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RELATIONSHIPS WITHIN AND BETWEEN THE 1960 STANFORD-BINET L-M AND THE GOODENOUGH INTELLIGENCE TEST WITH INTELLECTUALLY

SUBAVERAGE CHILDREN

A Thesis

Submitted to

The Faculty of the Department of Psychology

The University of Omaha

In Partial Fulfillment

of the Requirement for the Degree

Master of Arts

By

Frederick M. Rudie

December 1963

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F.M.R.

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CHAPTER I

INTRODUCTION AND BACKGROUND

The examination of relationships within and between the 1960 Stanford-Binet and the Goodenough Intelligence Test provided the basis for this thesis. Before considering the relationships, some background on the testing movement, Binet Scales, Stanford-Binet, Goodenough Draw-A-Man Test, Stanford-Binet vocabulary subtest, Stanford-Binet scatter, and the slow learner would prove useful.

Testing Movement. While it is somewhat arbitrary to associate a person and date with the founding of psychology, it appears equally arbitrary to assign a name and date to the beginning of intelligence testing. For convenience however, it is necessary to select some person to represent the beginning of the new science and the beginning of movements within the science. Wundt is generally designated as the "father" of psychology mainly because he was the first man in the history of psychology to be totally committed to the new subject. Other likely candidates such as Helmholtz and Fechner contributed to psychology but were more accurately classified as physiologists, physicists and philosophers. Analogously, Alfred Binet could be considered the "founder" of intelligence testing. Certainly others such as Galton and James Mc-Keen Cattell contributed to the field, but Binet introduced individual intelligence testing as it is in its present form (Boring, 1950, p. 573).

<u>Binet Scales.</u> Just prior to the turn of the century, Binet began his work on tests of intelligence or, more properly, subtests or items. With characteristic thoroughness he investigated a wide range of areas including graphology and palmistry (Peterson, 1925, p. 160-161). In 1904 a commission in Paris decided that intellectually subaverage children should be removed from regular schools and given instruction in special schools. It was to meet the obvious need to separate the average from subaverage that Binet, with Simon, set about constructing his first intelligence scale. The scale was published in 1905. Three years later, a second scale appeared. Peterson (1925, p. 196) argues that the 1908 revision should be considered the first real scale of intelligence since the 1905 test was used to demonstrate that a test could be constructed rather than as a test itself.

Fortunately, the 1908 scale was published with much greater accuracy and detail. Interest in the new test grew. Translations by Goddard and Huey helped promote it in America. Kuhlmann condensed and modified the English version and at the same time included in it the various comments Binet and Simon had made regarding administration and scoring.

In 1911 the test was again revised by Binet. In the same year Terman published his impressions of the 1908 revision. Though he found the need for radical changes, particularly at the upper and lower ends of the scale, Terman felt the test could be of great value. His interest and appreciation of Binet's method set the stage and form for the testing movement in America. In 1916 Terman published his revision of the Binet-Simon scales. The test, called the Stanford-Binet (S-B), became the accepted criterion for measuring the intelligence of children. Indeed, the Stanford-Binet and its subsequent revisions remain to this day unchallenged as an intelligence test for children.

The 1916 revision was standardized on a sample of about 1,000 children and 400 adults. An attempt was made to obtain a representative

sample of the general population. While providing a good measure of intelligence, the scale was perhaps equally important in establishing sampling procedures which were refined and used in the construction and restandardization of the 1937 scale (Terman & Merrill, 1960, pp. 5-6).

The 1937 revision required nearly ten years of research. The literature was surveyed for evaluations of the 1916 Binet and for test items which could be used in the new S-B. Thousands of items were tried out on 1,000 children who had earlier been given the 1916 revision. In addition 500 preschool children were tested. Items with high discriminating power were selected for further evaluation. Two forms of the test, form L containing 209 items, and form M containing 199 items, were used in the final standardization. Item validity was checked by an increase in percent-passing for successive ages and biserial correlation of each item with the total score. The sample composed of 3,184 native born white subjects tested in 17 communities in 11 widely separated states. There were approximately 100 subjects for each half-year interval from 1 1/2 to 5 1/2 years, 200 subjects at each age for years 6 through 14, and 100 at each age from 15 to 18. The final forms of L and M contained 129 items each. The mean of the scale was slightly above 100, and the S.D.'s for the age levels varied around a median value of approximately 16.

The 1960 revision of the S-B was not a restandardization but an attempt to improve and modernize its predecessor. Only the most discriminating items of forms L and M were included and many of these had to be relocated. The items were chosen and relocated on the basis

of performance of 4,498 subjects ages 2 1/2 to 18. The subjects were from 6 states of the Northeast, Midwest, and West coast.

<u>Goodenough Draw-A-Man Test</u>. The Stanford-Binet had a number of disadvantages. It was mainly a verbal test, it had to be administered individually, it required a fair amount of training to administer, and it required at least 40 minutes to administer and frequently longer. Thus psychologists continued to investigate other means of assessing intelligence. One such investigator was Florence Goodenough. In 1926 she published the Goodenough Draw-A-Man Intelligence Test (DAM Test). The test was characterized by Terman (Goodenough, 1926, p.x) as requiring only a child's drawing of a man, being nonverbal, taking but 10 minutes to test an entire class plus a few minutes scoring time per child, particularly useful between mental ages of 4 to 10, having a reliability coefficient for an unselected age group of between .80 and .90 and giving an average correlation of .76 with the S-B for separate age groups.

According to Goodenough (1926, pp. 1-11) the idea of using the drawings of children to study their development was not a new one. In 1885 Ebenezer Cooke noted successive stages in development in children's drawings. Two years later Corrado Ricci, working with Italian children, published an account of their drawings. The attention Cooke's work drew an increased interest in child study and stimulated a great deal of research which reached a peak between 1900 and 1915. The studies, which contained 2 international undertakings, included collecting thousands of drawings from different children in various schools and observing the drawings of individual children in a biographical fashion as they progressed from one age to the next. From this research Goodenough (1926, pp. 12-13) drew a number of conclusions which appeared to be the rationale for her test. They are as follows:

1. In young children a close relationship is apparent between concept development as shown in drawing, and general intelligence.

2. Drawing, to the child, is primarily a language, a form of expression, rather than a means of creating beauty.

3. In the beginning the child draws what he knows, rather than what he sees (Verworn's "ideoplastic stage"). Later on he reaches a stage in which he attempts to draw objects as he sees them. The transition from the first stage to the second one is a gradual and continuous process.

4. The ideoplastic basis of children's drawings is shown most conspicuously in the relative proportions given to the separate parts. The child exaggerates the size of items which seem interesting or important; other parts are minimized or omitted.

5. The order of development in drawing is remarkably constant, even among children of very different social antecedents. The reports of investigators the world over show very close agreement, both as regards the method of indicating the separate items in a drawing and the order in which these items tend to appear. This is especially true as regards the human figure, probably because of its universal familiarity.

6. The earliest drawings made by children consist almost entirely of what may be described as a graphic enumeration of items. Ideas of number, of the relative proportions of parts, and of spatial relationships are much later in developing.

7. In drawing objects placed before them young children pay little or no attention to the model. Their drawings from the object are not likely to differ in any important respect from their memory drawings.

8. Drawings made by subnormal children resemble those of younger normal children in their lack of detail and in their defective sense of proportion. They often show qualitative differences, however, especially as regards the relationship of the separate parts to each other. Not frequently the same drawing will be found to combine very primitive with rather mature characteristics.

9. Children of inferior mental ability sometimes copy well, but they rarely do good original work in drawing. Conversely, the child who shows real creative ability in art is likely to rank high in general mental ability.

10. There is much disagreement among investigators regarding the relationship between children's drawings and those made by primitive or prehistoric races. Until more careful study has been made of the many factors involved in such comparison, the legitimacy of drawing conclusions appears to be very doubtful.

11. Marked sex differences, usually in favor of the boys, are reported by several investigators, especially by Kerschensteiner and Ivanoff.

12.. Up to about the age of ten years children draw the human figure in preference to any other subject. (Goodenough, 1926, pp. 12-13)

It was Goodenough's hope that each child should be allowed to choose the subject he wished to draw. The plan had to be abandoned however since the relative difficulty of the various subjects presented seemingly insurmountable problems in scoring. Thus the subject matter had to be selected. The human figure was chosen since it was equally familiar to all children. It was also simple enough for the very young children and complex enough to challenge the adult. Moreover, it had universal appeal and varied little in essential characteristics. It was further decided that the subject matter should be restricted to "a man" since the clothing of men show greater uniformity than that of women or children.

In 1920 Goodenough obtained almost 4,000 drawings from New Jersey kindergarten through fourth grade children. One hundred drawings were selected for a preliminary analysis. Characteristic differences between the drawings of younger and older children were noted. In this manner an initial scale of 40 points was devised. The point of item validity was established by an increase in percent passing with successive ages. The first scale showed some obvious defects. More items were added and changes in scoring were made. The drawings were rescored and the resulting curves plotted. Then another set of drawings were scored, curves plotted and more changes in scoring were made. Five such revisions were necessary before the present form, containing 51 items, was developed. The final standardization was based on drawings from 3,593 children ranging in age from 4 to 10 years. A vast majority of the children were not of American white parentage, but rather Southern European and Negro descent.

With completion of the final revision, the problems of reliability and validity were again considered. Reliability was checked by the testretest and split-half methods. A correlation of .937 was found for 194 first graders with a one day interval between test and retest. The splithalf correlation was .77 for the separate ages 5 to 10 years. The validity of the test was established in 2 main ways: (1) by an increase in the percentage of children passing a point with successive ages as has already been indicated, and (2) by correlations with grade placement and other test scores. The correlations between the Stanford-Binet mental ages and the Draw-A-Man mental ages will be given in Chapter 2.

<u>Stanford-Binet Vocabulary</u>. Vocabulary has long been considered an important aspect of intelligence. The 1905 Binet-Simon Scale included an item which required defining abstract terms. The 1908 revision again required the definition of abstract words (age 11, item 4), (Freeman, 1962, pp. 188-191). Terman's 1916 revision placed an increased emphasis on the importance of vocabulary. Items calling for the definitions of words were found at ages 5, 8, 10, 12, 14, average adult and superior adult (Freeman, 1962, pp. 201-203). Quite naturally the vocabulary subtest remained an important part of both the 1937 and 1960 Stanford-

Binet revisions.

McNemar (1942, p. 140), after indicating several product-moment correlations between the number of words passed and the composite MA, stated "the magnitude of these correlations indicate that the vocabulary test alone contributes a good rough measure of intelligence." At another point in the same book McNemar (1942, p. 151) said, "the vocabulary test alone yields a fairly adequate measure of the kind of intelligence measured by the New Revision" (New Revision referred to the 1937 revision). These quotations lead to two obvious conclusions. First, that there would be some justification in using the vocabulary subtest as a gross screener. Secondly, it would appear that the Stanford-Binet is highly loaded with verbal items.

<u>Stanford-Binet Scatter</u>. Another aspect of the Stanford-Binet is the scatter of performance shown by individuals. Scatter may be defined as the difference between the basal and maximal; the difference is expressed in months. Basal refers to "....that level at which all tests are passed which just precedes the level where failure occurs" (Terman & Merrill, 1960, p. 60). The maximal or ceiling level refers to the first level at which all tests are failed beyond the last level in which success occurred. While Terman and Merrill (1960, pp. 59-60) reject the idea that diagnostic significance can be attached to scatter, they recognize uneven manifestations of intelligence in individuals. Thus scatter is likely to be the result of individual patterns of abilities. For example, a particular youngster might have exceptional verbal abilities and therefore score relatively high on the vocabulary subtest. Since the vocabulary test occurs at a number of age levels, it will likely increase the extent of his scatter.

<u>Slow Learner</u>. Classification systems are needed to make interpretations of IQ possible. Two well known classifications are provided by Terman and Merrill (1937; 1960) and Wechsler (1949; 1958). It is not the intent of the present investigator to present the classifications, but rather to discuss the group of individuals whose IQ's fall approximately between 70 and 84. The American Association on Mental Deficiency (1959) has termed the level of intelligence in this area as "borderline." This investigation is concerned with a broader range of intelligence than the slow learner, i.e., with Binet IQ's extending below 70 into the 60's. Nevertheless, the majority of subjects fell into the slow learner class and therefore a discussion of the slow learner would appear appropriate.

Unlike some forms of mental deficiency such as mongolism, the slow learner cannot be distinguished by merely looking at him (Johnson, 1963, pp. 30-32). Eut, as a group, slow learners are slightly below the average in height, weight, and motor abilities although a thorough physical examination will indicate they are "normal." This is somewhat contrary to the popular concept of the retarded being"ällbrawn and no brain."

IQ"merely states that a person's intelligence at any given time is defined by his relative standing among his age peers" (Wechsler, 1958, p.33). With this in mind, the slow learner may be seen as an individual whose relative intellectual standing is between the first and second standard deviation of the lower half of the intelligence distribution. Depending on what test is used, they represent about 14% of the population with a percentile rank ranging from 2 or 3 to about 16 (Wechsler, 1949, p. 15). That is, from 84 to 98 percent of their age peers will

be brighter in varying degrees than they. This intellectual retardation strongly affects the slow learner's educational development.

From a theoretical point of view a child with an IQ of 75 would not have the minimum mental age for reading readiness until he is 8 years old. Theoretically, he would be capable of learning many "basic skills," but some would always be beyond his grasp. Theoretically, he must be taught at a slower pace since each year his mental age drops further behind the average. For example, at age 6 he will be 18 months below the average in mental age but at age 15 he will be 45 months behind. However, the problem is neither this simple nor this clear-cut. The theoretical point of view overlooks important factors such as interest, aptitude, adjustment, temperament, motivation, and study habits. In real life prediction of academic performance for the individual remains very difficult.

There is no evidence to indicate that the slow learner's emotional and social characteristics differ from the average. The possible exception to this rule is that slower learner's interests are more constricted than the average or bright. Slower learners have the same basic needs, wants, and desires as other children. An example of some of these needs are: a need to belong, to be part of the group, and to be accepted by the group; a need for the feeling of self worth, and the need for love, attention, affection, and understanding.

Consider for a moment the plight of the slow learner in an average school. Does he belong to a group? Does he feel he contributes to class projects? Is he reinforced or rewarded for his academic efforts?

Can he identify with a group? Can he take pride in his work in the light of external criterion? Are teachers as positive in their attention, affection, and understanding as they are to the average or bright? Usually the answer to these questions is a resounding no. Frequently the slow learner is older and larger than his classmates due to failure to be promoted. For this reason he may stand out or apart from the group. Other children who are promoted yearly maintain their established friendships whereas the slow learner must look for new friends each time he fails to be promoted. The almost inevitable failures in promotion are likely to damage his already weak self-esteem. His daily work is often regarded as inferior and graded as such by the teacher. His contributions to class projects are met with ridicule rather than praise. The attention he gets is rarely for his good efforts but rather for his misbehavior. Teachers become annoyed with him because he can not "keep up" with the class. Tolerance, not understanding, is all that he gets. Frequently, he becomes apathetic rather than interested in learning and all too often becomes the playground bully or troublemaker. Viewed by the school faculty he becomes a discipline problem.

The rather gloomy picture painted above is not true for all slow learners, at least to the degree indicated. To be sure, there are many happy, well adjusted slow learners. Nevertheless, there is a disproportionately large number of problem children in the slow learner group (Johnson, 1963, p. 48) and consequently special attention has been given to this problem in this paper.

In summary, slow learners are normal in physical appearance but probably slightly below average in size and motor ability. The slow learner is an individual whose relative intellectual standing ranges from the 2nd and 3rd to about the 16th percentile. Quite naturally

their intellectual retardation has a severe effect on their academic proficiency. If forced to compete with average students, they almost inevitably fail. The constant failure frequently leads to varying degrees of emotional and behavioral problems.

The Problem of This Study. The major problem of this study revolves around the efficacy, value, or merit of using the Goodenough Draw-A-Man Intelligence Test as an estimate of intelligence with a subaverage population. To test this problem it was necessary to assume the Binet could act as a criterion of intelligence. Phrased another way, the S-B became the external, independent, objective, observable, referent to which the DAM could be compared. Without this assumption the statistical techniques, correlations, and tests of mean differences between the Binet and Goodenough would have been meaningless.

A second problem in this study was to examine the relationship between the Binet vocabulary subtest and the total Binet and between the Binet vocabulary and Goodenough test scores. A third problem was to investigate the relationship between test scatter on the S-B and the difference in scores made on the Binet and Goodenough tests. As in the first, the second and third problems were checked with a subaverage population.

CHAPTER II

REVIEW OF THE LITERATURE

Relationship of Goodenough Draw-A-Man Test to the Stanford-Binet. As was mentioned earlier, part of Goodenough's attempt to validate her test was done by correlating it with the 1916 Stanford-Binet. Her subjects were 334 children, ages 4 through 10 years. The correlations between the MA of the tests for each age are as follows: age 4, .863; age 5, .699; age 6, .832; age 7, .716; age 8 .557; age 9 .728; and age 10, 1849. The overall correlation for the IQ's of the group was .741. Tests for mean differences between the Binet and Goodenough were conspicuous by their absence.

Yepsen (1929, pp. 448-451) in an attempt to determine the reliability of the Goodenough test used 37 feeble-minded subjects from the Vineland training school. The subjects were all boys and ranged in ages from 9.0 to 18.2 years. The test was administered 3 times with a 4 day interval between test and retest. The testing was carried out in accord with Goodenough's instructions. The drawings were scored and rescored to eliminate errors. The resulting test scores (MAs) were correlated. The correlation between the first and second administration was .89, between the second and third .91, and between the first and third .91. Binet MA scores were also correlated with the DAM test. The resulting correlation was .60. It was concluded that the test "appears to measure something not entirely covered by the Binet" (Yepson, 1929, p. 451). No tests for mean differences were made.

McElwee (1932, pp. 217-218) used 45 subnormal 14 year olds as subjects in her study. The children, from ungraded classes of New York City, were given the Goodenough and Binet tests at the same time. It was found that the Binet median mental age was 8-0, while that of the Goodenough was 7-3. Whether the 9 month difference is significant was not pointed out. The product-moment correlation between the MA of the two tests was $.717 \stackrel{+}{=} .048$. It was concluded that the test was as equally satisfactory for subnormal children over 12 years as with younger children.

Earl (1933, pp. 305-327) began his study with 420 drawings from mental defective patients. Three hundred and seven subjects were eliminated from the sample since they did not fall between the ages of 16 and 40 years and/or their Binet mental ages did not lie between 5 and 9. Others were eliminated, for example, because of clinical psychosis, speech defects and physical disabilities. The final sample included 113 mentally defective subjects who passed the above criterion.

The Goodenough was usually given after the Einet test. Earl varied the instructions with the mental age of the subject and used 7 3/8 by 5 3/8 paper. When the 2 tests were correlated, a coefficient of .48 - .07 was found. Unfortunately Earl failed to say whether this correlation was for MA or IQ. Moreover, no information was given as to the sex of the subjects, their CAs except for range, their average IQs or MAs on the Binet or Goodenough, means of scoring the Goodenough and so on.

In an attempt to throw further light on scoring and test reliability Williams (1935, pp. 653-656) had 5 upper division education students independently score the drawings of 100 children. The subjects ranged in chronological age from 3 to 15 years. Einet tests were also given the children and it was determined that their mental ages ranged from 4 to 12.

It should be noted that the chronological age range is far beyond that suggested by Goodenough for use of the test. The scores of the 5 raters were intercorrelated. The resulting intercorrelations ranged from .80 to .96. Correlations of .90 to .97 were obtained between the totals of separate raters and mid-scores of rating.

The mid-scores of the rater and chronological age were also correlated. The resulting coefficient was .491 \pm .051. A relatively high coefficient of .801 \pm .024 was obtained between the mid-scores and the Binet mental age while a coefficient of .651 \pm .038 resulted in correlating Goodenough IQs with Binet IQs.

Williams found no sex differences on the Goodenough when considered in relation of Binet mental age. Williams concluded that relatively inexperienced persons could reliably score Goodenough drawings but that a brief period of supervision is advisable before independent rating is undertaken. Further, it was felt that this study gave added support to the Goodenough Test's validity and reliability.

In a series of studies carried out by McHugh (1943, 1945a, 1945b) the relationship between the 1937 Stanford-Binet and Goodenough test was examined. The subjects were 90 public school kindergarten children, 43 of which were boys and 47 were girls. All the children were initially tested with both tests during the 2 weeks prior to the beginning of school. They were retested shortly after school began (mean of 30.2, SD \pm 12.2 school day). The average age at the time of the second test was 64 months, SD \pm 3.97 months. The Goodenough test was given twice in succession after the subjects had finished the Binet. This necessitated modifying Goodenough's instructions slightly. The drawings were scored and rescored by others to check for accuracy. The highest of the 2 test scores were selected for final analysis. The relationship between the 2 tests was checked by way of correlations. The resulting correlation between the Goodenough and Binet MAs was .45, PE $\stackrel{+}{}$.06. Using the Goodenough and Binet IQs a correlation of .41, PE $\stackrel{+}{}$.06 was obtained. McHugh pointed out that correlations were probably somewhat depressed since half the subjects were given form L and half of the subjects were given form M of the S-B. Biserial correlations were also computed for the class B items of the Goodenough with the Binet IQs. It was found that only 30 of the 51 items yielded a positive correlated. The remaining 21 were either 0 or slightly negatively correlated. The highest correlation with the Binet was obtained by using 9 items with a biserial correlation of .30 or better. Tests for significant differences were not made.

As part of a control group for the study of Indian children, Havighurst, Gunther, and Pratt (1946) used white children from a small mid-western city. Fifty-eight of the 66 subjects were 10 years old while 8 were between 11 and 11 years 3 months. The size of the town was small and the 58 10 year olds represented nearly all the children in that age group. The 66 subjects included 28 boys and 38 girls. Among other tests, the Stanford-Binet and the Draw-A-Man were given. The IQs of the 2 tests were correlated and yielded a coefficient of $.50 \stackrel{+}{=} .06$. Apparently because the study was concerned mainly with Indian children, little other information was given concerning the white group.

Birch (1949, pp. 218-224) investigated the relationship between the Goodenough test and the 1937 Binet with borderline and mental defectives. The life age of the subjects ranged from 10 years 6 months to 16 years 3 months. This is somewhat beyond the age Goodenough (1926) intended the test to go. It should be remembered, however, that Goodenough was referring to mental age. Forty-three boys and 25 girls all having Binet IQs of 70 or less composed the 68 subjects in this sample.

As was true of other studies, care was taken to check the scoring of the drawings by someone other than the examiners. Using productmoment correlations, a correlation of .69 was found between the Binet and Goodenough MAs, .62 between Binet and Goodenough IQs, .37 between CA and Binet MA, .38 between CA and Goodenough MAs, and .64 between the Binet and Goodenough MAs after CA had been partialed out.

Birch (1949) noted that the mean DAM MA of 85.8 was significantly higher (at the 5% level) than the S-B MA mean of 80.5. The standard deviation of the Goodenough was also significantly (1% level) larger than the Binet. Birch attributed the significant differences in both factors to the truncated Binet scores, that is to the fact that no Binet scores above 70 were included in the sample. Birch concluded that the Goodenough test is a valid measure of mental ability for children ages 10 years 6 months to 16 years 3 months with Einet IQ's of 70 or below.

Using children from the Dixon State Hospital, Johnson, Ellard, and Lahey (1950) obtained a correlation of .48 between the Binet and Goodenough. It was felt the correlation was quite high considering the sample included feeblemindedness, epilepsy, post-encephalitis, and brain damage. West in a study with 48 4th and 5th grade children in 1960 obtained a correlation of .45 between the Goodenough and Binet. (West made no mention of which revision of the Binet was used so it was assumed to be the 1937).

In a recent study Rohrs and Haworth (1962) examined the relationship between the Goodenough and Stanford-Binet. The subjects were 46 mental defectives. The correlation (.28) between the IQs of the 2 tests was nonsignificant. In discussing the nonsignificant correlation the authors pointed out that many of the performance items of the 1937 Binet have been omitted in the 1960 revision. This might, in part, account for their results. No significant difference was found between the mean DAM IQ of 56.46 and mean S-B IQ of 56.91.

With the exception of Goodenough's standardization, the studies presented were concerned mainly with an overall relationship rather than the degree of association for separate age groups. Presenting an overall correlation when multiple age groups have been used has distinct drawbacks. The ages which contribute the most and least to the relationship remain unknown. For example, it is possible that in the lower age level there is no correlation between the 2 tests, but in the upper age levels the correlation is high. An overall correlation will reflect a compromise of the 2 extremes rather than the true picture. Thus the ages which show the greatest correlation are not known and the overall correlation is in question. Equally serious is the possibility of a high correlation existing between the means at the various age levels while within the age groups the correlation may be low. In such a case the total correlation would be spuriously high and again not reflect the true nature of the relationship.

A second aspect of the same problem is to determine whether the scores rendered by the S-B and DAM are equal. Birch (1949) found the mean Goodenough MAs were a significant 5.3 months higher than the Binet's while Rohrs and Haworth (1962) found no difference in IQ means in their sample. Here again there are advantages in considering each age separately. It is possible for example, that at the lower age levels the mean Goodenough score would be significantly higher than the Binet; in the middle range they could be equal; and in the upper age levels the Goodenough would be lower. In such a case, a test with overall means might find no significant differences. This thesis, in large part, considers the problem of significant or nonsignificant correlations and mean differences between Binet and Goodenough scores. (MA).

<u>Stanford-Binet Vocabulary</u>. In 1918 Terman made a searing attack against criticism of the vocabulary mental test. The offensive was based on research with 631 school children; with 482 adults composed of 150 "hobos", 150 prisoners, 150 deliquents, and 32 business men; and 65 university students. The correlations between the vocabulary test and Stanford-Binet MA for the 631 children, Terman pointed out, was a creditable .91. Even with the 482 adults a coefficient of .81 was obtained. Further evidence of the validity is the constant, regular, and almost straight line of the vocabulary growth curves for successive mental ages.

In discussing vocabulary in the manual for the 1937 revision, Terman reiterated "we have found the vocabulary test to be the most valuable single test in the scale....It agrees to a high degree with the mental age rating on the scale as a whole; correlations for single age groups range from .65 to .91 with an average of .81."

Shakow and Goldman (1938) reported a correlation of .64 between Binet vocabulary and education. They felt the degree of the relationship

was due mainly to the indirect effect of mental level. A year later, Elwood (1939) found an exceedingly high correlation (978⁺,0009) between the Einet mental age and vocabulary scores with a large number of Pittsburgh school children. Using part of the standardization data for the 1937 revision, McNemar (1942, pp. 139-140) reported productmoment correlations of .71, .83, .86, and .83 for ages 8, 11, 14, and 18 between vocabulary and composite MAs. There were better than 200 subjects at each age level except year 18 which had 101 subjects. Cureton (1954) used McNemar's data to provide mental age equivalents for vocabulary scores. In doing so, he acknowledged the usefulness and validity of the vocabulary test.

Lewinski (1948), in a lengthy review of the literature on vocabulary and mental tests, pointed out the acceptance of vocabulary tests in general and the Binet vocabulary in particular. Tests such as the Columbia Vocabulary Test, Wide Range Vocabulary Test, Knauber Art Vocabulary Test, and the Michigan Vocabulary Profile Test all point to the widespread use and acceptance of vocabulary tests. Tests of deterioration such as the Shipley-Hartford Retreat Scale, Babcock and the Hunt-Minnesota Test for Organic Brain Damage, assume vocabulary to be a measure of intellect.

Levinson (1958) attempted to find the relationship between the Binet MA and Binet vocabulary with foreign and native born American subjects. With age groups of 4 to 5-11, 6 to 7-11, and 8 to 9-11, he obtained correlations of .64,.44, and .61 for the foreign group, and .62, .70, and .70 for the native born group. He felt that the vocabulary over-estimated the MA of native born children and underestimated the MA of foreign children. Nevertheless, the high correlations indicate the test still provides a good index of intelligence.

Evidence from the literature is so conclusive there can be little doubt that a strong relationship exists between the Binet and the Binet vocabulary. Many feel the relationship is so high that the vocabulary alone provides a good rough measure of intelligence. Thus the Goodenough test and the Binet vocabulary are both proported to correlate well with the total Stanford-Binet. Do they (Goodenough test and S-B vocabulary) correlate well with each other? This question and the relationship between the S-B vocabulary and the total Binet became the basis for the second problem of this thesis.

<u>Stanford-Binet Scatter</u>. Scatter has long been of interest to psychologists. Binet and Simon (1916) thought it was a characteristic of the defective child. In 1937 Harris and Shakow (1937, pp. 134-150) reviewed the literature on the significance of scatter. A number of contradictions were found. For example, 5 studies reported feebleminded subjects scatter more than normals while 4 studies indicated that this was not true. Two studies reported greater scatter among neurotic children, and 1 study indicated that this was not the case. In 3 studies children of superior intelligence scattered more than the average and in 2 studies they did not. These and a number of other studies of S-B scatter lead Harris and Shakow to conclude:

1. Feebleminded, deliquent and neurotic children scatter little, if more, than normal children, so far as numerical measures of scatter are concerned.

2. Scatter is probably a little greater in bright than the average children, but not sufficiently so to be of diagnostic value.

3. Results vary somewhat with the test used and with the measure of scatter used.

4. At least some measures of scatter are systematically related to mental age. The results of studies which do not control this relationship allow only an ambiguous interpretation.

5. The relative merits of the various measures of scatter have not yet been satisfactorily determined.

6. In order to draw correct inferences about the clinical importance of numerical scatter in test results from adults, normal adults rather than children must be used as a standard for comparison. No such study has yet been reported. (Marris & Shakow, 1937, p. 148).

The following year the same writers (Harris & Shakow, 1938, pp. 100-111) checked 154 schizophrenic patients, 133 normal adults, and 138 delinquent adults with 4 scatter measures. The test was administered in the conventional manner except for some minor changes to make the test more suitable for adults. The results were negative, only mental age was found to be related to the amount of scatter.

After a brief review of the literature Hunt and Cofer (1944, pp. 548-550) concluded.... "the scatter approach appears now to be a blind alley."

A year later, in 1945, Mayman (pp. 548-551) concluded his review of literature by stating...."numerical measures of scatter on the Stanford-Binet have proved to be virtually useless as aides in clinical diagnosis; nevertheless, the clinical impression that the extent of scatter on the Stanford-Binet may be indicative of maladjustment persisted."

Two more recent opinions were expressed by Cronbach (1960, p. 186) and Freeman (1962, p. 326). The former feels that "after many studies of scatter, investigators now agree that it has no value as a score", and no diagnostic worth. The latter qualified his opinion by stating that "in view of inconsistent data, we must conclude that numerical measures of general scatter on the Stanford-Binet scales are, at present, of limited use as clinical aides, so far as most individual cases are concerned."" The evidence against scatter being a meaningful diagnostic aide is overwhelming. In light of this, the belief that scatter is due to differing patterns of abilities is much more acceptable (See earlier comments of Terman and Merrill in section on scatter). Assuming this, scatter would be due, in part, to individuals excelling on verbal and not performance items or the reverse. Since the Binet tends to be primarily a verbal test and the Goodenough a performance test, it would appear that the difference in scores on the tests would be related to scatter, that is, the greater the unevenness in verbal and performance abilities in the individual, the greater will be his scatter and the greater will be the difference between his Goodenough and Binet scores. The investigation of this relationship became the third problem of this thesis.

CHAPTER III

SUBJECTS, MATERIALS, AND STATISTICAL METHODS

The files of the Child Study Service of the University of Omaha provided the source from which the data for this investigation was abstracted. The files contained the records of all children having had intellectual evaluations in a 9 1/2 month period. The children came from the Omaha Public Schools as referrals in need of special programming. They were often suspected by their teachers and principals of being slow learners (as evidenced by low grades, for example) or not working up to their capacity. The evaluations were not limited to, but always included, the administration of the 1960 Stanford-Binet. Primarily for research purposes the Goodenough DAM test was also contained in the test battery.

Only those children 5 through 12 years, who were tested between October 1, 1961 and July 15, 1962, and whose IQ scores were below 85 on either the Binet or Goodenough were selected as subjects. Of those who met these requirements, 11 were not included for the following reasons: 5 did not have Goodenough records, 2 drew heads rather than complete figures, 1 produced the figure of a woman, 1 became ill during testing, 1 was uncooperative during testing, and 1 did not achieve a basal. It was later decided that only Class E (recognizable) Goodenough drawings should be considered, thus eliminating 15 subjects with Class A drawings. The final sample included 226 boys and 119 girls for a total of 345 subjects (See Table I). For the distribution of S-B and DAM IQ scores see the appendix.

The actual testing of the children was done by 4 psychometrists

Ages	S	ę	7	89	6	10	11	12	Total
2	24	59	47	40	43	57	38	37	345
CA Wean	66.958	77.738	80.543	101.875	112 476	175,196	137.324	150.243	107.889
ß	2.605	3.100	3,651	3,465	3.410	3,486	3.119	3.275	26.034
Binet MA				•	•	•	9		• •
Mean	56.000	61,423	69,085	84.525	97.162	105.421	110.947	120.108	88.240
SD	8.698	6.481	9.642	13,626	13.923	19.965	17.117	21.760	26.286
Goodenough	s. Fra			•					
Mean	53.875	67.644	71.361	80.325	86.534	93.578	100.894	99.486	82.379
ß	8.053	15.657	14.980	12.665	15.786	13,997	14.557	15.528	20,305
Binet IQ									
Mean	81.500	77.474	74.744	80.700	84.046	82.859	81.131	81.243	80.272
ß	12.780	8.332	10.587	12.898	11.603	14.567	11.112	13.090	12.260
Goodenough	P4								
Mean	80.625	87.423	79.617	78.975	76.976	74.701	73.526	66.270	77.704
SD	12.240	19.176	17.123	12.404	13.754	11.003	11.011	10.574	15.328

Chronological Ages and Binet and Goodenough Mental Ages for the 345 Subjects in Sample

Table I

Total Einet MA Median 86.000; total Goodenough MA Median 84.000.

under the supervision of Dr. D. T. Pedrini, the director of the Child Study Service. The 4 examiners were studying for Master's degrees in psychology and had undergraduate majors in psychology. All were enrolled in a graduate course in individual mental testing, and, prior to October 1, all had undergone an intensive 6 week training period in the administration and scoring of the 1960 Stanford-Binet. They were also given instruction in the administration and scoring of the Goodenough Intelligence test.

The procedure involved in testing a child followed a relatively stable pattern. The referred child was usually brought to the Child Study Service by his parent(s). He or she was introduced to the examiner who made every effort to put the child at ease, and, in general, establish rapport. When sufficient rapport was reached, the testing was begun. The Goodenough was usually given first, with the Binet following immediately after. With one exception both tests were administered according to the specific instructions of their respective authors. The lone exception was the use of 8 1/2 x 11 yellow paper instead of the test blank suggested by Goodenough.

Tests. The 1960 Stanford-Binet and Goodenough Intelligence tests are too well known to warrant more than a cursory description. The Binet test covers the range from age II through 3 Superior Adult levels. Half-year intervals are found at ages II through V, yearly intervals from V to XIV with the remaining levels designated as Average Adult and Superior Adult I, II, and III. There are a total of 142 subtests in the Binet, 6 plus an alternate at each level except the Average Adult which contains 8 plus an alternate.

The sub-tests at the lower age level frequently require eye-hand coordination and the identification of common objects and minimize the need for a great deal of verbal response. Examples are manipulating a 3-hole form board, identifying parts of the body on a paper doll, building a 4 block tower, requiring the child to point to a cup when the examiner says, "Show me what we drink out of," and the drawing of a vertical line.

In the middle range of the scale, a wide variety of subtests are found. They include such areas as comprehension, memory, recall, and spatial orientation. In the upper levels of the Binet, the subtests are almost entirely verbal as opposed to non-verbal. Vocabulary, abstract reasoning, and concept formation account, in a large part, for the type of abilities tapped in the Superior Adult levels.

The new Binet, like its predecessor, is an age scale; that is, the subtests are grouped and arranged in terms of various age levels. Each subtest passed earns credits towards the mental age score. Between Binet ages II and V, the subject is credited 1 month for each subtest passed; from years VI through XIV he is credited 2 months for each subtest success. Up through age XIV, a maximum of 12 months can be earned at each year level. At age AA (Average Adult) a maximum of 16 credits can be earned. Superior Adult levels I, II, and III are credited with a maximum of 24, 30, and 36 months respectively. The MA score is computed simply by totalling the months credited at each year. With the subjects MA and CA, his IQ can be found by referring to Pinneau's Revised IQ tables in the back of the manual. The tables provide deviation IQ's with a theoretical mean of 100 and a standard deviation of 16. This means the IQs are, in effect, standard scores and therefore comparable at all ages. The use of deviation IQs is perhaps the most important single innovation of the 1960 Stanford-Binet compared to the 1916 and 1937 S-B.

The Goodenough, a point scale, is a relatively simple test contrasted with the Binet. A blank sheet of paper and a pencil is all the equipment required. The subject is asked to draw a picture of a man, the very best picture he can. The drawing is then scored on the basis of passing or failing each of the 51 items of the test. In essence, the items represent different details or aspects of the drawing. For example, a point is given for the presence of a head; in the same manner, points are given for showing a neck, eyes, hair, clothing, fingers, mouth and legs. Points are also given for adequate body proportions, and for varying degrees of motor coordination depicted in the drawing.

Mental age is determined by totalling the scores and referring to the table of MA equivalents provided by Goodenough. Beginning at MA 3 years 3 months, each score is equivalent to 3 months. Thus a score of 4 is converted to 4 years 0 months, a score of 5 to 4 years 3 months, 6 to 4 years 6 months, and up to the MA level of 13 years 0 months. No mental age equivalents are given beyond 13. Intelligence quotients are obtained by dividing the MA by chronological age and multiplying by 100.

Theoretically the mean IQ is supposed to be 100. The standard deviations, however, are not equal at all levels so IQ's are not comparable at all ages.

Statistical Procedures. The statistical procedures employed are,

quite naturally, directly related to the questions raised. In this section major statistical techniques and the questions it was hoped they would answer will be discussed.

What was the relationship of the Goodenough to the Binet? The inquiry was not merely to determine whether a relationship existed but to ascertain the degree with which it occurred. Further, this information was desired for separate age levels as well as the sample as a whole. Lastly, information was desired relative to the possible effect of the correlation of the means of the age groups. (See Chapter II for the advantages of considering separate age levels and the possible effect of correlation of means.)

The statistical technique which would encompass the above problems was the Analysis of Covariance for simple factor experiments as described by Winer (1962, pp. 578-594). Minor changes in computation were necessary due to the unequal cell frequencies in the sample. Winer (1962, p. 594) provided the necessary computational formulae for the transformations, and a number of checks for the assumptions underlying the Analysis of Covariance. It was thus feasible to decide whether regression coefficients within each treatment class were homogeneous, whether regression coefficients within equalled regression coefficients between classes, whether the between class regression was linear, and whether the sample had overall linear regression. Intrinsic in the testing of regression effect is the formulation of regression equations. Consequently the prediction of the most likely score on one test from knowledge of a score on the other was made possible. The main importance of the Analysis of Covariance however, was to yield within class, pooled

within class, between class, and overall correlations between the covariate and the criterion. That is to say, the primary purpose of the analysis was to give correlations within each age group, an average correlation of the age groups, a correlation between age groups or means of the age groups, and an overall correlation between the MA scores of the Goodenough and Binet.

Were there significant differences between the MA's of the Goodenough and the MAs of the Binet? The answers sought for this question appeared to fall in the realm of the Analysis of Variance (ANOV). The data lent itself to a two-factor experiment with repeated measures on one factor (Winer, 1962, pp. 298-312). The general case is represented schematically below (Winer, 1962, p. 302).

	^b 1	• • •	^b j		b q
^a 1	G ₁	• • •	G ₁	• • •	G ₁
a	G _i	•••	G _i	•••	G i
• • •	Gp		G	•••	G p
¥	р		Р		р

The repeated measurements are made on factor B. In this study there were only two levels of factor B, b_1 symbolizing the Binet test and b_2 the Goodenough test scores. In like manner, factor A represented the age levels, i.e., a_1 being the 5 year olds, a_2 the 6 year olds, up to a_8 the 12 year olds. The subjects are nested under the various levels of factor A. The symbol G_1 refers to a group of n_1 subjects of level a_1 of factor A. It will be noted that the group of subjects from any level of factor A is observed under all levels of factor B. The design assumes an equal n at each level of factor A. To adapt the data (with n not equal) to the design, major computational transmutations were required but the general form remained the same. Since the unequal group size did not appear to represent "different strata within the specific population" (Winer, 1962, p. 374), the unweighted-means solutions were considered appropriate.

The use of the ANOV enabled the testing (F test) of overall significance of differences between the 2 levels of factor B on the Binet and Goodenough. Non-chance variations among the means led to further testing with the pairs of means at each level of factor A or age groups. An F test was also used to determine whether means at the age levels varied more than that expected by chance. It was assumed they would since mental age generally increases with chronological age. Because the F was significant, it meant individual t tests were in order. Tests were made for interaction, that is, for the joint effects of factor A and B acting together, or, more simply, whether significant differences between the Binet and Goodenough were related to the age levels.

In the foregoing ANOV discussion, the subjects had been grouped by chronological age. To cast further light on the subject, the subjects were regrouped by Binet mental age, completely disregarding CA. All children with a Binet MA of 5 years to 5 years 11 months formed one level, those from 6 to 6 years 11 months formed another and so on. All cases of MA 12 years and above were considered as one group as the n became small. Each year level was examined for significant differences

between the Binet and Goodenough MA scores. A two-tailed t test for correlated observations was used (Winer, 1962, pp. 39-43).

The relationship of Binet vocabulary to the total Binet and the vocabulary to the Goodenough was checked with the Pearson Product Moment Correlation. The r between the vocabulary and the total Binet MA score was in small part, spuriously higher because the vocabulary is part of the total score. McNemar (1942, p. 140) however, has pointed out that the degree of spuriousness is not great since the vocabulary subtest contributes less than 5 percent of the total score.

First order partial correlation technique was used to investigate the relationship between test scatter on the Stanford-Binet and the difference in MA scores made on the Binet and Goodenough tests. From a scatter diagram, it was observed that the difference in the Binet and Goodenough scores tended to increase with chronological age. Therefore, in correlating scatter and the difference, it became essential to hold age constant by partialling it out. The formula used may be found in either McNemar (1962, p. 166) or Guilford (1956, p. 316).

The statistical procedures thus far discussed are directly related to questions this study attempted to answer. One test was carried out however, which was not related to any of the original hypotheses. The sole purpose in making the test was simply to gain a clearer picture of the sample. To learn whether the children at each age level were of equal brightness, the Binet IQ scores were examined by a singlefactor ANOV for unequal sizes (Winer, 1962, pp. 96-104). This was followed by the Newman-Keuls test for differences between all ordered pairs of means. In the ANOV the 8 age groups were considered treatments. The means were the mean Binet IQ scores of each of the 8 age groups.

In this section only the more complex statistical operations have been touched upon. It was felt there was no need to discuss the descriptive statistics such as mean, median, and standard deviation. The formulas used for those statistics were the standard ones found in most statistic textbooks. Also omitted were discussions of minor questions such as percent passing the Binet vocabulary at various ages.

CHAPTER IV

RESULTS AND DISCUSSION

The results given in this chapter will follow the same order as the statistical procedures presented in the previous chapter. The results are discussed as they are given. A unified discussion of the results as a whole has been reserved for Chapter V.

<u>Correlations</u>. The Pearson correlations (r) between the Binet and Goodenough MAs are given in Table II. Of the 8 age groups only 1 r was significant. It was particularly interesting to note that

Table II

Correlation Between 1960 Stanford-Binet Mental Age and Goodenough Draw-A-Man Mental Age

Ages	5	6	7	8	9	10	11	12
						57 •310*		

*significant at the .05 level

between the 2 highest correlations, at ages 8 and 10, a negative r was found at age 9.

McNemar (1962, pp. 119-135) suggested 5 methods of interpreting r. One is that r is "associated with the rate at which one variable changes with another." Viewed in this manner, the low and nonsignificant correlations in Table II led to the conclusion that the rate at which the Goodenough changes has little to do with the rate of change in the Binet. Regardless of what interpretation is placed upon r, the degree of relationship in this case remains low.

The pooled within-class correlation was . 140. Basically this is an

average r of the 8 separate age groups. Contrasted with the almost negligible degree of association indicated by an r of .140 is the overall r of .623 significant at well beyond the .01 level. The 2 correlations, one of moderate size and one low, would appear to be a contradiction. Actually the overall correlation is strongly influenced by the between class r or correlation of the means, which in this case was .978. Consider for a moment the 8 ages plotted on a scatter diagram. The tally marks for each age taken separately appear almost randomly placed on the scattergram with the exception that the 5 year olds tend to fall at the lower left of the diagram, the 8 year olds more toward the middle and the 12 year olds nearer the upper right hand margin. The 8 separate age groups viewed together or taken as a whole appear to resemble a normal but somewhat fan shaped scatter plot. The overall r is affected by a variable other than the relation of the covariate to criterion. One method of controlling this is to hold the uncontrolled variable constant via the partial correlation. Another method to make the groups comparable is by the use of the Analysis of Covariance. The latter method was used to make each group comparable with respect to CA, resulting in a pooled within r of .140. It is this pooled within-class r, not the spuriously high overall r, which best suggests the degree of relationship between the MAs of the Binet and Goodenough tests.

In the preceding paragraph it was assumed there was a difference between the treatment means. Tables III and IV leave little doubt that this was true.

Source	df	MS	F
CA level	7	9,990.294	46,560
Error	337	214,567	-
Total	344	•	

Analysis of Variance Summary Table

The critical Valuesis F.99(7,337) 2.64

A critical value of $F_{.99}(7,337)$ 2.64 indicates statistically significant differences in the treatment means. Phrased another way, there were significant differences between the 8 age levels.

Table IV

Analysis of Covariance Summary Table

Source	df	MS	F
CA level	7	2,309.926	10,949
Error	336	210,960	
Total	343	·	

The critical value 15 F.99(7,336) 2.64

After an adjustment was made for the linear trend in the relationship between the criterion and covariate, the differences of course remained significant $F_{.99}(7,336)$ 2.64.

F ratios were used in testing the assumptions underlying the Analysis of Covariance. One of the fundamental assumptions is that the regression coefficients within each of the treatment classes are homogeneous. The hypothesis that they were equal was accepted when the F of .953 was compared with the critical value of $F_{.99}(7,329)$ 2.64. The assumption that the between class regression was linear was also accepted, F_{obs} 2.385 and $F_{*99}(6,336)$ 2.80. The assumption that the regression coefficient within-class equalled the regression betweenclass was not met, F_{obs} 62.330 and $F_{*99}(1,336)$ 6.63. Also not met was the assumption of linearity of the overall regression, F of 5.946 compared with a critical value of $F_{*99}(14,329)$ 2.14.

The fact that the last 2 assumptions were not satisfied does not necessarily negate the value of the Analysis of Covariance. The lack of overall linearity however, does suggest the overall r was not accurate. To correct for the possible effect of curvilinearity correlation ratios: were computed. The resulting etas for Y on X (Goodenough on Binet) and X on Y were .7012 and .8258 respectively. Unfortunately the usefulness of the correlation ratios was extremely limited as they are also influenced by the r between the means. They did, however, prove curvilinearity, $F_{\rm obs}$ of 11.45 and $F_{\cdot 95}(6,337)$ 2.09.

<u>Regressions</u>. The overall regression equation was undoubtedly also affected by the lack of overall linearity. The regression coefficient was .4811 for predicting the Goodenough (Y) from the Binet (X) and .8063 for predicting the Binet from the Goodenough. The regression equation for the prediction of the Goodenough (Y) from the Binet (X) is Y''=.4811X + 39.9272, for predicting X from Y is X''=.8063Y + 21.8778.

Since the within-class regression coefficients have been demonstrated to be homogeneous, it is possible to obtain a single pooled estimate for the pooled within-class regression. The pooled within-class regression coefficient was .1368 for predicting the Goodenough from the Binet and .1437 for predicting the reverse. The value of the regression equation is in predicting one variable with knowledge of another. The efficiency of such forecasting lies in the degree of relation which exists between the 2 variables. In this case the correlation is almost negligible and therefore accurate prediction is highly unlikely. For this reason, no regression equations will be given for the separate age groups. Moreover, it is felt that the equations could give a false sense of knowledge regarding the prediction of the Binet from the Goodenough.

<u>Mean Differences</u>. It is conceivable that the Binet MAs and Goodenough MAs could be highly correlated and still yield MA scores which are significantly different. It has already been shown that the tests are not highly correlated but the question of significant differences remains.

Table V shows that there are significant differences between levels of factor A (CA levels). This should be no surprise since it would be expected that the mental age of the 5 year olds would be much lower than the MA of, say, 12 year olds. Comparing the F ratio of 131.15 with the critical value of $F_{.95}(7,337)$ 2.01, indicated the differences were considerable. This information has already been given in discussing the results of the Analysis of Covariance.

Table V

Analysis of Variance Summary Table

Source	df	MS	F
Between subjects			
A	7	32,866,406	131,1513
Subj. w. groups	337	250,599	-
Within subjects			
B	1	6,539.796	34,6220
AB	7	1,497.358	7,9271
(subjects		·	
w. groups	337	188.891	

Normally a significant F as that obtained on CA levels, would call for individual comparison with all possible means. In this case it was not necessary, for it could be assumed that a significant difference between ages 5 and 6 would also mean that differences existed between 5 and 7, 8, 9, etc., so long as a significant difference was found between ages 6 and 7, 8, 9, etc. Stated another way, since the MA increases with CA, significant difference between 5 and 6 would automatically indicate a significant difference between 5 and 7, 5 and 8, 5 and 9 and so on. Therefore tests between each age and the age level next to it were made. The F tests shown in Table VI point out significant differences at each age.

Table VI

an a	F Tests f	or Signi	ficant Di	ferences	Between Ag	e Levels	
Ages F rati	566 .os12.568	647 6.811	768 25 . 732	889 14.677	9810 11.518	10&11 7.544	11612 2.258
The cr	ILICAL VAL	He d's F		2.01			

The critical value us F .95(7,377) 2.01

While it was expected that significant differences would occur between age levels, it was not necessarily anticipated that differences would occur between the Binet and Goodenough (levels of factor B). Comparing the ANOV F ratio of 34.622 with the critical value of $F._{95}(1,7)$ 3.84 leaves little doubt that differences exist. To check this, a separate F test was made for individual comparisons of factor B. The resulting F ratio was 34.684 and the critical value was $F._{95}(1,337)$ 3.84. The slight difference in F values was probably due to a rounding error. Thus the investigator was forced to conclude that the means of the Binet and Goodenough for the overall sample were decidedly significantly different. Are they significantly different at all age levels?

The F ratio for AB interaction suggests the differences are related to age levels (levels of factor A). That is, when the F ratio of 7.927 was compared with the critical value of $F_{.95}(7,337)$ 2.01, it was apparent the difference in factor B (the Binet and Goodenough) was not solely an attribute of factor B but varied at the separate levels of factor A (age levels). It was thus necessary to test each separate age level for a significant difference between the Binet and Goodenough. Twotailed t tests for the difference between 2 means with correlated observations were run. Table VII reveals significant differences at ages 6, 9, 10, 11, and 12. No differences were found at ages 5, 7, and 8.

Table VII

t Tests for Significant Differences Between the Binet and Goodenough at Separate Chronological Age Levels

Ages	5	6	7	8	9	10	11	12
tobs t.975± df	.963 2.07	-2.961 2.01	.959 2.02	1.706 2.02	3.1 15 2.02	4.357 2.01	3.089 2.04	4.778 2.04
df	23	58	46	39	42	56	37	36

To examine this problem further, the subjects were regrouped by Binet mental age levels. The results furnished in Table VIII, disclose significant differences at all Einet mental ages but years 6 and 7. The regrouping has produced markedly similar results. In general, equal

Table VIII

t Tests for Significant Differences Between the Binet and Goodenough at Separate Binet Mental Age Levels

Ages	below 5	5	6	7	8	9	10	11	12 plus
tobe	3.451	2.693	,375	.813	3,855	6.715	9.881	14.063	14.632
t.6751	3.451 2.02 42	2.00	2.02	2.01	2.02	2.31	2,09	2.14	2.31
df	42	74	40	52	43	44	19	14	8

means are limited to the lower ages regardless of grouping while significant differences tend to be found in the upper half of the age levels.

<u>Correlations with Binet Vocabulary</u>. In Table IX information and correlations (Pearson) relative to the Binet vocabulary is given. The vocabulary scores are raw scores. The r between the Binet and Goodenough IQ are also included.

All of the correlations between the S-B vocabulary and S-B MA or IQ were significant beyond the .01 level; none of the correlations between the S-B vocabulary and Goodenough MA or IQ (for separate ages) reached significance at that level. The correlations between vocabulary and Binet MA fell approximately in the range Terman found for the 1937 revision. They are also in close agreement with the other studies reported in the review of the literature. The total r between vocabulary and Binet MA is slightly higher than the individual correlations and probably reflects the influence of the correlations between the means. The total r between the vocabulary and Binet IQ, on the other hand, is slightly lower than most of the separate age correlations. This is probably due, in part, to the vocabulary score increasing with age while the IQ score remained nearly constant. Thus the vocabulary score would correlate moderately high with IQ at any one age but less well with the overall sample.

The correlations between the Goodenough MAsand vocabulary are so low, it can reasonably be concluded that no significant relationship exists between them as far as this sample was concerned. The moderately high total r is likely to be revealing the r between means rather than

Ages	'n	6	7	8	Q	10	<u>ب</u>
2	24	59	47	40	43	57	3 8
Means	2.625	4.050	5.276	7.275	9.093	9.771	9.736
ខ	2.324	2.204	2.130	2.280	2.310	3.366	1.665
	.746	.623	.725	.692	.802	.783	+653
r with Binet IQ	.735	.560	.724	.678	.791	.783	.674
r with Good. MA	.024	.015	.189	.268	095	.279	.238
	082	040	.119	.221	141	.258	.258

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IX

Binet Vocabulary Means. Standard Deviations, and Correlations with Binet and Goodenough MAs and IOs

the relationship of the two variables.

<u>Correlation of Stanford-Binet Scatter</u>. The correlation of S-B scatter and chronological age was .59. The r was significant well beyond the .01 level. Assuming that children's abilities differentiate with increasing CA and assuming that scatter represents differing abilities, the magnitude of the above r would not be unusual. Following this logic and assuming the difference in Binet and Goodenough MA scores represents differing abilities, the obtained r of .35 between the difference in individual S-B and DAM MA scores and CA would also not be unusual. Scatter and the difference were correlated and yielded an r of .34. Both the .35 and .34 correlations were significant at the .01 level. Part of the latter correlation could undoubtedly be attributed to CA. Partialling out CA gave an r between scatter and the difference in Binet and Goodenough MA scores of .23. While the r is low, it is significant beyond the .01 level.

The practical value of an r of .23 even when statistically significant is almost nil. However, from a theoretical point of view, the correlation has established beyond a reasonable doubt, the existence of a relationship between scatter and the difference in MA scores. From this relationship, we may speculate that Binet scatter is in part due to differences in verbal and performance abilities or more generally to unevenness of abilities.

<u>Mean Binet IQ Differences</u>. The means of the Binet IQ's for the 8 age groups were examined for significant differences. The results of the ANOV are presented in Table X. The F ratio of 2.95 compared with the critical value of $F_{.95}(7,337)$ 2.01 suggests significant differences.

Means74.744677.4745 80.7000 81.2432 81.5000 82.8596 74.74462.7299 5.9554 6.3869 6.4986 6.7554 8.1150 77.4745 2.7299 5.9554 6.3869 6.4986 6.77554 8.1150 80.7000 81.315 3.2255 3.6570 3.7687 4.0255 5.3851 81.1315 81.117 2.7296 4.315 5.7687 4.0255 5.3851 81.2432 81.5000 82.8596 1.1177 2568 1.6164 81.5000 82.8596 1.6164 1.3596 1.5268 1.6164 81.5000 82.112 2.10 9 2.568 1.6164 81.0465 8.112 5.10 9 1.3596 76 112 5 10 9	Treatments	ints	7	Q	œ	11	12	ъ	10	0
74.7446 2.7299 5.9554 6.4986 6.7554 8 77.4745 2.7299 5.9554 6.4986 6.7554 8 80.7000 81.1315 3.2255 3.6570 3.7687 4.0255 5 81.1315 3.2255 3.6570 3.7687 4.0255 5 8000 2 81.1315 81.1315 .3.2255 3.6570 3.7687 4.0255 5 8000 2 81.2432 81.2432 .81.2432 .81000 2 .8000 2 81.2000 82.8596 .1117 .2568 1 .2568 1 82.8596 84.0465 .112 5 10 9 .2568 1 7 6 1 12 5 10 9		Means	74.7446	77.4745	80.7000	81.1315	81.2432	81.5000	82,8596	84.0465
77.4745 3.2255 3.7687 4.0255 5 80.7000 81.1315 .4315 .5432 8000 2 81.1315 .4315 .4315 .5432 8000 2 81.1315 .1117 .5432 .8000 2 81.2432 81.5000 81.5000 .1117 .5685 1 81.5000 82.8596 84.0465 .1117 .2568 1 7 6 11 12 5 10 9 7 6 11 12 5 10 9	~	74.7446		2.7299	5.9554	6.3869	6.4986	6.7554	8.1150	9.3019
80.7000 .4315 .5432 .8000 2 81.1315 .1117 .3685 1 81.2432 81.2432 .8000 2 81.2432 81.2600 .1117 .3685 1 81.5000 82.8596 .1117 .2568 1 82.8596 84.0465 .1112 5 1 .2568 1 7 6 8 1 12 5 16 9	9	77.4745			3.2255	3.6570	3.7687	4.0255	5.3851	6.5720
81.1315 81.2432 81.2432 81.5000 82.8596 84.0465 84.0465 7 6 8 11 12 5 10 9 7 6 8 11 12 5 10 9	20	80.7000				.4315	.5432	. 8000	2,1596	3.3465
81.2432 81.5000 82.8596 84.0465 7 6 8 11 12 5 10 9 7 6 8 11 12 5 10 9	11	81.1315			Antiperior Allocations		1117	.3685	1.7281	2.9150
81.5000 82.8596 84.0465 7 6 8 11 12 5 10 9 * *	12	81.2432						.2568	1.6164	2.8033
82.8596 84.0465 7 6 8 11 12 5 10	IJ	81.5000							1.3596	2,5465
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Table XI

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TABLE X	TA	BL	E	Х
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Source	df	MS	F
A Levels	7	428.3432	2.9543
rror	337	144.9851	
Fotal	344		

ANOV of Binet IQ's for the CA Groups

The critical value is $F_{.95}(7,337)$ 2.01

An examination of all pairs of ordered means reveaged significant differences between CA's 10 and 7 and between 9 and 7. These results are given in Table XI. The reasons for these differences in the sample are not known. Perhaps it is because the slow learners are not recognized until they are almost through the 1st grade. Perhaps the 5 year olds are high on the ordered means because they were tested for early admittance to school. Such explanations, of course, are sheer conjecture.

Boys vs. Girls in Sample. The Binet and Goodenough MA scores for boys and girls are given in Table XII. Two-tailed t tests were again

TA	EL.	F	Х	T	T
			**	**	-

Means	and	Standard	Deviations	of	S-B	and	DAM MA	Scores 1	or	Boys (4 Gir)	IS.

	N	Bi	net	Goodenough		
		Mean	SD	Mean	SD	
Boys	226	92.853	26.934	83.247	19.997	
Girls	119	79.478	22.545	80.731	20.777	
GITIS	119	19.470	24.545	80.731	20.77	

used to investigate significant differences between the means (Winer, 1962, pp. 36-43). When the mean Binet score made by boys was compared with the mean Binet score made by girls, a significant difference was found, t_{obs} of 4.889 and a critical value of $t_{.975}(279)^{\pm}1.97$. Making the same comparison with the Goodenough test, no significant difference between the boys and girls was found, t_{obs} of 1.083 and $t_{.975}(225)^{\pm}1.97$. A t_{obs} of 8.386 and a critical value of $t_{.975}(225)^{\pm}1.97$ was found when the mean scores made by the boys on the Binet and Goodenough were compared. This of course, indicated a significant difference. However, the means made by the girls on both tests did not show a significant difference, t_{obs} of .793 and $t_{.975}(118)^{\pm}1.98$.

It has been pointed out that Binet and Goodenough MA means were equal at age 5, 7, and 8, i.e., at the lowermage levels. Whether girls (they were younger than the boys) caused the equal means at the lower ages or whether the lower ages produced equal means in girls is debatable. Further research with this problem could lead to some interesting and worthwhile results.

CHAPTER V

SUMMARY AND CONCLUSIONS

The primary purpose of the study was to investigate the relationship between the Goodenough Intelligence Test and the 1960 Stanford-Binet with mentally subaverage children. A secondary problem was to examine the relationship between the Binet vocabulary subtest and the total Binet and Goodenough tests. Further, the difference between individual Binet and Goodenough MA scores and scatter on the Binet was studied. The subjects were 345 mentally subaverage (IQ's of below 85 on either the Binet or Goodenough) children years 5 through 12 tested at the University of Omaha Child Study Service.

Previous research with the Stanford-Binet and Goodenough was mainly in the form of obtaining an overall correlation between the two tests. The reported correlations were generally of moderate size, between .40 and .70 with a few in the .80s. However, in the sole study (Rohrs & Haworth, 1962) using the 1960 Binet the r was low and nonsignificant.

Only 2 investigations reported tests of mean differences. Birch (1949) found a difference, significant at the .05 level, between the Binet and Goodenough MA means. An apparent contradictory finding was made by Rohrs & Haworth (1962) who reported no significant difference between the Binet and Goodenough IQ means.

In the present investigation, Analysis of Variance, Analysis of Covariance, t tests, partial correlations and correlations were the major statistical techniques used. Derived from the Analysis of Covariance was pooled within-class r of .140 between the Binet and

Goodenough MAs. Although extremely low, the r was significant at the .01 level with a sample size of 345. It was guite evident the size of this correlation was not in line with prior research. However, the correlation is not unlike the r yielded in the 1960 Binet study by Rohrs and Haworth (1962). They suggested the omitting of many performance items found in the 1937 revision and not in the 1960 revision might account for the low correlation. There are other possible explanations to reconcile this study's correlation with past research. Already mentioned is the danger of obtaining a spuriously high overall correlation due to the correlation of means. Many of the studies reported in the review of the literature used multiple age groups and reported only an overall correlation. Perhaps the correlations were spuriously high due to the between mean r. Another feasible explanation lay in changes in drawings made by children in 1926 as compared with the present. In the 37 years which have elapsed since Goodenough standardized and published her test the authors of the Stanford-Binet have felt it necessary to revise the S-B test twice. Perhaps if the Goodenough scoring criteria were restandardized, the correlations between it and the Binet would be higher.

The fact remained, the r obtained in this study with this sample is almost negligible. Interpreting r as either "the rate at which one variable changes with another" or "how accurately we can predict by a regression equation" (McNemar, 1962, p. 134) led to the conclusion that changes in the Binet were not reflected by similiar changes in the Goodenough and that prediction was unpractical and unwarranted. Assuming the Binet could be considered a criterion of intelligence, the fact that it couldn't be predicted places the Drawing test in an awkward position.

An F test from the ANOV determined the existence of an overall significant difference between the Binet and Goodenough MAs. A significant interaction F led to the conclusion that the differences did not exist throughout the entire age range. The individual age groups were checked and significant differences were found at chronological ages 6, 9, 10, 11, and 12 while no differences were observed at ages 5, 7, and 8.

Attempts to account for these results on the basis of previous research is complicated by many earlier researchers not having made tests of mean differences. Johnson et al., (1950) typified the character of the research when, in referring to the Binet and Goodenough IQ means, he stated they "were numerically close to each other." In her original publication, Goodenough (1926) made the means of the Binet and Goodenough conspicuous by their absence, which, <u>ipso</u> <u>facto</u>, implied they may not have been equal. The results of 2 studies (Birch, 1949: Rohrs & Haworth, 1962) made tests of mean differences apparently contradict each other. Had they tested for interaction, they may have obtained results similar to this study.

If the Binet is an adequate criterion of intelligence for children, then the low correlations and significant differences between it and the Goodenough would place the latter in a seemingly untenable position as a children's intelligence test. Unfortunately the matter is not that simple. Intelligence is a complex concept and it would be flagrantly presumptuous to assume the Stanford-Binet has embodied all its many facets and ramifications. Yet, the Binet has demonstrated a remarkable

ability to sample many of the important aspects of what is generally considered intelligence. With this in mind, it was concluded that the Goodenough Drawing Test could, in no way be considered adequate as an individual test of intelligence. Further, the advisability of using the Goodenough as a group intelligence test is questionable and for children beyond age 8 years it is not recommended. The questionable use of the DAM as a group test is based on MA and not IQ. How well the DAM compares with other group tests was not considered in this investigation. These conclusions are, of course, limited to children similiar in characteristics to those found in this sample.

Previous research with the Binet vocabulary and the total Binet MA indicated a high degree of relationship. The vast majority of product moment correlation coefficients fell in the .60 to .90 range. The coefficients yielded in this investigation ranged from .62 to .86 with a median r of .74 for the 8 age groups. These correlations are quite in accord with previous findings and suggest that the vocabulary subtest would make a satisfactory estimate of the kind of intelligence measured by the 1960 Binet. All the correlation coefficients cited, however, were probably somewhat inflated since the vocabulary subtest contributes to the total Binet score and no correction was applied.

No studies were found which attempted to correlate the Binet vocabulary with Goodenough mental age. The results of the attempt made here showed none of the correlations attained significance at the .01 level. Further discussion is therefore not offered.

Nothing in the literature was found regarding the relationship of Binet scatter to the difference in Binet and Goodenough MA scores.

Endeavors to attach diagnostic significance to scatter were inconclusive. Terman postulated that scatter was a function of uneven manifestations of intelligence in individuals. Perhaps the difference in Binet and Goodenough MA scores was also a function of uneven manifestations of abilities (primarily verbal and performance) in individuals. If so, scatter and the difference would show some degree of relationship or association. To test this, the 2 variables were correlated. After partialling the effect of chronological age, an r of .23 was obtained. The correlation was significant beyond the .01 level. Although low, the r did give evidence of a definite relationship between scatter and differences in Binet and Goodenough MA scores. Indirectly this may seem as lending support to Terman's belief that scatter is due to differing patterns of ability.

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IQ	Binet	Goodenough
145 - 149		1
125 - 129		2
120 - 124		3
115 - 119	2	5
110 - 114	2	3
105 - 109	10	8
100 - 104	10	9
95 - 99	23	. 8
90 - 94	27	17
85 - 89	33	25
80 - 84	64	59
75 - 79	66	58
70 - 74	50	50
65 - 69	25	38
60 - 64	21	24
55 - 59	9	18
50 - 54	2	13
45 - 49	1	3
40 - 44		1
35 - 39		

Distribution of Binet and Goodenough IQs

The differences between S-B and DAM IQ scores ranged from 0 to 73.