



BANCA D'ITALIA
EUROSISTEMA

Temi di discussione

(Working papers)

Innovation and productivity in SMEs.
Empirical evidence for Italy

by Bronwyn H. Hall, Francesca Lotti and Jacques Mairesse

June 2009

Number

718



BANCA D'ITALIA
EUROSISTEMA

Temi di discussione

(Working papers)

Innovation and productivity in SMEs.
Empirical evidence for Italy

by Bronwyn H. Hall, Francesca Lotti and Jacques Mairesse

Number 718 - June 2009

The purpose of the Temi di discussione series is to promote the circulation of working papers prepared within the Bank of Italy or presented in Bank seminars by outside economists with the aim of stimulating comments and suggestions.

The views expressed in the articles are those of the authors and do not involve the responsibility of the Bank.

Editorial Board: PATRIZIO PAGANO, ALFONSO ROSOLIA, UGO ALBERTAZZI, ANDREA NERI, GIULIO NICOLETTI, PAOLO PINOTTI, ENRICO SETTE, MARCO TABOGA, PIETRO TOMMASINO, FABRIZIO VENDITTI.

Editorial Assistants: ROBERTO MARANO, NICOLETTA OLIVANTI.

INNOVATION AND PRODUCTIVITY IN SMEs. EMPIRICAL EVIDENCE FOR ITALY*

Bronwyn H. Hall,[#] Francesca Lotti,[§] and Jacques Mairesse[♦]

Abstract

Innovation in SMEs exhibits some peculiar features that most traditional indicators of innovation activity do not capture. Therefore, in this paper, we develop a structural model of innovation which incorporates information on innovation success from firm surveys along with the usual R&D expenditures and productivity measures. We then apply the model to data on Italian SMEs from the “Survey on Manufacturing Firms” conducted by Mediocredito-Capitalia covering the period 1995-2003. The model is estimated in steps, following the logic of firms’ decisions and outcomes. We find that international competition fosters R&D intensity, especially for high-tech firms. Firm size, R&D intensity, along with investment in equipment enhances the likelihood of having both process and product innovation. Both these kinds of innovation have a positive impact on firm’s productivity, especially process innovation. Among SMEs, larger and older firms seem to be less productive.

Keywords: R&D, innovation, productivity, SMEs, Italy.

JEL Classification: L60, O31, O33.

Contents

1. Introduction	5
2. Previous studies of the innovation – productivity link	7
3. Data and main results	9
3.1 Descriptive statistics	9
3.2 Data and model specification	11
3.3 The R&D equations	12
3.4 Innovation equations	15
3.5 The productivity equation	16
3.6 Investment and innovation	18
3.7 Further robustness checks	19
3.8 Comparison to Griffith et al. (2006)	20
4. Conclusions and policy discussion	22
References	25
Tables and figures	28
Appendix	36

[#] Department of Economics, University of California at Berkeley; Maastricht University - The Netherlands; NBER - Cambridge, Mass. USA; IFS – London, UK.

[§] Bank of Italy, Department for Structural Economic Analysis.

[♦] CREST (ENSAE, Paris) – France; UNU-MERIT, Maastricht University - The Netherlands; NBER - Cambridge, Mass.

1. Introduction*

In the past decade, labor productivity growth in Italy has been one of the lowest in the EU; low growth has been particularly strong in manufacturing, where the growth rate even turned negative in the period from 2000 to 2005 (see Figure 1). Such a poor performance raises unavoidable policy concerns about the underlying reasons for it. Is the labor productivity slowdown due to the decline in total factor productivity (see Daveri and Jona-Lasinio 2005)? Or, more precisely, is it a consequence of the exhaustion of the so-called “capital deepening” phase that supported labor productivity growth during the Eighties (as documented by Pianta and Vaona, 2007)? Alternatively, is it simply due to input reallocation following a change in the relative price of labor with respect to capital after the labor market reforms of the early 1990s (Brandolini et al., 2007)? Or does the explanation lie in the evergreen *motto* that Italian firms exhibit insufficient R&D investment (European Commission, 2006)?

The latter aspect has been largely explained by the unquestionable fragmentation of the Italian production system. According to the latest available data from the Census, more than 99 % of active firms (out of 4 million) have fewer than 250 employees (95 % have fewer than 10 employees, see Figure 2). If there were a positive relationship between innovation activity – including R&D – and firm size, the size distribution of Italian firms could help to explain why Italy is lagging behind in terms of aggregate R&D investment.

Nevertheless, many scholars have argued that small firms are the engines of technological change and innovative activity, at least in certain industries (see the series of works by Acs and Audretsch, 1988, 1990). But at the same time, innovation in small

* We would like to thank the Unicredit (former Mediocredito-Capitalia) research group for having kindly supplied firm level data for this project, in particular Attilio Pasetto. We thank also Susanto Basu, Ernie Berndt, Piergiuseppe Morone, Stéphane Robin, Mike Scherer, Enrico Santarelli, Alessandro Sembenelli, Marco Vivarelli, and participants at the NBER Productivity Seminars, at the workshop “Drivers and Impacts of Corporate R&D in SMEs” held in Seville at IPTS and seminars at the Bank of Italy. The views expressed by the authors do not necessarily reflect those of the Bank of Italy.

and medium enterprises exhibits some peculiar features that most traditional indicators of innovation activity would not capture, incurring the risk of underestimating their innovation effort. In fact, innovation often occurs without the performance of formal R&D, and this is particularly true for SMEs. Despite the existence of a large number of policies designed to promote and facilitate the operation of the innovation process within SMEs, especially in Italy, the knowledge about how SMEs actually undertake innovative activities remains quite limited, causing a significant bias in the treatment of the R&D – innovation relationship (see Hoffman et al, 1998 for a literature review on this topic in the UK).

This paper is not an attempt to verify or disprove the Schumpeterian hypothesis, i.e. to study the relationships between firm size and innovative activity at the firm level; instead it investigates how and when innovation takes place in SMEs and whether – and how – innovation outcomes impact SME firms' productivity. We caution the reader that because we rely mainly on dummy variables for the present of innovation success, we are in fact unable to say very much about the size-innovation relationship *per se*. In general larger firms have more than one innovative activity, which implies a higher probability that one of them at the least is successful and that the innovation dummy is one.

The remainder of the paper is organized as follows. In Section 2 we put our modeling approach and our results into perspective by giving a summary – far from being exhaustive – of the previous empirical studies on the R&D innovation–productivity relationships. In Section 3, we first explain our data and how we bring it into play in our modeling approach; we then present in turn our main results on the R&D investment equations, the innovation equations and the productivity equation; and finally we discuss and give evidence on the robustness of these results and compare them to the comparable findings of Griffith and al. (2006). In Section 4 we conclude with a discussion of the results and with directions for further research.

2. Previous studies of the innovation – productivity link

Measuring the effects of innovative activities on firms' productivity has been an active area for research for several decades both as a policy concern and as a challenge for econometric applications. Notwithstanding a large number of empirical studies available, measuring the effect of innovation (product and process) on productivity at the firm level (see Griliches, 1995), the literature still does not provide a unique answer in terms of the magnitude of this impact. Because of the variability and uncertainty that is inherent in innovation, this fact is not unexpected: at best, economic research should give us a distribution of innovation outcomes and tell us how they have changed over time. Recent firm level studies, Lichtenberg and Siegel (1991) on the U.S., Hall and Mairesse (1995) and Mairesse and Mohnen (2005) on France, Harhoff (1998) and Bönte (2003) on Germany, Klette and Johansen (1996) on Norway, Van Leeuwen and Klomp (2006) on the Netherlands, Janz *et al* (2004) on Germany and Sweden, Lööf and Heshmati (2002) on Sweden, Lotti and Santarelli (2001) and Parisi *et al* (2006) on Italy, find that the effect of R&D on productivity is positive,¹ although some have suggested that the returns to R&D have declined over time (Klette and Kortum, 2004). The majority of the empirical analyses rely on an extended production-function approach, which includes R&D (or alternative measures of innovation effort) as another input to production.

However, it is widely recognized that R&D does not capture all aspects of innovation, which often occurs through other channels. This is particularly true for small and medium-size firms, and could lead to a severe underestimation of the impact of innovation on productivity. In order to overcome this problem, subsequent studies have moved from an input definition of innovation activities to an *output approach*, by including in the regressions the outcome of the innovation process rather than its input. The rationale behind this line of reasoning is simple: if it is not possible to measure the innovative effort a firm exerts because of the presence of latent and unobservable

¹ For a survey of previous empirical results, see Mairesse and Sassenou (1991) and Griliches (1998).

variables, one should look at the results of R&D investment: training, technology adoption, sales of products new to the market or the firm. All these activities may be signs of successful innovative effort, but if one considers R&D only, a lot of this informal activity is going to be missing from the analysis (Blundell et al., 1993, Crépon, Duguet and Mairesse, 1998). As suggested by Kleinknecht (1987), official R&D measures for SMEs may underestimate their innovation activities, and the underestimate is likely to be larger at the left end of the firm size distribution.

Crépon, Duguet and Mairesse (1998) take a further step in this literature combining the aforementioned approaches. They propose and estimate a model – CDM model hereinafter - that establishes a relationship among innovation input (mostly, but not limited, to R&D), innovation output and productivity. This structural model allows a closer look at the black box of the innovation process at the firm level: it not only analyzes the relationship between innovation input and productivity, but it also sheds some light on the process in between the two.

The CDM approach is based on a simple three-step modeling of the logic of firms' innovation decisions and outcomes. The first step corresponds to firm decision whether to engage in R&D or not and on how much resources to invest. Given the firm's decision to invest in innovation, the second step consists in a knowledge production function (as in Pakes and Griliches, 1984) which relates innovation output to innovation input and other factors. In the third step, an innovation augmented Cobb-Douglas production function specifies the effect of innovative output on the firm's productivity. The model is tailored to take advantage of innovation survey data, which provide measures of other aspects of innovation and not only on R&D expenditures. Given the increased diffusion of this type of micro data across countries and among scholars, many empirical explorations of the impact of innovation on productivity have relied on the CDM framework.²

² See Hall and Mairesse (2006) for a comprehensive survey. Recent papers based on the CDM model include Benavente (2006) on Chile, Heshmati and Löf, (2006) on Sweden, Jefferson et al. (2006) on China, Klomp and Van Leeuwen (2001) on the Netherlands, Mohnen, Mairesse and Dagenais (2006) on seven European countries, and Griffith, Huergo, Mairesse and Peters (2006) on four European countries.

In particular, Parisi, Schiantarelli and Sembenelli (2006) apply a modified version of the CDM model to a sample of Italian firms (using two consecutive waves of the Mediocredito-Capitalia survey, the same source we are using in our empirical analysis), enriching the specification with a time dimension.³ They find that process innovation has a large and significant impact on productivity and that R&D is positively associated with the probability of introducing a new product, while the likelihood of having process innovation is directly linked to firm's investment in fixed capital. In comparing those results to the ones we obtain in this paper, one has to keep in mind that, due to the design of the survey itself, the panel used by Parisi, Schiantarelli and Sembenelli is tilted towards medium and large firms much more than the original Mediocredito-Capitalia sample.

To our knowledge, none of the empirical papers investigating the relationship between innovation and productivity has dealt specifically with small and medium-sized firms. On one hand, this paper is aimed at filling this gap, since innovation in SMEs is even more difficult to measure; on the other, like Griffith et al (2006), we try to improve on the CDM original specification by considering separately both product and process innovation.

3. Data and main results

3.1. Descriptive statistics

The data we use come from the 7th, 8th and 9th waves of the "Survey on Manufacturing Firms" conducted by Mediocredito-Capitalia (an Italian commercial bank). These three surveys were carried out in 1998, 2001, and 2004 respectively, using questionnaires administered to a representative sample of Italian manufacturing firms. Each survey

³ Although the Mediocredito-Capitalia survey is not a panel itself, it contains repeated observation for a number of firms which is enough to allow the estimation of a dynamic framework. See Section 3 of this paper for further information on the data.

covered the three years immediately prior (1995-1997, 1998-2000, 2001-2003) and although the survey questionnaires were not identical in all three of the surveys, the questions providing the information used in this work were unaffected. All firms with more than 500 employees were included, whereas smaller firms were selected using a sampling design stratified by geographical area, industry, and firm size. We merged the data from these three surveys, excluding firms with incomplete information or with extreme observations for the variables of interest.⁴ We focus on SMEs, which represent nearly 90 % of the whole sample, imposing a threshold of 250 employees, in line with the definition of the European Commission; we end up with an overall unbalanced panel of 9,674 observations on 7,375 firms, of which only 361 are present in all three waves. Table 1 contains some descriptive statistics, for both the unbalanced and the balanced panel. Not surprisingly, in both cases, the firm size distribution is skewed to the right, with an average of respectively 50 and 53 employees and a median of respectively 32 and 36. Firms in the low-tech sector tend to be slightly smaller, with average employment of 47 and median employment of 30 (Table 2).⁵ In the unbalanced sample, 62% of the firms have successful product and/or process innovation, but only 41% invest in R&D. Such difference is evidence for the importance of informal innovation activities. Although a sizeable share of firms invests on R&D, only a small fraction seems to do it continuously: out of 361 firms in our balanced panel, 34% invest in R&D in every period under examination. For 21% of the firms product and process innovations go together, while 27% are process innovators only. Concerning competition, more than 42% of the firms in the sample have national competitors, while 18% and 14% have European and international competitors, respectively. Interestingly,

⁴ We require that sales per employee be between 2000 and 10 million euros, growth rates of employment and sales of old and new products between -150 % and 150 %, and R&D employment share less than 100 %. We also replaced R&D employment share with the R&D to sales ratio for the few observations where it was missing. For further details, see Hall, Lotti and Mairesse (2008). In addition, we restrict the sample by excluding a few observations with zero or missing investment.

⁵ We adopt the OECD definition for high- and low-tech industries. *High-tech industries*: encompasses high and medium-high technology industries (chemicals; office accounting & computer machinery; radio, TV & telecommunication instruments; medical, precision & optical instruments; electrical machinery and apparatus, n.e.c.; machinery & equipment; railroad & transport equipment, n.e.c.). *Low-tech industries*: encompasses low and medium-low technology industries (rubber & plastic products; coke, refined petroleum products; other non-metallic mineral products; basic metals and fabricated metal products; manufacturing n.e.c.; wood, pulp & paper; food, beverages & tobacco products; textile, textile products, leather & footwear).

low-tech firms tend to compete more within the national boundaries, while almost half of the high-tech firms operate in European or international markets, in line with Janz *et al.* (2004).

For comparability with the samples used by Griffith *et al.* (2006) for France, Germany, Spain and the UK, in Table A1 of the appendix we show the means for our entire sample, including non-SMEs and excluding firms with fewer than 20 employees. Even if the share of innovators – product and process – are not dissimilar, Italian firms display a significantly lower R&D intensity but roughly comparable investment intensities. These figures can be partially explained by the different firm size distribution within each country: around 60 of the firms in the Italian sample for the year 2000 belong to the smaller class size (20-49 employees), a figure much larger than that for other countries.⁶ Interestingly, labor productivity is somewhat higher for the Italian firms.

3.2 *Data and model specification*

As discussed earlier, in order to analyze the relationship between R&D, innovation and productivity at the firm level, we rely on a modified version of the model proposed by Crépon, Duguet and Mairesse (1998). This model - specifically tailored for innovation survey data and built to take into account the econometric issues that arise in this context - is made up by three building blocks, following the sequence of firms' decisions in terms of innovation activities and outcomes. The first one concerns R&D activities, i.e. the process that leads the firm to decide whether to undertake R&D projects or not, and how much to invest on R&D. The second one consists of a two-equations knowledge production function in which R&D is one of the inputs and process and product innovation are the two outputs. The third consists of a simple extended production function in which knowledge (i.e., process and product innovation) is an input.

⁶ We do not yet know how much of the difference is due to differences in sampling strategy across the different countries.

We perform our analysis for the whole sample of firms, and for high- and low- tech firms, since the effect of R&D on productivity can vary a lot with the technological content of an industry (see Verspagen, 1995 for a cross country, cross sector study and, more recently, an analysis based on micro data by Potters *et al*, 2008).

Because of the way our data, and innovation survey data in general, are collected, our analysis here is essentially cross-sectional. Although there are three surveys covering 9 years, the sampling methodology is such that few firms appear in more than one survey (as we saw in Table 1, fewer than 5 % of the firms and about 10 % of the observations are in the balanced panel). Due to the resulting small sample size and very limited information in the time series dimension we found that controlling for fixed firm effects was not really possible in practice. Other difficulties arise from the fact that the process and product innovation indicators are defined over three year periods, while the income statement data, when available, are on a yearly basis. As a robustness check we estimated the same three-equation model using lagged R&D intensity instead of contemporaneous R&D intensity in order to account for a plausible delay between R&D and innovation output. Given the low volatility of R&D investment over time, the results were very similar to those reported below.⁷

3.3 *The R&D equations*

The firm R&D decisions can modeled in terms of two equations: a selections equation and an intensity equation. The selection equation can be specified as:

$$RDI_i = \begin{cases} 1 & \text{if } RDI_i^* = w_i\alpha + \varepsilon_i > \bar{c} \\ 0 & \text{if } RDI_i^* = w_i\alpha + \varepsilon_i \leq \bar{c} \end{cases} \quad (1)$$

where RDI_i is an (observable) indicator function that takes value 1 if firm i has (or reports) positive R&D expenditures, RDI_i^* is a latent indicator variable such that firm i

⁷ Although we did not include these results in the paper for the sake of brevity, they are available from the authors.

decides to perform (or to report) R&D expenditures if they are above a given threshold \bar{c} , w_i is a set of explanatory variables affecting R&D and ε_i the error term.

The R&D intensity equation can be specified as:

$$RD_i = \begin{cases} RD_i^* = z_i\beta + e_i & \text{if } RDI_i = 1 \\ 0 & \text{if } RDI_i = 0 \end{cases} \quad (2)$$

where RD_i^* is the unobserved latent variable accounting for firm's innovative effort, z_i is a set of determinants of R&D expenditures. Assuming that the error terms in (1) and (2) are bivariate normal with zero mean and variance equal to unity, the system of equation (1) and (2) can be estimated by maximum likelihood. In the literature, this model is sometimes referred to as a Heckman selection model (Heckman, 1979) or Tobit type II model (Amemiya, 1984).

Before estimating the selection model, we performed a non parametric test for the presence of selection bias in the R&D intensity equation (see Das, Newey and Vella, 2003, and Vella, 1998 for a survey). In so doing, we first estimate a probit model in which the presence of positive R&D expenditures is regressed on a set of firm characteristics: firm size, age and their squares, a set of dummies indicating competitors' size and location, dummy variables indicating (i) whether the firm received government subsidies, and (ii) whether the firm belongs to an industrial group; the results are reported in Table A2 in the appendix. From this results, we recover for each firm the predicted probability of having R&D and the corresponding Mills' ratio, and then we estimate a simple linear regression (by OLS) for R&D intensity, including in this regression the predicted probabilities from the R&D decision equation, the Mills' ratio, their squares and interaction terms. The presence of selectivity bias is tested for by looking at the significance of those "probability terms".⁸ The results are reported in Table A2 in the appendix. As one can see, the probability terms are never significant, either singly or jointly. Therefore we adopted the linear regression (OLS) specification for the R&D intensity decision without any correction for selectivity bias.

⁸ Note that this is a generalization of Heckman's two step procedure for estimation when the error terms in the two equations are jointly normally distributed. The test here is valid even if the distribution is not normal.

In Table 3 we report the estimates performed using the pooled overall high- and low-tech samples, and including in the regression year and 2-digit industry dummies as well as “wave dummies” as controls. Wave dummies are a set of indicators for firm’s presence or absence in the three waves of the survey.⁹

Table 3 shows that the presence of EU and international competitors is strongly positively related to R&D effort: engaging in exporting activity implies investing more in R&D (see Baldwin, Beckstead, and Caves, 2002, and Baldwin and Gu, 2003, for an exploration using Canadian data), and this effect is particularly strong for high-tech firms, where competing internationally is associated with a doubling of R&D intensity. Non-exporting firms, i.e. those operating in a market that is mainly local, have, on average, lower R&D intensity.

We also find that having received a subsidy boosts significantly R&D intensity, as could be expected.¹⁰ Being part of an industrial group increases R&D intensity, but the coefficient is barely statistically significant.

We also included age classes dummies in the regression (the base group are younger firms, defined as those with less than 15 years): although the coefficients are not statistically significant, they seem to indicate that older firms may have a slightly lower incentive to do R&D than younger firms. We find also that “other things be equal” larger firms tend to do relatively less R&D per employee than small firms (the 11-20 size class), and this is particularly true for low-tech firms (for a discussion of the relationship between size and R&D investment at the firm level, see Cohen and Klepper, 1996).

⁹ For instance, a firm present in all the three waves will have a “111” code, “100” if present in the first only, “110” if in the first and in the second only, and so forth. These codes are transformed into a set of six dummies ($2^3 = 8$ minus the 000 case and the exclusion restriction).

¹⁰ Due to the large number of missing observation, we could not use a narrower definition of subsidies.

3.4 Innovation equations

In order to account for firm innovations that are not necessarily based on formal R&D activities, which are likely to be especially important for SMEs, we do not restrict estimation to R&D performing firms only. Following the original CDM model, we thus specify the innovation equation in terms of the latent R&D intensity variable, and not the observed R&D intensity. Also, like in Griffith and al. (2006), we specify separately an equation for product innovation and one for process innovation, which can thus be written as:

$$\begin{cases} PROD_i = RD_i^* \gamma + x_i \delta + u_{1i} \\ PROC_i = RD_i^* \gamma + x_i \delta + u_{2i} \end{cases} \quad (3)$$

where RD_i^* is the latent innovation effort proxied by the predicted value of R&D intensity from the first step model, x_i a set of covariates and u_{1i} and u_{2i} the error terms such that $Cov(u_{1i}, u_{2i}) = \rho$. Including the predicted R&D intensity in the regression accounts for the fact that all firms may have some kind of innovative effort, although only some of them invest in R&D and report it. Using the predicted value instead of the realized value is also a sensible way to instrument the innovative effort in the knowledge production function in order to deal with simultaneity problem between R&D effort and the expectation of innovative success.

Equation (3) is estimated as a bivariate probit model, assuming that most of the firm characteristics which affect product and process innovation are the same, although of course their impact may differ. The only exception is the investment rate, which is assumed to be related to process innovation but not to product innovation. Table 4 (as in Table 3) reports the results from the overall, and the high-tech and low- tech sample. The estimated correlation coefficient ρ is always positive and significant, which implies that process and product innovation are influenced to some extent by the same unobservable factors. Marginal effects are reported in square brackets. For an example of how to interpret these effects, the first two columns say that a doubling of predicted R&D intensity is associated with a 0.19 increase in the probability of process innovation and a 0.25 increase in the probability of product innovation.

As expected, the R&D intensity predicted by the first equation has a positive and sizeable impact on the likelihood of having product and process innovation, higher for product innovation, for all three groups of firms. Interestingly, the impact of R&D on process innovation in low tech firms is more than double that for high-tech firms (0.24 versus 0.10). Firms in low tech industries, on average, have lower R&D intensity, but their R&D effort leads to a higher probability of having at least one process innovation when compared to high-tech firms. A number of interpretations suggest themselves: one possibility is that innovating in this sector takes less R&D because it involves changes to the organization of production that are not especially technology-linked. A second related interpretation is provided by the dual role of R&D (Cohen and Levinthal, 1989): investment in research is fundamental for product innovation, but at the same time, it increases firm's ability to absorb and adopt those technologies developed somewhere else which are likely to become process innovation.

As suggested in the introduction, firm size is strongly associated with innovative success, especially among low-tech firms. Note that this result does not contradict that for R&D intensity, because innovation is measured by a dummy variable. Although larger firms may have somewhat lower R&D effort given their size, in absolute terms they do more R&D, so they have a higher probability of innovative success. Finally, with the exception of product innovation in firms older than 25 years, the age of the firm is not particularly associated with innovation of either kind.

We also note that investment intensity is positively associated with process innovation in both high and low tech firms. We defer a fuller discussion of the issues associated with the presence of investment in the process innovation equation (and also in the product innovation equation) until after we present the productivity results.

3.5 *The productivity equation*

The productivity equation is specified as a simple Cobb-Douglas technology with constant returns to scale, and with labor, capital, and knowledge inputs, which can be written as:

$$y_i = \pi_1 k_i + \pi_2 PROD_i + \pi_3 PROC_i + v_i \quad (4)$$

where y_i is labor productivity (real sales per employee, in logs), k_i is investment intensity, our proxy for physical capital, $PROD_i$ and $PROC_i$ are knowledge inputs, proxied by the predicted probability of product and process innovation. Using these predicted probabilities instead of the observed indicators is a way to address the issue of potential endogeneity (and measurement errors in variables) of the knowledge inputs. We thus generate the two predicted probabilities of innovation from the two estimated innovation equation as being respectively the probability of process innovation alone and the probability of product innovation, whether or not it is accompanied by process innovation.¹¹ Results are reported in Table 5 for specifications with and without investment as a proxy for capital; as before, estimates are reported separately for all firms, and for firms in high- and low-tech industries. Our preferred specifications, in columns (1a), (2a) and (3a), include investment intensity.

When investment is not included in the regression, i.e. in columns (1a), (2a) and (3a), process innovation displays a sizeable and positive impact on productivity. Process innovators have a productivity level approximately two and one half times that of non-innovators, *ceteris paribus*. On the contrary, when investment is included, the coefficients of process innovation are not significant. These differences clearly arise from the inclusion of the same investment variable in the process innovation equation with the consequence that process innovation in the productivity equation already encompasses the effect of investment in new machinery and equipment. However, since the investment rate is a better measure than the process innovation dummy, when both variables are included its effect tends to dominate.

Product innovation enhances productivity considerably, although to a lesser extent than process innovation when investment is included in the productivity equation. The impact is slightly stronger for firms in the high-tech firms than in the low-tech industries. Because in particular, much of product innovation is directed towards higher quality products and product differentiation, it is not surprising that it shows up quite

¹¹ The first is estimated probability of process and not product from the bivariate probit model in Table 4, and the second is the marginal probability of product innovation from the same model.

differently than process innovation in the productivity relation. Table A3 in the appendix confirms that the contribution of product innovation to productivity is much more robust to the inclusion of investment intensity in the productivity equation, included in the product and process innovation equations.

Another interesting and robust finding is that, among SMEs, relatively larger firms seem to be significantly less productive than smaller ones. It is also noteworthy, age impacts productivity negatively for firms in the high-tech industries.

3.6 *Investment and innovation*

In our preferred specification in Table 4 , we assumed that capital investment – which to a great extent means the purchase of new equipment – should contribute significantly to process innovation, but not to product innovation. In fact, we found a small marginal impact of investment on process innovation that was approximately the same for high and low-tech industries (0.05).

Because the assumption that investment is associated with process and not with product innovation may be somewhat arbitrary, we performed some robustness checks reported in Table A3 in the appendix, experimenting with different alternatives. Columns (1)-(4) of that Table reports all the possible combinations in the second step: whether investment is devoted to process innovation only (column 1), to product innovation only (column 2), to both (column 3) or to none (column 4). In the same columns we show the productivity equation, estimated using each of these different models to predict the probability of process and product innovation. In the bottom panel of the Table, we also report an alternative specification of the productivity equation without investment. Although column (1) still represents our preferred specification, column (3) suggests that physical investment has a small (0.02) positive impact on product innovation as well. Turning to the productivity equation, it can be noted that the inclusion of investment wipes out the significance of process innovation, since investment is one of its main determinants, but not of product innovation, which is more dependent on R&D investment. Excluding investment from the productivity equation reveals that the process innovation associated with investment is more

relevant for productivity than predicted product innovation (compare the process innovation coefficients for step 3 in columns 1 and 3).

3.7 Further robustness checks

The estimation method used in the body of the paper is sequential, with three steps: 1) the R&D intensity equation estimated only on firms that report doing R&D continuously; 2) a bivariate probit for process and product innovation that contains R&D predicted by the first step for all firms; and 3) a productivity equation that contains the predicted probabilities for process innovation alone and product innovation with or without process innovation. Because the last two steps contain fitted or predicted values, their standard errors will be underestimated by our sequential estimation method. In order to assess the magnitude of this underestimation, we re-estimated our preferred model specification (1a) on all firms simultaneously using maximum likelihood.

The likelihood function consists of the sum of a normal density for the R&D equation, a bivariate probit for process and product innovation, and a normal density for the productivity equation; it does not allow for correlation of the disturbances between the three blocks, although the resulting standard errors are robust to such correlation. In this likelihood function, the equation for R&D is directly entered into the innovation equations, and that for innovation probability directly into the productivity equation, so the coefficient standard errors take account of the estimation uncertainty in the first two stages.

The results of estimation on the pooled model are shown in column (2) of Tables A4 in the appendix, with the sequential estimation results in column (1) for comparison.¹² Although the results for the key coefficients are similar and have approximately the

¹² The sample size in this table is 9,014, reduced from 9,674 in the main tables of the paper due to the absence of lagged capital (beginning of year capital) for some of the observations.

same significance, the standard errors from pooled maximum likelihood are considerably larger, especially for the predicted process and product probabilities. So this should be kept in mind when interpreting the results in the tables of the paper.

Table A4 also shows the results of another experiment -- in this paper we chose to proxy capital intensity by investment intensity, in order to be comparable to the results in Griffith et al. 2006. However, in our data we also have a measure of capital available, constructed from investment using the usual declining balance method with a depreciation rate of 5 % and an initial stock from the balance sheet of the firm in 1995 or the year it entered the survey. Columns (3)-(5) of the table show the results of estimating specifications containing capital stock at the beginning of the period and using the pooled maximum likelihood method. Column (3) simply replaces investment with capital stock, while column (4) uses investment as an instrument for process innovation, but capital in the production function. Column (5) includes both investment and capital in both equations.

The results are somewhat encouraging: capital stock is clearly preferred in the production function. In fact, when it is included, investment enters only via its impact on process innovation. On the other hand, investment is a better predictor of process innovation, although capital still plays a role. However, recall that innovation is measured over the preceding three years, so that some of the investment associated with process innovation is likely to be already included in beginning of year capital. Our conclusion is that there is a strong association of process innovation with capital investment, and that such process innovation has a large impact on productivity.

3.8 Comparison to Griffith et al. 2006

The results shown in the previous section can help in shedding some light on the R&D–innovation–productivity relationship in Italian firms. Interesting insights can be gained from the differential impact of R&D on process and product innovation, as well as their different impact on productivity. Nevertheless, at this point, it is worth asking a further question: is the R&D–innovation–productivity link different for Italian firms when they

are compared to other European countries? In order to answer this question, we built a slightly different sample of firms from our data that removed firms with fewer than 20 employees and included firms with more than 250 employees.¹³ Using this sample, we are able to compare our estimates to those for France, Germany, Spain and the UK (Griffith et al, 2006). Table 6 presents the results of this comparison.¹⁴ The last two columns are for Italy, the column before the last being for the same period (1998-2000), and the last one for our overall sample for the three periods 1995-1997, 1998-2000, and 2001-2003 pooled together.

The Table shows that the results for Italy are roughly comparable with those for the other countries, but that the period (1998-2000) seems to be a bit of an outlier. We do not have an explanation for this fact other than to point out that this period corresponds to the introduction of the euro. We therefore focus on the results for the overall sample. R&D intensity is somewhat more strongly associated with process innovation than in the UK, and much less strongly than in the other countries. Investment intensity is more strongly related to process innovation than in the other countries. Also noteworthy is that for Italy, the explanatory power of the innovations equations is considerably lower.

In the productivity equation, only investment intensity enters, although product innovation has a large but insignificant impact, larger than that for any of the other countries. Together with the results for the innovation equations, this suggests that the variability in the R&D-innovation-productivity relationships is much greater for Italy than for the other countries. However, there is nothing obviously different about the relationship itself when compared to its peers in Europe, apart from the fact that R&D appears less closely linked to process innovation in Italian firms.

¹³ The overlap of this sample with the sample used in the main body of the paper is 75 %.

¹⁴ For precise comparability with the earlier paper, in this table we estimated the process and product innovation equations using single probits rather than a bivariate probit. This is consistent, but not efficient, given the correlation between the two equations.

4 Conclusions and policy discussion

In this paper we have proposed and estimated a structural model that links R&D decisions, innovation outcomes and productivity at the firm level. Based upon a modified version of the model earlier developed by Crépon, Duguet and Mairesse (1998), we were able to take into account also those firms which do not do (or report) explicitly R&D. Innovation activity, especially among small firms, can operate along several dimensions besides formal R&D.

Although preliminary, our results indicate that firm size is negatively associated with the intensity of R&D, but positively with the likelihood of having product or process innovation. We have argued that these two findings are not inconsistent, given the nature of the variables. Having received a subsidy boosts R&D efforts – or just the likelihood of reporting, more in high tech industries, even if the share of targeted firms is roughly the same in high and low tech industries (46 vs. 45 %). Given firms' unwillingness to reveal more details about the subsidies received, we can only speculate about the possibility that high tech firms are more likely to receive funding for innovation and R&D than low tech firms. International (including European) competition fosters R&D intensity, especially in high-tech firms. We find that R&D has a strong and sizeable impact on firm's ability to produce process innovation, and a somewhat higher impact on product innovation. Investment in new equipment and machinery matters more for process innovation than for product innovation.

While interpreting these results, one should keep in mind the dual nature of R&D. In fact, R&D investments contribute to develop the firm's ability to identify, assimilate, and exploit knowledge from other firms and public research organizations (Cohen and Levinthal (1989)). In other words, a minimum level of R&D activity is a necessary condition to benefit from spillovers and to appropriate public knowledge. On the other hand, more recent studies have suggested the emergence of a different knowledge paradigm, i.e. the one of innovation without research, particularly well suited for SMEs (Cowan and van de Paal, 2000), based on “the recombination and re-use of known practices” (David and Foray, 1995).

Finally, we find that product innovation has a positive impact on firms' labor productivity, but that process innovation has a larger effect via the associated

investment. Moreover, larger and older firms seem to be, to a certain extent, less productive, *ceteris paribus*.

With respect to the broader questions that motivated this investigation, we note that in most respects Italian firms resemble those in other large European countries. However that they do somewhat less R&D, and their R&D is less tightly linked to process innovation, but they are no less innovative, at least according to their own reports. Surprisingly, the firms in our sample, are more rather than less productive than firms in other countries. Like Italian industry as a whole, they experienced a negative labor productivity growth during the 2000-2003 period, but apparently with no consequences on innovation activity and its estimated impact on productivity. Thus, we do not find any strong evidence of innovation “underperformance”, other than the observation that those firms in our sample which do R&D do somewhat less on average than firms in comparable European countries.

In general, “underinvestment” relative to others may be due to demand factors (perceived market size, consumer tastes, etc.) and supply factors (high costs of capital or other inputs, availability of inputs, and the regulatory environment). Stepping outside traditional economic analysis, factors such as having goals other than profit maximization, limited information about opportunities, or even social and cultural norms can also influence investment in innovation. Choosing among these alternatives definitively is beyond the scope of this paper, but we can offer a few tentative thoughts.

There is limited evidence that lower rates of R&D investment in larger Italian firms is due to the fact that they face a higher cost of capital than other firms in continental Europe. In a comparative analysis, Hall and Oriani (2006) find high marginal stock market values for Italian R&D investment in large firms that do not have a majority shareholder, which suggests a high required rate of return and therefore a high cost of capital. However for the other firms (closely held), R&D is not valued at all, which carries the implication that investment in these firms may not be profit driven. These conclusions suggests that a “bank-centered” capital market system, such as the Italian one, with a shortage of specialized suppliers like venture capitalists (Rajan and Zingales, 2003), is less capable of valuing R&D projects (Hall, 2002). Smaller firms and family-controlled with a pyramidal structure, which are quite common in Italy, are likely to be affected by credit rationing problems and/or to have goals other than profit-

maximization. But this is to some extent speculative, and we hope to explore the question further in the future.

References

- Acs, Z. J., and D. B. Audretsch (1988), "Innovation in Large and Small Firms: an Empirical Analysis", *American Economic Review*, Vol. 78(4), pp. 678-69.
- Acs, Z. J., and D. B. Audretsch (1990), *Innovation and Small Firms*, MIT Press, Cambridge, MA.
- Amemiya, T. (1984), "Tobit Models: a Survey", *Journal of Econometrics*, Vol. 24(1-2), pp. 3-61.
- Baldwin, J. R., D. Beckstead, and R. Caves (2002), "Changes in the Diversification of Canadian Manufacturing Firms (1973-1997): a Move to Specialization. *Research Paper Series 11F0019MIE2002179*, Statistics Canada.
- Baldwin, J. R. and W. Gu (2003), "Participation in Export Markets and Productivity Performance in Canadian Manufacturing", *Canadian Journal of Economics*, Vol. 36(3), pp. 634-657.
- Benavente, J. M. (2006), "The Role of Research and Innovation in Promoting Productivity in Chile", *Economics of Innovation and New Technology*. Vol. 15(4/5), pp. 301-315.
- Blundell, R., R. Griffith, and J. Van Reenen, (1993), "Knowledge Stocks, Persistent Innovation and Market Dominance: Evidence from a Panel of British Manufacturing Firms", *Institute for Fiscal Studies*, working paper N. W93/19.
- Bönte, W (2003), "R&D and Productivity: Internal vs. External R&D – Evidence from West German Manufacturing Industries", *Economics of Innovation and New Technology*, Vol. 12, pp. 343-360.
- Brandolini, A., P. Casadio, P. Cipollone, M. Magnani, A. Rosolia, and R. Torrini (2007), "Employment Growth in Italy in the 1990s: Institutional Arrangements and Market Forces", in *Social Pacts, Employment and Growth*, N. Acocella and R. Leoni Editors, Physica-Verlag HD.
- Cohen, W. M., and S. Klepper (1996), "A Reprise of Size and R&D", *The Economic Journal*, Vol. 106(437), pp. 925-951.
- Cohen W. M., and D. A. Levinthal (1989), "Innovation and Learning: The Two Faces of R&D", *The Economic Journal*, Vol. 99(397), pp. 569-596.
- Cowan, R. and G. van de Paal (2000), "Innovation Policy in the Knowledge-Based Economy", *European Commission DG-Enterprise*, Brussels.
- Crépon, B., E. Duguet, and J. Mairesse (1998), "Research, Innovation and Productivity: an Econometric Analysis at the Firm Level", *Economics of Innovation and New Technology*, Vol. 7(2), pp. 115-158.
- Das, M., W. K. Newey, and F. Vella (2003), "Nonparametric Estimation of Sample Selection Models", *Review of Economic Studies*, Vol. 70(1), pp. 33-58.
- Daveri, F., and C. Jona-Lasinio (2005), "Italy's Decline: Getting the Facts Right", *Giornale degli Economisti*, Vol. 64(4), pp. 365-410.
- David, P., and D. Foray (1995), "Assessing and Expanding the Science and Technology Knowledge Base", *STI Review*, n. 16, pp. 16-38.

- European Commission (2006), *European Trend Chart on Innovation*. Country Report, Italy. Enterprise Directorate-General.
- Griffith, R., E. Huergo, J. Mairesse, and B. Peters (2006), “Innovation and Productivity across Four European Countries”, *Oxford Review of Economic Policy*, Vol. 22(4), pp. 483-498.
- Griliches, Z. (1995), “R&D and Productivity: Econometric Results and Measurement Issues”, in: Stoneman, P. (ed.), *Handbook of the Economics of Innovation and Technological Change*, Oxford, 52-89.
- Griliches, Z. (1998), *R&D and Productivity: The Econometric Evidence*, Chicago University Press.
- Hall, B. H. (2002). The Financing of Research and Development. *Oxford Review of Economic Policy*, 18(1), pp. 35-51.
- Hall, B. H., and J. Mairesse (2006), “Empirical Studies of Innovation in The Knowledge-Driven Economy”, *Economics of Innovation and New Technology*. Vol. 15(4/5), pp. 289-299.
- Hall, B. H., and J., Mairesse (1995), “Exploring the relationship between R&D and productivity in French manufacturing firms”, *Journal of Econometrics*, Vol. 65(1), pp. 263-293.
- Hall, B. H., F. Lotti, and J. Mairesse (2008), “Employment, Innovation and Productivity: Evidence from Italian Microdata”, *Industrial and Corporate Change*, Vol. 17, pp. 813-839.
- Hall, B. H., and R. Oriani (2006), “Does the Market Value R&D Investment by European Firms? Evidence from a Panel of Manufacturing Firms in France, Germany, and Italy”, *International Journal of Industrial Organization*, Vol. 24(5), pp. 971-993.
- Harhoff, D., (1998), “R&D and Productivity in German Manufacturing Firms”, *Economics of Innovation and New Technology*, Vol. 6(1), pp. 29-49.
- Heckman, J. (1979), “Sample Selection Bias as a Specification Error”, *Econometrica*, Vol. 47(1), pp. 153-161.
- Heshmati, A., and H. Lööf (2006), “On the Relation between Innovation and Performance: a Sensitivity Analysis”, *Economics of Innovation and New Technology*. Vol. 15(4/5), pp. 317-344.
- Hoffman, K., M. Parejo, J. Bessant, and L. Perren (1998), “Small Firms, R&D, Technology and Innovation in the UK: a Literature Review”, *Technovation*, Vol. 18(1), pp. 39–55.
- Janz, N., H. Lööf, and A. Peters (2004), “Firm Level Innovation and Productivity – Is there a Common Story across Countries?”, *Problems and Perspectives in Management*, Vol. 2, pp. 1-22.
- Jefferson, G. H., B. Huamao, G. Xiaojing, and Y. Xiaoyun (2006), “R&D Performance in Chinese Industry”, *Economics of Innovation and New Technology*. Vol. 15(4/5), pp. 345-366.
- Kleinknecht, A. (1987), “Measuring R&D in Small Firms: How Much are we Missing?”, *The Journal of Industrial Economics*, Vol. 36(2), pp. 253-256.

- Klette, T. J., and S. Kortum (2004), "Innovating Firms and Aggregate Innovation", *Journal of Political Economy*, Vol. 112(5), pp. 986-1018.
- Klette, T. J., and F. Johansen (1996), "Accumulation of R&D Capital and Dynamic Firm Performance: A Not-so-fixed Effect Model", Discussion Papers 184, Research Department of Statistics Norway.
- Klomp, L., and G. Van Leeuwen (2001), "Linking Innovation and Firm Performance: A New Approach", *International Journal of the Economics of Business*, Vol. 8(3), pp. 343-364.
- Lichtenberg, F. R., and D. Siegel (1991), "The Impact of R&D Investment on Productivity: New Evidence using Linked R&D-LRD Data", *Economic Inquiry*, Vol. 29(2), pp. 203-228.
- Lotti, F., and E. Santarelli (2001), "Linking Knowledge to Productivity: a Germany-Italy Comparison Using the CIS Database", *Empirica*, Vol. 28, pp. 293-317.
- Lööf, H., and A. Heshmati (2002), "Knowledge Capital and Performance Heterogeneity: a Firm Level Innovation Study", *International Journal of Production Economics* Vol. 76(1), pp. 61-85.
- Mairesse, J., and M. Sassenou (1991), "R&D Productivity: A Survey of Econometric Studies at the Firm Level", NBER Working Paper n. 3666.
- Mairesse, J., and P. Mohnen (2005), "The Importance of R&D for Innovation: A Reassessment Using French Survey Data", *The Journal of Technology Transfer*, special issue in memory of Edwin Mansfield, Vol. 30(1-2), pp. 183-197.
- Mohnen, P., J. Mairesse, and M. Dagenais (2006), "Innovativity: a Comparison across Seven European Countries", *Economics of Innovation and New Technology*. Vol. 15(4/5), pp. 391-413.
- Pakes, A., and Z. Griliches (1984), "Patents and R&D at the Firm Level: A First Look", in: Griliches, Z. (ed.), *R&D, Patents, and Productivity*, Chicago, pp. 55-71.
- Parisi, M. L., F. Schiantarelli, and A. Sembenelli (2006), "Productivity, Innovation and R&D: Micro Evidence for Italy", *European Economic Review*, Vol. 50, pp. 2037-2061.
- Pianta, M., and A. Vaona (2007), "Innovation and Productivity in European Industries", *Economics of Innovation and New Technology*, Vol. 16(7), pp. 485-499.
- Potters, L., R. Ortega-Argilés, and M. Vivarelli (2008), "R&D and Productivity: Testing Sectoral Peculiarities using Micro Data", *IZA Discussion Paper N.* 3338.
- Rajan, R. G., and L. Zingales (2003), "Banks and Markets: The Changing Character of European Finance", NBER Working Paper n. 9595.
- Van Leeuwen, G. and L. Klomp (2006), "On the Contribution of Innovation to Multi-Factor Productivity Growth", *Economics of Innovation and New Technology*. Vol. 15(4/5), pp. 367-390.
- Vella, F. (1998), "Estimating Models with Sample Selection Bias: A Survey", *The Journal of Human Resources*, Vol. 33(1), pp. 127-169.
- Verspagen, B. (1995), "R&D and Productivity: a Broad Cross-Section Cross-Country Look", *Journal of Productivity Analysis*, Vol. 6, pp. 117-135

Tables and Figures

Table 1 – Descriptive statistics, unbalanced and balanced sample

<i>Period: 1995-2003</i>	<i>Unbalanced sample</i>	<i>Balanced sample</i>
Number of observations (firms)	9,674 (7,375)	1,083 (361)
Continuous R&D engagement (in %)	41.49	26.04
R&D intensity (for R&D doing firms, in logs) ^(a)	1.08	1.02
Innovator (process and/or product, in %)	62.05	66.39
Process innovation (in %)	50.75	53.65
Product innovation (in %)	34.85	40.63
Process & product innovation (in %)	20.94	25.39
Process innovation only (in %)	27.21	25.76
Share of sales with new products (in %)	22.16	22.98
Labor productivity: mean/median ^(a)	4.99 / 4.94	4.94 / 4.85
Investment intensity: mean/median ^(a)	7.90 / 4.05	6.92 / 4.01
Public support (in %)	45.49	50.51
Regional competitors (in %)	16.84	14.87
National competitors (in %)	42.24	41.37
European competitors (in %)	17.53	18.10
International (non EU) competitors (in %)	13.56	17.17
Large competitors (in %)	36.18	34.16
% of firm in size class (11-20)	30.04	19.67
% of firm in size class (21-50)	38.85	44.04
% of firm in size class (51-250)	31.11	36.29
% of firm in age class (<15 yrs)	32.45	24.10
% of firm in age class (15-25 yrs)	30.48	31.12
% of firm in age class (>25)	37.07	44.78
Number of employees: mean/median	49.45 / 32	53.48 / 36
Group (in %)	20.07	16.25

^(a) Units are logs of euros (2000) per employee.

Table 2 – Descriptive statistics, high tech and low tech industries.

<i>Period: 1995-2003</i>	<i>High tech firms</i>	<i>Low tech firms</i>
Number of observations (firms)	2,870 (2,165)	6,804 (5,210)
Continuous R&D engagement (in %)	58.75	34.22
R&D intensity (for R&D doing firms, in logs) ^(a)	1.20	0.98
Innovator (process and/or product, in %)	69.41	58.95
Process innovation (in %)	54.25	49.28
Product innovation (in %)	43.80	31.06
Process & product innovation (in %)	25.57	18.72
Process innovation only (in %)	26.20	27.90
Share of sales with new products (in %)	22.63	21.88
Labor productivity: mean/median ^(a)	4.93 / 4.89	5.02 / 4.96
Investment intensity: mean/median ^(a)	6.22 / 3.36	8.62 / 4.38
Public support (in %)	46.27	45.16
Regional competitors (in %)	12.30	18.75
National competitors (in %)	36.45	44.68
European competitors (in %)	25.40	14.21
International (non EU) competitors (in %)	19.86	10.91
Large competitors (in %)	42.54	33.50
% of firm in size class (11-20)	27.25	31.22
% of firm in size class (21-50)	36.86	39.68
% of firm in size class (51-250)	35.89	29.10
% of firm in age class (<15 yrs)	32.79	32.30
% of firm in age class (15-25 yrs)	31.67	29.98
% of firm in age class (>25)	35.54	37.71
Number of employees: mean/median	54.17 / 35	47.46 / 30
Group (in %)	25.26	17.89

^(a) Units are logs of euros (2000) per employee.

Table 3 – R&D intensity (STEP 1): OLS model. Dependent variable, R&D intensity

<i>R&D expenditure per employee (in logarithms)</i>	<i>All firms</i>	<i>High Tech</i>	<i>Low Tech</i>
D(Large firm competitors)	0.062 (0.073)	0.197 (0.109)	-0.028 (0.098)
D(Regional competitors)	0.094 (0.167)	0.548 (0.320)	-0.049 (0.197)
D(National competitors)	0.138 (0.147)	0.638* (0.290)	-0.037 (0.172)
D(European competitors)	0.511*** (0.154)	0.834** (0.287)	0.448* (0.187)
D(International competitors)	0.570*** (0.159)	1.034*** (0.296)	0.357 (0.195)
D(Received subsidies)	0.389*** (0.072)	0.619*** (0.111)	0.213* (0.095)
D(Member of a group)	0.198* (0.084)	0.247 (0.128)	0.165 (0.114)
Size class (21-50 empl.)	-0.271** (0.104)	-0.141 (0.164)	-0.349** (0.134)
Size class (51-250 empl.)	-0.271* (0.109)	-0.123 (0.167)	-0.379** (0.145)
Age class (15-25 yrs)	-0.009 (0.094)	0.032 (0.141)	-0.032 (0.127)
Age class (>25 yrs)	-0.061 (0.090)	-0.147 (0.135)	-0.003 (0.120)
R-squared	0.061	0.065	0.038
Number of observations	4,015	1,687	2,328

Coefficients and their standard errors are shown. The standard errors are robust to heteroskedasticity and clustered at the firm level. * = significant at 10%, ** = significant at 5%, *** = significant at 1%. Industry, wave, and time dummies are included in all equations. Reference groups: D(provincial competitors), Size class (11-50 empl), Age class (<15 yrs).

Table 4 – A bivariate probit for process and product innovation dummies (STEP 2): all firms, high- and low- tech firms

	All firms		High-tech firms		Low-tech firms	
	(1)	(1a)	(2)	(2a)	(3)	(3a)
	Process Inno.	Product Inno.	Process Inno.	Product Inno.	Process Inno.	Product Inno.
Predicted R&D intensity (in logs)	0.483*** [0.193] (0.045)	0.686*** [0.250] (0.045)	0.256*** [0.102] (0.056)	0.499*** [0.196] (0.056)	0.602*** [0.240] (0.069)	0.749*** [0.261] (0.069)
Investment per employee (in logs)	0.125*** [0.050] (0.011)		0.120*** [0.047] (0.021)		0.129*** [0.051] (0.013)	
Size class (21-50 empl.)	0.255*** [0.101] (0.033)	0.310*** [0.115] (0.035)	0.159* [0.063] (0.062)	0.126* [0.050] (0.063)	0.350*** [0.139] (0.043)	0.431*** [0.153] (0.046)
Size class (51-250 empl.)	0.446*** [0.175] (0.037)	0.504*** [0.189] (0.038)	0.276*** [0.108] (0.068)	0.299*** [0.118] (0.067)	0.606*** [0.237] (0.048)	0.679*** [0.248] (0.049)
Age class (15-25 yrs)	0.009 [0.004] (0.034)	0.050 [0.018] (0.034)	0.004 [0.001] (0.061)	0.036 [0.014] (0.061)	0.020 [0.008] (0.040)	0.058 [0.020] (0.042)
Age class (>25 yrs)	-0.003 [-0.001] (0.033)	0.129*** [0.047] (0.034)	0.094 [0.037] (0.062)	0.157* [0.062] (0.062)	-0.067 [-0.026] (0.039)	0.094* [0.033] (0.041)
Rho	0.400***		0.345***		0.430***	
Pseudo R-squared	0.10	0.08	0.09	0.08	0.10	0.06
Number of obs. (firms)	9,674 (7,375)		2,870 (2,165)		6,804 (5,210)	

Coefficients, marginal effects and standard errors are shown. Marginal effects in square brackets. The standard errors are robust to heteroskedasticity and clustered at the firm level. * = significant at 10%, ** = significant at 5%, *** = significant at 1% . Industry, wave, and time dummies are included in all equations. Reference groups: D(provincial competitors), Size class (11-50 empl), Age class (<15 yrs).

Table 5 – Production function (STEP 3): all firms, high- and low-tech firms

Dep. variable: labor productivity (sales per employee in logs)	All firms		High-tech firms		Low-tech firms	
	(1)	(1a)	(2)	(2a)	(3)	(3a)
Predicted probability of process innovation only	2.624*** (0.146)	0.193 (0.267)	2.742*** (0.304)	0.664 (0.512)	2.797*** (0.171)	0.063 (0.391)
Predicted probability of product innovation	0.961*** (0.083)	0.597*** (0.093)	1.314*** (0.149)	0.700*** (0.200)	0.900*** (0.118)	0.708*** (0.122)
Investment per employee (in logs)		0.099*** (0.010)		0.073*** (0.015)		0.109*** (0.015)
Size class (21-50 empl.)	-0.184*** (0.016)	-0.136*** (0.017)	-0.140*** (0.029)	-0.085** (0.031)	-0.204*** (0.020)	-0.163*** (0.021)
Size class (51-250 empl.)	-0.313*** (0.023)	-0.243*** (0.024)	-0.177*** (0.037)	-0.116** (0.038)	-0.391*** (0.031)	-0.321*** (0.032)
Age class (15-25 yrs)	-0.006 (0.016)	-0.017 (0.016)	-0.0579* (0.026)	-0.064* (0.026)	0.0174 (0.020)	0.005 (0.020)
Age class (>25 yrs)	0.008 (0.016)	-0.038* (0.016)	-0.0764** (0.027)	-0.069** (0.027)	0.0469* (0.020)	-0.036 (0.022)
R-squared	0.209	0.219	0.194	0.201	0.227	0.226
Number of observations (firms)	9,674 (7,375)		2,870 (2,165)		6,804 (5,210)	

Coefficients and their standard errors are shown. The standard errors are robust to heteroskedasticity and clustered at the firm level. * = significant at 10%, ** = significant at 5%, *** = significant at 1%. Industry, wave, and time dummies are included in all equations. Reference groups: D(provincial competitors), Size class (11-50 empl), Age class (<15 yrs).

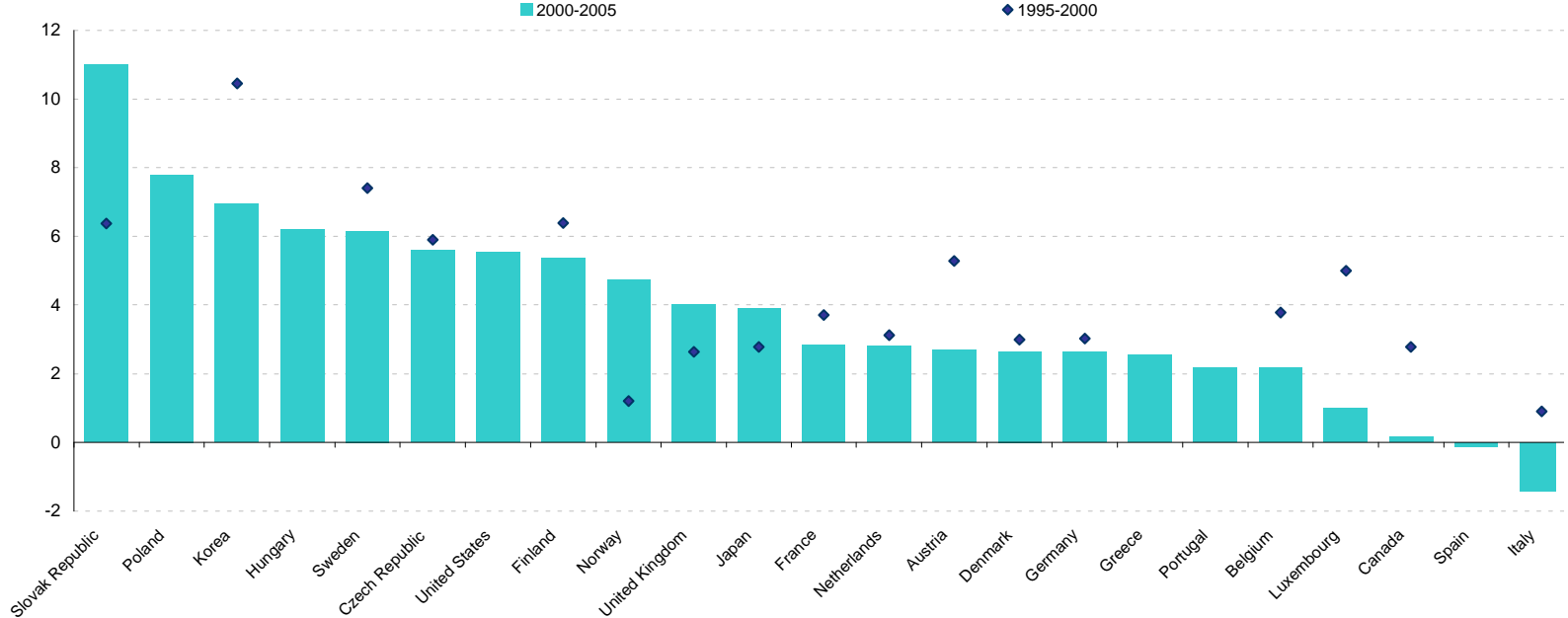
Table 6 – Comparison with Griffith et al. (2006)

<i>Period: 1998-2000</i>	<i>France</i>	<i>Germany</i>	<i>Spain</i>	<i>UK</i>	<i>Italy</i>	<i>Italy</i> ^(a)
Number of observations	3,625	1,123	3,588	1,904	2,594	8,377
<i>Process innovation equation</i>						
R&D intensity ^(b)	0.303 ***	0.260 ***	0.281 ***	0.161 ***	0.146 ***	0.192 ***
Investment intensity ^(b)	0.023 ***	0.022 ***	0.029 ***	0.037 ***	0.054 ***	0.049 ***
Pseudo R-squared	0.213	0.202	0.225	0.184	0.050	0.091
<i>Product innovation equation</i>						
R&D intensity ^(b)	0.440 ***	0.273 ***	0.296 ***	0.273 ***	0.192 ***	0.303 ***
Pseudo R-squared	0.360	0.313	0.249	0.258	0.058	0.081
<i>Labor Productivity equation</i>						
Investment intensity ^(b)	0.130 ***	0.109 ***	0.061 ***	0.059 ***	0.129	0.109 ***
Process Innovation	0.069 **	0.022	-0.038	0.029	-0.874	0.011
Product Innovation	0.060 ***	-0.053	0.176 ***	0.055 ***	1.152	0.384
R-squared	0.290	0.280	0.180	0.190	0.166	0.227

This table is based on tables in Griffith et al. 2006. Data are from the third Community Innovation Survey (CIS 3) for France, Germany, Spain, and the U. K. Results for Italy come from Tables 3-5 of this paper. (a) This column shows data for all 3 periods in Italy (1995-1997, 1998-2000, 2001-2003).

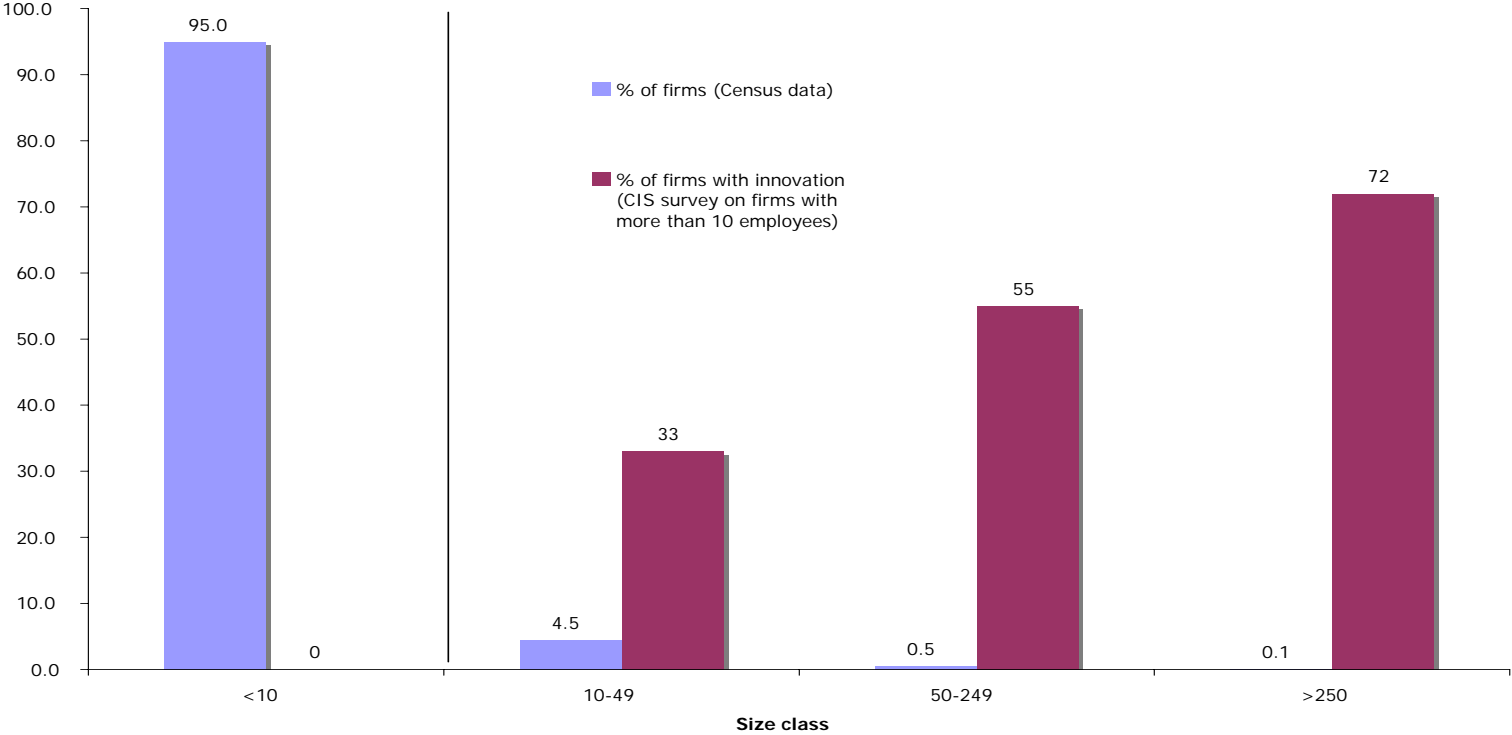
(b) Units are logs of euros (2000) per employee. * Significant at 10%, ** significant at 5%, *** significant at 1%

Figure 1 – Value added per employee. Percentage change, annual rate (1995-2000 and 2000-2005). Total manufacturing.



Source: OECD Factbook, April 2008. Permanent link <http://dx.doi.org/10.1787/271772787380>

Figure 2 – Size distribution of Italian firms (2001) and share of firms with innovation by size class (2002-2004).



Source: National Institute of Statistics (ISTAT). Census of Manufacturing and Services (2001) for the size distribution. Community innovation survey (CIS) for the presence of innovation activity (2002-2004).

Appendix

Variable Definitions

R&D engagement: dummy variable that takes value 1 if the firm has positive R&D expenditures over the three year of each wave of the survey.

R&D intensity: R&D expenditures per employee, in real terms and in logs.

Process innovation: dummy variable that takes value 1 if the firm declares to have introduced a process innovation during the three years of the survey.

Product innovation: dummy variable that takes value 1 if the firm declares to have introduced a product innovation during the three years of the survey.

Innovator: dummy variable that takes value 1 if the firm has process or product innovation.

Share of sales with new products: percentage of the sales in the last year of the survey coming from new or significantly improved products (in percentage).

Labor productivity: real sales per employee, in logs.

Investment intensity: investment in machinery per employee, in logs.

Public support: dummy variable that takes value 1 if the firm has received a subsidy during the three years of the survey.

Regional – National – European –International (non EU) competitors: dummy variables to indicate the location of the firm's competitors.

Large competitors: dummy variable that takes value 1 if the firm declares to have large firms as competitors.

Employees: number of employees, headcount.

Age: firm's age (in years).

Size classes: [11-20], [21-50], [51-250] employees.

Age classes: [<15], [15-25], [>25] years.

Industry dummies: a set of indicators for a 2-digits industry classification.

Time dummies: a set of indicators for the year of the survey.

Wave dummies: a set of indicators for firm's presence or absence in the three waves of the survey

High-tech firms: encompasses high and medium-high technology industries (chemicals; office accounting & computer machinery; radio, TV & telecommunication instruments; medical, precision & optical instruments; electrical machinery and apparatus, n.e.c.; machinery & equipment; railroad & transport equipment, n.e.c.).

Low-tech firms: encompasses low and medium-low technology industries (rubber & plastic products; coke, refined petroleum products; other non-metallic mineral products; basic metals and fabricated metal products; manufacturing n.e.c.; wood, pulp & paper; food, beverages & tobacco products; textile, textile products, leather & footwear).

Capital stock: fixed capital stock, in real terms, computed by a perpetual inventory method with constant depreciation rate ($\delta=0.05$). The starting value is the accounting value as reported in firm's balance sheets.

Appendix

Table A1 – A comparison of selected variables for France, Germany, Spain, UK and Italy.

Period: 1998-2000	<i>France</i>	<i>Germany</i>	<i>Spain</i>	<i>UK</i>	<i>Italy</i>	<i>Italy (b)</i>
Number of observations (firms)	3,625	1,123	3,588	1,904	2,594	8,377
Continuous R&D engagement (in %)	35.0	39.5	20.9	26.7	49.8	48.9
R&D per employee (for R&D-doers, mean)†	6.9	5.2	4.3	3.6	2.9	2.4
Innovator (process and/or product, in %)	52.9	65.8	51.2	41.5	54.7	66.9
Process innovation (in %)	32.3	42.3	34.7	27.1	44.7	55.4
Product innovation (in %)	44.6	54.7	33.6	28.6	33.3	39.9
Share of sales with new products for firms with product innovation (in %)	16.5	29.5	32.7	30.8	32.2	22.5
Labor productivity (mean)†	165.3	145.6	137.7	143.4	173.8	187.1
Investment per employee (mean)†	6.0	8.3	8.3	6.3	8.0	7.9
Public support for innovation (in %)						
Local	5.5	15.8	14.0	4.5		
National	15.4	21.2	12.5	3.6	49.9 (a)	50.6 (a)
EU	5.1	8.1	3.3	1.7		
% of firm in size class (20-49)	30.4	28.8	47.8	38.6	60.6	44.9
% of firm in size class (50-250)	39.6	42.8	37.5	39.3	27.8	36.7
% of firm in size class (>250)	30.0	28.5	14.7	22.1	11.1	18.4

This table is a slightly modified version of Table 3 in Griffith et al. 2006. Data are from the third Community Innovation Survey (CIS 3) for France, Germany, Spain, and the UK. Data for Italy are from the Mediocredito Surveys. Among the several variables included in the original table, we selected only those comparable to our data. Data are not population-weighted. (a) This figure encompasses all the subsidies, regardless their source. (b) This column shows data for all 3 periods in Italy (1995-1997, 1998-2000, 2001-2003). †Units are logs of euros (2000) per employee.

Table A2 – A non-parametric selectivity test

Dependent variable	Prob(R&D>0)	R&D expend. per employee
D(Large firms)	0.150*** (0.030)	0.305 (0.436)
D(Regional)	-0.138* (0.056)	-0.230 (0.408)
D(National)	0.012 (0.051)	0.0879 (0.085)
D(European)	0.339*** (0.057)	0.826 (0.988)
D(International)	0.391*** (0.060)	0.927 (1.142)
D(Public subsidies for innovation)†	0.324*** (0.028)	0.761 (0.943)
Group	0.145*** (0.037)	0.339 (0.423)
Size class (21-50 empl.)	0.147*** (0.035)	0.200 (0.431)
Size class (51-250 empl.)	0.482*** (0.040)	0.759 (1.402)
Age class (15-25 yrs)	0.022 (0.036)	0.0258 (0.089)
Age class (>25 yrs)	0.064 (0.036)	0.0684 (0.197)
Constant	-0.563*** (0.163)	499.4 (424.583)
	Predicted Pr(R&D>0)	157.1 (130.890)
	Inverse Mill's ratio	92.21 (81.214)
	Square Predicted Pr(R&D>0)	-399.9 (336.616)
	Square Inverse Mill's ratio	183.7 (152.908)
	Predicted Pr(R&D>0) * Inverse Mill's ratio	499.4 (424.583)
Industry, Time & Wave dummies	Yes	Yes
R-squared or pseudo R-squared	0.114	0.143
Number of observations	9,674	9,674

Standard errors are robust and clustered at the firm level. * = significant at 10%, **=significant at 5%, ***=significant at 1%. From this probit model we computed, for each observation in the sample, the inverse Mills' ratio, the predicted probability of having positive R&D and their quadratic and interaction terms. †This figure encompasses all the subsidies, regardless their source.

Table A3 – Robustness check for step 2 and 3.

	(1)	(2)	(3)	(4)
Step 2 - Process Innovation				
Predicted R&D intensity	0.483*** [0.193] (0.045)	0.544*** [0.217] (0.045)	0.476*** [0.190] (0.045)	0.547*** [0.218] (0.045)
Investment intensity	0.125*** [0.050] (0.011)	-	0.137*** [0.055] (0.011)	-
Step 2 - Product Innovation				
Predicted R&D intensity	0.686*** [0.250] (0.045)	0.677*** [0.247] (0.046)	0.660*** [0.241] (0.046)	0.691*** [0.252] (0.045)
Investment intensity	-	0.021* [0.008] (0.011)	0.055*** [0.020] (0.011)	-
Step 3 - Productivity including investment in the equation				
Predicted process inno	0.193 (0.267)	-0.395 (0.275)	0.010 (0.255)	-0.432 (0.277)
Predicted product inno	0.597*** (0.093)	0.554*** (0.087)	0.599*** (0.095)	0.538*** (0.086)
Investment intensity	0.099*** (0.010)	0.099*** (0.006)	0.093*** (0.009)	0.105*** (0.006)
Step 3 - Productivity without investment in the equation				
Predicted process inno	2.624*** (0.146)	-1.318*** (0.279)	2.286*** (0.168)	-0.171 (0.280)
Predicted product inno	0.961*** (0.083)	0.895*** (0.087)	1.133*** (0.079)	0.773*** (0.087)

Coefficients, marginal effects for step 2 in square brackets, and standard errors are shown. The standard errors are robust to heteroskedasticity and clustered at the firm level. * = significant at 10%, ** = significant at 5%, *** = significant at 1%. Industry, wave, and time dummies are included in all equations. Reference groups: D(provincial competitors), Size class (11-50 empl), Age class (<15 yrs). Specifications (1)-(4) encompass alternative assumptions for investment, whether it is devoted to process or product innovation, neither, or both.

Table A4 - Robustness check using lagged capital and ML estimation (9014 observations)

	(1)	(2)	(3)	(4)	(5)
	with investment	with investment	with capital	investment in process, capital in productivity	with both
<i>Method of estimation:</i>	<i>Sequential</i>	<i>pooled ML</i>	<i>pooled ML</i>	<i>pooled ML</i>	<i>pooled ML</i>
Step 2 - Process Innovation					
Predicted R&D intensity	0.440*** (0.048)	0.400*** (0.074)	0.416*** (0.078)	0.399*** (0.073)	0.389*** (0.075)
Log investment per employee	0.131*** (0.011)	0.142*** (0.013)		0.145*** (0.011)	0.120*** (0.013)
Log capital stock† per employee			0.098*** (0.013)		0.041*** (0.014)
Step 2 - Product Innovation					
Predicted R&D intensity	0.652*** (0.047)	0.656*** (0.094)	0.655*** (0.094)	0.658*** (0.093)	0.661*** (0.095)
Step 3 - Productivity equation					
Predicted process inno without product inno	0.517* (0.279)	0.712* (0.443)	0.902* (0.515)	1.108*** (0.159)	0.855* (0.477)
Predicted product inno	0.677*** (0.108)	1.081*** (0.310)	0.792** (0.337)	0.881*** (0.314)	0.830** (0.370)
Log investment per employee	0.081*** (0.011)	0.072*** (0.017)			0.018 (0.015)
Log capital stock† per employee			0.108*** (0.016)	0.111*** (0.007)	0.101*** (0.010)
Log likelihood	-27,119.9	-27,110.0	-26,979.0	-26,908.5	-26,901.3

The method of estimation in the last three columns is pooled maximum likelihood applied to the 3 steps, with the coefficient constraints imposed but without allowing for correlation among their disturbances. This method yields standard errors that account for the use of predicted variables in steps 2 and 3.

Coefficients and their standard errors are shown. The standard errors are robust to heteroskedasticity and clustered at the firm level. Marginal effects in square brackets.

* = significant at 10%, ** = significant at 5%, *** = significant at 1% .

Industry, wave, and time dummies are included in all equations.

Reference groups: D(provincial competitors), Size class (11-50 empl), Age class (<15 yrs).

† Capital measured at the beginning of the period.

RECENTLY PUBLISHED “TEMI” (*)

- N. 690 – *Nonlinearities in the dynamics of the euro area demand for M1*, by Alessandro Calza and Andrea Zaghini (Settembre 2008).
- N. 691 – *Educational choices and the selection process before and after compulsory schooling*, by Sauro Mocetti (Settembre 2008).
- N. 692 – *Investors' risk attitude and risky behavior: a Bayesian approach with imperfect information*, by Stefano Iezzi (Settembre 2008).
- N. 693 – *Competing influence*, by Enrico Sette (Settembre 2008).
- N. 694 – *La relazione tra gettito tributario e quadro macroeconomico in Italia*, by Alberto Locarno and Alessandra Staderini (Dicembre 2008).
- N. 695 – *Immigrant earnings in the Italian labour market*, by Antonio Accetturo and Luigi Infante (Dicembre 2008).
- N. 696 – *Reservation wages: explaining some puzzling regional patterns*, by Paolo Sestito and Eliana Viviano (Dicembre 2008).
- N. 697 – *Technological change and the demand for currency: an analysis with household data*, by Francesco Lippi and Alessandro Secchi (Dicembre 2008).
- N. 698 – *Immigration and crime: an empirical analysis*, by Milo Bianchi, Paolo Buonanno and Paolo Pinotti (Dicembre 2008).
- N. 699 – *Bond risk premia, macroeconomic fundamentals and the exchange rate*, by Marcello Pericoli and Marco Taboga (Gennaio 2009).
- N. 700 – *What determines debt intolerance? The role of political and monetary institutions*, by Raffaella Giordano and Pietro Tommasino (Gennaio 2009).
- N. 701 – *On analysing the world distribution of income*, by Anthony B. Atkinson and Andrea Brandolini (Gennaio 2009).
- N. 702 – *Dropping the books and working off the books*, by Rita Cappariello and Roberta Zizza (Gennaio 2009).
- N. 703 – *Measuring wealth mobility*, by Andrea Neri (Gennaio 2009).
- N. 704 – *Oil and the macroeconomy: a quantitative structural analysis*, by Francesco Lippi and Andrea Nobili (Marzo 2009).
- N. 705 – *The (mis)specification of discrete duration models with unobserved heterogeneity: a Monte Carlo study*, by Cheti Nicoletti and Concetta Rondinelli (Marzo 2009).
- N. 706 – *Macroeconomic effects of grater competition in the service sector: the case of Italy*, by Lorenzo Forni, Andrea Gerali and Massimiliano Pisani (Marzo 2009).
- N. 707 – *What determines the size of bank loans in industrialized countries? The role of government debt*, by Riccardo De Bonis and Massimiliano Stacchini (Marzo 2009).
- N. 708 – *Trend inflation, Taylor principle and indeterminacy*, by Guido Ascari and Tiziano Ropele (May 2009).
- N. 709 – *Politicians at work. The private returns and social costs of political connection*, by Federico Cingano and Paolo Pinotti (May 2009).
- N. 710 – *Gradualism, transparency and the improved operational framework: a look at the overnight volatility transmission*, by Silvio Colarossi and Andrea Zaghini (May 2009).
- N. 711 – *The topology of the interbank market: developments in Italy since 1990*, by Carmela Iazzetta and Michele Manna (May 2009).
- N. 712 – *Bank risk and monetary policy*, by Yener Altunbas, Leonardo Gambacorta and David Marqués-Ibáñez (May 2009).
- N. 713 – *Composite indicators for monetary analysis*, by Andrea Nobili (May 2009).

(*) Requests for copies should be sent to:
 Banca d'Italia – Servizio Studi di struttura economica e finanziaria – Divisione Biblioteca e Archivio storico – Via Nazionale, 91 – 00184 Rome – (fax 0039 06 47922059). They are available on the Internet www.bancaditalia.it.

2006

- F. Busetti, *Tests of seasonal integration and cointegration in multivariate unobserved component models*, Journal of Applied Econometrics, Vol. 21, 4, pp. 419-438, **TD No. 476 (June 2003)**.
- C. Biancotti, *A polarization of inequality? The distribution of national Gini coefficients 1970-1996*, Journal of Economic Inequality, Vol. 4, 1, pp. 1-32, **TD No. 487 (March 2004)**.
- L. Cannari and S. Chiri, *La bilancia dei pagamenti di parte corrente Nord-Sud (1998-2000)*, in L. Cannari, F. Panetta (a cura di), *Il sistema finanziario e il Mezzogiorno: squilibri strutturali e divari finanziari*, Bari, Cacucci, **TD No. 490 (March 2004)**.
- M. Bofondi and G. Gobbi, *Information barriers to entry into credit markets*, Review of Finance, Vol. 10, 1, pp. 39-67, **TD No. 509 (July 2004)**.
- W. Fuchs and Lippi F., *Monetary union with voluntary participation*, Review of Economic Studies, Vol. 73, pp. 437-457 **TD No. 512 (July 2004)**.
- E. Gaiotti and A. Secchi, *Is there a cost channel of monetary transmission? An investigation into the pricing behaviour of 2000 firms*, Journal of Money, Credit and Banking, Vol. 38, 8, pp. 2013-2038 **TD No. 525 (December 2004)**.
- A. Brandolini, P. Cipollone and E. Viviano, *Does the ILO definition capture all unemployment?*, Journal of the European Economic Association, Vol. 4, 1, pp. 153-179, **TD No. 529 (December 2004)**.
- A. Brandolini, L. Cannari, G. D'Alessio and I. Faiella, *Household wealth distribution in Italy in the 1990s*, in E. N. Wolff (ed.) *International Perspectives on Household Wealth*, Cheltenham, Edward Elgar, **TD No. 530 (December 2004)**.
- P. Del Giovane and R. Sabbatini, *Perceived and measured inflation after the launch of the Euro: Explaining the gap in Italy*, Giornale degli economisti e annali di economia, Vol. 65, 2, pp. 155-192, **TD No. 532 (December 2004)**.
- M. Caruso, *Monetary policy impulses, local output and the transmission mechanism*, Giornale degli economisti e annali di economia, Vol. 65, 1, pp. 1-30, **TD No. 537 (December 2004)**.
- L. Guiso and M. Paiella, *The role of risk aversion in predicting individual behavior*, In P. A. Chiappori e C. Gollier (eds.) *Competitive Failures in Insurance Markets: Theory and Policy Implications*, Monaco, CESifo, **TD No. 546 (February 2005)**.
- G. M. Tomat, *Prices product differentiation and quality measurement: A comparison between hedonic and matched model methods*, Research in Economics, Vol. 60, 1, pp. 54-68, **TD No. 547 (February 2005)**.
- L. Guiso, M. Paiella and I. Visco, *Do capital gains affect consumption? Estimates of wealth effects from Italian household's behavior*, in L. Klein (ed), *Long Run Growth and Short Run Stabilization: Essays in Memory of Albert Ando (1929-2002)*, Cheltenham, Elgar, **TD No. 555 (June 2005)**.
- F. Busetti, S. Fabiani and A. Harvey, *Convergence of prices and rates of inflation*, Oxford Bulletin of Economics and Statistics, Vol. 68, 1, pp. 863-878, **TD No. 575 (February 2006)**.
- M. Caruso, *Stock market fluctuations and money demand in Italy, 1913 - 2003*, Economic Notes, Vol. 35, 1, pp. 1-47, **TD No. 576 (February 2006)**.
- R. Bronzini and G. De Blasio, *Evaluating the impact of investment incentives: The case of Italy's Law 488/92*. Journal of Urban Economics, Vol. 60, 2, pp. 327-349, **TD No. 582 (March 2006)**.
- R. Bronzini and G. De Blasio, *Una valutazione degli incentivi pubblici agli investimenti*, Rivista Italiana degli Economisti, Vol. 11, 3, pp. 331-362, **TD No. 582 (March 2006)**.
- A. Di Cesare, *Do market-based indicators anticipate rating agencies? Evidence for international banks*, Economic Notes, Vol. 35, pp. 121-150, **TD No. 593 (May 2006)**.
- R. Golinelli and S. Momigliano, *Real-time determinants of fiscal policies in the euro area*, Journal of Policy Modeling, Vol. 28, 9, pp. 943-964, **TD No. 609 (December 2006)**.

- S. SIVIERO and D. TERLIZZESE, *Macroeconomic forecasting: Debunking a few old wives' tales*, Journal of Business Cycle Measurement and Analysis, v. 3, 3, pp. 287-316, **TD No. 395 (February 2001)**.
- S. MAGRI, *Italian households' debt: The participation to the debt market and the size of the loan*, Empirical Economics, v. 33, 3, pp. 401-426, **TD No. 454 (October 2002)**.
- L. CASOLARO and G. GOBBI, *Information technology and productivity changes in the banking industry*, Economic Notes, Vol. 36, 1, pp. 43-76, **TD No. 489 (March 2004)**.
- G. FERRERO, *Monetary policy, learning and the speed of convergence*, Journal of Economic Dynamics and Control, v. 31, 9, pp. 3006-3041, **TD No. 499 (June 2004)**.
- M. PAIELLA, *Does wealth affect consumption? Evidence for Italy*, Journal of Macroeconomics, Vol. 29, 1, pp. 189-205, **TD No. 510 (July 2004)**.
- F. LIPPI and S. NERI, *Information variables for monetary policy in a small structural model of the euro area*, Journal of Monetary Economics, Vol. 54, 4, pp. 1256-1270, **TD No. 511 (July 2004)**.
- A. ANZUINI and A. LEVY, *Monetary policy shocks in the new EU members: A VAR approach*, Applied Economics, Vol. 39, 9, pp. 1147-1161, **TD No. 514 (July 2004)**.
- D. JR. MARCHETTI and F. Nucci, *Pricing behavior and the response of hours to productivity shocks*, Journal of Money Credit and Banking, v. 39, 7, pp. 1587-1611, **TD No. 524 (December 2004)**.
- R. BRONZINI, *FDI Inflows, agglomeration and host country firms' size: Evidence from Italy*, Regional Studies, Vol. 41, 7, pp. 963-978, **TD No. 526 (December 2004)**.
- L. MONTEFORTE, *Aggregation bias in macro models: Does it matter for the euro area?*, Economic Modelling, 24, pp. 236-261, **TD No. 534 (December 2004)**.
- A. NOBILI, *Assessing the predictive power of financial spreads in the euro area: does parameters instability matter?*, Empirical Economics, Vol. 31, 1, pp. 177-195, **TD No. 544 (February 2005)**.
- A. DALMAZZO and G. DE BLASIO, *Production and consumption externalities of human capital: An empirical study for Italy*, Journal of Population Economics, Vol. 20, 2, pp. 359-382, **TD No. 554 (June 2005)**.
- M. BUGAMELLI and R. TEDESCHI, *Le strategie di prezzo delle imprese esportatrici italiane*, Politica Economica, v. 23, 3, pp. 321-350, **TD No. 563 (November 2005)**.
- L. GAMBACORTA and S. IANNOTTI, *Are there asymmetries in the response of bank interest rates to monetary shocks?*, Applied Economics, v. 39, 19, pp. 2503-2517, **TD No. 566 (November 2005)**.
- P. ANGELINI and F. LIPPI, *Did prices really soar after the euro cash changeover? Evidence from ATM withdrawals*, International Journal of Central Banking, Vol. 3, 4, pp. 1-22, **TD No. 581 (March 2006)**.
- A. LOCARNO, *Imperfect knowledge, adaptive learning and the bias against activist monetary policies*, International Journal of Central Banking, v. 3, 3, pp. 47-85, **TD No. 590 (May 2006)**.
- F. LOTTI and J. MARCUCCI, *Revisiting the empirical evidence on firms' money demand*, Journal of Economics and Business, Vol. 59, 1, pp. 51-73, **TD No. 595 (May 2006)**.
- P. CIPOLLONE and A. ROSOLIA, *Social interactions in high school: Lessons from an earthquake*, American Economic Review, Vol. 97, 3, pp. 948-965, **TD No. 596 (September 2006)**.
- L. DEDOLA and S. NERI, *What does a technology shock do? A VAR analysis with model-based sign restrictions*, Journal of Monetary Economics, Vol. 54, 2, pp. 512-549, **TD No. 607 (December 2006)**.
- F. VERGARA CAFFARELLI, *Merge and compete: strategic incentives for vertical integration*, Rivista di politica economica, v. 97, 9-10, serie 3, pp. 203-243, **TD No. 608 (December 2006)**.
- A. BRANDOLINI, *Measurement of income distribution in supranational entities: The case of the European Union*, in S. P. Jenkins e J. Micklewright (eds.), *Inequality and Poverty Re-examined*, Oxford, Oxford University Press, **TD No. 623 (April 2007)**.
- M. PAIELLA, *The foregone gains of incomplete portfolios*, Review of Financial Studies, Vol. 20, 5, pp. 1623-1646, **TD No. 625 (April 2007)**.
- K. BEHRENS, A. R. LAMORGESE, G.I.P. OTTAVIANO and T. TABUCHI, *Changes in transport and non transport costs: local vs. global impacts in a spatial network*, Regional Science and Urban Economics, Vol. 37, 6, pp. 625-648, **TD No. 628 (April 2007)**.
- M. BUGAMELLI, *Prezzi delle esportazioni, qualità dei prodotti e caratteristiche di impresa: analisi su un campione di imprese italiane*, v. 34, 3, pp. 71-103, *Economia e Politica Industriale*, **TD No. 634 (June 2007)**.
- G. ASCARI and T. ROPELE, *Optimal monetary policy under low trend inflation*, Journal of Monetary Economics, v. 54, 8, pp. 2568-2583, **TD No. 647 (November 2007)**.

- R. GIORDANO, S. MOMIGLIANO, S. NERI and R. PEROTTI, *The Effects of Fiscal Policy in Italy: Evidence from a VAR Model*, European Journal of Political Economy, Vol. 23, 3, pp. 707-733, **TD No. 656 (January 2008)**.
- B. ROFFIA and A. ZAGHINI, *Excess money growth and inflation dynamics*, International Finance, v. 10, 3, pp. 241-280, **TD No. 657 (January 2008)**.
- G. BARBIERI, P. CIPOLLONE and P. SESTITO, *Labour market for teachers: demographic characteristics and allocative mechanisms*, Giornale degli economisti e annali di economia, v. 66, 3, pp. 335-373, **TD No. 672 (June 2008)**.
- E. BREDA, R. CAPPARIELLO and R. ZIZZA, *Vertical specialisation in Europe: evidence from the import content of exports*, Rivista di politica economica, numero monografico, **TD No. 682 (August 2008)**.

2008

- P. ANGELINI, *Liquidity and announcement effects in the euro area*, Giornale degli Economisti e Annali di Economia, v. 67, 1, pp. 1-20, **TD No. 451 (October 2002)**.
- P. ANGELINI, P. DEL GIOVANE, S. SIVIERO and D. TERLIZZESE, *Monetary policy in a monetary union: What role for regional information?*, International Journal of Central Banking, v. 4, 3, pp. 1-28, **TD No. 457 (December 2002)**.
- F. SCHIVARDI and R. TORRINI, *Identifying the effects of firing restrictions through size-contingent Differences in regulation*, Labour Economics, v. 15, 3, pp. 482-511, **TD No. 504 (June 2004)**.
- L. GUIISO and M. PAIELLA, *Risk aversion, wealth and background risk*, Journal of the European Economic Association, v. 6, 6, pp. 1109-1150, **TD No. 483 (September 2003)**.
- C. BIANCOTTI, G. D'ALESSIO and A. NERI, *Measurement errors in the Bank of Italy's survey of household income and wealth*, Review of Income and Wealth, v. 54, 3, pp. 466-493, **TD No. 520 (October 2004)**.
- S. MOMIGLIANO, J. HENRY and P. HERNÁNDEZ DE COS, *The impact of government budget on prices: Evidence from macroeconomic models*, Journal of Policy Modelling, v. 30, 1, pp. 123-143 **TD No. 523 (October 2004)**.
- L. GAMBACORTA, *How do banks set interest rates?*, European Economic Review, v. 52, 5, pp. 792-819, **TD No. 542 (February 2005)**.
- P. ANGELINI and A. GENERALE, *On the evolution of firm size distributions*, American Economic Review, v. 98, 1, pp. 426-438, **TD No. 549 (June 2005)**.
- R. FELICI and M. PAGNINI, *Distance, bank heterogeneity and entry in local banking markets*, The Journal of Industrial Economics, v. 56, 3, pp. 500-534, **No. 557 (June 2005)**.
- S. DI ADDARIO and E. PATACCHINI, *Wages and the city. Evidence from Italy*, Labour Economics, v.15, 5, pp. 1040-1061, **TD No. 570 (January 2006)**.
- M. PERICOLI and M. TABOGA, *Canonical term-structure models with observable factors and the dynamics of bond risk premia*, Journal of Money, Credit and Banking, v. 40, 7, pp. 1471-88, **TD No. 580 (February 2006)**.
- E. VIVIANO, *Entry regulations and labour market outcomes. Evidence from the Italian retail trade sector*, Labour Economics, v. 15, 6, pp. 1200-1222, **TD No. 594 (May 2006)**.
- S. FEDERICO and G. A. MINERVA, *Outward FDI and local employment growth in Italy*, Review of World Economics, Journal of Money, Credit and Banking, v. 144, 2, pp. 295-324, **TD No. 613 (February 2007)**.
- F. Busetti and A. HARVEY, *Testing for trend*, Econometric Theory, v. 24, 1, pp. 72-87, **TD No. 614 (February 2007)**.
- V. CESTARI, P. DEL GIOVANE and C. ROSSI-ARNAUD, *Memory for prices and the Euro cash changeover: an analysis for cinema prices in Italy*, In P. Del Giovane e R. Sabbatini (eds.), The Euro Inflation and Consumers' Perceptions. Lessons from Italy, Berlin-Heidelberg, Springer, **TD No. 619 (February 2007)**.
- B. H. HALL, F. LOTTI and J. MAIRESSE, *Employment, innovation and productivity: evidence from Italian manufacturing microdata*, Industrial and Corporate Change, v. 17, 4, pp. 813-839, **TD No. 622 (April 2007)**.

- J. SOUSA and A. ZAGHINI, *Monetary policy shocks in the Euro Area and global liquidity spillovers*, International Journal of Finance and Economics, v.13, 3, pp. 205-218, **TD No. 629 (June 2007)**.
- M. DEL GATTO, GIANMARCO I. P. OTTAVIANO and M. PAGNINI, *Openness to trade and industry cost dispersion: Evidence from a panel of Italian firms*, Journal of Regional Science, v. 48, 1, pp. 97-129, **TD No. 635 (June 2007)**.
- P. DEL GIOVANE, S. FABIANI and R. SABBATINI, *What's behind "inflation perceptions"? A survey-based analysis of Italian consumers*, in P. Del Giovane e R. Sabbatini (eds.), *The Euro Inflation and Consumers' Perceptions. Lessons from Italy*, Berlin-Heidelberg, Springer, **TD No. 655 (January 2008)**.
- B. BORTOLOTTI, and P. PINOTTI, *Delayed privatization*, Public Choice, v. 136, 3-4, pp. 331-351, **TD No. 663 (April 2008)**.
- R. BONCI and F. COLUMBA, *Monetary policy effects: New evidence from the Italian flow of funds*, Applied Economics, v. 40, 21, pp. 2803-2818, **TD No. 678 (June 2008)**.
- M. CUCCULELLI, and G. MICUCCI, *Family Succession and firm performance: evidence from Italian family firms*, Journal of Corporate Finance, v. 14, 1, pp. 17-31, **TD No. 680 (June 2008)**.
- A. SILVESTRINI and D. VEREDAS, *Temporal aggregation of univariate and multivariate time series models: a survey*, Journal of Economic Surveys, v. 22, 3, pp. 458-497, **TD No. 685 (August 2008)**.

2009

- S. MAGRI, *The financing of small innovative firms: the Italian case*, Economics of Innovation and New Technology, v. 18, 2, pp. 181-204, **TD No. 640 (September 2007)**.
- S. MAGRI, *The financing of small entrepreneurs in Italy*, Annals of Finance, v. 5, 3-4, pp. 397-419, **TD No. 640 (September 2007)**.
- R. GOLINELLI and S. MOMIGLIANO, *The Cyclical Reaction of Fiscal Policies in the Euro Area. A Critical Survey of Empirical Research*, Fiscal Studies, v. 30, 1, pp. 39-72, **TD No. 654 (January 2008)**.
- L. ARCIERO, C. BIANCOTTI, L. D'AURIZIO and C. IMPENNA, *Exploring agent-based methods for the analysis of payment systems: A crisis model for StarLogo TNG*, Journal of Artificial Societies and Social Simulation, v. 12, 1, **TD No. 686 (August 2008)**.
- A. CALZA and A. ZAGHINI, *Nonlinearities in the dynamics of the euro area demand for M1*, Macroeconomic Dynamics, v. 13, 1, pp. 1-19, **TD No. 690 (September 2008)**.
- L. FRANCESCO and A. SECCHI, *Technological change and the households' demand for currency*, Journal of Monetary Economics, v. 56, 2, pp. 222-230, **TD No. 697 (December 2008)**.

FORTHCOMING

- L. MONTEFORTE and S. SIVIERO, *The Economic Consequences of Euro Area Modelling Shortcuts*, Applied Economics, **TD No. 458 (December 2002)**.
- M. BUGAMELLI and A. ROSOLIA, *Produttività e concorrenza estera*, Rivista di politica economica, **TD No. 578 (February 2006)**.
- M. PERICOLI and M. SBRACIA, *The CAPM and the risk appetite index: theoretical differences, empirical similarities, and implementation problems*, International Finance, **TD No. 586 (March 2006)**.
- R. BRONZINI and P. PISELLI, *Determinants of long-run regional productivity with geographical spillovers: the role of R&D, human capital and public infrastructure*, Regional Science and Urban Economics, **TD No. 597 (September 2006)**.
- E. IOSSA and G. PALUMBO, *Over-optimism and lender liability in the consumer credit market*, Oxford Economic Papers, **TD No. 598 (September 2006)**.
- U. ALBERTAZZI and L. GAMBACORTA, *Bank profitability and the business cycle*, Journal of Financial Stability, **TD No. 601 (September 2006)**.
- A. CIARLONE, P. PISELLI and G. TREBESCHI, *Emerging Markets' Spreads and Global Financial Conditions*, Journal of International Financial Markets, Institutions & Money, **TD No. 637 (June 2007)**.

- V. DI GIACINTO and G. MICUCCI, *The producer service sector in Italy: long-term growth and its local determinants*, Spatial Economic Analysis, **TD No. 643 (September 2007)**.
- L. FORNI, L. MONTEFORTE and L. SESSA, *The general equilibrium effects of fiscal policy: Estimates for the euro area*, Journal of Public Economics, **TD No. 652 (November 2007)**.
- Y. ALTUNBAS, L. GAMBACORTA and D. MARQUÉS, *Securitisation and the bank lending channel*, European Economic Review, **TD No. 653 (November 2007)**.
- P. DEL GIOVANE, S. FABIANI and R. SABBATINI, *What's behind "Inflation Perceptions"? A survey-based analysis of Italian consumers*, Giornale degli Economisti e Annali di Economia, **TD No. 655 (January 2008)**.
- A. ACCETTURO, *Agglomeration and growth: the effects of commuting costs*, Papers in Regional Science, **TD No. 688 (September 2008)**.
- L. FORNI, A. GERALI and M. PISANI, *Macroeconomic effects of greater competition in the service sector: the case of Italy*, Macroeconomic Dynamics, **TD No. 706 (March 2009)**.