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Gender Differences in Child Health-care Practices: Evidence from the Pakistan Demographic and Health Survey, 1990-91

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INTRODUCTION

Among other factors, health care utilisation is important in determining the health status and survival chances of children. The patterns of childhood mortality, in general, indicate that deaths of male children have consistently exceeded those of females, with a much greater difference in the first month of birth (NNR). This has largely been attributed to differences in the genetic and biological factors between the sexes [Lopez and Ruzika (1983)]. The mortality level, thereafter, is influenced more by the socio-economic, environmental, and health care factors, indicating a mortality disadvantage for females in some populations. It has therefore been postulated that gender-based differences in health care practices partly explain the sex differentials in child mortality in some countries of South Asia, where health-seeking behaviour of parents discriminates against female children [Chen, *et al.* (1981); Das Gupta (1987); Sathar (1987); Ahmed (1990)].

Using data from Bangladesh, Chen, Haq, and D'Souza (1981) found that girls' mortality risk was nearly 60 percent higher than that for boys after the neonatal period, and that girl children suffered more malnutrition and received lesser treatment for various infections. Das Gupta (1987) and Muhuri and Preston (1991) also explained the excess mortality of girls with a surviving elder sister in terms of conscious, selective neglect of the second daughter. Waldron (1983) in her extensive review of child mortality patterns in developing countries concluded that besides relative contributions of specific causes of death with different impact by sex, the variability in discrimination by gender, primarily in nutrition and health care utilisation, also contributes to excess female child mortality (1–4 age group).

However, a study on Karachi [Alizai and Zia (1990)] found no perceived differentiation in taking a sick child for treatment to a doctor/hospital, and little awareness of gender bias in the attitudes or behaviour of parents and professionals regarding the health care of children. Porapakkam (1983), looking at selected case

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studies, found for Thailand that the incidence rate of morbidity was higher for female than male children, but no gender differentials were apparent in the use and non-use of specific health facilities.

Recognising the fact that child health care practices are the proximate factors in child mortality, the objectives of this paper are to ascertain whether sex differences exist in under-five mortality, as well as to determine whether gender differences are apparent in the health status, morbidity incidence, and preventive and curative treatment of disease among children. The paper also explores some of the potential gender dimensions of female mortality disadvantage during childhood.

DATA AND METHODS

For the purposes of this analysis, we use data from the Pakistan Demographic and Health Survey of 1990-91, which has collected extensive information on fertility, child mortality, and child health indicators. The survey provides a full birth history of each woman with the date of birth and the sex and age at death of each child, making it possible to estimate child mortality and health care indicators by sex. Among the health care indicators, data on factors such as immunisation, morbidity rates for fever, cough, respiratory infections, and diarrhea are available, with information on whether any treatment to the child was given or not during sickness.

The sample selected for the present analysis is the number of all children born during the five years preceding the survey date (all births occurring since January 1986), for which information on each child's survival, health care, and morbidity status is available.

For mortality measures, neonatal, infant, and child mortality rates are used to examine changes in the levels of mortality over time. The proportion of children dead in a particular category is also described to indicate the differentials by sex in relation to selected socio-economic factors. For health care indicators, the measure used is generally the absolute differences between the percentages of female and male children (under five years of age) with a particular characteristic.

PAST EVIDENCE ON SEX DIFFERENTIALS IN INFANT/CHILD MORTALITY

The estimates of child mortality from various data sources indicate that the overall levels of childhood mortality have declined in Pakistan¹ (Table 1). Girls' infant mortality rate has considerably declined from 102 during 1984–88 to 80 per thousand live births during the 1986–91 period. The corresponding rates for boys are from 116 to 101 per thousand live births. The change in infant mortality is largely

¹This is also reflected in terms of improvements in the expectation of life at birth in recent years. It appears that health interventions have increased the rate of decline of female mortality more than that of males.

attributable to improvements in postneonatal mortality, while neonatal mortality disfavours boys, who have nearly 30 percent higher deaths than females in the 1986-91 period. However, child mortality (1–4 years) remains higher among girls than boys over the past years, indicating a female disadvantage in child mortality. This may be suggestive of some gender-related differences in child-rearing and health care practices that favour boys more than girls.

Table 1

Infant and Child Mortality Rates by Gender from Various Surveys

Childhood	1965	5–79 ^a	197	6–79 ^b	198	4–88°	198	6–91 ^d
Mortality	Boys	Girls	Boys	Girls	Boys	Girls	Boys	Girls
Neonatal								
(<1 Month)	79	70	_	_	67	54	59	44
Infant Mortality (0–11 Months)	120	117	102	86	116	102	101	80
Child Mortality (1–4 Years)	44	48	39	46	27	29	22*	37*

Source: *The rates are for the ten-year period preceding the survey, (1981–91).

RESULTS FROM THE PDHS

Sex Differences in Child Mortality

A more detailed examination of the PDHS data shows that out of the total births during the five years preceding the survey (6558), 8.6 percent died, and 91.4 percent survived. Of the total number of deaths (563), the proportion of boys is higher (53 percent) than of girls (47 percent). The distribution of deaths is given in Table 2, which shows that nearly 86 percent of the deaths occurred during infancy, while 14 percent between the ages of 1–4 years. The proportion of deaths is the highest for boys in the first month of life (50.2 percent as compared to 46.6 percent for girls). In the later months (1–11 and 12–48), the proportion is higher for girls than for boys.

Results from research studies and the National Health Survey indicate that the major causes of death during infancy are acute respiratory infections, diarrhea, and malnutrition [Ahmed (1990)]; these also suggest that better facilities in the household and health care practices tend to improve the health status and survival chances of children [Government of Pakistan (1986); Mahmood and Kiani (1994)].

^aPopulation, Labour Force and Migration Survey, as quoted in Sathar (1987).

^bPopulation Growth Survey, as quoted in Afzal (1988).

^cPakistan Demographic Survey, as quoted in Sathar (1994).

^dPakistan Demographic and Health Survey, 1990-91.

Table 2

Percentage Distribution of Child Deaths by Gender and by Age at Death, and Social Status Indicators: PDHS, 1990-91

	Во	ys	G	irls	Total D	eaths
Total	%	(N)	%	(N)	%	(N)
Total Deaths	53.1	(299)	46.9	(264)	100.0	(563)
Age at Death						
<1Month	50.2	(150)	46.6	(123)	48.5	(273)
1–11 Months	36.5	(109)	39.0	(103)	37.7	(212)
12-48 Months	13.4	(40)	14.4	(38)	13.9	(78)
Residence						
Urban	47.2	(141)	50.4	(133)	48.7	(274)
Rural	52.8	(158)	49.6	(131)	51.3	(289)
Women's Age						
<30	69.9	(209)	68.9	(182)	69.4	(391)
30 +	30.1	(90)	31.1	(182)	30.6	(172)
Women's						
Education						
None	81.6	(244)	81.8	(216)	81.7	(460)
Some	18.4	(55)	18.2	(48)	18.3	(103)

High risks of infant and child deaths are generally found among families of low socio-economic status. Higher family income, mother's education, and urban residence usually taken as indicators of better nutrition, and greater awareness and access to health facilities, are likely to reduce child mortality risks [Alam and Cleland (1984); Ahmed (1992)]. Based on the PDHS data, the proportion of child deaths by selected social status indicators, such as urban residence and mother's education, is shown in Table 2. The figures support the negative association of urban residence and education with child mortality. Nearly 18 percent of the deaths are experienced by educated mothers as compared to 82 percent for women with no education.

Gender Differentials in Child Morbidity and Curative Measures

The two diseases most prevalent among children in Pakistan are acute respiratory infections, indicated by cough with fast or difficult breathing, and diarrhea. The PDHS provided information on these specific diseases as cough,

fever, and diarrhea prevalent among children under five years of age in some reference period, typically two weeks before the survey. It should be noted that the data on morbidity and the use of health services are for surviving children only, who are 91.4 percent of the total births selected for this study. Hence, even with the deaths removed, the distribution of survivors would quite reasonably reflect health status differentials.

Table 3 presents the percentages of the girls and boys sick with fever, cough, and diarrhea during the past two weeks preceding the survey date. Overall, it appears that girls are slightly less likely to be sick (as reported sick by their mothers) than boys. Out of the total children sick with fever (1683), and cough (1911), nearly 52 percent are boys and 48 percent are girls. The number of children reported as sick with diarrhea is much lower than the incidence of fever and cough. However, their distribution by gender again shows a slightly higher proportion of boys with diarrheal infections both during the past two weeks and within 24 hours of the survey date. The table shows that nearly 54 percent of the boys are reported sick with diarrhea within the past 24 hours as compared to 45 percent of the girls.

Table 3

Morbidity Prevalence among Children (under five) Two Weeks Preceding the Survey, and the Curative Measures by Gender: PDHS, 1990-91

	Bo	ys	G	irls	Tot	al
Type of Disease	%	(N)	%	(N)	%	(N)
Fever	51.6	(869)	48.4	(814)	100.0	(1683)
Cough	51.6	(986)	48.4	(925)	100.0	(1911)
Diarrhea	52.8	(417)	47.2	(373)	100.0	(790)
Diarrhea						
(within 24 hours)	53.6	(241)	46.4	(209)	100.0	(450)
Treatment Given						
(for fever, cough)	52.1	(868)	47.9	(799)	100.0	(1677)
Treatment Given						
(for diarrhea)	50.1	(240)	49.9	(239)	100.0	(479)
ORS Given						
(for diarrhea)	52.0	(173)	48.0	(160)	100.0	(333)

The pattern of such gender differences in sickness finds support from the findings of another UNICEF study [Ahmed (1990)], which shows higher proportions of male children recorded in various hospitals and health care units as sick with respiratory infections, diarrhea, and other diseases. One possible explanation given for these differences in morbidity is the mothers' greater likelihood of using bottle-feeding for male infants, perhaps because it is perceived as being superior to breastfeeding or as a source of providing their male children some extra and

supplemental form of nutrition.² However, evidence from the analysis of 38 DHS Surveys for 35 developing countries indicates that there is no significant association between sickness differentials and female mortality disadvantage [Hill and Upchurch (1994)].

In terms of the question of differentials in the treatment given to boys and girls during sickness, Table 3 shows a slight advantage to male children in terms of the treatment given for fever, cough, and diarrhea. But the absolute differences in percentages are quite small. For example, 52 percent of the boys were given ORS treatment for diarrhea as compared to 48 percent of the girls. A review of the records of patients brought for treatment in different health care units/hospitals, as undertaken in a UNICEF study, indicates a gender bias in favour of male children (see Appendix Table A). It appears that although parents' attitudes and perceptions about their child health are not gender-biased, their actual behaviour in terms of disease treatment has shown a bias in favour of male children. However, the study [Ahmed (1990)] concludes that there appears to be no strong gender bias in the areas of nutrition and health care when mothers are asked about taking care of the nutritional and health needs of their children.

Use of Preventive Health Care

Another factor that might account for the differences in child mortality is a gender differential in the use of preventive health care practices. The indicators of preventive measures used from the survey data are immunisation coverage of surviving children, child ever-breastfed, and the age at which solid or supplementary food was given to the child every day. Table 4 shows the percentages of girls and boys for these three indicators. It may be noted here that of the total surviving children, nearly 96 percent of the children are ever breastfed and about 58 percent received immunisation. Differences in immunisation coverage between boys and girls are of small magnitude (51 percent against 49 percent), which may be partly due to the extensive coverage given to the expanded programme of immunisation during the 1990s. There is some evidence that in certain places the girls being immunised actually exceeded the boys by a small percentage [Ahmed (1990)], implying effects of the heavy publicity campaign by the government and the strong outreach of the immunisation programme in recent years (see Appendix Table 2).

One area where gender differences are most neutralised is breastfeeding, a strong indicator of child nutrition and prevention from diarrheal sickness. Table 4 shows that almost equal proportions of girls and boys were breastfed. As for the intake of solid food by age of children, we find no difference in the proportions of boys and girls upto 12 months. For later ages of 1–4 years, the proportion of boys

²The PDHS data show that about one-quarter of the breastfed children under one year of age were given a bottle with a nipple the day before the interview [NIPS (1992)].

given solid food regularly is higher (53 percent) than that of girls (47 percent). This difference may partially support the contention that boys get a preferential treatment in allocation of food relative to girls, an argument supporting the association between girls' nourishment and female mortality disadvantage during childhood (1–4 years).

Table 4

Use of Health-care Practices for Children (under five) by Gender:
PDHS, 1990-91

Health Care	Во	ys	G	irls	Tot	al
Factors	%	(N)	%	(N)	%	(N)
Child Immunised	51.2	(1381)	48.8	(1318)	100.0	(2699)
Child Ever-						
breastfed	50.5	(2894)	49.5	(2837)	100.0	(5731)
Age at which Solid/						
Mushy Food						
Given Daily						
< 6 Months	49.5	(799)	50.5	(815)	100.0	(1614)
6–12 Months	50.0	(1187)	50.0	(1187)	100.0	(2374)
13 +	53.0	(415)	47.0	(368)	100.0	783

The survey also collected information on the weight and height of children, making it possible to calculate anthropometric indices of nutritional status. Such indicators, in particular height for age, provide a good indication of the overall health status of children that is independent of mothers' perception. The PDHS report shows that nearly 30 percent to 50 percent of the children under five years of age are stunted, indicating chronic malnutrition (measured by height-for-age standard), and about 40 percent are classified as underweight (by weight-for-age standard). However, gender differences in these indices appear to be negligible and thus are not investigated in this paper. It may be noted here that the evidence for 35 DHS country reports indicates no significant association between stunting and wasting differentials by sex and female mortality disadvantage [Hill and Upchurch (1994)].

CONCLUDING OBSERVATIONS

An examination of past trends in infant and child mortality indicates that the overall levels of mortality have declined in Pakistan, even though the exact placement of the infant mortality rate is still questioned when viewed from different data sources.

The sex differentials in infant/child mortality still persist with more male deaths in the neonatal and infancy period, an age when genetic factors are more

important in determining mortality risks, and a female disadvantage in child mortality, suggesting the importance of gender factors in girl mortality.

A simple comparison of the proportion dead by sex from the PDHS data suggests that boys experience the highest number of deaths in the first month of life. Greater risks of child death are experienced by younger, uneducated, and rural mothers. An association of these factors is found in the expected direction.

The data available to investigate the general hypothesis that sex differentials in childhood mortality reflect the discriminatory behaviour of parents against girls is limited. The simple and a preliminary analysis undertaken does not reflect any profound differences in child sickness and preventive and health care practice (at least as reported by mothers) in the PDHS data. Gender differences in health-related indicators are only weak, if at all related to gender differences in child mortality.

Among the health-related variables examined in this paper, some gender differences in diarrhea treatment and a slight difference in immunisation coverage are noted, both indicators significantly related to male-female mortality differentials (as is evident from Hill and Upchurch's analysis of 35 DHS reports of developing countries). This suggests that immunisation programmes are likely to be effective in reducing mortality in general, and that widespread outreach of immunisation and its publicity may be effective in reducing excess female mortality. However, the health-related variables we have used in this paper do not shed any light on the cultural and societal processes that contribute to the persistence of gender mortality differentials.

Nevertheless, gender differences in mortality appear to be associated clearly with one indicator which reflects the social status of women in society, i.e., the proportion of educated women. Our analysis does recommend increasing education for women and enhancing maternal and child health care facilities for younger mothers. In addition, the programmes aimed at reducing the leading causes of infant and child mortality such as acute respiratory diseases and diarrheal infections need to be improved. To achieve this end, educational campaigns for the promotion of breastfeeding and immunisation coverage appear to be important in the overall context of policy development.

This study has been limited to a small sample of the survey-children born during the five years preceding the survey-because health-related information is available only for these children. We feel that in order to capture the effects of sociocultural factors, which account for gender differences in health factors and their relation to child mortality differentials, much better and more specific data are required. The type of data we have used in this paper give us the relative differences in health care factors between boys and girls, but not the specific cultural values or attitudes that contribute to the existence of gender differences in the health and nutritional intake of children.

Appendix A

Reception Desk Data Islamabad Children's Hos

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(ears	H	1535	1560	1505	2218	1736	2355	2067	2150	1850	1383	1294	1312	1633	1565	1257	1288	1829	1717	1458	1265	1168	
5+7	×	2377	2146	1266	2632	2121	2460	2478	2967	2576	1969	1837	1905	2357	2129	1923	1840	2344	2291	1918	1588	1243	
ears	ഥ	1333	1351	1268	1765	1543	1707	1244	1222	1424	1274	1175	1025	884	913	1043	1085	1879	1649	1258	1785	1412	00000
1-5 Y	¥	1366	1765	1861	2193	1926	1919	1413	1855	2382	1860	1562	1353	1384	1492	1579	1437	2195	2113	1490	1553	1465	26163
1 Year	L	774	623	662	1170	986	1063	986	911	929	821	822	623	554	620	999	807	1242	1032	915	942	786	17034
1 Month -	M	1184	696	1036	1367	1127	1147	1114	1260	1473	1323	1178	1126	1074	887	1064	1158	1975	1578	1247	1101	506	26202
onth	ц	57	21	35	4	34	20	39	4	27	· 82	224	210	224	269	202	50	43	80	8	181	170	2157
0-1 M	Z	52	33	51	99	28	71	52	94	72	166	317	347	405	447	356	121	401	94	83	233	222	3414
	Month	1/89	5/89	3/89	4/89	2/89	68/9	68//	68/8	68/6	10/89	11/89	12/89	1/90	2/90	3/90	4/90	2/90	06/9	2/90	06/8	06/6	
	0-1 Month 1 Month - 1 Year 1-5 Years	th Month 1 Month - 1 Year 1 F Month - 1 Year 5 + Years 5 + Years	O-1 Month 1 Month - 1 Year 1 Lears 5 + Years In Month - 1 Year M F M F M 52 57 1184 774 1366 1333 2377	th M F Month – 1 Year 1 Fears M F was M 5 + Years 52 57 1184 774 1366 1333 2377 33 21 969 623 1765 1351 2146	th 0-1 Month 1 Month 1 Year 1 Month 1 Year 1 F M M F M M F M M S + Years 5 + Years 5 + Years 52 57 1184 774 1366 1333 2377 33 21 969 623 1765 1351 2146 51 35 1036 662 1861 1268 1266	th M F M F M F + Years 52 57 1184 774 1366 1333 2377 33 21 969 623 1765 1351 2146 51 35 1036 662 1861 1268 1266 66 44 1367 1170 2193 1765 2632	h 0-1 Month – I Year M I Month – I Year F I –5 Years 5 + Years 52 57 1184 774 1366 1333 2377 33 21 969 623 1765 1351 2146 51 35 1036 662 1861 1268 1266 66 44 1367 1170 2193 1765 2632 28 34 1127 986 1926 1543 2121	th 0-1 Month Month F I Month I Year I F M F M 5 + Years 52 57 1184 774 1366 1333 2377 33 21 969 623 1765 1351 2146 51 35 1036 662 1861 1268 1266 66 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2460 2 54 44 1160 911 1855 1222 2478 2 72 27 1473 929 2382 1424 2576 11 166 85 1323 821 1860 1274 1969 11 347 210 1178 822 1562 1444 <

Appendix B
E.P.I. Data
Islamabad Children's Hospital

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	0-1 Month	nth	1 Month -	h – 1 Year		-	-5 Years		5 + Years	ears	
Month	×	<u>H</u>	Σ	II.		M		F	M	山	
1/89	33	29	245	270		09		89	0	166	
2/89	33	16	249	275		72			0	150	
3/89	30	28	228	269		91		108	0	155	
4/89	41	78	228	260		78		11	0	180	
68/5	22	22	228	256	•	69		69		6	
68/9	5 6	61	247	267		84		. 75	0	110	
68/L	21	14	237	259		78		79	0	6	
68/8	21	23	219	212		73		70	0	52	
68/6	29	56	226	237		120		81	4	75	
10/89	35	25	157	172	•.	2		54	·, 0	51	
11/89	24	23	237	263	÷	69		72	0	9	
12/89	32	29	309	307		79		9/	2	95	
1/90	22	21	320	318		107		88	0	82	
2/90	21	17	295	267		92		. 62	-	62	
3/90	23	17	368	332		95		98	0	89	
4/90	.35	53	274	247		80		83	0	34	•.
2/90	25	24	319	319		96		68	0	72	
06/9	29	21	219	219		98		72	0	73	
06/L	28	20	226	226		69		09	0	43	
06/8	24	22	269	276		89		65	0	45	
06/6	32	33	229	304		85		92	0	21	
	576	489	5329	5555		1715	1	614	∞	1827	

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