



## Internet traffic dynamics

Madden, Gary G and Coble-Neal, Grant Curtin University of Technology, Perth, Australia, Curtin University of Technology, Perth, Australia

2004

Online at http://mpra.ub.uni-muenchen.de/10827/ MPRA Paper No. 10827, posted 29. September 2008 / 09:28

### Internet traffic dynamics

GARY MADDEN AND GRANT COBLE-NEAL



Gary Madden is a Professor at CEEM, Curtin University of Technology, Perth, Australia

The telecommunications industry has evolved at unprecedented rates with current estimates suggesting that seven percent of the world's population now has access to the Internet. However, such growth has stimulated vigorous competition in national and international telecommunications markets leading to a price-cost margin squeeze and unsustainable rates of network expansion. This study demonstrates the reliability of established extrapolation methods for forecasting bandwidth demand and provides network managers with the opportunity to observe Internet traffic dynamics. The ability to anticipate periods of peak use and surplus capacity is likely to pay dividends in terms of a more targeted approach to network expansion plans.



Grant Coble-Neal is a PhD student

### I Introduction

Telecommunications bandwidth has grown at an unprecedented rate in recent years with current estimates suggesting that seven percent of the world's population now has access to the Internet. Indeed, while North America still leads the world in terms of adoption, Table I shows that nearly half of all users now reside outside the Unites States (US). Given the proliferation of telecommunication applications such as Internet browsing, email, Voice over Internet Protocol (VoIP) and video broadband, as well as strong volume growth in the traditional Public Switched Telephone Network (PSTN), it is likely that the growth exhibited in Figure I is likely to continue in the near future. In 1999, for example, standard international telephone traffic grew by over 15 % in 1999 to 107.8 billion minutes. Although VoIP still accounts for only a small fraction of the total voice market, traffic grew tenfold to over 1.7 billion minutes, with the fastest growth occurring in US-developing country outgoing routes (TeleGeography, 2001b). Not surprisingly, such growth has stimulated vigorous competition in both national and international telecommunications markets. At the national level, countries such as Germany and Israel have experienced spectacular returns to deregulation with long-distance calling market prices dropping 91 % and 94 %, respectively (Newton 2000). Similarly, Tele-Geography (2000) global trends suggest that call volume growth has been stimulated largely by successive price cuts. Technology has played a substantial role, initially by least-cost routing arrangements such as callback and traffic refile, and more recently by routing voice and facsimile transmissions through the Internet, thus providing competitors with the means of reducing or avoiding international settlements.

In response, incumbent carriers have sought to increase their scale so as to defend revenues and deter entry by new competitors. According to TeleGeography (2001b), submarine cables increased the aggregate trans-Atlantic bandwidth by a factor of 12 to

Region	Hosts ('000s)	%	Users ('000s)	%
Africa	305.3	0.25	3,337.6	0.77
Asia	10,280.3	8.28	81,733.9	18.73
Europe	23,365.6	18.81	103,827.0	23.80
Oceania	2,297.1	1.85	19,995.1	4.58
Central America	496.0	0.40	1,643.7	0.38
South America	1,367.6	1.10	18,001.1	4.13
North America	86,098.5	69.32	207,734.0	47.62
Total	124,210.3	100.00	436,272.4	100.00

Table I Internet hosts and users by region (Source: Telcordia Technologies, http://www.netsizer.com/)

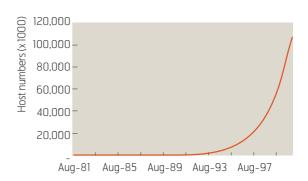


Figure I Internet host growth 1981–2001 (Source: Internet Software Consortium, http://www.isc.org/)

<sup>1)</sup> A terabit is one million million bits.

over two terabits per second in just one year. <sup>1</sup> Overall, telecommunication's three basic building blocks, fibre, digital signal processors and routers, are improving their capacity for throughput ten times faster than the mainstream computer industry (Newton, 2000). <sup>2</sup> High-speed routers, for example, are now switching terabits of information each second. In addition, laboratory tests show that a fibre strand the width of a human hair can transmit three trillion bits per second, enough to transmit the entire world's Internet (Newton 2000).

However, network expansion is expensive. Construction costs can range from USD 4,000 to USD 3 million per kilometer depending on the choice of upgrade level of dense wavelength-division multiplexing (TeleGeography, 2001a). Similarly, submarine cable installation costs range from USD 0.5 billion for a 10,000 kilometer cable to USD 2.0 billion for 30,000 kilometers. Meanwhile, carriers' mainstream business continues to be cannibalized by the proliferation of Internet Service Providers purchasing flat rate access to upstream network only to offer VoIP to the incumbent carriers' own customer base.

Thus, while telecommunications traffic continues to grow at a rapid rate, networks are expanding at economically unsustainable rates. Such long-term impacts of technological change are always hard to forecast, but that task is especially difficult in the case of e-commerce, where markets are currently very far from equilibrium. In the 'land rush' to secure Internet real estate, to gain first-mover market position and other advantages, many firms are pursuing strategies that are properly interpreted as the payment of one-time, largely sunk entry costs (Borenstein and Saloner 2001).

In this environment, common carriers will need to develop improved forecast models to accurately predict bandwidth demand and target network expansion. This paper uses Internet Traffic Report as a data source that measures Internet bandwidth loads and availability on a continuous basis.<sup>3)</sup> The data is generated by a test called a "ping", which measures round-trip travel time along major paths on the Internet. Several servers in different areas of the globe perform the same ping at the same time and an index

based on average response times across test servers is calculated.

The traffic index produces a score in the ranges [0, 100]. A zero score is 'slow' and 100 is 'fast' by comparing the current response of a ping echo to all previous responses from the same router over the past seven days. Response time in reference to Internet traffic is how long it takes for data to travel from point A to point B and back (round trip). A typical response time on the Internet is 200 milliseconds. To be continually accurate and useful, statistics are gathered at many geographically diverse routers and many geographically diverse 'satellite' locations to test from.

This study obtains alternative forecasts of broadband capacity using ARMA, ARARMA, Holt, Holt-D exponential smoothing, Naïve, Robust Trend, as well as a deterministic trend model. The ARMA method is the well-established Box-Jenkins approach to model systematically recurring patterns in stationary data. The ARARMA model, proposed by Parzen (1982), is designed to model long memory processes, using an initial autoregressive specification to filter potentially non-stationary data. Holt's exponential smoothing filters random noise and extrapolates the underlying linear trend contained in the data while Holt's-D models time series as a linear trend decaying towards a constant. Robust Trend essentially models a time series as a stochastic trend with an outlier filter. Thus, the trend is allowed to adapt as observations accumulate while providing a restrained reaction to sudden unexpected pulses in the data. Introduced by Grambsch and Stahel (1990), this technique has been shown to perform best for homogenous telecommunications data by Fildes et al. (1998). Naïve is the simple random walk extrapolation and Trend provides a deterministic alternative to Holt, Holt-D and Robust Trend. Both Naïve and Trend are included as indicative benchmarks with which to compare forecast accuracy of the alternative methods.

The paper is organised as follows: Section II describes sample data, and a discussion of the various forecast models is contained in Section III. Model results are presented in Section IV and concluding remarks are presented in Section V.

<sup>2)</sup> A digital signal processor (DSP) is a specialized micro-processor that performs calculations on digitized signals that were originally analogue (e.g. voice). DSPs are used extensively for echo cancellation, call progress monitoring, voice processing, and voice and video signal compression. Routers are the central switching offices of the Internet and are the interface devices between different network architectures such as x.25, Frame Relay and Asynchronous Transfer Mode. These intelligent devices decide which backbone network to transmit data, monitor the condition of the network and redirect traffic to avoid congestion.

<sup>3)</sup> The Internet Traffic Report URL is http://www.internettrafficreport.com/index.html.

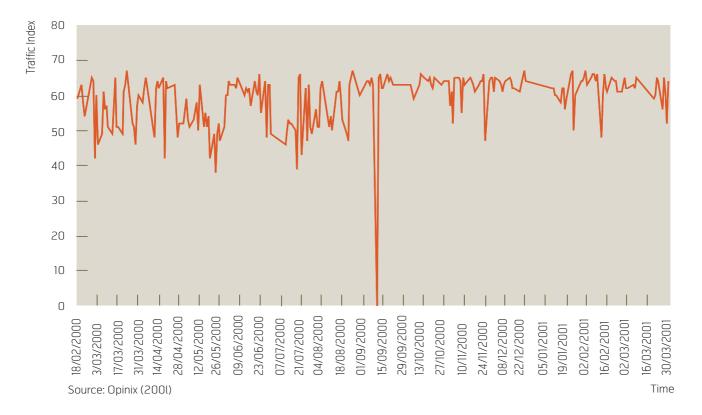


Figure II Japan dm-gw1.kddnet.ad.jp

### II Data

The data set described and analyzed in this paper is comprised of 59 time-series, each containing 232 observations. These data are sampled from a continuous data generating process and sampled daily at 7 AM Australian Eastern Standard Time weekdays for the period February 18, 2000 to March 30, 2001. A representative specimen of these data is shown in Figure II. As described, the data oscillate between zero and 100 and appear to exhibit characteristics typical of stationary series. Another feature, which is common to many of the series in this data set, is the sudden downward spike in the series. These spikes indicate brief periods of unusually high congestion and, depending on the motivation for generating forecasts, can either be treated as outliers which are atypical of the series or incorporated in the model as an infrequent but important characteristic of the data generating process.

Summary statistics, reported in Table II, highlight the frequency of the downward spikes with 28 of the 59 routers reporting zero minimum values. Regions represented include East Asia, Australia, Western Europe, Israel, Russia, North America and South America. Absent regions include Antarctica, Africa

and most parts of the Middle East. The Denver denver-br2.bbnplanet.net router is recorded as providing the fastest response while AOL1 pop1-dtc.atdn.net has the lowest response time. On average, the Perth1 opera.iinet.net.au router provides the consistently fastest response while Yahoo fe3-0.cr3.SNV.global-center.net is typically the slowest.<sup>4)</sup>

Following Fildes (1992) we analyze the series in terms of frequency of outliers, strength of trend, degree of randomness and seasonality, with the results shown in Figure III through Figure V. An observation  $(X_t)$  is classed as an outlier if  $X_t < L_x - 1.5(U_x - L_x)$  or  $X_t > L_x + 1.5(U_x - L_x)$ , where  $L_x$  denotes the lower quartile and  $U_x$  the upper quartile. The strength of trend is measured by the correlation between the series (with outliers removed) and a time trend, with the absolute value of the trend indicating its strength. Randomness is measured by estimating the regression:

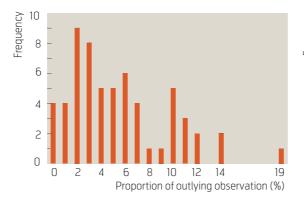
$$X_{t}^{'} = \alpha + \beta t + \delta_{1} X_{t-1}^{'} + \delta_{2} X_{t-2}^{'} + \delta_{3} X_{t-3}^{'}, \tag{1}$$

where  $X_t$  denotes the series  $X_t$  with outliers removed. The adjusted  $R^2$  is used to measure the variation explained by the model. A high  $R^2$  indicates low

<sup>4)</sup> Time of day effects and scale of demand may have an impact on router performance. For example, the Perth router services a small market and is likely to have relatively low congestion early in the morning, while in real time, the Yahoo router may be at peak demand in the mid-late afternoon.

Router	Average	Std.deviation	Min.	Max.
China2 beijing-bgw1-lan.cernet.net	57.68	10.23	22.00	87.00
HK1 hkt004.hkt.net	58.01	7.54	16.00	72.00
India cust-gw.Teleglobe.net	61.60	8.98	13.00	81.00
Japan dm-gw1.kddnet.ad.jp	58.96	7.54	0.00	67.00
Malay fe1-0.bkj15.jaring.my	56.05	11.53	4.00	79.00
Phil3 tridelinc.Sacramento.cw.net	60.98	5.60	31.00	67.00
Sing1 pi-s1-gw1.pacific.net.sg	58.11	8.39	4.00	69.00
Sing2 gateway.ix.singtel.com	60.81	7.81	27.00	72.00
Taiwan cs4500-fddi0.ficnet.net.tw	59.03	9.06	26.00	77.00
Bris FddiOcore1.Brisbane.telstra.net	61.15	7.92	0.00	69.00
Canb FddiO-O.civ3.Canberra.telstra.net	61.02	7.55	0.00	68.00
Gosfor EthernetO.gos2.Gosford.telstra.net	61.44	7.82	0.00	69.00
Melb mc5-a2-0-4.Melbourne.aone.net.au	60.28	7.43	14.00	70.00
Perth1 opera.iinet.net.au	62.63	6.67	0.00	72.00
Perth2 Fddi0-0.wel1.Perth.telstra.net	61.32	8.10	0.00	68.00
Syd1 sc2-exch-fe0.Sydney.aone.net.au	61.04	6.19	27.00	74.00
Syd2 FastEthernet0 Sydney. telstra.net	60.88	8.03	0.00	70.00
Terri terrigal-gw.terrigal.net.au	52.56	12.99	0.00	69.00
Denmar albnxi3.ip.tele.dk	56.29	9.20	0.00	68.00
Fran1 isicom-gw.iway.fr	53.72	10.68	0.00	67.00
Fran2 rbs2.rain.fr	54.90	8.82	0.00	69.00
Greece athens1.att-unisource.net	60.69	9.79	0.00	71.00
Holl1 amsterdam3.att-unisource.net	58.31	5.83	41.00	70.00
Holl2 hvs01.NL.net	56.23	9.22	3.00	68.00
Ice Reykjavik14ASI.isnet.is	57.81	10.86	0.00	68.00
Israel haifa-rtr.actcom.co.il	61.07	9.26	0.00	75.00
Italy Pa6.seabone.net	60.34	7.47	0.00	70.00
Norway ti09a95.ti.telenor.net	56.32	13.64	0.00	68.00
Russ1 ru-msk-en-1.teleport-tp.net	58.65	11.63	0.00	72.00
Swed1 apv-i1-pos1stockholm.telia.net	57.26	10.82	0.00	68.00
Swed2 mlm1-core.swip.net	58.91	6.71	28.00	67.00
UK1 atm0-0-x.lon2gw1.uk.insnet.net	57.92	10.93	0.00	66.00
UK2 access-th-3-e0.router.technocom.net	59.97	8.62	31.00	73.00
AOL1 pop1-dtc.atdn.net	50.04	8.76	15.00	63.00
AOL2 pop1-rtc.atdn.net	50.74	9.22	25.00	64.00
Atlant atlanta1-br1.bbnplanet.net	49.68	10.20	11.00	66.00
Bost1 cambridge1-br1.bbnplanet.net	55.29	11.33	7.00	73.00
Bost2 core3-hssi5-0.Boston.cw.net	49.53	10.58	0.00	65.00
Canad1 core-fa5-0-0.ontario.canet.ca	55.01	8.03	23.00	67.00
Canad2 border6.toronto.istar.net	51.11	15.82	0.00	74.00
Chical FddiO.AR1.CHI1.Alter.Net	53.09	9.19	18.00	69.00
Dallas dallas1-br2.bbnplanet.net	53.25	12.63	0.00	67.00
Denver denver-br2.bbnplanet.net	49.85	13.85	0.00	90.00
Detroi eth1-0-0.michnet1.mich.net	47.91	11.84	2.00	65.00
LA1 borderx2-fddi-1.LosAngeles.cw.net	54.41	9.21	0.00	66.00
LA2 la32-0-br1.ca.us.ibm.net	57.37	6.97	32.00	67.00
Mex4 core2-mexico.uninet.mx	54.87	13.09	0.00	69.00
Mex5 dgsca-cs.core-atm.unam.mx	48.60	14.02	2.00	67.00
Mex6 rr1.mexmdf.avantel.net.mx	57.76	10.13	8.00	67.00
NY p2-0-0.nyc4-br1.bbnplanet.net	49.12	9.53	19.00	70.00
Sacram border70.Sacramento.cw.net	57.30	6.74	28.00	67.00
SanFrn core1.SanFrancisco.cw.net	56.72	7.02	33.00	66.00
Seattl border3-fddi-0.Seattle.cw.net	54.46	9.07	7.00	66.00
Yahoo fe3-0.cr3.SNV.globalcenter.net	46.63	19.09	0.00	67.00
Brazil routrjo07.embratel.net.br	59.67	8.61	15.00	70.00
Chile bwl-gw-net3.rdc.cl	59.67	14.20	0.00	
	57.08	14.20	5.00	67.00 73.00
Colom1 gip-bogota-1-ethernet0-1.gip.net				
Colom2 impsat.net.co	58.85	8.95	13.00	71.00
Venez cha-00-lo0.core.cantv.net	58.63	8.96	0.00	81.00

Table II Summary statistics (Source: Opinix, 2001)





randomness while a low  $R^2$  reveals high randomness. Deterministic seasonality is estimated by regressing the series on an intercept and dummy variables which equal one when t = s, where t denotes observation  $X_t$ 's position in time and s corresponds to the frequency of the seasonality. For example, to test the hypothesis that Mondays are statistically different to bandwidth capacity for the rest of the week,  $t = \{1, 2, 3, 4, 5, ..., T\}$ ,  $s = \{1, 5, 10, 15, ..., T\}$  and dummy variable  $D_{Monday} = 1$  for t = s, zero otherwise.

Figure III reveals that half the series contain between one and five percent outliers. In percentage terms these data appear slightly more heterogeneous than Fildes' (1992) telecommunications data. As indicated in the specimen displayed in Figure II, Figure III shows that the data are generally uncorrelated with time. This contrasts with Fildes (1992) where the data there exhibit strong negative trends. Moreover, the histograms in Figure IV and Figure V reveals the

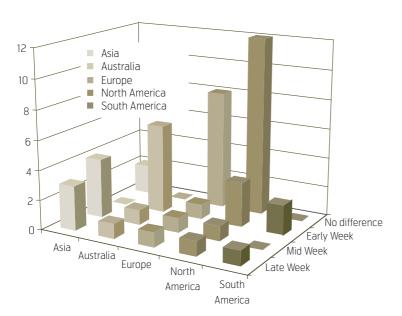
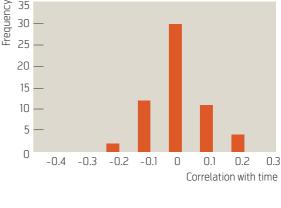


Figure VI Daily variation in capacity utilization



2 -0.02 -0.01 0 0.01 0.02 0.03 0.04 0.05 0.07 Adjusted r squared

Figure V Variation explained by linear/AR

variation in the data presents a high degree of randomness with virtually no serial correlation.

Finally, Figure VI presents some evidence of regularity in weekly capacity variation aggregated by region. As shown, there appear to be regular dips occurring on different days across regions. Asia generally experiences lower traffic volumes across the later part of the week, while the majority of Australian routers have excess capacity in the early part of the week. By contrast, Europe and North America experience relatively smooth traffic flows, possibly reflecting more sophisticated capacity pricing regimes and/or advanced network management systems. Finally, variations in South American Internet traffic are tied to specific routers.

In addition to daily variations in traffic volumes, regressions are conducted to test for regularity in weekly and monthly patterns. Weekly variations are virtually non-existent with only six routers revealing regular spikes across weeks. Surprisingly, given the short time series, significant monthly variation was found in 95 % of routers. Although the sustained increase in traffic is too haphazard across routers to discern a cyclical pattern, most experience statistically significant increases for an average of two

months with some routers showing surges of up to three months. Given the nature of the index calculations, this possibly reflects the average lagged response time required before routers are expanded to cope with the increased traffic. Once routers are expanded, the Internet traffic index for the router is likely to increase, reflecting the permanently increased capacity.

Overall, the data series exhibit a high degree of randomness and regular spikes in index scores. Compared to the telecommunications data analyzed in Fildes (1992) and Fildes et al. (1998), these data appear considerably more heterogeneous and less predictable.

# III Forecast models and accuracy measures

Forecast models considered are univariate ARMA, ARARMA, Holt, Holt-D exponential smoothing, Robust Trend, with Naïve and Trend benchmarks. All of these forecast methods have been shown to be reliable by Makridakis et al. (1982), Fildes (1992), Fildes et al. (1998) and Makridakis and Hibon (2000) and consistently perform in the annual M-Competition. Implicit in these analyses however, is that the data are nonstationary, while the data analysed here are believed to be stationary. Given this fundamental difference in assumption some of the forecast techniques have been modified to avoid problems associated with over-differencing. For example, the ARMA method is applied rather than ARIMA. ARARMA explicitly questions the practice of differencing to achieve stationarity and has the advantage of utilising information contained in the data normally lost when differencing. Moreover, the approach outlined in Parzen (1982) contains a formal method of determining when it is appropriate to apply the AR filter and hence, the method is adopted intact. Holt and Holt-D methods are techniques for extrapolating the underlying trend that may be present in the data. Although the deterministic trend correlations are mostly zero, short-run trends may prevail and therefore Holt and Holt-D may be appropriate given their simplicity and reliability. However, to ensure the opportunity for accuracy is maximised, the parameter is optimised (rather than being arbitrarily set once) at each time origin as recommended in Fildes et al. (1998). Robust Trend, however, is modified by not differencing the

data before calculating the stochastic trend. The perceived advantage in adopting this method is the outlier filter and its use of the median rather than mean in the estimator, which may provide some advantage over the simple random walk extrapolation. Thus for direct comparative purposes, Naïve is included as a benchmark model. If the outliers do not bias the estimates, the forecasts will be hard to improve on, given the reported properties of the data.

The choice of accuracy measures used in this analysis is guided by the recommendations of Armstrong and Collopy (1992). For the reasons outlined in that paper, the Mean Absolute Percentage Error (MAPE), Median Absolute Percentage Error (MdAPE), % Better, Geometric Mean Relative Absolute Error (GMRAE) and Median Relative Absolute Error (MdRAE) are used. Both GMRAE and MdRAE are Winsorized as recommended by Armstrong and Collopy. Mean square error measures are avoided since these statistics are scale dependent and sensitive to outliers.

### **IV** Forecast results

In order to identify forecast methods that perform well four sets of forecasts are created by dividing the data into overlapping time intervals, with each forecast method using 114 observations to forecast over the next 60 observations.<sup>5)</sup> In effect, this approach uses a rolling window beginning at the first observation and steps forward 10 days, re-estimating the forecasts over the next 114 observations. The overall result is 295 forecasts per method with which to judge forecast performance. In evaluating the reliability of the alternative methods, forecasts are compared with actual values retained in the post-sample data.

Table III presents the main results, measuring fore-cast accuracy in terms of the average absolute error. In general, the various trend extrapolation methods performed better than the more sophisticated ARMA and ARARMA methods while both Holt and Holt-D consistently performed worst.<sup>6)</sup> As shown at the bottom of Table III, the modified Robust Trend method produced the most accurate forecasts approximately 65 % of the time with an average 7.5 % error. Holt-D performed best on a number of occasions, which is probably due to the occasional appearance of weak trends in these data.

<sup>5) 60</sup> day forecasts are necessary due to the existence of standard capacity contracts.

<sup>6)</sup> Note that ARMA often provides the same accuracy as the Naïve trend forecast. This is due to the general-to-specific modeling approach that uses the Akaike Information Criterion to identify the best fitting model from a grid of up to six autoregressive and moving average lags. In many cases, this algorithm identified Naïve as the optimal model.

Time	Method	Forecasting horizon									
origin		1	21	41	61	1-6	1-26	1-46	1-61		
1	ARARMA	0.04	0.17	0.06	0.17	0.25	0.20	0.14	0.23		
	ARMA	0.18	0.13	0.17	0.11	0.12	0.10	0.00	0.05		
	Filtered-T	0.02	0.18	0.48	0.02	0.23	0.20	0.14	-		
	Holt	0.23	0.23	0.31	0.34	0.13	0.41	0.49	0.63		
	Holt-D	0.07	0.02	0.00	0.10	0.00	0.19	0.09	0.13		
	Naïve	0.18	0.13	0.17	0.11	0.12	0.10	0.00	0.0		
	Robust-T	0.03	0.04	0.01	0.05	0.23	0.02	0.10	0.0		
10	ARARMA	0.26	0.14	0.04	0.19	0.14	0.09	0.08	0.2		
	ARMA	0.18	0.13	0.17	0.11	0.12	0.10	0.01	0.0		
	Filtered-T	0.06	0.16	0.06	0.07	0.04	0.20	0.00	0.0		
	Holt	0.05	0.14	0.24	0.52	0.70	0.85	0.92	0.9		
	Holt-D	0.07	0.02	0.00	0.10	0.04	0.16	0.06	0.11		
	Naïve	0.18	0.13	0.17	0.11	0.12	0.10	0.01	0.0		
	Robust-T	0.03	0.04	0.01	0.05	0.23	0.02	0.10	0.0		
20	ARARMA	0.15	0.04	0.04	0.09	0.31	0.23	0.00	0.12		
	ARMA	0.18	0.12	0.16	0.11	0.13	0.10	0.01	0.0		
	Filtered-T	0.02	0.06	0.05	0.07	0.25	0.20	0.12	0.0		
	Holt	141.54	177.14	193.97	233.32	0.97	0.77	1.25	1.3		
	Holt-D	2.34	2.61	2.48	2.76	0.13	0.09	0.03	0.0		
	Naïve	0.16	0.10	0.14	0.09	0.13	0.10	0.01	0.0		
	Robust-T	0.03	0.04	0.01	0.05	0.23	0.02	0.10	0.0		
30	ARARMA	0.03	0.02	0.60	0.04	0.13	0.24	0.14	0.2		
	ARMA	0.16	0.10	0.14	0.09	0.12	0.10	0.01	0.0		
	Filtered-T	0.21	0.02	0.00	0.05	0.00	0.11	0.07	0.2		
	Holt	2.19	2.94	3.61	4.71	311.88	292.82	376.52	392.7		
	Holt-D	0.10	0.02	0.04	0.05	0.07	0.15	0.05	0.10		
	Naïve	0.17	0.12	0.16	0.10	0.13	0.10	0.01	0.0		
	Robust-T	0.03	0.04	0.01	0.05	0.23	0.02	0.10	0.0		
Mean	ARARMA	0.12	0.09	0.19	0.12	0.21	0.19	0.09	0.2		
	ARMA	0.17	0.12	0.16	0.10	0.12	0.10	0.01	0.0		
	Filtered-T	0.08	0.10	0.15	0.05	0.13	0.18	0.08	0.0		
	Holt	36.00	45.11	49.54	59.72	78.42	73.71	94.79	98.9		
	Holt-D	0.65	0.66	0.63	0.75	0.06	0.15	0.06	0.0		
	Naive	0.17	0.12	0.16	0.10	0.12	0.10	0.01	0.0		
	Robust-T	0.03	0.04	0.01	0.05	0.23	0.02	0.10	0.0		

Table III Mean absolute percentage error. Note: Bolded minimum MAPE statistic indicates best performing method

Further evaluation is provided in Table IV, which presents the GMRAE and MdRAE forecast error measures. Both GMRAE and MdRAE compare each method to a no change benchmark forecast for comparative purposes. Thus, a score less than one indi-

cate the forecast method is at least more reliable than the simplest extrapolation. Using these criteria, it is apparent that both Filtered Trend and Robust Trend consistently outperform the alternatives.

A factor often considered important is the variation in forecast accuracy over the forecast horizon. For example, evidence from M-Competition results indicates that some methods are better for short-term forecasts, while others perform best over a longer horizon. Examination of Figure VII, which shows forecast errors for the time period with the least number of outliers, indicates that Holt-D reliably forecast variation in bandwidth capacity for period one through 41, closely followed by Robust Trend. Interestingly, ARMA proved most resistant to the disturbance experienced for periods 46 through 60. A possible explanation for this is the ability of the ARMA method to better model periodic spikes in congestion while both Robust and Filtered Trend provide a muted adaptation to sudden large disturbances.

Figure VIII presents MdRAE statistics calculated across all time origins. This statistic provides a measure that is less susceptible to distortion than the MAPE for series where actual values frequently take zero values. As shown, this measure more clearly distinguishes the performance of the alternatives. Holt-D and Holt (omitted due to substantially larger error measures) are by far the worst performers. By contrast, ARMA, Filtered and Robust Trend are clustered closely together ranging between 0.5 and one. Not surprisingly, ARMA indicates greater variability with occasional brief spikes above one and below 0.5 while both trend models produce a more consistent estimate. Of interest is the robustness of these methods with little deterioration as the forecast horizon increases.

Finally, Table V reveals the proportion of times each forecast performed better than the random walk extrapolation across 295 forecasts. Clearly, both Naïve and Robust Trend are the most consistent with the results showing that forecasters can expect these methods to perform better than random walk extrapolation 60 % of the time. As a comparison of best to worst, Robust Trend is on average six times more accurate than Holt.

Overall, the results show that bandwidth capacity can be reliably forecast. The MAPE statistics show that Robust Trend tracks the actual index value with average variation of 7.5 % while ARMA is capable of corroborating long horizon forecasts. The inherent stationarity of these data may explain the relative failure of Holt and Holt-D. Both models work best with non-stationary data with a substantial noise-to-signal ratio. Implicit in the implementation of these models is that model parameters are optimized by first- and second-differencing series. The consequence of over-differencing data is the introduction of a unit-root in the error term and estimation of spurious trends.

	Geometric Mean RAE							
Method	1	12	24	36	48	60		
ARARMA	1.24	1.78	1.87	1.24	1.65	1.54		
ARMA	0.99	0.80	0.79	1.00	0.93	1.05		
Filtered-T	0.98	0.57	0.75	0.70	0.63	0.77		
Holt	4.50	5.39	4.79	6.17	5.76	5.61		
Holt-D	1.37	1.23	1.11	1.34	1.23	1.24		
Naïve	0.99	0.80	0.79	0.98	0.93	1.05		
Robust-T	1.10	0.50	1.13	0.67	0.50	1.00		
	Median RAE							
ARARMA	1.37	2.10	1.92	1.31	2.02	1.53		
ARMA	1.00	0.74	0.80	0.99	0.94	1.01		
Filtered-T	0.97	0.65	0.87	0.91	0.77	0.83		
Holt	7.54	8.19	7.32	8.24	8.20	8.28		
Holt-D	2.81	2.88	2.77	3.01	3.14	2.92		
Naïve	1.00	0.73	0.81	0.98	0.96	1.01		
Robust-T	1.25	0.56	1.17	0.90	0.64	1.08		

Table IV Geometric mean RAE and median RAE. Note: Bolded statistic indicates best performing method

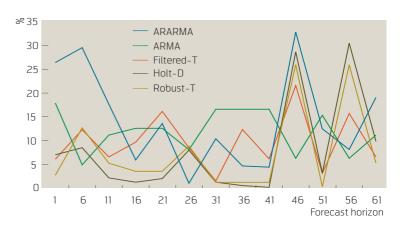


Figure VII Mean absolute percentage error for time origin 10

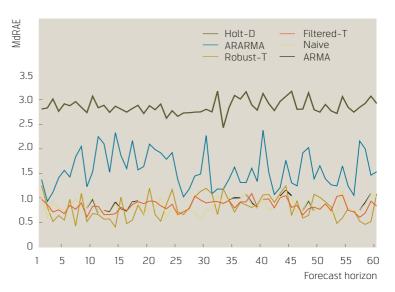


Figure VIII Median relative absolute error across all time origins Note: Holt omitted form chart to show detail

Method	Forecast horizon										
	1	6	12	18	24	30	36	42	48	54	60
ARARMA	37	41	28	43	35	26	44	48	38	41	31
ARMA	51	58	55	54	64	64	51	61	52	46	49
Filtered-T	47	53	62	53	51	41	53	55	59	47	56
Holt	11	12	8	14	10	7	6	6	8	6	8
Holt-D	43	48	47	44	55	51	43	53	45	42	43
Naïve	51	58	55	54	65	64	51	61	51	46	49
Robust-T	32	51	63	56	32	37	53	54	68	63	36

Table V Percent better. Note: Bold indicates best performing method

### **V** Conclusion

Telecommunications bandwidth has grown recently at an unprecedented rate with current estimates suggesting that seven percent of the world's population now has access to the Internet. However, globalisation of the telecommunications industry has led to unsustainable network expansion. In the future, carriers will need to develop accurate forecasts as an aid to a carefully targeted approach to expansion plans. This study demonstrates that relatively simple extrapolation techniques can provide a useful input into explaining broadband traffic movements.

The forecast techniques adopted here are extrapolation methods that have performed well in the M-Competition and are easily implemented. This study also highlights the need to better understand data generation characteristics, at least in a broad sense, and suggests that mechanically differencing data without reference to the characteristics exhibited data can yield substantially inferior results. Finally, despite the high degree of randomness and the high frequency of outliers, Robust Trend again performed best for telecommunications data.

In general, however, univariate extrapolation techniques can at best provide systematic benchmarks on observed data. For more insightful analysis, it is necessary to develop structural economic models using price, income data and traffic data. Among the benefits of such models are the ability to anticipate cyclical fluctuations due to economic factors external to the telecommunications industry, the estimation of price and income elasticities and as a means of determining the degree of reaction and interaction between competitors. The important distinction in adopting this approach is that economic analysis relates to the market for the service that generates these traffic flows. The release of such competitive intelligence would likely provide carriers with substantially great benefits and help to ensure maximal returns to their increasingly scarce investment funds.

### References

Armstrong, J S, Collopy, F. Error measures for generalizing about forecast methods: Empirical comparisons. *International Journal of Forecasting*, 8, 69–80, 1992.

Borenstein, S, Saloner, G. Economics and electronic commerce. *Journal of Economic Perspectives*, 15, 3–12, 2001.

Fildes, R. The evaluation of extrapolative forecasting methods. *International Journal of Forecasting*, 8, 81–98, 1992.

Fildes, R, Hibon, M, Makridakis, S, Meade, N. Generalising about univariate forecasting methods: Further empirical evidence. *International Journal of Forecasting*, 14, 339–358, 1998.

Grambsch, P, Stahel, W A. Forecasting demand for Special telephone services. *International Journal of Forecasting*, 6, 53–64, 1990.

Hampel, F R, Ronchetti, E M, Rousseeuw, P J, Stahel, W A. *Robust Statistics: The Approach Based On Influence Functions*. New York, Wiley, 1986.

Makridakis, S, Andersen, A, Carbone, R, Fildes, R, Hibon, M, Lewandowski, R, Newton, J, Parzen, E, Winkler, R. The accuracy of extrapolation (time series) methods; Results of a forecasting competition. *Journal of Forecasting*, 1, 111–153, 1982.

Makridakis, S, Hibon, M. The M3-Competition: results, conclusions and implications. *International Journal of Forecasting*, 16 (4), 451–476, 2000.

Newton, H. *Newton's Telecom Dictionary*. New York, CMP Books, 2000.

Parzen, E. ARARMA models for time series analysis and forecasting. *Journal of Forecasting*, 1, 67–82, 1982.

TeleGeography Inc. *TeleGeography 2000: Global Telecommunications Traffic Statistics and Commentary.* Washington, TeleGeography Inc., 2000.

TeleGeography Inc. (2001a) *TeleGeography 2001:* Global Telecommunications Traffic Statistics and Commentary. Washington, TeleGeography Inc., 2001.

TeleGeography Inc. (2001b) *International Bandwidth.* Washington, TeleGeography Inc., 2001.

Router	Location	Current index	Response time (ms
Asia			
beijing-bgw1-lan.cernet.net	China	66	287
hkt004.hkt.net	HongKong	61	267
cust-gw.Teleglobe.net	India	65	331
tlv-L1.netvision.net.il	Israel	66	471
haifa-rtr.actcom.co.il	Israel	67	262
hfa-L1.netvision.net.il	Israel	61	544
gsr-ote1.kddnet.ad.jp	Japan	66	201
doji-alp2-2-1-3-1.mcnet.ad.jp	Japan	65	216
POSO-2.oskg2.idc.ad.jp	Japan	65	201
fe1-0.bkj15.jaring.my	Malaysia	66	263
pi-s1-gwl.pacific.net.sg	Singapore	66	334
gateway.ix.singtel.com	Singapore	66	285
cs4500-fddi0.ficnet.net.tw	Taiwan	66	264
ntt-pc-communications.Tokyo.cw.net	Tokyo	66	211
Australia			
GigabitEthernet5-1.chabrisbane.telstra.net	Brisbane	66	418
Pos6-0.woo-core1.Brisbane.telstra.net	Brisbane	66	427
FddiO-O.civ3.Canberra.telstra.net	Canberra	64	417
border-gw03-atm301.powertel.net.au	Gold Coast	57	373
EthernetO.gos2.Gosford.telstra.net	Gosford	66	397
mc5-a2-0-4.Melbourne.aone.net.au	Melbourne	0	0
Pos5-0.exi-core1.Melbourne.telstra.net	Melbourne	66	401
So-O-O-1.XR1.MEL1.ALTER.NET	Melbourne	64	301
opera.iinet.net.au	Perth	63	351
FddiO-O.wel1.Perth.telstra.net	Perth	66	344
c3600.elink.net.au	Perth	66	328
sc2-exch-fe0.Sydney.aone.net.au	Sydney	65	282
FastEthernet0-0-0.pad8.Sydney.telstra.net	Sydney	66	403
So-3-3-1.XR2.SYD2.ALTER.NET	Sydney	63	296
FastEthernet0-0-0.pad13.Sydney.telstra.net	Sydney	66	400
bb2-gige5-0.rdc1.nsw.excitehome.net.au	Sydney	66	261
terrigal-gw.terrigal.net.au	Terrigal	64	554
Europe			
albnxi3.ip.tele.dk	Denmark	67	189
r3-AT2-0-1-Pas5.Hel.Fl.KPNQwest.net	Finland	67	228
isicom-gw.iway.fr	France	64	206
rbs2.rain.fr	France	53	239
feth-0-1-0.cr1.Stuttgart.seicom.NET	Germany	67	243
athens5.gr.eqip.net	Greece	66	264
amsterdam51.nl.eqip.net	Holland	66	180
194.atm1-0-0.hr1.ams6.nl.uu.net	Holland	66	194
Reykjavik14ASI.isnet.is	Iceland	65	196

Pa6.seabone.net	Italy	66	255
core1-pos8-0.telehouse.ukcore.bt.net	London	66	205
core2-6.csc-1.ldn5.psie.net	London	63	190
zcr1-so-1-0-0.LondonInt.cw.net	London	65	185
core1-gig2-0.bletchley.ukcore.bt.net	Milton Keynes	65	177
r2-SeO-1-0.0.ledn-KQ1.NL.kpngwest.net	Netherlands	65	209
ti09a95.ti.telenor.net	Norway	67	257
cisco0.Moscow.ST.NFT	Russia	65	241
bgw-ser5-0-0.Moscow.Rostelecom.ru	Russia	65	249
apv-i1-pos1-0-0-int-stockholm.telia.net	Sweden	66	212
mlm1-core.swip.net	Sweden	65	200
atm0-0-x.lon2gw1.uk.insnet.net	UK	63	188
access-th-3-e0.router.technocom.net	UK	65	180
pos3-0.cr1.lnd5.gbb.uk.uu.net	UK	64	180
North America	Anahaina		112
pos4-1-0-622M.cr1.ANA2.gblx.net	Anaheim	66	113
pop1-dtc.atdn.net	AOL	65	137
popl-rtc.atdn.net	AOL	65	137
atlanta1-br1.bbnplanet.net	Atlanta	66	120
cambridge1-br1.bbnplanet.net	Boston	41	178
core3-hssi5-0.Boston.cw.net	Boston	0	0
pos1-0-0-155M.ar1.BOS1.gblx.net	Boston	65	115
core-fa5-0-0.ontario.canet.ca	Canada	66	123
chi-core-03.inet.qwest.net	Chicago	66	80
FddiO.AR1.CHI1.Alter.Net	Chicago	65	83
c1-pos2-0.chcgil1.home.net	Chicago	66	80
router.mitchell.edu	Connecticut	61	138
dallas1-br2.bbnplanet.net	Dallas	65	105
dllstx1wcx2-oc48.ipcc.wcg.net	Dallas	65	115
denver-br2.bbnplanet.net	Denver	64	106
so-1-0-0-3.mp1.Denver1.level3.net	Denver	0	0
eth1-0-0.michnet1.mich.net	Detroit	53	112
borderx2-fddi-1.LosAngeles.cw.net	Los Angeles Los Angeles	64	119
	MAE West	63	109 0
mae-west.wenet.net ar8.mexmdf.avantel.net.mx	Mexico		154
		65	
core2-mexico.uninet.net.mx	Mexico	67	149
rr1.mexmdf.avantel.net.mx	Mexico	66	152
inet-mex-roma-3-g5-0-0.mex.uninet.net.mx	Mexico	65	147
if-9-0.core2.Montreal.Teleglobe.net	Montreal	65	136
vsnl-c-o-cwc.NewYorknyr.cw.net	New York	65	542
p2-0-0.nyc4-br1.bbnplanet.net	New York	64	107
ge12-0-0.access1.hud-ny.us.xo.net	New York	66	100
sl-gw9-nyc-8-0.sprintlink.net	New York	65	114
TelecomItaliaMumbi1.soNYC2.gblx.net	New York	65	330
ix-10-0-1.bb6.NewYork.Teleglobe.net	New York	66	318

pos2-0-622M.cr2.PHI1.gblx.net Phila pos1-0-0-155M.ar1.PHI1.gblx.net Phila j01-ge-0-1-0-0.phx.opnix.net Pho	adelphia 65 adelphia 66 adelphia 66 enix 63 enix 65 ramento 66	113 112 110 109 114
pos1-0-0-155M.ar1.PHI1.gblx.net Phila j01-ge-0-1-0-0.phx.opnix.net Pho	enix 65	110 109 114
j01-ge-0-1-0-0.phx.opnix.net Pho	enix 63 enix 65	109 114
1 3	enix 65	114
n2 1 phnyaz2 cr2 hhpplanot not Dhor		
pz-1.prirryazz-crz.ouriptariet.net	ramento 66	
border7-fddi-0.Sacramento.cw.net Sacr		113
core1.SanFrancisco.cw.net San	Francisco 66	112
main1-core5-oc12.sjc1.above.net San	Jose 65	175
bbr01-p3-0.sntc04.exodus.net	ta Clara 65	113
border3-fddi-0.Seattle.cw.net	ttle 65	113
198.ATM6-0.XR2.SEA1.ALTER.NET Seat	ttle 66	121
pos4-0.core1-ott.bb.attcanada.ca Torc	onto 63	99
dcr01-g6-0.trnt01.exodus.net	onto 64	95
299.ATM7-0.XR1.VAN1.ALTER.NET Van	couver 65	124
fa-1-1-0.a04.vinnva01.us.ra.verio.net Virgi	inia 63	109
br1-a3120s8.wswdc.ip.att.net Was	shington DC 65	113
wdc-core-02.inet.qwest.net Was	shington DC 65	105
111.at-6-0-0.TR2.DCA6.ALTER.NET Was	shington DC 63	112
so2-1-0-622M.br1.WDC2.gblx.net Was	shington DC 66	110
pos2-0-155M.cr1.WDC2.gblx.net Was	shington DC 67	109
South America		
rcorelma1-rcoreats1.impsat.net.ar Arge	entina 53	321
multicanal-atm1.prima.com.ar Arge	entina 66	267
gsr01.spo.embratel.net.br Braz	zil 66	246
fast5-cr2-net5.attla.cl Chile	e 65	268
telefonica-mundo-chile-no-rev-dns Chile	e 65	218
gip-bogota-1-ethernet0-1.gip.net Colo	ombia 66	223
cha-00-lo0.core.cantv.net Veni	ezuela 65	190

Gary Madden is a Professor in the Department of Economics and Director of the Communication Economics and Electronic Markets Research Centre (CEEM) at Curtin University of Technology in Perth, Australia. His research is primarily focussed on examining empirical aspects of electronic, information and communications markets. His research funding, which supports CEEM, has come from industry, government and the Australian Research Council. Gary is Editor of the Edward Elgar series, the International Handbook of Telecommunications Economics, Volumes I-III. He is a Member of the Board of Management of the Internatinal Telecommunications Society and a Member of the Editorial Board of the Journal of Media Economics.

email: MaddenG@cbs.curtin.edu.au

Grant Coble-Neal was a Research Associate at CEEM, Curtin University from 1999 to 2003. He is currently completing his PhD under Gary Madden's supervision.

email: CobleG@cbs.curtin.edu.au