Financial Development, Property Rights, and Growth

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Abstract

Claessens and Laeven analyze how property rights affect the allocation of firms' available resources among different types of assets. In particular, they investigate empirically for a large number of countries whether firms in environments with more secure property rights allocate available resources more toward intangible assets and consequentially grow faster.

The authors find that improved asset allocation due to better property rights has an effect on growth in sectoral value added equal to improved access to financing arising from greater financial development. The results are robust, using various samples and specifications, including controlling for growth opportunities.

This paper—a product of the Policy Division, Financial Sector Strategy and Policy Department—is part of a larger effort in the department to study the link between finance and growth. Copies of the paper are available free from the World Bank, 1818 H Street NW, Washington, DC 20433. Please contact Rose Vo, room MC9-624, telephone 202-473-3722, fax 202-522-2031, email address hvo1@worldbank.org. Policy Research Working Papers are also posted on the Web at http://econ.worldbank.org. The authors may be contacted at stijn@fee.uva.nl or llaeven@worldbank.org. November 2002. (47 pages)

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Recently, a large number of papers have established that financial development fosters growth and that a country's financial development is related to its institutional characteristics, including its legal framework. The financial development and growth literature has established that finance matters for growth both at the macro-economic and micro-economic level (King and Levine 1993 and Levine 1997). The law and finance literature has found that financial markets are better developed in countries with strong legal frameworks. These well-developed financial markets make it easier for firms to attract financing for their investment needs (Rajan and Zingales 1998, La Porta et al. 1998, Demirgüç-Kunt and Maksimovic 1998). Related work has established that debt structures of firms differ across institutional frameworks (Rajan and Zingales 1995, Demirgüç-Kunt and Maksimovic 1999, and Booth et al. 2000).

Thus far the literature has not paid much attention to differences across countries in terms of firms' asset structure, i.e., differences in the allocation of investible funds by firms across various types of assets. However, these differences are large as well. Demirgüç-Kunt and Maksimovic (1999) find that firms in developing countries have higher proportions of fixed assets to total assets and less intangible assets than firms in developed countries. This is surprising as the literature on firms' optimal capital structure (Harris and Raviv 1991) suggests that a lack of long-term financing – typical in a developing country – would make it more difficult to finance fixed assets. Why is it that firms in developing countries have more fixed assets? Is it that they need more fixed collateral to attract external financing? Or does the preference for fixed assets and a corresponding lower share of intangible assets arise in countries with worse property

¹ In particular, it has been established that firms in developing countries have a smaller fraction of their total debt in the form of long-term debt.

rights because the returns on fixed assets are easier to secure from the firm's point of view than the returns on intangible assets? More generally, what is the role of property rights in terms of affecting investment patterns of firms?

In this paper, we empirically explore the role of property rights in influencing the allocation of investable resources. We start from the well-established proposition that greater financial sector development increases the availability of external resources and thereby enhances firm investment. We also acknowledge the literature demonstrating the importance of a good legal framework and well-established property rights for overall economic growth. In terms of channels through which property rights affect firm growth, we argue that the degree of property rights protection affects the allocation of investable resources. At the firm level, we can use the term property rights as referring to the protection of entrepreneurial and other investment in firm assets. We argue that a firm operating in a market with weaker property rights may be led to invest more in fixed assets relative to intangible assets as it finds it relatively more difficult to secure returns from intangible assets than from fixed assets.

The argument goes as follows. A firm is always at risk of not getting the returns from its assets (tangible or intangible) due to actions by its own employees, other firms, or the government. For the firm's employees and other firms (in particular, powerful competitors) it is relatively easy to steal the intangible assets of a firm if property rights are not secure. In a narrow sense, this is because the value of many intangible assets – patents (property rights to inventions and other technical improvements), copyrights (property rights to authors, artists, and composers), and trademarks (property rights for distinctive commercial marks or symbols) – purely derive from the existence of (intellectual) property rights. Without property rights protection, employees can simply

walk away with many a firm's intangible assets and competitors can easily copy them. As such, property rights in a narrow sense are very important for securing returns to intangible assets. Stealing physical property, such as buildings and machinery, in contrast is more difficult, particular for competing firms, even when general property rights are not secure. In a broader sense therefore, property rights matter more to secure returns from intangible assets than from tangible assets. Since there is no apparent reason to expect that the risk of expropriation by the government is higher for tangible assets than for intangible assets, it follows that property rights matter more for intangible assets than for tangible assets. More generally, we argue that the degree to which firms allocate resources in an optimal way will depend on the strength of a country's property rights, with the allocation effect being important for consequent firm growth.

Across countries, firm growth will also be affected by the development of financial markets. As such, there are two effects to consider in a cross-country study, a finance effect and an asset allocation effect. The finance effect will determine the available resources for investment and thus affect firm growth. The asset allocation effect will determine the efficiency of firm investment and thus also affect growth. We empirically investigate the importance of the finance and asset allocation effects for different industries in a large number of countries. We find less growth in countries with a lower level of financial development because firms lack access to finance and thus underinvest. And in countries with less secure property rights, there is less growth because the allocation of firms' investment is inefficient as firms underinvest in intangible assets. Empirically, the two effects are equally important drivers of growth in sectoral value added. The results are robust to using different country samples and estimation techniques, including instrumental variables and variations in country controls.

The paper is structured as follows. Section I reviews the related literature, develops the finance and asset allocation effects, and presents our methodology to separate the two effects empirically. Section II presents the data used in our empirical application. Section III presents the empirical results concerning the relationships between growth in value added and the finance and asset allocation effects. Section IV presents a number of robustness tests. Section V concludes.

I. Related Literature and Hypothesis

Our work is related to several strands of literature. The starting point is the work by King and Levine (1993), Levine and Zervos (1998), and Beck, Levine and Loayza (2000) that has established an empirical link between financial development and economic growth. Also related is the law and finance literature initiated by La Porta et al. (1997).² The law and finance literature has established that financial sector development is higher in countries with better legal systems and stronger creditor rights since such environments increase the ability of lenders to collateralize their loans and finance firms.

The second strand we draw on is the capital structure literature (Myers 1977, Titman and Wessels 1988, and Harris and Raviv 1991). This literature relates firms' liability structure to firm asset choices, among others. It has established that real, tangible assets, such as plant and equipment, can support more debt than intangible assets. In particular, fixed assets can support more long-term debt as they have greater liquidation and collateralizable value. Holding other factors constant, debt ratios will be lower the

² This literature focuses on the relationship between the institutional framework of a country and its financial development (see also La Porta et al. 1998, Rajan and Zingales 1998, and Demirgüç-Kunt and Maksimovic 1998).

larger the proportion of firm values represented by intangible assets (Myers 1977). Bradley et al. (1984) provide empirical support for the argument that a larger amount of intangible assets reduces the borrowing capacity of a firm.³

The third strand of literature relates to the role of property rights in affecting overall investment and investment patterns. Besley (1995) shows the role of property rights for investment incentives and provides evidence of the importance of property rights in the context of land ownership by farmers in Ghana. Johnson, McMillan and Woodruff (2002) show for a sample of firms in post-communist countries that weaker property rights discourage the reinvestment of firm earnings, even when bank loans are available, suggesting that secure property rights are both a necessary and sufficient condition for entrepreneurial investment. The role of property rights in affecting investment patterns has also been acknowledged, although less explicitly studied. Mansfield (1995) hints that there may be a relationship between protection of property rights and the allocation of investable resources between fixed and intangible assets. Using a survey of firm managers, he states that "most of the firms we contacted seemed to regard intellectual property rights protection to be an important factor" ... "[influencing] investment decisions". Stern, Porter and Furman (2000) show that the strength of a country's intellectual property rights affect its innovative capacity, as measured by the degree of international patenting. In developing countries, the lower degree of investment in intangible assets may relate to the weaker protection of property rights. More generally, the institutional economics literature (North 1990) suggests that investment in particular types of assets will be higher the more protected the property rights of the assets are.

³ Work by Rajan and Zingales (1995) and Demirgüç-Kunt and Maksimovic (1999) confirms that debt maturity and asset structures for cross-sections of countries are related in this way, with firms with more fixed assets being able to support a greater amount of long-term debt.

Empirically, these three strands have not yet merged in investigating the effects of institutions on both firm financing and asset allocation, and consequently growth. The law and finance literature has already established that firms in a country with a better legal framework and more developed financial markets find it easier to attract external financing. Empirical investigation of how a country's property rights protection affects firms' asset allocation has not yet occurred. Here we want to test two hypotheses: firms in countries with better developed financial systems will have more access to finance and will therefore be able to invest more overall; and, firms in countries with better property rights will invest more efficiently across types of assets. In turn, both aspects will be reflected in higher growth rates.

For our empirical test, we use the setup of Rajan and Zingales (1998, RZ hereafter) to assess the relationship between financial and legal development and growth. Let there be m countries, each indicated by index k, and n industries, each indicated by index j. The RZ-model then relates the growth in real value added in a sector j in a particular country k to a number of country and firm-specific variables. In case of RZ, the specific test focuses on financial development and the argument of RZ is that financially dependent firms can be expected to grow more in countries with a higher level of financial development. In addition to including country indicators and industry indicators, they overcome some of the identification problems encountered in standard cross-country growth regressions by interacting a country characteristic (financial development of a particular country) with an industry characteristic (external financial dependence of a particular industry). This approach is less subject to criticism regarding an omitted

⁴ Other papers that use this approach include Cetorelli and Gambera (2001), which investigates the effects of bank concentration on sectoral growth, and Fisman and Love (2002a), which investigates the effects of trade credit usage on sectoral growth.

variable bias or model specification than traditional approaches and allows them to isolate the impact of financial development on growth. In the regression results explaining sectoral growth, RZ find a positive sign for the interaction between the external financial dependence ratio and the level of financial development. They also find a similar effect when including an interaction term between the typical external dependence variable for the particular sector and the quality of a country's legal framework.

The results of RZ provide support for the law and finance effect. We expand the RZ model to test for the asset allocation effect. We add to the basic model in RZ a variable that is the interaction of the typical ratio for each industrial sector of intangible-to-fixed assets and an index of the strength of countries' property rights. We then test whether industrial sectors that typically use many intangible assets grow faster (slower) in countries with more (less) secure property rights. If intangible-intensive sectors grow more in countries with better property rights, then we have indirect evidence that property rights affect firms' asset choices and consequently through that channel growth. We also perform a number of robustness tests on the importance of controlling for country-specific factors and using instrumental variables to control for the possible (residual) endogeneity of some variables.

In line with RZ, we use US firm data to construct proxies at the industry level for the typical external dependence for a particular industrial sector and the typical ratio of intangible-to-fixed assets for a particular industry. The presumption here is that the well-developed financial markets and the well-protected property rights in the U.S. should allow US firms to achieve the desired financial and asset structures for their respective industrial sector. This approach offers a way to identify the desired extent of external

dependence and the optimal asset mix of an industry anywhere in the world.⁵ It assumes that there are technological and economic reasons why some industries depend more on external finance and intangible assets than others, and that these differences, to a large degree, prevail across countries. This does not mean that we assume a sector in two countries with the same degree of property protection to have exactly the same optimal mix of intangibles and tangible assets. Local conditions such as growth opportunities are allowed to differ between countries. We only assume the rank order of optimal asset mixes across industries to be similar across countries. Furthermore, we explicitly conduct tests for the importance of this assumption.

Following RZ, the regressions include the industry's market share in total manufacturing in the specific country to control for differences in growth potential across industries. Industries with large market shares may have less growth potential than industries with small initial market shares when there is an industry-specific convergence. The initial share may also help to control for other variations between countries, such as in their initial comparative advantage among certain industries based on factors other than financial development and property rights protection. Finally, in lines with RZ, we use country and industry dummies. The final specification is as follows:

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Growth _{j,k} = Constant + \beta_{1...m} · Country indicators + \beta_{m+1...m+n} · Industry indicators + \beta_{m+n+1} · (Industry j share of manufacturing in country k in 1980) + \beta_{m+n+2} · (External dependence US industry j · Financial development country k) + \beta_{m+n+3} · (Intangible intensity US industry j · Property rights country k) + \varepsilon_{j,k}.
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⁵ The advantage of this approach is that we do not need information on the actual asset mix for industries in different countries. The comparability of such data would be limited as accounting practices, in particular with respect to intangible assets, differ greatly around the world.

II. Data

We use industry-specific and country-specific data from a variety of sources. Table 1 presents an overview of the variables used in the empirical analysis and their sources. Most of the variables are self-explanatory and have been used in other cross-country studies of firm financing structures and firm growth.

In line with RZ, we use the ratio of private credit to GDP as proxy for financial development. As proxies for the level of protection of property rights, we use three broad indexes of property rights, two indexes of intellectual property rights, as well as a specific index of patent rights. These indexes of property rights come from different sources, each having some advantages and disadvantages. Our main property rights index is the rating of protection of property rights from the Index of Economic Freedom constructed by the Heritage Foundation. This is a relatively broad index of property rights, is available for a large set of countries and has been used by other researchers (for example, Johnson et al. 1998, and La Porta et al. 1999, 2002). A second index of property rights rates the protection of intellectual property rights in particular by using data on the "Special 301" placements of the Office of the US Trade Representative (USTR). "Special 301" requires the USTR to identify those countries that deny adequate and effective protection for intellectual property rights or deny fair and equitable market access for persons that rely on intellectual property protection. Countries can be placed on different lists depending on their relative protection of intellectual property. For example, countries which have the most onerous or egregious acts, policies or practices and which have the greatest adverse impact on relevant US products are designated "Priority Foreign Countries". As such, the index weights the degree of property rights protection with the economic impact that protection deficiencies have on US trade. We use these qualifications to construct an index of intellectual property rights protection. The third index is the patent rights index constructed by Ginarte and Park (1997). This index focuses more specifically on the protection of patents. A fourth index is the property rights index of the World Economic Forum, which measures the general legal protection of private property in a country. The fifth index is the intellectual property rights index of the World Economic Forum, which measures the protection of intellectual property in a country. The two World Economic Forum indexes are only available for the year 2001. The sixth index is the property rights index constructed by Knack and Keefer (1995) using data from the International Country Risk Guide (ICRG). This index measures property rights in a broad sense and includes five measures: quality of the bureaucracy, corruption in government, rule of law, expropriation risk and repudiation of contracts by the government. Table 1 presents more details on these six indexes of property protection.

Our main index of protection of property rights covers the 1995-99 period; the Special 301 index of protection of intellectual property rights covers the 1990-99 period; the World Economic Forum indexes refer to 2001; and the Knack and Keefer index covers the 1982-95 period. The growth regressions, however, include data for the period 1980-89, as in RZ. Ideally, one would want to use property rights indexes for the period 1980-89 as well; however, this is not possible for the property rights indexes available to us due to data limitations. The one exception is the Ginarte and Park patent rights index, for which we do have data for the period 1980-89. Therefore, this index does not suffer from the non-overlapping time period problem and we can use the patent rights index for the year 1980, the beginning of the period 1980-89, in the regressions. For the other indexes, we use index values as of their first available date.

Although the indexes of property protection are from different sources and for different time periods, they appear quite related and are highly positively correlated (Annex Table 1). The correlation between our main property rights index and the other five indexes of protection of (intellectual) property rights ranges, for example, from 0.49 to 0.78. The fact that the property rights indexes relate to different time periods could nevertheless raise concerns in our specification, in part because property rights may have evolved in response to economic performance. We believe these concerns to be small, most importantly, because measures of institutional frameworks have been found to be stable over long periods of time (Acemoglu, Johnson and Robinson 2001, 2002). Also, RZ show that the sample means of the accounting standards variable they use do not differ significantly between 1983 and 1990.

This stability also applies to our property rights indexes, which do not change much over the time for which they are available. Table 2 shows that the mean property rights index for countries sampled in the first and last available year is not statistically significantly different for any of the three indexes. Importantly, the sample mean of the Ginarte and Park patents rights index – the only index for which we have data for the period 1980-89 – for countries sampled in 1980 does not differ statistically significantly from the sample mean in 1990 for the same set of countries. In addition, we find that the relative ordering of the different property rights indexes does not change much over time, as the Spearman rank order correlations of the respective indexes are high. A t-test of independence further confirms that the property rights indexes in the first and last available year are not independent. As a further robustness check, we also do our regressions instrumenting the property rights indexes with variables that predate the

period 1980-89, using the methodology used by Beck, Levine and Loayza (2000) and Acemoglu, Johnson and Robinson (2001).

Table 3 presents the summary statistics of the country-specific variables grouped by developing and developed countries (Annex Table 2 presents the same summary statistics, but by individual country). We only use the developing versus developed countries classification to illustrate the differences in the various variables by institutional settings. The country summary statistics show that, as a group, developing countries have less developed financial systems, weaker law and order systems, worse protection of (intellectual) property rights, and fewer patents per capita. All variables except for the stock market capitalization-to-GDP ratio and the accounting standards show a statistically significant difference between the two groups of countries. Other work has documented extensively the differences in the degree of law and order between developed and developing countries. This difference in legal frameworks partly relates to the difference in the credit-to-GDP ratio between these two groups of countries, where low contract enforcement environments have hindered the development of financial systems in developing countries.

The degree of financial development and the protection of property rights tend to go together and are both related to the overall level of development of the country. As such, analyzing the differential effects of financial development and property rights on the level of external financing available and the allocation of investment across different assets could be difficult; however, the correlation between the two concepts is not perfect. That is, there exist countries with good property rights and underdeveloped financial systems. Chile, for example, scores high on the protection of property rights (property rights index equals 5) but its level of financial development is only average

(private credit-to-GDP is 36 percent). France, on the other hand, has a well-developed financial system (reflected by a level of private credit-to-GDP of 54 percent) but the protection of its property rights is only average (with a property rights index of 4). Calculating the simple correlation between the property rights index and the level of financial development, 0.59, confirms that the relationship between the two concepts is high but not perfect. The correlations of the interaction variables are even less perfect, less than 0.20 (see further Annex Table 3).

Our dataset includes 45 countries.⁶ For the growth regressions, as in RZ, we need to drop the benchmark country, the United States, and we are therefore left with 44 countries. As we collected additional data, the number of countries included in our dataset somewhat exceeds that in RZ, who use data on 41 countries. For robustness, we also estimated the model using the subset of countries in RZ and results did not change.

Similarly to RZ, we construct benchmark data on an industry basis. We use the benchmark data from RZ for all of our industry variables, but add the intangible-to-fixed-assets ratio. We assume that the intangible-to-fixed-assets ratio for each industry in the U.S. forms a good benchmark (similar to RZ who use the US external financial dependence ratio as a benchmark). We refer to the ratio of intangible-to-fixed assets as the intangible intensity. In the same way RZ calculates the external financial dependence ratios by industry, we calculate the intangible intensity benchmark using Compustat-data on US firms for the years 1980-89. We measure intangibles by the net value of intangible assets, i.e., Compustat item 33. Generally, intangibles are assets that have no physical

⁶ The countries include Australia, Austria, Bangladesh, Belgium, Brazil, Canada, Chile, Colombia, Costa Rica, Denmark, Egypt, Finland, France, Germany, Greece, India, Indonesia, Israel, Italy, Jamaica, Japan, Jordan, Kenya, Korea, Malaysia, Mexico, Morocco, Netherlands, New Zealand, Nigeria, Norway, Pakistan, Peru, Philippines, Portugal, Singapore, South Africa, Spain, Sri Lanka, Sweden, Turkey, United Kingdom, United States, Venezuela, and Zimbabwe.

existence in themselves but represent rights to enjoy some privilege. In Compustat, this item includes: blueprints or building designs, patents, copyrights, trademarks, franchises, organizational costs, client lists, computer software patent costs, licenses, and goodwill (except on unconsolidated subsidiaries). Intangibles in the Compustat-data excludes goodwill on unconsolidated subsidiaries, which are included in investments and advances under the equity method (Compustat item 31). We measure tangibles by net fixed assets, i.e., Compustat item 8. This represents net property, plant and equipment, which equals gross property, plant and equipment (Compustat item 7) less accumulated depreciation, depletion and amortization (Compustat item 196).

Table 4 reports the intangible intensity benchmarks for US firms in different industrial sectors on a two-digit SIC level. The total number of firms used to calculate these benchmarks is 5,241. The average intangible intensity ratio during the 1980s for US manufacturing firms is 77 percent. The variation of the intangible intensity across industries is large: it ranges from as low as 2.0 percent for the petroleum and coal products industry to as high as 454 percent for the printing and publishing industry. The variation concurs with notions of what constitute relatively capital-intensive versus more knowledge-intensive industries. The stone, clay, glass and concrete products industry, for example, relies mainly on fixed assets for production, as would be expected as the technology used in this sector is well-established and embodied in the fixed assets. It has an intangible intensity ratio of 5 percent. The chemical and allied products industry and the electrical and electronic industry, in contrast, rely heavily on intangible assets, such as patents and licenses, as inputs. It has an intangible intensity ratio of 77 percent. The data show that the various technical and economic reasons that make various types of products require different input mixes can be benchmarked well at the industry level.

III. Empirical Results

In this section, the results of the regression model of section I are presented. The dependent variable is the growth in real value added in a particular sector in a particular country over the 1980-89 period. The basic regression specification is OLS and the results are presented in Table 5. We find that industrial sectors that rely relatively more on external finance develop disproportionately faster in countries with better developed financial markets because the coefficient for the interactive variable credit-to-GDP times external financial dependence is statistically significant (at the one percent level, column 1). Hence, consistent with the findings of RZ we find that financial development facilitates economic growth through greater availability of external financing. As noted by Beck, Levine and Loayza (2000) and others, the quality of the legal system influences financial sector development and overall growth. Interacting the external financial dependence variable with the index of the quality of the legal framework used by La Porta et al. (1998), instead of the financial development variable, also leads to a positive coefficient (not reported). The regression result confirms the law and finance view that increased availability of external financing and better legal systems enhance firm growth.

In terms of the asset allocation effect, we find that industrial sectors using relatively more intangible assets develop faster in countries with better protection of property rights because the coefficient for the interactive variable property rights times intangible intensity is statistically significant and positive (column 2). Hence, better property rights

facilitate economic growth as they favor growth through better asset allocation, i.e., in firms that would naturally choose a higher share of investment in intangible assets.⁷

The asset allocation effect on growth appears to be in addition to the increase in firm growth due to greater external financing since, in the regressions where both the external financial dependence and the intangible intensity variables are included (column 3), both interactive variables are statistically significant. Additionally, the coefficients in these regressions for both effects are of similar magnitudes as in the two regressions where each of them was included separately (columns 1-2), suggesting that the two variables measure complementary effects.⁸

The effects of external financial development and property protection on firm growth are not only both statistically significant but are also equally economically important. We can use the regression coefficient estimates of Table 5 to infer how much higher the growth rate of an industry at the 75th percentile of intangible-intensity would be compared to an industry at the 25th percentile level when the industries are located in a country at the 75th percentile of property protection rather than in a country at the 25th percentile. The industry at the 75th percentile, Instruments and related products, has an intangible intensity ratio of 0.90. The industry at the 25th percentile, Textile mill products, has an intangible intensity ratio of 0.21. The country at the 75th percentile of property protection has a value of 5 for the property rights index and the country at the 25th

⁷ Exclusion of sectors with relatively high estimated usage of intangible assets, such as Printing and publishing and/or Miscellaneous manufacturing industries, does not qualitatively alter the results (not reported).

⁸ The two interacted variables, external financial dependence and intangible intensity interacted with financial development and property rights indexes, do appear to measure different concepts as the correlation between these variables is low. The correlation between the external financial dependence variable interacted with the financial development measure and the intangible intensity measure interacted with the property rights index is 0.149. Similar correlations are found when the other four property rights indexes are used (see Annex Table 3).

percentile a value of 3. The estimated coefficient for the interaction term in model 2 of Table 5 equals 0.103 and we set the industry's initial share of manufacturing at its overall mean. The regression coefficient estimates therefore predict the difference in growth rates between the 75th and 25th percentile intangible intensive industry to be 1.4 percent per year higher in a country with a property rights index of 5 compared to an index of 3. For comparison, the average growth rate is 3.4 percent per year. Therefore, a differential rate of 1.4 percent due to an improvement in the property rights index from 3 to 5 represents a large increase.

The effect of financial development on differential real firm growth can be calculated in a similar way using the estimated coefficient for the interaction term of model 1 in Table 5 of 0.140. The coefficient estimate predicts the difference between the growth rate of the 75th and 25th percentile external financial dependence industry to be 1.4 percent higher in a country at the 75th percentile of financial development compared to at the 25th percentile. Thus, the effects of property protection and financial development on differential firm growth are not only both statistically significant, but also of similar economic importance. In other words, the asset allocation effect is economically as important as the finance effect.

Thus far, our specifications have focused on the differential effect on growth of property rights across industries with different asset mixes (captured by the interaction term of property rights and the intangible-intensity measure). To avoid possible biases caused by any omitted country-specific regressors, we have included country dummies to

⁹ RZ used the same approach to compute the effect of financial development on differential real firm growth. Our estimated effect differs somewhat from the differential growth rate effect estimated in RZ, 1.3 percent, because our sample is slightly larger and because we use private sector credit instead of total capitalization as our measure of financial development.

capture any institutional or other differences, such as comparative advantage or general level of development, affecting growth. Since we are less interested in the importance of general country differences, we use this approach rather than a vector of specific country control variables. Still, the use of country dummies could introduce a misspecification to the extent that any omitted institutional differences important for growth are correlated with our two interaction variables. Examples of such country-specific variables that have been used in the general growth literature include, besides financial depth and property rights, the level of per capita GDP, human capital, and other institutional variables (Romer 1990, Barro 1991, and Levine and Zervos 1998, among others). Furthermore, we want to analyze the first-order country effects of property rights to investigate whether property rights affect firm growth mainly through the asset allocation channel or as well in any other ways. We therefore replace our country dummies with country-specific institutional and other variables and thus perform a robustness check on whether any of our earlier results are affected if we control in other ways for country differences.

We start by documenting the fact that the effects of better property rights on growth work mostly through improved asset allocation as opposed through, for example, an improvement in the overall business environment increasing growth opportunities. We illustrate this by including in our basic regression specification the property rights index (and private credit-to-GDP) directly in addition to the interacted variable. The results are reported in column 4 of Table 5, where we exclude country dummies. We do not find a direct, statistically significant effect of the quality of a country's property rights on industrial sector growth. Most importantly, including the property rights index directly does not change the magnitude or the significance of the coefficients for the interaction variables in any meaningful way. Both the financial dependence and the asset mix

interaction variables remain statistically significant and neither change much in terms of magnitude. This suggests that the major effect of improved property rights on sectoral growth operates through improvements in asset allocation and that the interaction variable does not capture any general effects of, for example, improvements in the business environment leading to greater growth opportunities.

For other country-specific variables, we use the ratio of private credit to GDP in 1980, stock market capitalization over GDP in 1980, a measure of the level of human capital in 1980, a measure of the quality of the legal system, an accounting standards indicator, and the logarithm of per-capita income in 1980. RZ and Cetorelli and Gambera (2001) have also used these variables in the same model. We expect a positive effect on growth of private credit to GDP and market capitalization to GDP as proxies for the development of the banking system and stock market respectively, and financial development more generally. The level of human capital is measured as the average of the years of schooling attained by the population over 25 years of age in 1980 (as in Barro and Lee 1993) and is expected to have a positive effect on growth in value added. The quality of the legal system is measured by the law and order tradition variable of La Porta et al. (1998) and is also expected to have a positive effect on growth. The accounting standards indicator is an index reflecting the quality of accounting and is taken from RZ. This variable is also expected to have a positive effect on growth since it proxies for the quality of information investors have regarding firm and firms regarding investment prospects. Per capita GDP is included to capture the convergence effects of the economy as a whole to a long-run steady state and is expected to have a negative coefficient (see, among others, Barro 1991). The model continues to include industry dummies to control for any sector-specific effects and the property rights indexes. Since

the country variables included in the two interaction terms – private credit to GDP and an index of property rights – are now also part of the country controls, we can assess both the overall effect of financial development and property protection on value added growth as well as the finance and asset allocation effects captured by the two interaction terms. Note that data on accounting standards is missing for some countries, reducing the sample of countries to 33.

The results of this specification are reported in column 5 of Table 5. Except for the human capital variable, the country controls have the expected relationships with growth. The direct effect of the quality of property rights on growth remains insignificant, however, which suggests that better property rights of themselves do not translate into higher growth rates of sectoral value added. The depth of the financial systems – measured by private credit to GDP and the size of the stock market as a ratio to GDP – has a positive and statistically significant influence on growth in sectoral value added. The degree of human capital in the country, proxied by the average years of schooling attained by the population over 25 years of age, and the degree to which the rule of law applies do not have a statistically significant effect on growth in sectoral value added. The accounting index, however, is statistically significantly positive. The general level of development, proxied by the log of income per capita, has a negative sign, confirming the convergence effect.

The focus of our attention, the interaction between property rights and the allocation of resources, is very robust to these changes in model specification. The coefficient on the interaction term between the property rights indexes and the intangible-intensity measure remains positive and statistically significant in both specifications. The size of the coefficient is also only somewhat less from those in the models with country

dummies, and the coefficient remains statistically significant at the one percent level. The general result on the importance of the asset allocation effect is thus not altered. Also, the interaction term between financial development and external financial dependence remains statistically significant positive. The regression results in columns 4 and 5 thus show that the effect of property rights on growth operates in an important way through asset allocation, and not through a first order effect on growth.

Another concern is that the quality of property rights is affected by the investment behavior of firms and the resulting growth patterns. At the macro-level, countries that grow faster may demand greater property rights protection as a larger share of economic output derives from more property rights' intensive investments. At the more micro-level, sectors that are more dependent on property rights may seek a higher degree of protection of property rights relevant to their industry. Due to these and other concerns about potential endogeneity, we instrument the property rights variable with a number of predetermined institutional variables. Following RZ, we use the colonial origin of a country's legal system (indicating whether the legal origin is English, French, German, or Scandinavian) as reported in La Porta et al. (1998) as one instrument. As also shown by La Porta et al. 1998, legal origin tends to have a long lasting effect on a country's institutional structure, whereas the legal origin of a country is largely determined by the country colonizing it. As such, legal origin is a good instrumental variable and has been used in several other papers. Following Acemoglu, Johnson and Robinson (2001), we also use the settler mortality rate of European bishops, soldiers and sailors stationed in colonies in the 17th, 18th and 19th century as an instrument. As argued by Acemoglu et al. (2001), the willingness of colonizing powers to settle and develop long-lasting institutions depended greatly on the ability of colonizers to survive physically. They

show that the settler mortality rate is a good instrumental variable for past institutional characteristics which last into today (in their application, the particular institutional characteristic is the risk of expropriation of private property). We do not have settler mortality rates for European countries, that is the colonizing countries themselves. Since the European countries had the institutions which they were exporting to their colonies, we set their mortality rates to zero to account for this.

The instrumental variables (IV) results based on the specification of column 2 are presented in columns 6-8, using respectively only legal origin, mortality rates, or both as instruments. The results are very robust to the use of instruments. We again find a statistically significant effect of property rights on growth in sectoral value added through the asset allocation of resources. Interestingly, the magnitude of the coefficients for the interaction variable increases when using mortality rates as an instrumental variable (columns 7 and 8). Next, we re-estimate the regression model in column 3 using instrumental variables. Following Beck et al. (2000) we use legal origin as an instrument for financial development today (as measured by private credit-to GDP). In line with Acemoglu et al. (2001) we use settler mortality as instrument for property rights today. The results are presented in column (9) and confirm our previous findings, i.e., that the protection of private property has a statistically significant effect on growth in sectoral value added through the asset allocation of resources.¹¹

As an additional investigation into the channels through which financial development and property rights affect firm growth, and following RZ, we analyze

¹⁰ The first-stage regressions show a strong relationship between the instrumented variables and the potentially endogenous variables, i.e., between settler mortality and legal origin and property rights and financial development, private credit-to GDP, today (not reported).

¹¹ The results presented in Table 5 are based on all available data (up to 44 countries). For robustness, we re-estimated the regression models using the subset of 41 countries used in RZ, which implies excluding Indonesia, Jamaica, and Nigeria. The results are very similar to those in Table 5 (not reported).

whether industries in countries with better financial development and property rights grow faster because new establishments are added to the industry or because existing establishments grow faster. There are two reasons why it is interesting to decompose the effects of access to financing and asset allocation in terms of number and average size of firms. First, highlighted by RZ, the creation of new establishments is more likely to require external funds, while the expansion of existing establishments may more easily rely on internal funds. Thus, the effect of financial development could be more pronounced for new firms than for the growth of existing firms. Second, new firms are often set up in reaction to and to take advantage of new technological developments, while established firms tend to grow through expansion of scale, perhaps also because they are slower in reacting to new developments.¹² Furthermore, existing firms may be able to preserve the value of their assets in ways other than resorting to formal property rights (for example, using their name recognition, distribution or supply networks, or general economic and political influence). Thus, the importance of property rights that protect the returns to (new) technology and help assure a good allocation of an economy's overall resources might be more pronounced for the emergence of new firms than for the growth of existing firms.

As before, we follow RZ and use data derived from the UN Industrial Statistics Yearbook database for the growth in the number of establishments and the growth in the average size of existing establishments. The growth in the number of establishments is calculated by RZ as the logarithm of the number of end-of-period establishments less the logarithm of the number of beginning-of-period establishments. The average size of

¹² In fact, many new firms that take advantage of new technological developments are spun off from existing firms that have developed some elements of these new technologies.

establishments in the industry is calculated by dividing the value added in the industry by the number of establishments, with the growth in average size again defined as the difference in logarithms. RZ report that in their sample of countries roughly two-thirds of the growth in value-added results from an increase in the average size of existing establishments, while the remaining one-third is accounted for by an increase in the number of establishments.

We use the same specification as for our basic regression but with the growth in number of establishments or the growth in average size as the dependent variable instead of the growth in total value added by sector. As Table 6 indicates, the external financial dependence interacted with the financial development variable is statistically significant in explaining both the growth in the number of establishments and the growth in average firm size. This contrasts with RZ who do not find any statistical significance, perhaps because they use accounting standards as a measure for financial development (see their Table 7) rather than private credit to GDP and do not include the asset allocation interaction variable.

Interestingly, the asset allocation variable interacted with the property rights variable is only significant when explaining the growth in the number of establishments and not when explaining the growth in the average size of firms. This finding is consistent across all of our measures of property rights (not reported). It is also not affected by using as instrumental variable either legal origin or settler mortality (columns 3-8). It suggests, in terms of affecting growth through asset allocation, that the protection of property rights is most important through stimulating the growth of new establishments. Well-protected property rights can thus indirectly influence growth by allowing new firms to come to market in those industries that typically rely less on

tangibles in their optimal production mix. For established firms relying more on intangible inputs, growth seems less affected by the strength of property rights in the country. This may be because such firms have other means of protecting their returns from investments.

IV. Further robustness tests

We have already shown that the results are robust to different control variables, the use of instrumental variables and changes in the sample of countries. We next present evidence that the results are also robust to the particular measure of protection of property rights chosen, to alternative means of controlling for country-differences, and to differences in growth opportunities related to the level of general development.

First, we use the five alternative measures of the degree to which countries protect property rights: "Special 301", the patent rights index of Ginarte and Park (1997), the property rights index and the intellectual property rights index of the World Economic Forum and the property rights index of Knack and Keefer (1995). The results are presented in Table 7 and are very similar to those of Table 5. Both with and without the interaction term between financial development and external dependence, we find statistically significant coefficients on the interaction term between the intangible-intensity measure and all of the five alternative property rights measures. The results with the alternative measures of the degree of property rights protection are also robust to the use of legal origin and European settler mortality as instruments (not reported).

Second, we want to investigate whether growth opportunities differ across industries and countries in such a way that they confound the relationships between our

interaction variables and growth in sectoral value added. In particular, it is possible that the external financial dependence and asset mix variables are proxies for growth opportunities at the sectoral level. Provided that financial development is high and property rights are protected, it may not be those industries with a particular external financial dependence or intangibles intensity that grow fast, but rather those with better growth opportunities. If these growth opportunities happen to be correlated with our financial development and property rights variables, then a bias in the estimations can arise. In particular, countries with similar levels of financial development or property rights may experience the same growth patterns across industries because their firms face similar patterns of growth prospects, not because their levels of financial sector development or quality of property rights protection imply a greater supply of resources for firms or a better allocation of resources by firms. Correspondingly, countries with different levels of financial development or property rights may have different growth opportunities and consequently grow differently, not because of differences in the supply of external financing or the protection of property rights.

In a recent paper, Fisman and Love (2002b) explore this hypothesis using the RZ-model, focusing on financial development. They use the actual US sales growth at the sectoral level as a measure for sectoral growth opportunities at a global level. When they substitute in the interaction term the industry's actual sales growth for the industry's external financial dependence ratio, they find a positive coefficient for this interaction variable. Furthermore, when including both the old and new interaction variable, i.e., the industries' external financial dependence times countries' financial development as well as the actual sales growth times countries' financial development, they find that the external financial dependence variable is no longer statistically significant. This suggests,

if indeed actual US sales growth rates are a good proxy for (global) growth opportunities, that it is the similarity (or difference) in growth opportunities for countries at similar (or different) levels of financial development which leads to the positive relationship between growth and the interaction variable external financial dependence times countries' financial sector development.

A similar possibility may arise with respect to the asset allocation hypothesis and our asset mix variable. If growth opportunities systematically vary across countries with the degree of property rights protection, then a statistically significant coefficient for our interaction variable could be inaccurately interpreted as support for the asset allocation hypothesis. To investigate this possibility, we use the same approach as Fisman and Love. Specifically, we interact both the external financial development and property rights variables with the US sectoral sales growth rates and include these two new interaction variables in the regressions. The columns 2-4 in Table 8 show the results of adding the interacted US sales growth variable in this way to the model, with column 1 repeating the results of column 3 of Table 5. Column 2 confirms the result of Fisman and Love, that is, the interaction term between financial development and US sales growth "dominates" the interaction term between financial development and external financial dependence in terms of sectoral growth, as the coefficient on the interaction term financial development and external financial dependence is no longer statistically significant. In Column 3 we add the interaction variable between property rights and US sales growth. Although this new interaction variable is also statistically significant, our main result – a positive relationship between sectoral growth and the interaction variable property rights and asset mix - is robust to this change in specification, although the statistical significance for our main result decreases somewhat. When we add both new

interaction variables, i.e., between US sales growth and financial development and between US sales growth and property rights, to the model (column 4), our main result still holds, but the RZ and Fisman and Love variables are no longer statistically significant. This suggests that the asset allocation effect remains an important explanation of firm growth.

The measure of growth opportunities used in Fisman and Love, i.e., the actual sales growth at the sectoral level, is an ex-post measure. It is therefore highly correlated with actual growth in value added, our dependent variable, and as such may not be the best measure to use for growth opportunities and could explain the reduced significance of the interaction variables in columns 3 and 4. As an alternative, more forward-looking proxy for growth opportunities, we use Tobin's Q ratio, i.e., the ratio of the market value of the firm to the book value of its assets. We use Compustat data to construct the industry-level median of the time-average Tobin's Q of US firms during the period 1980-89. The results of using this alternative measure of growth opportunities in the interaction variables are presented in columns 5-7 of Table 8. In contrast to the actual sales growth measure, we find that the interaction variables with Tobin's Q do not enter significantly in any of the regressions, showing that results are dependent on the proxy used for growth opportunities. Our main result is strengthened, however, as the coefficients for the interaction variable property rights and asset mix become more statistically significant. This suggests that growth opportunities, as measured by firms' Tobin's Q, do not vary across countries in such a systematic way with the degree of property rights protection as to affect the relationship between property rights and actual growth, through asset allocation.

As a third robustness test, we investigate whether using US sectoral data biases our results in some way. It could be the case, for example, that investment opportunities in poorer countries are different from those in the U.S. due to differences in the general level of a country's development rather than differences in property rights. For a poor country with the same property rights as a rich country, for example, the asset mix across sectors variable may not relate in the same way to relative growth rates as growth opportunities differ because of its general lower level of development. Any relationship between growth and our interaction variable property rights times asset mix may then be spurious as it reflects differences in growth opportunities, and not the asset allocation effect. We test for this possibility by adding an interaction variable between the US sectoral asset mix and countries' per capita GDP to the regression. We use the level of per capita GDP as a measure of the overall level of a country's economic development and of corresponding country-level investment opportunities. The same robustness test was performed by RZ, but then interacting external dependence with per capita GDP. If investment opportunities relate systematically to a country's level of development and affect the ability of sectors with different asset mix to grow, rather than a country's property rights affecting growth through the asset mix chosen, then this new interaction variable should be significant and our old interaction variable no longer. Controlling for differences in the level of development in this way does not alter our main result as the new interaction variable is not statistically significant while our old interaction variable still is (column 8 in Table 8). Thus, variations in property rights across countries leading to different growth patterns do not seem to be due to simple differences in investment opportunities related to level of development, but rather to differences in the asset mix chosen in response to variations in property rights.

As an alternative robustness test along the same lines, we test whether, for countries with the same level of property rights, investment opportunities differ in a systematic way with income levels such as to confound the relationship between assets mix and growth. If investment opportunities across sectors do not vary in a systematic way with income level, then for the same level of property rights we should not find an effect across countries of the income level variable interacted with the asset mix variable. Columns 9-11 in Table 8 show the results of regressions for three subsamples of countries with each the same degree of protection of property rights (as measured by our main property rights index), but different levels of per capita GDP. Using this specification, we do not find an income level effect since the coefficients for the variable that is the interaction between per capita GDP and asset mix are insignificant in each of the three cases.

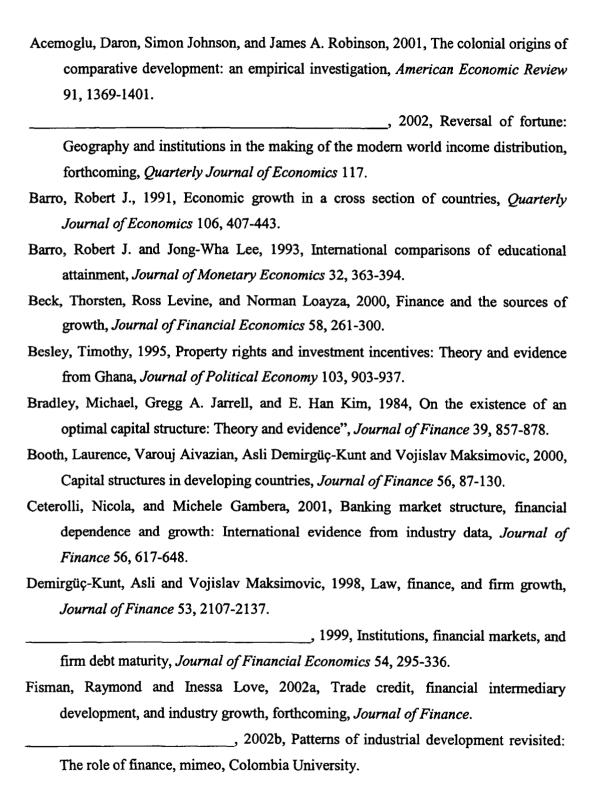
V. Conclusions

Countries differ from each other in many ways. Two aspects are the degree of their financial sector development and the quality of their property rights. This paper argues that the existence of an environment with poorly developed financial systems and weak property rights has two effects on firms: first, it reduces the access of firms to external financing; and, second, it leads firms to allocate resources in a suboptimal way. The importance of the lack of financing effect has already been shown in the law and finance literature. We investigate the importance of property rights for firm growth by studying its impact on firms' allocation of investable resources. We show that the effect of insecure property rights on the asset mix of firms, the asset allocation effect, is

economically as important as the lack of financing effect as it impedes the growth of firms to the same quantitative magnitude. Furthermore, the asset allocation effect seems to be particularly important in hindering the growth of new firms.

While we use the ratio of tangibles and intangible assets as a measure of asset mix, the implications of our results likely go beyond this particular asset choice and indicate that an efficient allocation of firm resources can be more generally impeded by weak property rights. Our results suggest that the degree to which firms allocate resources in an optimal way will depend on the strength of a country's property rights and that the allocation effect is an important channel of the effect of property rights on firm growth. Thus, our results have the important policy implication that, equally important as the establishment of a good financial system, requiring in turn a functioning legal system, is assuring the protection of returns to different type of assets. To the extent that the emergence of the "new economy" has increased the economic returns to assets on which yields are more difficult to secure, then our results would even underestimate the overall costs of weak property rights. If indeed new economy assets and future growth opportunities are more related to intangible assets, then any underallocation of investable resources towards intangible assets is likely to impede the future growth of firms and economies more generally even more so going forward.

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Table 1 The Variables

This table describes the variables collected for our study. The first column gives the names of the variable as we use it.

The second column describes the variable and provides the source from which it was collected.

Property (Freedom) A rating of property rights in each country (on a scale from 1 to 5). The more protection private property receives, the higher the score. The score is based, broadly, on the degree of legal protection of private property, the probability that the government will expropriate private property is the score. The score is based, broadly, on the degree of legal protection to private property. The indicates can be received the original order of the index. Intellectual Property (Special 30) Solid	The second column d	escribes the variable and provides the source from which it was collected.
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higher the score. I indicates that assets are poorly delineated and not protected by Jaw. while 7 indicates that assets are clearly delineated and protected by Jaw. Source: Global Competitiveness Report, World Economic Forum. An index of intellectual property protection is each country (on a scale from 1 to 7). The more protection intellectual property protection is even to intellectual property protection is even to the bureaucracy, corruption in government, rule of Jaw. expropriation risk and repudation of contracts by the government. Private credit Market cap Accounting Private credit Market cap Accounting Accounting Human capital Rule of Law Assessment of the law and order tradition in the country. Average of the months of April and October of the monthly index between 1982 and 1995. Scale from 0 to 10, with higher scores indicating more disclosure. Source: Center for International Financial Analysis and Research. Taken from Rz. Human capital Rule of Law Assessment of the law and order tradition in the country. Average of the months of April and October of the monthly index between 1982 and 1995. Scale from 0 to 10, with lower scores for less tradition for law and order. Source: Barro and Lec (1993). Legal origin Legal origin Legal origin Legal origin European settler mortality GDP per capital Growth in value added External financial of an added External financial State of the Law and order tradition in the country. Average of the months of April and October of the monthly index between 1982 and 1995. Scale from 0 to 10, with lower scores for less tradition for law and order. Source: International Country Risk Guide. Taken from La (1997). Legal origin European settler mortality and the score in value added by SDP in the company Law or Commercial Code of each country. There are four possible origins: (1) English Common law; (2) French Commercial Code; (3) German Commercial Code; (4) Scandinavian Commercial Code. Source: La Forta et al. (1999). European settler mortality in value added by	Patent rights (GP)	An index of patent rights (on a scale from 0 to 5) in 1980. The more protection patents receive, the higher the score. The
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Table 2 Stability of property rights measures over time

This table reports for each of the three property rights indexes the sample mean and standard deviation for the first year and the last year of the sample period across all sampled countries, the t-statistic for a test of difference in the sample means assuming unequal variances, the rank order correlation coefficient and a test of independence of the property rights indexes in the first year and the last year of the sample period. The null hypothesis of the test of independence is that the property rights indexes are independent. The sources and definitions of the data are reported in Table 1. Significance levels ^a, ^b and ^c correspond to one percent, five percent and ten percent respectively.

		Statis	stics across	countries:	Test of difference in means	Rank order correlation	Test of independence
Property rights index	Year	Mean	Standard deviation	Number of observations	t-statistic	Spearman's rho	•
Property (Freedom)	1995	3.93	0.96	44			
Property (Freedom)	2000	3.89	0.97	44	-0.22	0.90	0.000a
Intellectual property (301)	1990	4.29	0.60	28			
Intellectual property (301)	2000	4.03	0.81	28	-1.36	0.76	0.000a
Patents (GP)	1980	2.69	0.91	44			
Patents (GP)	1990	2.74	1.00	44	0.29	0.97	0.000a

Table 3 Summary statistics of institutional variables

This table reports summary statistics of the variables used in our study. For each variable, we report the mean across all sampled countries, across developing countries and across developed countries. To classify countries as developing or developed, we use the World Bank classification of countries. For comparison purposes, we also present t-statistics of tests of differences in the means of the variables across developing and across developed countries. The sources and definitions of the data are reported in Table 1. Significance levels ^a, ^b and ^c correspond to one percent, five percent and ten percent respectively.

	Ме	ans across count	ries:	t-Tests of difference in means
	Developed countries	Developing countries	All countries	Developed versus Developing countries (t-statistics)
Property (Freedom)	4.68	3.42	3.96	7.10 ^a
Intellectual property (301)	4.47	3.74	4.12	3.97ª
Patents (GP)	3.33	2.20	2.67	5.44ª
Property (WEF)	6.11	4.69	5.33	7.66°
Intellectual property (WEF)	5.74	3.47	. 4.51	10.64ª
Property (ICRG)	9.14	5.42	7.03	11.82ª
Private credit-to-GDP	0.49	0.26	0.36	4.37 ^a
Market cap-to-GDP	0.24	0.17	0.20	0.64
Law and order	9.23	4.40	6.67	11.74ª
Accounting standards	0.65	0.66	0.65	-0.12
Settler mortality rate	0.55	4.36	2.48	-11.19ª
Human capital	7.92	4.07	5.84	5.72°
GDP per capita	9.04	6.84	7.79	10.28 ^a
Number of countries	19	25	44	

Table 4 Benchmark US Intangible-to-Fixed assets ratio

The table reports intangible-to-net fixed assets ratios for each sector are averages for all US firms in the Compustat f(US) database for the period 1980-89. For external financial dependency benchmarks across sectors we refer to the original source: Table 1 in Rajan and Zingales (1998). The table also reports the number of US firms used to construct the benchmark for each industrial sector. As in Rajan and Zingales (1998) we focus on manufacturing firms and use 1980s data to construct the benchmarks. The total number of firms is 5,241.

SIC Code	Industrial sectors	Intangibles-to-fixed assets	Number o	f firms
20	Food and kindred products		0.75	304
21	Tobacco manufactures		0.49	2
22	Textile mill products		0.21	13
23	Apparel and other textile products		0.53	139
24	Lumber and wood products		1.20	9.
25	Furniture and fixtures		0.49	81
26	Paper and allied products		0.20	130
27	Printing and publishing		4.54	202
28	Chemicals and allied products		0.96	550
29	Petroleum and coal products		0.02	86
30	Rubber and miscellaneous plastics		0.46	19
31	Leather and leather products		0.33	4
32	Stone, clay, glass, and concrete products		0.05	90
33	Primary metal industries		0.11	19
34	Fabricated metal products		0.31	27
35	Industrial machinery and equipment		0.25	7 9:
36	Electrical and electronic equipment		0.77	81:
37	Transportation equipment		0.24	26
38	Instruments and related products		0.90	66
39	Miscellaneous manufacturing industries		2.29	16
	Mean		0.76	
	Median		0.48	
	Standard deviation		1.03	

Table 5 Growth, financial dependence, property rights and intangible assets

Dependent variable is the real growth in value added of a particular sector in a particular country. Table 1 describes all variables in detail. As measure for protection of property rights we use the property rights index from the Index of Economic Freedom from the Heritage Foundation. All regressions include industry dummies and a constant but these are not reported. Models (1)-(3) and models (6)-(8) include country dummies but these are not reported. Models (4) and (5) include country-specific variables rather than country dummies. Model 6 uses legal origin as IV for property rights. Model 7 uses European settler mortality as IV for property rights. Model 8 uses legal origin and European settler mortality as IV for property rights. Robust standard errors are shown below the coefficients. United States is dropped as it is the benchmark. Significance levels a, b and c correspond to one percent, five percent and ten percent respectively.

	(1)	(2)	(3)	(4)	(5)	(6) IV legal origin	(7) IV mortality	(8) IV legal origin and mortality	(9) IV legal origin and mortality
Fraction of sector in value added of manufacturing in 1980	-1.041 ^a (.2454)	9721 ^a (.2482)	-1. 076 * (.2491)	-1.040° (.2210)	4511* (.1028)	9672* (.2480)	9720 * (.2762)	9646 ^a (.2743)	-1.036 ^a (.2765)
Private credit-to-GDP * External financial dependence (US)	.1401 ^a (.0383)	(.1354° (.0376)	.1376 ^a (.0380)	.0509 ^a (.0204)	, ,	` '	, ,	.1023 ^a (.0379)
Property *Intangible-to-fixed assets (US)	` ,	.0103 ^a (.0029)	.0092ª (.0028)	.0091 ^a (.0033)	.0067 ^a (.0024)	.0090ª (.0033)	.0182ª (.0054)	.0166 ^a (.0049)	.0162° (.0051)
Private credit-to-GDP		,	- ,	0213 (.0163)	.0488ª (.0151)				
Property				0004 (.0050)	.0030 (.0058)				
Stock market capitalization-to-GDP					.0253° (.0068)				
Human Capital (Schooling)					0008 (.0017)				
Rule of law					.0019				
Accounting standards					.0428 ^b (.0180)				
Log of per-capita GDP					0205 ^a (.0043)				
R ²	.2711	.2548	.2757	.1028	.2386	.2547	.2188	.2201	.2277
N Number of countries	1242 44	1277 44	1242 44	1242 44	830 33	1277 44	1071 37	1071 37	1040 37

Table 6 Growth in average size and number of establishments

The dependent variable is either the growth in average size or the growth in the number of establishments of a particular sector in a particular country. Table 1 describes all variables in detail. All regressions include industry dummies, country dummies and a constant but these are not reported. Models 3 and 4 use legal origin as IV for property rights. Models 5 and 6 use European settler mortality as IV for property rights. Models 7 and 8 use legal origin as IV for private credit-to-GDP and European settler mortality as IV for property rights. Robust standard errors are shown below the coefficients. United States is dropped as it is the benchmark. For Costa Rica, France, Indonesia, Italy, Jamaica, Netherlands, South Africa, and Zimbabwe we do not have data on the growth of the average size and the number of establishments. Significance levels a, b and c correspond to one percent, five percent and ten percent respectively.

	(1) Growth average size	(2) Growth number	(3) Growth average size IV legal origin	(4) Growth number IV legal origin	(5) Growth average size IV mortality	(6) Growth number IV mortality	(7) Growth average size IV legal origin and mortality	(8) Growth number IV legal origin and mortality
Fraction of sector in value added of manufacturing in 1980	8687ª	3399 ^b	8396*	3038°	-1.020ª	2466	-1.039ª	2610
	(.3131)	(.1702)	(.3143)	(.1624)	(.3304)	(.1582)	(.3278)	(.1652)
Private credit-to-GDP * External financial dependence (US)	.0856"	.0480 ^b					.0969*	.0287
	(.0289)	(.0220)					(.0322)	(.0290)
Property *Intangible-to-fixed assets (US)	.0001	.0069ª	0007	.0082 ^b	.0057	.0127*	.0030	.0127*
	(.0021)	(.0022)	(.0036)	(.0034)	(.0040)	(.0050)	(.0034)	(.0049)
\mathbb{R}^2	.4329	.3656	.4164	.3619	.4208	.3805	.4580	.3815
N	1071	1104	1100	1133	912	943	887	918
Number of countries	36	36	36	36	30	30	30	30

Table 7 Growth, financial dependence, property rights and intangible assets – Alternative measures of property rights

The dependent variable in all regressions is the real growth in value added of a particular sector in a particular country. Table 1 describes all variables in detail. We use five alternative measures for protection of property rights. First, we use a measure for protection of intellectual property rights which is calculated using the "Special 301" placements of Office of the US Trade Representative. We use the median rating during 1990-1999. Second, we use the patent rights index by Ginarte and Park (1997). We use the rating for the year 1980. A higher rating of the patent rights index indicates more protection of patent rights. Third, we use the property rights index of the World Economic Forum. We use the rating for the year 2001. Fourth, we use the intellectual property rights index of Knack and Keefer (1995). Average over 1982-95. All regressions include industry dummies, country dummies and a constant but these are not reported. Robust standard errors are shown below the coefficients. United States is dropped as it is the benchmark. Significance levels ^a, ^b and ^c correspond to one percent, five percent and ten percent respectively.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
Fraction of sector in value added of manufacturing in 1980	5225ª	9592*	-1.053°	-1.055°	9802ª	5708ª	-1.064ª	-1.139ª	-1.141ª	-1.082ª
	(.1561)	(.2449)	(.2655)	(.2659)	(.2493)	(.1625)	(.2458)	(.2652)	(.2656)	(.2503)
Private credit-to-GDP * External financial dependence (US)						.0740*	.1357*	.1355ª	.1360°	.1353
						(.0252)	(.0382)	(.0389)	(.0390)	(.0376)
Intellectual property (301) *Intangible-to-fixed assets (US)	.0062*					.0052 ^b				
	(.0023)					(.0021)				
Patents (GP) *Intangible-to-fixed assets (US)		.0074*					.0066*			
D. CHIPPINAT . III . G. L. CHIP		(.0026)					(.0026)			
Property (WEF) *Intangible-to-fixed assets (US)			.0109ª					.0093*		
Latella de al constant (MEEN de la constant CMEEN d			(.0029)					(.0027)		
Intellectual property (WEF) *Intangible-to-fixed assets (US)				.0072*					.0062*	
December (ICDC) *Interweible to Errod content (IC)				(.0019)	00.428				(.0018)	002#8
Property (ICRG) *Intangible-to-fixed assets (US)					.0043*					.0037*
					(.0012)					(.0012)
R ²	.3269	.2521	.2581	.2575	.2548	.3592	.2734	.2789	.2786	.2755
N	1119	1277	1211	1211	1277	1090	1242	1179	1179	1242
Number of countries	36	44	42	42	44	36	44	41	41	44

Table 8 Growth, financial dependence, property rights and intangible assets – Robustness tests

The dependent variable in all regressions is the real growth in value added of a particular sector in a particular country. Table 1 describes all variables in detail. All regressions include industry dummies, country dummies and a constant but these are not reported. Robust standard errors are shown below the coefficients. Models (9)-(11) include only those observations for which Property rights index takes value 3, resp. 4, resp. 5. United States is dropped as it is the benchmark. Significance levels ^a, ^b and ^c correspond to one percent, five percent and ten percent respectively.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9) PROP=3	(10) PROP=4	(11) PROP=5
Fraction of sector in value added of manufacturing in 1980	-1.076°	-1.071ª	-1.074°	-1.072°	-1.068ª	-1.064°	-1.066°	-1.077ª	-1.466ª	9445ª	2194°
	(.2491)	(.2496)	(.2471)	(.2478)	(.2510)	(.2522)	(.2528)	(.2503)	(.2255)	(.3819)	(.1178)
Private credit-to-GDP * External financial dependence (US)	.1354°	.0649	.0896*	.0617	.1176ª	.1124*	.1183°	.1353ª			
	(.0376)	(.0458)	(.0338)	(.0457)	(.0364)	(.0324)	(.0364)	(.0376)			
Private credit-to-GDP * Sales growth (US)		1.170°		.5671							
		(.6806)		(.5426)							
Private credit-to-GDP * Tobin's Q (US)					.0318		0136				
					(.0430)		(.0363)				
Property *Intangible-to-fixed assets (US)	.0092*	.0075*	.0048°	.0046°	.0088°	.0071*	.0071*	.0086b			
	(.0028)	(.0025)	(.0026)	(.0026)	(.0028)	(.0028)	(.0028)	(.0038)			
Property *Sales growth (US)			.3377 ^b	.2915°							
			(.1731)	(.1612)							
Property * Tobin's Q (US)						.0185	.0198				
						(.0129)	(.0133)				
Per capita GDP 1980 *Intangible-to-fixed assets (US)								.0005	0049	.0027	.0056
								(.0022)	(.0046)	(.0023)	(.0045)
R ²	.2757	.2793	.2832	.2839	.2761	.2783	.2784	.2757	.3030	.3781	.4546
N	1242	1242	1242	1242	1242	1242	1242	1242	387	381	471
Number of countries	44	44	44	44	44	44_	44	44	14	13	15

Annex Table 1 Correlation matrix of property rights indexes

p-values are in parentheses. Significance levels a, b and c correspond to one percent, five percent and ten percent respectively.

	Property (Freedom)	Intellectual property (301)	Patents (GP)	Property rights (WEF)	Intellectual property (WEF)	Property (ICRG)
Property	1.000ª					
(Freedom)	(0.00)					
Intellectual property	0.4870a	1.000°				
(301)	(0.00)	(0.00)				
Patents	0.6260ª	0.4437ª	1.000a			
(GP)	(0.00)	(0.00)	(0.00)			
Property	0.7668ª	0.5394a	0.6852a	1.000 ^a		
(WEF)	(0.00)	(0.00)	(0.00)	(0.00)		
Intellectual property	0.7814ª	0.6506a	0.7170 ^a	0.9487 ^a	1.000°	
(WEF)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	
Property	0.7821a	0.5622ª	0.6868ª	0.8690°	0.9308 ^a	1.000a
(ICRG)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)

Annex Table 2 Country-specific data

This table reports several variables for the countries studied. Countries are sorted in ascending alphabetical order. n.a. is not available. More detail on the definitions and sources of the variables can be found in Table 1.

Country	Developing	Property	Intellectual property (301)	Patents (GP)	Property (WEF)	Intellectual property (WEF)	Property (ICRG)	Private credit	Market cap	Accounting	Human capital	Rule of law	Legal origin	European Settler Mortality	GDP per capita
Australia	0	5.00	4.00	3.23	6.20	6.00	9.30	0.28	0.38	0.70	10.08	10.00	1.00	2.15	9.20
Austria	0	5.00	5.00	3.81	6.40	6.20	9.45	0.77	0.03	0.48	6.22	10.00	3.00	0.00	9.10
Bangladesh	1	2.00	n.a.	1.99	3.70	2.20	2.85	0.07	0.00	n.a.	1.68	n.a.	1.00	4.27	4.79
Belgium	0	5.00	5.00	3.38	5.90	5.50	9.58	0.29	0.09	0.63	8.79	10.00	2.00	0.00	9.33
Brazil	1	3.00	3.00	1.85	5.00	4.10	6.64	0.23	0.05	0.69	2.98	6.32	2.00	4.26	7.41
Canada	0	5.00	4.00	2.76	6.20	5.80	9.73	0.45	0.46	0.68	10.16	10.00	1.00	2.78	9.26
Chile	1	5.00	4.00	2.41	5.60	4.20	6.44	0.36	0.34	0.60	5.99	7.02	2.00	4.23	7.84
Colombia	1	3.00	4.00	1.12	4.30	3.00	5.54	0.14	0.05	0.39	4.23	2.08	2.00	4.26	7.05
Costa Rica	1	3.00	n.a.	1.94	5.20	3.70	6.47	0.26	0.04	n.a.	4.81	n.a.	2.00	4.36	7.68
Denmark	0	5.00	5.00	3.62	6.40	6.30	9.80	0.42	0.09	0.62	10.14	10.00	4.00	0.00	9.41
Egypt	1	3.00	3.00	1.99	5.60	4.10	4.96	0.21	0.01	n.a.	2.16	4.17	2.00	4.22	6.33
Finland	0	5.00	5.00	2.95	6.50	6.40	9.76	0.48	0.06	0.71	9.61	10.00	4.00	0.00	9.23
France	0	4.00	5.00	3.90	6.40	6.60	9.37	0.54	0.10	0.76	5.97	8.98	2.00	0.00	9.34
Germany	0	5.00	5.00	3.86	6.50	6.30	9.55	0.78	0.09	0.68	8.46	9.23	3.00	0.00	9.42
Greece	0	4.00	3.00	2.46	5.00	3.90	6.56	0.44	0.08	0.44	6.56	6.18	2.00	0.00	8.25
India	1	3.00	3.00	1.62	4.90	3.00	5.80	0.24	0.05	0.71	2.72	4.17	1.00	3.88	5.48
Indonesia	1	3.00	4.00	0.33	3.80	2.90	4.38	0.20	0.00	n.a.	3.09	3.98	2.00	5.14	6.21
Israel	1	4.00	4.00	3.57	6.30	4.90	7.22	0.67	0.35	n.a.	9.14	4.82	1.00	n.a.	8.18
Italy	0	4.00	4.00	3.71	6.20	5.70	8.07	0.42	0.07	0.69	5.83	8.33	2.00	0.00	8.77
Jamaica	1	4.00	n.a.	2.86	4.90	3.50	5.05	0.15	0.02	n.a.	3.60	n.a.	1.00	4.87	7.11
Japan	0	5.00	4.00	3.94	6.10	5.50	9.34	0.86	0.30	0.67	8.17	8.98	3.00	n.a.	9.20
Jordan	1	4.00	4.50	1.86	5.80	4.60	5.15	0.54	0.50	n.a.	2.93	4.35	2.00	n.a.	7.01
Kenya	1	3.00	n.a.	2.57	n.a.	n.a.	5.58	0.20	0.00	n.a.	2.44	5.42	1.00	4.98	6.03
Korea, Rep. of	f 1	5.00	3.00	3.28	4.70	4.00	6.90	0.50	0.08	n.a.	6.85	5.35	3.00	n.a.	7.25
Malaysia	1	4.00	4.00	2.57	5.20	3.50	7.09	0.48	0.65	0.78	4.49	6.78	1.00	2.87	7.43
Mexico	1	3.00	4.00	1.40	4.60	3.60	5.76	0.16	0.07	n.a.	3.51	5.35	2.00	4.26	7.88
Morocco	1	3.50	n.a.	2.38	n.a.	n.a.	5.05	0.16	0.02	n.a.	n.a.	n.a.	1.00	4.36	6.69
Netherlands	0	5.00	5.00	4.24	6.50	6.50	9.87	0.60	0.19	0.73	8.20	10.00	2.00	0.00	9.32
New Zealand	0	5.00	4.00	3.32	5.90	5.30	9.80	0.19	0.33	0.61	12.14	10.00	1.00	2.15	8.92
Nigeria	1	3.00	n.a.	3.05	3.80	2.50	3.85	0.12	n.a.	0.62	n.a.	2.73	1.00	7.60	6.81

N		£ 00	5.00	2.20	£ 00	£ 20	0.00	0.24	0.00	0.71	10.33	10.00	4.00	0.00	0.61
Norway	0	5.00	5.00	3.29	5.90	5.30	9.69	0.34	0.06	0.71	10.32	10.00	4.00	0.00	9.51
Pakistan	1	4.00	4.00	1.99	n.a.	n.a.	4.21	0.25	0.03	0.69	1.74	3.03	1.00	3.61	5.67
Peru	1	3.00	4.00	1.02	4.10	3.00	4.19	0.11	0.06	n.a.	5.44	2.50	2.00	4.26	6.74
Philippines	1	4.00	4.00	2.67	4.30	2.90	3.62	0.28	0.10	0.63	6.00	2.73	2.00	n.a.	6.59
Portugal	0	4.00	5.00	1.98	5.30	4.90	7.94	0.52	0.01	0.52	3.23	8.68	2.00	0.00	7.74
Singapore	0	5.00	4 00	2.57	6.50	5.60	8.69	0.57	1.62	0.73	3.69	8.57	1.00	2.87	8.45
South Africa	1	3.00	4.00	3.57	5.30	4.50	7.50	0.26	1.20	0.81	4.61	4.42	1.00	2.74	7.97
Spain	0	4.00	4.00	3.29	5. 9 0	5.30	7.99	0.76	0.09	0.42	5.15	7.80	2.00	0.00	8.53
Sri Lanka	1	3.00	n.a.	2.79	4.20	3.10	4.64	0.21	0.06	n.a.	5.18	1.90	1.00	4.25	5.53
Sweden	0	4.00	4.00	3.47	5.90	5.80	9.80	0.42	0.11	0.81	9.47	10.00	4.00	0.00	9.57
Turkey	1	4.00	3.00	1.80	4.20	3.10	5.76	0.14	0.01	n.a.	2.62	5.18	2.00	n.a.	6.99
UK	0	5.00	5.00	3.57	6.30	6.10	9.40	0.25	0.38	0.80	8.35	8.57	1.00	0.00	9.17
Venezuela	1	3.00	4.00	1.35	3.80	3.00	5.82	0.30	0.05	n.a.	4.93	6.37	2.00	4.36	8.29
Zimbabwe	1	3.00	n.a.	2.90	3.90	2.90	5.09	0.30	0.45	n.a.	2.40	3.68	1.00	n.a.	6.09
Average	0.57	3.96	4.12	2.67	5.33	4.51	7.03	0.36	0.20	0.65	5.84	6.67	1.91	2.48	7.79

Annex Table 3 Summary statistics and correlation matrix of main explanatory variables

Summary statistics

Variable	Mean	Standard deviation	Country-industry Observations
Private Credit *	0.1136	0.1767	1,504
External Dependency			·
Property (Freedom)*	2.5328	3.3133	1,547
Intangible-to-fixed assets			•
Intellectual property (301)*	2.6138	3.3641	1,289
Intangible-to-fixed assets			
Patents (GP)*	1.7085	2.3510	1,547
Intangible-to-fixed assets			
Property (WEF)*	3.3983	4.3887	1,454
Intangible-to-fixed assets			
Intellectual property (WEF)*	2.8697	3.8636	1,454
Intangible-to-fixed assets			
Property (ICRG)*	4.4914	6.0588	1,547
Intangible-to-fixed assets			_

Correlation matrix of main explanatory variables

	Private Credit* External Dependency	Property (Freedom)* Intangible-to- fixed assets	Intellectual property (301)* Intangible-to-fixed assets	Patent rights (GP)* Intangible-to- fixed assets	Property (WEF)* Intangible-to- fixed assets	Intellectual property (WEF)* Intangible-to-fixed assets	Property (ICRG)* Intangible-to- fixed assets
Private Credit*	1.000ª						
External Dependency	(0.00)						
Property	0.1494ª	1.000 ^a					
(Freedom)* Intangible-to- fixed assets	(0.00)	(0.00)					
Intellectual	0.1336a	0.9746ª	1.000a				
property (301)* Intangible-to- fixed assets	(0.00)	(0.00)	(0.00)				
Patent rights	0.1620a	0.9514a	0.9264a	1.000ª			
(GP)* Intangible-to- fixed assets	(0.00)	(0.00)	(0.00)	(0.00)			
Property (WEF)*	0.1483ª	0.9846ª	0.9819ª	0.9539ª	1.000 ^a		
Intangible-to- fixed assets	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)		
Intellectual	0.1644ª	0.9761a	0.9714ª	0.9589ª	0.9866ª	1.000ª	
property (WEF)* Intangible-to- fixed assets	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	
Property (ICRG)* Intangible-to- fixed assets	0.1611 ^a (0.00)	0.9743 ^a (0.00)	0.9638 ^a (0.00)	0.9547 ^a (0.00)	0.9799 ^a (0.00)	0.9914 ^a (0.00)	1.000 ^a (0.00)

p-values are in parentheses. Significance levels a, b and c correspond to one percent, five percent and ten percent respectively.

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