

1 **Climate modulates the effects of tree diversity on forest productivity**

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9 **Running headline:** Biodiversity and wood production in forests

10

11 **Summary**

- 12 **1.** Despite growing evidence that, on average, diverse forests tend to be more productive
13 than species poor ones, individual studies often report strongly contrasting relationships
14 between tree species richness and above-ground wood production (AWP). In the attempt
15 to reconcile these apparently inconsistent results, we explored whether the strength and
16 shape of AWP – diversity relationships shifts along spatial and temporal environmental
17 gradients in forests across Europe.
- 18 **2.** We used tree ring data from a network of permanent forest plots distributed at six sites
19 across Europe to estimate annual AWP over a 15 year period (1997–2011). We then
20 tested whether the relationship between tree species richness and AWP changes (i) across
21 sites as a function of large-scale gradients in climatic productivity and tree packing
22 density, and (ii) among years within each sites as a result of fluctuating climatic
23 conditions.
- 24 **3.** AWP – species richness relationships varied markedly among sites. As predicted by
25 theory, the relationship shifted from strongly positive at sites where climate imposed a
26 strong limitation on wood production and tree packing densities were low, to weakly
27 negative at sites where climatic conditions for growth were most suitable. In contrast, we
28 found no consistent effect of interannual fluctuations in climate on the strength of AWP
29 – species richness relationships within sites.
- 30 **4.** *Synthesis.* Our results indicate that the shape and strength of the relationship between tree
31 diversity and forest productivity depends critically on environmental context. Across
32 Europe, tree diversity shows the greatest potential to positively influence forest

33 productivity at either end of the latitudinal gradient, where adverse climatic conditions
34 limit productivity and lead to the development of less densely packed stands.

35 **Key-words:** above-ground wood production; biodiversity – ecosystem functioning; context
36 dependency; FunDivEUROPE project; plant-climate interactions; stress gradient hypothesis;
37 tree packing density; tree ring data

38 **Introduction**

39 As evidence that biodiversity promotes key ecosystem functions and services continues to
40 amass (Cardinale *et al.* 2012), the argument for conserving biodiversity is increasingly being
41 framed in terms of developing natural capital (Naeem, Duffy & Zavaleta 2012). In the
42 context of forests, the consensus is that diverse stands are generally more productive than
43 species-poor ones (Paquette & Messier 2011; Morin *et al.* 2011; Vilà *et al.* 2013; Jucker *et al.*
44 2014a). Consequently, promoting tree diversity is seen as a promising strategy for increasing
45 timber yields and carbon sequestration rates across forest landscapes, while also ensuring a
46 host of additional ecological co-benefits (Scherer-Lorenzen 2014). Yet despite synthesis
47 work showing that the overall relationship between diversity and forest productivity is
48 positive, both the strength and shape of this relationship vary considerably among individual
49 studies (Zhang, Chen & Reich 2012). Recent work has helped resolve some of these context
50 dependencies, highlighting how the effects of diversity on productivity are scale-dependent
51 (Chisholm *et al.* 2013) and can change during stand development (Cavard *et al.* 2011; Lasky
52 *et al.* 2014). A critical next step is to understand how environmental conditions influence the
53 relationship between species richness and above-ground wood production (AWP) in forests,
54 so that we may identify where and when tree diversity has the greatest potential to positively
55 influence forest productivity (Paquette & Messier 2011; Jucker *et al.* 2014a).

56 Predicting how the relationship between AWP and diversity is likely to change along spatial
57 environmental gradients requires an understanding of how species interactions are influenced
58 by environmental conditions. Species interactions are central to explaining positive
59 biodiversity – ecosystem functioning relationships in plant communities, as they underpin the
60 role of niche complementarity (Loreau & Hector 2001). A key feature of species interactions

61 is that they are not constant in space, but instead shift in importance and intensity along
62 environmental gradients (Brooker *et al.* 2008). This is conceptualized in the stress gradient
63 hypothesis, which predicts that under increasingly stressful environmental conditions
64 competitive interactions weaken and give way to facilitative processes (see Maestre *et al.*,
65 2009 for a review). Following this reasoning, it has been hypothesized that biodiversity –
66 ecosystem functioning relationships should also vary along environmental gradients (Fig. 1a),
67 becoming progressively stronger as conditions for growth become less favourable (Jucker &
68 Coomes 2012; Gessner & Hines 2012; Wang *et al.* 2013; Forrester 2014). In forests,
69 competitive interactions among neighbouring trees have been shown to vary in strength along
70 environmental gradients (Kunstler *et al.* 2011; He, Bertness & Altieri 2013; Prior & Bowman
71 2014), and evidence suggests that positive relationships between diversity and productivity
72 are in fact more common where environmental conditions are most limiting (Paquette &
73 Messier 2011; Wu *et al.* 2014; Toigo *et al.* 2015). A possible explanation for these patterns is
74 that the relationship between diversity and productivity shifts in strength as a result of
75 changes in forest structure which unfold along environmental gradients (Condés, Del Rio &
76 Sterba 2013). For instance, Potter & Woodall (2014) showed that across the USA forest
77 biomass is most strongly related to species richness at sites where unfavourable climatic
78 conditions result in low stem packing densities.

79 In addition to focusing on spatial environmental gradients, it has also been suggested that
80 interannual variation in climate may influence the strength of the relationship between
81 diversity and productivity (Forrester 2014). In most cases, the expectation seems to be that
82 temporal responses to environmental stress should mirror those observed along spatial
83 environmental gradients. Specifically, years characterized by unfavourable climatic

84 conditions are predicted to exhibit stronger relationships between diversity and AWP (Fig.
85 1b). However, evidence that the relationship between diversity and forest productivity shifts
86 consistently between years as a function of climate is far from clear cut, as studies have
87 reported both stronger and weaker diversity effects in stressful years (del Río, Schütze &
88 Pretzsch 2013; Jucker *et al.* 2014b; Grossiord *et al.* 2014). Part of the problem is that most
89 studies so far have only compared the response of two contrasting years (e.g., normal vs
90 drought year). Consequently, it is hard to know whether underlying patterns are going
91 unnoticed because of the incomplete temporal resolution of most studies (i.e., comparisons
92 made exclusively for a select number of years, usually two), or if instead the effects of tree
93 diversity on productivity simply do not vary systematically in response to interannual
94 fluctuations in climate.

95 Here we take advantage of the FunDivEUROPE permanent plot network – a novel research
96 platform designed specifically to quantify the functional significance of biodiversity in
97 mature forests (Baeten *et al.* 2013) – to test how the effects of tree diversity on productivity
98 change along both spatial and temporal environmental gradients. We use tree ring data
99 collected at six sites across Europe to estimate the annual AWP of 209 forest plots over a 15
100 year period (1997 – 2011). In a first step, we relate trends in productivity to climate data with
101 the aim of identifying the primary climatic drivers of AWP both across sites (i.e., spatial
102 patterns) and within them (i.e., temporal patterns). We then test whether the relationship
103 between tree species richness and productivity changes (i) across sites as a function of large-
104 scale gradients in climatic productivity and tree packing density (Fig. 1a), and (ii) among
105 years within each site as a result of fluctuating climatic conditions (Fig. 1b).

106 **Materials and Methods**

107 STUDY DESIGN

108 The study was conducted in the FunDivEUROPE permanent forest plot network, the design
109 of which is described in detail in Baeten *et al.* (2013). Here we summarize the key features of
110 the FunDivEUROPE network, which we supplement with a detailed description of the
111 rationale and design of the project in Appendix S1 of Supporting Information. The plot
112 network consists of 209 permanent forest plots (30 × 30 m in size) distributed in six countries
113 across Europe (hereafter referred to as “sites”). Field sites were chosen to represent major
114 European forest types, and include boreal forests in Finland, hemiboreal mixed forests in
115 Poland, beech forests in Germany, mountainous beech forests in Romania, thermophilous
116 deciduous forests in Italy and Mediterranean mixed forests in Spain. At each site, plots
117 contain different combinations of locally dominant tree species, and range in species richness
118 from 1-3 in Finland, 1-4 in Romania, Germany, Italy and Spain, and 1-5 in Poland (see Table
119 S1 for a full species list). Each target tree species is found in monoculture and is represented
120 in all species richness levels, with the majority of species’ combinations being replicated at
121 least twice (59 of 91 combinations). This nested design approach was modelled around that
122 of the BIOTREE, ORPHEE and FORBIO tree diversity experiments (see Baeten *et al.*, 2013
123 and references therein). Plots were selected following a set of strict criteria designed to allow
124 robust comparisons among species richness levels. Specifically, species in mixture share
125 similar relative abundances (i.e., high species evenness; Fig. S3) and the presence of non-
126 target species is minimal (< 5% of the total basal area). In addition, plots were established
127 exclusively in mature forest stands (i.e., those at least in the mid-to-late stages of stem
128 exclusion) with similar management histories. Lastly – and critically for the purposes of this

129 study – great care was taken to ensure that (i) plots within a site share similar environmental
130 conditions and (ii) that stand attributes known to influence productivity (e.g., stem density
131 and stand age; Magnani, Mencuccini & Grace 2000; Kadmon & Benjamini 2006) were not
132 confounded with the species richness gradient (see Appendix S1 and Fig. S4). Consequently,
133 while sites differ markedly from one another in terms of climate and forest structure (e.g.,
134 basal area, stem density, maximum height), plots within a site are spatially clustered and have
135 similar elevation, topography and soil type.

136 ALLOMETRIC DATA

137 In each plot, all stems ≥ 7.5 cm in diameter were identified to species and permanently
138 marked (12939 stems in total). We recorded the diameter (measured to the nearest 0.1 cm at a
139 height of 1.3 m off the ground using diameter tape) and height (to the nearest 0.1 m using a
140 vertex hypsometer, Haglöf AB, Sweden) of each stem, and used these measurements to
141 estimate the above-ground biomass of each tree based on published biomass functions (see
142 Jucker *et al.*, 2014a for a complete list of equations used). We also characterized the crown
143 dominance of each tree using the crown illumination index, which consists in scoring trees
144 from 1 (suppressed crown receiving no direct light) to 5 (fully exposed dominant crown)
145 (Clark & Clark 1992). Crown illumination scores are an effective way of gauging the degree
146 to which a tree is shaded by its neighbours (Jennings, Brown & Sheil 1999; Jucker *et al.*
147 2014b), and were used to model the effects of competition for light on tree growth (see
148 following section).

149 CLIMATE DATA

150 We obtained data on air temperature (T; °C), precipitation (P; mm), solar radiation (R; J cm⁻²), actual evapotranspiration (AET; mm), potential evapotranspiration (PET; mm) and the
151 ratio between precipitation and potential evapotranspiration (P/PET; a measure of water
152 availability) in the form of daily values covering the period between 1997 and 2011 for each
153 of the six study sites. With the exception of AET, which was derived from the water balance
154 model BILJOU (Granier *et al.* 1999; <https://appgeodb.nancy.inra.fr/biljou/>), all climatic data
155 were obtained directly from the CGMS database of interpolated meteorological data
156 (AGRI4CAST; <http://mars.jrc.ec.europa.eu/mars>). The AGRI4CAST system combines
157 meteorological observations, remote sensing imagery and agro-meteorological modelling to
158 generate climate surfaces for Europe (25 × 25 km resolution).
159

160 ESTIMATING ANNUAL ABOVE-GROUND WOOD PRODUCTION FROM TREE RING DATA

161 We used tree ring data to obtain a time series of annual above-ground wood production
162 (AWP; Mg C ha⁻¹ yr⁻¹) for all 209 plots covering the 15 year period between 1997 and 2011.
163 The approach used to estimate AWP from tree ring data is outlined in detail in Jucker *et al.*,
164 (2014a). Here we summarize the main analytical steps of this workflow.

165 *Measuring individual tree growth from wood cores*

166 We used 5.15 mm diameter increment borers (Haglöf AB, Sweden) to extract wood cores
167 from 2950 trees across the plot network following a size-stratified random sampling approach
168 (Nehrbass-Ahles *et al.* 2014). Once mounted and sanded, wood cores were scanned using a
169 high resolution flatbed scanner (2400 dpi). From the scanned images we measured annual
170 radial growth increments for all cored trees using CDendro (Cybis Elektronik & Data,

171 Saltsjöbaden, Sweden). Finally, radial increments for each year between 1997 and 2011 were
172 converted to annual biomass growth (kg C yr^{-1}) using the biomass functions described above.
173 Biomass growth was expressed in units of carbon by applying the standard conversion of 0.5
174 g C per gram of biomass.

175 *Modelling individual tree biomass growth*

176 For each year between 1997 and 2011, we used linear mixed-effects models coded in R
177 (3.0.1; R Core Development Team, 2013) to model the biomass growth (G) of each target
178 species as a function of tree size, competition for light, species richness and a random plot
179 effect:

$$180 \quad \log(G_i) = \beta_{0j[i]} + \beta_1 \log(D_i) + \beta_2 CI_i + \beta_3 SR_j + \varepsilon_i \quad (\text{eqn 1})$$

181 where D_i and CI_i are, respectively, the stem diameter and crown illumination index of tree i
182 growing in plot j (as measured in 2011); SR_j is the species richness of plot j ; $\beta_{0j[i]}$ is a
183 species' intrinsic growth rate for a tree growing in plot j (*sensu* Rüger *et al.*, 2012); β_{1-3} are,
184 respectively, a species' growth response to size, light availability and species richness; and ε_i
185 is the residual error. Equation 1 effectively captured variation in biomass growth within and
186 among species [average R^2 across species = 0.80; for further information see Jucker *et al.*
187 (2014a)].

188 *Scaling up from tree-level biomass growth to plot-level AWP*

189 Equation 1 was used to estimate the annual biomass growth of all trees that had not been
190 cored for each year between 1997 and 2011. The annual AWP of each plot was then
191 quantified by summing the biomass growth of all trees within a plot. This step was repeated
192 for each year between 1997 and 2011 to generate a 15-year time series of AWP for all 209

193 plots. This approach to estimating AWP was chosen after having ruled out potential biases
194 associated with temporal autocorrelation in the AWP time series (see Appendix S2).
195 Similarly to previous diversity – productivity studies in forests (e.g., Paquette & Messier
196 2011; Vilà *et al.* 2013), estimates of AWP are based exclusively on the growth of trees
197 present in 2011, with the implicit assumption that mortality rates are unrelated to tree species
198 diversity (Liang *et al.* 2007).

199 RELATING VARIATION IN AWP ACROSS AND WITHIN SITES TO CLIMATE

200 We combined data on AWP and climate to identify which climatic predictors best explain
201 variation in wood production across sites (i.e., spatial variation in AWP) and within them
202 (i.e., interannual variation in AWP). To understand how differences in climate among sites
203 shape geographic patterns of wood production, we calculated mean annual values of T, P, R,
204 AET, PET and P/PET for each study site and related these to mean annual AWP using
205 univariate regressions (Appendix S3). Following the same approach, we also tested the ability
206 of climatic predictors to explain interannual variation in AWP within sites. For each site, we
207 aggregated daily climate values to obtain both yearly and seasonal (spring = March, April,
208 May; summer = June, July, August; autumn = September, October, November) estimates of
209 T, P, R, AET, PET and P/PET for each year between 1997 and 2011. We then fitted
210 univariate regressions relating interannual variation in AWP within sites to each climatic
211 predictor in turn (Appendix S3).

212 STATISTICAL MODELLING OF AWP – DIVERSITY RELATIONSHIPS

213 We devised a two-step approach to test whether AWP – diversity relationships change
214 predictably (i) across sites and (ii) among years within a site (Fig. 1). We first used a linear

215 regression model to estimate the effects of species richness on AWP (i.e., the slope of the
 216 AWP – species richness relationship) for (i) each site and (ii) for every year within a site. We
 217 then related the slope estimates obtained from the model to (i) differences in climate and
 218 packing density among sites, and (ii) interannual variation in climate within sites (e.g., Reich
 219 *et al.*, 2012; Wu *et al.*, 2014). We favoured this approach over one in which climatic effects
 220 on AWP were modelled directly, as it allows the strength of the species richness effect on
 221 AWP to be intuitively compared both among and within sites (Nakagawa & Cuthill, 2007).

222 *Step 1: Estimate AWP – species richness slopes across and within sites*

223 We modelled AWP as a function of (i) species richness, (ii) plot basal area (to account for the
 224 effects of stem packing density on productivity), (iii) site (i.e., accounting for variation in
 225 AWP among study sites), (iv) year nested within site (i.e., allowing AWP to vary among
 226 years within each site), (v) an interaction term between species richness and site (testing
 227 whether species richness effects on AWP vary among sites), and (vi) an interaction term
 228 between species richness and year nested within site (to determine whether species richness
 229 effects on AWP vary among years within sites):

$$\begin{aligned}
 230 \quad \log(AWP_{ijk}) = & \beta_0 + \beta_1 \log(BA_{ij}) + \beta_2 \log(SR_{ij}) + \beta_3 \text{Site}_j + \beta_4 \text{Year}_{jk} + \\
 231 \quad & \beta_5 [\log(SR_{ij}) \times \text{Site}_j] + \beta_6 [\log(SR_{ij}) \times \text{Year}_{jk}] + \varepsilon_{ijk} \quad (\text{eqn 2})
 \end{aligned}$$

232 where AWP_{ijk} is the above-ground wood production of plot i in site j in year k ; SR_{ij} and BA_{ij}
 233 are, respectively, the species richness and basal area of plot i in site j ; Site_j and Year_{jk} are
 234 categorical grouping variables which capture how AWP varies among the j sites and among
 235 the k years within each of the j sites, respectively; $\beta_0 - \beta_6$ are parameters to be estimated from
 236 the data (i.e., intercept and slope coefficients); and ε_{ijk} is the residual error.

237 In equation 2, both “Site” and “Year” were modelled as fixed effects so that we could
238 estimate uncertainties for all parameters. To verify the robustness of the parameter estimates
239 obtained from equation 2 (particularly β_5 and β_6 , which define how the effect of tree diversity
240 on AWP varies across and within sites), we first compared them with those estimated using a
241 mixed-effects modelling approach. In addition to this, we also fitted an additional model in
242 which we accounted for a number of potentially confounding factors which could
243 conceivably alter the relationship between tree diversity and AWP (elevation, terrain slope,
244 soil depth, rock cover and stand age). As all three approaches yielded almost identical results
245 (see Appendix S4), here we focus on the parameter estimates obtained through equation 2.
246 Note that support for the interaction terms in equation 2 was assessed through comparison
247 with simpler models lacking interacting effects.

248 *Step 2: Relate variation in slopes to climate and packing density*

249 From equation 2 we obtained the slope of the AWP – species richness relationship for (i)
250 each site (i.e., β_5) and (ii) for all years within a site (i.e., β_6). We then used Pearson’s
251 correlation coefficients (r) to explore whether (i) variation in the strength of the species
252 richness effect among sites is related to differences in macroclimate and/or basal area (a
253 measure of tree packing density which captures both the mean size and number of trees
254 within a plot) (Curtis & Marshall 2000), and (ii) whether the influence of species richness on
255 AWP changes over time within a site depending on how suitable climatic conditions are in a
256 given year.

257 **Results**

258 INFLUENCE OF CLIMATE AND PACKING DENSITY ON AWP

259 AWP varied considerably among sites (M2 vs M1 in Table 1), differing almost fivefold
260 between Spain ($0.65 \text{ Mg C ha}^{-1} \text{ yr}^{-1}$) – the least productive site – and Romania ($3.11 \text{ Mg C ha}^{-1}$
261 yr^{-1}), where AWP rates were highest. Geographic variation in AWP was strongly related to
262 differences in evapotranspiration among sites (Fig. 2a; see Appendix S3 for comparison with
263 other climatic drivers). Sites where annual AET was low – either because of insufficient
264 rainfall (e.g., Spain) or low temperatures (e.g., Finland) – had much slower rates of AWP
265 than those where AET was high (e.g., Germany and Romania). However, AET was only a
266 weak predictor of interannual variation in AWP within sites (see Table S2 and following
267 paragraph). Instead, we found that increased annual AET was associated with the
268 development of forests with greater basal areas (Fig. S7). This in turn resulted in markedly
269 faster rates of AWP – as basal area was a key driver of AWP across the plot network (Fig. 2b
270 and Table 1) – suggesting that at least in part the link between evapotranspiration and
271 productivity is mediated through changes in stand packing density.

272 In addition to varying across sites, AWP also fluctuated strongly among years within each
273 site (M3 vs M2 in Table 1). Climate explained much of this interannual variation in
274 productivity (Fig. 3), with a clear distinction emerging between Finland – where AWP was
275 co-limited by temperature and solar radiation – and all other sites – where annual growth was
276 instead primarily water-limited (Table 2). In Finland, the single best climatic predictor of
277 annual AWP was spring PET (Fig. 3b and Table 2), with productivity peaking in years
278 characterized by high solar radiation and mild spring temperatures which together contributed

279 to high PET rates (Table S2). For all other sites, fluctuations in annual AWP was best
280 captured by P/PET (Table 2), with productivity declining sharply in years when evaporative
281 demands (i.e., PET) were not met by precipitation (Fig. 3a,c-f). The timing of drought proved
282 equally important in explaining AWP patterns. For both Mediterranean sites, low P/PET
283 values during spring months were associated with strong reductions in AWP (Fig. 3a,c),
284 while summer drought had a much less pronounced influence on productivity (Table S2). In
285 contrast, for Germany and Poland the best predictor of AWP was P/PET integrated over the
286 entire year (Fig. 3d-f), while in Romania AWP was influenced by autumn P/PET (Fig. 3e).
287 Note that for all sites, AWP – climate relationships were best captured by linear functions
288 (see Table 2 for estimated regression coefficients).

289 AWP – DIVERSITY RELATIONSHIPS ACROSS SITES

290 We found strong support for the inclusion of the interaction term between species richness
291 and “site” in equation 2 (M4 vs M3 in Table 1), indicating that the effect of species richness
292 on AWP varies substantially across sites. In accordance with the stress gradient hypothesis
293 (Fig. 1a), the slope of the AWP – species richness relationship decreased progressively along
294 the AET climatic productivity gradient ($r = -0.88$, $P = 0.020$, $n = 6$; Fig. 2c), ranging from
295 strongly positive in Spain (0.24 ± 0.06) – where evapotranspiration and productivity were
296 low – to weakly negative in Romania (-0.05 ± 0.07) – where AET and AWP were greatest.
297 This shift in the strength of the diversity effect on productivity was equally well captured by
298 accounting for differences in basal area across sites ($r = -0.90$, $P = 0.013$, $n = 6$), with
299 diversity effects being strongest at sites where environmental conditions led to the
300 development of less densely packed stands (Fig. 2d and Fig. S7).

302 In contrast to patterns across sites, we found little evidence to suggest that species richness
303 effects on AWP varied strongly among years within sites (M5 vs M4 in Table 1).
304 Furthermore, we found no consistent relationship between the strength of the species richness
305 effect on AWP and climatic conditions within a given year when looking across the six sites
306 (Fig. 4). The only site to show a relationship between the magnitude of the species richness
307 slope and climate was Finland (Fig. 4b), where species richness effects weakened in years
308 when spring PET was high and climatic conditions for growth were favourable ($r = -0.73$, P
309 < 0.01 , $n = 15$). A similar response was found for Romania, although the pattern between
310 diversity effects and climate (autumn P/PET) was weaker ($r = -0.44$, $P = 0.10$, $n = 15$; Fig.
311 4f). Instead, in the case of Poland (Fig. 4d), and to a much lesser extent Spain (Fig. 4a), we
312 observed the opposite trend, with diversity effects strengthening during non-drought years (r
313 $= 0.49$, $P = 0.06$ and $r = 0.21$, $P = 0.45$, respectively; $n = 15$). These observations do not
314 support the predictions of the stress gradient hypothesis (Fig. 1b). Lastly, in the case of Italy
315 (Fig. 4c) and Germany (Fig. 4e) we found no discernible pattern linking diversity effects and
316 climatic conditions across years ($r = 0.03$, $P = 0.93$ and $r = 0.05$, $P = 0.87$, respectively; $n =$
317 15).

318 **Discussion**

319 In line with a number of recent studies, we found a generally positive relationship between
320 tree diversity and forest productivity across Europe (Morin *et al.* 2011; Vilà *et al.* 2013;
321 Pretzsch *et al.* 2015). Importantly, however, our results also highlighted that the effects of
322 species richness on productivity are strongly context dependent, varying in space and – to a
323 lesser extent – time. Specifically, AWP – species richness relationships shifted predictably in
324 strength across sites, becoming progressively stronger under harsher environmental
325 conditions (Fig. 2). In contrast, we found that interannual variation in the strength of species
326 richness effects within sites was rather weak and could not be consistently explained by year-
327 to-year fluctuations in climate (Fig. 4).

328 **AWP – DIVERSITY RELATIONSHIPS ARE STRONGEST IN STRESSFUL ENVIRONMENTS**

329 Macro-scale patterns in wood production were best captured by differences in
330 evapotranspiration among sites (Fig. 2a), confirming what has previously been reported in the
331 literature (Stephenson 1998; Oberle, Grace & Chase 2009). Evapotranspiration integrates the
332 effects of water availability and temperature on plant growth (Stephenson 1998; Boisvenue &
333 Running 2006): in order for trees to grow, temperatures need to be sufficient to initiate
334 photosynthesis and allow the microbially driven mobilization of soil nutrients, and at the
335 same time enough water needs to be available to meet evaporative demands. Our results also
336 suggest that the link between evapotranspiration and productivity is largely mediated through
337 changes in stand basal area which unfold across AET gradients (Fig. S7), as opposed to a
338 direct influence of AET on AWP (Stephenson 1998; Coomes *et al.* 2014; Michaletz *et al.*
339 2014). While interannual fluctuations in AET were not a particularly strong predictor of

340 variation in productivity within sites (Tables S2), sites with higher mean annual AET also
341 had greater basal areas, which in turn drove faster rates of wood production (Fig. 2b).

342 The strength of the species richness effect on productivity also shifted along the
343 evapotranspiration gradient, and was strongest at sites where AET was low – either as a result
344 of low annual rainfall (e.g., Spain) or because of low temperatures (e.g. Finland) – and
345 conditions for growth were poor (Fig. 2c). Conversely, at sites where conditions for growth
346 were more favourable, we saw a drastic reduction in the importance of diversity as a driver of
347 wood production. These results are broadly consistent with the predictions of the stress
348 gradient hypothesis (Fig. 1a), as well as the findings of a number of recent papers (Paquette
349 & Messier 2011; Jucker & Coomes 2012; Wu *et al.* 2014; Toïgo *et al.* 2015). For instance,
350 using forest inventory data from France, Toïgo *et al.* (2015) showed that overyielding was
351 more frequent at low productivity sites. Similarly, Wu *et al.* (2014) found that the strength of
352 the relationship between tree diversity and biomass weakened when transitioning from
353 boreal, to temperate and subtropical forests in China.

354 In addition to the stress gradient hypothesis, a number of alternative frameworks have also
355 been proposed to explain how environmental conditions shape the relationship between
356 diversity and productivity. For instance, Forrester (2014) noted that the outcome of species
357 interactions will depend on the type of limiting resource or resources (e.g., water, light,
358 nutrients), and whether – on average – interactions among neighbouring trees improve the
359 availability of those resources (also see Pretzsch *et al.* 2015). What our results suggest is that
360 the net outcome of species interactions shifts along environmental gradients and that – on
361 balance – competition tends to decrease in importance in more stressful environments
362 allowing greater room for complementarity (Kunstler *et al.* 2011; Prior & Bowman 2014).

363 Climate can influence interactions among neighbouring trees directly by modulating the
364 relative performance and strategy of species within a community (Brooker *et al.* 2008;
365 Maestre *et al.* 2009). In addition to this, climate can influence interactions among
366 neighbouring trees indirectly through its effect on forest structure. Climate plays an important
367 role in determining local species pools through environmental filtering (e.g., Stahl *et al.*,
368 2014), as well as shaping species demographic rates (Lines, Coomes & Purves 2010;
369 Vanderwel, Lyutsarev & Purves 2013; Coomes *et al.* 2014) and influencing how trees
370 allocate carbon above and below-ground (Lines *et al.* 2012; Reich *et al.* 2014), all of which
371 interact to determine forest structure. As our results suggest that positive AWP – diversity
372 relationships are much more likely in forests with low packing densities (Fig. 2d), the effect
373 of climate on forest structure appears to be particularly important in modulating the effects of
374 diversity on productivity (Condés *et al.* 2013; Potter & Woodall 2014). This may explain why
375 in Mediterranean forests, where drought impedes the development of densely packed stands,
376 the effects of diversity on productivity tend to be strongly positive (Vilà *et al.* 2007; Ruiz-
377 Benito *et al.* 2014; Jucker *et al.* 2014b).

378 INTERANNUAL FLUCTUATIONS IN CLIMATE HAVE WEAK AND IDIOSYNCRATIC EFFECTS ON AWP 379 –DIVERSITY RELATIONSHIPS WITHIN SITES

380 AWP responses to interannual fluctuations in climate within sites closely matched
381 expectations based on published reports, with most forest types in Europe being primarily
382 limited by water availability, while productivity in boreal forests depends strongly on
383 temperature and solar radiation (Boisvenue & Running 2006; Babst *et al.* 2013). In terms of
384 AWP – diversity relationships, we generally found that variation within sites was
385 considerably weaker than variation among them (Table 1). Moreover, we found no evidence

386 of consistent pattern linking variation in diversity effects across years to climate (Fig. 4).
387 Certain functional responses match expectations, such as the decreased strength of diversity
388 effects during warmer years in Finland (Fig. 4b). In boreal systems, warmer springs are
389 associated with earlier leaf-out dates (Polgar & Primack 2011). While this positively
390 influences productivity by increasing the length of the growing season (Chen *et al.* 1999;
391 Polgar & Primack 2011), it is also likely to reduce the degree of phenological mismatch (i.e.,
392 temporal complementarity) between evergreen conifers and deciduous broadleaf species
393 (Sapijanskas *et al.* 2014). In contrast, we found no consistent effect of drought on the strength
394 of diversity effects across forest types (Grossiord *et al.* 2014). For Spain, our results generally
395 confirm those of a previous study comparing responses in two years with contrasting climate
396 (a drought year and a wet year), which showed that drought tends to exacerbate competition
397 among neighbouring trees (Jucker *et al.* 2014b). However, in the case of Poland, Germany
398 and Romania we find strongly contrasting patterns despite the fact that these sites share
399 similar species compositions and have broadly similar climates.

400 Differences between responses across sites – which strongly matched theoretical predictions
401 – and those within sites – which instead exhibited weak and contrasting patterns – suggest
402 that in addition to climate itself, a key factor in shaping the strength of diversity effects on
403 productivity is how densely trees pack in space. If the relationship between diversity and
404 productivity was directly influenced by climate, diversity effects would be expected to shift
405 consistently both across and within sites. Instead, under the assumption that packing density
406 is what ultimately determines the strength of the relationship between diversity and
407 productivity (Condés *et al.* 2013; Potter & Woodall 2014), temporal patterns would be
408 unlikely to match spatial trends. The reason for this is that although basal area varies strongly

409 among sites (Fig. 2b), changes in packing density are decoupled from interannual variation in
410 climate in forest ecosystems (i.e., basal area does not increase or decrease markedly from one
411 year to next based on climatic suitability). Recent work in coastal dunes plant communities
412 has shown that shifts from competition to facilitation among years depend critically on
413 differences in vegetation cover: during years characterized by harsh environmental conditions
414 vegetation cover decreases, resulting in an increase in facilitation, while in more favourable
415 years increased vegetation cover exacerbates competition (Doxford, Ooi & Freckleton 2013).
416 However, because trees are long-lived and do not regenerate the bulk of their above-ground
417 biomass each year, a similar response in terms of basal area cannot be observed in forests
418 (Stephenson & Mantgem 2005). As a result, the way in which diversity effects change among
419 years may simply be more context dependent in the case of forests [e.g., highly dependent on
420 species' traits or on the type of limiting resource; Forrester (2014)].

421 IMPLICATIONS FOR FOREST CONSERVATION AND MANAGEMENT

422 Identifying where and when tree diversity has the greatest potential to positively influence
423 forest productivity has important implications for forest management and conservation
424 practises, as well as efforts to mitigate climate change (Zhang *et al.* 2012; Cardinale *et al.*
425 2012; Scherer-Lorenzen 2014). We found that in terms of maximising rates of wood
426 production, the benefits of maintaining diverse forests are most pronounced in systems where
427 environmental conditions strongly limit productivity. While our study provides a useful
428 framework for predicting under which conditions tree diversity is likely to matter most, there
429 are however several reasons why practises aimed at maintaining diverse forests should not
430 necessarily be limited to specific ecological contexts or geographic regions. For instance, in
431 addition to promoting forest productivity, tree diversity has also been shown to help stabilize

432 wood production over time across a range of forest types, highlighting the fact that mixed-
433 species forests are able to remain productive under a wider range of environmental conditions
434 than monocultures (Jucker *et al.* 2014a; Morin *et al.* 2014). Secondly, although we detected
435 clear differences in the importance of tree diversity as a driver of productivity among forest
436 types, only one site showed any indication of a negative association between diversity and
437 productivity. Consequently, even though gains in productivity may be negligible for certain
438 forest types, maintaining diverse forests is unlikely to adversely affect wood production and
439 has the advantage of delivering a number of added ecological and economic co-benefits (e.g.,
440 reduced risk of pest and pathogen outbreaks, increased associated biodiversity, greater soil
441 carbon storage; Scherer-lorenzen, 2014).

442 In addition to highlighting under which circumstance tree diversity is currently most
443 important for forest productivity, our study also provides a number of clues as to how climate
444 change is likely to influence AWP – diversity relationships in future forests. By the end of
445 this century Mediterranean forests in Europe are expected to suffer more frequent and
446 prolonged periods of drought, while boreal systems are predicted to warm considerably and
447 experience longer growing seasons (Jacob *et al.* 2014). As a result, diversity effects may
448 weaken in strength, particularly in the case of boreal forests where spring warming is
449 expected to reduce the degree of phenological mismatch among coexisting tree species
450 (Polgar & Primack 2011). More importantly, however, our results suggest that longer-term
451 responses of forests to climate change – such as changes in species composition and forest
452 structure – are what will ultimately determine how the relationship between tree diversity and
453 forest productivity will look like in the future.

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461 **Data Accessibility**

462 Above-ground wood production (AWP) data: uploaded as online supporting information

463 **References**

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673

674 **Supporting information**

675 Additional supporting information may be found in the online version of this article:

676 **Appendix S1:** FunDivEUROPE project

677 **Appendix S2:** Temporal autocorrelation in AWP

678 **Appendix S3:** Relating variation in AWP to climate

679 **Appendix S4:** Estimating diversity effects across and within sites

680 **Appendix S5:** Above-ground wood production (AWP) data archive

681

682 **Tables**

683 **Table 1:** Model comparison statistics for a set of candidate models testing whether diversity
 684 effects on above-ground wood production (AWP) vary across and within sites. In the baseline
 685 model (M1), AWP is expressed as a function of plot basal area (BA) and species richness
 686 (SR). “Site” indicates a model which accounts for variation in AWP among sites, while
 687 “Year” tests whether AWP varies among years within each site. Model M4 tests whether the
 688 effects of SR on AWP varies among sites (i.e., interaction between SR and Site), while model
 689 M5 further allows species richness effects to vary among years within a site (i.e., interaction
 690 between SR and Year) and corresponds to equation 2 in the text. Models of increasing
 691 complexity were compared to simpler ones (e.g., M5 vs M4) using F-tests to determine
 692 whether the reduction in the residual sum of squares between models was statistically
 693 significant ($\alpha = 0.05$). In addition to this we also report the R^2 and AIC of each model. Note
 694 that AWP, SR and BA were log-transformed prior to model fitting.

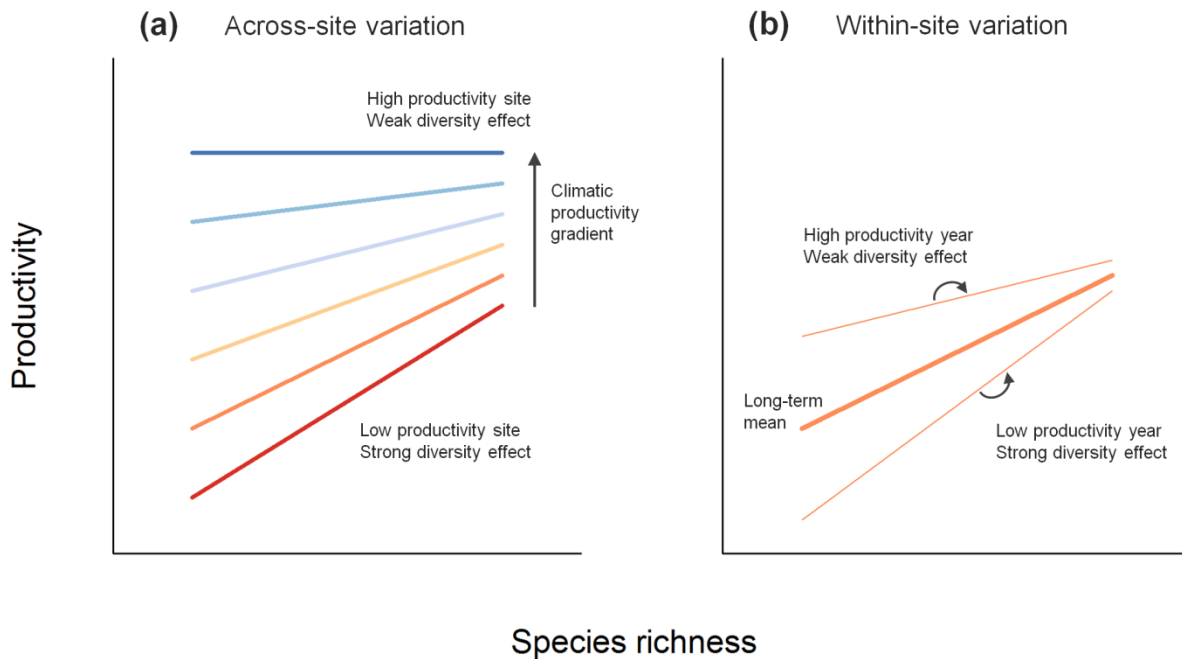
Model	Model structure	<i>P</i> (>F)	R^2	AIC
M0	AWP ~ BA	–	0.40	4306
M1	AWP ~ BA + SR	<0.0001	0.42	4276
M2	AWP ~ BA + SR + Site	<0.0001	0.74	1738
M3	AWP ~ BA + SR + Site + Year	<0.0001	0.78	1291
M4	AWP ~ BA + SR + Site + Year + SR×Site	<0.0001	0.79	1248
M5	AWP ~ BA + SR + Site + Year + SR×Site + SR×Year	n.s.	0.79	1384

695

696 **Table 2:** Best climatic predictors of interannual variation in above-ground wood production
697 (AWP) for each study site. Regression equations (with 95% confidence intervals for
698 parameter estimates in brackets) correspond to fitted relationships illustrated in Fig. 3. To
699 facilitate the interpretation of regression coefficients, climatic predictors were cantered prior
700 to model fitting (i.e., for each site, intercepts indicate the mean AWP across years). PET =
701 potential evapotranspiration; P/PET = precipitation/PET.

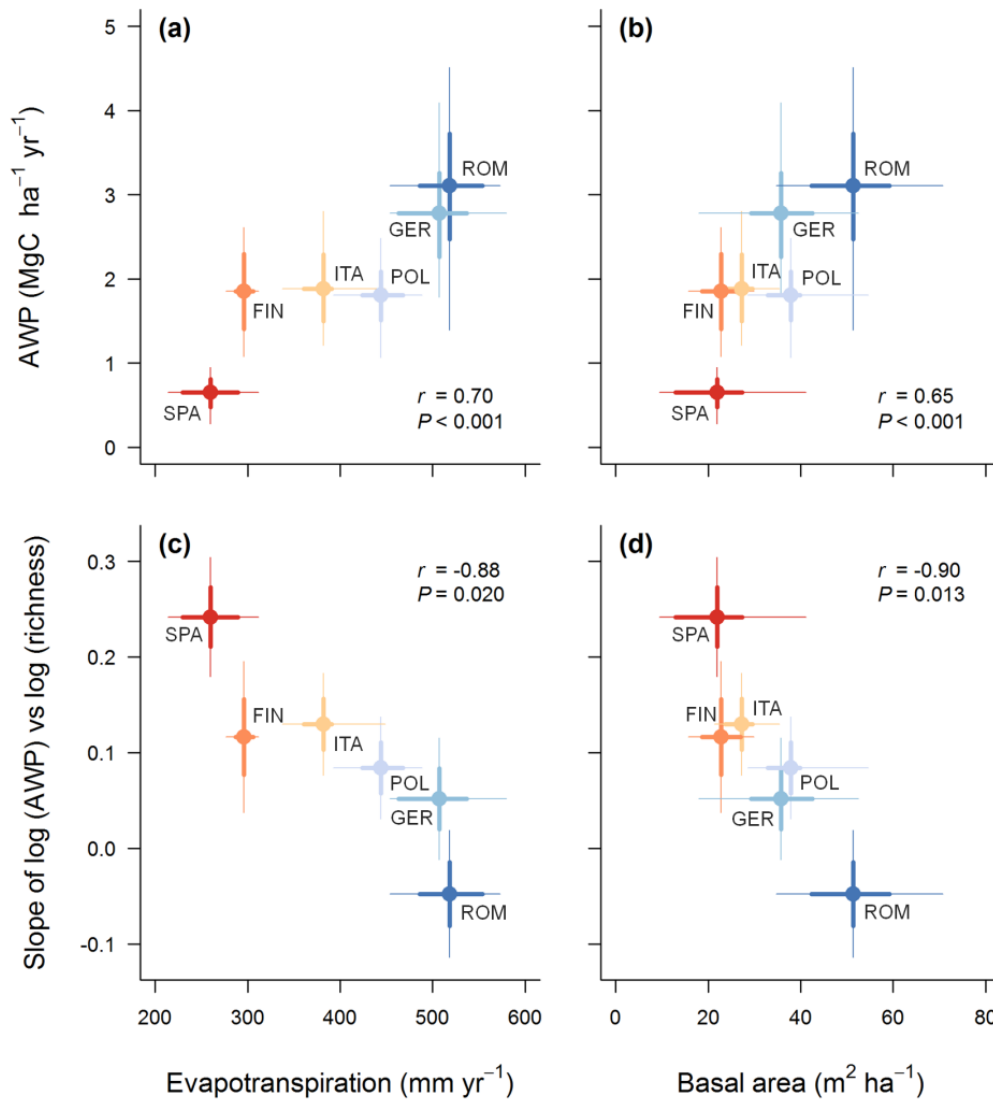
Site	Climatic predictor	Regression equation
Finland	Spring PET	$AWP = 1.85 (0.06) + 0.012 (0.003) \times PET$
Poland	Annual P/PET	$AWP = 1.81 (0.04) + 0.87 (0.23) \times (P/PET)$
Germany	Annual P/PET	$AWP = 2.78 (0.07) + 1.11 (0.44) \times (P/PET)$
Romania	Autumn P/PET	$AWP = 3.11 (0.11) + 0.40 (0.14) \times (P/PET)$
Italy	Spring P/PET	$AWP = 1.88 (0.05) + 0.48 (0.24) \times (P/PET)$
Spain	Spring P/PET	$AWP = 0.65 (0.02) + 0.28 (0.09) \times (P/PET)$

702



704

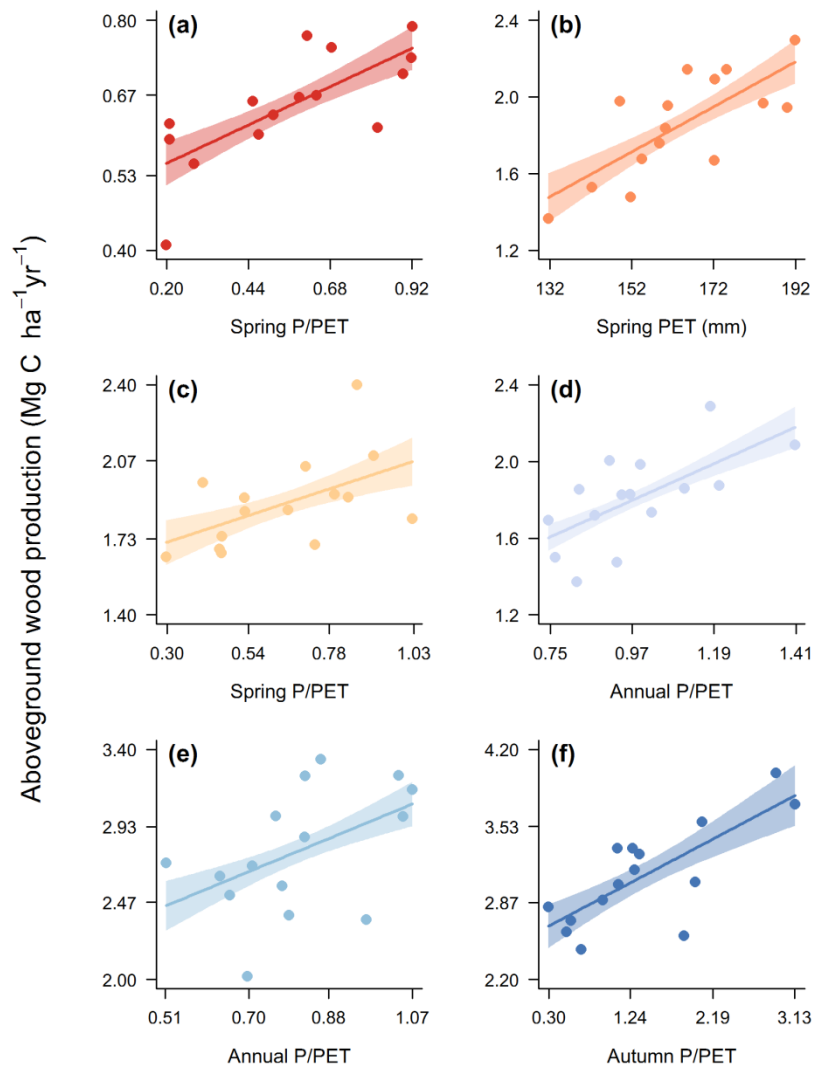
705 **Fig. 1:** Schematic diagram illustrating how the relationship between productivity and
 706 diversity is expected to change (a) across sites along an environmental gradient and (b)
 707 among years within a site. Panel (a) depicts different sites along an environmental gradient.
 708 Sites where productivity is low (e.g., as a result of strong environmental limitations on
 709 growth) exhibit strong positive effects of diversity on productivity, while sites where growing
 710 conditions are more favourable show little or no benefit from species mixing. Panel (b)
 711 illustrates how the effects of diversity on productivity vary among years within a given site.
 712 The expectation is that the slope of the diversity – productivity relationship will be steeper
 713 than average in low productivity years, weakening instead when conditions for growth
 714 improve.



715

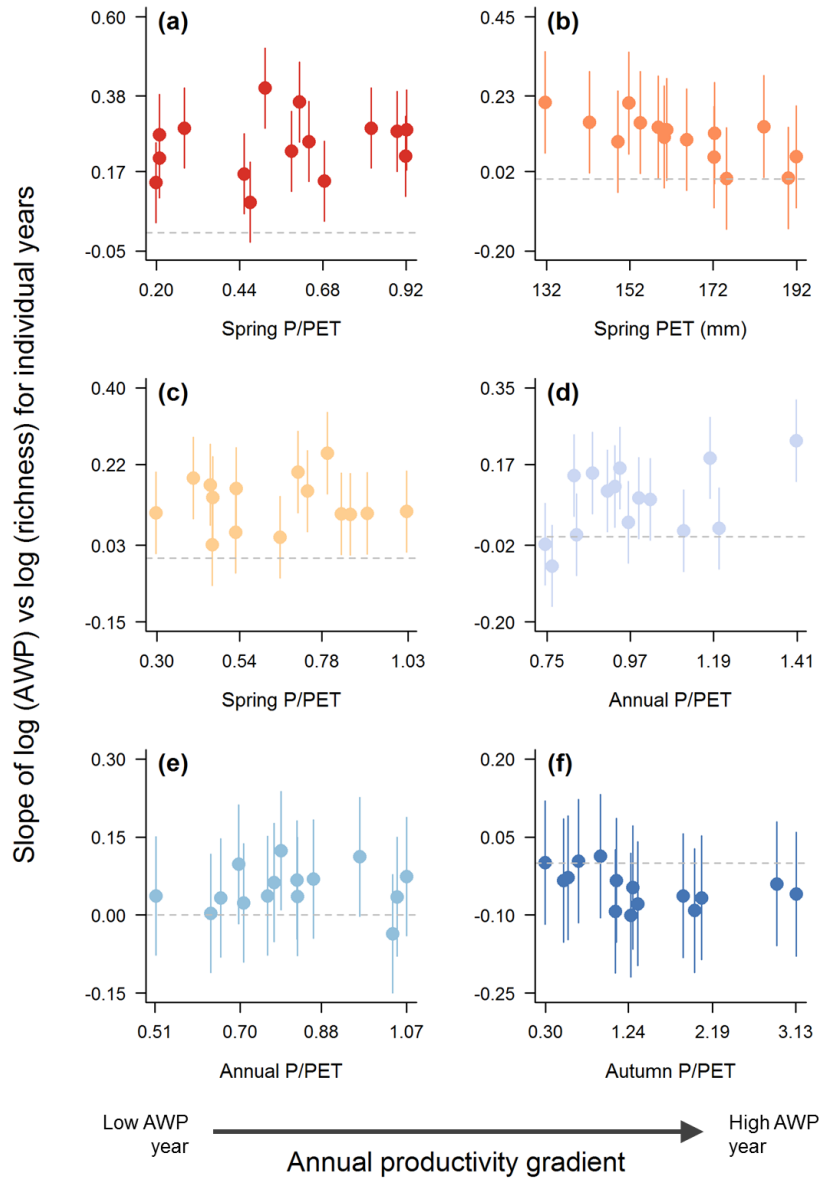
716 **Fig. 2:** Variation in (a–b) above-ground wood production (AWP) and (c–d) the slope of the
 717 AWP – species richness relationship among sites. Panels on the left show how
 718 evapotranspiration (AET) influences (a) AWP and (c) the strength of the AWP – species
 719 richness relationship, while right-hand panels (b and d) illustrate variation in response to plot
 720 basal area. For AWP and basal area, points represent mean values across all plots within a
 721 site, while AET is the long-term mean annual evapotranspiration registered at each site
 722 (between 1997 and 2011). Error bars delimit the interquartile range (thick lines) and 95%
 723 limits (thin lines) of the data. Slopes of log (AWP) vs log (species richness) were obtained

724 from equation 2 in the main text (thick lines: ± 1 SE; thin lines: ± 2 SE). Points are labelled
725 using the first three letters of each site. Pearson's correlation coefficients (r) and
726 corresponding P -values reported in the top panels (a–b) were calculated for $n = 209$
727 (corresponding to the number of FunDivEUROPE forest plot), while for the bottom panels
728 (c–d) $n = 6$ (corresponding to the number of FunDivEUROPE sites).



729

730 **Fig. 3:** Variation in annual above-ground wood production (AWP) as a function of climate
 731 for (a) Spain, (b) Finland, (c) Italy, (d) Poland, (e) Germany and (f) Romania. Points
 732 correspond to the mean AWP of each year (1997 to 2011; calculated across all plots within a
 733 site). Fitted relationships from linear regression models (with shaded 95% confidence
 734 intervals) are shown for each panel (see Table 2 for regression equations). PET = potential
 735 evapotranspiration; P/PET = precipitation / PET.



736

737 **Fig. 4:** Interannual variation in the slope of the AWP – species richness relationship as a
 738 function of climate in (a) Spain, (b) Finland, (c) Italy, (d) Poland, (e) Germany and (f)
 739 Romania. Slopes (± 1 SE) of log (AWP) vs log (species richness) for each year between 1997
 740 and 2011 are estimates from a linear regression model (equation 2 in the main text). Dashed
 741 grey lines mark a slope of zero. PET = potential evapotranspiration; P/PET = precipitation /
 742 PET.