# Seasonal Production Smoothing 

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Working Paper 1999-004A
http://research.stlouisfed.org/wp/1999/99-004.pdf

PUBLISHED: Federal Reserve Bank of St. Louis Review, 81(5), September/October 1999, pp 21-40

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## Introduction

Inventory investment dynamics appear to dominate the economy's movement around its long-run path. Blinder (1981) noted that changes in inventory investment can explain 60 to 101 percent of changes in Gross National Product during post-war recessions. Because inventory fluctuations play such a major role in business cycles (and possibly seasonal fluctuations), it is important to understand the theoretical motivation for inventory holdings and the implied dynamics. The received view, established by Holt, Modigliani, Muth and Simon (1960), is that inventories are used to smooth production in the presence of increasing marginal cost (convex costs). An empirical stylized fact of inventory investment, however, is that production is more volatile than sales. The failure to confirm production smoothing empirically has been explained by inadequacies of the data or exceptions to the abstraction of convex costs.

Intuition suggests that even with convex costs, firms may not be likely to smooth production over periods longer than a year. Production horizons are likely to be shorter than a year and inventory holding costs may make it uneconomical to hold inventory for as long as a year. Many industries have well documented seasonal patterns in demand allowing them to plan production in concert with available capacity, required lead times, and labor market flexibility. In addition, evidence has been uncovered suggesting that seasonal fluctuations in output can also be affected by inventory/production decisions. For example, Carpenter and Levy (1998) use frequency domain analysis and find a large and statistically significant average squared coherence between inventory investment and the change in output in the manufacturing sector at both seasonal and business cycle
frequencies. It seems appropriate, therefore, to focus some attention on inventory decisions at seasonal frequencies.

In this paper we look for evidence of seasonal production smoothing in seasonally unadjusted, monthly data on manufacturing and retail inventories and sales. Using detrended, seasonally unadjusted data we find that the variance of production is less than the variance of sales for 23 out of 35 industries. The equivalent test using seasonally adjusted data found none with production varying less than sales. We interpret this as stronger evidence of production smoothing than found in previous literature.

The Fourier series of the inventory-to-sales (I/S) ratio of the industries with the lowest variance of production relative to sales revealed strong seasonal components (annual, six months and three months). A strong seasonal component in the I/S ratio suggests a possible negative seasonal correlation between sales and inventory and is an intuitive indication that smoothing occurs at higher frequencies. ${ }^{1}$ The results confirm Ghali's (1987) finding of seasonal smoothing using detrended, seasonally unadjusted data for the cement industry. The results also suggest that a model other than production smoothing may be more appropriate for explaining trend movements in production relative to sales.

## Background and Literature Survey

Holt, Modigliani, Muth and Simon (1960) established the analytical framework demonstrating that optimizing firms facing convex production costs and uncertain demand are motivated to smooth production and use inventories to buffer demand

[^0]shocks. If the marginal cost of production is increasing, then storing output during periods of low demand is prudent as long as storage costs are sufficiently low. (See shaded insert). Much of the research in inventory since Holt et al., has focussed on the efficacy of using the production-smoothing paradigm at the macroeconomic level. If firms use inventories to smooth production, then production should vary less than sales. Empirical testing of this hypothesis has yielded mixed results. Using a simple test of the ratio of the variance of production to the variance of sales, a majority of researchers have found a ratio greater than 1.0, contradicting the theory.

As a rule, data on production are not available. However, production can be readily estimated by adding current period sales to the change in inventory from last period. If production exceeds (is less than) sales in a given period then the difference must go to increasing (decreasing) inventories. This can be represented by the following equation:

$$
P_{t}=S_{t}+\Delta I_{t}
$$

where $P$ is production, $S$ is sales and $\Delta I$ is the change in inventory. This fundamental equation implies an important relationship among the variances and covariance of $P, S$ and $\Delta I$ :

$$
\operatorname{Var}(P)=\operatorname{Var}(S)+\operatorname{Var}(\Delta I)+2 \operatorname{Cov}(S, \Delta I)
$$

For the variance of production to be less than the variance of sales, the covariance of sales and the change in inventories, $\operatorname{Cov}(S, \Delta I)$, must be negative and greater in absolute value than half the variance of inventories.

Testing this covariance relationship, Miron and Zeldes (1988) find no support for production smoothing using both seasonally adjusted and unadjusted data after removing
an estimated linear trend from monthly data. Blinder (1986) also finds little empirical support for the basic production smoothing model. However, he identifies conditions under which the stylized facts could be compatible with production smoothing, to wit: if cost shocks are present, if firms see demand shocks before they make their production decisions, if demand shocks build before they decay, or if technological parameters dictate a rapid speed of adjustment.

If firms do not face convex production costs, production smoothing is not optimal, and Ramey (1991) finds indication of non-convex costs in some industries. Blinder and Maccini (1991) observe that wholesale and retail trade, and the materials and supplies portion of manufacturing inventory make up a large portion of total inventories and are likely to face nontrivial "quasi-fixed" cost of ordering. This type of cost structure makes an (S,s) inventory rule more economical. That is, firms will wait until inventory falls below a trigger point (s) then order sufficient stocks to raise inventory to an upper bound (S). This way the "quasi-fixed" costs are spread over a larger quantity. This behavior, sometimes called "bunching" will result in a higher volatility of production than sales. This leads Blinder and Maccini to conclude that the (S,s) paradigm is more consistent with the empirical evidence.

Another source of empirical failure may be the data. Ghali (1987) demonstrated that seasonal adjustment and aggregation will remove evidence of seasonal smoothing, and Lai (1991) shows that aggregation can distort the data sufficiently to negate production-smoothing tests. Some researchers, using disaggregated physical product data, find some support for production smoothing. Fair (1989), and Krane and Braun (1991) use disaggregated physical product data for the United States and confirm
smoothing in several industries, while Beason (1993) has similar success with Japanese data. Dimelis and Ghali (1994) detect statistically significant evidence of smoothing in disaggregated physical product data for three out of five industries, using the variance bounds test introduced by West (1986) and generalized by Kollintzas (1995).

Physical unit information is more appropriate for testing the implications of inventory management. It makes more sense, when discussing the motivation for holding inventory, to talk about the number of cars in stock than the value. Unfortunately, the most readily accessible data, particularly at an aggregate level, are the nominal values of inventory. One way of getting closer to physical quantities is to remove the effects of price changes. Finding the appropriate price deflator to convert the nominal values to real values is not always an easy task. And even when data are converted to remove price increases, trend growth in the real level of sales can also disguise smoothing. If sales are trending up (down), then production will also trend up (down). If firms smooth production annually and adjust the target level of smoothed production each year, then the variance induced by the trend growth will also distort the smoothing measure.

Miron (1996) finds noticeably less seasonal variation in price variables than quantity variables.

Seasonal movements in both real and nominal price variables are noticeably smaller than those in quantity variables. For example, the standard deviation of the seasonals in the growth rates of prices is 0.2 percent, and seasonal dummies explain only 3.1 percent of the total variation. The same conclusions hold qualitatively for nominal interest rates, real interest rates, nominal wages, and real wages. Miron (1996) page 18.

This observation means that if we remove the trend from seasonal unadjusted data, the high frequency movements are more likely to reflect changes in quantities. This provides
justification for the data transformation that we discuss in the next section.
\{Shaded Insert \}

## The Production Smoothing with Trend

Figure A illustrates the production smoothing motivation when increasing marginal

## Figure A



## Quantity

costs exist. If $\mathrm{Q}_{1}$ and $\mathrm{Q}_{2}$ represent the demand in periods 1 and 2 respectively, then the point A represents the average cost, $\left(C_{1}+C_{2}\right) / 2$, if $\mathrm{Q}_{1}$ is produced in period 1 and $\mathrm{Q}_{2}$ is produced in period 2. Point B represents the average cost, $\operatorname{Cbar}$, if $\left(\mathrm{Q}_{1}+\mathrm{Q}_{2}\right) / 2$ is produced in both periods, with the excess produced in period 1 carried over to period 2. The trade off is between the cost of storage for one period versus the saving from smoothing. The difference between A and B must be greater than the cost of holding inventory to justify
smoothing. Note also that if mean demand is expected to decrease below current production for an extended period (i.e. $\mathrm{Q}_{2}$ is current demand and $\mathrm{Q}_{1}$ is next period's expected demand), then it becomes optimal to reduce production and serve part of current demand from inventory. Thus production-smoothing motivation can lead to level changes if forecast sales change direction.

Figure B illustrates how periodic adjustments to production to match trend growth can result in lumpy movements in production even with production smoothing. The curved

## Figure B Periodic Production Smoothing


line indicates trend growth in sales with a seasonal component. If sales are forecast and production planned at the beginning of each period, then P1 represents the production level for the first period, P2 the second period, and P3 the third period. In the first half of each period, production exceeds sales and the difference goes into inventory. During the second
half of the period, production is less than sales and the difference is made up out of inventory. Each period, production is smoothed. However, because of the trend growth in sales, production jumps at the start of each new period. If data over all three periods are used, the variance of production may exceed the variance of sales.
\{End shaded insert \}

## Data Source and Transformation

The data used are from the Census Bureau's monthly data on manufacturing and retail inventories and sales, seasonally unadjusted and adjusted. Production is computed by adding the change in inventories to sales each period. A total of 35 series, 25 manufacturing and 10 retail, were analyzed. Table 1 lists the series and the years of data used. Most manufacturing data covered the period 1958 to 1998. Retail data covered the period 1987-1998.

## HP Filtering

After taking logs of the data, a nonlinear trend was removed using a HodrickPrescott (HP) filter with penalty set to 14,400 as recommended for monthly data. This method removes the "low frequency" components from the data, whether due to price increase or trend growth. Figure 1 illustrates the transformation of the data for the most recent 10 years for the Stone, Clay and Glass manufacturing industry. The smooth line shows the trend that is extracted to get the filtered data.

## Frequency Domain

Looking at the data in the frequency domain highlights the effect of the seasonal adjustment and the HP filter. The Fourier series decomposes the data into the
contribution of various frequencies to the total. If there is a trend present, there will be a large contribution from the low frequencies. If there is a strong contribution at a particular frequency compared to others, there will be a noticeable spike at that frequency. Figure 2 compares the seasonally adjusted and unadjusted Fourier series of sales for Stone, Clay and Glass Industries with the Fourier series for the HP filtered data. The spikes in the unadjusted data occur at cycles of 12 months, 6 months, 4 months and 3 months, reflecting harmonics of the seasonal cycle. The seasonally adjusted data have no spikes. The appearance of harmonics in the data may reflect the aggregation of individual firms with seasonal cycles that are offset, (i.e. some may have peak sales in winter while others peak in summer. The HP filtered data show the absence of the low frequency components while the high frequency contributions appear to be intact. Figure 3 shows the I/S ratio of selected industries, and Figure 4 shows the Fourier series of the ratios. In the next section, we report the results of the variance ratio test, then compare this to the frequency spectra of the I/S ratios of the sectors.

## Results

The typical measure of production smoothing is the ratio of the variance of production to the variance of sales. ${ }^{2}$ A ratio more than 1.0 implies that production is more volatile than sales and therefore contradicts the smoothing hypothesis. A negative correlation between sales and the change in inventory may be insufficient to produce a lower variance in production than in sales. Tables 2,3 and 4, summarize the results, showing the variance of sales, inventories and the covariance of sales and the change in inventories. Of the 35 seasonally unadjusted series, there are only three manufacturing
industries with positive covariances between sales and the change in inventories, -
Nonferrous and Other Primary Metals (a sub-category of Primary Metals), Paper and Allied Products and the Petroleum and Coal Products. By contrast, the covariances of all but three manufacturing series with seasonally adjusted data are positive. The seasonally adjusted retail data indicated some with negative covariance of inventory investment and sales, but none sufficiently negative to result in variance ratio less than 1.0.

## Manufacturing

The variance ratio of the detrended seasonally unadjusted data for all manufacturing is less than 1.0, but only barely; leaving unanswered the question of whether the production-smoothing model is adequate at this level of aggregation. At the two-digit SIC level of aggregation, SIC codes 33, 34, and 36 have variance ratios greater than 1.0 for the detrended $\log$ data, while SIC codes $32,35,37$, and 38 , as well as the "all other durable goods" category have variance ratios less than 1.0. The implication is that most durable goods industries smooth production over high frequency periods. ${ }^{3}$ The seasonally adjusted data do not show smoothing, indicating that removing the higher frequencies from the data masks evidence of smoothing.

In the nondurable goods category in Table 2, only Textile Mill Products (SIC 22) and Chemical and Allied Products (SIC 28) have variance ratio less than 1.0 for the detrended log seasonally unadjusted data. The aggregate nondurable goods industries has a variance ratio greater than 1.0. Intuitively we would expect production of some nondurables to be less amenable to storage. For instance, Tobacco Products may be

[^1]largely influenced by crop size rather than by demand, while demand may be less elastic seasonally. Of the nine "other" manufacturing sectors, which reflect a lower level of aggregation, evidence of smoothing is revealed in seven when seasonally unadjusted data are used. (See Table 3). This suggests that aggregation may be playing a role as well.

## Retail Sector

The seasonally unadjusted data for the retail sector reveals smoothing by most industries, suggesting that some retail firms may accumulate inventory in anticipation of seasonal increases in sales. Retail Food Stores (SIC 54) and aggregate of Retail Durable Goods stores are the only two of the ten series that have a variance ratio higher than 1.0. Given that fixed costs associated with transportation are likely to induce (S,s) behavior at the retail level, detecting smoothing may appear to be contradictory. However, here again the frequency of observation influences the detection of the underlying decision rule. It is likely that adjustments to inventory based on the ( $\mathrm{S}, \mathrm{s}$ ) rule takes place at frequencies less than one month. So, on average, inventory moves between upper and lower bounds within a month. Consequently, monthly data reveals seasonal movements in the bandwidth, while obscuring higher frequency ( $\mathrm{S}, \mathrm{s}$ ) movements. Seasonal smoothing at the retail level does not preclude ( $\mathrm{S}, \mathrm{s}$ ) behavior at higher frequency. In addition, aggregation over a large number of establishments is likely to dampen high frequency movements.

[^2]
## Fourier Series

Figures $1 \mathrm{~A}-5 \mathrm{~A}$ in the appendix shows plots of the Fourier spectra of the I/S ratios, and detrended sales and inventory movements of all 35 series. The horizontal axes of the plot of the spectra are labeled in multiples of $\pi$. Annual cycles are at $\pi / 6$, cycles of 6 months are at $\pi / 3$ and so on. The magnitude of the spike at each frequency gives an indication of the dominant cycles. In Figure 1A, the I/S ratio of the Stone, Clay and Glass Products sector has a high annual component (compared to 6-month). Figure 2A shows for that industry a negative correlation between detrended inventory and sales. By comparison, the Instruments and Related sector shows a high 3-month (quarterly) component (compared to annual) in Figure 1A. The corresponding chart in Figure 2A shows the high frequency composition of sales in this sector while inventory shows more of an annual cycle. Whereas the seasonally unadjusted variance ratio of Stone, Clay and Glass Products was 0.6864, the variance ratio for Instruments and Related Products was 0.9359 .

For the three industries with positive covariance between sales and the change in inventory, Nonferrous and Other Metals and Paper and Allied Products show seasonal spikes in the I/S spectra, while no significant seasonality is depicted for Petroleum and Coal Products. The positive co-movement between sales and inventory for all three is observable in Figure 2A. For the Petroleum and Coal industry the positive co-movement between them eliminates all seasonal components from the I/S ratio while the other two industries show higher swings in sales than inventory, leaving some seasonality in the ratio.

## Summary and Conclusions

Empirical evidence of production smoothing has been relatively elusive. Part of the problem appears to be the tendency to use seasonally adjusted data. This paper finds that smoothing takes place in a large proportion of manufacturing industries at seasonal frequencies or higher. Seasonal adjustment of the data removes this evidence. Removing the trend from the data allows us to exclude changes in production associated with trend growth in sales. This confirms empirical results of Allen (1997B), which suggest that inventory management at the firm level reflects planned and unplanned changes. The trend component of production reflects planned additions to inventory levels based on trend movement in sales, while the higher frequency component of production reflects smoothing over shorter horizons. Evidence of seasonal smoothing in the retail sectors suggests that retail firms also manage inventory to smooth seasonal fluctuations in sales. Although smoothing is not generally associated with retail inventory management it is not inconsistent with ( $\mathrm{S}, \mathrm{s}$ ) behavior at frequencies higher than the observed data.

In summary, we find evidence of production smoothing at relatively high frequencies when trend is removed from seasonally unadjusted data. We interpret this to mean that using data that has been seasonally adjusted and includes trend growth, limits the ability to extract the underlying motivation for holding inventories. To the extent that seasonal cycles mimic business cycles, analysis of production/inventory behavior at seasonal frequencies may provide insights into business cycle dynamics.

## References

Allen, Donald S. "A Multi-Sector Inventory Model," Journal of Economic Behavior and Organization, (January 1997A), pp. 55-87.
$\qquad$ . "Do Inventories Moderate Fluctuations in Output?" this Review, (July/August 1997B), pp. 39-50.

Beason, Richard. "Tests of Production Smoothing in Selected Japanese Industries," Journal of Monetary Economics (June 1993), pp. 381-94.

Blinder, Alan S. "Retail Inventory Behavior and Business Fluctuations," Brookings Papers on Economic Activity, 2 (1981), pp. 443-505.
$\qquad$ . "Can the Production Smoothing Model of Inventory Behavior be
Saved?" Quarterly Journal of Economics (August, 1986), pp. 431-53.
$\qquad$ ., and Louis J. Maccini. "The Resurgence of Inventory Research: What Have We Learned?" Journal of Economic Surveys (1991), pp. 291-328.

Carpenter, Robert E., and Daniel Levy. "Seasonal Cycles, Business Cycles, and the Comovement of Inventory Investment and Output," Journal of Money, Credit, and Banking (August 1998), pp. 331-46.

Dimelis, Sophia P., and Moheb A. Ghali. "Classical and Variance Bounds Tests of the Production Smoothing Hypothesis," International Journal of Production Economics (June 1994), pp. 15-22.

Fair, Ray C. "The Production-Smoothing Model is Alive and Well," Journal of Monetary Economics (November 1989), pp. 353-70.

Ghali, Moheb A. "Seasonality, Aggregation and the Testing of the Production Smoothing Hypothesis," The American Economic Review (June 1987), pp. 464-69.

Holt, Charles C., Franco Modigliani, John F. Muth, and Herbert A. Simon. Planning Production, Inventories, and Work Force (1960), Prentice-Hall Inc., Englewood Cliffs, NJ.

Kollintzas, Tryphon. "A Generalized Variance Bounds Test with an Application to the Holt et al. Inventory Model," Journal of Economic Dynamics and Control (Jan/Feb 1995), pp. 59-89.

Krane, Spencer D., and Stephen N. Braun. "Production Smoothing Evidence from Physical-Product Data," Journal of Political Economy (June 1991), pp. 558-81.

Lai, Kon S. "Aggregation and Testing of the Production Smoothing Hypothesis," International Economic Review (May 1991), pp. 391-403.

Miron, Jeffrey A., and Stephen P. Zeldes. "Seasonality, Cost Shocks, and the Production Smoothing Model of Inventories," Econometrica (July 1988), pp. 877-908.

Miron, Jeffrey A. The Economics of Seasonal Cycles, (1996), Massachusetts Institute of Technology, Cambridge, Massachussets.

Ramey, Valerie A. "Nonconvex Costs and the Behavior of Inventories," Journal of Political Economy (April 1991), pp. 306-34.

West, Kenneth D. "A Variance Bounds Test of the Linear Quadratic Inventory Model," Journal of Political Economy (April 1986), pp. 374-401.

| Table 1: Industries Analyzed | Years of |
| :---: | :---: |
| All Manufacturing Industries | 1/58 to 12/98 |
| Stone, Clay and Glass (SIC 32) | 1/58 to 12/98 |
| Primary Metals (SIC 33) | 1/58 to 12/98 |
| Nonferrous and other Primary Metals | 1/58 to 12/98 |
| Fabricated Metal Products (SIC 34) | 1/58 to 12/98 |
| Industrial Machinery and Equipment (SIC 35) | 1/58 to 12/98 |
| Electrical Machinery (SIC 36) | 1/58 to 12/98 |
| Transportation Equipment (SIC 37) | 1/58 to 12/98 |
| Instruments/Related Products (SIC 38) | 1/58 to 12/98 |
| All Other Durable Goods | 1/58 to 12/98 |
| Nondurable Goods Manufacturing Industries | 1/58 to 12/98 |
| Tobacco Products (SIC 21) | 1/58 to 12/98 |
| Textile Mill Products (SIC 22) | 1/58 to 12/98 |
| Paper and Allied Products (SIC 26) | 1/58 to 12/98 |
| Chemical and Allied Products (SIC 28) | 1/58 to 12/98 |
| Petroleum and Coal Products (SIC 29) | 1/58 to 12/98 |
| Automotive Equipment | 1/58 to 12/98 |
| Home Goods and Apparel | 1/58 to 12/98 |
| Consumer Staples | 1/58 to 12/98 |
| Machinery and Equipment * | 1/68 to 12/98 |
| Business Supplies | 1/58 to 12/98 |
| Construction Materials/Supplies /Intermediate | 1/58 to 12/98 |
| Capital Goods Industries | 1/58 to 12/98 |
| Producers' Durable Equipment * | 1/68 to 12/98 |
| Household Durable Goods | 1/58 to 12/98 |
| All Retail | 1/87 to 12/98 |
| Retail: Durable Goods Stores | 1/87 to 12/98 |
| Retail: Bldg Matls/Hdwre/Garden Supply/Mobile Home Dealers (SIC 52) | 1/87 to 12/98 |
| Retail: Automotive Dealers (SIC 55) | 1/87 to 12/98 |
| Retail: Furniture, Home Furnishings \& Eqpt Stores (SIC 57) | 1/87 to 12/98 |
| Retail: Nondurable Goods Stores | 1/87 to 12/98 |
| Retail: General Merchandise Group Stores (SIC 53) | 1/87 to 12/98 |
| Retail: Department Stores ex Leased Departments (SIC 531) | 1/87 to 12/98 |
| Retail: Food Stores (SIC 54) | 1/87 to 12/98 |
| Retail: Apparel and Accessory Stores (SIC 56) | 1/87 to 12/98 |
| * Starts in 1968. |  |


| Table 2 Manufacturing |  | Variance | Variance | Covariance | Variance | Variance Ratio |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| All Manufacturing Industries |  |  | Inventory | Sales and |  |  |
|  |  | Sales | Investment | Inven. Invest. | Prod |  |
|  | NSA | 0.00272 | 0.00029 | -0.00015 | 0.00270 | 0.99554 |
|  | SA | 0.00061 | 0.00014 | 0.00009 | 0.00092 | 1.50964 |
| Stone, Clay and Glass (SIC 32) | NSA | 0.00883 | 0.00076 | -0.00176 | 0.00606 | 0.68635 |
|  | SA | 0.00117 | 0.00023 | 0.00000 | 0.00141 | 1.20013 |
| Primary Metals (SIC 33) | NSA | 0.01106 | 0.00072 | -0.00015 | 0.01148 | 1.03841 |
|  | SA | 0.00731 | 0.00054 | 0.00035 | 0.00855 | 1.16922 |
| Nonferrous and other | NSA | 0.00729 | 0.00059 | 0.00010 | 0.00808 | 1.10774 |
| Primary Metals | SA | 0.00462 | 0.00037 | 0.00021 | 0.00541 | 1.17196 |
| Fabricated Metal Products | NSA | 0.00399 | 0.00112 | -0.00027 | 0.00456 | 1.14233 |
| (SIC 34) | SA | 0.00104 | 0.00065 | 0.00011 | 0.00190 | 1.83141 |
| Industrial Machinery and | NSA | 0.00721 | 0.00123 | -0.00089 | 0.00667 | 0.92451 |
| Equipment (SIC 35) | SA | 0.00133 | 0.00050 | 0.00029 | 0.00240 | 1.80271 |
| Electrical Machinery (SIC 36) | NSA | 0.00496 | 0.00086 | -0.00020 | 0.00542 | 1.09311 |
|  | SA | 0.00104 | 0.00034 | 0.00022 | 0.00182 | 1.75383 |
| Transportation Equipment | NSA | 0.01422 | 0.00118 | -0.00148 | 0.01244 | 0.87488 |
| (SIC 37) | SA | 0.00388 | 0.00039 | -0.00014 | 0.00399 | 1.02722 |
| Instruments/Related Products | NSA | 0.00418 | 0.00121 | -0.00074 | 0.00391 | 0.93591 |
| (SIC 38) | SA | 0.00072 | 0.00058 | 0.00003 | 0.00137 | 1.89900 |
| All Other Durable Goods | NSA | 0.00634 | 0.00048 | -0.00046 | 0.00590 | 0.92985 |
|  | SA | 0.00155 | 0.00024 | 0.00012 | 0.00204 | 1.31248 |
| Nondurable Goods Manufacturing Industries | NSA | 0.00164 | 0.00019 | -0.00006 | 0.00171 | 1.03902 |
|  | SA | 0.00038 | 0.00011 | 0.00005 | 0.00058 | 1.52003 |
| Tobacco Products (SIC 21) | NSA | 0.01631 | 0.02009 | -0.00220 | 0.03199 | 1.96182 |
|  | SA | 0.00396 | 0.00469 | -0.00005 | 0.00855 | 2.15891 |
| Textile Mill Products (SIC 22) | NSA | 0.00668 | 0.00099 | -0.00073 | 0.00620 | 0.92822 |
|  | SA | 0.00139 | 0.00025 | 0.00012 | 0.00188 | 1.35705 |
| Paper and Allied Products (SIC 26) | NSA | 0.00227 | 0.00023 | 0.00008 | 0.00265 | 1.16834 |
|  | SA | 0.00109 | 0.00011 | 0.00014 | 0.00148 | 1.35517 |
| Chemical and Allied Products (SIC 28) | NSA | 0.00338 | 0.00038 | -0.00036 | 0.00304 | 0.89854 |
|  | SA | 0.00096 | 0.00019 | 0.00010 | 0.00134 | 1.40595 |
| Petroleum and Coal Products (SIC 29) | NSA | 0.00534 | 0.00080 | 0.00032 | 0.00678 | 1.27019 |
|  | SA | 0.00423 | 0.00058 | 0.00036 | 0.00552 | 1.30557 |


| Table 3: "Other" Manufacturing |  | Variance | Variance | Covariance | Variance | Variance Ratio |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Automotive Equipment |  | Sales | Inventory | Sales and | Production |  |
|  |  |  | Investment | Inv. Invest. |  |  |
|  | NSA | 0.02848 | 0.00097 | -0.00213 | 0.02518 | 0.88435 |
|  | SA | 0.00898 | 0.00027 | -0.00012 | 0.00901 | 1.00339 |
| Home Goods and Apparel | NSA | 0.00898 | 0.00138 | -0.00195 | 0.00647 | 0.72049 |
|  | SA | 0.00088 | 0.00024 | 0.00014 | 0.00140 | 1.58834 |
| Consumer Staples | NSA | 0.00133 | 0.00035 | 0.00004 | 0.00177 | 1.32663 |
|  | SA | 0.00032 | 0.00014 | 0.00002 | 0.00049 | 1.52705 |
| Machinery and Equipment | NSA | 0.00722 | 0.00130 | -0.00142 | 0.00567 | 0.78549 |
|  | SA | 0.00096 | 0.00041 | 0.00012 | 0.00161 | 1.68163 |
| Business Supplies | NSA | 0.00183 | 0.00033 | -0.00015 | 0.00186 | 1.01557 |
|  | SA | 0.00051 | 0.00019 | 0.00004 | 0.00078 | 1.54868 |
| Construction Materials/Supplies | NSA | 0.00563 | 0.00041 | -0.00052 | 0.00499 | 0.88617 |
| /Intermediate | SA | 0.00124 | 0.00020 | 0.00012 | 0.00167 | 1.35123 |
| Capital Goods Industries | NSA | 0.00694 | 0.00144 | -0.00163 | 0.00513 | 0.73934 |
|  | SA | 0.00080 | 0.00048 | 0.00010 | 0.00148 | 1.85593 |
| Producers' Durable Equipment | NSA | 0.00667 | 0.00083 | -0.00066 | 0.00619 | 0.92711 |
|  | SA | 0.00120 | 0.00033 | 0.00010 | 0.00173 | 1.43817 |
| Household Durable Goods | NSA | 0.00840 | 0.00149 | -0.00133 | 0.00723 | 0.86042 |
|  | SA | 0.00114 | 0.00047 | 0.00018 | 0.00196 | 1.72187 |


| Table 4: Retail |  | Variance | Variance | covariance | Variance | Variance ratio |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Sales | Invent. Invest | Sales and Inv Invest. | Production |  |
| All Retail | NSA | 0.00630 | 0.00222 | -0.00206 | 0.00439 | 0.69691 |
|  | SA | 0.00011 | 0.00009 | -0.00001 | 0.00018 | 1.68785 |
| Retail: Durable Goods Stores | NSA | 0.00618 | 0.00264 | -0.00116 | 0.00650 | 1.05145 |
|  | SA | 0.00046 | 0.00038 | -0.00004 | 0.00077 | 1.67262 |
| Retail: Bldg Matls/Hdwre/Garden | NSA | 0.02100 | 0.00199 | -0.00361 | 0.01576 | 0.75081 |
| Supply/Mobile Home Dealers (SIC 52) | SA | 0.00065 | 0.00034 | 0.00003 | 0.00104 | 1.60598 |
| Retail: Automotive Dealers (SIC 55) | NSA | 0.00805 | 0.00402 | -0.00241 | 0.00727 | 0.90217 |
|  | SA | 0.00075 | 0.00087 | -0.00016 | 0.00130 | 1.72889 |
| Retail: Furniture, Home Furnishings \& | NSA | 0.01237 | 0.00749 | -0.00483 | 0.01021 | 0.82485 |
| Eqpt Stores (SIC 57) | SA | 0.00041 | 0.00058 | 0.00004 | 0.00107 | 2.63839 |
| Retail: Nondurable Goods Stores | NSA | 0.00809 | 0.00270 | -0.00325 | 0.00430 | 0.53181 |
|  | SA | 0.00006 | 0.00004 | 0.00000 | 0.00009 | 1.45016 |
| Retail: General Merchandise Group | NSA | 0.04588 | 0.01966 | -0.02139 | 0.02276 | 0.49608 |
| Stores (SIC 53) | SA | 0.00011 | 0.00037 | -0.00004 | 0.00041 | 3.60050 |
| Retail: Department Stores ex | NSA | 0.04906 | 0.02083 | -0.02225 | 0.02539 | 0.51750 |
| Leased Departments (SIC 31) | SA | 0.00013 | 0.00043 | -0.00005 | 0.00046 | 3.70666 |
| Retail: Food Stores (SIC 54) | NSA | 0.00190 | 0.00019 | -0.00006 | 0.00198 | 1.04026 |
|  | SA | 0.00008 | 0.00002 | 0.00000 | 0.00010 | 1.28350 |
| Retail: Apparel and Accessory Stores | NSA | 0.03814 | 0.01931 | -0.01861 | 0.02023 | 0.53035 |
| (SIC 56) | SA | 0.00030 | 0.00048 | 0.00002 | 0.00082 | 2.73757 |

Figure 1 Stone, Clay and Glass Products (Monthly Sales, Seasonally Unadjusted, millions current dollars)


Figure 2 Fourier Spectrum of Stone, Clay and Glass Industries Sales Seasonally Adjusted and Unadjusted

## Seasonally Adjusted

Unadjusted Seasonally


Figure 3 Inventory-to-Sales Ratios Selected Industries


Figure 4 Fourier Spectrum of Inventory-to-Sales Ratios Selected Industries


## Appendix

Figure 1A: Fourier Series of Inventory to Sales Ratio Table 2 Series


Figure 2A: Detrended Log of Inventory (solid) and Sales(dotted) Table 2 Series


Figure 3A: Fourier Series of Inventory to Sales Ratio Table 3 Series


Figure 4A: Detrended Log of Inventory (solid) and Sales (dotted) Table 3 Series






Figure 5A Fourier Spectrum of Inventory-to-Sales Ratio and Detrended Sales (dotted) and Inventory (solid)



[^0]:    ${ }^{1}$ If inventories are high when sales are low and vice versa, then the I/S ratio will fluctuate accordingly. Obviously, if inventory remained constant and sales varied seasonally, the I/S ratio would also fluctuate seasonally, so this is not an exact metric of smoothing. However, the seasonality of the ratio does suggest the frequency over which smoothing is taking place.

[^1]:    ${ }^{2}$ A more appropriate test of production smoothing would be a comparison of the variance of production to the variance of forecasted sales. Disentangling anticipated and unanticipated changes in sales is troublesome at best.

[^2]:    ${ }^{3}$ At least if we interpret production smoothing as meaning that the growth rate of production varies less than the growth rate of sales.

