

Modelling Long Memory Volatility in Agricultural Commodity Futures Returns

Roengchai Tansuchat

Faculty of Economics
Maejo University
Chiang Mai, Thailand

Chia-Lin Chang

Department of Applied Economics
National Chung Hsing University
Taichung, Taiwan

Michael McAleer

Econometrics Institute
Erasmus School of Economics
Erasmus University Rotterdam
and
Tinbergen Institute
The Netherlands
and
Center for International Research on the Japanese Economy (CIRJE)
Faculty of Economics
University of Tokyo

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Abstract

This paper estimates the long memory volatility model for 16 agricultural commodity futures returns from different futures markets, namely corn, oats, soybeans, soybean meal, soybean oil, wheat, live cattle, cattle feeder, pork, cocoa, coffee, cotton, orange juice, Kansas City wheat, rubber, and palm oil. The class of fractional GARCH models, namely the FIGARCH model of Baillie et al. (1996), FIEGARCH model of Bollerslev and Mikkelsen (1996), and FIAPARCH model of Tse (1998), are modelled and compared with the GARCH model of Bollerslev (1986), EGARCH model of Nelson (1991), and APARCH model of Ding et al. (1993). The estimated d parameters, indicating long-term dependence, suggest that fractional integration is found in most of agricultural commodity futures returns series. In addition, the FIGARCH $(1,d,1)$ and FIEGARCH $(1,d,1)$ models are found to outperform their GARCH $(1,1)$ and EGARCH $(1,1)$ counterparts.

Keywords: Long memory, agricultural commodity futures, fractional integration, asymmetric, conditional volatility.

JEL Classifications: Q14, Q11, C22, C51.

1. Introduction

Accurate modeling of volatility in asset returns is one of the major issues of concern in financial economics. Poon and Granger (2003) have mentioned that, even though volatility is not the same as risk, when it is interpreted as uncertainty, it becomes a key input to many important financial applications, such as investment, portfolio construction, option pricing, hedging and risk management. Research on volatility models has focused on such different properties of the returns series as its time-varying conditional moments, volatility clustering, asymmetric patterns and long persistence, among others.

Derivative markets, particularly commodity futures markets, have become more sophisticated since the Chicago Board of Trade commenced futures trading in 1848. The futures price depends on the flow of information around the world. Small changes in prices could have tremendous effects on trading results across futures markets. This distinction implies that the futures market is more volatile and has high risk. This feature is also particularly important in the agricultural commodity futures market, where factors such as drought, natural disaster, deforestation, and debt default can have a major impact on demand and supply of commodities, and hence on the present and futures prices of the commodity.

In modern time series modelling, following the seminal work of Engle (1982), a group of time series models named Autoregressive Conditional Heteroskedasticity (ARCH), and later generalized by Bollerslev (1986) as Generalized Autoregressive Conditional Heteroskedasticity (GARCH), has been used to model time-varying conditional volatility. The ARCH and GARCH models explain time series behaviour by allowing the conditional variance to evolve over time and to respond to previous price changes. These models consider non-linearity in the conditional mean equation, and are also able to explain volatility clustering and volatility persistence. A considerable empirical literature in commodity cash and futures markets has used a variety of GARCH models to estimate expected price and returns volatility.

The GARCH model assumes that negative and positive shocks of equal magnitude have identical impacts on the conditional variance. In order to accommodate differential impacts on conditional variance between positive and negative shocks of equal magnitude, Glosten et al. (1992) proposed the asymmetric GARCH, or GJR model. Since the positive and negative

shocks on conditional volatility, called leverage effect, are asymmetric, Nelson (1991) proposed the Exponential GARCH (EGARCH) model.

In terms of volatility persistence, a GARCH model features an exponential decay in the autocorrelation of conditional variances. However, a shock in the volatility series seems to have very “long memory” and impacts on future volatility over a long horizon. Baillie et al. (2007) explained that the long memory refers to the presence of very slow hyperbolic decay in the autocorrelations and impulse response weights. Therefore, econometrically, the long memory is between the usual exponential rates of decay associated with the class of stationary and invertible ARMA models, and the alternative extreme of infinite persistence associated with integrated, unit root processes. Therefore, Baillie, Bollerslev and Mikkelsen (1996) proposed the FIGARCH(p, d, q) model, and Bollerslev and Mikkelsen (1996) proposed the FIEGARCH(p, d, q) model, where a full description of the properties of the process and the appropriate quasi-maximum likelihood estimation (QMLE) method can be found.

A number of previous papers have observed and provided application of fractional integrated models in many fields, namely stock returns (Bollerslev and Mikkelsen (1996), Degiannakis (2004) and Ñiguez (2007), Lux and Kaizoji (2007), Kang and Yoon (2007), Jefferis and Thupayagale (2008), Ruiz and Veiga (2008)); exchange rate (Baillie et al. (1996), Davidson (2004), and Conrad and Lamla (2007)) and inflation rate (Baillie et al. (2002)). However, in the literature to date, there have been few applications of the fractionally integrated GARCH class models to commodity futures markets. Barkoulas et al. (1997) examined the fractional structure of commodities spot prices, namely aluminum, cocoa, coffee, copper, rice and rubber. They found that some commodity spot price time series display a fractal structure, and the fractional orders vary among these commodities because the processes involved in the price movements of each commodity varies.

Crato and Ray (2000) investigated long-term memory in the returns and volatility of commodity futures market, namely five currencies, twelve agricultural commodities, three metals and heating oil, and five currencies futures markets. They found that commodity futures volatilities are typically more persistent than currency futures volatilities. However, they do not explicitly estimate the FIGARCH model. Jin and Frechette (2004) examined the presence of fractional integration in the volatility of fourteen agricultural futures prices series. The results show that the volatility series exhibit strong long-term dependence, which is an

indicator of fractional integration. In addition, FIGARCH(1, d ,1) performs significantly better than a traditional volatility model, GARCH(1,1), in modelling agricultural price volatility.

Baillie et al. (2007) examined long memory in volatility properties of both daily and high-frequency intraday futures returns for six important commodities. They found that the volatility processes were found to be accurately described by FIGARCH models, with statistically significant long memory parameter estimates.

Recently, Hyun-Joung (2008) explored a long memory conditional volatility model on international grain markets, namely wheat, corn and soybeans, and compared the performance of the models in capturing dependence of the price volatility, and also emphasized suitability of the student- t density intended to account for non-normal, fat-tailed properties of the data. The empirical results showed that grain cash price volatilities exhibit long memory and that the memory is adequately modelled by a fractionally integrated process and implemented by FIGARCH models. In addition, the suitability of the FIGARCH models is under the student- t distribution and the competitiveness of the parsimonious FIGARCH(1, d ,0) model. Therefore, it is desirable to use long memory conditional variance models for analysis of grain price volatility dynamics.

The fractionally integrated multivariate conditional volatility model of Brunetti and Gilbert (2000) applied the univariate volatility (FIGARCH) model to multivariate GARCH models by estimating and testing cointegrated bivariate FIGARCH models using NYMEX and IPE crude oil markets. They found a common order of fractional integration for the two volatility process, and confirmed that they are fractionally cointegrated. An estimated error correction FIGARCH model indicated that the predominant adjustment is the IPE toward NYMEX.

Coakley et al. (2008) explored the relationship between basis long memory and hedging effectiveness measures with error-correction and the multivariate GARCH (FIEC-BEKK) model, employing spot daily data and their corresponding futures contracts 1995-2005 for five commodities, namely soybeans, cocoa, heating oil, gold and live cattle. The results presented a long memory component that should theoretically affect hedging effectiveness.

The aims of the paper are to model the volatility in agricultural commodity futures returns series using a variety of conditional volatility models, namely GARCH, EGARCH and

APARCH, and fractionally integrated conditional volatility models, namely FIGARCH, FIEGARCH and FIAPARCH. The remainder of the paper is as follows. Section 2 discusses the methodological approach used in the paper. Section 3 describes the commodity futures prices time series. Section 4 presents the results from empirical modelling, and Section 5 gives some concluding comments.

2. Econometric Models

2.1 Univariate Conditional Volatility Models

This section presents the volatility models in commodity futures returns, namely the GARCH model of Bollerslev (1986), EGARCH model of Nelson (1991) and APARCH model of Ding et al. (1993), and fractionally integrated conditional volatility models, namely FIGARCH model of Baillie et al. (1996), FIEGARCH model of Bollerslev and Mikkelsen (1996), and FIAPARCH model of Tse (1998).

Following Engle (1982), consider the time series $y_t = E_{t-1}(y_t) + \varepsilon_t$, where $E_{t-1}(y_t)$ is the conditional expectation of y_t at time $t-1$ and ε_t is the associated error. The generalized autoregressive conditional heteroskedastity (GARCH) model of Bollerslev (1986) is given as follows:

$$\varepsilon_t = \sqrt{h_t} \eta_t \quad , \quad \eta_t \square N(0,1) \quad (1)$$

$$h_t = \omega + \sum_{j=1}^p \alpha_j \varepsilon_{t-j}^2 + \sum_{j=1}^q \beta_j h_{t-j} = \omega + \alpha(L) \varepsilon_t^2 + \beta(L) \sigma_t^2 \quad (2)$$

where $\omega > 0$, $\alpha_j \geq 0$ and $\beta_j \geq 0$ are sufficient condition to ensure that the conditional variance $h_t > 0$, L is the lag operator, $\alpha(L) = \alpha_1 L + \alpha_2 L^2 + \dots + \alpha_p L^p$ and $\beta(L) = \beta_1 L + \beta_2 L^2 + \dots + \beta_q L^q$. In (2) the parameter α_j represents the ARCH effect, or the short-run persistence of shocks to returns, β_j represents the GARCH effects, and $(\alpha_j + \beta_j)$ measures the persistence of the contribution of shocks to return i to long-run persistence. If the roots of $[1 - \alpha(L) - \beta(L)]$ and $[1 - \beta(L)]$ lie outside the unit circle, then $\{\varepsilon_t^2\}$ exhibits stability and covariance stationarity. The volatility shocks decay at a geometric rate.

Equation (2) assumes that the conditional variance is a function of the magnitudes of the lagged residuals and not their sign, such that a positive shock ($\varepsilon_t > 0$) has the same impact on the conditional variance as a negative shock ($\varepsilon_t < 0$) of equal magnitude. In order to accommodate differential impacts on the conditional variance of positive and negative shocks, Glosten et al. (1992) proposed the asymmetric GARCH, or GJR model, as given by

$$h_t = \omega + \sum_{j=1}^r (\alpha_j + \gamma_j I(\eta_{t-j})) \varepsilon_{t-j}^2 + \sum_{j=1}^s \beta_j h_{t-j} \quad (3)$$

where

$$I_{it} = \begin{cases} 0, & \varepsilon_{it} \geq 0 \\ 1, & \varepsilon_{it} < 0 \end{cases} \quad (4)$$

is an indicator function to differentiate between positive and negative shocks. Bollerslev (1986) showed the necessary and sufficient condition for the second-order stationarity of GARCH is $\sum_{j=1}^r \alpha_j + \sum_{j=1}^s \beta_j < 1$. For the GARCH(1,1) model, Nelson (1990) obtained the log-

moment condition for the strict stationary and ergodicity as $E(\log(\alpha_1 \eta_t^2) + \beta_1) < 0$, which is important in deriving the statistical properties of the QMLE.

In an alternative model that accommodates asymmetry between positive and negative shocks, and possibly also leverage, Nelson (1991) proposed the Exponential GARCH (EGARCH) model, interpreting as ARMA-type models for the logarithm of the conditional variance, namely:

$$\log h_t = \omega + \sum_{i=1}^p \alpha_i |\eta_{t-i}| + \sum_{i=1}^p \gamma_i \eta_{t-i} + \sum_{j=1}^q \beta_j \log h_{t-j}. \quad (4)$$

In (4), $|\eta_{t-i}|$ and η_{t-i} capture the size and sign effects, respectively, of the standardized shocks. Unlike the GARCH model, EGARCH in (4) uses the standardized residual rather than the unconditional shocks. As EGARCH also uses the logarithms of conditional volatility, there are no restrictions on the parameters in (4). As the standardized shocks have finite moments, the moment conditional of (4) are straightforward. The distinctions between EGARCH and the previous two GARCH models are discussed in McAleer (2005) and McAleer et al. (2007)

Alternatively, Bollerslev and Mikkelsen (1996) proposed expressing the EGARCH model as follows:

$$\ln(\sigma_t^2) = \omega + [1 - \beta(L)]^{-1} [1 + \alpha(L)] g(z_{t-1}). \quad (5)$$

The value of $g(z_t)$ depends on several elements. Following Nelson (1991), in order to accommodate the asymmetric relation between returns and volatility changes, the value of $g(z_t)$ must be a function of both the magnitude and sign of z_t , which yields the function $g(z_t)$ expressed as

$$g(z_t) = \underbrace{\gamma_1 z_t}_{\text{sign effect}} + \underbrace{\gamma_2 [|z_t| - E|z_t|]}_{\text{magnitude effect}}. \quad (6)$$

The parameter γ_1 captures the leverage effect. If $\gamma_1 < 0$, the futures conditional variances will increase proportionally more as a result of a negative shock than for a positive shock of the same absolute magnitude.

Ding, Granger and Engle (1993) proposed an asymmetric power GARCH (APARCH) model, whereby the power of the standard deviation, σ_t^δ , where $\delta > 0$, is a parameter to be estimated. The APARCH(p, q) is defined as:

$$\sigma_t^\delta = \omega + \sum_{j=1}^p \alpha_j (|\varepsilon_{t-j}| - \gamma_j \varepsilon_{t-j})^\delta + \sum_{j=1}^q \beta_j \sigma_{t-j}^\delta \quad (7)$$

and $-1 < \gamma_i < 1$, ($i = 1, \dots, q$). This model nests a least seven ARCH-type models, namely the ARCH model of Engle (1982), GARCH model of Bollerslev (1986), Taylor/Schwert GARCH in standard deviation of Taylor (1986) and Schwert (1990), GJR model of Glosten et al. (1993), TARARCH of Zakoian (1994), NARCH of Higgins and Bera (1992), and log-ARCH of Geweke (1986) and Pantula (1986). Following Ding et al. (1993), if $\omega > 0$ and $\sum_{i=1}^p \alpha_i E(|z| - \gamma_i z)^\delta + \sum_{j=1}^q \beta_j < 1$, a stationary solution for equation (7) exists and is given by

$$E(\sigma_t^\delta) = \frac{\alpha_0}{1 - \sum_{i=1}^p \alpha_i E(|z| - \gamma_i z)^\delta + \sum_{j=1}^q \beta_j}$$

In order to estimate the parameters of model (1)-(7), maximum likelihood estimation is used with a joint normal distribution of η_t . However, when the process for η_t does not follow a normal distribution, or when the conditional distribution is not known, the solution to maximizing the likelihood function is the quasi-MLE (QMLE) approach.

2.2 Univariate fractional integrated conditional volatility models

The long memory property can be defined through the properties of the autocorrelation function, which is defined as $\rho_k = \text{cov}(x_t, x_{t-k}) / \text{var}(x_t)$ for integer lag k . A covariance stationary time series process is expected to have autocorrelations such that $\lim_{k \rightarrow \infty} \rho_k = 0$. Most of the well-known class of stationary and invertible time series processes have autocorrelations that decay at the relatively fast exponential rate, so that $\rho_k \approx |m|^k$, where $|m| < 1$, and this property is true, for example, for the well-known stationary and invertible ARMA(p, q) process. For long memory processes, the autocorrelations decay at an hyperbolic rate which is consistent with $\rho_k \approx c_1 k^{2d-1}$ as k increases without limit, where c_1 is a constant and d is the long memory parameter.

In applications, it often occurs that the estimated sum of parameters α_1 and β_1 in GARCH(1,1) is close to unity, that is $\sum_{i=1}^p \alpha_i + \sum_{j=1}^q \beta_j \approx 1$, or for GARCH(p, q), the process exhibits strong persistence. If $\sum_{i=1}^p \alpha_i + \sum_{j=1}^q \beta_j < 1$, the process (ε_t) is second-order stationary, and a shock to the conditional variance h_t has a decaying impact on h_{t+h} , when h increases, and is asymptotically negligible. However, if $\sum_{i=1}^p \alpha_i + \sum_{j=1}^q \beta_j \geq 1$, the effect on h_{t+h} does not die out asymptotically.

This property is called persistence. Under the restriction $\sum_{i=1}^p \alpha_i + \sum_{j=1}^q \beta_j = 1$, Engle and Bollerslev (1986) called an integrated GARCH (IGARCH) model, meaning that current information remains of importance when forecasting the volatility for all horizons:

$$\phi(L)(1-L)\varepsilon_i^2 = \omega + [1 - \beta(L)]\nu_i \quad (8)$$

where $\nu_i = \varepsilon_i^2 - \sigma_i^2$ is the “innovation” in the conditional variance process or martingale difference process with respect to σ_i^2 or h_i , and has mean 0 and no serial correlation, $\phi(L) = [1 - \alpha(L) - \beta(L)](1-L)^{-1}$, and all the roots of $\phi(L)$ and $[1 - \beta(L)]$ lie outside the unit circle.

However, volatility tends to change quite slowly over time and, as shown in Ding, et al. (1993), the effects of a shock can take a considerable time to decay. Therefore, the distinction between $I(0)$ and $I(1)$ processes seems to be too restrictive. Indeed, the propagation of shock in an $I(0)$ process shocks dies out at an exponential rate (so that it only captures short memory) and, for an $I(1)$ process, the persistence of shocks is infinite and there is no mean reversion, whereas $0 < d < 1$ shocks die out at a slow hyperbolic rate. Baillie et al. (1996) (hereafter denoted BBM) introduced the fractionally integrated GARCH (FIGARCH) model in order to capture the long memory effect in volatility, allows a hyperbolic decay of the coefficient β_j , which is positive, summable, and satisfies the unit root condition. This model mimics the ARFIMA framework of the conditional mean equation.

The FIGARCH(p, d, q) process is defined as:

$$\phi(L)(1-L)^d \varepsilon_i^2 = \omega + \{1 - \beta(L)\}\nu_i, \quad (9)$$

where all the roots of $\phi(L)$ and $\{1 - \beta(L)\}$ lie outside the unit circle. Analogously to (9), the FIGARCH process can be represented as:

$$h_i = \underbrace{\omega [1 - \beta(L)]^{-1}}_{\omega^*} + \underbrace{\{1 - [1 - \beta(L)]^{-1} \phi(L)(1-L)^d\}}_{\lambda(L)} \varepsilon_i^2, \quad (10)$$

or $h_i = \omega^* + \lambda_i(L) \varepsilon_i^2 = \omega^* + \lambda(L) \varepsilon_i^2$, when $0 < d < 1$. $(1-L)^d$, where $0 < d < 1$, is the fractional differencing operator, and its value depends on the decay rate of a shock to conditional volatility. It is also most conveniently expressed in term of the hypergeometric function:

$$(1-L)^d = \sum_{k=0}^{\infty} \frac{\Gamma(d+1)}{\Gamma(k+1)\Gamma(d-k+1)} L^k \quad (11)$$

or
$$(1-L)^d = \sum_{k=0}^{\infty} \binom{d}{k} (-L)^k = 1 - dL - \frac{1}{2}d(d-1)L^2 - \frac{1}{6}d(d-1)(2-d)L^3 - \dots \quad (12)$$

It is easy to show that $\omega > 0$, $\beta_1 - d \leq \phi_1 \leq \frac{2-d}{2}$ and $d\left(\phi_1 - \frac{1-d}{2}\right) \leq \beta_1(\phi_1 - \beta_1 + d)$ are sufficient to ensure that the conditional variance of the FIGARCH(1,d,1) is positive almost surely for all t . FIGARCH nests the GARCH model when $d = 0$, and the IGARCH model when $d = 1$. Approximate maximum likelihood estimates of the parameters of the FIGARCH(p,d,q) process in (9) can be obtained by QMLE.

Chung (1999) argued that the method of parameterization of FIGARCH model of Baillie et al. (1996) may have a specification problem, and underscores some drawbacks in the BBM model. There may be a structural problem in the BBM specification in parallel with the ARFIMA framework of the conditional mean equation, thereby leading to difficult interpretations of the estimated parameters. Indeed, the fractional differencing operator applies to the constant term in the mean equation (ARFIMA), while it does not do so in the variance equation (FIGARCH). Therefore, Chung (1999) redefines the FIGARCH model as:

$$\phi(L)(1-L)^d (\varepsilon_t^2 - \sigma^2) = [1 - \beta(L)](\varepsilon_t^2 - \sigma^2), \quad (13)$$

where σ^2 is the unconditional variance of ε_t . If we retain the same definition of $\lambda(L)$ as in (13), we can formulate the conditional variance as:

$$h_t = \sigma^2 + \left\{1 - [1 - \beta(L)]^{-1} \phi(L)(1-L)^d\right\} (\varepsilon_t^2 - \sigma^2) \quad (14)$$

or
$$h_t = \sigma^2 + \lambda(L)(\varepsilon_t^2 - \sigma^2). \quad (15)$$

In order to accommodate asymmetries between positive and negative shocks, called the leverage effect, Bollerslev and Mikkelsen (1996) extend the FIGARCH process to FIEGARCH, to correspond with Nelson's (1991) Exponential GARCH model to allow for asymmetry. The FIEGARCH(p,d,q) model is given as

$$\ln(h_t) = \omega + \phi(L)^{-1} (1-L)^{-d} [1 + \alpha(L)] g(z_{t-1}), \quad (16)$$

where $g(z_t) = \theta z_t + \gamma [|z_t| - E|z_t|]$, the first term (θz_t) is the sign effect, and the second term ($\gamma [|z_t| - E|z_t|]$) is the magnitude effect. All the roots of $\phi(L)$ and $\lambda(L)$ are an autoregressive polynomial and a moving average polynomial in the lag operator L and lie

outside the unit circle, and both polynomials do not have a common root. When $d = 0$, the FIEGARCH(p, d, q) process reduces to EGARCH of Nelson (1991), and when $d = 1$, the process becomes integrated EGARCH (IEGARCH). Bollerslev and Mikkelsen (1996) presented evidence on the efficiency of QMLE applied to estimate the parameters of the FIEGARCH process.

Tse (1998) proposed a model which combines the fractionally integrated GARCH formulation of Baillie et al. (1996) with the asymmetric power ARCH specification of Ding, Granger and Engle (1993) (see Ling and McAleer (2002) for the theoretical properties of the model). This model increases the flexibility of the conditional variance specification by allowing (a) an asymmetric response of volatility to positive and negative shocks, (b) the data to determine that power of returns for which the predictable structure in the volatility pattern is the strongest, and (c) long range volatility dependence. The FIAPARCH(p, d, q) model can be written as:

$$\sigma_t^\delta = \omega + \left\{ 1 - [1 - \beta(L)]^{-1} \phi(L)(1-L)^d \right\} (|\varepsilon_t| - \gamma \varepsilon_t)^\delta, \quad (17)$$

where γ is the leverage coefficient, and δ is the parameter for the power term that takes (finite) positive values. When $d = 0$, the FIAPARCH(p, d, q) process reduces to APARCH of Ding et al. (1993). When $\gamma = 0$ and $\delta = 2$, the process in (13) reduces to the FIGARCH(p, d, q) specification, which includes Bollerslev's (1986) model when $d = 0$, and the integrated specification when $d = 1$, as special cases.

3. Data

The data are daily synchronous closing futures prices of agricultural futures on different major US commodity futures markets, specifically, the Chicago Board of Trade (CBOT) for corn, oats, soybeans, soybean meal, soybean oil and wheat; the Chicago Mercantile Exchange (CME) for cattle feeder, live cattle and pork, the New York Board of Trade (NYBOT) for cocoa, cotton, coffee and orange juice, and Kansas City Board of Trade (KCBOT) for wheat. The 7,889 price observations from 4 January 1979 to 16 April 2009 are obtained from the DataStream database service.

This paper also focuses on two important commodity futures prices of tropical rain plants, namely rubber (RSS3), trading on the Tokyo Stock Exchange (TOCOM), which 5,012 price

observations started from 23 January 1990 to 8 April 2009, and palm oil, trading on the Malaysia Derivatives Exchange (MDEX), which 7,425 price observations started from 23 October 1980 to 8 April 2009. These two commodity futures prices are expressed in local currencies and are obtained from Reuters. The returns of agricultural futures prices i of commodity j at time t in a continuous compound basis are calculated as $r_{ij,t} = \log\left(\frac{P_{ij,t}}{P_{ij,t-1}}\right)$, where $P_{ij,t}$ and $P_{ij,t-1}$ are the closing prices of the agricultural futures prices i of commodity j for day t and $t-1$, respectively.

The descriptive statistics for the agricultural commodity futures returns series are summarized in Table 1. The sample mean is quite small, but the corresponding variance of returns is much higher. Surprisingly, 4 of 16 return series have negative average returns, namely cocoa, coffee, cotton and orange juice. The normal distribution has a skewness statistic equal to zero and a kurtosis statistic of 3, but these agricultural commodity futures returns have high kurtosis, suggesting the presence of fat tails, and 9 of 16 return series have negative skewness, signifying the series have a longer left tail (extreme loss) than right tail (extreme gain). The Jarque-Bera (J-B) test lagrange multiplier statistics of the agricultural commodity futures return series are statistically significant, thereby signifying that the distributions for these returns are not normal, which may be due to the presence of extreme observations.

[Insert Table 1 here]

Figures 1-2 present the plots of synchronous agricultural commodity futures returns. These indicate volatility clustering, or periods of high volatility followed by periods of tranquility, such that agricultural commodity futures returns fluctuate in a range smaller than the normal distribution. However, there are some circumstances where agricultural commodity futures returns oscillate in a much wider scale that is permitted by a normal distribution.

[Insert Figures 1-2 here]

The unit root test for all commodity futures returns are summarized in Table 2. All unit root tests are conducted with EViews 6 econometric software package. The Augmented Dickey-Fuller (ADF) and Phillips-Perron (PP) tests were used to explore the existence of unit roots in

the individual returns series. The ADF test accommodates serial correlation by specifying explicitly the structure of serial correlation in the error, but the PP test allows fairly mild assumptions that do not assume a specific type of serial correlation and heteroskedasticity in the disturbances, and can have higher power than the ADF test under a wide range of circumstances. These results are checked by also performing the KPSS test. The null hypothesis of the ADF and PP tests is that the series have a unit root, while the null hypothesis of the KPSS test is that the series are stationary.

In Table 2, based on the ADF and PP test results, the large negative values in all cases indicate rejection of the null hypothesis at the 1% level. In addition, based on the KPSS test, the results indicate that the null hypotheses is not rejected at the 1% level, such that all agricultural commodity futures returns series are stationary.

[Insert Table 2 here]

4. Empirical Results

This section investigates a relevant framework of the conditional variance model through comparison among different specifications. The univariate conditional volatility model, namely the GARCH model of Bollerslev (1986), EGARCH model of Nelson (1991) and APARCH model of Ding et al. (1993), and fractionally integrated class of models, namely the FIGARCH model of Baillie et al. (1996), FIGARCH model of Chung (1999), FIEGARCH model of Bollerslev and Mikkelsen (1996), FIAPARCH model of Tse (1998), and FIAPARCH model of Chung (1996) with Gaussian errors, are estimated by QMLE, which allows for asymptotically valid inferences when the standardized innovations are not normally distributed. Corresponding estimates are obtained using the BFGS algorithm. The computations are performed using the Ox/G@RCH 4.2 econometrics software package of Laurent and Peters (2006).

[Insert Table 3 here]

The univariate estimates of the conditional volatilities, GARCH(1,1), EGARCH(1,1) and APARCH(1,1) of each agricultural commodity futures returns are given in Tables 3 to 5. Their respective estimates and robust t -ratios of each parameters are presented including

information criteria, namely AIC and SIC. Table 3 presents the estimates of the GARCH(1,1) models from equation (1) and (2) for commodity futures returns. The coefficients in the conditional variance equation are all significant, but with corn, cotton and wheat (Kansas City wheat), only in the long run. The details of the univariate estimates relating to the structural properties, namely the second moment and log-moment conditions, based on agricultural commodity futures returns, are available from the authors upon request.

[Insert Table 4 here]

Table 4 shows the estimates of the EGARCH(1,1) models from equations (5) and (6) for commodity futures returns. Most commodity futures returns show that the estimates of γ_2 are statistically significant, meaning that these returns have an asymmetric effect of negative and positive shocks on the conditional variance. Surprisingly, only for cattle (cattle feeder) and pork are both estimates of γ_1 and γ_2 statistically significant, and $\gamma_1 < 0$, indicating that the conditional variance has a leverage effect. However, for the remainder, namely live cattle, cotton, soybeans, soy bean oil and palm oil), the estimates of γ_1 and γ_2 are not statistically significant, meaning that an asymmetric effect of negative and positive shocks on conditional variance is not present. Therefore, the GARCH model is preferred to EGARCH for live cattle, soybeans, soybean oil and palm oil.

[Insert Table 5 here]

Table 5 presents the estimates of the APARCH(1,1) model from equation (7) for agricultural commodity futures returns. The power (δ) term estimated for APARCH is statistically significant for each of these commodity futures returns, ranging from 0.700 for Chicago wheat to 1.779 for soybean oil. The asymmetric (γ) term in the APARCH(1,1) model is statistically significant in 7 of 16 cases, whereas only 5 commodities, namely coffee, soybeans, soybean meal, soybean oil, and Kansas City wheat, has $\gamma < 0$, which means that these conditional volatilities have leverage effects. These results suggest that GARCH is not appropriate for commodity futures returns.

[Insert Table 6 here]

The parameters estimated for the FIGARCH-BBM(1,1) and FIGARCH-Chung(1,1) models are summarized in Tables 6-7. Table 6 presents the estimated FIGARCH-BBM(1,1) from equation (9). The ARCH effects are statistically significant in 10 of 16 agricultural commodity futures returns, while the GARCH effects are statistically significant in 15 of 16 agricultural commodity futures returns. However, the sum of the ARCH(1) and GARCH(1) effects is greater than one in 6 commodities, namely live cattle, cattle feeder, cocoa, coffee, corn and cotton, which indicates nonstationarity. The estimated d parameters in FIGARCH in all commodity futures returns lie between 0 and 1, indicating the stability of the process, but for cotton the estimated d parameter is not statically different from 0, so it reduces to the GARCH model.

[Insert Table 7 here]

The results for the FIGARCH-Chung model from equation (13) are reported in Table 7, and mirror those in Table 6. The GARCH effects are statistically significant for 15 of 16 commodity futures returns, but the ARCH effects are statistically significant for 12 of 16 commodity futures returns. There are 6 commodities, namely live cattle, cattle feeder, cocoa, coffee, corn and cotton, for which the sum of the ARCH(1) and GARCH(1) effects is greater than 1, indicating nonstationarity. However, the estimated d parameters in the FIGARCH-Chung model in all agricultural commodity futures returns are statistically significant, and lie between 0 and 1, thereby indicating the stability of the process. Therefore, the FIGARCH-Chung model is preferred to FIGARCH-BBM.

[Insert Table 8 here]

Table 8 presents the estimates of the FIEGARCH(1,1) model from equation (16). The estimated γ_2 parameters are statistically significant in 9 of 16 agricultural commodity futures returns, meaning these returns have asymmetric effects of negative and positive shocks on the conditional variance. However, only for cattle feeder are both the estimates of γ_1 and γ_2 statistically significant, and $\gamma_1 < 0$, indicating that the conditional variance has a leverage effect. For the remainder, namely live cattle, coffee, cotton, soybeans, soybean oil, palm oil and rubber, the estimates of γ_1 and γ_2 are not statistically significant, such that an

asymmetric effect of negative and positive shocks of equal magnitude on the conditional variance is not observed. Thus, for these agricultural commodity futures return series, the FIGARCH model is preferred to FIEGARCH. The estimated d parameters in FIEGARCH are statistically significant in 9 of 16 cases and lie between 0 and 1, thereby indicating the stability of the process.

[Insert Tables 9 and 10 here]

Tables 9 and 10 show the estimates of the FIAPARCH(1,1) model of Tse (1998) and FIAPARCH(1,1) model of Chung (1999), respectively. Table 9 presents the estimates of the FIAPARCH(1,1) model of Tse (1998) for agricultural commodity futures returns. The power parameter estimates (δ) of all agricultural commodity futures returns are statistically significant, and range from 0.570 for live cattle to 2.184 for orange juice. The asymmetric (γ) term in the FIAPARCH(1,1) model is statistically significant in 9 of 16 cases, but only 6 commodities, namely coffee, soybeans, soybean meal, soybean oil, Chicago wheat and Kansas City wheat, have $\gamma < 0$, so that these conditional volatilities have leverage effects. In addition, the estimated d for all agricultural commodity futures returns is statistically significant. Therefore, the FIGARCH model is not appropriate for modelling agricultural commodity futures returns.

Table 10 presents the estimates of the FIAPARCH(1,1) model of Chung (1999) for agricultural commodity futures returns. The power parameter estimates (δ) of all agricultural commodity futures returns are statistically significant, and range from 0.570 for live cattle to 2.184 for orange juice. The asymmetric (γ) term in the FIAPARCH(1,1) model is statistically significant in 7 of 16 cases, but only for 4 commodities, namely coffee, soybeans, soybean oil, and Kansas City wheat, is $\gamma < 0$, so that these conditional volatilities have leverage effects. In addition, the estimated d parameters for all agricultural commodity futures returns are statistically significant, which leads to the same conclusion as FIAPARCH(1,1)-Chung, namely that FIGARCH is not appropriate for modelling commodity futures returns.

5. Concluding Remarks

The paper estimated the long memory volatility model in 16 agricultural commodity futures returns from different futures markets, namely CBOT for corn, oats, soybeans, soybean meal, soybean oil and wheat; CME for live cattle, cattle feeder and pork; NBOT for cocoa, coffee, cotton, and orange juice; KCBT for wheat; TOCOM for rubber (RSS3); and MDEX for palm oil. The class of fractional GARCH models, namely FIGARCH of Baillie et al. (1996), FIEGARCH of Bollerslev and Mikkelsen (1996), and FIAPARCH of Tse (1998), were estimated and compared with the GARCH model of Bollerslev (1986), EGARCH of Nelson (1991), and APARCH of Ding et al. (1993).

The empirical results showed that, following the outcomes of the unit root tests, all agricultural commodity futures returns series were found to be stationary. The EGARCH (1,1) model out-performed GARCH(1,1), and the APARCH model was also preferred to GARCH(1,1). The robust t statistics of the estimated d parameters, indicating long term dependence, suggested evidence of fractional integration in most agricultural commodity futures markets. Consequently, the fractionally integrated models, namely FIGARCH(1, d ,1) and FIEGARCH(1, d ,1), performed significantly better than traditional conditional volatility models, such as GARCH(1,1) and EGARCH(1,1).

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Table 1. Descriptive Statistics for Agricultural Commodity Futures Returns

Commodity	Mean	Max	Min	S.D.	Skewness	Kurtosis	Jarque-Bera
cattlef	0.001	2.950	-2.610	0.408	-0.047	5.818	2613.18
cattlet	0.002	3.962	-4.224	0.473	-0.392	10.139	16955.27
cocoany	-0.002	5.455	-5.435	0.843	0.086	6.168	3307.76
coffee	-0.0009	10.324	-8.337	0.994	0.164	10.908	20589.09
corn	0.003	4.256	-9.403	0.649	-0.362	14.890	46640.28
cotton	-0.002	7.257	-32.538	0.785	-9.469	388.23	48891741
oats	0.002	5.618	-8.693	0.867	0.030	8.557	10152.27
orange	-0.003	10.378	-6.001	0.800	0.592	15.693	53412.20
pork	0.002	24.890	-14.843	1.111	3.606	76.760	1805221
soybean	0.002	3.276	-5.391	0.627	-9.350	6.882	5114.587
soymeal	0.003	3.855	-6.507	0.675	-0.330	7.819	7776.195
soyoil	0.002	3.786	-3.912	0.672	0.093	4.981	1301.018
wheatc	0.003	4.055	-6.949	0.717	-0.085	9.085	12179.78
wheatk	0.003	3.776	-6.339	0.620	-0.171	9.189	12627.68
palm	0.008	6.638	-7.359	0.781	0.031	11.503	22369.98
rubber	0.005	8.580	-9.684	0.998	-0.547	11.176	14207.00

Table 2. Unit Root Tests for Returns

Commodity	Augmented Dicky-Fuller			Phillip-Peron			KPSS	
	N	C	C&T	N	C	C&T	C	C&T
cattlef	-85.464	-85.459	-85.454	-85.409	-85.404	-85.400	0.036	0.027
cattlet	-85.863	-85.858	-85.853	-85.820	-85.815	-85.810	0.011	0.012
cocoany	-88.423	-88.418	-88.435	-88.423	-88.418	-88.438	0.206	0.019
coffee	-88.719	-88.714	-88.709	-88.736	-88.730	-88.725	0.036	0.029
corn	-84.407	-84.404	-84.400	-84.394	-84.391	-84.387	0.057	0.039
cotton	-87.648	-87.643	-87.638	-87.640	-87.635	-87.630	0.020	0.015
oats	-84.353	-84.348	-84.343	-84.355	-84.351	-84.345	0.025	0.025
orange	-64.808	-64.805	-64.802	-87.357	-87.33	-87.348	0.035	0.028
pork	-84.964	-84.959	-84.954	-85.010	-85.005	-85.000	0.014	0.011
soybean	-88.635	-88.630	-88.629	-88.682	-88.677	-88.675	0.063	0.026
soymeal	-86.635	-86.631	-86.629	-86.637	-86.629	-86.626	0.049	0.024
soyoil	-87.113	-87.108	-87.106	-87.139	-87.134	-87.131	0.058	0.028
wheatc	-87.264	-87.259	-87.255	-87.353	-87.348	-87.345	0.051	0.032
wheatk	-84.906	-84.903	-84.899	-84.909	-84.906	-84.902	0.055	0.038
palm	-35.215	-35.221	-35.219	-84.210	-84.195	-84.191	0.023	0.022
rubber	-68.064	-68.060	-68.055	-68.043	-68.038	-68.033	0.073	0.052

Note: Entries in bold are significant at the 1% level.

Table 3. Estimated GARCH(1,1) Models

Commodity	ω	α	β	AIC	SIC
cattlef	0.001	0.029	0.967	0.919	0.917
	2.503	5.082	135.1		
cattlel	0.001	0.009	0.988	1.287	1.285
	2.334	4.411	334.6		
cocoa	0.004	0.027	0.969	2.427	2.425
	2.382	5.368	155.9		
coffee	0.008	0.060	0.936	2.620	2.619
	2.566	6.275	91.10		
corn	0.016	0.090	0.874	1.806	1.805
	1.029	1.263	8.702		
cotton	0.112	0.022	0.801	2.351	2.350
	0.9963	0.6170	6.106		
oats	0.052	0.124	0.815	2.451	2.450
	2.972	4.170	18.12		
orange	0.004	0.044	0.953	2.280	2.278
	1.380	2.075	43.40		
pork	0.016	0.136	0.877	2.875	2.874
	1.739	5.628	51.05		
soybean	0.004	0.062	0.929	1.707	1.706
	4.085	10.95	146.9		
soymeal	0.006	0.053	0.936	1.897	1.896
	2.985	7.478	101.3		
soyoil	0.008	0.048	0.935	1.935	1.934
	3.736	6.625	85.83		
wheatc	0.009	0.036	0.947	2.072	2.070
	1.805	3.281	48.51		
wheatk	0.011	0.085	0.891	1.703	1.702
	0.664	0.960	7.306		
palm	0.009	0.102	0.886	1.839	1.837
	3.626	8.824	66.72		
rubber	0.035	0.110	0.861	1.686	1.685
	2.633	3.960	24.53		

Notes: (1) The two entries for each parameter are their respective parameter estimates and robust t -ratios.
(2) Entries in bold are significant at the 5% level.

Table 4. Estimated EGARCH (1,1) Models

Commodity	ω	α	β	γ_1	γ_2	AIC	SIC
cattlef	-1.503	0.028	0.993	-0.037	0.051	0.903	0.901
	-12.25	0.063	501.5	-2.371	2.312		
cattlet	-1.117	3.372	0.994	-0.002	0.007	1.061	1.059
	-4.969	0.678	296.4	-0.564	0.852		
cocoany	0.147	-0.078	0.991	0.015	0.082	2.424	2.422
	0.753	-0.204	339.8	1.404	2.394		
coffee	0.909	-0.244	0.989	0.042	0.188	2.290	2.288
	2.612	-1.661	302.4	3.215	5.804		
corn	-0.542	-0.616	0.981	-0.029	0.286	1.790	1.788
	-3.806	-5.060	165.5	-0.880	3.192		
cotton	-0.432	0.206	0.839	0.045	0.043	2.349	2.347
	-3.004	0.156	4.973	0.561	1.141		
oats	-0.001	-0.428	0.965	-0.009	0.257	2.440	2.439
	-0.008	-2.290	61.82	-0.429	5.713		
orange	0.415	-0.329	0.986	-0.011	0.163	2.272	2.270
	1.337	-1.028	112.2	-0.625	4.624		
pork	0.519	0.240	0.992	-0.027	0.086	2.735	2.734
	1.040	0.468	236.5	-2.088	2.064		
soybean	-0.540	1.435	0.987	0.011	0.059	1.701	1.699
	-3.060	1.090	336.3	1.716	1.832		
soymeal	-0.417	-0.027	0.984	0.024	0.120	1.892	1.890
	-2.781	-0.092	230.9	2.586	3.336		
soyoil	-0.582	1.184	0.982	0.007	0.050	1.939	1.937
	-6.521	0.780	185.8	0.991	1.466		
wheatc	-0.421	-0.410	0.983	0.028	0.159	2.058	2.056
	-3.221	-1.907	128.7	0.9882	3.302		
wheatk	-0.657	-0.457	0.979	0.057	0.238	1.687	1.685
	-5.355	-3.794	126.2	2.655	5.799		
palm	-0.054	0.248	0.974	0.010	0.169	1.847	1.845
	-0.324	0.903	169.2	1.403	4.833		
rubber	0.343	-0.583	0.971	0.001	0.387	1.680	1.678
	1.938	-6.038	85.46	0.018	6.549		

Notes: (1) The two entries for each parameter are their respective parameter estimates and robust t -ratios.
(2) Entries in bold are significant at the 5% level.

Table 5. Estimated APARCH (1,1) Models

Commodity	ω	α	β	γ	δ	AIC	SIC
cattlef	0.002	0.027	0.972	0.624	1.176	0.903	0.901
	2.362	4.276	160.6	4.953	8.747		
cattlel	0.003	0.016	0.983	0.337	0.936	1.063	1.062
	1.563	3.605	161.3	1.544	6.264		
cocoany	0.005	0.037	0.965	-0.090	1.245	2.423	2.421
	2.442	6.695	163.0	-1.028	7.124		
coffee	0.010	0.071	0.935	-0.159	1.427	2.615	2.613
	2.806	6.776	94.03	-2.627	10.10		
corn	0.017	0.080	0.914	0.158	1.095	1.794	1.792
	1.852	2.176	20.99	1.514	4.737		
cotton	0.002	0.033	0.974	-0.130	1.021	1.967	1.965
	1.133	4.773	122.5	-0.633	4.221		
oats	0.045	0.108	0.869	0.049	1.032	2.440	2.439
	2.899	4.560	25.43	0.748	9.013		
orange	0.008	0.058	0.950	0.068	1.202	2.272	2.2707
	1.425	2.466	44.09	0.641	6.630		
pork	0.006	0.070	0.945	0.444	0.858	2.825	2.824
	1.636	5.849	113.4	3.518	12.04		
soybean	0.006	0.071	0.935	-0.186	1.240	1.701	1.699
	3.700	10.75	162.6	-3.487	7.972		
soymeal	0.010	0.061	0.941	-0.226	1.039	1.889	1.888
	3.479	7.960	112.4	-2.758	6.768		
soyoil	0.008	0.049	0.939	-0.127	1.779	1.935	1.933
	3.543	6.549	81.92	-2.814	10.34		
wheatc	0.014	0.051	0.945	-0.200	0.700	2.055	2.053
	2.869	5.665	82.52	-1.779	5.667		
wheatk	0.012	0.073	0.925	-0.223	1.074	1.687	1.685
	2.814	5.053	52.11	-2.621	7.527		
palm	0.010	0.106	0.890	-0.057	1.780	1.838	1.836
	3.720	8.975	69.52	-1.574	11.09		
rubber	0.037	0.117	0.870	-0.015	1.507	1.685	1.684
	2.941	4.777	28.75	-0.220	7.973		

Notes: (1) The two entries for each parameter are their respective parameter estimates and robust t - ratios.

(2) Entries in bold are significant at the 5% level

Table 6. Estimated FIGARCH(1,1)-BBM Models

Commodity	ω	d	α	β	AIC	SIC
cattlef	0.008	0.317	0.383	0.648	0.9172	0.916
	2.866	8.382	7.048	10.16		
cattlel	0.025	0.186	0.526	0.681	1.293	1.291
	2.771	5.887	5.062	7.428		
cocoany	0.008	0.317	0.383	0.648	0.917	0.916
	2.866	8.382	7.048	10.16		
coffee	0.019	0.431	0.404	0.716	2.611	2.609
	3.204	7.607	7.822	16.40		
corn	0.031	0.278	0.452	0.555	1.791	1.789
	1.916	5.865	2.310	2.892		
cotton	0.085	-0.009	0.874	0.849	2.351	2.350
	1.489	-0.723	9.280	7.167		
oats	0.085	0.283	0.191	0.321	2.440	2.439
	1.798	6.228	0.686	1.108		
orange	0.057	0.295	0.226	0.439	2.269	2.267
	2.300	8.201	1.259	2.450		
pork	0.035	0.935	-0.119	0.817	2.867	2.866
	3.229	12.99	-1.981	21.66		
soybean	0.003	0.982	-0.014	0.924	1.709	1.708
	2.683	11.36	-0.222	34.93		
soymeal	0.019	0.375	0.338	0.631	1.898	1.897
	2.031	4.903	4.841	5.719		
soyoil	0.023	0.344	0.276	0.592	1.932	1.931
	3.958	9.458	7.788	12.93		
wheatc	0.060	0.216	0.402	0.533	2.068	2.066
	2.054	5.171	2.630	3.347		
wheatk	0.038	0.309	0.153	0.342	1.693	1.691
	2.639	5.378	1.006	2.634		
palm	0.025	0.451	0.074	0.419	1.835	1.834
	2.961	7.399	0.860	3.411		
rubber	0.052	0.335	0.471	0.600	1.677	1.676
	1.666	4.345	2.463	3.063		

Notes: (1) The two entries for each parameter are their respective parameter estimates and robust t -ratios.
(2) Entries in bold are significant at the 5% level.

Table 7. Estimated FIGARCH(1,1)-Chung Models

Commodity	ω	d	α	β	AIC	SIC
cattlef	0.177	0.341	0.373	0.661	0.918	0.917
	4.886	7.784	7.009	10.09		
cattlel	0.229	0.194	0.521	0.683	1.294	1.293
	8.738	5.603	4.965	7.369		
cocoany	0.713	0.340	0.405	0.696	2.429	2.427
	5.801	8.789	9.284	15.63		
coffee	0.753	0.402	0.417	0.701	2.611	2.609
	2.374	8.556	8.338	15.80		
corn	0.423	0.251	0.465	0.554	1.792	1.790
	4.566	6.872	2.300	2.845		
cotton	0.452	0.387	0.359	0.714	1.985	1.984
	3.680	12.73	10.10	21.33		
oats	0.815	0.259	0.189	0.304	2.441	2.440
	5.644	6.752	0.612	0.953		
orange	0.579	0.271	0.232	0.427	2.270	2.270
	4.572	9.713	1.246	2.265		
pork	1.949	0.483	0.245	0.673	2.717	2.716
	2.290	9.418	4.970	14.34		
soybean	0.490	0.472	0.233	0.665	1.711	1.710
	2.875	7.383	5.852	8.618		
soymeal	0.426	0.341	0.343	0.607	1.899	1.898
	4.573	6.086	4.663	6.094		
soyoil	0.466	0.344	0.277	0.593	1.932	1.931
	5.327	10.11	8.007	13.16		
wheatc	0.517	0.207	0.401	0.526	2.068	2.067
	7.706	5.272	2.545	3.194		
wheatk	0.354	0.259	0.167	0.318	1.694	1.693
	3.743	8.001	1.090	2.165		
palm	1.024	0.449	0.083	0.425	1.836	1.834
	1.972	9.195	1.020	3.928		
rubber	1.112	0.310	0.486	0.599	1.678	1.677
	3.082	5.927	2.390	2.995		

Notes: (1) The two entries for each parameter are their respective parameter estimates and robust t -ratios.

(2) Entries in bold are significant at the 5% level.

Table 8. Estimated FIEGARCH(1,1) Models

Commodity	ω	d	α	β	γ_1	γ_2	AIC	SIC
cattlef	-1.115	0.671	0.199	0.594	-0.036	0.041	0.897	0.896
	-6.773	12.08	0.2733	3.440	-2.179	2.054		
cattlel	-0.749	0.712	2.348	-0.871	-0.019	0.033	1.268	1.266
	-3.099	7.814	1.252	-8.347	-1.044	1.361		
cocoany	0.157	0.102	-0.264	0.986	0.015	0.082	2.425	2.422
	0.9041	0.305	-0.4279	44.32	1.386	2.412		
coffee	0.967	0.329	-0.641	0.946	0.041	0.200	2.613	2.612
	3.828	3.622	-5.960	28.95	2.953	6.074		
corn	-0.302	0.236	-0.794	0.958	-0.031	0.306	1.791	1.790
	-1.373	2.238	-8.174	40.68	-0.902	3.438		
cotton	-0.439	0.086	2.979	-0.261	0.057	0.063	2.343	2.341
	-3.396	0.259	1.558	-0.7980	1.438	1.347		
oats	0.372	0.361	-0.348	0.749	-0.010	0.262	2.435	2.433
	1.821	3.877	-0.600	2.501	-0.456	5.408		
orange	0.719	0.473	0.202	0.532	-0.0001	0.150	2.263	2.262
	3.034	10.82	0.3359	4.882	-0.014	3.784		
pork	0.590	0.301	-0.326	0.967	-0.027	0.088	2.734	2.733
	1.325	2.052	-0.999	52.04	1.904	2.026		
soybean	-0.409	0.195	0.785	0.969	0.011	0.056	1.701	1.699
	-2.024	1.836	0.7207	66.44	1.701	1.781		
soymeal	-0.179	0.300	-0.363	0.937	0.028	0.119	1.893	1.891
	-0.5855	1.387	-1.376	15.18	2.552	3.078		
soyoil	-0.283	0.454	0.859	0.827	0.007	0.045	1.936	1.934
	-1.863	7.069	0.6240	15.01	1.085	1.449		
wheatc	-0.298	0.254	-0.666	0.949	0.037	0.166	2.057	2.055
	-2.103	1.967	-4.007	26.73	1.513	3.562		
wheatk	-0.420	0.398	-0.295	0.685	0.060	0.242	1.683	1.681
	-2.394	5.843	-0.6191	3.003	2.835	5.762		
palm	-0.207	-0.226	0.966	0.992	0.008	0.160	1.845	1.844
	-1.050	-2.284	1.905	181.0	1.153	4.323		
rubber	0.873	0.321	-0.771	0.921	0.0001	0.386	1.678	1.676
	2.368	2.737	-9.088	23.89	0.005	6.538		

Notes: (1) The two entries for each parameter are their respective parameter estimates and robust t -ratios.
(2) Entries in bold are significant at the 5% level.

Table 9. Estimated FIAPARCH(1,1)-BBM Models

Commodity	ω	d	α	β	γ	δ	AIC	SIC
cattlef	0.023	0.302	0.367	0.632	0.605	1.438	0.904	0.902
	2.163	3.141	6.439	9.932	3.202	5.574		
cattlel	0.105	0.282	0.343	0.633	0.532	0.570	1.269	1.268
	3.575	5.968	6.794	9.319	3.236	2.666		
cocoany	0.044	0.487	0.329	0.764	-0.139	1.441	2.427	2.425
	3.433	2.317	3.226	6.631	-1.621	2.427		
coffee	0.016	0.405	0.405	0.696	-0.132	2.042	2.609	2.607
	1.431	5.897	6.663	14.78	-2.385	14.31		
corn	0.051	0.414	0.378	0.626	0.079	1.422	1.787	1.785
	1.957	5.685	2.756	4.862	0.7948	4.603		
cotton	0.288	0.342	0.333	0.625	0.049	2.092	1.920	1.918
	2.849	9.604	8.016	16.41	0.774	24.29		
oats	0.129	0.406	0.246	0.510	0.030	1.105	2.432	2.430
	3.155	6.810	2.168	3.831	0.3664	8.080		
orange	0.041	0.260	0.206	0.387	0.034	2.184	2.269	2.267
	1.415	5.016	0.9544	1.681	0.4134	8.244		
pork	1.197	0.385	0.225	0.605	0.771	1.605	2.846	2.844
	1.398	5.705	2.548	4.757	3.255	17.34		
soybean	0.009	0.943	-0.002	0.910	-0.189	1.326	1.700	1.698
	2.355	16.78	-0.059	45.53	-3.487	5.955		
soymeal	0.007	1.062	-0.039	0.957	-0.221	1.088	1.889	1.888
	1.760	15.55	-0.772	47.86	-2.476	5.124		
soyoil	0.483	0.339	0.295	0.601	-0.134	1.940	1.930	1.928
	3.151	10.19	8.113	13.91	-2.752	22.75		
wheatc	0.071	0.417	0.379	0.707	-0.281	0.866	2.056	2.054
	3.354	5.543	5.924	9.658	-2.305	4.950		
wheatk	0.059	0.380	0.275	0.535	-0.252	1.397	1.682	1.680
	2.529	6.148	1.818	3.019	-3.300	7.279		
palm	0.029	0.463	0.087	0.448	-0.055	1.917	1.835	1.833
	2.611	7.587	1.098	3.803	-1.504	17.45		
rubber	0.063	0.430	0.445	0.667	-0.021	1.707	1.677	1.675
	2.309	2.914	3.865	4.619	-0.2819	5.942		

Notes: (1) The two entries for each parameter are their respective parameter estimates and robust t -ratios.

(2) Entries in bold are significant at the 5% level

Table 10. Estimated FIAPARCH(1,1)-Chung Models

Commodity	ω	d	α	β	γ	δ	AIC	SIC
cattlef	0.428	0.297	0.369	0.629	0.604	1.434	0.903	0.901
	4.844	9.006	9.125	13.61	4.719	15.14		
cattlel	0.738	0.187	0.361	0.540	0.867	1.071	1.272	1.271
	5.080	6.670	5.371	7.315	2.727	11.21		
cocoany	0.735	0.342	0.399	0.696	-0.100	1.977	2.428	2.426
	3.984	8.679	9.713	16.40	-1.475	22.29		
coffee	0.612	0.394	0.415	0.693	-0.129	2.036	2.608	2.606
	1.796	7.953	7.770	14.62	-2.334	26.45		
corn	0.416	0.254	0.410	0.506	0.080	2.004	1.791	1.789
	2.646	6.456	2.122	2.770	0.947	14.38		
cotton	0.289	0.341	0.333	0.625	0.049	2.092	1.920	1.918
	2.849	9.604	8.016	16.41	0.774	24.29		
oats	0.850	0.264	0.133	0.271	0.118	1.961	2.440	2.438
	2.722	7.394	0.520	1.032	1.595	14.71		
orange	0.247	0.237	0.189	0.350	0.052	2.239	2.267	2.265
	2.232	7.340	0.798	1.444	0.670	20.58		
pork	1.197	0.385	0.225	0.605	0.771	1.605	2.846	2.844
	1.398	5.705	2.548	4.757	3.255	17.34		
soybean	0.495	0.487	0.242	0.699	-0.196	1.943	1.707	1.705
	2.478	8.227	7.835	12.23	-3.672	26.02		
soymeal	0.422	0.354	0.348	0.631	-0.147	1.964	1.897	1.895
	2.821	6.844	5.817	7.847	-1.906	19.05		
soyoil	0.483	0.339	0.295	0.601	-0.134	1.940	1.931	1.929
	3.151	10.19	8.113	13.91	-2.752	22.75		
wheatc	0.785	0.235	0.526	0.670	-0.383	1.591	2.064	2.061
	4.331	6.732	5.958	7.911	-1.785	14.88		
wheatk	0.180	0.212	0.238	0.355	-0.346	2.024	1.685	1.682
	1.875	6.443	1.103	1.674	-3.389	17.29		
palm	1.058	0.443	0.088	0.428	-0.053	1.962	1.836	1.834
	1.883	9.618	1.094	4.027	-1.471	23.26		
rubber	1.058	0.310	0.492	0.604	-0.038	2.034	1.678	1.676
	2.161	5.355	2.255	2.761	-0.5954	15.82		

Notes: (1) The two entries for each parameter are their respective parameter estimates and robust t -ratios.

(2) Entries in bold are significant at the 5% level.

Figure 1. Logarithm of Daily Agricultural Commodity Futures Returns

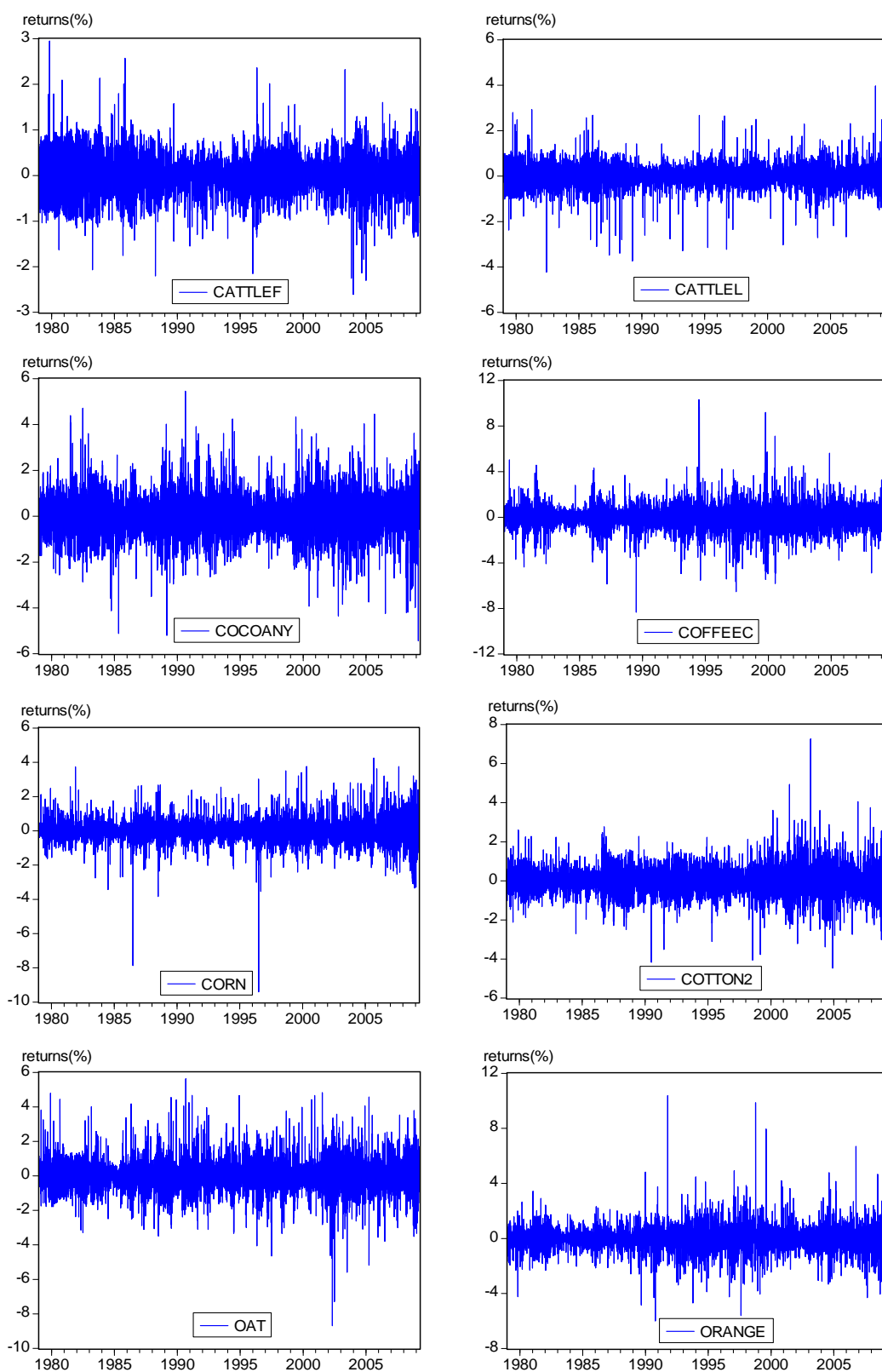


Figure 2. Logarithm of Daily Agricultural Commodity Futures Returns

