

# Does diversification improve the performance of German banks? Evidence from individual bank loan portfolios

Evelyn Hayden

(Oesterreichische Nationalbank)

**Daniel Porath** 

(University of Applied Sciences Mainz)

Natalja von Westernhagen

(Deutsche Bundesbank)



Discussion Paper Series 2: Banking and Financial Studies No 05/2006

Discussion Papers represent the authors' personal opinions and do not necessarily reflect the views of the Deutsche Bundesbank or its staff.

Editorial Board: Heinz Herrmann

Thilo Liebig

Karl-Heinz Tödter

Deutsche Bundesbank, Wilhelm-Epstein-Strasse 14, 60431 Frankfurt am Main, Postfach 10 06 02, 60006 Frankfurt am Main

Tel +49 69 9566-1

Telex within Germany 41227, telex from abroad 414431, fax +49 69 5601071

Please address all orders in writing to: Deutsche Bundesbank, Press and Public Relations Division, at the above address or via fax +49 69 9566-3077

Reproduction permitted only if source is stated.

ISBN 3-86558-175-7 (Printversion)

ISBN 3-86558-176-5 (Internetversion)

#### **Abstract**

Should banks be diversified or focused? Does diversification indeed lead to enhanced performance and, therefore, greater safety for banks, as traditional portfolio and banking theory would suggest? This paper investigates the link between banks' profitability (ROA) and their portfolio diversification across different industries, broader economic sectors and geographical regions measured by the Herfindahl Index. To explore this issue, we use a unique data set of the individual bank loan portfolios of 983 German banks for the period from 1996 to 2002. The overall evidence we provide shows that there are no large performance benefits associated with diversification since each type of diversification tends to reduce the banks' returns. Moreover, we find that the impact of diversification depends strongly on the risk level. However, it is only for moderate risk levels and in the case of industrial diversification that diversification significantly improves the banks' returns.

Keywords: focus, diversification, monitoring, bank returns, bank risk

JEL Classification: G21, G28, G32

## **Non-Technical Summary**

Should banks be diversified or focused? Does diversification indeed lead to enhanced performance and therefore greater safety for banks as traditional portfolio and banking theory would suggest? In this paper we try to shed some light on these questions by empirically investigating the situation for German banks. By exploiting a unique data set of individual bank loan portfolios for the period from 1996 to 2002, we analyse the link between banks' profitability measured by ROA and their portfolio diversification measured by the Herfindahl Index across different industries, broader economic sectors and geographical regions. To the best of the authors' knowledge, this is the first paper to study the effect of all three types of diversification based jointly on micro-level data on German banks.

The relevant academic literature puts forward two conflicting theories concerning the optimal degree of diversification. While traditional banking and portfolio theory recommends that banks should be as diversified as possible to reduce their risks of suffering a costly bank failure, corporate finance theory suggests that a bank should focus so as to obtain the greatest possible benefit from management's expertise and to reduce agency problems.

Our results clearly support the latter theory, as the evidence we present indicates that each kind of diversification tends to lower German banks' returns, ie focusing generally increases profitability. Furthermore, the impact of any diversification on banks' return changes in line with the risk level. While the effect of sectoral focus on return declines monotonously with increasing risk, there is mixed evidence to suggest either a monotonously decreasing or a U-shaped relationship for regional focus as well as a rather distinct indication of a U-shape with respect to industrial focus. In addition, our data shows that diversification significantly improves banks' profits only in the case of moderate risk levels and industrial diversification. Hence, from a policy point of view, regarding the decision on whether to increase the degree of their industrial, sectoral or geographical diversification banks should evaluate carefully the actual riskiness of their activities.

# Nicht-Technische Zusammenfassung

Sollen Banken ihr Portfolio diversifizieren oder sich besser auf bestimmte Bereiche konzentrieren? Ist es empirisch belegbar, dass ein höheres Maß an Diversifikation die Ertragskraft der Banken stärkt und somit ihre Sicherheit erhöht, wie es von der traditionellen Portfolio- und Bankentheorie postuliert wird? Die vorliegende Arbeit versucht mittels einer empirischen Studie über die Situation bei deutschen Banken näheren Aufschluss über diese Fragen zu geben. Auf der Grundlage der gemeldeten Kreditportfolios einzelner Banken im Zeitraum 1996 - 2002 wird der Zusammenhang zwischen der Rentabilität der Banken und der Diversifikation ihrer Portfolios über verschiedene Branchen, breiter gefasste Wirtschaftssektoren und geographische Regionen hinweg analysiert. Soweit den Autoren bekannt ist, handelt es sich hierbei um die bislang einzige Arbeit, welche all drei Diversifikationsarten gleichzeitig untersucht, und ebenfalls um die erste Studie, die sich auf Mikrodaten deutscher Banken stützt.

In der akademischen Literatur finden sich zwei widersprüchliche Theorien über den optimalen Diversifizierungsgrad von Banken. Während die traditionelle Portfolio- und Bankentheorie empfiehlt, dass Banken ihre Aktivitäten möglichst breit streuen sollen, um das Risiko eines Bankkonkurses zu reduzieren, schlägt die Corporate Finance Literatur vor, Banken sollten sich auf ausgewählte Bereiche spezialisieren, um möglichst stark Management Expertise aufbauen und Anreizprobleme reduzieren zu können.

Die Ergebnisse der vorliegenden empirischen Untersuchungen unterstützen nun eher die zweite Hypothese, da alle drei Diversifikationsarten die Ertragslage deutscher Banken tendenziell negativ beeinflussen, mithin eine höhere Spezialisierung in der Regel mit einer größeren Rentabilität einhergeht. Hinzu kommt, dass sich bei jeder Diversifikationsart der Einfluss auf die Ertragslage der Banken mit dem Risikoprofil verändert. Während bei der sektoralen Konzentration der positive Effekt auf die Ertragslage bei steigendem Risiko monoton abnimmt, gibt es im Hinblick auf die regionale Fokussierung sowohl für einen monoton abnehmenden als auch für einen U-förmig verlaufenden Zusammenhang Hinweise. Die Ergebnisse bei der Branchenfokussierung wiederum sprechen ziemlich deutlich für einen U-förmigen Zusammenhang, so dass sich stärkere Spezialisierungen besonders bei sehr geringem und sehr hohem Risiko ertragssteigernd auswirken. In Summe weist bei den untersuchten deutschen Daten eine größere Diversifizierung nur bei moderaten Risikoprofilen und einer Diversifikation nach Branchen einen signifikanten positiven Effekt auf die Ertragslage der Banken aus, weshalb Banken bei der Wahl ihres optimalen Diversifizierungsgrades unbedingt ihre aktuelle Risikosituation berücksichtigen sollten.

# Content

1 l	Introduction	2
2 1	Data	4
2.1	Data sources	4
2.2	MEASUREMENT OF DIFFERENT TYPES OF DIVERSIFICATION	6
2.3	BALANCE-SHEET VARIABLES	7
2.4	SUMMARY STATISTICS	9
3 1	Empirical framework	10
3.1	AVERAGE IMPACT OF DIVERSIFICATION	10
3.2	CONSISTENCY WITH PORTFOLIO THEORY	12
3.3	DIVERSIFICATION, MONITORING EFFECTIVENESS AND RETURNS	14
4	Conclusions	21
Refer	rences	23

# Does diversification improve the performance of German banks? Evidence from individual bank loan portfolios<sup>1</sup>

Evelyn Hayden<sup>2</sup>
Banking Analysis and Inspections Division
Oesterreichische Nationalbank
Otto-Wagner-Platz 3, POB 61, A-1011 Vienna
Tel: +43 1 40420 3224

E-mail: Evelyn.Hayden@oenb.at

Daniel Porath
Department of Statistics
University of Applied Sciences Mainz
An der Bruchspitze 50
D-55122 Mainz
Tel: +49 6131 628 143

E-mail: **Daniel.Porath@wiwi.fh-mainz.de** 

Natalja v. Westernhagen<sup>2</sup>
Banking and Financial Supervision Department
Deutsche Bundesbank
Postfach 10 06 02, D-60006 Frankfurt am Main
Tel: +49 69 9566 8437

E-mail: Natalja.von-Westernhagen@bundesbank.de

<sup>&</sup>lt;sup>1</sup> The authors would like to thank Thilo Liebig and the Department of Banking and Financial Supervision of Deutsche Bundesbank for research support. In addition, we thank Peter Raupach and Christoph Memmel, the participants of the Austrian Working Group on Banking and Finance 2004, Vienna, the participants of the seminar "Banking Supervision and the Stability of the Banking System" at the Deutsche Bundesbank 2004, Frankfurt am Main, the participants of the 10<sup>th</sup> Symposium on Finance, Banking and Insurance 2005, Karlsruhe, the participants of the conference "Managing Risk in Financial Institutions: From Theory to Practice" 2006, New York, as well as the participants of the European Financial Management Association 2006, Madrid for helpful comments. All errors and inaccuracies are solely in the authors' responsibility.

<sup>&</sup>lt;sup>2</sup> The opinions expressed in this paper are those of the authors and do not necessarily reflect the views of Oesterreichische Nationalbank or Deutsche Bundesbank or of their staff.

#### 1 Introduction

Should banks diversify their portfolios across different industries or even broader economic sectors and geographical regions, or should they focus on a few related fields? Does diversification indeed lead to enhanced performance and, therefore, greater safety for banks, as traditional portfolio and banking theory would suggest? In this paper, we try to shed some light on these questions by investigating empirically the situation of German banks.

The focus versus diversification issue is important in the context of banks as they are affected by several regulations that create incentives either to diversify or to focus their portfolios, ie the imposition of capital requirements tied to the risk of the banks' assets or asset investment restrictions. Hence, policymakers should be especially interested to see whether or not banks benefit from diversification.

Experts on financial institutions generally argue that banks – which are typically highly leveraged – should diversify to reduce their risks of suffering a costly bank failure. In addition, several models of intermediation theory suggest that diversification makes it cheaper for institutions to achieve credibility in their role as screeners or monitors of borrowers (see, for example, Diamond (1984) and Boyd and Prescott (1986)). However, corporate finance theory suggests that firms should, if anything, focus in order to avoid value-reducing diversification effects due to agency problems as pointed out by Denis et al (1997) or in order to avoid inefficient allocation of resources due to internal power struggles between divisions as discussed by Rajan et al (2000) or in order to protect themselves from significant losses in the firm value associated with corporate diversification strategies as documented by Berger and Ofek (1995). Since real-world cases can be found to support either view, the question arises as to the circumstances which call for one strategy or the other to be applied.

Winton (1999) presents a theoretical framework to investigate the above issue. In his model, the incentives of a bank to monitor loans and a bank's risk of failure strongly affect a bank's decision on whether to diversify or not. Diversification benefits appear to be greatest for banks with moderate levels of downside risk and when the banks' monitoring incentives need strengthening. Specialised banks with low exposure to downside risk have only minor incentives to diversify since diversification might increase their probability of failure, which is considered to be low at present. For diversified banks with sufficiently high downside risk, diversification can actually increase the probability of default since a diversified bank is exposed to more sectors than a specialised one and a downturn in one sector is enough to make a bank fail. Furthermore, banks which expand into new economic sectors or geographical regions have less expertise and, therefore, a lower monitoring effectiveness in these areas, at least initially. Some papers also suggest that a bank entering a sector with several established banks faces increased adverse selection in its pool of borrowers (see, for example, Gehrig (1998) and

Shaffer (1998)). Seen in this light, diversification may prove to be unattractive, particularly when the bank's loans have either a low or high downside risk.

Although the issue of focus versus diversification has a long history in corporate finance literature, it has not been addressed thoroughly in an empirical context for financial institutions and banks. The existing literature focuses mainly on geographical diversification and US data, and also provides mixed results. Hughes et al (1996) and Berger and DeYoung (2001), for example, use more aggregated measures of bank diversification to examine geographical diversification for US banks, while Caprio and Wilson (1997) consider cross-country evidence of a relationship between concentration and bank insolvency. In addition, Dahl and Logan (2003) and Buch at al (2004) suggest that international diversification offers benefits while, according to Klein and Saidenberg (1998) and Morgan and Samolyk (2003), the geographical diversification of US banks is not necessarily associated with an increase in profitability. DeLong (2001) finds that geographically-focused bank mergers in the US result in superior performance, while Stiroh and Rumble (2003) and Stiroh (2004) show that a shift towards non-interest income does not offer large diversification benefits.

Therefore, there is clearly a need for more empirical evidence on the effects of diversification on banks' performance based on individual bank-level data from European countries. The leading study in this respect is probably the one by Acharya et al (2006), which examines the impact of sectoral and industrial loan diversification on the performance of Italian banks. The results of this study are consistent with Winton's theory of a deterioration in the effectiveness of banks' monitoring activities at high levels of risk. In addition, Acharya et al find that both industrial and sectoral loan diversification reduces banks' returns while endogenously producing riskier loans for high-risk banks in their sample. This implies that a diversification of banks' assets is not guaranteed to result in a superior return performance and/or greater safety for Italian banks.

The question now arises as to whether the Italian results are valid for other European countries, too. Our study attempts to fill this gap by studying the situation of the German banking industry. Based on a unique data set of Deutsche Bundesbank involving data on individual bank loan portfolios disaggregated at micro level for the period from 1996 to 2002, we assess the impact of sectoral, industrial and geographical diversification on banks' profitability by looking at three major aspects. Firstly, we are interested in the average effect on banks' returns of banks' portfolio diversification across industries, sectors and regions. Secondly, we try to gain an insight into whether diversification is used as an instrument to induce shifts in banks' risk-return efficiency. Thirdly, we test how monitoring effectiveness on the part of low, medium and high-risk banks impacts on the relationship between banks' portfolio diversification and banks' returns. Here – in contrast to previous studies – we apply a Value at Risk approach to measure banks' risk and derive unexpected losses for each individual bank as, in our opinion, unexpected losses are better suited to capturing banks' riskiness

than is the more common proxy of expected losses.

Our main findings are as follows. Firstly, we find that portfolio diversification across different sectors, industries and regions tends to have a negative effect on banks' profitability rather than to lead to improved returns. Larger benefits are associated with geographical focus, whereas benefits from industrial focus appear to be only moderate. Secondly, there is evidence that, instead of operating at a constant risk-return efficiency level, banks use diversification in order to change their risk-return profile. As banks with highly risky credit portfolios are not systematically more profitable than banks with low risk portfolios, it seems that, overall, banks are not risk-return efficient. Thirdly, the profitability benefits associated with diversification are strongly dependent on the banks' risk level. In addition, the type of focus plays a crucial role. While the effect of sectoral focus on return declines monotonously with increasing risk, there is mixed evidence to suggest either a monotonously decreasing or a U-shaped relationship for regional focus as well as a rather distinct indication of a U-shape with respect to industrial focus. Therefore, our results confirm, at least partly, Winton's theory that diversification benefits are greatest for moderate risk levels. Finally, our data shows that diversification significantly improves banks' profitability only in the case of moderate risk levels and industrial diversification.

The remainder of this paper is structured as follows: in Section 2 we describe our data and in Section 3 we present the empirical results before reaching a conclusion in Section 4.

#### 2 Data

#### 2.1 Data sources

The main data source for our analysis originates in the database of the credit register for loans of 1.5 million euro (formerly 3 million Deutsche Mark) or more at the Deutsche Bundesbank<sup>3</sup>. German banks have to provide quarterly reports on all claims exceeding the threshold of 1.5 million euro. Bank

-

<sup>&</sup>lt;sup>3</sup> Since the data in the credit register have been primarily gathered for regulatory rather than research purposes, the credit register may overstate and even double-count exposures (Deutsche Bundesbank 1998a). Actual exposures of firms to banks are overstated since they include both on-balance sheet and off-balance sheet positions. Some off-balance sheet exposures do not appear to be direct exposures to a firm but e.g. guarantees for this firm's loans from another bank. Therefore, the inclusion of off-balance sheet exposures, which in reality represent guarantees, leads to actual exposure being overstated. Since in this study we have used only on-balance sheet positions this problem does not imply to our case. Actual exposures are double-counted in some rare cases, though, when partners of companies under civil law (*Gesellschaft bürgerlichen Rechts*, GBR) are jointly accountable for losses. In this case, the exposure of the GBR is reflected in the position of each partner with the same amount. Unfortunately, we cannot adjust the data for this double-counting, however, the error should be neglectable in size, as companies under civil law are typically small firms whose debt does not exceed the threshold of 1.5 million euro and hence is not included in the credit register.

claims are defined fairly broadly and cover details of types of claims<sup>4</sup>, types of borrowers by industries and sectors, international claims by individual foreign countries and regions<sup>5</sup>. In addition to balance sheet bank activities, claims also incorporate information on off-balance-sheet activities<sup>6</sup>. This credit register data set on the exposures of individual banks is combined with financial data from the second Bundesbank data source, namely BAKIS (BAKred<sup>7</sup> Information System). BAKIS incorporates information derived from the bank balance sheets and supervisory reports of all German banks. Our dependent variable and control variables stem from BAKIS. Since the data on bank balance sheets is mostly of annual frequency, we used annual data for the period from 1996 to 2002. Both the credit register and BAKIS represent unique data sources never before exploited in order to investigate the relationship between the diversification and performance of German banks.

Our data sample covers not only banks but also their subsidiaries and includes 3,760 individual entities. However, as small banks usually grant only very few large loans, the loans reported to the credit register sometimes cover only a rather small fraction of the total credit volume outstanding according to the banks' balance sheets. This implies that it might be misleading to analyse the diversification structure of these small banks on the basis of the information from the credit register, as the breakdown of the total portfolio could differ significantly from that of one of the large loans. Therefore, our study focuses only on those banks where the ratio of the reported balance-sheet loans from the credit register to the total amount of assets according to the balance sheet exceeds 50%. We also exclude affiliates of German banks abroad, mortgage banks and special-purpose banks from our analysis. This reduces the number of eligible banks to 983.

Over the period 1996 to 2002 many bank mergers took place. To handle mergers, we separate the two pre-merger banks from the merged bank. At the end, we thus have three banks, which are treated independently. We repeat this procedure as often as a merger took place. Additionally, in order to avoid the double counting of banks in the year of the merger, we dropped target banks in the year of the acquisition.

<sup>&</sup>lt;sup>4</sup> For example, lease receivables, mortgage loans, publicly guaranteed loans, inter-bank loans (with a residual maturity of up to one year) are listed separately under on balance sheet activities.

<sup>&</sup>lt;sup>5</sup> The following items are deemed not to be credit exposures: shares in other enterprises irrespective of how they are shown in the balance sheet and securities in the trading portfolio (Deutsche Bundesbank, 1998b).

<sup>&</sup>lt;sup>6</sup> Off balance sheet items include derivatives (other than written option positions), guarantees assumed in respect of these and other off balance sheet transactions (Deutsche Bundesbank, 1998b).

<sup>&</sup>lt;sup>7</sup> The former Federal Banking Supervisory Authority, now BaFin (*Bundesanstalt fürFinanzdienstleistungsaufsicht*), ie the Federal Financial Supervisory Authority.

<sup>&</sup>lt;sup>8</sup> For these banks, the average coverage rate is about 70%.

### 2.2 Measurement of different types of diversification

The data from the credit register provide considerable details about the industrial, broader sectoral and geographical breakdown of German bank claims. On an individual bank basis, the following information on the portfolio breakdown is available.

- 1. The disaggregated industrial sector breakdown includes (1) agricultural, forestry and fishing products, (2) energy products, (3) iron and non-iron material and ore, (4) ores and products based on non-metallic minerals, (5) chemicals, (6) metal products, apart from machinery and means of conveyance, (7) agricultural and industrial machinery, (8) office, EDP machinery and others, (9) electrical material, (10) transport, (11) food products, beverages and tobacco-based products, (12) textiles, leather, shoes and clothing products, (13) paper, publishing and printing products, (14) rubber and plastic products, (15) other industrial products, (16) construction, (17) services trade and similar, (18) hotel and public firms' products, (19) internal transport services, (20) sea and air transport, (21) transport-related services, (22) communication services and (23) other sales-related services. It should be noted that, in aggregate, these exposures (collectively defined in the data as non-financial and household exposures) constitute the dominant part of most banks' portfolios.
- 2. The broader sectoral breakdown includes (1) financial institutions and banks, (2) non-financial corporations, (3) households, (4) the public sector and (5) other counter-parties.
- 3. The geographical breakdown includes (1) Germany, six regions according to the IMF classification: (2) industrial countries, (3) Asia, (4) Africa, (5) the Middle East, (6) the Western hemisphere, (7) Emerging Europe, and (8) others.<sup>9</sup>

To measure diversification (or, alternatively, focus), we use the Hirschmann-Herfindahl Index. It is calculated as the sum of the squares of exposures as a fraction of total exposure under a given classification and is represented by the following formula

$$H = \sum_{i=1}^{n} \left(\frac{X_i}{X}\right)^2,$$

where n is the number of groups and  $X_i$  measures exposure to industry, sector or region i. The smallest and largest possible values for the Herfindahl Index are given by  $1/n \le H \le 1$ . Hence, lending is more concentrated the closer the Herfindahl Index is to one and is perfectly diversified if H equals 1/n.

In our case, we constructed three different kinds of Herfindahl Indices: an industrial (and household) sector Herfindahl Index (HI), a broad asset type (or sectoral) Herfindahl Index (HT) and a regional (or geographical) Herfindahl Index (HR).

•

<sup>&</sup>lt;sup>9</sup> For further details, see Nestmann et al. (2003).

#### 2.3 Balance-sheet variables

We employed the following (annual) variables obtained from the balance sheet data for the banks in our sample in the period from 1996 to 2002.

#### Return measures

"Operating Profit/Assets" serves as the principal measure of return. All of the results displayed are based on this measure. However, we also performed robustness checks using other measures, such as "Operating Profit/Equity". We found that, overall, the results are robust with respect to the return measures employed.

#### Risk measures

The simplest method of measuring risk would be to look at a balance sheet ratio such as "Doubtful and Non-Performing Loans/Total Loans", which could be interpreted as capturing the level of expected losses. However, we consider that risk is more accurately represented by unexpected losses, which is the reason why we focused on a Value at Risk (VaR) measure.

Value at Risk is the most widespread method of determining a bank's loan portfolio risk. The Value at Risk of bank i in period t,  $VaR_{it}$ , is the maximum loss over a target horizon such that, with a prespecified probability p, the realised loss will be smaller. The unexpected loss can be determined from the distribution of the portfolio losses at the target horizon as the difference between the mean of the portfolio value and the value at the p-percentile. In our calculations, p is 99.9%. This is based on the observation that banks typically work with percentiles higher than 99.5%. Since the following estimations are fixed-effects panel models where the levels of the variables are differenced out, the exact level of p will not affect our results. The values for the VaR have to be taken from the distribution of the portfolio value. We estimated the portfolio's value distribution using a simplified version of CreditMetrics. The basic assumptions of CreditMetrics are that the returns of a debtor's assets are normally distributed and that a default occurs when the returns of a debtor fall below a certain threshold. The default threshold is determined from the probability of default (PD).

As our data set does not comprise rating information for individual loans, we used the average insolvency rate of the industry associated with the loan in order to proxy the default probability for a loan and to calculate its return threshold. We further assumed that the correlation between the returns

\_

<sup>&</sup>lt;sup>10</sup> J.P. Morgan (1997)

of debtors can be approximated by the correlation between the industries' insolvency rates. <sup>11</sup> Using equal probabilities of default for each bank, however, may bias the results since, for example, focused banks may have more effective monitoring systems and, therefore, grant loans with lower PDs than diversified banks. Therefore, as no information on the risk of loans at an industry and individual bank level is available for German banks, we had to adjust the (observed) industry insolvency ratios by bank-specific factors. To do so, we defined the industry insolvency ratio multiplied by a scale parameter which is related to a bank's loan loss provisions as a bank-specific PD. <sup>12</sup> As a result, banks with large loan loss provisions (divided by the amount of total loans) are assigned higher PDs for loans to a specific industry than are banks with lower provisions. It should be noted that bank loans to all industries are adjusted using the same scale factor because, unfortunately, the data does not allow for a more precise adjustment.

The current value of a bank's overall portfolio at the beginning of a period is given by the sum of the bank's individual exposures to each industry, which we took from the credit register as described above. We then simulated returns using a multivariate normal distribution with mean zero and the correlation matrix from the insolvency data. Defaults occur when the simulated returns fall below the threshold given by the critical values derived from the industries' annual insolvency rates. The simulated value of the portfolio at the end of the period is equal to the value at the beginning of the period less 45% of the loans defaulting in the simulations, which means that we assume a loss given default (LGD) of 45% in line with the Basel II proposal (see Basel Committee on Banking Supervision (2004)). We then repeated this exercise 50,000 times in order to obtain the simulated loss distribution of a single bank in a specific period. Using the loss distribution, we calculated the

\_

$$Scale_{it} = \frac{Loan \ Loss \ Provisions_{it}}{Total \ Loans_{it}} / \frac{\sum_{j} Insolvency \ Rate_{jt} Exposure_{jit}}{\sum_{j} Exposure_{jit}}$$

where *i*, *t* and *j* index the bank, the period and the industry. *Loan Loss Provisions* and *Total Loans* are taken from the balance sheet data and *Exposure* is derived from the credit register. In line with Moody's KMV Credit Monitor, we introduce a cap of 20% for the resulting PD; see Bohn et al (2005).

<sup>&</sup>lt;sup>11</sup> The insolvency data used were those of the Federal Statistical Office (*Statistisches Bundesamt*). The industry codes of the insolvency data correspond to the industry codes of the credit register. The insolvency rate of a specific industry is calculated as the number of insolvencies divided by the total number of companies in the industry. The probability of default of a specific industry is then calculated as the average of the annual data from 1994 to 2002. The correlation between insolvencies of the industries is calculated using monthly data for the same period.

 $<sup>^{12}</sup>$  More precisely, the scale parameter for bank i is defined as

<sup>&</sup>lt;sup>13</sup> Insolvency data are used as a proxy for asset correlations since the latter are not observable. As a result, VaR<sub>it</sub> might be negatively biased as asset correlations are usually higher than insolvency correlations. However, with the assumption that the differences between the correlations of assets and insolvency are constant over time, the bias will difference out in the fixed-effects estimation.

<sup>&</sup>lt;sup>14</sup> As mentioned above, the VaR level differences out in the fixed effects estimations, which means that the value of the LGD will not affect our results.

unexpected loss as the difference between the 99.9% quantile and the mean. Finally, the variable  $Risk_{it}$  was calculated as

$$Risk_{it} = Unexpected\ Loss_{it}\ / Total\ Exposure_{it}$$

In order to obtain a panel of observations for  $Risk_{it}$ , we repeated the simulations for each bank and each period of our sample.

#### Control variables

Banks' returns might not only be dependent on the respective banks' diversification and risk; they are also likely to differ as a result of other criteria. In the estimation, we controlled for unobservable individual and time effects by using dummy variables. The bank-specific dummies check for all effects which do not change for individual banks over time. These effects include, for example, characteristics which differ between banking groups, such as regional constraints on the part of German savings or cooperative banks or different ownership structures. In addition to these fixed effects, we also monitored characteristics which may change over time.

```
Personnel_{it} = Personnel\ Costs_{it} / Assets_{it}
Size_{it} = Ln(Assets_{it}).
```

In line with Acharya et al (2006) we used the variable  $Personnel_{it}$  to proxy cost efficiency. The rationale is that banks with different cost efficiency levels may transform the benefits from diversification in a different way. The variable  $Size_{it}$  captures the possible effects of scale on return.

The banks' equity ratio is a common control variable in many empirical studies (Acharya et al (2006), Demirgüc-Kunt/Huizinger (1998), Quagliariello (2004), Barth et al (2002)).

$$Equity_{it} = Equity Capital_{it} / Assets_{it}$$

The amount of capital banks hold depends on the banks' risks and risk preferences. Accordingly,  $Equity_{it}$  depends on  $Risk_{it}$  and the Herfindahl Indices. We tried to avoid the emergence of bias from this dependency and thus estimated the influence of diversification on returns without controlling for equity. However, in order to compare our results with those from other studies (for example, Acharya et al (2006)), we report results which include  $Equity_{it}$ , too.

#### 2.4 Summary statistics

Table 1 presents univariate descriptive statistics for the variables used in the following estimations. Note that the mean (median) bank's size in the sample is about 4.2 billion euro (0.9 billion euro). The average industrial and sectoral focus measures (HI and HT) are quite low, which suggests a significant degree of diversification in these areas. However, the average geographical focus HR is rather high, which captures the fact that most German banks primarily do business with domestic counter-parties.

Table 1 Summary Statistics (3,529 observations)

	$HT_{it}$	$HR_{it}$	$HI_{it}$	Return <sub>it</sub>	$Risk_{it}$	$Personnel_{it}$	$Size_{it}$	Equity <sub>it</sub>
Mean	0.569	0.929	0.291	0.004	0.034	0.013	20.623	0.052
Median	0.514	0.972	0.204	0.004	0.026	0.013	20.626	0.044
Standard Deviation	0.156	0.076	0.241	0.028	0.041	0.015	1.366	0.054
Minimum	0.299	0.284	0.066	-0.018	0.000	0.001	15.626	0.001
Maximum	1.000	1.000	1.000	0.054	0.493	0.073	26.472	0.305
			(	Correlation				
$HT_{it}$	1.000							
$HR_{it}$	0.034	1.000						
$HI_{it}$	-0.250	-0.243	1.000					
Return <sub>it</sub>	0.030	0.006	0.026	1.000				
$Risk_{it}$	-0.227	-0.085	-0.160	-0.025	1.000			
Personnel <sub>it</sub>	0.085	0.048	0.044	-0.164	-0.082	1.000		
$Size_{it}$	-0.438	0.013	0.152	0.016	0.151	-0.283	1.000	
Equity <sub>it</sub>	0.130	-0.200	0.200	0.085	-0.104	0.484	-0.302	1.000

In addition, Table 1 presents the correlation matrix among the explanatory variables of the following estimations. As it illustrates, the three Herfindahl Indices are not highly correlated. This suggests that the effects of industrial, sectoral and regional diversification on the banks' return might be different.

## 3 Empirical framework

Our aim is to assess the impact of diversification on banks' profitability in the case of German banks. We address this question by looking at three aspects. First of all, we are interested in the average effect of diversification on return. Subsequently, we try to answer the question of whether the link between return and diversification is consistent with portfolio theory. Finally, we test how monitoring effectiveness affects the relationship between diversification and return.

#### 3.1 Average impact of diversification

We investigated the average impact of diversification on banks' performance in a panel regression where we regressed return on the Hirschman-Herfindahl Indices. More precisely, we estimated the following equation.

$$Re turn_{it} = \beta_0 + \beta_1 H T_{it} + \beta_2 H R_{it} + \beta_3 H I_{it} + \sum_{n=4}^{6} \beta_n X_{nit} + \delta z_t + \varepsilon_{it.}$$
 (1)

where  $Return_{it}$ ,  $HT_{it}$  and  $HI_{it}$  are measured as described in the previous section.  $X_{nit}$  is the set of control variables such as  $Personnel_{it}$ ,  $Size_{it}$  and - for reasons of comparability with other studies -  $Equity_{it}$ . The error term is given by  $\varepsilon_{it} = \eta_i + \omega_{it}$ , where  $\eta_t$  represents bank-specific fixed effects and  $\omega_{it}$  is a disturbance term.  $\omega_{it}$  is iid with mean zero and a constant variance. Moreover, we add a set of year dummy variables,  $z_t$ , in order to capture omitted macroeconomic developments. The empirical

model given in equation (1) is estimated using the fixed effects estimation techniques. The coefficients  $\beta_1$ ,  $\beta_2$  and  $\beta_3$  capture the average impact of focus on bank performance, which means that they are not conditioned by the banks' risk levels.

We estimate (1) with several restrictions. The results are reported in Table 2.<sup>15</sup> In all specifications, the coefficients for the Herfindahl Indices are positive; in most cases they are also highly significant. The results are remarkably stable for the estimations (1a) – (1d); however, they change when  $Equity_{it}$  is added to the equation; see specification (1e). In specification (1e), the coefficients for the Herfindahl Indices are considerably lower in terms of both absolute magnitude and significance level; at the same time,  $Equity_{it}$  is highly significant. Hence, the inclusion of  $Equity_{it}$  reduces the impact of the Herfindahl Index on  $Return_{it}$ . This is consistent with our assumption that  $Equity_{it}$  is determined by banks' risk preferences and that the coefficients  $\beta_1$ ,  $\beta_2$  and  $\beta_3$  in (1e) thus no longer reflect the average impact of focus on  $Return_{it}$ . Aside from this assumption, the results of all specifications confirm a positive impact of geographical focus at a 1% confidence level and a positive impact of sectoral focus at a level of at least 10%. Furthermore, with regard to the magnitude of the coefficients, all estimations reveal the same order with  $HR_{it}$  having the highest and  $HI_{it}$  having the lowest coefficient.

Table 2 Two-way fixed effects estimation of Equation (1) with alternative restrictions Dependent variable: *Return*<sub>ii</sub>,

الم	penaent variable	. Itelli iili,			
	(1a)	(1b)	(1c)	(1d)	(1e)
$\overline{HT_{it}}$	0.018***	0.020***			0.005*
	(4.87)	(5.74)			(1.65)
$HR_{it}$	0.030***	•	0.039***		0.017***
	(4.38)		(5.83)		(2.81)
$HI_{it}$	0.006**			0.005**	0.001
	(2.50)			(2.00)	(0.47)
$Equity_{it}$					0.185***
•					(24.66)
$Size_{it}$	-0.022***	-0.023***	-0.023***	-0.022***	-0.006***
	(14.75)	(-15.69)	(-15.13)	(-14.92)	(-4.27)
$Personnel_{it}$	-1.439***	-1.433***	-1.434***	-1.408***	-1.780***
	(-15.02)	(-14.89)	(-14.91)	(-14.54)	(-20.55)
Constant	0.429***	0.480***	0.438***	0.476***	0.124***
	(13.39)	(15.66)	(13.77)	(15.22)	(3.99)
No. of obs.	3,529	3,529	3,529	3,529	3,529

T-values in brackets, \*, \*\*, \*\*\* denotes significance at 10%, 5% and 1%. Blanks indicate that the coefficient of the variable is restricted to zero.

The positive coefficients of the Herfindahl Indices can be interpreted as confirmation that (at least on average) the mean profits arising from focusing loan portfolios exceed the mean profits achievable through diversification. The greatest benefits seem to be attainable through geographical focus,

<sup>&</sup>lt;sup>15</sup> Due to lack of space the coefficients of the year dummies are not reported with the other estimation results.

whereas the benefits from industrial focus appear to be only moderate.

In the following subsections, we will analyse whether the results are consistent with portfolio theory and/or how the quality of monitoring influences the link between diversification and returns.

#### 3.2 Consistency with portfolio theory

Portfolio theory describes the relationship between diversification, expected returns and risk in a liquid portfolio. In general, different portfolios display different levels of risk and expected return. Portfolio theory assumes that among all possible portfolios at a certain risk level investors will always select the one with the highest expected return, which is called risk-return efficient. The efficient portfolios for different risk levels constitute an efficient frontier, which displays a positive slope in a risk-return diagram. The positive slope implies that when moving on the efficient frontier there is a trade-off between risk and expected returns.

In our context portfolio theory implies that a bank can change the diversification of its credit portfolio either to reach the efficient frontier (and therefore increase risk-return efficiency) or to move along the efficient frontier. In order to test which policy is prevalent, we added the variable  $Risk_{it}$  (measured as described above) to Equation (1).

Re 
$$turn_{it} = \beta_0 + \beta_1 H T_{it} + \beta_2 H R_{it} + \beta_3 H I_{it} + \beta_4 R i s k_{it} + \sum_{n=5}^{7} \beta_n X_{nit} + \delta z_t + \varepsilon_{it}.$$
 (2)

Here, the coefficients  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$  capture the impact of a variation in focus on return conditioned by the banks' risk level. If all banks operated on the efficient frontier the conditional coefficients  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$  should be zero because banks with the same risk level should realize the same return on average. In other words, the effects of diversification would be entirely captured by risk. Besides,  $\beta_4$  should be positive due to the trade-off between risk and return. If, however, banks did not operate on the efficient frontier,  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$  would be different from zero and/or  $\beta_4$  would be zero or negative. This would imply that banks could use a change in their degree of diversification in order to change their risk-return profile, i.e. reach the efficient frontier.

It should be noted that risk-return efficiency is linked to expected returns. Since expected returns are not observable we have to replace them by observed returns in (2). Deviations of expected returns from observed returns are captured by  $\varepsilon_{it}$ , which is assumed to be Gaussian. Furthermore, following portfolio theory,  $Risk_{it}$  is endogenous in  $HT_{it}$ ,  $HR_{it}$  and  $HI_{it}$  and exogenous to  $Return_{it}$ . Therefore, (1) can be interpreted as the reduced form of (2).

Table 3 shows the estimated coefficients. Interestingly, conditioning by risk does not change the results from (1) since the conditional coefficients  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$  are almost equal to the average coefficients  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$  in Table 2. At the same time,  $Risk_{it}$  is significantly negative in specifications

(2a), (2c) and (2d). In the other specifications, the coefficient for  $Risk_{it}$  is insignificant. It should be noted that the outcome does not seem to result from a potential multicollinearity between  $Risk_{it}$  and the Herfindahl Indices as  $\beta_4$  remains stable when the Herfindahl Indices are excluded from Equation (2); see specification (2e). When  $Equity_{it}$  is added to the equation (see specification (2f)), the coefficients of both  $Risk_{it}$  and the Herfindahl Indices become insignificant. Again, we believe that this finding is induced by the fact that  $Equity_{it}$  depends on the banks' risk preferences.

Since there is no evidence of a positive relationship between risk and return, it appears that banks have not chosen their degree of diversification to operate at a constant risk-return efficiency level. This is confirmed by the non-zero coefficients of the Herfindahl Indices. Banks with highly risky credit portfolios were not systematically more profitable than banks with low risk portfolios.

Table 3 Two-way fixed effects estimation of Equation (2) with alternative restrictions Dependent variable:  $Return_{it}$ ,

	(2a)	(2b)	(2c)	(2d)	(2e)	(2f)
$\overline{HT_{it}}$	0.017***	0.020***				0.005
	(4.77)	(5.63)				(1.621)
$HR_{it}$	0.030***		0.039***			0.017***
	(4.38)		(5.82)			(2.81)
$HI_{it}$	0.007***			0.006**		0.001
	(2.71)			(2.28)		(0.55)
$Risk_{it}$	-0.018*	-0.014	-0.018*	-0.022**	-0.019*	-0.006
	(-1.67)	(-1.27)	(-1.67)	(-2.02)	(-1.71)	(-0.56)
$Equity_{it}$						0.109***
						(24.60)
$Size_{it}$	-0.021***	-0.022***	-0.022***	-0.022***	-0.022***	-0.006***
	(-14.29)	(-15.35)	(-14.76)	(-14.40)	(-15.09)	(-4.17)
$Personnel_{it}$	-1.431***	-1.429***	-1.427***	-1.398***	-1.406***	-1.777***
	(-14.92)	(-14.81)	(-14.83)	(-14.43)	(-14.52)	(-20.48)
Constant	0.421**	0.475***	0.431***	0.466***	0.480***	0.122***
	(13.00)	(15.36)	(13.45)	(14.72)	(15.45)	(3.89)
No. of obs.	3,529	3,529	3,529	3,529	3,529	3,529

T-values in brackets, \*, \*\*, \*\*\* denotes significance at 10%, 5% and 1%. Blanks indicate that the coefficient of the variable is restricted to zero.

To sum up, the positive Herfindahl Indices in Table 3 indicate that banks with a higher level of focus tend to be more profitable than diversified banks and, at the same time, banks with a higher risk level seem to be less profitable. Accordingly, instead of operating at a constant risk-return efficiency level, banks appear to have used diversification as an instrument to change their risk-return profiles. However, note that these interpretations of (2) in the sense of risk-return efficiency only hold under the assumption that the deviations of observed returns from expected returns average out in the time-window that we consider.

#### 3.3 Diversification, monitoring effectiveness and returns

Finally, we analysed how monitoring effectiveness affects the link between diversification and banks' returns. In Winton's (1999) model, effective loan monitoring is the force that prevents banks from failure by catching problem loans before the situation deteriorates too far. Therefore, the monitoring of loans allows banks to improve their loan returns and reduce their default probability. When deciding whether to diversify or not, banks take into account the impact of diversification on their incentives to monitor their loans and their probability of failure. Specialised banks which are exposed to sectors with low downside risk will derive only moderate benefits from diversification as they have a low default probability anyway. Alternatively, in the case of diversified banks with loans of sufficiently high downside risk, bank owners (equity holders or managers) have only few incentives to monitor as, on an expected basis, most of the benefits from monitoring will accrue only to the bank's creditors (uninsured depositors and providers of borrowed funds) and diversification could actually increase the banks' default probability. Accordingly, the benefits from diversification are greatest if banks' loans have moderate levels of downside risk and if banks' monitoring incentives need to be strengthened.

In terms of empirically testable hypotheses, Winton's theory implies that the relationship between return and focus (or, alternatively, diversification) should be expected to be non-linear and U-shaped in risk. To try to capture this, we first of all reproduced the tests proposed by Acharya et al. (2006). They expanded Equation (2) by non-linear terms.

$$Re turn_{it} = \beta_{0} + \beta_{1}HT_{it} + \beta_{2}HR_{it} + \beta_{3}HI_{it} + \beta_{4}Risk_{it} + \sum_{n=5}^{7} \beta_{n}X_{nit}$$

$$+\alpha_{11}HT_{it} * RISK_{it} + \alpha_{12}HT_{it} * RISK^{2}_{it}$$

$$+\alpha_{21}HR_{it} * RISK_{it} + \alpha_{22}HR_{it} * RISK^{2}_{it}$$

$$+\alpha_{31}HI_{it} * RISK_{it} + \alpha_{32}HI_{it} * RISK^{2}_{it} + \delta z_{t} + \varepsilon_{it}.$$
(3)

By calculating the first derivative of return on focus, it is easy to see that a U-shape in risk is given if

$$\alpha_{11} < 0$$
,  $\alpha_{12} > 0$ ,  $\alpha_{21} < 0$ ,  $\alpha_{22} > 0$ ,  $\alpha_{31} < 0$  and  $\alpha_{32} > 0$ ,

see Acharya et al. (2006).

The estimated coefficients (see Table 4) are consistent with the patterns associated with a U-shaped form, the only exception being the specification which contains  $Equity_{it}$ . In all other equations, the coefficients of the Herfindahl Indices interactant with  $RISK_{it}$  are negative, whereas they are positive when interactant with  $RISK_{it}^2$ . Most of the coefficients are significant at a 5% or even at a 1% confidence level. Thus, the results could be interpreted as strong evidence of a U-shaped relationship between focus and return depending on the level of risk.

Table 4: Two-way fixed effects estimation of Equation (3) with alternative restrictions Dependent variable: *Return*<sub>it</sub>

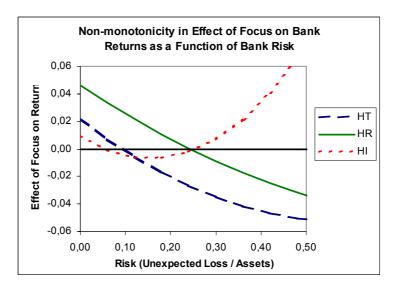
- ·r ·	(3a)	(3b)	(3c)	(3d)	(3e)
$HT_{it}$	0.023***	0.022***			0.008**
	(5.64)	(6.02)			(2.12)
$HT_{it}*RISK_{it}$	-0.324***	-0.254***			-0.255
	(-3.25)	(-3.22)			(0.25)
$HT_{it}*RISK^{2}_{it}$	0.293**	0.215**			0.430
	(1.96)	(2.15)			(0.74)
$HR_{it}$	0.029***	,	0.046***		0.022***
	(3.89)		(6.32)		(3.19)
$HR_{it}*RISK_{it}$	-0.123		-0.217***		-0.076
	(-1.16)		(-2.77)		(-0.79)
$HR_{it}*RISK^{2}_{it}$	0.140**		0.113**		-0.197
	(2.11)		(2.02)		(-0.63)
$HI_{it}$	0.011***		, ,	0.010***	0.001
	(3.58)			(3.40)	(0.48)
$HI_{it}*RISK_{it}$	-0.197*			-0.220***	0.025
	(-1.78)			(-2.89)	(0.25)
$HI_{it}*RISK^{2}_{it}$	0.603*			0.702***	-0.057
	(1.79)			(3.86)	(-0.19)
$Risk_{it}$	0.166**	0.066*	0.129*	-0.053***	0.173**
	(2.05)	(1.74)	(1.84)	(-2.68)	(2.38)
Equity <sub>it</sub>	,	,	,	,	0.184***
1 , ,					(24.06)
$Size_{it}$	-0.021***	-0.022***	-0.022***	-0.022***	-0.006***
**	(-14.12)	(-15.26)	(-14.84)	(-14.41)	(-4.14)
$Personnel_{it}$	-1.458***	-1.435***	-1.437***	-1.424***	-1.778***
	(-15.23)	(-14.92)	(-14.94)	(-14.71)	(-20.48)
Constant	0.415***	0.471***	0.428***	0.469***	0.117***
	(12.72)	(15.23)	(13.32)	(14.75)	(3.67)
No. of obs.	3,529	3,529	3,529	3,529	3,529

T-values in brackets, \*, \*\*, \*\*\* denotes significance at 10%, 5% and 1%. Blanks indicate that the coefficient of the variable is restricted to zero.

However, to better understand the economic significance of this potential U-shaped relationship, Figure 1 plots the marginal effect d(return)/d(focus) for different values of risk for all three types of diversification based on the estimated coefficients from (3b), (3c) and (3d). The range of risk is taken to be between 0% and 50%, which represents the minimum and the maximum values over our entire sample period. It should be noted that the mean (median) risk is about 3.4% (2.6%), while the 90th percentile is about 9%.

As can be seen in Figure 1, in our sample a small increase in industrial focus ( $HI_{it}$ ) has a rather minor and positive effect on return in the case of the mean (median) bank. For risk levels above 10%, the effect becomes slightly negative, but returns to a positive and sharply rising curve at a risk level of about 22% (corresponding to the 99th percentile of risk). Hence, we conclude that, within the range of risk levels observed in our sample, the marginal effect of industrial focus on return might indeed be U-shaped.

Figure 1 The marginal effect of focus on return for different risk values for HT, HR and HI based on the estimated coefficients from (3b), (3c) and (3d)



However, the result of the graph analysis is different for sectoral and geographical focus ( $HT_{it}$  and  $HR_{it}$ ). Here, an increase in focus also leads to rising returns for banks with risk levels below 12% and 27% respectively, but then the effect of focus stays negative and decreases for all of the risk levels observed. In fact, the effect becomes positive again only at hypothetical risk levels as high as 110% and 160%. Therefore, we suspect that the true impact of sectoral and geographical focus on return might be a linear or, at least, monotonous decrease with risk rather than a U-shaped relationship.

To further explore this issue, we have to overcome the drawback of the above test, ie the restrictions imposed by the parameterisation of the non-linearities between diversification, risk and return in (3). Richer patterns of non-linearity can be detected with non-parametric methods. To this end, we follow the example of Acharya et al. (2006) and define a set of dummy variables which measure different risk levels. The dummy variables are defined as follows:

$$D_{I} = 1$$
 if  $Risk^{[10]} < Risk_{it} \le Risk^{[25]}$  and zero otherwise,  
 $D_{2} = 1$  if  $Risk^{[25]} < Risk_{it} \le Risk^{[50]}$  and zero otherwise,  
 $D_{3} = 1$  if  $Risk^{[50]} < Risk_{it} \le Risk^{[75]}$  and zero otherwise,  
 $D_{4} = 1$  if  $Risk^{[75]} < Risk_{it} \le Risk^{[90]}$  and zero otherwise,  
 $D_{5} = 1$  if  $Risk_{it} \ge Risk^{[90]}$  and zero otherwise,

where  $Risk^{[p]}$  is the  $p^{th}$  percentile of  $Risk_{it}$ . We then interact the dummies with the Herfindahl Indices and regress the resulting variables on risk. Hence, our model has the following form:

$$\operatorname{Re} turn_{it} = \beta_{0} + \beta_{1}HT_{it} + \beta_{2}HR_{it} + \beta_{3}HI_{it} + \sum_{n=4}^{6} \beta_{n}X_{nit} + \sum_{k=1}^{5} \beta_{1k}D_{kit}HT_{it} + \sum_{k=1}^{5} \beta_{2k}D_{kit}HR_{it} + \sum_{k=1}^{5} \beta_{3k}D_{kit}HI_{it} + \delta z_{t} + \varepsilon_{it}$$

$$(4)$$

The estimation results are shown in Table 5.

Table 5 Two-way fixed effects estimation of Equation (1) with interaction terms for different risk levels, alternative restrictions, dependent variable: *Return*<sub>it</sub>

	,	· •			
	(4a)	(4b)	(4c)	(4d)	(4e)
$HT_{it}$	0.026***	0.025***	, ,	, ,	-0.001
	(2.89)	7.08			(-0.10)
$D_I*HT_{it}$	-0.003	-0.011***			0.015*
	(-0.40)	(-6.56)			(1.93)
$D_2*HT_{it}$	-0.017*	-0.012***			0.002
	(-1.86)	(-6.15)			(0.02)
$D_3*HT_{it}$	-0.018*	-0.013***			0.006
<i>.</i>	(-1.86)	(-5.68)			(0.73)
$D_4*HT_{it}$	-0.013**	-0.014***			0.004
,	(-2.53)	(-5.43)			(0.44)
$D_5*HT_{it}$	-0.014	-0.014***			0.012
<i>J u</i>	(-1.28)	(-4.38)			(0.20)
$HR_{it}$	0.038***	( 110 0)	0.046***		0.046***
	(3.73)		(6.85)		(5.00)
$D_I*HR_{it}$	-0.004		-0.010***		-0.016**
	(-0.61)		(-7.08)		(-2.47)
$D_2*HR_{it}$	0.005		-0.010***		-0.006
2 111111	(0.67)		(-6.60)		(-0.82)
$D_3*HR_{it}$	0.006		-0.011***		-0.007
_ 3 <sub>li</sub>	(0.81)		(-6.19)		(0.33)
$D_4*HR_{it}$	0.011		-0.012***		-0.005
2 4 111111	(1.30)		(-5.92)		(0.52)
$D_5*HR_{it}$	0.004		-0.012***		-0.009
_ <i>y</i> !i	(0.51)		(-5.23)		(-1.13)
$HI_{it}$	0.016***		( - 1 - 2 )	0.023***	-0.025***
ıı	(3.00)			(4.23)	(-4.97)
$D_I*HI_{it}$	-0.076***			-0.081***	-0.055***
- 1u	(-7.22)			(-8.15)	(-5.80)
$D_2*HI_{it}$	-0.033***			-0.038***	-0.073***
2 11	(-4.30)			(-4.41)	(-10.35)
$D_3*HI_{it}$	-0.015**			-0.024***	0.029***
<i>5 11</i>	(-2.22)			(-3.59)	(4.62)
$D_4*HI_{it}$	-0.017***			-0.025	0.023***
2 4 11111	(-2.60)			(-4.00)	(3.86)
$D_5*HI_{it}$	-0.011			-0.019	0.031***
- 3 <i>u</i>	(-1.56)			(-2.84)	(4.77)
Equity <sub>it</sub>	(1.50)			( 2.0 1)	0.185***
<b>_q</b> y <sub>ll</sub>					(24.22)
$Size_{it}$	-0.019***	-0.022***	-0.021***	-0.020***	-0.009***
~.2011	(-13.77)	(-14.90)	(-14.47)	(-12.94)	(-6.36)
Personnel <sub>it</sub>	-1.382***	-1.500***	-1.505***	-1.288***	-1.670***
1 C. Somment	(-14.65)	(-15.63)	(-15.71)	(-13.20)	(-19.58)
Constant	0.373***	0.463***	0.422***	0.424***	0.151***
Consum	(12.19)	(15.07)	(13.31)	(13.27)	(5.22)
No. of obs.	3,529				
110. OJ 00s.		0,345 poleste * ** *** d			

T-values in brackets, \*, \*\*, \*\*\* denotes significance at 10%, 5% and 1%. Blanks indicate that the coefficient of the variable is restricted to zero.

It should be noted that the coefficients of the Herfindahl Indices which are not interactant with the dummy variables capture the impact of focus on return for the group of banks with the lowest risk level. Therefore, banks with the lowest risk level represent our benchmark group and other coefficients has to be interpreted as a deviation from this benchmark group. The coefficients of the Herfindahl Indices which are not interactant are generally positive (again the only exception being the equation which includes equity) and highly significant. In the specifications where the Herfindahl Indices are analysed separately, the coefficients of sectoral and geographical focus are significantly negative when interactant with the dummies. At the same time, they exhibit a slight decrease in magnitude with rising risk (see (4b) and (4c)). This pattern confirms the hypothesis that the benefits from focus are greater for low-risk banks than for banks with higher levels of risk.

Furthermore, similar to the parametric analysis above, the overall influence of sectoral and geographical focus on return stays positive for all levels of risk as the absolute magnitude of the (negative) coefficients of the terms interactant with the risk dummies is lower than the (positive) baseline coefficient of the respective Herfindahl Index without interaction. However, in the case of industrial focus, the overall impact on return is negative for moderate levels of risk (compare the coefficients for  $HI_{it}$  and  $D_1*HI_{it}$ ,  $D_2*HI_{it}$  and  $D_3*HI_{it}$  in (4d)). Additionally, a U-shaped relationship with return can be detected since the overall impact of industrial focus increases to slightly positive (though insignificant) levels for the banks with the greatest risk.

Although these patterns are less pronounced for the estimations (4a) and (4e), in all cases the coefficients reveal evidence of a non-linear relationship between diversification and risk, with a strong positive impact of Herfindahl Indices in the case of low-risk banks and a moderate or insignificant impact for higher-risk banks. The less pronounced results for sectoral and geographical diversification in specification (4a) as compared with (4b) and (4c) may be attributed to the lower degree of freedom in the estimation.

To sum up, the dummy variable approach provides strong evidence that the impact of a bank's portfolio diversification on its return depends on the bank's risk level. Industrial, sectoral and geographical focus yields the highest benefits when risk is low. The benefits from focus decrease and, hence, the benefits from diversification increase with rising risk levels. For industrial focus, the impact becomes insignificant for high risk levels. The findings in Tables 2 and 3, namely that industrial focus, on average, has a smaller impact on returns than does sectoral and geographical focus, can be attributed mostly to banks with moderate risk.

However, it is still difficult to test the hypothesis of a U-shaped form, since the classes which define the dummy variables are fixed heuristically and may be too rough to detect the underlying structure of the relationship between diversification, risk and return. In order to gain a more precise picture of the shape of the non-linearities, we performed a second non-parametric procedure. We classified the data set according to risk level. We then estimated (1) with a window of 1,000 observations shifting from the lowest risk level to the highest risk level. More precisely, we first of all used a sub-sample of 1,000 observations with the lowest risk level to estimate (1), then shifted the sample by one observation and repeated the estimation. The result is a series of roughly 2,500 estimations for  $\beta_1$ ,  $\beta_2$  and  $\beta_3$  of Equation (1), which are classified according to the risk level. Plotting the series provides information about the impact of risk on the relationship between focus and return. Figures 2, 3 and 4 represent estimations for the specifications (1b), (1c) and (1d).

As expected, all of the charts reveal that the coefficients for the Herfindahl Indices ( $\beta_i$ ) vary according to the risk level. Although some of the coefficients of the Herfindahl Indices fluctuate considerably, there is some evidence to suggest that the relationships are either U-shaped or monotonously decreasing. In addition, taken all together, the influences depicted are comparable to those derived using the parametric and dummy variables approaches. Sectoral focus ( $HT_{it}$ ), for example, has a positive coefficient for a low risk level. However, in the case of increasing risk values,  $\beta_1$  decreases and becomes slightly (and insignificantly) negative. Therefore, as with the former approaches, the effect of sectoral focus on return seems to monotonously decrease with risk.

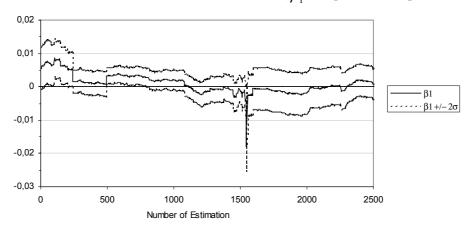
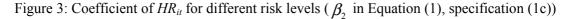


Figure 2: Coefficient of  $HT_{it}$  for different risk levels ( $\beta_1$  in equation (1), specification (1b))



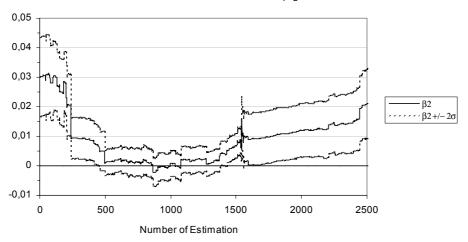
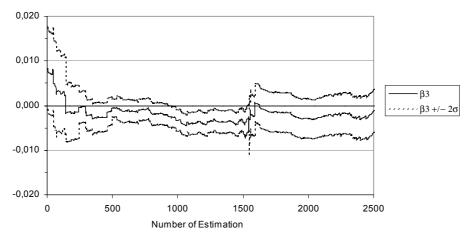


Figure 4: Coefficient of  $HI_{it}$  for different risk levels ( $\beta_3$  in Equation (1), specification (1d))



On the other hand, the coefficient of regional focus ( $HR_{ii}$ ) now shows – in contrast to former results – a U-shaped form. It is highly positive for low risk levels, decreases for moderate risk levels (although it remains positive) and then rises again for high-risk banks. Finally, industrial focus ( $HI_{ii}$ ) has a positive coefficient only at low risk levels. For moderate risk levels, it is slightly but significantly negative, while the coefficient becomes insignificantly negative for the highest risk. As such, the rolling window approach provides less distinct evidence of a U-shaped relationship between industrial focus and banks' returns than do the above results.

To sum up, in order to assess the impact of banks' portfolio diversification or focus on their returns at different risk levels, we applied and compared three different approaches. We first of all introduced non-linear terms in the base specification and then applied two non-parametric tests by interacting the Herfindahl Indices with dummies for different risk levels and using a rolling window approach for each type of diversification. Table 6 goes some way towards summarising the different results.

Table 6 clearly demonstrates that the benefits from industrial, sectoral and geographical diversification systematically and noticeably vary according to banks' risk levels. Therefore, banks' decisions on whether or not to diversify their loan portfolio should be closely linked to their current risk level. Moreover, the type of focus plays a crucial role. According to all three approaches, sectoral focus, for example, is moderately beneficial for low-risk banks, while its influence on return decreases monotonously for higher risk levels. This effect either stays positive for all banks or becomes only (insignificantly) negative for rather high risk levels. In the case of geographical focus, however, all of the results indicate a positive effect on return for all risk profiles. However, while the parametric and dummy variables approaches reveal a monotonous decline in this positive relationship for higher-risk banks, the rolling window approach clearly depicts a U-shaped form. Furthermore, for industrial focus, we found evidence of a U-shaped link to return, as the results show a positive influence for low risk, a significantly negative impact for moderate risk and almost no effect for very high-risk banks.

Hence, our analyses at least partly confirm Winton's theory that the diversification benefits are greatest at moderate risk levels.

Table 6 Comparison of the results of the impact of focus on return for different risk levels

Diele nargantila	0-1 <sup>th</sup>	-10 <sup>th</sup>	-25 <sup>th</sup>	-50 <sup>th</sup>	-75 <sup>th</sup>	-90 <sup>th</sup>	-99 <sup>th</sup>	-100 <sup>th</sup>
Risk percentile	perc.							
Sectoral focus (HTit)								
Parametric approach	(+)**	(+)**	(+)**	(+)**	(+)**	(+)**	(+/-)**	(-)**
(see Table 4, 4b)	(d)							
Dummy var. approach		(+)***	(+)***	(+)***	(+)***	(+)***	(+)***	
(see Table 5, 5b)			(d)	(d)	(d)	(d)	(c)	
Rolling window approach			(+)#	(+)#	(-)	(-)		
(see Figure 2)			(d)	(d)	(d)	(i)		
Geographical focus (HRit)								
Parametric approach	(+)**	(+)**	(+)**	(+)**	(+)**	(+)**	(+)**	(-)**
(see Table 4, 4c)	(d)							
Dummy var. approach		(+)***	(+)***	(+)***	(+)***	(+)***	(+)***	
(see Table 5, 5c)			(d)	(c)	(d)	(d)	(c)	
Rolling window approach			$(+)^{\#}$	(+)	(+)	(+)#		
(see Figure 3)			(d)	(d)	(i)	(i)		
Industrial focus (HIit)								
Parametric approach	(+)***	(+)***	(+)***	(+)***	(+)***	(+)***	(-)***	(+)***
(see Table 4, 4d)	(d)	(d)	(d)	(d)	(d)	(d)	(d/i)	(i)
Dummy var. approach		(+)***	(-)***	(-)***	(-)***	(-)	(+)	
(see Table 5, 5d)			(d)	(i)	(i)	(d)	(i)	
Rolling window approach			(+)	(-)	(-)#	(-)		
(see Figure 4)			(d)	(d)	(d)	(i)		

d/i/c indicate a decreasing / increasing / constant level for the respective risk interval. 
\*, \*\*, \*\*\* denote significance at 10%, 5% and 1%. 
# highlights that the interval  $\beta$ +/-  $2\sigma$  does not intersect the x-axis.

#### 4 Conclusions

Should banks diversify across different geographical regions and industrial sectors, or should they specialise in a few related fields? In this paper, we tried to shed some light on this question by empirically investigating the situation of German banks. By exploiting a unique data set of individual bank loan portfolios for the period from 1996 to 2002, we analysed the link between banks' profitability and their portfolio diversification across different industries, broader economic sectors and geographical regions. To the best of the authors' knowledge, this is the first paper to study the effect of all three types of diversification jointly based on micro-level data on German banks.

The relevant academic literature puts forward two conflicting theories concerning the optimum degree of diversification. While traditional banking theory based on a delegated monitoring argument (see, for example, Diamond (1984) and Boyd and Prescott (1986)) recommends that the optimum organisation of a bank is one where it is as diversified as possible, corporate finance theory suggests

that a firm should focus so as to obtain the greatest possible benefit from management's expertise and to reduce agency problems (see Jensen (1986), Denis et al. (1997) and Rajan et al. (2000)). Our results clearly support the latter theory, as the evidence we present indicates that each kind of diversification tends to lower German banks' returns, ie focusing generally leads to greater profitability benefits.

However, the impact of all types of diversification on banks' returns changes according to the risk level. While the effect of sectoral focus on return declines monotonously with increasing risk, there is mixed evidence to suggest either a monotonously decreasing or a U-shaped relationship for regional focus as well as a rather distinct indication of a U-shape with respect to industrial focus. Therefore, our results at least partly confirm Winton's theory regarding poor monitoring incentives for high-risk banks, which – in terms of empirically testable hypotheses – implies that the relationship between return and focus should be non-linear and U-shaped in risk.

Furthermore, in our data, diversification improves banks' profitability only in the case of moderate risk levels and industrial diversification. Hence, from a policy point of view, regarding the decision on whether to increase the degree of their industrial, sectoral or geographical diversification banks should evaluate carefully the actual riskiness of their activities.

#### References

Acharya, V., I. Hasan and A. Saunders (2006), Should Banks be Diversified? Evidence from Individual Bank Loan Portfolios, Journal of Business, 79(3), 1355-1413.

Barth, J.R. and D.E. Nolle and T. Phumiwasana and G. Yago (2002), A Cross-Country Analysis of the Bank Supervisory Framework and Bank Performance, Office of the Comptroller of the Currency Economic and Policy Analysis Working Paper 2002-2.

Basel Committee on Banking Supervision (2004), International Convergence of Capital Measurement and Capital Standards: A Revised Framework, June.

Berger, A. and R. DeYoung (2001), The Effects of Geographic Expansion on Bank Efficiency, Journal of Financial Services Research, 19, No 2/3, 163-185.

Berger, P.G. and E. Ofek (1995), Diversification's Effect on Firm Value, Journal of Financial Economics, 37, 39-65.

Bohn, J., N. Arora, and I. Korablev (2005), Power and Level Validation of the EDFTM Credit Measure in the U.S. Market, White Paper, Moody's KMV Company, March.

Boyd, J., and E. Prescott (1986), Financial Intermediary Coalitions, Journal of Economic Theory, 38, 211-232.

Buch, C.M., J.C. Driscoll and C. Ostergaard (2004), Cross-Border Diversification in Bank Asset Portfolios, Working Paper No. 04-26, Federal Reserve Board of Washington.

Caprio, G. and B. Wilson (1997), On Not Putting All the Eggs in One Basket: the Role of Diversification in Banking, Working Paper, World Bank.

Dahl, D. and A. Logan (2003), Granularity and International Diversification: An Empirical Analysis of Overdue Claims at Banks, mimeo, Bank of England.

DeLong, G. (2001), Stockholder Gains from Focusing Versus Diversifying Bank Mergers, Journal of Financial Economics, 59(2), 221-252.

Demirgüc-Kunt, A. and H. Huizinga (1998), Determinants of Commercial Bank Interest Margins and Profitability: Some International Evidence, Working Paper, World Banks.

Denis, D., D. Denis, and A. Sarin (1997), Agency Problems, Equity Ownership, and Corporate Diversification, Journal of Finance, 52, 135-160.

Deutsche Bundesbank (1998a), Instruction Sheet for the Reporting of Large Exposures of 3 million

Deutsche Mark or More Pursuant to Sections 13 to 14 of the Banking Act, in: Banking Regulations 7.

Deutsche Bundesbank (1998b), Monthly Report, August 1998b.

Diamond, D. (1984), Financial Intermediation and Delegated Monitoring, Review of Economic Studies, 59; 393-414.

Gehrig, T. (1998), Screening, Cross-Border Banking, and the Allocation of Credit, Review in Economics, 52, 387-407.

Hughes, J., W. Lang, L. Mester and C. Moon (1996), Safety in Numbers? Geographic Diversification and Bank Insolvency Risk, Working Paper No. 96-14, Federal Reserve Bank of Philadelphia.

J. P. Morgan (1997), "CreditMetrics-Technical Document", New York, April.

Jensen, M.C. (1986), Agency Costs of Fee Cash Flow, Corporate Finance, and Takeovers, American Economic Review, 76, 323-329.

Klein and Saidenberg (1998), Diversification, Organization, and Efficiency: Evidence from Bank Holding Companies, Center for Financial Institutions Working Papers 97-27, Wharton School Center for Financial Institutions, University of Pennsylvania.

Morgan, D.P. and K. Samolyk (2003), Geographic Diversification in Banking and Its Implications for Bank Portfolio Choice and Performance, the paper was presented at the BIS Workshop "Banking and Financial Stability", 20-21 March 2003

Nestmann, T., M. Wedow and N. v. Westernhagen (2003), A Micro Data-Set on Foreign Claims of German Banks, Deutsche Bundesbank, mimeo.

Quagliariello, M. (2004), Banks' Performance Over the Business Cycle: A Panel Analysis on Italian Intermediaries, Discussion Paper 04/17, University of York.

Rajan, R. and H. Servaes and L. Zingales (2000), The Cost of Diversity: The Diversification Discount and Inefficient Investment, The Journal of Finance, LV (1), 35-80.

Shaffer, S. (1998), The Winner's Curse in Banking, Journal of Financial Intermediation, 4, 158-187.

Stiroh, K. (2004), Diversification in Banking: Is Noninterest Income the Answer?, Journal of Money, Credit and Banking, 36(5), 853-882.

Stiroh, K. and A. Rumble (2003), The Dark Side of Diversification for U.S. Bank Holding Companies, unpublished paper, Federal Reserve Bank of New York.

Winton, A. (1999), Don't Put All Your Eggs in One Basket? Diversification and Specialization in Lending, Working Paper No. 00-16, University of Minnesota.

# The following Discussion Papers have been published since 2005:

# **Series 1: Economic Studies**

2	2005	Financial constraints and capacity adjustment in the United Kingdom – Evidence from a large panel of survey data  Common stationary and non-stationary	Ulf von Kalckreuth Emma Murphy
		factors in the euro area analyzed in a large-scale factor model	Sandra Eickmeier
3	2005	Financial intermediaries, markets, and growth	F. Fecht, K. Huang, A. Martin
4	2005	The New Keynesian Phillips Curve	
		in Europe: does it fit or does it fail?	Peter Tillmann
5	2005	Taxes and the financial structure of German inward FDI	Fred Ramb A. J. Weichenrieder
6	2005	International diversification at home and abroad	Fang Cai Francis E. Warnock
7	2005	Multinational enterprises, international trade, and productivity growth: Firm-level evidence from the United States	Wolfgang Keller Steven R. Yeaple
8	2005	Location choice and employment decisions: a comparison of German and Swedish multinationals	S. O. Becker, K. Ekholm, R. Jäckle, MA. Muendler
9	2005	Business cycles and FDI: evidence from German sectoral data	Claudia M. Buch Alexander Lipponer
10	2005	Multinational firms, exclusivity, and the degree of backward linkages	Ping Lin Kamal Saggi

11	2005	Firm-level evidence on international stock market comovement	Robin Brooks Marco Del Negro
12	2005	The determinants of intra-firm trade: in search for export-import magnification effects	Peter Egger Michael Pfaffermayr
13	2005	Foreign direct investment, spillovers and absorptive capacity: evidence from quantile regressions	Sourafel Girma Holger Görg
14	2005	Learning on the quick and cheap: gains from trade through imported expertise	James R. Markusen Thomas F. Rutherford
15	2005	Discriminatory auctions with seller discretion: evidence from German treasury auctions	Jörg Rocholl
16	2005	Consumption, wealth and business cycles: why is Germany different?	B. Hamburg, M. Hoffmann, J. Keller
17	2005	Tax incentives and the location of FDI: evidence from a panel of German multinationals	Thiess Buettner Martin Ruf
18	2005	Monetary Disequilibria and the Euro/Dollar Exchange Rate	Dieter Nautz Karsten Ruth
19	2005	Berechnung trendbereinigter Indikatoren für Deutschland mit Hilfe von Filterverfahren	Stefan Stamfort
20	2005	How synchronized are central and east European economies with the euro area? Evidence from a structural factor model	Sandra Eickmeier Jörg Breitung
21	2005	Asymptotic distribution of linear unbiased estimators in the presence of heavy-tailed stochastic regressors and residuals	JR. Kurz-Kim S.T. Rachev G. Samorodnitsky

22	2005	The Role of Contracting Schemes for the Welfare Costs of Nominal Rigidities over	Mali Davi
		the Business Cycle	Matthias Paustian
23	2005	The cross-sectional dynamics of German business cycles: a bird's eye view	J. Döpke, M. Funke S. Holly, S. Weber
24	2005	Forecasting German GDP using alternative factor models based on large datasets	Christian Schumacher
25	2005	Time-dependent or state-dependent price setting? – micro-evidence from German metal-working industries –	Harald Stahl
26	2005	Money demand and macroeconomic uncertainty	Claus Greiber Wolfgang Lemke
27	2005	In search of distress risk	J. Y. Campbell, J. Hilscher, J. Szilagyi
28	2005	Recursive robust estimation and control without commitment	Lars Peter Hansen Thomas J. Sargent
29	2005	Asset pricing implications of Pareto optimality with private information	N. R. Kocherlakota Luigi Pistaferri
30	2005	Ultra high frequency volatility estimation with dependent microstructure noise	Y. Aït-Sahalia, P. A. Mykland, L. Zhang
31	2005	Umstellung der deutschen VGR auf Vorjahres- preisbasis – Konzept und Konsequenzen für die aktuelle Wirtschaftsanalyse sowie die ökono- metrische Modellierung	Karl-Heinz Tödter

32	2005	Determinants of current account developments in the central and east European EU member states – consequences for the enlargement of the euro erea	Sabine Herrmann Axel Jochem
33	2005	An estimated DSGE model for the German economy within the euro area	Ernest Pytlarczyk
34	2005	Rational inattention: a research agenda	Christopher A. Sims
35	2005	Monetary policy with model uncertainty: distribution forecast targeting	Lars E.O. Svensson Noah Williams
36	2005	Comparing the value revelance of R&D reporting in Germany: standard and selection effects	
37	2005	European inflation expectations dynamics	J. Döpke, J. Dovern U. Fritsche, J. Slacalek
38	2005	Dynamic factor models	Sandra Eickmeier Jörg Breitung
39	2005	Short-run and long-run comovement of GDP and some expenditure aggregates in Germany, France and Italy	Thomas A. Knetsch
40	2005	A"wreckers theory" of financial distress	Ulf von Kalckreuth
41	2005	Trade balances of the central and east European EU member states and the role of foreign direct investment	Sabine Herrmann Axel Jochem
42	2005	Unit roots and cointegration in panels	Jörg Breitung M. Hashem Pesaran

1	2006	The dynamic relationship between the Euro overnight rate, the ECB's policy rate and the term spread	Dieter Nautz Christian J. Offermanns
2	2006	Sticky prices in the euro area: a summary of new micro evidence	Álvarez, Dhyne, Hoeberichts Kwapil, Le Bihan, Lünnemann Martins, Sabbatini, Stahl Vermeulen, Vilmunen
3	2006	Going multinational: What are the effects on home market performance?	Robert Jäckle
4	2006	Exports versus FDI in German manufacturing: firm performance and participation in international markets	Jens Matthias Arnold Katrin Hussinger
5	2006	A disaggregated framework for the analysis of structural developments in public finances	Kremer, Braz, Brosens Langenus, Momigliano Spolander
6	2006	Bond pricing when the short term interest rate follows a threshold process	Wolfgang Lemke Theofanis Archontakis
7	2006	Has the impact of key determinants of German exports changed? Results from estimations of Germany's intra euro-area and extra euro-area exports	Kerstin Stahn
8	2006	The coordination channel of foreign exchange intervention: a nonlinear microstructural analysis	Stefan Reitz Mark P. Taylor
9	2006	Capital, labour and productivity: What role do they play in the potential GDP weakness of France, Germany and Italy?	Antonio Bassanetti Jörg Döpke, Roberto Torrini Roberta Zizza
10	2006	Real-time macroeconomic data and ex ante predictability of stock returns	J. Döpke, D. Hartmann C. Pierdzioch

11	2006	The role of real wage rigidity and labor market frictions for unemployment and inflation dynamics	Kai Christoffel Tobias Linzert
12	2006	Forecasting the price of crude oil via convenience yield predictions	Thomas A. Knetsch
13	2006	Foreign direct investment in the enlarged EU: do taxes matter and to what extent?	Guntram B. Wolff
14	2006	Inflation and relative price variability in the euro area: evidence from a panel threshold model	Dieter Nautz Juliane Scharff
15	2006	Internalization and internationalization under competing real options	Jan Hendrik Fisch
16	2006	Consumer price adjustment under the microscope: Germany in a period of low inflation	Johannes Hoffmann Jeong-Ryeol Kurz-Kim
17	2006	Identifying the role of labor markets for monetary policy in an estimated DSGE model	Kai Christoffel Keith Küster Tobias Linzert
18	2006	Do monetary indicators (still) predict euro area inflation?	Boris Hofmann
19	2006	Fool the markets? Creative accounting, fiscal transparency and sovereign risk premia	Kerstin Bernoth Guntram B. Wolff
20	2006	How would formula apportionment in the EU affect the distribution and the size of the corporate tax base? An analysis based on German multinationals	Clemens Fuest Thomas Hemmelgarn Fred Ramb

21	2006	Monetary and fiscal policy interactions in a New		
		Keynesian model with capital accumulation	Campbell Leith	
		and non-Ricardian consumers	Leopold von Thadden	
22	2006	Real-time forecasting and political stock market anomalies: evidence for the U.S.	Martin Bohl, Jörg Döpke Christian Pierdzioch	
23	2006	A reappraisal of the evidence on PPP: a systematic investigation into MA roots in panel unit root tests and their implications	Christoph Fischer Daniel Porath	
24	2006	Margins of multinational labor substitution	Sascha O. Becker Marc-Andreas Mündler	

# **Series 2: Banking and Financial Studies**

1	2005	Measurement matters – Input price proxies and bank efficiency in Germany	Michael Koetter
2	2005	The supervisor's portfolio: the market price risk of German banks from 2001 to 2003 – Analysis and models for risk aggregation	Christoph Memmel Carsten Wehn
3	2005	Do banks diversify loan portfolios?  A tentative answer based on individual bank loan portfolios	Andreas Kamp Andreas Pfingsten Daniel Porath
4	2005	Banks, markets, and efficiency	F. Fecht, A. Martin
5	2005	The forecast ability of risk-neutral densities of foreign exchange	Ben Craig Joachim Keller
6	2005	Cyclical implications of minimum capital requirements	Frank Heid
7	2005	Banks' regulatory capital buffer and the business cycle: evidence for German savings and cooperative banks	Stéphanie Stolz Michael Wedow
8	2005	German bank lending to industrial and non- industrial countries: driven by fundamentals or different treatment?	Thorsten Nestmann
9	2005	Accounting for distress in bank mergers	M. Koetter, J. Bos, F. Heid C. Kool, J. Kolari, D. Porath
10	2005	The eurosystem money market auctions: a banking perspective	Nikolaus Bartzsch Ben Craig, Falko Fecht
11	2005	Financial integration and systemic risk	Falko Fecht Hans Peter Grüner

12	2005	Evaluating the German bank merger wave	Michael Koetter
13	2005	Incorporating prediction and estimation risk in point-in-time credit portfolio models	A. Hamerle, M. Knapp, T. Liebig, N. Wildenauer
14	2005	Time series properties of a rating system based on financial ratios	U. Krüger, M. Stötzel, S. Trück
15	2005	Inefficient or just different? Effects of heterogeneity on bank efficiency scores	J. Bos, F. Heid, M. Koetter, J. Kolatri, C. Kool
01	2006	Forecasting stock market volatility with macroeconomic variables in real time	J. Döpke, D. Hartmann C. Pierdzioch
02	2006	Finance and growth in a bank-based economy: is it quantity or quality that matters?	Michael Koetter Michael Wedow
03	2006	Measuring business sector concentration by an infection model	Klaus Düllmann
04	2006	Heterogeneity in lending and sectoral growth: evidence from German bank-level data	Claudia M. Buch Andrea Schertler Natalja von Westernhagen
05	2006	Does diversification improve the performance of German banks? Evidence from individual bank loan portfolios	Evelyn Hayden Daniel Porath Natalja von Westernhagen

# Visiting researcher at the Deutsche Bundesbank

The Deutsche Bundesbank in Frankfurt is looking for a visiting researcher. Among others under certain conditions visiting researchers have access to a wide range of data in the Bundesbank. They include micro data on firms and banks not available in the public. Visitors should prepare a research project during their stay at the Bundesbank. Candidates must hold a Ph D and be engaged in the field of either macroeconomics and monetary economics, financial markets or international economics. Proposed research projects should be from these fields. The visiting term will be from 3 to 6 months. Salary is commensurate with experience.

Applicants are requested to send a CV, copies of recent papers, letters of reference and a proposal for a research project to:

Deutsche Bundesbank Personalabteilung Wilhelm-Epstein-Str. 14

D - 60431 Frankfurt GERMANY