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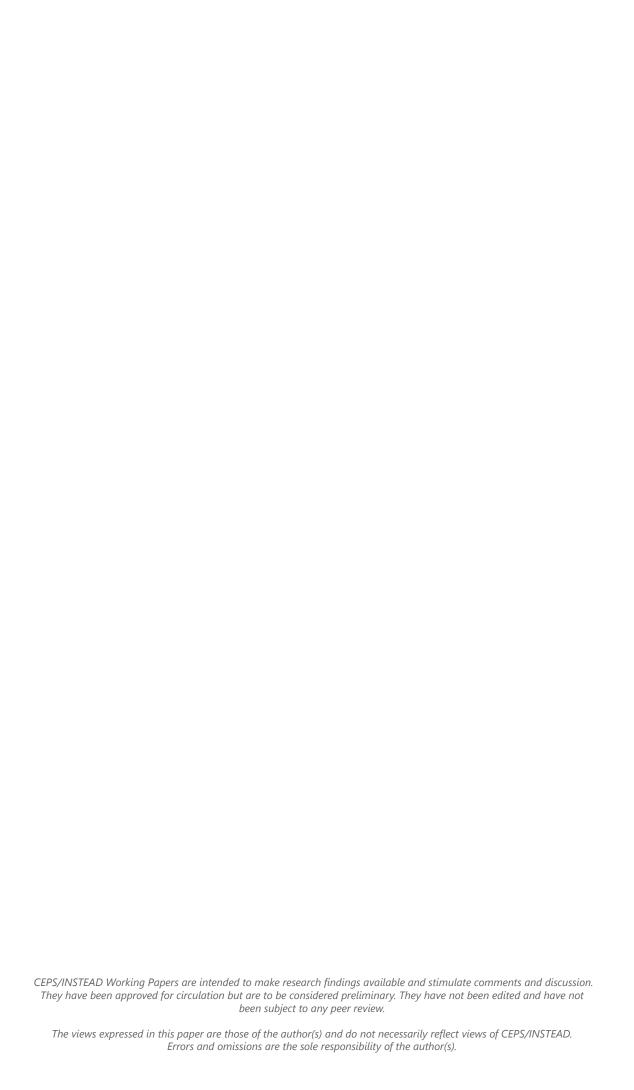


WORKING PAPERS

Macro Determinants of Individual Income Poverty in 93 Regions of Europe

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Abstract: The analysis of the at-risk-of-poverty determinants can be improved by taking into account factors at macro (regional) level. This hypothesis has already been made in previous research, at country-level, on cross-sectional data. We use longitudinal data in this analysis in order to get more precise estimated parameters, and we test if the regional unemployment rate and the regional GDP affect the individual at-risk-of-poverty status. The countries taken into account are those present in the Statistics on Income and Living Conditions (EU-SILC) dataset.

Key words: income poverty, EU-SILC, multilevel models, longitudinal data

JEL: I32

Avril 2010

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We would like to thank David Brady, Jacques Brosius, Alessio Fusco, Tony Atkinson, Eric Marlier and Philippe Van Kerm for useful discussions led during this project. This paper has been presented at the Net-Silc International Conference, 25-26 March 2010, Warsaw. Of course, remaining errors are ours.

1. Introduction

Tackling poverty by 2010 was one of the European objectives defined by the Lisbon European Council in 2000. Ten years later, 2010 is the European year for combating poverty and social exclusion. Poverty continues therefore to be at the heart of social policy in most European Member States. Ideally, social policies aimed at reducing poverty need to be based on an in-depth understanding of the underlying processes at work. A first step towards such an understanding consists in shedding some light on the main determinants of poverty.

Early descriptive studies have checked for relationships between poverty status and different characteristics taken in turn (Bradshaw, 1999; Bradubury et al., 1999; UNICEF, 2000; Mejer et al., 2000). This has given some insight on the factors involved, but these studies have only partially allowed to understand how these factors work. Other researches (see for example Cappellari and Jenkins, 2002; Fertig and Tamm, 2007; Brady et al., 2009) have extended this initial approach by reasoning all other things being equal, checking the effect on poverty of factors such as educational attainment, age, employment status, family structure – all of these factors having been calculated at the individual³ level. Simultaneously another stream of studies (see Moller et al., 2003; Wiepking and Maas, 2005; Brady et al., 2009; Tai and Treas, 2008) has emphasized the analysis of the role of macro characteristics in a cross-national context. These analyses have shown that the macro factors could well have an effect on the poverty probability. Indeed, the generosity of social benefits (and especially of family benefits) proves to have a significant negative effect on the odds of poverty (see Brady et al. 2009, Moller et al. 2003).

For all of these three types of analyses, one major improvement has consisted in taking into account the longitudinal feature of poverty, using panel data⁴. The indicator of persistent poverty⁵, for example, allows to figure out whether poverty is a temporary or rather a long-term phenomenon. Furthermore, developments in the econometrics of panel data have allowed researchers to further investigate important topics such as poverty duration or unobserved heterogeneity.

However, to our knowledge, no study has yet dealt simultaneously with all EU-countries, longitudinal data and factors at both individual and macro levels using a relevant specification. Brady et al. (2009) study the effect of macro-determinants on the probability of being poor using a GEE model⁶ applied to 15 EU (plus some non-EU) countries but using cross-sectional data. In this paper we will extend this kind of analysis to 93 EU-regions (26 countries) and, contrary to previous work, we will use a longitudinal dataset (EU-SILC 2005 and 2006).

This paper has the following objective: it aims at disentangling the role of micro and macro factors in explaining the poverty status, by using detailed information about different

³ As the same poverty status is, by the European definition, affected to all individuals belonging to the same household, some authors have defined the factors exclusively at that level (see for example Andriopoulou et al., 2008).

⁴ See Ray and Jeandidier (2003) for a comprehensive review of the French literature on this subject.

⁵ This indicator belongs to the set of common indicators for the social protection and social inclusion process adopted by the Social Protection Committee in 2006.

⁶ A Generalized Estimated Equations model can be used to estimate marginal or population-averaged effects taking into account the dependence among units nested in clusters (Rabe-Hesketh and Skrondal, 2008).

regions in Europe⁷. Indeed, we would like to test if there is a genuine effect of macro factors such as the unemployment rate on the poverty probability, and especially if these factors can affect the impact of individual characteristics such as the education level on this probability.

In Section 2 we present the definition of income poverty that we will apply. In Section 3 we then develop the different econometric methods available to deal with the question and data at hand. Section 4 gives a detailed description of our dataset. The results and comments of our own model are then presented in Section 5. Final conclusions are to be found in Section 6.

2. The definition of income poverty in Europe

In Europe, poverty is officially defined in relative terms, as the percentage of individuals living in a household whose equivalent income is below the poverty threshold. This threshold is defined in each country (equal to 60% of the national median equivalent income), aiming at taking into account the national income inequalities. As a consequence, two countries with very different standards of living (and thus very different median equivalent income and different poverty thresholds) can have the same poverty rate.

Seemingly contradictory results due to this definition do not matter as long as one is aware of the conventions they are based on, and when the at-risk-of-poverty rates are interpreted together with the threshold values. But, in our case, the main objective is to figure out to what extent the poverty status is explained by some macro factors such as the unemployment rate. It is thus necessary, in order to allow that kind of relationship to appear, that the poverty indicator ranks the countries as the macro variables do. As a consequence, we have chosen to keep defining poverty in a relative way (60% of a certain threshold) but to calculate a new threshold, allowing this kind of ranking. With this objective in mind, we calculate a unique European poverty threshold by considering all individuals to belong to a same big country, which is Europe⁸.

By doing this, we move away from the official EU-definition of the at-risk-of-poverty status (which is rather a measure of intra-country inequality), and we consider Europe as an integrated entity. In the same spirit, we exclude neither Iceland nor Norway: while not part of the EU-27 in 2009, they could be expected to join.

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⁷ Indeed it is between regions, and not between individuals, that some variance of macro factors can be found. This is probably also true between countries but because our data is limited with respect to the number of available countries, we prefer the analysis at the regional level.

⁸ Using the usual at-risk-of-poverty threshold (which gives similar at-risk-of-poverty rates for countries with very different economic situations) could lead to this kind of situation: the characteristics of the countries would not be relevant to explain their own at-risk-of-poverty rate because the latter is more a measure of income inequalities than a measure of economic performance. In other words, macro economic factors, which can be considered from a theoretical point of view to be associated with poverty, can not be linked with the individual poverty status such as defined when using the official definition. Note that this definition of a European at-risk-of-poverty threshold is supported by Marlier et al. (2007).

We then determine, for each country, the fraction of individuals⁹ situated below this new European threshold (see Appendix 1, where these figures can be compared with the official at-risk-of-poverty rates, based on the national thresholds).

3. Methodology

Two major approaches have been used to study the determinants of poverty. One consists in explaining the transitions into and out of poverty (probability of staying poor, and probability of entering poverty). The second approach focuses on the poverty status at a specific point in time.

The first approach takes into account the initial conditions problem¹⁰ by using longitudinal data. This problem refers to the fact that the poverty status during the first period may not be exogenous because of observed and unobserved characteristics, which would affect the probability of being poor afterwards. However, papers running that kind of analysis do not introduce macro factors in the analysis. This is either because they are interested in only one country (see Cappellari and Jenkins, 2002, 2004; Van Kerm, 2004; Buddelmeyer and Verick, 2007; Ayllon, 2008¹¹), or because the different countries are treated separately, with as many models as there are countries (Andriopoulou et al., 2008).

On the other hand, some authors estimate the probability of being poor at a specific point in time (see Wiepking and Maas, 2005; Brady et al., 2009; Tai and Treas, 2008). All of these authors use cross-sectional data of 22 countries from the Luxembourg Income Study (LIS) and integrate macro variables in the analysis (such as the unemployment rate or the welfare generosity), stressing that the welfare system could play a role in allowing individuals to escape from poverty¹².

As for our own work, it focuses on the poverty status and integrates macro factors as well. Our method is original in two ways: first, it uses longitudinal data and second, it takes into account the fact that some variability can be found at the regional level. In fact, two reasons have led us to do the analysis at the regional rather than at the country-level: first because the situation the individuals face (in terms of unemployment rate for example) could be very different from one region to another, within the same country, and second because there are more regions than countries (93 versus 26), which is better from a statistical point of view¹³.

(1999) and Cappellari (2004) on earning, Poggi (2007) on social exclusion persistence.

⁹ These figures concern working age individuals (aged 25-55) because the behaviours and thus the factors at work can be very different for both other groups (children and retired people).

Some authors also control for the retention probability (see for example Cappellari, 2002; Cappellari and Jenkins, 2002, 2004; Ayllon, 2008). The idea is that the probability of being observed during two consecutive years could depend on unobserved characteristics of the individuals that should thus be controlled for.

11 Other authors have used the same kind of models, on related subjects but not poverty: Stewart and Swaffield

Moller et al. (2003) work on these data as well, but at the macro level. Indeed, they link the national at-riskof-poverty rate to macro variables such as GDP or the employment rate in the agricultural sector. Their sample is quite small (61 observations, nested in 14 countries).

While having 93 higher-level groups is technically much better than using only 26 higher-level groups, especially as far as variances/covariances estimation is concerned, a problem remains here: regions are nested within countries. This suggests to move from a three-level analysis (observations over time nested within individuals, themselves nested within regions) to a four-level analysis (adding the country-level at the top of the

In other words, we estimate a model of poverty probability, using two years of observations for each individual (in order to increase the estimation accuracy). Some of these individuals live in the same region ¹⁴. From an econometric point of view, this data setup leads to a problem concerning the independence of observations: individuals being observed over two years and/or living in the same region share their own time-invariant characteristics and/or the characteristics of the area and can therefore no longer be considered to be independent. As a consequence, using traditional techniques would give consistent estimates but heavily ¹⁵ under-estimated standard errors.

In order to cope with that statistical problem, we have chosen one of the many available techniques: we run a multilevel model, which treats the upper levels (the individuals and the regions here) not as unique entities but as units primarily characterised by factors calculated at their level (e.g. characteristics of the individual, or the unemployment rate of the region). These models explicitly take into account the hierarchical structure of the data, thereby allowing us to analyse — first to measure, then to explain — the fraction of the variability of the poverty rate which is attached to each (nested) level. Contrary to the fixed effects models¹⁶, multilevel models make use of the between variance, and are therefore especially useful when this variance is quite high. Some authors have already stressed that the use of this kind of models would be relevant in this framework (Cappellari and Jenkins 2004; Brady et al. 2009) but they have underlined the complexity of these models, whose convergence status is often out of reach.

The model we estimate is a binary logistic regression, where the probability of being at-risk-of-poverty is explained. This multilevel model takes into account three levels: time (measured in years), individuals and regions. It can be written as follows in its structural form ¹⁷:

$$logit(P_{ijk}) = \beta_{0k} + \beta_{1k} x_{1ijk} + \beta_2 x_{2ijk}$$

$$\beta_{0k} = \gamma_0 + \gamma_1 z_{0k} + U_{0k}$$

$$\beta_{1k} = \delta_0 + \delta_1 z_{1k} + U_{1k}$$

hierarchy). But models with four levels and so many observations prove to have convergence problems and can thus not be estimated. We have therefore chosen to restrict our analysis to three levels, and to consider the third as being either the regions (in which case the country variable is introduced in the model as an explanatory variable) or the countries (the regional level being then put aside – see Appendix 2).

Some of them also live in the same household. Ideally, this fact should be taken into account as well, but again it would increase the complexity of the analysis: we would have to deal with 3 sources of non independence of the observations (individuals observed several years, living in the same household, and in the same region). In this paper, we cope with two of them.

According to what Angrist and Pischke (2009) name "the Moulton factor" (p. 310), we calculate to what extent the macro factors regression coefficients would be overestimated by ignoring intraregional correlation. We use the general formula 8.2.5 (ibidem p. 311) which allows for various cluster sizes. With an intra-class correlation coefficient (given by the empty model – see Appendix 6) of 0.551, a (non weighted) average regional sample size of 3206.20 individuals, a (non weighted) variance of regional sample size amounting to 15,550,017.33 and an intra-class correlation of macro factors equal to 1 by definition, we get 66.63. Note that this very large impact of clustering on standard errors is here due to the conjunction of big discrepancies between the size of regions (from 19 to 19941) and a high intra-class correlation coefficient (0.551). Even if the Moulton formula should be considered here only as a very rough approximation (because this formula ignores weighting and repeated observations), it clearly suggests that clustering effects must be taken into account.

¹⁶ Yet these models have a strong advantage: they control for group-invariant factors, measured or unmeasured. But this advantage has a price: the inability to estimate regression coefficients for these group-invariant factors, and thus to allow the analyst to conclude in terms of the effect of these factors.

¹⁷ The notation we use here is the same as Snijders and Bosker's (2004).

where:

i indexes time

i indexes the individuals

k indexes the regions

 P_{ijk} is the probability of being at-risk-of-poverty

 x_{1ijk} is a vector of independent factors defined at the individual/household level whose effects are assumed to be random

 x_{2ijk} is a vector of independent factors defined at the individual/household level whose effects are assumed to be fixed

 z_{0k} is a vector of independent factors defined at the regional level, which are supposed to have an impact on the average P_{ijk} in region k

 z_{1k} is a vector of independent factors defined at the regional level, which are supposed to moderate the effect of the x_{1ijk} on P_{ijk}

 β_{0k} is a random intercept

 β_{1k} is a vector of random slopes

 β_2 is a vector of fixed slopes

 γ_0 measures the average value of P_{ijk} across regions, when each independent variable is 0

 γ_1 measures the impact of z_{0k} on P_{ijk}

 δ_0 measures the average impact, across regions, of x_{1ijk} on P_{ijk} , when each z_{1k} is 0

 δ_1 measures the impact of z_{1k} on the effect of the x_{1ijk} on P_{ijk}

 U_{0k} and U_{1k} are error terms assumed to follow a multinormal distribution $\mathbb{N}(0,0;\Omega)$, Ω being the variance-covariance matrix¹⁸.

The reduced form is thus:

$$logit(P_{ijk}) = \gamma_0 + \gamma_1 z_{0k} + \delta_0 x_{1ijk} + \delta_1 z_{1k} x_{1ijk} + \beta_2 x_{2ijk} + U_{0k} + U_{1k} x_{1ijk}$$

This formula refers to a random slope model, meaning that the intercept and at least one of the explanatory variables have a random coefficient.

4. Data

The EU-SILC longitudinal dataset provides information at both individual and household levels and covers at most 5 years (from 2003 to 2007) depending on the country: data are not available for some countries in 2003, 2004 and 2007. Had we used all five waves to calculate the poverty threshold as we define it (i.e. at the European level), it would have increased or decreased over time just because some countries (e.g. Germany) are absent for some years – it means, without any link with the economic situation. As a consequence, we work with data from two waves only, 2005 and 2006, where all 26 countries are present.

The unit of analysis is the individual: as stated by an OECD report (2001), this is the usual choice for poverty analysis with longitudinal data because individuals can be followed

¹⁸ In our main model, we specify an unstructured form of the variance-covariance matrix (allowing covariances between random effects to be non zero) because the covariance between the error terms of the intercept and the number of employed people in the household appears to be highly significant.

over time whereas households cannot¹⁹. The sample contains 131 891 working age adults (25-55) for the first wave, and 166 379 for the second²⁰, split between 26 countries (see Appendix 3). These countries²¹ are in turn divided into 93 regions (see Appendix 4).

The explanatory variables have been chosen in order to control for different determinants of the poverty status. Some are related to the demographic characteristics of the household ²² (number of children and number of adults), others to the labour market (presence of at least one adult with an upper level of education, number of employed people), others still to the health status (presence²³ of at least one adult with chronic disease, or hampered by illness in his/her daily activities). Two additional variables are measured at the regional level: the GDP and the unemployment rate²⁴.

Some descriptive statistics, for both the whole sample and each country, can be found in Appendix 5 (but not for each region: as there are 93 regions²⁵, it proves not to be sensible to show the descriptive statistics for each of them).

Beyond the usual hypotheses concerning all the control variables²⁶, we make two further ones concerning our variables of interest: first we assume that the negative effect of the level of education on the probability of being at-risk-of-poverty could be weaker in richer areas (where the probability of being poor is quite low, whatever the level of education). Second we assume that the negative effect of the number of employed people in the

¹⁹ As usual, the poverty threshold is calculated at the individual level but the poverty status is defined at the household level.

²⁰ Note that the sample is not balanced: 41 % of the individuals have only one observation (only 5% in Denmark, up to 59% in Czech Republic).

²¹ The latest release of SILC longitudinal data (August 2009) contains only 22 countries, Germany, Ireland, Greece and Denmark being absent. But we wanted both to work on this release, the data of which have been cleared of previous problems, and to keep these four countries in the analysis. We have therefore added to these 22 countries the other 4 from the previous release (March 2009).

²² All of the explanatory variables, apart from gender and age, are calculated at the household level, because the poverty status is defined at that level.

²³ Several variables have originally been defined at the individual level (e.g. having a chronic disease). In order

²³ Several variables have originally been defined at the individual level (e.g. having a chronic disease). In order to have all variables defined at the same (household) level, we have tried to build aggregated variables such as the number of adults in the household suffering from a chronic disease. Unfortunately, this information was available only for one individual per household in the register countries (which use information from administrative datasets when available and interview only one individual per household for the remaining questions to be asked). Keeping this kind of definition would thus have led to a big loss of information. As a consequence, we have defined a much less precise indicator, such as "presence in the household of at least one adult suffering from a chronic disease". Note that even this imprecise indicator could be ill measured in those register countries, as only one household member is interviewed (and the construction of the variable would then only rest on that member). The level of education is defined for all household members aged 16 or more, but the variable contains a lot of missing values in some countries (13% in Portugal, 14% in Spain and up to 16% in the UK in 2005). We have thus adopted the same definition in order to construct a variable at household level.

²⁴ Unfortunately, we were not able to integrate in our main model potential important macro factors such as the social expenditures (expressed as percentages of the GDP): unemployment compensation, public health expenditures, expenditures with respect to inclusion. Indeed, they are not yet known at regional level.

²⁵ In fact, 16 countries out of the 26 countries at hand only have one region, either because they are quite small (12 of them), or because the region code is not available in the dataset (4 of them).

²⁶ Note that income (and thus the poverty status) and individual/household demographic characteristics have not been measured at the same time: income refers to the year prior to the survey, whereas demographic characteristics to the time of the survey. We could have dealt with this by lagging all non-income variables (e.g. demographic characteristics given for year 2006 should be linked to the income declared in 2007), yet at the price of losing some countries since not all of them have available data for 2007.

household on the probability of being at risk of poverty could be attenuated when the unemployment rate is high due to the downward pressure on wages.

5. Results and Comments

The results of the model are shown in Table 1. All our analyses (descriptive and econometric) use weighted data 27 .

²⁷ Moon and Stotsky (1993) state that, if the data come from a stratified and clustered random sampling, it is reasonable to treat the sample as a simple random sample, thus ignoring weights. And Poggi (2007) states that it is more efficient, from an econometrical point of view, not to weigh the data. That way of doing has been adopted by studies on panel data (as stated by Ayllon, 2008, or Andriopoulou et al., 2008). But the question of weighting the data is still open). We have chosen to weigh them, as the SILC dataset is known not to be representative of the population.

Table 1. Probability of being at-risk-of-poverty in 93 European regions. Estimation with a multilevel model²⁸

Table 1. Probability of being at-risk-of-poverty in 93 European region	ons. Estimation with a m	nultilevel model ²⁸	
	Parameter estimate	Standard error	Odds Ratios
Variables ²⁹			
intercept	0,2625	0,3806	
woman	0,01578	0,0139	1,016
age centered (around the average : 41.27)	0,002428 *	0,000848	1,002
age centered squared	0,000453 ***	0,000105	1,000
chronic disease in the household	-0,1214 ***	0,01692	0,886
activity hampered by disease in the household	0,1612 ***	0,01754	1,175
upper education level in the household	-1,5241 ***	0,1568	0,218
number of children in the household	0,3555 ***	0,01532	1,427
number of children in the household squared	-0,0127 *	0,004298	0,987
number of adults in the household	-0,2607 ***	0,02786	0,770
number of adults in the household squared	0,08983 ***	0,004659	1,094
number of employed people in the household	-1,2964 ***	0,07399	0,274
regional annual GDP per capita (in 10 ³ Euros)	-0,04818 ***	0,007353	0,953
regional unemployment rate (expressed in %)	0,02103	0,01122	1,021
upper education level * regional GDP per capita	0,02427 **	0,006689	,
number of employed people * regional unemployment rate	0,01392 *	0,005996	
wave 2005	-0,2055 ***	0,01488	0,814
wave 2006	ref.	ref.	ref.
country BE	-0,3208	0,3616	0,726
country CZ	1,9401 ***	0,3278	6,960
country DK	-0,6816	0,4419	0,506
country DE	-0,2740	0,4252	0,760
country EE	3,1916 ***	0,4460	24,328
country IE	-0,2453	0,4468	0,782
country EL	0,9511 *	0,3431	2,589
country ES	0,9605 *	0,3115	2,613
country FR	-0,2864	0,3107	0,751
country IT	0,2918	0,3311	1,339
country CY	-0,2265	0,4889	0,797
country LV	3,5168 ***	0,4476	33,678
country LT	3,7270 ***	0,4445	41,556
country LU	-0,3031	0,7205	0,739
country HU	3,0947 ***	0,3581	22,081
country NL	-0,2287	0,4267	0,796
country AT	-0,3297	0,3566	0,719
country PL	2,5769 ***	0,3475	13,157
country PT	1,4376 **	0,4321	4,210
country SI	0,1796	0,4516	1,197
country SK	3,2642 ***	0,4415	26,160
country FI	-0,5984	0,3536	0,550
country SE	-0,3938	0,4318	0,675
country UK	ref.	ref.	ref.
country IS	-0,4673	0,6692	0,627
country NO	-0,4939	0,4646	0,610
Regional-level error terms variances	-,	-,	-,010
intercept	0,3818	0,08275	
upper education level in the household	0,2077	0,04486	
number of employed people in the household	0,2077	0,03756	
Regional-level error terms covariances	0,2170	0,03730	
COV (intercept, upper education level in the household)	-0,04609	0,04281	
COV (intercept, upper education level in the household) COV (intercept, number of employed people in the HH)	-0,2342	0,05093	
COV (upper education level, nb. of employed people in the HH)	-0,2342	0,03109	
Other parameters	-0,0437	0,03109	
Rho coefficient of AR(1)	0,3138	0,002817	
Residual	0,9936		
	1832580	0,002727	
Fit measure: -2 Log Pseudo Likelihood	authors' computations		

Source: EU-SILC data, longitudinal file ³⁰, 1.08.2009 UDB release, authors' computations.

Level of significance for independent variable coefficients: *: p-value < 0.05; **: p-value < 0.01; ***: p-value < 0.001

We have used the SAS GLIMMIX command. Useful SAS code examples can be found in Allison (2008) and were kindly given by David Brady.

See appendix 5 for a description of the variables.

22 countries from this release plus 4 countries from the March release (see above).

Concerning the error terms, SAS/PROC GLIMMIX does not offer a statistical test indicating the level of significance of the variances and covariances of the error terms. But compared to their standard errors, the estimated variances are quite high, which suggests their high level of significance. This in turn justifies on the one hand the choice of the multilevel model, and on the other hand our choice of allowing these variables to have random rather than fixed coefficients. Looking at the empty model (see Appendix 6), we can see that the intra-class correlation (calculated according to the second formula given by Snijders and Bosker, 1999, page 224) is equal to 0.55, meaning that the between-variance is substantial.

Let us examine now the effects of our variables of interest. Recall that our objective is to measure the specific effect of the regional GDP per inhabitant and the regional unemployment rate on the probability of being at-risk-of-poverty. This effect could be either direct or indirect since these macro factors can act through other individual variables on the probability of being poor (such as the education level and the number of employed people in the household).

As expected, the regional GDP per capita has a strong (and highly significant) direct negative effect on the risk of poverty: for individuals living in households where nobody has an upper education level, the odds of being poor (probability of being poor divided by probability of not being poor) decreases by 4.7% (1-0.953=0.047) for an increase of annual GDP per capita by a 1000 Euros. This direct effect is supplemented by an indirect effect: the regional GDP per capita moderates the negative impact of upper education on the poverty risk. In fact, in the average region in terms of GDP per capita (about 24260 Euros/year), the presence of an adult with upper education level decreases the poverty odds by 61% (odds ratio = 0.39^{31}). In a rich region such as Luxembourg (GDP per capita = 60150 Euros/year), it decreases the odds by only 6.2%; in a quite disadvantaged region like Estonia (GDP per capita = 14547 Euros/year) it decreases the odds by 69 %.

For individuals living in households without any employed people, an additional percentage point of the regional unemployment rate increases by 2.1% the poverty odds, but this effect is only slightly statistically significant (p-value=6%) ceteris paribus (especially when GDP per capita is controlled for). But there is an indirect effect of the regional unemployment rate on the poverty risk, even if rather small: in the average region in terms of the unemployment rate (unemployment rate = 8.52 %), the presence of an additional employed individual decreases the poverty odds by 69%, when controlling for the number of adults in the household (odds ratio = 0.31^{32}). When the unemployment rate is much lower, such as in Ireland (unemployment rate 4.35%), it decreases these odds by 71%, and by 66% when the unemployment rate is quite high (for example in Slovenia - unemployment rate 14.85%).

To summarize, the moderating effect of the regional unemployment rate on the impact on poverty risk of the number of employed people does exist but it is marginal. By contrast, the moderating effect of the regional GDP per capita on the impact on poverty risk of the presence of highly educated people is quite large.

³¹ Odds ratio = $0.39 = \exp(-1.5241 + 0.02427*24.260)$

Besides and not surprisingly (given the sample size), almost all control variables have an effect on the poverty probability, except for gender³³: women do not have a higher risk of being poor than men. This can easily be explained by the fact that the poverty status is a household characteristic, which can hardly be influenced by a strictly individual feature³⁴. One interesting question is the extent to which these control variables have an impact on the poverty risk (even if our study focuses on the possible impact of macro determinants on the effect some factors of interest can have on the poverty risk):

- if the activity of at least one household member is hampered by disease, the odds of being poor increase by 18%
- controlling for especially the fact that the activity of at least one household member is hampered (or not) by disease, the chronic character of this disease decreases by 11% the odds of being poor (this could be a consequence of the social benefits people with chronic diseases are entitled to)
- the number of children in the household has a slightly concave effect on the odds of being poor: as the number of children increases, the effect of an additional child decreases progressively up to a value (14 children) lying beyond the observed maximum in the sample (12 children). But the effect of each additional child remains substantial. As an example, while a first child increases by 41% the odds of being poor, a fourth child still increases them by 31%
- the number of adults in the household also proves to have a clear non linear effect on the poverty risk: ceteris paribus (and especially when controlling for the number of employed people), the odds of being poor are virtually the same if there are one or two adults in the household, but they increase by 21% with the third adult and by 45% with the fourth one.

6. Conclusion

Analysing the determinants of the monetary poverty probability has already been attempted by many studies. But few of them have simultaneously used panel data, considered factors at the macro level, and used the right techniques to deal with all these elements.

As for our results, they show that both the regional GDP per capita and the regional unemployment rate do have an effect on poverty risk.

³³ Since the poverty status is defined at the household level, it could be considered as irrelevant to introduce variables measured at individual level such as age and gender. We have added these two variables to our model for comparison's sake, as almost all studies do this as well. As far as age is concerned, it could be argued that, even though age is, like gender, a factor measured at the individual level, it might have some signification as an household characteristic: due to frequent endogamy, the age of one adult belonging to the household offers some clue about the age of other adults in the household, at least for non single adult households. With these limitations in mind, we can interpret the (highly significant) quadratic effect of age: ceteris paribus, the poverty risk is first decreasing more than linearly with age, reaching a minimum at age 44, and then increasing more than linearly. For example, at age 25, an additional year results in a decrease of the poverty risk by 1.2%; and at age 50, an additional year of age results in an increase of 1.1%.

³⁴ This is the very reason why the different common indicators agreed upon in the context of the OMC on Social Protection and Social Inclusion which focus on gender differences are calculated on single person households.

In terms of economic and social policy implications, it means that:

- policies oriented towards higher economic growth rates in disadvantaged European regions are able to alleviate the risk of poverty even if poverty is defined in relative terms
- this kind of economic policies, if successful in its effort to sustain the economic well-being of families in poor regions, will, as a side-effect, diminish the anti-poverty effect of the presence in the household of higher educated people. We suspect that this indirect effect is associated to the choice of defining poverty as a relative concept a European view, which countries like the US do not share
- as for the regional unemployment rate, its direct positive impact on the poverty risk is essentially a confirmation of what could be expected and of what is already known, even if the weakness of this effect is quite surprising. Still more surprising is the fact that the regional unemployment rate does not moderate to a large extent the impact on the poverty risk of the number of employed people in the household. This implies that policies aimed at combating high regional unemployment rates will, as such, unfortunately not lower to a large extent the regional poverty rates.

However, our analysis faces two types of limitations. The first one results from methodological choices we have made, the second is due to the data.

First, because we needed an indicator differentiating and thus ranking the 93 regions in terms of poverty rates, we have made use of a European poverty threshold, which has proven quite relevant in terms of ability to estimate the econometric model. However, precisely because the European regions are quite dispersed around the average at-risk-of-poverty rate, we were not able to check the consistency of the results in using alternative measures of the European threshold³⁵ (such as 50% or 70% of the European median equivalent income).

Concerning then the data, and with the results at country-level in mind, we would have liked to test the effect of other macro characteristics at the regional level, such as the expenses in unemployment or social benefits (expressed in percentage of the GDP). But these were not available at the regional level. In future analyses, we would thus be interested in adding some variables of that kind, once they are available. And, as for the regions, the variable defined in the EU-SILC dataset is missing in some countries (even in quite big countries such as Germany and the United Kingdom). In order to keep these countries in the analysis, we have defined each of them as a unique – and quite large – region.

³⁵ In fact, with a poverty threshold equal to 60% of the European median equivalent income, the at-risk-of-povery rates of the different countries range between 1% and 82% in 2005, and between 2% and 75% in 2006. Changing this threshold for a lower (higher) one would lead to even lower (higher) rates in the richest (poorest) countries, making the analysis impossible to run.

References

ALLISON Paul D. (2008), Longitudinal Data Analysis Using SAS, October, 100 pages.

ANDRIOPOULOU Eirini, **TSAKLOGLOU Panos** (2008), "Once poor, always poor? Do initial conditions matter? Evidence from the ECHP", IRISS Conference, Oct., 33 pages.

ANGRIST Joshua D., PISCHKE Jörn-Steffen (2009), "Mostly Harmless Econometrics. An Empiricist's Companion", Princeton University Press, 373 p.

AYLLON Sara (2008), "Modelling poverty transitions in Spain: Do attrition and initial conditions really matter?", IRISS Working Paper Series, 2008-08, Oct., 36 pages.

BRADBURY Bruce, **JÄNTTI Markus** (1999) "Child Poverty across Industrialised Nations", Innocenti Occasional Papers, Economic and Social Policy Series, UNICEF, n°71.

BRADSHAW Jonathan (1999), "Child Poverty in Comparative Perspective", *European Journal of Social Security*, Vol.1, pp. 383-406.

BRADY David, FULLERTON Andrew S., MOREN-CROSS Jennifer (2009), "Putting Poverty in Political Context: A Multi-Level Analysis of Adult Poverty Across 18 Affluent Countries", *Social Forces*, Vol. 88(1), September, pp. 271-300.

BUDDELMEYER Hielke, **VERICK Sher** (2007), "Understanding the Drivers of Poverty Dynamics in Australian Households", IZA DP No. 2827, 29 pages.

CAPPELLARI Lorenzo (2004), "Earnings Mobility Among Italian Low Paid Workers", *Journal of Population Economics*, Vol. 20(3), pp. 465-482.

CAPPELLARI Lorenzo, **JENKINS Stephen P**. (2002a), "Who Stays Poor? Who Becomes Poor? Evidence from the British Household Panel Survey", *The Economic Journal*, Vol. 112, No. 478, Conference Papers (Mar., 2002), pp. C60-C67.

CAPPELLARI Lorenzo, **JENKINS Stephen P.** (2002b), "Modelling Low Income Transitions", *ISER Working Papers Number* 2002-8, 42 pages.

CAPPELLARI Lorenzo, **JENKINS Stephen P.** (2004), "Modelling Low Income Transitions", *Journal of Applied Econometrics*, Vol. 19, pp. 593–610.

FERTIG Michael, TAMM Marcus (2007), "Always Poor or Never Poor and Nothing in Between? Duration of Child Poverty in Germany", IZA Discussion Paper N° 2645, February, 29 pages.

MARLIER Eric, **ATKINSON A.B.**, **CANTILLON Bea**, **NOLAN Brian** (2007), The EU and Social Inclusion – Facing the challenges, The Policy Press, The University of Bristol, 303 pages.

MEJER Lene, **SIERMANN Clemens** (2000), "La pauvreté monétaire dans l'Union européenne : la situation des enfants, les différences entre les sexes et l'écart de pauvreté", Eurostat, *Statistiques en bref*, Population et Conditions sociales, 12/2000, 7 pages.

MOLLER Stephanie, HUBER Evelyne, STEPHENS John D., BRADLEY David, NIELSEN François (2003), "Determinants of Relative Poverty in Advanced Capitalist Democracies", *American Sociological Review*, Vol. 68, No. 1, February, pp. 22-51.

MOON Choon-Geol, STOTSKY Janet G. (1993), "The Effect of Rent Control on Housing Quality Change: A Longitudinal Analysis", *The Journal of Political Economy*, Vol. 101, No. 6, December, pp. 1114-1148.

OECD (2001), "When Money is Tight: Poverty Dynamics in OECD Countries", in *Employment Outlook*, Chapter 2, pp. 37-87.

POGGI Ambra (2007), "Does persistence of social exclusion exist in Spain?", *Journal of Economic Inequality*, Vol. 5, pp. 53-72.

RAY Jean-Claude, **JEANDIDIER Bruno** (2003), "Depuis une décennie, que nous ont appris les données longitudinales à propos de la pauvreté en France? Une première synthèse", Communication aux X^{es} Journées du Longitudinal, Caen, mai, 61 pages.

SNIJDERS Tom A.B., **BOSKER Roel J.** (2004), *Multilevel Analysis*. *An introduction to basic and advanced multilevel modeling*, Sage Publications, 266 pages.

STEWART Mark B., **SWAFFIELD Joanna K.** (1999), "Low Pay Dynamics and Transition Probabilities", *Economica*, Vol. 66, pp. 23-42.

TAI Tsui-o, **TREAS Judith** (2008), "Poverty, Household Composition, and Welfare States: A Multi-level Analysis of 22 Countries", Luxembourg Income Study Working Paper Series, Working Paper No. 492, September, 33 pages.

UNICEF (2000), "A League Table of Child Poverty in Rich Nations", Innocenti Report Card, Issue n°1, June, 28 pages.

VAN KERM Philippe (2004), "Une évaluation économétrique des flux vers et hors de la pauvreté en Belgique", IRISS Working Paper Series, 2004-04, 17 pages.

WIEPKING Pamala, **MAAS Ineke** (2005), "Gender Differences in Poverty: A Cross-National Study", *European Sociological Review*, Vol. 21(3), July, pp. 187-200.

Appendix 1. At-risk-of-poverty rates of working aged adults (25-55) in Europe

At-risk-of-poverty rates of working age adults (25-55) in Europe (European threshold = 60% of the European median equivalent income)

Country	2005	2006
AT	5	6
BE	6	8
CY	7	6
CZ	42	41
DE	7	6
DK	3	3
EE	70	64
ES	20	19
FI	5	6
FR	8	9
EL	24	26
HU	73	69
IE	8	8
IS	3	3
IT	14	17
LT	79	75
LU	1	2
LV	81	74
NL	6	3
NO	3	3
PL	74	73
PT	42	40
SE	5	5
SI	13	13
SK	77	72
UK	7	8

Source: EU-SILC data, longitudinal file, 1.08.2009 UDB release, authors' computations.

Reading note: with the European poverty threshold calculated for the whole population ³⁶, 5% of individuals aged 25-55 in Austria were atrisk-of-poverty in 2005.

We can notice that the countries face very different situations in terms of the at-risk-of-poverty rate, a conclusion that cannot be drawn from the figures shown in the table below (based on national poverty thresholds).

The

³⁶ The European poverty threshold is calculated by taking into account all individuals living in the 26 countries under study. In other words, children and elderly people are not excluded from this calculation, even if they are not kept in the analyses afterwards.

At-risk-of-poverty rates of working aged adults (25-55) in Europe (national thresholds)

Country	2005	2006
AT	11	11
BE	11	11
CY	10	10
CZ	10	9
DE	11	12
DK	9	9
EE	16	15
ES	16	15
FI	8	9
FR	10	11
EL	16	17
HU	14	15
IE	14	14
IS	9	8
IT	16	17
LT	19	18
LU	13	14
LV	18	19
NL	10	8
NO	8	8
PL	21	19
PT	15	15
SE	7	10
SI	10	9
SK	13	11
UK	14	14

Source: EU-SILC data, cross-sectional file, 1.08.2009 UDB release. Reading note: with the poverty threshold calculated at national level, 11% of individuals aged 25-55 in Austria were at-risk-of-poverty in 2005.

These official figures are based on the cross-sectional file, which means that all individuals are taken into account (whereas the longitudinal file concerns only those present at least two years).

Appendix 2. Poverty probability determinants at country-level

		Response	Profile			
	Ordered			Total		
	Value	pov_indi 1	icator F	requency 72218		
	2			211983		
The G	LIMMIX procedure is mo	deling the p	probability	that pov_i	ndicator=	'1'.
		Dimens				
		Cov. Parame		3 2		
		s in X	sters	21		
		s in Z per S	Subject	3		
		ts (Blocks i		26		
	Max Ob	s per Subjec	ST.	30703		
		ptimization n Technique				
		in Optimizat		Jon-Kapiisoi	1	
	Lower Bound	-	4			
	Upper Bound		1	511.1		
	Fixed Effec Residual Va		Pro:			
	Starting Fr			estimates		
/ 32 iterations		torion (DCO)	TT-1 11000F	0) antiafi	ad	
	Convergence cri	Fit Stat		-0) Satisii	.ea.	
		g Pseudo-Li				
		ed Chi-Squar		291051.3		
	Gener. Cn	i-Square / I)F	1.02		
	Cova	riance Param	neter Estima	ates	Q1 1 1	
	Cov Parm Sub	ject	E.s	stimate	Standard Error	
		ntry		0.3242	0.1091	
		ntry		0.2754 0.1406	0.08766	
	nbemployedHH cou AR(1) ID	ntry unique_UE27	(country)	0.1406	0.04498	
	Residual	dirique_obz,	(councry)	1.0242	0.002770	
	Asymptotic Covarianc	o Matrix of	Comariance	Paramotor	Estimatos	
Cov Parm	Subject	CovP1	CovP	20.7	7P3 (CovP4 CovP5
Intercept	country	0.01190	0.000089	-0.000	19 -6.9	96E-7 -5.7E-7
uppereducHH nbemployedHH	country	0.000089	0.007684	0.0000	132 -6.5	59E-7 -3.03E-7
AR(1)	country ID_unique_UE27(countr	v) -6.96E-7	-6.59E-	7 3.937E	124 3.90 1-7 7.70	07E-6 2.001E-6
Residual	_ • _ •	-5.7E-7	-3.03E-	7 2.3091	2.00	01E-6 7.903E-6
	Asymptotic Correlatio	n Matrix of	Covariance	Parameter	Estimates	
Cov Parm	Subject	CovP1		2 Cov	7P3 (CovP4 CovP5
Intercept uppereducHH	country country	1.0000	0.009333	0.03	776 -0.0	J0230 -0.00186
nbemployedHH	country	-0.03776	0.008229	9 1.00	0.00	00230 -0.00186 00271 -0.00123 03152 0.000183
AR(1)	ID_unique_UE27(countr	y) -0.00230	-0.00271	L 0.0031	.52 1.	.0000 0.2564
Residual		-0.00186	-0.00123	3 0.0001	.83 0.	.2564 1.0000
	Sol	utions for H		s		
Effect	wave	Estimate	Standard Error	DF	t Value	Pr > t
Intercept	wave	1.4742	0.4733	25	3.11	0.0046
wave	2005	-0.1846	0.02712	284E3	-6.81	<.0001
wave GDPhabnuts	2006	0 -0.1103	0.01437	284E3	-7.68	<.0001
GDPNabhuts	on	0.1560	0.01437	284E3	7.97	<.0001
unempraten		0.1386	0.01449	284E3	9.56	<.0001
socexclexp		-0.7600	0.1292	284E3	-5.88	<.0001
unemployme: woman	ntexpenses	-0.2225 0.01188	0.1042 0.01397	284E3 284E3	-2.14 0.85	0.0327 0.3949
age_center	ed	0.001467	0.000852	284E3	1.72	0.0849
age_center	ed2	0.000469	0.000105	284E3	4.46	<.0001
chronicdis		-0.1419	0.01692 0.01753	284E3 284E3	-8.38 9.81	<.0001 <.0001
activityha uppereducH		0.1720 -1.2942	0.01753	284E3 25	-9.54	<.0001
socexclexp	enses*uppereducHH	0.7339	0.1826	284E3	4.02	<.0001
nbchildren		0.3795	0.01529	284E3	24.82	<.0001
nbchildren nbadultsHH		-0.01734 -0.2065	0.004273 0.02767	284E3 284E3	-4.06 -7.46	<.0001 <.0001
nbadultsHH		0.08637	0.004631	284E3	18.65	<.0001
nbemployed	НН	-1.0282	0.1132	25	-9.08	<.0001
unemployme	ntexpenses*nbemployedH	н -0.1153	0.06613	284E3	-1.74	0.0812

Appendix 3. Sample size, by country and year

Number of individuals aged 25-55, in each country, for each year (sample size, unweighted cases)

unweighted cases)								
Country	2005	2006						
AT	3882	5469						
BE	2408	4002						
CY	2357	3373						
CZ	4391	7367						
DE	10529	9285						
DK	3549	3451						
EE	2665	4269						
ES	8472	11501						
FI	4021	5528						
FR	5876	7324						
EL	4642	4211						
HU	4353	6609						
IE	2872	1992						
IS	1584	2243						
IT	12955	18048						
LT	2533	3713						
LU	4322	4601						
LV	2704	3438						
NL	8574	9287						
NO	4088	3926						
PL	10627	14644						
PT	2671	3662						
SE	3541	4751						
SI	8671	11271						
SK	3447	4898						
UK	6224	7594						
C FILCH C 1 4 1 '4	ourse; EU SU C data longitudinal file 1 09 2000 UDD release authors' computations							

Source: EU-SILC data, longitudinal file, 1.08.2009 UDB release, authors' computations.

Appendix 4. List of available regions³⁷ in the dataset

Austria		ES62	Región de Murcia
AT1	Ostösterreich	ES63	Ciudad Autónoma de Ceuta
AT2	Südösterreich	ES64	Ciudad Autónoma de Melilla
AT3	Westösterreich	ES70	Canarias
Belgium		Finland	
BE1	Région de Bruxelles-Capitale /	FI13	Itä-Suomi
	Brussels Hoofdstedelijk Gewest	FI18	Etelä-Suomi
BE2	Vlaams Gewest	FI19	Länsi-Suomi
BE3	Région Wallonne	FI1A	Pohjois-Suomi
Republic	of Cyprus	France	
$\overrightarrow{\text{CY0}}$	Kypros / Kibris	FR10	Île de France
		FR21	Champagne-Ardenne
Czech Re	public	FR22	Picardie
CZ01	Praha	FR23	Haute-Normandie
CZ02	Stredni Cechy	FR24	Centre
CZ03	Jihozapad	FR25	Basse-Normandie
CZ04	Severozapad	FR26	Bourgogne
CZ05	Severovychod	FR30	Nord - Pas-de-Calais
CZ06	Jihovychod	FR41	Lorraine
CZ07	Stredni Morava	FR42	Alsace
CZ08	Moravskoslezsko	FR43	Franche-Comté
		FR51	Pays de la Loire
Germany		FR52	Bretagne
	n variable is not available in the	FR53	Poitou-Charentes
	nal dataset for Germany.	FR61	Aquitaine
DE	Germany	FR62	Midi-Pyrénées
22	Commany	FR63	Limousin
Denmark		FR71	Rhône-Alpes
DK0	Denmark	FR72	Auvergne
DIO	Dominark	FR81	Languedoc-Roussillon
Estonia		FR82	Provence-Alpes-Côte d'Azur
EE0	Estonia	FR83	Corse
LLO	Listoma	1103	Corsc
Spain		Greece	
ES11	Galicia	GR1	Voreia Ellada
ES12	Principado de Asturias	GR2	Kentriki Ellada
ES13	Cantabria	GR3	Attiki
ES21	País Vasco	GR4	Nisia Aigaiou, Kriti
ES22	Comunidad Foral de Navarra		
ES23	La Rioja	Hungary	
ES24	Aragón	HU1	Kozep-Magyarorszag
ES30	Comunidad de Madrid	HU2	Dunantul
ES41	Castilla y León	HU3	Alfold Es Eszak
ES42	Castilla-La Mancha		
ES43	Extremadura	Ireland	
ES51	Cataluña		ariable is not available in the
ES52	Comunidad Valenciana	longitudinal	dataset for Ireland.
ES53	Illes Balears	IE0	Ireland
ES61	Andalucía		
		Iceland	
		IS	Iceland

Details from EU-SILC documentation and http://ec.europa.eu/eurostat/ramon/nuts/codelist_en.cf m?list=nuts

Italy **Poland** ITC Nord-Ovest PL1 Region Centralny Region Poludniowy PL2 ITD Nord-Est Region Wschodni ITE Centro (I) PL3 Region Polnocno-Zachodni Region Poludniowo-Zachodni ITF Sud PL4 ITG Isole PL5 PL6 Region Polnocny

Sweden

Lithuania

LT0 Lietuva **Portugal** PT Portugal

Luxembourg

LU0 Luxembourg (Grand-Duché) Sweden SE

Latvia

LV0 Latvija Slovenia SI Slovenia

The Netherland

The region variable is not available in the Slovakia longitudinal dataset for the Netherlands. SK0 Slovenska

The Netherlands NL

The United Kingdom

The region variable is not available in the Norway NO0

longitudinal dataset for UK Norway

United-Kingdom UK

Appendix 5. Description of the explanatory variables and descriptive statistics

Description of the explanatory variables

Description of the expi	-	T
Name of the variable	Label of the variable	Description of the variable
woman	woman	EU-SILC variable RB090; woman=1 if RB090=2
age_centered	age centered (around the	EU-SILC variable RX020, centered (age in the year
	average: 41.27)	prior to the survey)
chronicdiseaseHH	chronic disease in the	Authors' calculations using the EU-SILC variable
	household	PH020: is there at least one household member who
		suffers from a chronic disease?
activityhamperedHH	activity hampered by disease in	Authors' calculations using the EU-SILC variable
	the household	PH030: is there at least one household member
		whose activities are hampered because of health
		problems?
иррегедисНН	upper education level in the	Authors' calculations using the EU-SILC variable
	household	PE040: is there at least one household member
		whose upper level of education is tertiary education
		(PE040=5)?
nbchildren	number of children in the	number of children (age 0-14) in the year prior to the
	household	survey
nbadultsHH	number of adults in the	number of adults (age 18 or more) in the year prior to
	household	the survey
nbemployedHH	number of employed people in	number of employed household members (authors'
	the household	calculations using the EU-SILC variable PL030 –
		codes 1 or 2)
country	country	EU-SILC variable RB020
wave	wave	EU-SILC variable RB010
GDPhabnuts	regional annual GDP per capita	Information from Eurostat
	$(in 10^3 Euros)$	
unempratenuts	regional unemployment rate	Information from Eurostat
	(expressed in %)	

Descriptive statistics for the whole sample

Wave	N Obs	Variable	N	NMiss	Mean	Std Dev	Minimum	Maximum
2005	131891	pov_indicator	131891	0	0.1965	0.4259	0	1.0000
		woman	131887	4	0.5130	0.5357	0	1.0000
		age centered	131891	0	-0.6425	9.1801	-16.2665	14.7335
		age centered2	131891	0	73.7935	76.2358	0.0710	264.6
		chronicdiseaseHH	131891	0	0.3885	0.5223	0	1.0000
		activityhamperedHH	131891	0	0.2931	0.4878	0	1.0000
		uppereducHH	126795	5096	0.3962	0.5182	0	1.0000
		nbchildren	131891	0	0.9077	1.1426	0	11.0000
		nbadultsHH	131891	0	2.3903	1.0792	1.0000	10.0000
		nbemployedHH	129384	2507	1.5131	0.8971	0	8.0000
		GDPhabnuts	131891	0	23.7538	7.1961	8.2000	57.1000
		unempratenuts	131891	0	8.8728	4.4381	2.5000	21.4000
2006	166379	pov_indicator	166379	0	0.2000	0.3757	0	1.0000
		woman	166371	8	0.5128	0.4695	0	1.0000
		age_centered	166379	0	-0.2041	7.9962	-16.2665	14.7335
		age_centered2	166379	0	72.5178	66.5149	0.0710	264.6
		chronicdiseaseHH	166379	0	0.3911	0.4584	0	1.0000
		activityhamperedHH	166379	0	0.2926	0.4273	0	1.0000
		uppereducHH	158892	7487	0.4059	0.4560	0	1.0000
		nbchildren	166379	0	0.9288	0.9917	0	12.0000
		nbadultsHH	166379	0	2.4239	0.9506	1.0000	11.0000
		nbemployedHH	162220	4159	1.6016	0.7713	0	8.0000
		GDPhabnuts	166379	0	24.7924	6.6164	8.7000	63.1000
		unempratenuts	166379	0	8.1568	3.0558	2.8000	21.0000

Descriptive statistics for at country level

wave	country	N Ob	s Variable		N N Miss	Mean	Std De	v Minimum	Maximum
2005	BE	2408	pov_indicator	2408	0	0.0609	0.2696	0	1.0000
			woman	2408	0	0.4920	0.5635	0	1.0000
			age_centered	2408	0	0.2142	9.5775	-16.2665	14.7335
			age_centered2	2408	0	72.2531	78.0715	0.0710	264.6
			chronicdiseaseHH activityhamperedHH	2408 2408	0	0.3267 0.3100	0.5286 0.5213	0	1.0000
			uppereducHH	2089	319	0.5100	0.5609	0	1.0000
			nbchildren	2408	0	0.9201	1.3651	0	7.0000
			nbadultsHH	2408	0	2.3255	1.1521	1.0000	7.0000
			nbemployedHH	2408	0	1.4716	0.9269	0	5.0000
			GDPhabnuts unempratenuts	2408 2408	0	26.6319 8.4371	10.0738	19.4000 5.4000	53.3000 16.3000
2006	BE	4002	nor indicator	4002	0	0.0789	0.2398	0	1.0000
2006	DE	4002	pov_indicator woman	4002	0	0.4979	0.2396	0	1.0000
			age centered	4002	0	0.0482	7.6695	-16.2665	14.7335
			age_centered2	4002	0	74.3560	62.3554	0.0710	264.6
			chronicdiseaseHH	4002	0	0.3407	0.4215	0	1.0000
			activityhamperedHH	4002	0	0.3010	0.4080	0	1.0000
			uppereducHH nbchildren	3634 4002	368 0	0.5390 0.8937	0.4439 1.0366	0	1.0000 7.0000
			nbadultsHH	4002	0	2.2926	0.8868	1.0000	7.0000
			nbemployedHH	3969	33	1.5155	0.7240	0	4.0000
			GDPhabnuts	4002	0	28.0513	8.4301	20.1000	55.1000
			unempratenuts	4002	0	8.3091	3.8080	5.0000	17.6000
2005	CZ	4391		4391	0	0.4156	0.4304	0	1.0000
			woman age centered	4391 4391	0	0.4986 -1.3598	0.4366 8.1510	0 -16.2665	1.0000 14.7335
			age_centered2	4391	0	88.9707	65.7152	0.0710	264.6
			chronicdiseaseHH	4391	0	0.3762	0.4230	0	1.0000
			activityhamperedHH	4391	0	0.3349	0.4121	0	1.0000
			uppereducHH	4391	0	0.2071	0.3539	0	1.0000
			nbchildren	4391	0	0.7598	0.8045	0	5.0000
			nbadultsHH nbemployedHH	4391 4391	0	2.5584 1.7005	0.8599 0.7878	1.0000	7.0000 4.0000
			GDPhabnuts	4391	0	17.0978	5.9279	13.3000	35.6000
			unempratenuts	4391	0	7.9891	3.1642	3.5000	13.9000
2006	CZ	7367	pov_indicator	7367	0	0.4136	0.3314	0	1.0000
			woman	7367	0	0.5002	0.3364	0	1.0000
			age_centered	7367	0	-1.2929	6.2442	-16.2665	14.7335
			age_centered2 chronicdiseaseHH	7367 7367	0	87.7967 0.3807	49.6113	0.0710	264.6 1.0000
			activityhamperedHH	7367	0	0.3379	0.3183	Ö	1.0000
			uppereducHH	7323	44	0.2118	0.2749	0	1.0000
			nbchildren	7367	0	0.7563	0.6111	0	7.0000
			nbadultsHH	7367	0	2.5852	0.6729	1.0000	8.0000
			nbemployedHH	7367	0	1.7118	0.6076	0	5.0000
			GDPhabnuts unempratenuts	7367 7367	0	18.2129 7.2759	4.8756 2.1418	14.2000 2.8000	38.4000 12.8000
2005	DK	3549	pov indicator	3549	0	0.0331	0.1237	0	1.0000
			woman	3549	0	0.4989	0.3458	0	1.0000
			age_centered	3549	0	-0.1867		-16.2665	14.7335
			age_centered2	3549	0	74.1813	48.3555	0.0710	264.6
			chronicdiseaseHH activityhamperedHH	3549	0	0.2248 0.1249	0.2887	0	1.0000
			uppereducHH	3549 3488	61	0.1249	0.2287	0	1.0000
			nbchildren	3549	0	0.9719	0.7503	0	7.0000
			nbadultsHH	3549	0	1.9228	0.4682	1.0000	7.0000
			nbemployedHH	3491	58	1.5338	0.4816	0	4.0000
			GDPhabnuts unempratenuts	3549 3549	0	27.8000 4.8000	0	27.8000 4.8000	27.8000 4.8000
2006	D.1/	2451	•						
2006	DK	3451	pov_indicator woman	3451 3451	0	0.0320 0.4933	0.1208	0	1.0000
			age centered	3451	0	-0.0156	5.7369	-16.2665	14.7335
			age centered2	3451	0	69.9331	46.3429	0.0710	264.6
			chronicdiseaseHH	3451	0	0.2375	0.2919	0	1.0000
			activityhamperedHH	3451	0	0.1280	0.2292	0	1.0000
			uppereducHH	3391	60	0.4392	0.3390	0	1.0000
			nbchildren nbadultsHH	3451 3451	0	0.9772 1.9430	0.7343 0.4724	1.0000	5.0000 5.0000
			nbadultshh	3382	69	1.5658	0.4724	1.0000	4.0000
			GDPhabnuts	3451	0	29.1000	0.4010	29.1000	29.1000
			unempratenuts	3451	0	3.9000	0	3.9000	3.9000
2005	DE	10529	pov_indicator	10529	0	0.0748	0.4145	0	1.0000
			woman	10529	0	0.5332	0.7863 13.1238	0 -16.2665	1.0000 14.7335
			age_centered age_centered2	10529 10529	0	-0.2368 69.3959	13.1238	0.0710	264.6
			chronicdiseaseHH	10529	0	0.4474	0.7836	0.0710	1.0000

			activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	10529 10364 10529 10529 10529 10529 10529	0 165 0 0 0 0	0.4231 0.5070 0.7408 2.1375 1.3698 26.3000 11.1000	0.7786 0.7878 1.5855 1.3513 1.2661 0	0 0 0 1.0000 0 26.3000 11.1000	1.0000 1.0000 6.0000 7.0000 5.0000 26.3000 11.1000
2006	DE	9285	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	9285 9285 9285 9285 9285 9285 9285 9285	0 0 0 0 0 0 0 0	0.0624 0.5655 1.8870 63.1927 0.4929 0.3678 0.5429 0.9117 2.2395 1.6796 27.4000 10.2000	0.3895 0.7983 12.4357 108.7 0.8051 0.7766 0.8022 1.6225 1.4165 1.2261	0 0 0 -16.2665 0.0710 0 0 0 1.0000 0 27.4000 10.2000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 6.0000 5.0000 27.4000
2005	EE	2648	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	2648 2648 2648 2648 2648 2648 2648 2648	0 0 0 0 0 0 0 0	0.6987 0.5190 -0.8065 81.8660 0.4745 0.4575 0.4219 0.8302 2.4462 1.6159 13.7000 7.9000	0.1792 0.1952 3.5205 28.6931 0.1951 0.1946 0.1929 0.3793 0.4073 0.3584	0 0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 13.7000 7.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 8.0000 6.0000 13.7000 7.9000
2006	EE	4257	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4257 4257 4257 4257 4257 4257 4257 4257	0 0 0 0 0 0 0 0 0 0 0	0.6417 0.5192 -0.8308 81.1909 0.4938 0.4280 0.4447 0.7761 2.4614 1.6337 15.4000 5.9000	0.1471 0.1533 2.7528 22.5456 0.1534 0.1518 0.1525 0.2906 0.3270 0.2754 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 15.4000 5.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 9.0000 8.0000 6.0000 15.4000 5.9000
2005	IE	2872	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	2872 2872 2872 2872 2872 2872 2872 2777 2872 2872 2872 2872 2872	0 0 0 0 0 0 0 95 0 0	0.0824 0.5097 -0.2720 75.3925 0.3453 0.2946 0.4414 1.1631 2.5971 1.6651 32.4000 4.3000	0.1751 0.3183 5.5261 44.0678 0.3027 0.2903 0.3168 0.7583 0.7238 0.6233	0 0 -16.2665 0.0710 0 0 0 1.0000 0 32.4000 4.3000	1.0000 1.0000 13.7335 264.6 1.0000 1.0000 9.0000 9.0000 6.0000 32.4000 4.3000
2006	IE	1992	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	1992 1992 1992 1992 1992 1992 1992 1992	0 0 0 0 0 0 0 63 0 0 0	0.0811 0.5266 0.5228 72.0837 0.3667 0.2901 0.4180 1.2204 2.5523 1.5983 34.8000 4.4000	0.2023 0.3700 6.2803 50.2126 0.3571 0.3363 0.3660 0.9358 0.7821 0.6942 0	0 0 -16.2665 0.0710 0 0 0 1.0000 34.8000 4.4000	1.0000 1.0000 13.7335 264.6 1.0000 1.0000 9.0000 7.0000 5.0000 34.8000 4.4000
2005	EL	4642	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4642 4642 4642 4642 4642 4642 4642 4642	0 0 0 0 0 0 0 196 0 0	0.2368 0.5011 -1.1845 77.5079 0.2578 0.2381 0.3357 0.7773 2.6662 1.5917 20.8883 9.9011	0.3385 0.3981 6.9457 57.6718 0.3483 0.3391 0.3777 0.7372 0.7512 0.6483 4.4502 0.9685	0 0 -16.2665 0.0710 0 0 0 1.0000 0 16.2000 8.2000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 10.0000 7.0000 28.3000 11.4000

2006	EL	4211	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4211 4211 4211 4211 4211 4211 4027 4211 4211 4211 4211 4211	0 0 0 0 0 0 0 184 0 0 0	0.2574 0.5024 -0.8390 76.5434 0.2626 0.2301 0.3405 0.7931 2.6953 1.6258 21.8582 8.9479	0.3462 0.3960 6.8965 57.1113 0.3485 0.3333 0.3772 0.7434 0.7519 0.6779 4.8696 0.5389	0 -16.2665 0.0710 0 0 0 1.0000 0 16.9000 7.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 0.0000 7.0000 30.5000 9.7000
2005	ES	8429	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	8429 8429 8429 8429 8429 7794 8429 8429 8429 8429 8429	0 0 0 0 0 0 635 0 0	0.2004 0.4927 -1.7718 78.3910 0.3623 0.4688 0.7216 2.6879 1.6190 23.1034 9.2296	0.5233 0.6537 11.3422 95.7980 0.6285 0.6131 0.6525 1.2211 1.4681 1.1830 6.0902 3.6424	0 0 -16.2665 0.0710 0 0 0 1.0000 0 15.6000 5.6000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 8.0000 6.0000 29.9000 19.7000
2006	ES	11438	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	11438 11438 11438 11438 11438 11438 11438 11438 11438 11438 11438 11438	0 0 0 0 0 0 833 0 0 0	0.1913 0.4098 -1.8626 78.3685 0.3500 0.3440 0.4680 0.7259 2.6490 1.6282 24.6383 8.6267	0.4309 0.5387 9.4802 80.1413 0.5225 0.5204 0.5469 0.9986 1.1568 0.9546 5.4181 2.7513	0 0 -16.2665 0.0710 0 0 0 1.0000 0 16.7000 5.3000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 7.0000 7.0000 32.1000 21.0000
2005	FR	5876	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	5876 5876 5876 5876 5876 5876 5876 5876	0 0 0 0 0 0 177 0 0 0 2 0	0.0798 0.5252 1.3631 67.8346 0.4439 0.2686 0.3917 1.0526 2.2903 1.5306 25.2810 8.8630	0.4672 0.8610 14.0051 116.5 0.8567 0.7642 0.8420 1.9390 1.5338 1.2917 11.6435 3.0632	0 0 -16.2665 0.0710 0 0 0 1.0000 0 19.6000 6.4000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 7.0000 4.0000 38.7000 13.2000
2006	FR	7324	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	7324 7324 7324 7324 7324 7324 7119 7324 7324 7324 7324 7324	0 0 0 0 0 0 0 205 0 0 0	0.0917 0.5262 1.2819 68.6154 0.4242 0.2657 0.3899 1.0192 2.2675 1.5391 26.1066 8.8601	0.4396 0.7608 12.4687 104.2 0.7530 0.6730 0.7433 1.6985 1.3719 1.1152 10.5634 2.5799	0 0 -16.2665 0.0710 0 0 0 1.0000 20.3000 6.1000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 7.0000 4.0000 40.1000 12.4000
2005	IT	12955	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	12955 12955 12955 12955 12955 12955 12955 12797 12955 12955 12955 12955	0 0 0 0 0 0 0 158 0 0 0	0.1430 0.4996 -0.6673 73.6183 0.2886 0.2261 0.2134 0.7360 2.5511 1.4667 23.6233 8.1620	0.4209 0.6012 10.2847 83.8161 0.5448 0.5030 0.4925 1.0747 1.2129 0.9832 7.0590 5.5057	0 0 -16.2665 0.0710 0 0 0 1.0000 15.6000 4.0000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 5.0000 7.0000 5.0000 28.8000 15.3000
2006	IT	18048	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH	18048 18048 18048 18048 18048 18048 17906 18048 18048	0 0 0 0 0 0 0 142 0 0	0.1718 0.4995 -0.4297 72.2801 0.2796 0.2694 0.2191 0.7350 2.5369 1.4865	0.3850 0.5103 8.6658 70.6271 0.4581 0.4588 0.4222 0.9174 1.0417 0.8410	0 0 -16.2665 0.0710 0 0 0 0 1.0000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 6.0000 7.0000 6.0000

			GDPhabnuts unempratenuts	18048 18048	0	24.5302 7.1662	6.1430 3.8868	16.3000 3.6000	29.8000 12.7000
2005	СҰ	2357	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	2357 2357 2357 2357 2357 2357 2357 2357	0 0 0 0 0 0 0 42 0 0 0	0.0719 0.5104 -1.1817 80.6322 0.3778 0.3540 0.4840 0.9912 2.7608 1.8203 20.4000 5.3000	0.0821 0.1588 2.8284 23.5761 0.1541 0.1520 0.1589 0.3387 0.3348 0.2727	0 0 -16.2665 0.0710 0 0 0 1.0000 0 20.4000 5.3000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 7.0000 6.0000 20.4000 5.3000
2006	CĂ	3373	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	3373 3373 3373 3373 3373 3373 3373 337	0 0 0 0 0 0 0 54 0 0 0	0.0552 0.5142 -1.0617 80.5458 0.4170 0.2706 0.4900 1.0275 2.8076 1.8817 21.3000 4.5000	0.0601 0.1316 2.3468 19.5748 0.1298 0.1170 0.1317 0.2843 0.2898 0.2335	0 0 -16.2665 0.0710 0 0 0 1.0000 0 21.3000 4.5000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 7.0000 6.0000 21.3000 4.5000
2005	LV	2704	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	2704 2704 2704 2704 2704 2704 2686 2704 2704 2704 2704 2704	0 0 0 0 0 0 0 18 0 0 0	0.8089 0.5195 -0.6736 78.6490 0.5020 0.4826 0.3015 0.7461 2.5481 1.6262 10.9000 8.9000	0.2019 0.2566 4.5409 37.5947 0.2568 0.2566 0.2357 0.4546 0.5490 0.4794	0 0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 10.9000 8.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 6.0000 5.0000 10.9000 8.9000
2006	LV	3438	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	3438 3438 3438 3438 3438 3437 3438 3438	0 0 0 0 0 0 0 21 0 0 0	0.7396 0.5180 -0.7024 79.0465 0.5185 0.4850 0.3494 0.7660 2.6926 1.7491 12.4000 6.8000	0.1993 0.2270 4.0261 32.8692 0.2270 0.2270 0.2165 0.4175 0.4952 0.4339 0	0 0 0 -16.2665 0.0710 0 0 0 1.0000 12.4000 6.8000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 6.0000 5.0000 12.4000 6.8000
2005	LT	2527	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	2527 2527 2527 2527 2527 2527 2455 2527 2527	0 0 0 0 0 0 0 72 0 0 0	0.7906 0.5165 -0.8742 75.0490 0.3968 0.3962 0.3853 0.9105 2.5087 1.6172 11.9000 8.3000	0.2636 0.3237 5.5828 46.8519 0.3169 0.3168 0.3162 0.6330 0.6445 0.5482	0 0 -16.2665 0.0710 0 0 0 1.0000 11.9000 8.3000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 9.0000 5.0000 11.9000 8.3000
2006	LT	3713	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	3713 3713 3713 3713 3713 3713 3613 3713 37	0 0 0 0 0 0 0 100 0 0	0.7549 0.5173 -0.6671 74.1260 0.4288 0.3475 0.4118 0.9010 2.5249 1.6651 13.1000 5.6000	0.2280 0.2649 4.5500 37.1506 0.2623 0.2524 0.2618 0.5201 0.5195 0.4300 0	0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 13.1000 5.6000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 9.0000 5.0000 13.1000 5.6000
2005	LU	4322	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH	4322 4322 4322 4322 4322 4322 43210	0 0 0 0 0 0	0.0106 0.4941 -0.5958 70.9201 0.3231 0.3213 0.3544	0.0191 0.0930 1.5620 12.6713 0.0870 0.0868 0.0895	0 0 -16.2665 0.0710 0 0	1.0000 1.0000 14.7335 264.6 1.0000 1.0000

			nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4322 4322 4322 4322 4322	0 0 0 0	0.9074 2.3891 1.5991 57.1000 4.5000	0.2003 0.1848 0.1397 0	0 1.0000 0 57.1000 4.5000	7.0000 7.0000 5.0000 57.1000 4.5000
2006	LU	4601	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4601 4601 4601 4601 4601 4601 4601 4601	0 0 0 0 0 0 0 107 0 0 0	0.0192 0.4979 -0.8068 73.1864 0.3191 0.3524 0.9019 2.3624 1.5974 63.1000 4.7000	0.0251 0.0917 1.5614 12.9800 0.0865 0.0885 0.0882 0.1980 0.1804 0.1353	0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 63.1000 4.7000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 6.0000 5.0000 63.1000 4.7000
2005	HU	4353	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4353 4353 4353 4353 4353 4353 4353 4353	0 0 0 0 0 0 0 0 0 0 0 0	0.7286 0.5111 -0.7273 86.0190 0.5287 0.4448 0.2556 0.8801 2.6492 1.5751 14.2595 7.3201	0.3793 0.4264 7.8870 62.2732 0.4258 0.4239 0.3721 0.9083 0.8968 0.7810 4.9600 1.4527	0 0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 9.3000 5.1000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 9.0000 5.0000 23.2000 9.2000
2006	HU	6609	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	6609 6609 6609 6609 6609 6609 6609 6609	0 0 0 0 0 0 0 0	0.6904 0.5155 -0.7190 85.6337 0.4654 0.3874 0.2955 0.8703 2.6738 1.5826 15.0909 7.5619	0.3197 0.3456 6.3795 50.2613 0.3449 0.3369 0.3155 0.7353 0.7254 0.6395 4.4141 1.4049	0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 9.7000 5.1000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 7.0000 5.0000 24.9000 9.9000
2005	NL	8574	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	8574 8574 8574 8574 8574 8574 8206 8574 8574 8407 8574 8574	0 0 0 0 0 0 368 0 0 167	0.0581 0.4921 -0.4148 74.2244 0.2497 0.1563 0.4600 0.9032 2.1546 1.4760 29.4000 4.7000	0.1868 0.3992 6.8706 56.6392 0.3456 0.2899 0.3967 0.8799 0.6870 0.6251	0 0 -16.2665 0.0710 0 0 0 1.0000 0 29.4000 4.7000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 7.0000 5.0000 29.4000 4.7000
2006	NL	9287	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	9287 9287 9287 9287 9287 9287 8949 9287 9287 9279 9287	0 0 0 0 0 0 338 0 0 0 8	0.0318 0.4975 -0.2668 73.4840 0.2542 0.1644 0.4702 0.9265 2.1678 1.6542 30.9000 3.9000	0.1310 0.3732 6.3946 53.0152 0.3249 0.2766 0.3716 0.8325 0.6369 0.5648	0 0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 30.9000 3.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 7.0000 5.0000 30.9000
2005	AT	3882	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	3882 3882 3882 3882 3882 3882 3882 3882	0 0 0 0 0 0 0 0	0.0522 0.5056 -0.1534 70.8171 0.2931 0.3432 0.3181 0.8440 2.3847 1.6527 28.0319 5.1660	0.1828 0.4107 6.9123 57.3363 0.3739 0.3900 0.3826 0.8552 0.8766 0.7373 1.8223 1.0814	0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 23.9000 3.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 8.0000 7.0000 29.6000 6.7000
2006	AT	5469	pov_indicator woman age_centered age_centered2	5469 5469 5469 5469	0 0 0	0.0555 0.5080 -0.3177 70.2327	0.1577 0.3443 5.7667 48.1919	0 0 -16.2665 0.0710	1.0000 1.0000 14.7335 264.6

			chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	5469 5469 5469 5469 5469 5469 5469	0 0 0 0 0 0	0.2754 0.3202 0.3149 0.8224 2.3748 1.6395 29.4096 4.7459	0.3076 0.3213 0.3198 0.7036 0.7198 0.6173 1.4577 0.9453	0 0 0 0 1.0000 0 25.3000 3.3000	1.0000 1.0000 1.0000 6.0000 8.0000 6.0000 30.8000 6.3000
2005	PL	10627	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	10627 10627 10627 10627 10627 10627 10561 10627 10564 10627 10627	0 0 0 0 0 0 66 0 0 63	0.7381 0.5039 -0.6632 85.8863 0.4761 0.2527 0.2646 0.8989 2.7400 1.4003 11.5603 17.7915	0.4729 0.5378 9.9424 79.9035 0.5372 0.4674 0.4748 1.1369 1.3407 1.0024 2.5551 2.0488	0 0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 8.2000 15.7000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 1.0000 9.0000 5.0000 15.7000 21.4000
2006	PL	14644	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	14644 14644 14644 14644 12928 14644 14644 14644 14644 14644	0 0 0 0 0 0 1716 0 0 1643	0.7258 0.5037 -0.6473 85.1495 0.4808 0.3260 0.2834 0.9003 2.9673 1.5990 12.3506 13.9915	0.4102 0.4598 8.4644 67.2797 0.4594 0.4310 0.4143 0.9688 1.1522 0.8760 2.4094 1.0485	0 0 -16.2665 0.0710 0 0 0 1.0000 0 8.7000 12.7000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 11.0000 9.0000 6.0000 17.0000 16.4000
2005	PT	2671	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	2671 2671 2671 2671 2671 2671 2198 2671 2671 2671 2671	0 0 0 0 0 0 473 0 0	0.4168 0.5072 -0.6899 79.3436 0.4695 0.4113 0.2303 0.8315 2.8063 1.8021 17.3000 7.6000	0.5441 0.5517 9.8001 79.7975 0.5507 0.4696 0.9949 1.1722 1.0142 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 17.3000 7.6000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 9.0000 8.0000 7.0000 17.3000 7.6000
2006	PT	3662	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	3662 3662 3662 3662 3662 3049 3662 3662 3662 3662 3662	0 0 0 0 0 0 0 613 0 0 0	0.4019 0.5112 -0.3565 79.3109 0.4380 0.3936 0.2419 0.7832 2.8086 1.8287 18.0000 7.7000	0.4575 0.4665 8.3039 66.5722 0.4630 0.4559 0.4037 0.8223 1.0182 0.8564	0 0 0 -16.2665 0.0710 0 0 0 1.0000 0 18.0000 7.7000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 9.0000 9.0000 7.0000 18.0000 7.7000
2005	SI	8671	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	8671 8667 8671 8671 8671 8671 8671 8671	0 4 0 0 0 0 200 0 0 13	0.1266 0.4899 -0.6833 79.6898 0.2501 0.2177 0.2149 0.7566 2.9380 1.7535 19.6000 6.5000	0.0915 0.1376 2.4495 19.9780 0.1192 0.1136 0.1128 0.2591 0.3047 0.2551 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 19.6000 6.5000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 9.0000 9.0000 6.0000 19.6000 6.5000
2006	SI	11271	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	11271 11263 11271 11271 11271 11271 11271 11189 11271 11271 11262 11271 11271	0 8 0 0 0 0 82 0 0 9	0.1306 0.4929 -0.7373 78.4043 0.2915 0.2126 0.2549 0.7634 2.9358 1.7851 20.7000 6.0000	0.0808 0.1200 2.1167 17.6044 0.1090 0.0981 0.1046 0.2254 0.2660 0.2174	0 0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 20.7000 6.0000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 9.0000 9.0000 6.0000 20.7000 6.0000

2005	SK	3447	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	3447 3447 3447 3447 3447 3447 3432 3447 3447	0 0 0 0 0 0 0 15 0 0	0.7661 0.5198 0.1174 81.3319 0.4441 0.3993 0.3085 0.8904 1.9700 13.5000 16.3000	0.3005 0.3547 6.4023 53.0850 0.3528 0.3477 0.3279 0.7283 0.8632 0.7508	0 0 -16.2665 0.0710 0 0 0 1.0000 0 13.5000 16.3000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 8.0000 8.0000 13.5000 16.3000
2006	sk	4898	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4898 4898 4898 4898 4898 4898 4898 4898	0 0 0 0 0 0 0 28 0 0 0	0.7157 0.5217 0.0502 82.3628 0.4327 0.4442 0.3248 0.8442 3.1409 2.0315 15.0000 13.4000	0.2680 0.2968 5.3919 44.6004 0.2944 0.2952 0.2783 0.5925 0.7288 0.6170 0	0 0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 15.0000 13.4000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 8.0000 8.0000 15.0000 13.4000
2005	FI	4021	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4021 4021 4021 4021 4021 4021 3927 4021 4021 4018 4021 4021	0 0 0 0 0 0 94 0 0 3	0.0529 0.4953 -0.3310 80.6732 0.3076 0.2930 0.5226 0.8869 2.1009 1.4707 25.6182 8.5095	0.1404 0.3137 5.6323 46.9628 0.2856 0.3133 0.7442 0.5226 0.4894 2.5938 1.1694	0 0 0 -16.2665 0.0710 0 0 0 0 1.0000 0 19.1000 6.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 7.0000 5.0000 29.6000 11.7000
2006	FI	5528	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	5528 5528 5528 5528 5528 5528 5528 5404 5528 5528 5528 5528	0 0 0 0 0 0 0 124 0 0 0	0.0622 0.4953 -0.3237 80.8606 0.2901 0.2894 0.5179 0.8898 2.0887 1.4683 27.1434 7.7955	0.1290 0.2669 4.7976 39.6984 0.2423 0.2421 0.2668 0.6325 0.4440 0.3940 2.3172 0.9975	0 0 -16.2665 0.0710 0 0 0 1.0000 20.2000 6.3000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 12.0000 4.0000 31.3000 11.3000
2005	SE	3541	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	3541 3541 3541 3541 3541 3541 3541 3541	0 0 0 0 0 0 127 0 0 386 0	0.0452 0.5016 -0.3763 78.1345 0.2827 0.1482 0.4394 0.9789 2.1022 1.5457 27.1000 7.5000	0.1750 0.4211 7.4383 60.4577 0.3793 0.2992 0.4189 0.9351 0.6935 0.5973	0 0 0 -16.2665 0.0710 0 0 0 1.0000 0 27.1000 7.5000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 7.0000 27.1000 7.5000
2006	SE	4751	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4751 4751 4751 4751 4751 4751 4751 4751	0 0 0 0 0 0 217 0 0 0 541 0	0.0527 0.5011 -0.4989 76.0001 0.2920 0.1593 0.4728 1.0203 2.0672 1.5782 28.7000 7.1000	0.1577 0.3531 6.1470 49.7601 0.3211 0.2585 0.3530 0.7818 0.5636 0.4843	0 0 -16.2665 0.0710 0 0 0 1.0000 -28.7000 7.1000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 7.0000 5.0000 28.7000 7.1000
2005	UK	6224	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH	6224 6224 6224 6224 6224 6224 4949 6224 4766	0 0 0 0 0 0 1275 0 0	0.0739 0.5260 -1.9273 67.7453 0.4031 0.2293 0.5177 1.2542 2.1928 1.5478	0.5272 1.0064 16.1286 142.1 0.9887 0.8473 1.0064 2.3990 1.7624 1.7058	0 0 -16.2665 0.0710 0 0 0 0 1.0000	1.0000 1.0000 13.7335 264.6 1.0000 1.0000 6.0000 7.0000 6.0000

			GDPhabnuts unempratenuts	6224 6224	0	27.4000 4.8000	0	27.4000 4.8000	27.4000 4.8000
2006	UK	7594	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	7594 7594 7594 7594 7594 7594 5998 7594 7594 6118 7594 7594	0 0 0 0 0 0 1596 0 0 1476	0.0837 0.5363 -1.5980 67.6428 0.4049 0.2199 0.5080 1.2305 2.1941 1.5570 28.4000 5.4000	0.4988 0.8983 14.5325 128.3 0.8842 0.7461 0.8985 2.1147 1.5651 1.4788 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 28.4000 5.4000	1.0000 1.0000 13.7335 264.6 1.0000 1.0000 6.0000 8.0000 5.0000 28.4000 5.4000
2005	IS	1583	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	1583 1583 1583 1583 1583 1583 1583 1318 1583 1583	0 0 0 0 0 0 224 0 0 265	0.0333 0.5066 -1.0735 76.5700 0.2143 0.1521 0.3835 1.1989 2.4249 1.8715 29.3000 2.5000	0.0418 0.1164 2.0227 16.9520 0.0956 0.0836 0.1150 0.2545 0.2303 0.1947 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 29.3000 2.5000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 5.0000 7.0000 6.0000 29.3000 2.5000
2006	IS	2240	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	2240 2240 2240 2240 2240 2240 2240 2021 2240 2054 2240 2240	0 0 0 0 0 0 219 0 0 186	0.0322 0.4993 -0.9177 78.2001 0.2180 0.1579 0.4147 1.1305 2.3563 1.7410 29.3000 2.8000	0.0348 0.0986 1.7342 14.4429 0.0814 0.0719 0.0983 0.2161 0.1875 0.1499	0 0 -16.2665 0.0710 0 0 0 1.0000 0 29.3000 2.8000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 5.0000 7.0000 5.0000 29.3000 2.8000
2005	NO	4088	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4088 4088 4088 4088 4088 4088 3887 4088 4088 3996 4088	0 0 0 0 0 0 201 0 0 92	0.0326 0.4919 -0.6817 74.7612 0.2545 0.1610 0.3970 1.0059 2.0123 1.4737 39.6000 4.4000	0.1001 0.2816 4.8555 40.0863 0.2454 0.2070 0.2753 0.6332 0.4476 0.4073	0 0 0 -16.2665 0.0710 0 0 0 1.0000 0 39.6000 4.4000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 8.0000 4.0000 39.6000 4.4000
2006	NO	3926	pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	3926 3926 3926 3926 3926 3926 3583 3926 3732 3926 3926	0 0 0 0 0 0 343 0 0 194	0.0312 0.4909 -0.5332 71.9978 0.2769 0.1713 0.4123 0.9872 2.0046 1.5183 43.4000 3.4000	0.0978 0.2814 4.7661 39.3175 0.2518 0.2120 0.2776 0.6287 0.4367 0.4010	0 0 -16.2665 0.0710 0 0 0 1.0000 0 43.4000 3.4000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 8.0000 5.0000 43.4000 3.4000

Appendix 6. Empty model (weighted)

		Response	Profile				
	Ordered			Total			
	Value	pov_indi	cator F	requency			
	1	1		76252			
	2	0		222018			
The GLIMMIX p	rocedure is mode	ling the p	robability	that pov_i	ndicator='1'.		
		Dimens					
		ov. Parame		1			
		ov. Parame	ters	2			
	Columns	in X		1			
	Columns	in Z per S	ubject	1			
		(Blocks i		93			
	Max Obs	per Subjec	t	19942			
		imization					
	Optimization Parameters in	Technique	New	ton-Raphson			
	Lower Boundar		2				
	Upper Boundar		1				
	Fixed Effects			filed			
	Residual Vari			filed			
/ 6 iterations	Starting From		GLM	estimates			
	onvergence crite	rion (PCON	V=1.11022E	-8) satisfi	ed.		
		Fit Stat	istics				
	-2 Res Log	Pseudo-Lik	elihood	1738618			
	Generalized			294644.1			
	Gener. Chi-	Square / D	F	0.99			
	Covaria	nce Parame	ter Estima				
					andard		
	Parm Subject		Est:	imate	Error		
	rcept region		1	.7720 .4571 0.	0.2639		
•		ue_UE(regi					
Resi	dual		0	.9878 0.	002734		
	tic Covariance M	atrix of C					
	Subject		CovP1	CovP2	CovP3		
Intercept	region ID_unique_UE(0.06965	-2.32E-7	-2.12E-7		
AR(1)	ID_unique_UE(region)	-2.32E-7	5.477E-6	2.254E-6		
Residual			-2.12E-7	2.254E-6	7.473E-6		
Asympto	tic Correlation	Matrix of	Covariance	Parameter	Estimates		
Cov Parm					CovP3		
Intercept			1.0000 -0.00038	-0.00038 1.0000	-0.00029		
AR(1)	ID_unique_UE(region)	-0.00038	1.0000	0.3524		
Residual			-0.00029	0.3524	1.0000		
	Solut	ions for F		ts			
		Standard			D. N. 111		
Effect Intercep	Estimate		DF	ts t Value -10.74	Pr > t <.0001		

The empty model contains only a random intercept. The dependent variable is thus explained by the overall mean, a random term at group level and a random term at individual level. The empty model allows to find out the relative parts of the variance between groups (regions here) and the variance within groups. It is thus possible to calculate the intraclass correlation coefficient (ICC) (here: 55.1% = 1.772/3.2169, i.e. the intercept variance divided by the total variance).



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