

Analysis of microdifferentiation in a Spanish cellar population of *Drosophila melanogaster*

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Abstract

Variation in *Adh* and α *Gpdh-1* gene frequencies has been used to check for microdifferentiation in Spanish samples of *Drosophila melanogaster* inside and outside a wine cellar. Flies were collected after vintage and after overwintering respectively; within each period samples were taken on up to five consecutive days each month. Variation of gene frequencies of *Adh* and α *Gpdh-1* can be considered random when samples collected each month are taken into account. When mean monthly frequencies are considered, α *Gpdh-1* does not show any significant variation all over the year; yet, variation of the frequency of *Adh*^S shows a cyclical pattern, its frequency being maximum at the end of the summer and minimum after overwintering. Due to the parallel change of the frequency of the inversion *In(2L)t* and the *Adh*^S allele, no decision can be made whether the *Adh* locus itself or the inversion are responsible for the changes.

Introduction

When collecting flies in nature drosophilists have often assumed that flies obtained from a given restricted area are members of a single panmictic unit. There is, however, some experimental evidence which does not support this assumption. In fact, Wallace (1966) detected some microdifferentiation in tropical populations of *Drosophila melanogaster* when studying lethal allelism. Krimbas and Alevizos (1973) also found evidence for spatial microdifferentiation in a Greek population of *Drosophila subobscura* when using inversion polymorphism as a genetic marker and so did Cabrera et al. (1985) in that same species when using the *Adh* locus as a marker. When cellar populations of *Drosophila melanogaster* are considered, the cellar population is often thought to be a single panmictic unit. One of the major features differentiating the cellar and its surroundings is the concentration of ethanol fumes; being ethanol fumes a possible selective

agent to maintain *Adh* polymorphism, different authors have compared samples taken inside and outside the cellar and contradictory results have been reported (Briscoe *et al.*, 1975; McKenzie & McKechnie, 1978).

In our previous collections of flies in Spanish cellars (Aguadé & Serra, 1980), a general feature emerged, namely their environmental heterogeneity, which made us think of a possible microdifferentiation inside the cellar. We have therefore tried to detect microdifferentiation both inside and outside the cellar in two periods of the year thought to be important for population structure – after vintage and after overwintering. To trace microdifferentiation we have chosen two enzyme loci – *Adh* and α *Gpdh-1* –, both because of their differential relationship to ethanol and their different degree of association with inversion *In(2L)t*. In addition, a polygenic trait – maximum wing length – was studied (data will be published separately).

Material and methods

A cellar located some 70 km northwest of Barcelona – “Cooperativa Vinícola d’Artés” – was chosen for its environmental heterogeneity and the relative high frequency of the slow allele of *Adh* found in its populations (Aguadé & Serra, 1980; Serra & Aguadé, 1982). Samples were taken in two different periods – after vintage and after overwintering –. Within each period samples were taken during up to five consecutive days each month (early October, November and December 1979 and early June and July 1980 for the second). Each day flies were collected at different sites inside and outside the cellar. Initially flies were taken at a filtering machine (ma), near the grape press (pr), on pomost (remains of pressed grapes) outside the cellar (po) and on pomost kept in a small building close to the cellar itself (pi). Some of these locations had to be changed during the project because no flies could be collected on them. A new sampling location was added in exchange – a leaking barrel (ba). In December due to general scarcity of flies, samples were collected on a single day. Flies were caught directly with a sucking device. No traps were used in order to avoid disturbance of microdifferentiation.

Specimens were classified upon arrival at the laboratory. Whenever possible the number of *Drosophila melanogaster* individuals analyzed per location and date was eighty, although this number could not always be attained. Males were used preferentially and females were only added to the sample whenever necessary to achieve the aimed sample number (homogeneity of gene frequencies between both sexes has been previously proven for both enzyme loci). Wild caught males or females were electrophoresed on starch gel and their genotype for *Adh* and α *Gpdh-1* established by differential staining of the gel (Ursprung & Leone, 1965; Grell, 1967).

Results and discussion

Genotypic frequencies

Genotypic frequencies were obtained directly from electrophoretic patterns of *Adh* and α *Gpdh-1* (Ta-

ble 1). Goodness-of-fit to random-mating proportions was tested for each locus independently (chi-squared values given in Table 2). For *Adh*, seven tests out of fifty-four show a significant departure from equilibrium and for α *Gpdh-1* five out of fifty-four. Within the significant cases there is no regularity for any particular location or date.

When data from different panmictic units are pooled, an excess of homozygotes is expected (Wahlund’s effect). Genotypic frequencies from samples obtained each day in different locations of the cellar were pooled and the pooled data tested for a possible departure from Hardy-Weinberg equilibrium. Only four of the eighteen tests performed for *Adh* evidenced a significant excess of homozygotes and only two out of eighteen did so for α *Gpdh-1* (Table 3). In all cases the excess of homozygotes can be explained by a significant departure from Hardy-Weinberg equilibrium from at least one of the pooled sets of data. One could therefore talk of a generalized absence of Wahlund’s effect, which could be considered as a first indication of the cellar being a single panmictic unit as far as the two loci studied are concerned.

Gene frequencies

More than 4500 flies were assayed for *Adh* and α *Gpdh-1* from collections made in 1979 and 1980. Gene frequencies for each collecting day and site are given in Table 1.

Homogeneity of gene frequencies for samples collected each month was tested by means of a non-balanced analysis of variance with the two variable site and data fixed (BMDP, 4V). No significant influence either of site or date was observed (Table 4), which would rule out any microdifferentiation in the population. In December, homogeneity of gene frequencies has been assessed by means of a χ^2 test (for *Adh* $\chi^2_1 = 0.57$, for α *Gpdh-1* $\chi^2_1 = 3.22$). When considering *Adh*, this homogeneity of gene frequencies for sites inside and outside the cellar and both for vintage and non-vintage periods would be in agreement with data by McKenzie and McKechnie (1978) and by Oakeshott and Gibson (1981). These authors did not find any difference in gene frequencies for *Adh* between samples taken inside and out-

Table 1. Gene frequency estimates of samples collected in different locations inside and outside the cellar (number of individuals in parentheses). Abbreviations: Ma, filtering machine; Pr, grape press; Pi, pomost in a small building; Po, pomost outside the cellar; Ba, leaking barrel.

| Month | Day | Ma | Pr | Pi | Po | Ba |
|---|-----|------------|------------|------------|-----------|------------|
| <i>Adh^s</i> | | | | | | |
| October | 1 | 0.279(95) | 0.265(97) | 0.240(77) | 0.266(31) | |
| | 2 | 0.210(62) | 0.150(10) | 0.244(88) | 0.274(32) | |
| | 3 | 0.298(67) | 0.233(75) | 0.134(82) | 0.313(32) | |
| | 4 | 0.186(51) | 0.333(81) | 0.222(81) | 0.159(41) | |
| | 5 | 0.223(92) | 0.150(90) | 0.237(78) | 0.277(83) | |
| November | 1 | 0.146(82) | 0.173(84) | 0.155(90) | | 0.237(78) |
| | 2 | 0.227(86) | 0.310(100) | 0.193(106) | | 0.190(87) |
| | 3 | 0.188(88) | 0.204(81) | 0.189(90) | | 0.135(89) |
| | 4 | 0.199(78) | 0.293(82) | 0.207(87) | | 0.163(89) |
| December | 1 | 0.228(46) | | | | 0.193(166) |
| June | 1 | 0.145(100) | | | | 0.117(90) |
| | 2 | 0.165(100) | | | | 0.134(138) |
| | 3 | 0.139(97) | | | | 0.149(87) |
| | 4 | 0.115(126) | | | | 0.125(84) |
| July | 1 | 0.097(98) | | | | 0.194(98) |
| | 2 | 0.170(97) | | | | 0.194(97) |
| | 3 | 0.250(117) | | | | 0.188(117) |
| | 4 | 0.176(88) | | | | 0.177(88) |
| <i>α-GPD1^S</i> | | | | | | |
| October | 1 | 0.552(86) | 0.461(51) | 0.558(77) | 0.518(27) | |
| | 2 | 0.450(80) | 0.600(10) | 0.563(79) | 0.586(29) | |
| | 3 | 0.440(67) | 0.493(75) | 0.507(77) | 0.328(30) | |
| | 4 | 0.560(50) | 0.422(81) | 0.512(81) | 0.634(41) | |
| | 5 | 0.414(87) | 0.500(90) | 0.500(78) | 0.470(83) | |
| November | 1 | 0.531(82) | 0.542(87) | 0.556(90) | | 0.436(78) |
| | 2 | 0.448(86) | 0.480(100) | 0.509(107) | | 0.477(87) |
| | 3 | 0.483(88) | 0.438(81) | 0.533(90) | | 0.517(89) |
| | 4 | 0.487(78) | 0.421(82) | 0.506(87) | | 0.449(89) |
| December | 1 | 0.348(46) | | | | 0.508(63) |
| June | 1 | 0.475(100) | | | | 0.450(90) |
| | 2 | 0.535(100) | | | | 0.522(138) |
| | 3 | 0.531(97) | | | | 0.500(87) |
| | 4 | 0.512(126) | | | | 0.547(84) |
| July | 1 | 0.541(98) | | | | 0.449(98) |
| | 2 | 0.474(97) | | | | 0.500(90) |
| | 3 | 0.521(117) | | | | 0.547(85) |
| | 4 | 0.517(88) | | | | 0.459(110) |

Table 2. χ^2 values (with one degree of freedom) for goodness-of-fit to random-mating proportions. See Table 1 for abbreviations.

| Month | Day | Adh | | | | | α -Gpdh-1 | | | | |
|----------|-----|--------|-------|-------|--------|-------|------------------|------|------|-------|-------|
| | | Ma | Pr | Pi | Po | Ba | Ma | Pr | Pi | Po | Ba |
| October | 1 | 8.17** | 1.28 | 0.12 | 0.19 | | 4.32* | 1.17 | 0.84 | 1.80 | |
| | 2 | 0.95 | 3.69 | 0.19 | 1.44 | | 2.21 | 0.28 | 1.97 | 0.68 | |
| | 3 | 0.36 | 3.54 | 5.77* | 0.52 | | 0.25 | 0.01 | 1.05 | 0.20 | |
| | 4 | 0.51 | 0.25 | 0.41 | 0.00 | | 0.15 | 3.53 | 0.32 | 5.59* | |
| | 5 | 2.40 | 6.71* | 1.02 | 7.42** | | 0.16 | 1.90 | 0.05 | 0.08 | |
| November | 1 | 8.23** | 3.64 | 0.43 | | 0.75 | 1.68 | 0.27 | 0.57 | | 0.29 |
| | 2 | 2.52 | 0.25 | 0.00 | | 0.62 | 1.46 | 0.26 | 1.57 | | 0.01 |
| | 3 | 0.40 | 0.06 | 2.32 | | 0.12 | 0.42 | 0.42 | 0.35 | | 0.01 |
| | 4 | 0.00 | 0.27 | 3.17 | | 1.13 | 0.05 | 1.27 | 0.35 | | 4.63* |
| December | 1 | 0.11 | | | | 0.17 | 0.08 | | | | 1.92 |
| June | 1 | 1.40 | | | | 5.65* | 1.90 | | | | 0.01 |
| | 2 | 0.41 | | | | 5.03* | 0.06 | | | | 0.69 |
| | 3 | 0.34 | | | | 0.28 | 0.91 | | | | 0.01 |
| | 4 | 0.04 | | | | 0.07 | 1.99 | | | | 1.53 |
| July | 1 | 0.29 | | | | 0.03 | 0.90 | | | | 0.01 |
| | 2 | 1.61 | | | | 0.02 | 1.68 | | | | 1.11 |
| | 3 | 2.77 | | | | 1.21 | 11.61*** | | | | 5.93* |
| | 4 | 0.96 | | | | 1.77 | 1.16 | | | | 0.70 |

* 0.01 <P<0.05

** 0.001 <P<0.01

*** P<0.001

side the cellar, although they observed a significant difference in ethanol tolerance. In contradiction, Briscoe *et al.* (1975) found that the frequency of the slow allele of *Adh* is higher in a neighbouring population than in the cellar.

This very homogeneity allowed us to pool the data and estimate mean gene frequencies for each month (Table 5). Variation in α Gpdh-1 throughout the year can be considered random, ($\chi^2_4 = 3.948$). For *Adh*, however, there are significant differences not only when a general test year-round is performed ($\chi^2_4 = 72.428$) but also when any two consecutive sampling months are compared – except November vs. December – (Table 6).

If we consider for *Adh* three additional samples collected in the same cellar early October 1978, late May 1979 (Aguadé & Serra, 1980) and early October 1980 (unpublished data), a cyclical change in gene frequencies becomes apparent (Fig. 1) – the frequency of *Adh*^S is minimum just after winter and

maximum at the end of the summer during vintage. This seasonal change would be in agreement with already published data on latitudinal (Pipkin *et al.*, 1973; Johnson & Schaffer, 1973; Vigue & Johnson, 1973; Voelker *et al.*, 1978; Oakeshott *et al.*, 1982) and altitudinal clines (Malpica & Vassallo, 1980), as well as with the kinetic properties of both alleles with regard to temperature (Johnson & Powell, 1974). The common conclusion of all these studies is that the slow allele of *Adh* seems to have a selective advantage at higher temperatures. When shifting from a spatial to a temporal dimension, one would expect an increase – if any – in the frequency of the slow allele in the warmer months and a decrease at the end of the summer. Both Johnson and Burrows (1976) and Langley *et al.* (1977) found such an increase in the frequency of *Adh*^S from spring to winter when surveying repeatedly the same population. This is in agreement with the before cited expectations and with our results. Yet, Cavener and Clegg (1981) found

Table 3. χ^2 values (with one degree of freedom) for goodness-of-fit to random-mating proportions when all samples are pooled (number of individuals in parentheses).

| Month | Day | Adh | | α -Gpdh-1 | |
|----------|-----|----------|-------|------------------|-------|
| October | 1 | 2.806 | (254) | 6.713** | (241) |
| | 2 | 0.437 | (191) | 0.303 | (178) |
| | 3 | 9.364** | (256) | 1.243 | (251) |
| | 4 | 0.653 | (254) | 2.691 | (253) |
| | 5 | 1.021 | (343) | 0.087 | (338) |
| November | 1 | 4.404* | (334) | 0.553 | (334) |
| | 2 | 1.056 | (379) | 0.040 | (379) |
| | 3 | 0.147 | (348) | 0.182 | (348) |
| | 4 | 0.619 | (336) | 3.159 | (336) |
| December | 1 | 0.001 | (109) | 2.253 | (109) |
| June | 1 | 10.265** | (198) | 0.603 | (198) |
| | 2 | 6.281 | (238) | 0.225 | (238) |
| | 3 | 1.705 | (184) | 0.366 | (184) |
| | 4 | 3.084 | (210) | 3.553 | (210) |
| July | 1 | 0.007 | (196) | 0.739 | (196) |
| | 2 | 1.919 | (187) | 0.044 | (187) |
| | 3 | 2.010 | (197) | 17.473*** | (202) |
| | 4 | 2.767 | (198) | 1.673 | (198) |

* 0.01 <P<0.05

** 0.001 <P<0.01

*** P<0.001

homogeneity of gene frequencies for *Adh* and α *Gpdh-1* when surveying a North American population for a period of two years.

One should point out the existence of a strong association of alleles *Adh^S* and α *Gpdh-1^F* with inversion *In(2L)t* and also some relevant differences between North American and Spanish cellar

Table 5. Monthly mean frequencies.

| Month, year | <i>Adh^S</i> | α - <i>Gpdh^S</i> |
|---------------|------------------------|------------------------------------|
| October 1979 | 0.235 \pm 0.008 | 0.501 \pm 0.010 |
| November 1979 | 0.201 \pm 0.007 | 0.490 \pm 0.009 |
| December 1979 | 0.201 \pm 0.019 | 0.450 \pm 0.066 |
| June 1980 | 0.136 \pm 0.008 | 0.510 \pm 0.012 |
| July 1980 | 0.167 \pm 0.009 | 0.500 \pm 0.013 |

Table 6. Homogeneity χ^2 values (with one degree of freedom) for *Adh* frequencies when two consecutive sampling months are compared.

| | |
|-----------------------|-----------------------|
| October vs. November | 9.247** |
| November vs. December | 1.69 $\times 10^{-4}$ |
| December vs. June | 11.212*** |
| June vs. July | 6.345** |
| July vs. October | 27.009*** |

* 0.01 <P<0.05

** 0.001 <P<0.01

*** P<0.001

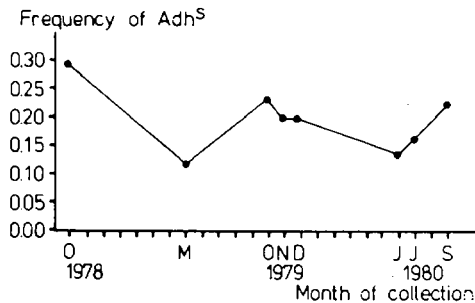


Fig. 1. Monthly gene frequencies.

Table 4. Analysis of variance of gene frequencies (using angular transformants).

| Month | <i>Adh</i> | | | | | | α - <i>Gpdh-1</i> | | | | | |
|----------|------------|--------|-------|-------|--------|-------|--------------------------|--------|-------|-------|--------|-------|
| | Day | | | Place | | | Day | | | Place | | |
| | F | d.f. | Prob | F | d.f. | Prob | F | d.f. | Prob | F | d.f. | Prob |
| October | 0.20 | (4,12) | 0.934 | 0.25 | (3,12) | 0.860 | 1.63 | (4,12) | 0.231 | 0.37 | (3,12) | 0.774 |
| November | 1.64 | (3,9) | 0.249 | 1.97 | (3,9) | 0.189 | 1.40 | (3,9) | 0.306 | 2.08 | (3,9) | 0.173 |
| June | 1.37 | (3,3) | 0.402 | 0.70 | (1,3) | 0.464 | 6.11 | (3,3) | 0.086 | 0.63 | (1,3) | 0.487 |
| July | 0.67 | (3,3) | 0.624 | 3.57 | (1,3) | 0.155 | 0.56 | (3,3) | 0.676 | 0.68 | (1,3) | 0.471 |

populations. Voelker *et al.* (1978) concluded that only 0.23 of the cline of *Adh^S* present on the east coast is due to association with *In(2L)t*, the remainder due to a cline in the frequency of *Adh^S* in the standard chromosome. Oakeshott *et al.* (1982) and Knibb (1983), on the other hand, stated that in Australia even a lower percentage of the *Adh^S* cline can be accounted for by the inversion. Samples collected in Artés in October 1978 and May 1979 – as well as samples collected in other Spanish cellars (Aguadé & Serra, 1980) – show a slightly different pattern. In fact, *Adh^S* shows the general strong association with *In(2L)t*, its conditional frequency within the inversion being close to 1.0, but in contrast with the American populations the frequency of *Adh^S* within the standard gene arrangement is very low, 0.08 and 0.0 respectively for the above mentioned collections in Artés. Even without direct information about inversions, there is no reason to believe that the situation has changed and one would expect a parallel change in the frequency of *In(2L)t* and *Adh^S*.

The assumed association between *In(2L)t* and *Adh^S* renders us unable to attribute the observed seasonal changes in the frequency of *Adh* either to the locus itself or to the inversion. The absence of a similar cyclical change in the frequency of *αGpdh-1^F*, which also shows association with the inversion and which is located within the inversion, does not help at all. The reason is first that the association with the inversion is weaker and second that the frequency of *αGpdh-1^F* is higher. Both these circumstances make that the frequency of *αGpdh-1^F* does not follow the frequency of inversion *In(2L)t*.

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