equations (4) and (5)) explains a large proportion of the variation in the trade intensity series, while the same overall common factor accounts for only a small fraction of interest rate correlation variability (Table 11). These observations reinforce the results shown in Tables 3 and 4 – trade intensity has a more dominant common factor than interest rate correlation.

The estimates derived from equation (4), which distinguishes between the effects of the overall common factor, the trade-specific common factor and the group-specific factor, are presented in Tables 12 and 13. The Chinese culture effect is the group factor in Table 12, while the ASEAN trade agreement is the group factor in Table 13. To assess the marginal effect of the trade-specific common factor, we regress F_t on the overall common factor, W_t, and use the resulting residuals, labelled ξ_t, as the trade-specific common factor in the regression. For the group-specific factor, we regress the group-specific principal component on W_t and F_t and use the resulting residuals, labelled U_t, as the group factor in the regression.

The results in Tables 12 and 13 are comparable to those in Tables 5 and 6. The trade intensity data are well explained by equation (4), as exemplified by the adjusted R^2 estimates. In general, the results support the notion that trade integration among these economies is driven by the three latent factors. Again, Chinese culture seems to have a stronger effect on trade intensity than the ASEAN trade agreement.

Tables 14 and 15 present the estimation results of equation (5). Table 14 shows the effect of the Chinese culture factor, Table 15 of the ASEAN trade agreement factor. Again, these results confirm that the latent factor model does not explain interest rate correlation as well as it explains trade intensity. Nonetheless, the results lend support to our findings about the effect of the three latent factors on interest rate correlation.

Conclusion

We propose a latent factor structure as an empirical device for studying the degree of integration. Data on selected Asian economies are used to illustrate the relevance of the proposed model in studying trade and financial integration. There is strong evidence that the integration of these economies is affected by an overall latent common factor that drives both trade and financial integration, a trade-specific integration factor, a financial-specific integration factor, a Chinese culture factor and an ASEAN trade agreement factor. These

results are indicative in general of the usefulness of the proposed model in analysing the integration of economies.

We recognise that the current exercise is an exploratory one and that the empirical strategy is not closely linked to any theory of integration. Indeed, in the paper we focus on fitting the data and are sketchy on the related economic interpretation. Currently, we are extending the exercise in several directions. First, we are considering dynamic factor models that allow the latent factors to have time-varying effects on the degree of integration. Obviously, a time-varying latent factor effect offers a means of capturing the possible temporal variation of the link between the latent factor and the degree of integration. Second, the choice of interest rate correlation as a proxy for financial integration may be controversial. We are examining alternative measures of financial integration, including price and interest rate parity conditions. Third, while the proposed factor approach offers a flexible way to study integration, the current framework does not provide much economic interpretation. It is our plan in the next stage to shed some light on the economic intuitions of the exercise by relating the latent common factors to observable economic variables.

Table 3

The proportion of trade intensity variability explained by the overall common factor F_t

1998M1 to 2006M12

	China	India	Japan	Korea	Singapore	Malaysia	Thailand	Indonesia	Philippines	Taiwan (China)	Hong Kong SAR	Viet-nam	Australia	New Zealand
China		13.12	6.48	15.30	15.43	24.45	14.62	27.82	24.55	11.62	57.79	31.42	11.47	45.27
India	86.88		56.93	28.66	20.13	69.14	27.95	8.75	92.78	35.57	56.48	8.92	33.63	86.47
Japan	93.52	43.07		26.71	44.18	53.15	7.52	15.93	40.75	26.39	15.01	12.89	24.41	25.85
Korea	84.70	71.34	73.29		99.54	90.98	91.81	100.00	63.61	32.07	72.56	57.64	97.00	94.25
Singapore	84.57	79.87	55.82	0.46		90.87	49.67	24.56	95.70	34.56	9.31	32.80	76.86	47.17
Malaysia	75.55	30.86	46.85	9.02	9.13		11.23	53.91	89.52	72.70	29.92	32.22	89.00	99.70
Thailand	85.38	72.05	92.48	8.19	50.33	88.77		47.48	75.12	44.16	11.26	17.30	31.89	50.70
Indonesia	72.18	91.25	84.07	0.00	75.44	46.09	52.52		99.29	58.52	82.26	89.45	97.49	88.91
Philippines	75.45	7.22	59.25	36.39	4.30	10.48	24.88	0.71		69.02	38.64	86.99	74.09	85.28
Taiwan (China)	88.38	64.43	73.61	67.93	65.44	27.30	55.84	41.48	30.98		84.30	18.91	99.87	68.44
Hong Kong SAR	42.21	43.52	84.99	27.44	90.69	70.08	88.74	17.74	61.36	15.70		20.40	70.38	92.48
Vietnam	68.58	91.08	87.11	42.36	67.20	67.78	82.70	10.55	13.01	81.09	79.60		54.31	99.59
Australia	88.53	66.37	75.59	3.00	23.14	11.00	68.11	2.51	25.91	0.13	29.62	45.69		99.07
New Zealand	54.73	13.53	74.15	5.75	52.83	0.30	49.30	11.09	14.72	31.56	7.52	0.41	0.93	
Mean	76.97	58.57	72.60	33.07	50.71	37.94	63.02	38.89	28.05	49.53	50.71	56.71	33.89	24.37

Model for each trading pair:

Above the diagonal:

Below the diagonal:

$$TI_t = C + \hat{\gamma}\hat{F}_t + \hat{\varepsilon}_t$$

$$\frac{\text{var}(\hat{\varepsilon}_t)}{\text{var}(TI_t)} * 100$$

$$\frac{\text{var}(\hat{\gamma}\hat{F}_t)}{\text{var}(TI_t)} * 100$$

Var is the sample variance.

Table 4

The proportion of interest rate correlation variability explained by the common factor F_t

2000M1 to 2006M12

	China	India	Japan	Korea	Singapore	Malaysia	Thailand	Indonesia	Philip-pines	Taiwan (China)	Hong Kong SAR	Australia	New Zealand
China		97.02	98.61	95.89	99.74	90.15	99.99	98.14	99.93	99.99	99.49	98.40	96.43
India	2.98		91.40	70.98	83.12	48.49	80.89	32.21	95.34	99.92	99.91	62.35	86.44
Japan	1.39	8.60		50.52	86.61	68.78	78.56	61.82	84.91	63.44	86.91	52.59	57.35
Korea	4.11	29.02	49.48		45.35	77.26	78.80	91.58	97.50	54.26	70.68	89.13	85.23
Singapore	0.26	16.88	13.39	54.65		59.76	96.01	83.45	69.71	75.93	98.01	78.05	80.58
Malaysia	9.85	51.51	31.22	22.74	40.24		54.57	99.20	99.89	36.22	55.36	80.80	57.35
Thailand	0.01	19.11	21.44	21.20	3.99	45.43		52.92	87.03	94.59	94.74	81.56	99.63
Indonesia	1.86	67.79	38.18	8.42	16.55	0.80	47.08		99.24	51.80	63.27	94.64	91.87
Philippines	0.07	4.66	15.09	2.50	30.29	0.11	12.97	0.76		79.07	71.03	84.72	78.72
Taiwan (China)	0.01	0.08	36.56	45.74	24.07	63.78	5.41	48.20	20.93		94.70	64.99	78.83
Hong Kong SAR	0.51	0.09	13.09	29.32	1.99	44.64	5.26	36.73	28.97	5.30		73.19	75.47
Australia	1.60	37.65	47.41	10.87	21.95	19.20	18.44	5.36	15.28	35.01	26.81		84.69
New Zealand	3.57	13.56	42.65	14.77	19.42	42.65	0.37	8.13	21.28	21.17	24.53	15.31	
Mean	2.19	21.00	26.54	24.40	20.31	31.01	16.73	23.32	12.74	25.52	18.10	21.24	18.95

Model for each interest rate correlation pair:

$$IR_t = C + \hat{\gamma}\hat{G}_t + \hat{\varepsilon}_t$$

Above the diagonal:

$$\frac{\text{var}(\hat{\varepsilon}_t)}{\text{var}(IR_t)} * 100$$

Below the diagonal:

$$\frac{\text{var}(\hat{\gamma}\hat{G}_t)}{\text{var}(IR_t)} * 100$$

Var is the sample variance.

Table 5

**Results of regressing trade intensity series on their
first principal component and the Chinese culture factor**

1998M1 to 2006M12

	China vs Singapore	China vs Taiwan (China)	China vs Hong Kong SAR	Singapore vs Taiwan (China)	Singapore vs Hong Kong SAR	Taiwan (China) vs Hong Kong SAR
Constant	1.14 (0.00)	1.37 (0.00)	9.48 (0.00)	2.22 (0.00)	7.15 (0.00)	6.84 (0.00)
First principal component (F_t)	0.06 (0.00)	0.13 (0.00)	0.21 (0.00)	0.06 (0.00)	0.27 (0.00)	0.05 (0.00)
U_t	0.18 (0.00)	0.21 (0.00)	2.26 (0.00)	0.14 (0.01)	–0.01 (0.93)	0.56 (0.00)
R^2	0.89	0.90	0.71	0.68	0.91	0.27
Adj. R^2	0.89	0.90	0.70	0.67	0.91	0.26

U_t is the residual series obtained from regressing the first Chinese culture principal component on the first principal component of the trading intensity series. ( ) contains the p-value of parameter estimate.

Table 6

**Results of regressing trade intensity series on their
first principal component and the ASEAN 5 factor**

1998M1 to 2006M12

	Singapore vs Malaysia	Singapore vs Thailand	Singapore vs Indonesia	Singapore vs Philippines	Malaysia vs Thailand	Malaysia vs Indonesia	Malaysia vs Philippines	Thailand vs Indonesia	Thailand vs Philippines	Indonesia vs Philippines
Constant	20.73 (0.00)	5.31 (0.00)	5.16 (0.00)	3.35 (0.00)	3.30 (0.00)	1.42 (0.00)	1.82 (0.00)	0.98 (0.00)	1.13 (0.00)	0.41 (0.00)
First principal component (F_t)	0.09 (0.00)	0.06 (0.00)	0.25 (0.00)	0.01 (0.01)	0.11 (0.00)	0.02 (0.00)	0.01 (0.00)	0.03 (0.00)	0.02 (0.00)	–0.001 (0.34)
U_t	1.45 (0.00)	0.29 (0.00)	0.02 (0.82)	0.23 (0.00)	0.12 (0.00)	0.09 (0.00)	0.08 (0.00)	0.06 (0.00)	0.08 (0.00)	0.04 (0.00)
R^2	0.58	0.73	0.75	0.27	0.91	0.61	0.18	0.57	0.37	0.17
Adj. R^2	0.57	0.73	0.75	0.26	0.91	0.60	0.16	0.56	0.36	0.16

U_t is the residual series obtained from regressing the first ASEAN trade agreement principal component on the first principal component of the trading intensity series.
( ) contains the p-value of parameter estimate.

Table 7

**Results of regressing interest rate correlation series on their
first principal component and the Chinese culture factor**

2000M1 to 2006M12

	China vs Singapore	China vs Taiwan (China)	China vs Hong Kong SAR	Singapore vs Taiwan (China)	Singapore vs Hong Kong SAR	Taiwan (China) vs Hong Kong SAR
Constant	0.04 (0.16)	−0.02 (0.65)	−0.14 (0.00)	0.54 (0.00)	0.56 (0.00)	0.62 (0.00)
First principal component (G_t)	−0.01 (0.40)	−0.001 (0.87)	−0.01 (0.46)	0.05 (0.00)	0.01 (0.11)	0.02 (0.01)
U_t	0.23 (0.00)	0.26 (0.00)	0.16 (0.00)	−0.13 (0.00)	−0.12 (0.00)	−0.14 (0.00)
R^2	0.70	0.59	0.27	0.52	0.39	0.44
Adj. R^2	0.70	0.58	0.25	0.51	0.37	0.43

U_t is the residual series obtained from regressing the first Chinese culture principal component on the first principal component of the interest rate correlation series.
( ) contains the p-value of parameter estimate.

Table 8

**Results of regressing interest rate correlation series on their
first principal component and the ASEAN 5 factor**

2000M1 to 2006M12

	Singapore vs Malaysia	Singapore vs Thailand	Singapore vs Indonesia	Singapore vs Philippines	Malaysia vs Thailand	Malaysia vs Indonesia	Malaysia vs Philippines	Thailand vs Indonesia	Thailand vs Philippines	Indonesia vs Philippines
Constant	0.12 (0.00)	0.40 (0.00)	–0.004 (0.92)	–0.08 (0.02)	0.11 (0.00)	0.04 (0.46)	0.39 (0.00)	0.32 (0.00)	0.19 (0.00)	0.14 (0.00)
First principal component (G_t)	0.09 (0.00)	0.02 (0.07)	–0.04 (0.00)	0.07 (0.00)	0.10 (0.00)	–0.01 (0.28)	–0.003 (0.77)	–0.08 (0.00)	0.05 (0.00)	–0.01 (0.14)
U_t	0.17 (0.00)	–0.01 (0.67)	0.06 (0.03)	0.18 (0.00)	0.23 (0.00)	0.26 (0.00)	0.04 (0.15)	0.06 (0.01)	0.25 (0.00)	0.23 (0.00)
R^2	0.63	0.04	0.21	0.64	0.82	0.46	0.03	0.51	0.75	0.73
Adj. R^2	0.62	0.02	0.19	0.63	0.82	0.44	0.002	0.50	0.75	0.72

U_t is the residual series obtained from regressing the first ASEAN trade agreement principal component on the principal component of the interest rate correlation series.
( ) contains the p-value of parameter estimate.

Table 9

Principal component analysis of trade intensity and interest rate correlation series

2000M1 to 2006M12

	First principal component	**Second principal component**	**Third principal component**	**Fourth principal component**	**Fifth principal component**
Eigenvalue	40.02	20.05	12.70	10.79	9.51
Cumulative value	40.02	60.07	72.77	83.56	93.06
Variance proportion	0.26	0.13	0.08	0.07	0.06
Cumulative proportion	0.26	0.39	0.47	0.54	0.60

The table presents eigenvalues of and proportions of variability explained by individual principal components.

Table 10

The proportion of trade intensity variability explained by the overall common factor W_t

1998M1 to 2006M12

	China	India	Japan	Korea	Singapore	Malaysia	Thailand	Indonesia	Philip-pines	Taiwan (China)	Hong Kong SAR	Australia	New Zealand
China		15.92	7.73	19.87	18.22	35.85	20.45	41.37	22.05	9.31	60.13	17.47	57.56
India	84.08		47.99	14.03	21.48	41.93	43.64	9.09	99.99	42.55	64.94	25.64	72.20
Japan	92.27	52.01		28.44	60.44	87.82	10.35	31.40	62.31	38.43	18.31	26.09	20.04
Korea	80.13	85.97	71.56		92.73	99.72	97.87	95.74	69.91	48.49	73.18	94.26	98.20
Singapore	81.78	78.52	39.56	7.27		99.97	74.61	12.25	99.50	57.75	11.20	91.40	42.84
Malaysia	64.15	58.07	12.18	0.28	0.03		17.08	57.60	86.97	93.63	39.97	97.86	98.37
Thailand	79.55	56.36	89.65	2.13	25.39	82.92		41.34	99.98	58.27	19.14	53.63	71.65
Indonesia	58.63	90.91	68.60	4.26	87.75	42.40	58.66		99.79	71.58	90.55	87.69	98.67
Philippines	77.95	0.01	37.69	30.09	0.50	13.03	0.02	0.21		97.52	47.77	60.08	99.57
Taiwan (China)	90.69	57.45	61.57	51.51	42.25	6.37	41.73	28.42	2.48		99.52	99.97	69.83
Hong Kong SAR	39.87	35.06	81.69	26.82	88.80	60.03	80.86	9.45	52.23	0.48		62.13	95.69
Australia	82.53	74.36	73.91	5.74	8.60	2.14	46.37	12.31	39.92	0.03	37.87		95.94
New Zealand	42.44	27.80	79.96	1.80	57.16	1.63	28.35	1.33	0.43	30.17	4.31	4.06	
Mean	72.84	58.38	63.39	30.63	43.13	28.60	49.33	38.58	21.21	34.43	43.12	32.32	23.29

Model for each trading pair:

$$TI_t = C + \hat{\beta}\,\hat{W}_t + \hat{v}_t$$

Var is the sample variance.

Above the diagonal:

$$\frac{\text{var}(\hat{v}_t)}{\text{var}(TI_t)} * 100$$

Below the diagonal:

$$\frac{\text{var}(\hat{\beta}\,\hat{W}_t)}{\text{var}(TI_t)} * 100$$

Table 11

The proportion of interest rate correlation variability explained by the overall common factor W_t

2000M1 to 2006M12

	China	India	Japan	Korea	Singapore	Malaysia	Thailand	Indonesia	Philip-pines	Taiwan (China)	Hong Kong SAR	Australia	New Zealand
China		96.45	62.91	81.68	99.92	98.50	93.21	55.47	94.17	99.72	89.69	83.84	90.31
India	3.55		96.91	99.79	88.93	97.49	71.62	87.29	79.06	98.68	99.55	92.95	95.37
Japan	37.09	3.09		99.96	90.07	99.99	99.08	99.98	76.90	99.77	99.41	98.97	99.99
Korea	18.32	0.21	0.04		87.69	98.77	99.79	90.63	99.33	94.74	92.87	63.49	83.64
Singapore	0.08	11.07	9.93	12.31		99.98	77.67	78.96	75.92	98.47	98.29	85.90	98.80
Malaysia	1.50	2.51	0.01	1.23	0.02		99.34	99.95	89.15	99.62	97.99	99.26	99.54
Thailand	6.79	28.38	0.92	0.21	22.33	0.66		96.04	84.88	59.79	64.91	99.27	53.83
Indonesia	44.53	12.71	0.02	9.37	21.04	0.05	3.96		100.00	88.21	57.69	99.68	76.48
Philippines	5.83	20.94	23.10	0.67	24.08	10.85	15.12	0.00		92.87	99.75	88.54	98.92
Taiwan (China)	0.28	1.32	0.23	5.26	1.53	0.38	40.21	11.79	7.13		98.51	97.81	94.72
Hong Kong SAR	10.31	0.45	0.59	7.13	1.71	2.01	35.09	42.31	0.25	1.49		91.58	100.00
Australia	16.16	7.05	1.03	36.51	14.10	0.74	0.73	0.32	11.46	2.19	8.42		63.05
New Zealand	9.69	4.63	0.01	16.36	1.20	0.46	46.17	23.52	1.08	5.28	0.00	36.95	
Mean	12.84	7.99	6.34	8.97	9.95	1.70	16.71	14.14	10.04	6.42	9.15	11.30	12.11

Model for each interest rate correlation pair:

$$IR_t = C + \hat{\beta}\,\hat{W}_t + \hat{\varepsilon}_t$$

Above the diagonal:

$$\frac{\text{var}(\hat{\varepsilon}_t)}{\text{var}(IR_t)} * 100$$

Below the diagonal:

$$\frac{\text{var}(\hat{\beta}\,\hat{W}_t)}{\text{var}(IR_t)} * 100$$

Var is the sample variance.

Table 12

**Results of regressing trade intensity series on W_t,
their first principal component and the Chinese culture factor**

2000M1 to 2006M12

	China vs Singapore	China vs Taiwan (China)	China vs Hong Kong SAR	Singapore vs Taiwan (China)	Singapore vs Hong Kong SAR	Taiwan (China) vs Hong Kong SAR
Constant	1.24 (0.00)	1.64 (0.00)	9.81 (0.00)	2.37 (0.00)	7.68 (0.00)	7.03 (0.00)
Overall common factor (W_t)	0.06 (0.00)	0.12 (0.00)	0.21 (0.00)	0.04 (0.00)	0.27 (0.00)	0.01 (0.48)
ξ_t	0.08 (0.00)	−0.07 (0.00)	0.58 (0.00)	0.15 (0.00)	0.21 (0.00)	0.34 (0.00)
U_t	0.21 (0.00)	0.35 (0.00)	2.21 (0.00)	0.12 (0.07)	−0.17 (0.22)	0.37 (0.02)
R^2	0.91	0.95	0.74	0.56	0.90	0.26
Adj. R^2	0.91	0.95	0.73	0.55	0.90	0.23

ξ_t is the residual series obtained from regression F_t on W_t. U_t is the residual series obtained from regressing the first Chinese culture principal component on W_t and F_t.
( ) contains the p-value of the parameter estimate.

Table 13

**Results of regressing trade intensity series on W_t,
their first principal component and the ASEAN 5 factor**

2000M1 to 2006M12

	Singapore vs Malaysia	**Singapore vs Thailand**	**Singapore vs Indonesia**	**Singapore vs Philippines**	**Malaysia vs Thailand**	**Malaysia vs Indonesia**	**Malaysia vs Philippines**	**Thailand vs Indonesia**	**Thailand vs Philippines**	**Indonesia vs Philippines**
Constant	21.06 (0.00)	5.47 (0.00)	5.53 (0.00)	3.42 (0.00)	3.58 (0.00)	1.46 (0.00)	1.84 (0.00)	1.04 (0.00)	1.22 (0.00)	0.40 (0.00)
Overall common factor (W_t)	–0.00 (0.78)	0.04 (0.00)	0.29 (0.00)	–0.01 (0.48)	0.09 (0.00)	0.02 (0.00)	0.02 (0.00)	0.03 (0.00)	0.00 (0.89)	0.00 (0.62)
ξ_t	1.23 (0.00)	0.24 (0.00)	–0.14 (0.06)	0.18 (0.00)	0.02 (0.40)	0.08 (0.00)	–0.02 (0.59)	–0.02 (0.26)	0.04 (0.01)	–0.00 (0.90)
U_t	1.02 (0.00)	0.23 (0.00)	0.25 (0.00)	0.12 (0.02)	0.17 (0.00)	0.09 (0.00)	0.08 (0.00)	0.09 (0.00)	0.03 (0.05)	0.05 (0.00)
R^2	0.72	0.69	0.90	0.20	0.91	0.72	0.22	0.72	0.11	0.33
Adj. R^2	0.71	0.68	0.89	0.17	0.90	0.71	0.19	0.71	0.08	0.30

ξ_t is the residual series obtained from regression F_t on W_t . U_t is the residual series obtained from regressing the first ASEAN trade agreement principal component on W_t and F_t. ( ) contains the p-value of the parameter estimate.

Table 14

Results of regressing interest rate correlation series on W_t, their first principal component and the Chinese culture factor

2000M1 to 2006M12

	China vs Singapore	China vs Taiwan (China)	China vs Hong Kong SAR	Singapore vs Taiwan (China)	Singapore vs Hong Kong SAR	Taiwan (China) vs Hong Kong SAR
Constant	0.04 (0.16)	−0.02 (0.66)	−0.14 (0.00)	0.54 (0.00)	0.56 (0.00)	0.62 (0.00)
Overall common factor (W_t)	−0.00 (0.65)	−0.00 (0.46)	−0.02 (0.00)	0.01 (0.10)	−0.01 (0.13)	−0.01 (0.14)
U_t	−0.01 (0.32)	−0.00 (0.71)	−0.02 (0.08)	0.06 (0.00)	0.01 (0.20)	0.02 (0.01)
ξ_t	0.23 (0.00)	0.27 (0.00)	0.15 (0.00)	−0.13 (0.00)	−0.13 (0.00)	−0.14 (0.00)
R^2	0.71	0.59	0.35	0.56	0.41	0.46
Adj. R^2	0.70	0.58	0.33	0.54	0.39	0.44

ξ_t is the residual series obtained from regression G_t on W_t. U_t is the residual series obtained from regressing the first Chinese culture principal component on W_t and G_t.
( ) contains the p-value of the parameter estimate.

Table 15

**Results of regressing interest rate correlation series on W_t,
their first principal component and the ASEAN 5 factor**

2000M1 to 2006M12

	Singapore vs Malaysia	Singapore vs Thailand	Singapore vs Indonesia	Singapore vs Philippines	Malaysia vs Thailand	Malaysia vs Indonesia	Malaysia vs Philippines	Thailand vs Indonesia	Thailand vs Philippines	Indonesia vs Philippines
Constant	0.12 (0.00)	0.40 (0.00)	−0.00 (0.92)	−0.08 (0.01)	0.11 (0.00)	0.04 (0.46)	0.39 (0.00)	0.32 (0.00)	0.19 (0.00)	0.14 (0.00)
Overall common factor (W_t)	0.00 (0.82)	0.04 (0.00)	0.03 (0.00)	−0.04 (0.00)	0.01 (0.01)	0.00 (0.79)	−0.02 (0.00)	0.01 (0.01)	−0.03 (0.00)	0.00 (0.94)
U_t	0.09 (0.00)	0.04 (0.00)	−0.03 (0.00)	0.06 (0.00)	0.11 (0.00)	−0.01 (0.30)	−0.01 (0.25)	−0.08 (0.00)	0.04 (0.00)	−0.01 (0.13)
ξ_t	0.17 (0.00)	−0.00 (0.95)	0.07 (0.01)	0.17 (0.00)	0.24 (0.00)	0.26 (0.00)	0.03 (0.20)	0.06 (0.01)	0.25 (0.00)	0.23 (0.00)
R^2	0.68	0.33	0.34	0.74	0.92	0.46	0.14	0.51	0.81	0.73
Adj. R^2	0.66	0.31	0.34	0.73	0.91	0.44	0.11	0.49	0.81	0.72

ξ_t is the residual series obtained from regression G_t on W_t. U_t is the residual series obtained from regressing the first ASEAN trade agreement principal component on W_t and G_t. ( ) contains the p-value of the parameter estimate.

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