

# A Note on Why Quarter of Birth is Not a Valid Instrument for Educational Attainment

Aliprantis, Dionissi University of Pennsylvania

05. October 2007

Online at http://mpra.ub.uni-muenchen.de/5168/ MPRA Paper No. 5168, posted 07. November 2007 / 04:29

## A NOTE ON WHY QUARTER OF BIRTH IS NOT A VALID INSTRUMENT FOR EDUCATIONAL ATTAINMENT

DIONISSI ALIPRANTIS

#### 1. Angrist and Krueger (1991)

The goal in Angrist and Krueger (1991) is to use variation in schooling attainment unrelated to ability to identify the effect of attending an additional year of school. The paper uses date of birth and compulsory education laws as a natural experiment to assign one group of students an extra year of education. While the paper rightly focuses much attention on demonstrating the effectiveness of compulsory education laws, it does not give sufficient consideration to the problems generated for the IV framework if there is significant redshirting<sup>1</sup> around entrance cutoff dates. This problem is distinct from those discussed in Bound and Jaeger (2000), which assumes "the typical compulsory school attendance law requires students to start first grade in the autumn of the calendar year in which they turn 6," precluding the possibility of redshirting.

1.1. Monotonicity. The implicit assumption needed to identify the effect of attending an additional year of school is that children enroll in school when first eligible. Without this implicit assumption, the explicit assumption of Monotonicity (Angrist and Imbens (1995)) does not hold. As reported in Aliprantis (2007), in 1998 about 30% of parents in the US chose to delay their children's enrollment in primary school if their child's birthday was within two months of their school's entrance cutoff date. This rate was almost one in four for the youngest quarter of eligibility to enroll. Furthermore, this selection was not random, as will be discussed in Section 2.

To illustrate the problem with this implicit assumption suppose, for the sake of simplicity, that all children born in quarter 1  $(Q_1)$  are born on February 15th, quarter 2  $(Q_2)$  on May 15th, quarter 3  $(Q_3)$  on August 15th, and quarter 4  $(Q_4)$  on November 15th. Assume that school starts on August 15th and the cutoff date for all children to turn six is December 31st of the academic school year. Then Table 1 shows the pattern of ages at entry of first grade we will observe.

Now assume that all children must attend school until they reach 16 years of age and that we are looking only at those children who would like to drop out once they turn 15 years old. Then Table 2 shows the years of completed schooling at which students are eligible to drop out conditional on redshirting and quarter of birth. The spirit of the identification scheme used in Angrist and Krueger (1991) is that since laws apply to a student's age, and not their years of completed schooling,

Date: October 5, 2007.

Email: jalipran@econ.upenn.edu

I acknowledge financial support from the Institute of Education Statistics.

<sup>&</sup>lt;sup>1</sup>Redshirting is the term given to parents voluntarily delaying their children's enrollment

variation in students' birth dates means they face different laws regarding the years of schooling they must complete. In our example, students born in  $Q_4$  must stay in school until they complete 10.25 years of schooling, while individuals born in  $Q_1$  only have to complete 9.5 years of schooling before being eligible to drop out. Thus comparing these groups will tell us the treatment effect (ie, the effect of attending an extra year of school against one's will).

There is a problem with this scheme if there is significant redshirting that is not random with respect to expected educational attainment. Let us begin with the framework in Angrist and Imbens (1995), comparing only those children born in the first and last quarters of the year. Define treatment as being eligible to enroll when young (ie, being born in  $Q_4$ ). Let  $S_1$  be the years completed for those exposed to treatment, and  $S_0$  the years of completed schooling for those not exposed to treatment. Then if all children enroll when eligible and comply with the compulsory education law, we have from Table 2 that  $S_1 = 10.25$  and  $S_0 = 9.5$ . Thus

$$E[Y_i|Q_4] - E[Y_i|Q_1] = E[Y_i|S_1] - E[Y_i|S_0] = E[Y_i|S_0 + 0.75] - E[Y_i|S_0].$$

However, consider what happens if not all parents enroll their child when first eligible. Suppose that there are two types, H and L. Assume that both types enroll when eligible when born in  $Q_1$ , but that only those of type H enroll when eligible if they are born in  $Q_4$ . Let  $B_j$  be the cohort of children born j months before their school's cutoff date. Table 3 shows that this assumption is reasonable in the ECLS-K data set. For those of type H,  $S_1 = S_0 + 0.75$ . However, for those of type L,  $S_1 = S_0 - 0.25$ . Thus for those of type H,  $S_1 - S_0 = 0.75 > 0$ , but for those of type L,  $S_1 - S_0 = -0.25 < 0$ . This violates the assumption of Monotonicity.

To illustrate why this creates a problem, consider what would happen if those who would drop out when 15 years of age are all of type L. In this case we would expect that those entering school when younger tend to complete more schooling precisely what is reported in Dobkin and Ferreira (2007). Then comparing outcomes of groups  $Q_4$  and  $Q_1$  would yield something very different from the parameter of interest:

$$E[Y_i|Q_4] - E[Y_i|Q_1] = E[Y_i|S_1] - E[Y_i|S_0] = E[Y_i|S_0 - 0.25] - E[Y_i|S_0].$$

This leaves us unable to interpret the IV estimates, similarly to other cases discussed in Rosenzweig and Wolpin (2000).

#### 2. Redshirting in the ECLS-K Data Set

I am not aware of any historical data on redshirting. Thus I cannot draw any conclusions from the above discussion about the validity of the estimates presented in papers such as Angrist and Krueger (1991), Acemoglu and Angrist (2000), or Lochner and Moretti (2004). Nevertheless, a first glance at this issue can be seen in Table 1, which shows the average age at enrollment for boys in the 1960 Census by quarter of birth (as reported in Angrist and Krueger (1991)), including the difference between the census average and the age predicted by our simple assumptions. Note that there is a large discrepancy in  $Q_4$ . This difference could be generated by schools with cutoff dates before January 1st, but it could also be evidence of significant redshirting in the fourth quarter.

In the ECLS-K data set we observe four groups of children whose birth dates are within one month of their school's cutoff date, as shown in Figure 1:

- **Cohort**  $C_1$ : Children eligible in the fall of 1997 but who chose to wait until the fall of 1998 to enroll.
- **Cohort**  $C_2$ : Children ineligible in the fall of 1997 who waited until the fall 1998 to enroll.

**Cohort**  $C_3$ : Children eligible in the fall of 1998 who chose to enroll.

**Cohort**  $C_4$ : Children ineligible in the fall of 1998 but who chose to enroll anyway.

As shown in Table 3, very few parents would choose to delay their child when first eligible if they are in Cohort  $C_2$ . In contrast, cohort  $C_1$  is made up entirely of children whose parents chose to delay their enrollment when eligible in 1997, while  $C_3$  is made up entirely of children whose parents chose to enroll them when eligible in 1998. This allows us to exploit the difference in the composition of groups  $C_1$ and  $C_3$  to obtain information about those children whose parents would choose to delay their enrollment. Tables 4–9 show how these cohorts differ. The children who delayed enrollment were disproportionately wealthy, white, male, English-speaking, and had better-educated parents. These characteristics are correlated with educational attainment, casting doubt on the validity of the instrument generated by entrance cutoff dates and compulsory education laws.

#### References

- Acemoglu, D. and J. Angrist (2000). How large are human-capital externalities? Evidence from compulsory schooling laws. *NBER Macroeconomics Annual 15*, 9–59.
- Aliprantis, D. (2007). When should children start school? Working Paper.
- Angrist, J. D. and G. W. Imbens (1995). Two-stage least squares estimation of average causal effects in models with variable treatment intensity. *Journal of the American Statistical Association* 90(430), 431–442.
- Angrist, J. D. and A. B. Krueger (1991). Does compulsory attendance affect schooling and earnings? The Quarterly Journal of Economics 106(4), 979–1014.
- Bound, J. and D. A. Jaeger (2000). Do compulsory school attendance laws alone explain the association between quarter of birth and earnings? In S. W. Polachek (Ed.), Worker Well Being, Volume 19, pp. 83–108. Research in Labor Economics.
- Dobkin, C. and F. Ferreira (2007, July). Do school entry laws affect educational attainment and labor market outcomes? *Working Paper*.
- Lochner, L. and E. Moretti (2004). The effect of education on crime: Evidence from prison inmates, arrests, and self-reports. *The American Economic Review* 94(1), 155–189.
- Rosenzweig, M. R. and K. I. Wolpin (2000). Natural "Natural Experiments" in Economics. Journal of Economic Literature 38, 827–874.

Tables

4

TABLES

Table 1: Entrance Age (in Years) by Quarter of Birth

Entrance Age (in Years)								
Entry $Q_1$ $Q_2$ $Q_3$ $Q_4$								
When Eligible Redshirted	$6.5 \\ 7.5$	$6.25 \\ 7.25$	6 7	$5.75 \\ 6.75$				
1960 Census Avg '60 – Eligible	6.45 -0.05	$\begin{array}{c} 6.28 \\ 0.03 \end{array}$	6.08 -0.08	6.07 -0.32				

Table 2: Years of Schooling Completed before Eligible to Drop Out by Quarter of Birth

Years of Schooling Completed							
Entry	$Q_1$	$Q_2$	$Q_3$	$Q_4$			
When Eligible Redshirted	$9.5 \\ 8.5$	$9.75 \\ 8.75$	$\frac{10}{9}$	$10.25 \\ 9.25$			

Table 3: Redshirting in the ECLS-K

Percent Redshirting												
Cohort	$B_1(C_1+C_3)$	$B_2$	$B_3$	$B_4$	$B_5$	$B_6$	$B_7$	$B_8$	$B_9$	$B_{10}$	$B_{11}$	$B_{12}(C_2)$
% Redshirting	28.8	29.8	13.9	14.0	11.3	7.7	8.0	6.8	6.1	5.2	5.4	3.7

## Tables

Different Variables Summarized by Cohort									
Variable	Statistic	ECLS-K	Cohort 1	Cohort 2	Cohort 3	Cohort 4			
Income	Mean Std. Dev. n	52,946 36,375 12,326	$61,098 \\ 40,389 \\ 241$	52,456 34,959 702	$\begin{array}{r} 48,\!531 \\ 34,\!466 \\ 619 \end{array}$	$61,912 \\ 40,146 \\ 57$			
# of Books	Mean Std. Dev. n	$74.35 \\ 59.87 \\ 14,910$	86.89 60.64 300	$75.11 \\ 57.57 \\ 849$	$69.74 \\ 57.45 \\ 744$	$76.84 \\ 65.18 \\ 63$			
# of Children's Records, Tapes, or CD's	Mean Std. Dev. n	$15.37 \\ 17.86 \\ 15,008$	$     \begin{array}{r}       18.26 \\       20.33 \\       303     \end{array} $	$16.09 \\ 18.81 \\ 851$	$     13.99 \\     17.42 \\     750   $	$17.70 \\ 17.68 \\ 63$			
Age at entry to K (in months)	Mean Std. Dev. n	$65.56 \\ 4.26 \\ 15,042$	$68.34 \\ 5.53 \\ 302$	70.63 2.91 851	$60.07 \\ 1.67 \\ 751$	$62.00 \\ 5.25 \\ 63$			
Mother's Age (At First Birth)	Mean Std. Dev. n	$23.98 \\ 5.44 \\ 14,054$	$25.26 \\ 5.74 \\ 288$	$23.93 \\ 5.58 \\ 793$	$23.64 \\ 5.37 \\ 696$	$24.12 \\ 5.49 \\ 57$			
Cost Nonparental Childcare (Pre-K, in \$)	Mean Std. Dev. n	$\begin{array}{c} 43.82 \\ 51.72 \\ 11,465 \end{array}$	$51.43 \\ 49.18 \\ 223$	$\begin{array}{c} 41.50 \\ 46.17 \\ 653 \end{array}$	$39.74 \\ 44.66 \\ 559$	59.38 77.77 54			
Age at first Nonparental Care (in Months)	Mean Std. Dev. n	$22.28 \\ 19.59 \\ 13,047$	$24.08 \\ 20.95 \\ 266$	$22.80 \\ 20.72 \\ 754$	$20.69 \\ 18.24 \\ 637$	$17.71 \\ 17.71 \\ 58$			
Hours Spent in Nonparental Care (Pre-K)	Mean Std. Dev. n	$25.23 \\ 21.73 \\ 14,858$	$22.05 \\ 19.79 \\ 298$	$26.66 \\ 23.31 \\ 839$	$25.33 \\ 22.15 \\ 744$	$27.08 \\ 20.32 \\ 63$			

# Table 4: Descriptive Statistics by Cohort

## Table 5: WIC Benefits

Did the Child Receive any WIC Benefits as a Child (in $\%)$								
	ECLS-K	Cohort 1	Cohort 2	Cohort 3	Cohort 4			
Yes	42.88	31.35	40.66	47.87	39.68			
No	56.00	67.66	58.05	51.06	60.32			
n	15,064	303	851	752	63			

# Table 6: Child's Language to Mother

Language to Mother in $\%$								
Speaks Non-English Language ECLS-K Cohort 1 Cohort 2 Cohort 3 Cohor								
Never	79.53	93.07	80.85	79.26	77.78			
sometimes, Often, or Very Often	18.53 15.064	5.28 303	851	752	63			

Table 7: Race

The Composition of Cohorts by Race (in $\%$ )								
Race	ECLS-K	Cohort 1	Cohort 2	Cohort 3	Cohort 4			
White, Non-Hispanic	56.31	75.29	60.95	53.94	63.01			
Black or African-American, Non-Hispanic	14.20	9.30	13.18	13.87	12.33			
Hispanic, Race Specified	8.52	3.49	7.51	6.58	1.37			
Hispanic, Race Not Specified	8.91	5.52	7.71	11.40	4.11			
Asian	6.35	2.03	5.68	7.17	15.07			
Native Hawaiian or Other Pacific Islander	1.14	_	1.12	2.47	_			
Native American or Native Alaskan	1.80	1.74	1.01	2.00	1.37			
More than one race, Non-Hispanic	2.55	2.62	2.64	2.59	2.74			
n	17,527	344	984	851	73			

Table 8: Gender

The Composition of Cohorts by Gender (in $\%$ )								
Gender	ECLS-K	Cohort 1	Cohort 2	Cohort 3	Cohort 4			
Male	51.15	63.66	53.25	49.82	32.88			
Female	48.78	36.34	46.75	50.18	67.12			
n	$17,\!554$	344	986	851	73			

Table 9: Parent's Highest Education Level

Parent's Highest Education Level by Cohort in $\%$								
	ECLS-K	Cohort 1	Cohort 2	Cohort 3	Cohort 4			
8th Grade or Below	3.09	0.95	2.76	3.05	1.43			
9th-12th Grade	6.55	3.15	6.62	6.62	4.29			
High School Diploma/Equivalent	23.80	17.67	25.69	27.61	21.43			
Voc/Tech Program	6.77	5.36	8.27	7.12	4.29			
Some College	26.50	27.13	25.69	27.86	28.57			
Bachelor's Degree	18.53	26.81	16.43	17.05	21.43			
Grad/Professional School - No Degree	2.56	3.15	2.65	1.65	2.86			
Master's Degree (MA, MS)	7.50	7.89	6.73	6.49	12.86			
Doctorate or Professional Degree	4.71	7.89	5.18	2.54	2.86			
n	$15,\!961$	317	907	786	70			

Figures

FIGURES



Figure 1: ECLS-K Cohorts

7