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BMJ Open Prognostic models in acute pulmonary embolism: a systematic review and meta-analysis

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ABSTRACT

Objective: To review the evidence for existing prognostic models in acute pulmonary embolism (PE) and determine how valid and useful they are for predicting patient outcomes.

Design: Systematic review and meta-analysis.

Data sources: OVID MEDLINE and EMBASE, and The Cochrane Library from inception to July 2014, and sources of grey literature.

Eligibility criteria: Studies aiming at constructing, validating, updating or studying the impact of prognostic models to predict all-cause death, PE-related death or venous thromboembolic events up to a 3-month follow-up in patients with an acute symptomatic PE.

Data extraction: Study characteristics and study quality using prognostic criteria. Studies were selected and data extracted by 2 reviewers.

Data analysis: Summary estimates (95% CI) for proportion of risk groups and event rates within risk groups, and accuracy.

Results: We included 71 studies (44 298 patients). Among them, 17 were model construction studies specific to PE prognosis. The most validated models were the PE Severity Index (PESI) and its simplified version (sPESI). The overall 30-day mortality rate was 2.3% (1.7% to 2.9%) in the low-risk group and 11.4% (9.9% to 13.1%) in the high-risk group for PESI (9 studies), and 1.5% (0.9% to 2.5%) in the low-risk group and 10.7% (8.8% to 12.9%) in the high-risk group for sPESI (11 studies). PESI has proved clinically useful in an impact study. Shifting the cut-off or using novel and updated models specifically developed for normotensive PE improves the ability for identifying patients at lower risk for early death or adverse outcome (0.5–1%) and those at higher risk (up to 20–29% of event rate).

Conclusions: We provide evidence-based information about the validity and utility of the existing prognostic models in acute PE that may be helpful for identifying patients at low risk. Novel models seem attractive for the high-risk normotensive PE but need to be externally validated then be assessed in impact studies.

INTRODUCTION

Venous thromboembolism (VTE), including pulmonary embolism (PE) and deep vein

Strengths and limitations of this study

- Comprehensive systematic review and meta-analysis of prognostic models in acute pulmonary embolism (PE) that was not restricted to only clinical prediction rules and derivation or validation studies, but was expanded to all available prediction/predictive models including update and impact studies to inform clinical decisions, with broad search strategy and predefined selection criteria, and no data or language restriction.
- Study quality assessed by using prognostic criteria more appropriate than diagnostic tools for prognostic studies and by using a domain approach, and full details on study characteristics provided.
- Quantitative analyses performed for 'both stable and unstable PE' and for 'stable PE', for each outcome/time point separately, and for every available risk cut-off for a model to assess how a model performs along the risk scale, and results provided in terms of summary estimates of proportion of risk groups and event rates within risk groups in 'absolute risk' more meaningful for clinicians and more appropriate for the study of prognosis, and summary estimates of sensitivity and specificity (accuracy).
- Not included in the systematic review because they deserve specific reviews, studies performed in selective populations such as in asymptomatic PE, unstable PE, patients with cancer, elderly patients, or studies restricted to only a single risk group (low-risk or high-risk group) with the exception of impact studies.

thrombosis (DVT), is a common and potentially fatal disorder, despite improvements in its management. The main short-term complications of PE are all-cause death, PE-related death, VTE events and bleeding.

In acute PE, there is a real clinical questioning and interest on how to choose the appropriate management for a specific patient.^{1–4} Usual care in the early phase is to treat patients in hospital and to use



anticoagulants in normotensive patients or rescue thrombolysis in patients with hypotension or cardiogenic shock. In normotensive patients, other options are available that might depend on the estimated risk for adverse outcomes for the individual patient and include home treatment for patients at low risk, or closer monitoring and escalating therapy for patients at high risk.

Prognostic tools seek to classify patients by their risk of adverse outcomes to help select the appropriate management for an individual patient. The European Society of Cardiology 2008 guidelines¹ suggest using the haemodynamic status based on systolic blood pressure (SBP) and signs of right ventricular dysfunction (RVD) and ischaemia. The American Heart Association (AHA) Scientific Statement² considers that patients with low-risk stable PE with no prognostic markers of RVD/ischaemia 'may still have significant rates of morbidity and mortality that are functions of older age and comorbidities'.

Systematic reviews of individual prognostic variables⁵⁻⁹ have shown inconsistent results across studies, leading to the development of prognostic models. The large number of studies of existing prognostic models and the differences between them suggest the need for a systematic review that could assess study quality, synthesise findings across studies, assess the validity of the original models and provide the best estimate from the model. Determining which prognostic models work, and which work best, may impact on clinical decisions and on research.

This systematic review assesses the characteristics and quality of studies of prognostic models in patients with acute PE. It investigates the performance and utility of the prognostic models, along their different phases of development in construction, validation, update and impact studies.

METHODS

Selection criteria

Criteria for considering studies for this review were derived from previous reviews of prognostic models.¹⁰⁻¹³

To be eligible, studies had to have developed a prognostic model to predict the outcome of patients with an acute symptomatic PE, with the specific aim of constructing a new prognostic model or validating, updating or studying the impact of an existing one. The models needed to contain a combination of at least two prognostic variables to predict patient outcome, incorporated from across the following characteristics: demographic/clinical, biological and imaging related. Patient outcomes were death, PE-related death or VTE (DVT or PE) recurrence. Only hospital-based studies were considered for inclusion.

Prognostic studies validating individual prognostic indicators, population-based studies, studies performed in selective populations (asymptomatic PE, unstable PE, patients with cancer, elderly patients, specific risk group except when assessed in impact studies) or considering

only surrogate outcomes were not included. As they include prognostic and practical variables for hospitalisation, checklists of exclusion criteria from early discharge or home treatment were not considered strictly prognostic. Thus, studies on safety of early discharge or home treatment based on these criteria were not included unless their objective was clearly to derive and validate the checklist or to compare it with existing prognostic models. Unpublished articles or those published only as conference abstracts, discussion articles on prognostic models or indicators or patient management were not included. If a study was published more than once, the first published article or the article combining cumulative results from different studies was included.

Search methods and search strategy

Studies were sought through electronic databases: OVID versions of MEDLINE and EMBASE, and The Cochrane Library, from inception to July 2014. Given that there are many conference presentations in this area, we corresponded with authors and further searched in Google Scholar and other databases to check for pending or recently published full articles. The search also covered various sources of 'grey literature' to identify published, unpublished and ongoing studies (see online supplementary protocol-S). Further searches included the reference lists of relevant articles and books, citation indexes and hand-search of issues in relevant journals. There were no date or language restrictions on the searches.

The following Medical Subject Headings (MESH) terms and text words were used: pulmonary embolism.tw., exp Pulmonary Embolism/, exp Venous Thromboembolism/, prognos*.tw., exp Prognosis/, predict*.tw., cohort stud*.tw., exp Cohort Studies/, course.tw., exp Incidence/, score.tw., model*.tw., index.tw., rule.tw., criteria.tw., tool*.tw., severity index*.tw., geneva.tw., davies.tw. These were combined as shown in online supplementary box-s.

Data collection

Selection strategy

Potentially eligible studies were identified by examining titles and abstracts or other summaries as available. Full articles were obtained to assess eligibility criteria, before critical appraisal. Study identification was performed by two independent and blinded assessors to avoid selection bias. Disagreements were resolved through discussion.

Data extraction content

The data extraction form (online supplementary protocol-S) had six detailed sections: information about the review and verification of study eligibility, information about the study, assessment of study quality, assessment of methods relevant to the development phase of the model, general findings and results of model performance.

Data extraction management

An electronic data collection form was created in Excel. An explanation for the data extraction items was available in corresponding cells, but there were no coding instructions. Contents were adjusted throughout the data extraction process. Data were extracted by one reviewer and checked by a second reviewer to obtain reliable information on study findings and study methodology. Disagreements were resolved by discussion.

Dealing with missing data

Study investigators were contacted by email for details not reported in the published reports. A standard data collection form was used to assist with this, when needed.

Assessment of study quality

All studies were assessed for methodological quality^{10 13–18} and risk of bias in regard to study design, analysis and reporting according to predefined criteria for type of study, sample of patients (proper formation of inception cohort, description of referral pattern), adequacy of prognostic factors, adequacy of outcome measures, blind assessment of outcomes, completeness of follow-up, sample size, treatment, missing data and adjustment for all potential confounders (online supplementary protocol-S).

Summary scores were not used to identify studies of high quality or low quality,^{17 19} but 'risk of bias' figures were produced to show the findings graphically.

Data analysis

The descriptive analysis addressed reporting on type and phase of development of the model, study characteristics, population and setting patterns, patient characteristics, prognostic information, prognostic modelling and other data analyses, and study quality.

Quantitative analyses were performed at each relevant model risk cut-off level for various outcomes and time points. These used data from model validation and model update studies, with and without the results of the model construction study. Outcomes were all-cause mortality, PE-related mortality, VTE recurrence, major or clinically relevant bleeding and composite outcomes. These were assessed for the duration of hospitalisation, at 1 and at 3 months. Summary estimates and their 95% CIs were calculated for the population event rates, for the proportion of patients in risk categories and the incidence of events within risk categories, and were obtained as weighted average by the inverse variance method. When data were available from at least four studies, summary estimates of sensitivity and specificity and summary receiver operating characteristic (sROC) curves were obtained using the bivariate random-effects model.^{20 21}

Studies that compared models either within the same article or in different articles or shared the same cohort in various model development phases were included once for the analysis of population event rates but as

often as the number of prognostic models that were assessed for the analysis of risk group distribution, the incidence of events within risk groups and the prognostic performance of a model. Homogeneity of study designs, differences because of the case mix and statistical heterogeneity (Cochran's Q test, Higgins I² statistic) were assessed to decide whether to combine the results of individual studies. When results were combined, a fixed-effects or a random-effects model was used depending on whether the effect was similar or variable across studies. The likely influence of the presence or absence of bias was examined in sensitivity analyses and funnel plots (observed and imputed studies, Egger's regression intercept). Subgroup analyses were performed for 'stable' and 'both stable and unstable' PE patient groups, 'prospective' or 'retrospective' studies and with regard to the phase of development (derivation, internal validation and external validation/update) of studies.

Analysis was performed using metandi²² and GLLAMM²³ modules in Stata/SE (V.13.1), and using Comprehensive Meta-Analysis (CMAV.2. 2.064).

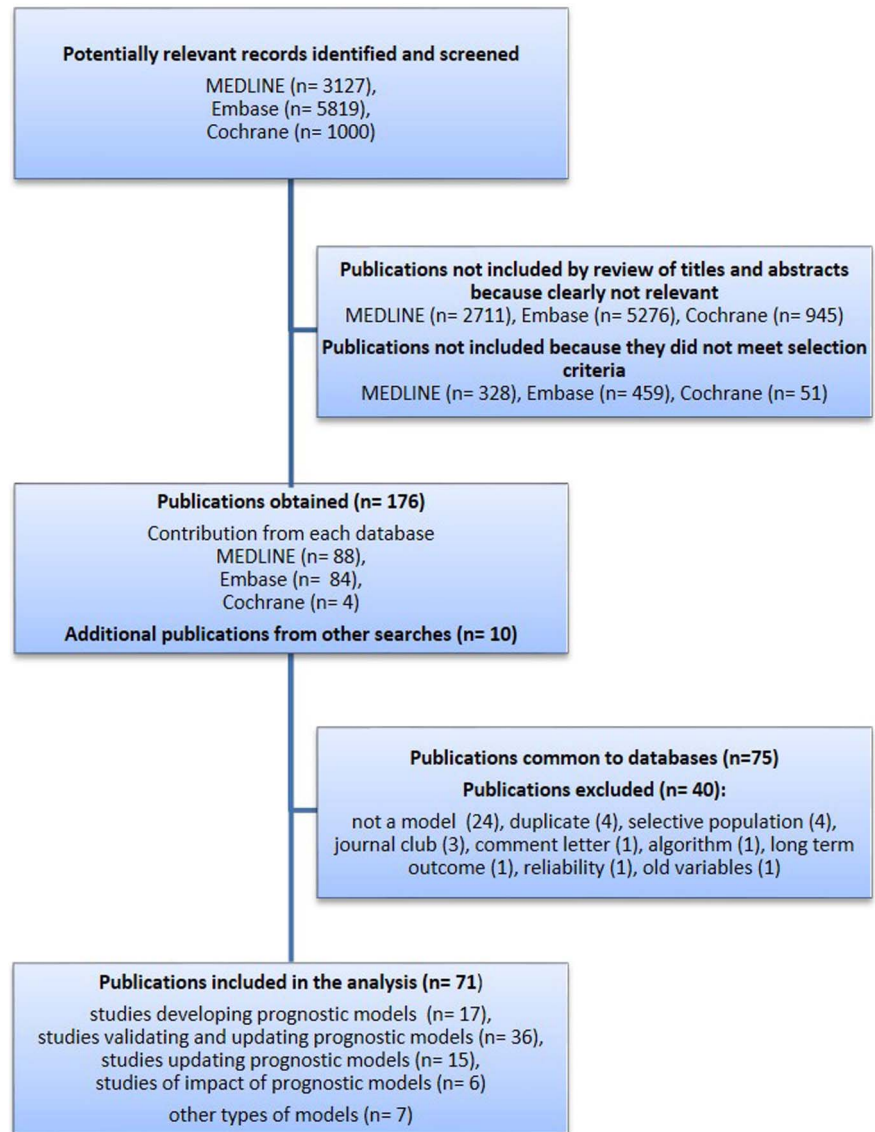
Data synthesis

To decide whether a prognostic model would be helpful for clinical practice, the following factors were required in external validation studies: direction and size of effect, and effect consistency across studies in the meta-analysis; and strength of evidence for the effect based on study quality and statistical measure of uncertainty.¹⁷ For a valid model to be acceptable, its utility should ideally have been demonstrated in an impact study, such as in a one-arm management study using the model or in a randomised trial comparing outcomes for patients in whom the model is used to inform decision-making and those for whom it is not used.

RESULTS

Search results

Figure 1 shows the flow of studies through the review following Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines.²⁴ The numbers of potentially relevant records identified and screened in each database were 3127 in MEDLINE, 5819 in EMBASE and 1000 in The Cochrane Library. Most of these were not eligible because they were clearly not relevant (2711, 5276 and 945, respectively) or because they did not meet selection criteria (328, 459 and 51, respectively). From 176 publications that were retained from these databases (88, 84 and 4, respectively) and from 10 additional publications from other sources, 75 were removed as duplicate records that had been found in more than one database and 40 were excluded for various reasons (figure 1): unmet criteria,^{25–48} duplicate,^{49–52} selective population,^{53–56} journal club,^{57–59} comment letter,⁶⁰ algorithm,⁶¹ long-term

Figure 1 Flow chart of search results.

outcome,⁶² reliability study⁶³ and very old variables in the model.⁶⁴

Finally, 71 studies were included in the review, 62 from databases^{65–126} and 9 from other sources that were mainly lists of references.^{127–135} Among studies included from databases, most were in MEDLINE and EMBASE, but four studies were not found in MEDLINE^{76 80 95 102} and one study was not in EMBASE.⁷⁷ Two studies were identified in The Cochrane Library,^{71 100} but these were also found in MEDLINE and EMBASE. Some studies were based on retrospective analysis of prospectively collected data from previous diagnostic studies,^{136–141} one¹³⁷ of which was used in two model construction studies^{68 69} and two external validation studies.^{70 97}

In total, among the 71 studies, 64 were found to be including variables specific to the domain of PE and its prognosis: 17 were identified as model construction studies,^{65 67–69 85 87 89 90 92 110 118 121–124 129 135} 41 as model external validation or model update studies^{70 72 74–78 80–84 86 88 91 93–97 99 101 102 104 106–109}

111–117 119 120 127 130 131 134 and 6 as measuring model impact.^{66 71 100 128 132 133} For the remaining seven studies, variables in the model were either originally not specific to PE^{98 103 105} or to its prognosis⁷³ or concerned a hospital checklist^{79 125 126} that is a combination of prognostic and practical variables for outpatient management.

Descriptive analysis

Prognostic models

We identified three types of models related to risk stratification of patients with PE (table 1 and see online supplementary table S1). Type 1 model includes variables specific to the domain and the prognosis of PE. Type 2 applies models that are used in other domains such as the Global Registry of Acute Coronary Events (GRACE)¹⁰³ or the Charlson Comorbidity Index,^{98 105} or applies diagnostic prediction rules⁷³ for the assessment of PE prognosis. Type 3 models incorporate hospital checklist variables as exclusion criteria from early discharge or home treatment.^{79 125 126}

Table 1 Variables used in the original prognostic models—prognostic variables included in the final models (only type 1 prognostic models presented) and points assigned (in numbers) to each variable when applied. Rule stands for scoring rule, score for simplified scoring rule or scoring system

Models	GPS	No acronym (Polish model)	PESI	PE Prognostic Algorithm	No acronym (Spanish model)	No acronym (Dutch model)	EMEP	No acronym (Japanese model)	No acronym (Chinese model)	Simplified PESI	PREP	UPS	No acronym	Simple Score FAST (Lankeit et al ⁸²) and (Dellas et al ⁸⁰)	SIRS-WCC (Jo et al ⁸⁹)	No acronym (Bova et al ¹³⁵)	eStiMaTe (PROTECT) (Jiménez et al ⁸⁷)
Model construction studies (derivation samples)	(Wicki et al ²²)	(Kostrubiec et al ⁹⁰)	(Aujesky et al ⁶⁸)	(Aujesky et al ⁶⁹)	(Uresandi et al ¹⁸)	(Agterof et al ⁶⁷)	(Volschan et al ¹²¹)	(Yamaki et al ²³)	(Zhu et al ¹²⁴)	(Jimenez et al ⁸⁵)	(Sanchez et al ¹¹⁰)	(Agterof et al ⁶⁵)	(Huang et al ¹²⁹)	(Lankeit et al ⁸²) and (Dellas et al ⁸⁰)	(Jo et al ⁸⁹)	(Bova et al ¹³⁵)	(Jiménez et al ⁸⁷)
N of candidate variables	14	NR	29	29	>30	7	20	16	NR	11	>30	7	38	NR	NR	Unclear	13
N of variables in final model	6	2	11	10	7	2	5	7	8	6	5	4	4	3	4	4	4
Proposed risk assessment method	Score	Rule	Score	Rule	Score	Rule	Score	Score	Equation	Score	Score	Rule	Score	Score	None	Score	Calculator
Type of selected variables in final model	Total N*																
Demographic and clinical characteristics																	
Age	6		Age, in year	x	1		3			1							x
Gender (male sex)	2							2						1.5			
Syncope	2								x								
Leg symptoms	1							1									
Immobilisation/bed rest	2				2		2										
No recent surgery	1				1												
Active cancer	8	2	30	x				2		1	6	x					x
Metastatic cancer	1				4												
Non-metastatic cancer	1				2												
Previous DVT/PE	2	1						2									
Heart failure	3	1	10	x													
Chronic lung disease	3		10	x									x				
Chronic cor pulmonale	1						4										
Heart failure or chronic lung disease	2									1							x
Chronic renal disease	1			x													
Cerebrovascular disease	1			x													
Recent bleeding	1				4												
Inadequate anticoagulation	1							1									
Clinical findings																	
Cardiogenic shock	2										6				x		
HR	12		20	x		x	2		x	1		x	x	2	x	1	x
SBP	6	2	30	x						1						2	x
RR	3		20				2										
Temperature	2		20														
Altered mental status	4		60	x							10				x		
PaO ₂ /SpO ₂	5	1	20	x						1							x
Accentuation of P2	1								x								
Laboratory findings																	
WCC	3												x		x		
Creatinine	1				3												
CPK	1								x								
cTnl	1																x
cTnT	1	x															
hscTn	0																
cTn	1															2	

Continued





Table 1 Continued

Model construction studies (derivation samples)	GPS (Wicki <i>et al</i> ²³)	No acronym (Polish model)	PE Prognostic Algorithm	No acronym (Spanish model)	No acronym (Dutch model)	EMEP (Voischan <i>et al</i> ²¹)	No acronym (Japanese model)	No acronym (Chinese model)	Simplified PESI (Jimenez <i>et al</i> ⁶⁵)	PREP (Sanchez <i>et al</i> ¹⁰)	UPS (Agterof <i>et al</i> ⁶⁹)	No acronym (Huang <i>et al</i> ³⁵)	Simple Score FAST (Lankeit <i>et al</i> ⁶²) and (Dellas <i>et al</i> ⁶⁰)	SIRS-WCC (Jo <i>et al</i> ⁶⁹)	No acronym (Bova <i>et al</i> ³⁵)	eStiMaTe (PROTECT) (Jiménez <i>et al</i> ⁶⁷)
BNP	1								x							x
NT-proBNP		x														
D-dimer																
H-FABP					x											
Imaging findings																
US-DVT	3															
US-proximal DVT	1						1									
TTE-RVD	1															
TTE-RV/LV ratio	2															
TTE-SPAP	1									x						
CT-RVD	1															
RVD (TTE or CT)	1													x		2

*Total N is the number of times a variable has been included across studies.

BNP, brain natriuretic peptide; CK, creatine kinase; cTnI, cardiac troponin I; cTnT, cardiac troponin T; DVT, deep vein thrombosis; EMEP, Estudio Multicéntrico de Embolia Pulmonar; GPS, Geneva prognostic score; H-FABP, heart-type fatty acid-binding protein; hscTnT, high-sensitive cardiac troponin; NR, not reported; NT-proBNP, N-terminal proBNP; PE, pulmonary embolism; PESI, PE Severity Index; PREP, Facteurs Pronostiques dans l'Embolie Pulmonaire; RR, relative risk; RVD, right ventricular dysfunction; SBP, systolic blood pressure; SIRS-WCC, systemic inflammatory response system—WCC; SPAP, systolic pulmonary artery pressure; TTE, transthoracic echocardiography; UPS, Utrecht Prediction Score; US-DVT, ultrasound-detected DVT; WCC, white cell count.

Among the 17 'type 1' prognostic models that were identified in model construction studies, 8^{68 69 85 87 92 110 118 122} underwent external validation or were updated; some of them in many studies. Three^{68 118 122} were assessed in impact studies.

There are two other models that were tested in impact studies,^{66 132} without being previously reported in a construction study. Table 1 shows that the variables that were mostly included in final models were in descending order: heart rate, active cancer, age, SBP, peripheral oxygen saturation and altered mental status.

Most of the studies identified for this review were external validation or update studies.

Study characteristics

Research objective, type of study and study design, setting, country, number of participating centres, patients (age, sex and haemodynamic status), outcomes, time points and corresponding event rates are displayed in online supplementary tables S2.1–S2.5.

In model construction studies (n=17) (see online supplementary table S2.1), the objective was usually to risk-stratify patients (13 studies) rather than to identify low-risk patients. Nine studies were prospective, seven reported on stable PE subgroup and seven were prospective and concerned stable PE as well. In 8 studies, patients were recruited in emergency departments (EDs) and in 13 studies, inception started on admission. Nine were multicenter studies. The PE Severity Index (PESI)⁶⁸ and Prognostic Algorithm⁶⁹ cohort included the largest derivation and internal validation sets, but the study was retrospective, had missing data and lack of information on PE diagnosis and on the type of treatment received. Four studies^{87 110 121 135} were sound methodologically and clinically relevant and present the following characteristics: prospective multicenter study, inception on admission, patients from ED, patients with stable PE reported, adequate early outcome and time point.

In model external validation studies (n=36) (see online supplementary table S2.2), the main objective was to identify low-risk patients. The other objectives were to assess added value of prognostic variables, head-to-head comparison or identification of patients with stable PE at high risk. Among the prospective studies (n=9) in stable and unstable PE, one study¹⁰² validated the European Society of Cardiology (ESC) model in terms of 30-day mortality and the other one compared Hestia checklist criteria with simplified PESI (sPESI) in terms of 7-day, 30-day and 3-month death; in stable PE, Geneva prognostic score (GPS),⁷⁴ PESI,^{104 111} sPESI,^{80 101} ESC,^{72 80} FAST⁸⁰ and eStiMaTe⁸⁷ models were validated in seven studies, of which one study⁸⁰ compared three models FAST, sPESI and ESC. Reliability was also assessed in two studies.^{75 106}

Model update studies (n=15) (see online supplementary table S2.3) assessed the value of adding a two-test or a three-test strategy to PESI, based on a combination of

RVD, cardiac troponin I (cTnI) and ultrasound (US)-detected DVT, to identify patients with high-risk PE;⁸⁶ assessed the additive value of CT-RVD to Hestia,¹²⁶ the additive value of transthoracic echocardiography (TTE)-RVD,¹¹¹ brain natriuretic peptide (BNP)¹¹¹ or cTnI to PESI,^{104 111} or TTE-RVD and BNP to Facteurs PRonostiques dans l'Embolie Pulmonaire (PREP),¹¹⁰ or the additive value of cardiac troponin T (cTnT),¹⁰¹ high-sensitive cTnT^{94 101} or N-terminal proBNP (NT-proBNP)⁹⁵ to sPESI, in order to identify patients at lower risk or at higher risk; or assessed the effect of cause of death classification on sPESI whether combined or not to cTnI.¹⁰⁹ Eight studies were prospective and specific to stable PE. Retrospective studies assessed the additive value of CT-RVD,¹¹² of white cell count (WCC) and systemic inflammatory response syndrome (SIRS),⁸⁹ of NT-proBNP, cTnI or CT-RVD⁷⁷ to PESI; the additive value of cTnI to PESI⁹⁶ or to sPESI;¹¹⁵ and the additive value of TTE-RVD to Shock Index.¹¹⁷

Six studies measured model impact (see online supplementary table S2.4). Four were performed in patients identified as low risk by the model, assessing the safety of early discharge or outpatient treatment. This was done in a one-arm management study for GPS¹²⁸ and NT-proBNP,⁶⁶ and in a randomised controlled trial (RCT) where outpatient management was compared with standard hospitalisation (inpatient treatment) either by using the Spanish model¹⁰⁰ or by using PESI in the Outpatient Treatment of Pulmonary Embolism (OTPE) study.⁷¹ The Pulmonary Embolism Thrombolysis (PEITHO) randomised double-blind trial¹³² assessed the role of fibrinolytic therapy tenecteplase in normotensive patients with PE considered at intermediate-high risk by TTE-RVD or CT-RVD, and myocardial injury as indicated by a positive test for cTnI or cTnT. Another study assessed the use and the impact of model-based risk stratification on treatment decisions and outcome in clinical practice.¹³³ All studies were multicenter and prospective except the latter, which was a single-centre and retrospective study.

Study quality

Quality of individual studies and a summary of study quality at different model development phases are shown in online supplementary figures S1.1 to S6.2.

Among model construction studies (see online supplementary figures S1.1 and S1.2), PREP,¹¹⁰ esTiMaTe⁸⁷ and Bova's study¹³⁵ models satisfied most of the study quality criteria. Criteria that were often met (in at least 70% of the included studies) are: inclusion criteria definition, characteristics description, length of follow-up and criteria about the outcome (objectiveness, full definition, appropriateness, proportion of patients with known outcome). Criteria that were least often met (in <30% of the included studies) are: sample completeness (selection bias), reasons for lost to follow-up reporting, similarity of outcome assessment for all study participants, differences with participants who did not

complete the study, availability of prognostic variables and justification of sample size.

In model validation studies (see online supplementary figures S2.1 and S2.2), including those that also updated the model, the most recent validation studies^{87 101 102 111} satisfied most of the study quality criteria. The most often met criteria concerned specification of inclusion criteria, sample selection, population characteristics description, follow-up, outcome (objective, fully defined, appropriate, known for a high proportion of patients), and full definition of prognostic variables and description of treatment. Major concerns were about sample completeness, reporting on reasons for lost to follow-up and differences with participants who did not complete the study, justification of sample size and statistical analysis.

In model update studies (see online supplementary figures s3.1 and s3.2), three studies^{95 101 111} satisfied most of the criteria. The summary of study quality shows that most of the criteria were often met, but there are concerns about sample completeness and reporting on reasons for lost to follow-up and differences with participants who did not end the study, justification of sample size and how lost to follow-up treated.

Model impact studies (see online supplementary figures S4.1 and S4.2) had the best compliance with study quality criteria, with most of the key factors provided.^{66 71 100 132}

For type 2 and 3 studies, the number of studies is too small to draw general conclusions. In type 2 studies (see online supplementary figures S5.1 and S5.2), although inclusion was well defined, there are problems with sample selection and sample completeness in all^{98 103 105} but one study,⁷³ in justification of sample size and data analysis. In type 3 studies (see online supplementary figures S6.1 and S6.2), the main concerns are about a lack of information on reasons for lost to follow-up, justification of sample size and data analysis (missing data, lost to follow-up, statistical adjustment).

Quantitative analysis

Population event rates

There were a total of 44 298 patients if patients in the cohorts account once, only regardless of how many studies reported those patients, and if impact studies that involved only low-risk or intermediate-risk patients are excluded.

Analysis based on inverse-variance method with random-effects model (see online supplementary table S3) showed overall in-hospital, 30-day and 3-month point estimate to be 6.5% (95% CI 4.9% to 8.5%) (20 studies), 7.4% (6.5% to 8.5%) (32 studies) and 6.8% (5.7% to 8.1%) (11 studies), respectively, for mortality; 3.1% (2.6% to 3.8%) (7 studies), 4.0% (3.2% to 5.0%) (15 studies) and 2.9% (2.2% to 3.7%) (6 studies), respectively, for PE-related mortality; and 0.2% (0.0% to 1.5%) (3 studies), 1.1% (0.7% to 1.8%) (6 studies) and 2.6% (1.6% to 4.4%) (4 studies), respectively, for VTE

recurrence. Composite event rates, major and non-major clinically relevant bleeding and fatal bleeding are also displayed.

Risk groups

All-cause mortality

The distribution of risk groups for a model (proportion of patients in low-risk or in high-risk groups) according to specified cut-off levels and the incidence of all-cause mortality within risk groups, expressed in % (95% CI), are shown in detail ([table 2](#)). The following paragraphs summarise the findings.

For $\text{GPS} \geq 3$ vs $\text{GPS} \leq 2$ cut-off, mortality was reported only in a single external validation study; the prevalence of the low-risk group is 82% (77.3% to 85.9%), and the incidence of 30-day mortality in the low-risk group is 9.8% (3.2% to 26.6%). The prevalence of high-risk group is 18% (14.1% to 22.7%) and the incidence of 30-day mortality is 20.1% (11.4% to 33%).

For $\text{PESI} \geq \text{III}$ vs $\text{PESI} \leq \text{II}$ cut-off, the overall prevalence of low-risk group is 43.1% (39.6% to 46.6%) and the overall in-hospital, 30-day and 3-month mortality rates within low-risk group are 1.4% (1.2% to 1.8%) (5 studies), 2.3% (1.7% to 2.9%) (9 studies) and 1% (0.5% to 1.8%) (6 validation studies), respectively. The overall prevalence of high-risk patients is 56.9% (53.4% to 60.4%) and the overall in-hospital, 30-day and 3-month mortality rates are 9.3% (8.4% to 10.4%), 11.4% (9.9% to 13.1%) and 13.0% (8.7% to 18.9%), respectively. Summary estimate of the prevalence and the mortality rates from external validation studies are consistent with those of the derivation and internal validation samples.

For sPESI, the overall prevalence of low-risk group is 36.3% (33.3% to 39.4%). The overall in-hospital, 30-day and 3-month mortality rates are 0.3% (0% to 2.3%) (2 external validation studies), 1.5% (0.9% to 2.5%) (11 studies) and 0.8% (0.3% to 2.2%) (3 external validation studies), respectively. The overall proportion of high-risk patients is 63.7% (60.6% to 66.7%) and the overall in-hospital, 30-day and 3-month mortality rates are 3.2% (1.6% to 6.4%), 10.7% (8.8% to 12.9%) and 13.6% (8.8% to 20.3%), respectively. Results are consistent across derivation, internal validation and external validation samples.

With the algorithm (4 studies), 22% (19.3% to 24.9%) of patients are classified as low risk. The in-hospital and 30-day mortality rates are 0.6% (0.4% to 1%) and 1.1% (0.5% to 2.1%), respectively. The results in the validation and the derivation samples are similar.

For the PREP model (mortality reported in 1 external validation study), 67.2% (61.7% to 72.3%) of patients are classified as low risk. The 30-day and 3-month mortality rates are 1% (0.2% to 3.9%) and 1.5% (0.5% to 4.5%), respectively, in the low-risk group, and 7.1% (3.4% to 14.1%) and 9.1 (4.8 to 16.6), respectively, in the high-risk group. In the original study, the incidence of a 30-day composite outcome is 2.5% (1.2% to 4.9%), 37.3% (33.2% to 41.5%) of patients are classified as

high risk and the incidence of events is 17.7% (12.9% to 23.8%).

Shock Index classifies 80.3% (68% to 88.7%) of patients in the low-risk group, but with a high 30-day mortality rate of 10.7% (6.1% to 17.9%) and 24.1% (13.6% to 39.1%) in the high-risk group.

With ESC, the low-risk group prevalence is 89.8% (72.7% to 96.7%); the in-hospital and 30-day mortality rates are 5% (3.6% to 7.1%) (3 studies) and 8.9% (4.7% to 16.5%) (1 study), respectively, in the low-risk group, and 41.1% (22.7% to 62.5%) and 26.5% (16.1% to 40.4%), respectively, in the high-risk group. These results apply for 'both stable and unstable' patients. For the stable PE patient subgroup, for 'intermediate-high' risk cut-off (intermediate-high-risk patients have RVD and myocardial injury, and the lower risk that includes intermediate-low-risk and low-risk groups either one of them or none), the proportion of the lower risk group is 74.4% (52.4% to 88.5%) (5 studies) and the in-hospital mortality rate is 3.1% (1.4% to 6.8%) in the lower risk group and 7.7% (4.7% to 12.2%) in the intermediate-high-risk group.

Given the high event rate within the 'low-risk' groups from Shock Index and ESC, these would not qualify as suitable for the identification of a low-risk but of a 'high-risk' patient group.

The performance of other models recently constructed is discussed in model update section.

Outcomes other than all-cause mortality

[Tables 3](#) and [4](#) show summary estimates with results from studies that report outcomes in terms of PE-related death, adverse outcome (clinical deterioration, haemodynamic collapse), VTE and their combinations in 'both stable and unstable' ([table 3](#)), and separately in 'stable' ([table 4](#)) patient groups. Although many studies were retrospective or were single studies, the results confirm the predictive ability for PESI and sPESI at the optimal cut-offs or less, preferably in combination with biomarkers, for identifying patients at low risk for these specific outcomes as well, mainly in stable PE. They also confirm the ability for some models at higher cut-offs for identifying patients at very high risk (30-day event rate over 20%) such as $\text{PREP—Clin} \geq \text{III}$ without or with BNP—TTE-RVD , $\text{PESI} \geq \text{IV}$, PESI V without or with TTE-RVD , ESC high, Shock Index, eStiMaTe high, FAST 3, Bova stage III. Most of the models combine biomarkers or imaging-based modalities to clinical variables or to pre-existing models. In most of the studies, the proportion of patients at very low or at very high risk is low.

Type 2 and type 3 model studies

[Table 5](#) shows the results of type 2 and type 3 model studies. Revised Geneva score⁷³ (RGS) and simplified Geneva score⁷³ (SGS) originally dedicated for the diagnosis of PE may provide at low cut-offs a small 3-month mortality rate (0.9% (0.1% to 12.3%)) in patients at low risk. Similarly, at the lowest cut-offs, GRACE ACS¹⁰³ and

Table 2 In-hospital, 30-day and 3-month all-cause mortality, and distribution of risk groups for relevant risk cut-offs

Outcome time point	Prognostic model Cut-off—risk groups	Prevalence of risk groups N (%) (95% CI)				Incidence of events within risk groups N (%) (95% CI)			
		Derivation	Internal validation	External validation	Overall	Derivation	Internal validation	External validation	Overall
In-hospital mortality	GPS-3			1 (18.9) (14.1 to 24.9)				1 (7.9) (2.6 to 21.8)	
		High						0.6 (0.1 to 4.2)	
	PESI-III	1	1	3	5	1	1	3	5
		High	(59.1) (58.1 to 60)	(59.3) (57.9 to 60.6)	(55.8) (41 to 69.7)	(58.4) (55.1 to 61.6)	(8.8) (8.1 to 9.5)	(9.8) (8.8 to 10.9)	(9.8) (6.5 to 14.4)
	sPESI			2 (53.9) (29.2 to 76.8)				2 (3.2) (1.6 to 6.4)	
		High						0.3 (0 to 2.3)	
	Algorithm	1	1	2	4	1	1	2	4
		High	(78.4) (77.6 to 79.2)	(78.4) (77.3 to 79.5)	(76.1) (55.1 to 89.2)	(78) (75.1 to 80.7)	(5.2) (4.7 to 5.7)	(6.1) (5.4 to 6.9)	(3.6) (0.7 to 16.2)
	ESC			3 (10.2) (3.3 to 27.3)				3 (41.1) (22.7 to 62.5)	
		High						5 (3.6 to 7.1)	
	Shock Index			1 (17.4) (12.2 to 24.2)				1 (22.2) (10.3 to 41.4)	
		High						3.9 (1.6 to 9)	
30-day mortality	GPS-3			1 (18) (14.1 to 22.7)			1 (20.1) (11.4 to 33)		
		High						9.8 (3.2 to 26.6)	
	PESI-III	1	1	7	9	1	1	7	9
		High	(59.1) (58.1 to 60)	(59.3) (57.9 to 60.6)	(56.1) (48.6 to 63.3)	(56.9) (53.4 to 60.4)	(14) (13.1 to 14.9)	(14.2) (13.1 to 15.5)	(9.8) (8.0 to 11.9)
	sPESI			10 (63.1) (59.8 to 66.3)				10 (10.6) (8.6 to 13)	
		High	(69.3) (66.4 to 72.1)					1.6 (0.9 to 2.7)	
	PREP Clin-II			1 (32.8) (27.7 to 38.3)				1 (7.1) (3.4 to 14.1)	
		High						1 (0.2 to 3.9)	
	Algorithm	1	1	2	4	1	1	2	4
		High	(78.4) (77.6 to 79.2)	(78.4) (77.3 to 79.5)	(76.1) (55.1 to 89.2)	(78) (75.1 to 80.7)	(11.5) (10.8 to 12.2)	(11.7) (10.7 to 12.7)	(7.9) (2.5 to 22.4)
	ESC			1 (10.1) (8 to 12.6)				1 (26.5) (16.1 to 40.4)	
		High						8.9 (4.7 to 16.5)	
Shock Index			2 (19.7) (11.3 to 32)				2 (24.1) (13.6 to 39.1)		
	High						10.7 (6.1 to 17.9)		
3-month mortality	GPS-3			2 (18) (14.1 to 22.7)			2 (23.6) (14.2 to 36.6)		
		High						4.9 (2.8 to 8.4)	
	PESI-III			6 (49.6) (43.3 to 56.0)				6 (13.0) (8.7 to 18.9)	
		High						1 (0.5 to 1.8)	
	sPESI			3 (53.8) (39.4 to 67.6)				3 (13.6) (8.8 to 20.3)	
		High						0.8 (0.3 to 2.2)	
	PREP Clin-II			1 (32.8) (27.7 to 38.3)				1 (9.1) (4.8 to 16.6)	
		High						1.5 (0.5 to 4.5)	

ESC, European Society of Cardiology; GPS, Geneva prognostic score; N, number of studies; PESI, Pulmonary Embolism Severity Index; PREP, Facteurs PRonostiques dans l'Embolie Pulmonaire; sPESI, simplified PESI.

Table 3 PE-related death, adverse outcome (death, clinical deterioration or haemodynamic collapse), VTE, major bleeding and their combinations in 'both stable and unstable' PE patient risk groups

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI)		Event rate (95% CI)	
							High-risk group	Low-risk group	High-risk group	Low-risk group
<i>Algorithm</i> _{≥1} (Jakobsson <i>et al</i> 2010) ¹³⁰	External validation	1	Retrospective	Not reported	PE-related death	30 days	83.7 (81.3 to 85.9)	16.3 (14.1 to 18.7)	6.6 (5.1 to 8.5)	0.6 (0.1 to 4.3)
<i>Agterof (Heraf rate—Ddimer)</i> (Agterof <i>et al</i> 2009) ⁶⁷	Derivation	1	Retrospective	Both stable and unstable	Death VTE major bleeding	In-hospital	67.1 (61.7 to 72.1)	32.9 (27.9 to 38.3)	5.3 (3.0 to 9.3)	0.5 (0 to 7.3)
<i>ESC high</i> (Vanni <i>et al</i> 2011) ¹¹⁹	External validation	1	Retrospective	Both stable and unstable	PE-related death	In-hospital	5.6 (3.8 to 8)	94.4 (92 to 96.2)	29.6 (15.6 to 49)	4.1 (2.7 to 6.4)
<i>ESC intermediate—high</i> (Vanni <i>et al</i> 2011) ¹¹⁹	External validation	1	Retrospective	Both stable and unstable	PE-related death	In-hospital	60 (55.6 to 64.3)	40 (35.7 to 44.4)	8.6 (5.9 to 12.4)	1 (0.3 to 4)
<i>GPS</i> _{≥1} (Wicki <i>et al</i> 2000) ¹²²	Derivation	1	Retrospective	Not reported	Death VTE major bleeding	3 months	80.6 (75.4 to 84.9)	19.4 (15.1 to 24.6)	12.5 (8.7 to 17.6)	0.9 (0.1 to 13.4)
(Nendaz <i>et al</i> 2004) ⁹⁷	External validation	1	Retrospective	Unclear	Death VTE major bleeding	3 months	71.4 (64.7 to 77.2)	28.6 (22.8 to 35.3)	12.0 (7.6 to 18.4)	3.5 (0.9 to 13.0)
	Overall	2			Death VTE major bleeding	3 months	76.3 (66.1 to 84.2)	23.7 (15.8 to 33.9)	12.3 (9.3 to 16.1)	2.7 (0.8 to 8.9)
<i>GPS</i> _{≥2} (Wicki <i>et al</i> 2000) ¹²²	Derivation	1	Retrospective	Not reported	Death VTE major bleeding	3 months	51.1 (45.1 to 57.1)	48.9 (42.9 to 54.9)	18.2 (12.6 to 25.6)	1.5 (0.4 to 5.9)
(Nendaz <i>et al</i> 2004) ⁹⁷	External validation	1	Retrospective	Unclear	Death VTE major bleeding	3 months	44.2 (37.5 to 51.2)	55.8 (48.8 to 62.5)	18.2 (11.4 to 27.6)	2.7 (0.9 to 8.0)
	Overall	2			Death VTE major bleeding	3 months	47.9 (41.2 to 54.7)	52.1 (45.3 to 58.8)	18.2 (13.7 to 23.8)	2.1 (0.9 to 5.1)
<i>GPS</i> _{≥3} (Wicki <i>et al</i> 2000) ¹²²	Derivation	1	Retrospective	Not reported	Death VTE major bleeding	3 months	32.8 (27.5 to 38.7)	67.2 (61.3 to 72.5)	26.1 (18.0 to 36.3)	2.2 (0.8 to 5.8)
(Nendaz <i>et al</i> 2004) ⁹⁷	External validation	1	Retrospective	Unclear	Death VTE major bleeding	3 months	20.1 (15.1 to 26.2)	79.9 (73.8 to 84.9)	27.5 (15.9 to 43.2)	5.0 (2.5 to 9.7)
	Overall	2			Death VTE major bleeding	3 months	26.2 (15.6 to 40.5)	73.8 (59.5 to 84.4)	26.6 (19.6 to 34.9)	3.6 (1.6 to 7.8)
<i>GPS</i> _{≥4} (Wicki <i>et al</i> 2000) ¹²²	Derivation	1	Retrospective	Not reported	Death VTE major bleeding	3 months	11.9 (8.6 to 16.4)	88.1 (83.6 to 91.4)	40.6 (25.3 to 58.1)	5.9 (3.5 to 9.8)
(Nendaz <i>et al</i> 2004) ⁹⁷	External validation	1	Retrospective	Unclear	Death VTE major bleeding	3 months	8.0 (5.0 to 12.7)	92.0 (87.3 to 95.0)	31.3 (13.6 to 56.7)	7.7 (4.6 to 12.5)

Continued

Table 3 Continued

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI)		Event rate (95% CI)	
							High-risk group	Low-risk group	High-risk group	Low-risk group
<i>GPS</i> ≥5 (Wicki <i>et al</i> 2000) ¹²² (Nendaz <i>et al</i> 2004) ⁹⁷	Overall	2			Death VTE major bleeding	3 months	10.2 (6.9 to 14.7)	89.8 (85.3 to 93.1)	37.6 (25.1 to 52)	6.7 (4.7 to 9.6)
	Derivation	1	Retrospective	Not reported	Death VTE major bleeding	3 months	3.7 (2.0 to 6.8)	96.3 (93.2 to 98.0)	70.0 (37.6 to 90.0)	7.8 (5.1 to 11.7)
	External validation	1	Retrospective	Unclear	Death VTE major bleeding	3 months	2.5 (1.0 to 5.9)	97.5 (94.1 to 99.0)	80.0 (30.9 to 97.3)	7.7 (4.7 to 12.4)
<i>GPS</i> =6 (Wicki <i>et al</i> 2000) ¹²² (Nendaz <i>et al</i> 2004) ⁹⁷	Overall	2			Death VTE major bleeding	3 months	3.3 (2.0 to 5.4)	96.7 (94.6 to 98.0)	73.0 (46.1 to 89.5)	7.7 (5.6 to 10.6)
	Derivation	1	Retrospective	Not reported	Death VTE major bleeding	3 months	1 (0.4 to 3.4)	98.9 (96.6 to 99.6)	87.5 (26.6 to 99.3)	9.1 (6.1 to 13.2)
	External validation	1	Retrospective	Unclear	Death VTE major bleeding	3 months	1 0.5 (0.1 to 3.5)	99.5 (96.5 to 99.9)	NA	9.6 (6.2 to 14.6)
<i>PESI</i> ≥II (Aujesky <i>et al</i> 2007) ⁷⁰	Overall	2			Death VTE major bleeding	3 months	0.9 (0.3 to 2.4)	99.1 (97.6 to 99.7)	87.5 (26.6 to 99.3)	9.3 (7.0 to 12.3)
	External validation	1	Retrospective	Both stable and unstable	PE-related death	3 months	79.5 (76.8 to 82)	20.5 (18 to 23.2)	2.9 (1.9 to 4.5)	0.3 (0 to 4.2)
<i>PESI</i> ≥III (Vanni <i>et al</i> 2011) ¹¹⁹ (Aujesky <i>et al</i> 2007) ⁷⁰	External validation	1	Retrospective	Both stable and unstable	PE-related death	In-hospital	68.7 (64.3 to 72.7)	31.3 (27.3 to 35.7)	7.5 (5.1 to 11)	0.7 (0.1 to 4.7)
	External validation	1	Retrospective	Both stable and unstable	PE-related death	3 months	52.6 (49.3 to 55.9)	47.4 (44.1 to 50.7)	3.8 (2.4 to 6)	0.7 (0.2 to 2.2)
<i>PESI</i> ≥IV (Choi <i>et al</i> 2014) ⁷⁷ (Aujesky <i>et al</i> 2007) ⁷⁰	External validation	1	Retrospective	Both stable and unstable	Adverse outcome	In-hospital	18 (15.3 to 21.2)	82 (78.8 to 84.7)	24.8 (17.8 to 33.4)	7.3 (5.4 to 9.9)
	External validation	1	Retrospective	Both stable and unstable	PE-related death	3 months	25 (22.3 to 28)	75 (72 to 77.7)	6.7 (4.1 to 10.8)	0.9 (0.4 to 2)
<i>PESI</i> =V (Aujesky <i>et al</i> 2007) ⁷⁰	External validation	1	Retrospective	Both stable and unstable	PE-related death	3 months	10.3 (8.5 to 12.5)	89.7 (87.5 to 91.5)	7.5 (3.6 to 15)	1.7 (1 to 2.9)
<i>PESI</i> ≥IV and other prognostic factors* (Choi <i>et al</i> 2014) ⁷⁷	External validation	1	Retrospective	Both stable and unstable	Adverse outcome	In-hospital				
NT-proBNP≥1136 pg/mL							34.2 (29.7 to 39.0)		22.4 16.1 to 30.2)	
cTnI≥0.05 ng/mL							65.8 (61.0 to 70.3)		4.7 (2.7 to 8.0)	
							39.1 (34.8 to 43.6)		21.6 (16.3 to 28.1)	
							60.9 (56.4 to 65.2)		5.2 (3.2 to 8.5)	

Continued

Table 3 Continued

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI) High-risk group Low-risk group	Event rate (95% CI) High-risk group Low-risk group
CT RV/LV diameter ratio \geq 1							36.1 (32.5 to 39.9) 63.9 (60.1 to 67.5)	20.1 (15.4 to 25.7) 5.1 (3.3 to 7.7)
PESI \geq IV—NT-proBNP \geq 1136 pg/mL							6.2 (4.2 to 9.1) 93.8 (90.9 to 95.8)	50.0 (31.0 to 69.0) 8.0 (5.6 to 11.3)
PESI \geq IV—cTnI \geq 0.05 ng/mL							10.3 (7.8 to 13.4) 89.7 (86.6 to 92.2)	41.7 (28.7 to 55.9) 8.1 (5.9 to 11.1)
PESI \geq IV—CT RV/LV ratio \geq 1							7.9 (6.0 to 10.2) 92.1 (89.8 to 94.0)	43.1 (30.4 to 56.9) 7.7 (5.8 to 10.2)
PESI \geq IV—NT-proBNP—cTnI							4.6 (2.9 to 7.3) 95.4 (92.7 to 97.1)	64.7 (40.4 to 83.2) 8.2 (5.8 to 11.6)
PESI \geq IV—NT-proBNP—RV/LV							4.2 (2.6 to 6.7) 95.8 (93.3 to 97.4)	62.5 (37.7 to 82.1) 8.4 (6.0 to 11.7)
PESI \geq IV—cTnI—RV/LV							6.0 (4.2 to 8.6) 94.0 (91.4 to 95.8)	60.7 (42.0 to 76.7) 8.5 (6.2 to 11.5)
PESI \geq IV—NT-proBNP—cTnI—RV/LV							3.8 (2.3 to 6.3) 96.2 (93.7 to 97.7)	71.4 (43.9 to 88.9) 8.5 (6.0 to 11.9)
<i>PREP</i> (Sanchez <i>et al</i> 2010) ¹¹⁰	Derivation	1	Prospective	Both stable and unstable	Death VTE	30 days		
PREP—Clin \geq II							37.3 (33.2 to 41.5) 62.7 (58.5 to 66.8)	17.7 (12.9 to 23.8) 1.9 (0.8 to 4.1)
PREP—Clin \geq III							7.2 (5.2 to 9.8) 92.8 (90.2 to 94.8)	43.2 (28.4 to 59.4) 5.4 (3.7 to 7.9)
PREP—Clin \geq II—BNP—RVD							34.8 (30.8 to 39) 65.2 (61 to 69.2)	18.4 (13.4 to 24.8) 2.1 (1 to 4.3)
PREP—Clin \geq III—BNP—RVD							5.6 (3.9 to 8) 94.4 (92 to 96.1)	44.8 (28.1 to 62.8) 5.6 (3.8 to 8)
<i>Shock Index</i> (Toosi <i>et al</i> 2008) ¹¹⁷	External validation	1	Retrospective	Not reported	Adverse outcome	In-hospital	17.6 (12.4 to 24.5) 82.4 (75.5 to 87.6)	25.9 (12.9 to 45.3) 10.3 (6.1 to 17)
<i>sPESI</i> \geq 1 (Righini <i>et al</i> 2011) ¹⁰⁶	External validation	1	Retrospective	Unclear	PE-related death	30 days	53.8 (48.6 to 58.9) 46.2 (41.1 to 51.4)	3.1 (1.4 to 6.8) 0.6 (0.1 to 4.2)
(Spirk <i>et al</i> 2011) ¹¹⁵	External validation	1	Retrospective	Both stable and unstable	Death VTE	30 days	70.8 (65.7 to 75.4) 29.2 (24.6 to 34.3)	8.1 (5.2 to 12.3) 1 (0.1 to 7)

Continued

Table 3 Continued

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI)		Event rate (95% CI)	
							High-risk group	Low-risk group	High-risk group	Low-risk group
(Righini <i>et al</i> 2011) ¹⁰⁶	External validation	1	Retrospective	Unclear	PE-related death	3 months	53.8 (48.6 to 58.9)	46.2 (41.1 to 51.4)	4.2 (2.1 to 8.1)	0.6 (0.1 to 4.2)
sPESI _{≥1} -cTn (Spirk <i>et al</i> 2011) ¹¹⁵	Update	1	Retrospective	Both stable and unstable	Death VTE	30 days	28.3 (65.7 to 75.4)	71.7 (66.6 to 76.3)	12.8 (7.4 to 21.1)	3.4 (1.7 to 6.6)
cTn							24.4 (20.1 to 29.3)	75.6 (70.7 to 79.9)	13.6 (7.7 to 22.9)	3.6 (1.9 to 6.7)

In the first column, the cut-off is indicated by the corresponding higher risk group for the cut-off.

*Information from authors.

BNP, brain natriuretic peptide; cTnI, cardiac troponin I; ESC, European Society of Cardiology; GPS, Geneva prognostic score; LV, left ventricle; NT-proBNP, N-terminal proBNP; PE, pulmonary embolism; PESI, Pulmonary Embolism Severity Index; PREP, Facteurs PRonostiques dans l'Embolie Pulmonaire; RV, right ventricle; RVD, right ventricular dysfunction; sPESI, simplified PESI; VTE, venous thromboembolism.

GRACE PE¹⁰³ risk scores show 1.3% (0.1% to 17.1%) and 1.1% (0.1% to 15.1%) 30-day mortality rate, respectively, within the low-risk patient groups. Of note is a large 95% CI for Geneva and GRACE scores due the small number of patients in the low-risk groups. Interestingly, the high event rates (23–30%) in the high-risk groups of GRACE scores whatever the cut-offs with a high proportion (49–81%) of patients in these risk groups. These models need to be validated in prospective studies with larger sample sizes in stable PE. Davies⁷⁹ and Hestia^{125 126} checklists (table 5 and online supplementary table S2.5) show a high proportion of patients at low risk (42–55%) with a small 3-month mortality rate: 1.9% (0.4% to 5.5%) and 1.2% (0.4% to 3.7%), respectively. These models need to be compared with PESI/sPESI models in impact studies.

Incidence of events along the risk scale

As expected, the event rates for different outcomes and time points increase along the risk scale. Tables 3–5 and online supplementary table S4 show how the predicted outcomes in the derivation samples compare with the observed outcomes in the validation samples (internal, external) at different cut-offs for GPS and PESI. Shifting the cut-off to a lower level or to a higher level provides more appropriate event rates in low-risk groups and in high-risk groups, respectively, but at the expense of a decrease in the proportion of patients within these risk groups. Thus, improvement in efficacy is associated with a decline in efficiency.

Prognostic accuracy

Summary estimates are displayed (see online supplementary tables S5 and S6) for different cut-offs as identified in studies to discriminate between low-risk and high-risk groups. They are provided only when at least four studies are identified for a specific cut-off, and a specific outcome and time point. The highest overall sensitivity estimates were obtained with the Prognostic Algorithm (4 studies) for predicting in-hospital mortality (97% (96% to 98%)) and 30-day mortality (98% (96% to 99%)), followed by PESI (94% (89% to 97%)) for 3-month mortality (6 external validation studies) and by sPESI (93% (90% to 95%)) for 30-day (11 studies) and for 3-month mortality (10 external validation studies) (see online supplementary table S5). All results are consistent and similar in external validation studies. The same applies for LR negative estimates. The specificity is at best 54% (46% to 62%) for PESI_{≥III} cut-off for 3-month mortality. Shifting the cut-off along the risk scale provides higher values for sensitivity or for specificity (see online supplementary table S6). In figures 2 and 3, sROC plots showing test accuracy of the two most validated models at relevant cut-offs PESI_{≥III} and sPESI_{≥1} associated with 30-day all-cause death for all studies and for external validation studies are given separately.

Table 4 PE-related death, adverse outcome (death, clinical deterioration or haemodynamic collapse), VTE, major bleeding and their combinations in 'stable' PE patient risk groups

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI)		Event rate (95% CI)	
							High-risk group	Low-risk group	High-risk group	Low-risk group
<i>Bova et al risk score</i> (Bova <i>et al</i> 2014) ¹³⁵										
	Derivation and internal validation	6 (IPD)	Both	Stable						
Stage III					PE-related death, collapse, VTE	In-hospital	5.8		28	
Stage II							18.6		9.7	
Stage I							75.5		3.6	
Stage III					PE-related death, collapse, VTE	30 days	5.8		29.2	
Stage II							18.6		10.8	
Stage I							75.5		4.2	
Stage III					PE-related death	30 days	5.8		15.5	
Stage II							18.6		5.0	
Stage I							75.5		1.7	
<i>ESC intermediate</i> (Vanni <i>et al</i> 2011) ¹¹⁹										
	External validation	1	Retrospective	Stable	PE-related death	In-hospital	57.6 (53.1 to 62.1)		6.4 (4.0 to 10.1)	
							42.4 (37.9 to 46.9)		1.0 (0.3 to 4.0)	
(Becattini <i>et al</i> 2013) ⁷²	External validation	1	Prospective	Stable	PE-related death	In-hospital	78.3 (75.4 to 80.9)		1.8 (1.0 to 3.1)	
							21.7 (19.1 to 24.6)		0.3 (0.00 to 4.1)	
	Overall	1	Both	Stable	PE-related death	In-hospital	68.9 (46.1 to 85.2)		3.4 (0.9 to 11.7)	
							31.1 (14.8 to 53.9)		0.8 (0.2 to 2.7)	
(Dellas <i>et al</i> 2014) ⁸⁰	External validation	1	Prospective	Stable	Adverse outcome	30 days	78.1 (72.2 to 83.1)		10.9 (7 to 16.4)	
							21.9 (16.9 to 27.8)		1 (0.1 to 14.1)	
<i>ESC intermediate-high</i> (Vanni <i>et al</i> 2011) ¹¹⁹										
	External validation	1	Retrospective	Stable	PE-related death	In-hospital	14.8 (11.9 to 18.4)		11.8 (6.0 to 21.8)	
							85.2 (81.6 to 88.1)		2.8 (1.6 to 5.0)	
(Becattini <i>et al</i> 2013) ⁷²	External validation	1	Prospective	Stable	PE-related death	In-hospital	41.8 (38.5 to 45.1)		1.9 (0.9 to 4.0)	
							58.2 (54.9 to 61.5)		1.0 (0.4 to 2.4)	
	Overall	1		Stable	PE-related death	In-hospital	26.2 (8.2 to 58.7)		4.9 (0.8 to 25.1)	
							73.8 (41.3 to 91.8)		1.8 (0.6 to 4.8)	
(Becattini <i>et al</i> 2013) ⁷²	External validation	1	Prospective	Stable	Death adverse outcome	In-hospital	41.8 (38.5 to 45.1)		8.8 (6.3 to 12.2)	
							58.2 (54.9 to 61.5)		3.2 (1.9 to 5.1)	
(Becattini <i>et al</i> 2013) ⁷²	External validation	1	Prospective	Stable	PE-related death adverse outcome	In-hospital	41.8 (38.5 to 45.1)		5 (3.1 to 7.7)	
							58.2 (54.9 to 61.5)		2.4 (1.4 to 4.1)	
<i>eStiMaTe high</i> (Jiménez <i>et al</i> 2014) ⁸⁷										
	Derivation	1	Prospective	Stable	Death adverse outcome VTE	30 days	3.7 (2.6 to 5.2)		25.8 (13.5 to 43.7)	
							96.3 (94.8 to 97.4)		3.4 (2.1 to 5.4)	
(Jiménez <i>et al</i> 2014) ⁸⁷	External validation	1	Prospective	Stable	Death adverse outcome	30 days	6.2 (4.5 to 8.6)		21.2 (10.5 to 38.3)	
							93.8 (91.4 to 95.5)		6.7 (5.2 to 8.7)	

Continued

Table 4 Continued

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI)		Event rate (95% CI)	
							High-risk group	Low-risk group	High-risk group	Low-risk group
(Jiménez <i>et al</i> 2014) ⁸⁷	Overall	2	Prospective	Stable	Death adverse outcome VTE	30 days	4.8 (2.8 to 8)	95.2 (92 to 97.2)	23.5 (14.7 to 35.4)	5 (2.5 to 9.4)
<i>eStiMaTe high-intermediate</i>										
(Jiménez <i>et al</i> 2014) ⁸⁷	Derivation	1	Prospective	Stable	Death adverse outcome VTE	30 days	74.5 (71.5 to 77.3)	25.5 (22.7 to 28.5)	9.7 (7.6 to 12.2)	0.9 (0.2 to 3.6)
(Jiménez <i>et al</i> 2014) ⁸⁷	External validation	1	Prospective	Stable	Death adverse outcome	30 days	63.5 (59.3 to 37.5)	36.5 (32.5 to 40.7)	7.1 (4.8 to 10.4)	0.3 (0 to 4)
(Jiménez <i>et al</i> 2014) ⁸⁷	Overall	2	Prospective	Stable	Death adverse outcome VTE	30 days	69.3 (57.6 to 79)	30.7 (21 to 42.4)	8.6 (6.5 to 11.4)	0.7 (0.2 to 2.4)
<i>FAST=3</i>										
(Dellas <i>et al</i> 2014) ⁸⁰	External validation	1	Prospective	Stable	Adverse outcome	30 days	28.4 (23.4 to 34.1)	71.6 (65.9 to 76.6)	22.1 (14.2 to 32.7)	1.5 (0.5 to 4.7)
<i>GPS_{≥3}</i>										
(Bova <i>et al</i> 2009) ⁷⁴	External validation	1	Prospective	Stable	PE-related death	In-hospital	18.9 (14.1 to 24.9)	81.1 (75.1 to 85.9)	2.6 (0.4 to 16.5)	0.3 (0.0 to 4.7)
<i>Kostrubiec et al (NT-proBNP—cTnT)</i>										
(Kostrubiec <i>et al</i> 2005) ⁹⁰	Derivation	1	Prospective	Stable						
	Intermediate				Death	30 days	72.0 (62.4 to 79.9)	28.0 (20.1 to 37.6)	20.8 (13.0 to 31.7)	1.7 (0.1 to 22.3)
	Intermediate				PE-related death	30 days	69.9 (59.8 to 78.3)	30.1 (21.7 to 40.2)	12.3 (6.3 to 22.7)	1.7 (0.1 to 22.3)
	Intermediate-high				Death	30 days	18.0 (11.6 to 26.8)	82.0 (73.2 to 88.4)	50.0 (28.4 to 71.6)	7.3 (3.3 to 15.3)
	Intermediate-high				PE-related death	30 days	16.1 (1.0 to 25.0)	83.9 (75.0 to 90.0)	40.0 (19.2 to 65.2)	2.6 (0.6 to 9.7)
<i>PESI_{≥II}</i>										
(Sanchez <i>et al</i> 2013) ¹¹¹	External validation	1	Retrospective (for PESI)	Stable	Adverse outcome	30 days	75.2 (71.4 to 78.7)	24.8 (21.3 to 28.6)	6 (4.1 to 8.8)	0.8 (0 to 5.2)
<i>PESI_{≥III}</i>										
(Palmieri <i>et al</i> 2008) ¹⁰⁴	External validation	1	Prospective	Stable	Death adverse outcome	In-hospital	69.7 (59.4 to 78.3)	30.3 (21.7 to 40.6)	53.2 (40.9 to 65.2)	11.1 (3.6 to 29.3)
(Sanchez <i>et al</i> 2013) ¹¹¹	External validation	1	Retrospective (for PESI)	Stable	Adverse outcome	30 days	37.8 (33.8 to 42)	62.2 (58 to 66.2)	9 (5.7 to 13.8)	2.1 (1 to 4.4)
(Vanni <i>et al</i> 2011) ¹¹⁹	External validation	1	Retrospective	Stable	PE-related death	In-hospital	67 (62.3 to 71.3)	33 (28.7 to 37.7)	6.1 (3.8 to 9.5)	0.7 (0.1 to 5)
<i>PESI_{≥IV}</i>										
(Sanchez <i>et al</i> 2013) ¹¹¹	External validation	1	Retrospective (for PESI)	Stable	Adverse outcome	30 days	17.2 (14.2 to 20.7)	82.8 (79.3 to 85.8)	8.8 (4.5 to 16.6)	3.9 (2.4 to 6.2)

Continued

Table 4 Continued

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI)		Event rate (95% CI)	
							High-risk group	Low-risk group	High-risk group	Low-risk group
<i>PESI</i> =V (Sanchez <i>et al</i> 2013) ¹¹¹	External validation	1	Retrospective (for <i>PESI</i>)	Stable	Adverse outcome	30 days	4 (2.6 to 6)	96 (94 to 97.4)	14.3 (4.7 to 36.1)	4.3 (2.9 to 6.5)
<i>PESI</i> +biomarkers (Sanchez <i>et al</i> 2013) ¹¹¹	External validation	1	Retrospective (for <i>PESI</i>)	Stable	Adverse outcome	30 days				
<i>PESI</i> I-II—BNP							32.7 (27.8 to 38.0)		4.7 (2.0 to 10.8)	
<i>PESI</i> I-II—cTnl							67.3 (62.0 to 72.2)		0.9 (0.2 to 3.6)	
<i>PESI</i> I-II—TTE-RVD							13.7 (10.3 to 17.9)		9.1 (3.5 to 21.8)	
<i>PESI</i> III-IV—BNP							86.3 (82.1 to 89.7)		1.1 (0.3 to 3.3)	
<i>PESI</i> III-IV—cTnl							12.3 (9.0 to 16.5)		10.8 (4.1 to 25.5)	
<i>PESI</i> III-IV—TTE-RVD							87.7 (83.5 to 91.0)		1.1 (0.4 to 3.5)	
<i>PESI</i> V—BNP							56.3 (48.8 to 63.4)		10.1 (5.5 to 17.8)	
<i>PESI</i> V—cTnl							43.8 (36.6 to 51.2)		6.5 (2.7 to 14.7)	
<i>PESI</i> V—TTE-RVD							24.0 (18.2 to 30.9)		16.7 (8.2 to 31.0)	
<i>PESI</i> +biomarkers (Jimenez <i>et al</i> 2011) ⁸⁶	Update	1	Unclear Prospective	Stable	PE-related death	30 days	76.0 (69.1 to 81.8)		5.3 (2.5 to 10.6)	
cTnl							22.8 (17.0 to 29.9)		10.8 (4.1 to 25.5)	
TTE-RVD							77.2 (70.1 to 83.0)		8.8 (4.9 to 15.2)	
US-DVT							76.2 (54.0 to 89.7)		18.8 (6.2 to 44.7)	
TTE-RVD—cTnl							23.8 (10.3 to 46.0)		8.3 (0.5 to 62.2)	
<i>PESI</i> ≥IV—TTE-RVD— cTnl							30.0 (14.1 to 52.7)		16.7 (2.3 to 63.1)	
cTnl—US-DVT							70.0 (47.3 to 85.9)		14.3 (3.6 to 42.7)	
							23.8 (10.3 to 46.0)		20.0 (2.7 to 69.1)	
							76.2 (54.0 to 89.7)		12.5 (3.1 to 38.6)	

Continued

Table 4 Continued

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI)		Event rate (95% CI)	
							High-risk group	Low-risk group	High-risk group	Low-risk group
PESI \geq IV—cTnI— US-DVT							18.0 (13.5 to 23.5)	82.0 (76.5 to 86.5)	24.4 (13.7 to 39.7)	6.4 (3.7 to 11.0)
TTE-RVD—US-DVT							8.6 (6.6 to 11.2)	91.4 (88.8 to 93.4)	19.6 (10.9 to 32.7)	5.0 (3.5 to 7.2)
TTE-RVD—cTnI— US-DVT							4.1 (2.7 to 6.0)	95.9 (94.0 to 97.3)	20.8 (8.9 to 41.3)	5.6 (4.0 to 7.9)
PESI \geq IV—TTE-RVD— US-DVT							10.5 (7.2 to 15.2)	89.5 (84.8 to 92.8)	25.0 (11.7 to 45.6)	7.8 (4.9 to 12.4)
PREP—Clin \geq II (Sanchez <i>et al</i> 2010) ¹¹⁰	Derivation	1	Prospective	Stable	Death VTE	30 days	32.3 (28.2 to 36.6)	67.7 (63.4 to 71.8)	12.3 (8 to 18.5)	2.5 (1.2 to 4.9)
PREP—Clin \geq III (Sanchez <i>et al</i> 2010) ¹¹⁰	Derivation	1	Prospective	Stable	Death VTE	30 days	1.9 (1 to 3.6)	98.1 (96.4 to 99)	22.2 (5.6 to 57.9)	4.9 (3.3 to 7.3)
sPESI \geq 1 (Dellas <i>et al</i> 2014) ⁸⁰	External validation	1	Prospective	Stable	Adverse outcome	30 days	66.1 (60.2 to 71.4)	33.9 (28.6 to 39.8)	11.2 (7.3 to 16.7)	0.5 (0 to 8)
(Lankeit <i>et al</i> 2011) ⁹⁴	External validation	1	Prospective	Stable	Death adverse outcome	30 days	62.4 (58.1 to 66.4)	37.6 (33.6 to 41.9)	8.8 (6.2 to 12.4)	1 (0.3 to 3.9)
sPESI \geq 1-hscTnT (Lankeit <i>et al</i> 2011) ⁹⁴	External validation	1	Prospective	Stable	Death adverse outcome	30 days				
hscTnT							59.3 (55.1 to 63.4)	40.7 (36.6 to 44.9)	8.7 (6.0 to 12.3)	1.9 (0.7 to 4.9)
sPESI \geq 1-hscTnT							75.9 (72.0 to 79.3)	24.1 (20.7 to 28.0)	7.8 (5.5 to 10.8)	0.4 (0.0 to 5.9)

In the first column, the cut-off is indicated by the corresponding higher risk group for the cut-off except for the Bova *et al*¹³⁵ and Sanchez *et al*¹¹¹ (PESI+biomarkers) studies which are displayed in risk categories. Bova's study combined IPD from six studies.

BNP, brain natriuretic peptide; cTnI, cardiac troponin I; cTnT, cardiac troponin T; DVT, deep vein thrombosis; GPS, Geneva prognostic score; hscTnT, high-sensitive cardiac troponin T; IPD, individual patient data; NT-proBNP, N-terminal proBNP; PE, pulmonary embolism; PESI, Pulmonary Embolism Severity Index; PREP, Facteurs PRonostiques dans l'Embolie Pulmonaire; sPESI, simplified PESI; TTE-RVD, transthoracic echocardiography-right ventricular dysfunction; US-DVT, ultrasound-detected DVT; VTE, venous thromboembolism.

Table 5 Studies reporting on derivation and external validation of models originally not specific to the prognosis or the domain of PE (type 2 model studies), and studies reporting on hospital criteria checklists (type 3 model studies)

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI)		Event rate (95% CI)	
							High-risk group	Low-risk group	High-risk group	Low-risk group
<i>RGS</i> (Bertoletti <i>et al</i> 2013) ⁷³	External validation	1	Prospective for RGS	Not reported	All-cause death	3 months				
RGS high							11.8 (8.8 to 15.5)	88.2 (84.5 to 91.2)	14.3 (6.6 to 28.3)	4.8 (2.9 to 7.7)
RGS high–intermediate							84.0 (79.9 to 87.5)	16.0 (12.5 to 20.1)	7.0 (4.6 to 10.5)	0.9 (0.1 to 12.3)
<i>SGS</i> (Bertoletti <i>et al</i> 2013) ⁷³	External validation	1	Retrospective for SGS	Not reported	All-cause death	3 months				
SGS high							7.8 (5.5 to 11.1)	92.2 (88.9 to 94.5)	17.9 (7.6 to 36.4)	4.9 (3.0 to 7.8)
SGS high–intermediate							84.0 (79.9 to 87.5)	16.0 (12.5 to 20.1)	7.0 (4.6 to 10.5)	0.9 (0.1 to 12.3)
<i>GRACE—ACS risk score</i> (Paiva <i>et al</i> 2013) ¹⁰³	External validation	1	Retrospective	Both stable and unstable	All-cause death	30 days				
GRACE—ACS high							56.8 (49.9 to 63.4)	43.2 (36.6 to 50.1)	25.6 (18.5 to 34.3)	10.1 (5.3 to 18.3)
GRACE—ACS high–intermediate							81.1 (75.1 to 85.9)	18.9 (14.1 to 24.9)	23.4 (17.6 to 30.4)	1.3 (0.1 to 17.1)
<i>GRACE—PE risk score</i> (Paiva <i>et al</i> 2013) ¹⁰³	External validation	1	Retrospective	Both stable and unstable	All-cause death	30 days				
GRACE—PE high							49.0 (42.3 to 55.8)	51.0 (44.2 to 57.7)	29.7 (21.6 to 39.3)	8.6 (4.5 to 15.7)
GRACE—PE high–intermediate							78.2 (72.0 to 83.3)	21.8 (16.7 to 28.0)	24.2 (18.2 to 31.4)	1.1 (0.1 to 15.1)
<i>Davies checklist</i> (Davies <i>et al</i> 2007) ⁷⁹	Derivation	1	Prospective	Not reported	All-cause death	3 months				
Unsuitable vs suitable							57.9 (51.0 to 64.5)	42.1 (35.5 to 49.0)	5.1 (2.3 to 10.9)	3.5 (1.1 to 10.4)

Continued

Table 5 Continued

Risk cut-off (author, year)	Development phase	N studies	Type of study	Haemodynamic status	Outcome	Time point	Proportion (95% CI) High-risk group Low-risk group	Event rate (95% CI) High-risk group Low-risk group
<i>Davies checklist</i> (Davies <i>et al</i> 2007) ⁷⁹	External validation and impact (management study)	1	Prospective	Stable meeting inclusion criteria (suitable) for home treatment	All-cause death PE-related VTE Major/fatal bleeding	3 months	Not reported	1.9 (0.4 to 5.5) None None None
<i>Hestia checklist</i> (Zondag <i>et al</i> 2013) ¹²⁵	External validation	1	Prospective	Both stable and unstable	All-cause death	30 days	46.9 (42.4 to 51.4) 53.1 (48.6 to 57.6)	4.1 (2.2 to 7.7) 0.8 (0.2 to 3.2)
						3 months	46.9 (42.4 to 51.4) 53.1 (48.6 to 57.6)	9.6 (6.4 to 14.3) 1.2 (0.4 to 3.7)
						7 days	46.9 (42.4 to 51.4) 53.1 (48.6 to 57.6)	1.8 (0.7 to 4.8) 0.2 (0.0 to 3.1)
<i>Hestia checklist</i> (Zondag <i>et al</i> 2013) ¹²⁶	External validation (and update)	1	Prospective	Both stable and unstable	All-cause death	30 days	44.6 (40.2 to 49.0) 55.4 (51.0 to 59.8)	4.1 (2.1 to 7.6) 0.7 (0.2 to 2.9)
				Both stable and unstable	Death adverse outcome	30 days	44.6 (40.2 to 49.0) 55.4 (51.0 to 59.8)	19.9 (15.2 to 25.7) 0.7 (0.2 to 2.9)
				Stable	All-cause death	30 days	41.5 (37.1 to 46.0) 58.5 (54.0 to 62.9)	3.6 (1.7 to 7.3) 0.7 (0.2 to 2.9)
				Stable	Death adverse outcome	30 days	41.5 (37.1 to 46.0) 58.5 (54.0 to 62.9)	10.3 (6.7 to 15.4) 0.7 (0.2 to 2.9)

ACS, acute coronary syndrome; GRACE, Global Registry of Acute Coronary Events; PE, pulmonary embolism; RGS, revised Geneva score; SGS, simplified Geneva score; VTE, venous thromboembolism.

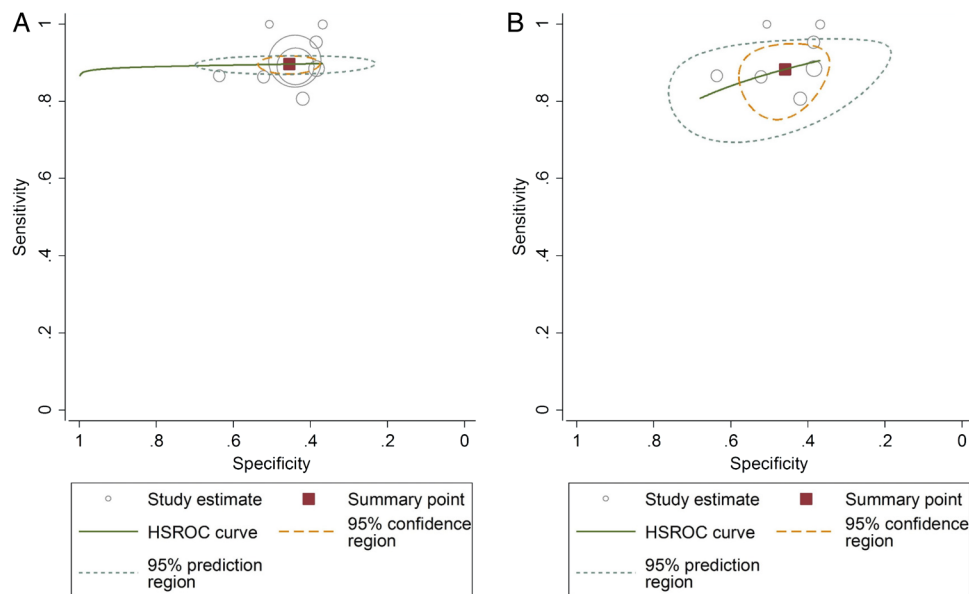


Figure 2 sROC plots showing test accuracy of $\text{PESI} \geq \text{III}$ cut-off associated with 30-day all-cause death: all studies (A) and external validation studies (B). HSROC, hierarchical summary receiver operating characteristic; PESI, Pulmonary Embolism Severity Index; sROC, summary receiver operating characteristic.

Model updating

For better identification of lower and higher risk groups, some studies assessed the effect on outcomes of adding one or more prognostic variables to an existing model. There is increasing evidence about the greater effectiveness of these ‘new models’, even though the existing ones such as PESI, sPESI and PREP are already highly effective.

In the study of Moores, the addition of non-elevated cTnI to low-risk PESI did not improve the negative predictive value compared with either test alone.⁹⁶ In the Lankeit study,⁹⁴ the combination of a negative high-

sensitive cTnT (hscTnT < 14 pg/mL) and a sPESI value of 0 at baseline improved risk classification and reduced the risk of mortality. The same applies for the combination of NT-proBNP (< 600 pg/mL) and low-risk sPESI.⁹⁵ In the study of Ozsu, hscTnT combined with sPESI provided better predictive information than cTnT.¹⁰¹ Adding cTn to the PESI¹¹³ and the sPESI¹¹⁵ models resulted in a higher area under the curve (AUC) value with no additive value for Ddimer to PESI and cTn.¹¹³ To sum up, there is an increased value in identifying lower risk patients by adding NT-proBNP and hscTnT to PESI or sPESI.

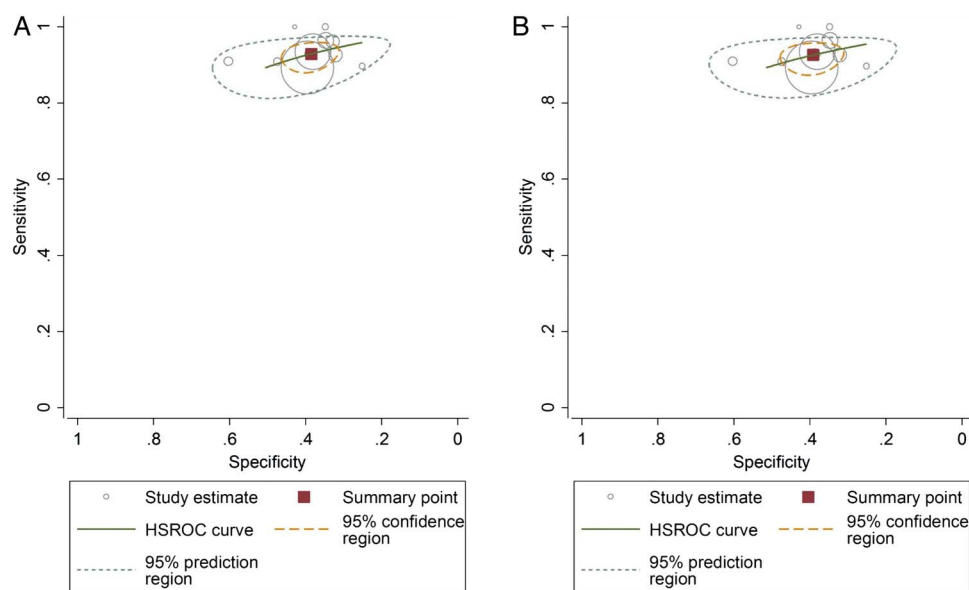


Figure 3 sROC plots showing test accuracy of $\text{sPESI} \geq 1$ cut-off associated with 30-day all-cause death: all studies (A) and external validation studies (B). HSROC, hierarchical summary receiver operating characteristic; PESI, Pulmonary Embolism Severity Index; sROC, summary receiver operating characteristic.

Other update studies tried to improve the identification of high-risk patients. Adding BNP and TTE-RVD to the PREP clinical model resulted in a significant difference in AUC.¹¹⁰ A significant increase in 'NRI: Net Reclassification Improvement', as defined by Pencina *et al*,¹⁴² was obtained in the original study which updated PREP model by adding BNP and TTE-RVD to the clinical model.¹¹⁰ Using the PREP cohort in patients with normotensive PE, Sanchez *et al* showed that biomarkers (cTnI, BNP) and echocardiography provided additional prognostic information to PESI.¹¹¹ In the Palmieri study¹⁰⁴ which included highly selected 'non-massive central PE', increased cTnI contributed to identifying patients with increased risk of development of haemodynamic instability, which was independent of, and in addition, to PESI. Novel models^{80 87 92 135} specifically developed for patients with normotensive PE and integrating biomarkers (cTn, NT-proBNP, heart-type fatty acid-binding protein (H-FABP)), CT or TTE-RVD and US-detected DVT with clinical variables (SBP, HR) or with sPESI, showed a high ability to identify patients at lower risk as well as at higher risk of early death or adverse outcome at the expense of a lower proportion of patients in these risk groups (tables 3 and 4).

Tests of heterogeneity, investigation for publication bias, subgroup analysis and sensitivity analysis

Results of tests for heterogeneity using Cochran's Q test and Higgins I² statistic, and results of subgroup analyses show major problems of heterogeneity. Comparing results between studies reporting on 30-day all-cause death for PESI optimal cut-offs (PESI \geq III, sPESI \geq 1), all conditions for comparison being optimal (all studies are retrospective and external validation studies, and include 'both stable and unstable' patients with PE), we found important between-study heterogeneity (for instance for PESI: Q value=39.69, df 5, p=0.000, I²=87) that might be related to only patient selection in studies. A high proportion of patients at low risk in study population (or a lower proportion of patients at high risk) results in less event rates in the population sample and less event rates within risk groups as shown on forest plots (see online supplementary figures S7–S12). Heterogeneity (I²>50%) was observed in most of the analyses for 30-day all-cause death with PESI \geq III and sPESI. At best I² test for heterogeneity was 39% (all studies combined) and 50% (external validation studies only) in low-risk PESI group.

Findings from the investigation for publication bias using the 'Funnel Plot' of SE and precision (=1/SE) by Logit event rate (with comparison of plots with observed studies against plots with observed and imputed studies), and the Egger's test of regression intercept show no evidence of publication bias except for studies using PESI \geq III and PESI=V cut-offs for 30-day all-cause death (see online supplementary figures S13.1 and S13.2).

Model impact

The clinical utility of a model is assessed by its effect on clinical decision-making and subsequent patient outcomes. However, few such studies have been performed of the PE models (see online supplementary table S2.4).

The safety of treating patients with low-risk PE as outpatients was examined in two management studies^{66 128} and in two randomised trials.^{71 100} The first management study¹²⁸ was a feasibility study among low-risk patients based on the GPS model in a very small number of patients. It showed the use of the GPS model to be safe. In the other one-arm management study, patients with stable PE with NT-proBNP<500 pg/mL were treated at home. Seven (4.6%) patients were readmitted within the first 10 days, but there were no deaths, no VTE events and no clinically relevant or major bleeding at 3-month follow-up. The first of the two randomised trials¹⁰⁰ did not demonstrate the clinical utility of Uresandi's model¹¹⁸ for outpatient management. The rate of short-term mortality was unexpectedly high in both management groups, as was the rate of VTE events and major bleeding. The second randomised trial⁷¹ was a non-inferiority study in highly selected patients with low-risk PE. It showed that outpatient care based on the PESI model can safely and effectively be used in place of inpatient care: 1 (0.6%) death in each group, with 1 (0.6%) VTE and 2 (1.2%) major bleeding events in the outpatient group at a 3-month follow-up. Owing to other eligibility criteria for outpatient management and study design, only 56% (152/271) with low NT-proBNP level,⁶⁶ 13% (132/1016) with low-risk Uresandi's model,¹⁰⁰ 18% (43/244) with low-risk GPS¹²⁸ and 44% (344/783) with PESI-II⁷¹ could be treated as outpatients.

The PEITHO study¹³² randomised patients with normotensive PE at 'intermediate-high' risk (with RVD and myocardial injury) to receive tenecteplase (fibrinolytics) or placebo. Death or hemodynamic decompensation occurred in 13/506 patients (2.6%) in the tenecteplase group as compared with 28/499 (5.6%) in the placebo group. Fibrinolytic therapy did not prevent in-hospital death and increased the risk of major haemorrhage and stroke. Models are less specific for predicting poor outcome in normotensive high-risk patients and might benefit from combining prognostic variables for better selection of patients, in particular those at risk of death from PE, as long as patients who are likely to respond to therapy are identified.

Finally, risk stratification was found to be frequently performed in patients admitted with acute PE and found to be stable during a 5-year period.¹³³ Its use was associated with assignment to higher levels of care and more intense treatment, but did not improve the outcome.¹³³

DISCUSSION

Summary of main results

This systematic review identified 17 prognostic models in acute PE and other types of models that were not



originally specific to PE or exclusively to its prognosis. Eight models were validated or updated and three were assessed in impact studies. Overall, PESI and sPESI are the models that have been most widely validated and updated, and PESI is the only one that has been assessed and found useful in a randomised trial for treating patients with low-risk PE as outpatients (level 1 of hierarchy of evidence according to McGinn *et al*¹⁴³). As expected, the event rates for different outcomes and time points increase along the risk scale in PESI model. The attempt to simplify PESI is attractive and successful in validation studies, but needs to be evaluated in impact studies. Agreement between the simplified PESI and the original PESI is fair.¹⁰⁶ PESI and sPESI models are now included in the risk stratification of patients with PE in the 2014 ESC guidelines.³ This new strategy needs to be validated and compared to the existing models. Other prognostic models have shown improvement in identifying low-risk and high-risk groups either on their own (algorithm for the low risk—Shock Index, 2008 ESC high, PREP—Clin III, PESI \geq IV, PESI V, eStiMaTe high, FAST 3, Bova stage III for the high risk), or by the addition of one or more clinical, biological and imaging-based markers of RVD and myocardial injury to the existing models or by incorporating these markers into new models (NT-proBNP and hscTn for the low risk—SBP, BNP, cTN, H-FABP, venous US and imaging-based RVD for the high risk). Large validation and impact studies are needed to assess these new and updated models.

Overall completeness and applicability of evidence

Most of the models are effective and provide a low event rate in low-risk groups, and appear reproducible and robust. However, using the model for identifying the low-risk groups may not be worthwhile if the aim is to make a choice to send the patients home; because the patients might still have an unacceptable high risk of a serious adverse outcome. It is important to define the incidence limit for a specific outcome, which should be clinically relevant and if models go beyond this they should not be viewed as of high performance. In agreement with the review of Vinson *et al*,¹⁴⁴ careful selection is needed for low-risk patients with acute PE who will be managed as outpatients. Broad implementation of this management strategy is controversial and varies across countries and across organisation of patient care and patient pathways. Furthermore, decisions about management might be modified by the availability of new oral anticoagulants being studied in large-scale clinical trials.¹⁴⁴ Hospital checklist criteria^{79 125} should also be helpful for patient selection and seem to be competitive with prognostic models.

Differences in predictive performance of the models may be due to differences in prognostic criteria and threshold for risk groups regardless of study design, development phase, population, outcome and time point. In two model construction studies^{68 69} using the

same cohort and the same outcome, a change in criteria and threshold resulted in change of risk group proportions and event rates within these risk groups. Shifting the cut-off to reduce the incidence of events in the low-risk group led to a smaller proportion of patients in the low-risk group, and selecting patients with lower risk PE makes the model more effective but less efficient. This also applies to the high-risk patient groups: the greater the proportion of patients classified as high risk, the lower the incidence of outcomes in this group and vice versa. Adding biomarker data may help to identify higher or lower risk patients, but the findings on the additive value of biomarkers are inconsistent across studies. This may be explained by the higher predictive performance of some models, heterogeneity in the populations arising from study design^{66 71 100} or a higher sensitivity of some biomarkers.⁹⁴

Quality of the evidence

In our review, we used a comprehensive search strategy, assessed the quality of the included studies to allow us to focus on those of good quality, and used statistical techniques to analyse their results and combine and compare the findings. However, the included studies contain some methodological flaws in the design, conduct and reporting. Many are retrospective or use prospectively collected data for diagnosis purposes, and one of the most common limitations is the lack of reporting of the case-ascertainment strategy used in the study. These problems might lead to selection bias that could have affected the event rates and influence the variables included in the final model. The time of inception of the cohorts (ie, whether it was at diagnosis or on admission) was ill defined and patient selection may have been different across studies. Other common methodological issues are the lack of a justification for the sample size, the absence of a full definition of prognostic variables and the lack of standardisation of treatments which might have been dictated by the prognostic variables that were assessed (and, therefore, may have affected the likelihood of certain outcomes), the lack of blinded assessment of the outcomes (to the prognostic variables), and possible problems with data analysis related to the statistical methods for missing data.

Potential biases in the review process

Our search was performed up to July 2014. Publications after that date were not included in the review. We believe there are no biases in our searches within the search period and decisions on study eligibility. Although we used a highly sensitive search to minimise the influence of reporting biases related to duplicate and cumulative full publication, time lag, language, location and reference list citation,¹⁷ it is not possible to know how a study's findings might influence the researchers' decision to publish it; unlike in randomised trials where positive results for new treatments are more

Table 6 Comparison of our study with existing systematic reviews and meta-analyses of clinical prediction rules

Author, year	Squizzato <i>et al</i> , 2012 ¹⁴⁶	Zhou <i>et al</i> , 2012 ¹⁴⁸	Kohn <i>et al</i> , 2015 ¹⁴⁷	Our study
Study design	Systematic review and meta-analysis	Meta-analysis	Systematic review and meta-analysis	Systematic review and meta-analysis
Years of search	Until August 2011	Up to June 2012	January 2000–March 2014	Inception to July 2014
Databases for search	MEDLINE, EMBASE	MEDLINE, EMBASE	MEDLINE, EMBASE	MEDLINE, EMBASE, The Cochrane Library
Language included	No language restriction	English	English	No language restriction
Number of studies in review	33	21	40	71
Type of studies included	Derivation, validation	Validation (external)	Derivation, validation	Derivation, validation, update and impact
Number of models in review	9	2	11	17 (+7 other types of models*)
Type of models included	Clinical prediction rules	PESI and sPESI	Clinical prediction rules	All prognostic models
Quality appraisal method used	3-point score for cohort studies	Newcastle-Ottawa Scale	QUADAS2 (diagnostic tool) Domain approach	Prognostic criteria Domain approach
Analysis of stable PE patients separately	No	No	No	Yes
Outcome	All-cause mortality, adverse outcomes	All-cause mortality, PE-related mortality, adverse outcomes	All-cause mortality	All-cause mortality, PE-related mortality, adverse outcomes, VTE, major bleeding
Analysis of different outcomes separately	Yes	Yes	Not applicable	Yes
Time point	14 days, 30 days, 3 months	Not reported	In-hospital, 7 days, 30 days, 3 months	In-hospital/7 days, 30 days, 3 months
Analysis of different time points separately	No	No	No	Yes
Risk groups	Low-risk group	Low-risk group vs high-risk group	Low-risk group	Low-risk group and high-risk group
Number of cut-offs accounted for	Single cut-off	Single cut-off	Single cut-off	All available cut-offs (analysis along the risk scale)
Summary estimate in meta-analysis	Event rate in patients at low risk for all models combined	Prognostic accuracy, [†] OR	Prognostic accuracy [†]	Prognostic accuracy, [†] proportion of patients in each risk group, event rates within each risk group (absolute risk), population event rate
Overall summary estimate from combined derivation and validation studies	Yes	No	Yes	Yes
Summary estimate from external validation studies	No	Yes	No	Yes
Analysis of update studies	No	No	No	Yes
Analysis of impact studies	No	No	No	Yes

Continued



Table 6 Continued

Author, year	Squizzato <i>et al</i> , 2012 ¹⁴⁶	Zhou <i>et al</i> , 2012 ¹⁴⁸	Kohn <i>et al</i> , 2015 ¹⁴⁷	Our study
Conclusion of the review	'Prognostic CPRs efficiently identify PE patients at a low risk of mortality'	'PESI has discriminative power to predict the short-term death and adverse outcome events in patients with acute pulmonary embolism, the PESI and the sPESI have similar accuracy'	'Numerous clinical prediction rules for prognosticating early mortality in patients with PE are available, but not all demonstrate the high sensitivity needed to reassure clinicians'	'We provide evidence-based information about the validity and utility of the existing prognostic models in acute PE that may be helpful for identifying patients at low risk. Novel models seem attractive for the high-risk normotensive PE but need to be externally validated than be assessed in impact studies'

*Seven other types of models include five models whose variables are originally not specific to PE or to its prognosis and two hospital criteria checklists.
†Prognostic accuracy (summary sensitivity, summary specificity and summary receiver operating characteristic).
CPR, cardiopulmonary resuscitation; PE, pulmonary embolism; PESI, Pulmonary Embolism Severity Index; sPESI, simplified PESI; VTE, venous thromboembolism.

likely to lead to full publication. To assess publication bias for proportion and event rates, we applied statistical tests to detect funnel plot asymmetry, but we found little evidence of publication bias. According to Macaskill *et al*,¹⁴⁵ 'applying such tests in systematic review of diagnostic test accuracy is likely to result in publication bias being incorrectly indicated by the test far too often'.

Data extraction was performed by two reviewers. The second reviewer was not blinded to the data extracted. To ensure quality, data were rechecked by the first reviewer. We assessed study quality using a domain approach¹⁷ rather than a scale. This is more transparent to the reader, by showing how each study is performed in each aspect of conduct. As shown by Juni *et al*,¹⁹ 'the use of summary scores to identify trials of high quality is problematic and relevant methodological aspects should be assessed individually and their influence on effect sizes explored'.

Our analyses were performed for distinct models, outcomes/time points, with subgroup and sensitivity analyses to deal with heterogeneity. We used multiple statistical techniques, which could introduce bias in the selection of results to present, but no discrepancies were found across the results and all analyses are available. We used absolute risk rather than relative risk because this is more appropriate for estimating the risks for individual patients.

Comparison with other reviews

This study was not limited to only clinical prediction rules and to our knowledge, it is the first broad systematic review and meta-analysis of prognostic models in patients with acute PE. There are two published systematic reviews on clinical prediction rules^{146 147} and a meta-analysis of PESI models,¹⁴⁸ but, as discussed below, our systematic review has important differences to all of these, expands their scope considerably and provides more comprehensive and up-to-date data (table 6).

We believe the key strengths of our review are well-defined selection criteria, broad search strategy, and presentation of results of all available predictive models with full details on study characteristics, population and prognostic information and study quality. We used a domain approach and prognostic criteria for the assessment of study quality. Quantitative analyses are provided to answer each aspect of model development in accordance with the recommendations of recently published guidelines for prognostic model research,¹⁴⁹ for different outcomes and time points in 'both stable and unstable' as well as in 'stable' PE, and at various cut-offs, with the aim of improving management of patients with acute PE who are either at low risk or at high risk. For the analysis of performance, we provide a summary estimate of sensitivity and specificity and also a summary estimate of event rates ('absolute risk'¹⁴⁹) within risk groups, which is more meaningful for clinicians and seems more appropriate for the study of prognosis.

CONCLUSIONS

Our systematic review provides useful information on the prediction ability and the utility of existing prognostic models in acute PE that might help clinicians and researchers for the identification of patients at low risk of events for safe early discharge or outpatient management and those at high risk who may need closer monitoring or more aggressive therapy. It shows the potential for improving the selection of lower risk and higher risk groups in patients with normotensive PE with novel and updated models that integrate biomarkers (cTn, BNP, NT-proBNP, H-FABP), CT or TTE-RVD and US-detected DVT into existing models or with other clinical variables (systolic blood pressure, heart rate). These findings provide a good direction for future research in validation and impact studies.

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REFERENCES

- Torbicki A, Perrier A, Konstantinides S, *et al*. Guidelines on the diagnosis and management of acute pulmonary embolism: the Task Force for the Diagnosis and Management of Acute Pulmonary Embolism of the European Society of Cardiology (ESC). *Eur Heart J* 2008;29:2276–315.
- Jaff MR, McMurtry MS, Archer SL, *et al*. Management of massive and submassive pulmonary embolism, iliofemoral deep vein thrombosis, and chronic thromboembolic pulmonary hypertension: a scientific statement from the American Heart Association. *Circulation* 2011;123:1788–830.
- Konstantinides SV, Torbicki A, Agnelli G, *et al*. 2014 ESC guidelines on the diagnosis and management of acute pulmonary embolism. *Eur Heart J* 2014;35:3033–69, 69a–69k.
- Kearon C, Akl EA, Ornelas J, *et al*. Antithrombotic therapy for VTE disease: CHEST guideline and expert panel report. *Chest* 2016;149:315–52.
- Becattini C, Vedovati MC, Agnelli G. Prognostic value of troponins in acute pulmonary embolism: a meta-analysis. *Circulation* 2007;116:427–33.
- Klok FA, Mos IC, Huisman MV. Brain-type natriuretic peptide levels in the prediction of adverse outcome in patients with pulmonary embolism: a systematic review and meta-analysis. *Am J Respir Crit Care Med* 2008;178:425–30.
- Sanchez O, Trinquart L, Colombet I, *et al*. Prognostic value of right ventricular dysfunction in patients with haemodynamically stable pulmonary embolism: a systematic review. *Eur Heart J* 2008;29:1569–77.
- Jimenez D, Uresandi F, Otero R, *et al*. Troponin-based risk stratification of patients with acute nonmassive pulmonary embolism: systematic review and metaanalysis. *Chest* 2009;136:974–82.
- Lega JC, Lacasse Y, Lakhall L, *et al*. Natriuretic peptides and troponins in pulmonary embolism: a meta-analysis. *Thorax* 2009;64:869–75.
- Altman DG. Prognostic models: a methodological framework and review of models for breast cancer. *Cancer Invest* 2009;27:235–43.
- Mallett S, Royston P, Dutton S, *et al*. Reporting methods in studies developing prognostic models in cancer: a review. *BMC Med* 2010;8:20.
- Mallett S, Royston P, Waters R, *et al*. Reporting performance of prognostic models in cancer: a review. *BMC Med* 2010;8:21.
- Williams C, Brunskill S, Altman D, *et al*. Cost-effectiveness of using prognostic information to select women with breast cancer for adjuvant systemic therapy. *Health Technol Assess* 2006;10:iii–iv, ix–xi, 1–204.
- Altman DG. Systematic reviews of evaluations of prognostic variables. *BMJ* 2001;323:224–8.
- Egger M, Juni P, Bartlett C, *et al*. How important are comprehensive literature searches and the assessment of trial quality in systematic reviews? Empirical study. *Health Technol Assess* 2003;7:1–76.
- Hayden JA, Cote P, Bombardier C. Evaluation of the quality of prognosis studies in systematic reviews. *Ann Intern Med* 2006;144:427–37.
- Higgins JPT, Green SE. *Cochrane handbook for systematic reviews of interventions version 5.1.0 [updated March 2011]*. The Cochrane Collaboration, 2011. <http://www.cochrane-handbook.org>
- Schulz KF, Grimes DA, eds. *Handbook of essential concepts in clinical research*. Elsevier Ltd, 2006.
- Juni P, Witschi A, Bloch R, *et al*. The hazards of scoring the quality of clinical trials for meta-analysis. *JAMA* 1999;282:1054–60.
- Harbord RM, Deeks JJ, Egger M, *et al*. A unification of models for meta-analysis of diagnostic accuracy studies. *Biostatistics* 2007;8:239–51.
- Harbord RM, Whiting P. Metandi: meta-analysis of diagnostic accuracy using hierarchical logistic regression. *Stata J* 2009;2:211–29.
- Harbord R. *Metandi: Stata module to perform meta-analysis of diagnostic accuracy*. Bristol, UK: MRC Health Services Research Collaboration, University of Bristol, 2006.
- Rabe-Hesketh S, Skrondal A, Pickles A. *GLLAMM Manual. U.C. Berkeley Division of Biostatistics Working Paper Series*. Working Paper 160, 2004. <http://www.bepress.com/ucbbiostat/paper160>
- Moher D, Liberati A, Tetzlaff J, *et al*. Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. *PLoS Med* 2009;6:e1000097.
- Al Otair HA, Al-Boukai AA, Ibrahim GF, *et al*. Outcome of pulmonary embolism and clinico-radiological predictors of mortality: experience from a university hospital in Saudi Arabia. *Ann Thorac Med* 2014;9:18–22.
- Almahameed A, Carman TL. Outpatient management of stable acute pulmonary embolism: proposed accelerated pathway for risk stratification. *Am J Med* 2007;120(Suppl 2):S18–25.
- Bircan A, Karadeniz N, Ozden A, *et al*. A simple clinical model composed of ECG, shock index, and arterial blood gas analysis for predicting severe pulmonary embolism. *Clin Appl Thromb Hemost* 2011;17:188–96.
- Bledsoe J, Hamilton D, Bess E, *et al*. Treatment of low-risk pulmonary embolism patients in a chest pain unit. *Crit Pathw Cardiol* 2010;9:212–15.
- Budaj-Fidecka A, Kurzyna M, Fijałkowska A, *et al*. In-hospital major bleeding predicts mortality in patients with pulmonary embolism: an analysis of ZATPOL Registry data. *Int J Cardiol* 2013;168:3543–9.
- Cameron A, Ogilvie C, Teckchandani S, *et al*. Outpatient imaging for pulmonary embolism may only be suitable for a minority. *Scott Med J* 2012;57:14–17.

31. Gallotta G, Palmieri V, Piedimonte V, *et al.* Increased troponin I predicts in-hospital occurrence of hemodynamic instability in patients with sub-massive or non-massive pulmonary embolism independent to clinical, echocardiographic and laboratory information. *Int J Cardiol* 2008;124:351–7.
32. Gul EE, Can I, Guler I, *et al.* Association of pulmonary artery obstruction index with elevated heart-type fatty acid binding protein and short-term mortality in patients with pulmonary embolism at intermediate risk. *Diagn Interv Radiol* 2012;18:531–6.
33. Lankeit M, Dellas C, Benz V, *et al.* The predictive value of heart-type fatty acid-binding protein is independent from symptom duration in normotensive patients with pulmonary embolism. *Thromb Res* 2013;132:543–7.
34. McCabe A, Hassan T, Doyle M, *et al.* Identification of patients with low-risk pulmonary embolism suitable for outpatient treatment using the pulmonary embolism severity index (PESI). *Ir J Med Sci* 2013;182:291–5.
35. Paiva L, Providencia R, Faustino A, *et al.* Manchester triage in acute pulmonary embolism: can it unmask the grand impersonator? *Emerg Med J* 2012;29:e6.
36. Palla A, Ribas C, Rossi G, *et al.* The clinical course of pulmonary embolism patients anticoagulated for 1 year: results of a prospective, observational, cohort study. *J Thromb Haemost* 2010;8:68–74.
37. Rydman R, Söderberg M, Larsen F, *et al.* d-Dimer and simplified pulmonary embolism severity index in relation to right ventricular function. *Am J Emerg Med* 2013;31:482–6.
38. Spirk D, Willenberg T, Aujesky D, *et al.* Use of biomarkers or echocardiography in pulmonary embolism: the Swiss Venous Thromboembolism Registry. *QJM* 2012;105:1163–9.
39. Taylor RA, Davis J, Liu R, *et al.* Point-of-care focused cardiac ultrasound for prediction of pulmonary embolism adverse outcomes. *J Emerg Med* 2013;45:392–9.
40. Tsai J, Abe K, Boulet SL, *et al.* Predictive accuracy of 29-comorbidity index for a diagnosis of venous thromboembolism. *PLoS ONE* 2013;8:e70061.
41. George E, Kumamaru KK, Ghosh N, *et al.* Computed tomography and echocardiography in patients with acute pulmonary embolism: part 2: prognostic value. *J Thorac Imaging* 2014;29:W7–12.
42. Goldhaber SZ, Visani L, De Rosa M. Acute pulmonary embolism: clinical outcomes in the International Cooperative Pulmonary Embolism Registry (ICOPER). *Lancet* 1999;353:1386–9.
43. Kanbay A, Kokturk N, Kaya MG, *et al.* Electrocardiography and Wells scoring in predicting the anatomic severity of pulmonary embolism. *Respir Med* 2007;101:1171–6.
44. Kasper W, Konstantinides S, Geibel A, *et al.* Management strategies and determinants of outcome in acute major pulmonary embolism: results of a multicenter registry. *J Am Coll Cardiol* 1997;30:1165–71.
45. Konstantinides S, Geibel A, Olschewski M, *et al.* Association between thrombolytic treatment and the prognosis of hemodynamically stable patients with major pulmonary embolism: results of a multicenter registry. *Circulation* 1997;96:882–8.
46. Kucher N, Wallmann D, Carone A, *et al.* Incremental prognostic value of troponin I and echocardiography in patients with acute pulmonary embolism. *Eur Heart J* 2003;24:1651–6.
47. Rodríguez-Cerrillo M, Alvarez-Arcaya A, Fernández-Díaz E, *et al.* A prospective study of the management of non-massive pulmonary embolism in the home. *Eur J Intern Med* 2009;20:598–600.
48. Sen HS, Abakay O, Tanrikulu AC, *et al.* Is a complete blood cell count useful in determining the prognosis of pulmonary embolism? *Wien Klin Wochenschr* 2014;126:347–54.
49. Hariharan P, Takayesu JK, Kabrhel C. Erratum to: Association between the Pulmonary Embolism Severity Index (PESI) and short-term clinical deterioration (Thromb Haemost 2011; 105: 706–711). *Thromb Haemost* 2012;107:599–9.
50. Aujesky D, Roy PM, Le Manach CP, *et al.* Validation of a model to predict adverse outcomes in patients with pulmonary embolism. *Eur Heart J* 2006;27:476–81.
51. Zondag W, Hiddinga BI, Crobach MJT, *et al.* Hestia criteria can discriminate high- from low-risk patients with pulmonary embolism. *Eur Respir J* 2013;41:588–92.
52. Zondag W, Mos ICM, Creemers-Schild D, *et al.* Outpatient treatment in patients with acute pulmonary embolism: the Hestia Study. *J Thromb Haemost* 2011;9:1500–7.
53. Barra S, Paiva L, Providencia R, *et al.* LR-PED rule: low risk pulmonary embolism decision rule—a new decision score for low risk pulmonary embolism. *Thromb Res* 2012;130:327–33.
54. Moores L, Zamarro C, Gómez V, *et al.* Changes in PESI scores predict mortality in intermediate-risk patients with acute pulmonary embolism. *Eur Respir J* 2013;41:354–9.
55. Otero R, Trujillo-Santos J, Cayuela A, *et al.* Haemodynamically unstable pulmonary embolism in the RIETE Registry: systolic blood pressure or shock index? *Eur Respir J* 2007;30:1111–16.
56. Stuck AK, Stone RA, Pugh N, *et al.* Comparison of the Pulmonary Embolism Severity Index to a simplified version: classification and clinical outcomes. *Thromb Res* 2014;133:515–16.
57. Dunn AS. ACP Journal Club. The simplified Pulmonary Embolism Severity Index and high-sensitivity troponin T predicted prognosis in PE. *Ann Intern Med* 2012;156:JC4–12.
58. Gould MK. A clinical prediction model was sensitive but not specific for predicting mortality in pulmonary embolism. *ACP J Club* 2006;145:53.
59. O'Flaherty M, Gandara E. A simple prediction rule identified patients with pulmonary embolism at low risk of short term death. *Evid Based Med* 2006;11:121.
60. Golpe R, Perez-de-Llano LA, Castro-Anon O. Prognostic value of the Charlson comorbidity index in pulmonary embolism. *Respiration* 2013;85:438.
61. Ahmad N, Srinivasan K, Moudgil H. Risk stratification in pulmonary embolism: an algorithmic tool approach. *Thorax* 2011;66:1098–9.
62. Chow V, Ng AC, Chung T, *et al.* Right atrial to left atrial area ratio on early echocardiography predicts long-term survival after acute pulmonary embolism. *Cardiovasc Ultrasound* 2013;11:17.
63. Fraga M, Taffe P, Mean M, *et al.* The inter-rater reliability of the Pulmonary Embolism Severity Index. *Thromb Haemost* 2010;104:1258–62.
64. Sequeira JF, Duarte PC, Madruga IM, *et al.* [Pulmonary embolism--mortality risk]. *Rev Port Cardiol* 1992;11:749–57.
65. Agterof MJ, Schutgens RE, Mounli N, *et al.* A prognostic model for short term adverse events in normotensive patients with pulmonary embolism. *Am J Hematol* 2011;86:646–9.
66. Agterof MJ, Schutgens REG, Snijder RJ, *et al.* Out of hospital treatment of acute pulmonary embolism in patients with a low NT-proBNP level. *J Thromb Haemost* 2010;8:1235–41.
67. Agterof MJ, Van Bladel ER, Schutgens REG, *et al.* Risk stratification of patients with pulmonary embolism based on pulse rate and D-dimer concentration. *Thromb Haemost* 2009;102:683–7.
68. Aujesky D, Obrosky DS, Stone RA, *et al.* Derivation and validation of a prognostic model for pulmonary embolism. *Am J Respir Crit Care Med* 2005;172:1041–6.
69. Aujesky D, Obrosky DS, Stone RA, *et al.* A prediction rule to identify low-risk patients with pulmonary embolism. *Arch Intern Med* 2006;166:169–75.
70. Aujesky D, Perrier A, Roy PM, *et al.* Validation of a clinical prognostic model to identify low-risk patients with pulmonary embolism. *J Intern Med* 2007;261:597–604.
71. Aujesky D, Roy PM, Verschuren F, *et al.* Outpatient versus inpatient treatment for patients with acute pulmonary embolism: an international, open-label, randomised, non-inferiority trial. *Lancet* 2011;378:41–8.
72. Becattini C, Casazza F, Forgiione C, *et al.* Acute pulmonary embolism: external validation of an integrated risk stratification model. *Chest* 2013;144:1539–45.
73. Bertoletti L, Le Gal G, Aujesky D, *et al.* Prognostic value of the Geneva prediction rule in patients with pulmonary embolism. *Thromb Res* 2013;132:32–6.
74. Bova C, Pesavento R, Marchiori A, *et al.* Risk stratification and outcomes in hemodynamically stable patients with acute pulmonary embolism: a prospective, multicentre, cohort study with three months of follow-up. *J Thromb Haemost* 2009;7:938–44.
75. Chan CM, Woods C, Shorr AF. The validation and reproducibility of the pulmonary embolism severity index. *J Thromb Haemost* 2010;8:1509–14.
76. Chan CM, Woods CJ, Shorr AF. Comparing the pulmonary embolism severity index and the prognosis in pulmonary embolism scores as risk stratification tools. *J Hosp Med* 2012;7:22–7.
77. Choi KJ, Cha SI, Shin KM, *et al.* Prognostic implications of computed tomographic right ventricular dilation in patients with acute pulmonary embolism. *Thromb Res* 2014;133:182–6.
78. Choi WH, Kwon SU, Jwa YJ, *et al.* The pulmonary embolism severity index in predicting the prognosis of patients with pulmonary embolism. *Korean J Intern Med* 2009;24:123–7.
79. Davies CWH, Wimperis J, Green ES, *et al.* Early discharge of patients with pulmonary embolism: a two-phase observational study. *Eur Respir J* 2007;30:708–14.
80. Dellas C, Tscheppe M, Seeber V, *et al.* A novel H-FABP assay and a fast prognostic score for risk assessment of normotensive pulmonary embolism. *Thromb Haemost* 2014;111:996–1003.

81. Donze J, Le Gal G, Fine MJ, *et al.* Prospective validation of the Pulmonary Embolism Severity Index: a clinical prognostic model for pulmonary embolism. *Thromb Haemost* 2008;100:943–8.
82. Erkens PMG, Gandara E, Wells PS, *et al.* Does the Pulmonary Embolism Severity Index accurately identify low risk patients eligible for outpatient treatment? *Thromb Res* 2012;129:710–14.
83. Ferrer M, Morillo R, Elias T, *et al.* Validation of two clinical prognostic models in patients with acute symptomatic pulmonary embolism. *Arch Bronconeumol* 2013;49:427–31.
84. Hariharan P, Takayesu JK, Kabrhel C. Association between the pulmonary embolism severity index (PESI) and short-term clinical deterioration. *Thromb Haemost* 2011;105:706–11.
85. Jimenez D, Aujesky D, Moores L, *et al.* Simplification of the pulmonary embolism severity index for prognostication in patients with acute symptomatic pulmonary embolism. *Arch Intern Med* 2010;170:1383–9.
86. Jimenez D, Aujesky D, Moores L, *et al.* Combinations of prognostic tools for identification of high-risk normotensive patients with acute symptomatic pulmonary embolism. *Thorax* 2011;66:75–81.
87. Jiménez D, Kopečna D, Tapson V, *et al.* Derivation and validation of multimarker prognostication for normotensive patients with acute symptomatic pulmonary embolism. *Am J Respir Crit Care Med* 2014;189:718–26.
88. Jimenez D, Yusen RD, Otero R, *et al.* Prognostic models for selecting patients with acute pulmonary embolism for initial outpatient therapy. *Chest* 2007;132:24–30.
89. Jo JY, Lee MY, Lee JW, *et al.* Leukocytes and systemic inflammatory response syndrome as prognostic factors in pulmonary embolism patients. *BMC Pulm Med* 2013;13:74.
90. Kostrubiec M, Pruszczyk P, Bochowicz A, *et al.* Biomarker-based risk assessment model in acute pulmonary embolism. *Eur Heart J* 2005;26:2166–72.
91. Kukla P, Bryniarski L, Dlugopolski R, *et al.* Acute pulmonary embolism registry in the Malopolska region—clinical course. *Kardiol Pol* 2009;67:735–41.
92. Lankeit M, Friesen D, Schafer K, *et al.* A simple score for rapid risk assessment of non-high-risk pulmonary embolism. *Clin* 2013;102:73–80.
93. Lankeit M, Gomez V, Wagner C, *et al.* A strategy combining imaging and laboratory biomarkers in comparison with a simplified clinical score for risk stratification of patients with acute pulmonary embolism. *Chest* 2012;141:916–22.
94. Lankeit M, Jimenez D, Kostrubiec M, *et al.* Predictive value of the high-sensitivity troponin T assay and the simplified pulmonary embolism severity index in hemodynamically stable patients with acute pulmonary embolism: a prospective validation study. *Circulation* 2011;124:2716–24.
95. Lankeit M, Jiménez D, Kostrubiec M, *et al.* Validation of N-terminal pro-brain natriuretic peptide cut-off values for risk stratification of pulmonary embolism. *Eur Respir J* 2014;43:1669–77.
96. Moores L, Aujesky D, Jimenez D, *et al.* Pulmonary Embolism Severity Index and troponin testing for the selection of low-risk patients with acute symptomatic pulmonary embolism. *J Thromb Haemost* 2010;8:517–22.
97. Nendaz MR, Bandelier P, Aujesky D, *et al.* Validation of a risk score identifying patients with acute pulmonary embolism, who are at low risk of clinical adverse outcome. *Thromb Haemost* 2004;91:1232–6.
98. Ng AC, Chow V, Yong AS, *et al.* Prognostic impact of the Charlson comorbidity index on mortality following acute pulmonary embolism. *Respiration* 2013;85:408–16.
99. Nordenholz K, Ryan J, Atwood B, *et al.* Pulmonary embolism risk stratification: pulse oximetry and pulmonary embolism severity index. *J Emerg Med* 2011;40:95–102.
100. Otero R, Uresandi F, Jimenez D, *et al.* Home treatment in pulmonary embolism. *Thromb Res* 2010;126:e1–5.
101. Ozsu S, Abul Y, Orem A, *et al.* Predictive value of troponins and simplified pulmonary embolism severity index in patients with normotensive pulmonary embolism. *Multidiscip Respir Med* 2013;8:34.
102. Ozsu S, Ozlu T, Şentürk A, *et al.* Combination and comparison of two models in prognosis of pulmonary embolism: results from Turkey Pulmonary Embolism Group (TUPEG) study. *Thromb Res* 2014;133:1006–10.
103. Paiva LV, Providencia RC, Barra SN, *et al.* Cardiovascular risk assessment of pulmonary embolism with the GRACE risk score. *Am J Cardiol* 2013;111:425–31.
104. Palmieri V, Gallotta G, Rendina D, *et al.* Troponin I and right ventricular dysfunction for risk assessment in patients with nonmassive pulmonary embolism in the Emergency Department in combination with clinically based risk score. *Intern Emerg Med* 2008;3:131–8.
105. Papaioannou AI, Kastanis A, Malli F, *et al.* Predicting short-term mortality in patients with pulmonary embolism: a simple model. *Am J Med Sci* 2013;345:462–9.
106. Righini M, Roy PM, Meyer G, *et al.* The Simplified Pulmonary Embolism Severity Index (PESI): validation of a clinical prognostic model for pulmonary embolism. *J Thromb Haemost* 2011;9:2115–17.
107. Roca IC, Aursulesei V, Roca M, *et al.* [Comparative evaluation of prognostic scores in pulmonary embolism]. *Rev Med Chir Soc Med Nat Iasi* 2012;116:69–74.
108. Sam A, Sanchez D, Gomez V, *et al.* The shock index and the simplified PESI for identification of low-risk patients with acute pulmonary embolism. *Eur Respir J* 2011;37:762–6.
109. Sanchez D, De Miguel J, Sam A, *et al.* The effects of cause of death classification on prognostic assessment of patients with pulmonary embolism. *J Thromb Haemost* 2011;9:2201–7.
110. Sanchez O, Trinquart L, Caille V, *et al.* Prognostic factors for pulmonary embolism: the PREP study, a prospective multicenter cohort study. *Am J Respir Crit Care Med* 2010;181:168–73.
111. Sanchez O, Trinquart L, Planquette B, *et al.* Echocardiography and pulmonary embolism severity index have independent prognostic roles in pulmonary embolism. *Eur Respir J* 2013;42:681–8.
112. Singanayagam A, Chalmers JD, Scally C, *et al.* Right ventricular dilation on CT pulmonary angiogram independently predicts mortality in pulmonary embolism. *Respir Med* 2010;104:1057–62.
113. Singanayagam A, Scally C, Al-Khairalla MZ, *et al.* Are biomarkers additive to pulmonary embolism severity index for severity assessment in normotensive patients with acute pulmonary embolism? *QJM* 2011;104:125–31.
114. Soares TH, de Bastos M, de Carvalho BV, *et al.* Prognostic value of computed tomographic pulmonary angiography and the pulmonary embolism severity index in patients with acute pulmonary embolism. *Blood Coagul Fibrinolysis* 2013;24:64–70.
115. Spirk D, Aujesky D, Husmann M, *et al.* Cardiac troponin testing and the simplified pulmonary embolism severity index the Swiss venous thromboembolism registry (SWIVTER). *Thromb Haemost* 2011;106:978–84.
116. Subramaniam RM, Mandrekar J, Blair D, *et al.* The Geneva prognostic score and mortality in patients diagnosed with pulmonary embolism by CT pulmonary angiogram. *J Med Imaging Radiat Oncol* 2009;53:361–5.
117. Toosi MS, Merlino JD, Leeper KV. Prognostic value of the shock index along with transthoracic echocardiography in risk stratification of Patients With Acute Pulmonary Embolism. *Am J Cardiol* 2008;101:700–5.
118. Uresandi F, Otero R, Cayuela A, *et al.* A clinical prediction rule for identifying short-term risk of adverse events in patients with pulmonary thromboembolism. [Spanish]. *Arch Bronconeumol* 2007;43:617–22.
119. Vanni S, Nazerian P, Pepe G, *et al.* Comparison of two prognostic models for acute pulmonary embolism: clinical vs. right ventricular dysfunction-guided approach. *J Thromb Haemost* 2011;9:1916–23.
120. Venetz C, Jimenez D, Mean M, *et al.* A comparison of the original and simplified pulmonary embolism severity index. *Thromb Haemost* 2011;106:423–8.
121. Volschan A, Albuquerque D, Tura BR, *et al.* Predictors of hospital mortality in hemodynamically stable patients with pulmonary embolism. [Portuguese, Spanish, English]. *Arq Bras Cardiol* 2009;93:135–40.
122. Wicki J, Perrier A, Perneger TV, *et al.* Predicting adverse outcome in patients with acute pulmonary embolism: q risk score. *Thromb Haemost* 2000;84:548–52.
123. Yamaki T, Nozaki M, Sakurai H, *et al.* Presence of lower limb deep vein thrombosis and prognosis in patients with symptomatic pulmonary embolism: preliminary report. *Eur J Vasc Endovasc Surg* 2009;37:225–31.
124. Zhu L, Wang C, Yang YH, *et al.* [Prognostic value of right ventricular dysfunction and derivation of a prognostic model for patients with acute pulmonary thromboembolism]. *Zhonghua Liu Xing Bing Xue Za Zhi* 2009;30:184–8.
125. Zondag W, den Exter PL, Crobach MJT, *et al.* Comparison of two methods for selection of out of hospital treatment in patients with acute pulmonary embolism. *Thromb Haemost* 2013;109:47–52.
126. Zondag W, Vingerhoets LMA, Durian MF, *et al.* Hestia criteria can safely select patients with pulmonary embolism for outpatient treatment irrespective of right ventricular function. *J Thromb Haemost* 2013;11:686–92.
127. Barra SNC, Paiva LV, Providência R, *et al.* Atrial fibrillation in acute pulmonary embolism: prognostic considerations. *Emerg Med J* 2014;31:308–12.



128. Beer JH, Burger M, Gretener S, *et al.* Outpatient treatment of pulmonary embolism is feasible and safe in a substantial proportion of patients. *J Thromb Haemost* 2003;1:186–7.
129. Huang CM, Lin YC, Lin YJ, *et al.* Risk stratification and clinical outcomes in patients with acute pulmonary embolism. *Clin Biochem* 2011;44:1110–15.
130. Jakobsson C, Jimenez D, Gomez V, *et al.* Validation of a clinical algorithm to identify low-risk patients with pulmonary embolism. *J Thromb Haemost* 2010;8:1242–7.
131. Labyk A, Ciurzynski M, Jankowski K, *et al.* Acute pulmonary embolism: analysis of consecutive 353 patients hospitalised in a single centre. A 3-year experience. *Kardiologia polska* 2012;70:15–22.
132. Meyer G, Vicaut E, Danays T, *et al.* Fibrinolysis for patients with intermediate-risk pulmonary embolism. *N Engl J Med* 2014;370:1402–11.
133. Stamm JA, Long JL, Kirchner HL, *et al.* Risk Stratification in acute pulmonary embolism: frequency and impact on treatment decisions and outcomes. *South Med J* 2014;107:72–8.
134. Vuilleumier N, Le Gal G, Cornily JC, *et al.* Is N-terminal pro-brain natriuretic peptide superior to clinical scores for risk stratification in non-massive pulmonary embolism? *J Thromb Haemost* 2010;8:1433–5.
135. Bova C, Sanchez O, Prandoni P, *et al.* Identification of intermediate-risk patients with acute symptomatic pulmonary embolism. *Eur Respir J* 2014;44:694–703.
136. Perrier A, Howarth N, Didier D, *et al.* Performance of helical computed tomography in unselected outpatients with suspected pulmonary embolism. *Ann Intern Med* 2001;135:88–97.
137. Perrier A, Roy PM, Aujesky D, *et al.* Diagnosing pulmonary embolism in outpatients with clinical assessment, D-dimer measurement, venous ultrasound, and helical computed tomography: a multicenter management study. *Am J Med* 2004;116:291–9.
138. Perrier A, Roy PM, Sanchez O, *et al.* Multidetector-row computed tomography in suspected pulmonary embolism. *N Engl J Med* 2005;352:1760–8.
139. Righini M, Le Gal G, Aujesky D, *et al.* Diagnosis of pulmonary embolism by multidetector CT alone or combined with venous ultrasonography of the leg: a randomised non-inferiority trial. *Lancet* 2008;371:1343–52.
140. Roy PM, Meyer G, Vielle B, *et al.* Appropriateness of diagnostic management and outcomes of suspected pulmonary embolism. *Ann Intern Med* 2006;144:157–64.
141. Vuilleumier N, Le Gal G, Verschuren F, *et al.* Cardiac biomarkers for risk stratification in non-massive pulmonary embolism: a multicenter prospective study. *J Thromb Haemost* 2009;7:391–8.
142. Pencina MJ, D'Agostino RB Sr, D'Agostino RB Jr, *et al.* Evaluating the added predictive ability of a new marker: from area under the ROC curve to reclassification and beyond. *Stat Med* 2008;27:157–72; discussion 207–12.
143. McGinn TG, Guyatt GH, Wyer PC, *et al.* Users' guides to the medical literature: XXII: how to use articles about clinical decision rules. Evidence-Based Medicine Working Group. *JAMA* 2000;284:79–84.
144. Vinson DR, Zehtabchi S, Yealy DM. Can selected patients with newly diagnosed pulmonary embolism be safely treated without hospitalization? A systematic review. *Ann Emerg Med* 2012;60:651–62.e4.
145. Macaskill P, Gatsonis C, Deeks JJ, *et al.* Chapter 10: analysing and presenting results. In: Deeks JJ, Bossuyt PM, Gatsonis C, eds. *Cochrane handbook for systematic reviews of diagnostic test accuracy version 1.0*. The Cochrane Collaboration, 2010. <http://srdta.cochrane.org/>
146. Squizzato A, Donadini MP, Galli L, *et al.* Prognostic clinical prediction rules to identify a low-risk pulmonary embolism: a systematic review and meta-analysis. *J Thromb Haemost* 2012;10:1276–90.
147. Kohn CG, Mearns ES, Parker MW, *et al.* Prognostic accuracy of clinical prediction rules for early post-pulmonary embolism all-cause mortality: a bivariate meta-analysis. *Chest* 2015;147:1043–62.
148. Zhou XY, Ben SQ, Chen HL, *et al.* The prognostic value of pulmonary embolism severity index in acute pulmonary embolism: a meta-analysis. *Respir Res* 2012;13:111.
149. Steyerberg EW, Moons KG, van der Windt DA, *et al.* Prognosis Research Strategy (PROGRESS) 3: prognostic model research. *PLoS Med* 2013;10:e1001381.

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