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14. March 2011

Online at http://mpra.ub.uni-muenchen.de/35166/ MPRA Paper No. 35166, posted 03. December 2011 / 13:09

# A Partial Differential Equation to Express a Business Cycle: Implications of Japan's Low Interest Rate Policy 

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#### Abstract

This study presents an equation of income derived from the Keynesian IS curve and the consumption Euler equation that explains the business cycle. Drawing on multi-period data from Japan, the model confirms the conventional wisdom that the appropriate policy response to an inflationary gap is to increase the interest rate when economic growth accelerates and decrease it when growth decelerates. However, the model indicates that to stabilize a deflationary gap, policymakers should decrease the interest rate when growth accelerates and increase it when growth decelerates. This prescription defies generations of conventional wisdom but fits the historical data remarkably well.


(JEL Nos. E32, E52)

Keywords:

1. Consumption Euler equation
2. Business Cycle
3. Capital
4. Interest rate
5. Elasticity of intertemporal substitution
6. Dynamic optimization problem
7. Variable separation method
8. Income elasticity of capital
9. Japan
10. Financial policy

Introduction
Thirty years ago, businesspeople and policymakers worldwide might have been reading a book to answer the question How can we become like Japan? Presently, they are more likely to ask, "How can we avoid becoming like Japan?" This study answers the latter question and argues that Japan's adoption of a low interest rate policy following the burst of the bubble economy has been the major reason for its inability to escape deflation.

Policymakers' doctrinaire response to a deflationary gap is to reduce interest rates in hopes of stimulating investment and escaping stagnation. Arguably, this response is based on a fundamental misconception. During a period of deflation, an economy inherently must suffer from productive overcapacity. If so, should you not discourage investment? When investment brings a higher interest rate, I presume that you achieve a higher growth rate.

What is needed is a reconsideration of previous equations of national income and theories of the business cycle. This paper proposes the following equation and discusses its implications:
$\frac{\frac{d^{2} Y}{d t^{2}}}{\frac{d r}{d t}}=\alpha$ (constant) $Y$ (income), $r$ (interest rate), $t$ (time)
Section 2 explains how this equation is derived. Section 3 organizes assumptions of this theory. Section 4 solves the equation, and Section 5 calculates the value of $\alpha$. Section 6 provides statistical proof, and Section 7 considers implications of the theory for financial policy. Section 8 presents the conclusion.

## 2. Derivation of the Equation

This section presents two ways of deriving Eq. (1-1): using the Old Keynesian IS curve and the consumption Euler equation familiar to New Keynesians. We begin with the Old Keynesian IS curve.

First, assume that the macro-economy model of Japan resembles the following, which includes credible figures for ease of illustration:
(trillion yen, interest rate is expressed in percent):

Consumption function $C=30+0.8 Y \ldots$
Investment function $I=75-3 r \ldots$
Income balance equation $Y=C+I \ldots$
We derive the IS curve from the following assumptions:
Substituting (2-1) and (2-2) into (2-3):

$$
\begin{align*}
& Y=30+0.8 Y+75-3 r \\
& 0.2 Y=105-3 r \\
& Y=525-15 r \ldots \tag{2-4}
\end{align*}
$$

Eq. (2-4) is the Old Keynesian IS curve. The IS curve can be expressed as a deferential equation as below. Differentiating both sides of (2-4) with respect to $r$,
$\frac{d Y}{d r}=-15$
when $r=0, Y_{0}=525$.
It is more generally expressed as the following equation:
$\frac{d Y}{d r}=\chi($ constant $) \ldots$
This means that IS curves can be rewritten universally as equations of the type expressed in (25).

We differentiate (2-5) with respect to $t$ and express it as $\alpha$ (constant) and find $\alpha=0$ in this case. In other words:
$\frac{d^{2} Y}{d r d t}=\chi^{\prime}=\alpha$ (constant) $\ldots$
Solving (2-6) as a partial differential equation and calculating the value of $\alpha$ yields $\alpha=0$.

Rewriting the equation for ease of solution:
$\frac{\frac{d^{2} Y}{d t^{2}}}{\frac{d r}{d t}}=\alpha$ (constant)
Next, we consider the derivation of Eq. (1-1) based on micro-economic theory consistent with ${ }^{i}$ New Keynesian ideas.
Household utility is defined below. First,

$$
\begin{equation*}
U\left(c_{t}, m_{t}\right) \equiv \frac{c_{t}^{1-\theta}}{1-\theta}+\frac{m_{t}^{1-\gamma}}{1-\gamma}, \tag{2-8}
\end{equation*}
$$

where U is household utility, $c_{t}$ is consumption in $t$, and $m_{t}$ is the quantity of money
held by households. Then utility is $\frac{m_{t}^{1-\gamma}}{1-\gamma}$.
There are two parameters, $\theta$ and $\gamma . \theta$ expresses relative risk aversion ${ }^{\mathrm{ii}}$ and is defined in (2-9)

$$
\begin{equation*}
\theta \equiv-\frac{U_{c c} C_{t}}{U_{c}}>0 \ldots \tag{2-9}
\end{equation*}
$$

$U_{c}$ and $U_{c c}$ are the first- and second-order derivatives, respectively, of U with respect to C .
The reciprocal is the elasticity of intertemporal substitution:
$\frac{1}{\theta} \equiv-\frac{U_{c}}{U_{c c} C_{\mathrm{t}}}>0 \ldots$
This definition plays an important role later.
The limiting condition is
$\left(1+\pi_{t}\right)\left(m_{t}+\mathrm{B}_{t}+C_{t}\right)=\left(1+i_{t-1}\right) \mathrm{B}_{t-1}+m_{t-1}$
$\mathrm{B}_{-1}, m_{-1}$
given that $\pi_{t}$ is the inflation rate in $t, \mathrm{~B}_{t}$ is a bond with a one-year maturity at $t$, and $i$ is the interest rate.

Then,
$a_{t} \equiv \mathrm{~B}_{t}+m_{t}$,
where $a_{t}$ is the real net asset in $t$
The limiting conditions become
$\left(a_{t}+C_{t}\right)=\frac{1}{1+\pi_{t}}\left(\left(1+i_{t-1}\right) a_{t-1}-\left(1+i_{t-1}\right) m_{t-1}+m_{t-1}\right)$
and
$\left(a_{t}+C_{t}\right)=\frac{1}{1+\pi_{t}}\left(\left(1+i_{t-1}\right) a_{t-1}-i_{t-1} m_{t-1}\right)$.
This is a dynamic optimization problem.
$\max : \mathrm{E}_{0} \sum_{\mathrm{t}=0}^{\infty} \beta^{\mathrm{t}} U\left(C_{t}, m_{t}\right)=E_{0} \sum_{t=0}^{\infty} \beta^{t}\left(\frac{C_{t}^{1-\theta}}{1-\theta}+\frac{m_{t}^{1-\gamma}}{1-\gamma}\right)$
Then the Lagrangian is
$\Gamma_{\mathrm{o}}=E_{0} \sum_{t=0}^{\infty} \beta^{t}\left(\frac{C_{t}^{1-\theta}}{1-\theta}+\frac{m_{t}^{1-\gamma}}{1-\gamma}-\phi_{t}\left(\left(a_{t}+C_{t}\right)-\frac{1}{1+\pi_{t}}\left((1+i) a_{t-1}-i_{t-1} m_{t-1}\right)\right)\right)$.
$\phi_{t}$ is a Lagrange coefficient.
The first-order optimization conditions are
$\mathrm{C}_{t}^{-\theta}=\phi_{t}$
$\mathrm{m}_{t}^{-\gamma}=E_{t} \frac{\beta i_{t}}{1+\pi_{t+1}} \phi_{t+1}$
$\phi_{t}=E_{t} \frac{\beta\left(1+i_{t}\right)}{1+\pi_{t+1}} \phi_{t+1}$
Thus, we obtain the following equation:
$C_{t}^{-\theta}=E_{t} \frac{\beta\left(1+i_{t}\right)}{1+\pi_{t+1}} C_{t+1}^{-\theta} \cdots$
This is the consumption Euler equation, which shows the relationship between the current and the following years' consumption.

Let us assume perfect foresight and that the subjective discount factor is 1 . Then $C_{t}^{-\theta}=\frac{\left(1+i_{t}\right)}{1+\pi_{t+1}} C_{t+!}^{-\theta}$.
Taking log of both sides, we get
$i n C_{t}-i n C_{t+1}=-\frac{1}{\theta}\left(\operatorname{in}\left(1+\mathrm{i}_{t}\right)-\operatorname{in}\left(1+\pi_{t+1}\right)\right)$.
So we assume the following relationships:
$i_{t} \cong$ in $\left(1+\mathrm{i}_{t}\right)$
$\pi_{t+1} \cong \mathrm{in}\left(1+\pi_{t+1}\right)$

Then
$i n C_{t}-i n C_{t+1}=-\frac{1}{\theta}\left(i_{t}-\pi_{t+1}\right)$
$\frac{d C_{t} / d t}{C_{t}}=\frac{1}{\theta}\left(i_{t}-\pi_{t+1}\right) \ldots$

New Keynesians construct their models on the bold assumption that $\mathrm{Y}=\mathrm{C}$. Because their models do not include capital, their assumption does not fit my model. Therefore, I assume propensity to consume is constant (c).
$\frac{C_{t}}{Y_{t}}=c$
If c is constant, by differentiating c with respect to $t$, we find that the value of $c^{\prime}$ is 0 .
$c^{\prime}=\frac{Y_{t}^{d C_{t} / d t-C_{t}^{d Y_{t}} / d t}}{Y_{t}^{2}}=0$.
Then

$$
\begin{equation*}
\frac{d Y_{t} / d t}{Y_{t}}=\frac{d C_{t} / d t}{C_{t}} \ldots \tag{2-13}
\end{equation*}
$$

Substitute (2-13) into (2-12)

$$
\begin{equation*}
\frac{d Y_{t} / d t}{Y_{t}}=\frac{1}{\theta}\left(i_{t}-\pi_{t+1}\right) \ldots \tag{2-14}
\end{equation*}
$$

Then we substitute (2-10) into $\theta$ :

$$
\frac{d Y_{t} / d t}{Y_{t}}=-\frac{U_{c}}{U_{c c} C_{t}}\left(i_{t}-\pi_{t+1}\right)
$$

$$
\frac{d Y_{t}}{d t}=-\frac{U_{c} Y_{t}}{U_{c c} C_{t}}\left(i_{t}-\pi_{t+1}\right)=-\frac{U_{c}}{U_{c c} c}\left(i_{t}-\pi_{t+1}\right)
$$

We define the real interest rate as $r_{t}=i_{t}-\pi_{t+1}$. Then

$$
\begin{equation*}
\frac{d Y_{t}}{d t}=-\frac{U_{c}}{U_{c c} c} r_{t} \ldots \tag{2-15}
\end{equation*}
$$

Differentiate both sides of (2-15) with respect to $r$ :

$$
\begin{equation*}
\frac{d^{2} Y_{t}}{d t d r_{t}}=-\frac{U_{c}}{U_{c c} c}=\alpha \ldots \tag{2-16}
\end{equation*}
$$

We assume $\alpha$ to be constant. Then we rewrite (2-16) as

$$
\begin{equation*}
\frac{\frac{d^{2} Y}{d t^{2}}}{\frac{d r}{d t}}=\alpha \ldots \tag{2-17}
\end{equation*}
$$

These are the two derivations of the equation shown.

## 3. Method of Analysis

Let us confirm the assumptions before solving the equation. The most important assumption is that the interest rate equals the marginal product of capital. Other assumptions that support this fact include perfect competition and the profit maximization principle.

We also assume Y is differentiable twice and that $r$ is differentiable once with respect to $t$ and that capital elasticity of real income is constant. The final assumption is that the income function is the product of the function of $t$ and the function of $k$. The equation has only two variables, $t$ (time) and $k$ (capital stock). Although it may be said that a two-variable model is too simple to explain business cycles, I suggest that the process by which the equation is derived confirms its explanatory power.

I believe it is unique to explain business cycles by only one parameter, but the explanation is robust because the parameters that explain economic phenomena influence each other, and these parameters can be aggregated into only one parameter (see statistical proof in Section 6).
4. Solving the Equation

When we solve this partial differential equation using a variable separation method, we automatically hypothesize the following function:

$$
\begin{equation*}
\mathrm{Y}=\mathrm{Y}(\mathrm{t}, \mathrm{k})=\mathrm{P}(\mathrm{t}) \mathrm{Q}(\mathrm{k}) \ldots \tag{4-1}
\end{equation*}
$$

Therefore, Y is a function of t (time) and k (capital), which is a product of $\mathrm{P}(\mathrm{t})$ (price) and Q (k) (quantity).

First, let us differentiate Y with respect to t once,

$$
\frac{d Y}{d t}=P^{\prime}(t) Q(k)+P(t) Q^{\prime}(k) .
$$

Then we differentiate Y a second time.

$$
\begin{align*}
& \frac{d^{2} Y}{d t^{2}}=P^{\prime \prime}(t) Q(k)+P^{\prime}(t) Q^{\prime}(k)+P^{\prime}(t) Q^{\prime}(k)+P(t) Q^{\prime \prime}(k) \\
& =P^{\prime \prime}(t) Q(k)+2 P^{\prime}(t) Q^{\prime}(k)+P(t) Q^{\prime \prime}(k) \ldots  \tag{4-2}\\
& =\frac{\partial^{2} P}{\partial t^{2}} Q(k)+2 \frac{\partial P}{\partial t} \frac{\partial Q}{\partial k} \frac{d k}{d t}+P(t) \frac{\partial^{2} Q}{\partial k^{2}}\left(\frac{d k}{d t}\right)^{2} \cdots \tag{4-3}
\end{align*}
$$

The interest rate is equal to the marginal product of capital $\left(r=\frac{d Y}{d k}=P(t) \frac{\partial Q}{\partial k}\right)$. Therefore,

$$
\begin{equation*}
\frac{d r}{d k}=d \frac{d Y}{d k d k}=\frac{d^{2} Y}{d k^{2}}=P(t) \frac{\partial^{2} Q}{\partial k^{2}} \ldots \tag{4-4}
\end{equation*}
$$

When we differentiate $r=\frac{d Y}{d k}=P(t) \frac{\partial Q}{\partial k}$ with respect to t , we get the denominator of (1-1) as follows:

$$
\begin{equation*}
\frac{d r}{d t}=\frac{\partial P}{\partial t} \frac{\partial Q}{\partial k}+P(t) \frac{\partial^{2} Q}{\partial k^{2}} \frac{d k}{d t} \ldots \tag{4-5}
\end{equation*}
$$

From (4-3) and (4-5), (1-1) becomes

$$
\begin{aligned}
& \frac{\partial^{2} P}{\partial t^{2}} Q(k)+2 \frac{\partial P}{\partial t} \frac{\partial Q}{\partial k} \frac{d k}{d t}+P(t) \frac{\partial^{2} Q}{\partial k^{2}}\left(\frac{d k}{d t}\right)^{2}=\alpha\left(\frac{\partial P}{\partial t} \frac{\partial Q}{\partial k}+P(t) \frac{\partial^{2} Q}{\partial k^{2}} \frac{d k}{d t}\right) \\
& \frac{\partial^{2} P}{\partial t^{2}} Q(k)+2 \frac{\partial P}{\partial t} \frac{\partial Q}{\partial k} \frac{d k}{d t}-\alpha \frac{\partial P}{\partial t} \frac{\partial Q}{\partial k}=\alpha P(t) \frac{\partial^{2} Q}{\partial k^{2}} \frac{d k}{d t}-P(t) \frac{\partial^{2} Q}{\partial k^{2}}\left(\frac{d k}{d t}\right)^{2} \\
& \frac{\partial^{2} P}{\partial t^{2}} Q(k)+\frac{\partial P}{\partial t} \frac{\partial Q}{\partial k}\left(2 \frac{d k}{d t}-\alpha\right)=P(t) \frac{\partial^{2} Q}{\partial k^{2}} \frac{d k}{d t}\left(\alpha-\frac{d k}{d t}\right) . \\
& \text { If } P(t)=P_{o} e^{\lambda t}, \text { so } \frac{\partial^{2} P}{\partial t^{2}}=\lambda \frac{\partial P}{\partial t} .
\end{aligned}
$$

Furthermore, we substitute $\frac{d k}{d t}$ for $b$.
Then

$$
\begin{align*}
& \frac{\partial P}{\partial t}\left(\lambda Q(k)+\frac{\partial Q}{\partial k}(2 b-\alpha)\right)=P(t) \frac{\partial^{2} Q}{\partial k^{2}} b(\alpha-b) \\
& \frac{\frac{\partial P}{\partial t}}{P(t)}=\frac{\frac{\partial^{2} Q}{\partial k^{2}} b(\alpha-b)}{\frac{\partial Q}{\partial k}(2 b-\alpha)+\lambda Q(k)}=\lambda(\text { Constant }) \ldots \tag{4-6}
\end{align*}
$$

Therefore,

$$
\begin{align*}
& \frac{\partial P}{\partial t}=\lambda P(t) \ldots  \tag{4-7}\\
& \frac{\partial^{2} Q}{\partial k^{2}} b(\alpha-b)=\lambda \frac{\partial Q}{\partial k}(2 b-\alpha)+\lambda^{2} Q(k) \ldots \tag{4-8}
\end{align*}
$$

From (4-7) we derive

$$
\begin{equation*}
P(t)=P_{o} e^{\lambda t} \ldots \tag{4-9}
\end{equation*}
$$

From (4-8) we derive

$$
b(\alpha-b) \frac{\partial^{2} Q}{\partial k^{2}}-\lambda(2 b-\alpha) \frac{\partial Q}{\partial k}-\lambda^{2} Q(k)=0 .
$$

Therefore,

$$
b(\alpha-b) Q^{\prime \prime}-\lambda(2 b-\alpha) Q^{\prime}-\lambda^{2} Q=0
$$

This is a second-order linear differential equation. The characteristic equation is therefore $b(\alpha-b) \zeta^{2}-\lambda(2 b-\alpha) \zeta-\lambda^{2}=0$.
The basic solution is determined by

$$
\begin{align*}
& \zeta_{+-}=\frac{\lambda(2 b-\alpha) \pm \sqrt{\lambda^{2}(2 b-\alpha)^{2}+4 \lambda^{2} b(\alpha-b)}}{2 b(\alpha-b)}=\frac{\lambda(2 b-\alpha) \pm \lambda \alpha}{2 b(\alpha-b)} \ldots  \tag{4-10}\\
& \zeta_{+}=\frac{\lambda}{\alpha-b} \ldots  \tag{4-11}\\
& \zeta_{-}=-\frac{\lambda}{b} \ldots \tag{4-12}
\end{align*}
$$

The content of the radical sign is

$$
\lambda^{2}(2 b-\alpha)^{2}+4 \lambda^{2} b(\alpha-b)=\lambda^{2}\left(4 b^{2}-4 \alpha b+\alpha^{2}+4 \alpha b-4 b^{2}\right)=\alpha^{2} \lambda^{2}
$$

The general solution follows.

$\begin{aligned} \text { If } \mathrm{D} & =\alpha \lambda>0 \\$\[

$$
\begin{equation*}
Q(k) \tag{4-13}
\end{equation*}
$$

\]$& =C_{1} \exp \left(\frac{\lambda}{\alpha-b} k\right)+C_{2} \exp \left(-\frac{\lambda}{b} k\right) \ldots\end{aligned}$

If $\mathrm{D}=\alpha \lambda=0$
$Q(k)=\left(C_{1}+C_{2} k\right) \exp \left(\frac{\lambda(2 b-\alpha)}{2 b(\alpha-b)} k\right) \ldots$
If $\mathrm{D}=\alpha \lambda<0$,
$Q(k)=\exp \left(\frac{\lambda(2 b-\alpha)}{2 b(\alpha-b)} k\right)\left(C_{1} \cos \left(\frac{\alpha \lambda}{2 b(\alpha-b)} k\right)+C_{2} \sin \left(\frac{\alpha \lambda}{2 b(\alpha-b)} k\right) \ldots\right.$
From (4-1), (4-9), (4-13), (4-14), and (4-15) we find the following:
When $\mathrm{D}=\alpha \lambda>0$

$$
\begin{equation*}
Y(t, k)=C_{1} \exp \left(\lambda t+\frac{\lambda}{\alpha-b} k\right) \ldots \tag{4-16}
\end{equation*}
$$

or

$$
\begin{equation*}
Y(t, k)=C_{1} \exp \left(\lambda t+\frac{-\lambda}{b} k\right) \ldots \tag{4-17}
\end{equation*}
$$

When $\mathrm{D}=\alpha \lambda=0$, it means that either $\alpha=0$ or $\lambda=0$.

Then if $\alpha=0$,

$$
\begin{align*}
& Y(t, k)=\left(C_{1}+C_{2} k\right) \exp \left(\frac{\lambda(2 b-\alpha)}{2 b(\alpha-b)} k\right) \ldots  \tag{4-18}\\
& Y(t, k)=\left(C_{1}+C_{2} k\right) \exp \left(\lambda t+\frac{-\lambda}{b} k\right) \ldots \tag{4-19}
\end{align*}
$$

If $\lambda=0$, from (4-9) and (4-19), we find

$$
\mathrm{Y}(\mathrm{t}, \mathrm{k})=P_{1}\left(C_{2}+C_{3} k\right) .
$$

Thus, $Y$ is a primary function of $k$.

If $\mathrm{D}=\alpha \lambda<0$,

$$
\begin{equation*}
Y(t, k)=P_{1} \exp \left(\lambda t+\frac{\lambda(2 b-\alpha)}{2 b(\alpha-b)} k\right)\left(C_{2} \cos \left(\frac{\alpha \lambda}{2 b(\alpha-b)} k\right)+C_{3} \sin \left(\frac{\alpha \lambda}{2 b(\alpha-b)} k\right) \ldots\right. \tag{4-20}
\end{equation*}
$$

5. Value of $\alpha$

In this section, we consider the value of $\alpha$.
From (4-1),

$$
\begin{equation*}
Y(t, \mathrm{k})=\mathrm{P}(\mathrm{t}) \mathrm{Q}(\mathrm{k}) \ldots \tag{4-1}
\end{equation*}
$$

So from (4-9) and (4-10), we can divide (4-1) into two functions, (5-1) and (5-2), as below.

$$
\begin{gather*}
P(t)=P_{o} e^{\lambda t} \ldots  \tag{5-1}\\
Q(k)=B \exp \left(\frac{\lambda(2 b-\alpha) \pm \alpha \lambda}{2 b(\alpha-b)} k\right)=B e^{\mu k} \ldots  \tag{5-2}\\
: \mu=\frac{\lambda(2 b-\alpha) \pm \alpha \lambda}{2 b(\alpha-b)}
\end{gather*}
$$

$\alpha$ is the quotient of (4-3) divided by (4-4). Therefore, $\alpha$ is
$\alpha=\frac{\frac{d^{2} Y}{d t^{2}}}{\frac{d r}{d t}}=\frac{P_{0} e^{\lambda t}}{P_{0} e^{\lambda t}} \frac{B e^{\mu k}}{B e^{\mu k}} \frac{\left(\lambda^{2}+2 \lambda \mu \frac{d k}{d t}+\mu^{2}\left(\frac{d k}{d t}\right)^{2}\right)}{\left(\lambda \mu+\mu^{2} \frac{d k}{d t}\right)}$
$\alpha=\frac{\frac{d^{2} Y}{d t^{2}}}{\frac{d r}{d t}}=\frac{\left(\lambda+\mu\left(\frac{d k}{d t}\right)\right)^{2}}{\mu\left(\lambda+\mu \frac{d k}{d t}\right)}=\frac{\left(\lambda+\mu \frac{d k}{d t}\right)}{\mu} \ldots$
If
$\mu=\frac{-\lambda}{b} \ldots$ (5-4),
then (5-3) is represented below.
$\alpha=\frac{\frac{d^{2} Y}{d t^{2}}}{\frac{d r}{d t}}=\frac{\left(\lambda+\mu \frac{d k}{d t}\right)}{\mu}=\frac{\left(\lambda+\frac{-\lambda}{b} b\right)}{\frac{-\lambda}{b}}=\frac{(\lambda-\lambda)}{\frac{-\lambda}{b}}=0 \cdots \cdot$
In Section 6, the statistical analysis shows that Japan has an equation of income such as in (4-17).

$$
\begin{equation*}
Y(t, k)=C_{1} \exp \left(\lambda t+\frac{-\lambda}{b} k\right) \ldots \tag{4-17}
\end{equation*}
$$

Eq. (5-4) indicates that $\alpha=0$. This implies that if $\alpha$ does not exist, the value of its mean is 0 , and that of its variance is huge. I calculate the actual value in Section 6. Indeed, the variance is large, but its mean does not equal 0 . Thus, I conclude that $\alpha$ exists. This means that when the correct investment is implemented, the acceleration of income and the fluctuation of the interest rate will cease, and the economy will have reached a new equilibrium.

Therefore, in the case where $\mathrm{D}=\alpha \lambda=0$, Eq. (4-19) shows equilibrium income at that time on the condition that $\alpha=0$.

$$
\begin{equation*}
Y^{*}(t, k)=\left(C_{1}+C_{2} k\right) \exp \left(\lambda t+\frac{-\lambda}{b} k\right) \ldots \tag{4-19}
\end{equation*}
$$

## 6. Statistical Proof Using Data

In this section, we verify that the equation solved in Section 4 can be reliably applied to actual cases. First, we consider (4-17).

$$
\begin{equation*}
Y(t, k)=C_{1} \exp \left(\lambda t+\frac{-\lambda}{b} k\right) \ldots \tag{4-17}
\end{equation*}
$$

Putting both sides into a natural logarithm,

$$
\operatorname{in} Y(t, k)=B_{0}+\lambda t+\frac{-\lambda}{b} k \ldots(6-1)\left(B_{0}=\operatorname{in} C_{1}\right) .
$$

Equation (6-1) shows that the logarithm value of the nominal GDP can be expressed as a linear regression model of time $(t)$ and capital $(k)$. Therefore, we fix 1980 as a standard and conduct a regression analysis on the nominal GDP shown in Reference List 1 by $t$ (1980 as 1 ) and k . The result is as follows.

Coefficient of determination: 0.89216685
Adjusted coefficient of determination by degree of freedom: 0.883872

Table I Results of Analysis(1)

|  | Coefficient | Standard Error | $t$ | P-value |
| :--- | ---: | ---: | :--- | :--- |
| $B_{0}$ | 4.90583638 | 0.1397002 | 35.1169 | $1.9 \mathrm{E}-23$ |
| $B_{1}$ | -0.0499785 | 0.0133574 | -3.74164 | 0.00091 |
| $B_{2}$ | 0.00234751 | 0.0004156 | 5.648127 | $6.1 \mathrm{E}-06$ |

Thus, a logarithm value of nominal income can be expressed as Eq. (6-2).

$$
\begin{equation*}
\operatorname{in} Y(t, k)=4.90583638-0.0499785 t+0.00234751 k \ldots \tag{6-2}
\end{equation*}
$$

The coefficient of determination is 0.89 , and the adjusted coefficient is 0.88 , which are sufficiently high. The negative value of $\lambda$, however, is unexpected. Therefore, I must question whether the following formula holds true:

$$
\begin{equation*}
B_{2}=\frac{-\lambda}{b} \ldots \tag{6-3}
\end{equation*}
$$

The results of calculating $\frac{-\lambda}{b}$ are found in Reference List 2.
The value of the mean is 0.002172 , which is very close to the value of $B_{2}$. The value of the variance is $1.91532 \times 10^{-6}$. Before we examine whether both means are equal, we must examine whether both variances are equal. We calculate the standard deviation from the standard error. Standard deviation $=$ Standard error $\times$ square root of the number of data. So, $\sigma=$ 0.002238 , and the variance $\sigma^{2}=5.00961 \times 10^{-6}$. Table II below shows the variance with $\frac{-\lambda}{b}$.

Table II Variance with $\frac{-\lambda}{b}$

|  | Mean | Variance | Data |
| :--- | :--- | :--- | :--- |
| $B_{2}$ | 0.002347511 | $5.00961 \mathrm{E}-06$ | 29 |
| $-\lambda / b$ | 0.002172425 | $1.91352 \mathrm{E}-06$ | 29 |

We must seek the $z$ value to examine whether both variances are equal. The $z$ value is giveb below.

$$
z=\frac{(n-1) s^{2}}{\sigma^{2}}
$$

in which $n$ is a number of data, $s^{2}$ is variance of $\frac{-\lambda}{b}$, and $\sigma^{2}$ is the variance of $B_{2}$.
That Z value has a chi-square distribution with $\mathrm{n}-1$ degrees of freedom.
The answer is as follows:

$$
\mathrm{z}=10.6951 .
$$

Rejection region of 0.005 left side: $z \leq 12.46134$.
Rejection region of 0.005 right side: $z \geq 50.99338$.
Thus, we can reject the hypothesis that the values of variances are equivalent with a $1 \%$ hazard ratio.

To test the mean value between sets in which each has different variances, we use the fact that t has a Student's t -distribution with $v$ degrees of freedom.
The value of $t$ is given below.

$$
t=\frac{\frac{-\bar{\lambda}}{b}-B_{2}}{\sqrt{\left(\frac{\vartheta_{1}^{2}}{n_{1}}+\frac{\vartheta_{2}^{2}}{n_{2}}\right.}}
$$

Here, $\frac{-\bar{\lambda}}{b}$ is the mean of $\frac{-\lambda}{b}, \vartheta_{1}^{2}$ and $\vartheta_{2}^{2}$ are variances, and $n_{1}$ and $n_{2}$ are numbers of the data of $\frac{-\lambda}{b}$ and $B_{2}$, respectively.

The value of $v$ is given below.

$$
v=\frac{\left(\frac{\vartheta_{1}^{2}}{n_{1}}+\frac{\vartheta_{2}^{2}}{n_{2}}\right)^{2}}{\left(\frac{\vartheta_{1}^{2}}{n_{1}}\right)^{2}} \frac{\left(\frac{\vartheta_{2}^{2}}{n_{2}}\right)^{2}}{n_{1}-1}+\frac{n_{2}-1}{n_{2}}
$$

The answers are below.
$\mathrm{T}=0.358342321$
$v=47$
Rejection region of $1 \%$ on both sides: $|t| \leq 2.945630052$.
Thus, we cannot reject the hypothesis that both means are equivalent.
Next, we calculate the actual value of $\alpha$. Reference List 3 shows the process of the calculation. The results follow below.

Table III
Calculation from Reference List 3

| Year | $\alpha$ |
| :---: | :---: |
| 1980 |  |
| 1981 |  |
| 1982 | -253.979 |
| 1983 | -306.041 |
| 1984 | -829.751 |
| 1985 | -1731.12 |
| 1986 | -1273.8 |
| 1987 | 247.5311 |
| 1988 | -3185.95 |
| 1989 | -214.671 |
| 1990 | 284.1922 |
| 1991 | 1069.932 |
| 1992 | -8213.74 |
| 1993 | 1024.186 |
| 1994 | 1033.595 |
| 1995 | -1493.78 |
| 1996 | -497.108 |
| 1997 | -34.6569 |


| 1998 | -1863.9 |
| :---: | :---: |
| 1999 | 441.8159 |
| 2000 | 4339.517 |
| 2001 | 2424.553 |
| 2002 | -325.836 |
| 2003 | -560.585 |
| 2004 | -5948.15 |
| 2005 | -2381.07 |
| 2006 | 2456.599 |
| 2007 | 723.4678 |
| 2008 | 1270.679 |

Basic statistical values

|  | $\alpha$ |
| :---: | :---: |
| MEAN | -511.039 |
| ST ERROR | 477.3842 |
| MEDIAN | -253.979 |
| MODE | \#N/A |
| ST |  |
| DEVIATION | 2480.561 |
| VARIANCE | 6153182 |
| KURTOSIS | 3.277517 |
| SKEWNESS | -1.25955 |

RANGE 12553.25
MIN $\quad-8213.74$
MAX 4339.517
TOTAL -13798.1
DATA 27
C-INTER
981.2772
(95.0\%)

The mean of $\alpha$ is -511 (trillion yen) ${ }^{2}$, which is very close to the actual GDP. Hence, $\alpha$ must be negative to maintain stable growth because $\lambda$ is negative.
The standard deviation (ST DEVIATION) is 2480.561 (trillion yen), which is quite large. Its value divided by 10 trillion is very close to $\frac{-\lambda}{b}$.

The skewness is negative. The mass of the distribution is concentrated on the right side of the figure. We find a consecutive series of negative figures from 1982 to 1989 and from 2002 to 2005 . We conduct a regression analysis by selecting only the years having negative $\alpha$. Then we examine whether the capability of this model becomes stronger. The following are the results:

The coefficient of determination increases to 0.93 from 0.89 .
The one adjusted by degree of freedom grows to 0.92 from 0.88 . We can confirm that this model is best applied to a condition with negative $\alpha$.

Coefficient of determination: 0.934773 .
Coefficient of determination adjusted by degree of freedom: 0.924738 .

Table IV Results of analysis (2)

|  | Coefficient | Standard | t | P-Value |
| :--- | :--- | :--- | :--- | :--- |
|  |  | deviation |  |  |
| $B_{0}$ | 4.909718 | 0.1489163 | 32.96964 | $6.46 \mathrm{E}-14$ |
| $B_{1}$ | -0.05517 | 0.0158882 | -3.47251 | 0.004126 |
| $B_{2}$ | 0.002447 | 0.0004736 | 5.166249 | 0.000181 |

Next, we verify that (4-19) can be reliably applied to an actual case. Eq. (4-19) calculates real income based on the assumption that real income is the equilibrium income on the condition that $\alpha=0$. I confirmed that this income is the real GDP, $Y^{*}(t, k)$.
$Y^{*}(t, k)=\left(C_{1}+C_{2} k\right) \exp \left(\lambda t+\frac{-\lambda}{b} k\right) \ldots$
Then $C_{1}+C_{2} k=\left(\frac{\mathrm{C}_{1}+C_{2} \mathrm{k}}{\mathrm{C}_{2} \mathrm{k}}\right) \mathrm{C}_{2} \mathrm{k}$.
If we define $C_{3}=\frac{C_{1}+C_{2} k}{C_{2} k}$,
$C_{1}+C_{2} k=C_{3} \mathrm{C}_{2} \mathrm{k}$.
So we rewrite $\lambda$ as $\lambda_{1}$ in (4-19)
Thus, (4-19) becomes (6-4).
$Y^{*}(t, k)=C_{3}\left(C_{2} k\right)^{2} \exp \left(\lambda_{1} t+\frac{-\lambda_{1}}{b} k\right) \ldots$
We assume real income is expressed as a well-known function, as in (6-5) below, whose capital elasticity of income is $a$.
$Y^{*}(k)=A k^{a} \ldots$
(6-5) (A: constant)

Hence, we can expect that $C_{2}^{2}$ is the following function:

$$
\begin{align*}
& C_{2}^{2}=C_{4}(\exp \eta) k^{a-2-\frac{\eta}{i n k} \cdots}  \tag{6-6}\\
& \eta=\frac{-\lambda_{1}}{b} k
\end{align*}
$$

We insert that function into (6-4).

$$
\begin{align*}
& Y^{*}(t, k)=C_{3} C_{4} k^{a-\frac{\eta}{i n k}}(\exp \eta)\left(\exp \lambda_{1} t+\frac{-\lambda_{1}}{b} k\right) \\
& Y^{*}(t, k)=C_{3} C_{4} k^{a-\frac{\eta}{i n k}} \exp \left(\lambda_{1} t+\frac{-2 \lambda_{1}}{b} k\right) \ldots \tag{6-5}
\end{align*}
$$

Putting both sides into the logarithm, we get:

$$
\begin{equation*}
i n Y^{*}(t, k)=i n C_{3}+i n C_{4} i n k k^{a-\frac{\eta}{i n k}}+\lambda_{1} t+\frac{-2 \lambda_{1}}{b} k \ldots \tag{6-6}
\end{equation*}
$$

We differentiate both sides with respect to $t$.

$$
\begin{aligned}
& \frac{d Y^{*} / d t}{Y^{*}}=a \frac{d k / d t}{k}-\eta^{\prime}+\lambda_{1}+\frac{-2 \lambda_{1}}{b} b \\
& \frac{d Y^{*} / d t}{Y^{*}}=a \frac{d k / d t}{k}-\frac{-\lambda_{1}}{b} b+\lambda_{1}+\frac{-2 \lambda_{1}}{b} b \\
& \frac{d Y^{*} / d t}{Y^{*}}=a \frac{d k / d t}{k}
\end{aligned}
$$

So we confirm that $a$ is the capital elasticity of income.
Additionally, (6-6) shows that in $Y^{*}$ can be brought into the regression analysis with respect to ink, $t$, and $k$. Therefore, we use the data of real GDP from Reference List 4, with the results that

Coefficient of determination: 0.989992444 .
Adjusted Coefficient of determination by degrees of freedom: 0.9887415 .

| Table V |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Results of Analysis (3) |  |  |  |
|  | Coefficient | St Error | t | P-value |
| $B_{0}$ | -0.51514704 | 0.480274 | -1.0726 | 0.2941 |
| $B_{1}$ | 1.119219197 | 0.090926 | 12.3091 | $7 \mathrm{E}-12$ |
| $B_{2}$ | 0.014294401 | 0.004325 | 3.30509 | 0.003 |
| $B_{3}$ | -0.00128167 | 0.000226 | -5.6828 | $7 \mathrm{E}-06$ |

That is, the logarithm of real income can be expressed as the following function:
in $Y^{*}(t, k)=-0.515154704+1.119219197$ ink $+0.014294401 t-0.00128167 k \ldots(6-7)$
The coefficient of determination is 0.98 , as is the adjusted coefficient. Thus, we can confirm the model works remarkably well.
Reference List 5 shows that $\frac{-2 \lambda_{1}}{b} \mathrm{~s}$. The value of the mean is $2 \frac{-\overline{\lambda_{1}}}{b}=-0.00117895$

This is very close to $B_{3}$. We test these values as we did earlier.

Table VI Variance with $\frac{-2 \lambda_{1}}{b}$

|  | MEAN | VARIANCE | DATA |
| :--- | :--- | ---: | ---: |
| B3 | -0.001282 | $1.4243 \mathrm{E}-06$ | 28 |
| $-2 \lambda_{1} / b$ | -0.001179 | $5.0555 \mathrm{E}-07$ | 27 |

First, we examine whether the variances are equal.
We calculate the $Z$ value as before: $Z=9.2287136$.
Rejection region of 0.005 left side: $Z \leq 11.167237$.
Therefore, we can reject the hypothesis that the values of variances are equivalent.
Then we calculate the $t$ value and $v$ degrees of freedom as before:

```
\[
t=0.3893929
\]
\[
v=44
\]
```

Rejection region of $1 \%$ both sides: $|t| \leq 2.995534$
Therefore, we cannot reject the hypothesis that both means are equivalent.

## 7. Implications for Financial Policy

For policymakers, the equation implies that the economy would not ascend on a stable course if the product of $\alpha$ and $\lambda$ is not positive. To maintain economic stability when $\lambda$ is positive, $\alpha$ should be positive; however, when $\lambda$ is negative, $\alpha$ should be negative. To sustain a positive value for $\alpha$, policymakers must increase the interest rate whenever income accelerates and reduce it whenever income decelerates. Yet, to sustain a negative value for $\alpha$, they must decrease the interest rate when income accelerates and increase the interest rate when income decelerates. In Japan's case, $\lambda$ is negative, so $\alpha$ should be negative.

Policy may be able to reduce the interest rate whenever there is acceleration, but it may seem impossible to increase the interest rate during deceleration. However, we may have an available strategy for such a policy. First, we assume that the current real growth rate is the equilibrium growth rate. Then, we decrease the interest rate and increase the real growth rate to be more than the equilibrium growth rate, followed by gradually raising the interest rate. We must continue to increase the interest rate until the real growth rate is equal to a new equilibrium growth rate. By this logic, we must quickly attempt to increase the interest rate when the real growth rate exceeds the current equilibrium growth rate.

If the economy shifts to a higher equilibrium as it continues to grow, the economy accelerates further. If so, when the economy reaches the new equilibrium, we will see the higher interest rate. Then we can decrease that interest rate to re-stimulate the economy. Thus, the interest rate will increase as the equilibrium growth rate rises.

We may increase much supply compared to demand on the condition that $\lambda$ is negative.

If we stimulate the economy by reducing the interest indiscreetly while maintaining a stable equilibrium growth rate, we will see smaller rates of both interest and growth. Currently, Japan's real growth rate is approximately $1 \%$ and its interest rate is approximately $1.5 \%$. Both may have reached their lower limits. Japan must soon seek an economic strategy for raising the equilibrium growth rate.

However, is there a way to increase income when interest rates rise? Generally speaking, we must reduce interest rates when we desire to increase the growth rate. However, Japan has pursued a low interest rate policy since the burst of the Bubble Economy, and the growth rate has not recovered. Why does the policy work not work? By way of an answer, I introduce the idea of production period ${ }^{\text {iii }}$ from Value and Capital by John Hicks, who wrote:

It follows at once from all this that if the average period of the stream of receipts is greater than the average period of the standard stream with which we are comparing it, a fall in the rate of interest will raise the capital value of the receipts stream more than that of the standard stream, and will therefore increase income. But if the average period of the stream of receipts is less than that of the standard stream, it is a rise in the rate of interest which will increase income.

To explain his statement briefly, let us assume that there exists the stream of inputs $R_{\mathrm{t}}\left(R_{0}, R_{1}, R_{2}, R_{3}, R_{4} \ldots\right)$ and the corresponding stream of earnings $\mathrm{S}_{\mathrm{t}}\left(\mathrm{S}_{0}, \mathrm{~S}_{1}, \mathrm{~S}_{2}, \mathrm{~S}_{3}, \mathrm{~S}_{4} \ldots\right)$.Then the present values of both discounted by interest rate are equal. So we find

$$
R_{0}+\frac{R_{1}}{(1+r)}+\frac{R_{2}}{(1+r)^{2}}+\ldots=\mathrm{S}_{0}+\frac{\mathrm{S}_{1}}{(1+\mathrm{r})}+\frac{\mathrm{S}_{2}}{(1+\mathrm{r})^{2}}+\ldots
$$

We calculate the capital elasticity of the interest rate as
$-\mathrm{A}_{s}=-\frac{\frac{\mathrm{S}_{1}}{(1+\mathrm{r})}+2 \frac{\mathrm{~S}_{2}}{(1+\mathrm{r})^{2}}+3 \frac{\mathrm{~S}_{3}}{(1+\mathrm{r})^{3}}+\cdots}{\mathrm{S}_{0}+\frac{\mathrm{S}_{1}}{(1+\mathrm{r})}+\frac{\mathrm{S}_{2}}{(1+\mathrm{r})^{2}}+\frac{\mathrm{S}_{3}}{(1+\mathrm{r})^{3}}+\cdots}$
This value is the average period of earnings weighted by the terms.
Then the average period of input is

$$
-\mathrm{A}_{R}=-\frac{\frac{R_{1}}{(1+\mathrm{r})}+2 \frac{R_{2}}{(1+\mathrm{r})^{2}}+3 \frac{\mathrm{R}_{3}}{(1+\mathrm{r})^{3}}+\cdots \cdot}{\mathrm{R}_{0}+\frac{\mathrm{R}_{1}}{(1+\mathrm{r})}+\frac{R_{2}}{(1+\mathrm{r})^{2}}+\frac{\mathrm{R}_{3}}{(1+\mathrm{r})^{3}}+\cdots \cdot}
$$

Then, if $\mathrm{A}_{s}>\mathrm{A}_{R}$, a decline in interest rates causes income to increase, and if $\mathrm{A}_{s}<\mathrm{A}_{R}$, an increase in the interest rate leads to an increase in income.

Now we calculate the average period when inputs or earnings are constant.
If

$$
R_{0}=R_{1}=R_{3}=R_{4}=\ldots
$$

or
$\mathrm{S}_{0}=\mathrm{S}_{1}=\mathrm{S}_{2}=\mathrm{S}_{3}=\mathrm{S}_{4}=\ldots$
So the average period of inputs or earnings resembles the following equation:

$$
\mathrm{A}=\frac{\frac{1}{(1+\mathrm{r})}+2 \frac{1}{(1+\mathrm{r})^{2}}+3 \frac{1}{(1+\mathrm{r})^{3}}+\ldots}{1+\frac{1}{(1+\mathrm{r})}+\frac{1}{(1+\mathrm{r})^{2}}+\frac{1}{(1+\mathrm{r})^{3}}+\ldots}=\frac{\frac{1}{(1+r)}}{\left(1-\frac{1}{(1+r)}\right)^{2}} / \frac{1}{\left(1-\frac{1}{(1+r)}\right)}=\frac{\frac{1}{(1+r)}}{1-\frac{1}{(1+r)}}=\frac{1}{r}
$$

That is equal to the reciprocal of the interest rate. If the stream has a crescendo, the average period of the stream is longer than the reciprocal of interest rate. If the stream has a decrescendo, its average period is shorter than the reciprocal of the interest rate. I think that the average period of inputs generally extends as the economy progresses, but that of earnings becomes longer; therefore, a fall in the interest rate causes an increase in income.

However, if we assume the stream of inputs is constant, then an average period of inputs is equal to the reciprocal of interest rate, and it is shorter than the average period of earnings. Then we find the equation

$$
\mathrm{A}_{s}<\mathrm{A}_{R}
$$

In that case, in general, the interest rate should rise to increase income. We must pay attention not to the absolute level of interest rate, but rather, to the magnitude of the relation between average periods of inputs and earnings.

If the interest rate is $1.5 \%$, the reciprocal is approximately 67 years. So to reduce interest rates further is unproductive if the stream of inputs is constant and the average period of earnings is less than 67 years. As Hicks demonstrated, one cannot expect income to increase by decreasing interest rates indiscreetly.

## 8. Conclusion

Nominal GDP can be expressed as the equation
If $\alpha \lambda>0$

$$
\begin{equation*}
Y(t, k)=C_{1} \exp \left(\lambda t+\frac{\lambda}{\alpha-b} k\right) \ldots \tag{4-16}
\end{equation*}
$$

or

$$
\begin{equation*}
Y(t, k)=C_{1} \exp \left(\lambda t+\frac{-\lambda}{b} k\right) \ldots \tag{4-17}
\end{equation*}
$$

If $\alpha \lambda<0$

$$
\begin{equation*}
Y(t, k)=P_{1} \exp \left(\lambda t+\frac{\lambda(2 b-\alpha)}{2 b(\alpha-b)} k\right)\left(C_{2} \cos \left(\frac{\alpha \lambda}{2 b(\alpha-b)} k\right)+C_{3} \sin \left(\frac{\alpha \lambda}{2 b(\alpha-b)} k\right) \ldots\right. \tag{4-20}
\end{equation*}
$$

If (4-16) can be applied, real income can be expressed as (6-8)

$$
\begin{aligned}
Y^{*}(t, k)=C_{3} C_{4} k^{a-\frac{\eta}{i n k}} \exp \left(\lambda t+\frac{-2 \lambda_{1}}{b} k\right) & \ldots \\
& \eta=\frac{-2 \lambda_{1}}{b} k \quad \text { (a: income elasticity of capital) }
\end{aligned}
$$

Statistical research shows Japan's GDP can be expressed by the following equations:

## Nominal income Y

$$
\begin{equation*}
\operatorname{in} Y(t, k)=4.90583638-0.0499785 t+0.00234751 k \ldots \tag{6-2}
\end{equation*}
$$

Real income $Y^{*}$

$$
\text { in } Y(t, k)=-0.515154704+1.119219197 \text { ink }+0.014294401 t-0.00128167 k \ldots \quad(6-10)
$$

From Section 4, we can show that a sign of $\alpha \lambda$ determines which type of equation can be applied:

If $\alpha \lambda>0$ :
If $\alpha \lambda<0$ :
The solution to this equation may provide a firm underpinning to economic policy by the following logic. Assuming that the trend of time is positive ( $\lambda>0$ ), policymakers must increase the interest rate whenever economic growth accelerates, and decrease it when growth decelerates. This action can lead the economy to a stable course because it preserves the condition that $\alpha>0$, as shown by (4-16) or (4-17). In other words, we must increase the interest rate rather than decrease it to keep acceleration positive whenever the real growth rate is lesser than the equilibrium growth rate.

However, if $\lambda<0$, the recommendation is reversed. We must decrease the interest rate whenever the economy accelerates and increase it in the event of deceleration. Thus, since Japan has negative $\lambda$, we may be able to decrease the interest rate whenever there is acceleration, but it may seem impossible to increase the interest rate in the event of deceleration.

Nonetheless, we may have a strategy available for such a policy. First, we assume the current real growth rate is the equilibrium growth rate. Then, we decrease the interest rate and increase the real growth rate making it more than the equilibrium growth rate, followed by gradually increasing the interest rate. We must continue to increase the interest rate until the real growth rate is equal to a new equilibrium growth rate. By this logic, we must quickly attempt to increase the interest rate when the real growth rate becomes higher than the current equilibrium growth rate.

If the economy shifts to a higher equilibrium as it continues to grow, the economy will accelerate further. If so, when the economy reaches the new equilibrium, we will see higher interest rates. Then we can reduce that interest rate to re-stimulate the economy. Thus, the interest rate will increase as the equilibrium growth rate rises.

We may increase much supply compared to demand on the condition that $\lambda$ is negative.
However, our condition that both $\lambda$ and $\alpha$ are negative to maintain economic stability can be considered a shrinking economy. If we stimulate the economy by reducing the interest rate indiscreetly while maintaining a stable equilibrium growth rate, we will achieve smaller rates of both interest and growth. Currently, Japan's real growth rate is approximately $1 \%$ and its interest rate is approximately $1.2 \%$. Both may have reached their lower limits. Japan's policymakers must soon seek an economic strategy for raising the equilibrium growth rate.

It is difficult to change basic beliefs about financial policy. But if policymakers hope to effect real-world change, they first must change their economic thought. John Maynard Keynes long ago noted the stubbornness of economic thought:

But besides this contemporary mood, the ideas of economists and political philosophers, both when they are right and when they are wrong, are more powerful than is commonly understood. Indeed the world is ruled by little else. Practical men, who believe themselves to be quite exempt from any intellectual influences, are usually the slaves of some defunct economist.

Japan has imposed a low interest rate policy for many years, and the result has been persistent deflation. If we maintain a low interest rate policy and do not fully consider the alternative presented in this research, we are acting unreasonably. Now is the time to try new theories and new economic thoughts.

## 9. References

## List 1

| Year | GDP <br> (Nominal) | Capital <br> Stock <br> (Begin) |
| :---: | :---: | :---: |
| 1980 | 242.8387 | 358.4012 |
| 1981 | 261.0682 | 382.292 |
| 1982 | 274.0866 | 405.8706 |
| 1983 | 285.0583 | 427.7703 |
| 1984 | 302.9749 | 447.2519 |
| 1985 | 325.4019 | 471.3132 |
| 1986 | 340.5595 | 524.3229 |
| 1987 | 354.1702 | 554.6213 |
| 1988 | 380.7429 | 595.5875 |
| 1989 | 410.1222 | 632.2497 |
| 1990 | 442.781 | 677.281 |
| 1991 | 469.4218 | 726.7462 |
| 1992 | 480.7828 | 786.1105 |
| 1993 | 483.7118 | 826.5722 |
| 1994 | 488.4503 | 857.0921 |
| 1995 | 495.1655 | 888.7079 |
| 1996 | 505.0118 | 918.2468 |
| 1997 | 515.6441 | 948.9018 |
| 1998 | 504.9054 | 980.756 |
| 1999 | 497.6286 | 1005.381 |


| 2000 | 502.9899 | 1026.533 |
| ---: | ---: | ---: |
| 2001 | 497.7197 | 1051.391 |
| 2002 | 491.3122 | 1068.489 |
| 2003 | 490.294 | 1082.417 |
| 2004 | 498.3284 | 1089.336 |
| 2005 | 501.7344 | 1120.987 |
| 2006 | 507.3648 | 1135.384 |
| 2007 | 515.5204 | 1162.551 |
| 2008 | 505.1119 | 1193.615 |

$¥$ Trillion

List 2

| YEAR | $-\lambda / \mathrm{b}$ |
| ---: | ---: |
| 1980 | 0.002092 |
| 1981 | 0.00212 |
| 1982 | 0.002282 |
| 1983 | 0.002565 |
| 1984 | 0.002077 |
| 1985 | 0.000943 |
| 1986 | 0.00165 |
| 1987 | 0.00122 |
| 1988 | 0.001363 |
| 1989 | 0.00111 |
| 1990 | 0.00101 |
| 1991 | 0.000842 |
| 1992 | 0.001235 |
| 1993 | 0.001638 |
| 1994 | 0.001581 |
| 1995 | 0.001692 |
| 1996 | 0.00163 |
| 1997 | 0.001569 |
| 1998 | 0.00203 |
| 1999 | 0.002363 |
| 10 |  |
| 10 |  |
| 10 |  |


| 2000 | 0.00201 |
| ---: | ---: |
| 2001 | 0.002923 |
| 2002 | 0.003588 |
| 2003 | 0.007224 |
| 2004 | 0.001579 |
| 2005 | 0.003472 |
| 2006 | 0.00184 |
| 2007 | 0.001609 |
| 2008 | 0.005744 |

## List 3

| Year | nGDP | $\Delta Y$ | $\Delta^{\wedge} 2 \mathrm{Y}$ | Prime rate | CPI | REAL | D | $\alpha$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1980 | 242.8387 |  |  | 9.163561644 | 0.077 | 0.014636 |  |  |
| 1981 | 261.0682 | 18.2295 |  | 8.636164384 | 0.049 | 0.037362 | 0.022726 |  |
| 1982 | 274.0866 | 13.0184 | -5.2111 | 8.587945205 | 0.028 | 0.057879 | 0.020518 | -253.979 |
| 1983 | 285.0583 | 10.9717 | -2.0467 | 8.356712329 | 0.019 | 0.064567 | 0.006688 | -306.041 |
| 1984 | 302.9749 | 17.9166 | 6.9449 | 7.919726027 | 0.023 | 0.056197 | -0.00837 | -829.751 |
| 1985 | 325.4019 | 22.427 | 4.5104 | 7.359178082 | 0.02 | 0.053592 | -0.00261 | -1731.12 |
| 1986 | 340.5595 | 15.1576 | -7.2694 | 6.529863014 | 0.006 | 0.059299 | 0.005707 | -1273.8 |
| 1987 | 354.1702 | 13.6107 | -1.5469 | 5.404931507 | 0.001 | 0.053049 | -0.00625 | 247.5311 |
| 1988 | 380.7429 | 26.5727 | 12.962 | 5.598082192 | 0.007 | 0.048981 | -0.00407 | -3185.95 |
| 1989 | 410.1222 | 29.3793 | 2.8066 | 5.890684932 | 0.023 | 0.035907 | -0.01307 | -214.671 |
| 1990 | 442.781 | 32.6588 | 3.2795 | 7.844657534 | 0.031 | 0.047447 | 0.01154 | 284.1922 |
| 1991 | 469.4218 | 26.6408 | -6.018 | 7.482191781 | 0.033 | 0.041822 | -0.00562 | 1069.932 |
| 1992 | 480.7828 | 11.361 | -15.2798 | 5.968219178 | 0.016 | 0.043682 | 0.00186 | -8213.74 |
| 1993 | 483.7118 | 2.929 | -8.432 | 4.844931507 | 0.013 | 0.035449 | -0.00823 | 1024.186 |
| 1994 | 488.4503 | 4.7385 | 1.8095 | 4.42 | 0.007 | 0.0372 | 0.001751 | 1033.595 |
| 1995 | 495.1655 | 6.7152 | 1.9767 | 3.487671233 | -0 | 0.035877 | -0.00132 | -1493.78 |
| 1996 | 505.0118 | 9.8463 | 3.1311 | 3.057808219 | 0.001 | 0.029578 | -0.0063 | -497.108 |
| 1997 | 515.6441 | 10.6323 | 0.786 | 2.489863014 | 0.018 | 0.006899 | -0.02268 | -34.6569 |
| 1998 | 504.9054 | -10.7387 | -21.371 | 2.436438356 | 0.006 | 0.018364 | 0.011466 | -1863.9 |
| 1999 | 497.6286 | -7.2768 | 3.4619 | 2.32 | -0 | 0.0262 | 0.007836 | 441.8159 |


| 2000 | 502.9899 | 5.3613 | 12.6381 | 2.211232877 | -0.01 | 0.029112 | 0.002912 | 4339.517 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2001 | 497.7197 | -5.2702 | -10.6315 | 1.772739726 | -0.01 | 0.024727 | -0.00438 | 2424.553 |
| 2002 | 491.3122 | -6.4075 | -1.1373 | 1.921780822 | -0.01 | 0.028218 | 0.00349 | -325.836 |
| 2003 | 490.294 | -1.0182 | 5.3893 | 1.560410959 | -0 | 0.018604 | -0.00961 | -560.585 |
| 2004 | 498.3284 | 8.0344 | 9.0526 | 1.708219178 | 0 | 0.017082 | -0.00152 | -5948.15 |
| 2005 | 501.7344 | 3.406 | -4.6284 | 1.60260274 | -0 | 0.019026 | 0.001944 | -2381.07 |
| 2006 | 507.3648 | 5.6304 | 2.2244 | 2.293150685 | 0.003 | 0.019932 | 0.000905 | 2456.599 |
| 2007 | 515.5204 | 8.1556 | 2.5252 | 2.342191781 | 0 | 0.023422 | 0.00349 | 723.4678 |
| 2008 | 505.1119 | -10.4085 | -18.5641 | 2.281232877 | 0.014 | 0.008812 | -0.01461 | 1270.679 |

* 1 Prime rates weighted averages by term
* 2 CPI stands for consumer price index
*3 Real is the real interest rate, that is, the difference between CPI and the prime rate
* 4 D is the difference in real interest rates

List 4

Real GDP

| YEAR | GDP * |
| :---: | :---: |
| 1980 | 284.375 |
| 1981 | 296.2529 |
| 1982 | 306.2562 |
| 1983 | 315.6299 |
| 1984 | 329.7193 |
| 1985 | 350.6016 |
| 1986 | 360.5274 |
| 1987 | 375.3358 |
| 1988 | 402.1599 |
| 1989 | 423.7565 |
| 1990 | 447.3699 |
| 1991 | 462.242 |
| 1992 | 466.0279 |
| 1993 | 466.8251 |
| 1994 | 470.8565 |
| 1995 | 479.7164 |
| 1996 | 492.3679 |
| 1997 | 500.0644 |
| 1998 | 489.8207 |
| 1999 | 489.13 |


| 2000 | 503.1198 |
| ---: | ---: |
| 2001 | 504.0475 |
| 2002 | 505.3694 |
| 2003 | 512.513 |
| 2004 | 526.5777 |
| 2005 | 536.7622 |
| 2006 | 547.7093 |
| 2007 | 560.8164 |

$¥$ Trillion

List 5

| YEAR | $-2 \lambda_{1} / b$ |
| :---: | :---: |
| 1980 |  |
| 1981 | -0.0012 |
| 1982 | -0.00121 |
| 1983 | -0.00131 |
| 1984 | -0.00147 |
| 1985 | -0.00119 |
| 1986 | -0.00054 |
| 1987 | -0.00094 |
| 1988 | -0.0007 |
| 1989 | -0.00078 |
| 1990 | -0.00063 |
| 1991 | -0.00058 |
| 1992 | -0.00048 |
| 1993 | -0.00071 |
| 1994 | -0.00094 |
| 1995 | -0.0009 |
| 1996 | -0.00097 |
| 1997 | -0.00093 |
| 1998 | -0.0009 |
| 1999 | -0.00116 |


| 2000 | -0.00135 |
| ---: | ---: |
| 2001 | -0.00115 |
| 2002 | -0.00167 |
| 2003 | -0.00205 |
| 2004 | -0.00413 |
| 2005 | -0.0009 |
| 2006 | -0.00199 |
| 2007 | -0.00105 |

〔1〕 Ryou Kato，Modern Macro Economic Lecture，53－57，Toyo Keizai Inc 2007.
〔2〕 Nobutyuki Oda and Murakami Jun，Natural Interest Rate，11－14 Bank of Japan Working Paper Series No $-\mathrm{J}-5$ ．

〔3〕John Hicks，Value and Capital，171－188，Oxford At The Clarendon Press． 1939.

