

MPRA

Munich Personal RePEc Archive

Measuring the recreational use value of migratory shorebirds on the Delaware Bay

Myers, Kelley H.; Parsons, George R. and Edwards, Peter E.T.

University of Delaware, National Oceanic and Atmospheric Administration

2010

Online at <http://mpra.ub.uni-muenchen.de/26126/>
MPRA Paper No. 26126, posted 09. January 2012 / 20:32

Measuring the Recreational Use Value of Migratory Shorebirds on the Delaware Bay

KELLEY H. MYERS

GEORGE R. PARSONS

University of Delaware

PETER E.T. EDWARDS

National Oceanic and Atmospheric Administration

Abstract *In this article we estimate the recreational use value of household trips to view shorebirds during the annual horseshoe crab/shorebird migration on the Delaware Bay. We use contingent valuation to estimate the value of day and overnight trips separately and use a discrete choice question followed by a payment-card question to generate our valuation data. Our best estimates for the value of a day trip are about \$66–\$90/household and for an overnight trip about \$200–\$425/household (2008\$). Our data are from the 2008 season, and our average household size is 1.66. For some context, estimates from four other studies report values that vary from \$63/trip/person to \$442/trip/person. These studies vary in method and specific birding populations studied and mix day and overnight trips.*

Key words Contingent valuation, discrete choice, bird watching, use value.

JEL Classification Codes Q5.

Introduction

According to the U.S. Fish and Wildlife Service, nearly 23 million residents aged 16 years or older traveled a mile or more from their home to observe or photograph wildlife in 2006 (Aiken 2009). The growth of wildlife-based recreation in the USA and the mounting concern over the preservation of our natural resources has increased the demand for information on the costs and benefits of activities related to wildlife. Individuals who engage in wildlife watching have measurable welfare resulting from their experience, and government action can play an active role in affecting this experience. To put recreation experiences in monetary terms, non-market valuation techniques, like the travel cost or contingent valuation methods, are used to estimate the value of wildlife watching to an individual or household.

Kelley H. Myers is a PhD candidate and George R. Parsons is a professor, College of Earth, Ocean, and Environment, University of Delaware, Newark, DE 19716 USA (email: kelleymyers@gmail.com; gparsons@UDel.edu). Peter E.T. Edwards is a natural resource economist, National Oceanic and Atmospheric Administration, 1315 East West Highway, SSMC3 Rm. 14709, Silver Spring, MD 20910 USA (email: Peter.Edwards@noaa.gov).

We thank Kevin Kalasz, state wildlife biologist with the Department of Natural Resources and Environmental Control (DNREC), for his contribution to our research. We also thank Dawn Webb of DNREC for providing us with a base location during our onsite surveying and Sally O’Byrne and Derek Stoner of the Delmarva Ornithological Society for introducing us to bird watching and assisting us with the development of our survey mechanism. Finally, we thank several graduate students in the Marine Policy program at the University of Delaware who assisted us during the onsite survey: Meredith Blades Lilley, John Lilley, Andy Krueger, and Kate Semmens.

In this article, we extend the non-consumptive wildlife valuation literature using a contingent valuation study of recreational bird watching on the Delaware Bay. Each spring, the Delaware Bay is considered the largest staging area for migratory shorebirds in eastern North America. The visiting shorebirds time their arrival with the spawning season of adult horseshoe crabs in early May. The birds stay for two to three weeks eating horseshoe crab eggs for up to 14 hours a day before flying north. This combination of migratory shorebirds and spawning horseshoe crabs is unseen elsewhere in the world, making the Delaware Bay a popular destination for wildlife enthusiasts and recreational birders.

Four species of migratory birds draw the most attention during the horseshoe crab/shorebird occurrence—red knot, ruddy turnstones, sanderlings, and semipalmated sandpipers. Over the last decade, scientists have documented a decline in both the horseshoe crab and shorebird populations. This may be attributed, in part, to the increased commercial harvest of the horseshoe crab in the mid 1990s. The red knot, in particular, has drawn attention because some scientists believe it is on a path to extinction. Some say it could be lost within 5 to 10 years (NJDEP 2007).

In our application, we focus on recreational use values and use survey data from 370 recreational birders to estimate the willingness to pay (WTP) for a household trip during the 2008 annual shorebird migration. We intercepted birders onsite and used a mail-back paper survey. We used a referendum-style discrete choice question followed by a payment card to generate our preference data. In the following section, we briefly discuss some other applications of contingent valuation to bird watching. Then we present our study design, model, and results.

Values for Bird Watching Trips from other Contingent Valuation Studies

There are a number of contingent valuation studies of bird watching and general wildlife viewing trips in the literature. A sampling is shown in table 1—four birdwatching studies and three wildlife viewing studies. There are also several travel cost studies of bird watching (e.g., Gürlik and Rehber 2007) and stated preference studies of the total value of protecting bird and other wildlife populations (Richardson and Loomis 2008; Loomis and White 1996). None of these are included in the table. For the purpose of this article, we will briefly review the studies that use contingent valuation to estimate the use value of bird watching trips. These applications are closest to ours.

Hvenegaard, Butler, and Krystofiak (1989) estimated the value of bird watching in Point Pelee National Park in Canada during the spring migration in 1987. They conducted random onsite personal interviews of over 600 birders during the peak spring birding season. Using an open-ended response format, respondents were asked to consider the highest increase in trip cost they would be willing to pay before canceling their most recent trip to Point Pelee. The average WTP, which mixed day and overnight trips, was \$442/trip/person.

Cooper and Loomis (1991) estimated the value of a bird watching trip with data from a mail survey of 3,000 randomly selected California households in 1987. The authors used a dichotomous choice referendum format that asked respondents to consider a hypothetical increase in the cost of their most recent trip. Bid values ranged from \$2 to \$200 and WTP was estimated under three different conditions: current, 50% increase in birds viewed, and 100% increase in birds viewed. The mean WTP for a trip, which again included day and overnight trips, under current conditions in California was about \$63/trip/household.

Eubanks, Ditton, and Stoll (1998) and Eubanks and Stoll (2000) estimated the consumer surplus of a bird watching trip during two annual migration events: the sandhill crane in Nebraska and shorebirds in New Jersey. Both studies used an open-ended question format and asked respondents to state the maximum they would be willing to pay before canceling their most recent trip. The sandhill crane study was a combination of

onsite intercepts and a mail-based survey of randomly selected bird watchers from seven Nebraska Wildlife organizations. The estimated values were \$77/trip/person. The New Jersey study used a data set that combined intercepted birders with a mail survey of members of the Cape May Bird Observatory and the New Jersey Audubon Society. The estimated values were \$324/trip/person. Again, day and overnight trips were mixed in both analyses.

The WTP estimates from these four studies then vary from \$63 to \$442/trip. All of the studies mix day and overnight trips in their analysis. Interestingly, all of the studies except Cooper and Loomis (1991) estimate the value of a bird watching trip during a migration event. The Eubanks and Stoll study is particularly interesting because it is applied to the same horseshoe crab/migratory shorebird resource we study, except that it is on the New Jersey side of the Bay.

Table 1
Contingent Valuation Studies of the Non-Consumptive Use Value of Wildlife

Reference	Resource	CVM Method	Survey Respondents	Mean WTP Per Trip or Per Day (2008\$)
Hvenegaard, Butler, & Krystofiak 1989	Bird watching	OE	Visitors to Point Pelee National Park, Canada	\$442/trip/person
Cooper & Loomis 1991	Bird watching	DC	Mail survey of CA residents	\$63/trip/household
Eubanks, Ditton, & Stoll 1998	Bird watching	OE	Visitors to Platte River and mail survey of NE birding associations	\$77/trip/person
Boyle, Roach, & Waddington 1998	Wildlife watching	DC	U.S. households	\$28/day/person
Clayton & Mendelsohn 1993	Wildlife watching	OE, DC	Visitors to McNeil River	\$366–\$420/trip/person
Eubanks & Stoll 2000	Bird watching	OE	Visitors to DE Bay in NJ and mail survey of NJ birding associations	\$324/trip/person
Aiken 2009	Wildlife watching	OE	U.S. households	\$61/day/person

OE = open ended; DC = dichotomous choice.

Study Design

Study Site and Sampling

There are approximately 20 different locations along the coast of the Delaware Bay on the Delaware side that are suitable for viewing shorebirds during the migration. During a field trip in 2007 and based on conversations with state wildlife biologists and major birding organizations in the area, we identified two major sites that most, if not all, birders

visit on a trip to the Delaware Bay: Port Mahon and Mispillion Harbor.¹ For this reason, we confined our sampling in 2008 to these sites.² This saved time, increased our sample and, in our judgment, introduced little bias. The sites are shown in figure 1. Port Mahon is about 26 miles north of Mispillion Harbor. There are no public facilities, nature centers, or trails to speak of at Port Mahon. Visitors view birds from their car (or just outside their car) along a mile-long stretch of road that runs parallel to the shore. At Mispillion Harbor, on the other hand, there is a new Nature Center, parking area, mounted scopes for viewing birds, and facilities. It is an ideal location for viewing the migration.

The spring shorebird migration takes place during the month of May and into early June. In order to capture visiting birders, we defined our sampling timeframe from May 1 to June 15, 2008. Our goal was to distribute a total of 600 surveys onsite. Respondents were asked to complete one survey per household at the end of the day and mail it back in the postage-paid return envelope. Sampling took place over the course of 11 days (4 of which were weekend days). Due to the low visitation rates per hour and the number of surveyors onsite, we were able to approach and distribute surveys to virtually all bird watchers during sampling.³



Figure 1. Delaware Bay Birding Sites

¹ We took up residence on the Bay in 2007 for seven days and visited each of the 20 proposed sites. We interviewed and gathered preliminary trip data on approximately 75 birders on this field trip. We also interviewed approximately 12 members of the Delmarva Ornithological Society—one of the largest bird watching clubs in Delaware.

² Initially, we did a small amount of sampling at the Prime Hook National Wildlife Refuge. We spent a short amount of time there in 2007 and believed that it might serve as a third major site. In 2008, we quickly learned that most of the visitors to Prime Hook were not in the area to view the horseshoe crab/shorebird migration. We also discovered that visitors to Prime Hook who were in the area to view the migration eventually made their way to Mispillion Harbor, so we stopped sampling Prime Hook.

³ People returning the survey were eligible to participate in a drawing for a wood carving of a red knot, one the four species targeted by almost every birder making a trip to the Bay during this migration. We included a separate postcard for respondents to return with their address. About half of the sample chose to participate.

Questionnaire

We designed our survey to obtain information about household trips on the day the respondent received the survey. The questions were organized in five main parts: *i*) introductory questions pertaining to where the trip began and ended, group size and type, and number of household members on the trip; *ii*) specific questions about the birding trip, such as total hours spent birding, what sites were visited, and other general activities; *iii*) detailed questions about the trip type (day or overnight), trip expenditures, and the economic valuation question; *iv*) follow-up questions about past and future trips to the Delaware Bay during the defined season; and *v*) birding intensity, experience, and demographic questions.

In part 3 of the survey, we asked respondents to classify their trip as either a day or overnight trip to view shorebirds on the day they were intercepted. Using a discrete choice question, we then asked if they would have made the trip on the day they were intercepted if trip cost had been higher by some hypothetical amount. Our response format allowed people to reply definitely yes, probably yes, probably no, or definitely no. Immediately after the discrete choice question, we asked respondents to specify their maximum WTP using a question in a payment card format. The choice questions are shown in figures 2 and 3. The bid amount, or cost for the discrete choice question, varied across 12 different versions of the survey. For day trip respondents, the bid amounts ranged from \$10 to \$300. For overnight respondents, it ranged from \$20 to \$2,000. In the payment card follow-up question, we asked respondents to consider the largest cost increase they would tolerate before they would no longer make the trip. They were asked to circle 1 of 15 different amounts from less than \$5 to greater than \$400 if on a day trip, and from less than \$10 to greater than \$2,000 if on an overnight trip.

We used the initial discrete choice question, in part, as a set up for the payment card question. The less precise and easier-to-answer discrete choice question was intended to get people thinking about the trip cost and what they might tolerate. Including definitely and probably responses made this question easier to answer than a simple yes-no dichotomous choice. The commitment is lower. The second, more precise and more difficult to answer payment card question then asked respondents to hone in on a specific amount.

26. Suppose the cost to you to make this trip possible had been \$300 more than it actually cost. Would you still have made the trip? Treat this as the increase in cost to your entire household if you traveled with household members.

- Definitely Yes
- Probably Yes
- Probably Not
- Definitely Not

Figure 2. Example of Discrete Choice Day Trip Question

27. In your best judgment, what is the largest increase in trip cost you would have tolerated and still made this day trip? Again, if you were traveling with household members, report the largest increase your household would tolerate. (*Circle One Amount*)

< \$5	\$30	\$75	\$175	\$300
\$10	\$40	\$100	\$200	> \$400
\$20	\$50	\$150	\$250	

Figure 3. Example of Payment Card Day Trip Question

Response Rate and Data

Of the 581 distributed surveys, 376 were returned by mid-August (2008) for a response rate of 65%. Of the 376 returned surveys, 6 were discarded due to incomplete information, respondent error, or the respondent being outside of the sample frame (e.g., person under the age of 18, volunteers of Delaware Shorebird Project, group leader of an educational trip). As mentioned earlier, we asked respondents to tell us whether or not they were on a day or overnight trip to the area. If a respondent did not indicate the type of trip, we used the information about their starting and ending location on the day of the intercept to make our decision. Using this information, we segmented the sample by day and overnight trips. We have 223 day trips and 147 overnight trips in our sample. Table 2 shows the sample characteristics by trip type.⁴ The characteristics do not vary significantly by trip type with the exception of group size, which is twice as large for households on an overnight trip. The average household size is slightly below two, and the average age is in the late 50s. It is an educated population and the income is more than twice the median household income in the USA (U.S. Census Bureau 2008). Average expenditures per trip were about \$50/household for day trips and about \$400/household for overnight trips. Table 3 shows the distribution of responses by offer price to the discrete choice question, and table 4 shows the distribution for payment card responses. The lower sample sizes shown in the tables are due to item non-response on the choice questions.

For the purpose of this study, a household was birding if they were viewing shorebirds for at least a half hour on the day of the intercept. For the entire sample, the average number of days spent birding in the past 12 months was 49.4. We also find that almost 70% of the sample keeps at least a modestly detailed lifetime list of all the birds they observe. The average market value for all birding related gear and equipment is \$3,652/household. We also find that 55% of the sampled population belongs to either a local or national birding society. These findings suggest that our sample population is both an active and avid population of recreational bird watchers. In addition to birding intensity, we ask about the respondent's activities on the day of the intercept. Almost 80% of the sample population indicated that they were in the area primarily to observe the horseshoe crab/shorebird occurrence. The same fraction of the population also indicated that they observed the red knot on the day of their trip.

⁴ Since these are onsite data, households taking more trips have a higher probability of being sampled. Since our interest is in estimating mean per-trip values over all trips taken, the onsite correction often done with travel cost models is inappropriate. If we were interested in mean per-trip values over respondents (and not trips), then the correction would be needed.

Table 2
Mean Sample Characteristics by Trip Type

Category	Day Trip (n=223)		Overnight Trip (n=147)	
	Mean	Std. Dev.	Mean	Std. Dev.
Household size	1.66	0.78	1.86	0.96
Group size	1.93	3.59	3.81	6.14
Hours birding on Delaware Bay during the trip	5.11	2.98	5.54	3.13
Days birding in past 12 months	52.25	37.27	47.59	37.70
Market value of birding equipment	\$3,662	6,135	\$3,622	6,599
Age (years)	57.31	12.51	58.44	10.26
Education (years)	16.53	2.68	17.31	3.09
Household income (2008\$)	\$106,132	67,244	\$122,591	74,779
Average trip expenditures (2008\$)	\$52.46	44.82	\$401.58	331.72

Table 3
Distribution of Responses to the Discrete Choice Question

Day Trips Bid Amount (\$)	Def_YES	Prob_YES	Prob_NO	Def_NO	Total
10	11	3	1	0	15
20	8	5	1	0	14
25	10	2	3	0	15
30	8	4	2	0	14
40	7	8	4	0	19
50	4	12	4	0	20
60	3	8	5	1	17
75	1	6	7	2	16
100	3	7	6	1	17
150	4	4	5	7	20
200	2	6	8	2	18
300	1	1	7	6	15
Total	62	66	53	19	200
Overnight Trips Bid Amount (\$)	Def_YES	Prob_YES	Prob_NO	Def_NO	Total
20	4	0	0	0	4
30	6	7	0	0	13
40	3	2	0	0	5
50	8	4	0	0	12
75	2	5	0	0	7
100	5	6	1	0	12
250	3	3	2	1	9
500	1	2	3	3	9
750	0	3	5	3	11
1,000	0	0	4	0	4
1,500	0	0	4	8	12
2,000	1	0	4	7	12
Total	33	32	23	22	110

Table 4
Distribution of Responses to Payment Card Question

Day Trips		Overnight Trips	
Bid Amount (\$)	Number of Responses	Bid Amount (\$)	Number of Responses
5	9	25	3
10	8	50	10
20	20	75	4
30	40	100	33
40	9	125	1
50	39	150	5
75	21	200	25
100	27	250	4
150	3	300	2
175	5	500	8
200	5	750	4
250	1	1,000	2
300	5	1,500	0
400	3	2,000	0
Total	195		101

Models

We estimate two models using the discrete choice data. In both cases we use a standard dichotomous choice random utility maximization (RUM) model. Model 1 treats the definitely yes and probably yes responses as yes responses. Model 2 treats only the definitely yes responses as yes. Assuming some of the probably yes responses are indeed yes and some are not, this approach should give us reasonable bounds on the actual value. For day trips, 31% of the respondents answered definitely yes, and 33% answered probably yes. In the overnight trip data, 30% answered definitely yes, and 29% answered probably yes. In both models the definitely and probably no responses are treated as no responses.

Our modeling approach follows a standard dichotomous choice analysis. The method is shown in Haab and McConnell (2002, pp. 24–36). The basic binary logistic model(s) where a yes response is a success and a no response is a failure has the form:

$$\Pr(\text{yes}_j) = [1 + \exp(-(\alpha z_j - \beta t_j))]^{-1}, \quad (1)$$

where α and β are normalized parameters to be estimated, z_j is a vector of individual characteristics believed to influence choice, and t_j is the offer amount the respondent has on his or her survey. Haab and McConnell show that the expected WTP from this model has the form:

$$E(\text{WTP}_j) = \alpha z_j / \beta. \quad (2)$$

Again, as noted above, two models are estimated: one treats definitely and probably yes responses as yes responses, and the other treats only the definitely yes responses as yes responses. Definitely and probably no responses are treated as no responses in both models.

Our follow-up choice question used a payment card format. Following Cameron and Huppert (1989), we use an interval data model to estimate the probability a household's WTP lies between the lower bound bid, t_{Lj} , and the upper bound bid, t_{Uj} :

$$\Pr(WTP_j \subseteq (t_{Lj}, t_{Uj})) = \Pr((t_{Lj} - \alpha z_j) < n_j < (t_{Uj} + \alpha z_j)), \quad (3)$$

where n_j is a standard normal variable, and α is a vector of normalized parameters to be estimated. The probability in equation (3) can be rewritten as the difference between two cumulative standard normal distribution functions and the parameters estimated using standard interval data regression. Mean WTP in the grouped data model has the form:

$$E(WTP_j) = \alpha z_j. \quad (4)$$

Our results for all models are presented in the following section.

Results

Discrete Choice Models

Table 5 presents results for the two discrete choice models. Model 1 treats the definitely yes and probably yes response as yes, and Model 2 treats only the definitely yes responses as yes. Definitely no and probably no responses are treated as no in both models. The recoding of the discrete choice responses in Model 2 was designed to address the issue of hypothetical bias. In Blumenschein *et al.* (1998) for example, the contingent valuation responses were recoded in a similar fashion and provided results that mirrored an actual payment treatment. The day and overnight models are shown in table 5. In both models we introduce covariates that we believed would have an effect on WTP. The variables WS, EQUIP, and CLUB are measures of birding intensity. WS is a dummy variable that indicates the respondent made a trip to Delaware to view a rare sighting of a wood sandpiper, EQUIP is the current market value of household birding equipment, and CLUB is a dummy variable that indicates whether or not the respondent belongs to a local birding society.⁵ PM is a dummy variable for surveys handed out at Port Mahon to test for sensitivity to sampling site, NIGHTS is the number of nights a person spent in the area while on an overnight birding trip, and INCOME is household income. Table 6 provides definitions and descriptive statistics for all of these variables.

The coefficient on BID is negative and significant in all models. As expected, the size of the constant term in Model 2 is lower than in Model 1 due to fewer responses coded as yes. The coefficients on CLUB, EQUIP, and INCOME show some statistical significance across the models, but not in all cases. There appears to be no difference in WTP between those sampled at Port Mahon versus Mispillion Harbor, since the coefficient on PM shows no statistical significance. And finally, the number of NIGHTS a person spends on a trip shows no significance in either overnight trip model.

The mean welfare estimates for each specification are shown at the bottom of table 5. The mean values for Model 1 are \$134/trip/household for day trips and \$563/trip/household for overnight trips. The values are considerably lower in Model 2 at \$24 and \$70. This is not surprising given that about 30% of the sample reported probably yes to the discrete choice question in both the day and overnight trip data.

The probably yes responses may be due to a number of factors—uncertainty, ambivalence, or perhaps an unwillingness to commit (Ready, Whitehead, and Bloomquist 1995). In any case, we take the estimates from these two models as bounds over the actual val-

⁵ Prior to (but not during) the 2008 horseshoe crab/shorebird season there was a sighting of a rare bird species in the area—a wood sandpiper. Many bird watchers made trips specifically to see the wood sandpiper. We decided to include a variable WS (WS = 1 if you had viewed the sandpiper) as a measure of a person's birding intensity. It had no significance in any of the day trip models, and no one in the overnight model had viewed the bird.

ues, believing that some of the probably yes responses are indeed yes and some are no.⁶ The follow-up payment card presses respondents to make more exact statements. We will now turn to that model.

Table 5
Parameter and Welfare Estimates for the Discrete Choice Models: Binary Logit

Variable	Model 1 Definitely Yes and Probably Yes = 1		Model 2 Definitely Yes Only = 1	
	Day Trip Model	Overnight Trip Model	Day Trip Model	Overnight Trip Model
Constant	0.68 (1.67)	1.96** (1.95)	-0.45 (1.03)	-0.95 (1.34)
WS ^a	-0.15 (0.55)	-	-0.32 (.60)	-
CLUB	0.89** (2.26)	-0.72 (0.912)	0.76** (2.07)	0.61 (1.14)
EQUIP	0.009 (0.262)	0.08 (0.84)	0.08** (2.23)	-0.14 (0.37)
PM	0.36 (0.913)	-0.50 (0.524)	-0.59 (1.48)	-0.28 (0.39)
NIGHTS	-	-0.02 (0.110)	-	0.18 (1.45)
INCOME	0.08** (2.42)	0.12** (2.02)	0.04 (1.29)	0.03 (1.17)
BID	-0.015*** (5.74)	-0.005*** (4.69)	-0.018*** (4.59)	-0.003*** (3.28)
Log-likelihood	-102.39	-29.22	-98.49	-49.92
McFadden R2	0.21	0.60	.20	0.26
E(WTP) per trip per household	\$133.96	\$563.11	\$24.22	\$70.34
95% CI ^β	[105.94, 161.99]	[385.79, 741.43]	[-2.46, 50,9]	[-317.42, 458.08]

t-statistics are in parentheses.

^a No overnight respondents viewed the wood sandpiper.

** Indicates significant at the 95% level of confidence; *** Indicates significant at the 99% level of confidence.

^β Represents confidence interval estimated with Krinsky-Robb procedure with 5,000 draws (Krinsky and Robb 1986).

⁶ We assume throughout that all of the probably no responses are no. If we applied the same analysis to these responses by introducing a scenario where they are treated as yes, our upper bound would be larger.

Table 6
Descriptive Statistics by Trip Type

Variable	Definition	Mean		Standard Deviation	
		Day	Overnight	Day	Overnight
WS	1 if the respondent viewed the wood sandpiper; 0 otherwise	0.125	–	0.331	–
CLUB	1 if the respondent belongs to a birding club/society; 0 otherwise	0.535	0.627	0.500	0.485
EQUIP	Current market value of birding equipment (divided by \$1,000)	3.809	4.382	6.326	7.173
PM	1 if the respondent received the survey at Port Mahon; 0 otherwise	0.310	0.200	0.464	0.402
NIGHTS	Number of nights spent in the area (overnight model only)	–	2.56	–	1.993
INCOME	Mid-point of income range (divided by \$10,000)	10.52	12.34	6.76	7.99
BID	Offered amount from survey	90.02	581.22	81.74	683.03

Payment Card Models

Table 7 presents the results from the interval data model using the payment card data. NIGHTS and INCOME show some statistical significance across the models. We also investigated whether or not anchoring might play a role in the payment card responses by using BID as an explanatory variable. A positive and significant coefficient suggests that there indeed is some evidence of anchoring in the day trip model—the higher the offer amount in the discrete choice question, the higher the circled payment card amount. The WTP estimates are shown at the bottom of table 7. The first row does not correct for anchoring, but the second does by setting DC_BID = 0 in the simulation across the sample. This correction gives the lowest possible anchor and thus a lower bound estimate. The values are \$78.71 per trip per household for day trips and \$228.04 for overnight trips without the correction. With the correction, the values drop to \$66.01 and \$201.54.

Table 8 shows the consistency between our discrete choice and payment card responses. The third column shows the percent of respondents, by type of discrete choice response, that had inconsistent payment card responses or, put differently, reversed their discrete choice response in their payment card response. One would expect these percentages to be low for the definitely yes and definitely no groups, and they are. They are also low for the probably no responses, but the probably yes group shows a large reversal. As we said at the outset, we expected that some of these probably yes responses would indeed be true yes responses and some would be true no responses. If you believe the payment card data, it appears as though 25% of the probably yes responses are no in the day trip data and 23% are no in overnight model. Table 9 presents a dichotomous choice model, where the probably yes responses to the discrete choice question are recoded to be consistent with the payment card response. The WTP values are \$96/trip/household for day trips and \$429/trip/household for overnight trips (2008\$). Again, these fall between the results for Models 1 and 2 in previous section. The adjusted day trip value is

close to the day trip payment card value, but the overnight trip value is almost double the overnight trip payment card value. Table 10 is a summary of the WTP values across our models, and figures 4 and 5 plot the estimated models probabilistically. The figures drive home the bounding effect of Models 1 and 2 on the payment card model and the discrete choice model adjusted with the payment card data. We also estimated a single-site travel cost model for day trips and found values on the higher side of our stated preference range for day trips. Our best travel cost estimates for day trips ranged from \$94 to \$190/trip/household (Edwards, Parsons, and Myers 2010).

Table 7
Parameter and Welfare Estimates from Payment Card Model: Interval Data Model

Variable	Day Trip (n=195) Coefficient	Overnight Trip (n=101) Coefficient
Constant	28.96 (2.12)	-33.74 (-0.55)
WS	-13.19 (-0.79)	-
CLUB	20.03 (1.71)	52.42 (1.18)
EQUIP	1.28 (1.30)	1.01 (0.36)
PM	11.49 (0.93)	45.97 (0.86)
DC_BID	0.14** (2.01)	0.04 (1.48)
NIGHTS	-	33.82*** (3.08)
INCOME	1.86** (2.07)	8.39*** (3.150)
Sigma	75.71*** (18.36)	201.78*** (13.07)
Log-likelihood	-520.85	-290.25
E(WTP) per trip per household	\$78.71	\$228.04
95% CI ^β	[67.95, 89.48]	[187.77, 268.31]
E(WTP) per trip per household if DC_BID = 0 for all observations	\$66.31	\$201.54
95% CI ^β	[55.55, 77.07]	[161.27, 241.81]

t-statistics are in parentheses.

** Indicates significant at the 95% level of confidence; *** Indicates significant at the 99% level of confidence.

^β Represents confidence interval estimated with Krinsky-Robb procedure with 5,000 draws (Krinsky and Robb 1986).

Table 8
Consistency of Payment Card and Discrete Choice Responses

Response to Discrete Choice	Total Number of Responses	Number of Reversals on Payment Card Response ¹	Percent of Total
Day Trip Model			
Definitely Yes	61	3	5
Probably Yes	63	16	25
Probably No	52	1	2
Definitely No	19	0	0
Total	195	20	10
Overnight Trip Model			
Definitely Yes	27	2	7
Probably Yes	31	7	23
Probably No	22	0	0
Definitely No	21	0	0
Total	101	9	9

¹ A reversal occurs when a payment card response is inconsistent with a discrete choice response. For example, if a person replies 'probably yes' to making a trip at a \$50 increase in trip cost on the discrete choice question, and then circles \$25 in the payment card question as the maximum amount she would pay and still make the trip, there is an inconsistency or reversal of responses.

Table 9
Parameter and Welfare Estimates from Adjusted Probably Yes Models: Binary Logit

Variable	Day Trip (n=195) Coefficient	Overnight Trip (n=101) Coefficient
Constant	0.44 (1.04)	0.98 (1.19)
WS	-0.85 (-1.55)	-
CLUB	1.30*** (3.15)	0.58 (0.92)
EQUIP	0.02 (0.68)	0.02 (0.44)
PM	0.19 (0.46)	-0.51 (-0.68)
NIGHTS	- (0.02)	0.003
INCOME	0.10*** (2.89)	0.05 (1.28)
BID	-0.02*** (-6.07)	-0.004*** (-4.29)
Log-likelihood	-92.11	-37.03
McFadden R ²	0.31	0.47
E(WTP) per trip per household	\$96.47	\$428.93
95% CI ^β	[79.36, 113.58]	[267.67, 590.19]

t-statistics are in parentheses. *** Indicates significant at the 99% level of confidence.

^βRepresents confidence interval estimated with Krinsky-Robb procedure with 5,000 draws (Krinsky and Robb 1986).

Table 10
Summary of Welfare Estimates (Mean Trip/Household in 2008\$)

Discrete Choice Models	Day Trip (\$)	Overnight Trip (\$)
Model 1: Definitely yes only = yes	24.41	70.29
Model 2: Probably yes & definitely yes = yes	133.96	563.11
Adjusted using payment card data	96.47	428.92
Payment card model (corrected for anchoring)	66.31	201.54
Single site ytravel-cost model ¹	94–190	NA

¹The travel cost model is reported in Edwards, Parsons, and Myers (2010).

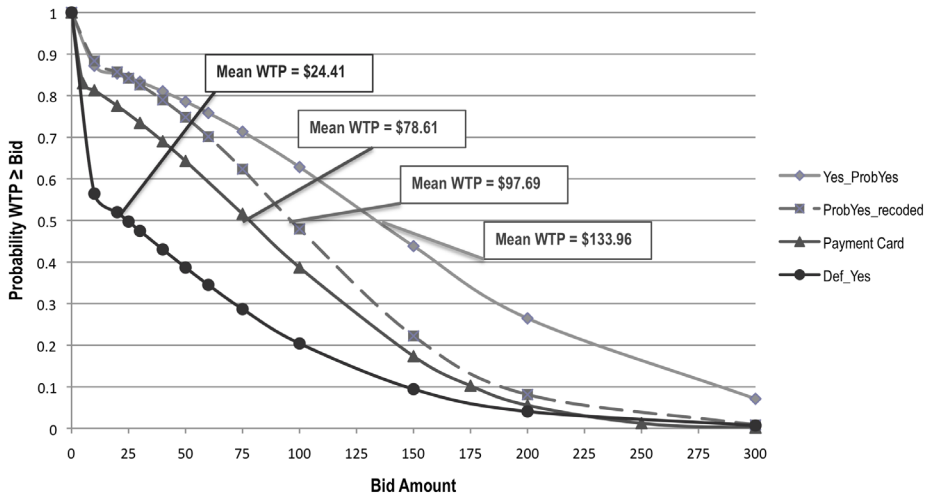


Figure 4. Probability Distribution Functions for WTP in Day Trip Models

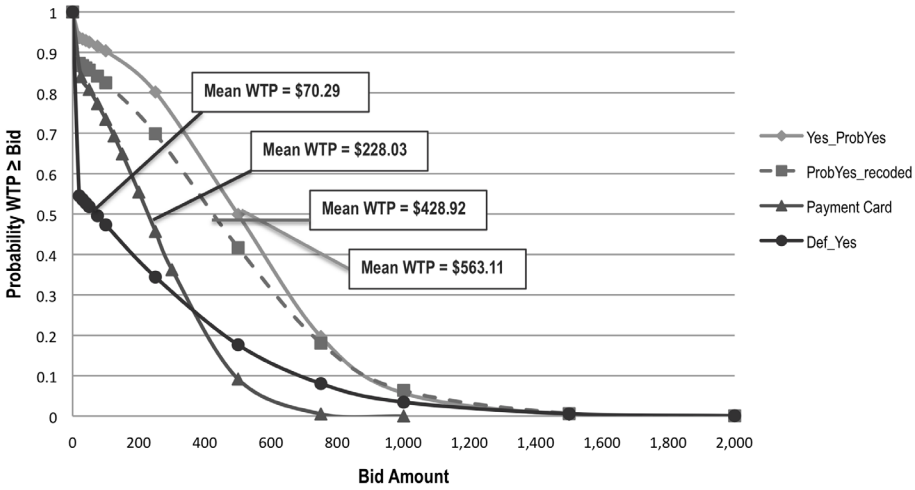


Figure 5. Probability Distribution Functions for WTP in Overnight Trip Model

Aggregate Welfare Estimate

To compute aggregate welfare estimates, we used information obtained onsite (number of surveys distributed, repeats, refusals, and hours sampled) during the sampling days to predict the total number of household visitation days over the season. We used two different approaches that produce similar results. The first assumes a constant hourly visitation rate equal to the average visitation rates we observed while sampling. We separate the rates by weekday and weekend and by the two primary sites. The weekday rates for Port Mahon and Mispillion Harbor are 1.6 and 1.9 households per hour. The weekend rates are 3.8 and 4.9. These rates are adjusted to account for people visiting both sites. Assuming 14 hours of daylight each day, 28 weekdays, and 7 weekend days at each site over the season, we estimate 3,092 household birding days to view the horseshoe crab/shorebird occurrence. This includes a day on site during a day trip or an overnight trip. Weather conditions were variable across the season, but different conditions were represented in our sample means.

In our second approach we used data gathered by the staff at the DuPont Nature Center at Mispillion Harbor. The DuPont staff keeps a daily log of the number of visitors to the center (bird watchers, family trips, stop-ins, etc.). If bird watchers are a constant share of the full visitation rate, we should be able to use the correlation between our count of visit days and their visitation rates to predict visit days to view birds on the days we did not sample. We use a simple linear regression of our household visit days ($HHDays_i$) regressed on the Nature Center count data (x_i) over the 11 days we sampled:

$$HHDays_i = \alpha + \beta x_i + \epsilon_j, \tag{5}$$

where i denotes one of our 11 sample days. The estimates are shown in table 11 for both sites. The Mispillion Harbor regression gives a sharper estimate than the Port Mahon regression. This is no surprise, since the DuPont Nature Center, where the visitation data are gathered, is located at Mispillion Harbor. For each day we did not sample, we used equation (5) to predict the number of household days at Mispillion Harbor and Port Mahon using the visit count reported by the Center for that day. Using this approach we

estimate that there were 3,363 household days to view the shorebird migration in 2008, slightly higher than our first approach that uses simple average rates.

Using the household per day trip WTP estimate from the payment card model (\$78) and assuming 3,363 household days (near the mean of the two approaches we used), we calculated the aggregate seasonal value of birding during the annual spring migration near \$263 thousand (2008\$). This assumes a day on site is valued the same if it is on a day or overnight trip. Assuming the estimated benefits accrue indefinitely over time with no growth or decline in visitation and using a discount rate of 3%, the estimated asset value of the sites for bird watching during the six weeks of the horseshoe crab/shorebird occurrence is approximately \$8.8 million (2008\$). This ignores any use value for households that visit other sites without visiting Port Mahon or Mispillion Harbor. We may overstate the value to some extent by using a somewhat longer average season (six weeks) than may actually be observed each year. We performed a sensitivity analysis over the discount rate (1, 3, 4, and 5%) and visitation growth rate (0, 4.5, and 6%) to generate a range of plausible asset values. The ranges of growth rates are based on findings in Leeworthy *et al.* (2005). The asset value ranges from as low as \$5 million to near \$95 million depending on the visitation growth rate and discount rate used. All of these values ignore nonuse and other uses of the resource areas.

Table 11
Prediction Equation for Days on Site by Site of Intercept

Coefficients	Mispillion Harbor	Port Mahon
α	19.28 (2.34)	27.77 (2.59)
β	0.36 (4.92)	0.22 (1.59)
N = days on site	11	11

t-statistics in parentheses.

Conclusion

Using an onsite intercept survey of 370 recreational birders who visited the Delaware Bay during the 2008 horseshoe crab/shorebird migration, we collected stated preference data and estimate household WTP for bird watching trips. Our best estimates are \$66–\$96/trip/household for a day trip and \$201–\$428/trip/household for an overnight trip. Using an average household size of 1.66 in our sample, our values per person are about \$40–\$60 for day trips and \$120–\$250 for overnight trips. For some context, estimates from four other studies report values that vary from \$63 to \$442/trip/person. These studies vary in method and specific birding populations studied and mix day and overnight trips. All values are in 2008\$. We also calculated the aggregate seasonal value of birding during the annual spring migration at near \$263 thousand (2008\$). Using different discount rates and growth rates for number of trips, the asset values varied from a low estimate of \$5 million to a high estimate of \$95 million. These ignore nonuse values and other values associated with using the resource area.

We used two valuation questions in our analysis—discrete choice followed by a payment card. The easier, less precise discrete choice question was used, in part, as a set up

for the more difficult, but more precise, payment card. The discrete choice response allowed people to respond definitely or probably yes and definitely or probably no. When we treated probably yes responses as certain yes responses in a dichotomous choice model, the results gave estimates greater than the payment card values. When we treated the probably yes responses as certain no responses, the results were lower than the payment card values. Assuming some of the probably yes responses were indeed yes and some no, this is the outcome one would expect. Finally, when we recoded the probably yes responses as a certain yes or certain no to be consistent with the payment card responses, the discrete choice logit models gave values close to the payment card results.

References

- Aiken, R. 2009. Net Economic Values of Wildlife-Associate Recreation: *Addendum to the 2006 National Survey of Fishing, Hunting and Wildlife-Associate Recreation*. U.S. Fish and Wildlife Service Report 2006–5. Washington, D.C.
- Blumenschein, K., G.C. Bloomquist, M. Johannassen, N. Horn, and P. Freeman. 1998. Experimental Results on Expressed Certainty and Hypothetical Bias in Contingent Valuation. *Southern Economic Journal* 118(1):114–37.
- Boyle, K.J., B. Roach, and D.B. Waddington. 1998. 1996 Net Economic Values for Bass, Trout and Walleye Fishing, Deer, Elk, and Moose Hunting and Wildlife Watching: *Addendum to the 1996 National Survey of Fishing, Hunting and Wildlife-Associated Recreation*. U.S. Fish and Wildlife Service Report 96–2. Washington, D.C.
- Cameron, T.A., and D.D. Huppert. 1989. OLS versus ML Estimation of Nonmarket Resource Values with Payment Card Interval Data. *Journal of Environmental Economics and Management* 17:230–46.
- Clayton, C., and R. Mendelsohn. 1993. The Value of Watchable Wildlife: A Case Study of McNeil River. *Journal of Environmental Management* 39:101–06.
- Cooper, J., and J. Loomis. 1991. Economic Value of Wildlife Resources in the San Joaquin Valley: Hunting and Viewing Values. *The Economics and Management of Water and Drainage in Agriculture*, A. Dinar and D. Zilberman, eds., pp. 447–62. Boston, MA: Kluwer Academic Publishers.
- Edwards, P., G. Parsons, and K. Myers. Under Review. The Economic Value of Migratory Shorebirds on the Delaware Bay: An Application of the Single Site Travel Cost Model Using Onsite Data. *Human Dimensions of Wildlife*.
- Eubanks, T., R.B. Ditton, and J. Stoll. 1998. The Economic Impact of Wildlife Watching on the Platte River in Nebraska. Report prepared by Fermata, Inc., Houston, TX.
- Eubanks, T., and J. Stoll. 2000. Wildlife-Associated Recreation on the New Jersey Delaware Bayshore. Report prepared by Fermata, Inc., Houston, TX.
- Gürlük, S., and E. Rehber. 2007. A Travel Cost Study to Estimate Recreational Value for a Bird Refuge at Lake Manyas, Turkey. *Journal of Environmental Management* 88:1350–60.
- Haab, T., and K. McConnell. 2002. *Valuing Environmental and Natural Resources: Econometrics of Non-Market Valuation*. Cheltenham, UK: Edward Elgar.
- Hvengaard, G.T., J.R. Butler, and D.K. Krystofiak. 1989. Economic Values of Bird Watching at Point Pelee National Park, Canada. *Wildlife Society Bulletin* 17:526–31.
- Krinsky, I., and A. Robb. 1986. On Approximating the Statistical Properties of Elasticities. *The Review of Economics and Statistics* 86:715–19.
- Leeworthy, R., J. Bowker, J. Hospital, and E. Stone. 2005. Projected Participation in Marine Recreation: 2005 & 2010 National Survey on Recreation and the Environment 2000. US Department of Commerce. National Oceanic and Atmospheric Administration, National Ocean Service. Special Projects. Silver Spring, MD.
- Loomis, J.B., and D.S. White. 1996. Economic Benefits of Rare and Endangered Species: Summary and Meta-analysis. *Ecological Economics* 18:197–206.

- New Jersey Department of Environmental Protection (NJDEP). 2007. Status of the Red Knot in the Western Hemisphere. Report prepared for U.S. Fish and Wildlife Service, Pleasantville, NJ.
- Ready, R., J. Whitehead, and G. Bloomquist. 1995. Contingent Valuation when Respondents are Ambivalent. *Journal of Environmental Economics and Management* 29:181–96.
- Richardson, L., and J.B. Loomis. 2008. The Total Economic Value of Threatened, Endangered and Rare Species: An Updated Meta-analysis. *Ecological Economics* 68:1535–48.
- U.S. Census Bureau, News Room: Press Releases (n.d.). 2008. Retrieved February 9 from <http://www.census.gov/PressRelease/www/releases/archives/income_wealth/012528.html>.