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Optimal solution of the nearest correlation matrix problem by minimization of the maximum norm

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1. Introduction

The *nearest correlation matrix problem* is to find a valid correlation matrix (positive semidefinite $R(m, m): -1 \leq r_{ij} = r_{ji} \leq 1; r_{ii} = 1; r_{ij} \in R \forall i, j = 1, 2, \dots, m; m \geq 3$) that is *nearest* to a given invalid (negative semidefinite) or pseudo-correlation matrix, Q with $-1 \leq q_{ij} = q_{ji} \leq 1; q_{ii} = 1; q_{ij} \in Q \forall i, j = 1, 2, \dots, m$. In the literature on this problem, '*nearest*' is invariably defined in the sense of the least Frobenius norm $\|\Delta\|_F = \|Q - R\|_F$. However, it is not necessary to define '*nearest*' in this conventional sense. The thrust of this paper is to define '*nearest*' in the least maximum norm (LMN), $\|\Delta\|_m = \|Q - R\|_m$ sense and to obtain R from Q . The LMN provides the minimum range of deviations.

2. Origins of pseudo-correlation matrices

Being the quadratic form, a valid *product moment* correlation matrix, R , is necessarily positive semidefinite (psd). All the successive principal minors of R are non-negative or stated differently, all the eigenvalues of R are non-negative. Each element $r_{ij} \in R$ is the cosine of angle θ_{ij} between the vectors x_i and x_j . An arbitrary *real* symmetric matrix, Q (defined above), is *not* a genuine product moment correlation matrix obtainable from some *real* X although it may appear to be so. Such negative semidefinite (nsd) or pseudo-correlation matrices may enter into empirical investigation due to several reasons. First, the coefficients of correlation may not be computed by the Karl Pearson's (product moment) formula. They might have been obtained by Spearman's formula (of rank correlation) or they could be the polychoric coefficients of correlation. Secondly, some of them might have been computed from variables different in sample size (observations). Suppose

$Q = \begin{bmatrix} Q_{11} & Q_{12} \\ Q_{21} & Q_{22} \end{bmatrix}$ such that Q_{11} is obtained from $X_1(n_1, m_1)$, Q_{22} is obtained from $X_2(n_2, m_2): n_1 > n_2$, and $Q_{12} = Q'_{21}$ is obtained from $[X_1(n_2, m_1), X_2(n_2, m_2)]$, while

$X = \begin{bmatrix} X_1 & X_2 \\ X_1 & \emptyset \end{bmatrix}$, \emptyset standing for '*information not available*'. Then Q could fail to be

positive semidefinite. Thirdly, when the off diagonal entries in Q are large (say ≥ 0.9) in magnitude, but recorded with substantial error or approximation, Q may fail to be positive semidefinite. Fourthly, when the elements of near-singular matrices are rounded off (for reporting in research papers, etc.) without a due care taken to the possible effects of rounding off on the status of the matrices regarding the properties such as positive

semidefiniteness etc, the reported matrices may lose the properties that they originally have had. A telling example of this is the psd matrix obtained by Higham (see Higham, 2002, p. 335 : the matrix was singular in the original). However, the reported matrix (rounded off at the fourth place after decimal) has its determinant = $-2.441038E-05$ (one of the eigenvalues is $-1.343337484E-05$, instead of zero). Surely, a negative value of the determinant is due to rounding off. Lastly, in simulation, especially when Q is an initial approximation to R large in dimension, the analyst has to arbitrarily fill in the values of q_{ij} ; $i \neq j \forall i, j = 1, 2, \dots, m$. The only restraint observed by the analyst is that $q_{ii} = 1$ and $-1 \leq q_{ij} = q_{ji} \leq 1 \forall i, j = 1, 2, \dots, m$. Such an arbitrary Q may often fail to be psd. It may also be noted that if a pseudo-correlation matrix has a non-negative determinant, it does not imply that it is psd, since the negative eigenvalues even in number may make the determinant positive.

3. A brief review of literature on the nearest correlation matrix problem

Rebonato and Jäckel (1999) proposed two methods to solve the nearest correlation matrix problem. The first method is based on a hypersphere decomposition of R (a trial matrix at every iteration). In this scheme the angular coordinates, $\theta_{ij} \in \theta \forall i, j$, are chosen on a trial basis and from these coordinates a matrix B is obtained such that $b_{ij} = \cos \theta_{ij} \prod_{k=1}^{j-1} \sin \theta_{ik}$ for $j = 1, \dots, m-1$ and $b_{ij} = \prod_{k=1}^{j-1} \sin \theta_{ik}$ for $j = m$. This is done for all $i = 1, 2, \dots, m$. From B we get $R = BB'$. Iteratively, the method searches for R such that $\|Q - R\|_F$ is minimized. Finally, after convergence, $\hat{R} = R$. The second method is based on a spectral decomposition and undergoes the five steps: (i) Calculate the eigenvalues λ_j and the eigenvectors s_j of Q ; (ii) set all negative λ_j to zero to obtain l_j ; (iii) multiply the vectors s_j by their associated "corrected" eigenvalues l_j and arrange as the columns of B^* ; (iv) obtain B from B^* by normalizing the row vectors of B^* to unit length; (v) obtain $\hat{R} = BB'$, which is a psd matrix and an approximation to Q , the given nsd matrix. It appears that the second method is quite crude but simple. The nearness of \hat{R} to Q will depend on the magnitude of the determinant of Q .

Higham (2002) proposed a method to obtain \hat{R} from Q such that $\|Q - \hat{R}\|_F$ is the least. The method is very general and allows for weights to be assigned to different elements of the distance matrix as desired by the analyst according to the level of confidence put in to the accuracy or (rationally justified) most probable value of q_{ij} . In that case, the weighted norm of difference is minimized. However, for larger matrices, the method is time consuming due to the linear convergence of the algorithm used by Higham.

Anjos et al. (2003) formulated the nearest correlation matrix problem as an optimization problem with a quadratic objective function and semidefinite programming constraints.

Using such a formulation they derived and tested a primal-dual interior-exterior-point algorithm designed especially for robustness and handling the case where Q is sparse. Instead of using the so-called normal equations to obtain search direction at each iteration, their algorithm eliminates the linear feasibility equations from the start, by maintaining exact primal and dual feasibility throughout and using a single bilinear equation to linearize for the search direction at each iteration. The search direction is found using an inexact Gauss-Newton method rather than a Newton method on a symmetrical system, and is computed using a preconditioned conjugate-gradient type method. The authors considered two types of preconditioner, an optimal diagonal preconditioner and a block diagonal preconditioner obtained from a partial Cholesky factorization. Once the current iterate is sufficiently close to the optimal solution, the algorithm applies a crossover technique that sets the barrier parameter to zero and does not maintain interiority of the iterates. This technique attributes robustness to the algorithm with asymptotic quadratic convergence and the ability to handle warm starts simply. Through the preliminary computational results, the authors demonstrated the robustness of the algorithm and showed that sparsity can be successfully exploited.

In Grubisic and Pietersz (2004) geometric optimization algorithms are developed that efficiently find the nearest low-rank correlation matrix. The algorithms are shown to be globally convergent to a stationary point, with a quadratic local rate of convergence. The connection with the Lagrange multiplier method is established, along with an identification of whether a local minimum is a global minimum. The proposed methods have additional benefits, first that any weighted norm can be applied, and second that neighborhood search can straightforwardly be applied. The authors showed numerically that their methods outperform the existing methods in the literature.

Pietersz and Groenen (2004) proposed a method based on majorization that finds a low-rank correlation matrix nearest to a given (pseudo) correlation matrix. The method is globally convergent and computationally efficient. Additionally, it is straightforward to implement and can handle arbitrary weights on the entries of the correlation matrix. A simulation study by the authors suggests that majorization compares favourably with competing approaches in terms of the quality of the solution within a fixed computational time.

Al-Subaihi (2004) proposed a modification of Kaiser-Dichman procedure (see Kaiser and Dichman, 1962) to generate normally distributed (correlated) variates from a given negative semidefinite Q , which, in the process, is approximated by a positive definite R^* matrix. The resulting variates satisfy the R^* matrix. It appears that Al-Subaihi's method does not guarantee that R^* is sufficiently close to Q .

We take an example from Al-Subaihi (2004, p. 11). The values of $q_{ii} = 1 \ \forall \ i = 1, 2, 3, 4, 5$. The value of $q_{12} = q_{21} = q_{13} = q_{31} = 0.5$. Other elements in the first row (as well as the first column) are all zero. The values of the off-diagonal elements $q_{ij} = q_{ji} = 0.84$ for $i, j = 2, 3, 4, 5 ; i \neq j$.

Al-Subaihi generated the first matrix (call it R^* , given in table 1) as an approximation to Q , while we have simply perturbed R^* to obtain R^{**} . We find that the second matrix, R^{**} , approximates Q more accurately than the first matrix, R^* , generated by Al-Subaihi. Note that neither of the two matrices (R^* and R^{**}) is optimally close to the given Q matrix.

4. The Chebyshev or maximum norm of deviations as a measure of proximity

Instead of the minimum Frobenius norm, one may opt for the Least Maximum Norm (LMN) such that the $\max_{i,j} |q_{ij} - \hat{r}_{ij}|$ is minimum. The LMN gives the minimum range in which \hat{R} (around Q) exists. This line of investigation may be useful since the LMN allows for the least substitutability among the off-diagonal elements of the distance matrix $\Delta : \delta_{ij} \in \Delta ; \delta_{ij} = |q_{ij} - \hat{r}_{ij}| \forall i, j$. We accomplish this task here and for the sake of comparison present some results. As an exercise we first take a matrix from Higham's (2002) paper. The results are presented in table 2. The $\max_{i,j}(\delta_{ij}) = \max_{i,j} |q_{ij} - \hat{r}_{Fij}|$ produced by Higham's estimated \hat{R}_F is 0.23931 and $\sum_{i,j} \delta_{ij}$ is 1.27186. On the other hand, the $\max_{i,j}(\delta_{ij}) = \max_{i,j} |q_{ij} - \hat{r}_{mij}|$ produced by LMN estimated \hat{R}_m is 0.21922 and $\sum_{i,j} \delta_{ij}$ is 1.31536. The determinants of the three matrices are : -1.0, 9.64694658E-06 and 1.28158791E-05 approximately. The eigenvalues of the three matrices are given in table 3.

Then we take a matrix from Al-Subaihi's (2004) paper. The values of $q_{ii} = 1 \forall i = 1, 2, 3, 4, 5$. The value of $q_{12} = q_{21} = q_{13} = q_{31} = 0.5$. Other elements in the first row (as well as the first column) are all zero. The values of the off-diagonal elements $q_{ij} = q_{ji} = 0.84$ for $i, j = 2, 3, 4, 5 ; i \neq j$. The results are presented in table 4. The first of the two matrices presented in table 4 is obtained by Al-Subaihi, while we obtain the second by minimizing the maximum norm of $\hat{\Delta}$.

The Erhardt-Schmidt (or Frobenius) norm $\|\Delta^*\|_F$ of $\Delta^* : \delta_{ij}^* \in \Delta^* ; \delta_{ij}^* = |q_{ij} - r_{ij}^*| \forall i, j$, where $r_{ij}^* \in R^*$ (Al-Subaihi's generated positive semidefinite matrix) and $q_{ij} \in Q$ (the negative semidefinite matrix from which R^* is generated) is 0.313057 against 0.096539, which is the $\|\hat{\Delta}\|_F$ of $\hat{\Delta} : \hat{\delta}_{ij} \in \hat{\Delta} ; \hat{\delta}_{ij} = |q_{ij} - \hat{r}_{ij}| \forall i, j$, while $\hat{r}_{ij} \in \hat{R}$, the psd matrix nearest to Q in the LMN sense. The corresponding maximum norms $\|\Delta^*\|_m$ and $\|\hat{\Delta}\|_m$ are 0.564 and 0.11185 respectively.

The $\max_{i,j}(\delta_{ij}) = \max_{i,j} |q_{ij} - r_{ij}^*|$ produced by Al-Subaihi's R^* is 0.1128 and $\sum_{i,j} \delta_{ij}$ is 0.9798.

The $\max_{i,j}(\delta_{ij}) = \max_{i,j} |q_{ij} - \hat{r}_{mij}|$ produced by \hat{R}_m is 0.02237 and $\sum_{i,j} \delta_{ij}$ is 0.430392. Thus,

\hat{R}_m is an indubitably better approximation than R^* . This shows that the R^* matrix generated from Q by Al-Subaihi is only sub-optimally close to Q .

As an example, let us take a matrix, Q , from Rebonato and Jäckel (1999, p.9). It is a symmetric (negative definite) matrix with unit diagonal elements; $q_{12} = q_{21} = 0.9$; $q_{13} = q_{31} = 0.7$ and $q_{23} = q_{32} = 0.3$. From this Q we obtain three different positive definite matrices, the first (\hat{R}_H) by the hypersphere decomposition, the second (\hat{R}_S) by the spectral decomposition and the third (\hat{R}_m) by the LMN procedure, presented in table 6. The first matrix has $\sqrt{\sum \sum (q_{ij} - \hat{r}_{Hij})^2} = 0.00972135$ and the $\min(\max(|q_{ij} - \hat{r}_{Hij}|)) = 0.00542$; the second matrix has $\sqrt{\sum \sum (q_{ij} - \hat{r}_{Sij})^2} = 0.009931042$ and the $\min(\max(|q_{ij} - \hat{r}_{Sij}|)) = 0.00598$ while the third matrix has the corresponding values 0.010371895 and 0.00484 respectively.

Thus we have two alternative approaches to obtain the nearest psd matrices from the given nsd Q : the one, used conventionally, that minimizes $\|\hat{\Delta}\|_F$ and the other, proposed by us in this paper, that minimizes $\|\hat{\Delta}\|_m$. Use of either norm has its own justification. The LMN minimizes the range of deviation and hence, does not allow any element $\hat{r}_{ij, i \neq j} \in \hat{R}$ to deviate too much from its corresponding q_{ij} . The min(Frobenius norm) may permit excessive deviation of a few elements if so required to bring other element of \hat{R} closer to their counterpart elements (of Q). However, to disallow any element $\hat{r}_{ij, i \neq j} \in \hat{R}$ to deviate too much from its corresponding q_{ij} amounts to place a high level of confidence on the elements of Q .

5. The proposed algorithm to obtain the LMN correlation matrix

Our proposed LMN algorithm that generates the nearest positive semidefinite correlation matrix from a given (fed by the user) nsd pseudo-correlation matrix, Q , runs as follows:

1. Let Q_0 be the given invalid correlation matrix. Set $Q = Q_0$.
2. Find all eigenvalues (L , a diagonal matrix) and eigenvectors (V) from Q . Each column v_j of V (associated with the eigenvalue l_{jj} = diagonal element of L) has unit Euclidean length.
3. Replace all negative values in L by zero.
4. Generate m uniformly distributed random numbers $U(0,1)$ and add them to the diagonal elements of L matrix. Normalize L such that its trace is equal to m .
5. By random walk method of optimization find the best possible L that characterizes trace = m , positive determinant and a positive definite $\hat{R} = VLV'$ closest to Q_0 . Nearness is defined in terms of the Chebyshev or maximum norm $\|\hat{\Delta}\|_m = \|Q_0 - \hat{R}\|_m$.

6. Check if all $\hat{r}_{ii} \in \hat{R}$ are approximately unity. It would depend on tolerance level chosen. If not, replace them by unity. Consider it as Q and go to step 2, else stop.

Note that up to step 3, our algorithm is identical to that of Rebonato and Jäckel (1999). The difference lies in steps 4 through 6 that make adjustments in the eigenvalues to minimize the maximum norm $\|\hat{\Delta}\|_m$.

6. The LMN computer program

We provide here the *source codes* of the computer program that implements the algorithm given above. The main program (LMN) checks if the Q matrix fed by the user is not an nsd matrix. If Q is not a psd matrix, it is best approximated by a positive definite matrix, \hat{R} . It is stored in a file named by the user. LMN invokes two *subroutine* subprograms and a *function* subprogram. Some procedures in the computer program (especially, the one that computes eigenvalues and eigenvectors) have been adapted from Krishnamurthy and Sen (1976), pp. 242-247. The program may be compiled by any suitable FORTRAN compiler. We have compiled the program by Microsoft FORTRAN Compiler.

7. Inputs to the Computer Program

When this program is run, it asks for the following parameters (and inputs). Although sufficiently explained in the program queries, they are explained here.

Before running the program, the Q matrix should be stored in some file. This can be done by some text editor such as EDIT.COM (of MICROSOFT). The name of this file is, say *inputfile*. When the program runs, it asks for the value of m (order of the matrix) and the *inputfile* name (in which Q is stored). The file name should be in single quotes '*inputfile*'. Then it asks for the seed to generate random numbers: with this seed the uniformly distributed random numbers lying between $(0, 1) = U(n,m)$ are generated. This number should lie between -32767 and 32767 , zero excluded. This is a suitable number for most personal computers.

The program runs and if Q is not negative semidefinite, it terminates. If so, the *inputfile* and the *outputfile* of LMN program are identical. If Q is negative semidefinite, the program obtains \hat{R} and asks for the *outputfile* name to store it. The file name should be in single quotes '*outputfile*'.

LMN should be run repeatedly on its own output file to ensure that the resulting matrix is psd. This is required because the output file stores correlation matrix with rounded off elements. Since the output matrix is almost always near-singular, rounding off may often make it negative definite. Note that an nsd pseudo correlation matrix, Q , is a problematic and pathological case. It has to be handled with care and patience. It may also be borne in mind that in numerical analysis a very small non-zero value and zero (exact) cannot be strictly discriminated and therefore, when an acute near-singularity is met with, the difference between singular and non-singular (regular) matrices is blurred in practice.

Presently, in the codes given here, the maximum permissible m is 10. This parameter can be increased. Accordingly, dimensions in the program may be changed before compilation.

8. Limitations and possibilities of improvement

Although theoretically there are no snags in minimizing the maximum norm of deviation of R from Q , our algorithm has clearly two weaknesses, (1) it fails if at any stage of iteration the intermediate \hat{R} turns out to be extremely near-singular, and, for some pathological cases of Q , LMN program may not converge; and (2) the random walk method is a very crude and slow method of optimization. It is easy to preclude extreme near-singularity of \hat{R} at any intermediate stage. But it would be a further research work to replace the random walk method of optimization by some more efficient method such as the *Genetic Algorithm* (see Holland, 1975; Goldberg, 1989; Wright, 1991).

9. Geometric programming and the min-max nearest correlation matrix

The Geometric Programming (GP) algorithm developed by Grubisic and Pietersz (2004) also can solve the nearest correlation matrix problem by minimization of the maximum (or Chebyshev) norm. After reading this paper that appeared on SSRN (Mishra, 2004), Pietersz wrote "... the idea of your paper, of using a maximum error function for rank reduction of correlation matrices, is good, novel and testifies of original work. ...Though you are the first to study precisely this min-max problem, other algorithms than your LMN algorithm are already available for solving it. For example, the geometric programming algorithm that I have developed with Igor Grubisic can already solve this problem. ..." Pietersz (2004), slightly modifying the GM algorithm of Grubisic and Pietersz (2004), showed that while the optimal solutions obtained by the LMN and GP algorithms are identical for Higham's matrix (see table 2, 3rd panel) and Rebonato & Jäckel's matrix (see table 6, 3rd panel), the GM solution is substantially better than the LMN solution in case of Al-Subaihi's matrix (see table 7). The minmax error, $\max_{i,j}(\delta_{ij}) = \max_{i,j} |q_{ij} - \hat{r}_{mij}|$ produced by GM algorithm is 0.0217 against 0.02237 obtained by the LMN algorithm.

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Table 1. Inoptimality of positive semidefinite matrix generated by Al-Subaihi’s method

	Al-Subaihi’s generated R^* matrix					A relatively better R^{**} matrix				
	x_1	x_2	x_3	x_4	x_5	x_1	x_2	x_3	x_4	x_5
x_1	1.0000	0.4964	0.5008	0.0011	0.0050	1.0000	0.4964	0.5008	0.0007	0.0010
x_2	0.4964	1.0000	0.8819	0.7317	0.7363	0.4964	1.0000	0.8819	0.7317	0.8400
x_3	0.5008	0.8819	1.0000	0.7272	0.7305	0.5008	0.8819	1.0000	0.7272	0.8200
x_4	0.0011	0.7317	0.7272	1.0000	0.8432	0.0007	0.7317	0.7272	1.0000	0.8400
x_5	0.0050	0.7363	0.7305	0.8432	1.0000	0.0010	0.8400	0.8200	0.8400	1.0000

Table 2. Nearest correlation matrices obtained by Higham’s and LMN methods

	Higham’s original matrix			Higham’s \hat{R}_F matrix			Min(max norm) \hat{R}_m matrix		
	x_1	x_2	x_3	x_1	x_2	x_3	x_1	x_2	x_3
x_1	1.0000	1.0000	0.0000	1.00000	0.76069	0.15731	1.00000	0.78077	0.21922
x_2	1.0000	1.0000	1.0000	0.76069	1.00000	0.76069	0.78077	1.00000	0.78077
x_3	0.0000	1.0000	1.0000	0.15731	0.76069	1.00000	0.21922	0.78077	1.00000

Table 3. Eigenvalues of Higham’s original matrix, his estimated \hat{R}_F matrix \hat{R}_m matrix

Eigenvalues	λ_1	λ_2	λ_3
Higham’s original matrix	2.4142135623731	1.0000	-4.1421356237309E-01
Higham’s estimated \hat{R} matrix	2.1573046934710	0.84269	5.3065290382026E-06
Min(max norm) \hat{R} matrix	2.2192126035928	0.78078	7.3964071641482E-06

Note: Higham’s estimated matrix (see Higham, 2002, p. 335) has turned negative definite. We perturbed it slightly on the fifth place after decimal to make it a positive definite matrix.

Table 4. A comparative view of nearest correlation matrices by Al-Subaihi's and min(max norm) methods

Al-Subaihi's generated R^* matrix						Ours min(max norm) \hat{R}_m matrix				
	x_1	x_2	x_3	x_4	x_5	x_1	x_2	x_3	x_4	x_5
x_1	1.0000	0.4964	0.5008	0.0011	0.0050	1.000000	0.477630	0.477630	0.018118	0.018118
x_2	0.4964	1.0000	0.8819	0.7317	0.7363	0.477630	1.000000	0.862370	0.817630	0.817630
x_3	0.5008	0.8819	1.0000	0.7272	0.7305	0.477630	0.862370	1.000000	0.817630	0.817630
x_4	0.0011	0.7317	0.7272	1.0000	0.8432	0.018118	0.817630	0.817630	1.000000	0.862370
x_5	0.0050	0.7363	0.7305	0.8432	1.0000	0.018118	0.817630	0.817630	0.862370	1.000000

Table 5. Difference matrices obtained by Al-Subaihi's and our proposed methods

Δ matrix from Al-Subaihi's R^* matrix					Δ matrix from min(max norm) \hat{R}_m matrix				
0	0.0036	0.0008	0.0011	0.0050	0	0.022370	0.022370	0.018118	0.018118
0.0036	0	0.0419	0.1083	0.1037	0.022370	0	0.022370	0.022370	0.022370
0.0008	0.0419	0	0.1128	0.1095	0.022370	0.022370	0	0.022370	0.022370
0.0011	0.1083	0.1128	0	0.0032	0.018118	0.022370	0.022370	0	0.022370
0.0050	0.1037	0.1095	0.0032	0	0.018118	0.022370	0.022370	0.022370	0

Table 6. Nearest correlation matrices obtained from Q of Rebonato and Jäckel , 1999, p. 9

The \hat{R}_H matrix			The \hat{R}_S matrix			The \hat{R}_m matrix		
1.00000	0.89458	0.69662	1.00000	0.89402	0.69632	1.00000	0.89513	0.69746
0.89458	1.00000	0.30254	0.89402	1.00000	0.30010	0.89513	1.00000	0.30486
0.69662	0.30254	1.00000	0.69632	0.30010	1.00000	0.69746	0.30486	1.00000

Table 7. A comparative view of nearest correlation matrices (obtained from Al-Subaihi's matrix) by LMN and Geometric Programming algorithms.

\hat{R}_m obtained by the LMN algorithm						\hat{R}_m obtained by the GP algorithm				
	x_1	x_2	x_3	x_4	x_5	x_1	x_2	x_3	x_4	x_5
x_1	1.000000	0.477630	0.477630	0.018118	0.018118	1.0000	0.4784	0.4786	0.0215	0.0217
x_2	0.477630	1.000000	0.862370	0.817630	0.817630	0.4784	1.0000	0.8616	0.8183	0.8185
x_3	0.477630	0.862370	1.000000	0.817630	0.817630	0.4786	0.8616	1.0000	0.8184	0.8183
x_4	0.018118	0.817630	0.817630	1.000000	0.862370	0.0215	0.8183	0.8184	1.0000	0.8617
x_5	0.018118	0.817630	0.817630	0.862370	1.000000	0.0217	0.8185	0.8183	0.8617	1.0000

The GP solution provided by Pietersz in his letter to the author, dated Aug. 26 & 27, 2004.

Updated and uploaded on <http://www.skmishra.owns1.com> on August 28, 2004.

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C -----PROGRAM LMN -----
C RANDOM WALK METHOD TO FIND Min(max norm) Nearest Positive
C definite Marix from a given Negative Semidefinite Matrix
C -----
C INTEGER *2 IU, IV
C DOUBLE PRECISION A(10), R(10), AR(10), XO(10,10), AA(10)
C DOUBLE PRECISION V(10,10), W(10,10), P(10), D, RH(10,10)
C DOUBLE PRECISION SUML, LAMBDA, EPS, F, VO, VR, VOO, RAND, CN, RNORM
C DIMENSION MM(10)
C CHARACTER *11 OFIL, IFIL
C PARAMETERS ----- MAY BE CHANGED -----
C EPS=ITERATIVE ACCURACY: EPSL=MAIN DIAGONAL ACCURACY
C NITR=NO. OF TRIALS FOR RANDOM WALK SEARCH
C ITMAX=MAX NO. OF ITERATION FOR CONVERGENCE
C EPS= 0.00001
C EPSL=0.00001
C NITR=3
C ITMAX=100
C Increase in NITR slows down program but gives better results
C WRITE(*,*) 'FEED M and INPUT FILE NAME'
C WRITE(*,*) '(M is the order of the input square matrix'
C WRITE(*,*) 'INPUT FILE NAME in single quotes)'
C READ(*,*) M, IFIL
C OPEN(7, FILE=IFIL)
C DO 1 I=1, M
1 READ(7, *) (XO(I, J), J=1, M)
C CLOSE(7)
C WRITE(*,*) 'FEED SEED TO GENERATE RANDOM NUMBERS'
C WRITE(*,*) '(SEED lies between -32767 AND 32767, avoid zero)'
C READ(*,*) IU
C WRITE(*,*) 'OUTPUT FILE ? '
C READ(*,*) OFIL
C VOO=10.0**10
C LTRY=0
C ICO=1
C CALL CONS(XO, P, M, ICO)
C WRITE(*,*) 'EIGEN VALUES OF THE ORIGINAL MATRIX XO ARE :'
C WRITE(*,*) (P(I), I=1, M)
C PAUSE 'STRIKE ENTER TO PROCEED'
C DO 70 I=1, M
C IF(P(I).LT.0.0) GOTO 90
70 CONTINUE
C WRITE(*,*) 'ALL EIGEN VALUES ARE NON-NEGATIVE'
C STOP
C =====
C 90 WRITE(*,*) 'SOME EIGEN VALUES ARE NEGATIVE'
C WRITE(*,*) 'THE R MATRIX AT THIS STAGE IS : '
C OPEN(8, FILE=OFIL, STATUS='NEW')
C DO 789 I=1, M
C P(I)=1.0
C 789 WRITE(8, 89) (RH(I, J), J=1, M)
C CLOSE(8)
C DO 788 I=1, M
C 788 WRITE(*, 89) (RH(I, J), J=1, M)
C ITEST=0
C PAUSE 'STRIKE ENTER TO PROCEED'
C F=0.0

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DO 71 I=1,M
IF(P(I).LE.0.0) P(I)=RAND(IU,IV)
F=F+P(I)
71 CONTINUE
DO 72 I=1,M
72 P(I)=DABS(P(I)/F*M)
WRITE(*,*) 'EIGEN VALUES ARE FORCED TO BE ALL POSITIVE'
SUML=0.0
DO 78 I=1,M
78 SUML=SUML+P(I)
WRITE(*,*) (P(I),I=1,M), ' SUM = ',SUML
DO 999 IIT=1,NITR
LAMBDA=20.0
C   Initialisation of decision variables
DO 7 I=1,M
7 A(I)=DABS(P(I))
ICO=2
CALL CONS(RH,P,M,ICO)
C   ----- FUNCTION EVALUATION -----
F=0.0
DO 11 I=1,M
DO 11 J=1,M
D=DABS(XO(I,J)-RH(I,J))
IF(D.GT.F) F=D
11 CONTINUE
VO=F
C   -----
IT=0
200 IT=IT+1
IF(IT.GT.1000) THEN
WRITE(*,*) 'NO CONVERGENCE IN 1000 ITERATIONS'
GOTO 1000
ENDIF
LAMBDA=LAMBDA/2.0
IMP=0
DO 100 II=1,ITMAX
C   GENERATE M UNIFORMLY DISTRIBUTED RANDOM NUMBERS (-1,1)
150 DO 2 I=1,M
2 R(I)=2.0*(RAND(IU,IV)-0.5)
C   NORMALISE THE RANDOM NUMBERS
RNORM=0.0
DO 3 I=1,M
3 RNORM=RNORM+R(I)**2
RNORM=DSQRT(RNORM)
IF(RNORM.GT.1.0) GOTO 150
DO 4 I=1,M
4 R(I)=R(I)/RNORM
C   ADD RANDOM NUMBERS TIMES LAMBDA TO A VECTOR
DO 5 I=1,M
AR(I)=A(I)+LAMBDA*R(I)
AR(I)=DABS(AR(I))
5 CONTINUE
C   ----- FUNCTION EVALUATION -----
CN=0.0
DO 73 I=1,M
CN=CN+AR(I)
73 CONTINUE

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    SUML=0.0
    DO 74 I=1,M
    AR(I)=AR(I)/CN*M
    SUML=SUML+AR(I)
74 CONTINUE
C   WRITE(*,*) 'SUML = ',SUML
    ICO=2
    CALL CONS(RH,AR,M,ICO)
    F=0.0
    DO 13 I=1,M
    DO 13 J=1,M
    D=DABS(XO(I,J)-RH(I,J))
    IF(D.GT.F) F=D
13 CONTINUE
    VR=F
C   -----
    IF(VR.LT.VO) THEN
    VO=VR
    DO 6 I=1,M
    A(I)=AR(I)
6 CONTINUE
    IMP=1
    ENDIF
100 CONTINUE
    IF((IMP.EQ.0).OR.(LAMBDA.GT.EPS)) GOTO 200
1000 CONTINUE
    IF(VOO.GT.VO) THEN
    VOO=VO
    DO 998 I=1,M
    AA(I)=A(I)
998 CONTINUE
    ENDIF
999 CONTINUE
    DO 997 I=1,M
    A(I)=AA(I)
997 CONTINUE
    VO=VOO
    SUML=0.0
    DO 77 I=1,M
    SUML=SUML+A(I)
77 CONTINUE
    WRITE(*,*) 'SMALLEST MAX DEVIATE = ',VO
    WRITE(*,*) ' '
    WRITE(*,*) ' TRIAL NUMBER = ',LTRY
    WRITE(*,*) ' '
    DO 152 I=1,M
    IF(DABS(RH(I,I)-1.00).GT.EPSL) THEN
    RH(I,I)=1.0
    ITEST=1
    ENDIF
152 CONTINUE
    IF(ITEST.EQ.1) THEN
    CALL CONS(RH,AR,M,1)
    LTRY=LTRY+1
    VOO=10.0**10
    GOTO 90
    ENDIF

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Write(*,*)' ----- Convergence achieved -----'
write(*,*)' '
WRITE(*,*)' NAME THE OUTPUT FILE TO STORE THE RESULT'
WRITE(*,*)' (OUTPUT FILE NAME IN SINGLE QUOTES)'
WRITE(*,*)'ESTIMATED MATRIX'
OPEN(8,FILE=OFIL, STATUS='NEW')
DO 75 I=1,M
C   RH(I,I)=1.00
   WRITE(*,89) (RH(I,J), J=1,M)
   WRITE(8,89) (RH(I,J), J=1,M)
75 CONTINUE
   CLOSE(8)
89 FORMAT(1X,6D13.5)
   WRITE(*,*)'RESULTING MATRIX STORED IN FILE = ',OFIL
   WRITE(*,*)'RUN THIS PROGRAM ONCE MORE ON ITS OWN OUTPUT FILE'
   WRITE(*,*)'UNTIL IT SAYS ALL EIGENVALUES ARE NON-NEGATIVE'
   END
C   -----
   SUBROUTINE CONS(A,P,M,ICO)
C   Constructs Matrix from its eigenvectors and values
   DOUBLE PRECISION A(10,10),B(10,10),V(10,10),W(10,10),P(10),F
   DIMENSION MM(10)
   IF(ICO.GT.1) GOTO 100
   NN=1
1000 NADJUST=0
   DO 10 I=1,M
   DO 10 J=1,M
   B(I,J)=A(I,J)
   10 CONTINUE
   CALL EIGEN(A,M,NN,V,W,P,MM)
C   =====
C   NORMALIZATION OF EIGEN VECTORS
   DO 50 I=1,M
   P(I)=0.0
   P(I)=A(I,I)
   F=0.0
   DO 51 J=1,M
51 F=F+V(J,I)*V(J,I)
   F=DSQRT(F)
   DO 52 J=1,M
   V(J,I)=V(J,I)/F
52 CONTINUE
50 CONTINUE
   DO 11 I=1,M
   DO 11 J=1,M
   A(I,J)=B(I,J)
   11 CONTINUE
   RETURN
C   =====
100 DO 34 J=1,M
   DO 341 JJ=1,M
341 W(J,JJ)=0.0
   W(J,J)=P(J)
   34 CONTINUE
C   WRITE(*,*)'NOW W IS THE DIAGONAL L MATRIX'
   DO 36 J=1,M
   F=0.0

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DO 37 I=1,M
37 F=F+V(I,J)*V(I,J)
DO 36 I=1,M
IF(P(J).EQ.0.0) THEN
V(I,J)=0.0
ELSE
V(I,J)=V(I,J)/DSQRT(F/P(J))
ENDIF
36 CONTINUE
DO 35 J=1,M
DO 35 JJ=1,M
A(J,JJ)=0.0
DO 35 I=1,M
A(J,JJ)=A(J,JJ)+V(J,I)*V(JJ,I)
35 CONTINUE
C WRITE(*,*) 'NOW A IS V*L*VT MATRIX'
RETURN
END
C -----
SUBROUTINE EIGEN(A,N,NN,V,W,P,MM)
C Computes eigenvalues and vectors of a real symmetric matrix
DOUBLE PRECISION A(10,10),V(10,10),W(10,10),P(10)
DOUBLE PRECISION PMAX,EPLN,TAN,SIN,COS,AI,TT,TA,TB
DIMENSION MM(10)
C ----- INITIALISATION -----
DO 50 I=1,N
DO 51 J=1,N
V(I,J)=0.0
51 W(I,J)=0.0
P(I)=0.0
50 CONTINUE
PMAX=0
EPLN=0
TAN=0
SIN=0
COS=0
AI=0
TT=0
EPLN=1.0D-310
C -----
IF(NN.NE.0) THEN
DO 3 I=1,N
DO 3 J=1,N
V(I,J)=0.0
IF(I.EQ.J) V(I,J)=1.0
3 CONTINUE
ENDIF
2 NR=0
5 MI=N-1
DO 6 I=1,MI
P(I)=0.0
MJ=I+1
DO 6 J=MJ,N
IF(P(I).GT.DABS(A(I,J))) GO TO 6
P(I)=DABS(A(I,J))
MM(I)=J
6 CONTINUE

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7 DO 8 I=1,MI
  IF(I.LE.1) GOTO 10
  IF(PMAX.GT.P(I)) GOTO 8
10 PMAX=P(I)
  IP=I
  JP=MM(I)
8 CONTINUE
C   EPLN=DABS(PMAX)*1.0D-09
  IF(PMAX.LE.EPLN) THEN
C   WRITE(*,*)'PMAX EPLN',PMAX, EPLN
C   PAUSE'CONVERGENCE CRITERION IS MET'
  GO TO 12
  ENDIF
  NR=NR+1
C   WRITE(*,*)'PMAX, EPLN',PMAX,EPLN
13 TA=2.0*A(IP,JP)
  TB=(DABS(A(IP,IP)-A(JP,JP))+
1DSQRT((A(IP,IP)-A(JP,JP))**2+4.0*A(IP,JP)**2))
C   WRITE(*,*) 'TA TB = ',TA,TB
  TAN=TA/TB
C   WRITE(*,*) 'TAN = ',TAN
  IF(A(IP,IP).LT.A(JP,JP)) TAN=-TAN
14 COS=1.0/DSQRT(1.0+TAN**2)
  SIN=TAN*COS
  AI=A(IP,IP)
  A(IP,IP)=(COS**2)*(AI+TAN*(2.0*A(IP,JP)+TAN*A(JP,JP)))
  A(JP,JP)=(COS**2)*(A(JP,JP)-TAN*(2.0*A(IP,JP)-TAN*AI))
  A(IP,JP)=0.0
  IF(A(IP,IP).GE.A(JP,JP)) GO TO 15
  TT=A(IP,IP)
  A(IP,IP)=A(JP,JP)
  A(JP,JP)=TT
  IF(SIN.GE.0) GO TO 16
  TT=COS
  GO TO 17
16 TT=-COS
17 COS=DABS(SIN)
  SIN=TT
15 DO 18 I=1,MI
  IF(I-IP) 19, 18, 20
20 IF(I.EQ.JP)GO TO 18
19 IF(MM(I).EQ.IP) GO TO 21
  IF(MM(I).NE.JP) GO TO 18
21 K=MM(I)
  TT=A(I,K)
  A(I,K)=0.0
  MJ=I+1
  P(I)=0.0
  DO 22 J=MJ,N
  IF(P(I).GT.DABS(A(I,J))) GO TO 22
  P(I)=DABS(A(I,J))
  MM(I)=J
22 CONTINUE
  A(I,K)=TT
18 CONTINUE
  P(IP)=0.0
  P(JP)=0.0

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DO 23 I=1,N
IF(I-IP) 24, 23, 25
24 TT=A(I, IP)
A(I, IP)=COS*TT+SIN*A(I, JP)
IF(P(I).GE.DABS(A(I, IP))) GO TO 26
P(I)=DABS(A(I, IP))
MM(I)=IP
26 A(I, JP)=-SIN*TT+COS*A(I, JP)
IF(P(I).GE.DABS(A(I, JP))) GO TO 23
30 P(I)=DABS(A(I, JP))
MM(I)=JP
GO TO 23
25 IF(I.LT.JP) GO TO 27
IF(I.GT.JP) GO TO 28
IF(I.EQ.JP) GO TO 23
27 TT=A(IP, I)
A(IP, I)=COS*TT+SIN*A(I, JP)
IF(P(IP).GE.DABS(A(IP, I))) GO TO 29
P(IP)=DABS(A(IP, I))
MM(IP)=I
29 A(I, JP)=-TT*SIN+COS*A(I, JP)
IF(P(I).GE.DABS(A(I, JP))) GO TO 23
GO TO 30
28 TT=A(IP, I)
A(IP, I)=TT*COS+SIN*A(JP, I)
IF(P(IP).GE.DABS(A(IP, I))) GO TO 31
P(IP)=DABS(A(IP, I))
MM(IP)=I
31 A(JP, I)=-TT*SIN+COS*A(JP, I)
IF(P(JP).GE.DABS(A(JP, I))) GO TO 23
P(JP)=DABS(A(JP, I))
MM(JP)=I
23 CONTINUE
IF(NN.EQ.0) GOTO 7
DO 32 I=1,N
TT=V(I, IP)
V(I, IP)=TT*COS+SIN*V(I, JP)
V(I, JP)=-TT*SIN+COS*V(I, JP)
32 CONTINUE
GO TO 7
12 RETURN
END

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C

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-----
FUNCTION RAND(IU, IV)
C Generates Rectangular (0,1) Random Numbers
DOUBLE PRECISION RAND
INTEGER *2 IU, IV
IV=IU*259
IF(IV.GE.0) GOTO 2
IV=IV+32767+1
2 RAND=IV
IU=IV
RAND=RAND*0.3051851E-04
RETURN
END

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