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**Multiple-interval timing in rats: Performance on two-valued
mixed fixed-interval schedules**

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Suggested running head: Mixed FI performance in rats

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Abstract

Three experiments studied timing in rats on two-valued mixed-FI schedules, with equally probable components, FI *S* and FI *L*. When the *L/S* ratio was greater than 4, two distinct response peaks appeared close to FI *S* and *L*, and data could be well-fitted by the sum of 2 Gaussian curves. When the *L/S* ratio was less than 4, only one response peak was visible, but nonlinear regression usually identified separate sources of behavioral control, by FI *S* and FI *L*, although control by FI *L* dominated. The data were used to test ideas derived from Scalar Expectancy Theory (SET), the Behavioral Theory of Timing (BeT), and Learning to Time (LeT).

For more than 60 years, since its development by Skinner (1938), the fixed-interval (FI) schedule of reinforcement has been used in the study of temporal control in animals. On a simple FI schedule, the first response occurring t s after the previous reinforcer delivery is itself reinforced, and this reinforcer delivery starts the next interval of the schedule. After a period of training, animals exhibit temporal control of responding, in the sense that their responding varies markedly as a function of elapsed time in the interval. Reinforcer delivery initiates a post-reinforcement pause (Lowe, Harzem, & Spencer, 1979), and when the pause finishes, subsequent responding either gradually accelerates to reach its maximum rate just before the next reinforcer delivery (Dews, 1978), or proceeds at a high steady rate until reinforcer delivery (Schneider, 1969).

When data from a number of intervals of FI are aggregated together, average response rate increases as a function of elapsed time in the interval in almost all animal species tested (Lejeune & Wearden, 1991), including mice, fish, and turtles, as well as the more conventional rats and pigeons. This response rate increase usually takes the form of the left half of a Gaussian curve with the peak located at the FI value (Lejeune & Wearden, 1991).

Although originally developed within the behavior-analytic framework of emphasizing relations between observed behavior and imposed environmental constraints rather than analysis in terms of internal processes, FI schedules and their variants like the peak-interval procedure (Roberts, 1981; Church, Meck, & Gibbon, 1994) have recently been used extensively to test theories of timing like *Scalar Expectancy Theory* (SET: Gibbon, Church, & Meck, 1984), the *Behavioral Theory of Timing* (BeT: Killeen & Fetterman, 1988), and an elaboration of ideas similar to those of BeT by Machado (1997), an account subsequently called the *Learning to Time* model (LeT: Machado & Guilhardi, 2000).

One example of the application of ideas derived from SET to simple FI schedules comes from Lejeune and Wearden (1991). As mentioned above,

they analyzed data from different species on FI schedules, and fitted Gaussian curves to the function relating response rate to elapsed time in the interval. The spread of the curve, that is, whether response rate grew gradually throughout the interval or was concentrated at the end of it, could be measured by the standard deviation of the fitted Gaussian curve, and the standard deviation can be divided by the mean (effectively the peak location) of the curve to yield a coefficient of variation statistic. SET requires that coefficients of variation remain constant as the interval timed varies (this is a form of the *scalar property* of variance which gives the theory its name), and Lejeune and Wearden (1991) found that there was a large duration range over which this was true for most animal species studied. In addition, the coefficient of variation, effectively a measure of temporal sensitivity where *lower* values indicate *higher* sensitivity, varied systematically with species, being lowest in cats, rats, and mice, higher in birds, and highest of all in fish and turtles. The more complex peak interval procedure has also been used frequently to test SET, see Church et al. (1994) for only one example among many.

The present article is concerned with the timing behavior of rats on single FI schedules but, more interestingly, on mixed FI schedules where the two FI values making up the mixed schedule are equiprobable. On a mixed FI schedule with two components, reinforcers are available either after the shorter value in the mixture, or the longer value, for example, after 30 s or 240 s, as in the present Experiment 1. Stimulus conditions in the interval remain constant, so at the start of the interval nothing signals to the animal which FI value is in force, and in our experiments, and most others, the different FI values are intermixed at random.

Why should performance on mixed FI schedules be of any particular theoretical interest? Leak and Gibbon (1995, pp. 3-6) provide an initial answer to this question, principally in terms of potential differences between the predictions that SET and BeT make about mixed-FI performance. SET explains timed behavior in terms of an interaction of clock, memory, and decision processes, and requires that behavior exhibit the scalar property. For

simple FI schedules, the coefficient of variation should remain constant as the interval timed varies. In the case of mixed FI schedules, the timing of the intervals in the mixture should also have scalar properties: for example, in mixed FI 30 s FI 240s, a condition presented in our Experiment 1, the timing of the 30 s component and the 240 s component should be controlled by a process with the *same* coefficient of variation.

In contrast, BeT derives timed behaviors which are experimentally-observed, such as key pecks and lever presses, from a sequence of adjunctive states, which are themselves usually unobserved. Transition from one state to another is governed by pulses from a Poisson pacemaker, but the scalar property normally found in behavior is reconciled with the principles of BeT by requiring that the pacemaker rate vary with the rate of reinforcement (effectively the FI value on simple FI schedules), such that the number of adjunctive states underlying different intervals remains the same. So, for example, if the animal goes traverses n states when timing 30 s, it should also traverse, considerably more slowly, the same n states when timing 240 s.

When a mixed-FI schedule is used (e.g. mixed FI 30 s FI 240 s), the number of states traversed in the entire interval presumably reflects the reinforcement rate in this condition (which is, of course, lower than that obtaining on FI 30 s alone, but higher than that on FI 240 s alone). The critical point is that both the shorter FI of the mixture and the longer one are timed by the *same* Poisson pacemaker, so the Poisson character of the process should manifest itself in terms of more precise timing at the longer FI than the shorter one (e.g. a smaller coefficient of variation for the 240 s component than for the 30 s one). In general, mixed FI schedules enable the timing of a consistent duration (e.g. a constant component in the mixture) to be observed while the other component, and thus the rate of reinforcement, changes. BeT relates pacemaker rate to the "temporal context" (Beam, Killeen, Bizo, & Fetterman, 1998) in which the animal finds itself, and it seems, at first sight, that, in mixed FI schedules, since all stimulus conditions remain constant, the temporal context for timing both components of the mixed FI must be the

same. However, as is usual in attempts to distinguish BeT and SET by "critical" experimental predictions, the situation may not be as clear as it first appears, as we will discuss later.

Leak and Gibbon (1995) used a number of mixed FI schedules, including those with two or three component schedules, to evaluate the relative predictions of SET and BeT, and generally found support for SET, particularly from their Experiment 3. Our experiments, while relevant to the SET versus BeT issue, have the broader aim of trying to understand more generally what factors govern timing on mixed FI schedules with two components. Some of the issues discussed are similar to those treated by Leak and Gibbon, for example, how does timing vary when the intervals making up the mixture are close together or far apart? Others, however, are rather different in particular our attempts to understand how the contributions of the components which make up the 2-valued mixed FI schedule are combined together to produced the behavior observed.

The general problem of timing of behavior on two-valued mixed-FI schedules is also discussed by Machado (1997), as part of the exposition of a complex model which deals with performance on a number of timing tasks. Machado's LeT not only predicts patterns of temporal control on FI-like schedules, but also deals with the question of how response rates should change as time passes in the schedule conditions. How our data relate to the predictions of LeT will be discussed later.

In the present article, we describe data from what was originally conceived of as a series of 5 experiments, although the studies are grouped together as 3 separate experiments in the present article to save space. In all experiments, rats timed either simple FI schedules, or two-valued mixed-FI schedules where the two components were equiprobable and presented at random within the experimental session.

Experiment 1 examined data from conditions timing FI 30s and FI 240 s separately, then using a mixed-FI condition with these schedules as

components, essentially a test of the principles for distinguishing between SET and BeT outlined by Leak and Gibbon (1995) in their Figures 1 and 2, and in an experiment similar to their Experiment 2.

Our Experiment 2 involved conditions where a single component (either FI 60 s in Experiment 2a or FI 120 s in Experiment 2b) was kept constant, while the other component in the mixture varied, from values below the constant component to values above it. This enabled us to examine performance on schedules where the two FI values were different (e.g. FI 20 s FI 60 s) or similar (e.g. FI 60 s FI 80 s). Leak and Gibbon (1995) also addressed the issue of what temporal control looks like from some schedules of this type.

Finally, Experiment 3 examined performance on schedules where the shorter component of the mixed FI was constant, while the longer value varied (Experiment 3a), or vice versa (Experiment 3b).

Although our data have not been published previously, and were derived from completely normal experimental procedures, they have the rather unusual characteristic that they were collected more than 20 years ago, as part of the doctoral work of one of the authors (Whitaker, 1980). This means that the data predate both the information-processing version of SET (first published in the early 1980s, e.g., by Gibbon et al., 1984, and Church & Gibbon, 1982), BeT (Killeen & Fetterman, 1988), and LeT (Machado, 1997) as well as other recent theories (e.g. Staddon & Higa, 1999). The only negative consequence of the age of our data is that we do not have some behavioral measures that might be desirable (e.g. the start-stop-spread analysis of individual trials, proposed for the first time by Gibbon and Church, 1990, which was used extensively by Leak and Gibbon for the analysis of their Experiment 3), nor can these more modern measures of performance now be derived from our data set. On the positive side, what we do have, as will be seen later, is the most extensive data set known to us on the timing of multiple intervals by animals, a data set which is not only very orderly, but which can be directly related to issues in modern theories of timing in animals.

Figure 1 about here

Finally, some terminological issues. When we discuss two-valued mixed FI schedules, we will consistently refer to the lower FI value in the mixture as FI *S* and the higher one as FI *L*: so, for example, we might consider effects of *L/S* ratio on timing behavior, as in Experiments 2 and 3. Leak and Gibbon (1995) refer to mixed-FI schedules as an example of *simultaneous* timing. It seems to us that this description carries considerable surplus meaning, via its implication that the animals are timing more than one interval at the same time. Some experiments where humans participate almost certainly involve simultaneous timing in this sense (e.g. Brown & West, 1990; Penney, Gibbon, & Meck, 2000, Experiment 1), but it is less clear that mixed FI does, or at least always does. Consider the data in the lower panel of Figure 1, which come from our Experiment 1 (and are similar to data presented in Leak & Gibbon, 1995, Figure 3, p. 6). Here, FI *S* is 30 s, FI *L* 240 s. At some elapsed time in the interval considerably greater than FI *S* (100 s, say), it is not necessarily the case that the animal is doing anything other than timing FI *L*, as the elapsed time is presumably highly discriminable from FI *S*, which has long passed. On the other hand, early in the interval, it is possible (although not logically necessary) that the animal is timing both FI *S* and FI *L* (e.g. at an elapsed time of 10 s), so timing may sometimes be simultaneous and sometimes not. We will just refer to timing on mixed FI schedules as *multiple* timing, without any implication of simultaneity. What the implications of this possible non-simultaneity of timing are for predictions about behavior will be discussed later.

Experiment 1

Experiment 1 is the simplest in our series. Rats received FI 30 s alone, FI 240 s alone, and mixed FI 30 s FI 240 s, with both intervals equiprobable, and randomly intermixed in the experimental session, as they were in all the experiments we report. We were interested in a number of issues. One was whether the scalar property of constant coefficient of variation would be

manifested in our single FI conditions. The second was what timing on the mixed schedule would look like: in particular, would the timing of the FI S and FI L also show scalar properties? If the scalar property was found in the mixed condition then, as mentioned above, this result is consistent with SET but contrary to the predictions of BeT which would require more precise timing (i.e. a lower coefficient of variation) at FI L than FI S.

Method

Animals

Four experimentally-naïve male hooded rats were housed individually with ad lib access to water. They were maintained at 80% of the weight reached in the last 5 days of free feeding, and were fed supplementary food approximately 1 hour after the experimental session to maintain the 80% weight.

Apparatus

Standard LeHigh Valley experimental chambers (model 142-25), enclosed in a sound-resistant housing (ambient noise level 60 dB) were used. Only the right lever was operative. Reinforcers were 45 mg Noyes pellets. The experiment was controlled by a Data General NOVA computer.

Procedure

In the first session the rats were trained to lever-press by successive approximation, then allowed to obtain 60 reinforcers, with each lever-press reinforced. Then they were exposed to FI 240 s (29 to 32 sessions for different rats), FI 30 s (28 or 29 sessions), mixed FI 30 s FI 240 s (50 to 53 sessions), then re-exposed to FI 240 s (32 to 34 sessions), then finally FI 30 s (23 to 26 sessions). Sessions normally lasted until 60 reinforcers were obtained, although they were sometimes longer.

Results and Discussion

Data were recorded in the form of response rate versus elapsed time in the interval. The total interval was always divided in 20 equal-length bins, so the bins were 1.5-s long for FI 30 s, and 12-s long for FI 240 s. The data of interest from the mixed FI condition come from intervals where responding according to FI L was reinforced, and bin-length in these cases was also 12 s. Data were averaged over rats, and over the two determinations of FI 30 s and FI 240 s, and were taken from the last 3 sessions of exposure to the schedule in force. Figure 1 shows the response rates versus elapsed time in the interval for FI 30 s (upper panel), FI 240 s (center panel), and the FI 240 s intervals of the mixed FI (bottom panel).

Inspection of the data from the single FI conditions (upper 2 panels of Figure 1) suggests that responding increased on average throughout the interval, from zero just after food to some maximum rate just before the time of food delivery. In the mixed FI condition, responding rose towards a peak located somewhere close to FI S , then declined, only to increase again to reach a peak at FI L .

The panels of Figure 1 also show fitted curves. The curves are Gaussian curves fitted to data by the nonlinear regression program of SPSS 10. A Gaussian curve is characterized by two parameters, its mean (or peak location), m , and its standard deviation, σ , so we will in general refer to this Gaussian function as $G(m, \sigma)$. To fit what is in fact a probability density function to response rates, G must be multiplied by some scaling constant, K .

When fitting Gaussian curves to the kind of data we have, a number of decisions need to be made about peak location. One possibility is to force m to be at the FI value in single FI schedules, so the only parameter derived from the fit to data from these schedules, apart from the scaling parameter, K , is σ , the standard deviation. Alternatively, the peak location, m , could itself be a parameter. When fitting data from mixed-FI schedules (for example data like those in the lower panel of Figure 1), the obvious solution is to sum two Gaussian functions together, so response rate, R , is fitted by

$$K_S * G(m_S, \sigma_S) + K_L * G(m_L, \sigma_L)$$

where K_S and K_L are scaling constants appropriate to the contribution of FI S and FI L to the response function, m_S , and m_L are peak locations appropriate to FI S and FI L , and σ_S , σ_L are standard deviations for the two Gaussian curves. This is essentially the method we used, with one proviso: for single FI schedules, we always forced m to be the FI value, and for fits to mixed FI schedules, we always forced the peak of the upper curve to be at FI L , although the position of the lower peak could vary as a parameter. The main reason this was done is that time values more than a few seconds higher than the FI value (or FI L on mixed FI schedules) were never experienced by the animal, so if such values are found by nonlinear regression it is impossible to interpret theoretically. In addition, the *worst* fit to any data sample we analyzed had an r^2 value of .96, so our decision did not do violence to the data, as will be obvious from inspection of data points and curves both in Figure 1, and presented subsequently. Table 1 shows the parameter values from the curve fits to the data in Figure 1.

Table 1 about here

Consider first a comparison of the single FI schedules, FI 30 s and FI 240 s. The fitted Gaussian curve accounted for 99% of data variance, and the coefficient of variation (σ derived from the curve fit divided by the FI value) was .25 for both schedules, supporting the scalar property in timing of single FI schedules, even when absolute duration timed varied 8-fold. Also shown are the peak heights derived from the curve fits, which correspond to the predicted response rate at the peak of the Gaussian curve. Peak response rate was nearly twice as high on FI 30 s than on FI 240 s, consistent with the fact that the rate of reinforcement is 8 times greater in the former case than the latter.

Consider next parameters derived from the fit to data from the mixed FI schedule. The peak of the Gaussian curve for the corresponding to the lower

FI was close to FI *S* (33.9 s compared to 30 s), and the coefficients of variation for the two Gaussian curves were .38 (FI *S*) and .35 (FI *L*). Peak heights from the two Gaussian curves were more similar than for the single FI conditions, although the peak was slightly higher at FI *S* than at FI *L*. The function produced by summing two Gaussian curves fitted data well ($r^2 = .96$).

Our data found scalar timing of behavior (i.e. constant coefficient of variation) in the two single FI schedules. This is consistent with both SET and BeT, as BeT produces the result by essentially varying pacemaker rate between the FI 30-s and FI 240-s schedules. In the mixed FI schedules, the relative constancy of the coefficient of variation for FI *S* and FI *L* appears more consistent with the predictions of SET than with those of BeT, which would predict that the longer FI component in the mixed schedule would be timed markedly more precisely than the shorter one. In fact, FI *L* was timed *slightly* more precisely than FI *S*, consistent with the direction of effect predicted by BeT. Exactly what the quantitative predictions of BeT might be in these cases will be considered in detail later.

Leak and Gibbon (1995) presented only one data set (their Figure 3, p. 3, which came from an unpublished study not fully described in their article) which used a 2-valued mixed FI with a large *L/S* ratio (12:1, from mixed FI 10 s FI 120 s), and their data resemble the results we obtained from mixed FI 30 s FI 240 s in all respects. Two peaks were clearly visible in the response rate versus time in the interval function obtained from the FI 120 s component of the mixed FI, with a period of low-rate responding between them; the peaks appeared to be centered approximately at FI *S* and FI *L* (although Leak and Gibbon did not present any quantitative analysis). The peak height at FI *S* was higher than that at FI *L*, but only about 50% higher.

Our Experiment 2 uses a more complex set of schedules, and consisted of two sub-experiments which are reported together as they were procedurally and conceptually almost identical. In Experiment 2a, an FI 60-s schedule was presented alone, or as a component of mixed FI. In the mixed FI conditions, the other component was varied over values of 20, 40, 80, and

100 s, producing L/S ratios of 3, 1.5, 1.33, and 1.67. In Experiment 2b, all FI values were twice as long, and the constant component was FI 120 s, presented either alone, or in a mixed schedule with values of 40, 80, 160, and 200 s, which produced the same L/S ratios as in Experiment 2a.

Leak and Gibbon (1995) discuss the conditions in which two distinct response peaks (like those shown in the bottom panel of our Figure 1) will be obtained from mixed FI schedules in terms of the L/S ratio. Catania and Reynolds (1968) reported double peaks when the L/S ratio was 4:1 or more, but not when it was less, and Leak and Gibbon (1995) reported double peaks when the ratio was 8:1 but not when it was 2.5:1. However, these conclusions seem to be based on visual inspection of the response rate versus elapsed time in the interval functions rather than quantitative analysis. In our Experiment 2, we allowed the nonlinear regression program to decide whether there were separate contributions to the response function observed on the mixed-FI schedule from FI S and FI L , rather than simple visual inspection. As will be seen later, whenever there were two peaks obvious on visual inspection, the nonlinear regression program identified them. In addition, however, the program sometimes indicated contributions from both FI S and FI L on mixed schedules with small L/S ratios.

Experiment 2

Method

Animals

8 male hooded rats were arbitrarily allocated to two groups of 4. One group served in Experiment 2a, one in Experiment 2b. Deprivation and housing conditions were as Experiment 1.

Apparatus

As Experiment 1.

Procedure

In Experiment 2a, the constant component of the mixed FI schedules was FI 60 s, which was also presented alone. In the first session, the rats received lever-press training, and then obtained 60 reinforcers, with each response reinforced. The rats were then exposed to FI 60 s (70 to 74 sessions), mixed FI 20-s FI 60-s (17 to 20 sessions), mixed FI 60-s FI 100-s (18 or 19 sessions), mixed FI 40 s FI 60 s (16 to 20 sessions), and finally mixed FI 60 s FI 80 s (16 to 19 sessions). Experiment 2b was almost identical, except that all schedule values were twice as long. After initial lever-press training, the rats received FI 120 s (72 to 75 sessions), mixed FI 120 s FI 200 s (18 to 20 sessions), mixed FI 40 s FI 120 s (23 or 24 sessions), mixed FI 120 s FI 160 s (32 to 36 sessions), and mixed FI 80 s FI 120 s (17 to 19 sessions).

Figures 2 and 3 about here

Results and Discussion

Figure 2 shows response rates versus elapsed time in the interval from the mixed FI schedules of Experiment 2a, and Figure 3 shows the same data from Experiment 2b. To improve the layout of the Figures, data from the single FI schedules (FI 60 s for Experiment 2a, FI 120 s for Experiment 2b) are not shown, but these were virtually identical in form to data from the single FI conditions of Experiment 1 (shown in the upper 2 panels of Figure 1).

Consider first data from Experiment 2a (Figure 2). Visual inspection of the data suggests that only the mixed FI 20-s FI 60-s schedule generated two obvious response peaks, but in fact the nonlinear regression program found that all the response functions could be well fitted by the sum of 2 Gaussian curves, even that from mixed FI 60 s FI 80 s, where the L/S ratio was 1.33. All 2-Gaussian fits were well-adjusted to the data (r^2 either .98 or .99 in all cases), and parameter values are shown in Table 1. The lower peak of the 2-Gaussian fit was allowed to vary, and Table 1 shows that it was always within 10% of FI S, as this varied across conditions from 20 s to 60 s. How the

contributions of FI *S* and FI *L* were combined when the *L/S* ratio varied will be considered further below.

Figure 3 shows analogous data from Experiment 2b. Once again, visual inspection identified two peaks when the *L/S* was large (mixed FI 40 s FI 120 s), although the regression produced two peak fits for two of the other three cases: only the mixed FI 120 s FI 160s condition could not be fitted by the sum of two Gaussian curves, and here a single curve with a peak at 160 s was used. Table 1 shows parameter values from the fits and, once again, r^2 values were high (.98 or .99), and the lower peak which was allowed to vary was within 10% of FI *S*.

Our data were similar to those of Leak and Gibbon (1995) in that when the *L/S* ratio was small (less than 3 in our case) two response peaks were not clearly visible on inspection of response functions. However, the fact that the nonlinear regression procedure identified two underlying sources of temporal control in all but one case (mixed FI 120 s FI 160 s) enables us to examine more precisely just what does happen when the *L/S* ratio is low in mixed FI schedules. The critical measure appears to be the peak height (i.e. the peak of responding predicted by the Gaussian curves).

Consider first data from Experiment 2a. In Figure 2, all the data were well-fitted by the 2-Gaussian function, but the peak heights reflect the relative contributions of the curve centered around FI *S* and that centered on FI *L*. In Experiment 1, with an 8:1 *L/S* ratio, peak height was higher for the FI *S* component than for the FI *L* one, but when the *L/S* ratio was small, the reverse was the case, and the peak at FI *S* was always lower than that at FI *L*. When the *L/S* ratio was particularly small (1.33 and 1.5: from the mixed FI 40 s FI 60 s and mixed FI 60 s FI 80 s conditions) the peak height associated with FI *S* component was also particularly small, and the peak height at FI *L* was much greater (more than 10 times as great in the 40/60 case).

Data from Experiment 2b (Figure 3 and Table 1) showed an almost identical picture. Peak heights at FI *L* were always more than twice as great

as those at FI *S*, and even data from the mixed FI 120 s FI 160 s condition followed this same rule trivially, as the peak at FI *S* disappeared altogether, leaving that at FI *L* infinitely greater.

In general, then, when the *L/S* ratio was small on mixed FI (less than 3 in our studies), not only did the two-peak form of the response function disappear, even though the nonlinear regression program usually identified separate sources of control by FI *S* and FI *L*, but the balance of control of responding shifted away from FI *S* towards FI *L*, with at a limit control by FI *L* alone.

Our final study, Experiment 3, consists of 2 sub-experiments, which differed procedurally but had the common theme of keeping one component of the 2-valued mixed-FI schedule constant, while varying the other. In Experiment 3a, an FI 30-s schedule was either presented alone, or as FI *S*, the lower component of the mixed schedule, while FI *L* varied over values from 45 s to 240 s. In Experiment 3b, FI *L* was fixed at FI 300 s (although this value was never presented alone), and FI *S* varied over the range from 15 s to 75 s.

Experiment 3

Method

Animals

8 male hooded rats were arbitrarily allocated to either Experiment 3a (4 rats) or Experiment 3b (4 rats).

Apparatus

Experiment 3a: Four LeHigh Valley RTC-028 operant chambers were used, all other details were as Experiment 1. Experiment 3b: all details were as Experiment 1.

Procedure

Experiment 3a: After initial lever-press training, rats received FI 30 s (30 sessions), mixed FI 30 s FI 120 s (25 sessions), mixed FI 30 s FI 60 s (25 sessions), mixed FI 30 s FI 240 s (25 sessions), and mixed FI 30 s FI 45 s (30 sessions). Experiment 3b: After initial lever-press training, rats received the following mixed FI schedules: mixed FI 30 s FI 300 s (30 sessions), mixed FI 60 s FI 300 s (25 sessions), mixed FI 15 s FI 300 s (25 sessions), mixed FI 45 s FI 300 s (25 sessions), mixed FI 15 s FI 300 s (25 sessions), and mixed FI 60 s FI 300 s (25 sessions).

Results and Discussion

Figures 4 and 5 about here

Figure 4 shows response rate versus elapsed time in the trial data from the longer intervals of the mixed FI conditions of Experiment 3a. Data from the single FI 30-s schedule are not shown, but were almost identical in form to those shown in the upper panel of Figure 1. Inspection of the data suggests that two response peaks were clearly present in all conditions except mixed FI 30 s FI 45 s, and that the peak of the lower response function was located close to FI *S*, 30 s. Table 1 shows results from the nonlinear regression analysis. As for Experiment 2a, the program found 2-Gaussian fits for data from all the mixed-FI schedules, including mixed FI 30 s FI 45 s. Inspection of the parameter values shows that (a) data were fitted well by the Gaussian (single FI) or 2-Gaussian (mixed FI) functions (smallest $r^2 = .97$), (b) the position of the lower response peak was close to FI *S*, 30 s, (c) peak heights for FI *S* were lower than those for FI *L* when the *L/S* ratio was less than 4 (with data from the FI 30-s FI 45-s condition showing a particularly small peak height at FI *S*), but higher when the *L/S* ratio was 4 and 8.

Figure 5 shows response rate versus elapsed time functions from the longer components of the mixed FI schedules of Experiment 3b. Data from the repetitions of the mixed-FI 15-s FI 300-s condition were averaged together. Inspection of the data suggests that two response peaks were

present in all cases. Table 1 shows the parameter values from the 2-Gaussian fits that the nonlinear regression program found. Examination of the parameters showed that (a) all data were well-fitted by the 2-Gaussian function (r^2 either .96 or .97), (b) the position of the lower peak was close to FI S (although consistently about 5 to 8 s above it), (c) peak heights for FI S were higher than for FI L in all cases but one (mixed FI 15 s FI 300s), but were often similar in spite of the considerable difference between FI S and FI L.

Figures 6 and 7 about here

Across-Experiment Analyses

As in Leak and Gibbon (1995) some trends in our data are seen most clearly when results from the different experiments are displayed together. Figures 6, 7, and 8 show some of these.

Consider first peak location, and peak height. The upper panel of Figure 6 shows the location of the lower peak of the 2-Gaussian function fitted to data from the different mixed FI schedules, determined by the nonlinear regression program, as a function of the FI value that was FI S. Obviously the peaks located by the program tracked the FI value closely. The line shown comes from linear regression of peak location against FI value. The (significant) slope was .98, the (non-significant) intercept 3.63 s, and r^2 was .98, confirming the impression that the peak location was located almost exactly at FI S, over the three experiments as a whole.

The lower panel of Figure 6 summarizes data on the relative peak heights from FI S and FI L in the different mixed FI schedules. An across-experiment analysis shows clearly that the relative peak heights at FI S and FI L were importantly determined by L/S ratio. When this was large (i.e. ≥ 4 , thus the FI schedules making up the mixed FI were very different), FI S had a greater peak height than FI L in 8/9 instances. In contrast, when the L/S ratio

was less than 4, all 8 instances produced greater peak height at FI *L*, and the differences were sometimes very large, as discussed above.

The data in the lower part of Figure 6 are particularly relevant to Machado's (1997) LeT account of timing. LeT is too complex to be anything other than sketched here, but like BeT it derives measured behavior from the proposed existence of a sequence of underlying behavioral states, which occur in sequence. Different states can have different levels of activation at any elapsed time since a "time marker" (e.g. the food delivery at the start of an interval of FI), and each state has some degree of association with the measured operant response, so controls its presence or absence. In LeT, states later in the sequence on FI-like schedules have higher associative strength (i.e. control the operant response more strongly) than earlier ones, and a consequence of this is that the peak response rate at FI *S* (to use our terminology) will always be lower than that on FI *L* on mixed FI schedules, even when FI *S* and FI *L* are equiprobable. The data in Figure 6 show that this prediction was almost always false when the *L/S* ratio was greater than 4, but was true when it was less. As Machado points out himself, the principles of LeT lead inevitably to the peak rate at FI *L* being greater than at FI *S*, and the reverse case (normal in our data when the *L/S* ratio was greater than 4) cannot be simulated by any simple manipulation of the parameters of the model.

Figures 7 and 8 are concerned with variability of responding around the peaks, measured by coefficient of variation. Figure 7 examines the coefficients of variation from the curve fits in two different ways. In the upper panel, coefficients of variation from all the conditions in the 3 experiments are plotted against the rate of reinforcement that the schedules provided overall (i.e. from the mixed schedules this was the rate of reinforcement on average, not that associated with each separate FI value). Data are separated according to whether they come from single FI conditions or mixed FI. Two questions were of interest. Firstly, were coefficients of variation from single and mixed FI conditions systematically different? Experiment 1 found higher coefficients of variation for FI *S* and FI *L* in the mixed FI schedules than for

the two components when presented alone, suggesting that multiple timing was associated with more timing variability than timing of a single interval. The upper panel of Figure 7 shows, however, that this was not generally true. It further suggests that the coefficient of variation was insensitive to the rate of reinforcement for the schedule: although some variability between conditions was evident, there was little to suggest systematic change in coefficient of variation as the rate of reinforcement changed.

Why is rate of reinforcement important? The answer comes from BeT, as according to this theory, the rate of adjunctive behaviors is determined by the reinforcement rate in the experimental "context" (a position for which Lejeune, Cornet, Ferreira, & Wearden, 1998, found some direct evidence), so higher rates of reinforcement in the single FI or mixed FI schedule, should, other things being equal, lead to more adjunctive states than lower rates of reinforcement, and thus more precise timing in terms of coefficient of variation. However, the upper panel of Figure 7 found no obvious effect of reinforcement rate on coefficient of variation.

The effect of reinforcement rate was examined another way, as shown in the lower panel of Figure 7 which presents selected coefficients of variation, with all data coming from constant components of mixed FI schedules. Recall that Experiments 2a, 2b, 3a, and 3b all involved one component of the mixed FI remaining constant while others changed in various ways. Suppose that we take the coefficient of variation from the constant component, and plot it against the rate of reinforcement for the condition from which it comes. The lower panel of Figure 7 shows the results when this is done. In some cases, like Experiment 2a and Experiment 3a, the changes of schedule across conditions produced marked changes in reinforcement rate: in the case of Experiment 3a the change was nearly 5-fold. According to SET, such a change in rate of reinforcement rate should have no effect overall, whereas according to BeT, coefficients of variation should decrease as the rate of reinforcement increases. Data from Experiment 3a do show declining coefficient of variation with increases in reinforcement rate, but data from other experiments either show the highest coefficient of variation at the

highest reinforcement rate (Experiments 2a and 2b) or unsystematic fluctuation (Experiment 3b, although the variation in reinforcement rate was very small in this experiment). Overall, the results were mixed, but there was little which clearly supports the assertion of a marked change in coefficient of variation with changing rates of reinforcement.

Figure 8 about here

However, although examination of effects of reinforcement rate on coefficient of variation offered little support for BeT overall, Figure 7 used data from all the experiments, and thus aggregated results from mixed-FI schedules with low L/S ratios, and those with longer ones, whereas a general theme of our article is that the contributions of the two FI values to responding on mixed FI is different in the two cases. In general, BeT requires more precise timing at FI L than FI S , for any L/S pair making up a mixed-FI schedule, as the constant rate of reinforcement in the situation will result in more adjunctive states in the FI L region than in the FI S one. In general, then, the coefficient of variation at FI L should be smaller than at FI S . SET, on the other hand, requires the coefficients of variation to be the same, or at least to vary unsystematically.

Figure 8 shows the ratios of the coefficients of variation for the FI L and FI S curves, plotted against L/S ratio. Here, coefficient of variation values less than 1.0 indicate that timing was more precise at FI L than FI S , values greater than 1.0 the reverse. The data were divided into those coming from mixed-FI schedules with L/S ratios greater than 4 (upper panel), and those with L/S ratios less than 4 (lower panel). In the former case, most values were less than 1.0 (average .80), supporting BeT's prediction of more precise timing at FI L , although the effect of L/S ratio did not appear to be systematic. On the other hand, when the L/S ratio was small, the reverse was generally true, and in most mixed-FI schedules, FI S was timed more precisely (average ratio 1.44). Once, again, the data provide another example of behavioral differences between mixed-FI schedules with small and larger L/S ratios. If this difference was ignored, and the ratios from all L/S ratios treated together,

then the average coefficient of variation ratio was 1.14, supportive of slightly more precise timing at the shorter FI value of the mixed schedule.

General Discussion

Our results might be summarized by considering an attempt to predict performance on the longer intervals of some mixed FI schedule mixed FI *S* FI *L*. If the *L/S* ratio was greater than about 4, then two peaks in the response function (one located at or near to FI *S*, the other located at FI *L*) would be obvious on visual inspection. The response curves around FI *S* and FI *L* would be of approximately Gaussian shape, with a coefficient of variation that was slightly lower for FI *L* than for FI *S*, and peak height at FI *S* would be greater (but sometimes not much greater) than that at FI *L*. On the other hand, if the *L/S* ratio was less than 4, visual inspection would not usually locate two distinct response peaks, although sources of control of responding by the separate FI *S* and FI *L* components would usually be identified by nonlinear regression. When two components were identified, the contribution of FI *S* to the overall response function would be smaller, sometimes much smaller, than that of FI *L*. Peak height at FI *S* would usually be a fraction of that at FI *L*, and the coefficient of variation of the FI *S* peak would usually be smaller than that from FI *L*.

The dependence of behavior on mixed FI schedules on *L/S* ratio is not particularly surprising either logically or on the basis of previous work. As FI *L* and FI *S* become more similar, the animals presumably find it more difficult to maintain separate sources of temporal control, because of overlap in their temporal representations of *S* and *L*, as discussed by Leak and Gibbon (1995), who reported a similar effect of *L/S* ratio (as did Catania & Reynolds, 1968, before them). What is not so obvious a priori, although becomes clearer in our data, is how the "conflict" between control by FI *S* and by FI *L* is resolved when the *L/S* ratio is small. It is clear that, in all cases, control is more powerfully exerted by FI *L* than FI *S*. Some evidence for this comes from the single case we have where 2 response peaks could not be identified (mixed FI 120 s FI 160 s), and here the data could be modelled by assuming

control only by the longer, rather than the shorter, component of the mixed-FI schedule. In addition, our analysis of peak heights (lower panel of Figure 6) shows that the peak height at FI *S* was smaller than that at FI *L*, when the *L/S* ratio was small, and the contribution of FI *S* to the response pattern on the longer components of mixed-FI schedules was correspondingly less than that of FI *L* in these cases.

In general, data from all our conditions suggest that FI *L* is in some senses more "important" than FI *S* in the mixed-FI schedule. Not only does FI *L* dominate in mixed-FI schedules with small *L/S* ratios, as discussed earlier, but even in others the relative peak heights of the components also suggest something similar. For example, in Experiment 1, the peak at FI *L* (240 s) was only slightly lower than the peak at FI *S* (30 s), in spite of the 8-fold difference in time to reinforcement between them. As is well-known, on aperiodic reinforcement schedules such as variable-interval (VI) response rate is highly sensitive to reinforcement rate, and this sensitivity can be embodied in Herrnstein's equation (Herrnstein, 1970), and related formulations. On VI, it would be surprising if two VI schedules with an 8-fold ratio of reinforcement rate supported nearly the same rate of responding, whereas in our mixed-FI schedules, with large *L/S* ratios, even disparate schedules produced similar peak heights.

One way of conceptualizing this effect is to suggest that the peak height at FI *L* is higher than it "should be", and the peak height at FI *S* lower than it "should be", on considerations of reinforcement rate alone. Why is this? One explanation for the general "dominance" of FI *L* in mixed FI schedules might be a role played by probability, or conditional probability, of reinforcement at FI *S* and FI *L*, rather than just time to reinforcement from the start of the interval, or reinforcement rate on the schedule overall. On a two-valued mixed FI schedule, reinforcer delivery at FI *S* is never certain until it arrives, whereas reinforcer delivery at FI *L* sometimes is. For example, at the start of the interval, the probabilities of reinforcement at FI *S* and FI *L* are equal, but when FI *S* has been passed, reinforcement at FI *L* becomes certain. Put a slightly different way, the conditional probability of reinforcement

at FI S and FI L given a certain elapsed time in the interval changes in different ways for FI S and FI L . Up to the time of FI S , the conditional probability of reinforcement at both FI S and FI L is the same at 0.5, whereas after S , the conditional probability of reinforcement at FI S drops to zero, whereas that at FI L increases to 1.0.

The animal's time discrimination is not of course perfect, so the conditional probabilities may be blunted to some extent by variability of temporal representations, but on mixed FI schedules with large L/S ratios, there will come a time in the interval which is longer than, and discriminable from, FI S , and at this point the animal can be certain that a reinforcer will be available at FI L . For example, it would be surprising if times greater than 100 s were not discriminable from 30 s on mixed FI 30 FI 240 s, so if the rat arrives at such a time, reinforcement at 240 s may be subjectively certain, and this certainty may boost response rate.

This kind of consideration is similar to ideas from Beam et al. (1998, p. 395), and may explain how BeT can be reconciled with some of the results we obtained, and some of those obtained by Leak and Gibbon (1995), which seem at first sight contrary to BeT. For example, if pacemaker rate is determined by arousal linked to expectancy of reinforcement, then on a mixed-FI schedule this arousal may change markedly as the interval elapses. For example, put informally, the animal may begin an interval on a mixed FI 30-s FI 240-s schedule in a high state of arousal (i.e., with high pacemaker rate), because reinforcement in a fairly short time is a 50% possibility. Later in the interval (e.g. at 60 or 100 s after the start), arousal falls because the animal now anticipates food for certain at 240 s into the interval, but this time is remote from the current elapsed time. So, in general, the assumption that the rate of the BeT pacemaker will necessarily remain constant on mixed FI schedules because the "temporal context" remains constant, which is the basis of Leak and Gibbon's (1995) argument that mixed FI schedules can distinguish between BeT and SET, may not be sustainable.

Such an argument may also help Machado's (1997) LeT deal with what for it is the "impossible" situation of a higher peak rate at FI *S* than FI *L*. LeT has a parameter which like BeT's "arousal", governs the rate of transition between underlying behavioral states and, like BeT, LeT reasonably assumes that this parameter is constant during an experimental trial. However, as noted above, this may not be true when the *L/S* ratio is large on the long interval of a mixed-FI schedule, at elapsed times that are discriminably above FI *S*. Changes in the rate of transition between behavioral states within trials may also change their subsequent associative strength, and some way might be found to predict higher peak rates at FI *S* than FI *L*. When the *L/S* ratio is low, on the other hand, the assumption of constancy of transition rates may hold, thus the prediction of LeT would be that the peak rate at FI *L* would be higher than that at FI *S* which, as the lower panel of our Figure 6 shows, was true in all the data sets we analyzed. However, whether LeT can predict the very large differences in peak height between FI *S* and FI *L* found in some of our conditions remains to be seen.

As our Figure 8 shows, when the *L/S* ratio was large on mixed FI schedules, FI *L* was usually timed slightly more precisely than FI *S*. BeT predicts the direction of this effect, but the differences appear to be much smaller than would be expected on the basis of BeT with constant pacemaker rate. Beam et al. (1998, p. 392) show that the coefficient of variation of the response rate function at some time value is given by the square root of $1/(n + 1)$, where n is the number of adjunctive states in the time period. For a constant pacemaker rate, n would just increase with FI value, so would be n for FI 30 s and $8n$ for FI 240 s. Working out the expected ratio of coefficients of variation depends on n , but data from Beam et al. (1998) suggest that the BeT pacemaker has inter-pulse times between 5 s and 20 s. The lower value would predict a coefficient of variation ratio of .37, for the FI 240-s and FI 30-s components of mixed FI 30 s FI 240 s, the higher one a value of .44, both values much less than the value found in data (0.92), but the discrepancy could be due, as mentioned above, to a change in pacemaker rate part-way through the mixed-FI interval.

In general therefore, although mixed-FI schedules produce orderly data, and are important in understanding the way that sources of temporal control are combined when there is more than one of them, they may not be as completely clear as tests of the different predictions of timing models as was supposed by work published even quite recently (e.g. Leak and Gibbon, 1995). In particular if, on long intervals of mixed-FI schedules with a large L/S ratio, "arousal" induced by expectancy of reinforcement changes within intervals, then certain findings which seem inconsistent with BeT and LeT might, in fact, be easily reconcilable with these theories.

SET has been criticised on the grounds that its complex structure makes it difficult to falsify (Staddon & Higa, 1999; see also Wearden, 1999, for a similar view from a position much more sympathetic to SET), but if pacemaker rate (or state transition rate) and the number of adjunctive states can change in complex ways both within and between conditions (sometimes so that changes in one can compensate for opposite changes in another, see Killeen, Hall, & Bizo, 1999, for an example), theories claiming a more "behavioral" character than SET, such as BeT and LeT, may also be prone to problems of testability.

Table 1

Schedule and parameter values from curve fitting conducted on data from Experiments 1, 2, and 3. **FI S**, **FI L**: upper and lower FI value in mixed-FI schedule (in s), single FI values are given as **FI L** only. **P S**: peak location, in s, of fitted Gaussian curve corresponding to FI S. **cv S**, **cv L**: coefficients of variation of fitted curves. **h S**, **h L**: peak heights (responses/s) of fitted curves. **r²**: proportion of variance accounted for by the curve fit.

Experiment	FI S	FI L	P S	cv S	h S	cv L	h L	r ²
1		30				.25	1.13	.99
		240				.25	.63	.99
	30	240	33.9	.38	.97	.35	.85	.96
2a		60				.19	1.66	.99
	20	60	21.4	.26	.51	.33	1.18	.98
	40	60	37.9	.14	.14	.29	1.42	.99
	60	80	54.4	.13	.20	.37	1.53	.99
	60	100	65.3	.28	.74	.15	.83	.99
2b		120				.19	.97	.99
	40	120	42.2	.32	.30	.43	.72	.99
	80	120	78.4	.20	.57	.23	1.30	.99
	120	160				.28	1.35	.98*
	120	200	120.9	.27	.32	.27	.86	.99
3a		30				.22	.79	.99
	30	45	31.0	.15	.18	.28	1.43	.99
	30	60	31.8	.22	1.02	.21	1.23	.99
	30	120	34.1	.22	1.01	.21	.84	.98
	30	240	32.0	.27	1.16	.33	.52	.97

3b	15	300	21.7	.57	.39	.21	.52	.96
	30	300	34.8	.40	.43	.39	.19	.97
	45	300	50.9	.33	.79	.24	.69	.97
	60	300	66.2	.32	.63	.26	.42	.97
	75	300	83.1	.33	.50	.32	.43	.96

* no two-peak fit

References

- Beam, J.J., Killeen, P.R., Bizo, L.A., & Fetterman, J.G. (1998). How temporal context affects temporal production and categorization. *Animal Learning and Behavior*, *26*, 388-396.
- Brown, S.W., & West, A.N. (1990). Multiple timing and the allocation of attention. *Acta Psychologica*, *75*, 103-121.
- Catania, A.C., & Reynolds, G.S. (1968). A quantitative analysis of responding maintained by interval schedules of reinforcement. *Journal of the Experimental Analysis of Behavior*, *11*, 327-383.
- Church, R.M., & Gibbon, J. (1982). Temporal generalization. *Journal of Experimental Psychology: Animal Behavior Processes*, *8*, 165-186.
- Church, R.M., Meck, W.H., & Gibbon, J. (1984). The application of scalar timing theory to individual trials. *Journal of Experimental Psychology: Animal Behavior Processes*, *20*, 135-155.
- Dews, P.B. (1978). Studies on responding under fixed-interval schedules of reinforcement: II: The scalloped pattern of the cumulative record. *Journal of the Experimental Analysis of Behavior*, *29*, 67-75.
- Gibbon, J., & Church, R.M. (1990). Representation of time. *Cognition*, *37*, 23-54.
- Gibbon, J., Church, R.M., & Meck, W. (1984). Scalar timing in memory. In J. Gibbon and L. Allan (Eds.), *Annals of the New York Academy of Sciences*, 423: Timing and time perception, (pp. 52-77). New York: New York Academy of Sciences.

- Herrnstein, R.J. (1970). On the law of effect. *Journal of the Experimental Analysis of Behavior*, 13, 243-266.
- Killeen, P.R. & Fetterman, G.J. (1988). A behavioral theory of timing. *Psychological Review*, 95, 274-295.
- Killeen, P.R., Hall, S., & Bizo, L.A. (1999). A clock not wound runs down. *Behavioural Processes*, 45, 129-139.
- Leak, T.M., & Gibbon, J. (1995). Simultaneous timing of multiple intervals: Implications of the scalar property. *Journal of Experimental Psychology: Animal Behavior Processes*, 21, 3-19.
- Lejeune, H., & Wearden, J.H. (1991). The comparative psychology of fixed-interval responding: Some quantitative analyses. *Learning and Motivation*, 22, 84-111.
- Lejeune, H., Cornet, S., Ferreira, J., & Wearden, J.H. (1998). How do mongolian gerbils (*Meriones unguiculatus*) pass the time? Adjunctive behavior during temporal differentiation in gerbils. *Journal of Experimental Psychology: Animal Behavior Processes*, 24, 325-334.
- Lowe, C.F., Harzem, P., & Spencer, P.T. (1979). Temporal control of behavior and the power law. *Journal of the Experimental Analysis of Behavior*, 31, 333-343.
- Machado, A. (1997). Learning the temporal dynamics of behavior. *Psychological Review*, 104, 241-265.
- Machado, A., & Guilhardi, P. (2000). Shifts in the psychometric function and their implications for models of timing. *Journal of the Experimental Analysis of Behavior*, 74, 25-54.

- Penney, T.B., Gibbon, J., & Meck, W.H. (2000). Differential effects of auditory and visual signals on clock speed and temporal memory. *Journal of Experimental Psychology: Human Perception and Performance*, 26, 1770-1787.
- Roberts, S. (1981). Isolation of an internal clock. *Journal of Experimental Psychology: Animal Behavior Processes*, 7, 242-268.
- Schneider, B.A. (1969). A two-state analysis of fixed-interval responding in the pigeon. *Journal of the Experimental Analysis of Behavior*, 12, 667-687.
- Skinner, B.F. (1938). *The behavior of organisms*. New York: Appleton-Century-Crofts.
- Staddon, J.E.R., & Higa, J. (1999). Time and memory: Towards a pacemaker-free theory of interval timing. *Journal of the Experimental Analysis of Behavior*, 71, 215-251.
- Wearden, J.H. (1999). "Beyond the fields we know...": Exploring and developing scalar timing theory. *Behavioural Processes*, 45, 3-21.
- Whitaker, J.S. (1980). *Temporal control of animal operant performance*. Unpublished Ph.D. thesis. University College of North Wales, Bangor.

Figure legends

Figure 1. Response rates versus elapsed time in the interval from FI 30 s (upper panel), FI 240 s (center panel), and the FI 240-s intervals of mixed FI 30 s FI 240 s (bottom panel). Data points are shown as unconnected filled circles, and the best fitting Gaussian or 2-Gaussian function (see text for details) is shown as a line. Parameter values for the curve fit are given in Table 1.

Figure 2. Data from the longer components of mixed-FI conditions of Experiment 2a. Other details as Figure 1.

Figure 3. Data from mixed-FI conditions of Experiment 2b. Other details as Figure 1.

Figure 4. Data from mixed-FI conditions of Experiment 3a. Other details as Figure 1.

Figure 5. Data from mixed-FI conditions of Experiment 3b. Other details as Figure 1.

Figure 6. Upper panel: Peak location (in s) of the Gaussian curve associated with FI S in the mixed-FI schedules used in Experiments 1 to 3, plotted against FI S value in s. Lower panel: Number of instances where the peak height at FI S was greater than at FI L (or lower). Data are group by L/S ratio (≥ 4 or < 4).

Figure 7. Upper panel: Coefficients of variation from the Gaussian curves fitted to data from single and mixed FI schedules used in Experiments 1 to 3, plotted against the programmed rate of reinforcement provided by the schedules. Lower panel: Coefficients of variation of curves fitted to the constant component of the mixed FI schedules used in Experiments 2 and 3, plotted against reinforcement rate for the schedule.

Figure 8. Ratio of coefficients of variation of curves fitted to FI L and FI S from mixed FI schedules, plotted against L/S ratio. Values greater than 1.0

indicate that the coefficient of variation at FI L was greater than that at FI S ; values below 1.0 the reverse. Data are shown separately from conditions with L/S ratio ≥ 4 (upper panel), or < 4 (lower panel).

Author note

The experimental work reported in this thesis was conducted in partial fulfilment of the requirements for a degree of Ph.D. awarded to the first author by the then-titled University College of North Wales, Bangor, in 1980. J.S. Whitaker is currently in the Learning Disabilities Unit University of Huddersfield, and C.F. Lowe at the School of Psychology, University of Wales, Bangor. Correspondence concerning this article should be sent to J.H. Wearden, Department of Psychology, Manchester University, Manchester, M13 9PL, U.K. email: wearden@psy.man.ac.uk.















