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*Research Article*

### **Family size and schooling throughout the Demographic Transition: Evidence from Brazil**

**Leticia Marteleto**

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## **Family size, adolescents' schooling and the Demographic Transition: Evidence from Brazil**

**Leticia Marteleto<sup>1</sup>**

### **Abstract**

The goal of this paper is to address whether and how the changing family sizes of cohorts born before and after the demographic transition are associated with increasing schooling in Brazil. Multivariate analyses of nationally representative data demonstrate that, although a higher proportion of adolescents born after the demographic transition live in smaller families, they also suffer a greater disadvantage from being in larger families than their peers born before the demographic transition. Additional case studies and comparative works are needed to disentangle the mechanisms behind the dynamic association between sibship size and adolescents' educational attainment found in this study.

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## 1. Introduction

Educational attainment is the result of the efforts of several institutions: individuals, families, schools, communities, and the state. A large body of research has examined each one of these dimensions in developing countries both separately and, to a smaller extent, simultaneously (Buchman and Hannum 2001). Social scientists have long been interested in the individual and family determinants of social stratification and education, emphasizing the role of parents' education and occupation in determining children's schooling (Blau and Duncan 1967; Shavit and Blossfeld 1993). Although parental education explains a large part of the variance in educational attainment, a significant portion of the variation also comes from additional family factors, such as family size.<sup>2</sup>

Past research has documented a negative association between family size and educational attainment in the West (for excellent reviews, see Lloyd 1994; Steelman et al. 2002). The *resource dilution* and *sibling rivalry* hypotheses offer frameworks to explain the nature of this negative relationship: a larger number of siblings translates into less capital (human, social, and cultural) for each child in the family. This negative association between family size and children's schooling has been challenged in both developed (Black, Devereux, and Salvanes 2005; Guo and VanWey 1999) and developing countries, where studies have found the association between family size and children's schooling to be positive, negative, and not statistically significant, depending on the period and context examined (Lu and Treiman 2008 for China; Maralani 2008 for Indonesia). The goal of this paper is to address whether and how the changing family sizes of cohorts born before and after the fertility transition are associated with children's schooling in Brazil. I examine the following specific questions: First, how important is sibship size for educational attainment; and, second, how has the association between sibship size and schooling changed across cohorts born before and after the demographic transition?

That sibship size varies substantially with the demographic transition and socioeconomic development—a cross-national phenomenon—suggests the need for case studies of how sibship size differentially relates to educational outcomes in a variety of demographic regimes and socioeconomic contexts. Brazil offers a particularly interesting case for examining this relationship that can advance our understanding of the process of educational attainment as it relates to family size in at least two important ways.

The first contribution of this paper is to introduce Brazil to the comparative template of research on how family size relates to educational attainment, thereby

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<sup>2</sup> I use the terms “sibship size” and “family size” interchangeably throughout this paper. Both refer to total number of siblings in the family.

adding variation to a literature that by now recognizes that this relationship is not homogenous across time or context (i.e., Lu and Treiman 2008). While a considerable body of research exists for developed countries, with a few exceptions (Anh et al. 1998; Lu and Treiman 2008; Maralani 2008; Parish and Willis 1993; Post and Pong 1998; Pong 1997), there is less coverage of empirical research that makes a serious attempt at identifying the association between siblings dynamics and educational outcomes outside the United States and Western Europe (Steelman et al. 2002).

The second contribution of this paper is to examine how the influence of family size on children's schooling varies in very different demographic and socioeconomic regimes within the same country, a necessary step toward the larger goal of understanding the patterns of this association. While past research has examined the relationship between family size and children's schooling, none has focused on differences between cohorts of children born before and after a demographic transition. Because demographic processes, socioeconomic development, intergenerational transmission of human capital, and opportunities jointly affect patterns of educational attainment, fertility decline may be a significant factor related to educational outcomes. Brazil's demographic transition has dramatically changed its family structure, particularly in regard to family size. For example, 14-year-olds born in 1963 had an average of 5.4 siblings, while their counterparts born in 1983 had 2.3 siblings. The focus on education is particularly relevant as the country has been characterized by persistent poor educational performance. Despite recent improvements, Brazil still has relatively low educational levels, particularly in comparison to other Latin American countries. For example, among adults 25 years and older, Brazilians averaged 4.9 years of schooling in 2000. In the same year, the average number of years of schooling was 8.8 for Argentines, 7.5 for Chileans, 7.2 for Mexicans, and 7.6 for Peruvians (UNESCO 2003). Brazil is also noteworthy to social scientists because the country has persistently had one of highest levels of social and economic inequality in the world, and education has been found to be the leading reason for these high levels of inequality (Barros and Lam 1996).

Understanding the determinants of children's educational attainment is a critical concern in developing countries such as Brazil, particularly because investments in the education of children and adolescents produce a skilled stock of human capital that helps eliminate constraints on national development. Knowing how the demographic transition and socioeconomic development affect investment in children's education may inform policies designed to promote universal school enrollment in secondary schooling and higher educational attainment and achievement; goals which the government of Brazil and of most other developing countries consider to be important.

## 2. Theoretical framework

The *sibling rivalry* and *resource dilution* theoretical frameworks offer insights into how the number of siblings is associated with children's outcomes. These theories posit that children with many siblings are generally worse off than their counterparts with fewer siblings in terms of several outcomes related to life chances and well-being, including nutrition, educational attainment, wealth, and mortality (Becker 1981; Blake 1981). According to the dilution of resources hypothesis (Blake 1985), a large number of siblings, or close spacing among siblings, dilutes the amount of capital (social, human, and cultural) parents invest in each child, which tends to negatively influence subsequent outcomes (Blau and Duncan 1967; Blake 1981). Simply put, parents with more children spend less time with each of them, thus making each of them less likely to develop tools to acquire greater capital than children who receive more support.

In contrast to the dilution of resources framework, which posits an equal allocation of resources among siblings, the sibling rivalry hypothesis assumes that parents invest in their children to maximize family utility, which often results in investment inequities within the family (Becker 1981). This hypothesis assumes that parents make human capital investments in their children based on assessments of their differential ability to contribute to the wealth of the entire family (Becker 1981). These parental investments improve the "quality" or the life prospects of children. However, even with inequities in resource allocation among children, the interaction between the quantity and quality of children means that education per child tends to be lower in families with more children. As a result, given a set pool of family resources, having more children means that fewer resources are available for each of the children, and that the children's life prospects are therefore reduced (Becker 1981).

The confluence theory also offers insights into how family size and children's outcomes are related: the presence of additional children lowers a family's average intellectual environment (Zajonc and Markus 1975). Children are born into intellectual environments that affect intelligence. First-born children are, for example, born into adult-only environments, and have the opportunity to tutor younger siblings, which is likely associated with higher IQ levels. In addition to family size, the confluence model emphasizes the role of birth order and birth spacing. Despite the persuasiveness of the model's arguments, early empirical analyses have not found much support for this theory (Hauser and Sewell 1985), although recent research has shown that birth order is an important factor in educational outcomes (Black, Devereux, and Salvanes 2005; Guo and Van Wey 1999), including IQ (Black, Devereux and Salvanes 2010).

Much of the empirical research on the U.S. and Western Europe has found the association between family size and several outcomes—from wealth accumulation (Keister 2003) to educational outcomes (Blake 1985)—to be negative. Additional

research has focused on age spacing (Powell and Steelman 1993), sex composition (Conley 2000; Powell and Steelman 1990), and the ordinal position of children in the family (Stelman et al. 2002). The magnitude of the effect of family size on educational outcomes seems to differ according to the type of educational outcome examined—school enrollment, educational attainment, educational achievement—but the negative sign of the relationship persists in these studies. Guo and VanWey (1999) have challenged the uniformity of the negative relationship between family size and intellectual development. When they implement change models to control for unobserved family factors, the effects of sibship size disappear. At the same time, using data from Norway and accounting for the endogeneity issues between family size and schooling, a recent study has shown that there is no causal effect of sibship size on educational attainment, and has argued that the important aspect to consider is birth order (Black, Devereux, and Salvanes 2005).

Less is known about the relationship between family size and education in developing countries than in the United States and Western Europe. Nonetheless, most of the empirical work in developing countries that has examined the associations between family size and schooling attainment confirms the negative association: children from larger families exhibit educational disadvantages compared to children from smaller families (Anh et al. 1998; Knodel and Wongsith 1991; Maralani 2008; Parish and Willis 1993; Patrinos and Psacharopoulos 1997; Pong 1997; Post and Pong 1998; Shavit and Pierce 1991). Past research has argued that the strength of the relationship varies with institutional and socioeconomic development (Knodel, Havanon, and Sittitrai 1990; Lu and Treiman 2008 for China; Maralani 2008 for Indonesia). Additional studies have yielded conflicting results, reporting negative but not statistically significant effects (Lloyd 1994), mixed results (Psacharopoulos and Arriagada 1989), or even a positive association (Chernichovsky 1985).

Among these studies, only a few have examined change over time: Post and Pong 1998 have found decreasing effects of sibship size on educational attainment in Hong Kong, while Pong (1997) has found increases in the negative impact of sibship size on secondary school attainment in Malaysia over a period of rising government regulation of school supply and employment opportunities (Pong 1997). Parish and Willis (1993) found an increasing effect of family size on educational attainment in Taiwan, despite declining average family size and increasing average income. They attributed this unexpected finding to the rising opportunity cost of school in recent decades. Most recently, Maralani (2008) found that the relationship between family size and children's schooling ranges from positive to negative in Indonesia, depending on the level of socioeconomic development.

These recent findings, particularly those on developing countries, raise questions as to whether the negative association between sibship size and educational attainment

holds throughout different periods and contexts, particularly for periods associated with marked decreases in family size. Socioeconomic and demographic trends shape the macro conditions in which the educational attainment process takes place. It is likely that the association between schooling and sibship size differs according to the demographic and social contexts inherent in high and low fertility regimes.

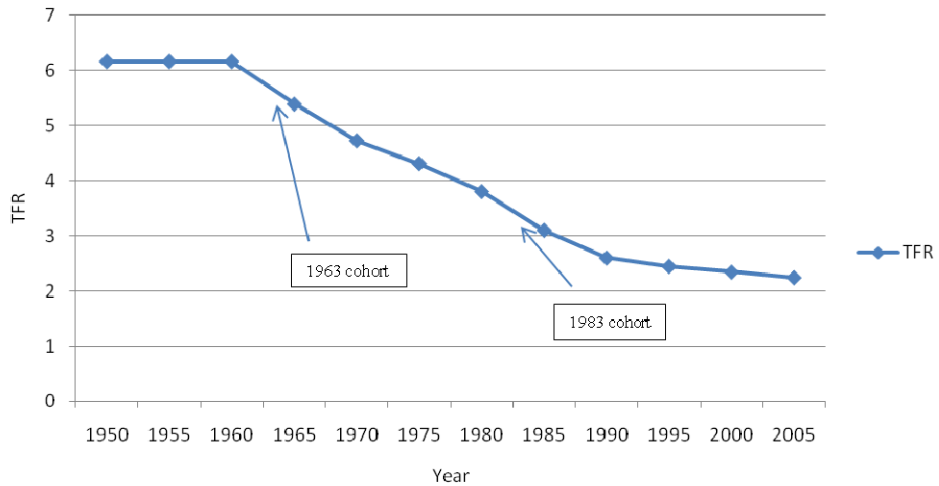
### **3. The Demographic Transition in Brazil**

Brazil's demographic transition is fairly typical of countries across the developing world, and is documented with excellent census data. As shown in Figure 1, the total fertility rate (TFR) began a rapid decline in the 1960s, falling to about 2.3 by 2000. The fertility decline occurred during a period of rapid social change that included periods of both economic growth and economic crisis (Martine 1996, Lam and Duryea 1999). There was a large degree of regional variation, with the fertility decline starting later in the poorer north and northeast regions than it did in the higher-income south and southeast. Brazil's rapid population growth in the 1960s and 1970s was also fairly typical.

The demographic transition in Brazil has been attributed to several factors, including ideational change accompanied by material change (Potter, Schmertmann, and Cavenaghi 2002); a dramatic increase in the use of contraception, including female sterilization, despite the complete lack of government-sponsored family planning programs (Martine 1996); and an increase in schooling (Lam and Duryea 1999). It is difficult to pinpoint a single reason for the fertility decline, as the processes of development, the spread of new ideas and values, increases in schooling, and higher levels of female labor force participation and of contraceptive use often take place simultaneously. Over the same period, Brazil underwent important political changes, with the end of a military coup in 1984 and the installation of a democratic regime and a new constitution in 1988.

The changes in fertility and mortality typical of the demographic transition led to substantial changes in family size. The distribution of family sizes within the Brazilian population has changed considerably: 15.29% of 14-year-olds born in 1963 came from families with one or two children, while 45.11% came from families of seven children or more. In contrast, among 14-year-olds born in 1983—after the bulk of the decline in fertility—50.94% came from families with one or two children, and only 11.86% came from families with seven or more children. Thus, in recent decades, growing proportions of Brazilian children have come from smaller families.



**Figure 1: Total Fertility Rate: Brazil, 1950-2005**

Lam and Marteleto (2008) have extensively examined the impact of the demographic transition on cohort size and family size in Brazil. They find that Brazil was in a stage of both increasing family size and cohort size between 1960 and 1970, even though fertility had already begun to decline. By 1980, Brazil was characterized by falling numbers of siblings and rising cohort size. With the largest cohort born in 1982, 1992 marks the year in which there is both a decline in the absolute number of 9-11-year-olds in the population and a (continuing) decline in the average number of siblings (Lam and Marteleto 2008). A recognition of this trend is important for this study, as an awareness of these developments may assist us in identifying the stages of the demographic transition in which the association between family size and educational attainment might vary. Children born between 1960 and 1970 experienced increasing family sizes, while children born after 1980 experienced a declining number of siblings. For this reason, I view birth cohort as an important component of this study (Easterlin 1980). Comparing birth cohorts allows us to address the question of whether the advantageous demographic conditions for children born after the demographic transition have contributed to improvements in children's schooling through smaller family sizes.

In light of these important changes, this paper addresses the following questions: First, how important is sibship size for educational attainment in Brazil; and, second, how has the association between sibship size and schooling changed across cohorts

born before and after the demographic transition? Few studies have examined the association between family size and educational attainment over time, and, to my knowledge, none has looked at cohorts born before and after the demographic transition in Brazil. Especially in countries like Brazil, where secondary school enrollment is not yet universal and educational attainment levels are far from satisfactory, understanding how family characteristics other than family wealth relate to education is essential in any broad effort to increase the level and quality of education.

#### **4. Data and analytical sample**

In this paper, I use data from the 1977 and 1997 *Pesquisa Nacional por Amostra de Domicílios* (PNAD), (National Research of Household Sample), which are annual household surveys conducted by the *Instituto Brasileiro de Geografia e Estatística* (IBGE), the Brazilian statistical bureau. The PNAD is a nationally representative survey with information on 498,679 individuals in 100,039 households for 1977, and on 365,870 individuals in 89,939 households for 1997. The samples are sufficiently large to permit sub-sampling of specific groups, such as of 14-year-olds. The 1977 PNAD provides data on 12,834 14-year-olds born in 1963, while the 1997 PNAD provides data on 7,861 14-year-olds born in 1983. Children's educational experiences at different ages are sufficiently diverse that it makes sense to analyze them separately. Fourteen-year-olds were chosen as the unit of analysis because this is the highest age at which school attendance is legally required in Brazil. Moreover, children who have been successful in school should be making the transition from primary to secondary education at 14 years of age.

The PNAD contains standard demographic and socioeconomic variables, such as sex, age, income, and schooling, for all members of the household. Data from 1977 and 1997 are comparable, although the 1977 PNAD does not contain information on race and ethnicity, making it impossible to compare cohort racial distributions. Moreover, neither the 1977 nor the 1997 PNAD covered the rural part of the northern region, which probably results in overestimates of the level of education in that region.

Because the PNAD is a household survey, it accounts for all family and non-family members who live in the household, but it cannot account for family members outside the household. Thus, PNAD data do not permit a detailed analysis of non-resident family members, particularly siblings, who may influence children's school enrollment and schooling. For the purposes of this investigation, however, it is fortunate that both the 1977 and 1997 PNAD provide the *total number of children born* to all women over the age of 15, and thus the total number of resident and non-resident siblings for the sampled 14-year-olds. Because these years of the PNAD survey include

information about the total number of children, the data for these years are more suitable than most household survey data used for this type of analysis (such as Conley and Glauber 2006 in the US and Li et al. 2008 in China), which often count only the children living in the household. However, the structure of the data does not allow for an analysis of the age or sex composition of all siblings, only of those living in the household.<sup>3</sup> It is important to note that the PNAD data does not have information on union history or on the support of relatives for children's education.

To accurately include the mother's education in the models, I restrict the sample of 14-year-olds to children of the head of the household. This selection does not create a selection bias. Most 14-year-olds live with at least one parent (87.81% in 1977 and 90.71% in 1997), which allows for the use of information on the mother's education, an important predictor of the child's educational level. Children of the head of the household are not significantly different from the full sample of children on their distribution across rural/urban location, family income, region, or sex (not shown). Moreover, there are no notable differences in levels of educational attainment among the full and analytical samples.

## **5. Methods**

In order to assess whether and how the number of siblings is related to educational attainment in pre- and post-demographic-transition Brazil, I estimate models of years of schooling for 14-year-olds born in 1963 and 1983. I model complete years of schooling by estimating ordinary least squares regressions. The number of siblings is a set of dummy variables, while the mother's education is coded as a continuous variable. I include a flag variable for whether the mother's education is missing. The log of family income was corrected to take account of currency changes and inflation across the period of study by using IPEA deflators, with 1997 as the base year (Corseuil and Foguel 2002). The omitted categories are male-headed household, rural and the state of São Paulo. The reference group is São Paulo because the state has one of the highest

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<sup>3</sup> One reviewer has suggested that I examine the differential associations between schooling and siblings living inside versus outside the household. The 1977 PNAD has information on the total number of children living inside and outside the household, and I have run models including these variables separately for this year (not shown). Although both coefficients are negative, the coefficient for the number of children inside the family is much larger, as expected. While this is an interesting finding, only a very small proportion of the adolescents in my sample have siblings living outside the household. Because of this fact, and also because this variable is not available in 1997, I decided not to include these models with number of siblings living inside versus outside the household. Indeed, it is not common for children to leave the household before marriage in Brazil, although this is likely to be changing in the metropolitan areas. Boarding schools are not common at all, and schools are available throughout the country.

levels of educational attainment for adolescents in both cohorts. Including a fixed-effect dummy variable for the birth state of each 14-year-old in the analysis (except for the reference state, São Paulo) reduces the threat of omitted variables bias. These birth state fixed effects represent the influence of characteristics we do not know or cannot measure about each birth state, such as the labor market conditions in the birth state when the adolescents were born. The number of siblings was coded as a set of dummy variables from zero to seven or more, and zero is the omitted category.

A caveat should be made at this juncture. It is possible that parents may make joint decisions about the quantity of children they wish to have in response to a preference for a higher "quality" of children with regard to education (Becker 1981). Family size may be correlated with unmeasured determinants of children's schooling, as parents may choose a combination of low fertility and high levels of schooling for their children. This contributes to a negative correlation between family size and schooling that tends to overstate the extent to which schooling would increase if parents were required to limit their number of children. Few studies using data from developing countries have attempted to account for such endogeneity by applying twins and same-sex methods as instrumental variables (Black, Devereux, and Salvanes 2005; Conley and Glauber 2006). Results from a recent study in Indonesia indicate that the negative relationship between family size and children's schooling is not sensitive to assumptions about the exogeneity of fertility (Maralani 2008). Although the conventional approach and an instrumental variable approach have yielded similar results in Indonesia, the results of this study should be interpreted with caution to avoid any implication of direct causality. Further research that takes the endogeneity between fertility and children's schooling into account is needed to examine whether this is the case for other countries.

## 6. Results

Table 1 provides comparisons of the distribution of 1963 and 1983 cohorts across socioeconomic and family characteristics. The general life conditions of these cohorts of adolescents differ in several ways. Although the majority of the adolescents in both cohorts live in the Southeast and Northeast combined, 66% of the adolescents in the older cohort lived in urban areas, while 80% of the adolescents in the younger cohort lived in urban areas. Brazil's increased urbanization across the 1970s and 1980s may imply changes in the overall value of children, and, consequently in their educational outcomes. Although the mother's education and the mother's age are coded as continuous variables in the regression models, I use a categorical version in Tables 1 and 2 because it provides more meaningful descriptive statistics. The distribution of

children by the mother's education has also changed dramatically across cohorts: while 36.12% of the adolescents in the older cohort had mothers with no formal education, 20.60% of the adolescents in the younger cohort had mothers with no formal education. While 44.85% of the 14-year-olds in the older cohort had seven or more siblings, only 11.86% of the 14-year-olds in the younger cohort had seven or more siblings. The consequences of the profound demographic changes that occurred in Brazil during the period that separates this study's cohorts are evident in the average number of siblings these 14-year-olds have. The average total number of siblings decreased from 6.13 in the older cohort to 3.12 in the younger cohort. Figure 2 illustrates the considerable change in the distribution of number of siblings for adolescents in the 1963 and 1983 cohorts.

**Figure 2: Cumulative proportion of 14-year-olds with X siblings or more: Brazil, 1963 and 1983 cohorts**

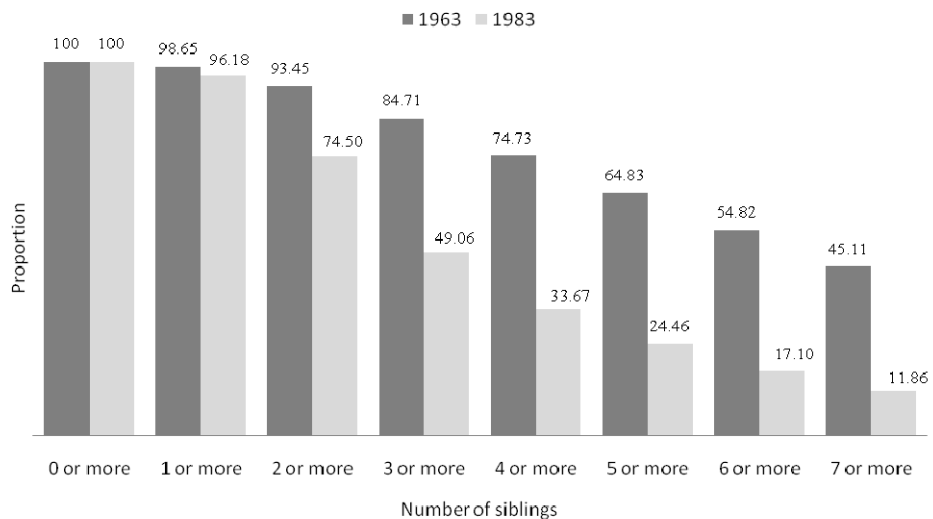


Table 2 shows complete years of schooling by family and socioeconomic characteristics and by cohort. The educational attainment of young people has increased dramatically in Brazil over the last 20 years: mean schooling grew from 3.43 for the older cohort to 4.70 for the younger cohort. Table 2 also provides evidence for gender differences in levels of educational attainment across and within cohorts. The schooling advantage for girls increases across cohorts: among adolescents in the older cohort, boys have 3.19 years and girls have 3.67 years of schooling; while in the younger

cohort, boys have 4.37 years and girls have 5.04 years of schooling. These gender differences favoring girls are not surprising. They are markedly different from findings in other developing countries (Knodel and Jones 1996), but are representative of the general trend in other Latin American countries. Table 2 also shows that the schooling levels are slightly higher among adolescents in male-headed households than among those in female-headed households in both cohorts.

**Table 1: Characteristics of 14-Year-Olds [%], 1963 and 1983 Cohorts, Brazil**

	Cohort of 1963	Cohort of 1983
<b>Household Headship</b>		
Female	11.94	19.73
Male	88.06	80.27
<b>Number of Siblings</b>		
0	1.43	3.82
1	5.33	21.68
2	8.85	25.44
3	9.95	15.39
4	9.91	9.21
5	10.01	7.36
6	9.69	5.24
7+	44.85	11.86
<b>Rural/Urban Residence</b>		
Urban	66.21	80.20
Rural	33.79	19.80
<b>Sex</b>		
Male	50.65	50.99
Female	49.35	49.01
<b>Mother's Education</b>		
No Education (0)	36.12	20.60
Attended First Primary (1-4)	43.97	36.49
Attended Second Primary (5-8)	12.81	23.01
Attended High School (9-11)	4.51	12.73
Attended University or more (12+)	2.59	7.17
<b>Household Size</b>	3.70	1.54
<b>Mother's Age</b>	41.63	40.97
<b>Family Income</b>	1,106	877
<b>[N]</b>	11,082	6,971

Source: From author's calculations of PNAD 1977 and 1997 data.

Note: The states of Tocantins and Roraima were created after 1977. They were coded as part of their original states in 1977 for comparison reasons. Weighted data. Household size is net of siblings.

**Table 2: Complete years of schooling by characteristics of 14-year-old children of the head of the family, 1963 and 1983 cohorts, Brazil**

	Mean Years of Schooling	
	Cohort of 1963	Cohort of 1983
<b>National average</b>	3.43	4.70
<b>Family Headship</b>		
Female	3.35	4.48
Male	3.43	4.75
<b>Total Number of Siblings</b>		
0	4.48	5.49
1	5.17	5.78
2	4.97	5.34
3	4.55	4.59
4	4.01	4.15
5	3.52	3.68
6	3.19	3.44
7+	2.56	3.09
<b>Rural/Urban Residence</b>		
Urban	4.16	5.00
Rural	1.98	3.48
<b>Sex</b>		
Male	3.19	4.37
Female	3.67	5.04
<b>Mother's Education</b>		
No Education (0)	2.08	3.08
Attended First Primary (1-4)	3.76	4.46
Attended Second Primary (5-8)	4.80	5.19
Attended High School (9-11)	5.52	6.06
Attended University or more	6.00	6.60
<b>Family Income (Quintiles)</b>		
First Quintile	1.86	3.24
Second Quintile	2.47	3.98
Third Quintile	3.28	4.74
Fourth Quintile	4.02	5.52
Fifth Quintile	5.09	6.20
<b>Mother's Age</b>		
< 40	3.50	4.80
40-50	3.45	4.78
>50	3.25	4.21
<b>[N]</b>	11,082	6,971

Source: From author calculations of PNAD 1977 and 1997 data. Weighted data.

Table 3 shows coefficients and standard errors of OLS regressions of complete years of schooling separately by cohort. There are several noteworthy results that can be seen in Table 3. First, girls have an educational advantage over boys in both cohorts, and this advantage increases among adolescents in the younger cohort. This is in contrast to some Asian and African countries, where girls are at an educational disadvantage relative to boys. However, an educational advantage for girls is fairly typical of Latin American countries, where, in general, girls are found to have higher levels of school enrollment rates, as well as higher educational attainment and educational achievement levels. Another interesting finding is the slight decrease in the impact of the mother's education on children's schooling—from 0.202 to 0.130—with a statistically significant difference of coefficients at the 0.05 level. Not surprisingly, adolescents in both cohorts living in families with higher levels of income also have higher levels of schooling. Family structure is another important factor related to children's schooling. Adolescents in families headed by women are at a greater educational disadvantage relative to their counterparts in male-headed families, and this disadvantage is higher for the younger than for the older cohort. Similarly, those adolescents in larger households, net of the number of siblings, are at a disadvantage when compared with those in smaller households. This disadvantage has increased across cohorts—from a coefficient of -0.018 to a coefficient of -0.098—although the difference is not statistically significant. Finally, older-cohort adolescents in urban areas had approximately 1.062 year more of schooling than their counterparts living in rural areas, controlling for all other covariates. The schooling disadvantage associated with living in rural areas has decreased, and was associated with 0.418 year less of schooling in 1997.

The coefficients of number of siblings in Table 3 are also plotted in Figure 3 for ease of visualization of the cohort trend. Figure 3 shows that adolescents with one or two siblings have higher levels of schooling than those with no sibling. This is an interesting finding that reveals that family size starts to have a negative impact children's schooling after three siblings. For example, older-cohort adolescents with five siblings have 0.290 year less of schooling than their peers who do not have siblings. Similarly, younger-cohort adolescents with five siblings have 0.662 year less of schooling than their counterparts who are only children.



**Table 3: Estimates of OLS Regressions of Years of Schooling, 14-Year-Olds, 1963 and 1983 Cohorts, Brazil**

	Cohort of 1963		Cohort of 1983	
	Coefficient	SE	Coefficient	SE
<b>Sex</b> (omitted: male)	0.437***	0.032	0.637***	0.039
<b>Number of Siblings</b> (omitted: 0)				
1	0.253	0.147	0.359***	0.103
2	0.259*	0.139	0.274***	0.102
3	0.074	0.138	0.063	0.107
4	-0.079	0.138	-0.286**	0.116
5	-0.292*	0.138	-0.662***	0.121
6	-0.385**	0.138	-0.738***	0.130
7 or more	-0.639***	0.131	-0.948***	0.117
<b>Mother's Education</b>	0.202***	0.007	0.130***	0.006
<b>Mother's Age</b>	0.019***	0.003	0.013***	0.003
<b>Family Headship</b> (omitted: male)	-0.027	0.054	-0.404***	0.052
<b>Household Size</b>	-0.018**	0.009	-0.098***	0.015
<b>Log Family Income</b>	0.413***	0.021	0.326***	0.026
<b>Urban vs. rural</b> (omitted: rural)	1.062***	0.040	0.418***	0.055
<b>Birth state</b> (omitted: São Paulo)				
Rondônia	-1.146**	0.485	-0.382	0.326
Acre	-1.664***	0.434	-0.993**	0.397
Amazonas	-1.360***	0.167	-0.987***	0.155
Roraima	0.353	0.748	-0.712	0.598
Pará	-0.997***	0.113	-1.166***	0.115
Amapá	-1.034**	0.466	-0.843**	0.376
Maranhão	-1.334***	0.106	-1.135***	0.141
Piauí	-1.154***	0.121	-1.426***	0.156
Ceará	-1.497***	0.085	-0.841***	0.099
Rio Grande do Norte	-1.027***	0.124	-0.571***	0.158
Paraíba	-1.271***	0.108	-1.128***	0.146
Pernambuco	-1.024***	0.077	-0.892***	0.100
Alagoas	-1.421***	0.131	-1.595***	0.166
Sergipe	-1.370***	0.166	-1.416***	0.191
Bahia	-1.280***	0.076	-1.097***	0.085

**Table 3: (Continued)**

	Cohort of 1963		Cohort of 1983	
	Coefficient	SE	Coefficient	SE
<b>Birth state</b> (omitted: São Paulo)				
Minas Gerais	-0.414***	0.064	-0.327***	0.082
Espírito Santo	-0.115	0.126	-0.178	0.163
Rio de Janeiro	-0.696***	0.073	-0.556***	0.094
Paraná	-0.311***	0.086	0.083	0.096
Santa Catarina	0.539***	0.114	-0.042	0.132
Rio Grande do Sul	0.146*	0.078	-0.018	0.091
Mato Grosso	-0.676***	0.137	-0.151	0.122
Goiás	-0.835***	0.101	-0.643***	0.102
Distrito Federal	0.001	0.113	-0.385***	0.140
Outside of Brazil	-0.775*	0.466	-1.096	0.705
<b>Constant</b>	-1.804***	0.253	1.362***	0.217
<b>[N]</b>	10,684		6,521	

Source: From author calculations of PNAD 1977 and 1997 data. Robust standard errors in parentheses \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Also controlling for flag for whether mother's education is missing. Weighted.

**Figure 3: Coefficients of Number of Siblings by 14-Year-Old Cohort: Brazil, 1963 and 1983 Cohorts**

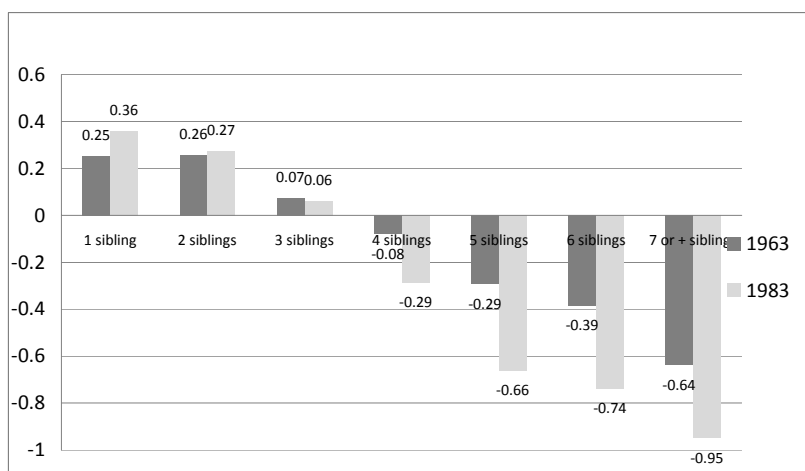


Figure 3 also shows that the educational disadvantages associated with family size have increased for the younger cohort. The figure makes clear that the coefficients of the number of siblings have changed considerably from the older to the younger cohort at larger parities, placing an additional burden on the schooling of younger-cohort children in larger families. For example, while older-cohort adolescents with six siblings had 0.440 year less of schooling than their only-child peers, younger-cohort adolescents had 0.738 year less of schooling than their only-child peers. I test for whether the difference in coefficients across cohorts is statistically significant in a pooled model in which I interact all variables with each cohort (not shown). I find that the difference in the coefficients of the number of siblings of the older versus the younger cohorts is statistically significant at higher-parity families (four siblings:  $p=0.0391$ ; five siblings:  $p=0.0205$ ; six siblings:  $p=0.0358$ ; seven or more siblings:  $p=0.000$ ). Adolescents in larger families in the younger cohort are at a greater educational disadvantage than adolescents in larger families in the older cohort.

Figure 4 shows adjusted and unadjusted mean years of schooling by the number of siblings for both younger- and older-cohort adolescents. I calculated the adjusted schooling using the estimates in Table 3. The baseline used for simulating the impact is a 14-year-old with all covariates to their mean values and varying number of siblings. When interpreting the unadjusted and adjusted years of schooling, it is worth noting that, if a 14-year-old had started school at the mandatory age of seven, and did not drop out or repeat a grade, she/he would have six or seven years of schooling. Figure 4 shows that 14-year-olds in both cohorts and in all family sizes are substantially behind this ideal level of educational attainment.

It is clear from the unadjusted lines in Figure 4 that, in both cohorts, adolescents with higher numbers of siblings have greater schooling disadvantages than their counterparts who have fewer siblings. For example, adolescents with six siblings in the older cohort had 3.2 years of schooling (calculated with a 95% confidence interval of 3.12 to 3.27), while their counterparts with two siblings had 5.0 years of schooling (calculated with a 95% confidence interval of 4.91 to 5.10). We see similar disadvantages among adolescents in the younger cohort: those with six siblings had 3.4 years of schooling (calculated with a 95% confidence interval of 3.33 to 3.52), while those with two siblings had 5.3 years of schooling (calculated with a 95% confidence interval of 5.30 to 5.40).

**Figure 4: Adjusted and unadjusted educational attainment of 14-year-olds by number of siblings: Brazil, 1963 and 1983 cohorts**

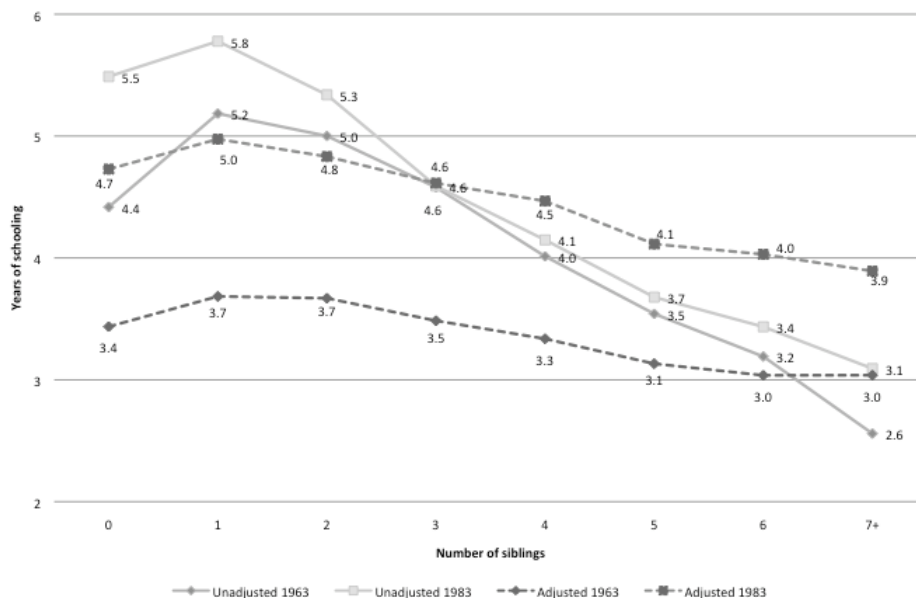


Figure 4 shows that the differences in schooling due to family size are smaller in the adjusted than in the unadjusted figure. This suggests that a large portion of the schooling differences by family size is in fact due to factors associated with family size included in the regression models. Figure 4 also shows that adjusted schooling has increased more than unadjusted schooling for children at higher parities. This implies that the schooling difference due to the other covariates is smaller among children in higher parities, while the schooling differences due to the negative impact of having larger family sizes has increased. While there are fewer children in larger families among younger-cohort adolescents, those who are in larger families are at a greater disadvantage than their peers in smaller families than among older-cohort adolescents. The figure for the adjusted schooling confirms that the association between family size and schooling has increased across cohorts. When viewed together, the adjusted and unadjusted lines in Figure 4 suggest a cohort change in how the number of siblings and educational attainment are related at higher parities, with the gap in schooling between those in high- and low-parity families becoming larger for adolescents in the younger than in the older cohorts.

Taken together, the findings from this paper suggest that the role of declining family size across cohorts born before and after the demographic transition in Brazil is important even when compared to the role of other covariates of educational attainment, such as the mother's education and family income. Although more adolescents in younger cohorts are in smaller families, the few who are in larger families should be the target of policies aimed at increasing the educational level of Brazilian adolescents.

## 7. Conclusions and discussion

The goal of this paper was to examine how the changing family sizes of cohorts born before and after the fertility transition are associated with increasing schooling in these two very different demographic regimes. Family size varies substantially throughout the demographic transition, suggesting the need for case studies of the association between family size and educational attainment in a variety of demographic regimes.

Findings from this study show that family size is an important factor associated with adolescents' schooling for cohorts of Brazilian adolescents born both before and after the demographic transition. Results of this analysis indicate that both the *association* between sibship size and schooling has become more negative at higher parities, even as the number of adolescents across larger families has declined. Although a higher proportion of younger-cohort adolescents are in smaller families, they also suffer a greater disadvantage from being in larger families than older-cohort adolescents. The demographic transition has benefited adolescents by producing a higher proportion of smaller families, but that advantage has been partly offset by the larger negative association between sibship size and adolescents' schooling. Recent cohorts of children in larger families should be the target of policies designed to improve the educational outcomes of youth in developing countries.

We should, however, exercise caution when using the results of this study to imply causality between family size and schooling. As discussed in the methods section, the conventional methodology used in this paper relies on assumptions of exogeneity about the relationship between family size and parents' preferences for children's schooling. Although other alternative analyses also require their own set of strong assumptions, it would be useful to compare results from conventional analysis, instrumental variable analysis, and twin data analysis using the same data to corroborate whether results are sensitive to assumptions about the exogeneity of fertility.

This study's findings show that sibship size is negatively associated with adolescents' schooling, and that this negative association has increased for adolescents born after the demographic transition. Although it is beyond the scope of this paper (and the data is limited in that regard), further research is needed to examine whether

this negative effect varies by siblings' age and sex composition in developing countries. The negative association between larger family sizes and schooling may be offset or even reversed by the support given to younger children by their older siblings, or by other family members in the household. The presence of others in the household might mediate the effect of sibship size, and, where the necessary data are available, should be the focus of future research on how family structure matters for children's schooling. Another potentially interesting extension of research in this area would be to investigate why only children have not fared consistently better than children with one sibling.

This study's findings show that the association between family size and schooling became stronger in Brazil after the demographic transition, and therefore lend support to the idea that this association varies across demographic regimes within the same context. Whether the same pattern found in this study holds in other contexts in which the demographic transition occurs at different times—for example, in some African countries that have not yet been through the demographic transition, or in countries that underwent the demographic transition a long time ago—remains an open question. Rather than seeking uniformity, I argue that more case studies and comparative studies are needed to disentangle the mechanisms behind the dynamic association between family size and educational attainment, particularly studies that seek to address the potential endogeneity between family size and educational attainment.

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