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#### SURVEY AND REVIEW OF FORECASTING MODELS IN INTERNAL GOVERNMENT REVENUES

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The views expressed in this study are those of the author and do not necessarily reflect those of the Institute.

SUMMARY DESCRIPTION OF PROJECT:

The need for good tax revenue forecasts cannot be overemphasized. It is a major input in budget planning and programming as well as a necessary guidepost against which to assess the tax collection effort of the government.

The purpose of this paper is to review and evaluate the existing works on tax revenue forecasting in the Philippines. This endeavor represents the initial step in an attempt to develop a tax revenue forecasting model for the Bureau of Internal Revenue. The review is intended to provide insights to the particular problems involved in the formulation and estimation of a revenue forecasting model and thus, set the stage for the development of such a model.

Earlier studies made on revenue forecasting are as follows: (1) the Kintanar-Mijares work; (2) the Jurado-Encarnacion government sector sub-model; (3) the Diokno public sector model; and (4) the various Bureau of Internal Revenue models.

The Kintanar-Mijares work suggests a procedure for forecasting corporate and individual income tax at a fairly high level of disaggregation. However, it has limiting assumptions arising from its use of a particular year sample data. Its forecasts for the other kind of taxes are aggregative in nature and are based on a simple time-trend. Likewise, the Jurado and Encarnacion is rather aggregative using only six types of taxes. Nevertheless, this study was among the first to relate different tax groupings to different explanatory variables.

The Diokno study considered three types of taxes and used one variable, GNP, to explain the variations in said taxes. In this sense, the work can be said to be very limited for BIR purposes.

Two of the BIR's forecasting approaches, the compound growth rate technique and the time-trend analysis, assume that tax collections are influenced only by time. Furthermore, the compound growth rate technique utilized only bench-mark figures of the data base. The Compertz curve time-trend analysis, on the other hand, has the tendency to give conservative forecasts in the long-run. The tax elasticity approach — the third used by the BIR-has the disadvantage of relating tax receipts to only one explanatory variable. Its advantage over the time-trend analysis, however, lies in the fact that the explanatory variable nost appropriate to the type of tax is used. In this last approach, four types of taxes are considered.

The above models leave much to be desired in terms of the level of disaggregation and the use of explanatory variables. For future modelling works, there appear five basic research directions, viz., the spatial or regional; categorical or into particular taxes; technical or methodological; behavioral, and the macroeconomic view, which will all be explained in the conclusion.

TECHNICAL REPORT: (see attached copies)

## PROBLEMS ENCOUNTERED AND RECOMMENDATIONS

As typical in many situations, there is lack of documentation on forecasting done by the BIR prior to recent years. With the turnover of people, it becomes difficult to know, much more to validate, the forecasting methods employed in the past. This suggests the full documentation of models that may be developed in the future - to include the data base. This should eliminate re-gathering of data already established and make the model implementable by Bureau personnel.

Submitted by:

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Date

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## SURVEY AND REVIEW OF FINDLASTILING FUNCTION ON INTERNAL COVERAMENT NEVERTHES

#### 1. INTRODUCTION

A lot of well-intentioned and feasible government projects have been stalled even in preliminary processes, not entirely because of the failure of the government to meet the cost of these programs per se, but because of its inability to forecast accurately the revenues that it can generate. Revenue forecasting then becomes a key factor in economic, social and political planning. For the Bureau of Internal Revenue, it is a tool to assess its tax collection efforts.

The purpose of this paper is to review and evaluate the existing works on tax reverses forecasting in the Phillippines. This endowor represents the initial step in an attempt to develop a tax reverses forecasting model for the Bureau of Internal Revenue. The review is intended to provide insights to the particular problems involved in the formulation and estimation of a reverse forecasting model and thus, set the stage for the development of such a model.

In sections 4, 5, 6 and 7, we review and synthesize the works that have been conducted in this area, namely: (1) the various Bureau of Internal Bovurue (BIN) methods and whels for forecascing; (2) the Kintanar-Mijares work; (3) the Junado-Encarnación government sector sub-model; and (4) the Diokno public sector model. Each section will include a summary of the functional specifications together with the respective definition of the variablesused in each model. The different functions specified by each of models reviewed were re-estimated using data for the period 1 the sudder It should be noted that all  $\bigvee_{\Lambda}$  studied (except those 1961-1979. of the BIR's) were originally estimated using the data the the sixties. It is to be expected that parameter values derived from these exercises will not hold for the more recent years which are of greater interest to the Bureau, in so far as revenue forecasting is concerned. "Historical" simulation over the new estimation period as was conducted to help assess the forecating capability of the various Confficient of ditermination, R<sup>2</sup>, and the The root mean square per cent error, model reviewed. und RMSE%, were in evaluating the said models. These statistic are discussed more fully in Section 2.

Section 3 provides a brief description of the data used in the analysis.

#### 2. METHODOLOGY

"In the case of the single equation regression model, there exists a set of statistic test (R<sup>2</sup>, F-test, ±-test, etc.) that be used to judge the sign in a statistical sence of the model and its individual estimated coefficients... The model's evaluation must also depend on the purpose for which the model was built.

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A model designed to test a specific hypothesis or to measuresome elasticity should have high t statistics<sup>#</sup>. A model designed for forecasting purposes should have as small a stan-<u>a medit disigned to test a specific fugation or to measure stand</u> dard error of forecast as possible while thus, the R<sup>2</sup> and the RMSE<sup>3</sup> were used to assess the different tax revenue models reviewed. Agreater weight is given to RMSE<sup>3</sup> since the present study is more concerned with the forecasting abilities of the said models.

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The coefficient of determination,  $R_{j}$  defined as the ratio of the regression sum of squares  $M_{1}RSS = (\hat{Y}_{1} - \bar{Y})^{2} J$  to the total sum squares  $M_{1}TSS = (Y_{1} - Y)^{2} J$ 

1 Pindych and Rubinfeld, p. 315

measures the proportion of the variation in the dependent variable which is "explained" by the regression equation or by the variations in the independent variables. The coefficinet of determination ranges from zero to one. A computed R<sup>2</sup> close to unity is indicative to fit. implies that the variations in the endogenous variable can be largely explained by the variations in its determinants.

are st. Stim , men Kin , men Kin variables le of The RMSE's evaluates the "fit" of the individual variables in a simulation context. First, a historical simulation is performed using the model. Then, the resulting or simulated figures of the endogenous variables are examined on how closely each tracks its corresponding historical data series. Finally, the RMSE% is computed,  $\mathcal{AUM}$ defined as follows

 $\frac{Y_t^s - Y_t^a}{v^a}$ Thus, the RMSE's the deviation of the simulated variables from its actual path, in percentage terms. As a rule of 10 per cent plassthumb, RMSE% value equivalent to or less than 10% is acceptable, i.e., an equation with such an RMSE value, for forecasting.

RMSE%

I The NBER set 10 % RMSE % as The standard for evaluating the porecasting capability of an equation.

"A single-equation regressin model can have significant t-statistics and a high  $R^2$  and still forecast very badly period after period. This might result from a structural (in the economy) occuring during the forecasts period and not explained by the model, good forecasts, on the other hand, may come from regression models which have relatively low  $R^2$ s and one or more insignificant regression coefficients. This may happen because there is very little variation in the dependent variable, so that although it is not explained well by the model, it is easy to forecast."<sup>2</sup>

#### 3. DATA

The data for the tax bases were obtained from the National Income Accounts Statistics of NEDA and Annual money wage rate figures attainable for the Central Bank Statistical Bulletin. The tax collections data came from the BIR Statistical Division. Tax data were available in fical year sples from 1961-1974 and in calendar year series from 1974-1979. Conversion of tax data from fiscal Year (FY) to Calendar Year (CY) follows the scheme given below:

Given: FY 2 = July (Year 1) - June (Year 2)
Subtract: July-December (Year 1)
To get: January-June (Year 2)
Add: July-December (Year 2)
= CY 2 = Jan. (Year 2) - Dec. (Year 2)

3) Pindych & Rubinfeld p. 161

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To Illustrate:	Income Tax-Collection.	et
Given:	FY 1964 (July 1963-June 1964)	= ₽419.7 M
Subtract:	July-December 1963	= <u>145.24 M</u>
To get:	January-June 1964	₽274.46 M
Add:	July-December 1964	<b>⊉</b> 173.7 M
=	CY 1964 (JanDec. 1964)	₽448.16 M

<sup>1</sup>The National Bureau of Economic Research set the 10% RMSE as the standard for evaluating the forecasting capability of an equation.

<sup>2</sup>Pindych & Rubinfeld, p. 161.

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#### 4. THE VARIOUS BIR TECHNIQUES MODELS

Over the years, the BIR has used various techniques to forecast tax collections namely: the compound growth rate technique, time-trend analysis and the tax elasticity approach. The first two of these approaches, though Tonger employed. Thus, in the simulations were not derived for these methods, but only for the elasticity approach.

#### 4.1 The Compound Growth Rate Technique

The initial effort to predict tax collections was rather simplistic and crude. The compound growth rate of actual collections in the preceding ten years were computed from income tax, business tax, specific tax and other taxes separately. The growth rates thus obtained were multiplied with the present years tax collections to yield the forecasts for the next year. This technique was used for the 1977-1978 projections.

#### 4.2 Time-Trend Analysis

In order to forecast tax revenue collections in 1979, time-trend analysis was employed. The revenue collections for the period 1970-1979 was plotted against time to get a rough idea of the general

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trend. Based on this, the rate of increase was found to be non-linear. Tax collection was increasing at a decreasing rate and a Compertz curve was fitted to the data. Income tax, business tax and other taxes were separately estimated with this curve. However, forecasts for the specific tax were based on a simple linear regression using consumption of commodities as explanatory variable.

This method shares the same basic drawback of the compound growth rate technique, i.e. it considers time as the sole factor explaining tax receipts. Its advantage over the latter lies in its ability to consider turning points in the pattern of growth.

#### 4.3. Tax Elasticity Approach

Forecasts of tax collection for 1980 onwards were based on the tax elasticity approach. The predicted increment (in absolute terms) in tax revenues in any given year is the product of tax collections in the previous year, the elasticity of the tax with respect to its base and the projected growth rate of the base. The different tax categories were related with different variables reflecting the appropriate tax base. Thus, the individual income tax was related with personal income; the corporate income tax was related to corporate income; the specific tax was related to value of manufacturing domestic product; license, business and other taxes treated as one, were related to industrial and services domestic product. The functional form used was that of the power curve, i.e.,  $Y = bX^m$ , where m is the slope, b is the y-intercept, when the function is estimated by the doublelog transformation. The parameter m is then interpreted as the elasticity of the tax with respect to the explanatory variable considered.

The re-estimated equations using the BIR's tax elasticity approach are as follows: 1.482 Individual Income Tax = 0.000056 (Personal Income) (1) $r^2 = 0.938$ MASE % = 13.7% 1.0174 Corporate Income Tax = 0.3336 (Corporate Income) (2)  $r^2 = 0.960$ RMSE % = 18.0%1.16 Specific Tax = 0.08179 (Manufacturing Domestic Product) (3) RMSE % = 17.1% $r^2 = 0.9696$ License, Business and other Taxes = 0.004989 (Indus-1.133 trial Service Domestic Product) (4)  $r^2 = 0.992$ RMSE % = 7.7% Table 1 presents the actual and simulated values of the various tax categories for the period 1961 -

1979 while Figures 1, 2, 3 and 4 provide a pictorial view of the same.

#### TABLE 1

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Actual and Simulated Values of Various Tax Categories Using BIR's Elasticity Approach, 1961-1979

YEAR	INDIVIDUAL I	NCOME TAX	CORPORATE TAX		SPECIFIC TAX .		LICENSE, BUSINESS & OTHER TAXES	
	Actual	Simulated	Actual	Simulated	Actual	Simulated	Actual	Simulated
	-0.1	70 5	174.4	171.5	259.8	252.4	134.9	142.8
1961	79.1	82.2	211.2	260.2	287.0	282.9	. 169 <b>.8</b>	160.6
1962	74.9	100 8	259.4	307.8	329.3	329.0	190.2	185.2
1963	94.0	113 /	271.6	325.0	371.3	341.9	219.9	208.0
1964	126.6	113.4	298.6	274.7	378.3	362.1	226.2	232.4
1965	140.5	1/7 1	288.4	374.3	438.1	397.3	265.6	257.6
1966	124.5	17/ 9	395.3	416.8	481,2	431.0	287.5	275.9
1967	216.6	174.0	527 9	540.2	544.5	470.0	332.2	298.4
1968	175.1	200.1	602 6	509.8	553.7	511.3	360.0	328.6
1969	252.9	235.2	763 0	725.5	579.7	665.15	373.6	411.5
1970	286.7	291.1	056 7	606.9	645.9	790.7	422.2	456.9
1971	371.9	379.2	950.7	648.6	663.4	924.6	464.2	516.3
1972	508.3	452.0	1957 7	1560.8	828.3	1217.3	599.4	657.8
1973	520.6	621.0	2391 3	1905.0	1623.3	1681.1	940.4	950.1
1974	788.9	1007.1	1954 1	2030.2	1935.1	1944.9	1033.7	1126.9
1975	1119.7	1212.0	2222 3	2479.7	2515.1	2212.3	1315.9	1375.5
1976	1483.0	1511.1	201.9 4	2212.7	3030.1	2663.7	1597.7	1658.9
1977	2473.3	1928.5	2040.0	3008.3	3614.2	3065.8	2041.2	1969.4
1978	2548.9	2397.1	2041,0	3563.6	4072.9	3795.9	3068.3	2537.8
1979	3185.0	3151.7	2012.3			· ·		

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Of the four equations considered, only the simulations for license, business and other taxes resulted in a favorable root mean square percentage error of less than 10%.

#### 5. KINTANAR-MIJARES TAX FORECASTING METHOD

Kintanar and Mijares (1965) suggested different frameworks for forecasting revenues from different tax categories. Based on these, they obtained revenue predictions for 1965-1970 using data from the earlier years.

In projecting revenue from the corporate income tax, the authors divided the work into three steps: (1) estimation of the total number of returns; (2) distributing the figure obtained in (1) to the various income taxe brackets; and (3) estimation of the mean net taxable income for the different tax brackets. The total number of returns was projected using a simple time-trend. This number was then apportioned to the different income brackets by assuming that percentage distribution follows the pattern exhibited by a subset of the total number of returns in 1962. The mean net taxable income for the various tax brackets was computed from the same subset mentioned above to come up with the middle estimate. Low and high estimates were obtained by utilizing the fiducial limits of the 95 per cent

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confidence interval of the mean est. Ste. These Low, middle and high estimates of the mean taxable income of the different tax brackets were multiplied with the appropriate tax rates to arrive at estimates of the average tax assessment for each taxable bracket. The product of the tax assessment valuation and the projected number of returns in each income bracket yields the reverse forecases for the corporate income tax by bracket.

The procedure followed in projecting the revenue from the individual income tax is similar to that of the corporate income tex. First, the total number of individual returns was projected by fitting a time trend to 1959-1962 data. Second, the percentage distribution of 1960 was applied to these projected values to yield the number of returns in each tax bracket. Third, estimates of the mean taxable income in each tax bracket was obtained from a sample of the 1960 data. Fourth, the appropriate tax rates were applied to the estimates obtained in the previous step to come up with the mean tax assessment for each bracket. Again, low and high valuation of the tax assessment were . computed using the 95 per cent fidurial limits of the confidence interval for the mean tax assessment. Finally, the mean tax assessment for each incone tax bracket was multiplied with the corresponding number of returns to estimate the total assessment. This five-step procedure is done for rettax utts from both married and single individuals; then the total

tax assessment for single individuals was added to that for married individuals to arrive at the aggregate estimate for these two groups for each income bracket.

Since the computation for the individual income tax and corporate income tax simulations over the estimation period 1961-1979 would require tedious effort, we simply estimated the RMSE % for the <u>ex ante</u> forecasts made by Kintanar and Mijares for 1963-1965. These estimates are presented in Table 2.

#### TABLE 2

#### Medium Projections from Kintanar

•	(FM)		
	<u>1963</u>	1964	RASE %
Projected Individual Tax Assessment	111.5	124.9	
Actual Individual Tax Assessment	94.0	126.5	13.2 %
	1964	1965	
Projected Corporation Tax Assessment	299.02	322.17	
Actual Corporation Tax Assessment	271.6	298.6	9.1 %

In forecasting the various components of license, bus ness and other taxes, the only explanatory variable considered was time. Tax collection on sale of non-essential or luxury goods as described under Section 184 of the National Internal Revenue Code (NIRC) denoted by  $T_L$ , compensating tax collections under Section 204 of the NIRC denoted by  $T_c$  and specific tax collected on sale of cigars and cigarettes denoted by  $T_s$ , were directly regressed on time. However, in forecasting the tax receipts from sales of semi-luxury and non-luxury items, their tax bases denoted by  $B_s$  and  $B_n$  respectively, were regressed on time and the projected tax bases were then multiplied by the relevant tax rates.

To arrive at  $B_s$  and  $B_n$  values, the ratios of each to the actual total sales tax were taken:

(a)	Tax on Semi-Luxury	;	Tax on Non-Librury
1	Total Sales Tax		Total Sales Tax

then multiplied by the manufacturing component of Gross Demostic Product (CDP)

(b) <u>Tax on Semi-Linknry</u> x  $CDP = B_s$ ; Total Sales Tax <u>Tax on Non-Linknry</u> x  $CDP = B_n$ Total Sales Tax (c)  $B_s$  x .35 (average of 30% and 40%, the tax rates prevalent during the estimating years)

= 
$$S_L$$
  
 $B_n \times .07 = N_L$   
where  $S_L$  represents tax receipts from semi-luxury  
items and  $N_L$  represents tax receipts from non-  
luxury items.

 $\rm B_{g}$  was multiplied by 35% (average of 30% and 40%, the tax rates prevalent during the estimating years); and  $\rm B_{n}$  was

multiplied by 7%

In re-estimating the specification of Kintanar and Mijares using more recent data, the data on sales taxes, compensating tax and specific tax on cigars and cigarettes from 1961 to 1976 were obtained from the Statistical Division of the BIR. Data for years later than 1976 were not available because the amendments made in the NIEC resulted in unspecified and questionable tax reports which have yet to be reconciled.

The resulting equations are as follows: .06155 (Plan) الن . 2164084t (5) 9.37655 RMSE % = 103.58.406 159 (6) -48.2251<sub>e</sub> RMSE % = 174.4 % .2067 168.6 (7) 204.58t 1347.94t + 5.94 % = 23.36 RMSE .977 -\$20. JZ (1000 (8) .14976 + .0185711t 89.28 RMSE % = .082 9566 t (9) 84.728 + 21.684t r<sup>2</sup> FIDE % = 20.09 .689 Note the very high RMSE % for equations (5), (6) and (8) indicating the inadequacy of these specifications in forecasting.

Table 3 presents the actual and simulated values of the non-income tax categories considered by Kintanar and Mijares while a graphical picture is given in Figures 5, 6, Actual and Simulated Values of Various Tax Categories Based on the Specifications of Kintanar & Mijares 1961-1976

	S <sub>L</sub> =Tax on Ser	i-luxury Itens	N_=Tax on Nor	-luxiry Iters	B <sub>S</sub> =Tax Base of Tax Co	f Seni-Luxury llections	B =Tax Base of Tax Coll	Non-Luxury ections
YEAR .	Actual	Simulated	Actual	Similated	Actual	Simulated	Actual	Simulated
1961	1.79	-16.8788	53.02	383.3158	75.33	- 48.23	3352.59	5475.94
1962	1.29	5,7975	67.82	303.2806	68.56	16.56	3611.76	4332.58
1963	(.73)	21.8867	80.5	251.8865	(40,68)	62.53	4476.79	3598.371
1964	2.63	34.3664	89.11	229.1331	132.76	98.19	4501.68	3273.33
1965	1 37	44.5630	92.23	235.0215	72.69	127.32	4895.66	3357.45
1966	83	53,1841	113.46	269.5504	39.09	151.96	5341.43	3850.72
1067	.03 96	60 6521	116.13	332.7205	48.63	173.29	5857.71	4753.15
1907		67.2393	139,44	424,5311	47.73	192.11	651.2.95	6064.7.
1969	8 16	73,1318	141.37	544.9836	383.10	208.95	6635.72	7785.43
1970	1 53	78,4622	182,83	694.0766	77.55	224.18	9261.89	9915.53
1971	1.0	83,3284	206.09	871,8108	54.80	238,08	11266.30	12:54.44
1972	1 76	87.8050	213.44	1078.1855	105.77	250.87	12774.83	154.02.0
1073	2.110	91.9496	257.03	1313.2021 -	161.21	262.71	17183.55	18760.03
107/	10.53	95 8081	418,24	1576.8592	600.44	273.74	23813.16	22526.56
1075	11.04	99 4175	388.59	1869-1575		- 284.05	27707.66	26702.25
1975	1.66	102,8081	472,62	2190,097	113.91	293.74	32346.43	31287.15
					-			

Actual and Simulated Values of Various Tax atesees Based on the Specifications of Kintarar & Mijares 1961-1976

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TS = Cigars and Cigarettes TC = Corpersating Tax  $T_{L} = Tax$  on Lixing Items (184) Simulated Actual YEAR Simulated Actual Similated Actual 106.412 143.82 .16833 .14 9.3765 2.25 117.254 1961 146.80 .17761 .18 8,4149 6.18 128.096 1962 159.74 .18691 .15 7.5519 4.30 138,938 1963 145.08 .19619 .26 6.7774 4.52 149.78 1964 115.76 .20548 .39 6.0824 3,53 160.622 1965 136.59 .21476 .06 5,4886 6.06 171.464 1966 147.69 .22405 .38 4,8988 4.93 182,306 1967 158,89 .23333 .26 4.3964 3.46 193.148 1968 167.98 ,24262 .09 3,9455 6.53 203.99 1969 167.28 .25191 .30 3.5409 4.63 214.832 1970 190.11 ,26119 .14 3.1777 1.76 225.674 1971 212.26 .27046 .14 2,8518 8.48 236.516 1972 231,98 .27976 .23 - **2,**5594 5.55 247.358 1973 285.61 .02891 .16 2,2969 3.42 258,200 1974 325,00 .29833 .62 2.0613-.71 269,042 1975 466.66 .30762 .37 1,8499 1.23 1976 1

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7, 8 and 9

6. JURADO-ENCARNACION COVERNMENT MODEL

Jurado and Encarnacion (1972) constructed a fiscal sector model employing ordinary least squares on data from 1955 to 1969. This model consists of eleven structural equations (ten of which are related to taxes) and two identities (one of which is tax related). The authors broke down total tax revenues into six tax categories. Revenue from corporate income tax was related with gross national product and the wage rate, as a proxy for corporate income, and a dummy variable to reflect tax rate changes. On the other hand, individual income tax receipts were regressed on personal income lagged one year. Indirect business tax collections (primarily from the sales or the percentage tax) were made a function of GNP. Tariff duties were related to the total value of imported goods and services and a variable that measures imports from the United States. Excise taxes, fees and penalties, charges on forest products, the franchise tax, wharfage and other fees were all classified under "other indirect taxes" and were explained by gross national product sold demestically and exports. Total taxes collected by the national government were related with the aggregate of all taxes received by the government as a whole. Finally, all other taxes of the National Government was obtained as a residual. The rest of the equations were devoted to explaining government expenditures and its components.

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(10)Tbd = f(Y, W, Utr)(11)Ibi = f 🚫

- (12) $Tp = f(\mathcal{Y}_{p}-1)$
- . (13) Toi = f(Y-X), x)

where.

Tod - direct income tax receipts from business enterprises which otherwise is known as corporate tax;

Tpd - direct income tax receipts from persons

- Thi indirect tax receipts from business enterprises sales Tax
- Toi receipts from other indirect taxes excise Tax; feest penaltes, Y Gross National Product Ya Gross Ya Gross

Yp - compensation of employees plus entrepreneurial

and property income of persons

- exports of goods and services Х

- annual money wage rate, computed as equal to the W daily wage rate of unskilled industrial workers in Manila multiplied by 250; in pesos

Utr - durmy variable for a change in tax rates;

= 1 for years beginning 1968, = 0 for years below 1968.

The model above was reestimated using data from 1961 to

1979. The resulting equations are as follows:

Tbd = 1137.10 + .716 + .618W + .800 Utr (10a) $R^2 =$ 

<sup>&</sup>lt;sup>2</sup>The import tax equation was not included in this report because so far, we are only interested in checking the models using BIR's internal data The other identitites were not also included in the reestimation.

Note that in (10a) the regression coefficient for wages is positive as opposed to the negative coefficient that Jurado and Encarnacion previously obtained. To resolve this error, we tried regressing (Y-W) as a proxy for corporate income, against Tbd with the following output:

Tbd = 
$$18.39 + .013$$
 (Y-W) + 415.69 Utr  
 $r^{2} = .921$  RMSE % = 20.86 (10b)  
Tpd =  $(367.91) + .025$  Y<sub>p-1</sub>  
 $r^{2} = .967$  RMSE % = 73.38 (11a)  
Tbi =  $(33.63) + .005$  Y  
 $r^{2} = .887$  RMSE % = 20.93 (12a)  
Tbi =  $(473.12) + .0399$  (Y-X) + .028X  
 $R^{2} = .986$  RMSE % = 30.01 (13a)

It can be observed that, not a single one of the four equations resulted in a computed root mean square percentage error lower than 10%. Although the equation for other indirect taxes (1\$a) yields a high r<sup>2</sup>, the turning point error must have been large so as to produce, at the same time, a Mint Mean high percentage error.

#### TABLE 4

#### Actual and Simulated VAlues of Various Tax Categories in Jurado and Encarnacion Model, 1961-1979

265.14

335,41

441.27

598.79

712.23

971,68

1499.66

1730.01

2107.82

2525.36

2985.25

YEAR

1961

1962

1963

1964

1965

1966

1967

1963

1969

1970

1971

1972

1973

1974

1975

1976

1978

1979

-1977--

849,1649

931,2244

1028,1429

1101.0082

1302.9277

1657.2966

1834.6670

2055.3386

2334.3776

2639.2008

3163.0132

602.6

763.9

956.7

867.1

1857.7

2391.3

1854.1

2222.3

2048.6

2641.6

2872.3

252.9

286.7

371.9

508.3

520.6

788.9

1119.7

1483.0

2473.3

2548.9

3185.0

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1000.46

1127.95

1228.32

2370.41

2885,11

2287.76

4345.19

5679.86

6005.76

8351.67

1096.5143

1395.8010

1624.9543

2194,9644

3250.1155

3833.1834

4514.8479

5334.2782

6277.2100

7829.4916

other indirect taxes indirect taxes direct taxes direct taxes Toi Thi-from Business Tpd-from persons Tod-from Business Simulated Actual Simulated Actual Simulated Actual Simulated Actual 111,4993 362.32 41,0637 56.46 79.1 82.00 192.4119 174.4 175.8013 411.74 50.8793 74.9 55.81 75.29 215.4614 211.2 277.5810 473.28 64.5917 22.73 84.09 249.3363 94.0 259.4 338.8348 530.54 96.26 72.4826 30.22 269,2849 126.6 271.6 547.99 411.7238 82.4083 . 97.13 140.5 61.42. 293.0850 298.6 498.0423 632.76 94.1306 120.35 124.5 109.26 288.4 321.3834 615.0813 697.05 108,9649 149.78 122,02 216.6 358.0433 395.3 739.3149 488,69 124,1361 175.1 207.57 143.92 869.7383 527.9 869.4383 861.82 140.1241

156.06

189.04

208.85

223.68

264.99

432.19

400.34

475.51

639.94

1144.67

1467.13

173.5561

212.5146

241.9295

221./019

462.3896

533,4429

621.1515

732.0366

852,6440

1060





## FIGURE 12



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#### 7. THE DIOKNO PUBLIC SECTOR MODEL

Diokno (1972) constructed a fiscal policy model aimed at evaluating the adequacy of the Philippine tax system in meeting the public revenue needs of the Philippine economy. The model consists of six behavioral equations and six definitions. Four variables were considered as exogenous and values for the said variables were obtained from the National Economic Council's target in the Feur-YEAR DEVELOPMENT PLAN. These variables are CNP, government investment expenditures, export tax revenue and the money wage rate. There are twelve endogenous variables; six of which are related to government revenue and the rest are related with government expenditures. The other two endogenous variables are government savings and government surplus.

Total government tax revenue was defined as the sum of direct taxes, internal indirect taxes and indirect taxes on the foreign trade sector. Each of these three tax categories was related to GNP. Similarly, non-tax government revenue was then defined as the sum of total tax collections and non-tax receipts of the government.

For our purposes, we made use of only two government tax revenues as endogenous variables:

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- (1) direct which include income taxes residence tax, transfer tax and real property tax, specific taxes on domestic products, license and business tax, DST, Franchise tax, charges on forest products.)
- (2) indirect taxes which include sales tax, fines and penlties,
   mining tax. These two tax classifications were
   related to GNP.

A simple model, i. e.,  $Y = b + mX_2$  for each tax classification was specified, where b is the y-intercept and m is the slope. The reestimated equations are as follows: *signal* 

(1) Direct Taxes = -219.0031 + 0.024 (GNP)  $r^2 = .994$  RMSE<sup>§</sup> = 11.91 (14) (2) Indirect taxes = -390.25 + 0.0473 (GNP)  $r^2 = .986$  RMSE<sup>§</sup> - 19.78 (15)

Note that RMSE% from (14) and (15) are relatively lower than those obtained from the specification of Kintanar and Mijares as well as those of Jurado and Encarnacion.

### TABLE 5

## Actual and Simulated Values of Direct and Indirect Taxes, 1961-1979

•

	DIREC	T TAXES	INDIREC	T TAXES Simulated	
YEAR	Actual	Simulated			
		0/0.05	421 8	228.7	
1961	289.7	242.35	421.0	305 1	
1962	340.5	299.2	494.0	200C	
1963	415.2	383.3	568.6	418.0	
1964	463.7	431.7	640.0	482.9	
1965	517.8	492.5	656.6	564+5	
1966	506.1	564.4	767.5	661.0	
1967	638.1	655.4	835.2	783.1	
1968	794.9	748.4	909.9	907.9	
1969	885.0	846.4	1017.7	1039.4	
1970	1082.3	1051.5	1189.5	1314.6	
1971	1366.6	.1290.3	1335.9	1635.0	
1972	1409.2	1470.7	1452.0	1877.0	
1973	1436.3	1960.3	2635.4	2534.0	
1974	3216.0	2822.4	3320.3	3690.9	
1975	3100.4	3258.1	3788.3	4275.5	
1976	3735,8	3795.9	5021.6	4997.1	
1977	4569.2	4476.0	6319.8	5909,8	
1978	5237.4	5215.3	7150.5	6901.8	
1979	6153.0	6489.0	9819.8	8610.8 \	







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#### 8. SUMMARY AND CONCLUSION

We have reestimated and tested the models suggested by the different works undertaken in tax forecasting using:

- (a) 1961-1979 as the estimation period;
- (b) r<sup>2</sup> and RMSE % statistics to test the model's forecasting capability;
- (c) BIR Statistical Division's data for the tax variables; and
- (d) the National Income Accounts revised series for the indicators of the tax bases.

The results of this exercise are summarized in the following tables:

TABLE 6 SUMMARY OF TEST RESULTS

Trable 6.1 R<sup>2</sup> and RMSE % for BIR Tax Elasticity Model

	Individual Income Tax	Corporate Tax	Specific Tax	License and Business Tax	Average
R <sup>2</sup>	.988	.960	.970	.992	.978
RMST %	13.68	17.99	17.15	7.73	14.14

-	Tudividua Income Tax	<u>13/</u> Corporate Tax	3/ Tax on Luxury Items	Base of Semi-Luxury Items	Base of Non-Luxu- ry Items	Compen- sating Tax	Cigars & Cigarette	s Average
R <sup>2</sup> RMSE	- % \3.2	- 9.1	.406 103.58	.209 174.4	.977 23.36	.082	.689 20.1	.473 61.86
' no re able 6.	$3 R^2$ and R	was cone (	ing used only			-		,
		MSE % for .	turado-Encarn	acion Model				
: 	Indivi Income 3	dual lax (Tpd)	Corporate Tax (Tbd)	Specific Taxes (Tol)	License Taxes	& Busines s (Tbi)	s Aver	age
	Indivi Income 9	MSE % for . dual tax (Tpd) .921	Corporate Tax (Tbd) .967	Specific Taxes (Tol) ,887	License Taxes	& Busines s (Tbi) .986	s Avera	age 940

Table 6.2  $\mathbb{R}^2$  and RMSE % for Kintanar and Mijares Model

Table 6.4 R<sup>2</sup> and RMSE % for Diokno's Model

, <u></u>	Indirect Taxes	Direct Taxes	Average
- •. - R <sup>2</sup>	.986	.988	.987
RMSF Z	19.78	11.91	15,84

Based on these, we make the following observations:

 Using tax bases as explanatory variable for the respective taxes is not a guarantee that the resulting forecasts will be highly accurate.

Although Kintanar and Mijares' work virtually used time to explain the fluctuations in taxes, the RMSE% averaged at 39% which is only a little higher than the average RMSE% of Jurado and Encarnacion's model (36.3%), using indicators of the tax bases as the determinants. Of course, Kintanar and Mijares did a lot of tedious disaggregation in projecting taxable income by brackets for the individual and corporate categories which probably offset the large mean squared errors derived from the purely simple trend regressions done with the other specific taxes.

2) The more aggregate the endogenous variable, the better the simulations. In other words, as we try to explain particular taxes in detail, the more difficulties we meet.

The low average RMSE% (15.85) in Diokno's work reflects this advangtage of aggregation more than anything else. On the other hand, the high RMSE% prevalent in Kintanar and Mijares' equations may be explained partly by the more volatile movements inherent in particular taxes, which are more difficult to capture.

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- 3) The elasticities approach provided the lowest average RMTE? error (14.14); indicating that there is a practical and theoretical foundation in assuming that taxes move in a power function fashion against time and the tax bases, rather than in a simple linear one.
- 4) Lastly, the general underestimation of almost all historical simulations is to be expected because of the inability of simulations to capture other explanatory variables which may have caused an upward shift in the values of the endogenous variables through time.

The future direction of the Bureau's forecasting efforts should focus on the following facets:

- I) Spatial or regional forecasting An earlier exercise using linear programming, tried to derive the regional implication of the tax forecast and determine tax goals for the same. This was later substituted by a houristic approach that was based on the regional gross domestic product. Even this approach has its problems knowing that the BIR has seventeen revenue regions as against the Whilippines' administrative delineation which has only twelve regions.
  - 2) Disaggregation forecasting It was found out that the more detailed the analysis of taxes, the more

difficult the required forecasting techniques become. This very difficulty represents a challenge to econometric forecasting as a whole.

- 3) Refinements in estimation methodology The methods used thus far, were generally limited because of the lack of cyclical analysis and the exclusion of lag variables. Any forecasting exercise is a search for the particular methodology most suited to give accurate projections.
- 4) Refinements in specification Other variables aside from the tax bases, should be considered in terms of their ability to explain the movements of particular taxes. Since a tax revenue represents an income for the government, it is not totally unwarranted to consider costs of tax collection as one of the explanatory variables. In other words, a function specified as: Tax revenue = f (tax base, cost of tax administration), where both determinants are expected to relate positively to tax revenue, is worth looking into. Other dumny variables representing tax annesty periods, new PD's should also be introduced.
  - 5) Widening of scope To the extent that tax receipts are determined by tax bases, it is imperative that we derive good forecast for the tax bases first before proceeding with the tax forecast proper.

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Essentially, this would require an econometric effort in a macro-scale i.e., calling for a wider and more comprehensive view of the economy.

Of course, the common thread woven into these factors is the need to establish a definite and stable data base. This calls for a close working effort with the Data Processing Center and the Statistical Division of the Bureau.

It is in line with the preceding analyses that we approach the second photo of this project. So far, we found out that the present Black elasticity approach produced the most acceptable historical simulations among all the other existing tax forecasting models and methods. This implies two things:

- (a) While we are still in the process of developing a more suitable forecasting model with the minimum standard error of forecasts, the Bureau may use the elasticity approach for its immediate need to forecast tax revenues; and
- (b) The RMSE % computed value for the BIR's elasticity approach should now serve as the benchmark in evaluating the structures to be formulated and estimated in the future.

We will also consider other means of evaluating the forecasting ability of the models particularly for those specifications whose forecasting errors do not deviate far from each other.

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### BIBLIOGRAPHY

Bahl, Roy. "A Representative Tax System Approach to Measuring Tax Effort in Developing Countries", <u>DMF Staff Papers, 1972</u>
Evans, Michael. <u>Macroeconomic Activity</u>. University of Pennsylvania, 1967

Diokno, Benjamin E. <u>The Philippine Tax System and the Public</u> <u>Revenue Needs of the Philippine Economy</u>, FY 1972-1975 Quezon City, 1972.

Jurado, Gonzalo, and Jose Encarnacion, Jr. <u>Covernment Submodel of</u> <u>the Philippine Economy, 1955-1969</u>. Philippine Economic Journal, Vol. XI, No. 2, 2nd semester, 1972.

Kintanar, Agustin Jr., and Mijares, Tito. <u>Estimation and Forecast-</u> ing of Tax Revenue in the Philippines, Q.C. IEDR, Discussion Paper p. 2-65.

Klein, Laurence. Econometrics. Maruzen Asian Edition 1953.

- Pindyck R. and Rubinfeld, D. <u>Econometric Models and Economic</u> Forecasts, 1976
- Sinay, Crean Cilda C. Buoyancy and Elasticity of the Philippine Tax System, CY 1961-1972. Q.C. 1974

Theil. Econometric Forecasts and Policy, Amsterdam, 1961 BIR Forecasting Models.