Performance Incentives with Award Constraints

Pascal Courty London Business School

Gerald Marschke State University of New York, Albany ^{1;2}

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Abstract: This paper studies the provision of incentives in a large U.S. training organization which is divided in about 50 independent pools of training agencies. The number and the size of the agencies within each pool vary greatly. Each pool distributes performance incentive awards to the training agencies it supervises, subject to two constraints: the awards cannot be negative and the sum of the awards cannot exceed an award budget. We characterize the optimal award function and derive simple predictions about how award prizes should depend on the number of agencies, on their sizes, and on their performances. Our results indicate that the constraints on the award distribution bind and reduce the overall $e \pm ciency$ of the incentive system. (JEL H72, J33, L14)

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¹The latest version of this paper is available at http://www.pcourty.com. Comments welcome at Pascal Courty, Department of Economics, London Business School, Regent's Park, NW1 4SA, London, UK, pcourty@london.edu; and Gerald Marschke, Department of Economics & Department of Public Administration and Policy, BA-110, SUNY-Albany, Albany, NY 12222, marschke@csc.albany.edu.

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1 Introduction

This paper studies the design of incentives in a large federal (U.S.) organization that provides job training to the economically disadvantaged. State boundaries segment the organization. Each state supervises the training agencies, or local decision makers, that are located within its boundaries. Training agencies are heterogeneous in the sense that they manage budgets of di®erent sizes. Training agencies' budgets are determined primarily by the density of the population of disadvantaged that live in their jurisdictions.

Each state distributes an award pot to provide incentives to the pool of training agencies it oversees subject to two constraints. First, the awards cannot be negative. Training agencies are guaranteed a ⁻xed budget and receive the awards on top of that budget. This constraint is similar to the limited liability constraint found in the incentive literature (Sappington, 1983). The second constraint is that the award function has to be fully-funded. By this, we mean that the sum of the rewards cannot be greater than a ⁻xed award pot. Tournaments are examples of fully-funded awards.

The main purpose of this paper is to investigate whether the limited liability and fully-funded constraints matter. From a theoretical point of view, there are good reasons to believe that they should. To see that, recall that the driving force behind performance incentives is that the way the principal stimulates $e^{\text{@ort}}$ is by creating a reward gap between high and low levels of agent performance. Under moral hazard, this implies that the agent will sometimes receive less and other times more than its contribution. Limited liability constraints, however, restrict the ability to give less to the agent than its contribution. Similarly, fully-funded constraints limit the principal's ability to give rewards that are greater than the agent's contribution. This upper-bound on the rewards together with the lower-bound on punishment due to limited liability may reduce the maximum award gap and the possibility to $e\pm$ ciently provide performance incentives.

In investigating whether these constraints matter, this paper proceeds in two steps. The ⁻rst step it to model the contractual features described above. The incentive literature has overlooked situations where fully funded and limited liability constraints interplay with the feature that agents are heterogeneous. The model asks three sets of questions. The ⁻rst set explores the relation between the agents' performances and their awards. What does the optimal incentive scheme look like? Should the awards be independent across agents as in a piece rate system or should the amount agents receive depend on the performance of other agents, as in a tournament incentive scheme? The second set of questions is speci⁻c to the feature that agents are heterogeneous. How does the optimal award function depend on the number of agents and on their relative sizes? Should awards be proportional to budget sizes? Or should smaller agents receive a disproportional fraction of the award? Third and most importantly, does the optimal contract achieve the e±cient level of e®ort? Do the limited liability and fully funded constraints bind?

The model predicts that the limited liability and the fully-funded constraints should bind and reduce the e[®]ectiveness of incentives. This should be even more pronounced in states where the agents' sizes are more heterogeneous. We show that when agents are very heterogeneous, the smaller agents will typically exert ine±ciently high level of e[®]orts. We also derive the optimal incentive contract and characterize its properties. Some of these properties suggest simple predictions on how budget sizes, award amounts and performance outcomes should vary within and across states. We also ⁻nd that the optimal award is characterized by group incentives. An agent's payo[®] is dependent on the performance of her peers even though their performances may not be statistically related. The reason for the optimality of group incentives here comes from the need to cross-subsidize awards in order to increase the award gap between high and low levels of performance.

The second step is to test if the predictions suggested by the optimal contract hold in the federal job training organization that is our case study. Our empirical strategy is to compare performance awards and performance outcomes across states that manage di®erent pools of agents. The empirical analysis uncovers three ⁻ndings. First, those agents that are small relative to their state average receive disproportionally larger awards. We also ⁻nd some mixed evidence that they perform better. Second, performance awards depend on absolute performance outcomes but also on performance outcomes relative to other agents in the state. Third, we ⁻nd some evidence that performance outcomes are lower in states that are more heterogeneous. The evidence is broadly consistent with the predictions of the model. It suggests that it is more di±cult to provide performance incentives in states that are more heterogeneous because the fully-funded and limited liability constraints are more costly in those states.

The theoretical part of this paper contributes to the contract literature. Following the early work of Lazear and Rosen (1981) on tournaments as a means to provide incentives, some authors have recently studied the speci⁻c problem of allocating ⁻xed award pots among contestants (e.g. Krishna and Morgan (1998) and Moldovanu and Sela (1999)) but these work do not assume limited liability on the part of the agent. As mentioned above, Sappington as well as Demski et al. (1988) study the restriction imposed by limited liability constraints but in a framework where the agent receives some private information after contracting. More recently, Innes (1990), and Kim (1997) considered the contractual restrictions imposed under limited liability but in a single agent framework and without the fully funded constraint.

On the empirical side, this work belongs to the empirical literature on the provision of incentive in organizations. See Prendergast (1999) for a recent survey of that literature. Another way to interpret our results is as a test of whether government bureaucrats write contracts that are consistent with the optimal incentive contracts predicted by incentive theory. There is some evidence that ⁻rms design optimal incentive contracts (Prendergast reviews studies of bonus, relative performance, and tournament) but to our knowledge, no one has yet asked whether government organizations also do so.

The paper is organized as follows. The next Section summarizes the key characteristics of the incentive system we study in the empirical application. This will be the starting point to motivate the model which is presented in Section 3. Section 4 derives some implications but the proofs are deferred to the Appendix. Section 5 tests some of the model's implication in a large training organization and Section 6 concludes.

2 The JTPA Incentive System

The Job Training Partnership Act (JTPA) of 1982 created what was until the late 90's the largest federal employment and training program serving the disadvantaged.¹ The core of our empirical work focuses on ⁻scal years 1985 and 1986. In these years, the JTPA annual budget was approximately \$4 billion and it was serving nearly one million people. JTPA is highly decentralized: job training is carried out by more than 600 semi-autonomous sub-state training agencies. The JTPA bureaucracy is unusual for many reasons but one will be of special interest for this study: Instead of a rigid, comprehensive set of rules that regulate bureaucratic conduct, the JTPA organization is driven by a set of incentive systems that in^o uence outcomes.²

JTPA gave the responsibility to individual states to design and administer the local incentive systems. There are 51 incentive systems in our data set corresponding to 50 individual states and the District of Columbia. Each incentive system rewards a pool of training agencies. In ⁻scal years 1985 and 1986, we have for each incentive system (read state) data on the number and the size of the training agencies, or more simply agents, and on the agents' performances outcomes and awards.³

To motivate the model, we present some basic statistics on the number of agents per state, and on the agents' budgets, awards and performances. The number of agents varies across states. In fact, there are on average 11.9 agents per incentive systems with a standard deviation of 11.0. The average agent's size also varies considerably across incentive systems. Agents manage on average a budget of \$3,084,309 but the standard deviation in average budget across states is \$3,254,630. This variation illustrates the fact that the JTPA funds are allocated to the states by formula on the basis of the relative size of their population that is eligible for training. Those states that have larger eligible

¹For a description of JTPA see Johnston, 1987.

²For a description of the JTPA incentives see Courty and Marschke, 2000.

³The data on the agents' performance outcomes and performance standards used in this study come from the JTPA Annual Survey Report (JASR). This report is compiled annually by the Department of Labor. The award and budget were collected by SRI, International (SRI) and Berkeley Planning Associates to evaluate for the National Council for Employment Policy the e±cacy of performance standards in JTPA. See Dickenson, et al. (1988) for a description of the data. We thank Carol Romero of the National Commission for Employment Policy for making these data available to us.

population manage more and/or larger training agencies. Agents' budgets within a state can also vary tremendously. The within-state variance in budgets is lower than \$1m in some states and as large as \$10m in others. This, again, is due to the fact that each training agency receives a share of its state's budget that is proportional to its fraction of the state population that is eligible for JTPA training. Most importantly for our study, this variation in the number of agents and in their budgets is exogenous since it depends on the local density of population in need.

As an aside, note that this feature of agent heterogeneity prevails in government organizations where the sizes of the basic managerial entities are largely determined by administrative boundaries. This implies that government organizations typically supervise pools of heterogeneous agents. In fact, this is the case in education (agents are schools), health (hospitals) and many other government service organizations where some experimentation with incentives has been tried (Dixit, 1999).

Next, we describe the performance outcomes. Before presenting some numbers, it may be useful to describe the concept of performance measures and performance standards in the JTPA organization. In ⁻scal years 1985 and 1986, there were seven performance measures and the U.S. Department of Labor (DOL) required that the States use them all. There were four measures for the adult participants, and three for the youth participants. Table 1 de⁻nes the seven performance measures.

Each state in JTPA develops an incentive system based on the DOL-de⁻ned measures to reward its pool of training agencies. The states have considerable latitude in the construction of the incentive scheme as long as awards are contingent on the achievement of numerical standards de⁻ning minimum acceptable level of performance. For non-cost measures (see Table 1), agents receive awards if their outcomes exceed the corresponding standards. For cost measures, on the other hand, agents receive awards if their outcomes are exceeded by the corresponding standards.

The DOL sets performance standard benchmarks for each performance measure based on the historic performance of other training centers in the system. For the non-cost (cost) measures, the DOL sets the benchmark at the 25th (90th) percentile of the agent performance nationwide for the previous two ⁻scal years; this means that 75 (90) percent of agents in the previous two years would have attained the standard. The DOL o®ers states a procedure for adjusting the each measure's benchmark by the characteristics of the local labor market (e.g., the local unemployment rate) and by characteristics of the agent's enrollee population (e.g., enrolee representation of welfare recipients). The purpose of the adjustment procedure is to level the playing ⁻eld so that agents are held to standards that are appropriate to their local economic conditions and the kinds of clients served. The states have discretion over the formulation of the standards, but most states during the period under investigation adopted the same DOL formulae to control for outside factors.⁴ Table 2 computes the fraction of agents who have exceeded the performance standard and the average performance in excess of the standard (that is, the actual performance outcome minus the standard) for the seven performance measures. Table 2 shows that while most agents exceed the standard, their excess performances vary considerably.

Finally, we present the award prizes. By mandate, a state's award pot is about seven percent of the training budgets it supervises.⁵ Table 3 presents the mean and standard deviation of the agents' awards, and of their awards per unit of budget. The award per unit of budget varies across agents suggesting that the award funds are not allocated only according to a proportional sharing rule. We also ⁻nd (not reported here) that the level of awards vary greatly across agents within a state rejecting a ⁻xed sharing rule where the award pots would be distributed equally across agents.

Although the awards vary greatly across agents, there are some important restrictions on the award distribution. First, the awards have to be positive, meaning that the states cannot reduce the agents' budgets following a poor performance. Second, the states cannot spend more than the award budget even if all agents do exceptionally well: the

⁴See Heckman et al. (1997) for a general discussion on the use of performance standards in government organizations.

⁵The JTPA funds are allocated in three sub-funds: 78 percent are set aside for training services, 6 percent are set aside for the incentive system and the remaining 16 percent are set aside for other special services. The award fund as a fraction of total training budget is 7.1 percent (6/(78+6)) if one assumes that all award funds are eventually distributed as training budget. The actual ⁻gure should typically be lower than 7.1 percent because some of the incentive set aside fund is spent to administrate the incentive funds. In our data, the award as a fraction of budget also varies across states because some agents are missing in some states. The fraction of award to budget will be greater than seven percent, for example, when poorly performing agents are missing.

award has to be fully-funded.⁶

3 The Model

The previous Section showed that budget sizes, performance outcomes, and award prizes varied greatly both within and across incentive systems. One goal of this paper is to investigate whether incentive theory can explains these variations. Our objective in this Section is to provide a framework for structuring and motivating the empirical analysis. In the core of this Section, we restrict to the simple design problem with only two agents. To establish a comparison benchmark, we will also ignore scale e[®]ects in budget size. Toward the end of this Section, we show how the main qualitative predictions generalize to multi-agents and non-linear budget e[®]ects.

Agent i 2 I = f1; 2g manages budget b_i with $b_1 \ b_2$. Agent i has reservation utility U(b_i) = b_i U, and exerts e®ort e_i at cost b_i c(e_i) with c^0 , c^0 , and c^{00} positive and c(0) = $c^0(0) = 0$. The principal values e®ort $\frac{1}{4}(b_i; e_i) = b_i e_i$ from agent i. Let W denote the award pot for agents b_1 and b_2 .⁷

Budget multiplies all the fundamental parameters of the model in a proportional fashion. The cost and pro⁻t functions say that e[®]ort is measured in e±ciency units. Under no scale e[®]ect, e[®]ort should be understood as a measure of quality of managerial decision. This framework suggests a simple comparison benchmark corresponding to the e±cient (or ⁻rst-best) levels of e[®]ort in the absence of moral hazard problems. The e±cient e[®]orts maximize the weighted sum of e[®]orts b₁e₁ + b₂e₂ subject to the participation constraints w_i i b_ic(e_i) , b_iU for i 2 I and the budget constraint W , w₁ + w₂ where w_i is the wage paid to agent i 2 I. The optimal level of e[®]ort is the same for both agents,

$$e^{e} = c^{i^{1}}(f_{i} U);$$

⁶States may be able to transfer some award fund from one ⁻scal year to the other although there are some constraints restricting the amount states can transfer. For simplicity, we will focus in the model section on the polar case where the amount they can transfer is zero.

⁷As a side comment, we assumed that the award fund was ⁻xed. This assumption simpli⁻es the analysis and does not really matter for our empirical application since the interest there is not on the optimal award pot (W) but rather on the optimal award function to be de⁻ned below. We could also solve for the optimal award pot. This would just add another decision variable without much supplementary insight for our application.

where $f = \frac{W}{b_1+b_2}$ represents the award as a fraction of budget and we assume $f_i \ U > 0$ to guaranty that $e^e > 0$. One should think of e^{\circledast} ort as an $e\pm$ ciency multiplier in the use of the budget. Both agents supply the same e^{\circledast} ort because they equally increase the $e\pm$ ciency of their budgets. Agent i's wage is equal to its relative share of total budget $\frac{b_i}{b_1+b_2}W$.

Next, consider the moral hazard case. In line with the moral hazard paradigm, we assume that the principal cannot directly observe the agents' e®orts but observes only an imperfect measure of performance. To simplify, we assume that the performance measure can only take high or low values. Four performance outcomes may occur that we will denote J = fhh; hl; lh; llg where performance outcome hl, for example, is interpreted as agent one performing high and agent two low. Outcome j 2 J occurs with probability $p^{j}(e_{1}; e_{2})$ and agent i 2 I then receives w_{1}^{j} . To focus on the main issues, we will assume a simple symmetric linear functional from for the joint probabilities. The symmetry and linearity assumptions in addition to the condition that the probability that an agent achieves a given level of performance does not depend on the other agent's e®ort (e.g. $\frac{d}{de_{2}}(p^{hh} + p^{hl}) = 0$) imply that $p^{hh}(e_{1}; e_{2}) = k^{hh} + @e_{1} + @e_{2}$, $p^{hl}(e_{1}; e_{2}) = k^{hl} + [e_{1}] @e_{2}$, $p^{lh}(e_{1}; e_{2}) = k^{lh} = [e_{2}, and p^{ll}(e_{1}; e_{2})] = k^{lh} = [e_{2}, and k^{j}]$ given constants such that $k^{hl} = k^{lh}$ and $p^{j} 2 [0; 1]$ within the relevant e®ort ranges.

De⁻ne agent i's expected award as,

$$W_i(e_1; e_2) = \frac{X}{\sum_{j \ge J} p^j(e_1; e_2) w_i^j}$$

To focus on the main issues, we will assume that the agents are risk neutral.⁸ Agent i's utility under the above award scheme is,

$$U_i(e_i) = W_i(e_1; e_2) i b_i c(e_i)$$
:

The incentive compatibility constraint for agent i says that she chooses the level of e[®]ort that maximizes her utility given the other agent's e[®]ort. The ⁻rst order condition

⁸Under the strong participation constraints to be introduced below this assumption is not very restrictive since the agents are guaranteed their reservation utilities anyway.

to agent i's maximization problem is,

$$\frac{d}{de_i}W_i(e_1;e_2) = b_ic^0(e_i) \quad (I \ C \ C_i):$$

The ⁻rst order condition is su±cient because the agent's maximization problem is convex. The next set of constraints says that the principal guaranties the agents their reservation utility under every performance outcome. Stretching the contract literature's terminology, we will call these constraints the strong participation constraints,

$$w_{i}^{j}$$
 | $b_{i}c(e_{i})$ | $b_{i}U$ (SPC_i^j);

for j 2 J and i 2 I. These participation constraints are stronger than the ones found in the incentive literature, or weak participation constraints, saying that the agents are better-o[®] participating on average,⁹

The ⁻nal set of constraints is new to this problem and will play an important role in the analysis. These constraints say that the total award payments in any performance outcome cannot exceed the total award pot. We call these constraints the strong budget constraints.

$$W_{y} w_{1}^{j} + w_{2}^{j} (SBC^{j});$$

for j 2 J. The strong budget constraints emerge, for example, when the incentive system has to be fully funded so that the principal cannot transfer award funds from one contract year to the other. They are the mirror image to the principal of what the strong participation constraints are to the agent. The strong budget constraints are stronger than the standard budget constraint found in the incentive literature, or weak budget constraint in this work, saying that the award cannot exceed on average the total award pot,

$$W = W_1(e_1; e_2) + W_2(e_1; e_2)$$
 (WBC):

⁹The SPC as modelled here are a strong version of the limited liability constraint found in the literature saying that the agent's utility has to be greater than a ⁻xed constant that could be lower than the agent's reservation utility. SPC occur in practice when the principal needs to overcome the agent's resistance to the introduction of explicit incentives. The principal uses SPC to reassure the agent that she will not lose-out under the new compensation contract (e.g. Lazear, 1999).

In the analysis Section, we will pay special attention to two incentive mechanisms that have received much attention in the contract literature and that are commonly used in practice: piece rate awards and tournaments. An issue of interest will be to investigate if the optimal mechanism can be implemented by these mechanisms. For clarity, we formally de⁻ne these two mechanisms. A piece rate award mechanism rewards each agent based on her performance outcome alone. Formally, agent b₁ is rewarded according to a piece rate if $w_1^{hh} = w_1^{hl}$ and $w_1^{ll} = w_1^{lh}$. A tournament mechanism ranks the agents and rewards them a prize that depends on their rankings alone. This implies that $w_1^{hl} = w_2^{lh}$ and $w_2^{hl} = w_1^{lh}$.¹⁰

4 Analysis

We analyze the problem gradually. First, we solve the incentive design problem under moral hazard with only the weak participation and budget constraints. The novel twist in this analysis is to revisit the standard incentive design problem with heterogeneous agents. Second, we investigate the problem with the strong version of these constraints. This is the main contribution of this theoretical section.

Moral Hazard with WBC and WPC Under moral hazard, the e±cient outcome can be achieved as long as the ICCs and the WPC hold at the e±cient level of e®ort. Then, the WBC is implied by the WPCs. The ICCs will hold at the optimal level of e®ort if the principal can create an award di®erential between high and low performances large enough to provide the right e®ort incentives. The principal will be able to bind the WPC if it can adjust the average level of performance by punishing the agent under low performance to compensate for the high rewards under high performance.

This will typically be the case as long as the principal has enough degrees of freedom on the 8 outcome dependent awards (w_i^j) to satisfy the 5 constraints $((ICC_i, WPC_i)_{i21}, WBC)$. Many mechanisms implement the e±cient outcome but the goal of this section is to focus on piece rate and tournament.

¹⁰We assume that when both agents achieve the same outcome, they are randomly ranked. Tournament then implies that the total award given when agents perform the same is equal to the total award when they perform di[®]erently, $w_1^{hh} + w_2^{hh} = w_1^{II} + w_2^{II} = w_1^{hI} + w_2^{hI}$.

To start, note that the principal cannot implement the e±cient outcome under a \pure" tournament. A tournament o[®]ers a ⁻xed prize schedule that is independent of the size of contestants. The tournament's winner then earns the same prize whether it is managing a large or a small budget. When $b_1 > b_2$, tournaments give too much incentive to the small agent relative to the large one. This result is similar to the result in the tournament literature that tournaments may not achieve the e±cient outcome when one agent has a comparative cost advantage (Lazear and Rosen, 1981). The solution in these models is to handicap the favorite agent. In our model, a simpler solution consists in a modi⁻ed tournament structure where the prize schedule is weighted by the sizes of the agents. De⁻ne a `weighted tournament' mechanism as a tournament where the winner earns $b_i w^W$ and the looser $b_{i,i} w^L$ where w^W and w^L are the prizes per unit of budget.

Proposition 1 Under WBC and WPC, the e±cient outcome can be implemented under a weighted tournament system where awards are proportional to budget sizes.

A similar analysis applies to piece rate system. Although the principal cannot implement the $e\pm$ cient level of e®ort with a single piece rate rewarding only high and low performances, she can implement the $e\pm$ cient levels of e®ort under a weighted piece rate system. Weighted tournament and weighted piece rate belong to a more general class of `weighted mechanisms' that satisfy $\frac{w_1^1}{w_2^1} = \frac{b_1}{b_2}$ for j = hh; II and $\frac{w_1^{h_1}}{w_2^{h_1}} = \frac{b_1}{b_2}$. There are many weighted mechanisms that implement the $e\pm$ cient outcome. The intuition is that under a weighted incentive scheme ICC₁ is equivalent to ICC₂ and similarly WPC₁ is equivalent to WPC₂. Therefore, the principal can achieve the $e\pm$ cient outcome because she has 4 degree of freedom (the four prizes) and must satisfy only two constraints (ICC and WPC). Note, however, that there are some mechanisms that do not satisfy the condition for a `weighted mechanism' and that still implement the $e\pm$ cient outcome.¹¹

Moral Hazard with SBC and SPC Let's now turn to the design problem with the

 $[\]frac{\left[(w_1^{hh}_i \ w_1^{lh}) + \left[(w_1^{hi}_i \ w_1^{ll}) + \left[(w_1^{hi}_i \ w_1^{ll})\right]\right]}{\left[(w_2^{hh}_i \ w_2^{hl}) + \left[(w_2^{hi}_i \ w_1^{ll})\right]\right]} = \frac{b_1}{b_2}.$ This condition says that prizes have to be weighted but only in an average sense.

strong budget and participation constraints. The incentive design problem is,

(ID)

$$Max \quad \frac{1}{4_1(e_1)} + \frac{1}{4_2(e_2)}$$

$$(e_i; w_i^j)_{i_{21}}^{j_{2J}}$$
s:t: (ICC_i; SPC_i^j; SBC^j)_{i_{21}}^{j_{2J}}

To start, we consider the relaxed incentive design problem (RID) where we take into account only the inequality $\frac{d}{de_i}W_i(e_1;e_2)$, $b_ic^0(e_i)$ from the ICCs. It will be easy to check that the principal can still implement the optimal RID pro⁻ts when the reverse inequalities are imposed.

Lemma 1 The optimal RID pro⁻ts can be implemented by an incentive system where SPC₁^{II}, SPC₂^{II}, SPC₂^{hI}, SPC₁^{hI}, SBC^{hI}, SBC^{hI}, and SBC^{II} bind (hold as equality) and SPC₁^{hI}, SPC₂^{II}, and SBC^{II} do not bind.

To provide e[®]ort incentives, the principal tries to create the largest award di[®]erential between high and low performances. This has straightforward implications for the states of the world where only one agent performs well. The agent who does not perform gets her reservation utility while the agent who does perform get the rest of the award pot. Similarly, when both agents perform poorly they get only their reservation utilities. Lemma 1 greatly simpli⁻es the incentive design problem. In fact, we can replace, or get rid of, most of the constraints and are left only with ICC₁, ICC₂, SPC₁^{hh}, and SPC₂^{hh}. De⁻ne the simpli⁻ed relaxed incentive design problem as,

where the all award prizes but w_2^{hh} have been replaced using Lemma 1. Let c^c represent the Lagrange multiplier associated with constraint c.

Lemma 2 $\]^{SPC_1^{hh}} = 0$ and $\]^{ICC_1} > 0$ in the optimal SRID contract.

The large agent is the one who is $di\pm cult$ to motivate. The incentive compatibility constraint will always bind for that agent. Similarly, that agent will always receive more than its reservation utility when both agents are performing high. The intuition is simple.

The small agent gets a disproportionaly large award when she is the only high performer. Therefore, the small agent is facing stronger incentives than the large one from the way the award pot is distributed when there is only one high performer. This has to be balanced if one wants the two agents to provide the same e[®]ort and the only opportunity to over reward the large agent is when both agents perform well. A ⁻nal result will help interpret the results.

Lemma 3 $SPC_2^{hh} = 0$, $e_1 = e_2$ and $SPC_2^{hh} > 0$, $e_1 < e_2$.

This Lemma says that the small agent supplies more e^{\otimes} ort than the large one if she just receives her reservation utility in the state of the world where both agents perform high. The optimal incentive scheme depends on which constraints out of ICC₂ and SPC₂^{hh} bind and this in turn depends on the parameters of the model. Three mutually exhaustive cases may occur. (A formal proof is presented in the Appendix.)

- 1. Contract (C1), $(\]^{1CC_2} > 0$, $\]^{SPC_2^{hh}} = 0$). The solution to SRID without SPC_2^{hh} does satisfy SPC_2^{hh} . Then, both agents supply the same $e^{@}$ ort $e_1 = e_2$. The optimal pair (e, w_2^{hh}) is obtained by solving the agents' $\]^{rst}$ order conditions.
- 2. Contract (C2), ($_1^{ICC_2} > 0$, $_3^{SPC_2^{hh}} > 0$). The small agent supplies more e®ort than the large one and is paid her reservation utility in the state of the world where both agents perform well $w_2^{hh} = b_2(U + c(e_2))$. The small agent's ICC binds.
- 3. Contract (C3), ($_{,}^{ICC_2} = 0$, $_{,}^{SPC_2^{hh}} > 0$). Again, the small agent supplies more e[®]ort than the large one and is paid her reservation utility in the state of the world where both agents perform well. The di[®]erence now is that the small agent's ICC does not bind. As a consequence the awards w^{II}₂ and w^{Ih}₂ are not uniquely determined.¹²

Note that the optimal contract is not uniquely determined only in contract (C3) for the small agent and for performance outcomes Ih and II. The intuition for this result is

 $^{^{12}}$ The optimal SRID award scheme violates ICC₂'s reverse inequality. To meet that constraint, it is necessary to lower w^{lh}₂ and/or increase w^{ll}₂. It is possible to do so because ICC₂ in SRID does not bind so SBC^{Ih} and SPC^{II}₂ do not have to bind. Any combination of w^{ll}₂ and w^{lh}₂ that binds ICC₂ and satis⁻es SPC^{II}₂ and SBC^{Ih} implements the SRID pro⁻ts and satisfy all the ID constraints.

simple. The small agent would be facing too powerful incentive if she would receive the entire leftover award pot (after giving the large agent her reservation utility) when she is the only high performer and only her reservation utility when both agents perform low. Under such powerful incentive, the small agent would supply too much e®ort relative to the large one. Therefore, $\int^{ICC_2} = 0$. One solution to lower the small agent's e[®]ort is to waste some award funds when the small agent is the only high performer. Another way to go is to increase the small agent's award when both agent perform poorly. The principal is indi[®]erent between these two options.

	The Optimal Award Prizes							
	C1	C2	C3					
W ₁ ^{hh}	W j W_2^{hh} (a)	$W_{i} b_{2}(U + c(e_{2}))$	$W_{i} b_{2}(U + c(e_{2}))$					
W_2^{hh}	W ₂ ^{hh}	$b_2(U + c(e_2))$	$b_2(U + c(e_2))$					
W ₁ ^{hl}	$W \mid b_2(U + c(e))$	$W_{i} b_{2}(U + c(e_{2}))$	$W \mid b_2(U + c(e_2))$					
W ₂ ^{hl}	$b_2(U + c(e))$	$b_2(U + c(e_2))$	$b_2(U + c(e_2))$					
W ₁ ^{lh}	$b_1(U + c(e))$	$b_1(U + c(e_1))$	$b_1(U + c(e_1))$					
W ₂ ^{lh}	$W i b_1(U + c(e))$	$W_{i} b_{1}(U + c(e_{1}))$	$W_2^{lh\ (b)}$					
W ₁	$b_1(U + c(e))$	$b_1(U + c(e_1))$	$b_1(U + c(e_1))$					
W_2^{II}	$b_2(U + c(e))$	$b_2(U + c(e_2))$	W ₂ ^{II}					

The Ontimal Award Prizes

 ${}^{a}w_{2}^{hh}$ solves ICC₁ and ICC₂ for $e_{1} = e_{2}$. ${}^{b}Any w_{2}^{lh}$ and w_{2}^{ll} that satisfy SPC₂^{II}, SBC^{Ih} and ICC₂ at the optimal levels of e[®]ort (e₁; e₂).

Table 1 presents the optimal award prizes under the three possible contracts. In contract (C1) when both agents are doing well, the large agent receives a larger award than the smaller one by a factor that overstates their sizes di[®]erence $\left(\frac{w_1^{hh}}{b_1} > \frac{w_2^{hh}}{b_2}\right)$. The intuition is that the small agent is already facing pretty strong incentives because she can be generously rewarded when she is the only high performer. Therefore, the small agent does not need to be rewarded as much as the large one does when both perform well. This result will also typically hold for contracts (C2) and (C3) as long as the small agent does not exert much more e®ort than the large one.

Table 1 shows that the principal does not always distribute the entire award pot. This will typically occur when performance is low across the board. Burning out some award money is the optimal punishment scheme to provide ex-ante incentives. The rational for this outcome is that the principal cannot carry award funds from one incentive contract to the other. Under contract (C3), the principal may even burn some award fund in the

state of the world where only the small agent performs well.

Another implication of Table 1 is that the optimal incentive scheme cannot be implemented under a (weighted) piece rate system. In fact, under a piece rate system the small agent would receive the same prizes when she is the only high performer and when both agents perform high. In the optimal contract, however, the small agent receives less when both agents perform high than when she is the only high performer ($w_2^{lh} > w_2^{hh}$). Similarly, a (weighted) tournament system cannot be optimal because it would entail to sometimes reward the large agent more than its reservation utility when both agents perform low.

The agents' awards depend not only on their performances but also on the performances of the other agent. The reason for the optimality of group incentive in this model with SPC and SBC constraints is distinct from the standard reason found in the incentive literature. The traditional reason is that group incentives allow the principal to better insure the agents against performance risk when the measures of performance are stochastically related across agents. This is also known as Holmstrom's (1979) informativeness principle. In this model, agents are risk-neutral and group incentives are optimal even when the performance outcomes are independent across agents. The reason for the optimality of group incentives here comes from the need to cross-subsidize performance rewards in order to increase the award di®erential in the presence of the strong budget constraint.

In the empirical section, we want to investigate how the optimal contract changes as agents are more heterogeneous and as total budget changes. To investigate this issue theoretically we assume that the budgets are $b_1 = \frac{1}{b} + C_b$ and $b_2 = \frac{1}{b}_i C_b$ with $\frac{1}{b} > C_b > 0$. To control for scale e®ects, we will assume that $W = f\frac{1}{b}$ so that the award pot increases proportionally with budget.

Proposition 2 There exist $0 < \mathbb{C}^{\pi} < \mathbb{C}^{\pi\pi} < 1$ such that (C1) is optimal for $\frac{\mathbb{C}_{b}}{b} < \mathbb{C}^{\pi}$, (C2) is optimal for $\mathbb{C}^{\pi} < \frac{\mathbb{C}_{b}}{b} < \mathbb{C}^{\pi\pi}$, and (C3) is optimal for $\frac{\mathbb{C}_{b}}{b} > \mathbb{C}^{\pi\pi}$.

This Proposition together with Lemma 3 implies that $e_1 = e_2$ for C^{μ} , $\frac{c_h}{b}$ while $e_1 < e_2$ for $\frac{c_h}{b}$, C^{μ} . When agents are heterogeneous enough, the small agent exerts more

e®ort than the large one. This will happen if the award system over-rewards the small agent when she is the only high performer so much that this cannot be compensated by under-rewarding her when both agents perform well. The principal will not be able to level incentives across agents when the agents' budgets are too heterogeneous. Note that the proper measure of agent heterogeneity is relative budget di®erence scaled by mean budget. Budget di®erences matter more when mean budget is lower. Put di®erently, the larger the average award pot, the easier it is for the principal to compensate for budget heterogeneity. The ⁻nal result regards the average level of prizes and the average level of performance.

Proposition 3 The small agent is more likely to perform high than the large one. The small agent earns more on average than the large one when $\frac{b_1}{b_2} > \frac{p^{h1} + p^{hh}}{p^{lh}}$.

We conclude with a comment on the welfare implications of the model. The SPC and SBC are source of two kinds of distortions. First, the optimal incentive system does not always allocate e[®]ort optimaly across agents. When agents are too heterogeneous, the large one exerts too little e[®]ort and the small one too much e[®]ort. Second, even contract C1 does not achieve the e±cient outcome although it does satisfy the condition that both agents supply the same level of e[®]ort ($e_1 = e_2$). There are two reasons for that. One reason is that the SBCs force the principal to throw away award funds when both agents perform poorly.¹³ Another reasons is that the agents receive more than their expected reservation utility under SPC. As a consequence, agents exert less e[®]ort under SPC and SBC than under WPC and WBC. Note that the ine±ciency of having the SBC and the SPC is not driven by one of these constraints alone. In fact, the principal would be better-o[®] with SBC and WPC or with WBC and SPC than with SBC and SPC. Both the limited liability and the fully funded constraints bind.

Extensions The most crucial assumption in the model is the assumption that there are no (dis)economies of scale in budget size. To investigate the role this assumption, we

¹³To the extent that the award money could be used for other activities than agent compensation, the e±ciency impact of this distortionary e[®]ect could be mitigated and really depends on the value of these other activities. Interestingly, the JTPA incentive system anticipated that potential problem and created a \technical assistance" fund. States are allowed to channel some of the award money to the technical assistance fund to help poorly performing training agencies.

assume that the cost function is not linear in budget $\left(\frac{@^2C}{@^2b} \notin 0\right)$ and consider how this would change the optimal contracts and the levels of e®ort. Assuming diseconomies of scale in budget $\frac{@^2C}{@^2b} > 0$ would add a force pushing toward requiring more e®ort from the small agent relative to the large one while assuming economies of scale in budget $\frac{@^2C}{@^2b} < 0$ would push toward relatively less e®ort from the small agent. Under moral hazard with SPC and SBC these e®ects would just add to the incentive e®ect we identi⁻ed in the analysis.

It is clear that it would be impossible to identify the incentive e[®]ect in a single contract environment without knowing anything about the cost function. This is not true, however, if one has access to a cross section of contracts that cover di[®]erent pools of agents. To illustrate this point, assume for example, that there are two pairs of agents where the large agent in one pair is the same size as the small agent in the other pair. Then, a simple extension of the model would predict that although these two agents are identical, they should receive di[®]erent awards when they perform well and their paired agent also do so. In such event, the agent that is paired with a larger agent should receive a smaller award than the one that is paired with a smaller one. All the other predictions of the model can also be identi⁻ed.

Another important assumption of the model is that there are only two agents. To simplify, consider the case of four agents corresponding to two identical pairs of agent and let's compare this four-agent case (two pairs) with the corresponding two-agent case (one pair). In the four-agent case, the distortion e[®]ect identi⁻ed in the two-agent case will be less pronounced because the principal will have more degree of freedom to smooth the award function across agents. In addition, the performance outcome where all agents perform poorly will occur less frequently implying that the principal will burn out the award pot less frequently. These two forces imply that average performance should increase with the number of agents.

5 Application to the JTPA Incentive System

In this section we test whether states implement the optimal award scheme. The theoretical model establishes how the agents' awards should depend upon their budgets and upon their performances. The theoretical model also makes predictions about how the award distribution and the performance outcomes should vary across states that supervise di®erent agent pools. We test the following predictions of the optimal incentive system:

- Award as a function of budget | An agent that is small relative to the average agent in the state receives disproportionally large awards, given its performance. States should distribute on average less than their entire award pot.
- Award as a function of performance | The agent's award should depend positively on its performances but negatively on the performance of other agents within the same incentive system.
- 3. Performance as a function of budget | Smaller agents should perform better on average than larger ones. States that are more heterogeneous should perform worse.

To test these implications we use data that contain information on performance outcomes on the seven DOL measures, on awards and on budgets. Depending on the prediction we are testing, our unit of observation is either a training agency or a state. Our two data sources were presented in footnote 3. From the SRI data set, we have ⁻nancial data for approximately 400 of the training agencies in ⁻scal years 1985 and 1986. For about 42 states, we have a signi⁻cant fraction of the agents. The sample we work with represents only about two thirds of the JTPA population of training agencies (recall that there are over 600 training agencies in JTPA) primarily because many training agencies failed to report their awards and/or their budgets. In addition, we have agency performance outcomes and standards for most agencies between 1984 and 1988 from the JASR data set. Broadly speaking, our testing strategy is to examine whether incentive theory predicts how awards are distributed and how agents respond to awards in JTPA. Award as a function of budget | We begin by testing how the agent's budget in^ouences its award. In this reduced form approach we focus on the predictions that (a) larger agents should receive larger awards and (b) agents who are relatively small in their states should receive disproportionately larger awards. Model I in Table 4 regresses award on budget. Model I shows that the award rises on average 4 cents for a 1 dollar increase in budget. (The coe±cient estimate is statistically signi⁻cant.) Model II in Table 4 adds to the right hand side of the regression the mean budget in the agent's state. The mean budget picks up the e[®]ect of the agent's relative size.

Several implications can be drawn from the two regressions in Table 4. First, note that the intercept, which is positive and signi⁻cant in Model I, is not signi⁻cantly di[®]erent from zero in Model II. This says that states do not give ⁻xed prizes independently of size. Second, the coe±cient estimate on mean budget is signi⁻cant and positive, indicating that agents that are large relative to their state peers earn less. These results are consistent with the thrust of the theory.

Award as a function of performance | In testing for budget e[®]ects, we concentrate on the determinants of scaled awards or awards per unit of budget. Table 5 explores the implications of the model by examining the e[®]ects of performance, and performance relative to the performance of other agents in the state, on the award as a fraction of the budget. The regressions in Table 5 include on the right hand side measures of excess performance, the agent's performance outcome minus the corresponding performance standard. The wage and cost measures in the excess performance calculations are denominated in dollars. The employment rate and youth positive termination rate measures are multiplied by 100.¹⁴

In these regressions, the right-hand side contains seven measures of agent excess performance.¹⁵ Recall that the incentive system rewards cost outcomes only when they are exceeded by the cost standard. For the sake of consistency, we compute excess performance for the two cost measures, CE and CEY, as the performance standard minus the

¹⁴For example, the excess adult employment rate measure for an agent who produces a year-end employment rate of 70 percent and faces a standard of 67 percent, is calculated as 70_{i} 67 = 3.

¹⁵Each performance measure must receives a positive weight in the determination of the agent's award.

outcome. That way, if the regression is correctly speci⁻ed, and each performance outcome matters for the award, we should ⁻nd that the coe±cient estimates on excess performance are positive.

Because we have a two-year panel for each agent, we estimate the relationship using a random e[®]ects model, i.e., with separate, agent-speci⁻c disturbances. All regressions reported include state dummies to control for state variation in other dimensions of the incentive system that a[®]ect award size. We build the model in two steps. We ⁻rst investigate the role of performance, and then investigate the role of relative performance in the determination of the award size.

Model I contains on the right-hand side only measures of excess performance. The coe±cient estimates for the average wage at placement measure, the adult cost measure, and the youth employment measure have the predicted signs and are statistically significant by conventional signi⁻cance criteria. To understand the impact of performance on the award implied by these point estimates, consider the average agent whose budget is equal to \$3 million (the approximate mean budget in our sample). A \$100 reduction in the cost per placement relative to the cost standard raises the agent's award by approximately \$3,300. A 10 cent increase in the wage at placement relative to the wage standard raises the agent's award by \$14,100. A 10 point increase in the agent's youth placement relative to the standard raises the agent's award by \$3,600. These ⁻gures correspond to arc award elasticities of .37, .97 and .25, respectively.¹⁶

Model II investigates whether awards are determined by relative performance. On the right hand side, we add to the agent's own excess performance the mean values of excess performance in the agent's state. Negative coe±cients on the mean values indicate that agents are paid more when the other agents in their state do worse. Here we are testing the model's predictions that the states construct group incentives. The coe±cient estimates on the mean values of excess performance in the average wage at placement and the youth cost measures both have the predicted sign and are signi⁻cant (the p values

¹⁶Another relevant measure is the budget elasticity to performance. These elasticities are about fourteen times smaller since the award represents only about eight percent of the budget. Although these elasticities may seem small, they are not when compared to similar measures estimated from executive compensation (Jensen and Murphy, 1990).

of the two-tailed tests of signi⁻cance are .004 and .001, respectively). Consider again the agent with a budget equal to the system average of \$3 million per year. Independent of the agent's absolute average wage at placement (youth cost per placement) outcome, the agent wins approximately an extra \$27,000 (\$8,400) when its wage (youth cost per placement) outcome relative to the state average increases by 10 cents (decrease by \$100).¹⁷

A surprising \neg nding that emerges from Table 5 is that not all performance measures are signi \neg cant. Related to that result, we also \neg nd that the explanatory variables do not explain much of the variation in award per unit of budget. The R² for Model I is about .256. As a benchmark, the state dummies alone (this regression is not reported) explain about 13 percent of the total variation in the award per unit of budget. Thus, while the R² is low in the model, excess performance accounts for nearly half of the explained variation in the award per unit of budget. The addition of the mean values of excess performance in Model II only modestly raise the R² (from 25.6 percent to 27.8 percent).

The statistical insigni⁻cance of some coe±cients on excess performance and more generally, their limited explanatory power, have three possible causes. First, most award policies are highly nonlinear and complex. The low R² may re[°]ect that the linear speci⁻cation imposed in the regressions does not capture well how performance determines the award. Second, an accurate measure of the relationship between award and performance may be di±cult to obtain due to measurement error. Administrative data from JTPA data sources are known to contain considerable error.¹⁸ Third, states may be using award funds to meet political objectives rather than incentive objectives as assumed in our model. For example, states may use award funds to redistribute resources to politically-favored agents, or from one geographical area to the other.

Performance as a function of budget | The model predicts that smaller agents should exert more e[®]ort and achieve higher levels of performance. The estimates reported in

¹⁷These ⁻ndings are consistent with the model but they are also consistent with the hypothesis that the contracts use relative performance to control for common shocks. Our data does not reject the hypothesis that performance is statistically related within states.

¹⁸For example, for ⁻scal year 1986, the JASR and the SRI data set contain measures of the same performance outcomes and standards. These measures are frequently di[®]erent, and in non-systematic ways.

Table 6 test this hypothesis. Table 6 presents estimates of the determinants of performance with respect to each of the seven performance measures. Table 6 is divided into 7 panels, a panel for each of the seven performance measures. As in Table 5, the dependent variable is de⁻ned as excess performance de⁻ned as the performance outcome minus the standard for the non-cost measures and the opposite for the cost measures.

To test whether small agents perform better than large ones, we construct a measure of relative size that is equal to the di[®]erence between the agent's budget and the mean budget for its state, normalized by the mean budget.¹⁹ We include the budget variable to control for scale e[®]ect in the production of the performance outcome. Having done so, we can be sure that the coe±cient on the relative budget measure picks up only the performance e[®]ect of the agent's size relative to the size of its peers in the state.

Consider ⁻rst Model I. In the adult employment rate regression (Panel A), the coef-⁻cient on the relative budget measure is negative and signi⁻cant by conventional criteria (the p value is .09). A negative and signi⁻cant estimate in this speci⁻cation is consistent with our hypothesis that because they receive stronger incentives, small agents will generate greater outcomes. We ⁻nd negative but insigni⁻cant coe±cients on the relative budget variable for the adult welfare employment rate (Panel B), the adult wage at termination (Panel D), and the youth positive termination rate (Panel G) regressions. The coe±cients on relative budget size are positive for both cost measure regressions (Panels C and E) and signi⁻cant for the youth cost measure (its p value is .04). This later ⁻nding is inconsistent with the predictions of the model.

A prediction of the model is that relative size should be more important in states where agents are more heterogeneous. Model II estimates separate $coe\pm cients$ on the relative budget size measure for agents in highly heterogeneous states. We use as a measure of state heterogeneity the standard deviation of budget divided by the total allocation of the state. We divide the standard deviation by the state allocation to capture the idea that the larger the agents in a state, the smaller the distortion caused by a given amount of

¹⁹All regressions include state dummies to control for state variation in the other dimensions of the incentive system (e.g., state-speci⁻c modi⁻cations to the construction of the performance standard).

spread in budget sizes.²⁰ For Model II, we de⁻ne three indicator variables: \pm^{I125} is equal to one if the heterogeneity measure of the agent's state falls in the lower 25th percentile of the distribution of state heterogeneity outcomes, and equal to zero otherwise. \pm^{2575} is equal to one if the heterogeneity measure falls between the 25th and 75th percentiles, and equal to zero otherwise. \pm^{gt75} is equal to one if the heterogeneity measure falls between the 25th and 75th percentiles, and equal to zero otherwise. \pm^{gt75} is equal to one if the heterogeneity measure exceeds the 75th percentile, and equal to zero otherwise. The theoretical model predicts that the relative size should have a more pronounced e®ect on performance the greater the heterogeneity in the state. Therefore, the coe±cient estimates on $\frac{B_1 B}{B} \notin \pm^{2575}$ and $\frac{B_1 B}{B} \notin \pm^{gt75}$ in Model II are more likely to be more negative than the coe±cient on $\frac{B_1 B}{B}$ alone in Model I.

We nd negative and signi cant coe±cient estimates for the variable $\frac{B_1 B}{B} (t_2^{2575})$ for both adult employment measures (the p values are .09 and .08, respectively; Panels A and B).²¹ Coe±cient estimates were negative and insigni cant for the adult wage measure (Panel D) and the youth positive termination rate measure (Panel F). Again, the coe±cients estimates from the cost regressions were positive, contradicting the model. (In the adult cost regression, the coe±cient estimate on $\frac{B_1 B}{B} (t_2^{175})$ is signi cant, and in the youth cost regression, the coe±cient estimate on $\frac{B_1 B}{B} (t_2^{175})$ is signi cant.) Taken at face value, the evidence that relatively small agents face stronger incentives is mixed.

In Table 7, we test whether states that are more heterogeneous perform worse. For our measure of state performance, we compute a weighted average of excess performance, where the weights are the agents' relative sizes. Our measure of heterogeneity is once again the standard deviation of budget size, normalized by the state's budget allocation. We enter on the right hand side the mean budget in the state, to control for any separate scale e[®]ect. We estimate the relationship between a state's size distribution and the weighted performance measures using a panel of between 40 to 50 states for ⁻scal years

²⁰The model does not clearly specify how one should measure heterogeneity when there are more than two agents. We chose to divide the standard deviation in budget by the sum of budgets rather than by the average budget to capture the idea that a greater number of agents will provide the state more degrees of freedom with which to smooth the award function. In any event, we tried di®erent measure of heterogeneity and they give similar results.

²¹The regression estimates shown in Table 6 suggest that relative size matters more in more heterogeneous states, that is, that the coe±cient on $\frac{B_{\perp}B}{B}$ interacted with the heterogeneity measure is negative. We have conducted this test formally. While the point estimate of such a test is more often than not negative, we always reject the hypothesis at conventional levels of signi⁻cance.

1984 through 1988.

As in Table 6 we estimate 7 separate regressions, one for each performance measure. Model I includes only the budget. Model II contains both the budget and the budget heterogeneity variable, de⁻ned as before. Considering Model II, the coe±cients on the heterogeneity variable in ⁻ve of the seven regressions are negative, as predicted. The coe±cient is both negative and signi⁻cant in the youth cost and youth employment regressions (Panels E and G). In the two adult employment rate regressions (Panels A and B) the coe±cients are positive, but insigni⁻cant. Table 7 therefore presents weak evidence consistent with the model: states with more heterogeneous sets of agents perform worse with respect to the performance measures.

To summarize, the evidence provides some con⁻rmation of the theory's implication for how awards should depend on budgets and performance and how performance should depend on pool composition. We nd the following. First, we nd that the scale prediction holds: larger agents receive larger awards. We also ⁻nd that relative size matter: agents that are small relative to their state average receive larger awards. Second, we -nd that while the relationship is not as strong as we would expect, an agent's award is determined by its performance. This ⁻nding implies that a real incentive exists, and that awards are not fully determined by political or equity concerns. Third, we -nd some evidence that a high-performing agent's award is even higher when the other agents in the state perform poorly although again this evidence is not as widespread as it could be. Thus the award function depends on relative performance in a way that is consistent with the theory. Fourth, we ⁻nd that for some performance measures, relatively small agents perform better than large ones. This ⁻nding is consistent with the major implication of the model: that smaller agents face stronger incentives than larger ones. This evidence is mixed, however. For cost measures, relatively larger agents appear to generate higher outcomes, even after controlling for scale e®ects. Fifth, we -nd some evidence that e®ort distortions are greater in states with greater size disparities among agents. We also nd that relatively heterogeneous states perform worse than relatively homogeneous states for some measures of performance.

6 Conclusion

This paper studies the provision of incentives in a large federal job training organization for the disadvantaged. In this organization, each state develops a ⁻nancially-backed incentive system, subject to the constraint that the individual awards cannot be negative (limited liability constraint) and the sum of the awards cannot exceed a ⁻xed award pot (fully funded constraint). With this pot, states reward a pool of training agencies that typically manage di®erent budgets. The training agencies are evaluated on the basis of their performance relative to a ⁻xed set of performance standards. The states have considerable discretion in the construction of the incentive schemes. Piece rates and tournaments, for example, are allowed.

We show that in the presence of the limited liability and fully funded constraints on the award distribution, the optimal award function will not in general elicit the unconstrained e±cient level of e®ort from the agents. The optimal award scheme `over rewards' small agents relative to large ones. Because small agents receive relatively large awards, they put forth ine±ciently high levels of e®ort. We ⁻nd strong evidence consistent with the prediction that smaller agencies receive greater rewards and mixed evidence that smaller agents exert more e®ort. As predicted, we ⁻nd some evidence that ine±ciencies are greater in states that are more heterogeneous. Our evidence suggests that constraints on the award distribution lower the overall e®ectiveness of performance incentives.

Our analysis suggests that the e[®]ectiveness of performance incentive depends on the constraints organizations face. Not all organizations have to distribute a ⁻xed award pot to a pool of agents. In many incentive relationships, there is a surplus to be shared (e.g. peasants and landlords share crops, executives and stockholders share stock market value creation, and ⁻rms and sales people share sales margins). It is only when there is nothing to be shared that the principal prefers to set aside an award pot rather than, for example, taking the risk of committing to a subjective award formula that may lead the incentive system to bankruptcy. In that respect, the ⁻xed award pot feature distinguishes the incentive design problem that prevails in government organizations.

From a positive point of view, the analysis suggests that the sorting of agents into

pools is an important step in the design of incentive systems. In the same way that grading on a curve works well only in large classes, the use of ⁻xed award pots works better in large and homogeneous pools. Along the same lines, note that incentives would be more e[®]ective if states could transfer some of the award pot from one year to the other, thereby relaxing the fully funded constraint, or if states could punish agents by lowering their budgets when they perform poorly, thereby relaxing the limited liability constraint. Any of these solutions to the design challenge we uncovered, however, introduces practical problems of their own.

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Proofs

Proof of Proposition 1: First I show by contradiction that pure tournament system cannot implement the $e\pm$ cient level of e^{\circledast} ort when $b_1 > b_2$. Assume the opposite. Under the $e\pm$ cient level of e^{\circledast} ort $e_1 = e_2$ while under tournament w^W and w^L the agent's ICC say,

$$\begin{array}{l} \frac{1}{2}(p_1^{hh} + p_1^{ll})(w^W + w^L) + p_1^{hl}w^W + p_1^{lh}w^L = b_1c^0(e_1) \\ \frac{1}{2}(p_2^{hh} + p_2^{ll})(w^W + w^L) + p_2^{lh}w^W + p_2^{hl}w^L = b_2c^0(e_2): \end{array}$$

These condition imply $e_2 > e_1$. A contradiction. QED

Next, de ne the per-unit of budget tournament prizes,

(

$$\frac{1}{2}(p^{hh} + p^{II})(w^{W} + w^{L}) + p^{hI}w^{W} + p^{Ih}w^{L} = U_{i} c(e^{fb})$$

 $\frac{1}{2}(p_{1}^{hh} + p_{1}^{II})(w^{W} + w^{L}) + p_{1}^{hI}w^{W} + p_{1}^{Ih}w^{L} = c^{0}(e^{fb})$

A solution to this system always exists. The tournament prizes where the winner earns $b_i w^W$ and the looser $b_i^{\ i} w^L$ satisfy the ICCs and WPCs at the e±cient level of e®ort. The WBC is implied by the WPCs.

Proof of Lemma 1: The proof goes by contradiction. Assume for example that SPC_1^I does not bind, i.e., $w_1^{II} > b_1(U + c(e_1))$. Consider a new contract where w_1^{II} is decreased such that SPC_1^{II} binds. SBC^{II} and ICC_1 still hold while all the other constraints are unchanged. A contradiction. The same reasoning applies to show that there is an optimal contract where SPC_2^{II}, SPC_2^{hI}, and SPC_1^{Ih} bind. Next, we show that SBC^{hh} also binds. Assume it does not. Consider a new contract where w_1^{hh} is increased such that SBC^{hh} binds. SPC_1^{hh} and ICC_1 still hold while all other constraints are unchanged. A contradiction. A similar argument shows that SBC^{hI} and SBC^{Ih} also bind. Next, we show that SPC_1^{hI} does not bind. Assume it does, that is, $w_1^{hI} = b_1(U + c(e_1))$. Because SPC_2^{hI} and SBC^{hI} bind $b_1(U + c(e_1)) + b_2(U + c(e_2)) = W$. This implies that $w_i^i = b_i(U + c(e_i))$ for i 2 I and j 2 J. ICC_i imply that $e_i = 0$ for i 2 I. But k i U > guaranties that there is a solution with positive e®orts. A contradiction. The same reasoning shows that SPC_2^{Ih} does not bind. Finally, we show that SBC^{III} does not bind. Assume it does not bind. SPC_2^{Ih} does not bind. Finally, we show that SBC^{III} does not bind. Assume it does not bind. Assume it does bind, that is, $w_1^{II} + w_2^{II} = W$. Since SPC_1^{II} and SPC_2^{II} bind, we have $w_i^j = b_i(U + c(e_i))$ for i 2 I and j 2 J. Again a contradiction. QED

Proof of Lemma 2: We focus on solutions with positive e[®]orts for both agents. The RSID problem is,

The ⁻rst order condition to RSID are

$$\begin{array}{cccccc} FOC_{e_{1}} & 1 + \frac{1}{2} \overset{ICC_{1}}{(p_{1}^{h} + p_{1}^{h})} c^{\emptyset}(e_{1}) \ i \ \frac{1}{2} \overset{ICC_{2}}{(e_{1})} i \ \frac{1}{2} \overset{ICC_{1}}{(e_{1})} c^{\emptyset}(e_{1}) \ i \ \frac{1}{2} \overset{SPC_{1}^{hh}}{(e_{2})} c^{\emptyset}(e_{2}) \ i \ \frac{1}{2} \overset{SPC_{1}^{hh}}{(e_{2})} \ i \ \frac{1}{2} \overset{SPC_{1}^{hh}}{(e_{2})} c^{\emptyset}(e_{2}) \ i \ \frac{1}{2} \overset{SPC_{1}^{hh}}{(e_{2})} c^{\emptyset}(e_{2}) \ i \ \frac{1}{2} \overset{SPC_{1}^{hh}}{(e_{2})} \ i \ \frac{1}{2} \overset{SPC_{1}^{hh}}{(e_$$

Consider <code>-rst</code> the case <code>_ICC_2 > 0</code>. We show by contradiction that <code>_SPC_1^hh > 0</code> and <code>_SPC_2^hh = 0</code> is impossible. Assume this is true. <code>_SPC_1^hh > 0</code> implies w_2^hh = W i b_1(c(e_1) + U): ICC_2 says that,

$$c^{0}(e_{2}) = \frac{p_{2}^{hh} + p_{2}^{lh}}{b_{2}}(W_{i} b_{1}(U + c(e_{1}))_{i} b_{2}(U + c(e_{2})))$$

But $\frac{p_2^{h_1}+p_2^{h_1}}{b_2}$ (W i $b_1(U + c(e_1))$ i $b_2(U + c(e_2))$) > $\frac{p_1^{h_1}}{b_1}$ (W i $b_1(U + c(e_1))$ i $b_2(U + c(e_2))$) > $c^{0}(e_1)$ where the last inequality holds by ICC₁. Therefore, $c^{0}(e_2) > c^{0}(e_1)$ and $e_2 > e_1$. Next, $s^{SPC_1^{h_1}} = i s^{ICC_1}p_1^{h_1} + s^{ICC_2}p_2^{h_1} > 0$ and the symmetry property saying that $p_1^{h_1} = p_2^{h_1}$ implies that $s^{ICC_2} > s^{ICC_1}$. Replace $s^{SPC_1^{h_1}}$ in FOC_{e1} and substract FOC_{e1} and FOC_{e2} gives after using the symmetry properties for the marginal probabilities,

$${}^{ICC_{1}}(p_{1}^{hl}(c^{0}(e_{2}) | c^{0}(e_{1})) | c^{0}(e_{1})) = {}^{ICC_{2}}((p_{1}^{hh} + p_{1}^{hl})(c^{0}(e_{1}) | c^{0}(e_{2})) | c^{0}(e_{2})):$$

Since $(p_1^{hh} + p_1^{hl})(c^{0}(e_1)_{j} c^{0}(e_2))_{j} c^{0}(e_2) < 0$, $|CC_2| > |CC_1|$ implies after simplications,

$$(2p_1^{hl} + p_1^{hh})c^{0}(e_2) + c^{00}(e_2) < (2p_1^{hl} + p_1^{hh})c^{0}(e_1) + c^{00}(e_1)$$

The above inequality implies $e_1 > e_2$. A Contradiction.

Next, we show by contradiction that $_{SPC_1^{hh}} > 0$ and $_{SPC_2^{hh}} > 0$ is impossible. This would imply that $W = (b_1 + b_2)U + b_1c(e_1) + b_2c(e_2)$. Then, $w_i^j = b_i(U + c(e_i))$ for i 2 I and j 2 J and $e_i = 0$. A contradiction.

Finally, we turn to the case $_{1CC_{2}}^{ICC_{2}} = 0$. Assume $_{SPC_{1}^{hh}} > 0$. $FOC_{w_{2}^{hh}}$ implies $_{SPC_{2}^{hh}} > 0$. But $_{SPC_{1}^{hh}} > 0$ and $_{SPC_{2}^{hh}} > 0$ imply $w_{i}^{j} = b_{i}(U + c(e_{i}))$ for i 2 I and j 2 J and $e_{i} = 0$. A contradiction.

This establishes the Lemma's rst claim, $\[\] ^{SPC_1^{hh}} = 0$. To establish the Lemma's second claim, plug $\] ^{SPC_1^{hh}} = 0$ in FOC_{wh} and assume $\] ^{ICC_1} = 0$,

$$\int^{ICC_2} p_2^{hh} + \int^{SPC_2^{hh}} = 0$$

A contradiction since $p_2^{hh} > 0$. Therefore, $\sum_{n=1}^{1CC_1} > 0$. QED

Proof of Lemma 3: This rst part of this Lemma says that the two agents supply the same e[®]ort when SPC₂^{hh} does not bind. $SPC_2^{hh} = 0$ imply $ICC_1 = ICC_2$. Taking the di[®]erence in FOC_{e1} and FOC_{e2} gives

$$(p_1^{lh} + p_1^{ll} | p_2^{lh})c^0(e_1) | c^0(e_1) = (p_2^{hl} + p_2^{ll} | p_1^{hl})c^0(e_2) | c^0(e_2)$$

After simpli⁻cation, the above equation implies that $e_1 = e_2$. The second claim in the Lemma naturally follows. QED

(C1)
$$p_{2}^{hh}(W_{i} w_{2}^{hh}) + (p_{1}^{lh} + p_{1}^{ll})b_{1}(U + c(e)) + p_{1}^{hl}(W_{i} b_{2}(U + c(e))) = b_{1}c^{0}(e)$$

$$p_{2}^{hh}w_{2}^{hh} + (p_{2}^{hl} + p_{2}^{ll})b_{2}(U + c(e)) + p_{2}^{lh}(W_{i} b_{1}(U + c(e))) = b_{2}c^{0}(e)$$

This is the solution to the optimal design problem if the optimal wage satisfy SPC₂^{hh}, that is, w_2^{hh} i $b_2c(e)$ b₂. Next we show that $\frac{w_1^{hh}}{b_1} > \frac{w_2^{hh}}{b_2}$. Using ICC₁ and ICC₂,

$$\frac{w_2^{hh}}{b_2} = \frac{c^{0}(e) \ i \ (p_2^{hl} + p_2^{ll})(U + c(e)) \ i \ p_2^{lh} \frac{W_i \ b_1(U + c(e))}{b_2}}{p_2^{hh}} > \frac{c^{0}(e) \ i \ (p_1^{lh} + p_2^{ll})(U + c(e)) \ i \ p_1^{hl} \frac{W_i \ b_2(U + c(e))}{b_1}}{p_1^{hh}} = \frac{w_1^{hh}}{b_1}:$$

Consider next the case where $_1^{CC_2} > 0$ and $_s^{SPC_2^{hh}} > 0$. $_s^{SPC_2^{hh}} > 0$ implies $w_2^{hh} = W_i \ b_2(U + c(e_2))$. The optimal levels of e®ort are given by solving for the agents' rst order conditions.

(C2)
$$\begin{pmatrix} (p_1^{hh} + p_1^{hl})(W_i \ b_2(U + c(e_2))) + (p_1^{lh} + p_1^{ll})b_1(U + c(e_1)) = b_1c^0(e_1) \\ (p_2^{hh} + p_2^{hl} + p_2^{ll})b_2(U + c(e_2)) + p_2^{lh}(W_i \ b_1(U + c(e_1))) = b_2c^0(e_2) \end{pmatrix}$$

The ⁻rst order condition to the design problem are equivalent to $c^{0}(e_1)$, $(p_1^{hh} + p_1^{hl})(c^{0}(e_2)_i c^{0}(e_1))$ and e_2 , e_1 .

The \neg nal case is $\Box^{ICC_2} = 0$ and $\Box^{SPC_2^{hh}} > 0$. After replacement, one can show that the optimal levels of e[®]ort are given by,

(C3)
$$(p_1^{hh} + p_1^{hl})(W_i \ b_2(U + c(e_2))) + (p_1^{lh} + p_1^{ll})b_1(U + c(e_1)) = b_1c^0(e_1) \\ c^{00}(e_1) = (p_1^{hh} + p_1^{hl})(c^0(e_2)_i \ c^0(e_1))$$

1

This contract does not satisfy ICC_2 's reverse inequality in ID. However, the optimal pro⁻ts can be implemented under ID by increasing w_2^{II} and/or decreasing w_2^{Ih} by the correct amounts so that ICC_2 holds at the optimal (C3) levels of e[®]orts.

Finally, from $FOC_{w_2^{hh}}$ we have $\int CC_2 = 0$, $\int SPC_2^{hh} > 0$ implying that the case $\int CC_2 = SPC_2^{hh} = 0$ is impossible. The three contracts C1, C2 and C3 are exhaustive and mutually exclusive.

Proof of Proposition 2: $De^{-}ne e^{\alpha}$ as the level of $e^{\otimes}ort$ that solves the agents' ICCs in (C1). Adding the two ICCs gives,

$$kp_1^{hh} + 2(p_1^{lh} + p_1^{ll})(U + c(e^{\alpha})) + 2p_1^{hl}(k_i (U + c(e^{\alpha}))) = 2c^{0}(e^{\alpha}):$$

The above equation shows that e^{α} does not depend on $\frac{\mathbb{C}_{h}}{b}$. De ne \mathbb{C}^{α} such that the agents' ICCs hold at e^{α} when $w_{2}^{hh} = \frac{1}{b}(1 + c(e^{\alpha}))$,

$$\Phi^{a} = \frac{p_{1}^{hh}}{2c^{0}(e^{a})}(f_{i} 2(U + c(e^{a}))):$$

For $\frac{\Phi_b}{b} = \Phi^{\pi}$, (C1) is obviously the solution to the design problem. Define $G(\Phi_b = \hat{b}) = \frac{w_2^{hh}}{b}i$ (1 i $\frac{\Phi_b}{b}$)(U + c(e)). Using ICC₁ one can show that $\frac{dG}{d(\Phi_b = \hat{b})} = i \frac{c^0(e)}{p_1^{hh}} < 0$ Therefore, the optimal w_2^{hh} in C¹ actually satisfies SPC₂^{hh} when $\frac{\Phi_b}{b} < \Phi^{\pi}$ and does not satisfy SPC₂^{hh} when $\frac{\Phi_b}{b} > \Phi^{\pi}$. Next, define ($e_1^{\pi}; e_2^{\pi}; \Phi^{\pi\pi}$) such that,

$$\begin{array}{l} \textbf{s} \\ \textbf{s} \\ (p_{1}^{hh} + p_{1}^{hl})(\textbf{W}_{i} \ (\overset{1}{b}_{i} \ \textbf{C}^{\pi\pi})(\textbf{U} + c(e_{2}^{\pi}))_{i} \ (\overset{1}{b} + \textbf{C}^{\pi\pi})(\textbf{U} + c(e_{1}^{\pi}))) = (\overset{1}{b} + \textbf{C}^{\pi\pi})c^{\emptyset}(e_{1}^{\pi}) \\ \textbf{s} \\ (p_{2}^{lh}(\textbf{W}_{i} \ (\overset{1}{b}_{i} \ \textbf{C}^{\pi\pi})(\textbf{U} + c(e_{2}^{\pi}))_{i} \ (\overset{1}{b} + \textbf{C}^{\pi\pi})(\textbf{U} + c(e_{1}^{\pi}))) = (\overset{1}{b}_{i} \ \textbf{C}^{\pi\pi})c^{\emptyset}(e_{2}^{\pi}) \\ \textbf{s} \\ c^{\emptyset}(e_{1}^{\pi}) = (p_{1}^{hh} + p_{1}^{hl})(c^{\emptyset}(e_{2}^{\pi})_{i} \ c^{\emptyset}(e_{1}^{\pi})) \end{array}$$

For $\frac{c_b}{b} = C^{\pi\pi}$, $(e_1^{\pi}; e_2^{\pi})$ solves both C2 and C3. Next, de ne $H(e_1; e_2) = c^{(0)}(e_1)_i (p_1^{hh} + p_1^{hl})(c^{(0)}(e_2)_i c^{(0)}(e_1))$. Using the ICCs, it is possible to show that $\frac{dH}{d(c_b=b)} < 0$. For $\frac{c_b}{b} < C^{\pi\pi}$, H > 0 measured at the optimal C2 level of e®orts and C2 is the optimal contract. For $\frac{c_b}{b} > C^{\pi\pi}$, H < 0 measured at the optimal C2 level of e®orts and C3 is the optimal contract.

Finally, we show that $0 < \mathbb{C}^{\pi} < \mathbb{C}^{\pi\pi} < 1$. The ⁻rst inequality holds because $f_i 2(U + c(e^{\pi})) > 0$. The second inequality holds because C1 and C3 are mutually exclusive. The third inequality follows after simplifying the ICCs de⁻ning $(e_1^{\pi}; e_2^{\pi})$,

$$\frac{p_1^{hh} + p_1^{hl}}{p_2^{lh}} (1_i \ \Phi^{\mu\mu}) c^{0}(e_2^{\mu}) = (1 + \Phi^{\mu\mu}) c^{0}(e_1^{\mu})$$

implying,

$$\mathbb{C}^{^{^{^{_{a_{a}}}}}} = \frac{\frac{p_{1}^{hh} + p_{1}^{hi}}{p_{2}^{h}} c^{\emptyset}(e_{2}^{^{^{_{a}}}}) i c^{\emptyset}(e_{1}^{^{^{_{a}}}})}{\frac{p_{1}^{hh} + p_{1}^{hi}}{p_{2}^{h}} c^{\emptyset}(e_{2}^{^{^{_{a}}}}) i c^{\emptyset}(e_{1}^{^{^{_{a}}}})} < 1: \text{QED}$$

Proof of Proposition 3: The small agent is more likely to perform high than the large agent if

$$p^{hh} + p^{lh} > p^{hh} + p^{hl}$$
:

This is equivalent to $(e_2 i e_1)(^{(0)} + ^{-}) > 0$ which is always true.

The small agent earns more on average than the large agent if,

$$\frac{\mathsf{W}_1}{\mathsf{b}_1} < \frac{\mathsf{W}_2}{\mathsf{b}_2}$$

After reordering terms, this is equivalent to,

 $p^{lh}b_1(W_i \ (b_1 + b_2)(U + c(e_1)))_i \ (p^{hl} + p^{hh})b_2(W_i \ (b_1 + b_2)(U + c(e_2))) + p^{ll}b_1b_2(c(e_2)_i \ c(e_1)) \] 0: \\ e_2 \] e_1 \text{ implies that the above inequality always holds when } \\ \frac{b_1}{b_2} > \frac{p^{hl} + p^{hh}}{p^{hh}}: \text{ QED}$

Performance Measure	Name	Definition
ERT	Employment Rate at Termination	Fraction of terminees employed at termination
WERT	Welfare Employment Rate at Termination	Fraction of terminees receiving welfare at date of application who were employed at termination
CE	Cost per Employment	Training agency's year's expenditures on adults divided by the number of adults employed at termination
AWT	Average Wage at Termination	Average wage at termination for terminees who were employed at termination
ERTY	Youth Employment Rate at Termination	Fraction of youth terminees employed at termination
YPTR	Youth Positive Termination Rate	Fraction of youth terminees either placed in a job or satisfying an educational objective (see note below)
CEY	Youth Cost per Employment	Training agency's year's expenditures on youths divided by the number of youths positively terminated

TABLE 1National JTPA Performance Measures in Effect in 1985-86

1. The date of termination is the date the enrollee officially exits training. A terminee is an enrollee after he has officially exited training.

2. A positive termination is entering un-subsidized employment, attaining youth employment "competencies" (through course-work, training and/or tests in work maturity, basic education, or job-specific skills), entering non-JTPA training, returning to school full-time, or completing a major level of education.

		Mean O	utcome	Mean S	Standard		Excess mance	Percentage of Tra Centers Meeting Star	-
Performa	ance Measure	1985	1986	1985	1986	1985	1986	1985	1986 (19
ERT	Employment Rate at Termination	69.25	71.6	54.7	60.9	14.6	10.7	92	97
		(13.9)	(13.1)	(9.5)	(9.5)	(10.2)	(8.5)		
WERT	Welfare Employment Rate at	60.0	63.1	44.8	51.1	15.3	12.1	89	92
	Termination	(15.0)	(14.4)	(10.2)	(8.4)	(13.4)	(11.4)		
CE	Cost per Employment (\$)	3059.1	2923.6	4806.6	4419.2	1747.5	1495.6	96	95
		(1250.5)	(1190.1)	(1340.2)	(1071.4)	(1211.8)	(1108.1)		
AWT	Average Wage at Termination (\$)	4.9	5.0	4.5	4.6	.4	.4	89	86
		(.9)	(.9)	(.7)	(.8)	(0.5)	(.4)		
ERTY	Youth Employment Rate at	50.6	52.1	33.1	39.8	17.5	12.3	84	94
	Termination	(16.1)	(17.0)	(10.5)	(10.3)	(12.9)	(13.2)		
YPTR	Youth Positive Termination Rate	77.4	80.0	73.8	73.2	3.6	6.8	84	72
		(14.6)	(13.3)	(11.5)	(10.9)	(11.5)	(8.9)		
CEY	Youth Cost per Employment (\$)	2516.0	2403.2	3865.0	3773.4	1349.0	1370.2	94	90
		(1250.5)	(936.5)	(1188.6)	(962.8)	(1405.1)	(1017.7)		

TABLE 2JTPA Performance Outcomes in 1985-86

Notes:

1. Rate measures defined as percentages.

2. Standard deviations in parentheses.

3. 600 and 623 observations used in the calculations for 1985 and 1986 respectively.

4. Excess performance is the performance standard subtracted from the performance outcome (times minus 1 for the cost measures).

-	1985	1986
Budget (\$)	N.A.	2,337,773
		(3,044,874)
Award (\$)	178,091	119,715
	(258,098)	(146,108)
Award/Budget (%)	N.A.	7
		(7)

TABLE 3JTPA Mean Budgets and Awards in 1985-86

Notes:

1. Standard deviations in parentheses.

2. 419 and 384 observations used in the calculations for 1985 and 1986 respectively.

3. The unit of analysis in the computation of Award/Budget is the training agency, not the state.

TABLE 4Determinants of Agent AwardsDependent variable = agent award (\$)

	Obs. = 802	
-	Ι	II
Variable		
Constant	47513	7165
	(6.19)	(.68)
Budget (\$)	.04	.04
	(21.50)	(15.29)
Mean budget (\$)		.02
		(5.52)
$\overline{R^2}$.366	.389

Notes:

1. T stat in parentheses.

2. Mean budget is the state average budget.

TABLE 5 Determinants of Agent Awards Dependent Variable = Agent Award/Budget

	Ι		II	
Variable	Coef.	p val.	Coef.	p val.
Constant	0156	.7002	0178	.7215
Constant	(385)	.7002	(357)	.7215
ERT_{Λ}	.4673E-3	.4072	.2831E-3	.6312
	(.829)	.1072	(.480)	.0312
WERT Δ	.3828E-3	.3166	.2801E-3	.4798
	(1.002)	.5100	(.707)	.+790
CE_{Δ}	.1091E-4	.0017	.1189E-4	.0013
eL_{Δ}	(3.141)	.0017	(3.226)	.0015
AWT_{Δ}	.0469	.0000	.0560	.0000
	(5.453)	.0000	(6.241)	.0000
$ERTY_{\Delta}$.1182E-2	.0001	.1136E-2	.0002
	(4.020)	.0001	(3.735)	.0002
$YPTR_{\Lambda}$	2006E-4	.9577	.1880E-4	.9365
ΠΠΔ	(053)		(.046)	.,505
CEY_{Λ}	.4415E-6	.9036	.3165E-5	.4227
	(.121)	.9050	(.802)	.7227
\overline{ERT}_{Λ}	(.121)		.1681E-3	.9474
			(.066)	.)+/+
$\overline{WERT}_{\Lambda}$.2469E-2	.1056
$W L R I \Delta$			(1.618)	.1050
\overline{CE}_{Λ}			6361E-5	.5568
CL_{Δ}			(588)	.5500
\overline{AWT}_{Λ}			0888	.0041
			(-2.870)	.0041
$\overline{ERTY}_{\Lambda}$.3262E-3	.8273
			(.218)	.0275
$\overline{YPTR}_{\Lambda}$			0178	.7215
			(.831)	.7215
\overline{CEY}_{Λ}			2825E-4	.0102
			(-2.570)	.0102
$\overline{R^2}$.2558		.2775	

Notes:

1. Data from NCEP-SRI and the JTPA Annual Status Report. Obs. = 802.

2. I and II estimated using a one-way random effects model, i.e., with separate agent-specific disturbances. All regressions include state dummies, whose coefficient estimates are omitted. T statistics are in parentheses.

3. ERT_{Δ} , $WERT_{\Delta}$, AWT_{Δ} , $ERTY_{\Delta}$, and $YPTR_{\Delta}$ are defined as the agent's performance outcome minus the performance standard. CE_{Δ} and CEY_{Δ} are defined as -1 times the performance outcome minus the performance standard. Variables with bars are measures of state means.

TABLE 6

Panels A-C

Determinants of Agent Performance

Dependent Variable = Perf. outcome - standard¹

		Model I		1	Model II			
A.		Determinants of	Excess Adult	$\begin{array}{cccccccccccccccccccccccccccccccccccc$				
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val		
Constant	1.543	(.354)	.724	.892	(2.273)	.023		
В	2.25E-07	(1.295)	.195	2.41E-07	(1.357)	.175		
$(B-\overline{B})/\overline{B}$	848	(-1.712)	.087					
$(B-\overline{B})/\overline{B}\cdot\delta^{lt25}$				801	(-1.068)	.285		
$(B-\overline{B})/\overline{B}\cdot\delta^{2575}$				-1.018	(-1.687)	.092		
$(B-\overline{B})/\overline{B}\cdot\delta^{gt75}$				774	(-1.406)	.160		
Obs.		2011			1973			
$\frac{R^2}{2}$.1785		.1709				
В.	Deter	minants of Exce	ss Adult Welf	t Welfare Employment Rate: WERT $_{\Delta}$				
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val		
Constant	8.608	(1.185)	.236	10.434	(1.979)	.048		
В	2.48E-07	(1.063)	.293	2.92E-07	(1.214)	.225		
$(B-\overline{B})/\overline{B}$	933	(-1.363)	.167					
$(B-\overline{B})/\overline{B}\cdot\delta^{lt25}$				911	(903)	.367		
$(B-\overline{B})/\overline{B}\cdot\delta^{2575}$				1.429	(-1.750)	.080		
$(B-\overline{B})/\overline{B}\cdot\delta^{gt75}$				669	(891)	.373		
Obs.		1963			1946			
R^2		.1496			.1486			
<u>С.</u>	D	eterminants of E	Excess Adult (Cost per Employm	ent: CE_{Δ}			
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val		
Constant	1013.963	(1.765)	.078	899.683	(1.936)	.053		
В	-3.83E-05	(1.652)	.099	-2.93E-05	(1.241)	.031		
$(B-\overline{B})/\overline{B}$	100.230	(1.575)	.130					
$(B-\overline{B})/\overline{B}\cdot\delta^{lt25}$				135.426	(1.375)	.169		
$(B-\overline{B})/\overline{B}\cdot\delta^{2575}$				8.380	(.104)	.917		
$(B-\overline{B})/\overline{B}\cdot\delta^{gt75}$				134.729	(1.840)	.066		
Obs.		1950			1912			
R^2		.1691			.1654			

(Continued)

		Par	nels D-F				
		Model I]	Model II		
D.		Determinants	of Excess Ad	ult Wage Rate: A	WT_{Δ}		
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val	
Constant	.3431	(.709)	.478	.564	(1.590)	.112	
В	8.75E-09	(.556)	.578	1.26E-08	(.783)	.434	
$(B-\overline{B})/\overline{B}$	046	(-1.039)	.299				
$(B-\overline{B})/\overline{B}\cdot\delta^{lt25}$				037	(.542)	.588	
$(B-\overline{B})/\overline{B}\cdot\delta^{2575}$				074	(-1.356)	.175	
$(B-\overline{B})/\overline{B}\cdot\delta^{gt75}$				065	(-1.301)	.193	
Obs.		1989		1952			
$\underline{R^2}$.0919			.0824		
E.	1	Determinants of L	Excess Youth	Employment Rate.	$: ERTY_{\Delta}$		
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val	
Constant	7.220	(1.213)	.225	20.268	(3.773)	.000	
В	-4.07E-07	(-1.618)	.106	-3.60E-07	(-1.397)	.163	
$(B-\overline{B})/\overline{B}$.651	(.918)	.358				
$(B-\overline{B})/\overline{B}\cdot\delta^{lt25}$				1.010	(.971)	.332	
$(B-\overline{B})/\overline{B}\cdot\delta^{2575}$.229	(.265)	.791	
$(B-\overline{B})/\overline{B}\cdot\delta^{gt75}$.714	(.908)	.364	
Obs.		1918			1882		
$\frac{R^2}{2}$.1713			.1701		
F.	Dete	rminants of Exce	ess Youth Pos	itive Termination	Rate: $YPTR_{\Delta}$		
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val	
Constant	-1.130	(181)	.856	1.882	(.413)	.680	
В	2.07E-07	(1.012)	.311	1.89E-07	(.905)	.365	
$(B-\overline{B})/\overline{B}$	125	(217)	.828				
$(B-\overline{B})/\overline{B}\cdot\delta^{lt25}$.374	(.426)	.674	
$(B-\overline{B})/\overline{B}\cdot\delta^{2575}$.156	(.220)	.826	
$(B-\overline{B})/\overline{B}\cdot\delta^{gt75}$				463	(722)	.470	
Obs.		1952			1918		
R^2		.1352			.1245		

TABLE 6 (Continued)

Panels D-F

(Continued)

TABLE 6 (Continued)

Donal	\mathbf{C}
Panel	U

		Model I		Model II		
G.	De	eterminants of E	Excess Youth	Cost per Employm	ent: CEY_{Δ}	
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val
Constant	2270.172	(4.416)	.000	2351.686	(5.204)	.000
В	-2.55E-05	(-1.393)	.164	-1.95E-05	(-1.043)	.297
$(B-\overline{B})/\overline{B}$	107.199	(2.044)	.041			
$(B-\overline{B})/\overline{B}\cdot\delta^{lt25}$				211.751	(2.695)	.007
$(B-\overline{B})/\overline{B}\cdot\delta^{2575}$ $(B-\overline{B})/\overline{B}\cdot\delta^{gt75}$				57.396	(.889)	.374
$(B-\overline{B})/\overline{B}\cdot\delta^{gt75}$				89.866	(1.539)	.124
Obs.		1763			1735	
R^2		.1727			.1657	

Footnote:

¹ For the cost measures CE_{Δ} and CEY_{Δ} , the dependent variable is defined as performance standard - outcome. Notes:

1. Data are from NCEP-SRI and the JTPA Annual Status Report.

2. Models I and II are estimated using a one-way random effects model, i.e., with separate agent-specific disturbances. All regressions include state dummies, whose coefficient estimates are omitted.

3. The variables *B* and \overline{B} are the agent's budget and state mean budget, respectively. δ^{tr25} is an indicator variable, equal to one if the heterogeneity measure of the agent's state falls in the lower 25th percentile of the distribution of state heterogeneity outcomes, and equal to zero otherwise; δ^{2575} is equal to one if the heterogeneity measure falls between the 25th and 75th percentiles, and equal to zero otherwise; and δ^{gr75} is equal to one if the heterogeneity measure exceeds the 75th percentile, and equal to zero otherwise.

TABLE 7

Panels A-D

Determinants of State Performance

Dependent Variable = State Mean Excess Performance¹

		Model I			Model II			
A.	De	terminants of State	Mean Excess	cess Adult Employment Rate: \overline{ERT}_{Δ}				
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val		
Constant	11.692	(4.770)	.000	7.693	(3.016)	0.003		
Mean budget	1.88e-07	(1.110)	.267	5.94e-07	(3.108)	.002		
Budget heterogeneity				3.391	(.818)	.413		
Obs.		254			215			
R^2		.4076			.3585			
B.	Determi	nants of State Mea	n Excess Adult	t Welfare Employment Rate: \overline{WERT}_{Δ} Coef. (t stat) p v				
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val		
Constant	14.167	(3.874)	.000	8.265	(2.544)	.011		
Mean budget	-1.09e-07	(-0.412)	.680	9.18e-07	(3.640)	.000		
Budget heterogeneity				3.926	(.718)	.473		
Obs.		254		215				
R^2		.3692		.4285				
C.	Dete	rminants of State N	Aean Excess Ad	dult Cost per Emplo	yment: \overline{CE}_{Δ}			
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val		
Constant	2028.693	(4.926)	.000	1674.351	(4.028)	.000		
Mean budget	-3.38E-05	(-1.042)	.297	3.85E-05	(1.063)	.288		
Budget heterogeneity				-499.8737	(-0.637)	.524		
Obs.		254			215			
R^2		.3568			.3581			
D.		Determinants of St	ate Mean Exce	cess Adult Wage Rate: \overline{AWT}_{Δ}				
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val		
Constant	2.684	(19.989)	.000	.705	(5.008)	.000		
Mean budget	-1.18e-07	(-10.830)	.000	-2.79e-08	(-2.296)	.022		
Budget heterogeneity				403	(-1.529)	.126		
Obs.		254			215			
R^2		.6816			.4849			

(Continued)

		Pane	s E-G				
		Model I			Model II		
E.	Determinants of State Mean Excess Youth Employment Rate: \overline{ERTY}_{Δ}						
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val	
Constant	4.84	(1.336)	.181	14.777	(3.926)	.000	
Mean budget	8.43E-07	(3.482)	.000	1.22E-06	(4.290)	.000	
Budget heterogeneity				-15.089	(-2.440)	.015	
Obs.	254			215			
R^2	.3831			.3835			
F.	Determinants of State Mean Excess Youth Positive Termination Rate: \overline{YPTR}_{Δ}						
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val	
Constant	7.059	(2.459)	.014	.097	(0.035)	.972	
Mean budget	-2.30E-07	(-1.075)	.283	7.03E-07	(3.119)	.002	
Budget heterogeneity				326	(-0.067)	.947	
Obs.		254			215		
R^2	.4160			.3228			
G.	Determinants of State Mean Excess Youth Cost per Employment: \overline{CEY}_{Δ}						
Variable	Coef.	(t stat)	p val	Coef.	(t stat)	p val	
Constant	972.1862	(1.812)	.070	1757.322	(3.907)	.000	
Mean budget	2.19E-05	(0.501)	.617	7.32E-05	(1.797)	.072	
Budget heterogeneity				-1829.808	(-2.074)	.038	
Obs.	254			215			
R^2	.3657			.3595			

TABLE 7 (Continued)

Panels E-G

Footnote:

¹ The cost-related excess performance measures, CE_{Δ} and CEY_{Δ} , are weighted averages of the agent's performance standard minus the performance outcome. All other measures of excess performance are weighted averages of the agent's performance outcome minus the performance standard. The agent's weight is its share of its state's allocation. Notes:

1. Data are from NCEP-SRI and the JTPA Annual Status Report.

2. Models I and II are estimated using a one-way random effects model, i.e., with separate agent-specific disturbances. All regressions include state dummies, whose coefficient estimates are omitted.

3. Mean budget is the state average budget. Budget heterogeneity is a measure of variation in agent size within the state. See text for definition.