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On the Accuracy of Latin American Trade Statistics: a Nonparametric Test for 1925*

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ABSTRACT

This paper proposes a nonparametric test in order to establish the level of accuracy of the foreign trade statistics of 17 Latin American countries when contrasted with the trade statistics of the main partners in 1925. The Wilcoxon Matched-Pairs Ranks test is used to determine whether the differences between the data registered by exporters and importers are meaningful, and if so, whether the differences are systematic in any direction. The paper tests for the reliability of the data registered for two homogeneous products, petroleum and coal, both in volume and value. The conclusion of the several exercises performed is that we cannot accept the existence of statistically significant differences between the data provided by the exporters and the registered by the importing countries in most cases. The qualitative historiography of Latin American describes its foreign trade statistics as mostly unusable. Our quantitative results contest this view.

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INTRODUCTION

The issue of the (in)accuracy of the foreign trade statistics remains in the economics, development and trade literature to the present day (Parniczky (1980); Rozansky and Yeats (1994); Makhoul and Otterstrom (1998)). This paper proposes a nonparametric test in order to establish the level of accuracy of the Latin American foreign trade statistics when contrasted with the trade statistics of the main trading partners.

The study of Federico and Tena (1991) showed that, in historical terms, the accuracy of foreign trade statistics seems to be more robust than previously thought. The results of this paper also point in such direction. Nevertheless, this paper departs from previous exercises regarding the (in)accuracy of foreign trade data in several aspects. First, the paper focuses in the trade of a particular region in a single year. That is the paper provides a test for the accuracy of the foreign trade statistics of 17 Latin American countries for the year 1925. Second, rather than testing for the accuracy of the overall trade figures, the test is performed on data registered for a couple of quite homogeneous products, petroleum products and mineral coal. Third, the test applies to the accuracy of both the volumes and values registered on the official statistics of the exporting and importing countries. Most previous exercises did only test for the accuracy of the values registered, since the aggregate trade figures were used. Most of the previous tests tended to compare figures provided international bodies (OECD, IMF, League of Nations, etc), whereas here the foundations of such figures, the official statistics of the individual countries, are contrasted. Fourth, the Wilcoxon Matched-Pairs Signed-Ranks test is used to determine whether the differences between the data registered by exporters and importers are meaningful, and if so, whether the differences are systematic in any direction. At the end of the day, the question addressed is whether the differences observed are statistically meaningful. In other words, whether the story told from the exporters' side is or it is not

substantially different from the story told using the importers' figures. The paper, therefore, is not concerned with the issues of why and where from do the differences arise, these are important questions on their own right, but exceed the aim of this paper.

The first section of the paper presents the problem and specifies the issues to be investigated. The second section contemplates the nature of the data proposed for the test and introduces the data set to be used. Section three offers some preliminary contrasts of the data offered by exporters and importers. In section four the choice of a nonparametric test is justified and the workings of the test are revealed. Section five summarizes the results obtained. The conclusions recapitulate the main findings and propose a research agenda.

1. THE PROBLEM

The general mistrust placed on trade statistics, particularly those of underdeveloped countries, represents a heavy burden on economic history research, since trade statistics are one of the oldest and most complete economic series available for analysis. For instance, a research project such as the described in Carreras et al. (2003) or Carreras et al. (2004) aimed at estimating the level of economic modernization in Latin American and Caribbean countries before World War II making systematic use of the trade statistics of these countries as well as of their principal trading partners in the developed world is immediately under suspicion.

From the seminal work of Morgernstern (1963) to the present day, the users of trade figures are aware of the divergence that exists between exporters' and importers' figures. The impression from the economic literature is that the researcher should be even more suspicious of the data the more underdeveloped the country. Among others, the studies of Naya and Morgan (1969), Yeats (1990), Rozansky and Yeats (1994) and, Makhoul and

Otterstrom (1998), show that the accuracy of trade statistics provided by developed countries is higher than that of the developing countries. For instance, Makhoul and Otterstrom (1998) found that the quality of the OECD trade statistics is much better than that provided by the non-OECD in a relatively recent period such as 1980 to 1994. Also Rozansky and Yeats (1994) found that discrepancies between importers' and exporters' reports appear especially important for the less developed countries.

That underdeveloped countries shall misreport statistics more often than developed nations comes as no much of a surprise. Allegedly many of the causes for misreporting have to do with lack of means for the collection of data, systematic distorted statistics for a specific purpose --improve credit worthiness; collect (or avoid) higher taxes--, simple corruption, smuggle, etc., all of which seem to occur more often in low income countries (see Yeats (1990)). Following such a line of reasoning the straightforward solution seemed to be to use the statistics of the more developed trade partners instead, which are expected to be of higher quality. However, Yeats (1995) concluded that 'the partner country gap filling procedures have little or no potential for improving the general coverage or quality of international trade data'. His final remark points at the need of 'improved procedures for data collection and reporting at the country level'.

In fact, there are a wide array of potential matters that would need to improve in order to reduce the differences between the quantities and, overall the values, annotated at the port of origin and that registered at destination: different accounting methods (CIF versus FOB, general versus special trade), different time of recording (goods movement versus money movement, fiscal versus calendar years), prices used (declared prices versus official prices), different units of measurement (currencies and exchange rates; units, dozens, weight, volume, length, etc), misclassification of products (thousand subcategories versus 'all others' type of categories), geographical misallocation (country of consignment

versus country of origin/destination), just to name the most relevant. A detailed explanation these and more reasons for discrepancies can be found in Allen and Ely (1953) and also Federico and Tena (1991). Given the list of issues, the ample pessimism about the accuracy and usefulness of international trade statistics for economic analytical purposes is comprehensible.

In historical terms, the view of Don (1968) that the 'comparison of trade statistics, for a historical analysis of economic relations between two countries, must be abandoned', and the cautions of Platt (1971) regarding the interpretation of the Latin American trade statistics before World War I, added to the overall wariness. Also McGreevey (1975) insisted in such direction when indicating that the trade data of industrialized countries may offer supplementary and alternative sources, especially for the Latin American countries, which only have incomplete and inexact commercial records.

Few exercises, however, challenge the general distrust on trade statistics. The work of Federico and Tena (1991) contested some of the above issues using international foreign trade statistics of the pre-World War II period and focusing on overcoming errors due to geographical assignment. Their results strengthened the trust on the accuracy of foreign trade statistics, at least at the aggregate level. In addition, they found no significant relationship between the level of development of the countries and the quality of the trade statistics produced.

Given the state of the art, any research based upon Latin American trade statistics first need to face the challenge of proving that useful and trustworthy interpretations can be extracted from the historical trade figures. This is the challenge of the present paper. Economic historians cannot hope for improvement of data collected many years ago, but renouncing to use trade statistics altogether is, to the say the least, inappropriate. The problems associated to trade statistics must be recognized, but also the magnitude of the

discrepancies observed must be placed within context. For some purposes a difference that in some metric looks large might in another metric be unimportant. How large is large in the present case depends, as usual, on the question asked. Differences that at the country level may look abysmal, placed in the context of the region will be a minor problem for the analysis at hand. The remaining of the paper is aimed at proving that the story told from the exporter side is almost identical to the story told using the importers' figures when analyzing the Latin American countries within the context of the region. The question addressed is whether the differences observed are statistically (and economically) meaningful for the interpretation of the imports of petroleum and coal of the Latin American countries relative to each other.

2. THE DATA

Before having a look at the data, lets reflect about the nature of the data proposed for the test. It has been already said that rather than testing for the accuracy of the overall trade figures, the paper concentrates on the reliability of the data registered for a couple of quite homogeneous products, petroleum products and mineral coal. There are some reasons to believe that specific product comparisons may be more fruitful than overall trade contrasts. To start with, contrast of homogeneous products had proven to yield better results. As referred by Federico and Tena (1991), the analysis of homogeneous commodities, such as wheat, provides a much less pessimistic view of the quality of the data (see the results of Ricci (1914)). Indeed, homogeneous products have some advantages at the time of contrasting figures at port of origin and that of destination. Homogeneous implies simpler standardization and classification. Fewer errors can be attributable to misclassification of products and measurement error. Even when petroleum and coal products were not totally free of such problems the truth is that still the most 'detailed' listings did not go beyond three categories for coal and up to a dozen for

petroleum products in 1925. The units of measurement although not completely standardized —long tones, short tones, kilograms, barrels, gallons, litters, cubic meters, were all in use— were straightforward to translate into common units (metric tones). In addition, the nature of the products made them difficult to smuggle in sizeable amounts. Finally, there was little incentive to systematically distort the figures in order to avoid taxes. Duties on imports of fossil fuels were small, if any, in the 1920s, according to the report by the U.S. Department of Commerce (by J.R.Bradley) (1931). As for export taxes, although existed, they were affordable compared to the additional set of taxes paid at origin (production taxes, royalties, handling taxes, etc).

Homogeneity and the nature of the products chosen gets rid of some of the problems listed in the previous section, especially when contrasting quantities rather than values traded. Yet some important ones remain on the list. It is the case of the geographical misallocation of trade. Where the final destination was taken to be the port of landing, the trade to all non-seaboard countries (Bolivia and Paraguay) would had been misallocated in the exporters' reports as corresponding to the intermediate countries (Argentina, Chile, Uruguay and Peru). The importance and possible ways of correcting this effect is investigated in Carreras-Marín and Badía-Miró (2005, forthcoming). Another issue, relevant only to the selected products, is the treatment given to bunkering. In some instances coal and petroleum arriving at Latin American ports would be right away loaded as fuel to departing ships. Some countries may record both the import and the export of such amounts, some other countries may record the importation but not the re-export, finally some other countries may not account at all the amounts unloaded and loaded at port for bunkering purposes. Regardless of the treatment given at the port of destination, the amounts were in all instances recorded at the port of origin as exports to the Latin American country.

More crucial to the differences between the values reported by the exporters and that reported by the importers is the so-called 'freight factor'. The 'freight factor' is the ratio between the CIF value of a commodity and the freight rate paid for its carriage, and can be taken as a proxy for the difference between CIF and FOB values. According to Moneta (1959), two main rules apply to the freight factor: (1) commodities of low unit value are relatively more expensive to ship than high-value ones and (2) the longer the distance a given commodity is shipped, the higher the freight rate and the higher the freight factor to be applied. Bulky commodities such as coal, petroleum, wheat, cement etc, are included in the first rule, they have a low unit value and are relatively expensive to ship. A telling illustration of the magnitude of the freight factor for bulky commodities is the method used by GATT still in the 1950s according to which the frequent procedure to adjust values from CIF to FOB was to reduce the value of fuel imports by 50 per cent and that of all other imports by 10 per cent (see Moneta (1959)). So in principle, imports of petroleum and coal by Latin America are the worse case scenario for contrasting the values traded according to importers and exporters since both rules apply. The large magnitude of the freight factor should make the values much larger at the port of destination than at port of origin, thus very different.

Having reflected on the nature of the products let's look into the dataset. By 1925, most Latin American countries were net importers of coal and petroleum products, mostly from the United Kingdom, the United States and, Germany; Mexico and Peru also supplied petroleum within the region. According to Rubio and Folchi (may,2005), the United States supplied little more than half of the oil and derivatives imported by the Latin American countries in 1925. Mexico, with a share of 40 per cent of the tonnage, and Peru, with seven percent, together supplied the remaining half. In the case of coal, the United Kingdom was the main supplier of the larger consumers --Argentina, Brazil and Uruguay--

while the United States was the chief supplier, sometimes even the only one, of coal to Central America and the Caribbean. In total, the United Kingdom supplied 68 per cent of the coal, the United States 26 per cent, while Germany and Belgium had much smaller shares (3.4 and 0.05 respectively). Altogether the G4 –United Kingdom, United States, Germany and Belgium-- provided 98 percent of the coal imported by the Latin American countries in 1925.

Of the 33 countries that constitute Latin America and the Caribbean at the dawn of the 21st century 18 elaborated trade statistics in 1925, although only 15 offer sufficient detail about the country of origin of the merchandise and the type of products imported.¹ These were Argentina, Bolivia, Brazil, Colombia, Costa Rica, Cuba, Chile, Dominican Republic, Ecuador, El Salvador, Guatemala, Haiti, Mexico, Nicaragua and, Peru. Apparently most of the smaller Caribbean islands, neither Honduras nor Panama elaborated trade statistics in 1925. Paraguay did, but the level of detail made them unusable. From the statistics of Venezuela and Uruguay it was not easy to detect the country of origin of the products but the total amounts imported were collected. According to the Société des Nations (1928) most Latin American republics used the Anglo-Saxon system for reporting their imports, that is, they reported 'general trade'.² The same source also informs that only Argentina, Guatemala, Paraguay and, Uruguay used official prices in the valuation of their imports; the rest applied the declared prices. From the exporters' side the official national trade statistics of the United States, the United Kingdom, Germany and Belgium (referred as G4 hereunder) were used, plus the export statistics of Mexico and Peru in the case of petroleum products. The problems of classification of

¹ Sources are quoted in the references under the heading 'official publications'.

² General trade accrued imports for home consumption plus the imports deposited in entrepôts, generally excluding trans-shipment trade and trade in transit. European countries tended to use the 'Continental system' and reported 'special trade' instead, which only includes the imports for domestic consumption. The few Latin American countries using the continental system were: Argentina, Bolivia, Chile, Paraguay and Peru. See Société des Nations (1928).

products and unification of the different units of measurement (volume in some cases, weight in others) were dealt with using the homogeneous criteria described in Folchi and Rubio (2004). Currencies were unified to the US dollar using the exchange rates in U.S. Department of Commerce (1925,1926). The result was a list of pairs of data, in volume and value, consisting of the figures of trade of petroleum products and mineral coal registered both at port of origin and at port of destination. Table 1 provides the pairs of data for the main trading partners of petroleum with the Latin American countries, while Table 2 does for coal.

{TABLE 1: Pairs of data, petroleum products}

{ TABLE 2: Pairs of data, coal }

Some issues must be clarified before any further examination of the data. It is possible to understand Tables 1 and 2 as before and after observations of the same shipment. Theoretically, what the tables report are cargoes to a specific destination measured at port of origin and the same cargo measured again when it reached its destination. In practice, the measures correspond to different sources, that is the official publications of the country of origin of the shipment (the amounts exported) and that of the country of destination (the amount imported). From the section above it is clear that the two measures are subject to sufficient hazards to make them differ from each other. Furthermore, exporters (especially the United Kingdom and Germany) did not report minor quantities sold to smallish countries, but these amounts show up in the Latin American home statistics. These cases, where the source was checked but no trade was reported, are represented in Tables 1 and 2 with ceros. It may also be the case that amounts reported in the Latin American country, or mere inexistence of the source. These cases, where the source was unavailable, are represented in Tables 1 and 2 with a dot.

3. SOME PRELIMINARY CONTRASTS

The first impression from Tables 1 and 2 is somehow hazy. Take the quantities traded first. In the case of petroleum products there are quite few remarkable matches: Brazil, Cuba, Chile, Haiti, Mexico and Peru report tonnages arriving from the United States that are pretty close to the amounts the United States reported as exported to such destinations; Argentina, Bolivia, Costa Rica, Ecuador and Nicaragua reported imports from Peru only kilos away from the Peruvian exports record. But then, on the contrary, less than half of tones reported as leaving the United States to destinations such as Colombia, Costa Rica or Ecuador were reported as imports in such countries, while Guatemala and Salvador declare receiving several times over the amount of oil that departed from the United States. Consider now the case of the quantities of coal imported from G4 shown in Table 2. The contrast of the volume imported according to both types of sources reveals that for a first group of countries a very close match: Argentina, Brazil, Colombia, Cuba, Ecuador, Nicaragua and, Dominican Republic. A second group of countries (Chile, Salvador and Peru) exhibits a less satisfactory correspondence, but still plausible. Finally, four countries show what appear to be irreconcilable differences between their statistics and the reported by the exporting countries. These are Bolivia, Costa Rica, Haiti and Mexico. Turning the attention to the values in dollars in Tables 1 and 2 does not improve the matters. There are observable differences of millions of dollars in Argentina and Brazil over the value reported at the ports of origin of the petroleum or coal such the United States, the United Kingdom, Mexico or Peru. A priori these differences could be thought to correspond to the 'freight factor'. Yet a closer look reveals thousands of dollars under valuations found in the same Latin American countries in the trade with Germany and Belgium. Furthermore, Chile and Cuba report values for millions of dollars below the value assigned at the ports of departure. For all the reasons already

described differences were expected to be greater in the values than in the quantities, but the sings are not the expected ones in all occasions. Besides, there are countries where the match comes down to few hundred dollars in transactions involving several thousands, as it is the case of the United States coal trade with Nicaragua or Ecuador. It is not clear to the naked eye in which cases the quantities and values reported by importers and exporters are close enough to each other's or just the opposite.

Of course, several measures might be deployed to quantify the error. One measure commonly used is the difference between the sources as percentage of either source, although it implies an arbitrary choice of the source that is to serve as denominator. An alternative practical indicator for the measurement of errors in data that are reported in two sources is shown in the last two columns of Tables 1 and 2. It is the *implicit minimal measurement error*, IMME, defined by van Bergeijk (1995), which here takes the form: {(destination source - origin source)/ (destination source + origin source)*100}. The IMME-indicator assumes implicitly that both sources are wrong, and offers a conservative estimate (indeed a lower limit only) for the measurement error in the data. It should be considered as an optimistic indicator of accuracy. Even with its help it is difficult to take an informed position over the accuracy of the data as a whole. The indicator ranges from 0 to 100 per cent, and takes both positive and negative sings.

Moreover, it is impossible to find out from this indicator whether the 100 per cent found between say the British exports of coal to Costa Rica is more, less or as relevant as the 78 per cent found for Bolivia in the same trade, or the 30 per cent corresponding to Haiti in its coal trade with the United States. All it responds to is to the fact that there are at least 320 tones, 4049 tones and 72 tones respectively misreported in each of these transactions. Even within the same country it is not clear why the 51 percent indicator obtained for the value of the petroleum trade from Mexico to Argentina is better or worse

than the indicator obtained for the petroleum trade from the United Kingdom (which obtained a minus 69 percent), when there were 9 million dollars in excess in the trade with Mexico and not even a quarter of a million mismatch in the trade with the British. Beyond informing of the existence of country-pairs differences, the simple contrast of individual country data does not help much to determine how important the discrepancies are for acquiring an accurate impression of the coal and petroleum trade in Latin America by 1925.

Federico and Tena (1991) argued that a better test for the reliability is the comparison between the total of each country's trade (according to its own statistics) and the sum of these flows as registered by its partner countries' statistics. That is rather than using country-pairs contrast, they advocate for comparisons of total trade flows. The trade flows from main trade partners of each country's trade (according to its own statistics) and the sum of these flows as registered by the main partner countries' statistics are reported in Table 3 for petroleum and Table 4 for mineral coal. Only the countries where sources could be checked at both ends are included in these tables.

{TABLE 3: Trade flows from principal partners, petroleum}

{TABLE 4: Trade flows from principal partners, coal }

Indeed, as prognosticated the comparability of the data improves when the aggregated flows are used instead than the bilateral trade. The range of the IMME indicator improves, since the ceros are not present in the aggregate trade flows, thus it gets rid of the automatic 100 percent indicators. Nevertheless the variation still goes from cero percent, as in the case of the value of the Mexican petroleum imports, to 47 percent (negative) in the quantity of oil traded by Colombia (which in fact respond to Colombia reporting over 5000 tones of oil coming from Costa Rica, oil that was almost certainly from the United States and so reported at origin, but not accounted here on the Colombian

side). In general the gaps are smaller in quantities than in values, but no general rule applies. Only two sings remain negative in the value of coal traded but in petroleum trade five countries report smaller values than the aggregated values at port of origin. But still at the aggregated level there are many millions of dollars over and underreported at destination. Argentina declared values are almost double of the values for oil at port of origin, while Cuba and Chile report lower values than their exporter partners.

This repeated issue gives ground to question one of the main assumptions of the literature: the exporters report values FOB, the importers report them CIF. In fact, while the former seems to be generally true, the later was not in the 1920s. According to the Société des Nations (1928), more than half of the 17 countries included in Tables 3 and 4, plus the United States, valued and reported their import figures FOB, at least in the period 1913 to 1926. The list includes Cuba, Chile, Guatemala (if imported by sea although added an arbitrary 25 to the value), Honduras, Mexico, Nicaragua, Peru (which added a 20 percent to the value), Dominican Republic and Venezuela. This fact together with the inclusion or exclusion of duties and taxes from the values reported at either end may help to explain the unexpected undervalues at port of destination and that the differences between importers and exporters values were in most cases smaller than the 50 percent rule generally applied for fuels.

Nevertheless, the most striking feature of Tables 3 and 4 is the accuracy of the aggregated figures for the region as a whole. The IMME-indicator obtained for the total tonnage comes down to 1 percent (negative) in the case of petroleum, and 0,2 percent in the case of coal. In absolute terms the implication is that of the over 4 million tones of petroleum and over 6 million tones of coal revealed by the exporters as sold to Latin American countries just over fifty thousand tones of petroleum and thirty thousand of coal got misreported somehow. The aggregate values for the region are not so exciting, for the

IMME-indicators remain at 11 percent for petroleum and 23 per cent for coal. In absolute terms the implication is that the Latin American countries reported a total value over 20 million dollars above the values the exporters declared they obtained from their sales to Latin America of each of these products. Most of the difference was solely explained by the Argentinean overvaluation in all cases.

If the analysis were brought to an end at this point, the conclusions could only be pessimistic at the country level. The differences seem irreconcilable for some countries. Whichever of the ample list of reasons of the previous section could be blamed for the discrepancies observed. At the aggregate level, however, the results of Federico and Tena still hold, but someone could argue it could be due to mere chance or self-cancellation of errors. Yet as van Bergeijk (1995) remind us 'absolute precision obviously is impossible to achieve and the improvement of the accuracy of economic measurement is in many cases not an optimal solution'. The important question is whether these gaps are sufficient to invalidate any inference extracted from this data. The economic historian would like to be able to take home some lessons from these data that withstand the mistrust on the original data sources. The trust placed on the answer depends very much of the question asked to the data. If the question addressed was about the precise quantity of coal imported by Guatemala, the over 3000 thousand tones discrepancy between the importer and the exporters' sources may be a problem (in per capita terms the difference comes down to 0.11 kg per capita versus 1.5 kg, see Rubio and Folchi (may,2005)). If the question, however, referred to the imports of Guatemala relative to the rest of the continent, the difference between sources may be trivial: Guatemala imported very little coal relative to most countries of the region whatever source used.

{FIGURE 1: Total petroleum trade flows, exporter vs importers data}
{FIGURE 2: Total petroleum trade flows, exporter vs importers data}

In fact, the story told from the exporter's side seems to be very similar to the story told using the importers' figures when analyzing the Latin American countries within the context of the region. Better than the naked eye or the IMME-indicator this can be grasped graphically. If the importers and exporters were issuing exactly the same reports on the quantities and values traded, when plotted against each others a perfect 45-degree diagonal will appear. Figures 1 and Figure 2 present the scatter plot of the data in Tables 3 and 4. The x-axis shows the exporter data in logs while the y-axis reflects the importer data also in a log scale. A data point below (above) the straight line indicates that the importer was under-reporting (over-reporting) trade with respect the partner's data. The quantities and values match quite nicely the diagonal line, which links identical values in both axes. From either source Argentina is the country making the larger expenditure in buying energy inputs. From either source Brazil and Cuba were next, though Uruguay spent just much as the latter in buying coal. Do not matter the sources used, Argentina bought more tones of coal than any other Latin American country, followed by Brazil, Cuba, Uruguay, Chile and Mexico. In the case of oil, Cuba bought more tones than any other country, followed by Chile, Argentina, Brazil, Mexico and Uruguay, regardless of whether the exporters or the Latin American nations give the answer. At the other end, Haiti used fewer tones of petroleum and coal than any of its neighbors. Bolivia, even admitting than the exporters could not see her as final destination, thus Bolivian statistics reported more tones than her trade partners, still remain at the bottom buying little coal and petroleum, and so on and so forth. This is not a complete picture of the energy intakes of the region -alternative suppliers and domestic production must be taken into account as in Rubio and Folchi (may,2005))-- but it is a good indication that useful and trustworthy interpretations can be extracted from the historical trade figures of Latin America. However, the graphical representation is still subject to the observer's interpretation, finding some statistical

reassurance confirming that the stories told from either side are sufficiently similar would strengthen the results.

4. A NON PARAMETRIC TEST

It has been already mentioned that is possible to think about the data presented as before and after observations of the same shipment. Theoretically, what the data report are cargoes to a specific destination measured at port of origin and the same cargo measured again when it reached its destination. Before/after matched-pairs tests are widely used in medical, biological, behavioral and engineering experiments, where a choice is to be made between parametric or nonparametric tests of paired data (see, for instance, Bland (1995) and Motulsky (1995)). Paired data means that the values in the two groups being compared are linked, that is both samples have some factor in common, it does not matter whether it is geographical location or before/after treatment. That is why they are also known as 'tests for correlated samples'. In studies that gather before and after measurements like this, interest focuses on the difference between each pair.

The choice between a parametric and a nonparametric test derives from the underlying assumptions about the data to be tested. Two assumptions are most relevant here. The parametric tests assume the data to be normally distributed and of equal-interval nature (that is someone who improved four points improved twice as much compared with someone who improved two points). In a nonparametric test, however, the assumption of being a normal distribution does not have to be met and, the data are assumed at an ordinal-metric level (i.e., that the original data can be validly ordered and that the difference between the two sets of data can be validly ordered). As stated by Lowry (1999) the choice is not simply a question of good manners or good taste. If there is one or more of these assumptions that cannot reasonably supposed to be satisfied, then the corresponding test for correlated samples cannot be legitimately applied. According to

Motulsky (1995) a nonparametric test is definitively preferred in situations where either the population is clearly not normal or some values are 'off the scale', that is, too high or too low. To these, most experts add other situations in which nonparametric tests will be preferred: when testing in small samples (<30) and when there are unequal variances across groups.

As it happens, it seems that most of these situations concur in the data presented here. On the one hand, the data presented is skewed. The Shapiro-Wilk normality test applied consecutively to data in Tables 1 through 4 rejects they are normally distributed. Neither they are the various subgroups of data in the tables or the differences between values and quantities in each table. The differences in volume for the trade originating in the United States and in Mexico are the only series at the edge of a normal distribution, but failed the test. With few data points, however, it is difficult to tell whether the data are Gaussian by inspection, and the formal test has little power to discriminate between normal and non-normal distributions. Nevertheless, nonparametric statistical test are to be used to assess variables that are skewed or which normality is doubted as in this case. On the other hand, it is clear that there are values off the scale, million dollars in the larger countries versus few thousands at the bottom end, which also applies to the quantities with million of tones versus few thousand kilograms. Besides, the equal-interval nature of the parametric tests it is difficult to assume here. A country with a 10,000 thousand dollars gap is not ten times as accurate as one with a gap of 100,000 thousand dollars, while the former may be a huge gap for Haiti, Bolivia or El Salvador, the latter may be a smallish gap for Chile, Uruguay or Argentina. In any event, when the data within two correlated samples fail to meet one or another of the assumptions of the parametric tests, the appropriate nonparametric alternative can be found in the Wilcoxon Matched-Pairs Signed-Ranks test (Wilcoxon-MPSR test, hereafter).

The Wilcoxon-MPSR test, named after the work of the chemist and statistician Frank Wilcoxon (1945), is one of the cornerstones of nonparametric inferential statistics.³ This test is mostly applied in biological and medical sciences, although it has its foundations on an earlier paper by the economist Milton Friedman (1937). The assumptions of the Wilcoxon test, according to Lowry (1999) are: (a) that the paired values of X_A and X_B are randomly and independently drawn (i.e., each pair is drawn independently of all other pairs); (b) that the measured variable (e.g., a subject's probability estimate) is intrinsically continuous, capable in principle, if not in practice, of producing measures carried out to the nth decimal place; and (c) that the measures of X_A and X_B have the properties of at least an ordinal scale of measurement, so that it is meaningful to speak of "greater than," "less than," and "equal to."

The Wilcoxon-MPSR test can be used to determine whether the differences between the data registered by exporters and importers are meaningful, and if so, whether the differences are systematic in any direction. The Wilcoxon-MPSR tests the null hypothesis that there is no systematic difference within pairs against alternatives that assert a systematic difference (either one-sided or two-sided). Ignoring zero differences, the differences between the values in each pair are ranked without regard to sign. Then the sums of the positive ranks (R^+) and of the negative ranks (R^-) are calculated. For a two-tail test, the smaller of R^+ and R^- is called T. This T is the statistic that may be compared with the critical values in the appropriate statistical table (see Table 7). For one-tailed tests, Twill take the value of R^+ or R^- , depending of the specification of the alternative hypothesis. In plain language, if the null hypothesis was true and there was no difference between the two series compared, then we would expect the rank sums for positive and negative ranks to be the same, i.e. to have as many large positive as negative differences

³ Do not confuse the Wilcoxon Matched-Pairs Signed-Ranks test with the other main test by the same author, the Wilcoxon Rank-Sum test, which compares one group with a hypothetical median.

and as many small positive and negative differences. For the difference to be significant (i.e. to reject the null hypothesis) the calculated *T* must be less than or equal to the tabulated value. Note that the Wilcoxon-MPSR *T* statistic has a sampling distribution that is approximately normal when the number of pairs is large – say, $n \ge 15$ --, close enough to allow for the calculation of a z-ratio, which can then be referred to the unit normal distribution, for the approximation formulae see Quang and Hong (2000).

Friedman (1937) when pondering nonparametric methods stated that 'it is evident that the method of ranks does not utilize all the information furnished by the data, since it relies solely on the order and makes no use of the quantitative magnitude of the variance. It is this very fact that makes it independent of the assumption of normality. At the same time, it is desirable to obtain some notion about the amount of information lost, that is about the efficiency of the method of ranks'. In the same paper, Friedman concluded that the loss of information in using the method of ranks is not very great. A positive aspect of the Wilcoxon-MPSR test is that it is a very powerful test. If all the assumptions for the parametric tests were met the Wilcoxon-MPSR has about 95 percent of the power of the parametric alternative. Further detail about the calculation and interpretation of the Wilcoxon-MPSR test can be found in Bland (1995) and Conover (1998).

5. THE RESULTS

Two different specifications of the Wilcoxon-MPSR test were designed for the data of Latin American imports of fossil fuels in 1925, one for the quantities and another one for the values. The test for quantities had no prior opinion regarding the direction of the mismatch: the data provided by the importers could either overvalue or undervalue the data registered at the ports of origin. Simply, this test goes along the line that the Latin American statistics are unusable and the errors could go in any direction. Thus in the first

test the null hypothesis is H_0 : there is no systematic difference between importers' and exporters' tonnage data, versus an alternative H_1 : importers' tonnage records are different (greater or smaller) than exporters' data. The second test matched the usual assumptions explaining the discrepancies in values: data provided by the importers may be larger in value due to the difference between CIF and FOB registrations. Therefore in the second tests, applied to the value data, tests the null hypothesis H_0 : there is no systematic difference between importers' and exporters' value data, versus an alternative H_2 : importers' value records are larger than the values registered by the exporters.

Table 5 starts by reporting the STATA results when the Wilcoxon-MPSR test was applied to the total flows of trade measured in quantities (tonnage data in Tables 3 and 4). The column on the left summarizes the results for petroleum while the coal results appear on the column on the right hand side. The results of this first test for the quantities performed over the total flows are very clear: the null hypothesis cannot be rejected either for petroleum or for coal data at the levels of *p* specified (*p*=0.05 and *p*=0.01). That is the Wilcoxon-MPSR test found no compelling evidence that the tonnage data offered by exporters and importers differ when the flows of trade to each country are considered.

Furthermore, Table 5 also extended the test to the data in Tables 1 and 2. That is testing for individual exporters (for instance, exports of Mexico registered at Mexico contrasted with every country's record of Mexican oil imports). Such a test presumes that the discrepancies among the pairs are independent within exporters, and only tests for those cases where data is known at both ends and sufficient pairs are available (N>5). Again the results are very encouraging, for only in one case, the trade of petroleum from Germany to the Latin American countries, the null hypothesis can be rejected with confidence. Nevertheless, the trade of oil from Germany was insignificant for the region in absolute levels. Actually, it is due to its small magnitude that the data recorded at both

ends differ: Germany did not report small trade while small countries did record this imports. In all other cases, for the tonnage of petroleum and coal imported, the null hypothesis cannot be rejected. Therefore, the Wilcoxon-MPSR test concludes that in the contrast of the Latin American imports of fossil fuels with each of the main trading partners, no systematic difference can be found in the tonnage data.

The expectation in the case of values was for rejecting the null hypothesis in most cases due to the assumptions of the literature: data provided by the importers may be larger in value due to the difference between CIF and FOB registrations. The results however point in a different direction. Table 6 present first the results of applying the Wilcoxon-MPSR test to the flows of trade in values given in Tables 3 and 4, and then extend the test to the pairs of values reported by each exporter (Tables 1 and 2). As in Table 5, the column on the left summarizes the results for petroleum while the coal results appear on the column on the right hand side.

The results for the contrast of the value of petroleum trade are striking. Again as in the case of quantities of petroleum traded, only the German data rejects the null hypothesis. For all the other data, including the total flows and the individual cross-checks for each exporter, the null hypothesis cannot be rejected. According to the Wilcoxon-MPSR test there is no systematic difference between importers' and exporters' reports of value traded in petroleum.

In the case of the value of coal traded, the null hypothesis is rejected in two relevant and related cases: in the aggregated flow of trade and in the trade from at the United Kingdom. The rejection on the null hypothesis, thus concluding that importers' value records are larger than the values registered by the exporters in the case of coal trade, seems more in accordance with the expectations of the literature. The issue then is why it is only found in these two cases and neither in the petroleum trade nor in the value of coal

traded from the United States or Germany. The answer may lay on the different composition of the trade flows and the characteristics of the trade statistics of the destination countries. While almost two thirds of the coal trade had as destination countries applying CIF valuations and official prices, not even half of the oil trade had such countries as destination. In fact, most of the trade originating in the United States had as destination countries using FOB valuations in their imports and declared prices, which explains the test result of no systematic differences between the United States exports data and the Latin American importing records. On the contrary, over sixty per cent of the trade in coal that originated in the United Kingdom had as destination one single country Argentina, a country where imports were valued CIF and official prices were used. In addition, the omission by the British of the petty trade added to the overvaluation on the Latin American side.

In summary, the result of the Wilcoxon-MPSP test for the value of coal traded can be explained by the distinct composition of the coal trade flows and the characteristics of the trade statistics of the main destination countries. Nevertheless, these finding do not invalidate the fact that Argentina was the larger importer of coal of the region in 1925, whichever source used. It simply adds a caveat over the value given to such imports at the port of destination, especially if coming from the United Kingdom. As it does for the rest of the imports of coal, which according to these results, were overvalued with respect the value at origin.

The conclusion of the several exercises performed in this section is that only in very few cases we can accept the existence of statistically significant differences between the data provided by the exporters and the registered by the importing countries, and these only in value, never in volume. Given the nature of the products traded, the results are not so surprising. It should be pointed out that some works mention the fact that imports

figures respond better to accuracy tests than export figures, for instance Federico and Tena (1991), Yeats (1995) and Kuntz (2002). Nevertheless, it should noticed that here the exports of petroleum of two Latin American countries, Mexico and Peru, have also passed the test. No significant difference can be found between the export reports of these two countries and the imports reported from them by the rest of the region, either in value or quantities.

CONCLUSIONS

The paper was aimed at proving that the story told from the exporter side is almost identical to the story told using the importers' figures when analyzing the Latin American countries within the context of the region. The question addressed, whether the differences observed are statistically (and economically) meaningful for the interpretation of the imports of petroleum and coal of the Latin American countries relative to each other found an answer making use of the Wilcoxon-MPSR test. The answer is that in quantities the differences are always unimportant statistically and economically. In values, the several exercises performed showed that only in very few cases we can accept the existence of statistically significant differences between the data provided by the exporters and the registered by the importing countries. The differences found are restricted to the values registered in coal trade, mostly trade from the United Kingdom, where the test rejected the null hypothesis thus concluding that the Latin American values were larger than the values registered at port of origin. It is worth remembering that 'how large is large' always depends on the question asked. Differences that at the country level may look abysmal, placed in the context of the region will be a minor problem for the analysis at hand.

This endorses the view that foreign trade statistics are more robust than previously thought for providing the basis of economic analysis. It is possible to extract some lessons

from the Latin American trade data that withstand the mistrust on the original data sources. The results also point that the trade data of industrialized countries are compatible with the Latin American trade statistics and may serve as a reasonable supplement and complement to the regions' trade data. The historiography of Latin American describes its foreign trade statistics as mostly unusable. Our research and quantitative results contest this view.

Further research shall expand these results to other products in order to prove whether accuracy was restricted to homogeneous products. The examination shall also extend the time frame in order to explore whether the 1920s were an exceptionally good period for trade reporting. It would also be good to be able to test individual importers with respect to their trade partners (say Argentina's imports of petroleum versus all its suppliers rather than just the main ones) in order to study individual biases of trade. Researchers of using historical trade statistics may find useful some of the challenges to the literature that the scrutiny of the Latin American trade statistics brought into light: while it seems generally true that the exporters report values FOB, not all the importers reported their imports CIF in the 1920s, actually less than half of the 17 countries examined here, plus the United States. Finally, the wider field of economic measurement (in)accuracy may also benefit from the nonparametric test used here, since in most cases, economic data fit better the assumptions of nonparametric inferential statistics than the most widely used parametric ones.

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Table 1 Pairs of registered trade by exporters and importers: Petroleum products imports of Latin America in 1925, by principal origin

| | | Topos ovported | Topos imported | U.S. dollars

 | U.S. dollars | | |
|---|--|--|--
--
--
--
--
--
---|--|--|---|
| Country of origin | Country of | Tones exported | rones importeu | exported

 | imported | | |
| Country of origin | destination | of origin) | of destination) | (source: country

 | (source: country of | IMML LONES | THINE Value |
| | | or origin) | or descributor) | of origin)

 | destination) | | |
| United States | Argentina | 338.920 | 271.606 | 14.023.793

 | 18.555.730 | -11% | 14% |
| United States | Bolivia | 5.408 | 7.262 | 207.594

 | 261.431 | 15% | 11% |
| United States | Brasil | 243,176 | 210.200 | 15.907.927

 | 15.216.036 | -7% | -2% |
| United States | Colombia | 11 144 | 3 600 | 048 844

 | 354 043 | -51% | -46% |
| United States | Contra Dise | 11.177 | 1.007 | 200.10

 | 224.730 | 760/ | 160/ |
| United States | Costa Rica | 11.//4 | 1.597 | 308.105

 | 224.728 | -70% | -10% |
| United States | Cuba | 382.192 | 390.112 | 8.318.516

 | 7.137.420 | 1% | -8% |
| United States | Chile | 766.852 | 700.859 | 8.655.373

 | 6.785.093 | -4% | -12% |
| United States | Ecuador | 16.660 | 1.459 | 337.918

 | 210.713 | -84% | -23% |
| Inited States | FI Salvador | 9.841 | 19,902 | 297.011

 | 385,263 | 34% | 13% |
| United States | Guatomala | 2 172 | 21.005 | 212 244

 | 277 007 | 87% | 28% |
| | Guaternala | 2.1/5 | 51.005 | 212.277

 | 377.007 | 20/ 70 | 2070 |
| United States | Haiti | 5.165 | 5.446 | 395.072

 | 419.584 | 3% | 3% |
| United States | Honduras | 63.042 | • | 1.109.116

 | • | | |
| United States | Mexico | 324.314 | 361.357 | 6.233.289

 | 6.226.786 | 5% | -0,1% |
| United States | Nicaragua | 9.140 | 7.246 | 582.948

 | 299.470 | -12% | -32% |
| United States | Panama | 612.898 | | 6.815.130

 | | | |
| United States | Paraquay | 191 | | 19,217

 | | | |
| United States | Peru | 6 743 | 7 443 | 850 100

 | 018 867 | 5% | 3% |
| United States | P Dominicana | 41.040 | 20.007 | 1 522 542

 | 042 512 | 220/- | 220/ |
| | R.Dominicana | 41.940 | 20.987 | 1.522.542

 | 943.512 | -33% | -23% |
| United States | Uruguay | 112.157 | • | 4.745.413

 | • | | |
| United States | Venezuela | 14.021 | | 771.165

 | | | |
| | <u> </u> | Tones exported | Tones imported | U.S. dollars

 | U.S. dollars | | |
| Country of origin | Country of | (source: country | (source: country | exported

 | imported | IMME tones | IMME value |
| -same, or origin | destination | of origin) | of destination) | (source: country

 | (source: country of | | |
| Marrian | A | | | of origin)

 | destination) | 0.401 | F14. |
| Mexico | Argentina | 125.847 | 204.880 | 4.306.868

 | 13.367.192 | 24% | 51% |
| Mexico | Brasil | 308.234 | 295.359 | 3.582.710

 | 5.773.082 | -2% | 23% |
| Mexico | Costa Rica | 23.698 | 37.155 | 218.143

 | 255.629 | 22% | 8% |
| Mexico | Cuba | 970.205 | 891.830 | 8.782.727

 | 5.960.089 | -4% | -19% |
| Mexico | Chile | 58 546 | 115 261 | 944 807

 | 1,906,701 | 33% | 34% |
| Mexico | El Salvador | 100 | 10 | 9.010

 | 1.500.751 | -83% | -77% |
| Mexico | Customala | 190 | 24.200 | 6.010

 | 1.050 | -03 /0 | -77/0 |
| Mexico | Guaternala | 47.090 | 34.335 | 610.993

 | 957.097 | -10% | 22% |
| Mexico | Honduras | 44.570 | • | 383.532

 | • | | |
| Mexico | Nicaragua | 0 | 4.650 | 0

 | 30.693 | 100% | 100% |
| Mexico | Panama | 208.446 | | 1.694.595

 | | | |
| Mexico | R.Dominicana | 4.969 | 9.797 | 46.850

 | 91.902 | 33% | 32% |
| Mexico | Uruquay | 71.526 | | 601.599

 | | | |
| Mexico | British Honduras | 863 | • | 29 327

 | • | | |
| Movico | Other Deltabal II | 70 000 | • | 20.027

 | • | | |
| Mexico | OtherBritishW.I.I. | 79.999 | • | 990.113

 | • | | |
| Mexico | Puerto Rico | 58./84 | | 502.281

 | · · · · | | |
| | Country of | Tones exported | Tones imported | U.S. dollars

 | U.S. dollars | | |
| Country of origin | Country of | (source: country | (source: country | exported

 | Imported | IMME tones | IMME value |
| , . | destination | of origin) | of destination) | (Source: country

 | (source: country of | | |
| Doru | Argontina | 202 1/2 | 200 112 | 6 736 409

 | 10 556 0/0 | -10/ | 400% |
| Peru | Argenuna | 203.143 | 200.112 | 6.736.498

 | 19.556.940 | -1% | 49% |
| D a sur s | | 2 000 | 4 077 | 110 (50

 | | - 1/1/ | 11.11/ |
| Peru | Bolivia | 2.099 | 1.977 | 118.652

 | 144.374 | -3% | 10% |
| Peru
Peru | Bolivia
Colombia | 2.099
693 | 1.977
456 | 118.652
29.981

 | 144.374
33.872 | -3%
-21% | 10%
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| Peru
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Peru | Bolivia
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Costa Rica | 2.099
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1.326 | 1.977
456
1.522 | 118.652
29.981
80.408

 | 144.374
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130.233 | -3%
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7% | 6%
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Sources: Official publications listed in the references Notes: a cero indicates no trade reported in the source of reference, a dot indicates source not available IMME={(destination source - origin source)/ (destination source + origin source)*100}.

Table 2Pairs of registered trade by exporters and importers:Coal imports of Latin America in 1925, by principal origin

Country of origin	Country of destination	Tones exported (source: country of origin)	Tones imported (source: country of destination)	U.S. dollars exported (source: country of origin)	U.S. dollars imported (source: country of destination)	IMME tones	IMME value
United Kingdom	Argentina	2.444.917	2.768.735	15.685.196	29.954.051	6%	31%
United Kingdom	Bolivia	564	4.614	6.688	53.359	78%	78%
United Kingdom	Brasil	1.195.484	1.081.395	7.367.011	9.732.788	-5%	14%
United Kingdom	Colombia	1.429	1.609	9.484	11.789	6%	11%
United Kingdom	Costa Rica	0	312	0	3.384	100%	100%
United Kingdom	Cuba	6.814	7.366	59.836	52.481	4%	-7%
United Kingdom	Chile	143.736	208.829	768.719	967.597	18%	11%
United Kingdom	Ecuador	0	77	0	878	100%	100%
United Kingdom	Mexico	0	380	0	9.324	100%	100%
United Kingdom	Nicaragua	0	357	0	382	100%	100%
United Kingdom	Peru	10.220	17.660	69.552	158.416	27%	39%
United Kingdom	Uruguay	343.207		2.049.007			
United Kingdom	Venezuela	19.355	•	130.624			
Country of origin	Country of destination	Tones exported (source: country of origin)	Tones imported (source: country of destination)	U.S. dollars exported (source: country of origin)	U.S. dollars imported (source: country of destination)	IMME tones	IMME value
United States	Argentina	149.417	150.569	669.079	1.626.988	0%	42%
United States	Bolivia	0	1.017	0	8.932	100%	100%
United States	Brasil	596.798	631.318	2.714.739	5.402.663	3%	33%
United States	Colombia	1.696	1.643	12.458	16.180	-2%	13%
United States	Costa Rica	78	496	1.610	4.252	73%	45%
United States	Cuba	694.893	652.023	3.453.898	3.062.379	-3%	-6%
United States	Chile	40.546	38.210	225.035	200.908	-3%	-6%
United States	Ecuador	1.116	1.095	7.131	7.392	-1%	2%
United States	El Salvador	113	154	1.952	7.068	15%	57%
United States	Guatemala	3.287	264	18.524	3.279	-85%	-70%
United States	Haiti	83	156	1.275	2.771	30%	37%
United States	Honduras	1.943	•	1.368			
United States	Mexico	118.393	65.324	505.847	555.681	-29%	5%
United States	Nicaragua	2.476	2.289	11.238	11.538	-4%	1%
United States	Panama	317.994	•	1.559.705	•		
United States	Paraguay	24		432		40/	20/
United States	Peru	19.4/3	17.813	111.891	105.665	-4%	-3%
United States	R.Dominicana	9.484	9.697	69.440	//.15/	1%	5%
United States	Uruguay	20.133	•	92.240			
United States	Venezuela	639	•	0.042	•		
Country of origin	Country of destination	Tones exported (source: country of origin)	Tones imported (source: country of destination)	U.S. dollars exported (source: country of origin)	U.S. dollars imported (source: country of destination)	IMME tones	IMME value
Germany	Argentina	324.750	192.450	1.606.976	2.082.097	-26%	13%
Germany	Bolivia	100	446	714	455	63%	-22%
Germany	Brasil	19.654	474	104.482	10.540	-95%	-82%
Germany	Chile	10.515	5.995	61.642	53.249	-27%	-7%
Germany	Ecuador	15	15	238	192	1%	-11%
Germany	El Salvador	0	0,02	0	3	100%	100%
Germany	Mexico	250	43	1.190	1.788	-71%	20%
Germany	Peru	2.850	2.916	15.946	17.752	1%	5%
Germany	Uruguay	22.116	•	105.910	•		
Germany	Venezuela	643	•	2.856	•		
Country of origin	Country of destination	Tones exported (source: country of origin)	Tones imported (source: country of destination)	U.S. dollars exported (source: country of origin)	U.S. dollars imported (source: country of destination)	IMME tones	IMME value
Belgium	Argentina	6.008	225	54.625	2.431	-93%	-91%
Belgium	Brasil	2.200	2.016	12.540	13.408	-4%	3%
Belgium	Chile	400	520	2.090	5.295	13%	43%

Sources: Official publications listed in the references

Notes: a cero indicates no trade reported in the source of reference, a dot indicates source not available

 $IMME = \{ (destination \ source \ - \ origin \ source) / \ (destination \ source \ + \ origin \ source) * 100 \}.$

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Country of origin	Country of destination	Tones exported (source: countries of origin)	Tones imported (source: country of destination)	U.S. dollars exported (source: countries of origin)	U.S. dollars imported (source: country of destination)	IMME tones	IMME value
G3+Mexico+Peru	Argentina	670.046	677.187	25.388.166	51.569.067	0,5%	34%
G3+Mexico+Peru	Bolivia	7.506	9.768	326.246	481.170	13,1%	19%
G3+Mexico+Peru	Brasil	552.147	505.753	19.591.826	21.019.438	-4,4%	4%
G3+Mexico+Peru	Colombia	11.838	4.189	978.825	408.144	-47,7%	-41%
G3+Mexico+Peru	Costa Rica	36.799	40.298	606.716	614.167	4,5%	1%
G3+Mexico+Peru	Cuba	1.352.397	1.281.949	17.101.243	13.098.023	-2,7%	-13%
G3+Mexico+Peru	Chile	923.112	906.540	12.359.821	10.117.282	-0,9%	-10%
G3+Mexico+Peru	Ecuador	27.838	13.100	628.118	661.074	-36,0%	3%
G3+Mexico+Peru	El Salvador	12.072	22.549	434.678	562.864	30,3%	13%
G3+Mexico+Peru	Guatemala	50.794	68.247	902.237	1.449.398	14,7%	23%
G3+Mexico+Peru	Haiti	5.165	5.446	395.072	419.694	2,6%	3%
G3+Mexico+Peru	Mexico	324.330	361.448	6.237.097	6.241.741	5,4%	0%
G3+Mexico+Peru	Nicaragua	11.639	14.615	705.560	491.826	11,3%	-18%
G3+Mexico+Peru	Peru	6.743	8.006	859.199	996.424	8,6%	7%
G3+Mexico+Peru	R.Dominicana	46.908	30.784	1.569.392	1.035.500	-20,8%	-20%
G3+Mexico+Peru	Uruguay	183.684	226.045	5.347.012	7.381.505	10,3%	16%
G3+Mexico+Peru	Venezuela	14.021	8.552	771.165	798.032	-24,2%	2%

Table 3	
Petroleum products imports of Latin America in 1925	totals from principal partners

Sources and notes: Table 1, exept for the total quantities and values imported by Uruguay and Venezuela, which correspond to the total imports reported in their respective trade statistics

G3 refers to United States, United Kingdom and Germany

IMME={(destination source - origin source)/ (destination source + origin source)*100}.

Coal imports of Latin America in 1925, totals from principal partners								
Country of origin	Country of destination	Tones exported (source: countries of origin)	Tones imported (source: country of destination)	U.S. dollars exported (source: countries of origin)	U.S. dollars imported (source: country of destination)	IMME tones	IMME value	
Total G4	Argentina	2.925.091	3.111.979	18.015.876	33.665.568	3%	30%	
Total G4	Bolivia	664	6.077	7.402	62.747	80%	79%	
Total G4	Brasil	1.814.136	1.715.203	10.198.772	15.159.399	-3%	20%	
Total G4	Colombia	3.125	3.252	21.942	27.970	2%	12%	
Total G4	Costa Rica	78	808	1.610	7.636	82%	65%	
Total G4	Cuba	701.707	659.389	3.513.734	3.114.860	-3%	-6%	
Total G4	Chile	195.197	253.554	1.057.486	1.227.050	13%	7%	
Total G4	Ecuador	1.131	1.187	7.369	8.462	2%	7%	
Total G4	El Salvador	113	154	1.952	7.071	15%	57%	
Total G4	Guatemala	3.287	264	18.524	3.279	-85%	-70%	
Total G4	Haiti	83	156	1.275	2.771	30%	37%	
Total G4	Mexico	118.643	65.746	507.037	566.793	-29%	6%	
Total G4	Nicaragua	2.476	2.646	11.238	11.920	3%	3%	
Total G4	Peru	32.542	38.389	197.389	281.834	8%	18%	
Total G4	R.Dominicana	9.484	9.697	69.440	77.157	1%	5%	
Total G4	Uruguay	385.457	352.531	2.245.157	3.468.769	-4%	21%	
Total G4	Venezuela	20.837	23.816	139.522	156.887	7%	6%	
	Total 17 countries	6.214.051	6.244.847	36.015.726	57.850.172	0,2%	23%	

Table 4

Sources and notes: Table 1, exept for the total quantities and values imported by Uruguay and Venezuela, which

correspond to the total imports reported in their respective trade statistics.

G4 refers to United States, United Kingdom, Germany and Belgium

IMME={(destination source - origin source)/ (destination source + origin source)*100}.

Rubio	Rubio & Folchi On the Accurracy of Latin American Trade Statistics							
	Table 5	A HUHPATAL			simporteu		lica	
H ₀ : there is no H ₁ : importers'	systematic difference tonnage records are	e between impo different (greate	rters and exporters er or smaller) than	s tonnage data exporters' data				
Two-ta	ail test:			•				
Te	est statistic: T =	• min (R-, R+) (in bold in the r	esults below)	,			
Кејес	ction region: $I \leq$	I crit for small	samples // foi	r large sample	es $z < -z_{\alpha l}$	$_{2} \text{ or } z > z_{\alpha/2}$		
			Quantit	ies (tones)				
	Petro	eum				Coal		
Wilcoxon si	Igned-rank test	origin: G3+M	lexico+Peru	Wilcoxon s:	igned-rank	test origin: To	orported	
positive	10	 73	 76.5	positive		89	 76.5	
negative zero	7 0	8 0 0	76.5	negative zero	5 0	64 0	76.5	
all	17	153	153	all	17	153	153	
Ho: TIMP =	TEXP z = -0.166			Ho: CTIMP :	= CTEXP z = 0.	592		
Prob >	z = 0.868	4	NO REJECT**	Prob >	z = 0.	5540	NO REJECT**	
Wilcoxon s	igned-rank tes	t origin: UM	ζ	Wilcoxon s:	igned-rank	test origin:UK		
sign	obs su	m ranks(R)	expected	sign	obs	sum ranks(R)	expected	
positive	12	80	60	positive negative	10 1	56 10	33 33	
zero	0 	40 0	0	zero	0	0	0	
all	15	120	120	all	11	66	66	
Ho: TIMP =	TEXP z = 1.136	0		Ho: CTIMP :	= CTEXP z = 2.	045		
Prob >	Z = 0.256	0	NO REJECT**	Prob >	z = 0.	0409	NO REJECT*	
Wilcoxon s	igned-rank tes	t origin: GB	ER	Wilcoxon s:	igned-rank	test origin:GEI	S	
sign	obs su	m ranks(R)	expected	sign +	200 	sum ranks(k)	 18	
positive negative	13 1	96 9	52.5 52.5	negative	4 0	25	18 0	
2010 + all	 14		 105	all	8	36	36	
Ho: TIMP =	TEXP	100	200	HO, CTIMP	- CTEXP			
Prob >	z = 2.731 z = 0.006	3	DE IECT**	Prob >	$\begin{vmatrix} z \\ z \end{vmatrix} = 0.$	980 3270		
Wilcoxon s	igned-rank tes	t origin: US	SA KEJECT	Wilcoxon s	igned-rank	test origin:USA	NO REJECT**	
sign	obs su	m ranks(R)	expected	sign	obs	<pre>sum ranks(R)</pre>	expected	
positive negative	7 8	43 77	60 60	positive negative	7 8	49 71	60 60	
zero	0	0	0	zero	0	0	0	
all No: TIMD -	15 TEVD	120	120	all	15	120	120	
Prob >	z = -0.966	3		Ho: CTIMP :	= CTEXP z = -0.	625		
1102 /	121 01001		NO REJECT**	Prob >	z = 0.	5321	NO REJECT**	
Wilcoxon s	igned-rank tes	t origi n: MB	EX					
sign + positive	ש צ צמט ק	ш гапкз(К) 27	expected 22 5					
negative zero	4 0	18 0	22.5					
all	9	45	45					
Ho: TIMP =	TEXP			Sources and	notes:			
Prob >	z = 0.535	0	NO REJECT**	Data in quant	tities from Tab	oles 1 to 4.		
Wilcoxon s	igned-rank tes	t origin:PEF	2	Variable label	S: moorted (not	roleum)		
sign	obs su	m ranks(R)	expected	TEXP: tones	exported (peti	roleum)		
positive	5 4	23 22	22.5	CTIMP: tones	s imported (co	al) al)		
zero	Ö	0	<u>0</u>	* result hold	ds at p=0.01 l	out reject at p=0.0	5	
all	9 TEVD	45	45	** result hold	ls at both leve	els p=0.05 and p=0	0.01	
HU: TIMP =	z = 0.059	8						
- LOD >	121 - 0.952	0	NO REJECT**					

Rubio & Folchi Table 6	On the Accurracy of Latin American Trade Statistics A nonparametric test for values imported by Latin America
H₀: there is no systematic difference H₂: importers' value records are lar Two-tail test: Test statistic: T = Rejection region: T≤	The between importers and exporters tonnage data to the values registered by exporters R = (in bold in the results below) T_{crit} for small samples // for large samples $z > z_{z}$ (or $z < -z_{z}$)
	Values (dollars)

Values	(dollars)
Petroleum	Coal
Wilcoxon signed-rank test origin: G3+Mexico+Peru	Wilcoxon signed-rank test origin: Total G4
sign obs sum ranks(R) expected	sign obs sum ranks(R) expected
positive 12 91 76.5 negative 5 62 76.5 zero 0 0 0	positive 15 131 76.5 negative 2 22 76.5 zero 0 0 0
all 17 153 153	all 17 153 153
Ho: VIMP = VEXP z = 0.686	Ho: $CVIMP = CVEXP$ z = 2.580
Prob > z = 0.4925 NO REJECT**	Prob > z = 0.0099 REJECT**
Wilcoxon signed-rank test origin: UK	
aign by obs gum ronks (B) expected	Wilcoxon signed-rank test origin: UK
sign obs sum ranks (k) expected	sign obs sum ranks(R) expected
positive 12 80 60	
zero 0 0 0	negative 1 5 33
all 15 120 120	zero 0 0 0
	all 11 66 66
HO: VIMP = VEXP z = 1.136	HO: $CVIMP = CVEXP$ z = 2.490
Prob > z = 0.2560	Prob > z = 0.0128 REJECT**
Wilcoxon signed-rank test origin: GER	Wilcoxon signed-rank test origin: GER
	sign obs sum ranks(R) expected
sign obs sum ranks(k) expected	nogitive / 18 18
positive 13 99 52.5 negative 1 6 52.5 zero 0 0 0	negative 4 18 18 zero 0 0 0
all 14 105 105	all 8 36 36
Ho: VIMP = VEXP	Ho: $CVIMP = CVEXP$
z = 2.919 Prob > $ z = 0.0035$	Prob > z = 1.0000
REJECT**	NO REJECT
wilcoxon signed-rank test origin: USA	wilcoxon signed-rank test origin: USA
sign obs sum ranks(k) expected	sign obs sum ranks(k) expected
positive 6 38 60	positive 11 79 60
$\begin{array}{c c} 1 & 1 \\ 1 & 2 \\ 2 \\ 2 \\ 2 \\ 2 \\ 1 \\ 2 \\ 2 \\ 1 \\ 2 \\ 2$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$
all 15 120 120	all 15 120 120
Ho: VIMP = VEXP $z = -1.250$	Ho: $CVIMP = CVEXP$
Prob > z = 0.2115	Prob > z = 0.2805
Wilcoxon signed-rank test origin: MEX	NO REJECT
sign obs sum ranks(R) expected	
positive 7 36 22.5	
negative 2 9 22.5	
all 9 45 45	
Ho: VIMP = VEXP	
z = 1.599 Prob > $ z = 0.1097$	Sources and notes:
Wilcoxon signed-rank test origin: PER	Data in values from Tables 1 to 4.
sign obs sum ranks(R) expected	Variable labels:
	VIMP: tones imported (petroleum)
negative 8 3/ 22.5	VEXP: tones exported (petroleum)
zero 0 0 0	CVIMP: tones imported (coal)
all 9 45 45	CVEXP: tones exported (coal) * result holds at p=0.01 but reject at p=0.05
Ho: VIMP = VEXP	** result holds at both levels p=0.05 and p=0.01
z = 1.718 Prob > $ z = 0.0858$	
NO REJECT*	

Table 7: Critical values of the Wilcoxon Matched Pairs Signed Rank Test

For any N (number of subject minus ties) the observed value is significant at a given level of significance if it is equal to or less than the critical value shown in table below.

(One-tailed Test	Two-taile	ed Test	
N	p≤0.05	p≤0.01	p≤0.05	p≤0.01
5	1	-	-	-
6	2	-	1	-
7	4	0	2	-
8	6	2	4	0
9	8	3	6	2
10	11	5	8	3
11	14	7	11	5
12	17	10	14	7
13	21	13	17	10
14	26	16	21	13
15	30	20	25	16
16	36	24	30	19
17	41	28	35	23
18	47	33	40	28
19	54	38	46	32
20	60	43	52	37

Source: Statistical Tables, School of Psychology, University of Nottingham, University Park, Nottingham, NG7 2RD, UK, available at:

http://www.psychology.nottingham.ac.uk/courses/modules/statsguides/StatisticalTables.htm. The first column also serves for the critical values of the two-tailed test at $p \le 0.10$

