

A Strong Invariance Principle for Associated Random Fields

R.M. Balan^{*†}

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Abstract

In this paper we generalize Yu's strong invariance principle for associated sequences to the multi-parameter case, under the assumption that the covariance coefficient $u(n)$ decays exponentially as $n \rightarrow \infty$. The main tools will be the Berkes-Morrow multi-parameter blocking technique, the Csörgő-Révész quantile transform method and the Bulinski rate of convergence in the CLT for associated random fields.

Keywords: strong invariance principle; associated random fields; blocking technique; quantile transform.

1 Introduction

Amongst various concepts introduced to measure the dependence between random variables, association deserves a special place because of its numerous applications and its relatively easy mathematical manipulation. A finite collection (X_1, \dots, X_m) of random variables is said to be **associated** (or satisfies the **FKG inequalities**) if for any coordinatewise non-decreasing functions f, g on \mathbf{R}^m , $\text{cov}(f(X_1, \dots, X_m), g(X_1, \dots, X_m)) \geq 0$, whenever the covariance is defined. An infinite collection of random variables is associated if every finite sub-collection is associated. This concept was introduced formally in [15], where one can also find some of its most important properties.

In the past few decades, a lot of effort has been dedicated to prove limit theorems for random fields $(X_j)_{j \in \mathbf{Z}_+^d}$ of associated random variables. In the case $d = 1$, this culminated with the strong invariance principle of Yu (see

^{*}*Postal address:* Department of Mathematics and Statistics, University of Ottawa, 585 King Edward Avenue, Ottawa, ON, K1N 6N5, Canada. *E-mail address:* rbala348@science.uottawa.ca

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[22]), from which one can easily deduce all the other major limit theorems, like the weak invariance principle and the functional law of iterated logarithm (FLIL). The present paper was motivated by the need for a similar result in the case $d \geq 2$, which arises in the context of higher dimensional models, like the percolation model of [11].

The first asymptotic result for zero-mean associated random fields was the central limit theorem (CLT) proved by Newman in [17] for the (strongly) stationary case. This result says that if the *finite susceptibility* assumption holds, i.e. $\sigma^2 := \sum_{i \in \mathbf{Z}^d} \rho(i) < \infty$, where $\rho(j-k) := \text{cov}(X_j, X_k)$, then

$$n^{-d/2} S_n \xrightarrow{d} N(0, \sigma^2) \quad (1)$$

where $S_n := \sum_{j_1 \leq n} \cdots \sum_{j_d \leq n} X_j$. This was generalized in [11] to the non-stationary case, under the assumption that $u(n) \rightarrow 0$ as $n \rightarrow \infty$, where

$$u(n) := \sup_{j \in \mathbf{Z}_+^d} \sum_{k: \|j-k\| \geq n} \text{cov}(X_j, X_k) \quad (2)$$

and $\|i\| := \max_{s=1, \dots, d} |i_s|$.

The weak invariance principle for (strongly) stationary associated random fields was proved by Newman and Wright in the case $d = 1$ and $d = 2$ (see [18], [19]), under the same finite susceptibility assumption. In [19], it was conjectured that the same principle holds for $d > 2$. A partial solution to this problem was given in [10] (in the stationary case) and in [16] (in the non-stationary case), under the *finite r -susceptibility* assumption:

$$E|S_N|^{2+r} \leq C[N]^{1+r/2}$$

where $S_N := \sum_{j \leq N} X_j$ and $[N] := \prod_{s=1}^d N_s$ for $N = (N_1, \dots, N_d) \in \mathbf{Z}_+^d$. (If $i, j \in \mathbf{Z}_+^d$, we use the notations $i \leq j$ if $i_s \leq j_s, \forall s = 1, \dots, d$ and $i < j$ if $i_s < j_s, \forall s = 1, \dots, d$.)

The conjecture was fully solved in [8], where it is proved that for a zero-mean (weakly) stationary associated random field $(X_j)_{j \in \mathbf{Z}_+^d}$ with uniformly bounded moments of order $s > 2$ and a power decay rate for the covariance coefficient $u(n)$, we have

$$W_n(\cdot) \xrightarrow{d} W(\cdot) \text{ in } D([0, 1]^d) \quad (3)$$

where $W_n(t) := n^{-d/2} \sum_{j_1 \leq nt_1} \cdots \sum_{j_d \leq nt_d} X_j$ and $W = (W(t))_{t \in [0, 1]^d}$ is a d -parameter Wiener process with variance σ^2 . (We note in passing that for $d = 1$, generalizations to the non-stationary case and to case of vector-valued random variables are given in [3], respectively [9].)

The FLIL for associated sequences was obtained in [13], under the finite r -susceptibility assumption with $r = 1$ and a condition which requires that $E(S_n^2)/n$ converges to 1 with a power decay rate.

The strong invariance principle proved by Yu in 1996, strengthened and unified all of these results in the case $d = 1$ and implied other asymptotic fluctuation results, like the Chung's type of FLIL for the maxima of partial sums (see Theorems A-E of [20]). More precisely, Yu showed that if $(X_j)_{j \in \mathbf{Z}_+}$ is a sequence of associated random variables such that the moments of order $s > 2$ are uniformly bounded, the variances are bounded below away from 0 and the covariance coefficient $u(n)$ decays exponentially as $n \rightarrow \infty$, then it is possible to redefine the original sequence on a richer probability space together with a standard Wiener process $W = (W(t))_{t \in [0, \infty)}$ such that, for some $\epsilon > 0$

$$S_n - W(\sigma_n^2) = O(n^{1/2-\epsilon}) \text{ a.s.}$$

where $\sigma_n^2 := E(S_n^2)$. As far as we know, there are no generalizations of this principle to the case $d \geq 2$. The purpose of the present paper is to fill this gap and to provide a powerful approximation tool that can be used in higher dimensions.

Unlike the case $d = 1$, the strong invariance principle for associated random fields in higher dimensions holds only for points $N \in \mathbf{Z}_+^d$ which are not “too close” to the coordinate axes. This is not at all surprising and a similar fact happens for mixing random fields (see [1]). The reason for this phenomenon is the irregular behavior of $E(S_N^2)$ close to the coordinate planes.

We proceed now to introduce the notations that will be used throughout this paper.

Let $(X_j)_{j \in \mathbf{Z}_+^d}$ be a weakly stationary associated random field with zero mean and $\rho(j - k) := E(X_j X_k), \forall j, k \in \mathbf{Z}_+^d$. Let $u(n)$ be the covariance coefficient defined by (2). Because of stationarity, we have $u(n) = \sum_{i \in \mathbf{Z}^d: \|i\| \geq n} \rho(i)$ for every $n \geq 0$. We will suppose that $\rho(0) > 0$ and $\sigma^2 := u(0) = \sum_{i \in \mathbf{Z}^d} \rho(i) < \infty$.

For any finite subset $V \subseteq \mathbf{Z}_+^d$, we let $|V|$ be the cardinality of V , $S(V) := \sum_{j \in V} X_j$, $\sigma^2(V) := E[S^2(V)]$ and $F_V(x) := P(S(V)/\sigma(V) \leq x)$, $x \in \mathbf{R}$. Note that for any finite subset $V \subseteq \mathbf{Z}_+^d$

$$r(0) \leq \frac{\sigma^2(V)}{|V|} \leq \sigma^2 \tag{4}$$

Most of the time we will work with “rectangles” $V \subseteq \mathbf{Z}_+^d$ of the form $V := (a, b] = \prod_{s=1}^d (a_s, b_s]$ with $a_s, b_s \in \mathbf{Z}_+ \cup \{0\}, a_s \leq b_s$; note that $|V| = [b - a]$. We denote with \mathcal{A} the class of all the subsets V of this form.

We will use the following conditions:

(C1) $\sup_{j \in \mathbf{Z}_+^d} E|X_j|^{2+r+\delta} < \infty$ for some $r, \delta > 0$

(C2) $u(n) = O(e^{-\lambda n})$ for some $\lambda > 0$

(C2') $u(n) = O(n^{-\nu})$ for some $\nu \geq 0$

We recall that a d -parameter Wiener process $W = \{W_t; t \in [0, \infty)^d\}$ with variance σ^2 is a Gaussian process with independent increments such that $W(R)$ has a $N(0, \sigma^2|R|)$ -distribution for any rectangle R ($|R|$ denotes the volume of R). Following [1], we put $G_\tau := \cap_{s=1}^d \{j \in \mathbf{Z}_+^d : j_s \geq \prod_{s' \neq s} j_{s'}^\tau\}$ for any $\tau \in (0, 1)$.

Here is the main result of this paper.

Theorem 1.1 *Let $d \geq 2, \tau \in (0, 1)$ and $(X_j)_{j \in \mathbf{Z}_+^d}$ be a weakly stationary associated random field with zero mean and $\rho(j-k) := E(X_j X_k)$ for any $j, k \in \mathbf{Z}_+^d$. Suppose that $\rho(0) > 0$ and $\sigma^2 := \sum_{i \in \mathbf{Z}_+^d} \rho(i) < \infty$.*

If (C1) and (C2) hold, then without changing its distribution we can re-define the random field $(X_j)_{j \in \mathbf{Z}_+^d}$ on a richer probability space together with a d -parameter Wiener process $\{W_t; t \in [0, \infty)^d\}$ with variance σ^2 such that

$$S_N - W_N = O([N]^{1/2-\epsilon}) \quad \text{a.s.}$$

for $N \in G_\tau$. Here ϵ is a positive constant depending on the field $(X_j)_{j \in \mathbf{Z}_+^d}$.

From the previous theorem one can easily deduce the following CLT:

$$[N]^{-1/2} S_N \xrightarrow{d} N(0, \sigma^2)$$

when $[N] \rightarrow \infty$ and $N \in G_\tau$ for some $\tau \in (0, 1)$; this is more general than (1) which was obtained only for $N = (n, \dots, n) \in \mathbf{Z}_+^d$. The non-functional version of LIL obtained in [21] for any multi-parameter process with independent increments (in particular for the Wiener process) allows us to conclude that

$$\limsup_{[N] \rightarrow \infty, N \in G_\tau} (2[N] \log \log [N])^{-1/2} S_N = \sigma \quad \text{a.s.}$$

We proceed now to the proof of Theorem 1.1. This is divided into several steps which are explained in Section 2. The remaining sections contain the developments that are needed to perform each step. To ease the exposition, we placed in the Appendix the proofs of some preliminary lemmas.

2 Description of the Method

In this section we will indicate what are the main ingredients that are needed for the proof of Theorem 1.1. More precisely, by blending the multi-parameter blocking technique of Berkes and Morrow with the quantile transform technique of Csörgő and Révész, we will be able to generalize to the multi-parameter case the method introduced by Yu in [22].

Let $\alpha > \beta > 1$ be integers to be chosen later and $n_0 := 0$. For $l \in \mathbf{Z}_+$ let

$$n_l := \sum_{i=1}^l (i^\alpha + i^\beta) \sim \frac{1}{\alpha+1} l^{\alpha+1}.$$

For each $k := (k_1, \dots, k_d) \in \mathbf{Z}_+^d$, we put $N_k := (n_{k_1}, \dots, n_{k_d})$. For all $k \in \mathbf{Z}_+^d$ we have $[N_k] \sim (\alpha + 1)^{-d} [k]^{\alpha+1}$.

Let $B_k := (N_{k-1}, N_k] = \prod_{s=1}^d (n_{k_s-1}, n_{k_s}]$. Note that $|B_k| = \prod_{s=1}^d (k_s^\alpha + k_s^\beta) \leq 2^d [k]^\alpha$. We define the ‘‘big’’ blocks H_k and the ‘‘small’’ blocks I_k by

$$H_k := \prod_{s=1}^d (n_{k_s-1}, n_{k_s-1} + k_s^\alpha], \quad I_k := B_k \setminus H_k.$$

Note that $|H_k| = [k]^\alpha$ and $(2^d - 1)[k]^\beta \leq |I_k| \leq (2^d - 1)[k]^\alpha$. We denote $u_k := S(H_k)$, $\lambda_k^2 := \sigma^2(H_k)$ and $v_k := S(I_k)$, $\tau_k^2 := \sigma^2(I_k)$. By (4)

$$C[k]^\alpha \leq \lambda_k^2 \leq C[k]^\alpha, \quad C[k]^\beta \leq \tau_k^2 \leq C[k]^\alpha. \quad (5)$$

The sums over the big blocks will be used to generate a Gaussian approximating sequence $(\eta_k)_k$ which will in turn be approximated by a Wiener process. In order to do this, we will need an upper bound for the covariance of the sums over two big blocks in terms of the distance between these blocks. The small blocks are introduced simply to give some space between the big blocks, i.e. to ensure that the distance between any two big blocks is non-zero.

H_{13}	H_{23}	H_{33}
H_{12}	H_{22}	H_{31}
H_{11}	H_{21}	H_{31}

If the distribution function \tilde{F}_k of u_k/λ_k is continuous, then one could use directly the quantile transform method of Csörgő and Révész [12], to approximate the variable u_k/λ_k by a $N(0, 1)$ -random variable. In general, this assumption may not be satisfied, and therefore one needs to employ a ‘‘smoothing’’ technique (see [22]). Without changing its distribution, we redefine the random field $(u_k)_{k \in \mathbf{Z}_+^d}$ on a rich enough probability space together with a random field $(w_k)_{k \in \mathbf{Z}_+^d}$ of independent random variables such that w_k is $N(0, \tau_k^2)$ -distributed and $(u_k)_k$ and $(w_k)_k$ are independent. Let

$$\xi_k := (u_k + w_k)/(\lambda_k^2 + \tau_k^2)^{1/2}, \quad k \in \mathbf{Z}_+^d$$

and F_k be the distribution function of ξ_k . By the CLT for associated random fields, $\tilde{F}_k(x) \rightarrow \Phi(x)$ as $k \rightarrow \infty$ and consequently $F_k(x) \rightarrow \Phi(x)$ as $k \rightarrow \infty$,

where $\Phi(x)$ denotes the $N(0, 1)$ distribution function. Therefore it is reasonable to consider the following $N(0, 1)$ -random variable

$$\eta_k := \Phi^{-1}(F_k(\xi_k))$$

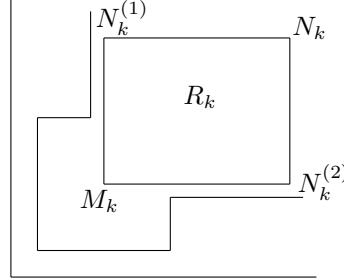
as an approximation for ξ_k . Let $e_k := \sqrt{\lambda_k^2 + \tau_k^2} (\xi_k - \eta_k)$.

In what follows we will adapt the method introduced by Berkes and Morrow for mixing random fields to suit the special needs of an associated random field.

Following [1] (p. 25), we let $\tau \in (0, 1)$ be arbitrary, $\rho := \tau/8$, L be the set of all indices i corresponding to the “good” blocks $B_i \subseteq G_\rho$, and H be the set of all points in \mathbf{Z}_+^d which fall in one of the good blocks. To each point $N \in H$ we associate the points $N^{(1)}, \dots, N^{(d)}$ which can be thought as the intersections of the hyperplanes $n_s = N_s, s = 1, \dots, d$ with the “boundary” of the domain H ; their precise definition is: $N_{s'}^{(s)} = N_{s'}, \forall s' \neq s$ and

$$N_s^{(s)} := \min_{n \in H; n_{s'} = N_{s'}, s' \neq s} n_s.$$

Unlike the authors of [1], we raise a small technical point by noting that H may not be a nice “L-shaped” region. This is why we consider the rectangles $R_k := (M_k, N_k] \subseteq H$, where $M_k := ((N_k^{(1)})_1, \dots, (N_k^{(d)})_d)$. We note that $L_k := \{i : B_i \subseteq R_k\} \subseteq L \cap \{i \leq k\}$.



If V is a rectangle in \mathbf{Z}_+^d and \tilde{V} is the rectangle in \mathbf{R}_+^d which corresponds to V , then we make an abuse of notation by writing $W(V)$ instead of $W(\tilde{V})$. This convention will be used throughout this work and will occasionally apply to finite unions of rectangles as well. We write

$$\begin{aligned} S_N &= (S_N - S_{N_k}) + S(R_k) + S((0, N_k] \setminus R_k) \\ W_N &= (W_N - W_{N_k}) + W(R_k) + W((0, N_k] \setminus R_k) \end{aligned}$$

and we use the following decomposition of $S(R_k)$, based on the definitions of ξ_i and e_i and the fact that $S(B_i) = u_i + v_i$:

$$S(R_k) = \sum_{i \in L_k} e_i + \sum_{i \in L_k} \sqrt{|B_i|} \left(\sqrt{\frac{\lambda_i^2 + \tau_i^2}{|B_i|}} - \sigma \right) \eta_i + \sum_{i \in L_k} \sigma \sqrt{|B_i|} \eta_i - \sum_{i \in L_k} w_i + \sum_{i \in L_k} v_i. \quad (6)$$

In Section 3, we will show that all the sums in the above decomposition, except the third one, can be made sufficiently small. The third sum will be treated separately in Section 4 and will be approximated by $W(R_k) = \sum_{i \in L_k} W(B_i)$, via a very powerful approximation result (Theorem 5 of [2]) and a carefully chosen procedure for counting the indices in L . Finally, in Section 5 we will show that the terms $S((0, N_k] \setminus R_k)$, $W((0, N_k] \setminus R_k)$ can be made sufficiently small if $N_k \in G_\tau$, and the differences $S_N - S_{N_k}$, $W_N - W_{N_k}$ are small if $N \in G_\rho$. This will conclude the proof of Theorem 1.1.

3 The “good” blocks

In this section we will show that all the sums in the decomposition (6) of S_{R_k} , except the third one, can be made sufficiently small.

In order to treat the first sum of this decomposition, we need to evaluate the precision of the approximation of ξ_k by η_k . This will be given by the rate of convergence in the CLT. In this paper we decided to use the rate obtained by Bulinski in [7], under the assumption that the covariance coefficient $u(n)$ decays exponentially as $n \rightarrow \infty$; under this assumption, this is the sharpest rate of convergence in the CLT (see [4]). We note in passing that in the case $d = 1$, a different rate of convergence in the CLT was developed and used in [22] for associated sequences with a power decay rate of the covariance coefficient; however, the exponential decay rate of $u(n)$ was eventually needed in [22] for the strong invariance principle. The problem of whether or not the strong invariance principle continues to hold for associated random fields with a power decay rate of covariances is still open even in the case $d = 1$, and we do not attempt to tackle it here.

Throughout our work we will use the letter C to denote a generic positive constant, independent of k .

Lemma 3.1 (Theorems 1, 2 of [7]) *Suppose that (C1) and (C2) hold and let $s := 2 + r + \delta$. Then for any finite subset $V \subseteq \mathbf{Z}_+^d$*

$$\sup_{x \in \mathbf{R}} |F_V(x) - \Phi(x)| \leq \begin{cases} C|V| \cdot (\sigma^2(V))^{-s/2} \cdot (\log(|V| + 1))^{d(s-1)} & \text{if } s \leq 3 \\ C|V| \cdot (\sigma^2(V))^{-3/2} \cdot (\log(|V| + 1))^d & \text{if } s > 3 \end{cases}$$

The next result is a generalization of Lemma 3.2 of [22] to the case $d \geq 2$, in the case of an exponential decay rate of $u(n)$. Its proof is routine and is given in the appendix.

Lemma 3.2 *If (C1) and (C2) hold and $2r_0r/(2+r) < \alpha/\beta < 2(1+r)/(2+r)$ with $r_0 := \max\{1, (r + \delta)^{-1}\}$, then*

$$\sup_{x \in \mathbf{R}} |F_k(x) - \Phi(x)| \leq C[k]^{-r\beta/(2+r)} \quad \text{and} \quad \sup_{x \in \mathbf{R}} |f_k(x) - f(x)| \leq C$$

where $f_k(x)$ is the density function of ξ_k and $f(x)$ is the $N(0, 1)$ density function.

Using Lemma 3.2 and an argument that was introduced by Csörgő and Révész (in the proof of Lemma 3 of [12]), we get the precision of the approximation of ξ_k by η_k .

Lemma 3.3 *Under (C1) and (C2), we have*

$$|\Phi^{-1}(F_k(x)) - x| \leq C[k]^{-\{r\beta/(2+r) - K^2/2\}}$$

provided that $|x| \leq K\sqrt{\log[k]}$, where $0 < K < \sqrt{2r\beta/(2+r)}$.

Next we give the precision of the approximation of ξ_k by η_k in terms of the L^2 -distance. For this we will need the following lemma which gives an upper bound for the moments of order $2+r$, generalizing an older result of Birkel in the case $d = 1$ (see [5]). In particular, this lemma shows that $(X_j)_{j \in \mathbf{Z}_+^d}$ has finite r -susceptibility (as defined in the introduction).

Lemma 3.4 (Corrolary 1 of [6]) *Suppose that (C1) and (C2') hold with $\nu \geq d\nu_0$, where $\nu_0 := r(2+r+\delta)/(2\delta) < (d-2)^{-1}$ if $d \geq 3$. Then for any $V \in \mathcal{A}$*

$$E|S(V)|^{2+r} \leq C|V|^{1+r/2}.$$

Using (5), Lemma 3.3 and Lemma 3.4, and employing the same technique that was used in the proof of Lemma 3.10 of [22], we get the following result.

Lemma 3.5 *Under (C1) and (C2), we have*

$$E[e_k^2] \leq C[k]^{\alpha - \epsilon_0}, \quad \forall k \in \mathbf{Z}_+^d$$

where $\epsilon_0 := 2r^2\beta/\{(2+r)(4+3r)\}$.

The next result will show us that the first sum in the decomposition (6) of $S(R_k)$ is small.

Lemma 3.6 *Suppose that (C1) and (C2) hold and $\beta > (1+2/r)(3+4/r)$. Then there exists $\epsilon_1 > 0$ such that for every $k \in \mathbf{Z}_+^d$ with $L_k \neq \emptyset$*

$$\sum_{i \in L_k} |e_i| \leq C[N_k]^{1/2 - \epsilon_1} \quad a.s.$$

Proof: Let $q > 0$ be such that $\alpha - \epsilon_0 + 1 < 2q < \alpha - 1$ (this is possible since $\epsilon_0 > 2$ by our choice of β). By the Chebyshev's inequality and Lemma 3.5, we have

$$P(|e_i| \geq [i]^q) \leq [i]^{-\{2q - (\alpha - \epsilon_0)\}}, \quad \forall i \in \mathbf{Z}_+^d.$$

By the Borel-Cantelli lemma, it follows that $|e_i| \leq C[i]^q, \forall i \in \mathbf{Z}_+^d$ a.s. and hence $\sum_{i \in L_k} |e_i| \leq C \sum_{i \in L_k} [i]^q \leq C[k]^{q+1} \leq C[k]^{(\alpha+1)/2 - \epsilon_1} \leq C[N_k]^{1/2 - \epsilon_1}$ a.s., where $0 < \epsilon_1 < (\alpha - 1)/2 - q$ and $\epsilon_1 := \epsilon_1/(\alpha + 1)$. \square

The proof of the following lemma is given in the appendix.

Lemma 3.7 *If (C2') holds with $d < \nu < 2d$ then*

$$\sigma^2 - \frac{\sigma^2(V)}{|V|} = O(|V|^{-\delta_0}) \quad (7)$$

where V is a finite union of rectangles in \mathcal{A} and $\delta_0 := \nu/d - 1$.

Remark: Relationship (7) is exactly Dabrowski's condition for the law of the iterated logarithm for associated sequences (see [13]).

Lemma 3.8 *Suppose that (C2') hold with $d < nu < 2d$ and $\beta > 3/\delta_0$, where $\delta_0 := \nu/d - 1$. Then for every $k \in \mathbf{Z}_+^d$ with $L_k \neq \emptyset$*

$$\sum_{i \in L_k} \sqrt{|B_i|} \left(\sigma - \sqrt{\frac{\lambda_i^2 + \tau_i^2}{|B_i|}} \right) |\eta_i| \leq C[N_k]^{1/2-\alpha_0} \text{ a.s.}$$

where $\alpha_0 := 1/\{2(\alpha + 1)\}$.

Proof: Note that $a_i := \sigma - \sqrt{(\lambda_i^2 + \tau_i^2)/|B_i|} > 0$, by (4) and the association property. Using (7), we have

$$\begin{aligned} a_i^2 &\leq \sigma^2 - \frac{\lambda_i^2 + \tau_i^2}{|B_i|} = \frac{|H_i|}{|B_i|} \left(\sigma^2 - \frac{\lambda_i^2}{|H_i|} \right) + \frac{|I_i|}{|B_i|} \left(\sigma^2 - \frac{\tau_i^2}{|I_i|} \right) \\ &\leq C(|H_i|^{-\delta_0} + |I_i|^{-\delta_0}) \leq C[i]^{-\beta\delta_0} \end{aligned}$$

and hence, by the Chebyshev's inequality

$$P(\sqrt{|B_i|} a_i \eta_i \geq [i]^{\alpha/2-1}) \leq [i]^{-(\alpha-2)} |B_i| a_i^2 \leq C[i]^{-(\beta\delta_0-2)}.$$

By the Borel-Cantelli lemma, it follows that $\sqrt{|B_i|} a_i \eta_i \leq C[i]^{\alpha/2-1}, \forall i \in \mathbf{Z}_+^d$ a.s. and hence $\sum_{i \in L_k} \sqrt{|B_i|} a_i \eta_i \leq C \sum_{i \in L_k} [i]^{\alpha/2-1} \leq C[k]^{\alpha/2} \leq C[N_k]^{1/2-\alpha_0}$ a.s. since $[k] \sim (\alpha + 1)^{d/(\alpha+1)} [N_k]^{1/(\alpha+1)}$. \square

The final result of this section shows that the last two sums in the decomposition (6) of $S(R_k)$ are small.

Lemma 3.9 *If $\alpha - \beta > 2 + 4/\rho$, then for every $k \in \mathbf{Z}_+^d$ with $L_k \neq \emptyset$ we have*

$$\sum_{i \in L_k} |v_i| \leq C[N_k]^{1/2-\alpha_0} \text{ a.s. and } \sum_{i \in L_k} |w_i| \leq C[N_k]^{1/2-\alpha_0} \text{ a.s.}$$

Proof: For the first inequality, we follow the lines of the proof of Lemma 8 of [1]. Note that $I_i = \cup_{s=1}^d I_i(s)$, where $I_i(s)$ are disjoint rectangles with $|I_i(s)| \leq C i_s^\beta \prod_{s' \neq s} i_{s'}^\alpha$. Hence $v_i = \sum_{s=1}^d v_i(s)$ with $v_i(s) := \sum_{j \in I_i(s)} X_j$.

By the Chebyshev's inequality and (4)

$$\begin{aligned} P(|v_i(s)| \geq [i]^{\alpha/2-1}) &\leq C[i]^{-(\alpha-2)} |I_i(s)| \leq C i_s^{-(\alpha-\beta-2)} \prod_{s' \neq s} i_{s'}^2 \\ &\leq i_s^{-(\alpha-\beta-2-2/\rho)} \leq C[i]^{-(\alpha-\beta-2-2/\rho)\rho/2} \end{aligned}$$

for every $i \in L_k$. (As in the proof of the above-mentioned lemma, we used the fact that $i \in L_k$ implies that $i_s \geq C \prod_{s' \neq s} i_{s'}^\rho$ and consequently $i_s \geq C[i]^{\rho/2}$.) Since $(\alpha - \beta - 2 - 2/\rho)\rho/2 > 1$, the result follows by the Borel-Cantelli lemma. A similar argument applies to w_i , since $E(w_i^2) = \tau_i^2 \leq C|I_i| = C \sum_{s=1}^d |I_i(s)|$. \square

4 The approximation theorem

In this section we will verify that the third sum in the decomposition (6) of $S(R_k)$ can be approximated by $W(R_k)$, where W is a d -parameter Wiener process with variance σ^2 . Some preliminary lemmas are needed.

The next result follows exactly as Theorem 2.1 of [22], using Lemma 3.2.

Lemma 4.1 *If (C1) and (C2) hold and $2r_0r/(2+r) < \alpha/\beta < 2(1+r)/(2+r)$ with $r_0 := \max\{1, (r+\delta)^{-1}\}$, then for any $0 < \theta < 1/2$ and all $i \neq j$*

$$E(\eta_i \eta_j) \leq C\{([i][j])^{-\alpha/2} E(u_i u_j)\}^{\theta/(1+\theta)}.$$

The next lemma gives a generalization of relationship (3.11) of [22] to the multi-parameter case.

Lemma 4.2 *If (C2) holds, then*

$$E(u_i u_j) \leq C e^{-\lambda M_{i,j}^\beta}$$

where $M_{i,j} := \max_{s: i_s \neq j_s} (M_s(i, j) - 1)$ and $M_s(i, j) := \max(i_s, j_s)$, $s = 1, \dots, d$.

Proof: Let $d := \min_{k \in H_i} d(k, H_j)$ be the distance between H_i and H_j , where $d(k, H_j) := \min_{k' \in H_j} \|k - k'\|$. Then $d_k := d(k, H_j) - d \geq 0 \forall k \in H_i$,

$$E(u_i u_j) = \sum_{k \in H_i} \sum_{k' \in H_j} E(X_k X_{k'}) \leq \sum_{k \in H_i} u(d + d_k) \leq C e^{-\lambda d} \sum_{k \in H_i} e^{-\lambda d_k} \leq C e^{-\lambda d}$$

and $d = \max_{s=1, \dots, d} \min_{k \in H_i, k' \in H_j} |k_s - k'_s| = \max_{s: i_s \neq j_s} \{m_s^\beta + \sum_{l=m_s+1}^{M_s-1} (l^\alpha + l^\beta)\} \geq M_{i,j}^\beta$, where $m_s = m_s(i, j) := \min(i_s, j_s)$ and $M_s = M_s(i, j)$. \square

In order to prove our approximation theorem, we need to be able to ‘‘count’’ properly the indices in L , i.e. to define a bijection $\psi : \mathbf{Z}_+ \rightarrow L$ satisfying certain properties. This will be given by the following lemma, whose proof can be found in the appendix.

Lemma 4.3 *There exists a bijection $\psi : \mathbf{Z}_+ \rightarrow L$ such that*

$$l < m \Rightarrow \exists s^* = s^*(l, m) \text{ such that } \psi(l)_{s^*} \leq \psi(m)_{s^*} \quad (8)$$

$$\exists m_0 \in \mathbf{Z}_+ \text{ such that } m \leq C[\psi(m)]^{\gamma_0} \quad \forall m \geq m_0 \quad (9)$$

for any $\gamma_0 > (1 + 1/\rho)(1 - 1/d)$.

We are now able to prove the desired approximation theorem.

Theorem 4.4 *Suppose that (C1) and (C2) hold, $\alpha > 3(1 + 1/\rho)(1 - 1/d)$, $\beta > (2/\rho)(1 + 1/\rho)(1 - 1/d)$ and $2r_0r/(2 + r) < \alpha/\beta < 2(1 + r)/(2 + r)$ with $r_0 := \max\{1, (r + \delta)^{-1}\}$. Then without changing its distribution we can redefine the random field $(X_j)_{j \in \mathbf{Z}_+^d}$ on a rich enough probability space together with a d -parameter Wiener process $W = (W_t; t \in [0, \infty)^d)$ with variance σ^2 , such that for every $k \in \mathbf{Z}_+^d$ with $L_k \neq \emptyset$*

$$\sum_{i \in L_k} \sigma \sqrt{|B_i|} \left| \eta_i - \frac{W(B_i)}{\sigma \sqrt{|B_i|}} \right| \leq C[N_k]^{1/2 - \alpha_0} \text{ a.s.}$$

where $\alpha_0 := 1/\{2(1 + \alpha)\}$.

Proof: Let $0 < \theta < 1/2$ be such that $\alpha\{(1 + 1/\rho)(1 - 1/d)\}^{-1} > 1 + 1/\theta$ and choose γ_0 such that $(1 + 1/\rho)(1 - 1/d) < \gamma_0 < \min\{\alpha\theta/(1 + \theta), \beta\rho/2\}$. Let $\psi : \mathbf{Z}_+ \rightarrow L$ be the bijection given by Lemma 4.3.

We will apply Theorem 5 of [2] to the sequence $Y_m := \eta_{\psi(m)}$, $m \in \mathbf{Z}_+$ of random variables and the probability distributions $G_m := N(0, 1)$, $m \in \mathbf{Z}_+$ and we will prove that for each $m \in \mathbf{Z}_+$, $m \geq 2$ there exists some $\rho_m > 0$ such that

$$\left| E \exp \left\{ i \sum_{l=1}^m t_l Y_l \right\} - E \exp \left\{ i \sum_{l=1}^{m-1} t_l Y_l \right\} E \exp \{ i t_m Y_m \} \right| \leq \rho_m \quad (10)$$

for all $t_1, \dots, t_m \in \mathbf{R}$ with $\sum_{l=1}^m t_l^2 \leq U_m^2$, where $U_m > 10^{32}$.

Then, by the above-mentioned theorem, without changing its distribution we can redefine the sequence $(Y_m)_{m \in \mathbf{Z}_+}$ on a rich enough probability space together with a sequence $(Z_m)_{m \in \mathbf{Z}_+}$ of independent $N(0, 1)$ -random variables such that

$$P(|Y_m - Z_m| \geq \alpha_m) \leq \alpha_m, \quad \forall m \in \mathbf{Z}_+$$

where $\alpha_m \leq C\{U_m^{-1/4} \log U_m + \exp(-3U_m^{1/2}/16)m^{1/2}U_m^{1/4} + \rho_m^{1/2}U_m^{m+1/4}\}$. We will prove next that

$$\alpha_m \leq Cm^{-2} \text{ for } m \text{ large} \quad (11)$$

Then, by the Borel-Cantelli Lemma, $|Y_m - Z_m| \leq C\alpha_m$, $\forall m \in \mathbf{Z}_+$ a.s. Using a straightforward d -parameter generalization of Lemma 4 of [12], without

changing its distribution we can redefine the sequence $(Z_m)_{m \in \mathbf{Z}_+}$ on a richer probability space together with a d -parameter Wiener process with variance σ^2 such that $Z_m = W(B_{\psi(m)})/(\sigma\sqrt{|B_{\psi(m)}|})$, $\forall m \in \mathbf{Z}_+$. Hence

$$\left| \eta_i - \frac{W(B_i)}{\sigma\sqrt{|B_i|}} \right| \leq C\alpha_{\psi^{-1}(i)} \quad \forall i \in L \quad \text{a.s.}$$

and because $|B_i| \leq |B_k| \leq C[k]^\alpha$, $\forall i \in L_k$ and $\sum_{l \in \mathbf{Z}_+} \alpha_l < \infty$, we have

$$\sum_{i \in L_k} \sigma\sqrt{|B_i|} \left| \eta_i - \frac{W(B_i)}{\sigma\sqrt{|B_i|}} \right| \leq C[k]^{\alpha/2} \sum_{i \in L_k} \alpha_{\psi^{-1}(i)} \leq C[k]^{\alpha/2} \leq C[N_k]^{1/2-\alpha_0}.$$

We proceed next to the verification of (10) and (11). By Lemma 4.1 and Lemma 4.2, we have

$$\begin{aligned} E(Y_l Y_m) &\leq C \left\{ ([\psi(l)][\psi(m)])^{-\alpha/2} E(u_{\psi(l)} u_{\psi(m)}) \right\}^{\theta/(1+\theta)} \\ &\leq C ([\psi(l)][\psi(m)])^{-\alpha\theta/(2+2\theta)} e^{-\lambda\theta M_{\psi(l), \psi(m)}^\beta / (1+\theta)} \\ &\leq C ([\psi(l)][\psi(m)])^{-\alpha\theta/(2+2\theta)} e^{-\lambda\theta [\psi(m)]^{\beta\rho/2} / (1+\theta)} \end{aligned}$$

(For the last inequality above we used (8) to obtain an $s^* = s^*(l, m)$ for which $M_{s^*}(\psi(l), \psi(m)) = \psi(m)_{s^*}$; since $\psi(m) \in L$, we have $M_{\psi(l), \psi(m)} \geq \psi(m)_{s^*} - 1 \geq C[\psi(m)]^{\rho/2}$.) By Lemma 2.2 of [14], the left-hand side of (10) is smaller than $2 \sum_{l=1}^{m-1} |t_l t_m| E(Y_l Y_m)$, which is in turn smaller than

$$\begin{aligned} &C e^{-\lambda\theta [\psi(m)]^{\beta\rho/2} / (1+\theta)} \sum_{l=1}^{m-1} 2|t_l t_m| ([\psi(l)][\psi(m)])^{-\alpha\theta/(2+2\theta)} \\ &\leq C e^{-\lambda\theta [\psi(m)]^{\beta\rho/2} / (1+\theta)} \left\{ \sum_{l=1}^{m-1} t_l^2 [\psi(l)]^{-\alpha\theta/(1+\theta)} + (m-1)t_m^2 [\psi(m)]^{-\alpha\theta/(1+\theta)} \right\} \\ &\leq C e^{-\lambda\theta [\psi(m)]^{\beta\rho/2} / (1+\theta)} \sum_{l=1}^m t_l^2 \leq C e^{-\lambda\theta [\psi(m)]^{\beta\rho/2} / (1+\theta)} U_m^2 := \rho_m \end{aligned}$$

for m large enough (In the second inequality above, we used the fact that $m \leq C[\psi(m)]^{\alpha\theta/(1+\theta)}$, which follows from Lemma 4.3 by our choice of γ_0 .)

Finally, relationship (11) follows if we take $U_m := m^q$ with $q > 8$. Clearly $U_m^{-1/4} \log U_m \leq m^{-2}$ and $\exp(-3U_m^{1/2}/16)m^{1/2}U_m^{1/4} \leq \exp(-2U_m^{1/2}/16) \leq m^{-2}$ for m large enough. We have

$$\rho_m^{1/2} U_m^{m+1/4} = e^{-\lambda\theta [\psi(m)]^{\beta\rho/2} / (2+2\theta)} m^{q(m+5/4)} \leq m^{-2}$$

since $\{2 + q(m + 5/4)\} \log m \leq C m^{1+\epsilon} \leq C [\psi(m)]^{(1+\epsilon)\gamma_0} \leq C [\psi(m)]^{\beta\rho/2}$, for m large enough. This concludes the proof of the theorem. \square

Remark: A similar argument can be used to give a simplified proof for Theorem 2.5 of [22] (in the case $d = 1$). More precisely, one can check directly the condition of Theorem 5 of [2] for the sequence $(\eta_k)_{k \geq 1}$ of random variables and the probability distributions $G_k = N(0, 1), k \geq 1$ (as we did above). We obtain in this manner a sequence $(Z_k)_{k \geq 1}$ of independent $N(0, 1)$ -random variables with $P(|\eta_k - Z_k| \geq \alpha_k) \leq \alpha_k$ and $\alpha_k \leq Ck^{-2}$. Without changing its distribution we can redefine the sequence $(Z_k)_{k \geq 1}$ on a richer probability space together with a standard Brownian motion $W = \{W_t; t \in [0, \infty)\}$ such that $Z_k = W(\hat{H}_k)/\sqrt{\lambda_k^2 + \tau_k^2}$, where $\hat{H}_k := (V_{k-1}, V_k]$ and $V_k := \sum_{i=1}^k (\lambda_i^2 + \tau_i^2)$. Since $\lambda_i^2 + \tau_i^2 \leq Ci^\alpha \leq Ck^\alpha$ for $i \leq k$ and $\sum_{i \geq 1} \alpha_i < \infty$, this gives immediately the desired approximation

$$\sum_{i=1}^k \sqrt{\lambda_i^2 + \tau_i^2} \left| \eta_i - \frac{W(\hat{H}_i)}{\sqrt{\lambda_i^2 + \tau_i^2}} \right| \leq Ck^{\alpha/2} \sum_{i=1}^k \alpha_i \leq CN_k^{1/2-\alpha} \quad \text{a.s.}$$

5 The remaining terms

In this section we show that the terms $S((0, N_k] \setminus R_k), W((0, N_k] \setminus R_k), S_N - S_{N_k}, W_N - W_{N_k}$ can be made sufficiently small if $N \in G_\tau$.

Note that $(0, N_k] \setminus R_k = \cup_{s=1}^d (0, N_k^{(s)})$. If we let $D_s(N) := \max_{n \leq N^{(s)}} |S_n|$ and $\hat{D}_s(N) := \max_{n \leq N^{(s)}} |W_n|$, for each $s = 1, \dots, d$ and $N \in H$, then

$$S((0, N_k] \setminus R_k) \leq \sum_{s=1}^d 2^{d-s} D_s(N_k), \quad W((0, N_k] \setminus R_k) \leq \sum_{s=1}^d 2^{d-s} \hat{D}_s(N_k).$$

On the other hand $(0, N] \setminus (0, N_k] = \cup_J I_k^{(J)}$, where $I_k^{(J)} := \prod_{s \in J} (n_{k_s}, N_s] \times \prod_{s \in J^c} (0, n_{k_s}]$ and the union is taken over all non-empty subsets J of $\{1, \dots, d\}$. Let $M_k^{(J)} := \max |S(I_k^{(J)})|$ and $\hat{M}_k^{(J)} := \sup |W(I_k^{(J)})|$, where the maximum and the supremum are taken over all N with $n_{k_s} < N_s \leq n_{k_{s+1}}, \forall s \in J$. We have

$$\max_{N_k < N \leq N_{k+1}} |S_N - S_{N_k}| \leq \sum_J M_k^{(J)}, \quad \sup_{N_k < N \leq N_{k+1}} |W_N - W_{N_k}| \leq \sum_J \hat{M}_k^{(J)}.$$

We note in passing that the arguments that are valid for the terms depending on the original random field $(X_j)_{j \in \mathbf{Z}_+^d}$ can be applied to the terms depending on the Wiener process W , since $W(V) = \sum_{j \in V} \hat{X}_j, \forall V \in \mathcal{A}$, where $\hat{X}_j := W((j-1, j])$ are independent $N(0, \sigma^2)$ -random variables. Clearly $(\hat{X}_j)_{j \in \mathbf{Z}_+^d}$ is a weakly stationary associated random field with zero mean and covariance coefficient $\hat{u}(n) = 0, \forall n \geq 1$.

Lemma 5.1 (a) *Suppose that (C1) and (C2') hold with $\nu \geq d\nu_0$ and $\nu_0 := r(2+r+\delta)/(2\delta) < (d-2)^{-1}$ if $d \geq 3$. Then there exists x_0 such that $\forall V \in \mathcal{A}, \forall x \geq x_0$*

$$P(M(V) \geq x|V|^{1/2}) \leq Cx^{-(2+r)}$$

where $M(V) := \max\{|S(Q)|; Q \subseteq V, Q \in \mathcal{A}\}$.

(b) *If (C1) and (C2) hold, then there exists $\gamma > 0$ such that $\forall V \in \mathcal{A}$*

$$P(\tilde{M}(V) \geq |V|^{1/2}(\log |V|)^{d+1}) \leq C|V|^{-\gamma}$$

where $\tilde{M}((a, b]) := \max\{|S(Q)|; Q = (a, c], a < c \leq b\}$.

Proof: **(a)** Using Lemma 1 of [8], the Markov inequality and Lemma 3.4, we have $P(M(V) \geq x|V|^{1/2}) \leq 2P(|S(V)| \geq x|V|^{1/2}/2) \leq Cx^{-(2+r)}|V|^{-(1+r/2)}E|S(V)|^{2+r} \leq Cx^{-(2+r)}$.

(b) This follows exactly as the second inequality of Lemma 7 of [1], using the moment inequality given by Lemma 3.4 and the rate of convergence in the CLT given by Lemma 3.1. This rate is sharper than the rate of Lemma 5 of [1]. To see this, we use (4) and we note that $\sup_{x \in \mathbf{R}} |F_V(x) - \Phi(x)|$ is either smaller than $C|V|^{-\{s/2-1-\epsilon d(s-1)\}}$ if $s \leq 3$, or smaller than $C|V|^{-(1/2-\epsilon d)}$ if $s > 3$; in both cases a suitable choice of $\epsilon > 0$ gives us the rate $C|V|^{-t}$ for some $t \in (0, 1)$. We also note that the requirement $|V| \in G_\tau$ is not needed. \square

The next result follows exactly as Lemma 6 of [1], using Lemma 5.1,(a).

Lemma 5.2 *If $\alpha > 16/(3\tau) - 1$, then*

$$\max_{s=1, \dots, d} D_s(N_k) \leq C[N_k]^{1/2-\epsilon} \text{ a.s.}, \quad \max_{s=1, \dots, d} \hat{D}_s(N_k) \leq C[N_k]^{1/2-\epsilon} \text{ a.s.}$$

for every $N_k \in G_\tau$ and $0 < \epsilon < \tau/32$.

The following result follows exactly as Lemma 9 of [1], using Lemma 5.1,(b).

Lemma 5.3 *Let γ be the constant given by Lemma 5.1,(b). If $\alpha > 2/\gamma$, then*

$$\max_j M_k^{(j)} \leq C[N_k]^{1/2-\epsilon} \text{ a.s.}, \quad \max_j \hat{M}_k^{(j)} \leq C[N_k]^{1/2-\epsilon} \text{ a.s.}$$

for every $N_k \in G_\rho$ and $0 < \epsilon < \rho/(8\alpha)$.

A Appendix

Proof of Lemma 3.2: Using Lemma 3.1 for $V = H_k$ and relationship (5), we obtain that $\sup_{x \in \mathbf{R}} |\tilde{F}_k(x) - \Phi(x)|$ is either smaller than $C[k]^{-\{\alpha s/2 - \alpha - \epsilon d(s-1)\}}$ if $s \leq 3$, or smaller than $C[k]^{-(\alpha/2 - \epsilon d)}$ if $s > 3$. If $\alpha/\beta \geq 2r_0r/(2+r)$, then a suitable choice of $\epsilon > 0$ allows us to conclude that $|\tilde{F}_k(x) - \Phi(x)| \leq C[k]^{-r\beta/(2+r)}, \forall x \in \mathbf{R}$. The first inequality follows by a change of variables.

For the second inequality we use a technique similar to that used to prove relationship (3.3) of [22]. Let $\varphi_k(t) := E[\exp(it\xi_k)]$, $\tilde{\varphi}_k(t) := E[\exp(itu_k/\lambda_k)]$ and $\varphi(t) = \exp(-t^2/2)$. Since $(\lambda_k^2 + \tau_k^2)/\lambda_k^2 \leq C$, we have for any $T > 0$

$$\begin{aligned} |f_k(x) - f(x)| &\leq \frac{1}{2\pi} \int_{-\infty}^{\infty} |\varphi_k(t) - \varphi(t)| dt \leq \frac{C}{2\pi} \int_{-\infty}^{\infty} |\tilde{\varphi}_k(t) - \varphi(t)| e^{-\frac{\tau_k^2 t^2}{2\lambda_k^2}} ds \\ &\leq \frac{C}{2\pi} \cdot 2T[k]^{-r\beta/(2+r)} + \frac{C}{\pi} \int_{|t| \geq T} \exp\left\{-\frac{t^2 \tau_k^2}{2\lambda_k^2}\right\} dt \\ &\leq C \cdot T[k]^{-r\beta/(2+r)} + \frac{C}{T} \cdot \frac{\lambda_k^2}{\tau_k^2} \exp\left\{-\frac{\tau_k^2 T^2}{2\lambda_k^2}\right\} \end{aligned}$$

Since $\lambda_k^2/\tau_k^2 \leq C[k]^{\alpha-\beta}$, the conclusion follows by choosing $T = C[k]^q$ with $\alpha - \beta < q < r\beta/(2+r)$. Such a choice is possible if $\alpha/\beta < 2(1+r)/(2+r)$. \square

Proof of Lemma 3.7: First we claim that it is enough to prove (7) for “squares”, i.e. for rectangles $V = (m, n] \in \mathcal{A}$ for which $n_s - m_s = l, \forall s = 1, \dots, d$. To see this we note that each rectangle V can be written as a finite union of disjoint squares: $V = \cup_{i=1}^p V_i$. By the association property $\sigma^2(V) \geq \sum_{i=1}^p \sigma^2(V_i)$ and

$$\sigma^2 - \frac{\sigma^2(V)}{|V|} \leq \frac{1}{|V|} \sum_{i=1}^p |V_i| \left(\sigma^2 - \frac{\sigma^2(V_i)}{|V_i|} \right) \leq \frac{1}{|V|} \sum_{i=1}^p C|V_i|^{1-\delta_0} \leq C|V|^{-\delta_0}$$

because $0 < \delta_0 < 1$. Let us now prove relationship (7) for a square $V = (m, n]$ with $n_s - m_s = l, \forall s = 1, \dots, d$. Note that $|V| = l^d$. By stationarity

$$\begin{aligned} \sigma^2(V) &= |V| \cdot r(0) + \sum_{\substack{-(n-m-1) \leq i \leq n-m-1, \\ i \neq 0}} \prod_{s=1}^d (l - |i_s|) \cdot r(i) \\ &= |V| \cdot \sum_{\|i\| \leq l-1} r(i) - \sum_{\emptyset \neq K \subseteq \{1, \dots, d\}} (-1)^{|K|-1} \sum_{\|i\| \leq l-1, i \neq 0} c(K, i) \cdot r(i) \end{aligned}$$

where $c(K, i) := l^{|K^c|} \cdot \prod_{s \in K} |i_s|$. Since $\sigma^2 - \sum_{\|i\| \leq l-1} r(i) = \sum_{\|i\| \geq l} r(i) = u(l)$ and $c(K, i) \leq |V|$ if $\|i\| \leq l-1$, we have

$$\begin{aligned} \sigma^2 - \frac{\sigma^2(V)}{|V|} &\leq u(l) + \sum_{\emptyset \neq K \subseteq \{1, \dots, d\}, |K| \text{ odd}} \frac{1}{|V|} \sum_{\|i\| \leq l-1, i \neq 0} c(K, i) \cdot r(i) \\ &\leq C|V|^{-\nu/d} + \sum_{\emptyset \neq K \subseteq \{1, \dots, d\}, |K| \text{ odd}} \sum_{\|i\| \leq l-1, i \neq 0} r(i) \\ &\leq C|V|^{-\nu/d} + C|V|^{-\nu/d+1} \end{aligned}$$

We used the fact that $u(l) \leq Cl^{-\nu} = C|V|^{-\nu/d}$ and $r(i) \leq u(\|i\|) \leq u(\lceil |i| \rceil) \leq C\lceil |i| \rceil^{-\nu/d}$ for any $i \in \mathbf{Z}^d$, where $\lceil |i| \rceil = \prod_{s: i_s \neq 0} |i_s|$. \square

Proof of Lemma 4.3: The idea of the proof is based on the following simple observation in the case $d = 2$. For each $m \in \mathbf{Z}_+, m \geq 2$ with $(m, m) \in L$, there exists a $k_1^*(m) \geq m$ such that $(k_1, m), (m, k_1) \in L$ for every $m \leq k_1 \leq k_1^*(m)$. Therefore, to each vertex $(m, m) \in L$ one can associate an “L-shaped” region $L(m)$ consisting of $2\{k_1^*(m) - m\} + 1$ points in L . In view of the desired property (8), we will count consecutively the indices in $L(2), L(3)$, etc. To verify property (9), we note that $k \in L(m)$ implies $[k] \geq m^2$.

We begin now the proof for arbitrary $d \geq 2$. Let $m \in \mathbf{Z}_+, m \geq 2$ be such that $(m, \dots, m) \in L$ and $k = (k_1, \dots, k_{d-1}, m) \in L$ be such that $k_s > m, \forall s < d$. This implies that all the vertices of B_k are in G_ρ , and in particular $n_m \geq n_{k_s}^\rho, \forall s < d$. Since m is fixed, this cannot happen for infinitely many k_s 's. It follows that for each $s = 1, \dots, d-1$, there exists a $k_s^*(m) \geq m$ such that $k_s \leq k_s^*(m)$. We note that $k_s^*(m) \leq Cm^{1/\rho}$, if m is large enough. This argument shows us that we have a maximum number of $k^*(m) := \prod_{s=1}^{d-1} \{k_s^*(m) - m\}$ points of the form (k_1, \dots, k_{d-1}, m) in L , with $k_s > m, \forall s < d$.

By symmetry, we can repeat this argument for each of the axes. We let $L_s(m) := \{k = (k_1, \dots, k_{s-1}, m, k_s, \dots, k_{d-1}); m < k_{s'} \leq k_{s'}^*(m), \forall s' < d\}$ for every $s = 1, \dots, d$. The “L-shaped” region corresponding to the index m is

$$L(m) := \cup_{s=1}^d L_s(m) \cup \{(m, \dots, m)\}.$$

Note that $|L(m)| = dk^*(m) + 1$ and that $k \in L(m)$ implies $[k] \geq m^d$. Clearly $L \subseteq \cup_m L(m)$ (note that in the case $d = 2$, we actually have $L = \cup_m L(m)$). Next we count consecutively the indices in $L(2), L(3)$, etc., i.e. we define a bijection $\varphi : \mathbf{Z}_+ \rightarrow \cup_m L(m)$ such that $\forall z \in \mathbf{Z}_+$

$$\sum_{l=2}^{m-1} |L(l)| < z \leq \sum_{l=2}^m |L(l)| \Rightarrow \varphi(z) \in L(m).$$

The bijection φ clearly satisfies condition (8). To verify (9), we note that

$$\begin{aligned} z &\leq d \sum_{l=2}^m \prod_{s=1}^{d-1} (k_s^*(l) - l) + m \leq d \prod_{s=1}^{d-1} \sum_{l=2}^m (k_s^*(l) - l) + m \leq d \prod_{s=1}^{d-1} \left(\sum_{l=2}^m k_s^*(l) \right) - \\ &d \left(\sum_{l=2}^m l \right)^{d-1} + m \leq d \prod_{s=1}^{d-1} \left(\sum_{l=2}^m k_s^*(l) \right) \leq Cm^{(1+1/\rho)(d-1)} \leq Cm^{d\gamma_0} \leq C[\varphi(z)]^{\gamma_0} \end{aligned}$$

for m large enough and $\gamma_0 > (1 + 1/\rho)(1 - 1/d)$ arbitrary. Finally, define the bijection $\psi : \mathbf{Z}_+ \rightarrow L$ such that $\psi^{-1}(k) \leq \varphi^{-1}(k), \forall k \in L$. The result follows since if $z_1, z_2 \in \mathbf{Z}_+$ are such that $\psi(z_1) = \varphi(z_2)$, then $z_1 \leq z_2$. \square

References

- [1] Berkes, I. and Morrow, G. J. (1981). Strong invariance principles for mixing random fields. *Z. Wahrsch. verw. Gebiete* **57**, 15-37.
- [2] Berkes, I. and Philipp W. (1979). Approximation theorems for independent and weakly dependent random vectors. *Ann. Probab.* **7**, 29-54.
- [3] Birkel, T. (1988). The invariance principle for associated sequences. *Stoch. Proc. Appl.* **27**, 57-71.
- [4] Birkel, T. (1988). On the convergence rate in the central limit theorems for associated processes. *Ann. Probab.* **16**, 1685-1698.
- [5] Birkel, T. (1988). Moment bounds for associated sequences. *Ann. Probab.* **16**, 1184-1193.
- [6] Bulinski, A. V. (1993). Inequalities for the moments of sums of associated multi-indexed variables. *Theory Probab. Appl.* **38**, 342-349.
- [7] Bulinski, A. V. (1995). Rate of convergence in the central limit theorem for fields of associated random variables. *Theory Probab. Appl.* **40**, 136-144.
- [8] Bulinski, A. V. and Keane, M. S. (1996). Invariance principle for associated random fields. *J. Math. Sciences* **81** (1996), 2905-2911.
- [9] Burton, R. M., Dabrowski, A. R. and Dehling, H. (1986). An invariance principle for weakly associated random vectors. *Stoc. Proc. Appl.* **23**, 301-306.
- [10] Burton, R. M. and Kim, T.-S. (1988). An invariance principle for associated random fields. *Pacific J. Math.* **132**, 11-19.
- [11] Cox, T. J and Grimmett, G. (1984). Central limit theorems for associated random variables and the percolation model. *Ann. Probab.* **12**, 514-528.
- [12] Csörgő, M. and Révész, P. (1975). A new method to prove Strassen type laws of invariance principle I. *Z. Wahrsch. Verw. Gebiete* **31**, 255-260.
- [13] Dabrowski, A. R. (1985). A functional law of the iterated logarithm for associated sequences. *Stat. Probab. Letters* **3**, 209-212.
- [14] Dabrowski, A. R. and Dehling, H. (1988). A Berry-Essen theorem and a functional law of the iterated logarithm for weakly associated random variables. *Stoc. Proc. Appl.* **30**, 277-289.
- [15] Esary, J. D., Proscahn, F. and Walkup, D. W. (1967). Association of random variables, with applications. *Ann. Math. Stat.* **38**, 1466-1474.

- [16] Kim, T.-S. (1996). The invariance principle for associated random fields. *Rocky Mount. J. Math.* **26**, 1443-1454.
- [17] Newman, C. M. (1980). Normal fluctuations and the FKG inequalities. *Commun. Math. Phys.* **74**, 119-128.
- [18] Newman, C. M. and Wright, A. L. (1981). An invariance principle for certain dependent sequences. *Ann. Probab.* **9**, 671-675.
- [19] Newman, C. M. and Wright, A. L. (1982). Associated random variables and martingale inequalities. *Z. Wahrsch. verw. Gebiete* **59**, 361-371.
- [20] Philipp, W. and Stout, W. F. (1975). Almost sure invariance principles for partial sums of weakly dependent random variables. *Mem. Amer. Math. Soc.* **161**.
- [21] Wichura, M. J. (1973). Some Strassen-type laws of the iterated logarithm for multiparameter stochastic processes with independent increments. *Ann. Probab.* **1**, 272-296.
- [22] Yu, H. (1996). A strong invariance principles for associated random variables. *Ann. Probab.* **24**, 2079-2097.