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APPENDIX C: DETAILED NOTES TO APPENDIX A AND B TABLES

TABLE A-1

Dates:

Through 1940 the estimates from which the rates of change were calculated refer to census or mid-census dates, as follows:

<i>Census Dates</i>	<i>Mid-Census Dates</i>
June 1, 1870	June 1, 1875
June 1, 1880	June 1, 1885
June 1, 1890	June 1, 1895
June 1, 1900	May 8, 1905
April 15, 1910	February 21, 1915
January 1, 1920	February 14, 1925
April 1, 1930	April 1, 1935
April 1, 1940	

For 1945 through 1959, the dates are approximately April 1. Because of the estimating technique used in the source, the date of reference is not precise. Since the effect on the rates of change was small, we continued to use the original source for the sake of consistency rather than shift to the more precisely dated Bureau of the Census estimates.

Source and method:

Through 1955 the basic estimates from which the rates of change were computed are from [27, p. 37, Table 1, cols. (1) and (3)]. In the source, the estimate for 1870 is adjusted for general underenumeration of the population and those for all other dates for underenumeration of children aged 0-4. The 1959 estimate was obtained by extrapolating the 1955 figure shown in the source on the basis of the Bureau of the Census estimates for total white population (including adjustment for underenumeration of children aged 0-4) for July 1, 1955 and 1959 [53, No. 146 (Nov. 12, 1956), p. 7, and No. 212 (Jan. 26, 1960), p. 9].

TABLE A-2

Dates:

The estimates refer to calendar years; thus "1855-59" refers to January 1, 1855 through December 31, 1859.

Source:

Column 1 [73].

Column 2 1910-1954 [16, p. 26, Table 11].

1955-59 [58, p. 52, Table No. 52].

In both sources birth rates are adjusted for undercount.

TABLE A-3

Dates:

For the fertility ratio, the precise intervals spanned are given by the mid-census and census dates bordering the periods shown in the notes to Table A-1. For the general fertility rate, the intervals covered are calendar years.

*Source and method:**Fertility ratio, 1865-69/1925-29*

Column 1 Table A-3a, col. (1) divided by Table A-3b, col. (1).

Column 2 Table A-3a, the sum of col. (3) and (7) divided by Table A-3b, col. (2).

Column 3 Table A-3a, the sum of col. (5), (8), and (9) divided by Table A-3b, col. (3).

Details on source and method are given in the notes to Tables A-3a and A-3b. For the native white population, the present estimates draw heavily on a valuable unpublished memorandum prepared by Everett S. Lee of the University of Pennsylvania Study of Population Redistribution and Economic Growth, which provides age and parentage detail underlying the quinquennial series for 1870 to 1940 published by Kuznets [27]. This memorandum is subsequently cited as U. of P. Memo. Some of the principal features of the present estimates relevant for evaluation and comparison with other series follow:

1. To correct for underenumeration of children under 5, native whites aged 0-4 at census and mid-census dates were estimated by applying reverse survival ratios to the population aged 5-9 and 10-14 at the subsequent census (U. of P. Memo, pp. 6-8). The implied underenumeration adjustment, which varied from census to census, was assumed by us to apply to foreign-born white children 0-4. (See below, notes to Table A-3a, col. (9).
2. Mid-census estimates of foreign-born white children aged 0-4 were derived from an estimate of the ratio of children to immigrant women in the reproductive ages. See below, notes to Table A-3a, col. (9).
3. To distribute native white children aged 0-4 by nativity of mother required at most dates an estimate of (a) the distribution of native white children of foreign or mixed parentage between the foreign and mixed categories, and (b) the distribution of native white children of mixed parentage by nativity of mother. While the census reports provided a satisfactory source for the estimate of (a), the basis for estimating (b) appeared somewhat less reliable. See below, notes to Table A-3a, cols. (5)-(8). However, the available estimates of (b)

TABLE A-3 (continued)

for three points in time did not suggest high variability, and the magnitudes involved are small enough so that the patterns of rates of change over time in the fertility ratios, our ultimate concern, would not be noticeably altered even by significant changes in the estimate of (b).

4. To obtain mid-censal estimates of native white females aged 20-44, it was necessary to apportion the decade total of deaths by quinquennium. For this purpose United States life tables at approximately decade intervals from 1900 through 1950 were used. For decades before 1900, the allocation of 1900-1910 was assumed to hold; and for 1910-20, a special adjustment to allow for the influenza epidemic was introduced (U. of P. Memo, pp. 3-6). In deriving mid-censal estimates of foreign-born white women aged 20-44, an apportionment by quinquennium of the decade total of both deaths and immigration was necessary. For deaths the proportionate allocation for native white women was used; for immigration, the underlying source was the annual immigration data (see below, notes to Table A-3b, col. 3).

Some evaluation of the present estimates can be obtained by comparing their consistency with alternative series. Unfortunately, this can be done only for the total white population, and thus at best only a minimum test of the estimates can be secured.

For the *decade* rates of change, two alternative series are available: Zelnick's estimates of the crude birth rate (Table A-2, cols. 1 and 3) and those of Grabill, Kiser, and Whelpton of the fertility ratio [16, p. 28]. The comparison for the total white population follows:

<i>Period</i>	Crude Birth Rate (annual average) (1)	<i>Fertility Ratio</i>		<i>Change Since Preceding Period</i> (per cent per quinquennium)		
		Grabill et al. (2)	Present (3)	Col. 1 (4)	Col. 2 (5)	Col. 3 (6)
1865-69	39.7	814	877	—	—	—
1875-79	38.0	780	812	-4.3	-4.2	-7.4
1885-89	35.3	685	744	-7.1	-12.2	-8.4
1895-99	31.2	666	665	-11.6	-2.8	-10.6
1905-09	29.4	631	632	-5.8	-5.3	-5.0
1915-19	26.9	604	614	-8.5	-4.3	-2.8
1925-29	21.5	506	505	-20.1	-16.2	-17.8

Conceptually, the estimates of Grabill, Kiser, and Whelpton are identical with ours, even to the extent of including an adjustment for census underenumeration of children under 5 years old, though, because of a difference in the method of adjustment, the figures for the two series differ, especially before 1895-99. Differences between the pattern shown by the crude birth rate and the fertility ratio would be expected on conceptual grounds because

TABLE A-3 (continued)

of variations in the ratio of women of reproductive age to the total population, in the mortality experience of children under 5, and in the dating of the estimates. Also, while the basic source for all three series is the federal census, the specific censuses used in deriving the estimates in col. (1) differ from those in col. (2) and (3).

The rates of change of all three series appear fairly consistent. The biggest exception, the disparate pattern shown by the Grabill, Kiser, and Whelpton series in the early part of the period, is accounted for by the fact that this series employs the same underenumeration adjustment at all dates, though it is recognized that in 1890 underenumeration was disproportionately bad [16, pp. 13, 54]. The only other case where the rates of change fail to show the same direction of movement is between the first and second decades of this century, when the crude birth rate series shows an increased rate of decline, while the other two show a smaller rate. This appears to reflect a conceptual difference—a much faster decline in infant mortality in the second decade causing the fertility ratio to fall noticeably less than the crude birth rate.

For the *quinquennial* rates of change, we may compare the pattern shown by the crude birth rate estimates (Table A-2, col. 3) with the present figures for the fertility ratio (Table A-3, col. 4):

<i>Period</i>	<i>Change Between Specified Periods</i> (per cent per quinquennium)	
	Crude Birth Rate (Zelnick)	Fertility Ratio (present)
1865-69/1870-74	0.0	-2.5
1870-74/1875-79	-4.4	-5.2
1875-79/1880-84	-5.1	-3.6
1880-84/1885-89	-2.2	-5.1
1885-89/1890-94	-3.8	-2.9
1890-94/1895-99	-8.6	-8.4
1895-99/1900-04	-8.0	-4.5
1900-04/1905-09	+2.1	-0.6
1905-09/1910-14	-4.2	-3.5
1910-14/1915-19	-4.7	+0.7
1915-19/1920-24	-6.5	-4.7
1920-24/1925-29	-15.8	-14.8

There are some differences in the first part of the period—and one would of course expect some disparity because of differences in concept and source, but after 1885-89/1890-94 the direction of change in the rates of change of the two series is identical, with one exception: the relative standing of the rate of change from 1910-14/1915-19 is noticeably different in the two series. This disparity appears primarily to be owing to the fact that in the present estimates no adjustment was made for the influenza epidemic of 1918, and, as a result, the 1915 estimate of children under 5 and therefore

TABLE A-3 (continued)

of the fertility ratio is understated. (This is probably true also of the 1910 estimate, but only to a minor extent.) Since this deficiency did not appear to have serious consequences for the analysis, no attempt was made to correct for it.

General fertility rate, 1920-24/1954-58

Column 1 Average of annual data in [55, Series B-23]. Recent-year data from [58, p. 56].

Column 2 Average of annual data in [55, Series B-24]. Recent-year data from [67, pp. 3-22].

TABLE A-3a

Dates:

The precise dates are the census or mid-census dates falling at the end of the intervals used in Table A-3, and are given in the notes to Table A-1.

Source and method:

Column 1 Col. (2) plus col. (9).

Column 2 U. of P. Memo., Table I, except 1870 and 1875, Table II-B.

Columns 3,4 U. of P. Memo., Table III. A small proportional adjustment was made to the 1920-30 figures so that their sum corresponded to the total native white figure of Table I, Estimates for 1880 and 1885 were obtained by applying the appropriate "reverse survival" ratios (given in Appendixes G and H of the memo) to the 1890 populations aged 5-9 and 10-14.

Columns 5,6 Since distributions of the "foreign or mixed parentage" category by component for each five-year age group are not available for census dates before 1910, the distribution was estimated for each sex separately for each date prior to 1910 on the basis of the 1910 distribution for the appropriate cohort. Thus the 1880 figure for native white males aged 0-4 of foreign or mixed parentage was divided into foreign and mixed components on the basis of the published distribution for the corresponding group aged 30-34 in 1910, and so on. The specific source used to obtain the distribution at each date was Bureau of Census, *1930 Census*, II, pp. 577, 579, 583.

Columns 7,8 The key figures used in distributing col. (6) between cols. (7) and (8) were estimates of the percentage of native white children of mixed parentage with native-born mother for 1890 and 1900 (58.4 per cent), 1910 and 1920 (62.2 per cent), and 1930 (67.2 per cent). For years prior to 1890, the 1890 distribution was assumed to apply, and for all other years the distribution was estimated by simple interpolation between the values for adjacent census dates. Sources used were:

TABLE A-3a (continued)

1890 and 1900 An estimate of native white children aged 0-4 by nativity of mother in Bureau of Census, *1900 Census. Special Report, Supplementary Analysis*, pp. 434, 436.

1920 A corresponding estimate in [57, p. 201].

1930 An estimate of the distribution of the mixed parentage group aged 0-4 by nativity of mother in Walter F. Willcox, *Studies in American Demography* (Ithaca, N. Y., 1940), p. 274.

The 1890-1900 and 1920 estimates were used in conjunction with our estimates above in cols. (5) and (6) to obtain the implied distribution of the mixed parentage group, aged 0-4, by nativity of mother. The 1910 figure was assumed the same as that for 1920. For both of these years our figure for the distribution of the mixed parentage group by nativity of mother appears to be the same as that implicit in the figures of the National Resources Committee [64, p. 30], though an exact check was not possible because the absolute figures underlying the committee estimates were not published. For 1930 the present estimate appeared slightly higher than that implicit in the National Resources Committee figures.

Column 9 Census dates: For each sex the census figure for foreign-born whites aged 0-4 [49, p. 16] was adjusted upward by the underenumeration ratio for native whites. (U. of P. Memo, Appendix F.) For 1870, the 1880 underenumeration ratio was used.

Mid-census dates: An estimate of the ratio of foreign-born white children aged 0-4 to female immigrants aged 20-44 during the preceding quinquennium was applied to the estimated mid-census figure for the latter. The ratio was estimated by simple interpolation between corresponding ratios for the census dates. Estimates of female immigrants aged 20-44 by quinquennium were obtained in connection with the derivation of Table A-3b.

TABLE A-3b

Dates:

See notes to Table A-3a.

Source and method:

Column 1 Col. 2 plus col. 3.

Column 2 Census dates, 1880-1930, from basic census reports.

All other dates: U. of P. Memo, Table I, except 1870 and 1875, Table IIB. (The 1870 figures have been adjusted for general underenumeration of the population.) Where necessary, ten-year age groups in the U. of P. Memo were subdivided on the basis of the average proportions of the five-year component groups in the same cohort at adjacent census dates.

TABLE A-3b (continued)

Column 3 Census dates from basic census reports.

Mid-census dates: For each five-year class within the 20-44 age group, the beginning of decade number in the cohort was increased by the number of net migrants of that age during the quinquennium and decreased by the estimated number of deaths in the cohort. Net migrants and deaths were estimated as follows for each five-year age group:

1. The decade total of net migrants was obtained by surviving the beginning of decade population to the end of the decade and subtracting the "expected" population thus obtained from the actual. Deaths during the decade were obtained as the difference between the beginning of decade population and end of decade "expected" population. Survival rates used were those in [31, p. 23].
2. The decade total of net migrants was distributed between quinquennia according to the estimated share of each quinquennium in the net migrants of a given sex (all ages) during the decade. For decades before 1910, this distribution was computed from worksheets underlying [29, Table B-6]. We are grateful to Professor Kuznets for making these worksheets available. For subsequent decades the figures of the U.S. Immigration and Naturalization Service, published regularly in the *Statistical Abstract of the United States*, were used.
3. Decade totals of deaths were distributed by quinquennia in the same proportions as for native whites of the corresponding age group (U. of P. Memo, Appendix B).

TABLE A-4

Dates:

For the periods through 1935-39, the precise intervals spanned are given by the mid-census and census dates bordering the periods shown in the notes to Table A-1. For 1945-49 and 1954-58, the precise dates are the five years preceding respectively April 1, 1950, and July 1, 1959.

Source and method:

Native white, 1885-89, 1895-99. Estimates of the present study, constructed so far as possible on a basis consistent with the National Resources Committee figures for the subsequent three dates [64, p. 30]. In general the aim was first to derive the rural native white fertility ratio from estimates of the ratios for the rural white and rural foreign-born white. The basis for the estimate for the rural foreign-born was less reliable than that for the rural total white, but since the foreign-born accounted for only 10 per cent of rural females aged 20-44, substantial variation in the estimate would affect that for the rural native white only slightly. The latter was then used together with the estimate in Table A-3 for the total native white (adjusted to the National Resources Committee underenumeration basis) to obtain

TABLE A-4 (continued)

the urban native white fertility ratio.

The basic sources used were the census reports and [45]. The figures in the latter are also based chiefly on the census.

At several points an extrapolation was made, using an index tested against the record of the following decades through 1925-29. Where this was done, figures are presented below for all five dates from 1885-89 through 1925-29. The detailed procedure follows (underscored figures are estimates made in the present study) :

1. The rural native white fertility ratio was calculated from the following formula (numbers in brackets indicate the step in which the specified item was estimated) :

$$\begin{aligned} \text{Fert.Rat.,} & \\ \text{Tot. Wh. Rur.} & = \left\{ \begin{array}{l} \text{Fert.Rat.,} \\ \text{Na. Wh. Rur.} \end{array} \right\} \frac{\text{Na. Wh.Rur.F. 20-44}}{\text{Tot.Wh.Rur.F. 20-44}} + \\ [2] & \qquad \qquad \qquad [3] \\ & \qquad \qquad \qquad \left\{ \begin{array}{l} \text{Fert.Rat.,} \\ \text{FBW Rur.} \end{array} \right\} \left\{ \frac{\text{FBW Rur.F. 20-44}}{\text{Tot.Wh.Rur.F. 20-44}} \right\} \\ & \qquad \qquad \qquad [4] \qquad \qquad \qquad [3] \end{aligned}$$

2. The rural total white fertility ratio was estimated by extrapolation via the fertility ratio for places of 25,000 population or less, and then adjusted for 5 per cent underenumeration of children under 5 years of age, as assumed in the National Resources Committee figures for the later dates.

Details follow :

Line	Fertility ratio of total white population in:	1885-89	1895-99	1905-09	1915-19	1925-29
2a	Places of 25,000 population or less	731	722	<u>708</u>	<u>680</u>	<u>587</u>
2b	Rural areas	<u>805</u>	<u>796</u>	782	744	653
2c	Rural areas, adjusted	<u>845</u>	<u>836</u>	<u>821</u>	<u>781</u>	<u>686</u>

Line 2a 1885-89, 1895-99 [45, p. 279].
 1905-09/1925-29 Calculated as weighted average of fertility ratios of white population in places of 2,500-25,000 population and in rural areas. Weights were the proportionate distributions of the white population aged 20-44 between these places [45, p. 130]. Figures for both sexes rather than females alone were used, since those for the latter were not readily available.

TABLE A-4 (continued)

Line 2b 1885-89, 1895-99 Extrapolated from 1905-09 on the basis of absolute change in line 2a.
1905-09/1925-29 [45, p. 279].

Line 2c 1885-89/1925-29 Line 2b \times 1.05. For the last three dates the figures check almost exactly against those implicit in the ratios for rural native white and rural foreign-born white published in [64, p. 30].

3. The percentage distribution of rural white females aged 20-44 by nativity was estimated by extrapolation via the corresponding nativity distribution for places of 25,000 population or less, as shown below:

<i>Line</i>	1890	1900	1910	1920	1930
3a Percentage foreign-born of white population (both sexes) aged 20-44 in places of less than 25,000 population	16.54	<u>14.67</u>	14.99	12.19	7.95
3b Percentage foreign-born of white females aged 20-44 in rural areas	<u>10.36</u>	<u>9.19</u>	9.39	8.41	6.34

Line 3a 1890, 1910-30 [45, pp. 130, 150, 370].
1900 Interpolated between 1890 and 1910 on the basis of change in percentage for corresponding populations aged 15-44 and 25-44, Bureau of Census, *1910 Census*, I, p. 433.

Line 3b 1890, 1900 Extrapolated from 1910 on basis of relative change in line 3a.
1910 *1910 Census*, I, 427; *1920 Census*, II, 371.
1920 *1920 Census*, II, 371.
1930 *1940 Census*, II, 24-25.

4. The rural foreign-born white fertility ratio was estimated by extrapolation via the movement in the fertility ratio for the total foreign-born white, as follows:

TABLE A-4 (continued)

Line	Fertility ratio of foreign-born white population:	1885-89	1895-99	1905-09	1915-19	1925-29
4a	Total	<u>845</u>	<u>860</u>	775	818	513
4b	Rural	<u>1085</u>	<u>1100</u>	1017	1048	694

Line 4a 1885-89, 1895-99 Table A-3, col. (3), readjusted to a constant 5 per cent allowance for underenumeration of children 0-4 used in National Resources Committee estimates for 1905-09 on. 1905-09/1925-29 [64 p. 30].

Line 4b 1885-89, 1895-99 Extrapolated from 1905-09 on basis of absolute change in line 4a. 1905-09/1925-29 [64 p. 30].

5. The urban native white fertility ratio was calculated from the following formula (numbers in brackets indicate the step in which the specified item was estimated) :

$$\frac{\text{Fert.Rat., Tot.Na.Wh.}}{[6]} = \left\{ \frac{\text{Fert.Rat.,}}{\text{Ur.Na.Wh.}} \right\} \left\{ \frac{\text{Ur.Na.Wh.F. 20-44}}{\text{Tot.Na.Wh.F. 20-44}} \right\} + \left\{ \frac{\text{Fert.Rat.,}}{\text{Rur.Na.Wh.}} \right\} \left\{ \frac{\text{Rur.Na.Wh.F. 20-44}}{\text{Tot.Na.Wh.F. 20-44}} \right\}$$

[1] [7]

6. The fertility ratio of the total native white population (Table A-3), which contains a variable allowance for underenumeration of children under 5, was adjusted to a straight 5 per cent allowance at all dates to obtain consistency with the National Resources Committee estimates for the later dates :

$$706 \times \left\{ \frac{1.05}{1.1049} \right\} = 671 \qquad 628 \times \left\{ \frac{1.05}{1.0447} \right\} = 631$$

7. The percentage distribution of native white females aged 20-44 by rural-

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TABLE A-4 (continued)

urban residence was computed from the following formula (all ratios refer to females aged 20-44):

$$\frac{\text{Rur.Na.Wh.}}{\text{Tot.Na.Wh.}} = \frac{\left\{ \frac{[7a]}{\text{Na.Wh.Rur.}} \right\} \left\{ \frac{[7b]}{\text{Tot.Wh.Rur.}} \right\}}{\left\{ \frac{\text{Tot.Na.Wh.}}{\text{Tot.Wh.}} \right\} [7c]}$$

Item 7a was estimated in step 3; item 7c was computed from the census and item 7b was derived by extrapolation via the movement in the corresponding ratio for both sexes in places of 25,000 population or less, as shown below:

Line	1890	1900	1910	1920	1930
7b Percentage of white population (both sexes) aged 20-44 in places of less than 25,000 population	72.25	<u>67.84</u>	62.53	57.79	53.80
7b> Percentage of white females aged 20-44 in rural areas	<u>54.50</u>	<u>50.09</u>	44.78	40.52	36.08

Line 7b 1890, 1910-30 [45, pp. 130, 150, 370].
 1900 Interpolated between 1890 and 1910 on basis of change in percentage for corresponding populations aged 15-44 and 25-44.

Line 7b> 1890, 1900 Extrapolated from 1910 on basis of absolute change in line 7b.
 1910-30 1940 *Census*, II, pp. 23-25, 28-9.

Native white, 1905-09/1925-29 [64, p. 30]. The estimates include an allowance at each date for 5 per cent underenumeration of children under 5 years of age.

Total white, 1925-29/1945-49 [16, p. 17]. The figures are not adjusted for underenumeration of children under 5.

Total white, 1954-58 Col. (1) is derived from [53, No. 212 (January

TABLE A-4 (continued)

26, 1960), p. 9]. Col. (2) was estimated on the assumption that the relative change for urban white is the same as that for total white shown in col. (1).

TABLE A-5

Dates:

The figures refer to the census dates listed in the notes to Table A-1.

Source:

Column 1 [49, p. 16].

Column 2 Computed from the census sources cited in [45, p. 397, note a].

TABLE A-6

Dates:

Column 1 The intervals spanned are given by the mid-census and census dates bordering the periods shown in the notes to Table A-1.

Columns 2,3 The ratios are for the mid-census or census dates at the end of the periods given in the notes to Table A-1.

Source and method:

Column 1 Table A-3, col. (3).

Columns 2,3 Census dates: From basic census reports.

Mid-census dates: Computed from absolute figures estimated by the procedure described in the notes to Table A-3b, col. (3).

TABLE A-7a

Dates:

Column 1 The precise dates of the intervals spanned are given by the mid-census and census dates bordering the periods shown in the notes to Table A-1. Thus 1885-89 refers to the period June 1, 1885, through June 1, 1890.

Column 2 The dates refer to the calendar years preceding by one those listed in the table. Thus 1885-89 refers to the period January 1, 1884, through December 31, 1888. The lead of col. (2) over col. (1) therefore varies slightly because of variations in the calendar intervals covered by the figures in col. (1). At its maximum (in 1885-89 and 1895-99) the lead is one year and five months; at its minimum (in 1915-19) it is slightly over one year.

TABLE A-7a (continued)

Source and method:

Column 1 See notes to Table A-4, step 2 of the explanation of the derivation of the figures for native white, 1885-89 and 1895-99. The adjusted figure for 1885-89 given in footnote a to Table A-7a was computed by assuming an underenumeration of children under 5 of 10 per cent, the approximate underenumeration in 1890 of total native white (urban and rural combined) aged 0-4 suggested by the U. of P. Memo used in the preparation of Table A-3.

Column 2 Annual estimates of real gross farm income per engaged were averaged for appropriate quinquennial intervals (see "Dates" above) and converted to index form on 1924-28 as 100. The annual estimates were calculated from separate figures for gross farm income in current prices, persons engaged in farming, and a consumer price index, as given below. A comparison for the overlap period 1910-14/1930-34 of the rates of change in quinquennial averages of real gross farm income per engaged with those in real net farm income per capita (used in Table A-7b) indicated that the former constituted a reasonable approximation to the latter.

Gross farm income in current prices [39, p. 24, Table 8]. Figures are for calendar years and are adjusted for inventory change.

Persons engaged in farming [23, Table A-VI, pp. A-115-116]. Annual estimates for the decade of the eighties were derived by linear interpolation between Kendrick's 1890 figure and an 1880 estimate, extrapolated from 1890 on the assumption that the decade rate of change was the same as that shown by Tostlebe's figures for the decade of the eighties [47, p. 46, Table 4].

Consumer price index [34, pp. 150-51]. A comparison for the overlap period 1910-58 of the movements in this price series, which technically refers to retail prices of goods and services purchased by city wage earners and clerical workers, with those in the Department of Agriculture index of prices paid by farmers for family living indicated that the former was a reasonable approximation to the latter.

TABLE A-7b

Dates:

Column 1 Dates refer to year beginning April 1. Thus 1920-24 refers to the period April 1, 1920, through March 31, 1925.

Column 2 The dates refer to the calendar years preceding by one those listed in the table. Thus 1920-24 refers to the period January 1, 1919, through December 31, 1923. The lead of col. (2) over col. (1) is therefore 1.25 years.

TABLE A-7b (continued)

Source and method:

Column 1 Annual estimates of the birth rate of the farm population were averaged for quinquennial intervals. Birth rates were computed by dividing the annual number of births by the mean of the beginning and end-of-year farm populations. For 1920-1949, [63, pp. 8-14]. For 1950-58, [62, p. 6].

Column 2 Annual estimates of real net farm income per head of farm population were calculated from separate figures for net income to persons on farms from farming, farm population, and prices paid by farmers for family living.

Net income to persons on farms from farming For 1920-54, [55, series K-128]. For 1955-58, [60, p. 488, Table 688].

Farm population Estimates of April 1 of each year given in the sources above for col. (1) were used to place the income figures on a per capita basis.

Prices paid by farmers for family living 1920-54, [55, series K-132]. For years since 1954, [60, p. 479, Table 682].

TABLE A-8

Dates:

Column 1 For the periods through 1935-39, the intervals spanned are given by the mid-census and census dates bordering the periods shown in the notes to Table A-1. For 1945-49 and 1954-58, the periods are the five years preceding respectively April 1, 1950, and July 1, 1959.

Column 2 The intervals cover the calendar years one year prior to the periods specified. Thus the entry for 1885-89 refers to the period January 1, 1884, through December 31, 1888, and so on.

Column 3 Through 1935-39, the intervals are given by the census and mid-census dates bordering the period shown in the notes to Table A-1. For the overlap value for 1935-39 and for 1945-49 and 1954-58, the intervals are respectively July 1, 1935—July 1, 1940; July 1, 1945—July 1, 1950; and July 1, 1954—July 1, 1959.

Source and method:

Column 1 Table A-4, col. (2).

Column 2 Annual estimates of percentage of civilian labor force unemployed were averaged for quinquennial intervals. For 1900-1948, [30, p. 215].

For 1944-48 (overlap value) and 1953-57, [59, p. 1, Table A-1]. For 1894-98, an unpublished extension of this series kindly provided by Stanley

TABLE A-8 (continued)

Lebergott was used. The 1884-88 estimate was made by the present writer on the basis of Frickey's annual index of industrial and commercial production [14, p. 128], adjusted for secular trend, as follows. In each of the three quinquennial periods 1899-1903, 1904-08, and 1909-13, Frickey's index was only slightly below the 1884-88 standing of 101. Lebergott's estimates of percentage unemployed in these quinquennia varied between 3.8 and 5.4 per cent, and we therefore chose a figure of 5.0 for 1884-88, preferring to err on the side of overestimating unemployment so as to obtain a conservative estimate of the deterioration from 1884-88 to 1894-98.

Column 3 Through 1935-39, computed from estimates for census and mid-census dates derived as follows: census dates, from basic census reports; mid-census dates, derived by summing separate estimates for native white (U. of P. Memo, Table I) and foreign-born white, the latter derived by the procedure explained in the notes to Table A-3b, col. (3). For 1935-39 (comparable to later years) on, figures from which rates of change were computed are from [53, No. 98, p. 16; No. 114, p. 6; No. 146, pp. 7, 10; No. 212, p. 9].

TABLES B-1 AND B-2

The underlying method consisted of holding the values of all components constant at their beginning-of-period levels, except for the component whose contribution was being assessed, and computing the change in the total that would have resulted from the change in this component alone. Variations in this procedure are of course possible, but in the present case the differences would be small and not significant for our purpose. The contribution of the interaction term shown here is derived as a residual but a direct computation was also made which checked within a point or two of the present figures.

An illustrative calculation is provided below for Table B-1:

	<i>Fertility Ratio</i>			<i>Percentage Distribution of White Women Aged 20-44</i>		
	Total White	Native White	Foreign- Born White	Total White	Native White	Foreign- Born White
1895-99	665	628	819	100.0	80.5	19.5
1900-04	636	606	768	100.0	81.1	18.9

Contribution of change in native white fertility ratio: $.805 (606-628) = -18$; of change in foreign-born white fertility ratio: $.195 (768-819) = -10$; of change in nativity distribution of white females aged 20-44: $(.189-.195) (819-628) = -1$.