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Toward a Bias Corrected Currency Equivalent Index

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Abstract

Measuring the economic stock of money, defined to be the present value of current and future monetary service flows, is a difficult asset pricing problem, because most monetary assets yield interest. Thus, an interest yielding monetary asset is a joint product: a durable good providing a monetary service flow and a financial asset yielding a return. The currency equilivant index provides an elegant solution, but it does so by making strong assumptions about expectations of future monetary service flows. These assumptions cause the currency equivalent index to exhibit significant downward bias. In this paper, we propose an extension to the currency equivalent index that will correct for a significant amount of this bias.

Key words: Currency Equilivant Index, Monetary Aggregation, Money Stock JEL classification codes: E49

1 Introduction

Measuring the economic stock of money (ESM), defined by Barnett (1991) to be the present value of current and future monetary service flows, is a difficult asset pricing problem, because most monetary assets yield interest. Thus, an interest yielding monetary asset is a joint product: a durable good providing a monetary service flow and a financial asset yielding a return. Barnett showed that the currency equivalent index (CE), first suggested by Hutt (1963) and formally derived Rotemberg et al. (1995), provides an elegant solution, but it does so by making strong assumptions about expectations of future monetary service flows. Barnett et al. (2005) found that these assumptions cause the CE

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to exhibit significant downward bias. In this paper, we propose an extension to the CE that will correct for a significant amount of this bias. We shall call this new index number the modified currency equivalent index (MCE).

Aggregation theoretic measures of the ESM, like the CE and the MCE, differ significantly from the Federal Reserve's official measure of the capital stock of money, which uses the usual simple sum aggregation methodology. Because interest yielding monetary assets are joint products, yielding both monetary and investment services, simple sum aggregation confounds the present value of the investment yield provided by current portfolio of monetary assets and the ESM. This confounding causes the simple sum aggregates to overstate and to smooth the dynamics of the true ESM, which may explain the poor performance of the official monetary aggregates in forecasting and policy applications.

2 The Economic Capital Stock of Money

2.1 Definition Under Perfect Foresight

Following Barnett (1978), let the representative consumer's current period inter temporal utility function, u_t , be weakly separable in each period's consumption of goods and monetary assets.¹ Let period t be the current time period, and let T be the length of the planning horizon, possibly infinity, such that the representative consumer plans for all periods, $s = t, t + 1, \ldots t + T$. Now define the following variable for period s:

- $C_s = (I \times 1)$ vector of planned per capital consumption of goods and services during period s,
- $p_s = (I \times 1)$ vector of goods and services expected prices and of durable goods expected rental prices,
- $p_s^* =$ the true cost of living index,
- $m_s = (N \times 1)$ vector of planned real balances of monetary assets,
- $M_s = (N \times 1)$ vector of planned nominal balances of monetary assets,
- $r_s = (N \times 1)$ vector of expected nominal holding period yields on monetary assets,
- R_s = the expected nominal one-period holding yield on the benchmark asset.

¹ It is worth noting that u_t is not an elementary utility function, since it contains monetary assets. However, Arrow and Hahn (1971) proved that if money has positive value, then there exists a derived utility function containing money.

Further assume that there exists $p_s^* = p_s^*(p_s)$, which is the true cost of living index that can be used to deflate nominal values in period s. Barnett (1991) recursively combines the T + 1-period budget constraints from the representative agents problem into the single discounted Fisherine wealth constraint. From this discounted Fisherine wealth constraint, Barnett defines the ESM, V_t , as

$$V_t \equiv \sum_{s=t}^T \sum_{n=1}^N \left[\frac{p_s^*}{\rho_s} - \frac{p_s^* (1+r_{n,s})}{\rho_{s+1}} \right] m_{ns}, \tag{1}$$

where the discount rate for period s is

$$\rho_s = \begin{cases} 1 & s = t \\ \prod_{u=t}^{s-1} (1+R_u) & s \neq t \end{cases}.$$
(2)

Following Barnett et al. (2005), (1) can be rewritten as

$$V_{t} = \sum_{s=t}^{\infty} \sum_{n=1}^{N} \left[m_{ns} \left(p_{s}^{*} \cdot \frac{R_{s} + r_{n,s}}{1 + R_{s}} \right) \frac{1}{\rho_{s}} \right] = \sum_{s=t}^{\infty} \frac{TE_{s}}{\rho_{s}},$$
(3)

where TE_s is the total nominal expenditure on monetary services in period s, and T is allowed to approach infinity.

2.2 Extension to Uncertainty

Barnett (1995) and Barnett et al. (1997) showed that, assuming inter-temporally strong separability, all the results on user cost and Divisia aggregation can be extended to the case of risk neutrality by replacing all random variables with their expectations. Thus, applying the consumption-based capital asset pricing model theory,² the formulas for the economics capital stock of money under inter-temporal strong separability becomes

$$V_t = E_t \left(\sum_{s=t}^{\infty} \Gamma_s T E_s \right), \tag{4}$$

where

$$\Gamma_s = \beta^{s-t} \frac{\partial u}{\partial C_s} / \frac{\partial u}{\partial C_t}$$
(5)

is the subjectively-discounted marginal rate of inter-temporal substitution between consumption in the current period t and the future period s.

 $[\]overline{^2$ See Blanchard and Fischer (1989) section 6.3 and Cochrane (2005).

3 Bias Correction of the Currency Equivalent Index

3.1 Source of Bias observed in Currency Equivalent Index

The CE is a special case of the ESM. To see this, follow Barnett (1991) by assuming that $R_s = R_t$, $r_{ns} = r_{nt}$, and $M_{ns} = M_{nt}$ for all $n \in N$ and s = t, t + 1, t + 2..., then (1) can be written

$$V_t = \sum_{s=t}^{\infty} \sum_{n=1}^{N} \left[\frac{R_t - r_{nt}}{\prod_{u=t}^{s} (1+R_t)} \right] M_{nt} = \sum_{n=1}^{N} \left[\frac{R_t - r_{nt}}{R^{s-t+1}} \right] M_{nt} = V_t^{CE}.$$
 (6)

Equation (6) is the CE derived by Rotemberg et al. (1995). Thus, to calculate the CE, we assume that total expenditure on monetary assets follows a Martingale process. Barnett et al. (2005) showed that this assumptions results in a small downward bias³ because total expenditure on monetary service trends upwards over time, see figure 1. Thus, as we move further into the future, the assumptions upon which the CE rests do not match an important characteristic of the data.



Figure 1. Total Expenditure on Monetary Service

3.2 Derivation of Bias Corrected Currency Equivalent Index

We begin by assuming that the expectation of the stochastic discount factor in time period t is

$$E_t\left(\Gamma_s\right) = \prod_{u=t}^s \left[1 + E_t\left(R_u\right)\right]^{-1},$$

 $^{^3}$ Figures 2, 4 and 6 shows the deviation of the CE from the perfect foresight ESM, calculated by Barnett et al. (2005).

then (4) becomes

$$V_t = \sum_{s=t}^{\infty} \frac{E_t \left(TE_s \right)}{\prod_{u=t}^{s} \left[1 + E_t \left(R_u \right) \right]} + cov \left(TE, \Gamma \right).$$

Now assume that expectations about the benchmark rate are stationary, so that $R_s = R_t$ for all $s \ge t$, and that $cov(TE, \Gamma) = 0$, then we have

$$V_{t} = \sum_{s=t}^{\infty} \frac{E_{t} \left(T E_{s} \right)}{\left(1 + R_{t} \right)^{s-t}}.$$
(7)

Finally, assume that

$$E_t (TE_s) = TE_t (1+g_t)^{s-t}, \qquad (8)$$

where g_t is the average growth rate of total expenditure on monetary service at time t. Combining (7) and (8) yields

$$V_t^{MCE} = \frac{TE_t (1 + R_t)}{R_t - g_t},$$
(9)

which we shall call the modified currency equivalent index (MCE). Note that if g_t is zero for all t, then V_t^{MCE} reduces to V_t^{CE} .

4 Results

In order to compare the performance of the MCE relative to the CE, we compare each to the perfect foresight ESM calculated by Barnett et al. (2005).⁴ Figures 2, 4, and 6 show the percentage deviation of the MCE from the perfect foresight ESM (PF) as compared to the CE. Table 1 summarizes these results. First we examine the mean percent error (MPE). At the M1 level of aggregation, the MCE overstates by an average of 3.015; and, at the M2 and M3 levels of aggregation, the MCE understates by an average of 7.254 percent, and 5.807 percent, respectively. Thus, at the M1, M2 and M3 levels of aggregation, the MCE eliminates on average 46.45 percent, 55.37 percent and 62.85 percent, respectively, of the systematic bias observed in the CE.

Table 1 also compares the performance of the MCE to that of the CE by measuring the mean absolute percent error (MAPE) and root mean squared error (RMSE) of both the MCE and the CE relative to the perfect foresight ESM. The MCE is found to be superior to the CE by all three criterion at all

 $[\]overline{^{4}$ The results are generated using Ox version 4.00 (Doornik, 2006). See http://www.doornik.com for further information.

three levels of aggregation. Figures 3, 5 and 7 plot the economic capital stock of money at M1, M2 and M3 levels of aggregation, respectively, as measured by the MCE, CE and perfect foresight ESM.



Figure 2. Percent Error of the MCE as Compared to the CE Relative to Perfect Foresight ESM (M1)



Figure 3. Economic Capital Stock of Money (M1)



Figure 4. Percent Error of the MCE as Compared to the CE Relative to Perfect Foresight ESM (M2)



Figure 5. Economic Capital Stock of Money (M2)





Figure 6. Percent Error of the MCE as Compared to the CE Relative to Perfect Foresight ESM (M3)

Figure 7. Economic Capital Stock of Money (M3)

Comparison of MCE to CE relative to perfect for esight ESM			
	MPE	MAPE	RMSE
MCE1	3.015%	3.539%	13.219
MCE2	-7.254%	9.633%	127.671
MCE3	-5.870%	9.187%	152.903
CE1	-5.630%	5.688%	21.055
CE2	-16.253%	16.330%	197.048
CE3	-15.802%	16.010%	224.892

Table 1

$\mathbf{5}$ Conclusion

Aggregation theoretic measures of the economic stock of money, require forecasting future monetary service flows. The currency equivalent index greatly simplifies the procedure of discounting future monetary service by assuming that future monetary service flows remain constant at current levels. This assumption has led to reluctance in accepting the currency equivalent index as a credible measure of the money stock. However, Barnett et al. (2005) showed that the resulting bias is relatively small when compared to the bias observed in the simple sum aggregates. Moreover, we find that by allowing total expenditure on monetary services to increase over time, the bias observed in the currency equivalent index can be reduced by nearly 63%.

While the modified currency equivalent index does exhibit a small bias, it is a significant improvement of the currency equivalent index. Moreover, both indices are vast improvements over the simple sum aggregates. Hence, we conclude that while more sophisticated forecasting methodology may lead to slightly improved measures of the money stock, the modified currency equivalent index provides an easily calculated, internally consistent measure of the economic stock of money that improves upon both the currency equivalent and the simple sum indices.

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