

NEXT YEAR ON THE U.S. FARMLAND MARKET: AN INFORMATIONAL APPROACH

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*Paper for presentation at the 2005 AAEA annual meeting
Providence, RI
July 24-27, 2005*

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Abstract: This paper formulates an information measure for changes in asset values and applies the formulation to farmland values in the United States for 1960-99. The results indicate that changes in asset values contained significant information following the Russian wheat sale in the early 1970s and the financial crisis in agriculture in the mid 1980s. Further, information about preceding year's asset value largely explains the regional distribution of current year's farmland values.

Keywords: information measurement, farmland values

1. Introduction

Measuring the informational content of asset values is not new. Theil and Leender (1965) and Fama (1965) both present methodologies for measuring the informational content of stock markets based on informational measures (Theil, 1967). Both studies focused on information in asset prices based on the number of stocks advancing, remaining constant or declining from one trading day to the next. Based on these measurements, Theil and Leender concluded that stock markets in Amsterdam were not informationally efficient while Fama concluded that the New York stock market was efficient. This study departs from the formulation in these studies to examine the information content of changes in relative asset values. Our formulation allows for a regional decomposition of the information in asset value changes. The informational measure is then applied to farmland values in the United States. The results indicate that significant information in farmland values in the United States occurred during the mid 1970s and mid 1980s.

2. Measurement of Information

The general formulation of the information measure is presented in detail in Theil (1967).

The information measure used by both Theil and Leenders and Fama can be expressed as

$$I = \sum_{i=1}^N p_i \ln \left(\frac{p_i}{q_i} \right) \quad (1)$$

where I is the information index (or the information inequality), p_i is the posterior probability, q_i is the prior probability, and N is number of observations. The information approach assumes that the two sets of probabilities embody the effect of a signal. As described in Theil and Leenders, I measures the information in that signal or, more precisely, the difference between the distribution functions resulting from observing some signal. In both Theil and Leenders and Fama, the difference in the number of stocks advancing, remaining constant, or declining between two time periods provided evidence of trends in the stock market. If the measure of information was small, the stock market possessed a memory.

The specific application of the information index in equation 1 is one of a host of uses to which information inequality has been applied. Theil (1967) demonstrates how the index could be applied to measure income inequality. Other studies such as Theil (1989) built on this application. In this application, q_i is defined as the share of income to group i (usually defined by a geographic region such as a state or country) and p_i is defined as the population share in region i . In addition to demonstrating the versatility of the information measure, these applications emphasize the measure's decomposability. Specifically, grouping the geographic regions into C different groups ($c = 1, \dots, C$) such that each state belongs to one group, it is possible to show that

$$\begin{aligned}
I &= \bar{I} + I_R \\
I_R &= \sum_{c=1}^C Q_c \ln \left(\frac{Q_c}{P_c} \right) \\
Q_c &= \sum_{i \in c} q_i, P_c = \sum_{i \in c} p_i \\
\bar{I} &= \sum_{c=1}^C Q_c I_c \\
I_c &= \sum_{i \in c} \frac{q_i}{Q_c} \ln \left(\frac{q_i / Q_c}{p_i / P_c} \right)
\end{aligned} \tag{2}$$

where \bar{I} is the average inequality in each region, I_R is the inequality across regions, Q_c is the relative share of q_i in region c , P_c is the relative share of p_i in region c , and I_c is the inequality within region c . Equation 2 allows for the decomposition of the information in the signal into an average within region measure and a measure of the information in the signal across regions.

3. Measurement of Information in Asset Values

We propose to measure the relative information in asset values by comparing the relative value in the share of asset values across time periods. Specifically, we define $v_{i,t}$ as the relative share of asset values in state i at time t :

$$v_i = \frac{V_{i,t}}{\sum_{i=1}^N V_{i,t}} = \frac{a_{i,t} l_{i,t}}{\sum_{i=1}^N a_{i,t} l_{i,t}} \tag{3}$$

where $V_{i,t}$ is the total value of farmland in state i at time t , $a_{i,t}$ is the acres of farmland in state i at time t and $l_{i,t}$ is the price of land per acre in state i at time t . We can then define a measure of relative change in asset prices over time as

$$I = \sum_{i=1}^N v_{i,t+1} \ln \left(\frac{v_{i,t+1}}{v_{i,t}} \right). \tag{4}$$

This inequality then measures the relative persistence in the spatial value of land prices. From equation 3 it is apparent that the changes in information index may be the result of either relative changes in farmland prices or spatial changes in the acres of farmland in each state. Accordingly, a similar measure as presented in equation 4 can be derived to analyze relative changes in the total acres in farmland in each state.

Table 1 presents the information measure utilizing the regional decomposition¹ presented in equation 2. Farmland values in each state were taken from the *Balance Sheet of the Farm Sector* published by the United States Department of Agriculture, Economic Research Service. The total inequality, average regional inequality, and across-region inequality are also presented graphically in figure 1. These results indicate two periods of significant information in the change in asset values. The first corresponds with the 1973-78 period and the second period is from 1983-89. The second period corresponds with the financial crisis in the agricultural sector during the 1980s. Harl (1990) and others have suggested that the combination of expansionist fiscal policy and tight monetary policy during this time period resulted in significant losses in agricultural equity. Particularly hard hit were the regions that were dependent on export markets for grains such as the Corn Belt, Lake States, and Northern Plains. The first period corresponds to several changes in the agricultural sector. First, the grain markets in 1973 experienced dramatic

¹ The regions used in this study are the 10 farm production regions used by the Economic Research Service. The ten regions include the Northeast (Connecticut, Delaware, Maine, Massachusetts, Maryland, New Hampshire, New Jersey, New York, Pennsylvania, Rhode Island, and Vermont), the Lake States (Michigan, Minnesota, and Wisconsin), the Corn Belt (Illinois, Indiana, Iowa, Missouri, and Ohio), the Northern Plains (Kansas, Nebraska, North Dakota, and South Dakota), Appalachia (Kentucky, North Carolina, Tennessee, Virginia, and West Virginia), the Southeast (Alabama, Florida, Georgia, and South Carolina), the Delta (Arkansas, Louisiana, and Mississippi), the Southern Plains (Oklahoma and Texas), the Mountain States (Arizona, Colorado, Idaho, Montana, Nevada, New Mexico, Utah, and Wyoming), and the Pacific States

increases as a result of increased exports of grain to Russia. Second, the period roughly corresponds with the oil crisis. Finally, the entire period 1973-78 is typically cited as a period of excessive inflationary pressure on farmland prices. Thus, the exact cause of the relative change in information cannot be assigned *a priori*.

The foregoing discussion implicitly attributes all the informational value to changes in farmland prices. However, as implied in equation 3 another possibility is that regional changes in farmland have affected farmland values. Figure 2 presents the information index applied to relative acres. This figure demonstrates that the information in relative changes of acreage is small except in 1975 and 1993. The realignment in 1993 can be traced to an adjustment in accounting for Indian land. One possible explanation of the anomaly in 1975 is the redefinition of the farm. In 1975, a farm was redefined as an entity with agricultural sales of \$1,000 or more. In either case, the results support the conclusion that the information fluctuations (relative persistence in the spatial values of land) presented in figure 1 are largely the result of fluctuations in farmland prices rather than to changing relative acres of farmland in each state.

4. Conclusions and Suggestions for Further Research

This study examines the persistence of asset values using an informational approach similar to that developed by Theil and Leender (1965). However, while Theil and Leender focus on the number of stocks that advance, stay constant, or decline, we examine the change in relative value of each asset. Specifically, we ask whether the share of the value of an asset in one year is a good indicator of the share of that asset in the next year. It is our contention that this formulation is more consistent with the original

(California, Oregon, and Washington). In addition, while Alaska and Hawaii are not typically

question posted by Theil and Leender and Fama (1965), that is whether information exists in the asset market from day to day (or in our application from year to year). While our application focuses on farmland markets in the United States, the procedure can be extended to broader capital markets. Further, the decomposition of the informational approach could be used to examine whether informational content exists for various market segments (i.e., does more information exist in technology stocks than in manufacturing stocks).

For the example developed in this study, information about the preceding year's farmland values largely explains the regional distribution of this year's farmland values. Therefore, the preceding year's farmland values are a good leading indicator for forecasting future farmland prices. The exceptions are the emergence of regional financial stress in the Corn Belt, Lake States, and Northern Plains during the 1980s and the regional changes that emanated from the Russian grain deal of 1973. Another spike in the informational inequality in 1993 can be largely attributed to statistical changes in the definition of farmland.

used because of idiosyncratic factors, we include these states as a separate region.

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Table 1. Information Inequality for Farmland Values over time

| Year | Northeast | Lake States | Cornbelt | Northern Plains | Appalachia | Southeast | Delta States | Southern Plains | Mountain States | Pacific States | Alaska & Hawaii | Average Inequality | Regional Inequality | Total Inequality |
|------|-----------|-------------|----------|-----------------|------------|-----------|--------------|-----------------|-----------------|----------------|-----------------|--------------------|---------------------|------------------|
| 1961 | 602 | 17 | 331 | 127 | 246 | 334 | 42 | 196 | 608 | 2 | 6 | 243 | 237 | 480 |
| 1962 | 32 | 102 | 99 | 216 | 183 | 2 | 191 | 0 | 465 | 16 | 18 | 121 | 102 | 222 |
| 1963 | 57 | 28 | 116 | 835 | 85 | 213 | 245 | 351 | 748 | 170 | 1161 | 269 | 138 | 407 |
| 1964 | 665 | 31 | 435 | 1087 | 532 | 337 | 315 | 75 | 1228 | 237 | 1534 | 478 | 322 | 799 |
| 1965 | 596 | 10 | 338 | 1234 | 547 | 700 | 214 | 1081 | 770 | 909 | 13 | 596 | 206 | 802 |
| 1966 | 940 | 24 | 716 | 395 | 297 | 121 | 540 | 183 | 478 | 2 | 37 | 399 | 190 | 589 |
| 1967 | 749 | 123 | 186 | 300 | 705 | 499 | 1059 | 54 | 595 | 241 | 1 | 374 | 211 | 585 |
| 1968 | 68 | 51 | 166 | 118 | 189 | 634 | 87 | 383 | 260 | 212 | 41 | 198 | 125 | 322 |
| 1969 | 385 | 61 | 314 | 75 | 164 | 6 | 142 | 16 | 581 | 57 | 1 | 189 | 116 | 306 |
| 1970 | 184 | 202 | 156 | 149 | 19 | 182 | 395 | 202 | 350 | 48 | 151 | 174 | 214 | 388 |
| 1971 | 81 | 104 | 584 | 174 | 69 | 355 | 527 | 150 | 292 | 21 | 115 | 270 | 228 | 498 |
| 1972 | 385 | 70 | 365 | 727 | 46 | 7 | 137 | 682 | 1524 | 84 | 0 | 406 | 309 | 715 |
| 1973 | 2258 | 1152 | 716 | 876 | 476 | 373 | 251 | 306 | 378 | 66 | 1 | 661 | 400 | 1062 |
| 1974 | 390 | 317 | 471 | 285 | 1972 | 2027 | 972 | 57 | 2697 | 132 | 887 | 747 | 1508 | 2254 |
| 1975 | 1184 | 2216 | 1147 | 1335 | 373 | 1325 | 1737 | 2 | 2516 | 265 | 2486 | 1178 | 1575 | 2753 |
| 1976 | 978 | 316 | 1298 | 444 | 359 | 877 | 1670 | 89 | 2189 | 441 | 1106 | 868 | 654 | 1522 |
| 1977 | 662 | 238 | 1142 | 3341 | 1475 | 653 | 21 | 2272 | 1952 | 152 | 31 | 1170 | 165 | 1336 |
| 1978 | 1164 | 530 | 1771 | 6865 | 83 | 61 | 510 | 2807 | 1947 | 36 | 1 | 1587 | 170 | 1757 |
| 1979 | 341 | 231 | 1128 | 4157 | 569 | 72 | 413 | 4 | 1682 | 866 | 203 | 1057 | 1062 | 2119 |
| 1980 | 1260 | 455 | 1012 | 2150 | 1446 | 1901 | 399 | 383 | 1575 | 167 | 1585 | 1018 | 856 | 1874 |
| 1981 | 704 | 481 | 2505 | 840 | 1413 | 2621 | 1773 | 1181 | 3154 | 188 | 1520 | 1465 | 607 | 2072 |

^aNumbers represent 100,000 times the actual index measure.

Source: Authors' computations based on the USDA/ERS Farmland Values from the Balance Sheet of the Farm Sector.

Table 1. Information Inequality for Farmland Values over time (continued)

| Year | Northeast | Lake States | Cornbelt | Northern Plains | Appalachia | Southeast | Delta States | Southern Plains | Mountain States | Pacific States | Alaska & Hawaii | Average Inequality | Regional Inequality | Total Inequality |
|------|------------------|-------------|----------|-----------------|------------|-----------|--------------|-----------------|-----------------|----------------|-----------------|--------------------|---------------------|------------------|
| 1982 | 653 ^a | 53 | 546 | 1418 | 441 | 368 | 923 | 1739 | 1414 | 8 | 298 | 684 | 255 | 939 |
| 1983 | 256 | 427 | 329 | 1300 | 419 | 557 | 491 | 1925 | 788 | 265 | 34 | 648 | 376 | 1025 |
| 1984 | 408 | 220 | 1283 | 1330 | 1359 | 511 | 572 | 317 | 821 | 69 | 74 | 733 | 1150 | 1883 |
| 1985 | 342 | 165 | 688 | 353 | 729 | 261 | 225 | 1 | 757 | 204 | 57 | 396 | 664 | 1059 |
| 1986 | 154 | 121 | 370 | 370 | 890 | 161 | 796 | 327 | 311 | 15 | 3 | 307 | 394 | 701 |
| 1987 | 351 | 11 | 144 | 400 | 543 | 358 | 316 | 29 | 1496 | 449 | 181 | 354 | 115 | 469 |
| 1988 | 430 | 1 | 412 | 431 | 437 | 68 | 301 | 127 | 422 | 129 | 62 | 271 | 401 | 672 |
| 1989 | 49 | 339 | 377 | 1928 | 49 | 685 | 828 | 202 | 654 | 112 | 12423 | 567 | 867 | 1434 |
| 1990 | 515 | 156 | 615 | 818 | 409 | 757 | 366 | 43 | 392 | 14 | 13015 | 454 | 514 | 968 |
| 1991 | 380 | 415 | 312 | 84 | 97 | 147 | 769 | 3 | 673 | 2127 | 484 | 505 | 577 | 1083 |
| 1992 | 367 | 799 | 599 | 487 | 215 | 171 | 208 | 0 | 378 | 519 | 51 | 395 | 631 | 1026 |
| 1993 | 251 | 203 | 2336 | 35 | 198 | 91 | 99 | 60 | 161 | 1 | 701 | 502 | 627 | 1129 |
| 1994 | 217 | 220 | 1502 | 243 | 202 | 297 | 888 | 250 | 183 | 145 | 1939 | 445 | 975 | 1420 |
| 1995 | 113 | 103 | 1944 | 516 | 68 | 379 | 712 | 183 | 1897 | 237 | 3211 | 734 | 1766 | 2500 |
| 1996 | 289 | 1611 | 1047 | 384 | 140 | 49 | 87 | 145 | 476 | 1 | 19 | 462 | 368 | 829 |
| 1997 | 438 | 361 | 152 | 983 | 174 | 24 | 338 | 1451 | 1044 | 354 | 46 | 469 | 1058 | 1526 |
| 1998 | 1710 | 165 | 1112 | 806 | 1288 | 509 | 441 | 479 | 508 | 52 | 0 | 675 | 1680 | 2355 |
| 1999 | 247 | 407 | 616 | 2287 | 225 | 75 | 280 | 59 | 258 | 318 | 52 | 542 | 308 | 850 |
| 2000 | 426 | 615 | 743 | 268 | 125 | 228 | 147 | 502 | 396 | 143 | 342 | 376 | 708 | 1084 |
| 2001 | 1138 | 404 | 979 | 192 | 369 | 405 | 320 | 726 | 781 | 37 | 300 | 542 | 430 | 972 |
| 2002 | 1420 | 588 | 247 | 241 | 398 | 1204 | 98 | 4 | 1703 | 48 | 1881 | 569 | 1241 | 1810 |

^aNumbers represent 100,000 times the actual index measure.

Source: Authors' computations based on the USDA/ERS Farmland Values from the Balance Sheet of the Farm Sector.

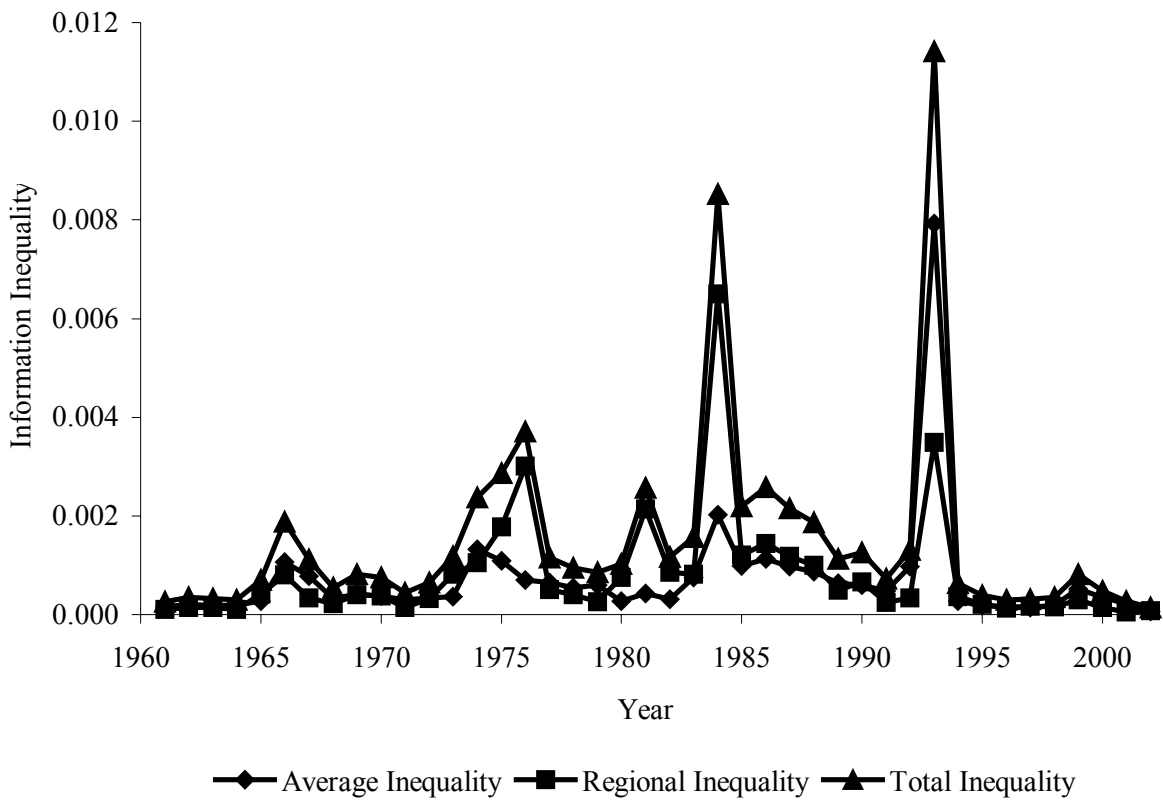


Figure 1. Total, Regional, and Average Information in Land Values over Time

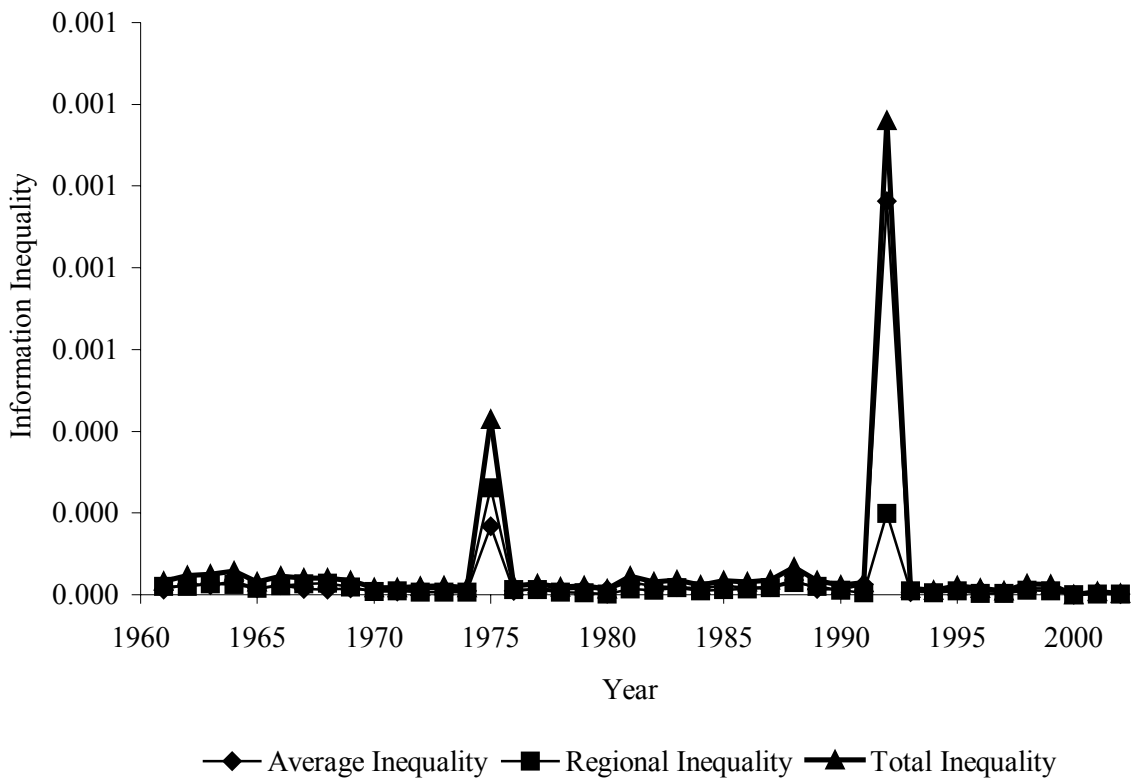


Figure 2. Total, Regional, and Average Information in Land Area over Time