

1 **The stress field in Europe: optimal orientations with confidence limits**

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9 Cite as: Carafa, M. M. C. & Barba, S., 2013. The stress field in Europe: optimal orientations with
10 confidence limits, *Geophys. J. Int.*, doi: 10.1093/gji/ggt024

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12 First published online: February 19, 2013.

13 **Summary**

14 In this study, we modify and extend a data analysis technique to determine the stress orientations
15 between data clusters by adding an additional constraint governing the probability algorithm. We
16 apply this technique to produce a map of the maximum horizontal compressive stress (S_{Hmax})
17 orientations in the greater European region (including Europe, Turkey and Mediterranean Africa).
18 Using the World Stress Map dataset release 2008, we obtain analytical probability distributions of
19 the directional differences as a function of the angular distance, θ . We then multiply the probability
20 distributions that are based on pre-averaged data within $\theta < 3^\circ$ of the interpolation point and
21 determine the maximum likelihood estimate of the S_{Hmax} orientation. At a given distance, the
22 probability of obtaining a particular discrepancy decreases exponentially with discrepancy. By
23 exploiting this feature observed in the World Stress Map release 2008 dataset, we increase the
24 robustness of our S_{Hmax} determinations. For a reliable determination of the most likely S_{Hmax}
25 orientation, we require that 90% confidence limits be less than $\pm 60^\circ$ and a minimum of three
26 clusters, which is achieved for 57% of the study area, with small uncertainties of less than $\pm 10^\circ$ for
27 7% of the area. When the data density exceeds 0.8×10^{-3} data/km², our method provides a means of
28 reproducing significant local patterns in the stress field. Several mountain ranges in the
29 Mediterranean display 90° changes in the S_{Hmax} orientation from their crests (which often
30 experience normal faulting) and their foothills (which often experience thrust faulting). This pattern
31 constrains the tectonic stresses to a magnitude similar to that of the topographic stresses.

32

33 **Keywords:** Neotectonics; Seismicity and tectonics; Fractures and faults; Intra-plate processes; Plate
34 motions; Dynamics: gravity and tectonics.

35 1. Introduction

36 This study aims to exploit the high data density of the World Stress Map release 2008 (hereinafter,
37 WSM08) (Heidbach *et al.*, 2008) in the greater European region (including Europe, Turkey and
38 Mediterranean Africa) to gain additional information on short-wavelength stress sources.
39 Knowledge of the local stress field helps to identify potential slip on pre-existing favourably
40 oriented fault planes (e.g., Sibson *et al.*, 2011; Syracuse *et al.*, 2012) and has an impact on seismic
41 hazard assessment, geodynamics, hydrocarbon exploitation, and global-to-regional-scale tectonics
42 (Carafa & Barba, 2011; Herget & Heidbach, 2011; Ghisetti & Sibson, 2012). However,
43 interpolating or smoothing the stress orientations of clustered and scattered data requires several
44 assumptions and approximations.

45 Rebai *et al.* (1992) studied the irregular spatial distribution of stress data by determining the
46 maximum horizontal compressive stress (S_{Hmax}) orientations on an irregular triangular element grid
47 with the locations of grid nodes dependent upon the spatial distribution of the data. For clusters, the
48 nodes were collocated at the centre, whereas for sparse data, the nodes were placed at data
49 locations. At each node, Rebai *et al.* (1992) computed an average S_{Hmax} orientation by smoothing
50 the S_{Hmax} data within a specified distance from the node. This distance was assumed to be
51 proportional to the size of the smallest triangle connected to the node. The S_{Hmax} average was
52 weighted by the quality factor assigned to the stress datum and by the scale factor associated with
53 the type of S_{Hmax} indicator (decimetres to metres for borehole breakouts, kilometres for focal
54 mechanisms). Within the triangular elements, the nodal values were linearly interpolated.

55 Lee & Angelier (1994) reconstructed the 2D regional principal palaeostress axis orientations using
56 two different approaches. In the first approach, the stress orientations were represented by a
57 polynomial function of flat-Earth coordinates. The coefficients of the polynomial were determined
58 using weighted least-squares to fit the measured S_{Hmax} orientations, with the weight of each S_{Hmax}
59 datum dependent upon its uncertainty. In the second approach, the S_{Hmax} data were also weighted
60 by a power-law function of the distance. The second approach provided better results, indicating
61 that proper distance weighting is essential in stress interpolation.

62 Coblenz & Richardson (1995) quantified long-wavelength trends in the global S_{Hmax} orientations
63 and stress regime patterns using the World Stress Map database release 1992 (hereinafter, WSM92;

64 Zoback, 1992). The authors binned the S_{Hmax} data into 582 $5^\circ \times 5^\circ$ cells. Of the 382 bins containing
 65 at least two S_{Hmax} orientations, Coblenz & Richardson (1995) estimated which bins had non-
 66 random S_{Hmax} orientations with standard deviation $< 25^\circ$ by employing a Rayleigh test at the 90%,
 67 95%, and 97.5% confidence levels. The test assumes that random S_{Hmax} orientations follow a von
 68 Mises normal distribution (Mardia, 1972). The hypothesis that stress orientations are random in
 69 WSM92 was rejected at the 97.5% confidence level for 197 out of 382 bins, suggesting that large
 70 portions of intra-plate regions have coherent S_{Hmax} orientations and are primarily in a state of
 71 compression.

72 Assuming two-point conditional probabilities, Bird & Li (1996) empirically determined the
 73 probability distributions of the angular differences (also called discrepancies) in the S_{Hmax}
 74 orientations based on WSM92 data. They used two approaches: one that accounted for the quality
 75 factor and one in which the clustered data were pre-averaged. Although both concepts are
 76 important, they have not yet been implemented simultaneously. Neglecting the quality factor results
 77 in the underweighting of the most accurate data, whereas when clustered data are used, the uneven
 78 data sampling causes a loss of precision that is not easily included in the error estimates.

79 Müller *et al.* (2003) determined the S_{Hmax} orientations using the nearest-neighbour regression
 80 method. The authors compared two approaches using either a fixed search radius or a fixed number
 81 of nearest neighbours. The smoothing was performed at grid points, with the data weighted by both
 82 quality and distance. In particular, the quality weight was 1 for rank A, 0.75 for rank B, 0.5 for rank
 83 C, and 0 for ranks D and E (see Zoback (1992) for a description of the quality factors). The distance
 84 weight was expressed as a tricube weight function:

$$85 \quad w(x, x_i) = \begin{cases} \left[1 - \left(\frac{\|x - x_i\|}{R} \right)^3 \right]^3 & \text{for } \|x - x_i\| < R, \\ 0 & \text{otherwise} \end{cases}$$

86 where x_i is the position of the i th data point and R is either the fixed search radius or the distance to
 87 the $(N+1)$ -th nearest neighbour. Müller *et al.* (2003) performed checkerboard-like synthetic tests
 88 and verified that using a fixed search radius is preferable for recovering the stress domains. To
 89 avoid interpolating stress orientations in areas of poor data coverage, minimum threshold values for
 90 the sum of the quality and distance weights were suggested. The approach of Müller *et al.* (2003)
 91 proved to be flexible and capable of providing an empirical model of the stress processes. However,
 92 this approach is sensitive to outliers and suffers when the data density is low.

93 Rather than using discrete weights, several groups (e.g., Hardebeck & Michael, 2006; Arnold &
 94 Townend, 2007) have developed methods to propagate observational errors as formal uncertainties
 95 in S_{Hmax} , which comprise mainly focal mechanisms. Hardebeck & Michael (2006) used a damped
 96 inversion method to determine stress on a grid, which minimises the difference between adjacent

97 points. The damped inversion removes the stress rotation artefacts exhibited by undamped
98 inversions and resolves stress rotations better than moving-window approaches. Arnold & Townend
99 (2007) developed a Bayesian formulation for estimating tectonic stress orientations, which
100 combines a geologically based prior stress model and focal mechanism observations, including their
101 precisions. Their technique produces the posterior density function of the principal components of
102 the stress tensor and of the stress-magnitude ratio, which have been used to determine S_{Hmax}
103 following the approach of Lund & Townend (2007).

104 Heidbach *et al.* (2010) calculated the global S_{Hmax} orientations by smoothing stress data that were
105 weighted by both quality and distance. The distance weighting function was $\min(1/D, 1/20 \text{ km})$,
106 where D is the distance between the data location and grid point and min represents the minimum of
107 the two arguments. The quality weights were 1/15, 1/20, and 1/25 for A-, B-, and C-quality data,
108 respectively. The S_{Hmax} orientations were defined if there were at least five WSM08 data within the
109 variable search radius and if the weighted standard deviation was $\leq 25^\circ$. Pierdominici & Heidbach
110 (2012) introduced an additional condition that the data be located in two diagonal quadrants, with at
111 least 10 WSM08 data within the search radius. No constraints were placed on the standard deviation
112 in their study.

113 In this paper, we aim to address three major issues that have not previously been resolved
114 simultaneously: the scatter in S_{Hmax} orientations, the uneven sampling of stress data, and the
115 correlation of stress orientations with distance. We modified and extended the clustered data
116 analysis technique used by Bird & Li (1996) by adding a constraint governing the probability
117 algorithm for estimating S_{Hmax} orientations. Given the intrinsically clustered nature of earthquakes
118 and the uneven sampling of boreholes, we pre-averaged clustered data and updated the empirical
119 probability distributions using the WSM08. In agreement with the data, we modelled the
120 probabilities as exponentially decreasing with discrepancy, thereby compensating for the low data
121 density. An S_{Hmax} orientation was assigned to each interpolation point if three conditions were
122 satisfied: a) a minimum of three clusters; b) 90% confidence limits less than $\pm 60^\circ$; and c) a
123 maximum search radius of 327 km.

124 The resulting stress maps have 90% confidence limits of less than $\pm 45^\circ$ for 38% of the study area
125 and uncertainties of less than $\pm 30^\circ$ for the 19%. The results presented in this work cover 57% of the
126 entire study area. Our S_{Hmax} orientations overlap with 28% of those in Heidbach *et al.* (2010); the
127 spatial distributions of the two stress fields are therefore complementary. Where the two maps
128 overlap, the results differ by more than 30° for only 2% of our study area. The stress orientations
129 determined using our method are generally consistent with the S_{Hmax} orientations obtained using
130 geological indicators, earthquakes, and numerical models. Our method is able to detect several

131 small-scale (20-50 km) stress perturbations, thereby better constraining and complementing the
132 stress field determinations available elsewhere in the literature.

133 **2. Data and method**

134 Based on previous studies of stress interpolation, the angular differences (discrepancies) between
135 two S_{Hmax} orientations tend to increase with distance or, alternatively, the closest S_{Hmax} datum to
136 an interpolation point is weighted most heavily. Of the various data weighting algorithms suggested
137 in the literature, we adopted the approach of Bird & Li (1996). In this approach, S_{Hmax} is
138 determined at each integration point using two-point conditional probabilities and assuming all of
139 the WSM08 data to be independent. The conditional probability of an event is the probability of that
140 event occurring given that another event has already occurred. In terms of conditional probabilities,
141 the S_{Hmax} orientation represents the azimuth with the highest probability of occurring given the
142 neighbouring WSM08 data. We defined the empirical two-point conditional probability based on
143 the stress data (see subsection 2.1) and the analytical probability using least-squares fitting
144 (subsection 2.2). We then used the declustering method of Bird & Li (1996) and the analytical
145 probabilities to determine the S_{Hmax} orientations. We applied our method to study short-wavelength
146 patterns over a wide area (total extent of 2.7×10^9 km²) from 30W to 40E longitude and 25N to 76N
147 latitude, roughly corresponding to the geographic location of the greater European region.

148 **2.1. Data**

149 We used the S_{Hmax} data records from the WSM08 dataset (<http://www.world-stress-map.org>). Each
150 stress datum was assigned a quality factor between A and E, with A being the highest quality and E
151 the lowest. The standard deviations of A-, B-, C-, and D-quality data records are within $\pm 15^\circ$, $\pm 20^\circ$,
152 $\pm 25^\circ$, and $\pm 40^\circ$, respectively. E-quality data records have standard deviations greater than 40° . The
153 WSM08 dataset consists of 21,750 data records, of which 16,961 have qualities of A-C; these data
154 form approximately 1.4×10^8 unique pairs, representing a significant improvement over the 1.8×10^7
155 data pairs obtained by Bird & Li (1996) using WSM92. The data density increased from 0.012×10^{-3}
156 data/km² in WSM92 to 0.041×10^{-3} data/km² (0.033×10^{-3} data/km² for A-, B-, and C-quality data) in
157 WSM08. North America, Europe, and China are the most densely sampled zones. Within Europe,
158 Italy and Switzerland have the highest data densities, with an average of 3×10^{-3} data/km² (Fig. 1).
159 However, many locations in Anatolia, the Hellenic Arc, the Pyrenees, Northwestern Europe, and the
160 North Sea have densities exceeding 0.7×10^{-3} data/km².

161

162 **2.2. Empirical probabilities**

163 The angle between any pair (r,s) of S_{Hmax} data points on Earth is the discrepancy β defined by Bird
 164 & Li (1996) as follows:

$$\beta = \min_{m=-3}^{+3} |\alpha_r - \gamma_r + \gamma_s - \alpha_s + m \times 180^\circ|, \quad (1)$$

165 where m is an integer between -3 and 3, γ_r and γ_s are the local azimuths of the great circle
 166 connecting the two data points, and α_r and α_s are the S_{Hmax} azimuths relative to North (Fig. 2).
 167 Our first step was to determine the number of WSM08 pairs as a function of the discrepancy β and
 168 range (or angular distance) θ . Following Bird & Li (1996), we used 30 3° -wide sectors for β and
 169 150 concentric annuli for θ . The annulus n is defined such that $\theta_{n-1} \leq \theta < \theta_n$, with

$$\theta_n \cong \cos^{-1}\{1 - 2(n/150)^{1/(1-\varepsilon)}\}. \quad (2)$$

170 The WSM08 dataset is 3.5 times larger than the WSM92 dataset available to Bird and Li (1996)
 171 (6,000 data points), for which they used $\varepsilon = 0.4$. The larger dataset can be exploited by increasing
 172 either the number of discrepancy bins (e.g., using a finer binning for β) or the number of annuli at
 173 short range (i.e., $\varepsilon > 0.4$). Based on a series of tests, we chose the latter option. We selected
 174 $\varepsilon = 0.6$, which enabled the investigation of short ranges while avoiding high uncertainties for some
 175 S_{Hmax} orientations.

176 We used 150 annuli for the range θ and 30 3° -wide sectors for the discrepancy angle β , resulting in
 177 a total of 4,500 bins. Each bin was identified by two indices, $n(\theta)$ and $i(\beta)$, such that the range
 178 angle θ lies within the bounds $\theta_{n-1} \leq \theta \leq \theta_n$ and the index i is given by $i = \frac{\beta}{3^\circ} + 1$, rounded down
 179 to the nearest integer. The number of data pairs within each bin can be represented by a matrix
 180 $C_{i(\beta),n(\theta)}$ of integers.

181 The probability that a discrepancy β at range θ falls into angular bin i is given by

$$D_{i(\beta),n(\theta)} = C_{i(\beta),n(\theta)} / \sum_{j=1}^{30} C_{j,n(\theta)}. \quad (3)$$

182 Given a datum r with azimuth α_r , we can define the two-point conditional probability $P(k(\alpha_r)|\alpha_s)$
 183 for any azimuth α_s at any range angle θ determined by the appropriate angular bin k ($0 \leq \alpha <$
 184 180 ; $k = 1 \dots 60$):

$$P(k(\alpha_r)|\alpha_s) = \frac{1}{2} D_{i(\beta),n(\theta)}, \quad (4)$$

185 where the factor of $1/2$ avoids the double counting of pairs (there are two possible azimuths for the
 186 same β value for each data pair).

187 Fig. 3a supports our choice of ε for the investigation of short ranges. The probabilities decrease
188 sharply for the first three annuli with $\varepsilon = 0.6$, whereas smaller values of ε fail to capture this
189 probability variation, resulting in a uniform probability distribution. For example, with $\varepsilon = 0.4$,
190 Annulus 1 contains the upper limit set to 1.76° , whereas for $\varepsilon = 0.6$, this upper limit falls in
191 Annulus 6. Due to the increased number of data points in WSM08 compared to the WSM92 dataset,
192 the finite sample effect for each bin was on the order of 2%, similar to the value found by Bird & Li
193 (1996) in their analysis of the WSM92 dataset.

194 We divided the WSM08 dataset into five classes according to the quality factor that was assigned to
195 each data record. The probability distributions were of comparable width for the A-, B-, and C-
196 classes (Fig. 3b) and substantially narrower than for the D- and E-classes for all ranges. The D- and
197 E-quality data exhibited a large scatter in the probability distribution at short ranges, indicating a
198 lack of correlation. We therefore discarded the 3,946 data records with D- or E-quality factors and
199 considered only WSM08 data with A-, B-, or C-quality factors in our analysis.

200 The high probability of finding small discrepancies at short ranges ($\theta \cong 0^\circ$) does not exclude local
201 scatter, as illustrated in Fig. 3c by the non-zero probabilities of high discrepancies in β . These
202 discrepancies are related to the repeatability of the measurements, depth variability, or
203 heterogeneities in the material properties. Following the approach of Barba *et al.* (2010), we
204 estimated the repeatability error at zero distance $\sigma_0 = \sigma(0)$ by interpolating the standard deviation
205 σ versus the range θ using the following equation:

$$\sigma(\theta) = a + b\theta + c\theta^{0.5}. \quad (5)$$

206 The error was $\sigma_0 = 4^\circ \pm 1^\circ$ for A-quality data (Fig. 4a), $\sigma_0 = 10.9^\circ \pm 0.6^\circ$ for B-quality data, and
207 $\sigma_0 = 13.2^\circ \pm 0.7^\circ$ for C-quality data. Combining the A-, B-, and C-quality data yields $\sigma_0 =$
208 $11.8^\circ \pm 0.5^\circ$ (Fig. 4a).

209 The inclusion of D-quality data would worsen the results, given the large scatter in the probability
210 distributions for all discrepancies (Figure 3b). The results are affected by both the larger
211 uncertainties for small discrepancies in D-quality data and the larger biases for larger discrepancies.
212 An alternative is weighting the A-, B-, and C-quality data proportionally to the error σ_0 (Eq. 5),
213 which essentially gives a higher weight to A-quality data. Although formally correct (Arnold &
214 Townend, 2007), given the relatively low number of A-quality data compared to B- and C-quality
215 data, the different weighting would have a negligible effect on the results while making the
216 procedure more complex. However, this point may be reconsidered in the future when a greater
217 number of A-quality data are available.

218 For $\theta < 3^\circ$, the correlation between data pairs is high and the probabilities of large discrepancies
 219 are small, whereas for $\theta > 6^\circ$, the discrepancy probabilities do not show any significant trends (Fig.
 220 3c). Following Bird and Li (1996), we define a scalar measure of the correlation as follows:

$$I_{cor}(WSM08) = \frac{\sum_{i=1}^{10} D_i}{\sum_{j=21}^{30} D_j}, \quad (6)$$

221 i.e., the ratio of D (Eq. 3) for “small” discrepancies ($0 < \beta \leq 30^\circ$) to D with “large” discrepancies
 222 ($60^\circ < \beta \leq 90^\circ$). Bird & Li (1996) reported a quasi-exponential falloff of $I_{cor}(WSM92)$ out to a
 223 range of 22° , whereas $I_{cor}(WSM08)$ exhibited a quasi-exponential decrease for all ranges (Fig. 4b).
 224 To improve the detail at short range, we used the empirical probabilities for $\theta < 3^\circ$ (Annuli 1 to 8)
 225 instead of $\theta < 22^\circ$, as adopted by Bird & Li (1996). The choice of the upper value of θ depends on
 226 the complexity of the study region. For example, compared to the Pacific Basin in the Tertiary
 227 period, the Mediterranean/Tethyan Basin is of similar complexity and has a similar number of
 228 subduction zones, yet much shorter along-strike distances. Therefore, the stress fields associated
 229 with subduction, mountain ranges and perhaps hanging slabs in the Mediterranean have relatively
 230 short length scales compared to those of the Pacific stress fields. Thus, using lower cutoff values
 231 than those used in global modelling allows the resolution of these smaller features.

232 The exponential trend of the probabilities confirms the hypothesis of Bird & Li (1996) that the
 233 “correlation horizon” at a range of $\theta = 22^\circ$ is due to uneven sampling of Earth in WSM92. The
 234 decreased ringing in $I_{cor}(WSM08)$ with respect to $I_{cor}(WSM92)$ is due to the eightfold increase in
 235 the number of unique pairs in WSM08.

236 2.3. Analytic probabilities and interpolation method

237 The probability distributions (Fig. 3a) follow an exponential law for all annuli. If few data are
 238 considered, the probabilities do not behave exponentially, which explains why this behaviour has
 239 not been observed previously. Based on highly successful regressions ($R^2 \geq 0.99$), we
 240 parameterised the empirical probabilities of Eq. (4) as follows:

$$P^*(k(\alpha_r)|\alpha_s) \equiv P^*(k_r|s) = P_0^* + P_1^* \exp(-\theta/\theta_0), \quad (7)$$

241 where P^* denotes the analytic probability and θ_0 , P_0^* and P_1^* are constants determined from a
 242 nonlinear least-squares fit to the empirical probabilities P (Eq. 4) within the annulus. In this
 243 equation, we simplified the notation by using the integer indices r and s in place of the azimuths α_r
 244 and α_s . The values of θ_0 , P_0^* , and P_1^* for Annuli 1-8 are listed in Table 1.

245 Let x be a point at which we wish to estimate the stress orientation, and let us select the WSM08
 246 data points within a range of θ_n (Eq. 2). The probability of the azimuth at x falling into bin k given

247 one datum r is $P^*(k_x|r)$ (Eq. 7). Although the presence of independent data is easily formalised
 248 through probability multiplication, clustered data require further analysis. A pair of stress data
 249 points r and s form a cluster if

$$P^*(s|r) > \max_{i=r,s} P^*(i|x), \quad (8)$$

250 i.e., the two-point conditional probability of r and s is larger than the highest possible conditional
 251 probability with respect to the interpolation point x . The opposite guarantees that the data are
 252 independent. Note that this definition of independence depends on distance; a pair of data points can
 253 both be independent for interpolation points located between the data and form a cluster when they
 254 are far away from the interpolation point. After the clusters have been defined, a two-pass
 255 procedure is applied. In the first pass, the clustered data are pre-averaged, resulting in a set of fully
 256 independent S_{Hmax} orientations. This set consists of pre-averaged data and data points determined
 257 to be independent according to Eq. 8. In the second pass, the S_{Hmax} orientation is interpolated to x .
 258 The pre-averaged values of the clustered data are obtained in the first pass using the two-point
 259 conditional probability distributions, assuming no azimuthal dependence between the data. For
 260 simplicity, we omit the cluster index. The S_{Hmax} orientation of the cluster is assigned to its
 261 geographical centre, hereinafter denoted by z , and is determined using the maximum likelihood
 262 method. For each trial azimuth, defined by an integer value of k_z ($k_z = 1, \dots, 60$ with 3° bins), we
 263 compute the probability as

$$P_N^*(k_z) = \frac{\prod_{i=1}^N P^*(k_z|i)}{\sum_{j=1}^{60} \prod_{i=1}^N P^*(j_z|i)}, \quad (9)$$

264 where N is the number of stress data forming the cluster. The maximum-likelihood estimate of the
 265 S_{Hmax} orientation at the cluster centre z is $\alpha_z = \left(k_z - \frac{1}{2}\right) 3^\circ$, where k_z is the integer that maximises
 266 $P_N^*(k_z)$. This operation is performed for each cluster.

267 In the second pass, the S_{Hmax} orientations are interpolated to x . To identify the now fully
 268 independent orientations, we call them “clusters” and label them as c regardless of whether they
 269 arose from pre-averaging. For each trial azimuth, defined by an integer value of k_x ($k_x =$
 270 $1, \dots, 60$ with 3° bins), we calculate the probability as

$$P_{N_c}^*(k_x) = \frac{\prod_{c=1}^{N_c} P^*(k_x|c)}{\sum_{j=1}^{60} \prod_{c=1}^{N_c} P^*(j_x|c)}, \quad (10)$$

271 where N_c is the number of clusters within range θ_n (Eq. 2). However, the index n is omitted in Eq.
 272 (10) for simplicity. The maximum likelihood estimate of the S_{Hmax} orientation at interpolation
 273 point x is

$$\alpha_x = \left(k_x - \frac{1}{2}\right) 3^\circ, \quad (11)$$

274 where k_x is the integer that maximises $P_{N_C}^*(k_x)$ in Eq. (10). The 90% confidence interval $\Delta\alpha$ is
 275 determined using the following equation:

$$\int_{\alpha_x - \Delta\alpha}^{\alpha_x + \Delta\alpha} p_N^*(\alpha'_x \bmod 180^\circ) d\alpha'_x = \mathbf{0.90}, \quad (12)$$

276 where $p_N^*(\alpha'_x \bmod 180^\circ)$ is the functional form corresponding to the discrete $P_N^*(k_x)$ and “mod”
 277 indicates the remainder of the integer division, used to account for the periodicity of α'_x . This
 278 procedure allows the uncertainties due to data scattering (shown in Fig. 3 and included in Eq. 10) to
 279 propagate into the posterior uncertainties $\Delta\alpha$ (Eq. 12).

280 We adopt two conditions to ensure a well-defined S_{Hmax} orientation: a) $N_C \geq 3$ within a specified
 281 range θ_n and b) $\Delta\alpha \leq 60^\circ$ for the N_C clusters. The upper limit of the 90% confidence interval
 282 ($\Delta\alpha \leq 60^\circ$, corresponding to $\sim 2\sigma$ in a normal distribution) accounts for the error associated with C
 283 qualities ($1\sigma = 30^\circ$), the most common quality factor in WSM08. The search began with Annulus 1
 284 ($\theta_1 < 0.22^\circ$) and increased stepwise up to Annulus 8 ($\theta_8 < 3^\circ$). We selected the first annulus n
 285 (and thus the smallest range) that satisfied both conditions. However, if the conditions were not met
 286 by Annulus 8 (the upper limit), then we did not assign an S_{Hmax} orientation to the interpolation
 287 point x .

288 3. Stress interpolation in Europe

289 We interpolated the WSM08 S_{Hmax} orientations using data-driven analytical weight distributions
 290 after pre-averaging the clustered data in the area from 30W to 40E longitude and 25N to 76N
 291 latitude. We weighted by distance and discrepancy and adopted thresholds for the number of
 292 clusters N_C and the 90% confidence limits $\pm\Delta\alpha$. The threshold values determine the interpolated
 293 stress orientations α_x (Eq. 11) and limit the area in which the stress orientations can be recovered
 294 (Table 2). We illustrate our results with examples in Europe.

295 Stress orientations were determined for 57% of the study area. The zones with the smallest $\Delta\alpha$
 296 values are the Apennines, the Dinarides, the Alps, Western Turkey, and the Aegean Sea. We found
 297 $\Delta\alpha < 15^\circ$ for 3% of the study area, $\Delta\alpha < 30^\circ$ for 19% of the area, and $\Delta\alpha < 45^\circ$ for 38% of the
 298 area (Fig. 5). This fraction of recovered stress orientations is satisfactory and comparable to those
 299 obtained in other studies (e.g., Heidbach *et al.*, 2010). The search radius was smaller than the upper
 300 limit of Annulus 3 ($\theta_3 < 0.86^\circ$, 96 km) for a large part of continental Europe and the Norwegian
 301 Sea (Fig. 6). Conversely, S_{Hmax} was not recovered in the Atlantic Ocean, Northern Africa, or
 302 Russia due to the scarcity of data in those locations. The S_{Hmax} orientations were interpolated using

303 $3 \leq N_C \leq 4$ (average number of data $N = 4.87$) for 28% of the study area, $5 \leq N_C \leq 8$ (average
304 $N = 8.89$) for 22% of the study area, $9 \leq N_C \leq 12$ (average $N = 16.05$) for 5% of the study area,
305 and $N_C > 12$ (average $N = 31.42$) for 2% of the study area (Fig. 7). Due to the high density of
306 uniform data, the stress orientations were calculated using large N_C values and short ranges in the
307 Pyrenees, Central Italy, and Romania. Large N_C values and long ranges were used in the Western
308 North Sea, Tyrrhenian Sea, Western Alps, and Ionian Sea due to locally scattered WSM08 S_{Hmax}
309 orientations. Our results are provided in both numeric and map form at a resolution of $0.1^\circ \times 0.1^\circ$ as
310 an electronic supplement to this article (Table S1 and Fig. S1).

311 The S_{Hmax} map (Fig. A1, available in the auxiliary material) displays the same main features
312 already described in the literature in relation to the European stress field (Müller *et al.*, 1992; Golke
313 & Coblenz, 1996). The stress pattern in Northwestern Europe (Figure 8a) is dominated by long-
314 wavelength features and follows the NW-SE pattern expected for a stress field produced by plate-
315 scale forces (Müller *et al.*, 1992; Golke & Coblenz, 1996; Müller *et al.*, 1997). Therefore, plate
316 boundary forces are likely the main factor controlling the stress orientation in Northwestern Europe
317 (Goes *et al.*, 2000). Scandinavia exhibits scattered stress orientations (Fig. 8b), likely due to a
318 combination of plate boundary forces, glacial isostatic adjustment, and local heterogeneities
319 (Bungum *et al.*, 2010). However, no clear radial pattern can be observed, indicating that rebound
320 stresses are either spatially limited or of relatively low magnitude. The S_{Hmax} orientations in Iberia
321 display substantial variability (Fig. 8c). The stress field has a complex pattern; first-order
322 compressive stresses, such as the Atlantic ridge push and Eurasia-Africa convergence, interact with
323 and are perturbed by local stress fields, such as the rifting of the Valencia Trough and topographic
324 loading (De Vicente *et al.*, 1996; Andeweg *et al.*, 1999; Liesa and Simon, 2009).

325 In Italy (Fig. 8d), the NNW- SSE S_{Hmax} orientations mark the extension in the Apennines (Boncio
326 & Lavecchia, 2000; Carminati *et al.*, 2001; Barba *et al.*, 2010), whereas the S_{Hmax} in the Adria
327 microplate corresponds with the active Adria-Eurasia compression (Kastelic & Carafa, 2012; Barba
328 *et al.*, 2013). The S_{Hmax} orientations in the Eastern Mediterranean rotate from NNW-SSE to NW-
329 SE in the Marmara Sea and to E-W in the Northern Aegean Sea (Fig. 8e). This pattern follows the
330 lateral extrusion of Anatolia and the N-S directed backarc spreading caused by the rollback of the
331 Hellenic subduction zone (Taymaz *et al.*, 1991; Herget & Heidbach, 2011; Petricca *et al.*, 2013).

332 Short-wavelength features are scattered throughout the map of the interpolated S_{Hmax} orientations.
333 Short-wavelength patterns typically arise from crustal heterogeneities, topographic effects, and
334 sharp variations in the kinematics; we describe a few examples associated with these different
335 causes.

336 Crustal-scale local stress sources, such as lateral density contrasts, rifting processes, and diapirism,
337 interfere with regional stresses and can result in S_{Hmax} reorientation depending on the stress
338 magnitudes and angular differences. Such a phenomenon occurs in Vrancea and the Southeastern
339 Carpathians (Fig. 8e) and has been extensively studied by Müller *et al.* (2010), who inferred that the
340 high variability in the stress orientations indicates that the stresses generated by local sources are of
341 similar magnitude to the regional stress. In the Central Adriatic (Fig. 8d), the S_{Hmax} orientations
342 also arise from an interplay between the regional stresses, caused by the Adria-Eurasia collision
343 (Kastelic & Carafa, 2012), wrench tectonics, and salt diapirism (Grandic & Markulin, 2000; Geletti
344 *et al.*, 2008; Korbar, 2009). A more extreme case is the easternmost part of the Po Plain, where
345 S_{Hmax} takes on an E-W orientation (Fig. 8d), which is in contrast to the regional N-S compression,
346 as determined from earthquakes and geodesy (Carminati & Vadacca, 2010; Caporali *et al.*, 2011).
347 The smoothed orientations are instead determined using the WSM08 data closest to the
348 interpolation point, which reflect the sedimentary infilling of the Po Plain. The rotation of these
349 stress data with respect to the regional trend can be ascribed to differential sediment compaction,
350 favoured by the presence of anticlines and large variations in the sediment thickness. The shallow
351 E-W orientations of S_{Hmax} due to the tensional, strike-slip stress field at a depth of approximately
352 1,200 m are perpendicular to the deeper N-S orientations of S_{Hmax} due to the compressive stress
353 field near the main detachment (Carminati *et al.*, 2010). In this specific case, the three closest data
354 points are iso-oriented and represent the sediments, and our algorithm captures the local and
355 shallow tensional stress field induced by the differential sediment compaction. However, applying
356 our algorithm to a different dataset (e.g., the regional moment tensors calculated using the method
357 of Herrmann *et al.*, 2011) yields N-S S_{Hmax} orientations.

358 In the case of topographically induced stress, S_{Hmax} can align with a mountain chain at the highest
359 elevations and rotate perpendicularly to the chain at the foot of the belt (Assameur & Mareschal,
360 1995; Carminati *et al.*, 2004). This short-wavelength pattern is evident in the interpolated stress
361 map for the Pyrenean domain (Fig. 8c) and the Eastern and Western Alps (Fig. 8d). However, the
362 pattern is less evident in the Southern Scandes (Fig. 8b) due to the relative scarcity of data there.

363 Parallel trends in S_{Hmax} and the isohypses are typically interpreted as the spreading of
364 topographical crests at high elevation in mountains and as gravitational collapse at mountain bases
365 (Decker *et al.*, 1993; Herraiz *et al.*, 2000; Sue *et al.*, 2007; Pascal & Cloething, 2009; Bungum *et al.*
366 *et al.*, 2010). In Southern Cyprus, S_{Hmax} follows a N-S orientation (Fig. 8d), supporting the presence
367 of a relatively small collision zone between Cyprus and the Eratosthenes Seamount, which is a
368 faulted block that has been uplifted by 1,500 m. Cyprus and its surroundings are characterised by

369 NE-SW S_{Hmax} orientations (Fig. 8e) that correspond to the collision between Sinai and Anatolia
370 (Wdowinski *et al.*, 2005).

371 Another observed short-wavelength feature is the sudden change in tectonic style. Moving from
372 extension to compression implies a rotation of the stress tensor and a consequent rapid reorientation
373 of S_{Hmax} . This phenomenon is observed in the Northern Apennines and Hellenic Arc. As we move
374 from the extension of the Apennines to the compression of the eastern coastline, the S_{Hmax}
375 orientations rotate by 90° within a few tens of kilometres and turn perpendicular to the External
376 Apennines thrust faults (Scrocca, 2006; Basili & Barba, 2007; Carafa & Barba, 2011). Along the
377 Hellenic Arc, similar transitions from extension to compression and subsequent abrupt reorientation
378 of stresses occur in the Central Epirus (Northern Hellenides) and the shallow marine zone between
379 the Ionian Islands and mainland Greece (in the Southern Hellenides) (Papanikolaou and Royden,
380 2007; Jolivet *et al.*, 2009; Papanikolaou *et al.*, 2011).

381 4. Validation

382 Independent validation is required to understand the degree to which the modelled stress field
383 represents genuine features in Earth's crust. We performed two comparisons: one with synthetic
384 data and one with similar studies available in the literature.

385 We ran a synthetic checkerboard test to evaluate the significance of short-wavelength stress
386 patterns. The a priori stress flow lines were arranged in 110-km squares, with the stress patterns
387 perpendicular to each other in adjacent squares. S_{Hmax} orientations were assigned to points at 0.2°
388 intervals (Fig. 9). This synthetic data distribution corresponds to a data density of 2.05×10^{-3}
389 data/km². To produce three additional synthetic datasets, S_{Hmax} orientations were randomly chosen
390 with assumed data densities of 1.21×10^{-3} data/km², 0.81×10^{-3} data/km², and 0.37×10^{-3} data/km².
391 These values are typical of the WSM08 data densities for much of Europe (Fig. 1), whereas the
392 dimension of the squares and perpendicular S_{Hmax} orientations represent possible stress targets. We
393 applied the interpolation procedure described in the previous section to each of the four datasets
394 (one equispaced and three generated randomly) using the same threshold values ($N_C \geq 3$, $\Delta\alpha <$
395 60° , and $\theta < 2.94^\circ$) and a spacing of 0.2° . To quantify the agreement between the a priori and
396 recovered checkerboards, we defined a correlation index as the fraction of the recovered S_{Hmax}
397 orientations differing by less than 5° from the initial orientation. As expected, the checkerboard was
398 fully recovered (correlation index equal to one) for the equispaced synthetic dataset (2.05×10^{-3}
399 data/km²). For the three random datasets, the correlation indices were 0.93 (for 1.21×10^{-3} data/km²),
400 0.82 (for 0.81×10^{-3} data/km²), and 0.70 (for 0.37×10^{-3} data/km²). The stress patterns were

401 satisfactorily recovered, both visually (Fig. 10) and according to the correlation index. This test
402 demonstrates that our interpolation procedure can resolve the short-wavelength stress pattern in the
403 parts of Europe with data densities similar to those considered here. In areas with very low data
404 densities (such as Scandinavia and Northern Africa), the S_{Hmax} orientations represent the regional
405 trend and can only be determined using longer ranges (Annulus $n = 6 - 8$, $2.05^\circ < \theta < 2.94^\circ$).

406 To compare our results with those obtained in similar studies in the literature, we selected the works
407 of Heidbach *et al.* (2010) and Pierdominici & Heidbach (2012). We also applied the Bird & Li
408 (1996) procedure (with pre-averaging and using the Bird & Li (1996) empirical two-point
409 conditional probabilities calculated in WSM92) to WSM08 data. The S_{Hmax} orientations obtained
410 with our method are significantly closer to those of Heidbach *et al.* (2010) than to those obtained by
411 applying the method of Bird & Li (1996) to WSM08 (Fig. 11). This difference arises partly from
412 the algorithm – the range cutoff was $\theta = 22^\circ$ for the global modelling of Bird & Li (1996) versus
413 $\theta = 2.94^\circ$ in our case – and partly from the new set of conditional probabilities based on WSM08.

414 Given the number of data points, the number of unique pairs in WSM08 is eight times larger than
415 that in WSM92. This increase in the number of data pairs permits the investigation of shorter ranges
416 and the resolution of greater spatial detail than in Bird & Li (1996). The S_{Hmax} orientations
417 presented in this work are in good agreement with those from Heidbach *et al.* (2010) for Europe
418 (Fig. 12). The criteria used in the smoothing procedure of Heidbach *et al.* (2010) requires more than
419 five data points to be within a specified search radius and for the standard deviation of these data to
420 be less than 25° . The search radius is initialised at 1,000 km and decreased in 50-km steps until the
421 standard deviation of the $N \geq 5$ S_{Hmax} orientations is less than 25° . Our approach and that
422 presented in Heidbach *et al.* (2010) are both capable of determining the S_{Hmax} orientations for a
423 large part of Europe; our results and theirs overlap by 28%. Within this overlapping region,
424 approximately 10% of the S_{Hmax} orientations (corresponding to 3% of the study area) differ by
425 more than 25° , i.e., more than the standard deviation of Heidbach *et al.* (2010). Approximately 3%
426 of the orientations (corresponding to 1% of the study area) differ by more than 45° , indicating a
427 lack of correlation (Fig. 12), which requires further investigation. Fig. 12 shows the area where the
428 discrepancy in the S_{Hmax} orientations exceeds 25° . Most of these zones (Northeastern Iberia, the
429 Western Alps, the Po Plain, the Northern Apennines, the Southern Tyrrhenian, and Cyprus) are
430 characterised by high data densities (greater than 1.2×10^{-3} data/km²) and have short-wavelength
431 stress sources and sharply rotating S_{Hmax} orientations, as described in the previous section. Based
432 on these findings, we conclude that our method highlights the presence of short-wavelength stress
433 patterns and sudden S_{Hmax} reorientations ($\theta < 100$ km) that cannot be identified with the minimum
434 search radius of 100 km used in Heidbach *et al.* (2010).

435 To further validate our method and underline the importance of both pre-averaging the data and
436 investigating short-range stress sources, we compared our results with those of Pierdominici &
437 Heidbach (2012), in which the search radius was decreased from 1,000 km to 100 km in 50-km
438 steps. Their smoothing algorithm requires that more than 10 data points are located in at least two
439 diagonal quadrants of the search circle around the grid point. The two methods display a good
440 overall agreement, with the largest discrepancies in the S_{Hmax} orientations occurring in the Western
441 Alps, Po Plain, Northern Apennines, Central Adriatic, and Southern Tyrrhenian. These zones,
442 characterised by high WSM08 data densities, are the same as those with large discrepancies
443 between our work and that of Heidbach *et al.* (2010) (Fig. 13). We argue that our method performs
444 better in terms of reproducing local patterns and sharp S_{Hmax} rotations, both of which are typical of
445 these high-discrepancy zones.

446 5. Discussion and conclusions

447 The method proposed here is appropriate for detecting short-wavelength patterns in S_{Hmax}
448 orientations for large parts of Europe. Our use of analytical probabilities strengthens the
449 contribution of the data close to the interpolation point, where fewer data are available. The data
450 with superior quality factors (A, B, and C) were included, while those with inferior quality factors
451 (D and E) were omitted, which overcomes the difficulties of introducing a weighting scheme into
452 the declustering algorithm.

453 Müller *et al.* (2003) suggested a search radius between 60 and 200 km for local stress field
454 investigations. We have shown that for data densities greater than 0.8×10^{-3} data/km², reliable results
455 can be obtained by pre-averaging the clustered data and choosing a search radius below 60 km
456 (upper limit of Annulus $n = 2$; $\theta_2 < 0.52^\circ$, 58 km). In fact, a few independent data located at short
457 distances tend to highlight local stress sources, whereas a large number of stress data (with a low
458 standard deviation) at greater distances can dominate over the local stress. Increasing the distance
459 increases the likelihood of mixing data affected by different sources of stress. In a few cases, the
460 large number of distant and clustered data could counterbalance the distance weighting. Thus, the
461 orientations of S_{Hmax} could be influenced more by the stress source sampled by the most data than
462 by the local source of stress closest to the interpolation point.

463 We discuss two examples in which changes in the tectonic regime and rotations in stress
464 orientations occur over short distances and the S_{Hmax} orientations presented here differ from those
465 reported in other studies (Heidbach *et al.*, 2010; Pierdominici & Heidbach, 2012). The first example
466 (Fig. 14) is the transition from the NW-SE-oriented S_{Hmax} field caused by the extension in the

467 Central Apennines (Bennett *et al.*, 2012) to the NE-SW offshore compression of the Apenninic
468 outer thrust (Boncio & Bracone, 2009; Maesano *et al.*, 2013). When a large search radius is
469 considered (e.g., 100 km), the clustered stress data in the Central Apennines (80% of the data)
470 dominate the stress interpolation along the Adriatic coastline, causing S_{Hmax} to be parallel to the
471 Marche coastline. In our interpolation scheme, 17 WSM08 data are declustered into 12 independent
472 data, or clusters. The search radius for the interpolation point P_m (Fig. 14) is 58 km (corresponding
473 to Annuli 1 and 2). Similar considerations also apply to the second example, concerning the
474 Southern Tyrrhenian (Fig. 15). Moving from west to east, the S_{Hmax} field rotates from NW-SE,
475 related to the Southern Tyrrhenian compression (Giunta *et al.*, 2004), to NE-SW in the Peloritani
476 area, which is undergoing extension. Between these two zones lies the Madonie-Nebrodi area,
477 which is characterised by a NW-SE S_{Hmax} (SgROI *et al.*, 2011; Palano *et al.*, 2012) that can only be
478 detected with a search radius of less than 60 km. This area is also characterised by a high data
479 density (greater than 0.8×10^{-3} data/km²), and the search radius for the interpolation point P_s (Fig.
480 15) is 24 km for the three WSM08 data (which are independent data). In this case, a search radius of
481 100 km (as in Pierdominici & Heidbach, 2012) captures data primarily related to the Peloritani
482 extension, which is spread over the eastern part of the search area, and results in a different
483 orientation than obtained when using a smaller search radius.

484 These examples demonstrate that the quantity of data available in WSM08 can be exploited with
485 analytical probabilities by applying a declustering algorithm before interpolating and smoothing the
486 data over a distance less than 60 km from the interpolation point. With future increases in the
487 number of stress data, this approach will help to identify short-wavelength stress patterns and locate
488 shifts in the S_{Hmax} orientations with higher spatial accuracy.

489 The distance weighting is a key component in the interpolation or smoothing of the data. The
490 various algorithms differ in the weighting function used (see Introduction and Fig. 16 for a
491 comparison). In this work, we used exponentials, which provide high-quality fits to the empirical
492 discrepancy-range relations for the WSM08 data. For the data located near an interpolation point,
493 Heidbach *et al.* (2010) adopted an inverse distance weighting that produces the maximum weight
494 among the methods reviewed in this paper. Bird & Li (1996) employed a more uniform weight at
495 small distances due to the scarcity of data in WSM92. The development of WSM08 (Zang *et al.*,
496 2012), other available data, and new empirical probabilities will permit further assessment of the
497 quality of the various distance-weighting schemes.

498 The approach presented here can be used to estimate fault activity through kinematic or dynamic
499 models and to predict the stress orientations of future earthquakes, e.g., for seismic hazard
500 assessment, even where instrumental earthquake recordings are unavailable. Interestingly, several

501 mountain ranges exhibit 90° rotations of the S_{Hmax} orientation from their crests (which are often in
502 horizontal tension and experience normal faulting) to their foothills (which often undergo thrust
503 faulting). This pattern, seen in the Pyrenees, Alps, Southern Scandes, Hellenic Arc and Apennines,
504 places a strong constraint on the magnitude of the tectonic stresses with respect to the topographic
505 load, a matter that has often debated (e.g., Bird, 1998; Bird et al., 2008). Varying topographical loads
506 are known to produce stress magnitudes of up to 20-30 MPa, which significantly influences the
507 stress orientations. This influence occurs not only in continental Europe (Golke & Coblenz, 1996),
508 where tectonic stresses are very small, but also in the Alpine-Dinaric belt and Pannonian Basin,
509 where 2-km changes in elevation generate tensional stresses ranging from 6 to 22 MPa in the
510 mountain belt and compressional stresses ranging from 3 to 12 MPa in the basin (Bada et al., 2001).
511 Tectonic stresses are often estimated to be high at plate boundaries due to ridge push or sinking
512 slabs (80-120 MPa; Whittaker et al., 1992), but these stresses are expected to produce observable
513 shear heating (e.g., Molnar and England, 1990). However, the influence of these stresses within
514 continents remains controversial. In fact, stresses on major faults have been found to be smaller
515 than 30 MPa (e.g., Mount and Suppe, 1987; Zoback et al., 1987), and basal tractions acting across
516 continents have been estimated to produce changes of up to 20 MPa in compressive stresses (e.g.,
517 Bird, 1998; Gosh et al., 2008). The 90° rotation of the S_{Hmax} suggests that, at least in the mountain
518 ranges where these rotations are observed, the magnitude of tectonic and topographic stress must be
519 comparable, thus supporting the hypothesis that tectonic stresses can be characterized by relatively
520 low magnitudes of 20-30 MPa. As an indirect consequence, this finding supports the idea that
521 topographic stress can cause an inversion, at a certain depth, of the expected pattern of differential
522 stress, allowing for the initiation of ruptures on low-angle thrust faults at shallow depths to
523 propagate downward (Carminati et al., 2004).

524 The derived S_{Hmax} orientations describe the present-day stress field in Europe. The proposed
525 approach produced acceptable uncertainties for nearly half of Europe. The modified Bird & Li
526 (1996) algorithm facilitates the identification of third-order features, such as topography effects and
527 crustal-scale density heterogeneities. The stress orientations modelled in this work are generally
528 compatible with those of previous studies. However, for areas with data densities greater than
529 0.8×10^{-3} data/km², our algorithm outperforms the existing methods in identifying short-wavelength
530 stress sources.

531 **Acknowledgements**

532 We thank the associate editor Saskia Goes, Peter Bird, and John Townend for their thorough
533 reviews. We thank D. Coblenz for his valuable comments on an early version of the manuscript, S.
534 Pierdominici for kindly supplying the results of her work, O. Heidbach for suggesting that we
535 perform synthetic tests, and again P. Bird for the NeoKinema.f90 code, which helped us to design
536 our work. This work was supported by the DPC-INGV 2008-2010 S1 project, the EU-FP7 project
537 “Seismic Hazard Harmonization in Europe” (SHARE; Grant agreement no. 226967), and project
538 MIUR-FIRB "Abruzzo" (code: RBAP10ZC8K_003). Some figures were created using Generic
539 Mapping Tools (Wessel & Smith, 1998).

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735 Table Captions

736 **Table 1.** Regression parameters and related standard error of Eq. (7) for Annuli 1 to 8. P_0^* , P_1^* , and
737 θ_0 : coefficients of Eq. (7) considering the associated standard errors. RMS: root mean square
738 between empirical and analytical probabilities. R^2 : coefficient of determination, which measures the
739 goodness of the regression fit.

740 **Table 2.** Percentage of the area in which the stress orientations can be recovered for various
741 threshold values of $\Delta\alpha$ and N_C , with the maximum search radius set to the upper limit of Annulus 8
742 ($\theta_8 < 2.94^\circ$; 327 km). The total extent of the study area is $2.71 \times 10^9 \text{ km}^2$.

743 Figure Captions

744 **Figure 1.** WSM08 data density in Europe. Densities were computed for $0.5^\circ \times 0.5^\circ$ cells, from 25N
745 to 76N latitude and 30W to 40E longitude. Cells with no data entries are not coloured.

746 **Figure 2.** Cartographic description of the discrepancy angle β . For each pair of stress data points, r
747 and s , we define the great circle (green line) between r and s for an oblique Mercator projection.
748 The discrepancy is the angle between the two S_{Hmax} orientations as expressed in Eq. (1). The
749 S_{Hmax} orientations of r and s are shown in black, and the S_{Hmax} orientation of s , translated onto r ,
750 is shown in grey.

751 **Figure 3.** Probability distributions as a function of discrepancy and range. (a) Analytical (lines) and
752 empirical (symbols) probabilities for Annuli (An) 1-3 (A-C quality factors); (b) empirical
753 probabilities for Annulus 1 for A-E quality factors; (c) empirical probabilities for various
754 discrepancy intervals. The lines connect the empirical values in (b) and (c).

755 **Figure 4.** (a) Standard deviation versus range (symbols) and analytical representation of Eq. (5)
756 (SV: lines, with confidence limits); (b) the ratio of D (Eq. 3) for “small” discrepancies ($0 < \beta \leq$
757 30°) to “large” discrepancies ($60^\circ < \beta \leq 90^\circ$) versus range.

758 **Figure 5.** Colour-coded map of the 90% confidence limits ($\pm\Delta\alpha$).

759 **Figure 6.** Minimum search radius for which $\Delta\alpha < 60^\circ$ and $N_C \geq 3$. The upper bound is set
760 to $\theta = 2.94^\circ$.

761 **Figure 7.** Number of clusters N_C used to interpolate S_{Hmax} .

762 **Figure 8.** S_{Hmax} orientations in Europe. (a) Northwestern Europe; (b) Scandinavia; (c) Iberia; (d)
763 Italy; (e) Southeastern Europe. Thick bars: WSM08 data; thin black bars: interpolated S_{Hmax} (this
764 work). Inset: colour-coded map of the 90% confidence limits ($\pm\Delta\alpha$).

765 **Figure 9.** Input synthetic checkerboard stress orientations.

766 **Figure 10.** Results of the checkerboard test. From left to right: (1's) S_{Hmax} orientations randomly
767 selected from the checkerboard shown in Figure 9; (2's) interpolated stress field with the same
768 threshold values used to calculate the S_{Hmax} orientations in Europe ($N_C \geq 3, \Delta\alpha < 60^\circ, \theta <$
769 2.94°); (3's) minimum search radius. From top to down: (a's) Data density of 1.21×10^{-3} data/km²;
770 (b's) data density of 0.81×10^{-3} data/km²; (c's) data density of 0.37×10^{-3} data/km².

771 **Figure 11.** (a) Discrepancies between the S_{Hmax} stress orientations presented in this work and those
772 obtained using the WSM08 and the probability distributions of Bird and Li (1996). The values are
773 calculated on the Heidbach *et al.* (2010) grid. (b) Frequency distributions of the discrepancies (grey
774 histogram: relative frequency; black line: cumulative frequency).

775 **Figure 12.** (a) Discrepancies between the S_{Hmax} orientations presented in this work and those of
776 Heidbach *et al.* (2010). The values are calculated on the Heidbach *et al.* (2010) grid. (b) Frequency
777 distributions of the discrepancies (grey histogram: relative frequency; black line: cumulative
778 frequency).

779 **Figure 13.** Interpolated S_{Hmax} orientations in Italy, calculated on the Pierdominici & Heidbach
780 (2012) grid. The colours indicate the discrepancies between the S_{Hmax} orientations obtained using
781 the algorithm presented here and the orientations of Pierdominici and Heidbach (2012). The
782 rectangles indicate the areas shown in Figures 14 and 15.

783 **Figure 14.** Comparison of the S_{Hmax} orientation obtained using our algorithm with that obtained by
784 Pierdominici & Heidbach (2012) for the external Apennines. See Fig. 13 for the location of the
785 area. Thick black bar (PH2012): smoothed S_{Hmax} orientation calculated by Pierdominici &
786 Heidbach (2012) with a search radius of 100 km; thick grey bar: interpolated S_{Hmax} orientation
787 obtained in the present work ($N_C = 12$; search radius $\theta = 0.52^\circ, 58$ km); P_m : interpolation point.
788 The background colours indicate the discrepancies between the S_{Hmax} orientations obtained using
789 the algorithm presented here and those of Pierdominici and Heidbach (2012). The multicolour lines
790 denote the S_{Hmax} orientations for different kinematic regimes (WSM08; Heidbach *et al.*, 2008).
791 Tectonic regime identifiers: NF, normal fault; NS, normal-oblique; SS, strike-slip; TS, thrust-
792 oblique; TF, thrust; U, unknown tectonic regime.

793 **Figure 15.** Comparison of the S_{Hmax} orientation obtained using our algorithm ($N_C = 13$; search
794 radius $\theta = 0.52^\circ, 58$ km; P_s : interpolation point) with that obtained by Pierdominici & Heidbach
795 (2012) for Northern Sicily. See Fig. 13 for the location of the area and the caption of Fig.14 for
796 further explanations.

797 **Figure 16.** Normalised distance weight of different algorithms for points at distances between 20
798 and 340 km with 20-km steps.

799 **Figure A1.** S_{Hmax} orientations in Europe with confidence limits. Red bars: interpolated orientations
800 (this work); black bars: confidence limits (this work). Relief model from ETOPO1 (Amante and
801 Eakins, 2009).

802

803 **Supplementary files**

804 The following files and figure are for the online download.

805

806 **Table_S1.csv:** contains the interpolated stress orientations in Europe (330W-40E;25N-76N)

807 Column "ANN" contains the searching radius (expressed as Annulus number). To convert in
808 degrees, use eq. (2).

809 Column "Lat" contains the latitude of the interpolation point (degrees).

810 Column "Lon" contains the longitude of the interpolation point (degrees).

811 Column "AZI" contains the azimuth of the maximum horizontal compressive stress
812 orientations (degrees).

813 Column "C90" contains the 90% confidence interval for the maximum horizontal
814 compressive stress orientations (degrees).

815 Column "N" contains the number data (from World Stress Map release 2008) falling into
816 the search circle.

817 Column "Nc" contains the number of clusters.

818

819 **Table_S1.csv.gz:** same as Table_A1.csv, in compressed (gzip) format.

820

821 **Figure_S1.pdf:** SHmax orientations in Europe with confidence limits.

822 Red bars: interpolated orientations (this work). Black bars: confidence limits (this work). Relief
823 model from ETOPO1 (Amante and Eakins, 2009).

824

825 **Table_S1_Carafa_Barba.kmz** contains Table S1 in Google Earth format. In order for the kmz to
826 work properly, the file icon.png must be located in the same directory of the kmz.

827

828 **Important notes.** The size of the kmz can slow down the computer and let Google Earth lock.

829 We recommend proceeding as follows:

830 1) verify that the kmz and the icon.png are located in the same directory

831 2) launch Google Earth (do not click open the kmz at this stage)

832 3) zoom in the area of interest

833 4) open the kmz (File...Open, or click open the kmz)

834

835

836 First published online: February 19, 2013

837

838 <http://gji.oxfordjournals.org/content/early/2013/02/19/gji.ggt024.abstract>

839

840 Note that the online version currently has two versions of the supplementary material, one (the

841 wrong version) has the DC1 in the directory name, whereas the correct version has DC2 in the

842 directory name. The full link to the accompanying material is

843 http://gji.oxfordjournals.org/content/suppl/2013/02/21/ggt024.DC2/ggt024-supplementary_files.zip

844

845 Comments to michele.carafa@ingv.it and salvatore.barba@ingv.it

846

847 Thanks to Gabriele Tarabusi for helping with the kml conversion.

Table 1

Table 1. Regression parameters and related standard error of Eq. (7) for Annuli 1 to 8. P_0^* , P_1^* , and θ_0 : coefficients of Eq. (7) considering the associated standard errors. RMS: root mean square between empirical and analytical probabilities. R^2 : coefficient of determination, which measures the goodness of the regression fit.

Annulus	$P_0^* \cdot 10^{-3}$	St. error $P_0^* \cdot 10^{-3}$	$P_1^* \cdot 10^{-3}$	St. error $P_1^* \cdot 10^{-3}$	θ_0 (degrees)	St. error θ_0 (deg)	RMS (10^{-3})	R^2
1	4.6	0.4	204	1	12.7	0.1	1.3	0.999
2	7.6	0.3	140.5	0.7	16.6	0.2	0.8	0.999
3	12.7	0.6	102	1	18.4	0.5	1.6	0.996
4	14.9	0.8	81	1	20.7	0.9	1.8	0.993
5	14.2	0.9	70	1	25	1	1.7	0.991
6	14	1	65	1	28	2	1.8	0.989
7	13.8	0.9	57.1	0.8	33	2	1.1	0.995
8	14.8	0.8	51.4	0.7	35	2	0.8	0.996

Table 2

Table 2. Percentage of the area in which the stress orientations can be recovered for various threshold values of $\Delta\alpha$ and N_C , with the maximum search radius set to the upper limit of Annulus 8 ($\theta_8 < 2.94^\circ$; 327 km). The total extent of the study area is 2.71×10^9 km².

Minimum number of clusters (N_C)	Maximum value of 90% confidence interval $\Delta\alpha$			
	20°	40°	60°	80°
1	18%	44%	59%	75%
3	18%	44%	57%	63%
5	18%	43%	52%	56%
7	18%	40%	46%	49%

Figure 1

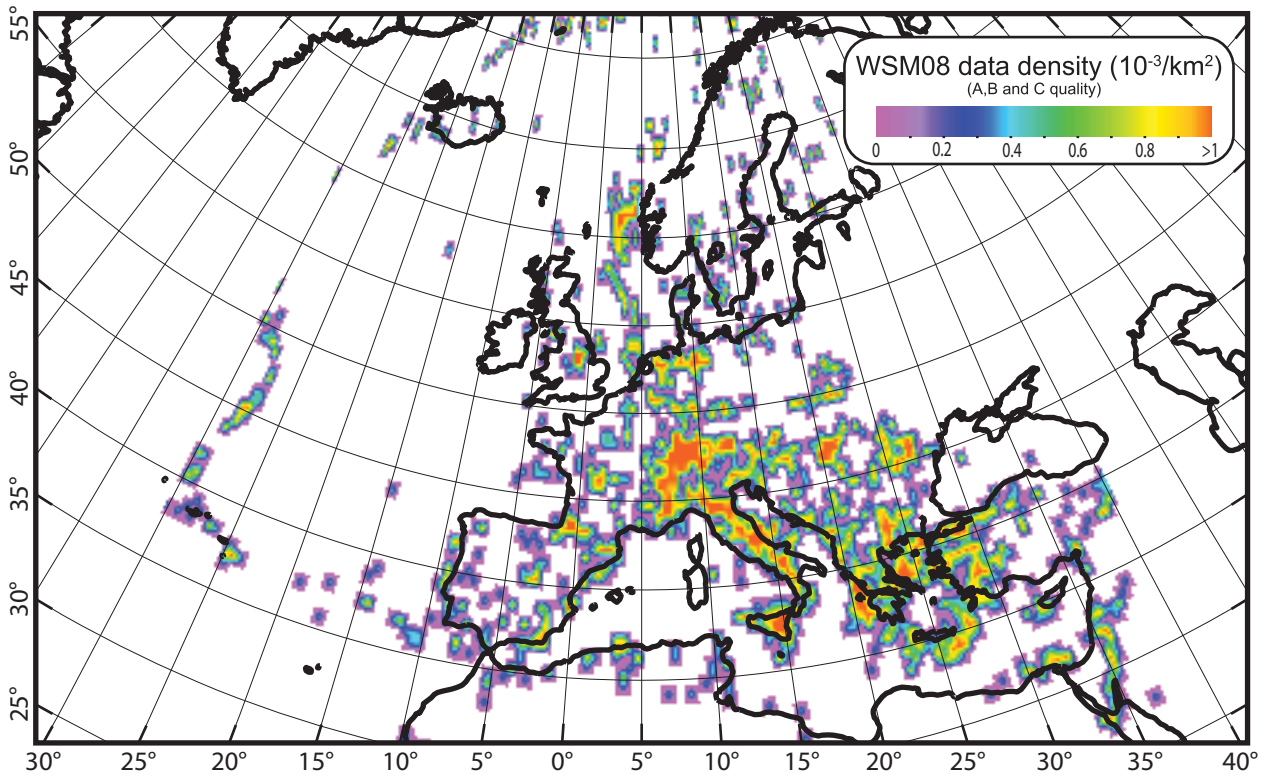
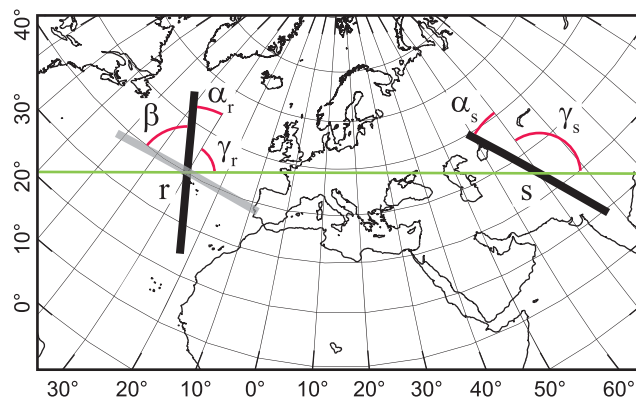


Figure 2



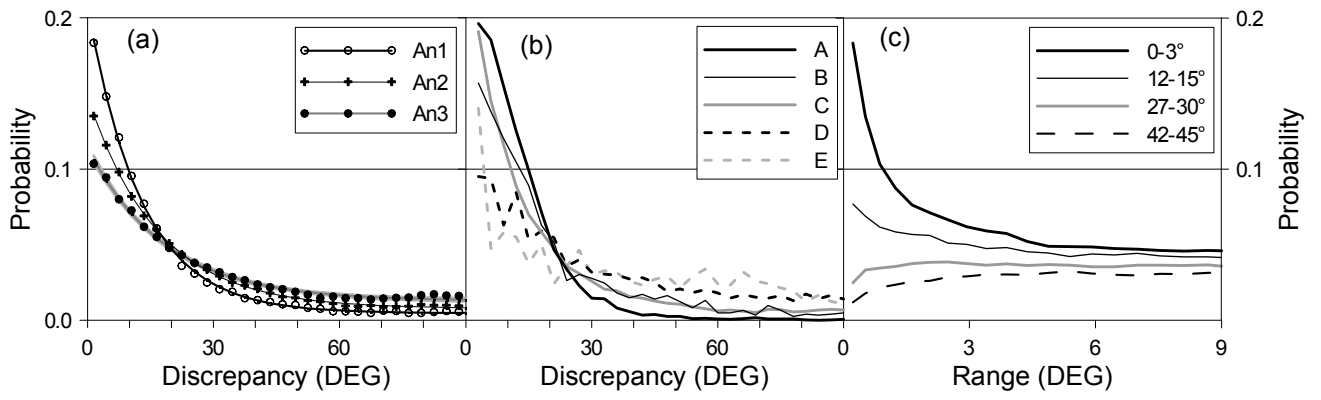


Figure 3

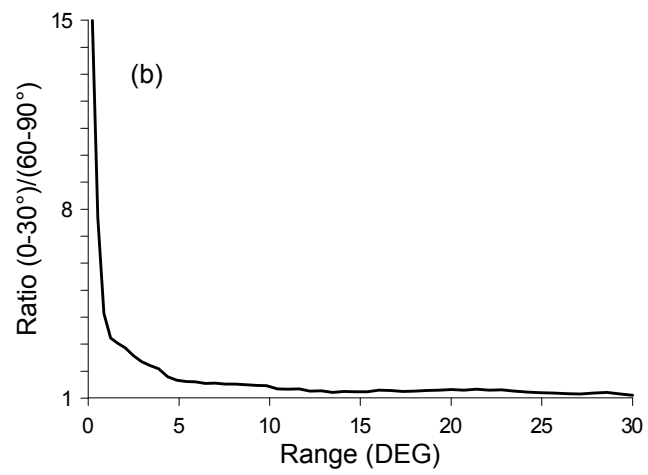
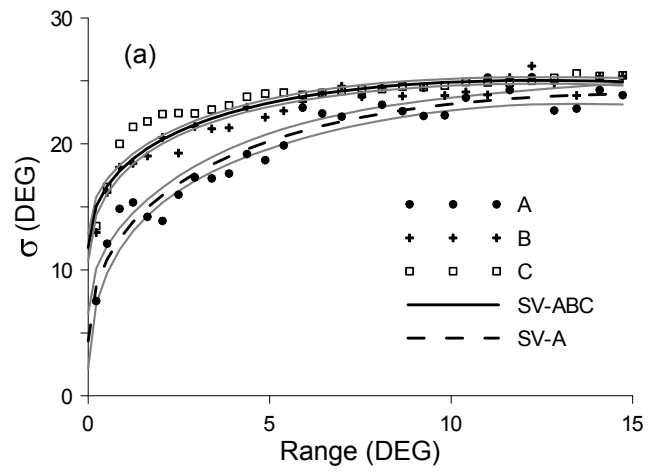


Figure 4

Figure 5

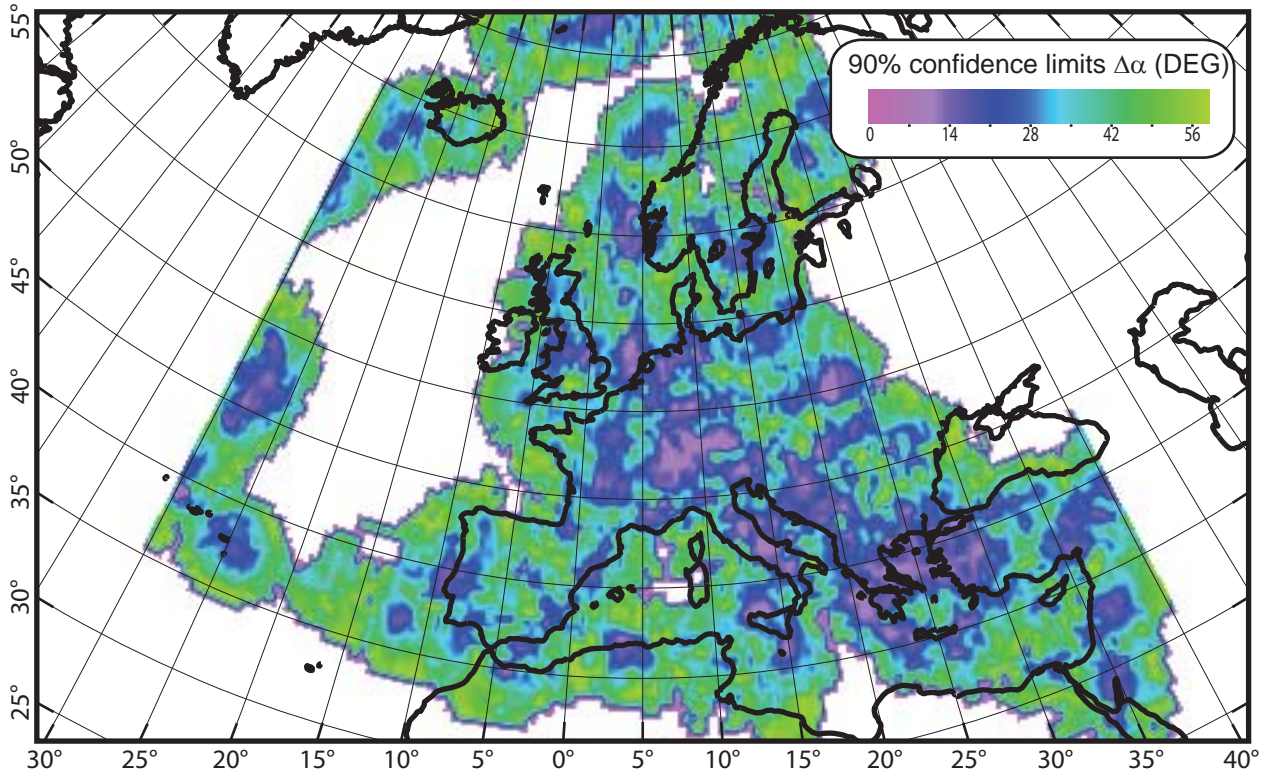


Figure 6

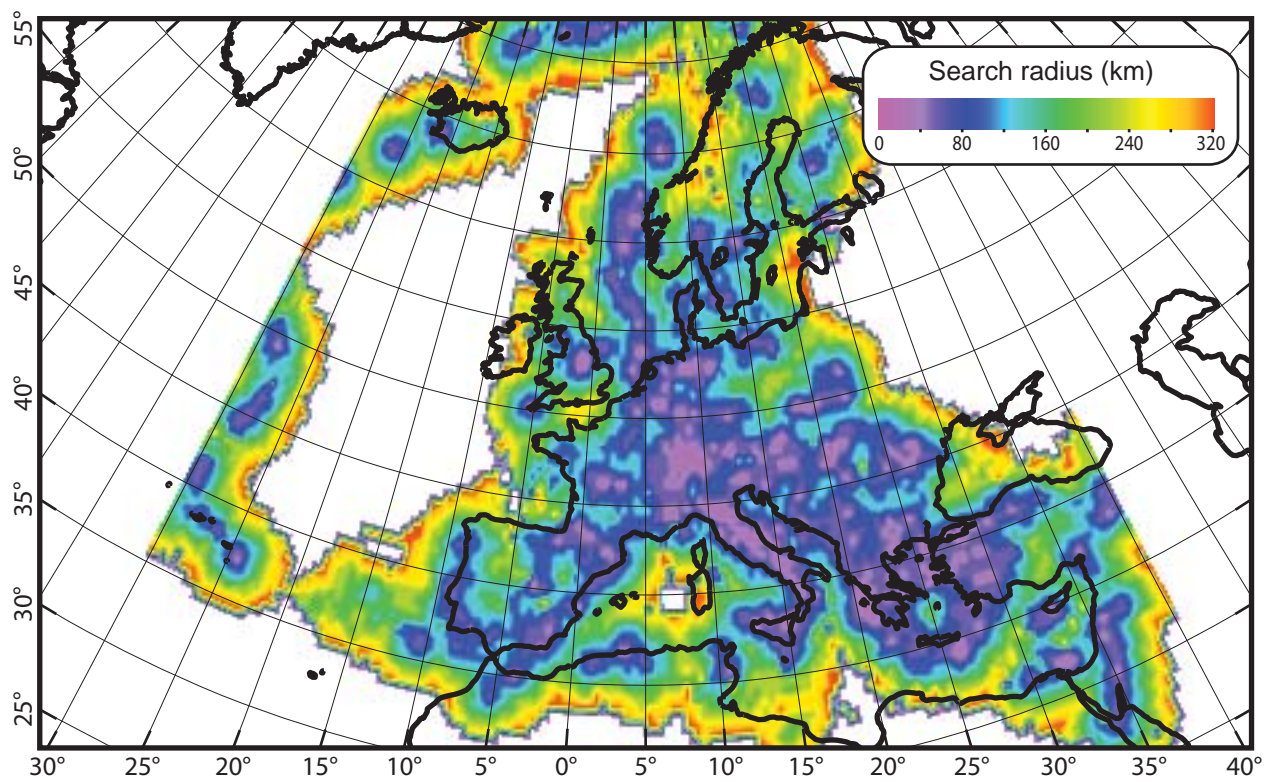


Figure 7

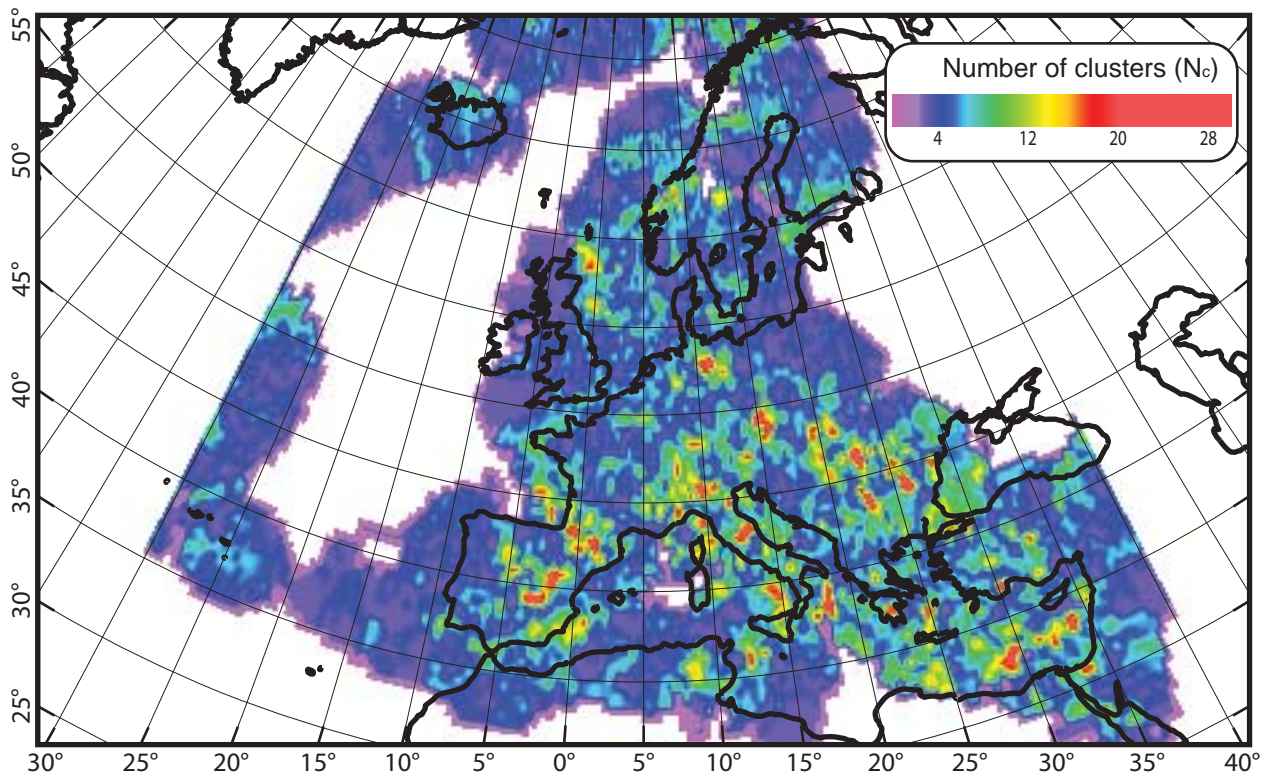


Figure 8a

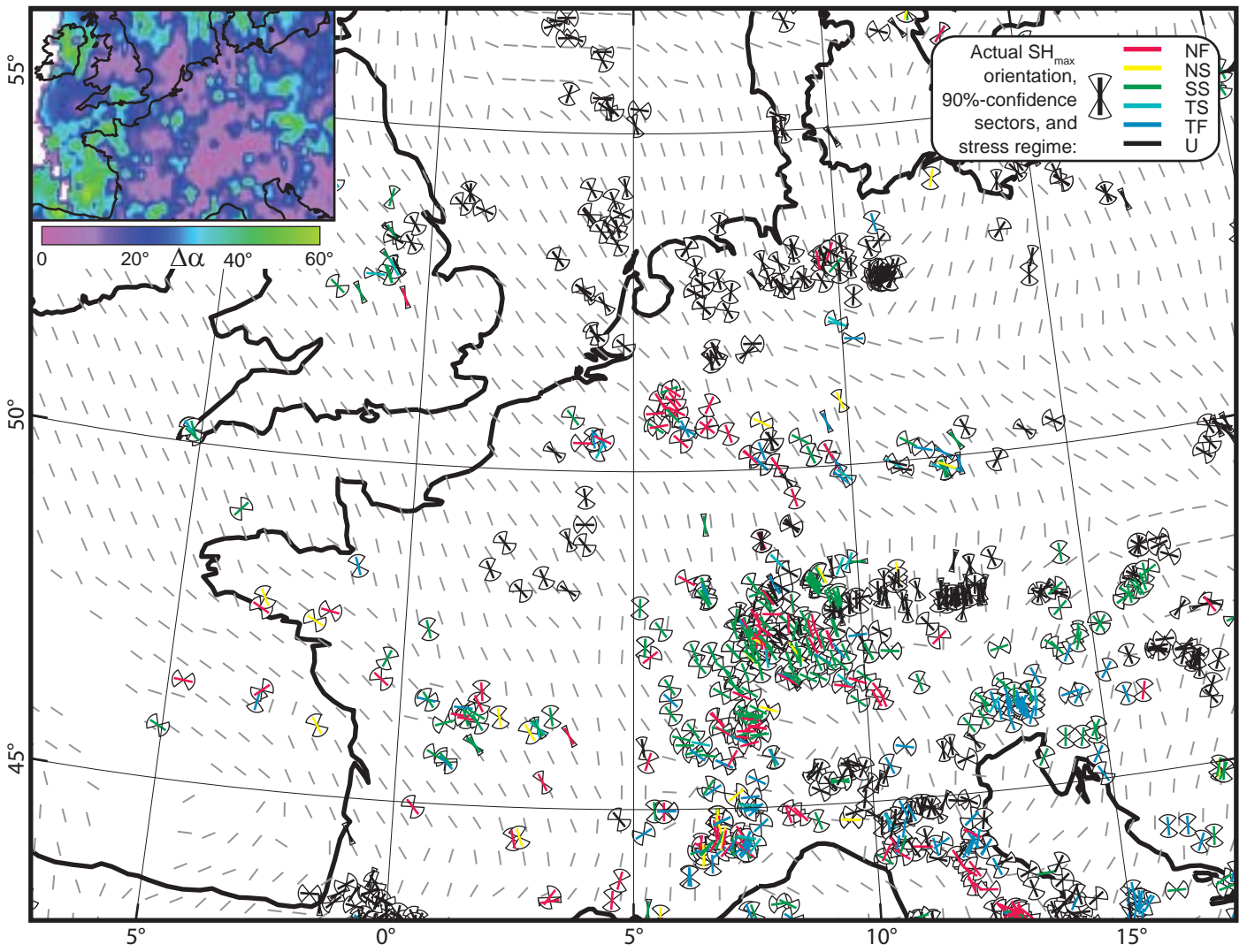


Figure 8b

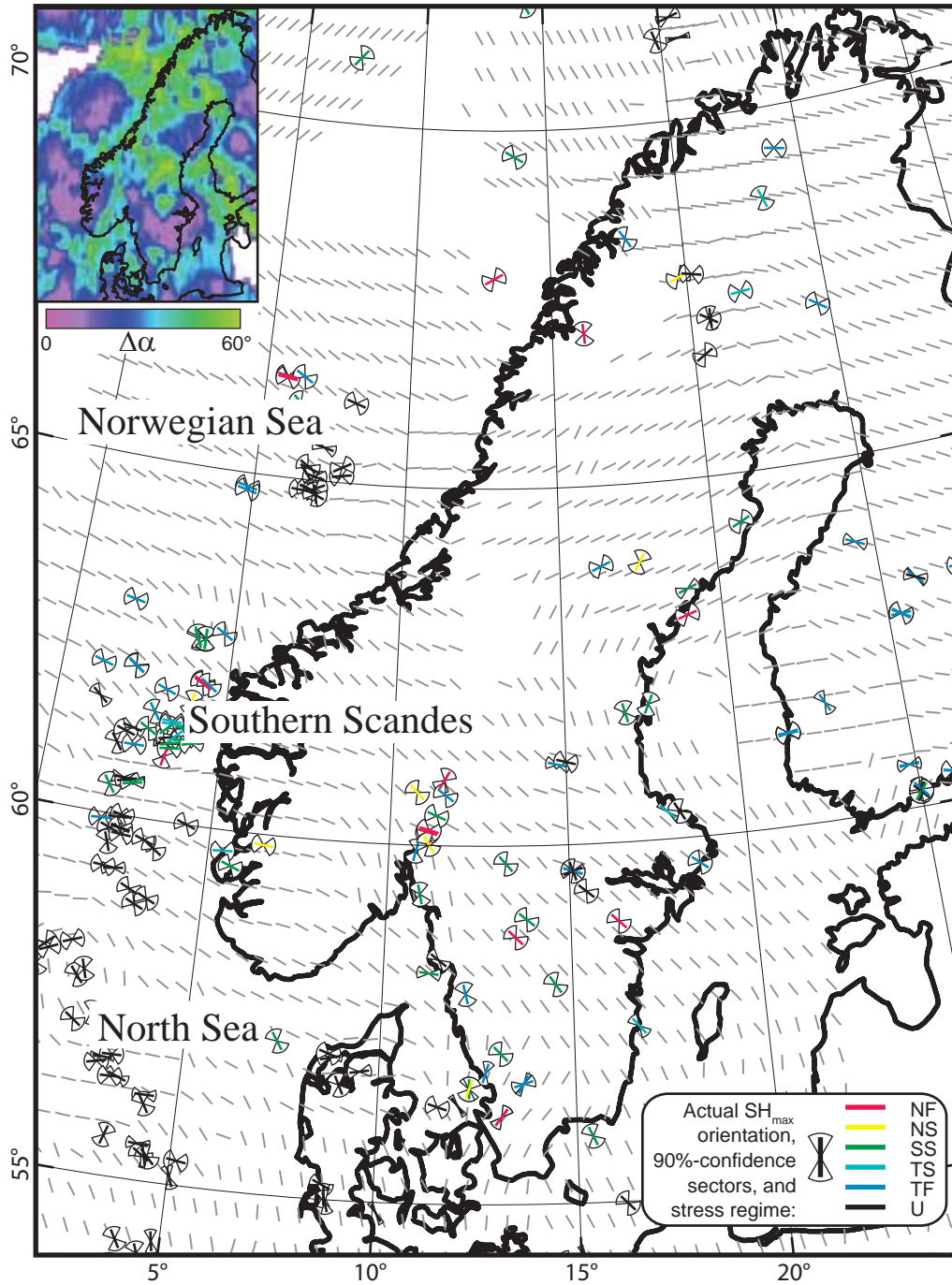


Figure 8c

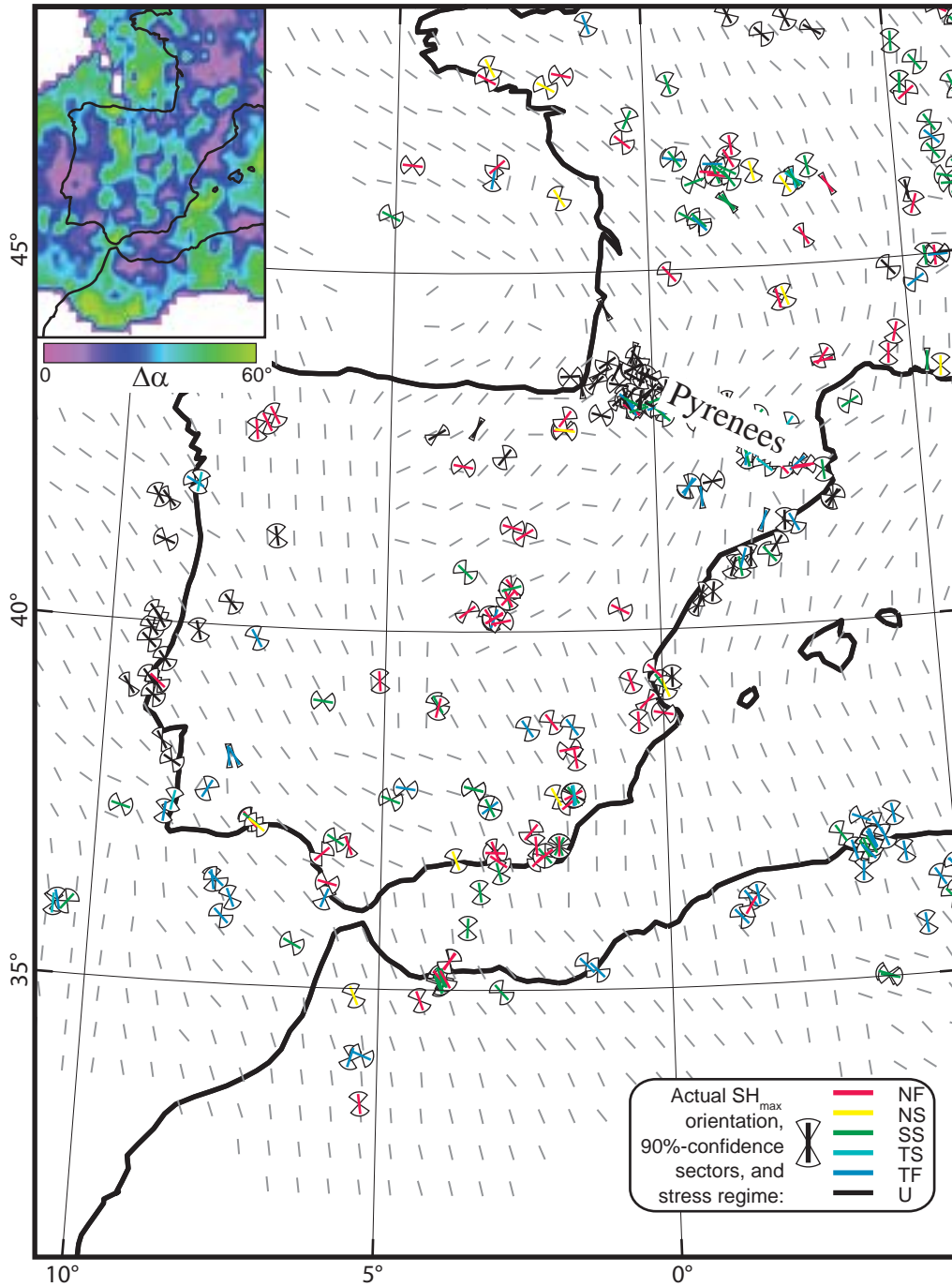


Figure 8d

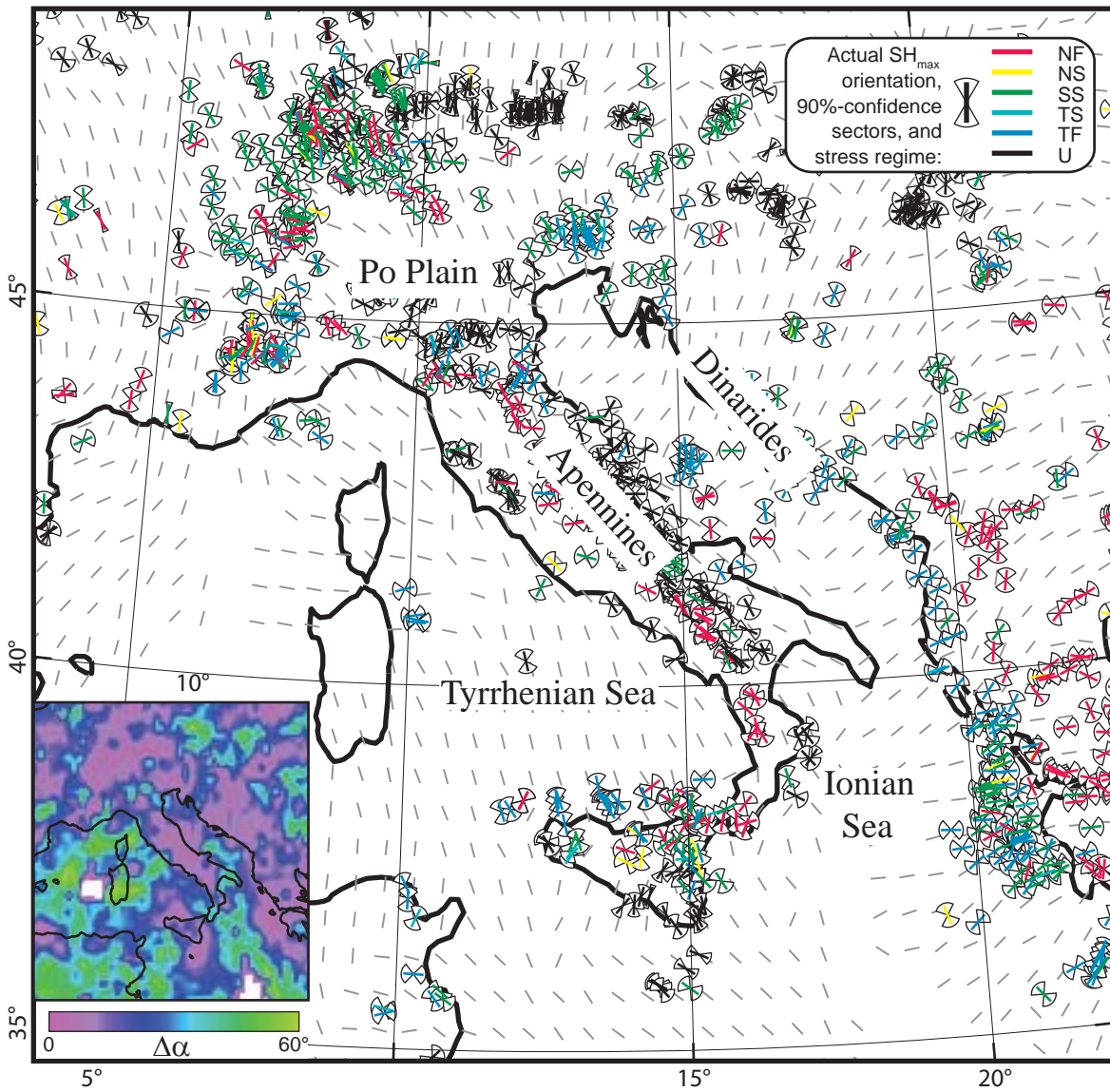


Figure 8e

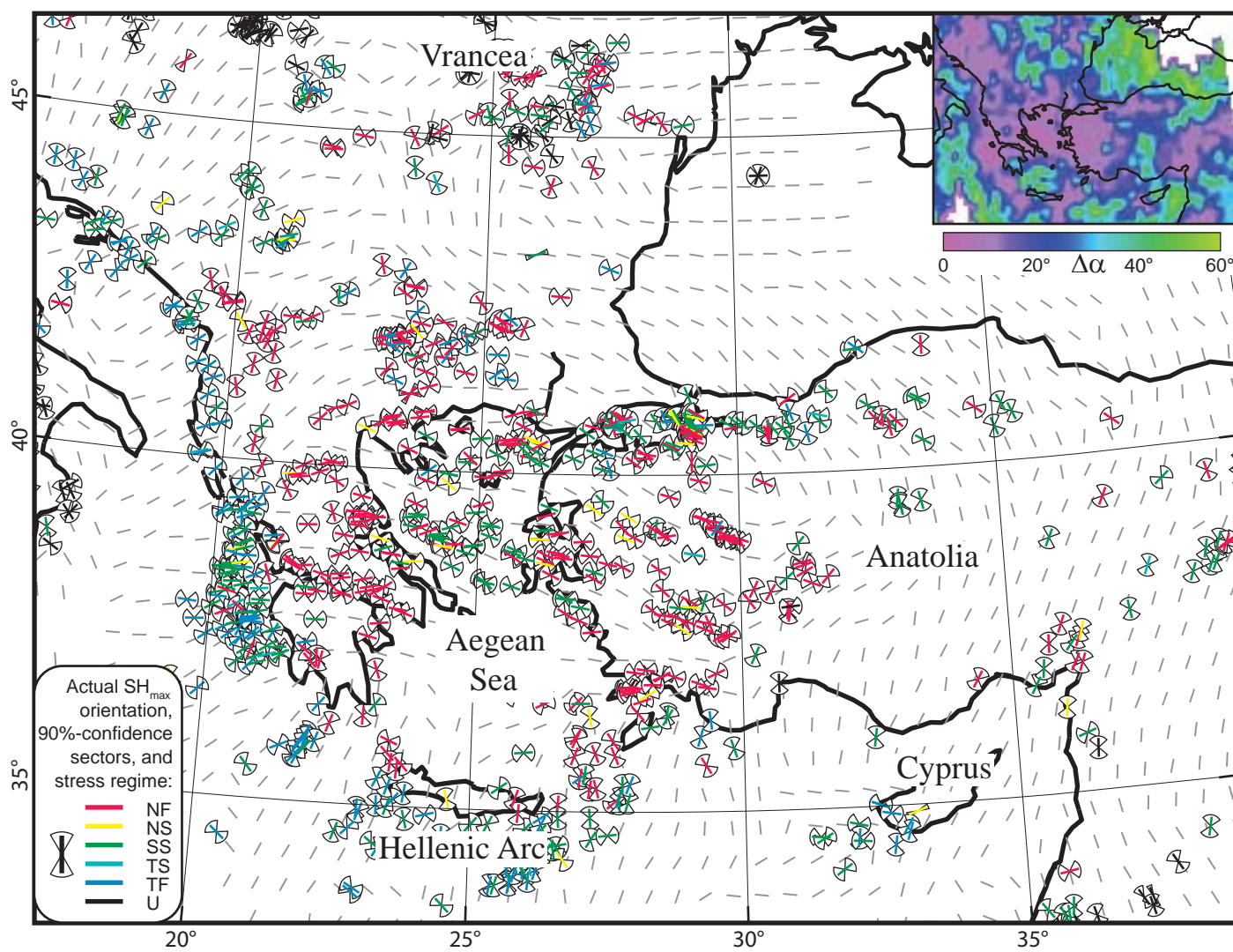


Figure 9

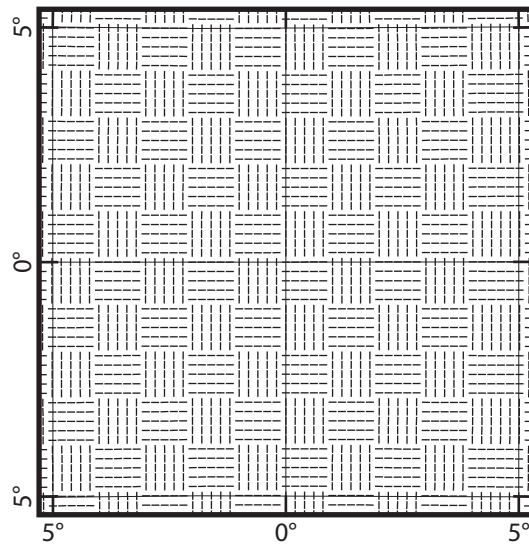


Figure 10

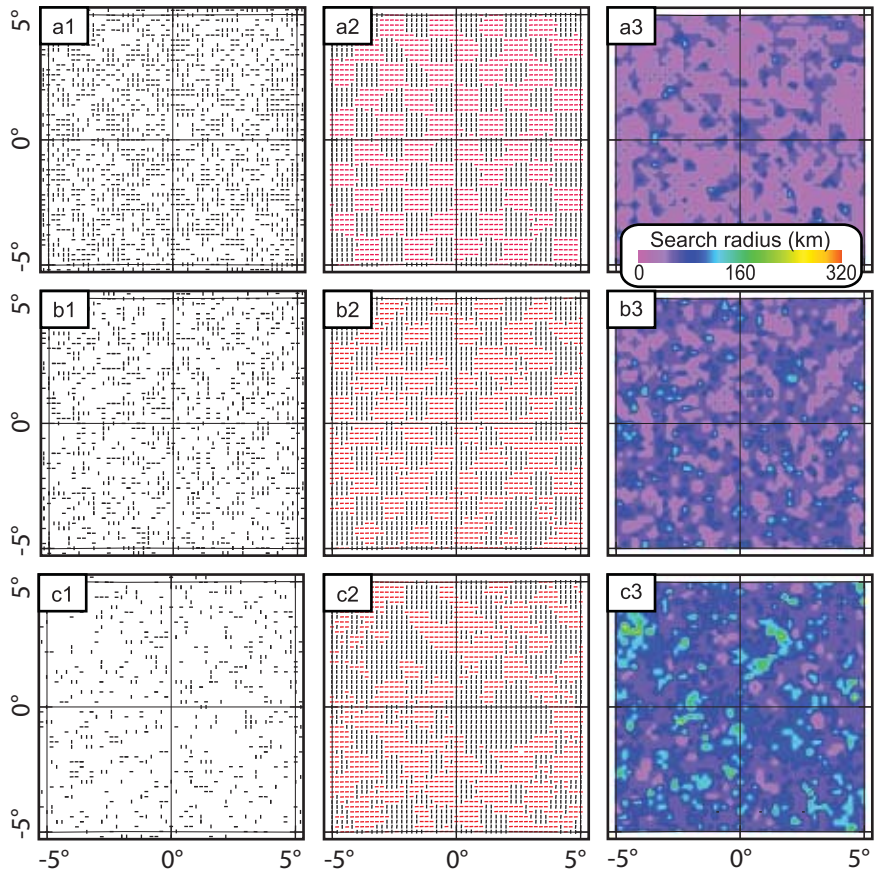


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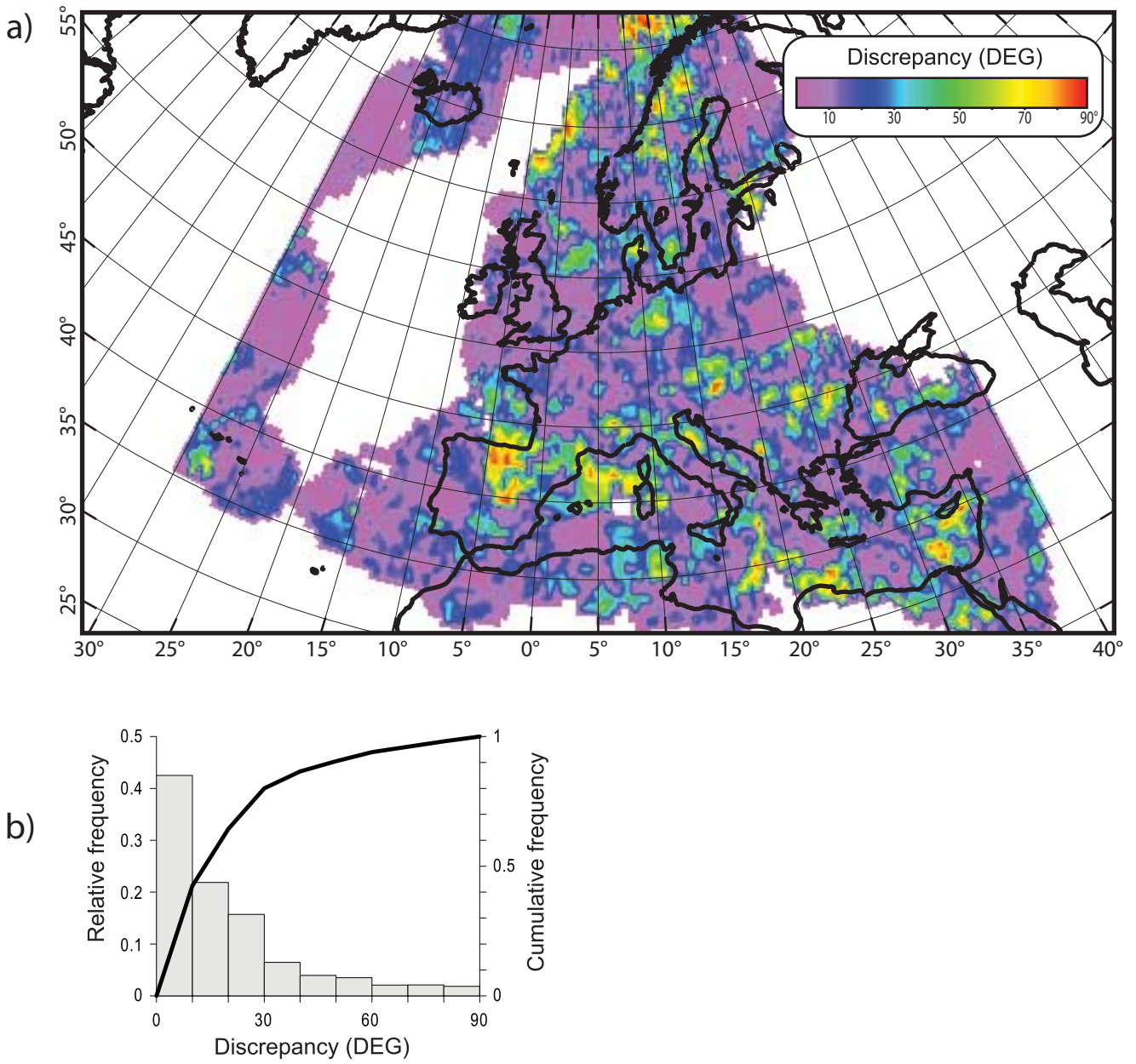


Figure 12

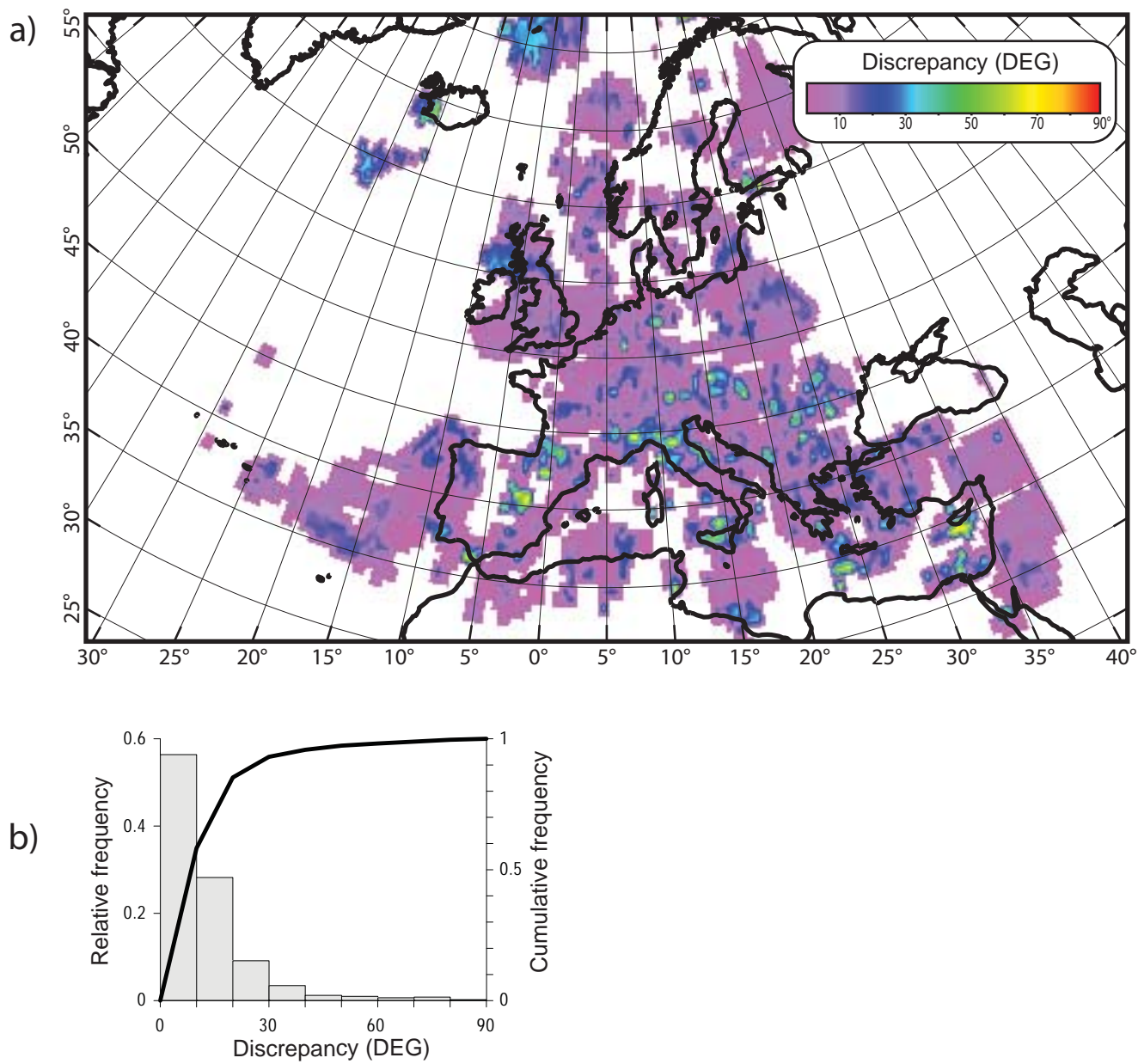


Figure 13

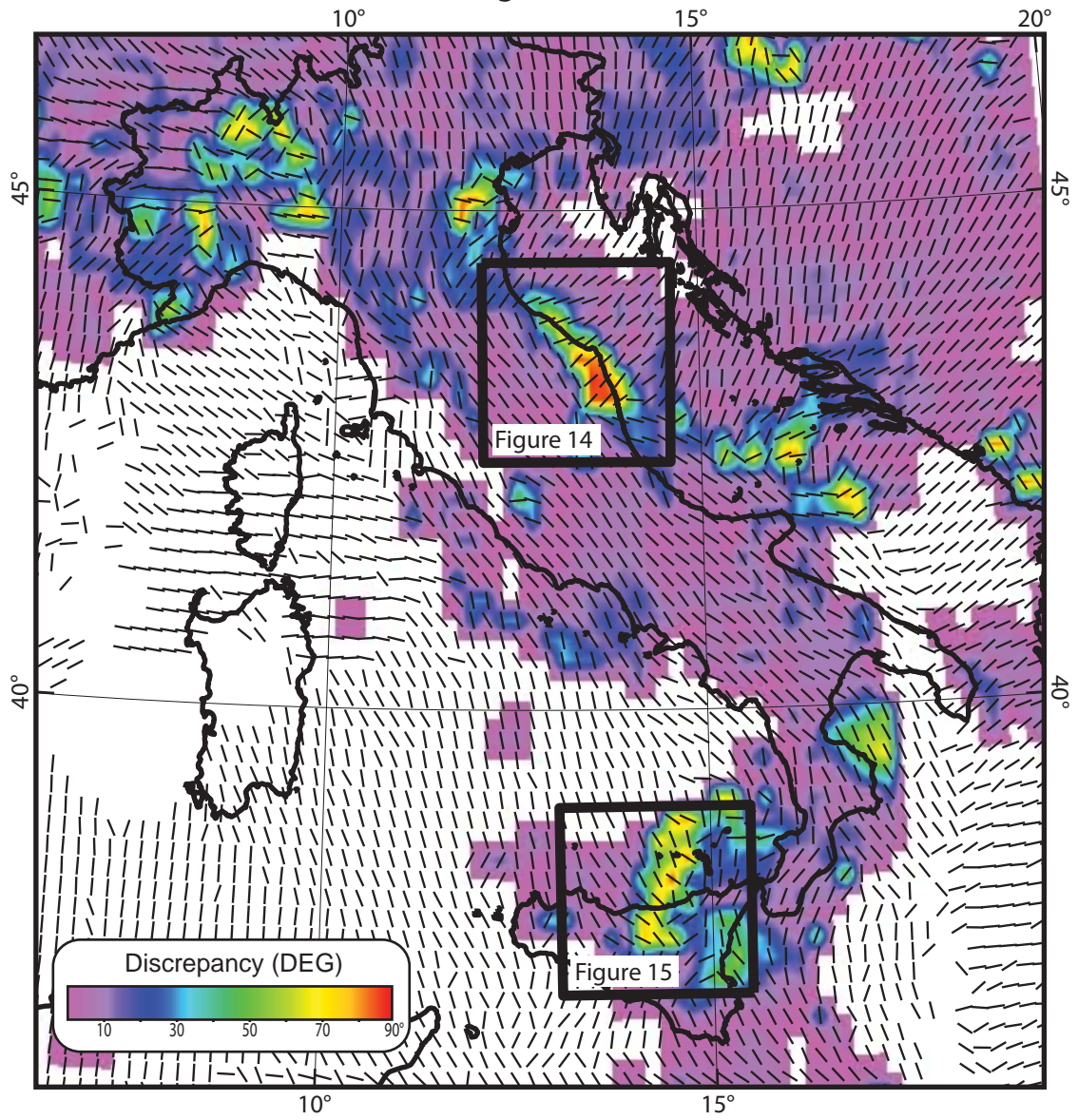


Figure 14

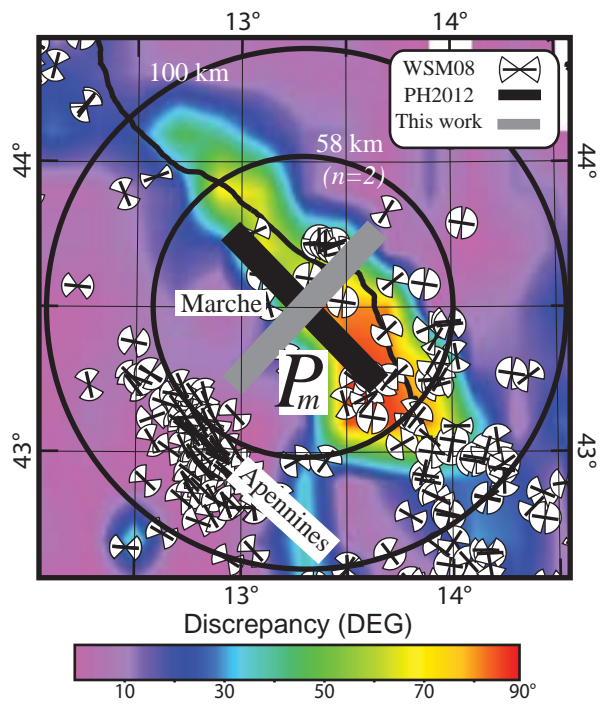
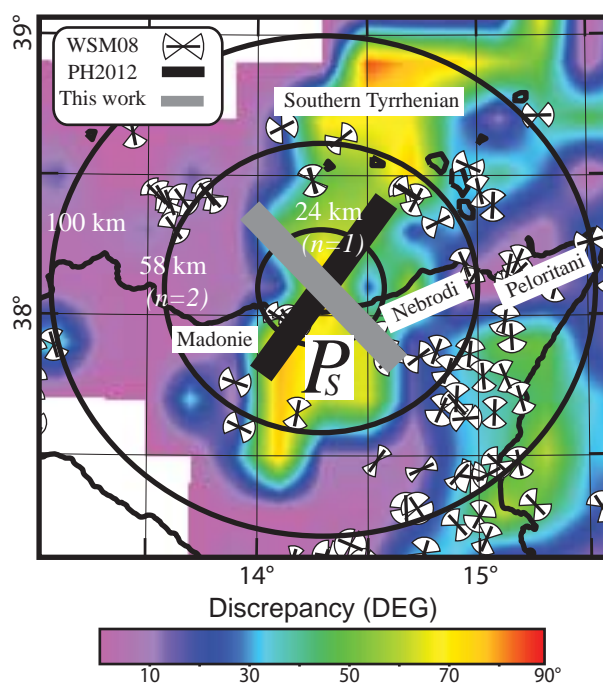


Figure 15



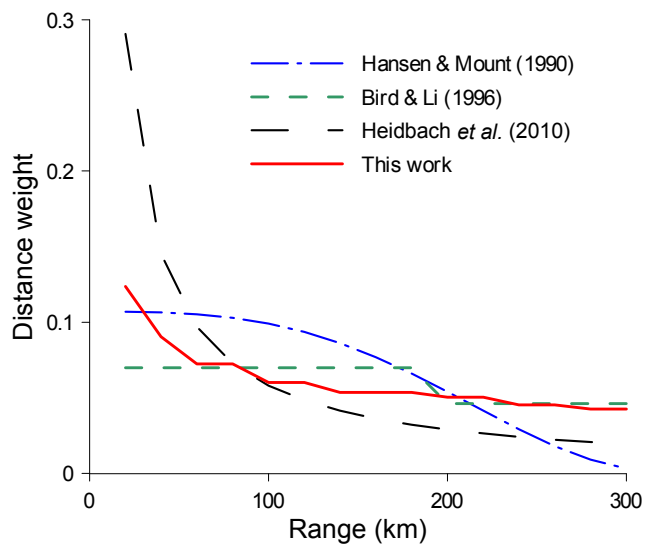


Figure 16

Figure A1

