FARM BENEFITS AND NATURAL RESOURCE PROJECTS IN HONDURAS AND EL SALVADOR

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1. Introduction

Peasant farmers in Central America typically grow subsistence crops -mainly corn and beans- in sloping and marginal uplands. These farmers often use production practices that are highly erosive and as a result the region contains large areas with significantly degraded land (Oldeman *et al.*, 1990; Barbier, 2000). The development path on the Central American agriculture has usually involved the migration of poor or landless farmers to public or open-access lands, who then clear forest areas, and cultivate staple crops for a few years, before moving on to clear new plots. As migration increases and the pool of available land declines, farmers use land more intensively. Those unable to purchase agrochemicals face declines in soil fertility and productivity, and farming becomes unsustainable (Neill and Lee, 2001). For many resource-poor farmers, soil degradation exacerbates the continuous struggle for food security.

To respond to this deteriorating situation, local governments and international organizations have undertaken a series of public investments focusing on poverty reduction and the promotion of conservation technologies. A major thrust of these projects has been to foster long-term environmental sustainability by enhancing management capabilities and decreasing vulnerability of communities to natural disasters.

Two of these natural resource management projects that deserve special attention due to their magnitude and scope are: (i) the Environmental Program for El Salvador (PAES); and (ii) the Natural Resource Management Program in the CAJON-Basin, Honduras. PAES is treated as three separate projects, PAES1, PAES2 and PAES3, since each one has been implemented in a different area by a different international consortium. These projects seek to generate both on-site and off-site benefits. The on-site benefits consist of increased household income through improved soil productivity, the adoption of conservation technologies, and product diversification. Off-site benefits are reflected on improved environmental conditions and positive externalities associated with water, soil and forestry resources. The CAJON project has concluded recently, while the PAES project will finish by the end of 2004. Most of the analyses done to examine the performance of these projects have focused on assessing physical targets and institutional aspects (for example, Barbier and Flores, 2003; PNUD, 1999; PAES/IICA2003; PAES/CARE, 2003; PAES/ABT–Winrock, 2001). So far, no attempt has been made to evaluate the off-site impacts of the projects. In turn, Bravo-Ureta *et al.* (2003) have performed an assessment of the on-site benefits, using farm-level data collected among beneficiaries.

The objective of this paper is to measure the success of the poverty reduction strategies promoted by the projects, by studying the relationships among technology adoption, product diversification and household income among participant farmers.

The rest of the paper is organized as follows. The second section presents the data set and the methodological framework. The third section discusses the econometric results and the fourth section summarizes and presents the main conclusions.

2. Data and Methodology

The data used consist of detailed farm-level information obtained from surveys applied to representative samples of project beneficiaries in El Salvador and Honduras. These data were collected and analyzed by Bravo-Ureta et al., (2003), as part of Technical Cooperation # 01-08-01-1-RS between the University of Connecticut and the Inter-American Development Bank. The sampling procedure in Honduras started with a grouping of the 240 communities participating in the CAJON project into three agroecological zones according to elevation: High, Medium, and Low. From these communities, 48 (20%) were randomly selected and distributed evenly by agroecological zone. Within each community, four farm households were randomly chosen. Overall, 210 beneficiary *campesinos* were interviewed, of which 35 are also contact farmers or extensionists. In El Salvador, data were gathered by surveying a sample of 530 farm households belonging to 102 communities within the regions of Resbaladero and Texistepeque (PAES1), San Juan Opico and Nueva Concepción (PAES2), and Tenancingo and Guazapa (PAES3). 175 respondents belong to PAES1, 177 to PAES2 and 178 to PAES3. The sample was stratified by region to reflect the geographical distribution of the three projects.

Table 1 includes descriptive statistics of the dependent variables used in the econometric model. The data are presented by project and aggregated for the whole sample. The last column of the table presents the results of an ANOVA test performed to verify whether the observed mean differences among projects are statistically significant. The bold figures indicate statistically significant mean differences at least at the 1% level.

The study of the relationships among technology adoption, product diversification and poverty reduction requires an understanding of the main sources of rural income, including farm and off-farm sources. These relationships can be stylized as a set of functional links, as illustrated in Figure 1, which is based on a framework presented by Minten and Zeller (2000)¹. In this framework, the household is taken as a single decision-making body; therefore, decisions that affect the way resources are allocated among household members, the so-called intra-household resource allocation, are not taken into account (Ellis, 1998).

The household's set of resources (assets) consists of natural capital (land, livestock, durables and environmental quality), human capital (education, experience, demographic attributes) and social capital (access to social networks and institutions). In addition, the household's decision-making process is also affected by a set of external factors, including the socio-economic and agro-ecological environment, input, output and financial markets, prices, wages and infrastructure. The household's set of resources determines the allocation of labor and effort between farm and off farm activities.

The mapping of assets to household income through both off and on farm activities can conceptually be considered as a production process, with assets corresponding to factors of production and income as the output (Barrett and Reardon, 2000). The allocation of assets to each activity is assumed to maximize household income subject to a set of constrains. Households will allocate assets in a manner that equates the marginal value product across activities or will allocate assets entirely to one

¹ A similar framework can be found in Mbaga-Semgalawe and Fomer (2000).

activity with higher returns. A key feature of this approach is that households simultaneously determine the allocation of assets to different activities (Winters *et al.*, 2002).

Moreover, the natural resource management projects under study motivate farmers to adopt soil conservation technologies and to diversify their product mix. These technological changes should improve farm production and productivity and, consequently, should be reflected in a greater household income, which is the sum of farm and off-farm income. Finally, income improvement is considered a necessary condition for sustainability of the changes introduced by the projects. The relationships depicted in Figure 1 are modeled considering that, for a given agricultural year, the household decisions that affect the allocation of labor between farm and off-farm activities, as well as the adoption of conservation practices and output mix, also determine total farm income.

A growing body of research speculates that farmers' land allocation decisions may have a significant influence on the pattern of technology adoption (Moreno and Sunding, 2003; Winters *et al.*, 1998; Barbier, 1990; Hopkins *et al*, 1999; Kruseman, 2000). For instance, Murray (1994) argues that farmers are more open to soil conservation measures when they are presented not as the principal component of rural extension, but rather as secondary items in a menu featuring innovations with good short-term income-generating potential.

Accordingly, if technology choice is influenced by the same factors explaining land allocation, then there is a simultaneity problem that must be addressed (Moreno and Sunding, 2003). In line with this reasoning, it is reasonable to assume that within a period of time, the choice of output mix and soil conserving investments are jointly determined along with the allocation of farm and off-farm income. Therefore, technology adoption and crop diversification are simultaneously estimated along with the income equations. The empirical model can be represented as:

(1)	Staple Income Equation	$Y^{S} = Y^{S} (NK, HK, SK, ACP, OD, VI)$
(2)	Cash Crop Income Equation	$Y^{C} = Y^{C} (NK, HK, SK, ACP, OD, VI)$
(3)	Off-Farm Income Equation	$Y^{NF} = Y^{NF}(NK, HK, SK)$
(4)	Conservation Equation	ACP = f(NK, HK, SK, FV)
(5)	Diversification Equation	OD = f(NK, HK, SK, FV),

where

ACP = Adopted Conservation Practices;

- OD = Number of items produced by the firm (output mix) over and above the subsistence crops (corn and beans);
- Y^{s} = Income obtained from corn and beans;
- Y^{c} = Income obtained from farm output other than corn and beans;
- Y^{NF} = Income obtained from off-farm employment;
- NK = Natural Capital (total land, tenure, slope, distance from house to parcel);
- HK = Human Capital (family size, gender of household head, agricultural experience, age of household head, education, perceives erosion, farmer is extensionist);
- *SK* = Social Capital (access to credit, access to output markets, access to labor markets, participates in social organizations);
- PV = Project Variables (adoption of conservation practices and structures, crop diversification, receives extension, receives training, frequency of extension visits, years with project, PAES1, PAES2, PAES3, CAJON); and
- *VI* = Variable Inputs (Total expenditures in variable inputs).

Equation (1) and (2) include the variable ACP as an explanatory variable, i.e., soil conservation practices are inputs in the production of staples and cash crops, as proposed by Thampapillai and Anderson $(1994)^2$.

Consideration must be given to the fact that the dependent variables staples, cash crops and off-farm income are observed only for a restricted, non-random sample of those farmers who decide to participate in those activities. Therefore, sample selection bias is likely to arise and Heckman's two-stage approach for the estimation of selectivity bias will be employed to address this issue (Heckman, 1976).

A key feature of the framework presented above is that households simultaneously determine the allocation of assets to different income-generating activities. In this context of simultaneity, the possible dependency between regressors and residuals implies that ordinary least squares (OLS) estimates are biased and inconsistent. The standard prescription for correcting such cases is the instrumental variable (IV) technique. If instrumental variables that are correlated with the explanatory variables but uncorrelated with the error terms are available, then IV regression yields consistent estimates (Deaton, 1997). An additional problem that needs to be addressed, given the cross-section nature of the household surveys under study, is the likely existence of heteroskedasticity (Deaton, 1997; Judge *et al.*, 1988). Heteroskedasticity affects not the consistency but the efficiency of the IV coefficient estimates. Moreover,

 $^{^2}$ in a cross-sectional sample of households, even without market imperfections, prices can be safely assumed to be constant across sections of farmers and therefore can be ignored (Holden *et al.*, 2001; Van Dusen and Taylor, 2003).

the IV estimates of the standard errors are inconsistent, preventing valid inference, since the usual forms of the diagnostic tests for endogeneity and overidentification are invalid. Using heteroskedasticity-consistent or "robust" standard errors can solve these problems and, since the seminal work by Hansen (1982), the Generalized Method of Moments (GMM) has become the standard remedy for heteroskedasticity in virtually every field of applied economics (Baum *et al.* (2003). GMM provides an inference procedure that is robust in the presence of heteroskedasticity (Kieffer and Vogelsang, 2002).

Since several explanatory variables have direct as well as indirect effects on income, it is necessary to calculate marginal effects in order to measure their overall effect on income-generating activities. Therefore, the total effect of an additional unit of a particular explanatory variable x_i is the sum of the direct effect of x_i on the income-generating activities plus the indirect effects on adoption and diversification. In the case of the variable x_i , the total effect $dY/d x_i$ is equal to:

$$\frac{d Y}{d x_{1}} = \frac{\delta \text{ staples}}{\delta x_{1}} + \frac{\delta \text{ cashcrops}}{\delta x_{1}} + \frac{\delta \text{ offfarm}}{\delta x_{1}} + \frac{\delta \text{ (ffarm}}{\delta x_{1}} + \frac{\delta \text{ (ffarm}}{\delta P})}{\delta P} + \frac{\delta \text{ (ffarm}}{\delta P}) \cdot \frac{\delta P}{\delta x_{1}} + \frac{\delta \text{ (ffarm}}{\delta D} + \frac{\delta \text{ cashcrops}}{\delta D} + \frac{\delta \text{ offfarm}}{\delta D}) \cdot \frac{\delta D}{\delta x_{1}}$$

where Y = income, P = practices, and D = diversification. In the case of dummy variables, the partial derivatives do not exist and the marginal effects are simply the value of the corresponding coefficients, since the estimated system does not include any logarithmic or exponential transformations.

We chose a linear functional form because of the presence of zero values in the dependent variables. Widely used alternatives are logarithmic or semi-logarithmic specifications, wherein to avoid the calculation of the logarithm of zero, each observation is transformed by adding a small positive constant. However, this technique precludes the possibility of some households being at a corner solution (Carson and Cameron, 2000). Moreover, there is evidence that the change of zero-values to facilitate the logarithmic transformation is extremely sensitive to the value of the constant chosen (Soloaga, 2000). Recent studies using linear specifications include Wilkins *et al.*, 2001; Finan *et al.*, 2004; Winters *et al.*, 2002; Taylor and Yúñez-Naude, 2000; de Janvry and Sadoulet, 2001; Dutilly-Diane *et al.*, 2003.

3. Results

This section reports the estimation results of the selectivity-corrected fiveequation system of conservation practices, farm-output diversification, staple crop income, cash crop income and off-farm income. The results are displayed in Table 3, while Table 4 shows the marginal effects computed using the coefficients from Table 3 and equation (6). Before proceeding with the parametric analysis, we explore the robustness of the results, considering the possibility of endogeneity (Hausman's specification test), heteroskedasticity (White and Breusch-Pagan tests), and the validity of the instruments used in the IV estimation (*J*-statistic or OID).

The Hausman test of endogeneity was performed comparing the coefficients estimated with OLS and 3SLS specifications. The Hausman statistic is significantly

different from zero (p-value= <.0001), indicating that a system estimation (IV) is preferred over OLS.

The Breusch-Pagan and White tests of heteroskedasticity show that the hypothesis of homoskedasticity cannot be rejected in the diversification and cash crops equations. Conversely, both tests indicate that the estimation of the conservation practices; staples and off-farm income equations exhibit heteroskedastic residuals. Given the presence of heteroskedastic residuals in three of the five equations, the system is estimated using the GMM procedure.

The test of overidentification (OID) makes it possible to determine if the null hypothesis that the over-identifying instruments are valid. This test is performed using the value of the objective function minimized by the GMM procedure multiplied by the number of observations (*J*-statistic), which is distributed as a chi-squared with r-p degrees of freedom. The value r is the product of the number of instruments (65) times the number of equations (5) and p is the number of parameters in the system (144). The OID test fails to reject the null that the instruments are statistically valid (p-value=0.486).

Entries in Table 3 are the estimated absolute effects of one-unit changes in the corresponding explanatory variables on income-generating activities. Coefficients with a level of significance greater than 90% are indicated in bold. Overall, 68 out of 114 coefficients (60%) are statistically significant. The following discussion will focus mainly on the results regarding the income-generating equations for staples, cash crops and off-farm activities.

Table 3 shows that in all equations the Inverse Mills Ratios are significant except in staples, indicating that self-selection is an important factor in considering the incomegenerated from a particular activity and failure to control for it would lead to biased results. All significant Mills have negative signs, and a similar pattern can be found in Winters *et al.* (2000).

The coefficients for expenditures on farm inputs, labor and land allocated to the production of staples and cash crops are all significant and, as expected, positive, indicating the presence of well-behaved production functions for both staples and cash crops. Farm size (a proxy for wealth) is positively associated with off-farm income. White (1991) found land-rich households receiving the largest returns from non-farm enterprises in Java. Similar results are reported for Ecuador (Lanjouw and Lanjouw, 2001), El Salvador (Lanjouw, 2000), Mexico (de Janvry and Sadoulet, 2001; Winters *et al.* 2000) and Brazil (Ferreira and Lanjouw, 2001). As Lanjouw and Lanjouw (2001) point out, the relatively disadvantaged appear to face barriers to employment in the most attractive non-agricultural occupations.

Table 4 reports the corresponding marginal effects, wherein the entries are the returns of an extra unit of each explanatory variable on income from staples, cash crops and off-farm. The total income column is the sum of the marginal effects by activity, while the last column shows the marginal effect on total income as percentage of the predicted household income displayed in the last row.

The marginal income effects displayed in Table 4 reveal that, besides the large effect of gender, the greatest gains in household income are associated with activities directly related with the allocation of land. The marginal income value of gender indicates that when the household head is a man, household income is expected to be 80.0% higher than when the household head is a woman. The large income effect of gender is consistent with the findings by Finan *et al.*, (2004) among Mexican farmers. The second greatest marginal income effect is land allocated to cash crops (14.8%), followed by diversification (14.0%), conservation practices (13.9%), and land allocated to staples (10.6%).

Output diversification significantly decreases income from staple crops and greatly increases cash crop income. The overall effect on household income of adding one extra production activity to the farm plan is \$307.4, or equivalently, a total income gain of 14% (Table 4). These results highlight the strategic role of diversification in fighting rural poverty and are consistent with those reported by Nerlove *et al.* (1996), Delgado and Siamwalla (1997), and Ruben and Clemens (2000). However, the income gain due to a more diversified income portfolio does not occur without cost, since an extra item added to the farm plan implies a reduction in the production of corn and beans (staples) of \$135.4. This trade-off between diversification and subsistence food production suggests that switching to a more market-oriented production pattern may increase household food insecurity, especially in environments where institutional or market failures are prevalent (Immink and Alarcón, 2003; Von Braun, 1995; Quiroz and Valdés, 1995).

The marginal income value of conservation practices is \$304.9, which amounts to a 13.9% increase over average total income. The practices included in this variable comprise ground-cover technologies such as crop-mulch/residue management, green manure and conservation tillage. Ground cover is increasingly recognized not only as a crucial soil conservation component but also in terms of its potential effects on land productivity. Therefore, technologies emphasizing ground cover are more likely to be profitable (Erenstein, 1999; López-Pereira *et al.*, 1994). As in the case of output diversification, conservation practices have a positive association with cash crops but negative with staples. In sum, the adoption of conservation practices and output diversification are positive correlated and denote a switching away from more traditional staple production.

The positive association between conservation practices and income contrasts sharply with the effects of conservation structures. The marginal income effect of an additional conservation structure is -\$184.3, or equivalently, an 8.4% drop in household income. A substantial body of literature contains empirical evidence of poor private economic returns associated with conservation structures (Erenstein, 1999; Wiggins, 1981; Blaikie and Brookfield, 1987). It is increasingly recognized that structures are expensive to build and maintain whereas they add little to the productivity of the land in the short run (Shaxson *et al.*, 1989; Douglas, 1993). Lutz *et al.*, 1994 reports several case studies in Central America and the Caribbean where physical structures seemingly lessen the available area for cultivation. Examples include construction of cutoff drains in Costa Rica and terraces in Guatemala that reduced the effective cultivation area by 14%, and 15%, respectively. Such drawbacks may clearly affect the profitability of these conservation technologies. Further, terracing often entails movements of earth that

brings unproductive soil to the surface (Erenstein, 1999; de Graaff, 1996). McIntire (1994) reviewed 20 conservation techniques in Mexico and found that cultivation and cropping practices, including vegetative barriers, were superior to structures in terms of profitability. On the other hand, a combination of diversion ditches and live barriers in Guatemala appears to be substantially more profitable than terraces, even if much less effective to control erosion (Lutz *et al.*, 1994).

Surprisingly, farmer experience exhibits a negative association with household income. However, as shown in Table 3, most of the coefficients measuring experience are not significantly different from zero. Similar results regarding farmer experience can be found in Taylor and Yúñez-Naude (2000).

Education plays a significant and positive role in all income-generating activities. Other things being equal, an extra year of average household schooling is associated with a gain of \$107.7 or 4.9% in total household income. This result is in line with the apparent robustness of the evidence on the returns to education at the micro-economic level. Estimates for developing countries find that each additional year of schooling is associated with a 6-10% increase in earnings and similar patterns are found in developed countries (Besley and Burgess, 2003; Krueger and Lindhal, 2001). In Latin America, a one-year increase in the average education of the adult population can lead to increases of 3-5% in real GDP (Lau *et al.*, 1990; Glewwe, 1996).

The returns from schooling are the highest in cash crop production (\$45.5) followed by off-farm activities (\$40.4), and the lowest in food staple production (\$21.8). These returns to education by source of income are consistent with other findings in the

rural development literature (Finan *et al.*, 2004; Winters *et al.*, 2000; López and Romano, 2000; López and Thomas, 2000; Ellis, 1998). As Taylor and Yúñez-Naude (2000) affirm "farm households may reap rewards from schooling by abandoning one activity (e.g., traditional agriculture, in which returns from schooling may be limited), in favor of a new activity (e.g., modern agriculture or non-crop production, in which the returns are high" (p. 288). Our results are fully compatible with this observation.

In addition to schooling, several variables representing human capital have a significant and positive association with household income. These variables are erosion perception, training, years with extension, and frequency of extension visits.

4. Summary and Conclusions

A model of conservation adoption, diversification and household income, including farm and off-farm sources was formalized, wherein households simultaneously allocate assets to different activities. The mapping of assets to household income through both off and on farm activities can conceptually be considered as a production process, with assets corresponding to factors of production and income as the output. Either adoption of conservation technologies and farm output diversification are influenced by participation in natural resource management programs. Therefore, these technological improvements should foster farm production and productivity and, consequently, should be reflected in a greater household income,. Finally, household income improvement is considered a necessary condition for sustainability of the changes introduced by the projects. Overall, the results indicate that the variables more directly reflecting land allocation, such as area with staples and cash crops, output diversification and conservation practices are associated with the greatest gains in household income. Output diversification significantly decreases income from staple crops and greatly increases cash crop income. These results reaffirm the strategic role of diversification in fighting rural poverty. However, gains stemming from a more diversified income portfolio do not occur without cost, since an extra item added to the farm plan implies a reduction in the production of corn and beans (staples). This trade-off between diversification and subsistence food production suggests that switching to a more marketoriented production pattern may increase household food insecurity.

The marginal income value of conservation practices is \$304.9, which amounts to a 13.9% increase over average total income. The practices included in this variable comprise ground-cover technologies such as crop-mulch/residue management, green manure and conservation tillage. Ground cover is increasingly recognized not only as a crucial soil conservation component but also in terms of its potential effects on land productivity. Therefore, technologies emphasizing ground cover are more likely to be profitable. As in the case of output diversification, conservation practices have a positive association with cash crops but negative with staples. Seemingly, the diffusion of conservation practices and cash crops mark a switching away from more traditional staple production.

Contrasting with the positive association between conservation practices and income we found evidence of poor economic performance of conservation structures such as terraces, stonewalls and drainage ditches. The income effect of an additional conservation structure is equivalent to an 8.4% drop in household income. This finding is consistent with a substantial body of literature that recognizes that structures are expensive to build and maintain, they add little to land productivity in the short run and reduce the available area for cultivation. Such drawbacks may clearly affect the profitability of these conservation technologies.

Human capital, measured by erosion awareness, average education level, training, years of participation with rural extension, and frequency of extension visits, have a significant and positive association with household income. Education plays a significant and positive role in all income-generating activities. An extra year of average household schooling is associated with a 4.9% gain in total household income. The returns from schooling are the highest in cash crop production and off-farm activities and the lowest in food staple production. These returns to education by economic activity suggest that farm households may reap rewards from schooling by abandoning traditional agriculture, where returns from schooling are limited, in favor of modern agriculture or non-crop production, in which the returns are higher.

Off-farm income is positively associated with farm size. Seemingly, the relatively disadvantaged appear to face barriers to employment in the most attractive non-agricultural occupations.

	ALL PROJECTS	PAES1	PAES2	PAES3	CAJON	Test of Mean Differences				
Household Income										
Annual Farm Income (US \$)	1646.0	1502.5	1781.0	1357.1	1878.6	0.395				
Staples (US \$)	807.3	814.2	1040.8	847.2	576.0	0.000				
Cash Crops (US \$)	838.7	688.3	740.1	509.9	1302.6	0.028				
Off-Farm Income (US \$)	592.3	525.6	639.0	559.9	631.9	0.763				
Off-Farm Income: 1 if member receives off-farm	57.0	53.0	51.0	51.0	70.0	0.000				
Crop Diversification										
Diversification (#)	1.17	0.92	0.96	1.12	3.28	0.000				
Entropy	0.13	0.10	0.09	0.11	0.21	0.008				
Adoption of Conservation Technologies										
Number of Structures (#)	2.10	2.34	2.15	2.83	1.29	0.000				
Number of Practices (#)	1.78	1.50	1.69	1.68	2.17	0.000				
Adoption of Conservation Practices (%)										
Live Barriers	70.0	62.0	69.0	70.0	76.0	0.034				
Crop Residue Mulching	84.0	90.0	89.0	75.0	82.0	0.001				
Green Manure	3.0	0.0	2.0	1.0	9.0	0.000				
Minimum Tillage	15.0	0.0	0.0	0.0	50.0	0.000				
Adoption of Conservation Structures (%)										
Live Fences	71.0	59.0	76.0	86.0	64.0	0.000				
Forest Plantations	82.5	79.6	82.8	91.6	77.1	0.002				
Stone Walls	8.0	2.0	2.0	1.0	24.0	0.000				
Terraces	34.0	42.0	37.0	60.0	3.0	0.000				
Ditches	27.0	48.0	10.0	27.0	23.0	0.000				
Number of Observations	719	167	175	167	210					

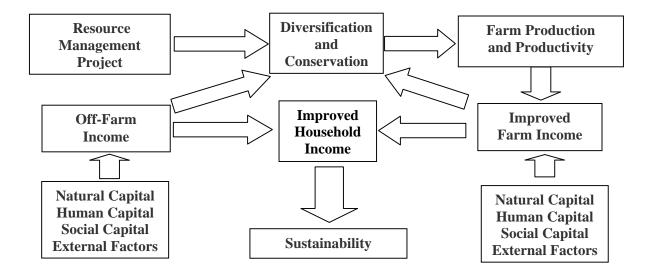
Table 1. Descriptive Statistics of Household Income, Output Diversification and Soil Conservation Technologies.

	ALL PROJECTS	PAES1	PAES2	PAES3	CAJON	Test of Mean Differences
Farm Characteristics						
Land: (Manzanas)	5.88	5.67	4.83	3.85	8.54	0.003
Tenure: 1 if owner (%)	74.0	79.0	75.0	80.0	65.0	0.002
Slope: 1 if average slope is greater than 15% (%)	61.0	54.0	59.0	50.0	77.0	0.000
Credit: 1 if farm uses credit (%)	18.0	5.0	11.0	15.0	38.0	0.000
Distance: Distance from house to parcel (Kms)	1.18	1.25	1.22	0.77	1.41	0.004
Farmer Characteristics		•			•	
Family Size: (#)	5.52	5.20	5.06	5.41	6.23	0.000
Education: Average household education (Years)	3.52	3.23	3.38	4.06	3.44	0.005
Gender: 1 if household head is a man (%)	86.0	87.0	78.0	83.0	96.0	0.000
Age: age of household head (Years)	47.80	48.78	48.83	48.34	45.74	0.107
Contact Farm: 1 if also an extension agent (%)	12.0	5.0	11.0	15.0	17.0	0.006
Erosion: 1 if perceives erosion as a problem (%)	88.0	84.0	86.0	87.0	93.0	0.056
Communal: 1 if member of a communal organization	64.0	47.0	54.0	56.0	94.0	0.000
Project Characteristics		1				
Years w/project	3.16	2.50	2.91	2.91	4.08	0.000
Extension: 1 if received extension visits (%)	88.0	80.0	82.0	88.0	98.0	0.000
Frequency of extension visits (#)	1.98	1.88	1.86	1.92	2.17	0.018
PAES1: 1 if belongs to PAES1 (%)		0.23				
PAES2: 1 if belongs to PAES2 (%)			0.24			
PAES3: 1 if belongs to PAES3 (%)				0.23		
CAJON: 1 if belongs to CAJON (%)					0.29	
Number of Observations	s 719	167	175	167	210	

Table 2. Descriptive Statistics of Farm, Farmer and Project Characteristics.

1 Manzana=0.7 Hectare

Figure 1. Relationship among Resource Management Projects, Household Income and Sustainability



	Prac	tices	Diversif	ication	Staples		Cash Crops		Off-Farm	
	Estimate	P-value	Estimate	P-value	Estimate	P-value	Estimate	P-value	Estimate	P-value
Practices					-94.11	0.00	399.05	0.00		
Structures					9.86	0.30	-194.12	0.00		
Diversification					-135.36	0.00	442.77	0.00		
Expenditures Staples					0.69	0.00				
Expenditures Cash Crops							0.93	0.00		
Labor Staples					3.43	0.00				
Labor Cash Crops							1.08	0.00		
Land w/Staples					232.03	0.00				
Land w/Cash Crops							325.68	0.00		
Farm size	0.00	0.00	0.01	0.00					16.24	0.00
Tenure	0.02	0.55	0.24	0.00	25.10	0.38	-153.95	0.01	-76.53	0.17
Education	-0.04	0.18	0.04	0.31	52.95	0.00	0.21	1.00	29.66	0.29
Education ²	0.00	0.55	-0.01	0.55	-9.18	0.00	15.24	0.06	3.05	0.64
Experience	-0.04	0.00	-0.01	0.47	8.86	0.15	23.71	0.11	-20.17	0.12
Experience ²	0.00	0.16	0.00	0.43	-0.09	0.27	-1.06	0.00	-0.14	0.42
Family size	0.01	0.43	0.02	0.16	-5.68	0.28	-32.15	0.01	2.54	0.79
Age	0.03	0.00	0.01	0.54	-5.81	0.38	16.27	0.30	30.72	0.01
Gender	-0.74	0.00	0.79	0.00	-58.95	0.49	1628.88	0.00	167.24	0.16
Hire Labor	-0.08	0.03	0.14	0.07					109.17	0.03
Market access	-0.10	0.00	-0.01	0.82					-122.75	0.00
Perception	0.13	0.05			-84.32	0.03	202.23	0.03	18.26	0.81
Credit	0.00	0.36	0.00	0.00	0.02	0.00	0.06	0.00		
Distance	-0.01	0.62	0.01	0.71					5.93	0.59
Slope	-0.04	0.29	0.03	0.62					121.78	0.00
Social	-0.02	0.64	0.00	0.96					-48.65	0.32
Contact Farm	0.15	0.00	0.08	0.47					-89.42	0.07
Training	0.18	0.00	0.05	0.56						
Years w/extension	-0.23	0.01	0.19	0.02						
Years w/extension ²	0.09	0.00	-0.06	0.06						
Visit frequency	0.04	0.01	-0.02	0.38						
PAES1	-0.71	0.00	-1.12	0.00	-516.35	0.00	3060.77	0.00	144.64	0.04
PAES2	-0.04	0.53	-1.87	0.00	-85.02	0.13	2047.25	0.00	369.48	0.00
PAES3	-0.27	0.00	-1.23	0.00	-206.08	0.00	2157.55	0.00	153.76	0.03
Mills	-1.48	0.00	-1.56	0.00	102.72	0.24	-477.03	0.00	-978.90	0.00
Constant	2.68	0.00	1.74	0.00	487.24	0.00	-4483.15	0.00	218.40	0.39
White Test for										
Heteroskedasticity		0.020		0.035		<.0001		0.263		0.000
Breusch-Pagan Test for										
Heteroskedasticity		0.020		0.035		<.0001		0.263		0.000

 Table 3. Selectivity-Corrected System Estimates of Conservation, Diversification and Income-Generating Equations

J-Statistics (OID): 209.6265 (p-value= 0.48623); Hausman's Test of Endogeneity: 226.6 (p-value= <.0001)

Number of Observations = 695; Coefficients with a level of significance greater than 90% are indicated in bold.

					% of Total
	Staples	_		Total Income	Income
Gender	-96.8	1685.2	167.2	1755.7	80.0
Land with Cash Crops		325.7		325.7	14.8
Divers	-135.4	442.8		307.4	14.0
Practices	-94.1	399.0		304.9	13.9
Land with Staples	232.0			232.0	10.6
Perception	-96.1	252.2	18.3	174.4	7.9
Hire Labor	-11.3	29.4	109.2	127.3	5.8
Slope	-0.9	-0.9	121.8	120.0	5.5
Education	21.8	45.5	40.4	107.7	4.9
Training	-23.4	93.3		69.8	3.2
Age	-10.2	33.4	30.7	53.9	2.5
Years w/extension	-6.3	26.6	0.0	20.2	0.9
Farm size			16.2	16.2	0.7
Distance	-0.4	0.9	5.9	6.4	0.3
Visit frequency	-0.7	6.2	0.0	5.4	0.2
Labor Staples	3.4			3.4	0.2
Labor Cash Crops		1.1		1.1	0.0
Expenditures Cash Crops		0.9		0.9	0.0
Expenditures Staples	0.7			0.7	0.0
Credit	0.0	0.0	0.0	0.1	0.0
Contact Farm	-24.5	93.8	-89.4	-20.2	-0.9
Family size	-8.8	-21.3	2.5	-27.6	-1.3
Experience	8.6	-32.6	-25.7	-49.6	-2.3
Social	1.4	-6.3	-48.6	-53.6	-2.4
Tenure	-9.7	-37.8	-76.5	-124.1	-5.7
Market access	11.4	-46.2	-122.7	-157.6	-7.2
Structures	9.9	-194.1		-184.3	-8.4
PAES1	-298.8	2284.6	144.6	2130.5	97.1
PAES2	172.0	1202.4	369.5	1743.9	79.5
PAES3	-13.6	1502.9	153.8	1643.1	74.9
Predicted Household Income	807.1	811.7	576.0	2194.7	

 Table 4. Marginal Effects on Income-Generating Activities (US \$)

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