

On the Causality between GDP and Health Care Expenditure in Augmented Solow Growth Model*

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ABSTRACT

This paper examines conditional convergence of OECD countries in gross domestic product (GDP) and health care expenditure (HCE) per capita. It presents estimation of the augmented Solow model suggested by Mankiw, Romer and Weil (1992) to explain variation in output and expenditure per capita across countries. The variation is due to different steady state growth paths resulting from differences in the countries savings rate, education, and population growth. This paper is an extension of the MRW model by incorporating health capital proxied by HCE to the augmented Solow model. The analysis is further related to the studies of health care expenditure where GDP per capita appear to be the main factor determining the level of expenditure on health care. The issue of causality relationship between GDP and HCE is investigated. The empirical analysis is based OECD countries' data for the period of 1970-1992. The results indicate that OECD countries converge at 3.7% per year to their steady state level of income per capita. The results show that HCE has positive effect on the economic growth and the speed of convergence. The speed of convergence is found to be sensitive to whether one imposes a constant or estimate the depreciation and technological growth components. With no restrictions imposed the convergence rate is 5.2%. Considering the rate of convergence in the HCE model the results show that OECD countries converge at 2.7% to their steady state of HCE per capita. In the HCE model a regression of the speed of convergence on variables determining the rate of convergence show close link to the variables characterizing the health care system of sample countries.

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Keywords: Solow growth model; health care expenditure; GDP; convergence; OECD.

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1. INTRODUCTION

Following the classic paper by Solow (1956), there has been a significant development in the theoretical and empirical literature of endogenous growth. Much attention is paid to the issues of the failure of countries to converge in per capita income. A large number of the studies use data from the Penn World Tables (Summer and Heston, 1991). Solow studied economic growth by assuming a neoclassical production function with decreasing returns to capital and the rates of saving and population growth considered as exogenous. He showed that the rates of saving and population growth determine the steady state level of income per capita across countries. The countries reach different steady states due to variations in the key factors determining the level of steady state. The higher the rate of saving (population growth) the richer (poorer) the country and the lower the rate of return to physical and human capital. More than half of the cross-country variations in income per capita can be explained by those two variables. A large body of literature found Solow's prediction of the direction of the effects to be consistent with the empirical evidence, but not the magnitudes.

The issue of convergence has developed in three different directions in the growth literature. The first type, beta convergence, considers the speed at which income per capita tends to a steady state value of income from some initial level. Depending on the type of data used, the estimate of speed of convergence is based on the coefficients of lagged income or some initial conditions. Convergence can be conditional or unconditional on some country specific variables and to a common or country specific steady state assuming a homogenous or heterogeneous growth rate (see Lee, Pesaran and Smith, 1997). The third type, sigma convergence, considers the behavior of cross-country variance of income over time. It assumes global technology and tastes that determine convergence to a common steady state of income at a same rate across countries (see Quah, 1993). In the sigma type, the income per capita is treated as an integrated variable. The objective is to determine whether sample countries share a common deterministic and /or stochastic trend (see, Durlauf (1996) and Evans (1996)). In this paper we consider the conditional convergence.

The Solow model was augmented by Mankiw, Romer and Weil, (1992) to include accumulation of human capital. According to this model, the convergence path is a linear time trend, the slope of which is determined by the rate of exogenous technical progress, while the intercept reflects the effects of factors characterizing the conditional convergence. Inclusion of human capital was motivated because accumulation of human capital may be correlated with saving and population growth rates resulting in biased estimated effects. Thus, the exclusion of human capital can explain the overestimation of the effects of saving and population growth on the level of income. The explanatory power of the human capital augmented model increases to about 80 percent of the variation in income providing explanation to the differences in the wealth among sample nations. Empirical results show evidence that countries converge given differences in their saving and population growth rates is taken into account. Furthermore, the model explains the magnitude of over-estimation of the influences of saving and population growth.

The objective of this study is to examine conditional convergence of OECD countries in gross domestic product (GDP) and health care expenditure (HCE) per capita. It presents estimation of the investment in physical, human and health capital augmented growth model to explain variation in output and HCE expenditure per capita across countries. Although the Solow model has been augmented in different ways¹, there are few studies that examine the effects of health capital on growth. Knowles and Owen (1995) examine the effects of incorporating health capital in the MRW model. Cross-sectional results suggest a strong and robust relationship between health capital and income. In Knowles and Owen (1997), the labor variable in an aggregate production function is education and health augmented. Their result suggests that, incorporating human and health capital as labor augmenting or as separate factors of production does not change the conclusions empirically. Again results suggest a strong and positive relationship between income and health. Unlike in the MRW model, the effect of human capital is insignificant. Temple (1999) also found the effects of human capital to be data specific and sensitive to the model specification and estimation methods used.

This paper is an extension of the previous literature in a number of ways. First, it augments the Solow model to health capital. Health capital is proxied by health care expenditure per capita.² Knowles and Owen (1995, 1997) used life expectancy as proxy for health. Second, the analysis is further related to the studies of health care expenditure where GDP per capita appears to be the main factor determining the level of expenditure on health care. The issue of causality relationship between GDP and HCE is investigated. Third, in the later model a regression of the speed of convergence on variables determining the speed of convergence show any link to the variables characterizing the health care system of sample countries. Fourth, we investigate how sensitive the speed of convergence is to the in the literature frequently imposed restriction of constant depreciation and technological growth components. Finally, the empirical analysis is based on a homogenous group of countries' (OECD) data for an extended period of 1970-1992.

The findings in this paper are in line with those of Mankiw et al (1992) and support the assumptions of decreasing return to capital, improved prediction performance of the model and countries convergence to different steady states. Results indicate that OECD countries converge at the rate of 2.7% per year to their steady state of income per capita with the usual Solow model. HCE has significant effect on the economic growth and the speed of convergence. When investment in health is explicitly taken into account in the model, the speed of convergence is increased to 3.7%. The speed of convergence is also found to be sensitive to various specifications of capital depreciation and technological growth components. At the absence of any assumptions about the sizes of those two components, the rate of convergence increases to 5.2%. Considering the rate of convergence in the HCE

¹ For other forms of augmentations see Cadoret and Tavera (1998) in the context of budget deficit, Fölster and Henrekson (1998) in the context of government expenditure and Knight, and Loayza and Villanueva (1993) in the context of degree of openness of the economy to foreign trade.

² Alternative proxies of health capital other than health care expenditure are the use of all causes, life expectancy, mortality rate, maternal mortality, and perinatal mortality. We are currently working on analyses of sensitivity of the rate of convergence when using different measures of health capital.

model the results show that OECD countries converge at 2.5% to their steady state of HCE per capita.

In Section 2 the growth model is outlined. The model is augmented to include investment in physical, human and health capital. The issues of variations in output per capita across countries, endogenous growth and convergence are discussed. The analysis is further extended to the health care literature by analyzing the issues of causality between GDP and health care expenditure. Section 3 presents the data from OECD countries. Section 4 discusses the issues of estimation and presents empirical results under the various specifications suggested in Section 2 and discuss their implications for the speed of convergence in income per capita or health care expenditure. Section 5 summarizes the results and concludes.

2. THE HEALTH CAPITAL AUGMENTED MODEL

2.1 The Augmented Growth Model

In the Solow's (1956) growth model the rates of saving, population growth and technical change are exogenous. Assuming a Cobb-Douglas functional form and two factor inputs of capital and labor, the labor-augmenting technological progress at time t is written as

$$(1) \quad Y_{it} = K_{it}^{\alpha} A_{it} L_{it}^{1-\alpha} \quad 0 < \alpha < 1$$

where Y is output, K capital, L labor, and A the level of technology. The subscripts i and t denote country and time periods, respectively. In order to simplify the notation we drop the subscript i . Labor and technology are assumed to grow exogenously at rates of n and g as

$$(2) \quad L_t = L_0 e^{nt}, \quad A_t = A_0 e^{gt}$$

that makes the effective units of labor ($A_t L_t$) to grow at the rate $n + g$. Defining $k_t = (K_t / A_t L_t)$ and $y_t = (Y_t / A_t L_t)$ as the stock of capital and the level of output per effective unit of labor, the evolution of capital is governed by

$$(3) \quad \dot{k}_t = s_t^k y_t - (n_t + g + \delta) k_t = s_t^k k_t^{\alpha} - (n_t + g + \delta) k_t$$

where a dot indicates change, s_t^k is a fraction of output invested in physical capital in period t , and δ is the rate of depreciation. The stock of capital (k_t) converges to a steady state value of capital (k_t^*) defined as

$$(4) \quad k_t^* = [s_t^k / (n_t + g + \delta)]^{1/(1-\alpha)}$$

which is positively related to the rate of saving but negatively to the growth rate of population.

The Solow model concerns the impact of saving and population growth on real income. Substitution of (4) into (1) and taking logs, the steady state income per capita is written as

$$(5) \quad \ln y_t = \beta_0 + \frac{\alpha}{1-\alpha} \ln s_t^k - \frac{\alpha}{1-\alpha} \ln(n_t + g + \delta) + \varepsilon_t$$

where $\beta_0 = (\ln A_0 + gt)$ denotes the technology factor and ε is stochastic country specific shock. The model in (5) has frequently been used as the basic model in empirical specifications (see e.g. Summers and Heston (1988), Barro and Sala-I-Martin (1992), Islam (1995), among others). The rate of saving and population growth are assumed to be independent of ε and the model is estimated with OLS.³

Assuming that the countries are currently in their steady states, Mankiew, Romer and Well (1992) used (5) to see how saving and population growth rates explains the difference in the current per capita income across countries. The coefficient of capital (α) was found to be high requiring a definition of capital in broad sense that incorporates human capital. Thus, human capital was included as another input of production (see Barro and Lee (1993), Benhabib and Spiegel (1994)). Augmentation of human capital to the process of growth showed to be useful concerning the performance of the model and the size of α . Ignoring human capital affects the coefficient on physical capital investment and population growth leading to incorrect conclusions. The production function in (1) is rewritten as

$$(6) \quad Y_t = K_t^\alpha H_t^\beta A_t L_t^{1-\alpha-\beta} \quad \alpha + \beta < 1$$

where H is the stock of human capital and in addition to growth in physical capital (3), the stocks of human capital growth is determined by

$$(7) \quad \dot{h}_t = s_t^h y_t - (n_t + g + \delta)h_t = s_t^h h_t^\beta - (n_t + g + \delta)h_t$$

where s_t^h is fraction of output invested in human capital in period t and $h_t = (H_t / A_t L_t)$ is human capital per effective unit of labor. The relation in (7) indicates that the stocks of physical and human capital grow only if new investment exceeds depreciation adjusted for population growth and technological progress. The steady state income per capita as a function of population growth and accumulation of physical and human capital is given as

$$(8) \quad \ln y_t = \beta_0 + \frac{\alpha}{1-\alpha-\beta} \ln s_t^k + \frac{\beta}{1-\alpha-\beta} \ln s_t^h - \frac{\alpha+\beta}{1-\alpha-\beta} \ln(n_t + g + \delta) + \varepsilon_t.$$

³ For reasons of making assumptions of independence see Mankiew, Romer and Well (1992) pp. 411-412.

Similar to the human capital augmentation, the Solow model can be augmented to investment in health as well. The evolution of the health care expenditure is determined by

$$(9) \quad \dot{e}_t = s_t^e y_t - (n_t + g + \delta)e_t = s_t^e e_t^\theta - (n_t + g + \delta)e_t$$

where s_t^e is the share of output invested in health capital in period t and $e_t = (E_t / A_t L_t)$ is effective investment in health per capita. The model in (8) is then rewritten as

$$(10) \quad \ln y_t = \beta_0 + \frac{\alpha}{1 - \alpha - \beta - \theta} \ln s_t^k + \frac{\beta}{1 - \alpha - \beta - \theta} \ln s_t^h + \frac{\theta}{1 - \alpha - \beta - \theta} \ln s_t^e - \frac{\alpha + \beta + \theta}{1 - \alpha - \beta - \theta} \ln(n_t + g + \delta) + \varepsilon_t$$

where the model can be estimated with OLS. The model in (10) indicates that the steady state path for the log of income per capita follows a linear time trend. The slope of this linear trend is exogenously determined by the rate of technical progress, while the intercept reflects the effects of the rate of population growth and investments in physical, human and health capital. Adding human and health capital improves the performance of the Solow model. Investment in human, physical and health capital are expected to increase the level of income per capita, while high population growth lowers income per capita. Human and health capital accumulations increase also the impact of physical capital accumulation on income. Human capital and healthiness are positively correlated with saving rate and negatively correlated with population growth.

2.2 The Endogenous Growth and Convergence

The endogenous growth models are characterized by the assumption of non-decreasing returns to the set of production factors implying countries with higher saving rate to grow faster. Hence, countries do not need to converge to a common level of income per capita even if they employ the same technology. The model predicts that countries reach different steady states. It does not predict convergence between countries, but only convergence within country or convergence to own steady state value of per capita income. The convergence is thus conditional on the determinants of the steady state, accumulation of various components of capital and population growth. Predictions about the speed of convergence is given by

$$(11) \quad d \ln y_t / dt = \lambda [\ln y_t^* - \ln y_t]$$

where $\ln y_t^*$ and $\ln y_t$ are the log of steady state level and actual values of per capita income in period t given by (10) and λ is the annual rate of convergence at which the economy moves to own steady state

$$(12) \quad \lambda_i = (n_i + g + \delta)(1 - \alpha - \beta - \theta).$$

Although the rate of population growth differs across countries and over time, in previous studies λ is assumed to be constant across countries at a value interpreted as the average speed of convergence. Equation (11) implies that

$$(13) \quad \ln y_i = (1 - e^{-\lambda\tau}) \ln y_i^* + e^{-\lambda\tau} \ln y_0$$

where $\ln y_0$ is log of income per capita at some initial year. According to the model in above, countries are different both in their levels of income per capita and their growth rates in income. The latter differences result from differences in the initial level of income per capita and the steady state value. For given initial level of income countries with higher rates of investment in physical, human and health capital or lower population growth will experience faster growth. On the other hand countries with higher level of initial income will experience slower growth. Subtracting $\ln y_0$ from both sides and substituting for $\ln y_i^*$ the model

$$(14) \quad \begin{aligned} \ln y_i - \ln y_0 &= \beta_0(1 - e^{-\lambda\tau}) - (1 - e^{-\lambda\tau}) \ln y_0 \\ &+ (1 - e^{-\lambda\tau}) \frac{\alpha}{1 - \alpha - \beta - \theta} (\ln k_i - \ln(n_i + g + \delta)) \\ &+ (1 - e^{-\lambda\tau}) \frac{\beta}{1 - \alpha - \beta - \theta} (\ln h_i - \ln(n_i + g + \delta)) \\ &+ (1 - e^{-\lambda\tau}) \frac{\theta}{1 - \alpha - \beta - \theta} (\ln e_i - \ln(n_i + g + \delta)) + \varepsilon_i \end{aligned}$$

is the health capital augmented Solow model in which the growth in income is a function of the determinants of steady state and the initial level of income. Thus, in the endogenous growth model outlined above there is no common steady state level of income and the difference in income per capita among countries can persist indefinitely for the same savings and population growth rates.

Solow's model predicts both the sign and the magnitude of production factors of saving and population growth on the standards of living. Real income is predicted to be higher in countries with higher saving rates and lower in countries with higher values of $(n_i + g + \delta)$. Since the capital share in income (α) is about 1/3, the elasticity of income per capita with respect to the saving rate ($\ln s_i^k$) is about 0.5 and the elasticity with respect to $\ln(n_i + g + \delta)$ about -0.5 . In empirical studies the sum of g (0.02) and δ (0.03) is assumed to be constant and equal to 0.05. MRW found that reasonable changes in this assumption have little effect on the estimates using US data. The model in (14) can then be estimated both with and without the constraint that the coefficients of $\ln s_i^k$ and $\ln(n_i + g + \delta)$ are equal in magnitude but opposite in sign, and tested. Alternatively, one can estimate μ from the following relation

$$\begin{aligned}
(15) \quad \ln y_t - \ln y_0 &= \beta_0(1 - e^{-\lambda\tau}) - (1 - e^{-\lambda\tau}) \ln y_0 \\
&+ (1 - e^{-\lambda\tau}) \frac{\alpha}{1 - \alpha - \beta - \theta} (\ln k_t - \ln(n_t + \mu)) \\
&+ (1 - e^{-\lambda\tau}) \frac{\beta}{1 - \alpha - \beta - \theta} (\ln h_t - \ln(n_t + \mu)) \\
&+ (1 - e^{-\lambda\tau}) \frac{\theta}{1 - \alpha - \beta - \theta} (\ln e_t - \ln(n_t + \mu)) + \varepsilon_t
\end{aligned}$$

and compare the estimate of μ with the imposed constraint ($g + \delta = 0.05$).

The analysis is further related to the studies of health care expenditure where GDP per capita appears to be the main factor determining the level of expenditure on health care

$$\begin{aligned}
(16) \quad \ln e_t - \ln e_0 &= \beta_0(1 - e^{-\lambda\tau}) - (1 - e^{-\lambda\tau}) \ln e_0 \\
&+ (1 - e^{-\lambda\tau}) \frac{\alpha}{1 - \alpha - \beta - \theta} (\ln k_t - \ln(n_t + g + \delta)) \\
&+ (1 - e^{-\lambda\tau}) \frac{\beta}{1 - \alpha - \beta - \theta} (\ln h_t - \ln(n_t + g + \delta)) \\
&+ (1 - e^{-\lambda\tau}) \frac{\theta}{1 - \alpha - \beta - \theta} (\ln y_t - \ln(n_t + g + \delta)) + \varepsilon_t
\end{aligned}$$

or alternatively, one can estimate μ from the following relation

$$\begin{aligned}
(17) \quad \ln e_t - \ln e_0 &= \beta_0(1 - e^{-\lambda\tau}) - (1 - e^{-\lambda\tau}) \ln e_0 \\
&+ (1 - e^{-\lambda\tau}) \frac{\alpha}{1 - \alpha - \beta - \theta} (\ln k_t - \ln(n_t + \mu)) \\
&+ (1 - e^{-\lambda\tau}) \frac{\beta}{1 - \alpha - \beta - \theta} (\ln h_t - \ln(n_t + \mu)) \\
&+ (1 - e^{-\lambda\tau}) \frac{\theta}{1 - \alpha - \beta - \theta} (\ln y_t - \ln(n_t + \mu)) + \varepsilon_t
\end{aligned}$$

and test it against the imposed constraint. The slope is exogenously determined by the rate of technical progress, while the intercept reflects the effects of the rate of population growth and investments in physical and human and GDP growth. Investment in human and physical capital and GDP growth are expected to increase the level of health care expenditure per capita, while high population growth lowers investment in health.

The issues of causality relationship between GDP and HCE is examined by regressing the log of GDP ($\ln y$) and HCE ($\ln e$) on their past values and testing for their joint significance as follows

$$(18) \quad \begin{aligned} \ln y_t &= \alpha_0 + \sum_{\tau=1}^n \alpha_\tau^1 \ln y_{t-\tau} + \sum_{\tau=1}^m \alpha_\tau^2 \ln e_{t-\tau} + \zeta_t \\ \ln e_t &= \beta_0 + \sum_{\tau=1}^n \beta_\tau^1 \ln e_{t-\tau} + \sum_{\tau=1}^m \beta_\tau^2 \ln y_{t-\tau} + \xi_t \end{aligned}$$

where non-zero values of α_τ^2 and β_τ^2 are indications of causality relationships between the two variables.

Using the parameter estimates from (13)-(17) estimates of the rate of convergence are obtained. The rate of convergence varies across sample countries as a result of growth in population. The sensitivity of the speed of convergence is examined with respect to the assumption of constant depreciation and technological progress. A regression of the speed of convergence on variables determining the rate of convergence is performed

$$(19) \quad \lambda_i = \gamma_0 + \sum_j \gamma_j x_{ij} + \varsigma_i$$

where x_j is a vector of variables characterizing the health care system of the sample countries.

3. THE DATA

The data are obtained from the Penn World Data (PWD)⁴ known as Summers-Heston (1991) data. The current data set is an updated version of the previous versions called Mark 5.0 and Mark 5.6. PWD allows access to online statistics covering 29 key variables on 151 major world economies for which data are available. The data are annual and cover the period of 1950-1992. The information includes population, various definitions of GDP, private and public consumption, investment, different components of capital, exchange rate, standard living index and measures of openness.

The PWD data was further completed with information on human capital obtained from the Barro and Lee (1996)⁵ and health care expenditure extracted from the OECD Health Data File (OECD, 1997). The Barro and Lee data contains information on educational attainment at various levels for male and female populations. The sample includes 129 countries observed quinquennially during 1960-1990. One hundred and fifteen of the sample countries are found in the sample countries of the Summers and Heston (1991).

For purposes of comparability of results with those of previous studies, the sample in this paper is restricted to OECD countries. Other factors determining the sample size are availability of information on human capital, health care expenditure and GDP per capita. The final sample used in this paper is similar to Summers and

⁴ <http://bized.ac.uk/dataserv/pennhome.htm>

⁵ <http://www.worldbank.org/html/prdmg/grthweb/ddbarle2.htm>

Heston (1988), Mankiw, Romer and Well (1992), and Islam (1995) subsample of OECD countries⁶ but updated to 1992.

The variables used are GDP (y), health care expenditure (e), population growth (n), investment in physical (k) and human (h) capital. The GDP is defined as real GDP per worker (RGDPW) in 1985 international prices. The investment in physical capital is measured as the percentage real investment share of the GDP in 1985 international prices. Population growth is defined as the year to year growth rate of population that includes all ages. Human capital is defined as average schooling years in total population above the age of 25. Following a number of studies the rate technological progress and depreciation ($g + \delta$) is taken to be constant across countries and over time. Hence, it is assumed to be equal to 0.05 in the constrained growth models, while it is estimated as a constant parameter in the unconstrained models.

Considering the period of 1970-1992 we have five data points for all of the sample countries. The growth rate in GDP, population growth, investment and human capital are with the exception of the last period measured as 5 years averages: 1970-74, 1975-79, 1980-84, 1985-89, and 1990-92. Health care expenditure is measured in per capita terms at constant 1990 international prices. For the transformation of HCE the national GDP price index and purchasing power parities for the base year 1990 as deflators are used.

A number of system variables are considered as determinants of speed of convergence. These variables represent the sample countries' health care system in 1991. Public reimbursement (PUBLREIMB), public contract (PUBLCONT), public integrated (PUBLINTEG) and gate keeper (GATEKEEP) are defined as 1 indicating whether a country has any of the above systems as dominant means of remunerations in the in-patient care, 0 otherwise. The data is a balanced panel data where countries are observed consecutively. A summary statistics of the data are found in the Table 1.

4. EMPIRICAL RESULTS

4.1 Specification Tests and Parameter Estimates

The GDP model (10) with various degrees of augmentation (human and health capital) was estimated. First, the model was estimated separately using five years average cross-sectional 1970 and 1990 data. This approach has the disadvantages of not providing any information about the dynamic process of convergence. Second, the model was estimated using 1970 as initial income to capture the dynamics of convergence.⁷ Third, the models were estimated both with and without imposing the constraint that the coefficients on investment and population growth are equal in magnitude but opposite signs. Fourth, later models are estimated assuming that the sum ($g + \delta$) is equal to the constant 0.05 (equation 14) or alternatively estimate the

⁶ Our sample differs from the sample used by Mankiw, Romer and Well (1992) and Islam (1995) by Iceland, Korea and Luxembourg are not observed and thus were excluded.

⁷ For standard and heterogeneous panel estimations of the rate of convergence see Islam (1995) and Lee, Pesaran and Smith (1997).

sum from the data (equation 15). Fifth, the procedure is applied to the HCE class of models (equations 16 and 17) again with various degrees of augmentation (human capital and GDP per capita) and restrictions imposed. Sixth, the issue of causality between GDP and HCE is examined (equation 18). Finally, the rate of convergence is regressed on the health care system variables (equation 19).

The models are estimated using both linear and iterative non-linear regression methods. To conserve space only estimation results from the general versions of the models are reported here.

4.2 The GDP per Capita Models

The results for the GDP models are presented in Table 2. Makiew et al. (1992) found that including human capital accumulation to the Solow model can potentially alter either the theoretical or empirical analysis of economic growth. At the theoretical level the changes are related to the nature of the growth process which at the empirical level, human capital can be considered as an omitted variable. Leaving out human capital affects the coefficients on investment and population growth. Analogously a disaggregation of capital into physical, human and health capital investments has both theoretical and empirical implications.

The coefficient of initial value of GDP is negative indicating positive relation between growth and the initial distance from the steady state. The coefficient of investment in capital is positive showing that growth is an increasing function of saving. The coefficient of human capital is unexpectedly negative and insignificant. Human capital can be accumulated through improvements in health (see Muskin, 1962) and capacity of work suggesting that improved nutrition and health status affect labor productivity positively (see e.g. Strauss and Thomas, 1995). Our results are in line with those found by Knowles and Owen (1995), who found that Human capital is not significant in any estimated equation when health capital and base year income per capita are included in MRW model. Moreover, the results are also sensitive to the influence of some sample countries. The insignificance effects of the human capital held even when the data was split into developed and less developed country subsamples.

Investment in health care is found to have positive effects on growth, while population growth has negative effects. These results are consistent with results found in previous studies. It should be noted that the proxy for investment in health capital in our study is the proportion of income spent on health, while Knowles and Owen used life expectancy to proxy health capital. Their choice of proxy was due to the fact that health spending can be associated with different actual outcomes across countries. The OECD countries are relatively homogenous group of countries. Hence, health care expenditure per capita is an appropriate proxy for investment in health.

4.3 The HCE per Capita Models

The results of the HCE models are presented in Table 3. A simple specification similar to the one appearing in the health literature show that GDP per capita is a major determinant of the investment in health (for a recent survey see,

Gerdtham and Jönsson, 1998). The coefficient of GDP is found to be positive and significant. GDP alone explains more than 85% of the total variations in the health care expenditure per capita in 1970. The corresponding explanatory power is 89% in 1990. The coefficient is larger than one indicating that larger fraction of GDP growth is invested in health care. There is decline in the size of coefficient of $\ln\text{GDP}$ when one uses the change in $\ln\text{HCE}$ between 1970 and 1990.

The coefficient of initial level of HCE is negative indicating positive relation between growth and the initial distance to the steady state level of HCE. The effect of population growth on HCE is negative. The effect of investment in human capital on HCE is unexpectedly negative and in most model specifications statistically insignificant. Unlike the GDP model, an augmentation of the HCE model to incorporate investment in physical and human capital and population growth, does not alter the sign and significance of the GDP coefficient. However, the model performance is not improved and most of the extra parameters are found to be insignificant.

4.4 The Speed of Convergence

The estimated rate of convergence from the GDP and HCE models are presented in Table 1 and at the bottom panels of Table 2 and 3, respectively. The GDP model predicts convergence to the steady state level of income per capita at annual rate of 3.7%. The corresponding rate where the technical change and depreciation rates are estimated, is 5.2%. Restrictions imposed on the size of $(\mu = g + \delta)$ results in underestimation of the speed of convergence. The parameter estimates are constant but the growth rate of population is country specific. Hence, the rate of convergence becomes country specific. Depending on specification of μ , the percentage convergence rate varies in the intervals 3.1-4.6% and 4.6-6.1%, respectively. A summary statistics of rate of convergence is given in Table 1.

The HCE model predicts a lower rate of convergence to the steady state level of health care expenditure. The rate is 2.5% and 2.7% depending on whether one imposes any restriction of the size of $(g + \delta)$ or not. Again the sum of $(g + \delta)$ is found to be statistically different from the 0.05. The lower convergence rate in investment in health compared to income per capita can be explained by the difference in the countries' preferences in public spending. The range of variations in the convergence rate in HCE models is much smaller, 2.1-3.1% and 2.4-3.2%, respectively.

It should be noted that results from cross-sectional analysis of growth and convergence should be treated with caution. A number of recent studies (Solow (1994), Lee, Pesaran and Smith (1997), Knowles and Owen (1997), among others) show potential sources of bias in cross-sectional, pooled and heterogeneous panel estimation of convergence coefficient. The magnitude of bias can be larger in the GDP model due to simultaneity of health and human capital inputs. Their significance might reflect the ability of countries with faster growth in devoting more resources to investment in health care and education.

4.5 Causality Relationship between GDP and HCE

The augmentation of HCE in the growth model and the specification of health care expenditure in the health literature as a function of GDP imply the issue of causality to be important in this respect. Granger's concept of causality is that a variable x causes a variable y if taking account of past values of x leads to improved predictions for y , all other things being equal. The most common approach to answer the question of relationship between x and y is to regress x on y and test the coefficient of y for significance. In the current case it is important to establish and test for the direction of causality. Using the relation in (18) for the test of causality between GDP and HCE the values of m and n were set to 1 respectively. The choice of minimum lag structure was due to few periods of observations.

The test results presented in Table 4 indicate the presence of a unidirectional causality from HCE to GDP. Contrary to the case in the health care literature this is interpreted as a rejection of the hypothesis that causality is unidirectional from GDP to HCE. Hence, the GDP and HCE models can be estimated as single equations as is done in this study. Although the sample countries are very homogeneous, the data set is very small and the results should be interpreted with caution.

4.6 Determinants of the Rate of Convergence

The results from regression of the rate of convergence obtained from the HCE model on a number of health care system variables that are considered as determinants of speed of convergence are reported in Table 5. These variables represent the sample countries' health care system in 1991. Public contract has positive effect on the rate of convergence while gate keeper system has negative effect compared to the reference group of public reimbursement. The coefficient of public integrated is found to be insignificant. Austria, Canada, Germany and Netherlands have mixed public contract and gate keeper systems with both positive and negative effects on the rate of convergence. Most of the sample countries including Denmark, Finland, Greece, Ireland, Italy, New Zealand, Norway, Portugal, Spain, Sweden, Turkey, and UK have a public integrated system with no effects on the convergence rate. The reference group with public reimbursement health care system includes Australia, Belgium, France, Japan, Switzerland and USA. The two models which differ by the size of μ give identical results.

5. SUMMARY AND CONCLUSIONS

This paper examines conditional convergence of OECD countries in gross domestic product (GDP) and health care expenditure (HCE) per capita for the period 1970-1992. It presents estimation of the augmented Solow model to explain variation in output per capita across countries. The variation is due to different steady state growth paths resulting from differences in the countries savings rate, education, and population growth.

This paper is an extension of the MRW paper to health care expenditure. The analysis is further related to the studies of health care expenditure where GDP per

capita appears to be the main factor determining the level of expenditure on health care. The causality tests show a unidirectional causality from HCE to GDP but not from GDP to HCE. The result indicates that OECD countries converge at 3.7% per year to their steady state of incomes per capita. HCE has positive effects on the economic growth and the speed of convergence. An inclusion of health care expenditure in the growth model results in an insignificant coefficient of human capital.

The speed of convergence is also found to be sensitive to whether one imposes a constant or estimates the depreciation and technological growth components. With no restrictions imposed the convergence rate is 5.2%. Considering the rate of convergence in the HCE model the results show that OECD countries converge at 2.5%-2.7% to their steady state level of HCE per capita. Again the assumption of constant rate of depreciation and technological growth underestimates the rate of convergence.

In the latter models a regression of the speed of convergence on variables determining the rate of convergence shows a close link to the variables characterizing the health care system of sample countries. Public contract has a positive effect on the rate of convergence while gate keeper system has a negative effect compared to the reference group of public reimbursement. The system of public integrated is found not to be significantly different from the public reimbursement.

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Table 1. Summary statistics of the OECD data, 1970-1992.

Variable	Definition	Mean	Std Dev	Minimum	Maximum
<i>1.A Causality test data set (N=22, T=5, NT=110):</i>					
PERIOD	1970-74, 75-79, 80-84, 85-89, 90-92	80.00	7.10	70.00	90.00
GDP	GDP per capita, 5 years averages	22896.78	6715.31	4841.00	36771.00
HCE	Health care expenditure per capita	942.18	468.51	75.78	2931.08
<i>1.B Growth Data (N=22):</i>					
GDP70	GDP/cap, average 1970-74, 1985 int. prices	19243.68	6396.82	4841.00	30468.00
GDP90	GDP/cap, average 1990-92, 1985 int. prices	26944.00	6324.61	8632.00	36771.00
HCE70	HCE/cap, average 1970-74, 1985 Int. prices	606.96	273.23	75.78	1253.70
HCE90	HCE/cap, average 1990-92, 1985 Int. prices	1268.47	554.13	177.07	2931.08
INV	Real Inv. % share of GDP, 1985 int. Prices	23.63	5.22	16.30	38.17
HUMAN	Average schooling year, population above 25	8.38	2.20	3.35	12.00
DPOP	Average annual population growth, 1970-92	0.77	0.59	-0.15	2.13
DGDP	Change in GDP/cap (GDP90-GDP70)	1.47	0.26	1.05	1.98
DHCE	Change in HCE per capita (HCE90-HCE70)	2.17	0.42	1.52	3.34
<i>1.C Health Care System Variables (N=22):</i>					
PUBREIMB	Public reimbursement system	0.27	0.46	0.00	1.00
PUBCONTR	Public contract system	0.18	0.39	0.00	1.00
PUBINTEG	Public integrated system	0.55	0.51	0.00	1.00
GATEKEEP	Gate keeper	0.55	0.51	0.00	1.00
<i>1.D Rate of Convergence ($\lambda = (1 - \alpha - \beta)(n + g + \delta)$) Data, (N=22):</i>					
λ_GDP1	Convergence rate GDP 1990-1970, $\mu = 0.05$	0.037	0.004	0.031	0.046
λ_GDP2	Convergence rate GDP 1990-1970, $\hat{\mu} = 0.0716$	0.052	0.004	0.046	0.061
λ_HCE1	Convergence rate HCE 1990-1970, $\mu = 0.05$	0.025	0.003	0.021	0.031
λ_HCE2	Convergence rate HCE 1990-1970, $\hat{\mu} = 0.0628$	0.027	0.002	0.024	0.032

List of OECD countries included:

Australia, Austria, Belgium, Canada, Denmark, Finland, France, Germany, Greece, Ireland, Italy, Japan, Netherlands, New Zealand, Norway, Portugal, Spain, Sweden, Switzerland, Turkey, United Kingdom, United States of America.

22 OECD countries observed during 1970-1992

1970 is average of 1970-74

1990 is average of 1990-92

Table 2. Augmented static and dynamic GDP per capita models.

Models	Static Models				Dynamic Models				
	lnGDP70		lnGDP90		Restricted μ		Unrestricted μ		
Dependent variable	Estimate	Std error	Estimate	Std error	Estimate	Std error	Estimate	Std error	
<i>2.A Linear Models:</i>									
B0	7.1680a	1.2405	6.2875a	0.7695	2.9594a	0.9054	-	-	
lnINV	-0.1734	0.1965	0.0510	0.1219	0.1315	0.0852	-	-	
lnPOP	0.5812	0.4082	-0.1003	0.2532	-0.3702c	0.1831	-	-	
lnHUM	0.3546c	0.1740	0.1444	0.1079	-0.0203	0.0823	-	-	
lnHCE	0.5822a	0.1045	0.4454a	0.0648	0.1751b	0.0745	-	-	
lnGDP70	-	-	-	-	-0.5357a	0.1028	-	-	
R2 adjusted	0.8399		0.8811		0.8167		-		
F-test lnHE=0	31.0194	0.0001	47.1836	0.0001	5.5246	0.0319	-	-	
<i>2.B Non-linear Models:</i>									
B0	4.8668a	1.2303	5.3627a	0.6762	5.7541a	0.7532	5.3628a	0.6404	
lnINV	α	-0.2253	0.1766	-0.0130	0.0823	0.1544	0.0953	0.1258	0.0865
lnHUM	β	0.2504b	0.1213	0.1017	0.0710	-0.0154	0.0913	0.0070	0.0826
lnHCE-lnPOP	θ	0.3319a	0.1126	0.2690a	0.0579	0.2194a	0.0624	0.2484a	0.0570
	λ	-	-	-	-	0.0409a	0.0090	-	-
(g + δ)	μ	-	-	-	-	-	-	0.0766a	0.0219
R2 adjusted	0.7598		0.8600		0.8260		0.7943		
Calculated λ	-	-	-	-	0.0370	0.0038	0.0522	0.0037	

Significant at the <1% (a), 1-5% (b), 5-10% (c) levels of significance.

Table 3. Augmented static and dynamic HCE per capita models.

Models	Static Models				Dynamic Models				
	lnHCE70		lnHCE90		Restricted μ		Unrestricted μ		
Dependent variable	Estimate	Std error	Estimate	Std error	Estimate	Std error	Estimate	Std error	
<i>3.A1 Restricted Linear Models:</i>									
B0	-12.5931a	1.6965	-10.6131a	1.3697	-4.0507c	2.0799	-	-	
lnGDP	1.8557a	0.1668	1.7354a	0.1347	0.7684b	0.2830	-	-	
lnHCE70	-	-	-	-	-0.4789a	0.1415	-	-	
R2 adjusted	0.8539		0.8871		0.3388		-	-	
<i>3.A2 Unrestricted Linear Models:</i>									
B0	-8.8904a	2.0869	-10.7171a	2.0142	-4.1441b	1.2256	-	-	
lnINV	-0.0270	0.2440	0.0496	0.2355	0.0696	0.1531	-	-	
lnPOP	0.2486	0.5040	-0.2298	0.4865	-0.4136	0.3418	-	-	
lnHUM	0.5518b	0.2256	0.0735	0.2177	-0.3345c	0.1385	-	-	
lnGDP	1.4565a	0.2490	1.6506a	0.2403	0.5738b	0.3626	-	-	
lnHCE70	-	-	-	-	-0.2607	0.2600	-	-	
R2 adjusted	0.8817		0.8694		0.4532		-	-	
F-test lnGDP=0	34.2256a	0.0001	47.1836a	0.0001	4.3092b	0.0544	-	-	
<i>3.B Non-linear Models:</i>									
B0	-10.0906a	2.8038	-11.5480a	2.3644	-15.2304a	4.9300	-16.8919a	3.9022	
lnINV	α	-0.0926	0.1331	-0.0389	0.1120	0.1144	0.2299	0.0811	0.1479
lnHUM	β	0.2845b	0.1107	0.0801	0.0991	-0.5040	0.4294	-0.2868	0.2281
lnGDP-lnPOP	θ	0.4210a	0.1356	0.5585a	0.1249	0.9515b	0.3874	0.8203a	0.2325
	λ	-	-	-	-	0.0167b	0.0075	-	-
($g + \delta$)	μ	-	-	-	-	-	-	0.0628b	0.0274
R2 adjusted	0.7821		0.8163		0.4839		0.4197		
Calculated λ	-	-	-	-	0.0253	0.0026	0.0272	0.0023	

Table 4. Test for causality between GDP and Health Care Expenditure per capita.

Model	GDP per capita		HCE per capita		
Parameter	Estimate	Std error	Parameter	Estimate	Std error
<i>GDP Model 1:</i>			<i>HCE Model 1:</i>		
Intercept	6.4797a	0.1802	Intercept	-8.0279a	1.4132
lnHCEt-1	0.5340a	0.0267	lnGDPT-1	1.4929a	0.0891
R2adjusted	0.8597		R2adj	0.8113	
F-test	399.3190a	0.0001	F-test	33.5330a	0.0001
<i>GDP Model 2:</i>			<i>HCE Model 2:</i>		
Intercept	1.6422a	0.3704	Intercept	0.6813	0.4622
lnGDPT-1	0.7708a	0.0572	lnGDPT-1	-0.0679	0.0713
lnHCEt-1	0.1090a	0.0343	lnHCEt-1	1.0224a	0.0429
R2adjusted	0.9633		R2 adjusted	0.9809	
F-test	854.8100a	0.0001	F-test	1669.0380a	0.0001
F-test lnGDPT-1=0	181.8590a	0.0001	F-test lnGDPT-1=0	0.9082	0.3442
F-test lnHCEt-1=0	10.0658a	0.0023	F-test lnHCEt-1=0	568.8330a	0.0001

Significant at the <1% (a), 1-5% (b), 5-10% (c) levels of significance.

Conclusions on causality test:

HCE causes GDP but GDP does not cause HCE.

Table 5. Determinants of rate of convergence obtained using HCE per capita (lnHCE90-lnHCE70) models.

	Restricted $\mu, \mu = 0.05$		Unrestricted $\mu, \hat{\mu} = 0.0628$	
	Estimate	Std error	Estimate	Std error
B0	0.0253a	0.0008	0.0272a	0.0007
Public contract	0.0060a	0.0018	0.0053a	0.0053
Public integrated	0.0019	0.0013	0.0017	0.0012
Gate keeper	-0.0039a	0.0013	-0.0034a	0.0011
R2 adjusted	0.3684		0.3684	
F-test	5.0830a	0.0101	5.0830a	0.0101
<i>Dependent variable:</i>				
Calculated λ	0.0253	0.0026	0.0272	0.0023

Public reimbursement is the reference system.

Significant at the <1% (a), 1-5% (b), 5-10% (c) levels of significance.