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**A note Hadamard differentiability and differentiability
in quadratic mean**

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A note on Hadamard differentiability and differentiability in quadratic mean

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Abstract

We prove that Hadamard differentiability in addition with usual assumptions on the loss function for M estimates implies differentiability in quadratic mean. Thus both concepts are exchangeable.

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1 Introduction

As it was pointed out by [Pollard, 1997] there is a connection between differentiability in an L^2 sense, differentiability in quadratic mean and local asymptotic normality. The idea of local asymptotic normality (LAN) goes back to [LeCam, 1970], who showed that differentiability in quadratic mean implies a quadratic approximation for the log-likelihood function, without using the 2nd derivative. We will show, how a different concept of differentiability, namely the Hadamard differentials, are under mild regularity conditions equivalent to the definition of differentials in quadratic mean, and therefore the results of [Pollard, 1997], [LeCam, 1970] and [Fernholz, 1983] will coincide for M estimates. Thus it can be seen, that it is also sufficient for M estimates to be Hadamard differentiable to enjoy the LAN property.

2 The equivalence Theorem

Throughout this article we regard an σ -finite probability space and a sampling X_i (not necessarily iid.) from a measurable random variable $X : \Omega \rightarrow \mathbb{R}$, where X has a distribution function, which is absolutely continuous w.r.t. some σ -finite measure. Define a M estimate as the solution of the following minimization problem

$$\hat{\theta}_n = \arg \min_{\theta \in \Theta} \frac{1}{n} \sum_{i=1}^n \rho(X_i, \theta), \quad (1)$$

where θ is some unknown parameter, where $\theta \in \Theta$, and Θ is some metric space with the uniform metric. Under standard conditions, that $\theta \rightarrow \rho(\theta)$ is continuous and Θ is compact as well as ρ is uniformly integrable w.r.t. θ , we get that $\hat{\theta}_n \xrightarrow{P} \theta_0$, where

$$\theta_0 = \arg \min_{\theta \in \Theta} E[\rho(X, \theta)]. \quad (2)$$

We suppress the dependence of ρ on $\omega \in \Omega$ from now on, and treat ρ as a function of the parameter(-vector) θ . If ρ is differentiable w.r.t. to θ , one may take a first-order Taylor series expansion around $\hat{\theta}_n$ in neighborhood of θ_0 , where (2) attains its minimum:

$$\rho(\hat{\theta}_n) = \rho(\theta_0) + (\hat{\theta}_n - \theta_0)^T \rho'_\theta(\theta_0) + Rem(\theta)(\hat{\theta}_n - \theta_0). \quad (3)$$

Suppose that $Rem \rightarrow 0$ in some proper sense. If $Rem \xrightarrow{L^2} 0$, i.e. $E[|Rem(\hat{\theta}_n - \theta_0)|^2] = 0$, we call ρ differentiable in quadratic mean (DQM). If ρ is Hadamard differentiable,

then $Rem \xrightarrow{P} 0$, see [Fernholz, 1983]. Under usual regularity conditions on ρ and its derivative we can conclude, that $\sqrt{n}(\hat{\theta}_n - \theta_0) \rightarrow N(0, \sigma^2)$, with $\sigma^2 = \text{var}(\phi(X))$, and ϕ is the influence function. Define Hadamard differentials as follows (see [van der Vaart, 1998]):

Definition 1 A function $\rho : D_\theta \rightarrow W$ is Hadamard differentiable (or compact differentiable) at $\theta \in D_\theta$ if there exists $\rho'_\theta \in L_1(D_\theta, W)$ such that for any compact set Γ of D_θ ,

$$\lim_{t \rightarrow 0} \left\| \frac{\rho(\theta + th_t) - \rho(\theta)}{t} - \rho'_\theta(h) \right\|_W \rightarrow 0, \quad (4)$$

for every $h_t \rightarrow h$, s.t. $\theta + th_t \in D_\theta$ for $t > 0$ and t small enough.

One may write instead of (4):

$$Rem_1(\hat{\theta}_n - \theta_0) = \rho(\hat{\theta}_n) - \rho(\theta_0) - (\hat{\theta}_n - \theta_0)^T \rho'_\theta(\theta_0) \quad (5)$$

as the remainder term of (3). Note that ρ'_θ indicates the Hadamard derivative. It is shown by [Fernholz, 1983], that if

$$\lim_{\theta \rightarrow \theta_0} \frac{Rem_1(\theta - \theta_0)}{|\theta - \theta_0|} = 0,$$

then: $\sqrt{n}Rem_1(\theta - \theta_0) = o_P(1)$. DQM can be defined using (3):

Definition 2 $\rho(\theta)$ is said to be DQM if there exists a function $D\rho_\theta \in L^2(P)$, s.t.

$$\lim_{\theta \rightarrow \theta_0} \frac{\|\rho(\theta) - \rho(\theta_0) - (\theta - \theta_0)^T D\rho_\theta(\theta_0)\|_2}{\|\theta - \theta_0\|} = 0 \quad (6)$$

Once again we write for (3):

$$Rem_2(\hat{\theta}_n - \theta_0) = \rho(\hat{\theta}_n) - \rho(\theta_0) - (\hat{\theta}_n - \theta_0)^T D\rho_\theta(\theta_0) \quad (7)$$

We now state the result:

Theorem 1 Suppose $\rho(\theta)$ is Hadamard differentiable w.r.t. θ in a neighborhood N_θ of θ_0 , $\sup E[|\rho(\theta)|] < \infty$ and $\sup E[\rho^2(\theta)] < \infty$ for $\theta \in N_\theta$ and $\hat{\theta}_n$ is a consistent estimate for θ_0 , then $\rho(\theta)$ is also differentiable in quadratic mean, and both derivative coincide on N_θ .

Proof Write $t = \hat{\theta}_n - \theta_0$, $\theta = \theta_0$ and $h_t \rightarrow 0$ in (4), so we can rewrite (4):

$$\lim_{\hat{\theta}_n \rightarrow \theta_0} \left\| \frac{\rho(\hat{\theta}_n) - \rho(\theta_0) - (\hat{\theta}_n - \theta_0)^T \rho'_\theta(0)}{\hat{\theta} - \theta_0} \right\|_W = 0 \quad (8)$$

Because $\hat{\theta}_n \xrightarrow{P} \theta_0$ is a Cauchy sequence we have for $\hat{\theta}_n \in N_\theta$, that $|\hat{\theta}_n - \theta_0| \rightarrow 0$ and we get expression (6). We now have to show, that both remainder terms converge in a L^2 sense, as well as in probability to zero. Then from (3) we can deduce that $Rem_1 = Rem_2$ a.s. and therefore both derivatives must coincide a.s.. From the assumptions imposed on ρ we can conclude that ρ is uniformly integrable (see corollary 6.22 in [Klenke, 2006]). If $\hat{\theta}_n \xrightarrow{P} \theta_0$ we have that $\sqrt{n}Rem_1(\hat{\theta}_n - \theta_0) = o_P(1)$ and $\sqrt{n}Rem_2(\hat{\theta}_n - \theta_0) \xrightarrow{L^2} 0$. As $\rho(\theta) \in L^2(P)$ it follows with continuity of $\rho(\theta)$ w.r.t. θ that $h = |\rho'_\theta| \in L^2$, where ρ'_θ is the Hadamard derivative. Therefore we have $Rem_1(\hat{\theta}_n - \theta_0) \in L^2$ for n large enough. We now have: Rem_1 is uniformly integrable, and $Rem_1 \xrightarrow{P} 0$ and thus

$$|\sqrt{n}Rem_1(\hat{\theta}_n - \theta_0)|^2 \xrightarrow{n \rightarrow \infty} |0|^2$$

Write $g = |\sqrt{n}Rem_1(\hat{\theta}_n - \theta_0) - 0|^2$, then $g_n \xrightarrow{P} 0$ and g_n is uniformly integrable, as

$$g_n = |\sqrt{n}Rem_1(\hat{\theta}_n - \theta_0)|^2 \leq 4 \cdot |o_P(1)|^2.$$

Therefore $g_n \in L^2(P)$ and thus $D\rho_\theta = \rho'_\theta$ a.s. in N_θ .

□

3 Conclusion

We proved the equivalence of Hadamard differentiability and differentiability in quadratic mean for M estimates usually considered in robust statistics. This result can be used to deduce alternative assumptions imposed on M estimators to be consistent and asymptotically normally distributed.

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