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Quaderno N. 08-05

Decanato della Facoltà di Scienze economiche
Via G. Buffi, 13 CH-6900 Lugano

New empirical evidence on local financial development and growth[§]

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This draft: 13 May 2008

[§] The authors would like to thank Alessandra Guariglia and David Roodman for help with the Stata module xtabond2. Andrea Vaona gratefully acknowledges financial support from the Italian National Research Council. The usual disclaimer applies.

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Abstract

In this paper, we show that the regional finance-growth nexus in Italy is robust to a series of innovations with respect to the existing literature on the topic. We use finer measures of economic and financial development, as well as instruments with a deeper economic content. We rely on state-of-the-art cross-sectional and panel estimation methods, and we offer a thorough investigation of the nonlinearities in the relation between finance and growth. Our results show that, while local financial development is a key factor for economic growth, in regions with inefficient courts more credit might translate into reduced growth due to opportunistic behaviour and the consequent misallocation of funds.

Jel codes: O18, O16, C31.

Keywords: Finance, Growth, Regions, Italy, Cross-Section Analysis, Panel Data Analysis.

1. Background

Since the groundbreaking contributions of King and Levine (1993a,b), economists have shown renewed interest in the finance–growth nexus. Results have often been conflicting across studies using a variety of indicators for financial development, different econometric methods, model specifications, and datasets. In this regard, a thorough review is offered by Levine (2004).

Focusing on cross-country studies using measures of financial development connected to the banking sector, Levine (1998), Levine et al. (2000) and Beck et al. (2000) designed an innovative research strategy. In order to overcome possible biases deriving either from the endogeneity of financial development indicators within a growth regression, or from the omission of possible relevant variables, they took two steps. In the first step, they resorted to the following model within a cross-sectional framework

$$G = \alpha + \beta F + \gamma X + \varepsilon, \quad (1)$$

where: G was either per capita GDP growth, or growth of capital stock per head, or productivity growth; F was different financial development indicators; X was a set of controls (income per capita, education, political stability, indicators of exchange rate developments, international trade, fiscal and monetary policy); α , β and γ were coefficients; and ε was the stochastic error. F was instrumented by indicators of origin of the legal system, which, according to the literature, affects the letter and the enforcement of national credit laws, protecting external investors and promoting financial development to different extents.

In the second step, they resorted to dynamic panel models, with the following specification

$$y_{i,t} = \alpha y_{i,t-1} + \beta' Z_{i,t} + \eta_i + \varepsilon_{i,t}, \quad (2)$$

where: $y_{i,t}$ was the log of real per capita GDP at time t in country i ; $Z_{i,t}$ was a set of controls including financial indicators; η_i was an unobserved country-specific effect; and $\varepsilon_{i,t}$ a stochastic error.

However, the earlier promising results of Levine and co-authors need to be qualified, since they rely on a strong assumption: the poolability of countries at different stages of development. This assumption has found scant empirical support in the empirical literature (Arestis et al., 2004, Rioja and Valev, 2004, Loyaza and Ranciere, 2005, Apergis et al., 2007, Schiavo and Vaona, 2007).

As a matter of consequence, Vaona (2008a), using Italian data, proposed to apply the research strategy designed by Levine and co-authors to a regional setting, where poolability is supported by statistical tests. In this context, financial development can have an impact on local growth due to market segmentation (Guiso et al., 2004a).

To this purpose, an Italian dataset appears particularly interesting, given the intensive monitoring activity implemented by the Bank of Italy on local credit market conditions, and given the considerable

attention that the international literature has devoted to Italian regions so far (Lucchetti et al., 2001; Guiso et al., 2004a, b, 2006; Usai and Vannini, 2005; Jappelli et al., 2005).

In order to show the robustness of our results, we follow the setup of Vaona (2008a), using a sample of Italian NUTS-3 regions, performing both cross-section and dynamic panel estimations, and controlling for endogeneity, spatial unobserved heterogeneity, spatial correlation and poolability. For the dynamic panel estimates we use three-year averages in order to overcome possible distortions deriving from business cycles, and we rely on Roodman (2005).

2. Model Specification and Innovations of the Paper

Regarding the model specification, on the one hand, our choices for G , F and y represent one of the elements of novelty of the paper, and are illustrated below. On the other hand, for X and Z , we follow Vaona (2008a). For X , we use: the number of crimes per head; the number of students attending secondary schools over resident population; the sum of exports and imports over GDP; the expenditure in finished public infrastructures over GDP; and the real GDP per head. All the variables are taken at the first year of observation, namely 1986. In Z , we include: the number of crimes per head;¹ the number of students attending secondary schools over resident population; and the sum of exports and imports over GDP. We could not include the expenditure in finished public infrastructures, since no relevant data is available after 2000.

With respect to Vaona (2008a), we introduce four important novelties.

First, the aforementioned study considers, as an indicator of economic development, the growth rate of real per capita value-added in the cross-section estimation, and the level of per capita value-added in the dynamic panel estimation. On the other hand, following Jappelli et al. (2005) and Fabiani and Pellegrini (1997), we use the growth rate and the level of real per capita GDP for G in model (1) and for y in model (2), respectively.

Second, we use a finer measure of financial development. Since Vaona (2008b) showed that credit to households and non-profit organisations, credit to non-financial enterprises with a public structure, and credit to the public administration are not connected to economic growth, we focus here only on credit supplied by private banks to private firms, taking – in a regional setting – a similar step to the one taken by Levine et al. (2000) in the cross-country literature. On the one hand, our measure of financial development is aggregate, so it is not a direct indicator of credit rationing based on micro data, as the one used by Guiso et al. (2004a). On the other hand, its availability for many years makes it possible to show not only cross-sectional but also longitudinal evidence, as in cross-country studies. Furthermore, Jappelli et al. (2005) showed that different measures of financial development in Italian regions tend to show similar geographical patterns.

¹ We also included its square term in order to control for possible nonlinear effects of crime on economic growth (see, for the case of Italy, Peri, 2004).

Third, the instruments used by Vaona (2008a) in his cross-section estimation have mainly a statistical nature, being regional dummies not related with future growth, which should capture regional disparities in the efficiency and efficacy of courts within a country. Here, similarly to Jappelli et al. (2005), we consider as instruments two direct measures of efficiency and efficacy of the judiciary system: the length of first instance and appeal civil trials. In particular, we consider both their un-weighted sum and the sum weighted by the number of trials in each instance. We do so in order to capture the possible effects of the presence of agents, who discount time in different ways and, therefore, could weight differently the length of first instance and appeal trials. In the instruments set of the dynamic panel data estimator, we include the lags in the levels and in the first differences of logs of the two variables above, as well as of private credit to private firms. Descriptive statistics for all the data are set out in Table 1.

Fourth, and finally, when considering dynamic panel data estimation, we manage to capture the nonlinearities underlying the finance-growth nexus, by fitting to the data a dynamic model of the (log of) real per capita GDP containing a third-order polynomial (in logs) of our indicator of financial development, besides the further controls sketched above. Consistently with Rioja and Valev (2004), we find that the higher is the degree of local financial development, the weaker is its impact on growth. However, consistently with De Gregorio and Guidotti (1995) and Guariglia and Poncet (2007) and similarly to Rioja and Valev (2004), we also find that, in provinces with an inefficient allocation of credit, a higher ratio between credit and local GDP can have a negative impact on growth. Our identification strategy allows us to propose a new explanation of why this might happen.² In regions with a highly inefficient judiciary system, opportunistic behaviour is widespread. Under these conditions, banks fund projects doomed to fail and to absorb more economic resources than their returns. Consequently, the growth rate of local economies may be damaged – rather than benefited – by a greater availability of credit.

The following section illustrates our results and concludes.

3. Results and Conclusions

Table 1 sets out our cross-sectional results. We group our sample of NUTS-3 provinces according to the NUTS-2 region they belong to. In this way, it is possible to use static panel estimators also in a cross-sectional setting, controlling better for unobserved heterogeneity. First, we fit a random-effects model and a fixed-effects model to the data. A Hausman test points to the latter as the one to be preferred. Subsequently, we drop insignificant regional dummy variables, and we implement a 2-stage least squares dummy variables (2SLSDV) estimation, using the instruments mentioned above. An F test supports our instruments, and the first stage results show that an increase of one day in both the

² Other possible theoretical justifications for this pattern can be found in Rioja and Valev (2004). They are based on economies of scale, minimum size requirements, and learning-by doing effects in banking.

weighted and un-weighted sums of the length of first instance and appeal civil trials decreases the ratio of private credit to private firms over GDP by 0.13 per cent.³ The model is supported also by a test for overidentifying restrictions. Though a Durbin-Wu-Hausman test points to the LSDV estimator as the one to be preferred, in both 2SLS and LSDV the main forces driving local economic growth in Italy appear to be financial development and convergence, which is consistent with the findings of Vaona (2008a). Southern regions like Puglia and Campania appear to have specific factors that make them lag behind the rest of the country. No spatial correlation is found in the residuals.

A similar picture emerges when considering panel data estimates (Table 3). Convergence is still a major factor affecting growth, given the significance of the first lag of real per capita GDP. However, the exposure to international markets is important as well, possibly reflecting the consequent better allocation of production factors. As anticipated above, financial development has a highly nonlinear effect on growth. A third-order polynomial in private credit to private firms turns out to be significant at a 5 per cent level. In order to better assess the connection between finance and growth, we simulate the model over the interval between the minimum and maximum values of private credit to private firms over GDP, keeping the significant controls at their average values, and forcing the coefficient of insignificant variables to zero. Figure 1 shows our results. As anticipated in the previous section, finance has a positive – though declining – effect on growth. However, in regions where the judiciary system is inefficient and opportunistic behaviour is more widespread, more finance might translate into less growth, as more resources are inefficiently allocated to poor projects.

In order to assess the poolability of provinces belonging to the South, the Centre, the North-West and the North-East of Italy, we interacted the third-order polynomial in our measure of financial development with dummies for each of the four Italian macro-regions.⁴ A Wald test for the equality of the coefficients across macro-regions could not reject the null. All the other specification tests support the model.

Summarizing, in this study we showed that the finance-growth nexus at the regional level in Italy is robust to a series of innovations in the analysis carried out. With respect to previous studies, we used more refined measures of economic growth and financial development. Moreover, we used instruments with a deeper economic content than earlier, and we offered a better assessment of the nonlinearities characterizing the finance-growth nexus. To conclude, the judiciary system is of paramount importance in understanding this economic issue. The more efficient are the courts located in a local economy, the better will its credit market work, and funds will be less likely to be allocated to projects doomed to fail.

³ We compute elasticities at the average value of regressors.

⁴ A picture of the Italian macro-regions is available in Vaona (2008a).

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Figure 1 – The effect of local financial development on real GDP per capita – simulation results

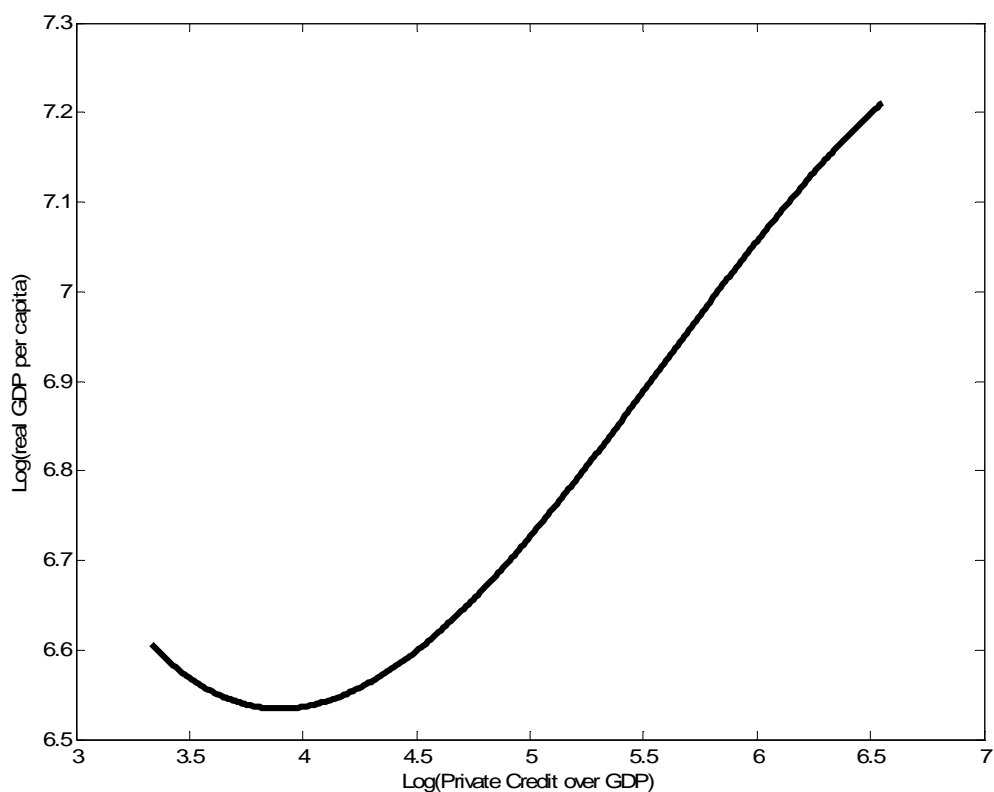


Table 1 – Descriptive statistics

| | | Mean | Standard Deviation | Min. | Max. |
|---------------------------|---|-------------|---------------------------|-------------|-------------|
| Cross-section data | Total percentage growth of real GDP per head between 1986 and 2006^a | 46.67 | 21.06 | -42.04 | 102.63 |
| | Private credit to private firms over GDP in 1986^b | 225.58 | 73.34 | 69.16 | 410.99 |
| | Sum of exports and imports over GDP in 1986^b | 271.95 | 150.70 | 14.10 | 667.41 |
| | Value of finished public infrastructures over GDP in 1986 | 18.19 | 12.80 | 3.49 | 73.04 |
| | Students attending secondary school over resident population in 1986^c | 45.66 | 5.79 | 33.95 | 58.32 |
| | Real GDP per head in 1986^{a,c} | 1584.297 | 323.05 | 832 | 2272 |
| | Crimes per head in 1986^d | 29.41 | 11.85 | 10.05 | 81.51 |
| | Un-weighted sum of the length of first instance and appeal civil trials in days | 718 | 135.791 | 577 | 1243 |
| | Weighted sum of the length of first instance and appeal civil trials in days | 1405.063 | 207.0759 | 1185 | 2202 |
| | | | | | |
| Panel data | Private credit to private firms over GDP^b | 211.58 | 114.37 | 27.96 | 703.27 |
| | Real GDP per head^{a,c} | 1966.37 | 441.44 | 644.67 | 3173.67 |
| | Sum of exports and imports over GDP^b | 327.51 | 195.77 | 14.197 | 1181.67 |
| | Students attending secondary school over resident population^c | 45.08 | 8.01 | 29.48 | 65.10 |
| | Crimes per head^d | 37.17 | 14.18 | 10.51 | 111.19 |
| | Un-weighted sum of the length of first instance and appeal civil trials in days | 792.27 | 174.62 | 385 | 1493 |
| | Weighted sum of the length of first instance and appeal civil trials in days | 1651.02 | 332.04 | 826.67 | 2554 |

Notes:

^a all real variables were deflated using the CPI of the main city of each province. The base year is 1985. ^b billions of eurolire over millions of eurolire (by eurolire we mean that for the years after the introduction of the Euro we converted the value of the variables into Lire by using the official exchange rate of €1 = £1936.27). To obtain percentages, divide by 10. ^c number of students over resident population in thousands. To obtain percentages divide by 10. ^d number of crimes over resident population in thousands. To obtain the number of crimes per inhabitant, divide by 1000. ^e in ten thousands of eurolire.

Table 2 – The effect of financial development on real economic growth in cross-section models

Dependent variable: total real growth rate of GDP per head between 1986 and 2006

| | LSDV | 2SLSDV |
|--|---------|---------|
| Private credit to private firms over GDP in 1986 | 7.36* | 21.17* |
| t-stat. | (2.23) | (2.43) |
| Sum of exports and imports over GDP in 1986 ^a | 0.05 | -0.01 |
| t-stat. | (0.27) | (-0.64) |
| Students attending secondary school over resident population in 1986 | 0.65 | 0.42 |
| t-stat. | (1.36) | (0.75) |
| Expenditure in finished public infrastructures over GDP in 1986 | 0.22 | 0.37 |
| t-stat. | (1.22) | (1.68) |
| GDP per head in 1986 | -39.88* | -44.48* |
| t-stat. | (-3.83) | (-3.63) |
| Crimes per head in 1986 | 0.24 | 0.18 |
| t-stat. | (1.14) | (0.75) |
| Constant | 53.28 | 43.76 |
| t-stat. | (1.67) | (1.18) |
| Dummy Campania ^b | -32.49 | -26.42 |
| t-stat. | (-1.80) | (-1.25) |
| Dummy Puglia | -43.37* | -30.20 |
| t-stat. | (-2.43) | (-1.86) |
| Dummy Sicilia | -31.38* | -41.91* |
| t-stat. | (-2.22) | (-2.04) |
| R ² | 0.47 | 0.30 |
| Moran's I ^c | 0.55 | -0.73 |
| Hausman test (p-value) ^d | 0.00 | - |
| Durbin – Wu – Hausman test (p-value) ^e | - | 0.96 |
| Instrumental variable F-test (p-value) ^f | - | 0.00 |
| Test for overidentifying restrictions (p-value) ^g | - | 0.74 |
| Observations | 64 | 64 |

Notes:

* denotes coefficients significant at the 5% level. T-statistics are shown in parentheses. Instruments in the 2SLSDV regression in the second column include the weighted and the un-weighted sums of the length of first instance and appeal civil trials. ^a exports and imports include only goods traded on international markets, since interregional trade is not registered. ^b this dummy was significant when including a full set of regional dummies. So we chose to keep it in the model. ^c the null is no spatial correlation. ^d the null is that the random-effects and the fixed-effects estimators are similar. If they are not similar, the latter will be preferred. ^e the null is no endogeneity in the comparison between the LSDV and the 2SLSDV estimators. ^f the null is that the instruments are not correlated with the instrumented variables. ^g the null is that instruments in excess with respect to instrumented variables are not correlated with 2SLS residuals.

Table 3 - The effect of financial development on real economic growth – dynamic panel estimates (1986-2003)

Dependent variable: log of real GDP per capita.

Method: System-GMM.

Frequency of the data: three year averages.

| | | | |
|--|---------|---|---------|
| Log(Private credit to private firms over GDP) | -3.40* | Dummy Puglia | -0.14* |
| t-stat. | (-2.07) | t-stat. | (-2.22) |
| [Log(Private credit to private firms over GDP)]² | 0.67* | Dummy for the years 1992-1994 | -0.08* |
| t-stat. | (2.09) | t-stat. | (-5.00) |
| [Log(Private credit to private firms over GDP)]³ | -0.04* | Dummy for the years 1995-1997 | -0.04* |
| t-stat. | (-2.11) | t-stat. | (0.00) |
| Log(real GDP per capita) _{t-1} | 0.81* | Test for first order serial correlation (p-value) ^b | 0.03 |
| t-stat. | (8.05) | Test for second order serial correlation (p-value) ^c | 0.59 |
| Log(Students attending secondary school over resident population) | 0.11 | Hansen test for overident. restrictions (p-value) ^d | 0.13 |
| t-stat. | (1.19) | Moran's I (p-value) ^e | 0.88 |
| Log(Sum of exports and imports over GDP) ^a | 0.05* | Wald test for poolability ^f | 0.51 |
| t-stat. | (2.79) | | |
| Log(Crimes per head) | 0.37 | | |
| t-stat. | (0.45) | Number of instruments | 53 |
| [Log(Crimes per head)] ² | -0.03 | Number of provinces | 77 |
| t-stat. | (-0.31) | Number of obs. | 359 |

Notes: the instruments are past first differences and past levels of (Private credit to private firms over GDP)_{t-1}, logs of the weighted and un-weighted sums of the length of first instance and appeal civil trials. * denotes coefficients significant at the 5% level. t-statistics are shown in parentheses. ^a exports and imports include only goods traded on international markets as interregional trade is not registered. ^b the null is absence of first order serial correlation in the differenced residuals. Presence of first order serial correlation in the differenced residuals does not affect the validity of estimates. ^c the null is absence of second order serial correlation in the differenced residuals. ^d the null is that instruments in excess with respect to instrumented variables are not correlated with GMM residuals. We preferred the Hansen test to the Sargan test, because the former is robust to heteroskedasticity, and we do not have a large number of instruments when compared to the cross-sectional dimension of the dataset. ^e the null is no spatial correlation. ^f the null is that the coefficient of the financial indicator is the same across provinces belonging to the South, the North-East, the North West and the Centre of Italy, respectively.

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