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# LABOUR FORCE PARTICIPATION IN THE EURO AREA A COHORT BASED ANALYSIS 

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Ramón Gómez-Salvador
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LABOUR FORCE PARTICIPATION
IN THE EURO AREA

# A COHORT BASED ANALYSIS' 

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#### Abstract

We use a cohort based model to analyse determinants of labour force participation for disaggregated groups of workers in the euro area and the five largest euro area countries. The model captures age and cohort effects as indicators of (unobserved) determinants of participation behaviour. We use these effects and observed determinants to construct trends and projections of labour supply. Our results suggest that age and cohort effects can account for a substantial part of the recent increase in participation. Cohort effects are particularly relevant for women with those born in the late 1960s and early 1970s more likely to participate over the life-cycle. There is substantial variation in the estimated age and cohort effects across countries. Looking forward, positive cohort effects for women are not large enough to compensate for the downward impact of population ageing on participation rates in the euro area.


Keywords: labour force participation, cohort analysis, labour market institutions
JEL Classification: J11, J21

## Non-technical summary

The euro area labour force participation rate, defined as the ratio between the labour force and the working age population, has increased from below $65 \%$ in the early 1980 s to $70.9 \%$ in 2007. The participation rate of females in the euro area has increased by more than 15 percentage points over this time period, to $63.3 \%$ in 2007, compared to the participation rate of $78.6 \%$ for males.

The large increase in the propensity of the euro area population to work or to search for and to be available for jobs has been one of the main drivers of the substantial increase in euro area labour supply that has accelerated since the mid-1990s. This has significantly reduced the gap in the use of labour input between the euro area and the United States, and has made a substantial positive contribution to output growth and welfare in the euro area. A number of factors could have contributed to this increase, including reforms in the labour market, changes in cultural attitudes towards work (particularly for women), as well as demographic factors. Looking forward, demographic factors will become less favourable with population ageing increasing the importance of positive participation trends within age and gender groups in sustaining potential growth in the euro area.

We use a cohort based model of labour force participation to analyse determinants of participation for disaggregated groups of workers in the euro area and the five largest euro area countries (Germany, France, Italy, Spain and the Netherlands). The model is used to decompose the evolution of time-series of age-specific participation rates into the impact of the business cycle, observed structural determinants of participation and other unobserved determinants captured by age and cohort specific effects.

We find that analysing participation behaviour both between (age and gender effects) and within (cohort effects) detailed age and gender groups is particularly useful for modelling trends in euro area aggregate participation rates and projecting them forward. Our results suggest that age and cohort effects can explain a substantial part of the recent increase in labour force participation rates in the euro area, although not the surge since early 2000s. Cohort effects are particularly relevant for women, with those born in the 1920s and 1930s less likely and those born in the late 1960s and early 1970s more likely to participate in the labour market over the life-cycle. There is substantial variation in cohort effects across the five largest euro area countries that we analyse. Depending on the country, the estimated cohort profiles suggest an increase of 10 to 30 percentage points in female
participation rates. We also find that a number of observed determinants, such as labour taxes, union density, unemployment benefits and the average number of children have had an impact on labour force participation rates, although the specific impact varies across age and gender groups and countries. Looking forward, while they continue to provide some upward support to participation rates of women in the euro area, positive cohort effects are not large enough to compensate for the downward impact of population ageing on labour force participation rates in the euro area.

## 1. Introduction

The euro area labour force participation rate, defined as the ratio between the labour force and the working age population, has increased from below $65 \%$ in the early 1980 s to $70.9 \%$ in 2007. The increase in the propensity to work or to search for and to be available for jobs has been the main driver of the substantial increase in euro area labour supply that has accelerated since the mid-1990s. However, the overall increase reflects substantial heterogeneity in the evolution of participation behaviour across population groups and across euro area countries. The participation rate of females in the euro area has increased by more than 15 percentage points over this time period, to $63.3 \%$ in 2007, compared to the participation rate of $78.6 \%$ for males (see Figure 1). The participation rate of the young (15-24 years old) declined markedly until the mid-90s and has stabilised to around $45 \%$ in the last decade, whereas, following a long period of stable participation rates, the participation rate of those 55-64 years old increased markedly in the last few years. Also, the extent of the increase in participation and its composition across worker groups varies across euro area countries, suggesting an important role for cross-country heterogeneity in the underlying factors that determine individual labour supply decisions. A number of factors could have contributed to the overall increase in participation: robust, employment intensive economic growth (in particular from the mid-1990s onwards), reforms targeted at groups with lower attachment to the labour market, changes in cultural attitudes towards work (particularly for women), as well as demographic factors, such as the larger share of the population in prime working age.

We use harmonised data from the EU Labour Force Survey (LFS) and a cohort-based model of labour force participation to analyse determinants of participation in the euro area and the five largest euro area countries (Germany, France, Italy, Spain and the Netherlands) over the last few decades. We refer to two euro area aggregates. The first (EA12) consists of the euro area 12 countries before Slovenia, and later Malta, Cyprus and Slovakia, entered the euro area and the second (EA5) an aggregation of the five largest euro area countries. The cohort-based model is used to decompose time-series of age-specific participation rates in euro area countries into the impact of the business cycle, observed structural determinants of participation (such as labour market institutions) and unobserved determinants captured by age and birth-year specific (i.e. cohort) effects.

The age and cohort effects are derived from the age-participation profile. The propensity to participate evolves over the life-cycle, as reflected in an inverted u-shape age-participation profile participation rates of younger and older people change from one age group to another, whereas the substantially higher participation rates for those in prime working age show a relatively flat profile between the ages of 30 and 50 . The age effects in the model capture this feature of the underlying age-participation profile. At the same time, the age-participation profile is continuously shifting. The cohort based model captures parallel shifts in the profile that are specific to a birth-year through the unobserved cohort effects. While cohort effects generally encompass any factor associated with a particular birth year, they are likely to reflect the impact of individual participation choices made early on in life (for example choices relating to starting a family, maternity leave and/or education) that persist throughout the life-cycle. They may also reflect crowding-out effects or slowly evolving preferences, cultural factors or institutions. ${ }^{2}$ Controlling for business-cycle effects, we first estimate age and cohort effects for the euro area aggregate (EA12) and individual euro area countries.

While cohort effects explain shifts in the age-participation profile, potential changes in the shape of this profile are captured through observed time-varying determinants, such as demographic trends and changing labour market institutions. We use the model with observed determinants to explain changes in trend participation rates over time and also to project them forward in the five largest euro area countries. We then aggregate these country trends and projections for the euro area (EA5). Projections that take age and cohort effects and the changing population structure into account provide a useful benchmark scenario for future labour supply in the euro area. In particular, a cohort based model takes into account the extent of the pass-through of participation behaviour from the young cohorts to the oldest cohorts. Looking forward, demographic factors will become less favourable with population ageing increasing the importance of positive participation trends within age and gender groups in sustaining potential growth in the euro area. As we estimate the model separately for individual euro area countries and aggregate the results, the results for the euro area also fully incorporate heterogeneity across countries.

Our paper is related to two main strands of literature. First, our main focus is on accurately estimating trends in participation based on both observed determinants and the unobserved age and cohort effects in the euro area. For this purpose we use a modified version of the cohort-based model

[^0]presented in Fallick and Pingle (2007) and applied in Aaronson et al. (2006) to data for the United States. By simultaneously estimating participation equations for single ages for each gender and taking advantage of cross-equation restrictions the model provides a detailed account of the role of age and cohort effects in explaining movements in the aggregate participation rate. Fallick and Pingle (2007) find that these effects provide additional insights compared to time series based trend/cycle decompositions. For example, they find that the levelling off of the increase in the propensity to participate at cohorts born around 1950 suggests that increased labour market attachment is less likely to support an increase in the participation of females in the United States. We are not aware of a cross-country study of European participation rates that accounts for these features. Euwals et al. (2007) find using microdata that cohort effects have played an important role in explaining the increase in female participation from 1992 to 2004 in the Netherlands. Fitzenberger et al. (2004) use an alternative age, cohort, and period accounting model to study participation and employment in Germany and find significant cohort effects for females. ${ }^{3}$

Second, a number of studies have documented the impact of labour market institutions on unemployment and employment in European countries (for a recent contribution and review of the literature, see Bassanini and Duval, 2006 and Bertola et al., 2007 for age-group specific analysis). Participation decisions have received less attention in this context. Blöndal and Scarpetta (1999) and Duval (2006) focus on older workers and their retirement decisions and Jaumotte (2003) on females. Genre et al. (2005 and 2008) focus on group specific participation rates in European countries. Using annual data for a panel of European Union countries, they estimate participation equations for age and gender groups in order to identify the impact of institutions in participation decisions. They find that labour market institutions indeed matter for labour supply: higher union density, more employment protection and more generous unemployment benefits lower participation rates. Genre et al. (2008) also find using lagged participation rate as a proxy, that a common (across countries) cohort effect is an important element for understanding participation rates of older women (those between 55 and 64) in European countries. We add to these studies by considering disaggregated groups and by evaluating age and cohort effects and possible observed determinants of participation in the same model. Instead of the cross-country focus of most previous studies, we exploit the time series dimension of the data and incorporate the impact of a broader set of factors through the unobserved age and cohort effects.

[^1]We find that analysing participation behaviour both between (age and gender effects) and within (cohort effects) detailed age and gender groups is particularly useful for modelling trends in euro area aggregate participation rates and projecting them forward. Our results suggest that age and cohort effects can explain a substantial part of the recent increase in labour force participation rates in the euro area, although not the surge since early 2000s. Cohort effects are particularly relevant for women, with those born in the 1920s and 1930s less likely and those born in the late 1960s and early 1970s more likely to participate in the labour market over the life-cycle. There is substantial variation in cohort effects across the five largest euro area countries that we analyse. Depending on the country, the estimated cohort profiles suggest an increase of 10 to 30 percentage points in female participation rates. We also find that a number of observed determinants, such as labour taxes, union density, unemployment benefits and the average number of children have had an impact on labour force participation rates, although the specific impact varies across age and gender groups and countries. Looking forward, while they continue to provide some upward support to participation rates of women in the euro area, positive cohort effects are not large enough to compensate for the downward impact of population ageing on labour force participation rates in the euro area.

The rest of the paper is organized as follows. We describe recent changes in euro area labour force participation by age and gender groups in Section 2. In Section 3 we describe the data and the cohort based model of participation. In Section 4 we present results from the model in three parts. We first illustrate the role of estimated age and cohort effects in determining participation. Second, we analyse the impact time-varying observed determinants of participation within a full model. Third, we present projections for participation rates up to 2030 based on the model and compare them with alternative scenarios. Finally, we summarise our results and conclude in Section 5.

## 2. Labour force participation by age and gender in the euro area

Participation behaviour and its determinants vary systematically by age and gender and changes in group-specific participation rates translate into the aggregate through an evolving population structure. ${ }^{4}$ As a result, analysing participation behaviour of detailed age and gender groups is essential for understanding aggregate participation developments. We therefore, first

[^2]describe recent developments in group-specific participation rates in more detail. We then assess the impact of the changing population structure on the aggregate participation rate.

Labour supply and participation rates evolve substantially over the life-cycle, tracing a wellknown overall inverted u-shape profile of participation rates that peaks around the prime working age. Figure 2 illustrates these profiles for euro area males and females in 2007. The participation rates of younger workers (those below 25) and older workers (those above 50) change substantially from one age group to another, whereas the substantially higher participation rates for those in prime working age show a relatively flat profile between the ages of 30 and 50 . The age participation profile for females is always below, peaking earlier, than the profile for men, and showing a smaller gap at younger than at older ages. These age participation profiles for males and females are continuously evolving as a result of changes both between and within age groups. Overall, since 1983 the female profile has been lifted up for those above 25 years old and has in particular tilted upwards for older women, with more women staying in the labour market after child-bearing. At the same time, the participation rates for the youngest women have declined. By contrast, a change in the age-participation profile of males is only visible for the youngest males, whose participation rate also declined somewhat.

Focussing on changes in group-specific participation rates helps paint a more precise picture of how the age-participation profiles in the euro area have evolved over time. Figure 3 plots the overall change in participation rates in two time periods, 1983-1995 and 1995-2007, for each single age. These two periods are comparable both in terms of length and in terms of economic developments (i.e. the business cycle). Starting with males, the change in trend observed in the mid1990s is mostly accounted for by the youngest age groups. The participation rates of those between 15-24 years old switched from declines between 5 and 25 percentage points in the pre-1995 period to a stabilisation and even small increases since then. At the same time, the participation rates of those between 60-64 years old switched from a decline of around 10 pp on average to an increase slightly below 10 pp on average (see the upper panel of Figure 3). This pattern is also shared by those between 55 and 59 years old. Moving to females, the ongoing increase in female participation over the two time periods considered appears to result mainly from the sustained increase recorded for those between 25 and 58 years old (see lower panel of Figure 3). This has been supported, since the mid-1990s by a much smaller decline in the participation rates of the youngest and by an increase for the oldest. For both males and females, the increases in participation rates for the youngest and oldest
workers may be related to the impact of labour market reforms that have focused on groups with a weaker attachment to the labour market.

Figure 3 also hints at possible birth year or cohort specific effects in participation behaviour. In particular, focussing on changes in female participation rates, the hump-shaped pattern of an increase in participation for those between 25 and 58 years old has shifted towards older age groups over time. For example, while the age group recording the highest increase in participation in 198395 was between 40 and 45 years old, for the next twelve year period (1995-2007) the highest increase took place for those between 52 and 56 years old. This pattern is reminiscent of cohort specific participation effects, i.e. female participation behaviour for a particular cohort persists over time. In terms of the age-participation profile, these cohort effects shift the whole profile up, as discussed in the previous section.

At the country level, we can observe somewhat different participation profiles in 2007, pointing to important heterogeneity in participation behaviour. For instance, the gap between male and female participation rates is more substantial in Italy and Spain than in France, Germany and the Netherlands, especially for those in prime age. While the participation rates of the youngest age groups are comparable between most countries (at levels around 10-30\%), they are substantially higher in the Netherlands (at around 60\%). Finally, for the oldest age groups (60-64 years old), differences are mostly concentrated in female participation, which varies from $10 \%$ in Italy to around $30 \%$ in Germany, while for males, participation rates are generally between $30-50 \%$, with the only exception of France (below 20\%).

In addition to changes in group-specific participation behaviour, the composition of the euro area population has changed over time. The impact of the change in composition can be measured by applying the change in the population composition between the two periods to the participation rates of the first period, by age and gender groups. Table 1 shows that the positive effect of the population composition observed in 1983-1995, around 1.3 percentage point, declined significantly in 19952007. This recent fall in the contribution of the population effect results from the decline in the share of the prime age population in favour of older groups with lower participation rates. In particular, while the contribution of those between 25-34 years old was very supportive to aggregate participation in 1983-1995 it turned negative in 1995-2007. At the same time, the table indicates that the increase in participation across age and gender groups was very significant between the two periods considered. Indeed, even though the population effect declined in 1995-2007, the overall participation rate increased much more in the second period than in 1983-1995. Looking forward,
population ageing implies that the older age groups within the working age population gain more weight: those above 55 years old that are expected to be around $20 \%$ of the working age population in 2015, compared with $17.6 \%$ in 2007. In contrast, the weight of the youngest age groups, i.e. below 24 years, is expected to decline by 1 percentage point over the same period to $16.5 \%$; and the weight of the group between 35 and 44 years old, i.e. those most attached to the labour market, is expected to decline by more than 2 percentage points to $21.4 \%$. The mechanical decomposition based on age specific participation rates and population shares therefore suggests a substantial decline (by 0.6 percentage point) in the aggregate participation rate, putting downward pressure on total labour supply and potential growth in the euro area. This downward pressure intensifies significantly (a decline of 2.4 percentage points in the aggregate participation rate) if the horizon is extended up to 2030. By 2030 the oldest group (those between 55 to 64 years) is expected to account for one fourth of the working age population.

## 3. Data and methodology

The source for data on population, employment and unemployment for detailed age and gender groups for euro area countries is the EU Labour Force Survey (LFS) compiled by Eurostat. ${ }^{5}$ The same LFS data are used by Eurostat to calculate official statistics on participation and unemployment for EU countries. The LFS data are harmonised across countries and therefore particularly well-suited for cross-country comparative analysis. The annual data from 1983 to 2007 are based on the spring (second quarter) results. Data are available for ages from 15 to those over $70 .{ }^{6}$

Constructing consistent data over time requires some adjustments. In the case of Germany, data prior to 1991 have been extrapolated backwards on the basis of the developments in West Germany. We refer to two euro area aggregates in the paper. The first consists of the euro area 12 countries before Slovenia, and later Malta, Cyprus and Slovakia, entered the euro area. While there is no information available in the LFS for the euro area countries before they joined the European Union (i.e. for Spain and Portugal prior to 1986, for Austria and Finland prior to 1995), this has been taken into account in the calculation of the euro area 12 aggregate. In particular, data for the euro

[^3]area 12 aggregate prior to 1996 have been obtained on the basis of the growth rate of the largest aggregate available (i.e. 12 countries in 1995 to 2006, 10 countries between 1986 and 1995 and 8 countries before 1986). The second aggregate (euro area 5) contains only the largest 5 countries of the euro area, namely Germany, France, Italy, Spain and the Netherlands. This aggregate is calculated from the actual and estimated participation rates of the single countries, weighted with their respective population shares.

We use the output gap to measure the business cycle. The output gap is calculated as a deviation of real GDP from an HP filtered trend. In line with Uhlig and Ravn (2002), the smoothing parameter in the HP-filter for annual data is set at lambda $=6.25$. The real GDP data for both the euro area 12 and the single countries is taken from the AMECO database.

The full model specifications include a number of indicators for key time-varying institutions. We include OECD indicators for union density, labour taxes, implicit tax for older workers, the unemployment benefit replacement rate, the share of youth in education and average number of children also used in Bassanini and Duval (2006) and a measure of life expectancy from Eurostat. When missing, data for the last few years has been extrapolated based on past trends. As noted before, we include institutions to control for changes in the shape of the age-participation profile. This means that we rely on time-series variation of institutions within a single country to identify the impact of institutions. Therefore, several important institutional determinants of labour supply that do not generally vary over time, such as the mandatory retirement age, are excluded from this analysis.

Finally, in order to construct a scenario for future labour supply, we use population projections from the New Cronos database by Eurostat (EUROPOP2008). EUROPOP2008 contains statistical information on population projections with reference to projected 1st of January population by sex and single year of age, projected vital events (births and deaths) and assumptions concerning fertility, life expectancy at birth by sex and international migration. In the projections, we have made use of two variants: the baseline projection includes migration, while an alternative scenario captures the population developments without migration.

The estimation strategy is based on the cohort-based model presented in Fallick and Pingle (2007). ${ }^{7}$ Specifically, we estimate a system of constrained least squares regressions for single ages 15 to 70 and over, separately for men and women:

[^4]$$
\ln \left(\frac{L F P R_{g, t}}{1-L F P R_{g, t}}\right)=\alpha_{g}+\sum_{b=1917}^{1992} C_{g, b, t} \beta_{b}+\lambda_{g} X_{g, t}+\varepsilon_{g, t}
$$

The dependent variable is the logistic transformation of the participation rate for males or females. We use the logistic transformation to ensure that predicted participation rates remain bounded between 0 and 100 and undo the transformation after estimation.

The coefficient $\alpha$ represents an age effect that is constant over time and measures the average propensity to participate in the labour market at a certain age. Alphas for all ages trace an underlying fixed age-participation profile. The coefficients $C_{g, b, t}$ represent dummies for the different birth years and are equal to one if the birth cohort $b$ appears in age $g$ at time $t$. Within each gender group and country, the coefficient $\beta$ is constrained to be the same across equations. This allows an identification of cohort effects separately from the age and business cycle effects. As a consequence, the coefficients $\beta$ represent cohort effects that are constant over time and may be interpreted as the average propensity to participate in the labour force when born in a particular year. The cohort effect shifts the underlying age participation profile up and down, depending on the propensity of the birth year cohort to participate in the labour market throughout their working lives. We include all cohorts in the estimation which results in considering persons born between the years 1917 and 1992. However, as the most recent birth cohorts are only observed when they are very young, we estimate the model without the last eight cohorts. ${ }^{8}$ Later, we assign a cohort effect to these cohorts after estimation by setting it equal to the last estimated cohort effect (equal to the cohort effect for those born in 1984).

Finally, $X$ contains other variables that have explanatory power for participation rates of particular age groups. In the baseline specification, this encompasses business cycle effects represented by the contemporaneous value and two preceding lags of the output gap. In addition, both the estimated age and cohort effects are potentially influenced by time-varying institutions. In the full model therefore $X$ includes also a set of indicators of observed determinants. Note that the institutions do not vary across ages, although some institutions are included only in the equations for

[^5]young (youth in education), female (number of children) or older workers (implicit tax and life expectancy). The coefficients of the observed determinants vary freely across ages and therefore allow the underlying age-participation profile to tilt.

The total system is estimated based on 1400 age-year observations, with 56 equations, resulting in 56 estimated age and 168 estimated business cycle parameters each and 72 (constrained) cohort parameters. The unconstrained model results in a regressor matrix that is of reduced rank. With the help of the restrictions on the cohort effects, the estimation is nevertheless possible as shown in Greene and Seaks (1991). Significance tests are based on robust (White-corrected) standard errors. The cross-equation constraints identify the cohort effects only up to a scale factor. As in Fallick and Pingle (2007), we therefore normalize the coefficient estimates by setting the parameter of one cohort (here 1969) to one.

## 4. Results

We present results in three parts. We first illustrate the role of age and cohort effects using a basic decomposition of participation rates into age, cohort and business cycle effects. Second, we add a number of indicators of time-varying observed determinants of participation behaviour, such as labour market institutions, in the model. Finally, we present projections for male and female participation rates up to 2030. While the focus is on the euro area, we use country results to illustrate and to account for cross-country heterogeneity in participation behaviour. In particular, the full model with time-varying observed determinants is estimated separately for the five largest euro area countries (Germany, France, Italy, Spain and the Netherlands) and the results, in terms of trends and projections, are then aggregated to form a euro area 5 aggregate.

### 4.1. Basic model

Figures 4 to 7 show the results from the decomposition to age and cohort effects using the basic model for the euro area (EA12) and the five largest euro area countries. The results suggest substantial and highly statistically significant age effects that show the familiar hump-shaped pattern. For males, underlying euro area participation rates increase until age 30, remain stable until age 50, before gradually declining again (see Figure 4). The overall level of the underlying euro area ageparticipation profile is lower for females. While male participation rates are close to $100 \%$ in primeage, female participation is highest at ages $40-50$ at around $45 \%$. In addition, for females the
estimated age participation profile features a pronounced dip around early 30s. The dip suggests that a number of women leave the labour market temporarily to start a family, returning to work in their late 30s. The overall pattern of the age effects is similar across euro area countries. However, both the level and shape of the underlying female age-participation profiles differ substantially across countries (see Figure 5). While female participation rates peak at close to $70 \%$ in France, they do not exceed $40 \%$ in Spain and the Netherlands. The post child-bearing-age increase in participation is visible in all countries except Italy. These differences point to significant underlying differences in the participation behaviour of European women that are likely to reflect a combination of timeinvariant cultural and institutional factors. The results also point to the need to model the female participation rate using a flexible functional form that varies across countries. While the male ageparticipation profile could be characterized by a second order polynomial of age, female ageparticipation profiles are more complex and cannot be adequately captured by simple polynomials.

The results from the basic model also confirm that cohort effects are statistically significant and robust to age and period effects as measured by an indicator of the business cycle. In line with the descriptive evidence, cohort effects appear more significant in size for females than males. The normalized cohort fixed effects for the euro area are plotted in Figure 6 for both males and females. The results show a broadly declining profile for males and an increasing profile for females. The overall pattern of declining cohort effects for men and decreasing cohort effects for women appears similar to that observed in the United States (see Figure 8 in Fallick and Pingle, 2007). This mix of a positive cohort effect for the middle female cohorts and a negative effect for the younger female cohorts has a large impact on overall labour supply and, as demographic change shifts the weight between birth cohorts, turns out to be a relevant factor for future euro area labour supply. Again country results broadly confirm the overall pattern of estimated cohort effects (see Figure 7). The relative decline in cohort effects for men varies most across countries, with a substantial decline in Italy contrasted with an increase throughout in the Netherlands.

Combining the age and cohort effects, and excluding the business cycle and the error term, provides a measure of trend participation rates for each age group. For both females and males the actual and trend participation rates show a decline in participation of the younger age groups (up to 20-24 for females and 25-29 for males). For those in prime age and for older workers the trends diverge somewhat. For females, actual and trend participation increase for those in prime working age, and post mid 1990s also for older workers. For males, actual and trend participation rates are either stable or declining for those in prime age, whereas a more recent increase in participation rates
for older workers results in a mild u-shaped pattern. Estimated trend participation rates from the simple decomposition capture actual developments reasonably well for most detailed age groups. The model does particularly well in explaining the increasing trend of female participation and the recent increase in the participation of older workers for both males and females (with different timing across genders). In comparison, the results for some age groups suggest that the simple model misses important determinants of participation. Aggregating results for both males and females shows that beyond the broad trends of increasing participation of females and decreasing participation of males, important medium term developments are not fully captured by the simple model. For both males and females this includes a mild slump in participation in the 1990s and the most recent increase beginning around 2004. For males, actual participation rates were also above trend rates as captured by the model in the early 1980s. Overall, while the simple decomposition does well in explaining broad trends in participation, in particular for females, for some groups, age and cohort effects alone are not sufficient to capture trend participation patterns in the euro area.

### 4.2. Model with observed determinants

Going beyond the basic model, it is likely that other factors, such as time-varying labour market institutions, may have influenced participation trends in the medium term. Therefore, in a second step, we estimate the cohort model for the five largest euro area countries with a number of indicators of observed determinants that may matter for participation decision. We include union density, labour taxes, implicit tax on retirement for older workers, unemployment benefit replacement rate, the share of highly educated in the youth population, average number of children and life expectancy. The list of indicators is suggested by previous empirical analysis on the impact of institutions on labour force participation (see Bassanini and Duval, 2006 and Genre et al. 2005 and 2008) and theoretical considerations. In addition, availability of comparable indicators with sufficient time variation limits the list of relevant institutional factors that are considered (excluding, for example, indicators of employment protection legislation or the retirement age).

A number of hypotheses about the likely impact of these institutional factors can be put forward. First, we expect that declining union density in a number of euro area countries may have contributed to increase participation through its positive impact on expectations about the availability of jobs to those that have been previously inactive. As unions tend to compress the wage distribution, the decline in unionisation may have more of an impact on those at the lower part of the wage distribution (more likely to be younger and older workers). Second, an increase in labour taxes
(observed in a number of euro area countries) over time is also expected to result in lower labour participation by making leisure relatively less expensive. However, from a household labour supply perspective an increase in labour taxes for the head of the household may also result in an increased propensity to participate for other members of the household (more likely to be women). For older workers the implicit tax on continued work, a summary measure of retirement incentives, is likely to be more relevant than the overall labour tax. A higher implicit tax rate is expected to lower incentives to retire early (for the ages 55-64 considered here) and therefore to increase the participation rate of older workers. Third, observed declines in the generosity of the unemployment benefits system, as measured by the replacement rate, in a number of euro area countries is likely to lower the incentive to participate in the labour market by lowering alternative income when unemployed relative to inactivity. By contrast, unemployment benefits may also have a positive impact on participation via wage bargaining, with lower generosity leading to weakening of the insider's position in the labour market relative to the outsiders, or as a proxy for the overall generosity of the welfare system. Fourth, longer life expectancy is likely to lead to higher participation for older workers as they remain active and may also anticipate a longer period of retirement. Fifth, the higher share of young in education relative to older workers is expected to lower participation of young workers. Finally, the number of children is expected to influence female participation rates, with more children lowering participation rates of women around the typical age for starting a family. While union density, unemployment benefits and labour taxes are included in the equations for all age and gender groups (in working age), variables relating to education are included only for the youngest workers, life expectancy and the implicit tax on continued work for the oldest workers and, finally, the number of children for females only.

Tables 2 and 3 show the aggregated coefficient estimates and their t-statistics of the observed determinants of participation for three main age groups: young (15-24), prime-aged (25-54) and older (55-64), for all five countries. To simplify comparisons of coefficient estimates across groups and indicators, the data on observed determinants has been standardised. Note that the identification of the impact of institutions here relies only on available within-country time variation, which is often limited for the indicators of labour market institutions considered here. As a result, relatively few indicators turn out to be statistically significant. With this caveat in mind, a number of institutional indicators seem to matter, although the magnitude, and in some cases the sign, varies across countries and age groups. Higher labour taxes tend to lower participation rates (as reflected in 16 out 19 statistically significant coefficients). This impact is estimated more consistently for males in all countries. Higher union density (in 11 out of 15 statistically significant coefficients) and more
generous unemployment benefits ( 17 out of 24 statistically significant coefficients) also tend to lower participation rates. The negative impact of unemployment benefits is consistent with the interpretation that unemployment benefits impact participation rates either via their impact on bargaining (with increased power for insiders leading to higher bargained wages and lower participation rates for outsiders) or via their role as a proxy of the overall generosity of the welfare state (more generous benefits tend to coincide with more generous welfare benefits for financing non-participation, lowering participation rates). Exceptions to this result occur mainly for young people, whose participation rates in some countries are positively associated with unemployment benefits. The results also suggest that unemployment benefits increase participation of all males in Germany. While not conclusive, these results are suggestive of negative incentive effects for the unemployed stemming from generous unemployment benefits that are also of relatively long duration. In this case, a decline in benefits over time would lead some unemployed workers (who may have not been actively looking for jobs) to leave the labour force altogether. Overall, the results for union density and unemployment benefits are broadly in line with panel regression results in Genre et al. (2005 and 2008), who also find that higher union density and more generous unemployment benefits lower participation rates.

For females, with the exception of young females in France and the Netherlands, higher number of children tends to lower participation. The decline in number of children in most euro area countries is therefore associated with an increase in female participation rates. This is also in line with the Genre et al. (2008) finding that the fertility rate is negatively associated with participation rate of prime-aged females. ${ }^{9}$ Other group specific variables appear to be estimated less consistently, with both the sign and statistical significance changing across age groups and countries. Higher implicit tax on retirement, in the few cases when it is statistically significant, increases participation of older workers. With few counterintuitive exceptions (older people in France) increased life expectancy also increases participation of older workers. Both the sign and statistical significance of the share of youth in education varies across countries, suggesting that investment in human capital may not be well captured in the model. ${ }^{10}$

[^6]As regards the business cycle, we find that the sum of the coefficients of current and two lags of the output gap for worker groups are often not statically significant (not shown). In addition, for a number of groups we find a negative business cycle effect. For some groups, such as young people and females, this result could reflect "added worker" effects. For example, for individuals in families with a main bread-winner, in good times labour income from the rest of the family members may not be needed, whereas additional income from a second job is needed in bad times. ${ }^{11}$ We tried other indicators of the business cycle (unemployment and employment gap measures) with similar results. We therefore conclude that the business cycle has little influence on participation decisions in these countries, in line with results that show that European unemployment and employment rates are mainly influenced by structural factors or interactions of structural factors and shocks (e.g. Bassanini and Duval, 2006 and Blanchard and Wolfers, 2000).

In order to illustrate the size of the total impact of observed determinants we compare the total change in trend participation rates as measured by our model with a scenario of keeping the observed determinants at their 1995 values. The scenario reflects the view that the acceleration of labour market reforms from the second half of the 1990s onwards has contributed positively to participation rates (see for example, Masuch et al., 2008). The results are shown in Figure 8. The positive impact of observed determinants on participation is most evident for older males, as reflected in the large gap between the two bars. The most relevant variable in this respect appears to be life expectancy. The increase in life expectancy since 1995 has had a positive impact in the participation rate of older males. Overall, the observed determinants have resulted only in small increase in participation rates for females, with most of the increase over this time period attributed to age and cohort effects instead. For young people, the impact of observed determinant has been to dampen participation rates. This is partly explained by the increase in the proportion of young people in education. At the country level, it is worth mentioning that the impact of the change in the institutional framework is broadly based for prime-age age males and females, and for females aged 55 and over. In contrast, for the young, the developments are strongly influenced by the results for Italy and France, and for males aged 55 and over by the results for Netherlands and Germany.

Both age and cohort effects remain jointly statistically significant in all models even after including business cycle indicators and other time-varying determinants of participation. These coefficients can be thought of as capturing the impact of other time-invariant cultural or institutional

[^7]factors (for the age coefficients) or slowly changing impact of factors that are specific to birth years (for the cohort coefficients). The latter may include factors such as cultural attitudes towards labour market participation (for women in particular) or institutional factors and reforms that are not captured by the observed determinants. Figure 9 plots the estimated cohort profiles based on the trend participation rates from the model, aggregated to the euro area five (EA5) level. For males, the lines indicating participation rates for specific cohorts are mostly overlapping. There is some indication that most recent cohorts enter the labour market later, reflecting the substantial increase in the average number of years spent in education. The same impact is visible also for the youngest female cohort. However, in addition, the cohort profile for females suggests a substantial shifting up of the age participation profile over time. For prime-aged women, those in their mid 30s and 40s, the participation rate has increased by more than 20 percentage points. Furthermore, while the cohort profile for those born between 1953 and 1962 shows a pronounced dip at child bearing age, this dip is not visible for the next cohort (those born between 1963 and 1973). The higher propensity to participate of females born in the late 1960s and early 1970s has therefore contributed to the increase in female participation in the euro area.

Country results show that participation behaviour differs across the largest five euro area countries and that again this is most evident for women. For women in their 30s and 40s, the estimated cohort profiles show that participation rates of most recent cohorts has increased most, by more than 30 percentage points, in the Netherlands and Spain and the least in France, with roughly 10 percentage points (see Figure 10 for women, results for men are available upon request). The disappearance of the dip at child bearing age is most pronounced in the Netherlands: while the participation rates of women in the late 20s and early 30 s for those born in 1950s dropped by as much as 20 percentage points, the more recent cohorts appear to have stayed in the labour market through the child bearing years.

### 4.3. Projections

In a third and final step we use the model results to project participation rates forward until 2030. The results of both trend and projection for the euro area (EA5) are obtained by aggregating the full model estimates for Germany, France, Italy, Spain and the Netherlands weighting the countries with their respective populations. We assume that age and cohort effects are fixed throughout the sample and keep observed determinants at their 2007 values. In addition, for the young cohorts, i.e. the last eight cohorts of our sample and those that enter the labour market after

2007, we fix their cohort effects at the level of the last cohort effect we estimate, namely those born in 1984. Figure 11 shows the trend from the estimation of the full model for males and females together with the actual participation rates for the euro area (EA5). The results clearly show that within the sample period the full model captures both trends and medium-term developments well. This is confirmed by results for individual ages shown in the Appendix.

The projected euro area participation rate decreases for males throughout the projection period. In contrast, the euro area participation rate for females increases before stabilising at about $70 \%$ in 2030. This pattern is in line with the waning impact of the positive cohort effects for females that continue to support participation rates looking forward. While the gap between male and female participation rates is expected to decline substantially, at the end of the projection horizon male participation rate remains 4.6 percentage points above the female participation rate. Overall, towards the end of the sample the negative impact of population ageing shifting the larger share of the population to older age groups with lower participation rates begins to dominate and dampen the overall participation rate. As a result, the overall participation rate is anticipated to increase slightly up to 2015, by 1 percentage point, but to decline thereafter. However, in 2030 it is still expected to remains at just above the 2007 level (see Table 4). The underlying country results from the baseline model are shown in Table 5. The results for all countries point to an ongoing increase in female participation and a decline in male participation. Indeed, in the Netherlands and France the gap between male and female participation closes by 2030. Reflecting the continued positive trend in female participation, the overall participation rate is expected to continue increasing in all countries except Germany.

In order to explore these results further and to evaluate robustness we calculate three additional scenarios for the euro area (see Table 4). First, we compare our results with a scenario that keeps participation rates by age and gender groups unchanged at their 2007 level, i.e. accounting only for population effects. The model based results imply a more positive outlook for participation than a scenario based on unchanged participation behaviour. Indeed, in the latter scenario, the overall participation rate declines already in 2015, with a gap of 3.9 percentage points in 2030 between the baseline model results and the alternative scenario. Second, we calculate a scenario that accounts for the impact of migration through the population structure. Specifically we compare the baseline model results with a scenario that assumes no migration and find that the impact of migration through the population structure is positive. On average, migrants tend to be younger and therefore to have higher participation rates than the native population. The impact is relatively small, but its
relevance grows over time - the gap in the participation rate between the baseline migration and nonmigration scenario is 0.1 percentage point in 2015, but reaches 0.8 percentage point in 2030. A comparison of the scenarios by gender shows that migration is only relevant for the male participation rate, while the impact on the female participation rate is negligible. Finally, we compare our results with those derived from the participation rate projections at the country level published by the European Commission (European Commission, 2008). We find that there is a significant gap between our baseline model results and the European Commission projections. According to the European Commission, the overall participation rate is expected to increase somewhat more, by 2.2 percentage points in 2015 and 3 percentage points in 2030. This gap reflects a substantially more positive outlook for male participation in the European Commission projections - for males the gap in participation is 2.4 and 5.8 percentage points in 2015 and 2030 respectively. While it is not straightforward to decompose the difference in terms of underlying determinants, the European Commission projections appear to incorporate more inertia from recent participation trends for males. Note that we keep the effect of institutional variables unchanged in the model based projections. Therefore, recent changes in observed determinants that persist or have lagged effects are not reflected in our model based scenario. In contrast, reflecting the important role of cohort effects in explaining past participation trends, our results suggest a somewhat more positive outlook for female participation.

## 5. Conclusions

We use a cohort based model of labour force participation to analyse determinants of participation for disaggregated groups of workers in European countries, with a focus on the euro area. The model identifies significant age and cohort effects for detailed worker groups as indicators of (unobserved) structural determinants. We use observed structural determinants and age and cohort effects to construct trend measures of labour supply and to disentangle the impact of structural and business cycle factors on labour force participation rates.

Our results suggest that age and cohort effects can explain a substantial part of the recent increase in labour force participation rates in the euro area, although not the surge since early 2000s. Cohort effects are particularly relevant for women, with those born in the 1920s and 1930s less likely and those born in the late 1960s and early 1970s more likely to participate in the labour market over the life-cycle. There is substantial variation in cohort effects across the five largest euro area countries that we analyse. While cohort effects generally encompass any factor associated with a
particular birth year, we speculate that the cohort effects that we observe reflect evolving preferences or social norms that vary across countries. Depending on the country, the estimated cohort profiles suggest an increase of 10 to 30 percentage points in female participation rates. We control for a number of observed time-varying institutions, such as labour taxes, union density, unemployment benefits and the average number of children and find that they have had an impact on labour force participation rates, although the specific impact varies across age and gender groups and countries. Looking forward, while they continue to provide some upward support to participation rates of women in the euro area, positive cohort effects are not large enough to compensate for the downward impact of population ageing on labour force participation rates in the euro area.

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Table 1. The contribution of population composition to changes in participation rates

|  | $1983-1995$ | $1995-2007$ | $2007-2015$ | $2007-2030$ |
| :--- | :---: | :---: | :---: | :---: |
| $15-19$ | -1.4 | -0.2 | -0.1 | 0.0 |
| $20-24$ | -0.4 | -0.8 | -0.3 | -0.1 |
| $25-34$ | 1.9 | -2.4 | -1.4 | -1.9 |
| $35-44$ | 1.1 | 1.9 | -1.9 | -3.5 |
| $45-54$ | -0.3 | 1.7 | 2.2 | 0.4 |
| $55-64$ | 0.4 | 0.2 | 1.0 | 2.8 |
|  |  |  |  |  |
| Total | 1.3 | 0.4 | -0.6 | -2.4 |
|  |  |  |  |  |
| Change in PR | 1.5 | 5.7 | -- | -- |

Note: percentage points. Sources: EU LFS (Eurostat) and own calculations.

Table 2. Impact of observed determinants: males

|  | LT | UD | UB | TR | LE | YE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Germany: |  |  |  |  |  |  |
| Young | $\begin{gathered} 0.00 \\ (-0.22) \end{gathered}$ | $\begin{gathered} -0.01 \\ (-0.33) \end{gathered}$ | $\begin{gathered} 0.10 \\ (5.07) \end{gathered}$ |  |  | $\begin{gathered} 0.05 \\ (3.68) \end{gathered}$ |
| Prime-aged | $\begin{gathered} -0.07 \\ (-7.03) \end{gathered}$ | $\begin{gathered} -0.03 \\ (-1.43) \end{gathered}$ | $\begin{gathered} 0.03 \\ (1.97) \end{gathered}$ |  |  |  |
| Older | $\begin{gathered} -0.12 \\ (-7.89) \end{gathered}$ | $\begin{gathered} -0.04 \\ (-1.02) \end{gathered}$ | $\begin{gathered} 0.07 \\ (3.34) \end{gathered}$ | $\begin{gathered} 0.03 \\ (0.54) \end{gathered}$ | $\begin{gathered} 0.40 \\ (7.29) \end{gathered}$ |  |
| France: |  |  |  |  |  |  |
| Young | $\begin{gathered} -0.16 \\ (-4.86) \end{gathered}$ | $\begin{gathered} 0.46 \\ (5.14) \end{gathered}$ | $\begin{gathered} 0.05 \\ (2.54) \end{gathered}$ |  |  | $\begin{gathered} -0.27 \\ (-5.00) \end{gathered}$ |
| Prime-aged | $\begin{gathered} -0.07 \\ (-5.10) \end{gathered}$ | $\begin{gathered} -0.04 \\ (-2.10) \end{gathered}$ | $\begin{gathered} -0.02 \\ (-2.28) \end{gathered}$ |  |  |  |
| Older | $\begin{gathered} 0.02 \\ (0.86) \end{gathered}$ | $\begin{gathered} -0.09 \\ (-2.06) \end{gathered}$ | $\begin{gathered} -0.06 \\ (-5.49) \end{gathered}$ | $\begin{gathered} 0.02 \\ (1.44) \end{gathered}$ | $\begin{gathered} -0.09 \\ (-2.47) \end{gathered}$ |  |
| Italy: |  |  |  |  |  |  |
| Young | $\begin{gathered} -0.14 \\ (-6.10) \end{gathered}$ | $\begin{gathered} -0.07 \\ (-2.00) \end{gathered}$ | $\begin{gathered} -0.32 \\ (-9.00) \end{gathered}$ |  |  | $\begin{gathered} 0.20 \\ (2.95) \end{gathered}$ |
| Prime-aged | $\begin{gathered} -0.11 \\ (-4.83) \end{gathered}$ | $\begin{gathered} 0.00 \\ (-0.15) \end{gathered}$ | $\begin{gathered} -0.10 \\ (-3.85) \end{gathered}$ |  |  |  |
| Older | $\begin{gathered} 0.00 \\ (0.08) \end{gathered}$ | $\begin{gathered} -0.15 \\ (-4.43) \end{gathered}$ | $\begin{gathered} 0.06 \\ (1.54) \end{gathered}$ | $\begin{gathered} 0.04 \\ (3.12) \end{gathered}$ | $\begin{gathered} 0.21 \\ (3.06) \end{gathered}$ |  |
| Spain: |  |  |  |  |  |  |
| Young | $\begin{gathered} -0.03 \\ (-3.18) \end{gathered}$ | $\begin{gathered} -0.04 \\ (-2.24) \end{gathered}$ | $\begin{gathered} -0.12 \\ (-4.27) \end{gathered}$ |  |  | $\begin{gathered} 0.05 \\ (0.85) \end{gathered}$ |
| Prime-aged | $\begin{gathered} -0.08 \\ (-6.52) \end{gathered}$ | $\begin{gathered} -0.01 \\ (-0.50) \end{gathered}$ | $\begin{gathered} -0.11 \\ (-7.65) \end{gathered}$ |  |  |  |
| Older | $\begin{gathered} -0.03 \\ (-1.78) \end{gathered}$ | $\begin{gathered} 0.01 \\ (0.62) \end{gathered}$ | $\begin{gathered} -0.07 \\ (-2.33) \end{gathered}$ | $\begin{gathered} -0.03 \\ (-0.59) \end{gathered}$ | $\begin{gathered} 0.05 \\ (1.49) \end{gathered}$ |  |
| Netherlands: |  |  |  |  |  |  |
| Young | $\begin{gathered} -0.08 \\ (-1.91) \end{gathered}$ | $\begin{gathered} -0.34 \\ (-5.21) \end{gathered}$ | $\begin{gathered} -0.08 \\ (-2.97) \end{gathered}$ |  |  | $\begin{gathered} -0.02 \\ (-0.25) \end{gathered}$ |
| Prime-aged | $\begin{gathered} -0.03 \\ (-1.37) \end{gathered}$ | $\begin{gathered} -0.18 \\ (-4.42) \end{gathered}$ | $\begin{gathered} -0.08 \\ (-4.21) \end{gathered}$ |  |  |  |
| Older | $\begin{gathered} -0.11 \\ (-2.66) \\ \hline \end{gathered}$ | $\begin{gathered} 0.05 \\ (0.71) \\ \hline \end{gathered}$ | $\begin{gathered} 0.04 \\ (1.30) \\ \hline \end{gathered}$ | $\begin{gathered} 0.42 \\ (4.49) \\ \hline \end{gathered}$ | $\begin{gathered} 0.14 \\ (2.20) \end{gathered}$ |  |

Note. T-statistics based on robust standard errors in parenthesis. LT is labour taxes, UD is union density, UB is unemployment benefits, TR is tax on retirement, LE is life expectancy, YE is youth education. For each age group, the coefficients and their standard errors have been aggregated from single ages using labour force weights in 2007.

Table 3. Impact of observed determinants: females

|  | LT | UD | UB | TR | LE | YE | NC |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Germany: |  |  |  |  |  |  |  |
| Young | 0.00 | 0.13 | 0.06 |  |  | 0.03 | -0.12 |
|  | $(0.35)$ | $(4.70)$ | $(4.22)$ |  |  | $(3.26)$ | $(-10.04)$ |
| Prime-aged | -0.04 | 0.02 | 0.01 |  |  |  | -0.01 |
|  | $(-8.08)$ | $(1.70)$ | $(1.01)$ |  |  |  | $(-1.51)$ |
| Older | -0.04 | 0.00 | -0.05 | 0.06 | -0.03 |  | -0.02 |
|  | $(-3.78)$ | $(-0.10)$ | $(-3.98)$ | $(2.12)$ | $(-0.95)$ |  | $(-2.07)$ |
| France: |  |  |  |  |  | -0.29 | 0.17 |
| Young | -0.14 | 0.51 | 0.07 |  |  | $(-5.11)$ | $(1.89)$ |
|  | $(-3.96)$ | $(4.18)$ | $(3.36)$ |  |  |  | 0.00 |
| Prime-aged | 0.03 | -0.02 | -0.01 |  |  |  | $(-0.11)$ |
|  | $(2.47)$ | $(-1.07)$ | $(-2.27)$ |  |  | 0.08 |  |
| Older | 0.06 | -0.19 | -0.02 | -0.03 | -0.09 |  | $(0.91)$ |
|  | $(1.99)$ | $(-3.64)$ | $(-1.49)$ | $(-1.50)$ | $(-2.22)$ |  |  |

Italy:

| Young | -0.11 | -0.09 | -0.22 |  |  | 0.59 | -0.04 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $(-5.79)$ | $(-2.14)$ | $(-6.94)$ |  |  | $(8.85)$ | $(-1.95)$ |
| Prime-aged | -0.07 | -0.01 | -0.10 |  |  |  | -0.02 |
|  | $(-6.18)$ | $(-0.84)$ | $(-8.17)$ |  |  |  | $(-2.55)$ |
| Older | -0.01 | -0.04 | -0.06 | 0.01 | 0.06 |  | 0.00 |
|  | $(-0.28)$ | $(-1.29)$ | $(-1.75)$ | $(0.86)$ | $(0.94)$ |  | $(-0.13)$ |

Spain:

| Young | 0.03 | 0.05 | -0.09 |  |  | -0.01 | 0.15 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $(2.31)$ | $(1.42)$ | $(-3.02)$ |  |  | $(-0.05)$ | $(1.03)$ |
| Prime-aged | 0.01 | 0.03 | -0.06 |  |  |  | -0.10 |
|  | $(1.28)$ | $(2.88)$ | $(-4.02)$ |  |  |  | $(-1.29)$ |
| Older | 0.00 | -0.04 | -0.05 | -0.01 | 0.01 |  | 0.06 |
|  | $(0.04)$ | $(-1.11)$ | $(-1.61)$ | $(-0.08)$ | $(0.35)$ |  | $(0.19)$ |

## Netherlands:

| Young | 0.03 | -0.15 | 0.05 |  |  | 0.04 | 0.15 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $(0.71)$ | $(-2.41)$ | $(2.72)$ |  |  | $(0.58)$ | $(4.73)$ |
| Prime-aged | -0.01 | -0.07 | -0.04 |  |  |  | -0.04 |
|  | $(-0.31)$ | $(-3.39)$ | $(-2.67)$ |  |  |  | $(-2.07)$ |
| Older | 0.02 | -0.11 | -0.02 | 0.20 | -0.03 | -0.05 |  |
|  | $(0.32)$ | $(-1.67)$ | $(-0.42)$ | $(2.20)$ | $(-0.52)$ | $(-0.76)$ |  |

Note. T-statistics based on robust standard errors in parenthesis. LT is labour taxes, UD is union density, UB is unemployment benefits, TR is tax on retirement, LE is life expectancy, YE is youth education and NC is number of children. For each age group, the coefficients and their standard errors have been aggregated from single ages using labour force weights in 2007.

Table 4. Alternative scenarios for future participation rates (EA5)

|  | 2007 | 2015 | 2020 | 2025 | 2030 |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Total participation rate |  |  |  |  |  |
| PR (baseline model) | 72.0 | 73.0 | 72.9 | 72.5 | 72.5 |
| PR (model - no migration) | 72.0 | 72.9 | 72.6 | 71.9 | 71.8 |
| PR (2007 level) | 72.0 | 71.3 | 70.3 | 69.1 | 68.6 |
| PR (EC) | 72.0 | 74.2 | 74.6 | 74.6 | 75.0 |
|  |  |  |  |  |  |
| Females participation rate |  |  |  |  |  |
| PR (baseline model) | 64.1 | 67.6 | 68.8 | 69.5 | 70.2 |
| PR (model - no migration) | 64.1 | 67.5 | 68.7 | 69.2 | 69.9 |
| PR (2007 level) | 64.1 | 63.1 | 62.1 | 60.9 | 60.5 |
| PR (EC) | 64.1 | 67.6 | 68.5 | 68.8 | 69.5 |
|  |  |  |  |  |  |
| Males participation rate |  |  |  |  |  |
| PR (baseline model) | 79.9 | 78.4 | 76.9 | 75.4 | 74.8 |
| PR (model - no migration) | 79.9 | 78.2 | 76.5 | 74.6 | 73.8 |
| PR (2007 level) | 79.9 | 79.4 | 78.4 | 77.1 | 76.5 |
| PR (EC) | 79.9 | 80.8 | 80.8 | 75.9 | 80.6 |

Note: Euro area obtained as the aggregation of Germany, Italy, France, Spain and the Netherlands. PR (2007 level) refers to a scenario based on unchanged participation rates at the 2007 level. PR (EC) refers to a scenario derived from European Commission (2008); it has been re-based to the 2007 level derived from the EU-LFS. Sources: EU LFS (Eurostat) and own calculations.

Table 5. Country projections

|  | 1987 | 1997 | 2007 | 2015 | 2020 | 2025 | 2030 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Germany |  |  |  |  |  |  |  |
| Total | 69.6 | 70.4 | 75.5 | 75.1 | 74.3 | 73.2 | 72.7 |
| Females | 57.6 | 61.8 | 70.0 | 71.4 | 71.5 | 71.1 | 70.8 |
| Males | 81.8 | 78.8 | 81.0 | 78.7 | 77.0 | 75.2 | 74.5 |
| France |  |  |  |  |  |  |  |
| $\quad$ Total | 67.1 | 68.0 | 68.6 | 68.8 | 68.8 | 68.7 | 68.8 |
| Females | 57.0 | 61.1 | 64.1 | 66.4 | 67.6 | 68.5 | 69.3 |
| $\quad$ Males | 77.6 | 75.2 | 73.3 | 71.2 | 70.0 | 68.8 | 68.4 |
|  |  |  |  |  |  |  |  |
| Italy |  |  |  |  |  |  |  |
| Total | 59.7 | 58.7 | 62.7 | 66.3 | 67.0 | 67.3 | 68.3 |
| Females | 41.9 | 44.0 | 50.9 | 57.7 | 60.3 | 62.1 | 64.1 |
| $\quad$ Males | 78.3 | 73.6 | 74.5 | 74.8 | 73.7 | 72.5 | 72.5 |
|  |  |  |  |  |  |  |  |
| Spain |  |  |  |  |  |  |  |
| Total | 58.3 | 62.8 | 72.5 | 74.2 | 74.0 | 73.3 | 72.8 |
| Females | 36.9 | 47.7 | 61.9 | 66.3 | 67.5 | 68.0 | 68.5 |
| Males | 80.3 | 78.0 | 82.7 | 82.0 | 80.3 | 78.4 | 76.9 |
| Netherlands |  |  |  |  |  |  |  |
| Total | 63.9 | 71.4 | 78.2 | 79.6 | 79.9 | 80.3 | 81.0 |
| Females | 48.4 | 61.2 | 71.8 | 76.7 | 78.7 | 80.4 | 81.9 |
| Males | 79.1 | 81.4 | 84.4 | 82.4 | 81.1 | 80.2 | 80.0 |

Sources: EU LFS (Eurostat) and own calculations.

Figure 1. Participation rates by worker groups in the euro area (EA12)


Sources: EU LFS (Eurostat), OECD and own calculations.

Figure 2. Age participation profiles by gender in the euro area (EA12), 2007


Sources: EU LFS (Eurostat) and own calculations.

Figure 3. Changes in participation rates by age and gender in the euro area (EA12) Male


Female


Note: Percentage points. Sources: EU LFS (Eurostat) and own calculations.

Figure 4: Estimated age-participation profiles in the EA12


Figure 5: Estimated age-participation profiles


Figure 6: Estimated cohort effects in the EA12


Figure 7: Estimated cohort effects


Figure 8: Total impact of observed determinants


Figure 9. Estimated cohort profiles in the EA5
Females


| - | $1974-1984$ | - | -- |
| :--- | :--- | :--- | :--- |
| $1963-1973$ |  |  |  |
| ----- | $1952-1962$ | $-\sim-$ | $1941-1951$ |
| .---- | $1930-1940$ | - | $1919-1929$ |

Males


|  | 1974-1984 | - | 1963-1973 |
| :---: | :---: | :---: | :---: |
|  | 1952-1962 | - | 1941-1951 |
| ------ | 1930-1940 | - - | 1919-1929 |

Figure 10. Estimated cohort profiles by country, females

Germany


| $\square-$ | $1974-1984$ | - | $1963-1973$ |
| :--- | :--- | :--- | :--- |
| -- | $1952-1962$ | - | $1941-1951$ |
| $-\cdots$ | $1930-1940$ | - | $1919-1929$ |

Italy


| - | $1974-1984$ | - | $1963-1973$ |
| :--- | :--- | :--- | :--- |
| --- | $1952-1962$ | - | $1941-1951$ |
| $-\cdot$ | $1930-1940$ | - | $1919-1929$ |



France


| - | $1974-1984$ | - | $1963-1973$ |
| :--- | :--- | :--- | :--- |
| --- | $1952-1962$ | - | $1941-1951$ |
| $-\cdots$ | $1930-1940$ | - | $1919-1929$ |

Spain


| - | $1974-1984$ | - | $1963-1973$ |
| :--- | :--- | :--- | :--- |
| --- | $1952-1962$ | - | $1941-1951$ |
| $-\cdots$ | $1930-1940$ | - | $1919-1929$ |

Figure 11. Trend participation and projections in the EA5 by gender, 1986-2030


Appendix A: Trend participation rates in the EA5 by age
Figure A1. Young females


Note: trend is a three year moving average of the estimated trend participation rate.

Figure A2. Prime-aged females


Note: trend is a three year moving average of the estimated trend participation rate.

Figure A3. Older females


Note: trend is a three year moving average of the estimated trend participation rate.

Figure A4. Young males


Note: trend is a three year moving average of the estimated trend participation rate.

Figure A5. Prime-aged males


Note: trend is a three year moving average of the estimated trend participation rate.

Figure A6. Older males


Note: trend is a three year moving average of the estimated trend participation rate.

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[^0]:    ${ }^{2}$ For example, Fernandez (2007) builds a model of female participation that is based on culture and learning. She argues that cultural factors can explain the increase (and the S-shaped time series pattern) of female participation rates in the United States. Antecol (2000) finds that the home country plays an important role in participation decisions of first generation female immigrants in the US, suggesting that culture matters for participation behaviour.

[^1]:    ${ }^{3}$ Other studies that use closely related methods include Beaudry and Lemieux (1999) for Canada and Fukuda (2006) for Japan. In addition, Carone (2005) and Burniaux et al. (2004) take advantage of cohort effects to project participation rates for EU and OECD countries (respectively).

[^2]:    ${ }^{4}$ Naturally, participation behavior varies also across other personal characteristics, such as education and skills, immigrant status etc. We focus on age and gender for reasons of data availability: in particular, LFS data by education categories is only available from the early 1990s onwards. That data shows that more educated workers tend to have higher participation rates and that an increase in overall educational attainment over time has coincided with an increase in participation rates, particularly for women.

[^3]:    ${ }^{5}$ A detailed description of the sampling methods and adjustment procedures used in the LFS can be found in "The European Union Labour Force Survey - Methods and Definitions, 2001", the available variables are listed and described in the "EU Labour Force Survey database - User guide". The change from annual to quarterly periods by Eurostat has resulted in breaks in the LFS survey in many euro area countries. Therefore we rely on the more consistent spring (second quarter) data throughout the sample period, except for France and Austria (first quarter).
    ${ }^{6}$ Except for Spain where data are available for those above 16.

[^4]:    ${ }^{7}$ Closely related models based on age, cohort and period accounting have a long tradition in sociological and demographic research and have been recently applied to analyse labour supply in Beaudry and Lemieux (1999) and

[^5]:    Fitzenberger et al. (2004). Articles in Mason and Fienberg (1985) provide an early discussion of basic accounting models and applications that rely on functional form assumptions. From an economic perspective the pure age, cohort and period accounting approach seems rather ad hoc in nature. The current model is therefore an attempt to move beyond a pure statistical decomposition by including observable variables that capture underlying factors that determine participation rates. See also Euwals et al (2007) for a discussion and comparison of different modelling strategies.
    ${ }^{8}$ We do this by replacing the values of the participation rate and the other explanatory variables of the ages affected with means from the rest of the sample. We also restrict the cohort effects of the last eight cohorts to equal the average of the remaining cohorts for the respective age.

[^6]:    ${ }^{9}$ We also experimented with other determinants of female participation, in particular, the tax rate on second earners and marriage rate. Previous literature has suggested that both are potentially important determinants (see Jaumotte, 2003). The tax rate on second earners was usually not statistically significant for these countries. We found some (counterintuitive) indication that the marriage rate is positively associated with participation. Both variables where therefore excluded from the final model. Jaumotte (2003) and Genre et al. (2008) exploit cross-section variation to establish other potential determinants of female participation (such as maternity leave) that we do not consider here.
    ${ }^{10}$ These effects could be captured better by changes in returns to education. However, we are not aware of comparable estimates of returns to education with a sufficiently long time-series that we could use.

[^7]:    ${ }^{11}$ Prieto-Rodriguez and Rodriguez-Gutierrez (2000) find these effects to be relevant for women in Spain, in line with our finding of negative business cycle effects for women of all ages.

