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**KNOXVILLE  
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Project Title

ADAPTIVE STATISTICAL PATTERN CLASSIFIERS  
FOR REMOTELY SENSED DATA

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## ABSTRACT

A new technique for the adaptive estimation of non-stationary statistics necessary for Bayesian classification is developed. The basic approach to the adaptive estimation procedure consists of two steps: (1) an optimal stochastic approximation of the parameters of interest and (2) a projection of the parameters in time or position. A divergence criterion is developed to monitor algorithm performance. Comparative results of adaptive and non-adaptive classifier tests are presented for simulated four dimensional spectral scan data.

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## LIST OF SYMBOLS

$\theta_n$  = true value of distribution parameter at time (position)  $n$ .

$Y_n$  = data sample classified into a particular class at time (position)  $n$ .

$X_n$  = "refined" estimate of  $\theta_n$  made after classification number  $n$  provides new data sample.

$X_n^*$  = "projected" estimate of  $\theta_{n+1}$  made at preceding time (position)  $n$ .

$\overline{e_n^2}$  = mean square error  $\overline{(X_n^* - \theta_{n+1})^2}$ .

$\gamma_{n-1}$  = weight used in stochastic approximation to specify weighted average of past estimate and present data (chosen to minimize mean-square error,  $\overline{e_n^2}$ ).

$\overline{E_n^2}$  = mean square error  $\overline{(X_n - \theta_n)^2}$ .

$N(u, \sigma^2)$  denotes normal distribution of mean  $u$  and variance  $\sigma^2$ .

## CHAPTER I

### INTRODUCTION

A Bayes classifier for  $M$  pattern classes is essentially a mechanization of  $M$  discriminant functions of the patterns  $\underline{x}$ . These functions are of the form

$$d_i(\underline{x}) = p(\underline{x}/\omega_i) p(\omega_i) \quad (1)$$

$$i = 1, 2, \dots, M$$

where  $p(\underline{x}/\omega_i)$  is the probability density function of the patterns of class  $\omega_i$  and  $p(\omega_i)$  is the *a priori* probability of this class, that is the probability of occurrence of class  $\omega_i$ . The maximum discriminant function will correspond to the minimum conditional risk. In other words, the Bayes classifier will minimize total expected loss, where loss represents classification error [1].

In order to make a decision on a particular pattern  $\underline{x}$ , the classifier computes  $d_1(\underline{x})$ ,  $d_2(\underline{x})$ , ...,  $d_M(\underline{x})$ , and assigns  $\underline{x}$  to class  $\omega_j$  if  $d_j(\underline{x})$  has the largest value. Ties are resolved arbitrarily. Because the Bayes classifier has found such wide acceptance in pattern recognition, this classifier will serve as the basis for an adaptive recognition system capable of adjusting itself to a changing environment.

The structure of a Bayes classifier is determined primarily by the conditional densities  $p(\underline{x}/\omega_i)$ . Of the various density functions that have been investigated, none has received more attention than the multivariate normal density. Although this attention is due largely to its analytical tractability, the multivariate normal density is also an appropriate model for an important situation: the case where the feature vectors  $\underline{x}$  for a given class  $\omega_i$  represent a single typical or prototype vector  $\underline{u}_i$ , mildly corrupted by zero mean sampling and measurement noise [2,3].

For M pattern classes the general multivariate normal density functions may be written as

$$p(\underline{x}/\omega_i) = \frac{1}{(2\pi)^{n/2} |C_i|^{1/2}} \exp[-1/2(\underline{x}-\underline{u}_i)' C_i^{-1} (\underline{x}-\underline{u}_i)] \quad (2)$$

$$i = 1, 2, \dots, M$$

$$n = \text{dimensionality of } x$$

where

$$\underline{u}_i = E[\underline{x}] \quad (3)$$

and

$$C_i = E[(\underline{x}-\underline{u}_i)(\underline{x}-\underline{u}_i)'] \quad (4)$$

For class  $\omega_i$  the Bayes decision function which minimizes probability of classifier error is found to be  $d_i(\underline{x}) = p(\underline{x}/\omega_i)p(\omega_i)$ . Due to the exponential form of the normal density function, it is more convenient to work with the natural logarithm of this decision function [1].

The decision function may therefore be written as

$$d_i(\underline{x}) = \text{Ln } p(\underline{x}/\omega_i)p(\omega_i) = \text{Ln } p(\underline{x}/\omega_i) + \text{Ln } p(\omega_i) \quad (5)$$

$$i = 1, 2, \dots, M.$$

Dropping the term  $n/2 \text{Ln}(2\pi)$  because it is common to all  $M$  decision functions being compared yields

$$d_i(\underline{x}) = \text{Ln } p(\omega_i) - 1/2 \text{Ln}|C_i| - 1/2[(\underline{x}-\underline{u}_i)' C_i^{-1} (\underline{x}-\underline{u}_i)] \quad (7)$$

$$i = 1, 2, \dots, M.$$

An examination of Equation (7) reveals that the changing environment to which the system must adapt is composed of the particular class mean vectors  $\underline{u}_i$  and covariance matrices  $C_i$ . In the context of a classification, to adapt means to provide the classifier optimal current estimates of parameters necessary for the classification. The parameters may vary with time or position.

In this thesis various stochastic approximation techniques are presented for adaptive estimation. A criterion is also suggested which may be used to detect the divergence of estimates of means.

The ability of these algorithms to accurately estimate the varying mean of a normal density has been tested by computer simulation.

These algorithms have been incorporated into a Bayes classifier to make it adaptive. Comparisons of the various adaptive classifiers, incorporating different estimation algorithms, to the ordinary (non-adaptive) Bayes classifier have been made revealing the desirability of adaptive recognition capability.

A practical application which has been implemented in this work is real-time classification and physical class boundary definition of synthetic multispectral scan data. These boundaries are those between classes in a truth table, and should not be confused with Bayes decision surfaces in pattern space.

The classifier developed here is to serve as a model or prototype; therefore, only the two class recognition problem has been considered. Extension to the more general multiclass case involves no more difficulty than would be involved with an ordinary Bayes classifier, once the stochastic approximation procedures are understood.

The data used in testing the classifiers was generated on the IBM 360/65 computer system [4]. Algorithm checkout and classifier testing have been performed on the IBM 360/65 and the PDP 11/40 computer systems. Results of classification and subsequent boundary definition have been displayed via the Data Disk video system in conjunction with the PDP 11/40 computer.

## CHAPTER II

### ESTIMATION ALGORITHMS

A typical sequence of events for classifying and subsequently estimating class statistics at the next time, assuming current estimates have been made, is as follows.

Step 1. The current data sample  $Y_n$  is classified into a particular class using current estimates of parameters for all classes.

Step 2. A "refined" estimate of the parameters of the class chosen in Step 1 is computed by stochastic approximation as

$$\theta_n \approx X_n = X_{n-1}^* + \gamma_{n-1}(Y_n - X_{n-1}^*).$$

This step is omitted for all other classes for lack of data  $Y_n$ .

Step 3. A "projected" estimate of  $\theta_{n+1}$ ,  $X_n^*$ , may be made by transforming  $X_n$  according to the way the algorithm assumes  $\theta$  is changing with  $n$ . If the change is due to time, this step is made for all classes; if the change is due to position (i.e., as when classifying pixels of a multispectral scan frame) within the current class being scanned, this step is performed only for that class chosen in Step 1 above.

Step 4. Increment  $n$  by 1 and return to Step 1.

Several notable contributions have been made to the problem of estimating the parameters for a classifier where the class statistics vary with time or space.<sup>1</sup> One such adaptive estimator gave larger weight to more recent samples, as specified by an empirically determined exponential weighting parameter; the consequent "limited memory" made the resultant average more up-to-date [5]. Intuitively the resulting estimates of parameters would be better than an unweighted average. Another adaptive estimation algorithm "projected" the current estimate to the next step by adding an amount of a certain form of anticipated change to the last estimate, and then combining it with the next data sample in a weighted average with weights chosen to minimize the mean square error [6]. This algorithm will subsequently be referred to as the CF algorithm, after the authors. The algorithm developed in this work consists of "refine" and "project" steps [7]. This algorithm differs from the previous one in the sense that the former (1) makes projections suitable for more complex variations with time, and (2) is arranged to operate as part of a Bayes classifier. It will be seen that in both these algorithms the "refine" step of combining previous estimate and new data is in the form of a stochastic approximation formulation shown previously in Step 2 of the typical sequence of events for adaptive classification.

The CF algorithm is essentially a two step algorithm designed to optimally estimate present values of interest rather than to project

---

<sup>1</sup>Time will henceforth denote true time or space (positional index), unless otherwise specified.

an estimate for future use. The two CF steps are defined as follows:

$$(1) \quad X_{n-1}^* = [1 + (n-1)^{-1}] X_{n-1} \quad (8)$$

$$(2) \quad X_n = X_{n-1}^* + \gamma_{n-1} (Y_n - X_{n-1}^*) \quad (9)$$

where  $X_{n-1}^*$  represents a projected parameter estimate,  $X_n$  represents the previous estimate of the parameter,  $\theta_n$ , and  $Y_n$  is the current data sample.  $\gamma_{n-1}$  is a sequence of positive numbers satisfying the conditions of Dvoretzky [8]

$$\lim_{n \rightarrow \infty} \gamma_{n-1} = 0, \quad \sum_{n=1}^{\infty} \gamma_{n-1} = \infty, \quad \sum_{n=1}^{\infty} \gamma_{n-1}^2 < \infty \quad (10)$$

and chosen to minimize the mean square error of the estimates. Because this algorithm is similar to the form required by an adaptive Bayes classifier, the incorporation of the technique in a classifier is justifiable. In contrast to the empirically derived algorithm discussed previously, this procedure produces optimal estimates.

Examination of equation (8) reveals that this technique assumes the estimated parameter to be time varying in a linear or nearly linear fashion, with zero initial value. The algorithm lacks compensation for an initial non-zero offset or bias of the parameter value.

A modification to the algorithm consisted of subtracting the initial parameter value from the classified sample, applying the CF algorithm to the result, and adding back the initial value to the algorithm estimate. In effect the modification allowed the algorithm to project estimates as if the initial value were zero.

The algorithm developed here produces true distribution parameter estimates for the class of interest at the next classification time. A "refine" step is made, then a "project" step is also made to the next time, because once the data has been classified, the classifier will require an estimate of the future parameter value, not the present. An optimum compromise between the present parameter estimate  $X_{n-1}^*$ , made at the previous step  $n-1$ , and the present data sample  $Y_n$  is made by the stochastic approximation in the "refine" step. The "project" operation then provides the classifier an estimate of the mean for the time when it is actually needed by the classifier. An input, unbiased by variation, is also provided for the next stochastic approximation by the "project" step. Therefore, the "project" operation should remove (in a statistical sense) the estimation bias.

A name considered appropriate for the algorithm is "polynomial fit," hereafter to be referred to as the PF algorithm. The particular algorithm presented was derived to make nonlinear estimates of degree two; however, PF actually represents a class of algorithms derivable for any finite degree. The second degree PF algorithm can be specified as follows.

The refine step (Step 2 of typical classification sequence) is denoted

$$X_n = X_{n-1}^* + \gamma_{n-1} (Y_n - X_{n-1}^*) \approx \theta_n \quad (11)$$

and the project step (Step 3 of typical classification sequence)

$$X_n^* = X_n + \hat{S} \approx \theta_{n+1} \quad (12)$$

where

$\theta_n \equiv$  true value at step  $n$

and

$$\begin{aligned} \hat{S} &= \{[i(i+1) - j(j+1)] Y_n - [i(i+1)] Y_{n-j} + [j(j+1)] Y_{n-i}\} / ij(i-j) \\ &\approx \theta_{n+1} - \theta_n \end{aligned} \quad (13)$$

and

$$\gamma_{n-1} = \frac{\overline{e_n^2} - K_1 \sigma_n^2}{\overline{e_n^2} + \sigma_n^2} \quad (14)$$

and the estimate of mean square error for use in the calculation of  $\gamma_n$  is

$$\begin{aligned} \overline{e_{n+1}^2} &\equiv \overline{(X_n^* - \theta_{n+1})^2} \\ &= \frac{\overline{e_n^2} \sigma_n^2}{\overline{e_n^2} + \sigma_n^2} (K_1 + 1)^2 + (K_2^2 + K_3^2) \sigma_n^2 \end{aligned} \quad (15)$$

the required terms for error calculation being

$$K_2 = -\frac{j+1}{i(i-j)} \quad (16)$$

and

$$K_3 = \frac{i+1}{j(i-j)} \quad (17)$$

with  $K_1$  defined as the sum of these two or

$$K_1 \equiv K_2 + K_3. \quad (18)$$

Here the variance of the density function from which samples  $Y_n$  are drawn is represented as  $\sigma_n^2$ .

Another form of the PF algorithm has been developed using previously projected estimates rather than previous data samples to

fit the polynomial assumed in derivation. This form of second degree algorithm may be specified as follows.

The refine and project steps are specified exactly as in equations (11) and (12) except with

$$\hat{S} = \{[i(i+1) - j(j+1)] x_n - [i(i+1)] x_{n+j} + [j(j+1)] x_{n-i}\} / ij(i-j) \quad (19)$$

and

$$\gamma_{n-1} = \frac{\overline{e_n^2}}{\overline{e_n^2} + \sigma_n^2} \quad (20)$$

and the estimate of mean square error for use in the calculation of  $\gamma_n$  is

$$\overline{e_{n+1}^2} = (1+K_1)^2 \overline{E_n^2} + K_2^2 \overline{E_{n-j}^2} + K_3^2 \overline{E_{n-i}^2} \quad (21)$$

where

$$\begin{aligned} \overline{E_n^2} &\equiv \overline{(X_n - \theta_n)^2} \\ &= (1-\gamma_{n-1})^2 \overline{e_n^2} + \gamma_{n-1}^2 \sigma_n^2 \end{aligned} \quad (22)$$

and the required constants for error calculation being

$$K_2 = \frac{(i+1)}{j(i-j)} \quad (23)$$

and

$$K_3 = \frac{-(j+1)}{i(i-j)} \quad (24)$$

and with  $K_1$  again defined as the sum of these two or

$$K_1 \equiv K_2 + K_3 . \quad (25)$$

The "project" operation of equation (12) takes a form suitable for the manner in which the mean is assumed to vary with time while in the CF algorithm, "projection" is accomplished as  $X_n^* = (1+1/n) X_n$ .  $\hat{S}$  of equation (12) is an estimate of anticipated change on the next time interval based on the assumption that the true value varies as a second degree function of time, which is in turn estimated by the values of  $Y_n$ ,  $Y_{n-i}$ , and  $Y_{n-j}$ , or  $X_n$ ,  $X_{n-i}$ , and  $X_{n-j}$ . Equations (13) or (20) give the optimum weight  $\gamma_{n-1}$  to minimize  $e_{n+1}^2$  for the two forms of the algorithm. The classifier then uses  $X_n^*$  as the best available value for  $\theta_{n+1}$  for the next classification, at step  $n+1$ .

Tests of both forms of the second degree PF algorithm revealed that each made equally reliable projections. A disadvantage of the

second form is that approximately twice as much memory is required in order to accommodate all previous estimates of the necessary error term  $\overline{E_i^2}$ ,  $i = 1, 2, \dots, n$ .

The ability of the CF and PF algorithms to "track" the varying mean of a Gaussian density has been tested by computer simulation. The data  $\{Y_n\}$  were drawn from a unit-variance, one dimensional Gaussian density with mean  $9(n-50)^2/2500 + 1$  for  $n = 1$  to 100, and the algorithms produced up-to-date estimates of this mean. Ten statistically independent runs were made for  $1 \leq n \leq 100$ ; the CF algorithm performance is shown in Figure 1, while the second degree PF algorithm performance is shown in Figure 2. For the sake of comparison the performance shown in Figure 3 is that resulting from a least mean square error fit of a second degree curve to the set  $\{Y_K\}$ ,  $K = 1, 2, \dots, 100$ .

Both the PF and CF algorithms have been applied to the problem of adapting to changing mean vectors and covariance matrices of normal class signatures. In order to adapt to changing covariance matrices, the problem addressed was that of estimating elements of the correlation matrices separately from the elements of the mean vectors and then combining these to form the particular covariance matrices [9]. A problem initially encountered using both algorithms was that of maintaining a positive-semidefinite covariance matrix.

Based on the assumption that covariance terms vary at a slower rate than mean vector components, satisfactory estimates of the covariance matrices of  $M$  classes may be obtained by updating the  $j^{\text{th}}$  class covariance matrix when  $P$  samples have been classified as members of that class in the following manner.

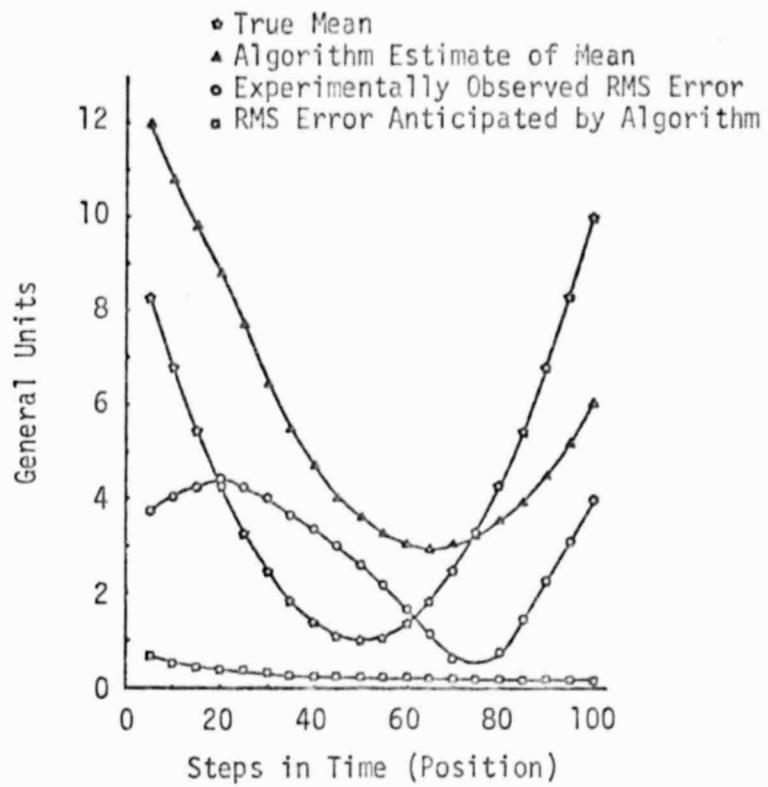


Figure 1. Performance of CF (Chien and Fu) algorithm.

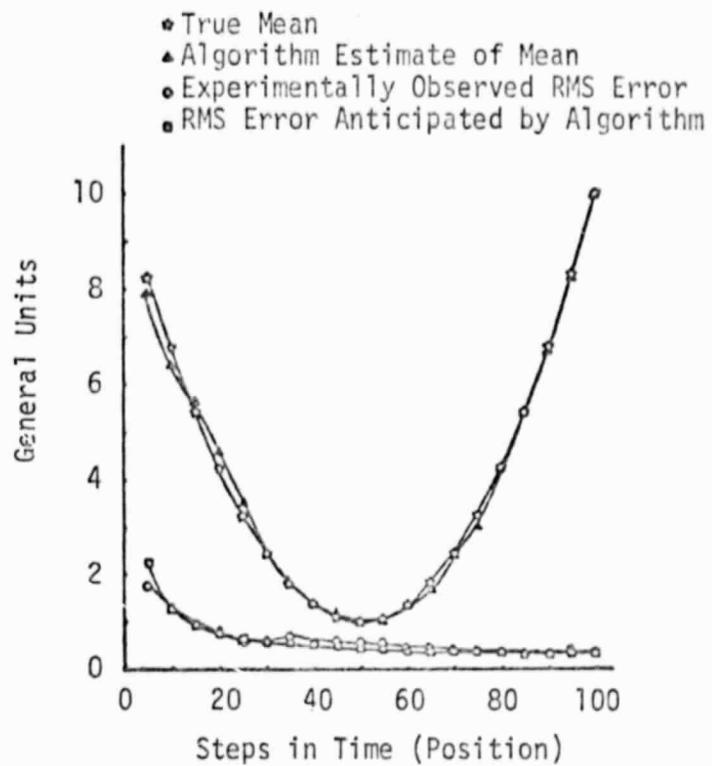


Figure 2. Performance of PF (polynomial fit) algorithm.

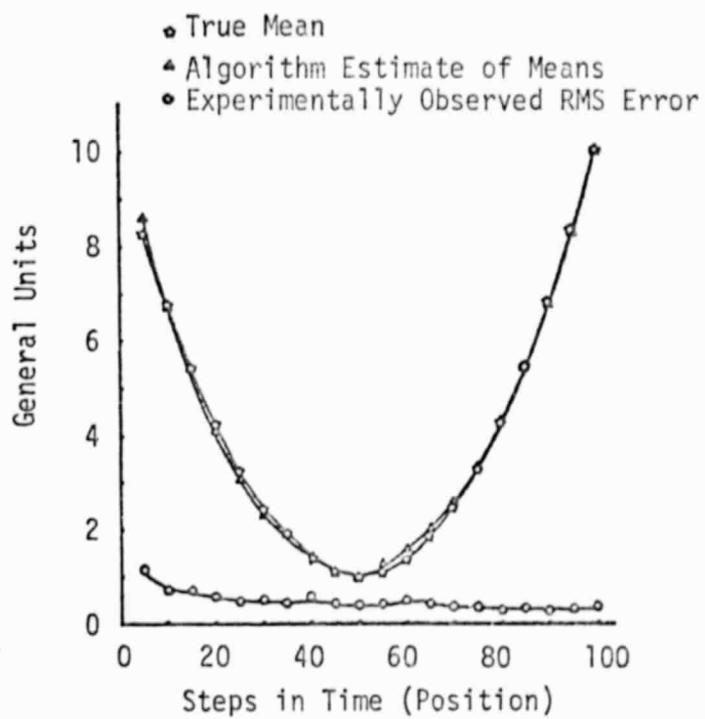


Figure 3. Performance of estimator operating as a least mean square error curve fit.

Step 1 involves specifying the initial covariance matrices for  $M$  classes  $C_i$ ,  $i = 1, 2, \dots, M$ , a zero matrix  $\phi_i$ ,  $i = 1, 2, \dots, M$  of equal dimensionality to the  $C$  matrices, and a counter  $N_i$ ,  $i = 1, 2, \dots, M$ . Each counter should be initialized to zero.

In step 2 specify the number of samples  $P$  (where  $P >$  dimensionality of pattern vectors) to be used in producing new estimates of the covariance matrix of each class.

At step 3 classify a pattern using the covariance estimate  $C_i$ ,  $i = 1, 2, \dots, M$  for the classification.

During step 4, if the pattern was classified into class  $j$ , update  $\phi_j$  according to the relation

$$\begin{aligned} \phi_i(N_i+1) = & \frac{1}{N_i+1} [N_i \phi_i(N_i) + N_i m_i(N_i) m_i^{\sim}(N_i) \\ & + Y(N_i+1) Y^{\sim}(N_i+1) - \frac{1}{(N_i+1)^2} (N_i m_i(N_i) \\ & + Y(N_i+1))] + (N_i m_i(N_i) + Y(N_i+1)^{\sim} \end{aligned} \quad (26)$$

where  $m_i$  represents a mean estimate.

Step 5, increment the counter  $N_j$  by one.

Step 6, if  $N_j$  is less than  $P$ , go to step 3, otherwise, go to step 7.

At step 7, replace  $C_j$  by  $\phi_j$ . Reset  $\phi_j$  to the zero matrix and rezero the counter  $N_j$ . Go to step 3.

$P$  is chosen greater than the pattern dimensionality in order to insure that the estimate of the covariance will possess an inverse given that the samples are drawn from a normal population [1]. Justification of equation (26) is given in Appendix A.

## CHAPTER III

### A DIVERGENCE CRITERION

A problem associated with the CF algorithm and also the PF class of algorithms is that their derivations assume the parameters to be estimated vary as some finite degree function. A PF algorithm of very high degree, and hence great flexibility is cumbersome to derive and to run; likewise, computer execution time increases as the degree of algorithm complexity is increased. If the parameter being estimated changes with time in a way more complex than assumed by the algorithm, the predictions of stochastic approximation techniques may diverge. Although the "weak memory" inherent in stochastic approximation will compensate somewhat for this problem, it would be desirable to more strongly limit the memory by restarting the algorithm at the point of divergence, resulting in a piecewise implementation of an estimator. A technique for detecting divergence and restarting the particular stochastic approximation algorithm in the area of divergence is necessary. This restart capability should be provided external to the function of the particular stochastic approximation technique being implemented. In other words, what is needed is a "monitor" for the operation of the algorithm.

Consider the problem of the estimation of the unknown mean of some distribution  $Y \sim N(\theta(n), \sigma^2)$ , where the mean  $\theta(n)$  varies with time. To assume that this function  $\theta(n)$  might be approximated by segments would not be unreasonable. The quantity  $X_n$  will be considered

the approximation to  $\theta(n)$  made by a stochastic approximation algorithm. Associated with each time interval is a random variable  $Y$  with variance  $\sigma^2$ . The  $n^{\text{th}}$  sample value of  $Y$  shall be referred to as  $y_n$ . If the particular stochastic approximation algorithm accurately estimates  $\theta(n)$  based on  $y_n$  in some region, it is then possible to define a new, time invariant random variable  $Z \sim N(u_z, \sigma^2)$ , where the samples  $z_n$  are given by

$$z_n = y_n - x_n . \quad (27)$$

However, if  $x_n$  is an accurate estimate of  $\theta(n)$ , it is clear that  $u_z$  will be zero. A statistical inference built around the notion of a "confidence interval" for a known statistic of the distribution function  $Z$  may now be made [10,11]. Let the average value of  $Z$  be calculated by the algorithm

$$\bar{z} \equiv \frac{1}{n} \sum_{i=1}^n (Y_i - X_i) . \quad (28)$$

It can be shown that

$$\Pr\left[ -\frac{3\sigma}{\sqrt{n}} < \bar{z} < \frac{3\sigma}{\sqrt{n}} \right] \approx 1 \quad (29)$$

(see Appendix B). Therefore, the statistic  $\bar{z}$ , which is nothing more than the average of the difference of the random sample patterns and

the corresponding estimates of their means, may serve as an indication of divergence.

The restart "monitor" may thus be implemented as follows. If the "confidence interval" condition

$$|\bar{z}| < \frac{3\sigma}{\sqrt{n}} \quad (30)$$

is violated, the algorithm should be restarted at that point ( $n$  should be reset to one).

The interval in which  $\bar{z}$  must lie is reduced in proportion to  $1/\sqrt{n}$ . The maximum rate at which a stochastic approximation of a quantity may converge to the true value is in proportion to  $1/n$ , the harmonic sequence, and still satisfy equation (10) of Chapter II [8]. If the  $\gamma$  sequence of equations (9) and (11) approaches the harmonic sequence in the limiting case, it would also be desirable to reduce the confidence interval around  $\bar{z}$  in proportion to  $1/n$ . However, were the interval around  $\bar{z}$  reduced in proportion to  $1/n$ , the probability of divergence would no longer remain approximately equal to one, nor would it remain constant for each value of  $n$ .

The effect of the divergence criterion developed above is to increase the sensitivity to divergence as much as possible while maintaining a constant probability of successfully detecting divergence. In particular this technique has the advantage that it may be used to monitor any estimation algorithm, no matter what degree of complexity was assumed in algorithm derivation.

One point of interest concerns the variance of  $Z$ . Since accurate estimation forms the basis for the confidence interval concept, inaccurate estimation will result in the variance associated with  $Z$  being larger than  $\sigma^2$ . The resulting divergence criterion may be stricter than anticipated. This problem may be circumvented by making the criterion more lax (for example, by increasing the interval length around  $\bar{z}$  to  $\pm 4\sigma/\sqrt{n}$ ).

An alternative method for testing for divergence would be to make use of the estimates of mean square error  $\overline{e_n^2}$  made by the algorithms as discussed in Chapter II. If  $\sqrt{\overline{e_n^2}}$  could be considered a measure of the error between  $x_n$  and the true value  $\theta(n)$  and  $\sigma$  a measure of the error between  $y_n$  and the true value  $\theta(n)$  then  $x_n$  and  $y_n$  should differ at most by  $\sqrt{\overline{e_n^2}} + \sigma$ . An algorithm restart, with  $n$  reset to one, could be made at the point where

$$K|x_n - y_n| > \sqrt{\overline{e_n^2}} + \sigma$$

with  $K$  a constant factor (for example,  $K = 2$ ).

Another possible variation on this idea might be to use both the original divergence criterion (confidence interval) together with this latter relation in combination.

## CHAPTER IV

### ADAPTIVE RECOGNITION AND BOUNDARY

#### DEFINITION PROGRAM

An adaptive Bayes classifier is realized by incorporating within the ordinary Bayesian classification program estimation operations which optimally estimate statistics for the next classification time. An application suggested was that the adaptive classifier might be useful in locating or defining spatial boundaries (not to be confused with the Bayes decision surface or boundary) between data classes. A physical example would be the definition of the shoreline between a body of water and a land mass; varying means would then correspond to spectral shifts of scan data caused by transition from deep water to shallow water near the shoreline. As a test, different data sets have been generated, each having two equally likely data classes. These data sets are composed of patterns synthetically produced to simulate a  $128 \times 128$  pixel frame of four dimensional Gaussian spectral scan data.

Adaptive classification and boundary definition programs have been developed which treat each of the 128 individual horizontal rows as a separate, independent classifier test. These programs utilize the CF and the second degree PF algorithms to adapt to changing class mean vectors. Updated estimates of the covariance matrix for each class are made using the recursive estimation technique discussed in Chapter II.

A general flow chart of program operation is shown in Figure 4. Program initialization is accomplished by specifying an appropriate disk file of input data for classification, specifying a disk file to contain output boundary results for video display, specifying initial estimates of the mean vectors and covariance matrices of the two classes, and inputting a decision variable. The process of classification and subsequent boundary definition then begins.

A 128 pattern row of data is read into memory from the input disk file, each pattern of which is four dimensional. Patterns are classified by a Bayesian classification subroutine. The classifier returns the variable ICLASS as a one or a two to indicate that the pattern has been assigned to class one or class two.

In order to determine whether or not a boundary between the two classes has been crossed in a row test, a stack, whose length is assigned by the specification of the decision variable at initialization, is used. ICLASS associated with the first classified pattern of a row is stored and also pushed onto the stack. The value of ICLASS associated with each successive classified pattern is pushed onto the stack. Only when the stack is full may a decision be made as to whether or not a boundary has been crossed. At that time, and subsequent times, each element of the stack is examined; if more than half of the members of the stack have values equal to that of the ICLASS of the first classified pattern of the row, the boundary definition algorithm decides no boundary has been crossed. If more than half of the members of the stack differ from the ICLASS of the

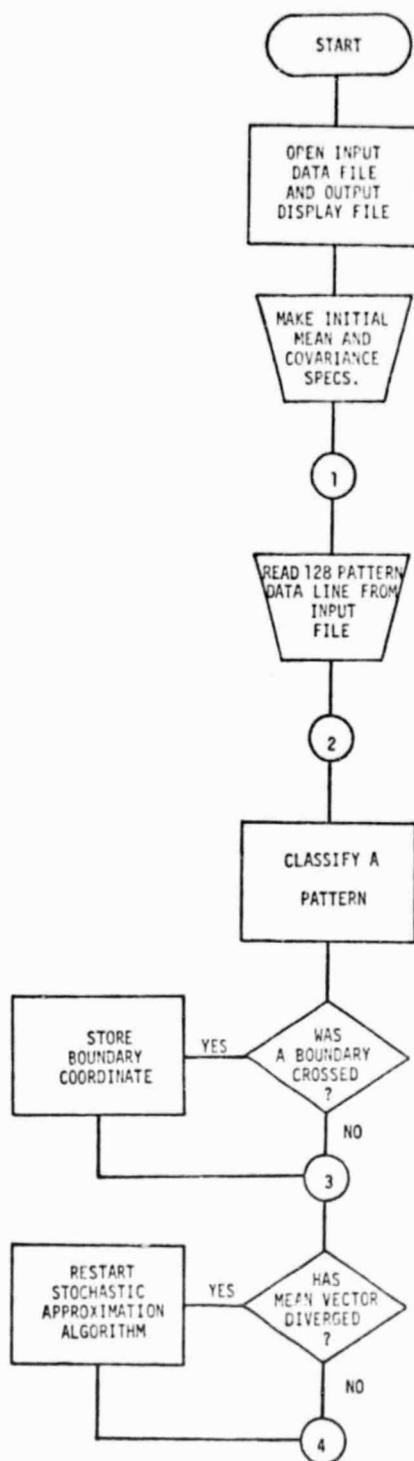


Figure 4. A general flowchart of classification and boundary definition program operation.

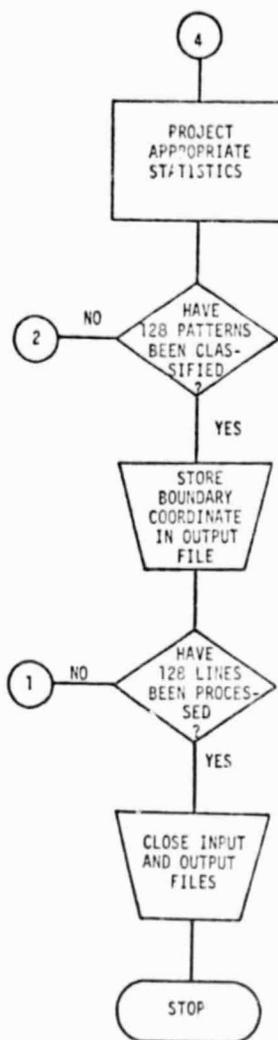


Figure 4. (continued)

first classified pattern of the row, the algorithm decides a boundary has been crossed and the value stored for the ICLASS of the first classified pattern of the row is replaced by the ICLASS of the new class which has been encountered. The appropriate boundary address is stored and the same process continues for the remainder of the row.

After classifying each pattern and performing the boundary test, the divergence criterion of Chapter III may be employed to determine whether or not the estimates of particular mean vector components have diverged. Divergence of a mean vector component requires a restart of the estimation algorithm for that component in the area of divergence.

As each pattern is classified, class statistics for the appropriate class must be projected ahead for the next classification by either the CF or the PF algorithms and the recursive form for the covariance estimation. Upon completion of a 128 pattern row test, boundary information is written into the disk output file, the next row of input data is read, and the process is begun on the unclassified row.

This procedure is repeated until classification and subsequent boundary definition of all 128 rows is accomplished. Upon completion, all input and output disk files are closed and program execution terminates.

Appendices C and D each contain a compiled Fortran IV program listing of two different version of an adaptive Bayes classifier. The numbers at the leftmost side of the listings correspond to the internal

sequence or statement numbers supplied by the Digital Equipment Corporation RT-11 Operating System FORTRAN Compiler. These statement numbers will be used in reference to particular statements.

The first version of the classifier (Appendix C) incorporates the modified CF, algorithm to adaptively estimate class mean vectors, the confidence interval divergence criterion to test for divergence of mean estimates, and the recursive form of covariance estimation. In order to adapt to class mean vectors only and check for their divergence, the statement corresponding to line 117 of the main program should be deleted. To adapt to mean vectors only and neglect the possibility of their divergence, statements corresponding to line numbers 62 through 115 as well as line 117 of the main program should be deleted. To implement the unmodified CF algorithm to adapt mean vectors only, statements corresponding to lines 6, 7, 12, 16, 17, and 22 of SUBROUTINE PROJECT and lines 62 through 115 and also line 117 of the main program should be deleted. An ordinary Bayes classifier (non-adaptive) may be implemented by deletion of lines 62 through 117 of the main program.

The second version of the classifier (Appendix D) incorporates the second degree PF algorithm to adaptively estimate class mean vectors and the recursive form of covariance estimation. In order to adapt to class mean vectors only, the statement corresponding to line 65 should be deleted. To implement an ordinary (non-adaptive) Bayes classifier, the statements corresponding to lines 64 and 65 may be deleted.

## CHAPTER V

### RESULTS

Five data sets have been synthesized to simulate five  $128 \times 128$  pixel multispectral scan data frames [4]. These data sets are each composed of two classes of four dimensional Gaussian data. A photograph depicting the true spatial boundary between the two classes is shown in Figure 5. The area to the left of this wedge shaped boundary is referred to as class one; similarly, the area to the right of the boundary is class two. The shortest and longest rows of data for each class are 32 and 96 patterns.

Individual rows of data were generated a row at a time from left to right. Data sets one and two were both generated with all four class one mean components varying according to the relation

$$\frac{5}{1024} (N-32)^2 + 5$$

from the left edge of the frame to the boundary (N is simply the position index having an initial value of zero at the left edge of the frame and incremented by one at each position to the right). A plot of this relation versus N is shown in Figure 6. Class two data was generated for the remainder of each row. Class two of data set one was generated having the constant mean vector

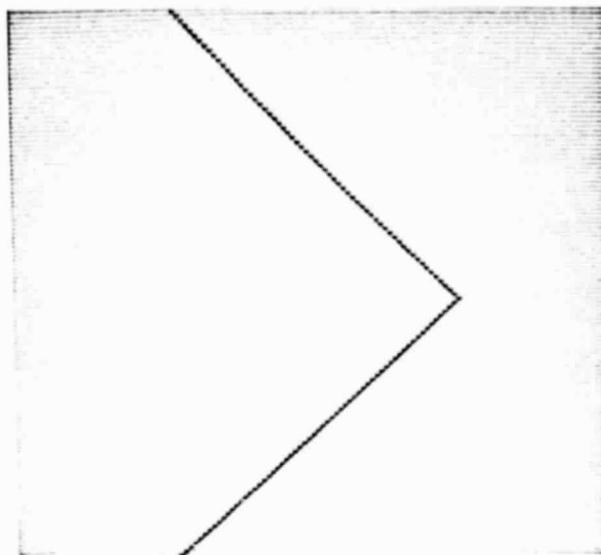


Figure 5. True spatial class boundary.

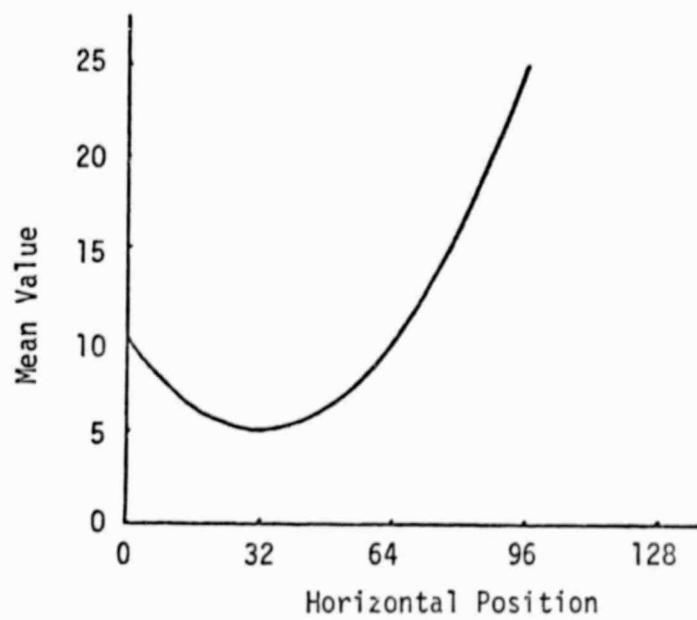


Figure 6. Variation of class one mean with position for data sets 1, 2, 4, and 5.

$$\begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

while the mean vector associate with class two of data set two is

$$\begin{bmatrix} 1.5 \\ 1.5 \\ 1.5 \\ 1.5 \end{bmatrix} .$$

The covariance matrices of both classes of data sets one and two are

$$\begin{bmatrix} 1 & .5 & .5 & .5 \\ .5 & 1 & .5 & .5 \\ .5 & .5 & 1 & .5 \\ .5 & .5 & .5 & 1 \end{bmatrix} .$$

Data set three was generated with the four class one mean components varying according to the relation

$$7.5 + 2.5 \cos (.1047N)$$

from the left edge of the frame to the boundary (N again denotes a positional index). A plot of this relation is shown in Figure 7.

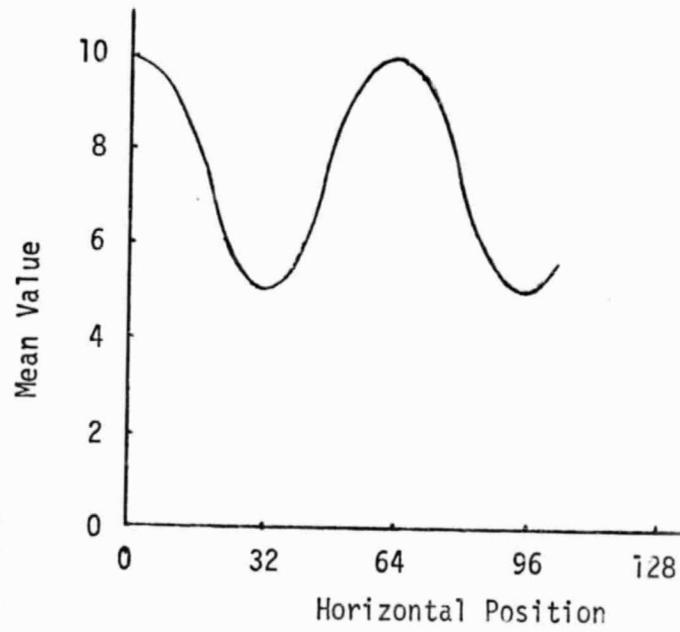


Figure 7. Variation of class one mean with position for data set 3.

Class two data have been generated for the remainder of each row having the constant mean vector

$$\begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \end{bmatrix} .$$

The covariance of both class one and two of data set three is the same as was specified for data sets one and two.

Data sets four and five were generated having class one and two means specified in exactly the same manner as data set one. In addition, however, each term of the covariance matrix of the class one data was changed in a linear manner according to the equation

$$c_{ij}(N) = c_{ij}(0) + m N$$

$$i = 1, \dots, 4$$

$$j = 1, \dots, 4$$

where  $m$  is simply a slope factor. In other words a linear scalar function of position is added to each term of the initial covariance. For data sets four and five the initial class one covariance was

$$c(0) = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} .$$

Covariance matrix elements of class one, data set four were varied with a slope  $m$  of 0.02 while like elements of data set five changed with a slope of 0.2. Covariance matrices for class two of data sets four and five were both specified as

$$\begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} .$$

Eight different classification and boundary definition programs have been applied to the problem of striking the boundary separating the two classes in each of the five data sets. Each program requires initial estimates of the mean vectors and covariance matrices of the two classes. Because the estimation algorithms predict well, initial mean vector estimates may be made by training over a small area. The effect is to hold sample scatter to a minimum while providing reasonable estimates of the mean.

The first program, referred to as BAYES1, implements an ordinary, non-adaptive Bayes classifier. The initial estimates of class mean vectors and covariance matrices are incorporated throughout classifi-

cation of a complete data set and the resulting data file containing boundary information may be displayed by the DATA-DISK video system.

The second program, BAYES2, employs the CF algorithm in its original form to adaptively estimate mean vectors for a Bayes classifier. Boundary data is subsequently deduced and stored for display.

Program number three, BAYES3, utilizes the modified CF algorithm discussed in Chapter II to produce up-to-date estimates of changing mean vectors for a Bayes classifier.

BAYES4 incorporates not only the modified CF algorithm, but also the confidence interval divergence criterion introduced in Chapter III to adaptively estimate class mean vectors for the classifier.

BAYES5 implements the second degree PF algorithm to adaptively project estimates of class mean vectors for a Bayesian classifier. BAYES6, BAYES7, and BAYES8 take the same form as BAYES3, BAYES4, and BAYES5, respectively, with the exception that BAYES6 through BAYES8 also employ the recursive covariance estimation technique.

Table I provides a cross-reference summary relating Figures 8 through 47 to the particular data sets and programs. Each figure is also individually identified by the program name and data set number used. These figures are photographs of boundaries defined by the various programs for each data set.

A comparison of the results obtained applying the various programs to the different data sets reveals that the ability to adapt to changing mean vectors is essential to successful classification. False boundaries have been generated in each case where the non-

TABLE I

A CROSS-REFERENCE OF FIGURES DEPICTING RESULTS OBTAINED  
UPON APPLICATION OF THE CLASSIFICATION AND  
BOUNDARY DEFINITION PROGRAMS TO THE  
VARIOUS DATA SETS

PROGRAM	DATA SETS				
	1	2	3	4	5
BAYES1	Figure 8	Figure 16	Figure 24	Figure 32	Figure 40
BAYES2	Figure 9	Figure 17	Figure 25	Figure 33	Figure 41
BAYES3	Figure 10	Figure 18	Figure 26	Figure 34	Figure 42
BAYES4	Figure 11	Figure 19	Figure 27	Figure 35	Figure 43
BAYES5	Figure 12	Figure 20	Figure 28	Figure 36	Figure 44
BAYES6	Figure 13	Figure 21	Figure 29	Figure 37	Figure 45
BAYES7	Figure 14	Figure 22	Figure 30	Figure 38	Figure 46
BAYES8	Figure 15	Figure 23	Figure 31	Figure 39	Figure 47

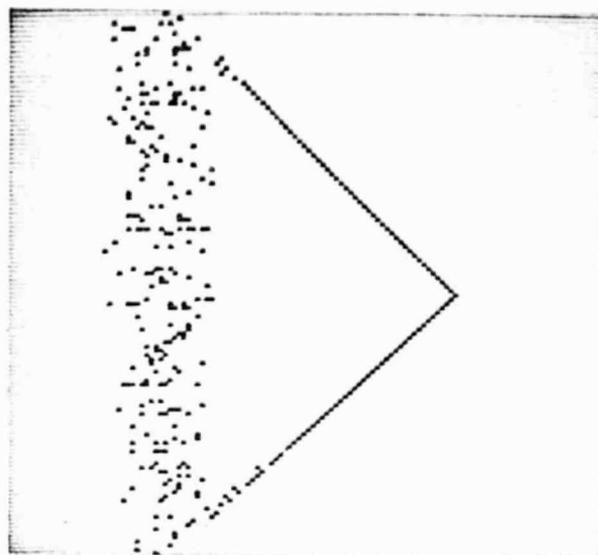


Figure 8. Spatial boundaries resulting from the application of an ordinary Bayes classifier (BAYES1) to data set 1. Note the false boundaries.

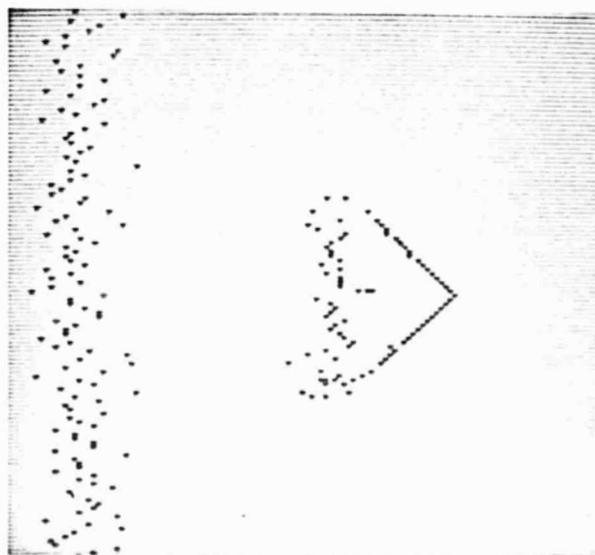


Figure 9. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES2) using the CF algorithm to data set 1. Note the false boundaries.

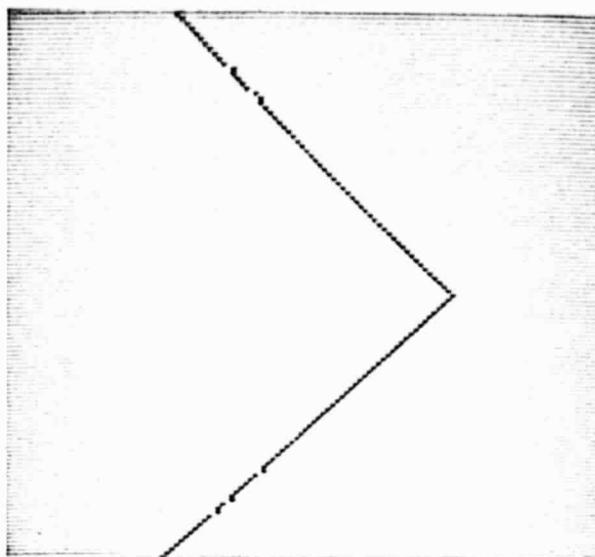


Figure 10. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES3) using the modified CF algorithm to data set 1.

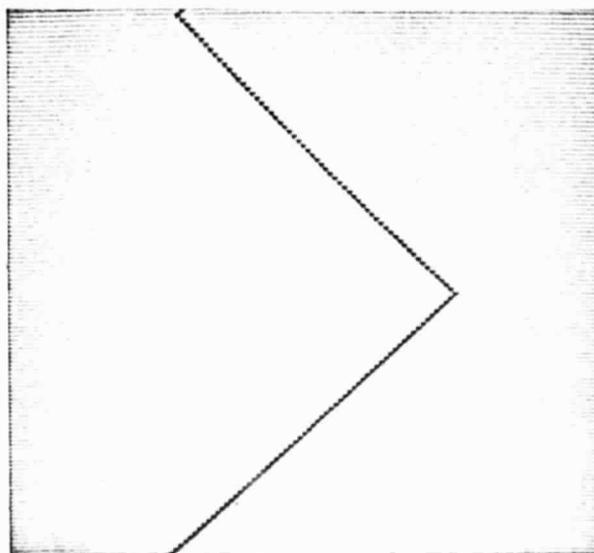


Figure 11. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES4) using the modified CF algorithm and the divergence criterion to data set 1.

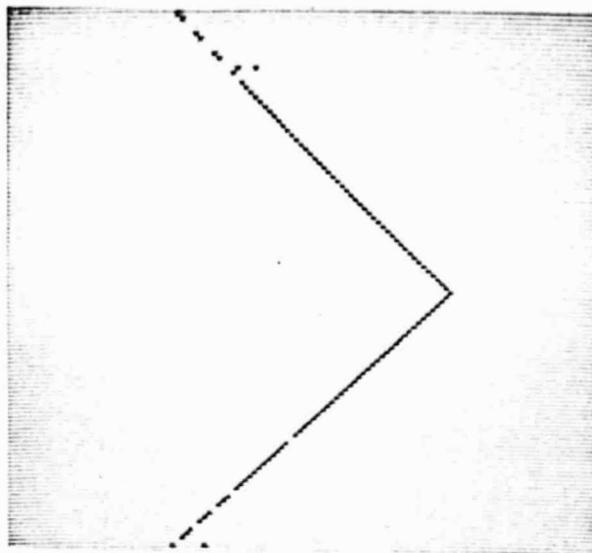


Figure 12. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES5) using the PF algorithm to data set 1.

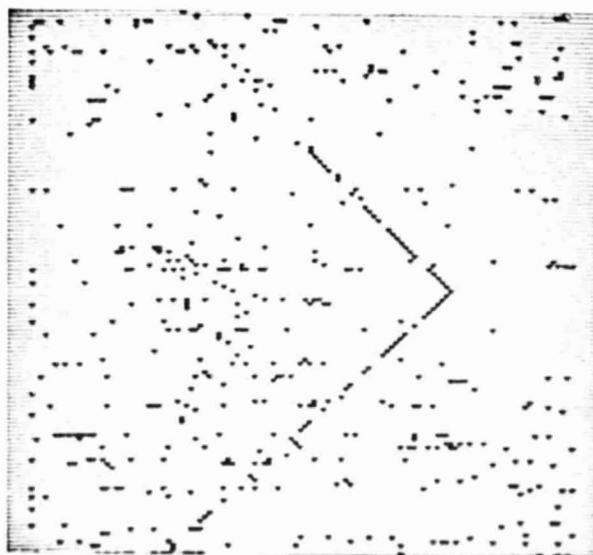


Figure 13. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES6) using the modified CF algorithm and recursive covariance estimation to data set 1.



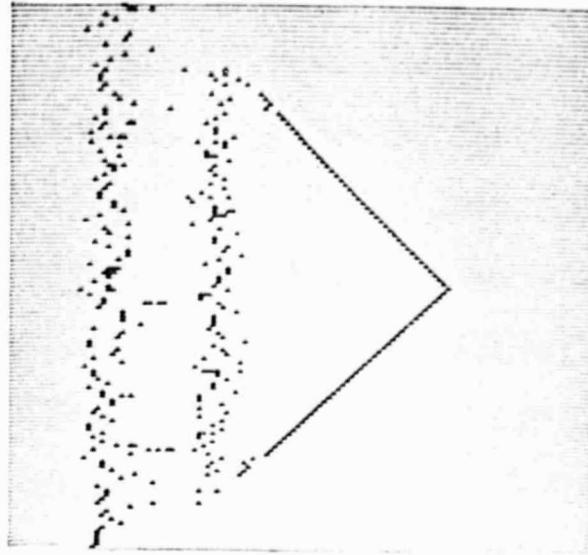


Figure 16. Spatial boundaries resulting from the application of an ordinary Bayes classifier (BAYES1) to data set 2. Note the false boundaries.

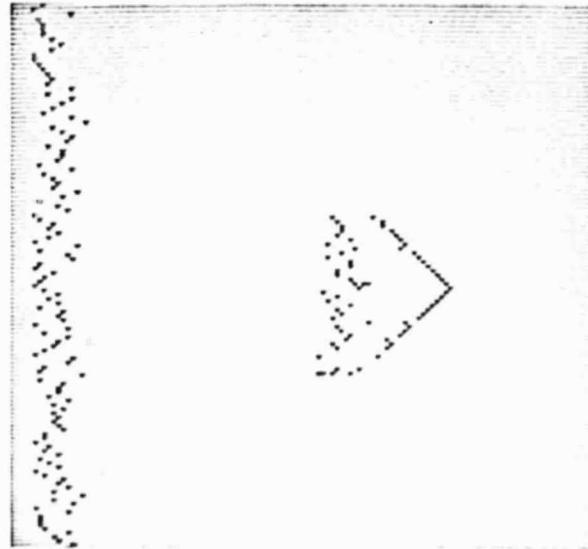


Figure 17. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES2) using the CF algorithm to data set 2. Note the false boundaries.



Figure 18. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES3) using the modified CF algorithm to data set 2.

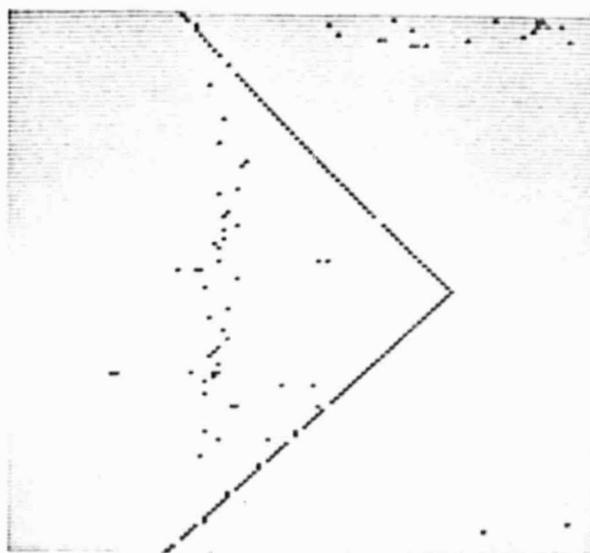


Figure 19. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES4) using the modified CF algorithm and the divergence criterion to data set 2.

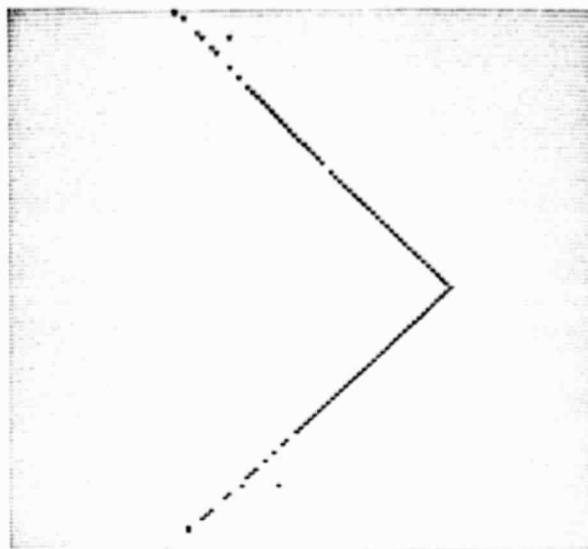


Figure 20. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES5) using the PF algorithm to data set 2.



Figure 21. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES6) using the modified CF algorithm and recursive covariance estimation to data set 2.

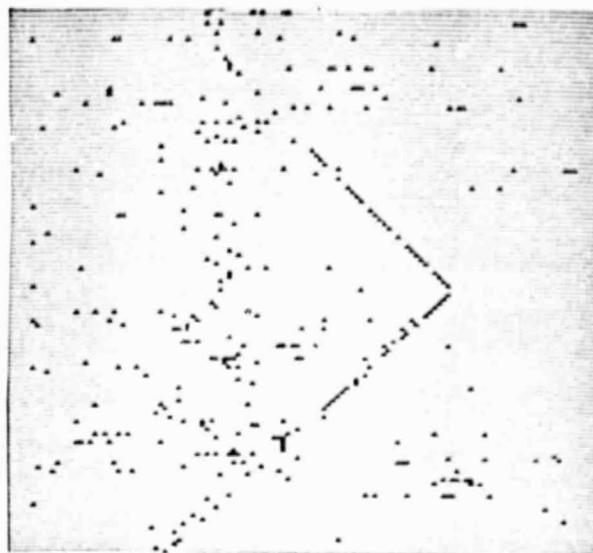


Figure 22. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES7) using the modified CF algorithm, divergence criterion, and recursive covariance estimation to data set 2.

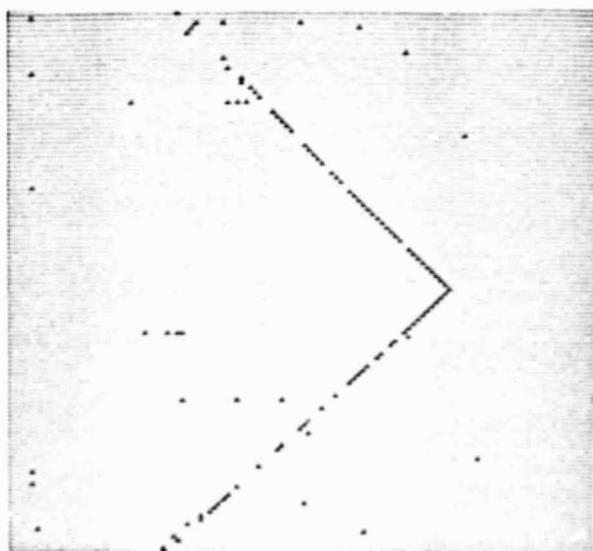


Figure 23. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES8) using the PF algorithm and recursive covariance estimation to data set 2.

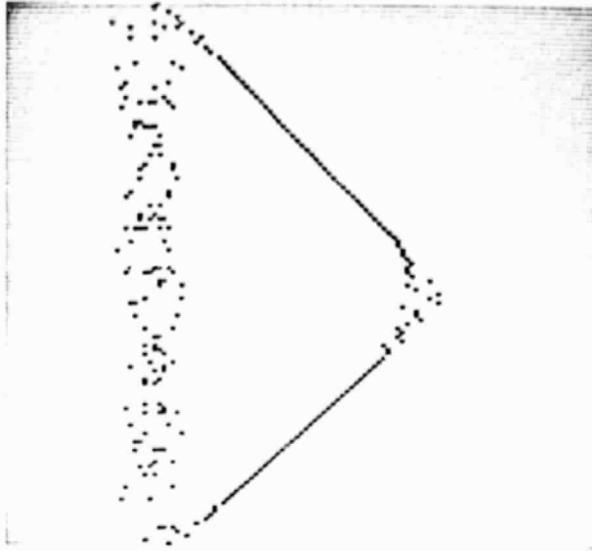


Figure 24. Spatial boundaries resulting from the application of an ordinary Bayes classifier (BAYES1) to data set 3. Note the false boundaries.



Figure 25. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES2) using the CF algorithm to data set 3. Note the false boundaries.

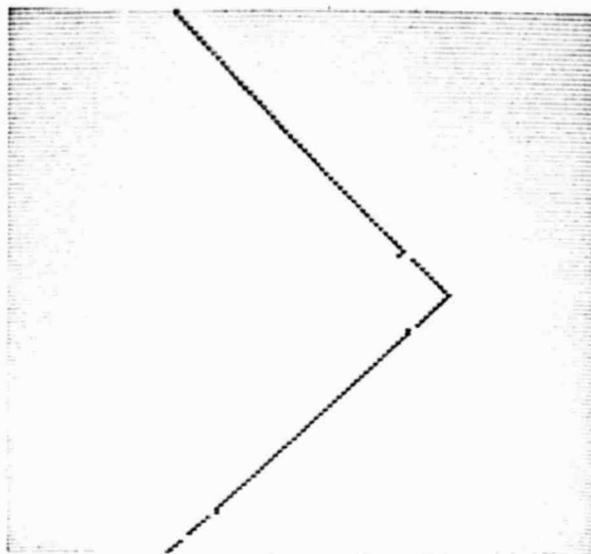


Figure 26. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES3) using the modified CF algorithm to data set 3.

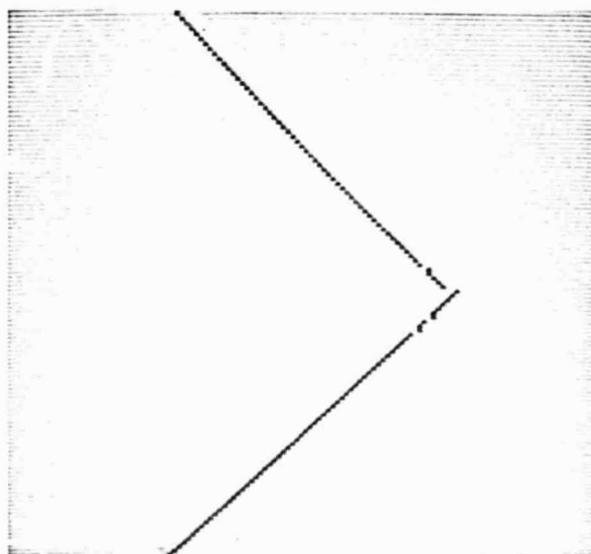


Figure 27. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES4) using the modified CF algorithm and the divergence criterion to data set 3.

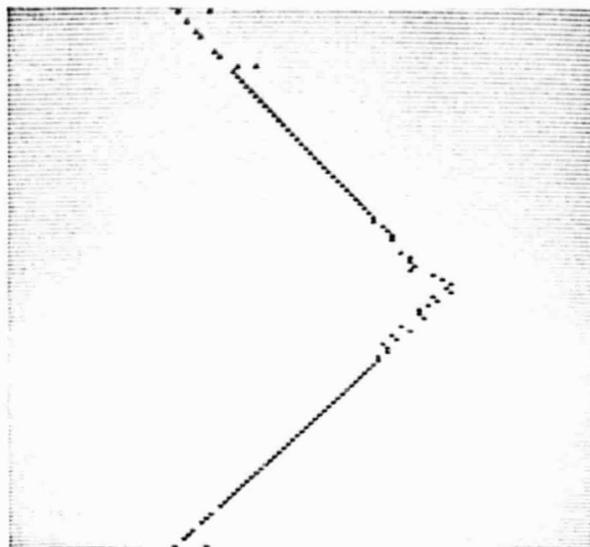


Figure 28. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES5) using the PF algorithm to data set 3.

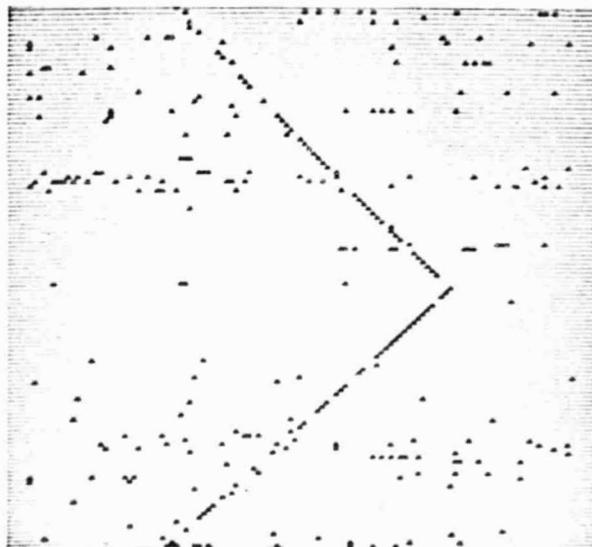


Figure 29. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES6) using the modified CF algorithm and recursive covariance estimation to data set 3.

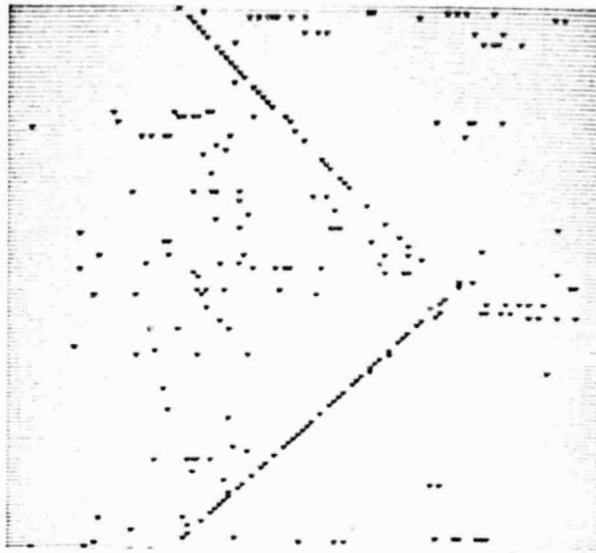


Figure 30. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES7) using the modified CF algorithm, divergence criterion, and recursive covariance estimation to data set 3.

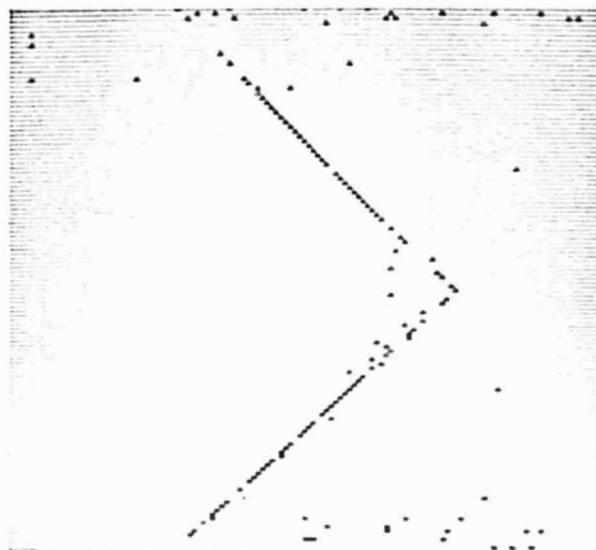


Figure 31. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES8) using the PF algorithm and recursive covariance estimation to data set 3.

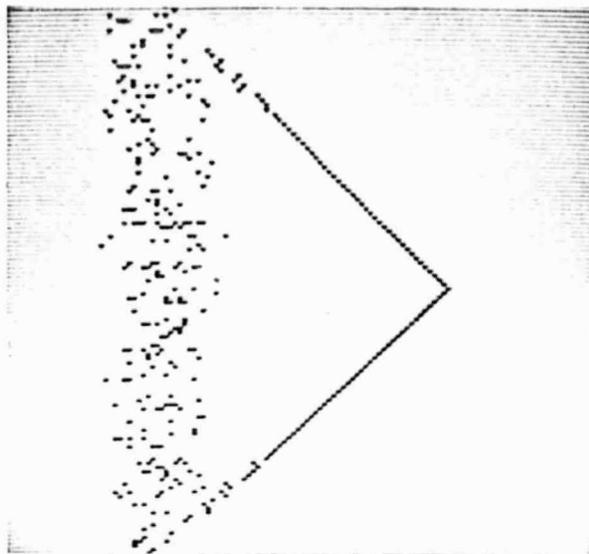


Figure 32. Spatial boundaries resulting from the application of an ordinary Bayes classifier (BAYES1) to data set 4. Note the false boundaries.

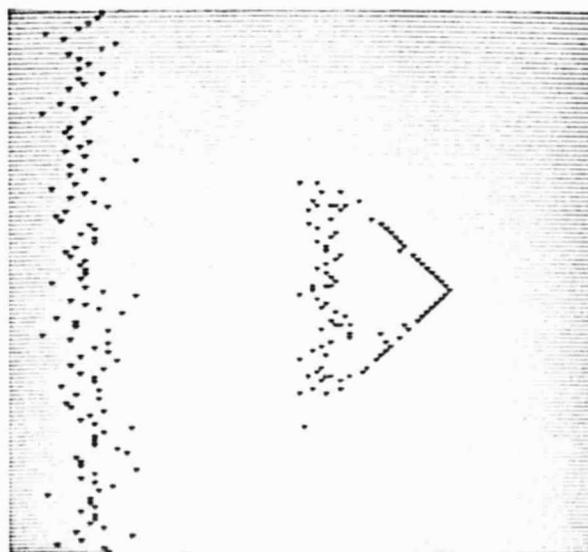


Figure 33. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES2) using the CF algorithm to data set 4. Note the false boundaries.

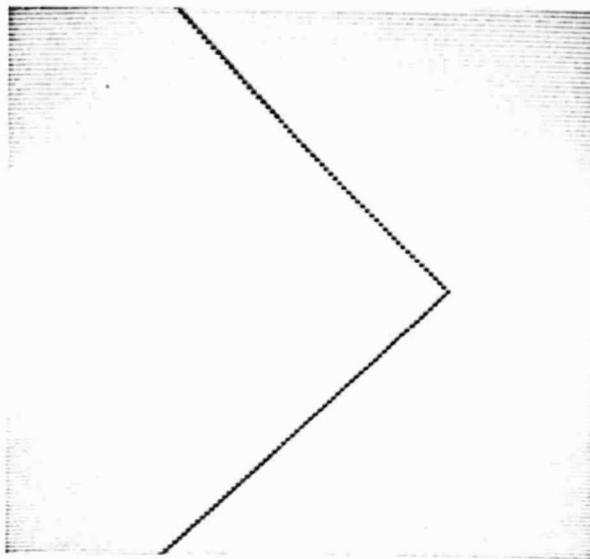


Figure 34. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES3) using the modified CF algorithm to data set 4.

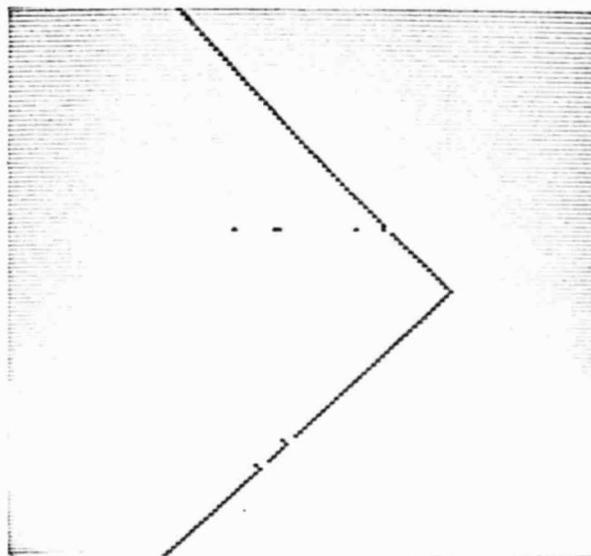


Figure 35. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES4) using the modified CF algorithm and the divergence criterion to data set 4.

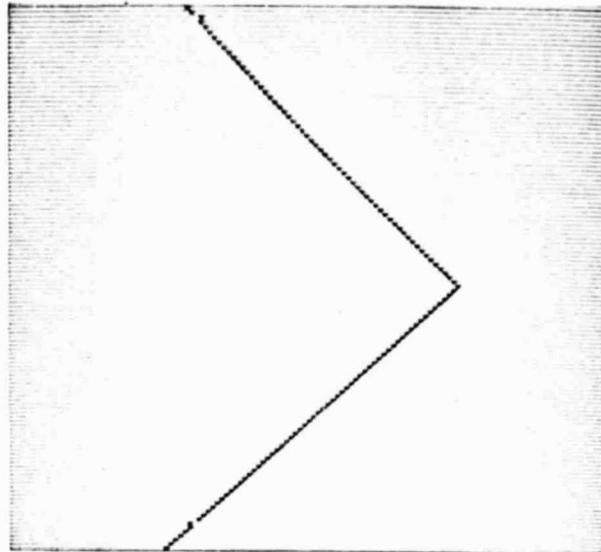


Figure 36. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES5) using the PF algorithm to data set 4.

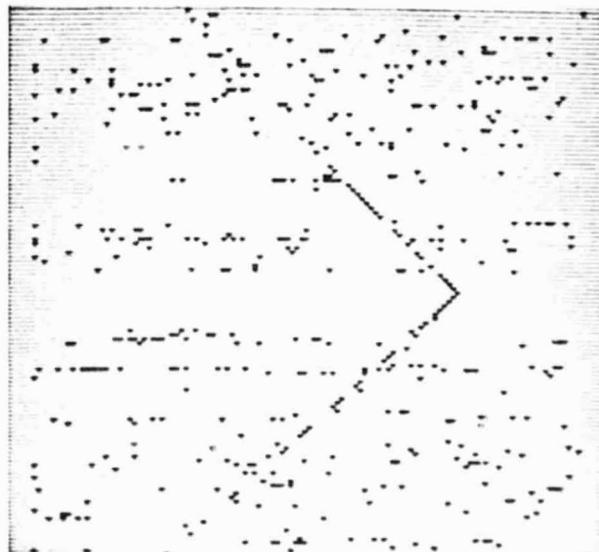


Figure 37. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES6) using the modified CF algorithm and recursive covariance estimation to data set 4.



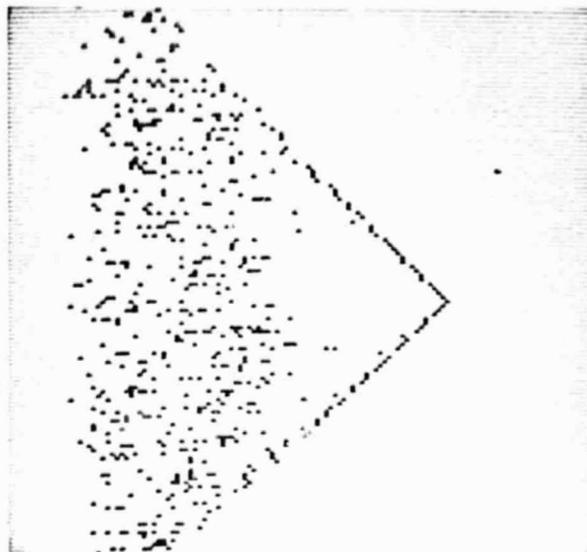


Figure 40. Spatial boundaries resulting from the application of an ordinary Bayes classifier (BAYES1) to data set 5. Note the false boundaries.

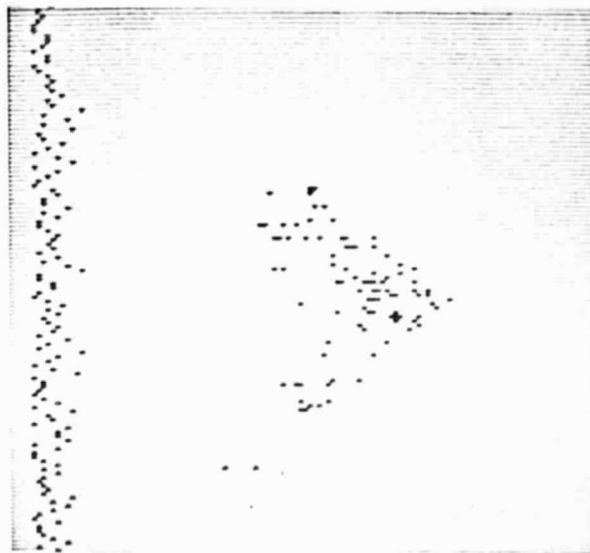


Figure 41. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES2) using the CF algorithm to data set 5. Note the false boundaries.

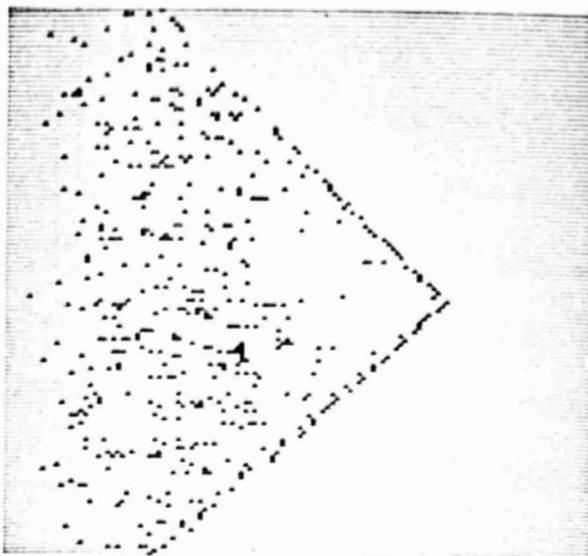


Figure 42. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES3) using the modified CF algorithm to data set 5. Note the false boundaries.



Figure 43. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES4) using the modified CF algorithm and the divergence criterion to data set 5. Note the false boundaries.

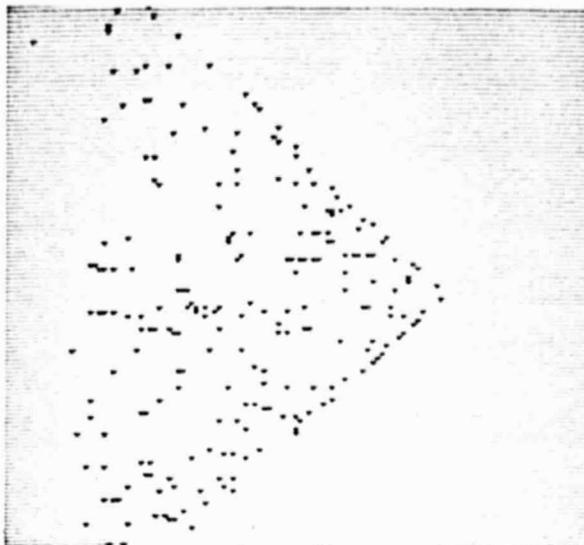


Figure 44. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES5) using the PF algorithm to data set 5. Note the false boundaries.

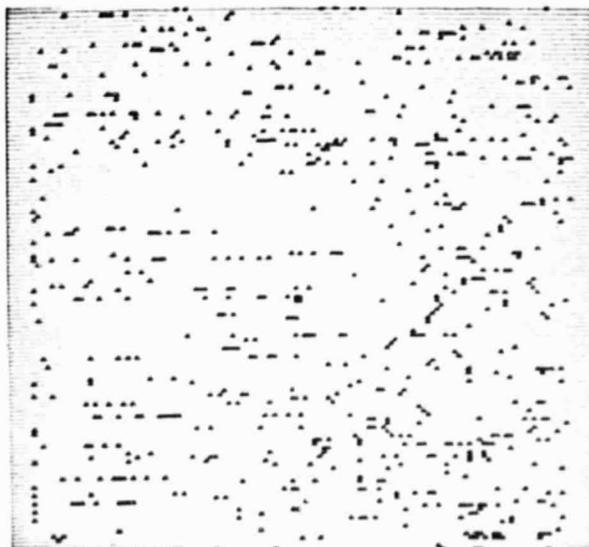


Figure 45. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES6) using the modified CF algorithm and recursive covariance estimation to data set 5. Note the false boundaries.



Figure 46. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES7) using the modified CF algorithm, divergence criterion, and recursive covariance estimation to data set 5. Note the false boundaries.

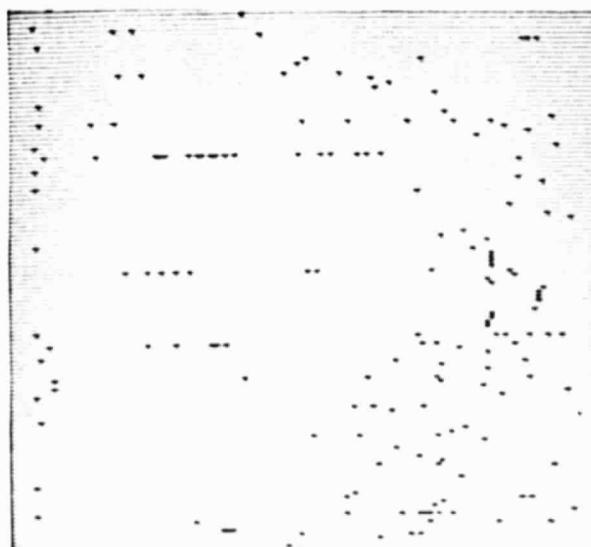


Figure 47. Spatial boundaries resulting from the application of an adaptive Bayes classifier (BAYES8) using the PF algorithm and recursive covariance estimation to data set 5. Note the false boundaries.

adaptive classifier, BAYES1, has been applied to each data set so as to obscure the form of the true boundary. No improvement in the boundary definition results in using the original form of the CF algorithm, implemented in BAYES2, as the adaptive estimator of mean vectors.

Significant improvement in boundary definition performance results have been achieved by BAYES3 which employs a modified CF algorithm to adaptively estimate class means. The modification (see Chapter II) consists of subtracting the initial value of the mean vector of the classified pattern from the pattern and current mean, applying the CF algorithm to the results, and finally adding to the projected mean estimate made by the CF algorithm the initial mean value.

Still more improvement resulted upon application of BAYES4 which incorporates the modified CF algorithm and also the confidence interval divergence criterion of Chapter III. The effect of employing this criterion was to restart the modified CF algorithm (as if  $n$  were 1 again) if the condition

$$\left| \frac{1}{n} \sum_{i=1}^n (X_i - Y_i) \right| < \frac{3\sigma}{\sqrt{n}}$$

was violated. Most of the erroneous boundary points appear at points where restarts were made, due to poor initial tracking when the algorithm is first started with little prior training.

Boundary definition resulting from the application of BAYES5 is very satisfactory, especially in the case of data set two where the degree of overlap is large.

In no case were the results obtained using BAYES6, BAYES7, and BAYES8 of equal quality to the results obtained using BAYES3, BAYES4, and BAYES5. Even in the case of data set four in which a class covariance was changing slowly, the programs which ignored the fact that a covariance matrix could be position dependent proved superior.

In data set five the covariance associates with class one grew at such a rapid rate that class one quickly overlapped class two resulting in an impossible situation for each of the eight techniques.

## CHAPTER VI

### CONCLUDING OBSERVATIONS

PF type algorithms represent an algorithm class that predicts well; a second degree PF algorithm was used as an example in this thesis but algorithms of this class can be derived (with different  $\hat{S}$  and  $\gamma$  formulas) for tracking parameters that vary with time as an  $n^{\text{th}}$  degree polynomial. In order to achieve the best results, the degree of polynomial assumed in algorithm derivation should be of approximately the same order as anticipated variation. A PF type algorithm can also track variations not of the exact form assumed because of the limited memory characteristic of the "refine" step [7]. Another advantage of the PF class of algorithms is that shifts between algorithms of different complexity can be effected in mid-operation since the output of any PF algorithm (the error estimate and projected parameter estimate) along with the next data sample may serve as the input to any other algorithm of the PF form.

Modifications of the CF algorithm have also been found suitable for tracking varying parameters. Although a PF algorithm may produce better estimates than modified versions of the CF algorithm, especially as the degree of class overlap is increased, the additional memory required for storage of past history may in some cases prohibit use of the PF form and warrant utilization of the modified CF algorithm.

Much work has been done in the area of state estimation in controls engineering. Kalman predictors are capable of producing

statistically optimum state estimates when measurements and inputs are stochastic in nature [12,13]. The Kalman predictor has been found similar in nature to estimation algorithms presented in this thesis. An area for future study lies in investigating the possibility of modifying the Kalman predictor to account for changing covariance.

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## LIST OF REFERENCES

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APPENDICES

## APPENDIX A

### COVARIANCE ESTIMATION

Letting  $C(N)$  represent the estimate of the covariance for  $N$  samples,

$$C(N) = \frac{1}{N} \sum_{j=1}^N y_j y_j' - m(N) m'(N)$$

where the expected value of  $y$  has been approximated by the sample average  $m(N)$  [1].

$$\begin{aligned} C(N+1) &= \frac{1}{N+1} \sum_{j=1}^{N+1} y_j y_j' - m(N+1) m'(N+1) \\ &= \frac{1}{N+1} \left( \sum_{j=1}^N y_j y_j' + y_{N+1} y_{N+1}' \right) - m(N+1) m'(N+1) \\ &= \frac{1}{N+1} (NC(N) + N m(N) m'(N) + y_{N+1} y_{N+1}') \\ &\quad - \frac{1}{(N+1)^2} (N m(N) + x_{N+1})(N m(N) + x_{N+1})' \end{aligned}$$

This expression provides a convenient method for estimating or updating the covariance matrix, starting with  $C(1) = y_1 y_1' - m(1) m'(1)$ . Since  $m(1) = y_1$ ,  $C(1) = 0$ , the zero matrix.

## APPENDIX B

### CONFIDENCE INTERVAL DERIVATION

Consider the statistic

$$\bar{z} = \frac{1}{n} \sum_{i=1}^n z_i .$$

If  $\bar{z}$  denotes the mean of a random sample of size  $n$  from a distribution  $Z \sim N(u_z, \sigma^2)$ , then  $\bar{z} \sim N(u_z, \sigma^2/n)$  [10,11]. Consider the probability that the interval  $(u_z - 3\sigma/\sqrt{n}, u_z + 3\sigma/\sqrt{n})$  includes the point  $\bar{z}$ . The event  $u_z - 3\sigma/\sqrt{n} < \bar{z} < u_z + 3\sigma/\sqrt{n}$  occurs when and only when the event  $-3 < \sqrt{n}(u_z - \bar{z})/\sigma < 3$  occurs, thus these two events have the same probability. However,  $\sqrt{n}(u_z - \bar{z})/\sigma$  is  $N(0,1)$ . Accordingly, the probability that the interval  $(u_z - 3\sigma/\sqrt{n}, u_z + 3\sigma/\sqrt{n})$  includes the point  $\bar{z}$  is equal to

$$\int_{-3}^3 \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = 0.998 .$$

This probability in no manner depends upon the values of  $\bar{z}$ ,  $\sigma^2$ , or the integer  $n$ . Consider next, the length of the interval. The length is seen to be  $6\sigma/\sqrt{n}$ . Note that this length is unknown until both  $\sigma^2$  and

$n$  are known; note also that for  $\sigma^2$  assigned, this length may be made as short as desired by taking  $n$  sufficiently large.

For the particular interval considered above, 0.998 was found to be the probability that the interval  $(u_z - 3\sigma/\sqrt{n}, u_z + 3\sigma/\sqrt{n})$  contains the statistic  $\bar{z}$ . That is,

$$\Pr[u_z - 3\sigma/\sqrt{n} < \bar{z} < u_z + 3\sigma/\sqrt{n}] = 0.998 .$$

Since  $\sigma$  is known, each of the variables  $u_z - 3\sigma/\sqrt{n}$  and  $u_z + 3\sigma/\sqrt{n}$  is a known quantity if  $u_z$  is known.

APPENDIX C

COMPILED FORTRAN IV PROGRAM LISTING OF ADAPTIVE BAYES CLASSIFIER INCORPORATING

MODIFIED CF ALGORITHM AND CONFIDENCE INTERVAL DIVERGENCE CRITERION

C THIS PROGRAM READS A 128 X 128 DATA POINT ARRAY FROM A DISK  
 C FILE, EACH POINT OF WHICH MAY BE CONSIDERED A FOUR DIMENSIONAL  
 C PATTERN. EACH PATTERN IS CLASSIFIED INTO ONE OF TWO CLASSES  
 C BY MEANS OF A BAYES CLASSIFIER. A STOCHASTIC APPROXIMATION  
 C TECHNIQUE (CF) IS USED IN THE ESTIMATION OF THE MEAN OF THE TWO CLASS  
 C SIGNATURES. UPDATED ESTIMATES OF COVARIANCE MATRICES ARE  
 C MADE FOR BLOCKS OF CLASSIFIED DATA. THE A PRIORI PROBABILITIES  
 C ASSOCIATED WITH THE TWO CLASSES ARE FURTHER ASSUMED TO BE  
 C EQUIVALENT. A BOUNDARY IS DEFINED SEPARATING THE TWO CLASSES  
 C AND IS STORED IN DISK MEMORY IN PROPER FORM FOR DISPLAY  
 C VIA THE DATA DISK VIDEO SYSTEM.

VARIABLES:

- NC -NUMBER OF CLASSES
- IC -CLASS CONSIDERATION INTERVAL
- IX -HORIZONTAL PICTURE ARRAY INDEX
- IY -VERTICAL PICTURE ARRAY INDEX
- ICLASS -EITHER 1 OR 2 INDICATING WHETHER LAST PATTERN  
 WAS CLASSIFIED A MEMBER OF CLASS ONE OR TWO
- INDEX -AN INDEX OF THE IADATA ARRAY



```

C IADATA -ARRAY CONTAINING ROW BOUNDARY POSITIONS
C ICDATA -A LOGICAL STAR*1 ARRAY USED IN BUILDING A DISK OUTPUT
C FILE FOR DISPLAY VIA DATA-DISK VIDEO SYSTEM
C IODATA -ARRAY CONTAINING TRUE BOUNDARY
C
C SUBROUTINES:
C
C INPUT -READS DATA FROM DISK INPUT FILE
C DECIDE -BAYES CLASSIFIER FOR TWO CLASSES
C BOUND -SUBROUTINE TO PLACE CLASS BOUNDARY
C AVERAGE -COMPUTES AVERAGES USED AS INDICATIONS OF DIVERGENCE
C OF STOCHASTIC APPROXIMATIONS
C INTERVAL -SUBROUTINE TO CALCULATE CONFIDENCE INTERVALS FOR
C STOCHASTIC APPROXIMATIONS
C PROJECT -STOCHASTIC APPROXIMATION SUBROUTINE
C PROCOV -MAKES ESTIMATES OF COVARIANCE MATRIX
C COVAR -COMPUTES COVARIANCE OF DATA SAMPLES IN RECURSIVE
C FORM
C MINV -INVERTS A MATRIX
C MPRD -FORMS THE PRODUCT OF TWO MATRICES
C OUTPUT -WRITES BOUNDARY DATA TO A DISK FILE FOR DISPLAY
C PURPOSES
C
C
C DIMENSION DATAI(512),C1(16),C2(16),C1I(16),C2I(16),
0001 1U1EST(4),U2EST(4),UA(4),UB(4),U1(4),U2(4),X(4),
0002 1U1BAR(4),U2BAR(4),TIME1(4),TIME2(4),CIA(4),CIB(4)
0003 DIMENSION IADATA(128),IODATA(128)
0004 EQUIVALENCE (INT,DUM)
0005 LOGICAL*1 ICDATA(128),DUM(2)
0006 COMMON /DUMB1/DATAI
COMMON /DUMB2/ICDATA

```

```

0007 DATA IAST/'*'/
0008 DATA IBLNK/' '/
0009 WRITE(7,10)
0010 10 FORMAT(' SPECIFY INPUT FILE'//)
0011 CALL ASSIGN(1,'SY0:ABCDEF.DAT',-14,'RD0','NC',1)
0012 DEFINE FILE 1(128,1024,U,I1)
0013 WRITE(7,20)
0014 20 FORMAT(' SPECIFY OUTPUT FILE'//)
0015 CALL ASSIGN(2,'SY0:ABCDEF.DAT',-14,'NEW','NC',1)
0016 DEFINE FILE 2(128,64,U,I2)
0017 WRITE(7,25)
0018 25 FORMAT(' SPECIFY CLASS CONSIDERATION INTERVAL'//)
0019 READ(7,26)IC
0020 26 FORMAT(I3)
0021 WRITE(7,30)
0022 30 FORMAT(' SPECIFY 4X4 EST. OF CLASS 1 COV. MATRIX'//)
0023 DO 35 NM=1,4
0024 READS(5,40)(C1(I),I=NM,NM+12,4)
0025 35 CONTINUE
0026 WRITE(7,36)
0027 36 FORMAT(' SPECIFY 4X4 EST. OF CLASS 2 COV. MATRIX'//)
0028 DO 37 NM=1,4
0029 READ(5,40)(C2(I),I=NM,NM+12,4)
0030 37 CONTINUE
0031 40 FORMAT(4F10.5)
0032 WRITE(7,50)
0033 50 FORMAT(' SPECIFY EST. OF CLASS 1 MEAN VECTOR'//)
0034 READ(5,40)(U1EST(I),I=1,4)
0035 WRITE(7,60)
0036 60 FORMAT(' SPECIFY EST. OF CLASS 2 MEAN VECTOR'//)
0037 READ(5,40)(U2EST(I),I=1,4)
0038 NC=2
0039 CALL PROCJCV(C1,C2,C11,C21,D1,D2,X,I,CLASS,IC,0)

```

```

0040 DO 70 IY=1,128
0041 INDEX=1
0042 DO 71 M=1,128
0043 IADATA(M)=129
0044 71 CONTINUE
0045 DO 75 NM=1,4
0046 UA(NM)=U1EST(NM)
0047 UB(NM)=U2EST(NM)
0048 UT(NM)=U1EST(NM)
0049 U2(NM)=U2EST(NM)
0050 TIME1(NM)=0.
0051 TIME2(NM)=0.
0052 75 CONTINUE

C
C BEGINNING OF BOUNDARY DEFINITION PROCEEDURE FOR THE 128 X 128
C PICTURE
C
0053 CALL INPUT(IY)
0054 DO 90 J=1,512,4
0055 IX=(J+3)/4
0056 X(1)=DATAI(J)
0057 X(2)=DATAI(J+1)
0058 X(3)=DATAI(J+2)
0059 X(4)=DATAI(J+3)
0060 CALL DECIDE(X,C1I,C2I,U1,U2,D1,D2,ICLASS)
0061 CALL BOUND(IADATA,ICLASS,IX,IC,INDEX)
0062 CALL AVERAGE(X,U1,U2,UTBAR,U2BAR,ICLASS,TIME1,TIME2)
0063 CALL INTERVAL(TIME1,TIME2,ICLASS,CIA,CIB)
0064 CI=FLOAT(IC)
0065 IF(ICLASS.EQ.2)GO TO 84
0067 IF(TIME1(1).LE.CI)GO TO 80

```

```
0069 IF(ABS(U1BAR(1)).LE.CIA(1))GO TO 80
0071 TIME1(1)=0.
0072 UA(1)=U1(1)
0073 80 IF(TIME1(2).LE.CI)GO TO 81
0075 IF(ABS(U1BAR(2)).LE.CIA(2))GO TO 81
0077 TIME1(2)=0.
0078 UA(2)=U1(2)
0079 81 IF(TIME1(3).LE.CI)GO TO 82
0081 IF(ABS(U1BAR(3)).LE.CIA(3))GO TO 82
0083 TIME1(3)=0.
0084 UA(3)=U1(3)
0085 82 IF(TIME1(4).LE.CI)GO TO 88
0087 IF(ABS(U1BAR(4)).LE.CIA(4))GO TO 88
0089 TIME1(4)=0.
0090 UA(4)=U1(4)
0091 GO TO 88
0092 84 IF(TIME2(1).LE.CI)GO TO 85
0094 IF(ABS(U2BAR(1)).LE.CIB(1))GO TO 85
0096 TIME2(1)=0.
0097 UB(1)=U2(1)
0098 85 IF(TIME2(2).LE.CI)GO TO 86
0100 IF(ABS(U2BAR(2)).LE.CIB(2))GO TO 86
0102 TIME2(2)=0.
0103 UB(2)=U2(2)
0104 86 IF(TIME2(3).LE.CI)GO TO 87
0106 IF(ABS(U2BAR(3)).LE.CIB(3))GO TO 87
0108 TIME2(3)=0.
0109 UB(3)=U2(3)
0110 87 IF(TIME2(4).LE.CI)GO TO 88
0112 IF(ABS(U2BAR(4)).LE.CIB(4))GO TO 88
0114 TIME2(4)=0.
0115 UB(4)=U2(4)
0116 88 CALL PROJECT(X,U1,U2,ICLASS,UA,UB,TIME1,TIME2)
0117 CALL PROCOV(C1,C2,C1I,C2I,D1,D2,X,ICLASS,IC,1)
```

```
0118 90 CONTINUE
0119 DO 91 M=1,128
0120 IODATA(M)=IBLNK
0121 91 CONTINUE
0122 M=1
0123 92 IAM=IODATA(M)
0124 IF(IAM.GT.128)GO TO 93
0126 IODATA(IAM)=IAST
0127 M=M+1
0128 GO TO 92
0129 93 CONTINUE
0130 DO 100 I=1,128
0131 IF(IODATA(I).EQ.IBLNK)IODATA(I)=0
0133 IF(IODATA(I).EQ.IAST)IODATA(I)=129
0135 INT=IODATA(I)
0136 ICDATA(I)=DUM(1)
0137 100 CONTINUE
0138 CALL OUTPUT(IY)
0139 70 CONTINUE
0140 WRITE(7,110)
0141 110 FORMAT(' CLASSIFICATION COMPLETE')
0142 ENDFILE 1
0143 ENDFILE 2
0144 STOP
0145 END
```

```
0001 SUBROUTINE INPUT (IY)
      C THIS SUBROUTINE READS ONE ROW OF FOUR-DIMENSIONAL DATA OF A 128 X 128
      C DATA POINT ARRAY FROM A PRESPECIFIED DISK FILE. THE ROW OF DATA
      C IS READ INTO THE ARRAY DATAI.
          DIMENSION DATAI(512)
          COMMON /DUMBI/DATAI
          READ(1,IY)DATAI
          RETURN
          END
0002
0003
0004
0005
0006
```

```
0001 SUBROUTINE DECIDE(X,CAI,CBI,UA,UB,DA,DB,ICLASS)
C THIS SUBROUTINE IMPLEMENTS A BAYES CLASSIFIER FOR TWO CLASSES
C OF EQUAL A PRIORI PROBABILITY.
0002 DIMENSION X(4),CAI(16,),CBI(16),UA(4),UB(4),
1YA(4),YB(4),RA(4),RB(4),A(1),B(1)
0003 DO 10 I=1,4
0004 YA(I)=X(I)-UA(I)
0005 YB(I)=X(I)-UB(I)
0006 10 CONTINUE
0007 CALL MPRD(YA,CAI,RA,1,4,4)
0008 CALL MPRD(YB,CBI,RB,1,4,4)
0009 CALL MPRD(RA,YA,A,1,4,1)
0010 CALL MPRD(RB,YB,B,1,4,1)
0011 F1=- (ALOG(DA)+A(1))
0012 F2=- (ALOG(DB)+B(1))
0013 ICLASS=1
0014 IF(F2.GT.F1)ICLASS=2
0015 RETURN
0016 END
0017
```

```

0001 SUBROUTINE BOUND(IDATA,ICLASS,IX,IC,INDEX)
0002 C THIS SUBROUTINE FORMS A BOUNDARY BETWEEN THE TWO DATA CLASSES.
0003 DIMENSION IDATA(128),ICON(32)
0004 IF(IX.GT.1)GO TO 10
0005 NCLASS=ICLASS
0006 DO 20 N=1,IC-1
0007 ICON(N)=ICON(N+1)
0008 20 CONTINUE
0009 ICON(IC)=ICLASS
0010 IF(IX.LT.IC)GO TO 30
0011 ICNT1=0
0012 ICNT2=0
0013 DO 40 II=1,IC
0014 IF(ICON(II).EQ.1)ICNT1=ICNT1+1
0015 IF(ICON(II).EQ.2)ICNT2=ICNT2+1
0016 40 CONTINUE
0017 IF(ICNT1.GT.ICNT2)MCLASS=1
0018 IF(ICNT2.GT.ICNT1)MCLASS=2
0019 IF(MCLASS.EQ.NCLASS)GO TO 30
0020 NCLASS=MCLASS
0021 IDATA(INDEX)=IX-IC/2
0022 INDEX=INDEX+1
0023 30 RETURN
0024 END
0025
0026
0027
0028
0029
0030

```

```

0001 SUBROUTINE AVERAGE(X,U1,U2,X1BAR,X2BAR,ICLASS,TIMEX,TIMEY)
C THIS SUBROUTINE KEEPS AN INDEPENDENT RUNNING TIME AVERAGE OF THE
C DEVIATION OF THE DATA FROM THE TRUE OR PROJECTED MEAN OF EACH
C CLASS.
0002 DIMENSION X(4),U1(4),U2(4),X1BAR(4),X2BAR(4),TIMEX(4),TIMEY(4)
0003 IF(ICLASS.EQ.2)GO TO 20
0004 DO 10 I=1,4
0005 X1BAR(I)=(1./(TIMEX(I)+1.))*(TIMEX(I)*X1BAR(I)+(X(I)-U1(I)))
0006 10 CONTINUE
0007 GO TO 30
0008 DO 30 I=1,4
0009 X2BAR(I)=(1./(TIMEY(I)+1.))*(TIMEY(I)*X2BAR(I)+(X(I)-U2(I)))
0010 30 CONTINUE
0011 RETURN
0012 END
0013

```

```
0001 SUBROUTINE INTERVAL(TIMEX,TIMEY,ICLASS,CIA,CIB)
0002 C THIS SUBROUTINE CALCULATES THE CONFIDENCE INTERVAL ASSOCIATED WITH
0003 C EACH COMPONENT OF THE MEAN VECTOR FOR THE CLASS OF INTEREST.
0004 C THE CONFIDENCE INTERVAL IS FOR USE AS AN INDICATION OF DIVERGENCE.
0005 DIMENSION TIMEX(4),TIMEY(4),CIA(4),CIB(4)
0006 IF(ICLASS.EQ.2)GO TO 20
0007 IF(TIMEX(1).EQ.0.)GO TO 10
0008 CIA(1)=3./SQRT(TIMEX(1))
0009 10 IF(TIMEX(2).EQ.0.)GO TO 12
0010 CIA(2)=3./SQRT(TIMEX(2))
0011 12 IF(TIMEX(3).EQ.0.)GO TO 14
0012 CIA(3)=3./SQRT(TIMEX(3))
0013 14 IF(TIMEX(4).EQ.0.)GO TO 40
0014 CIA(4)=3./SQRT(TIMEX(4))
0015 GO TO 40
0016 20 IF(TIMEX(1).EQ.0.)GO TO 30
0017 CIB(1)=3./SQRT(TIMEX(1))
0018 30 IF(TIMEX(2).EQ.0.)GO TO 32
0019 CIB(2)=3./SQRT(TIMEX(2))
0020 32 IF(TIMEX(3).EQ.0.)GO TO 34
0021 CIB(3)=3./SQRT(TIMEX(3))
0022 34 IF(TIMEX(4).EQ.0.)GO TO 40
0023 CIB(4)=3./SQRT(TIMEX(4))
0024 40 RETURN
0025 END
0026
0027
0028
0029
0030
0031
```

```

0001 SUBROUTINE PROJECT(X,UX,UY,ICLASS,UA,UB,TIMEX,TIMEY)
C THIS SUBROUTINE PROJECTS EACH COMPONENT OF THE MEAN VECTOR OF THE
C CLASS OF INTEREST. THE PROJECTION IS MADE USING A STOCHASTIC
C APPROXIMATION. THE TWO COMPONENTS ARE PROJECTED INDEPENDENTLY OF
C ONE ANOTHER. (CHIEN AND FU)
DIMENSION X(4),UX(4),UY(4),UA(4),UB(4),TIMEX(4),TIMEY(4)
IF(ICLASS.EQ.2)GO TO 20
DO 10 I=1,4
UX(I)=UX(I)-UA(I)
X(I)=X(I)-UA(I)
USTAR=(1.+1./((TIMEX(I)+1.))*UX(I)
GAMMA=6.*(TIMEX(I)+2.)/((TIMEX(I)+3.)*(2.*(TIMEX(I)+1.))+3.))
UX(I)=USTAR+GAMMA*(X(I)-USTAR)
TIMEY(I)=TIMEX(I)+1.
UX(I)=UX(I)+UA(I)
10 CONTINUE
DO 30 I=1,4
UY(I)=UY(I)-UB(I)
X(I)=X(I)-UB(I)
USTAR=(1.+1./((TIMEY(I)+1.))*UY(I)
GAMMA=6.*(TIMEY(I)+2.)/((TIMEY(I)+3.)*(2.*(TIMEY(I)+1.))+3.))
UY(I)=USTAR+GAMMA*(X(I)-USTAR)
TIMEY(I)=TIMEY(I)+1.
UY(I)=UY(I)+UB(I)
30 CONTINUE
RETURN
END
0002
0003
0005
0006
0007
0008
0009
0010
0011
0012
0013
0014
0015
0016
0017
0018
0019
0020
0021
0022
0023
0024
0025

```

```

0001 SUBROUTINE PROCOV(CA,CB,CAINV,CBINV,DA,DB,X,ICLASS,IC,IFLAG)
C THIS SUBROUTINE MAKES A NEW ESTIMATE OF THE COVARIANCE OF A
C CLASS AFTER IC POINTS HAVE BEEN CLASSIFIED INTO THAT PARTICULAR
C CLASS.
0002 DIMENSION CA(16),CB(16),CAINV(16),CBINV(16),
1PHIA(16),PHIB(16),X(4)
IF(IFLAG.GT.0)GO TO 10
ICNTA=1
ICNTB=1
DO 5 I=1,16
CAINV(I)=CA(I)
CBINV(I)=CB(I)
5 CONTINUE
CALL MINV(CAINV,4,DA)
CALL MINV(CBINV,4,DB)
GO TO 50
10 IF(ICLASS.EQ.2)GO TO 30
CALL COVAR(PHIA,X,ICNTA,4)
ICNTA=ICNTA+1
IF(ICNTA.LE.IC)GO TO 50
DO 20 I=1,16
CA(I)=PHIA(I)
CAINV(I)=PHIA(I)
20 CONTINUE
CALL MINV(CAINV,4,DA)
ICNTA=1
GO TO 50
30 CALL COVAR(PHIB,X,ICNTB,4)
ICNTB=ICNTB+1
IF(ICNTB.LE.IC)GO TO 50
DO 40 I=1,16
CB(I)=PHIB(I)
CBINV(I)=PHIB(I)
0033

```

0034  
0035  
0036  
0037  
0038

40 CONTINUE  
CALL MINV(CBINV,4,DB)  
ICNTB=1  
50 RETURN  
END

```

0001 SUBROUTINE COVAR(COV,X,INDEX,NDIM)
C THIS SUBROUTINE IMPLEMENTS THE RECURSIVE FORM OF COVARIANCE
C ESTIMATION.
0002 DIMENSION COV(4,4),X(4),XM(4)
0003 X1=FLOAT(INDEX)
0004 X0=X1-1.
0005 IF(INDEX.NE.1)GO TO 5
0007 DO 4 I=1,NDIM
0008 XM(I)=X(I)
0009 DO 3 J=1,NDIM
0010 COV(I,J)=0.
0011 3 CONTINUE
0012 4 CONTINUE
0013 GO TO 20
0014 5 DO 10 I=1,NDIM
0015 DO 11 J=1,NDIM
0016 COV(I,J)=(1./X1)*(X0*COV(I,J)+X0*XM(I)*XM(J)+X(I)*X(J))
1-(1./X1*X1)*X0*XM(I)+X(I))*(X0*XM(J)+X(J))
0017 11 CONTINUE
0018 10 CONTINUE
0019 DO 15 I=1,NDIM
0020 XM(I)=(1./X1)*(X0*XM(I)+X(I))
0021 20 RETURN
0022 END

```

```

0001 SUBROUTINE MINV(A,N,D)
C THIS SUBROUTINE FINDS THE INVERSE OF A GENERAL MATRIX
C 'A', WHICH IS DESTROYED IN COMPUTATION. THE INVERSE
C IS RETURNED IN 'A'. THE DETERMINANT IS CALCULATED,
C 'D'. 'N' IS THE ORDER OF THE MATRIX, 'L' IS A WORK
C VECTOR OF LENGTH 'N', AND 'M' IS A WORK VECTOR OF
C LENGTH 'N'.
      DIMENSION A(16),L(4),M(4)
0002 D=1.0
0003 NK=-N
0004 DO 80 K=1,N
0005 NK=NK+N
0006 L(K)=K
0007 M(K)=K
0008 KK=NK+K
0009 BIGA=A(KK)
0010 DO 20 J=K,N
0011 IZ=N*(J-1)
0012 DO 20 I=K,N
0013 IJ=IZ+I
0014 IF (ABS(BIGA)-ABS(A(IJ))) 15,20,20
0015 15 BIGA=A(IJ)
0016 L(K)=I
0017 M(K)=J
0018 20 CONTINUE
0019 J=L(K)
0020 IF(J-K) 35,35,25
0021 KI=K-N
0022 DO 30 I=1,N
0023 KI=KI+N
0024 HOLD=-A(KI)
0025 JI=KI-K+J
0026 A(KI)=A(JI)
0027 30 A(JI)=HOLD
0028

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35 I=M(K)
   IF(I-K)45,45,38
38 JP=N*(I-1)
   DO 40 J=1,N
   JK=NK+J
   JI=JP+J
   HOLD=-A(JK)
   A(JK)=A(JI)
40 A(JI)=HOLD
45 IF(BIGA) 48,46,48
46 D=0.
   RETURN
48 DO 55 I=1,N
   IF(I-K) 50,55,50
50 IK=NK+I
   A(IK)=A(IK)/(-BIGA)
55 CONTINUE
   DO 65 I=1,N
   IK=NK+I
   HOLD=A(IK)
   IJ=I-N
   DO 65 J=1,N
   IJ=IJ+N
   IF(I-K) 60,65,60
60 IF(J-K) 62,65,62
62 KJ=IJ-I+K
   A(IJ)=HOLD*A(KJ)+A(IJ)
65 CONTINUE
   KJ=K-N
   DO 75 J=1,N
   KJ=KJ+N
   IF(J-K) 70,75,70
70 A(KJ)=A(KJ)/BIGA
75 CONTINUE

```

```
0063 D=D*BIGA
0064 A(KK)=1.0/BIGA
0065 80 CONTINUE
0066 K=N
0067 100 K=K-1
0068 IF(K) 150,150,105
0069 105 I=L(K)
0070 IF(I-K) 120,120,108
0071 108 JQ=N*(K-1)
0072 JR=N*(I-1)
0073 DO 110 J=1,N
0074 JK=JQ+J
0075 HOLD=A(JK)
0076 JI=JR+J
0077 A(JK)=-A(JI)
0078 110 A(JI)=HOLD
0079 120 J=M(K)
0080 IF(J-K) 100,100,125
0081 125 KI=K-N
0082 DO 130 I=1,N
0083 KI=KI+N
0084 HOLD=A(KI)
0085 JI=KI-K+J
0086 A(KI)=-A(JI)
0087 130 A(JI)=HOLD
0088 GO TO 100
0089 150 RETURN
0090 END
```

```

0001 SUBROUTINE MPRD(A,B,R,N,M,L)
      C THIS SUBROUTINE FORMS THE PRODUCT OF TWO GENERAL MATRICES
      C 'A' AND 'B', AND RETURNS THE PRODUCT IN 'R'. 'N' IS THE NUMBER
      C OF ROWS IN 'A', 'M' IS THE NUMBER OF COLUMNS IN 'A' AND ROWS
      C IN 'B', AND 'L' IS THE NUMBER OF COLUMNS IN 'B'.
      DIMENSION A(16),B(16),R(16)
      IR=0
      IK=-M
      DO 10 K=1,L
      IK=IK+M
      DO 10 J=1,N
      IR=IK+1
      JI=J-N
      IB=IK
      R(IR)=0
      DO 10 I=1,M
      JI=JI+N
      IB=IB+1
      10 R(IR)=R(IR)+A(JI)*B(IB)
      RETURN
      END
0002
0003
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0016
0017

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```

0001 SUBROUTINE OUTPUT (IY)
      C THIS SUBROUTINE WRITES ONE 128 POINT ROW (BYTE DATA POINTS) OF A
      C 128 X 128 POINT ARRAY FOR BOUNDARY DISPLAY VIA VIDEO DISPLAY SYSTEM.
0002 LOGICAL*1 IODATA(128)
0003 COMMON /DUMB2/ICDATA
0004 WRITE(2,IY)ICDATA
0005 RETURN
0006 END

```

C.2





```

C CIB -ARRAY CONTAINING CONVERGENCE INFORMATION FOR USE IN
C STOCHASTIC APPROXIMATION OF CLASS TWO MEAN
C E1 -ARRAY CONTAINING AN ESTIMATE OF THE MEAN SQUARE ERROR
C OF THE CLASS ONE MEAN VECTOR
C E2 -ARRAY CONTAINING AN ESTIMATE OF THE MEAN SQUARE ERROR
C OF THE CLASS TWO MEAN VECTOR
C IADATA -ARRAY CONTAINING ROW BOUNDARY POSITIONS
C ICDATA -A LOGICAL STAR*1 ARRAY USED IN BUILDING A DISK OUTPUT
C FILE FOR DISPLAY VIA DATA-DISK VIDEO SYSTEM
C IODATA -ARRAY CONTAINING TRUE BOUNDARY
C
C SUBROUTINES:
C
C INPUT -READS DATA FROM DISK INPUT FILE
C DECIDE -BAYES CLASSIFIER FOR TWO CLASSES
C BOUND -SUBROUTINE TO PLACE CLASS BOUNDARY
C PROJECT -STOCHASTIC APPROXIMATION SUBROUTINE
C PROCOV -MAKES ESTIMATES OF COVARIANCE MATRIX
C COVER -COMPUTES COVARIANCE OF DATA SAMPLES IN RECURSIVE
C FORM
C MINV -INVERTS A MATRIX
C MPRD -FORMS THE PRODUCT OF TWO MATRICES
C OUTPUT -WRITES BOUNDARY DATA TO A DISK FILE FOR DISPLAY
C PURPOSES
C
C DIMENSION DATAI(512),C1(16),C2(16),C1I(16),C2I(16),
0001 TUIEST(4),U2EST(4),UA(4),UB(4),U1(4),U2(4),X(4),
TUIBAR(4),U2BAR(4),TIME1(4),TIME2(4),CIA(4),CIB(4),
TIE1(4),E2(4)
C DIMENSION IADATA(128),IODATA(128)
C EQUIVALENCE (INT,DUM)
0002
0003 LOGICAL*1 ICDATA(128),DUM(2)
0004

```

```

0005 COMMON /DUMB1/DATA1
0006 COMMON /DUMB2/ICDATA
0007 DATA IAST/'*'/
0008 DATA IBLNK/' '/
0009 WRITE(7,10)
0010 10 FORMAT(' SPECIFY INPUT FILE'//)
0011 CALL ASSIGN(1,'SYO:ABSDEF.DAT',-14,'RDO','NC',1)
0012 DEFINE FILE 1(128,1024,U,I1)
0013 WRITE(7,20)
0014 20 FORMAT(' SPECIFY OUTPUT FILE'//)
0015 CALL ASSIGN(2,'SYO:ABCDEF.DAT',-14,'NEW','NC',1)
0016 DEFINE FILE 2(128,64,U,I2)
0017 WRITE(7,25)
0018 25 FORMAT(' SPECIFY CLASS CONSIDERATION INTERVAL'//)
0019 READ(7,26)IC
0020 26 FORMAT(I3)
0021 WRITE(7,30)
0022 30 FORMAT(' SPECIFY 4X4 EST. OF CLASS 1 COV. MATRIX'//)
0023 DO 35 NM=1,4
0024 READ(5,40)(C1(I),I=NM,NM+12,4)
0025 35 CONTINUE
0026 WRITE(7,36)
0027 36 FORMAT(' SPECIFY 4X4 EST. OF CLASS 2 COV. MATRIX'//)
0028 DO 37 NM=1,4
0029 READ(5,40)(C2(I),I=NM,NM+12,4)
0030 37 CONTINUE
0031 40 FORMAT(4F10.5)
0032 WRITE(7,50)
0033 50 FORMAT(' SPECIFY EST. OF CLASS 1 MEAN VECTOR'//)
0034 READ(5,40)(U1EST(I),I=1,4)
0035 WRITE(7,60)
0036 60 FORMAT(' SPECIFY EST. OF CLASS 2 MEAN VECTOR'//)
0037 READ(5,40)(U2EST(I),I=1,4)
0038 NC=2

```

```

0039          CALL PROCOV(C1,C2,C11,C21,D1,D2,X,ICLASS,IC,0)
C
C BEGINNING OF BOUNDARY DEFINITION PROCEDURE FOR THE 128 X 128
C PICTURE
C
0040          DO 70 IY=1,128
0041             INDEX=1
0042          DO 71 M=1,128
0043             IADATA(M)=129
0044          71 CONTINUE
0045          DO 75 NM=1,4
0046             UA(NM)=U1EST(NM)
0047             UB(NM)=U2EST(NM)
0048             U1(NM)=U1EST(NM)
0049             U2(NM)=U2EST(NM)
0050             TIME1(NM)=0.
0051             TIME2(NM)=0.
0052             E1(NM)=0.
0053             E2(NM)=0.
0054          75 CONTINUE
C
C READ A ROW AND PLACE BOUNDARIES WHERE NECESSARY
C
0055          CALL INPUT(IY)
0056          DO 90 J=1,512,4
0057             IX=(J+3)/4
0058             X(1)=DATAI(J)
0059             X(2)=DATAI(J+1)
0060             X(3)=DATAI(J+2)
0061             X(4)=DATAI(J+3)
0062          CALL DECIDE(X,C11,C21,U1,U2,D1,D2,ICLASS)
0063          CALL BOUND(IADATA,ICLASS,IX,IC,INDEX)
0064          88 CALL PROJECT(X,U1,U2,ICLASS,E1,E2,TIME1,TIME2)

```

```
0065 CALL PRGOV(C1,C2,C1I,C2I,D1,D2,X,ICLASS,IC,1)
0066 CONTINUE
0067 DO 91 M=1,128
0068 IODATA(M)=IBLNK
0069 91 CONTINUE
0070 M=1
0071 IAM=IADATA(M)
0072 IF(IAM.GT.128)GO TO 93
0074 IODATA(IAM)=IAST
0075 M=M+1
0076 GO TO 92
0077 93 CONTINUE
0078 DO 100 I=1,128
0079 IF(IODATA(I).EQ.IBLNK)IODATA(I)=0
0081 IF(IODATA(I).EQ.IAST)IODATA(I)=129
0083 INT=IODATA(I)
0084 ICDATA(I)=DUM(1)
0085 100 CONTINUE
0086 CALL OUTPUT(IY)
0087 70 CONTINUE
0088 WRITE(7,110)
0089 110 FORMAT(' CLASSIFICATION COMPLETE')
0090 ENDFILE 1
0091 ENDFILE 2
0092 STOP
0093 END
```

```
0001 SUBROUTINE INPUT (IY)
      C THIS SUBROUTINE READS ONE ROW OF FOUR-DIMENSIONAL DATA OF A 128 X 128
      C DATA POINT ARRAY FROM A PRESPECIFIED DISK FILE. THE ROW OF DATA
      C IS READ INTO THE ARRAY DATAI.
      DIMENSION DATAI(512)
      COMMON /DUMBI/DATAI
      READ(1,IY)DATAI
      RETURN
      END
0002
0003
0004
0005
0006
```

```

0001 SUBROUTINE DECIDE(X,CAI,CBI,UA,UB,DA,DB,ICLASS)
C THIS SUBROUTINE IMPLEMENTS A BAYES CLASSIFIER FOR TWO CLASSES
C OF EQUAL A PRIORI PROBABILITY.
0002 DIMENSION X(4),CAI(16),CBI(16),UA(4),UB(4),
1YA(4),YB(4),RA(4),RB(4),A(1),B(1)
DO 10 I=1,4
0003 YA(I)=X(I)-UA(I)
0004 YB(I)=X(I)-UB(I)
0005
0006 10 CONTINUE
0007 CALL MPRD(YA,CAI,RA,1,4,4)
0008 CALL MPRD(YB,CBI,RB,1,4,4)
0009 CALL MPRD(RA,YA,A,1,4,1)
0010 CALL MPRD(RB,YB,B,1,4,1)
0011 F1=- (ALOG(DA)+A(1))
0012 F2=- (ALOG(DB)+B(1))
0013 ICLASS=1
0014 IF(F2.GT.F1) ICLASS=2
0016 RETURN
0017 END

```

```
0001 SUBROUTINE BOUND(IDATA,ICLASS,IX,IC,INDEX)
0002 C THIS SUBROUTINE FORMS A BOUNDARY BETWEEN THE TWO DATA CLASSES.
0003 DIMENSION IDATA(128),ICON(32)
0004 IF(IX.GT.1)GO TO 10
0005 NCLASS=ICLASS
0006 DO 20 N=1,IC-1
0007 ICON(N)=ICON(N+1)
0008 20 CONTINUE
0009 ICON(IC)=ICLASS
0010 IF(IX.LT.IC)GO TO 30
0011 ICNT1=0
0012 ICNT2=0
0013 DO 40 II=1,IC
0014 IF(ICON(II).EQ.1)ICNT1=ICNT1+1
0015 IF(ICON(II).EQ.2)ICNT2=ICNT2+1
0016 40 CONTINUE
0017 IF(ICNT1.GT.ICNT2)MCLASS=1
0018 IF(ICNT2.GT.ICNT1)MCLASS=2
0019 IF(MCLASS.EQ.NCLASS)GO TO 30
0020 NCLASS=MCLASS
0021 IDATA(INDEX)=IX-IC/2
0022 INDEX=INDEX+1
0023 30 RETURN
0024 END
0025
0026
0027
0028
0029
0030
```

```

0001 SUBROUTINE PROJECT(X,UX,UY,ICLASS,E1,E2,TIMEX,TIMEY)
C THIS SUBROUTINE PROJECTS EACH COMPONENT OF THE MEAN VECTOR OF THE
C CLASS OF INTEREST. THE PROJECTION IS MADE USING A STOCHASTIC
C APPROXIMATION. THE FOUR COMPONENTS ARE PROJECTED INDEPENDENTLY OF
C ONE ANOTHER. (SECOND DEGREE POLYNOMIAL FIT)
0002 DIMENSION X(4),UX(4),UY(4),E1(4),E2(4),TIMEX(4),TIMEY(4),
1 X1(128,+),X2(128,4),IT1(4),IT2(4)
0003 REAL K,K1,K2,K3
0004 IF(ICLASS.EQ.2)GO TO 30
0006 DO 10 I=1,4
0007 IT1(I)=IFIX(TIMEX(I))+1
0008 J=IT1(I)
0009 X1(J,I)=X(I)
0010 IF(TIMEX(I).LT.3.)GO TO 15
0012 II=IFIX(TIMEX(I)/2.)+1
0013 B=FLOAT(II)
0014 IB=IFIX(B)
0015 ID=J-1
0016 D=FLOAT(ID)
0017 K=(D+B+1.)/(D*B)
0018 GAMMA=(E1(I)-K)/(E1(I)+1.)
0019 UX(I)=UX(I)+GAMMA*(X(I)-UX(I))
0020 S=(X(I)*(D*(D+1.)-B*(B+1.))-X1(J-IB,I)*(D*(D+1.))+X1(1,I)
1 *(B*(B+1.)))/((D-B)*D*B)
0021 UX(I)=UX(I)+S
0022 K2=- (D+1.)/(B*(D-B))
0023 K3=(B+1.)/(D*(D-B))
0024 K1=(-(K2+K3)+1.)**2
0025 E1(I)=(E1(I)/(E1(I)+1.))*K1+K2**2+K3**2
0026 TIMEX(I)=TIMEX(I)+1.
0027 10 CONTINUE
0028 GO TO 60
0029 30 DO 40 I=1,4
0030 IT2(I)=IFIX(TIMEX(I))+1
0031 J=IT2(I)

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```

0032 X2(J,I)=X(I)
0033 IF(TIMEY(I).LT.3.)GO TO 45
0035 II=IFIX(TIMEY(I)/2.)+1
0036 B=FLOAT(I)
0037 IB=IFY(B)
0038 ID=J-1
0039 D=FLOAT(ID)
0040 K=(D+B+1.)/(D*B)
0041 GAMMA=(E2(I)-K)/(E2(I)+1.)
0042 UY(I)=UY(I)+GAMMA*(X(I)-UY(I))
0043 S=(X(I)*(D*(D+1.)-B*(B+1.))-X2(J-IB,I)*(D*(D+1.))+X2(1,I)
      1*(B*(B+1.)))/((D-B)*D*B)
0044 UY(I)=UY(I)+S
0045 K2=- (D+1.)/(B*(D-B))
0046 K3=(B+1.)/(D*(D-B))
0047 K1=(-(K2+K3)+1.)**2
0048 E2(I)=(E2(I)/(E2(I)+1.))*K1+K2**2+K3**2
0049 TIMEY(I)=TIMEY(I)+1.
0050 45 CONTINUE
0051 60 RETURN
0052 END

```

```

0001 SUBROUTINE PROCOV(CA,CB,CAINV,CBINV,DA,DB,X,ICLASS,IC,IFLAG)
C THIS SUBROUTINE MAKES A NEW ESTIMATE OF THE COVARIANCE OF A
C CLASS AFTER IC POINTS HAVE BEEN CLASSIFIED INTO THAT PARTICULAR
C CLASS.
0002 DIMENSION CA(16),CB(16),CAINV(16),CBINV(16),
1PHIA(16),PHIB(16),X(4)
IF(IFLAG.GT.0)GO TO 10
ICNTA=1
ICNTB=1
DO 5 I=1,16
CAINV(I)=CA(I)
CBINV(I)=CB(I)
5 CONTINUE
CALL MINV(CAINV,4,DA)
CALL MINV(CBINV,4,DB)
GO TO 50
10 IF(ICLASS.EQ.2)GO TO 30
CALL COVAR(PHIA,X,ICNTA,4)
ICNTA=ICNTA+1
IF(ICNTA.LE.IC)GO TO 50
DO 20 I=1,16
CA(I)=PHIA(I)
CAINV(I)=PHIA(I)
20 CONTINUE
CALL MINV(CAINV,4,DA)
ICNTA=1
GO TO 50
30 CALL COVAR(PHIB,X,ICNTB,4)
ICNTB=ICNTB+1
IF(ICNTB.LE.IC)GO TO 50
DO 40 I=1,16
CB(I)=PHIB(I)
CBINV(I)=PHIB(I)
40 CONTINUE
CALL MINV(CBINV,4,DB)
0003
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0010
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0012
0013
0014
0016
0017
0018
0020
0021
0022
0023
0024
0025
0026
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0028
0029
0031
0032
0033
0034
0035

```

ICNTB=1  
50 RETURN  
END

0036  
0037  
0038

```

0001 SUBROUTINE COVAR(COV,X,INDEX,NDIM)
C THIS SUBROUTINE IMPLEMENTS THE RECURSIVE FORM OF COVARIANCE
C ESTIMATION.
0002 DIMENSION COV(4,4)X(4),XM(4)
0003 X1=FLOAT(INDEX)
0004 X0=X1-1.
0005 IF(INDEX.NE.1)GO TO 5
0006 DO 4 I=1,NDIM
0007 XM(I)=X(I)
0008 DO 3 J=1,NDIM
0009 COV(I,J)=0.
0010 3 CONTINUE
0011 4 CONTINUE
0012 GO TO 20
0013 5 DO 10 I=1,NDIM
0014 DO 11 J=1,NDIM
0015 COV(I,J)=(1./X1)*(X0*COV(I,J)+X0*XM(I)*XM(J)+X(I)*X(J))
0016 1-(1./X1*X1)*(X0*XM(I)+X(I))*(X0*XM(J)+X(J))
0017 11 CONTINUE
0018 10 CONTINUE
0019 DO 15 I=1,NDIM
0020 XM(I)=(1./X1)*(X0*XM(I)+X(I))
0021 20 RETURN
0022 END

```

```

0001 SUBROUTINE MINV(A,N,D)
C THIS SUBROUTINE FINDS THE INVERSE OF A GENERAL MATRIX
C 'A', WHICH IS DESTROYED IN COMPUTATION. THE INVERSE
C IS RETURNED IN 'A'. THE DETERMINANT IS CALCULATED,
C 'D'. 'N' IS THE ORDER OF THE MATRIX, 'L' IS A WORK
C VECTOR OF LENGTH 'N', AND 'M' IS A WORK VECTOR OF
C LENGTH 'N'.
      DIMENSION A(16),L(4),M(4)
      D=1.0
      NK=-N
      DO 80 K=1,N
      NK=NK+N
      L(K)=K
      M(K)=K
      KK=NK+K
      BIGA=A(KK)
      DO 20 J=K,N
      IZ=N*(J-1)
      DO 20 I=K,N
      IJ=IZ+I
      10 IF (ABS(BIGA)-ABS(A(IJ))) 15,20,20
      15 BIGA=A(IJ)
      L(K)=I
      M(K)=J
      20 CONTINUE
      J=L(K)
      IF(J-K) 35,35,25
      25 KI=K-N
      DO 30 I=1,N
      KI=KI+N
      HOLD=-A(KI)
      2025 JI=KI-K+J
      2026 A(KI)=A(JI)
      2027 30 A(JI)=HOLD
      2028 35 I=M(K)
      2029

```

```

0030 IF(I-K) 45,45,38
0031 JP=N*(I-1)
0032 DO 40 J=1,N
0033 JK=NK+J
0034 JI=JP+J
0035 HOLD=-A(JK)
0036 A(JK)=A(JI)
0037 A(JI)=HOLD
0038 40 IF(BIGA) 48,46,48
0039 D=0.
0040 RETURN
0041 DO 55 I=1,N
0042 IF(I-K) 50,55,50
0043 IK=NK+I
0044 A(IK)=A(IK)/(-BIGA)
0045 55 CONTINUE
0046 DO 65 I=1,N
0047 IK=NK+I
0048 HOLD=A(IK)
0049 IJ=I-N
0050 DO 65 J=1,N
0051 IJ=IJ+N
0052 IF(I-K) 60,65,60
0053 IF(J-K) 62,65,62
0054 KJ=IJ-I+K
0055 A(IJ)=HOLD*A(KJ)+A(IJ)
0056 65 CONTINUE
0057 KJ=K-N
0058 DO 75 J=1,N
0059 KJ=KJ+N
0060 IF(J-K) 70,75,70
0061 A(KJ)=A(KJ)/BIGA
0062 75 CONTINUE
0063 D=D*BIGA

```

```
0064 A(KK)=1.0/BIGA
0065 80 CONTINUE
0066 K=N
0067 100 K=K-1
0068 IF(K) 150,150,105
0069 105 I=L(K)
0070 IF(I-K) 120,120,108
0071 108 JQ=N*(K-1)
0072 JR=N*(I-1)
0073 DO 110 J=1,N
0074 JK=JQ+J
0075 HCLD=A(JK)
0076 JI=JR+J
0077 A(JK)=-A(JI)
0078 110 A(JI)=HOLD
0079 120 J=M(K)
0080 IF(J-K) 100,100,125
0081 125 KI=K-N
0082 DO 130 I=1,N
0083 KI=KI+N
0084 HOLD=A(KI)
0085 JI=KI-K+J
0086 A(KI)=-A(JI)
0087 130 A(JI)=HOLD
0088 GO TO 100
0089 150 RETURN
0090 END
```

```
0001 SUBROUTINE MPRD(A,B,R,N,M,L)
      C THIS SUBROUTINE FORMS THE PRODUCT OF TWO GENERAL MATRICES
      C 'A' AND 'B', AND RETURNS THE PRODUCT IN 'R'. 'N' IS THE NUMBER
      C OF ROWS IN 'A', 'M' IS THE NUMBER OF COLUMNS IN 'A' AND ROWS
      C IN 'B', AND 'L' IS THE NUMBER OF COLUMNS IN 'B'.
      DIMENSION A(16),B(16),R(16)
      IR=0
      IK=-M
      DO 10 K=1,L
      IK=IK+M
      DO 10 J=1,N
      IR=IR+1
      JI=J-N
      IB=IK
      R(IR)=0
      DO 10 I=1,M
      JI=JI+N
      IB=IB+1
      10 R(IR)=R(IR)+A(JI)*B(IB)
      RETURN
      END
0002
0003
0004
0005
0006
0007
0008
0009
0010
0011
0012
0013
0014
0015
0016
0017
```

```
0001 SUBROUTINE OUTPUT (IY)
      C THIS SUBROUTINE WRITES ONE 128 POINT TOW (BYTE DATA POINTS) OF A
      C 128 X 128 POINT ARRAY FOR BOUNDARY DISPLAY VIA VIDEO DISPLAY SYSTEM.
0002 LOGICAL*1 ICDATA(128)
0003 COMMON /DUMB2/ICDATA
0004 WRITE(2,IY)ICDATA
0005 RETURN
0006 END
```