



ON SOME PROBLEMS OF OPTIMIZATION IN SAMPLE SURVEYS

**ABSTRACT
OF THE
THESIS**

SUBMITTED FOR THE AWARD OF THE DEGREE OF

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IN

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BY

NAJMUSSEHAR

UNDER THE SUPERVISION OF

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ALIGARH MUSLIM UNIVERSITY
ALIGARH (INDIA)**

2002

ABSTRACT

This thesis entitled “On some problems of optimization in sample surveys” is submitted to the Aligarh Muslim University, Aligarh, INDIA, to supplicate the degree of Doctor of Philosophy in Statistics. It embodies of research work carried out by me in the Department of Statistics and Operation Research, Aligarh Muslim University, Aligarh.

In the development of theory underlying statistical methods, one is frequently faced with optimization problems. Attempts have therefore been made to find optimization techniques that have wider applicability and can easily be implemented with the available computing power. One such technique that has the potential for increasing the scope of application of statistical methodology is mathematical programming. In this thesis an attempt has been made to formulate and solve some problems arising in sample surveys using classical optimization techniques such as Lagrange multipliers technique as well as using mathematical programming techniques.

This thesis consists of five chapters. Chapter-I provides an introduction to sample surveys with some basic results in Simple Random Sampling, Stratified Sampling, Non-Response and Double Sampling.

Chapter-II deals with the problem of allocation of a sample to strata in multivariate stratified sampling. In this chapter two new compromise allocations are proposed and compared with the already available

compromise allocations in sampling literature. As assumed by Cochran (1977), here no assumption about the correlation between the different characteristics is made. It has been shown through numerical illustrations that the proposed allocations are more precise than the already existing compromise allocations in sampling literature. This chapter is based on my paper entitled "Allocation of a sample to strata: The multivariate case" to be presented in the "National Seminar on Recent Development in Statistical Methods and Operation Research" organized by Department of Statistics, Dibrugarh University, Assam (India), during March 20-21, 2003.

In Stratified sampling the sampler has to decide about the sample sizes from various strata before drawing a sample. In sampling literature this problem is known as the problem of allocation. The equal, proportional and optimum allocations are well known allocations. In practice any one type of allocation is selected according to the prevailing situation in the population and is applied to all strata. However, there are practical situations in which the prevailing circumstances markedly differ from one group of strata to other. Hence the use of the same allocation in all the strata may not be advisable. In such situations it is proposed to divide the group of strata into non-overlapping and exhaustive subgroups according to some reasonable criterion. The use of particular allocation is advised in a particular subgroup depending upon the characteristic of the subgroup. Since different allocations are to be used in different subgroups, the proposed allocation is named as a "Mixed allocation". Chapter-III of this

thesis discusses the “Mixed allocation in Stratified Sampling”. It is assumed that the population mean is of interest. The problems of finding the mixed allocation for fixed cost and for fixed variance of the estimator of the population mean, based on a stratified sample are formulated and solved as nonlinear programming problems. The variance of the estimator under mixed allocation is worked out and compared with the variance under the overall optimum allocation. The relative increase in the variance due to the use of the mixed allocation is studied to decide whether a mixed allocation is advisable or not in a given situation. This chapter is based on my research paper entitled “ Mixed Allocation in Stratified Sampling ” to be presented in the International Conference on Statistics, Combinatorics and Related Areas Organized by Department of Statistics, University of Allahabad (India) going to be held during December 21-23, 2002.

In Chapter-IV the problem of optimum allocation in Double Sampling for stratification (DSS) with subsampling the non-respondents is formulated as a mathematical programming problem (MPP). When strata weights are not known, double sampling technique may be used to estimate them. A large simple random sample from the unstratified population is drawn and the units belonging to each stratum (in the sample) is obtained. A second stratified sample is then obtained from which a simple random subsample is drawn out of the previously selected units of the stratum. If the problem of non-response is there, then the subsamples are divided into respondents and non-respondents respectively. A second sub-sample of non-respondents

units is selected out of non-respondents and information is obtained on second attempt. The objective of the problem is to find the optimum sizes of the subsamples of non-respondents. For this in the first phase of solution the optimum values of the sample sizes are obtained for which the variance of the estimated population mean for double sampling is minimum for a fixed sample size. In the second and the final phase of solution, the optimum values of subsamples of non-respondents are obtained for fixed total cost of the survey. A solution procedure using dynamic programming technique is developed to solve the resulting MPP. The computational details of the procedure are illustrated through a numerical example.

This chapter is based on my research paper entitled “Double sampling for stratification for subsampling the non-respondents” published in Aligarh Journal of Statistics (see Najmussehar and Abdul Bari (2002)).

Chapter-V deals with the problem of optimum stratification. For stratified sampling to be efficient the strata should be as homogeneous as possible with respect to the main study variable. In other words the stratum boundaries are so chosen that the stratum variances are as small as possible. This could be done effectively when the frequency distribution of the main study variable is known. Usually this frequency distribution is unknown but it is possible to approximate it from the past experience and prior

knowledge about the population. In this chapter the problem of optimum stratification and formulated as a Nonlinear Programming Problem (NLPP) assuming exponential frequency distribution of the main study variable. The formulated NLPP is separable with respect to the decision variables and is treated as a multistage decision problem. A procedure is developed using dynamic programming technique to work out the optimum stratum boundaries. These stratum boundaries are optimum in the sense that they minimize the sampling variance of the stratified sample mean under Neyman allocation. A computer program in 'Java SDK 2' is also developed for the procedure. This computer program is executed to work out the optimum strata boundaries for a given exponential distribution to provide a numerical example.

This chapter is based on my research paper entitled "The problem of optimum stratification for exponential study variable under Neyman allocation: A Mathematical Programming Approach" to be presented in the International Symposium on Optimization and Statistics to be held in the Department of Statistics and Operations Research, Aligarh Muslim University, Aligarh (India), during December 28-30, 2002.

A comprehensive list of references, arranged in alphabetical order is also provided at the end of the thesis.



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T-5998

DEPARTMENT OF STATISTICS & OPERATIONS RESEARCH
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2002



T5998

Dedicated

To My

Parents



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CERTIFICATE

I certify that the material contained in this thesis entitled
“ON SOME PROBLEMS OF OPTIMIZATION IN SAMPLE
SURVEYS”, submitted by Miss Najmussehar for the award of
the Degree of Doctor of Philosophy in Statistics is original.

This work has been done under my supervision. In my
opinion the work is sufficient for consideration for the award of the
Ph. D. degree in Statistics to the candidate.

**(DR. MOHAMMAD JAMEEL AHSAN)
SUPERVISOR**

PREFACE

PREFACE

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CHAPTER-I

INTRODUCTION

CHAPTER-I

INTRODUCTION

1A: SAMPLING

1A.1 INTRODUCTION TO SAMPLING

Sampling is used in all kinds of investigations in our every day life. It is the selection and study of a part of an aggregate material to represent the whole. Deming (1950) describes sampling as “The science and art of controlling and measuring reliability of useful statistical information through the theory of probability”.

A sampling method is a scientific and objective procedure of selecting units from a population and provides a sample that is representative of the population.

The sampling procedures discussed in this thesis are procedures of random sampling or probability sampling. All random sampling procedures satisfy the following properties.

- (i) A set of distinct samples are available which the procedure is capable of selecting if applied to a specific population.

- (ii) Each possible sample has assigned to it known probability of selection.
- (iii) A selection procedure is available in which each sample receives its assigned probability of selection.
- (iv) The method of constructing the estimate from the sample must lead to a unique value for a specified sample. (See Cochran (1977)).

Other sampling procedures that do not possess the above properties are called non-random or non-probability sampling procedures and are out of the scope of this thesis.

1A.2 USES OF SAMPLING

Sample surveys are widely in use in all most all walks of life in a variety of ways all over the world. The objective of a sample survey may be to obtain some measure with respect to the characteristic of the whole population under study. For example for national planning and socio-economic development the governments need information about agricultural production, utilization of land and water resources, industrial production, unemployment, labor force, whole sale and retail prices of various commodities, income and expenditure per household, number of literate persons and school going children, health status of people

etc, which can be obtained efficiently through sample surveys.

Sampling methods are also used in census or complete enumerations. In fact, except for certain basic information, all other data in a census are collected on sampling basis, which results in much earlier publication of the census report and substantial savings in terms of money and time.

Sampling methods are used extensively in business and industry to increase operational efficiency. Market research is heavily dependent on sample surveys. Manufacturers and retailers can have an idea of the reactions of people to new products, their complaint about old products and the reasons for preferring one product to another, through sample surveys.

Sampling methods are also used in experimental investigations. For example in determining the quality of milk, response of fertilizers to various crops, the composition of the soil etc.

1A.3 SAMPLING DESIGNS

Various random sampling procedures that can be applied to the population under study according to the aims and objectives of the survey and the nature and variation in the population are also termed as sampling designs. The commonly used sampling designs

are:

- (i) Simple Random Sampling (SRS)
- (ii) Stratified Sampling
- (iii) Cluster Sampling
- (iv) Systematic Sampling
- (v) Two-Stage Sampling
- (vi) Sub-Sampling or Multistage Sampling etc.

A sampling procedure or design may be carried out with replacement or without replacement. In with replacement (WR) sampling the selected unit is replaced before the next draw, whereas in without replacement (WOR) sampling the unit once selected is not considered for further draws. In a WR sample a population unit may appear more than once, while in a WOR sample all the selected units are distinct. Obviously a WOR sample contains more information about the population as compared to a WR sample. In this thesis the discussions are limited to WOR sampling.

Some times the population characteristic under study is strongly correlated to another characteristic called the auxiliary characteristic. The data on this auxiliary characteristic is either available or can be easily collected for all the units in the population. This auxiliary information (data on auxiliary

characteristic) may be used to improve the quality (precision) of the estimates of the population parameters obtained from a sample. Some methods that uses the auxiliary information are:

- (i) Ratio Method
- (ii) Regression Method
- (iii) Double Sampling or Two-Phase Sampling Method

Out of the sampling designs pointed out in this section the first two namely Simple Random Sampling and Stratified Sampling are the most commonly used sampling designs. In the following two sections the basic results of these sampling design are stated.

1A.4 SIMPLE RANDOM SAMPLING WITHOUT REPLACEMENT (SRSWOR)

It is the simplest method of random sampling. It is a method of selecting n units out of N such that every possible distinct sample of size n has an equal chance of being selected.

Let y_i be the value of the characteristic under study for the i th unit of the population/ sample.

Further, let

$$\bar{Y} = \frac{1}{N} \sum_{i=1}^N y_i \quad \text{the population mean}$$

$$S^2 = \frac{1}{N-1} \sum_{i=1}^N (y_i - \bar{Y})^2 \quad \text{the population mean square}$$

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i \quad \text{the sample mean}$$

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2 \quad \text{the sample mean square}$$

The following are the basic results in SRSWOR.

- (1) The sample mean \bar{y} is an unbiased estimate of the population mean \bar{Y} .
- (2) The sample mean square s^2 is an unbiased estimate of the population mean square S^2 .
- (3) The sampling variance of the sample mean \bar{y} is

$$V(\bar{y}) = \left(\frac{1}{n} - \frac{1}{N}\right) S^2.$$

- (4) $v(\bar{y}) = \left(\frac{1}{n} - \frac{1}{N}\right) s^2$ is an unbiased estimate of $V(\bar{y})$.

1A.5 STRATIFIED SAMPLING

It is the most popular and widely used sampling design. In this sampling procedure the population is divided in non-overlapping and exhaustive groups of units. These groups are called strata. Independent WOR simple random samples are then

drawn from each stratum. Let there be L strata. The following symbols refer to the stratum h (h=1,2,...,L):

N_h	number of units in the stratum (the stratum size)
n_h	number of units in the sample (the sample size)
y_{hi}	value obtained from the ith unit
$W_h = \frac{N_h}{N}$	stratum weight
$\bar{Y} = \frac{1}{N_h} \sum_{i=1}^{N_h} y_{hi}$	stratum mean

$$S_h^2 = \frac{1}{N_h - 1} \sum_{i=1}^{N_h} (y_{hi} - \bar{Y}_h)^2 \quad \text{stratum variance}$$

Also let

$$\begin{aligned} \bar{Y} &= \frac{1}{N} \sum_{h=1}^L \sum_{i=1}^{N_h} y_{hi} \\ &= \sum_{h=1}^L W_h \bar{Y}_h \end{aligned} \quad \text{the over all population mean}$$

$$\bar{y}_{st} = \sum_{h=1}^L W_h \bar{y}_h \quad \text{the stratified sample mean}$$

The following are the basic results.

(1) \bar{y}_h and s_h^2 are the unbiased estimates of \bar{Y}_h and S_h^2 respectively.

(2) The sampling variance of \bar{y}_h is

$$V(\bar{y}_h) = \left(\frac{1}{n_h} - \frac{1}{N_h}\right) S_h^2.$$

(3) An unbiased estimate of $V(\bar{y}_h)$ is

$$v(\bar{y}_h) = \left(\frac{1}{n_h} - \frac{1}{N_h}\right) s_h^2.$$

(4) \bar{y}_{st} is an unbiased estimate of \bar{Y} .

(5) The sampling variance of \bar{y}_{st} is

$$V(\bar{y}_{st}) = \sum_{h=1}^L \left(\frac{1}{n_h} - \frac{1}{N_h}\right) W_h^2 S_h^2.$$

(6) An unbiased estimate of $V(\bar{y}_{st})$ is

$$v(\bar{y}_{st}) = \sum_{h=1}^L \left(\frac{1}{n_h} - \frac{1}{N_h}\right) W_h^2 s_h^2.$$

The use of stratified sampling design involves the following four decision-making design operations.

- (1) The choice of the stratification variable.
- (2) The choice of the number L of strata.
- (3) The choice of the stratum boundaries.

- (4) The choice of the size n_h of the sample from the h th stratum.

The discussion of all the operations is beyond the scope of the thesis. However, the last two operations are discussed in detail in the subsequent chapters of this thesis.

A chapter of this thesis is devoted to the use of double sampling to deal with the problem of non-response. In the next two sections these topics are introduced in brief.

1A.6 DOUBLE SAMPLING

As given in section 1.3, a number of sampling techniques use the auxiliary information. Ratio and regression methods require knowledge of the population mean \bar{X} of the auxiliary variable x . If the auxiliary variable x is to be used as the stratification variable its frequency distribution must be known. When such information is not available the technique of double sampling (or two-phase sampling) may be used to obtain the auxiliary information. In double sampling a large preliminary sample is taken in which the auxiliary variable x alone is measured and a reasonably good estimate of the auxiliary information required is obtained.

The double sampling may be appropriate when the cost of

measuring the auxiliary variable is significantly low as compared to the main variable.

1A.7 ERRORS IN SURVEYS

In sample surveys there will always be a difference between the population value and its estimate. This error is due to the sampling itself, that is due to the none-enumeration of the entire population and thus is called sampling error.

Other errors in the surveys arising in the collection, processing, compiling and analysis of the data are called non-sampling errors.

Non-sampling errors can further be classified into response error and non-response error. Errors of measurement on a unit due to the use of faulty or biased measuring device and errors introduced in editing coding and tabulating the results are called response errors. Whereas the error due to the failure in measuring some units selected in the sample, which results in an incomplete sample data, are called non-response error.

1A.8 NON-RESPONSE

The non-response refers to the failure to measure some of the units selected in the sample. In surveys it is commonly experienced that complete data from the sampling units is often

not obtainable for various reasons. For example, in an opinion survey, the selected family might have shifted to some other place, selected person might have died. In mailed questionnaire, some of the selected addresses may be wrong or they do not reply. Such a problem of incomplete sample data is known as the problem of non-response in sampling literature.

One way to deal with the problem of non-response is to assume that the population consists of two strata, one of the respondents, on which the information is available and the other of the non-respondents, on which the information is not available at first attempt. A sub sample is drawn out of the sampled units falling in non-respondents stratum and a second and extensive attempt is made to obtain information on these units. The information obtained on the two attempts is then pooled to construct the required estimate.

1B: OPTIMIZATION

1B.1 INTRODUCTION

Optimization is the act of obtaining the best result under given circumstances. The efforts required or the benefits desired in any practical situation can often be expressed as a function of some decision variables. The ultimate goal of such decision is

either to maximize the benefit desired or to minimize the loss or cost incurred or efforts required. Mathematically, optimization is the maximization or minimization of a function of several variables. These variables may be unconstrained or subjected to certain constraints in the form of equations or/and inequalities. There is no single method available for solving all optimization problems. A number of optimization methods are developed for solving different types of optimization problems. The constrained optimization techniques are also known as mathematical programming methods or techniques.

1B.2 A BRIEF HISTORICAL SKETCH

The existence of optimization methods can be traced back to the days of Newton, Lagrange and Cauchy. But in spite of these early contributions very little progress has been made until the middle of the nineteenth century, when the high-speed digital computers made the implementation of the optimization procedures possible and stimulated further research on new methods.

Constrained optimization or mathematical programming has developed rapidly during and after World War II as a new field of study dealing with applications of the scientific method of

business operations and management decision-making. Mathematical programming problems can be broadly classified as (i) Linear Programming Problems (LPP) when all the involved functions are linear and (ii) Nonlinear Programming Problems (NLPP), when all the involved functions are not linear.

In 1947 the United States Air Force team SCOOP (Scientific Computation of Optimum Programs) started intensive research on some optimum resource allocation problem that led to the development of the famous simplex method by George B. Dantzig for solving a linear programming problem (LPP). The simplex method is an iterative procedure, which yields an exact optimal solution in a finite number of steps. But the method was not available until it was published in the Cowles Commission Monograph No. 13 in 1951.

One of the earliest enterprises undertaken by the exponents of mathematical programming grew out of the problems involved in the war mobilization program of the 1940's. The problems of planning and co-coordinating among various project and optimum allocation of limited resources to obtain the desired result were emerged as the basic problems.

Kuhn, H.W. and Tucker, A.W. (1951) derived the necessary conditions for the optimal solution of a constrained optimization

or mathematical programming problem. These conditions (popularly known as K-T conditions) laid the foundation of a great deal of later research and development in the area of non-linear programming.

No single technique (like simplex method for solving LPP) is available till date for solving NLPP. However different methods are available for solving some special types of NLPP. Beale (1959) developed a method for solving convex quadratic programming problem (CQPP). Wolfe (1959), using the K-T conditions, transformed the CQPP into equivalent LPP with an additional non-linear restriction to which simplex method could be applied. Other authors who gave the technique for solving QPP are Markowitz (1956), Hilderth (1957), Houthaker (1960), Lemke (1962), Van de Panne and Whisnton (1964a, 1964b, 1966), Graves (1967), Fletcher (1971), Aggarwal (1974a, 1974b), Finkbeiner and Kall (1978), Arshad, Khan and Ahsan (1981). Ahsan, Khan and Arshad (1983), Todd (1985), Fukushima (1986), Yuan (1991), Wei (1992), Benzi (1993), Anstreicher, Den Hertog and Terlaky (1994) and Several others.

Rosen (1960, 1961), Kelly (1960), Goldfarb (1969), Du, Wu and Zhag (1990), Lai, Gao, and He (1993) developed Gradient projection methods for solving NLPP with linear and nonlinear

constants. This is an iterative procedure in which at each step we move from one feasible solution to another in such a way that the value of the objective function is improved.

A linear fractional programming technique was proposed by Charnes and Cooper (1962). The algorithms for solving non-linear fractional programming were developed by Dinkelbach (1967) and Mangasrian (1969).

Geometric programming provides a systematic method for formulating and solving the class of optimization problems that tend to appear mainly in engineering designs. This technique was first developed by Duffin, Peterson and Zener (1967). Ermer (1971) used geometric programming for optimization of the constrained machinery economic problem. His work was further extended by Dembo (1982), Kortanek and Hoon (1992), Yeh (1993) and several others.

Dantzig (1959), Charnes and Cooper (1959, 1960) developed stochastic programming techniques. Some other authors who worked on stochastic programming are Shapiro (1990), Weintraub and Vera (1991), Flam and Schult (1993), Schoen (1994) and Bahn et al. (1995) etc.

A technique known as goal programming for solving multi-objective linear and non-linear programming problems was

developed by Charnes and Cooper (1977). Other authors who made contribution for solving multiobjective linear and non-linear programming problems are Sherali (1982), Roy and Wallenius (1992), Arbel (1993, 1994), Bit, Biswal and Alam (1993) and Okada (1993) etc.

Dynamic programming technique, based on the principle of optimality, was developed by Richard Bellman (1957). This technique is applicable to mathematical programming problems having some special features. Several others who contributed significantly to this area are Bellman and Dreyfus (1962), Wachs (1989), Li (1990), Li and Haimes (1990), Wang (1990a, 1990b) Wang and Xing (1990), Lin (1994), Badinelli (2000) etc.

Developments of new techniques for solving mathematical programming are still going on. To cover all of them is beyond the scope of this thesis.

In this thesis dynamic programming technique is used for solving some of the optimization problems arising in sampling. The following section gives a brief account of the dynamic programming technique.

1B.3 DYNAMIC PROGRAMMING TECHNIQUE

The problems requiring sequential decision-making at different stages may be called multistage decision problems. The problem of making a set of optimal decisions may be formulated as an MPP. The dynamic programming technique is a procedure, which can handle the problem of optimal decision-making at various stages of a multistage decision problem. The general nature of the MPP that can be handled by this technique may be described as follows.

- (i) The MPP can be treated as a multistage decision problem. At each stage the value(s) of one or more decision variables are to be determined.
- (ii) The MPP must have the same structure at every stage irrespective of the number of stages.
- (iii) At every stage the values of the decision variables and the objective function must depend on a specified set of parameters describing the state of system. These parameters are called the state parameters.
- (iv) Same set of state parameters must describe the state of the system irrespective of the number of stages.
- (v) The decision at any stage must have no effect on the decisions to be made at the remaining stages except in

changing the values of the state parameters.

In solving an MPP by dynamic programming technique we start with a one-stage problem, moving on to a two-stage problem, to a three-stage problem and so on until all stages are included. The final solution is obtained by adding the n th (final) stage to the solution of $(n-1)$ stage. For this a relation between the two successive stages is defined. This relation is called the “Recurrence Relation” of dynamic programming.

The computational efficiency of the dynamic programming technique as compared to the complete enumeration is very impressive. But unfortunately the computational efforts involved in solving an MPP by dynamic programming technique multiply incredibly fast with the increase in the number of state parameters (number of constraints). The number of state parameters is called the dimensionality of the MPP. The problem of handling the great bulk of computation in dynamic programming technique is termed as the “Problem of Dimensionality” or the “Curse of Dimensionality” to dynamic programming.

Bellman and Dreyfus (1962) suggested a procedure to reduce the dimensionality of the problem.

However, as far as the problems discussed in this thesis are concerned dimensionality poses no threat to the convergence of

computational procedures developed using dynamic programming technique.

1B.4 APPLICATIONS OF OPTIMIZATION TECHNIQUES

During the last five decades attempts have been made to develop suitable and efficient optimization techniques that can be easily implemented with the available computing power to solve various optimization problems. The early applications of optimization techniques were limited to problems involving military operations. Later on they are widely used in dealing with the optimization problems in almost every walk of life. In recent past the optimization or the mathematical programming techniques (as they are popularly known) are successfully used in solving a variety of constrained optimization problems arising in Planning, Business, Industry, Economics, Commerce, Biological and Medical Services, Agriculture, Environmental Protection, Artificial Intelligence, Space Research, Engineering, Information Technology, Statistics etc etc.

1B.5 OPTIMIZATION TECHNIQUES IN STATISTICS

According to C.R. Rao (See Arthanari and Dodge (1981)) all statistical procedures are, in the ultimate analysis, solutions to suitably formulated optimization problems. Whether it is

designing a scientific experiment or planning a large scale survey for collection of data, or choosing a stochastic model to characterize observed data, or drawing inference from available data, such as estimation, testing of hypothesis and decision making, one has to choose an objective function and minimize or maximize it subject to given constraints on unknown parameters and inputs such as the cost involved. The classical optimization methods based on differential calculus are too restrictive and are either inapplicable or difficult to apply in many situations that arise in statistical work. This together with the lack of suitable numerical algorithms for solving optimizing equations has placed several limitations on the choice of objective functions and constraints and led to the development and use of some inefficient statistical procedures.

Attempts have therefore been made during the last five decades to find other optimization techniques that have wider applicability and can be easily implemented with the available computing power. One such technique that has the potential for increasing the scope for application of efficient statistical methodology is mathematical programming. Although endowed with a vast literature, this method has not come into regular use in statistical practice mainly because of lack of good expositions

integrating the techniques of mathematical programming with statistical concepts and procedures.

A few successful applications of optimization or mathematical programming techniques to the problems arising in statistical analysis are given below.

Jesen (1969), Rao (1971), Buhler et al (1975), Littschwager and Wang (1978) in cluster analysis.

Foody and Heydayat (1977), Arthanari and Dodge (1978), Whitaker, Thriggs and John (1990) in construction of BIB designs.

Barankin (1951), Dantzig and Wald (1951) Francis and Wright (1969), Kraft (1970), Meeks and Francis (1973), Pukelshein (1978), Kabe (1989), Ozturk (1991) in testing of statistical hypothesis.

Neuhardt, Bradely and Henning (1973) in optimal design of multifactor experiments.

Chakraborty (1986, 1988, 1990, 1991), Gosh (1989), Seidel (1991), Crowder (1992) in quality control.

Tillman, Hwang and Kuo (1977) in reliability theory etc etc.

1B.6 OPTIMIZATION TECHNIQUES IN SAMPLING

The basic need of present day society is the need of reliable data to understand better the world in which we live. Such data can

only be collected through sample surveys. The fundamental problem in sample surveys is to choose a sampling design that either gives the maximum precision within available budget or minimizes the cost of survey for a prefix level of tolerance regarding the precision. Thus the base of sample survey methodology is an optimization problem. The cost of the sample survey and the precision of estimates are function of sample size. Thus the problem of deriving statistical information on population characteristics based on sample data can be formulated as an optimization problem. In stratified sampling the problem of determining the optimum number of strata, the problem of fixing optimum stratum boundaries, the problem of obtaining optimum allocations to sample sizes from various strata are optimization problems that can be formulated and solved as mathematical programming problems.

In multivariate surveys where more than one characteristic are to be measured on each and every unit of the selected sample the problem of working out optimum sample size (or sizes in case of stratified sampling) can be formulated as a multi objective optimization problem.

When two or more sample surveys are conducted on the same population, the same population unit may be assigned

different probabilities for different surveys. In such situations we may want to maximize the expected number of common units in the selected sample for different surveys for the given probabilities of selection. This is called integration of surveys. Thus the problem of optimum integration of surveys is also an optimization problem.

Some successful applications of optimization techniques in the problems arising in sample surveys are due to:

Stock and Frankel (1939), Ghosh (1958), Aoyama (1963), Kokan (1963), Folk and Anle (1965), Ericson (1967), Kokan and Khan (1967), Kish (1967), Chatterjee (1966, 1967, 1968, 1972), Murthy (1967), Raj (1969), Chaddha et al (1971), Ahsan (1975,1978), Ahsan and Khan (1977,1982), Cochran (1977), Omule (1985), Bethal (1989), McCallion (1992), Sheela and Unnithan (1992), Kreinbrock (1993), Rahim and Currie (1993), Jahan et.al. (1994), Mandowara (1994), Jahan and Ahsan (1995), Csenki (1997), Khan et al (1997, 2002a), Clark and Steel (2002), Bretthauer, Ross and Shetty (1999), etc in optimum allocation of sample sizes.

Dalenius and Gurney (1951), Dalenius and Hodge (1959), Ghosh (1963), Sethi (1963), Hartley (1965), Herleker (1967), Serfling (1967), Buhler, Aachen and Mannhein (1975), Singh

(1977), Unnithan (1978), Jarque (1981), Khare (1987), Miles et al (1987), Miles and Robert (1989), Rahim and Jocelyn (1994), Chernayak and Starytskyy (1998), Chernayak and Chornous (2000), Jahan et al (2001), Khan et.al. (2002b) etc in optimum stratification.

Dalenius (1957), Cochran (1963), Serfling (1968), Khan et al (1995) etc in determining optimum number of strata.

Allredge and Amstrong (1974) in estimation of overlap size created by interlocking sampling process.

Kefitz (1951), Lahiri (1954), Murthy (1967), Raj (1969), Arthanari and Dodge (1981), Mitra and Pathak (1984), Aragon and Pathak (1990), Fahim and Pathak (1992) etc in optimum integration of surveys.

CHAPTER-II

SOME NEW COMPROMISE ALLOCATIONS IN MULTIVARIATE STRATIFIED SAMPLING DESIGNS

CHAPTER II

SOME NEW COMPROMISE ALLOCATIONS IN MULTIVARIATE STRATIFIED SAMPLING DESIGNS

2.1 INTRODUCTION

In stratified random sampling the value of the sample sizes for various strata are to be chosen in advance. In sampling literature, the problem of selecting the sample sizes for various strata is termed as an allocation problem. The sample sizes may be chosen to minimize the variance of the estimate for a fixed total cost of the survey or to minimize the total cost of the survey for a given precision of the estimate. Equal, proportional and optimum allocations are well known in sampling literature.

When several characteristics (say 'p') are to be measured on each selected unit of the sample, the problem of optimum allocation becomes more complicated because there is no single optimality criterion through which we can attack the allocation problem. In such situations we need a suitable compromise criterion to workout a usable allocation which is optimum in some sense for all characteristics. This allocation may be called a compromise allocation because it is based on a compromise criterion.

In this chapter the already existing compromise allocations are

discussed and two new compromise allocations are proposed that are more precise than their existing counter parts.

2.2 STRATIFIED SAMPLING

Let a population of size N is be divided into L strata. The following symbols refer to stratum h ($h= 1,2, \dots,L$).

N_h	stratum size (number of units in the stratum)
n_h	sample size (number of units selected in the sample)
y_{hi}	value obtained for the i th units
$W_h = \frac{N_h}{N}$	stratum weight
$\bar{Y}_h = \frac{\sum_{i=1}^{N_h} y_{hi}}{N_h}$	stratum mean
$\bar{y}_h = \frac{\sum_{i=1}^{n_h} y_{hi}}{n_h}$	sample mean
$S_h^2 = \frac{\sum_{i=1}^{N_h} (y_{hi} - \bar{Y}_h)^2}{N_h - 1}$	stratum variance

If the estimation of the population mean per unit $\bar{Y} = \frac{\sum_{h=1}^L \sum_{i=1}^{N_h} y_{hi}}{N} = \sum_{h=1}^L W_h \bar{Y}_h$ is of

interest then it is well known that the stratified sample mean

$\bar{y}_{st} = \sum_{h=1}^L W_h \bar{y}_h$ serves as an unbiased estimate of \bar{Y} with a sampling variance:

$$V(\bar{y}_{st}) = \sum_{h=1}^L \frac{W_h^2 S_h^2}{n_h} - \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} \quad (2.1)$$

The total cost C of the survey may be express as

$$C = c_o + \sum_{h=1}^L c_h n_h$$

$$\text{or } C_o = \sum_{h=1}^L c_h n_h \quad (2.2)$$

where $C_o = C - c_o$, c_o represents the overhead cost and c_h represents the per unit measurement cost for the hth stratum.

2.3 PROPORTIONAL ALLOCATION

The allocation in which n_h are proportional to N_h is called the proportional allocation and was originally proposed by Bowley (1926). Under proportional allocation

$$n_h \propto N_h$$

$$\text{or } n_h = KN_h$$

where K is the constant of proportionality.

Substituting this value of n_h in (2.2)

$$C_o = \sum_{h=1}^L Kc_h N_h$$

$$\text{or } K = \frac{C_o}{\sum_{h=1}^L c_h N_h}$$

$$\text{Thus } n_h = \frac{C_o N_h}{\sum_{h=1}^L c_h N_h} = \frac{C_o W_h}{\sum_{h=1}^L c_h W_h}; h = 1, 2, \dots, L \quad (2.3)$$

If $c_h = c$ for all h then (2.3) gives

$$n_h = nW_h; h = 1, 2, \dots, L \quad (2.4)$$

where $n = \frac{C_o}{c}$, is the total sample size.

Expression (2.4) gives the proportional allocation for fixed total sample size.

Under proportional allocation fixed cost the sampling variance of \bar{y}_{st} is given by

$$V(\bar{y}_{st})_{prop} = \frac{(\sum_{h=1}^L W_h S_h^2)(\sum_{h=1}^L c_h W_h)}{C_o} - \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} \quad (2.5)$$

and for fixed total sample size

$$V(\bar{y}_{st})_{prop} = \frac{\sum_{h=1}^L W_h S_h^2}{n} - \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} \quad (2.6)$$

2.4 OPTIMUM ALLOCATION:

Stuart (1954) used Cauchy-Schwarz inequality to show that the

$$\text{expression: } n_h = n \cdot \frac{W_h S_h / \sqrt{c_h}}{\sum_{h=1}^L W_h S_h / \sqrt{c_h}}; h = 1, 2, \dots, L \quad (2.7)$$

gives the values of n_h that minimize (i) $V(\bar{y}_{st})$ when the total cost C is fixed and (ii) C when the variance $V(\bar{y}_{st})$ is fixed.

Allocation given in (2.7) is known as optimum allocation.

If the total cost is fixed then the total sample size n is given by

$$n = \frac{(C - c_o) \sum_{h=1}^L (W_h S_h / \sqrt{c_h})}{\sum_{h=1}^L (W_h S_h \sqrt{c_h})} \quad (2.8)$$

The expression (2.8) is obtained by substituting the values of n_h from (2.7) in (2.2). On the other hand if the variance is fixed then n is given by

$$n = \frac{N \left(\sum_{h=1}^L W_h S_h \sqrt{c_h} \right) \sum_{h=1}^L (W_h S_h / \sqrt{c_h})}{NV + \sum_{h=1}^L W_h S_h^2} \quad (2.9)$$

where V is the fixed value of the variance $V(\bar{y}_{st})$.

The expression (2.9) is obtained by substituting the values of n_h from (2.7) in (2.1) (See Cochran (1977)).

If $c_h = c$ for all h , that is the per unit measurement cost is same in each stratum then total cost C given by (2.2) becomes $C = c_o + cn$. In this case the optimum allocation for fixed cost reduces to the optimum allocation for fixed sample size n and we have the allocation problem as

“Minimize $V(\bar{y}_{st})$

subject to $\sum_{h=1}^L n_h = n$ ”

Neyman (1934) showed that $V(\bar{y}_{st})$ is minimum for fixed n if n_h are given by

$$n_h = n \cdot \frac{W_h S_h}{\sum_{h=1}^L W_h S_h}; h = 1, 2, \dots, L. \quad (2.10)$$

Therefore, n_h given by (2.10) is sometimes called the Neyman allocation.

The variance under Neyman allocation is given as:

$$V(\bar{y}_{st})_{\min} = \frac{(\sum_{h=1}^L W_h S_h)^2}{n} - \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} \quad (2.11)$$

Usually the total cost of the survey is fixed in advance. Hence here in after by an allocation we mean the allocation that minimizes $V(\bar{y}_{st})$ for fixed total cost of the survey.

2.5 PROBLEM OF ALLOCATION: THE MULTIVARIATE CASE

When several characteristics (say ‘p’) are to be measured on each selected unit of the sample the problem of optimum allocation becomes more

complicated. In such cases S_h^2 and c_h may vary from stratum to stratum as well as from characteristic to characteristic and the optimum allocation given by (2.5) becomes

$$n_{hj} = n \frac{W_h S_{hj} / \sqrt{c_{hj}}}{\sum_{h=1}^L W_h S_{hj} / \sqrt{c_{hj}}}; h = 1, 2, \dots, L; J = 1, 2, \dots, p. \quad (2.12)$$

where

n_{hj} = sample size for measuring j th characteristic; $J = 1, 2, \dots, p$ in h th stratum;

$h = 1, 2, \dots, L$

S_{hj}^2 = stratum variance of the j th characteristic in the h th stratum.

c_{hj} = per unit cost of measuring the j th characteristic in the h th stratum.

For different characteristics there are different sets of optimum allocations. In such cases n_{hj} given by (2.10) can be arranged as an $(L \times p)$ matrix whose j th column represents the optimum allocation with respect to the j th characteristic. Hence there is no unique set of values of n_h that minimizes all the variances

$V(\bar{y}_{jst})$, $J = 1, 2, \dots, p$ simultaneously, where

$$V(\bar{y}_{jst}) = \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_{hj}} - \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{N_h} \quad (2.13)$$

In such situations we need a single representative for each row of the matrix $((n_{hj}))$.

There are two ways to deal with this situation. One way is to select a single

representative for each row according to some reasonable criterion. And the other way is to reformulate and solve the problem of allocation in which the objective is to find n_h that minimize, some function of $V(\bar{y}_{jst})$ for fixed total cost.

In the remaining sections of this chapter the problem of optimum allocation in multivariate stratified random sampling is studied in detail and allocations proposed by various authors are discussed. Two new allocations are proposed and compared with the already existing allocations in the sampling literature through numerical examples.

2.6 COMPROMISE ALLOCATION BASED ON THE ROW REPRESENTATIVES

Since the optimum allocation with respect to different characteristics are different there is no unique set of values of $n_h; h = 1, 2, \dots, L$ that minimize every $V(\bar{y}_{jst})$, $J = 1, 2, \dots, p$, simultaneously. Therefore, for practical purposes some compromise must be reached in a multivariable survey regarding the sample sizes from various strata.

An allocation based on some compromise criterion may be called a compromise allocation (See Cochran (1977)). If the correlations between the characteristics are sufficiently high the individual optimum allocations may vary relatively little. In such situations Cochran (1977) proposed the compromise allocation based on the averages of the individual optimum

sample sizes with respect to different characteristics. If n_{hj}^* ; $h = 1, 2, \dots, L$; $j = 1, 2, \dots, p$ denote the individual optimum allocation for j th characteristic in the h th stratum then by formula (2.10).

$$n_{hj}^* = n \cdot \frac{W_h S_{hj}}{\sum_{h=1}^L W_h S_{hj}}; h = 1, 2, \dots, L; j = 1, 2, \dots, p \quad (2.14)$$

where the optimum allocation is for a fixed total sample size n . As suggested in Section 2.5 n_{hj}^* given by (2.14) can be arranged as an $(L \times p)$ matrix whose j th column represents the optimum allocation with respect to the j th characteristic.

Let $n_{h(a)}$ denote the compromise allocation based on averages, as suggested by Cochran (1977) then

$$n_{h(a)} = \frac{1}{p} \sum_{j=1}^p n_{hj}^*; h = 1, 2, \dots, L \quad (2.15)$$

where the symbol (a) stands for the average.

For the j th characteristic using (2.1) and ignoring finite population correction (fpc) the variances $V(\bar{y}_{jst})$ under this compromise allocation are given by

$$V_{J(a)} = V(\bar{y}_{jst}) = \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_{h(a)}}; J = 1, 2, \dots, p \quad (2.16)$$

Using (2.11) the variances $V(\bar{y}_{jst})$ under individual optimum allocations

ignoring fpc are given by

$$V_J^* = V(\bar{y}_{Jst})_{opt} = \frac{(\sum_{h=1}^L W_h S_{hj})^2}{n}; J = 1, 2, \dots, p \quad (2.17)$$

Using (2.4) the variances under proportional allocation ignoring fpc are given as

$$V(\bar{y}_{Jst})_{prop} = \frac{\sum_{h=1}^L W_h S_{hj}^2}{n}; J = 1, 2, \dots, p. \quad (2.18)$$

Using the data given by Jessen (1942), Cochran showed that the average allocation gives results almost as precise as if it were possible to use individual optimum allocations.

In working out the compromise allocation he assumed all characteristics equally important. The author suggests that a more precise compromise allocation may be obtained if weighted averages are used instead of simple averages of n_{hj}^* . As regards the selection of weights for various characteristics it would be reasonable to take them proportional to the respective individual optimum variances given by (2.17) that is:

$$a_j \propto V_j^*$$

$$\text{or } a_j = KV_j^*; J = 1, 2, \dots, p \quad (2.19)$$

where $a_j > 0; j = 1, 2, \dots, p$ denote the weights assigned to the individual optimum allocations n_{hj}^* and K is the constant of proportionality.

$$\text{or } \sum_{j=1}^p a_j = K \sum_{j=1}^p V_j^*$$

$$\text{or } K = \frac{\sum_{j=1}^p a_j}{\sum_{j=1}^p V_j^*} = \frac{1}{V_j^*} \quad (2.20)$$

(by putting the sum of weights $\sum_{j=1}^p a_j$ equal to 1)

where $V_j^* = V(\bar{y}_{jst})_{opt}$; $J = 1, 2, \dots, p$ are as given in (2.15).

Substituting the value of K from (2.20) in (2.19) we get

$$a_j = \frac{V_j^*}{\sum_{j=1}^p V_j^*}, J = 1, 2, \dots, p \quad (2.21)$$

where $a_j > 0$ and $\sum_{j=1}^p a_j = 1$

The weighted averages of n_{hj}^* ; $h = 1, 2, \dots, L$ as the proposed compromise allocation are thus given as:

$$\begin{aligned} n_{h(w)} &= \sum_{j=1}^p a_j n_{hj}^* \\ &= \sum_{j=1}^p \frac{V_j^*}{\sum_{j=1}^p V_j^*} n_{hj}^* \end{aligned}$$

$$= \frac{(\sum_{j=1}^p V_j^* n_{hj}^*)}{\sum_{j=1}^p V_j^*}; h = 1, 2, \dots, L \quad (2.22)$$

Now using (2.14) and (2.17) we get

$$V_j^* n_{hj}^* = \frac{1}{n} \left(\sum_{h=1}^L W_h S_{hj} \right)^2 \left(n \cdot \frac{W_h S_{hj}}{\sum_{h=1}^L W_h S_{hj}} \right)$$

$$= W_h S_{hj} \left(\sum_{h=1}^L W_h S_{hj} \right) \quad (2.23)$$

By (2.22) and (2.23)

$$n_{h(w)} = \frac{\sum_{j=1}^p W_h S_{hj} \left(\sum_{h=1}^L W_h S_{hj} \right)}{\frac{1}{n} \sum_{j=1}^p \left(\sum_{h=1}^L W_h S_{hj} \right)^2}$$

$$= n \frac{\left(\sum_{j=1}^p W_h S_{hj} \left(\sum_{h=1}^L W_h S_{hj} \right) \right)}{\sum_{j=1}^p \left(\sum_{h=1}^L W_h S_{hj} \right)^2}; h = 1, 2, \dots, L. \quad (2.24)$$

The variances ignoring (fpc) under this allocation may be obtained by substituting $n_{h(w)}$ given by (2.24) for n_h in (2.1) as

$$\begin{aligned}
V_{J(w)} = V(\bar{y}_{Jst}) &= \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n \sum_{j=1}^p (W_h S_{hj} \sum_{h=1}^L W_h S_{hj}) / \sum_{j=1}^p (\sum_{h=1}^L W_h S_{hj})} \\
&= \left(\frac{\sum_{j=1}^p (\sum_{h=1}^L W_h S_{hj})^2}{n} \right) \left(\sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{\sum_{j=1}^p W_h S_{hj} (\sum_{h=1}^L W_h S_{hj})} \right); J = 1, 2, \dots, p \quad (2.25)
\end{aligned}$$

In practice usually the values of S_{hj}^2 are not known. In such situations their usual unbiased sample estimates s_{hj}^2 may be used. All the above expressions will be exactly same in this situation except that S_{hj} is replaced by s_{hj} .

2.7 THE MINIMUM DEVIATION COMPROMISE ALLOCATION

Chatterjee (1967) used, the compromise criterion of minimizing the total proportional increase in individual optimum variances due to the use of a non-optimal allocation for obtaining a compromise allocation. He worked out the expression for the sample size n_h for the h th stratum for a fixed total sample size n as

$$n_{h(c)} = n \cdot \frac{\sqrt{\sum_{j=1}^p n_{hj}^{*2}}}{\sum_{h=1}^L \sqrt{\sum_{j=1}^p n_{hj}^{*2}}}; h = 1, 2, \dots, L \quad (2.26)$$

where the symbol (c) stands for Chatterjee. (see Cochran (1977)).

The corresponding variances are

$$V_{J(c)} = \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_{h(c)}}; j = 1, 2, \dots, p \quad (2.27)$$

We may obtain a more precise compromise allocation if instead of minimizing the total proportional increase in the individual optimum variances due to the use of a non-optimal allocation, we minimize the total deviation 'D' from the individual optimum variances.

Using (2.11) ignoring fpc and (2.17) the total deviation D may be expressed as

$$D = \sum_{j=1}^p (V_j - V_j^*)$$

$$= \sum_{j=1}^p \left(\sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_h} - \frac{(\sum_{h=1}^L W_h S_{hj})^2}{n} \right) \quad (2.28)$$

where $V_j = \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_h}$ denote the sampling variance of \bar{y}_{jst} under any

general allocation n_h and $V_j^* = \frac{(\sum_{h=1}^L W_h S_{hj})^2}{n}$ denote the sampling variance

of \bar{y}_{jst} under optimum allocation for fixed total sample size n ignoring fpc.

As $V_j \geq V_j^*; j = 1, 2, \dots, p$, the quantity inside () in (2.26) is always positive

and $D = \sum_{j=1}^p (V_j - V_j^*)$ will present the true magnitude of the total deviation of

the sampling variances of \bar{y}_{jst} from V_j^* for not using the individual optimum allocations. Thus a reasonable compromise criterion for working out the values of the compromise allocation n_h would be to minimize D subject to

$\sum_{h=1}^L n_h = n$, that is by solving the optimization problem:

$$\text{“Minimize } D \text{ given by (2.28) subject to } \sum_{h=1}^L n_h = n \text{”} \quad (2.29)$$

The problem (2.29) can be solved easily by using Lagrange multipliers technique as follows. Define the Lagrangian function ϕ as

$$\begin{aligned} \phi(n_h, \lambda) &= D + \lambda \left(\sum_{h=1}^L n_h - n \right) \\ &= \sum_{j=1}^p \left(\sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_h} - \frac{\left(\sum_{h=1}^L W_h S_{hj} \right)^2}{n} \right) + \lambda \left(\sum_{h=1}^L n_h - n \right) \end{aligned}$$

Differentiating ϕ with respect to $n_h; h=1,2,\dots,L$ and λ and equating the partial derivatives thus obtained to zero we get the following $L+1$ simultaneous equations.

$$= - \sum_{j=1}^p \frac{W_h^2 S_{hj}^2}{n_h^2} + \lambda = 0; h = 1, 2, \dots, L \quad (2.30)$$

$$\text{and } \frac{\partial \phi}{\partial \lambda} = \sum_{h=1}^L n_h - n = 0 \quad (2.32)$$

(2.32) gives

$$\lambda = \sum_{j=1}^p \frac{W_h^2 S_{hj}^2}{n_h^2}$$

$$\text{or } n_h^2 = \frac{1}{\lambda} \sum_{j=1}^p W_h^2 S_{hj}^2; h = 1, 2, \dots, L$$

$$\text{or } n_h = \frac{1}{\sqrt{\lambda}} \sqrt{\sum_{j=1}^p W_h^2 S_{hj}^2}; h = 1, 2, \dots, L \quad (2.33)$$

Substitution of the value of n_h from (2.33) in (2.32) gives

$$\frac{1}{\sqrt{\lambda}} \sum_{h=1}^L \sqrt{\sum_{j=1}^p W_h^2 S_{hj}^2} - n = 0$$

$$\text{or } n = \frac{1}{\sqrt{\lambda}} \sum_{h=1}^L \sqrt{\sum_{j=1}^p W_h^2 S_{hj}^2}$$

$$\text{or } \frac{1}{\sqrt{\lambda}} = \frac{n}{\sum_{h=1}^L \sqrt{\sum_{j=1}^p W_h^2 S_{hj}^2}} \quad (2.34)$$

Substituting the value of $\frac{1}{\sqrt{\lambda}}$ from (2.34) in (2.33) we get the

compromise allocation $n_{h(d)}$ based on minimum total deviation as

$$n_{h(d)} = n \frac{\sqrt{\sum_{j=1}^p W_h^2 S_{hj}^2}}{\sum_{h=1}^L \sqrt{\sum_{j=1}^p W_h^2 S_{hj}^2}}; h = 1, 2, \dots, L \quad (2.35)$$

where the symbol (d) stands for deviation.

The variances $V_{j(d)}$ (ignoring fpc) under this allocation can be obtained by substituting $n_h = n_{h(d)}$ in (2.1). Thus

$$V_{j(d)} = \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_{h(d)}}; j = 1, 2, \dots, p \quad (2.36)$$

2.8 NUMERICAL COMPARISONS

Example 1: Data used in the example are from Jessen (1942). The state of Iowa was divided into five geographic regions, each denoted by its major agricultural enterprise. These regions are to be used as strata in survey on dairy farming. The three items of most interest are the number of cows milked per day, the number of gallons of milk per day, and the total annual cash receipts from daily products. From a survey made in 1938, the estimated standard deviations s_{hj} within strata are shown in Table 2.1. It has been decided to fix the total sample size n as 1000. .

The proposed compromise allocation given by (2.24) based on weighted averages along with the corresponding expected variances given by (2.25) for the values given in Table 2.1 are worked out as

$n_{1(w)} = 236$, $n_{2(w)} = 246$, $n_{3(w)} = 194$, $n_{4(w)} = 115$ and $n_{5(w)} = 209$
with $V_{1(w)} = 0.0130$, $V_{2(w)} = 0.0811$, $V_{3(w)} = 76.9$ respectively.

Table 2.1

Standard deviations within strata

Stratum No. h	W_h	Cows Milked s_{h1}	Gallons of Milk s_{h2}	Receipts for Dairy Products s_{h3}
1.	0.197	4.6	11.7	332
2.	0.191	3.4	9.8	357
3.	0.219	3.3	7.0	246
4.	0.184	2.8	6.5	173
5.	0.208	3.7	9.8	279

The proposed compromise allocation based on minimum total deviation given by (2.34) and the corresponding expected variances given by (2.33) are worked out as:

$n_{1(d)} = 236$, $n_{2(d)} = 246$, $n_{3(d)} = 194$, $n_{4(d)} = 115$ and $n_{5(d)} = 209$
with $V_{1(d)} = 0.0130$, $V_{2(d)} = 0.0811$, $V_{3(d)} = 76.9$ respectively.

The sample sizes for a fixed total of 1000 under different allocations discussed in Sections 2.3 to 2.7 are summarized in Table 2.2. Table 2.3 shows the expected variances of \bar{y}_{jst} under the allocations given in Table 2.2.

If $T(\underline{n})_K$ denote the trace of the variance-covariance matrix of $\bar{y}_{jst}; j=1,2,\dots,p$ for a given allocation $(\underline{n})_K = (n_1, n_2, \dots, n_L)_K$. It is to be noted that this variance-covariance matrix will be a diagonal matrix when the

Table 2.2

Sample sizes within strata ($n=1000$)

Stratum No. h	Allocation							
	Proportional $n_{h(p)}$	Optimum for			Average $n_{h(a)}$	Chatterjee's $n_{h(c)}$	Proposed $n_{h(w)}$	Proposed $n_{h(d)}$
		Cows n_{h1}^*	Gallons n_{h2}^*	Receipts n_{h3}^*				
1	197	254	258	236	205	249	236	236
2	191	182	209	246	212	213	246	246
3	219	203	171	194	189	189	194	194
4	184	145	134	115	131	132	115	115
5	208	216	228	209	218	217	209	209

characteristics are mutually uncorrelated. The relative efficiency of the allocation $(\underline{n})_K$ with respect to another allocation $(n)_{K'} = (n_1, n_2, \dots, n_L)_{K'}$ may be defined as the ratio:

$$T(\underline{n})_{K'} / T(\underline{n})_K \quad . \text{ (see Sukhatme et.al.(1994)).}$$

$(\underline{n})_K = (n_1, n_2, \dots, n_L)_K$ denotes an L-component vector such that $n_h > 0; h = 1, 2, \dots, L$ and $\sum_{h=1}^L n_h = n$ (the total sample size).

The last column of Table 2.3 gives the relative efficiencies of various allocations with respect to the proportional allocation.

Table 2.3

Expected Variances of the estimated mean

Type of allocation	Cows	Gallons	Receipts	Trace	R.E. w.r.t. Proportional allocation
Optimum n_{hj}^*	0.0127	0.0800	76.9	76.9927	1.0520
Average $n_{h(a)}$	0.0128	0.0802	77.9	77.6930	1.0425
Chatterjee $n_{h(c)}$	0.0128	0.0800	77.5	77.5928	1.0438
Proposed $n_{h(w)}$	0.0130	0.0811	76.9	76.9941	1.0520
Proposed $n_{h(d)}$	0.0130	0.0811	76.9	76.9941	1.0520
Proportional $n_{h(p)}$	0.0130	0.0837	80.9	80.9968	-

It is observed that all the compromise allocations are more efficient than the

proportional allocation. However, the proposed compromise allocation based on weighted averages and the minimum deviation are equally good and most efficient. The percentage gain in efficiency in using the proposed allocations over the proportional allocation is 5.2% where as the corresponding value for average allocation is 4.2% and for Chatterjee's allocation is 4.4%. Thus the proposed allocations are more precise than other compromise allocations.

Example 2: The data are from a farm survey in Iowa reported by Jessen (1942) (see Sukhatme et al., (1984)). The relevant data with respect to three characteristics (i) number of hogs bought during the year (ii) number of cattle bought during the year and (iii) number of cows milked during the year, are shown in Table 2.4

Table 2.4

Estimated strata mean squares

Stratum No. h	W_h	Hogs bought s_{h1}^2	Cattle bought s_{h2}^2	Cows Milked s_{h3}^2
1	0.197	12	56	41.3
2	0.191	80	2,132	23.1
3	0.219	1,113	565	10.9
4	0.184	84	355	11.5
5	0.208	247	68	38.8

It has been decided to fix the total sample size as $n = 1000$

Table 2.5

Sample sizes within strata ($n=1000$)

Stratum h	No	Proportional $n_{h(p)}$	Optimum for			Average n_h	Chatterjee's $n_{h(c)}$	Proposed $n_{h(w)}$	Proposed $n_{h(d)}$
			Hogs n_{h1}^*	Cattle n_{h2}^*	Milk n_{h3}^*				
1		197	46	71	262	126	142	70	73
2		191	117	426	190	244	248	318	323
3		219	499	252	150	300	299	328	323
4		184	115	168	129	138	125	149	141
5		208	223	83	269	192	186	135	140

The proposed compromise allocation based on weighted averages given by (2.24) and the corresponding expected variances given by (2.25) are:

$n_{1(w)} = 69$, $n_{2(w)} = 318$, $n_{3(w)} = 328$, $n_{4(w)} = 150$ and $n_{5(w)} = 135$ with $V_{1(w)} = 0.2766$, $V_{2(w)} = 0.4606$ and $V_{3(w)} = 0.0424$ respectively.

The proposed compromise allocation based on minimum deviation given by (2.34) and the corresponding expected variances given by (2.35) are

$n_{1(d)} = 73$, $n_{2(d)} = 323$, $n_{3(d)} = 323$, $n_{4(d)} = 141$ and $n_{5(d)} = 140$ with $V_{1(d)} = 0.2772$, $V_{2(d)} = 0.4607$ and $V_{3(d)} = 0.0409$ respectively.

Table 2.5 gives the different allocations. The optimum expected variances under various allocations are shown in Table 2.6. The last column shows the relative efficiency of different allocations with respect to proportional allocation based on the ratio of traces of the variance-covariance matrices of \bar{y}_{jst} under different allocations.

It is observed that, all the compromise allocations are more efficient than the proportional allocation. However the two proposed compromise allocations are more efficient than other compromise allocations. The compromise allocation based on minimum deviation is the most efficient. The percentage gain in efficiency is about 24.4% where as the same figure corresponding to average allocation is only 11.7% and for Chatterjee's allocation is merely 5.9%.

Table.2.6

Expected variances of the estimated mean

Type of allocation	Hog	Cattle	Milk	Trace	R.E. w.r.t. Proportional allocation
Optimum n_{hj}^*	0.2147	0.4277	0.0233	0.6657	1.4556
Average $n_{h(a)}$	0.2694	0.5291	0.0293	0.8278	1.1170
Chatterjee $n_{h(c)}$	0.2739	0.5314	0.1097	0.9150	1.0590
Proposed $n_{h(w)}$	0.2768	0.4607	0.0408	0.7783	1.2450
Proposed $n_{h(d)}$	0.2772	0.4607	0.0409	0.7788	1.2442
Proportional $n_{h(p)}$	0.3260	0.6180	0.0250	0.9690	—

2.9 ALLOCATION WITH VARIABLE COST OF MEASUREMENT

Let $c_{hj}; h=1,2,\dots,L; j=1,2,\dots,p$, denote the per unit cost of measuring the j th characteristic in h th stratum. Also let, out of the total budget 'C', $n_{h(d)}$, denote the cost allocated for measuring the j th characteristic. The individual optimum allocations using (2.7) are given as

$$n_{hj}^* = n \frac{W_h S_{hj} / \sqrt{c_{hj}}}{\sum_{h=1}^L W_h S_{hj} / \sqrt{c_{hj}}}; h=1,2,\dots,L; j=1,2,\dots,p \quad (2.36)$$

where n given by (2.8) is

$$n = \frac{C_j \sum_{h=1}^L (W_h S_{hj} / \sqrt{c_{hj}})}{\sum_{h=1}^L (W_h S_{hj} \sqrt{c_{hj}})} \quad (2.37)$$

Substitution of the value of n from (2.37) in (2.36) gives

$$n_{hj}^* = \frac{C_j W_h S_{hj} / \sqrt{c_{hj}}}{\sum_{h=1}^L W_h S_{hj} \sqrt{c_{hj}}}; h = 1, 2, \dots, L; j = 1, 2, \dots, p \quad (2.38)$$

where overhead cost c_o is ignored, that is the cost functions for individual allocations are taken as

$$C_j = \sum_{h=1}^L c_{hj} n_{hj}; j = 1, 2, \dots, p \quad (2.39)$$

The optimum value of the variance $V(\bar{y}_{jst})$ (ignoring fpc) of the estimate \bar{y}_{jst} of the population mean \bar{Y}_j of the j th characteristic under the optimum allocation is given by

$$V_j^* = V(\bar{y}_{jst})_{opt} = \frac{(\sum_{h=1}^L W_h S_{hj} \sqrt{c_{hj}})^2}{C_j}; j = 1, 2, \dots, p \quad (2.40)$$

(see Cochran (1977)).

For working out a compromise allocation we have to restructure the cost setup as below.

Let $c_h = \sum_{j=1}^p c_{hj}$; denote the per unit cost of measuring all the 'p'

characteristics in the h th stratum. Then for any compromise allocation

$\underline{n} = (n_1, n_2, \dots, n_L)$, we have

$$C = \sum_{h=1}^L c_h n_h \quad (2.41)$$

as the cost constraint, where $C = \sum_{j=1}^p C_j$ is the total fixed budget. The

variances $V_j; j=1,2,\dots,p$ (ignoring fpc) under a compromise allocation

$n_h; h=1,2,\dots,L$ can be worked out directly by using

$$V_j = \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_h}; j=1,2,\dots,p \quad (2.42)$$

Due to this restructured cost setup it is not advisable to use the Cochran's average allocation or the allocation based weighted averages proposed in Section 2.6 because these compromise allocations either do not utilize the cost fully or become infeasible by violating the cost constraint in (2.41).

Chatterjee's allocation discussed in Section 2.7 can be used to work out compromise allocation for fixed total cost. It gives the compromise allocation as:

$$n_{h(c)} = \frac{C \sqrt{\sum_{j=1}^p n_{hj}^{*2}}}{\sum_{h=1}^L c_h \sqrt{\sum_{j=1}^p n_{hj}^{*2}}}; h=1,2,\dots,L \quad (2.43)$$

The corresponding variances can be obtained by putting $n_h = n_{h(c)}$ in (2.42)

Thus

$$V_{j(c)} = \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_{h(c)}}; j = 1, 2, \dots, p \quad (2.44)$$

where the symbol (c) stands for Chatterjee. In fact Chatterjee's allocation given in Section 2.7 is a special case ($c_h = c$) of his general allocation given in (2.43)

2.11 THE MINIMUM DEVIATION COMPROMISE ALLOCATION FOR FIXED COST

As discussed in Section 2.9 the total deviation D for fixed cost is given

$$\begin{aligned} \text{as } D &= \sum_{j=1}^p (V_j - V_j^*) \\ &= \sum_{j=1}^p \left(\sum_{h=1}^L \frac{W_h S_{hj}^2}{n_h} - \frac{(\sum_{h=1}^L W_h S_{hj} \sqrt{c_{hj}})^2}{C_j} \right) \end{aligned} \quad (2.45)$$

where V_j^* is given by (2.40). The problem of allocation thus become to find $n_h; h = 1, 2, \dots, L$ that minimize D given by (2.45) subject to the cost constraint in (2.41)

Defining the Lagrangian function as

$$\phi(n_h, \lambda) = D + \lambda \left(\sum_{h=1}^L c_h n_h - C \right) \quad (2.46)$$

and equating the partial derivatives $\frac{\partial \phi}{\partial n_h}$ and $\frac{\partial \phi}{\partial \lambda}$ equal to zero and

solving the

$(L+1)$ simultaneous equations thus obtained we get the minimum deviation compromise allocation $n_{h(d)}$ as

$$n_{h(d)} = \frac{C \sqrt{\sum_{j=1}^p W_h S_{hj}^2}}{\sqrt{c_h} \sqrt{\sum_{j=1}^p W_h S_{hj}^2}}; h = 1, 2, \dots, L \quad (2.47)$$

where the symbol (d) indicates that the compromise allocation is based on minimum deviation.

The corresponding variances can be obtained by using

$$V_{j(d)} = \sum_{h=1}^L \frac{W_h^2 S_{hj}^2}{n_{h(d)}}; j = 1, 2, \dots, p \quad (2.48)$$

2.11 A NUMERICAL ILLUSTRATION

The data of this example are from Jahan, N. et al (1994). In a two-variate survey ($p=2$) the population is stratified into two strata ($L=2$). The following information are available.

Also

$$\left((c_{hj}) \right) = \begin{pmatrix} 4 & 20 \\ 5 & 22 \end{pmatrix} \Rightarrow \left((c_h) \right) = (24, 27)$$

$$((C_j)) = (400, 2000) \Rightarrow C = 2400$$

Table 2.7

Data for two strata and two characteristics

Stratum No. h	W_h	S_{h1}	S_{h2}
1	0.4	4.6	332
2	0.6	2.8	173

The individual optimum allocations are worked out using formula (2.36)

$$\text{as } ((n_{hj}^*)) = \begin{pmatrix} 49 & 55 \\ 40 & 41 \end{pmatrix}$$

$$\text{with } V_1^* = 0.1384 \text{ and } V_2^* = 584.0261$$

The average allocation is given as

$$((n_{h(a)})) = \begin{pmatrix} 52 \\ 41 \end{pmatrix}$$

The cost associated with this allocation is

$$52 \times 24 + 41 \times 27 = 2355$$

which is less than the available cost 2400 and hence it is not advisable to use this allocation. However, for the sake of comparison the variances under this allocation are worked out as

$$V_{1(a)} = 0.1339 \text{ and } V_{2(a)} = 601.9420.$$

The weighted average allocation proposed by author in Section 2.6 is

given as $((n_{h(w)})) = \begin{pmatrix} 55 \\ 41 \end{pmatrix}$

The cost associated with this allocation is

$$55 \times 24 + 41 \times 27 = 2427$$

which is more than the available cost 2400. Hence this allocation is infeasible and cannot be used. However, for the sake of comparison the variances under this allocation are worked out as

$$V_{1(w)} = 0.1315 \text{ and } V_{2(w)} = 589.38308 ,$$

after adjusting the $n_{h(w)}$ to maintain the feasibility by multiplying it by an adjustment factor of $2400/2427=0.9889$.

Thus the adjusted $((n_{h(w)})) = \begin{pmatrix} 54 \\ 41 \end{pmatrix}$.

Compromise allocation given by Chatterjee's, using formula (2.43) is

$$((n_{h(c)})) = \begin{pmatrix} 53 \\ 41 \end{pmatrix} \text{ with corresponding variances}$$

$$V_{1(c)} = 0.1327 \text{ and } V_{2(c)} = 595.5429 .$$

The proposed minimum deviation allocation using formula (2.47) is

$$((n_{h(d)})) = \begin{pmatrix} 55 \\ 40 \end{pmatrix} \text{ with corresponding variances}$$

$$V_{1(d)} = 0.1322 \text{ and } V_{2(d)} = 590.0126 .$$

The proportional allocation for a fixed cost C is given by

$$n_{h(p)} = \frac{CW_h}{\sum_{h=1}^L W_h c_h}$$

(see Sukhatme et al (1984)).

For $C = 2400$, $W_1 = 0.4$, $W_2 = 0.6$, $c_1 = 24$, $c_2 = 27$

The proportional allocation $n_{h(p)}$; $h = 1, 2$ is worked out as

$$n_{1(p)} = \frac{2400 \times 0.4}{0.4 \times 24 + 0.6 \times 27} = 37.2093 \cong 37$$

$$\text{and } n_{2(p)} = \frac{2400 \times 0.6}{0.4 \times 24 + 0.6 \times 27} = 55.8139 \cong 56$$

Thus $((n_{h(p)})) = \begin{pmatrix} 37 \\ 56 \end{pmatrix}$. The corresponding variances are

$$V_{1(p)} = 0.1419 \text{ and } V_{2(p)} = 669.0450 .$$

Where the symbol (p) stands for proportional.

These results arranged in a tabular form are given in Tables 2.8 and 2.9

Table 2.8

Sample sizes within strata

Stratum No. h	Compromise allocations (n_h)				
	Average $n_{h(a)}$	Weighted average $n_{h(w)}$	Chatterjee's $n_{h(c)}$	Proposed $n_{h(d)}$	Proportional $n_{h(p)}$
1	52	55	53	55	37
2	41	40	41	40	56

Table 2.9

Variances under different compromise allocations

Allocations	V_1	V_2	Trace	R.E. w.r.t. Proportional
Average $n_{h(a)}$	0.1339	601.9420	602.0359	1.1115
Chatterjee's $n_{h(c)}$	0.1327	595.5429	595.6756	1.1234
Proposed $n_{h(w)}$	0.1315	589.3808	589.5123	1.1351
Proposed $n_{h(d)}$	0.1322	590.0126	590.1448	1.1339
Proportional $n_{h(p)}$	0.1419	669.0450	669.1869	-

The last column of Table 2.9 shows the relative efficiencies with respect to the proportional allocation based on the ratio of traces of the variance-covariance matrices.

It is observed that all the compromise allocations are more efficient than the proportional allocation. However, both the proposed allocations are more efficient than other compromise allocations. The compromise allocation based on weighted averages is the most efficient allocation for this example with the percentage gain in efficiency over proportional allocation as 13.5%.

2.12 CONCLUSION

The three numerical examples worked out in Sections 2.9 and 2.12 indicate that the compromise allocations based on (i) weighted averages (the weights a_j ; proportional to individual optimum variances V_j^* ; $j = 1, 2, \dots, p$)

and (ii) minimizing total deviation $D = \sum_{j=1}^p (V_j - V_j^*)$ are more efficient than other compromise allocations existing in the sampling literature. The criterion for working out the relative efficiency is the ratio of the trace of the variance-covariance matrix of $\bar{y}_{jst}; j = 1, 2, \dots, p$ under proportional allocation to the trace under a given compromise allocation. Thus the proposed compromise allocations are an improvement over the compromise allocations already existing in the sampling literature.

CHAPTER-III

MIXED ALLOCATION IN STRATIFIED SAMPLING

CHAPTER – III

MIXED ALLOCATION IN STRATIFIED SAMPLING

3.1 INTRODUCTION

It is stratified sampling where the population of size N is divided into L strata of sizes N_1, N_2, \dots, N_L ($\sum_{h=1}^L N_h = N$) the variance of the stratified sample

mean $\bar{y}_{st} = \sum_{h=1}^L W_h \bar{Y}_h$ is given by

$$V(\bar{y}_{st}) = \sum_{h=1}^L \frac{W_h^2 S_h^2}{n_h} - \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} \quad (3.1)$$

The total cost 'C' of the survey may be given as

$$C = c_o + \sum_{h=1}^L c_h n_h$$

$$\text{or } C - c_o = \sum_{h=1}^L c_h n_h = C_o \quad (\text{say}) \quad (3.2)$$

where all the symbols have the same meaning as defined in Section 2, Chapter 2 of this thesis. The cost structure of the survey may be more complicated than given in (3.2) (See Hansen et .al. (1953) and Groves (1989)). For example the cost function may be of the form

$$C = c_o + \sum_{h=1}^L t_h \sqrt{n_h} + \sum c_h n_h$$

where t_h denote the traveling cost between the selected units of the h th stratum. Csenki (1977) used the cost function of the form

$$C = c_o + \sum_{h=1}^L c_h n_h^\delta$$

where $\delta > 0$ is a known constant.

In spite of all the above discussed cost functions the cost function given in (3.2) is often serves as an adequate approximation for practical purposes.

The fixed cost allocation that minimizes $V(\bar{y}_{st})$ is well known in sampling literature as optimum allocation is given as

$$n_h = n \frac{W_h S_h / \sqrt{c_h}}{\sum_{h=1}^L W_h S_h / \sqrt{c_h}}; h = 1, 2, \dots, L \quad (3.3)$$

where the total sample size n for fixed cost C is given as

$$n = C_o \frac{\sum_{h=1}^L (W_h S_h / \sqrt{c_h})}{\sum_{h=1}^L (W_h S_h / \sqrt{c_h})}; h = 1, 2, \dots, L \quad (3.4)$$

where $C_o = C - c_o$

Using (3.3) and (3.4), we get

$$n_h = C_o \frac{(W_h S_h / \sqrt{c_h})}{\sum_{h=1}^L (W_h S_h \sqrt{c_h})}; h = 1, 2, \dots, L \quad (3.5)$$

Substituting n_h given by (3.5) in (3.1) the value of the variance $V(\bar{y}_{st})$ under optimum allocation comes out to be

$$V^* = V(\bar{y})_{opt} = \frac{\left(\sum_{h=1}^L W_h S_h \sqrt{c_h} \right)^2}{C_o} - \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} \quad (3.6)$$

(see Cochran (1977)).

The optimum allocation can also be worked out to minimize the cost for fixed variance. Using Cauchy -Schwarz inequality Stuart (1945) showed that in terms of total sample size n the expression of the sample sizes $n_h; h = 1, 2, \dots, L$ that minimize the cost for fixed variance $V(\bar{y}_{st})$ can also be given by (3.3). The value of n , the total sample size, in this case is given by

$$n = \frac{\left(\sum_{h=1}^L W_h S_h \sqrt{c_h} \right) \left(\sum_{h=1}^L W_h S_h / \sqrt{c_h} \right)}{V + \frac{1}{N} \sum_{h=1}^L W_h S_h^2} \quad (3.7)$$

where $V = V(\bar{y}_{st}) + \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} = \sum_{h=1}^L \frac{W_h^2 S_h^2}{n_h}$, is fixed .

Substituting the value of n from (3.7) in (3.3) we get

$$n_h = \frac{(W_h S_h / \sqrt{c_h})(\sum_{h=1}^L W_h S_h \sqrt{c_h})}{V + \frac{1}{N} \sum_{h=1}^L W_h S_h^2}; h = 1, 2, \dots, L \quad (3.8)$$

The resulting minimum cost is

$$C_{opt} = c_o + \frac{(\sum_{h=1}^L W_h S_h \sqrt{c_h})^2}{V + \frac{1}{N} \sum_{h=1}^L W_h S_h^2} \quad (3.9)$$

obtained by substituting n_h given by (3.8) in (3.2) (see also Kish (1967)).

The practical difficulty in using optimum allocation is that usually S_h^2 are not known, thus we can only approximate this allocation by using estimated values of S_h^2 . They may be the values computed on some previous occasion or values obtained by a pilot survey. Other allocations that are less precise than optimum allocation are proportional and equal allocations. In proportional allocation the sample sizes n_h from various strata are proportional to the corresponding stratum weights W_h . This gives

$$n_h \propto W_h$$

$$\text{or } n_h = KW_h. \quad (3.10)$$

where K is the constant of proportionality.

The proportional allocation may be worked out for fixed cost or for fixed total sample size. For fixed cost, under proportional allocation

$$n_h = \frac{C_o W_h}{\sum_{h=1}^L c_h W_h}; h = 1, 2, \dots, L \quad (3.11)$$

$$\text{with } V(\bar{y}_{st})_{prop} = \frac{(\sum_{h=1}^L W_h S_h^2)(\sum_{h=1}^L c_h W_h)}{C_o} - \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} \quad (3.12)$$

for fixed total sample size, under proportional allocation

$$n_h = n W_h; h = 1, 2, \dots, L \quad (3.13)$$

$$\text{with } V(\bar{y}_{st})_{prop} = \frac{\sum_{h=1}^L W_h S_h^2}{n} - \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h} \quad (3.14)$$

(see Section 2.2, Chapter 2).

Practical implementation of proportional allocation is easy because usually the strata sizes N_h and thus strata weights W_h are known. In case W_h are unknown they can also be estimated from a pilot survey.

In the absence of the true value of W_h , if other situations permit one can use equal allocation. To implement equal allocation only knowledge of the total sample size n and the number strata L are required. The sample sizes n_h are given by

$$n_h = \frac{n}{L}; h = 1, 2, \dots, L \quad (3.15)$$

The variance $V(\bar{y}_{st})$ under equal allocation is given as:

$$V(\bar{y}_{st})_{equal} = \frac{\sum_{h=1}^L W_h^2 S_h^2}{n} - \frac{\sum_{h=1}^L W_h S_h^2}{N} \quad (3.16)$$

In this chapter the problem of allocation of a sample to strata is discussed in general conditions. There are sometimes valid reasons due to which only a particular type of allocation is advisable in a particular part of a stratified population. Under this situation it would be reasonable to divide the group of strata into subgroups and use a particular type of allocation in one group. Clark and Steel (2000) used a similar idea in two-stage stratified sampling.

Such an allocation, which uses different type of allocations for different subgroups of strata, may be called a “Mixed allocation”.

3.2. THE MIXED ALLOCATION

Let the group of L strata is divided into k subgroups G_1, G_2, \dots, G_k , where the group G_j consists of $L_j; j = 1, 2, \dots, k$ strata such that $\sum_{j=1}^k L_j = L$.

Without loss of generality we can assume that the first L_1 strata constitute the first subgroup G_1 , the next L_2 strata constitute the second subgroup G_2 , and so on and the last L_k strata constitute the last subgroup G_k . Under this scheme, the j th subgroup $G_j; j = 1, 2, \dots, k$ will consists of

$$\left(\sum_{i=1}^{j-1} L_i + 1\right)\text{th}, \left(\sum_{i=1}^{j-1} L_i + 2\right)\text{th}, \dots, \text{ and } \left(\sum_{i=1}^{j-1} L_i + L_j\right) = \left(\sum_{i=1}^j L_i\right)\text{th strata. .}$$

Further let due to the prevailing circumstances in a particular subgroup a particular type of allocation is to be used. This could be done by letting

$$n_h = \alpha_j \beta_h; h \in I_j; j = 1, 2, \dots, k \quad (3.17)$$

where $I_j; j = 1, 2, \dots, k$ is the set of indices of the strata that constitute the j th subgroup G_j ,

$\beta_h; h \in I_j; j = 1, 2, \dots, k$ are known constants depending upon the type of allocation to be used in the j th subgroup, and $\alpha_j; j = 1, 2, \dots, k$ are to be determined.

For example if in any particular subgroup, say G_p , equal allocation is to be used then

$$\beta_h = 1; h \in I_p$$

Proportional allocation in the q th subgroup G_q is characterized by

$$\beta_h = W_h; h \in I_q$$

To use optimum allocation in the r th subgroup G_r , β_h is given as

$$\beta_h = \frac{W_h S_h}{\sqrt{c_h}}; h \in I_r \text{ and so on.}$$

Two other allocations that are used sometimes are allocation proportional to $W_h \bar{Y}_h$ and allocation proportional to $W_h R_h$, where

$R_h; h = 1, 2, \dots, L$ denote the range of the h th stratum (see Murthi (1967)). If any of the above allocations is to be introduced in a subgroup we may take $\beta_h = W_h \bar{Y}_h$ or $\beta_h = W_h R_h$ accordingly.

It can be seen that

$$I_j = \left\{ \sum_{i=1}^{j-1} L_i + 1, \sum_{i=1}^{j-1} L_i + 2, \dots, \sum_{i=1}^{j-1} L_i + L_j = \sum_{i=1}^j L_i \right\}; j = 1, 2, \dots, k \quad (3.18)$$

where $I_r \cap I_s = \phi; r \neq s$ and $\bigcup_{j=1}^k I_j = \{1, 2, \dots, L\}$

The mixed allocation defined in (3.17) may be computed for minimizing $V(\bar{y}_{st})$ given by (3.1) for a fixed cost or for minimizing the total cost given by (3.2) for a fixed value of $V(\bar{y}_{st})$.

These optimization problems can be formulated as the following two nonlinear programming problems (NLPP)

NLPP1: (Minimizing $V(\bar{y}_{st})$ for fixed cost)

$$\text{Minimize } V(n_h) = \sum_{h=1}^L \frac{W_h^2 S_h^2}{n_h} \quad (3.19)$$

$$\text{subject to } \sum_{h=1}^L c_h n_h = C_o \quad (3.20)$$

$$n_h = \alpha_j \beta_h; h \in I_j; j = 1, 2, \dots, k \quad (3.21)$$

$$n_h \geq 0; h = 1, 2, \dots, L \quad (3.22)$$

where from the expression (3.1) of $V(\bar{y}_{st})$ the terms independent of n_h are

$$\text{dropped and } C_o = C - c_o \quad (3.23)$$

NLPP2: (Minimizing the cost for fixed value of $V(\bar{y}_{st})$)

$$\text{Minimize } C(n_h) = \sum_{h=1}^L c_h n_h \quad (3.24)$$

$$\text{subject to } \sum_{h=1}^L \frac{W_h^2 S_h^2}{n_h} = v \quad (3.25)$$

$$n_h = \alpha_j \beta_h; h = 1, 2, \dots, k \quad (3.26)$$

$$\text{and } n_h \geq 0; h = 1, 2, \dots, L \quad (3.27)$$

where from the expression (3.2) of C the term c_o which is independent of n_h is dropped

$$\text{and } v = V + \sum_{h=1}^L \frac{W_h^2 S_h^2}{N_h}; V \text{ being the fixed value of } V(\bar{y}_{st}).$$

Using constraints $n_h = \alpha_j \beta_h; h \in I_j; j = 1, 2, \dots, k$ the expression $\sum_{h=1}^L \frac{W_h^2 S_h^2}{n_h}$

and $\sum_{h=1}^L c_h n_h$ in NLPP 1 and 2 may be expressed as :

$$\sum_{h=1}^L \frac{W_h^2 S_h^2}{n_h} = \sum_{j=1}^k \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j \beta_h} \quad (3.28)$$

$$\text{and } \sum_{h=1}^L c_h n_h = \sum_{j=1}^k \sum_{h \in I_j} \alpha_j c_h \beta_h \quad (3.29)$$

respectively, where $I_j; j = 1, 2, \dots, k$ are given by (3.18)

Using (3.24) and (3.25) the two NLPP can thus be restated as:

$$\text{NLPP 1: Minimize } F_1(\alpha_j) = \sum_{j=1}^k \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j \beta_h} \quad (3.30)$$

$$\text{subject to } \sum_{j=1}^k \sum_{h \in I_j} \alpha_j c_h \beta_h = C_o \quad (3.31)$$

$$\text{and } \alpha_j \geq 0; j = 1, 2, \dots, k \quad (3.32)$$

$$\text{NLPP2: Minimize } F_2(\alpha_j) = \sum_{j=1}^k \sum_{h \in I_j} \alpha_j c_h \beta_h \quad (3.33)$$

$$\text{subject to } \sum_{j=1}^k \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j \beta_h} = v \quad (3.34)$$

$$\text{and } \alpha_j \geq 0; j = 1, 2, \dots, k \quad (3.35)$$

Ignoring restrictions $\alpha_j \geq 0; j = 1, 2, \dots, k$ both the NLPP1 and 2 can be solved by using Lagrange multipliers technique. If the solutions thus obtained satisfy the restrictions $\alpha_j \geq 0; j = 1, 2, \dots, k$ and thus $n_h \geq 0; h = 1, 2, \dots, L$ also then the NLPP1 and 2 are solved completely, otherwise some nonlinear programming technique may be used to solve them.

3.3 THE SOLUTION

The NLPP1 after ignoring restrictions in (3.32) can be described as “Find $\alpha_j; j = 1, 2, \dots, k$ that minimize $F_1(\alpha_j)$ given by (3.30) subject to the constraint (3.31)”.

The Lagrangian function $\phi_1(\alpha_j, \lambda_1)$ for this problem is

$$\phi_1(\alpha_j, \lambda_1) = \sum_{j=1}^k \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j \beta_h} + \lambda_1 \left(\sum_{j=1}^k \sum_{h \in I_j} \alpha_j c_h \beta_h - C_o \right) \quad (3.36)$$

where λ_1 is the Lagrange multiplier.

Differentiating ϕ_1 with respect to α_j and λ_1 and equating the partial derivatives equal to zero we get the following $(k+1)$ simultaneous equations

$$\frac{\delta \phi_1}{\delta \alpha_j} = - \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j^2 \beta_h} + \lambda_1 \sum_{h \in I_j} c_h \beta_h = 0; j = 1, 2, \dots, k \quad (3.37)$$

$$\text{and } \frac{\delta \phi_1}{\delta \lambda_1} = \sum_{j=1}^k \sum_{h \in I_j} \alpha_j c_h \beta_h - C_o = 0 \quad (3.38)$$

$$\text{From (3.37) } \lambda_1 \sum_{h \in I_j} c_h \beta_h = \frac{1}{\alpha_j^2} \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\beta_h}$$

(because $\alpha_j; j = 1, 2, \dots, k$ is constant within a particular subgroup)

$$\text{or } \alpha_j^2 = \frac{1}{\lambda_1} \frac{\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h}{\sum_{h \in I_j} c_h \beta_h}$$

$$\text{or } \alpha_j = \frac{1}{\sqrt{\lambda_1}} \sqrt{\frac{\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h}{\sum_{h \in I_j} c_h \beta_h}}; j = 1, 2, \dots, k \quad (3.39)$$

Substituting the value of α_j from (3.39) in (3.38) we get

$$\sum_{j=1}^k \sum_{h \in I_j} \frac{1}{\sqrt{\lambda_1}} \sqrt{\frac{\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h}{\sum_{h \in I_j} c_h \beta_h}} c_h \beta_h = C_o$$

$$\text{or } \frac{1}{\sqrt{\lambda_1}} = \frac{C_o}{\sum_{j=1}^k \left(\sum_{h \in I_j} c_h \beta_h \right) \sqrt{\frac{\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h}{\sum_{h \in I_j} c_h \beta_h}}}$$

$$\text{or } \frac{1}{\sqrt{\lambda_1}} = \frac{C_o}{\sum_{j=1}^k \sqrt{\left(\sum_{h \in I_j} c_h \beta_h \right) \left(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h \right)}} \quad (3.40)$$

From (3.39) and (3.40)

$$\alpha_j = C_o \frac{\sqrt{\left(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h \right) / \left(\sum_{h \in I_j} c_h \beta_h \right)}}{\sum_{j=1}^k \sqrt{\left(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h \right) \left(\sum_{h \in I_j} c_h \beta_h \right)}}; j = 1, 2, \dots, k \quad (3.41)$$

The values of the sample sizes n_h for the strata belonging to a particular subgroup say G_p , that is for $h \in I_p$ can be obtained by substituting the value of α_j given by (3.41) for $j=p$, in (3.17), where $p \in \{1, 2, \dots, k\}$

The resulting variance (ignoring fpc) is

$$V_{(mixed)} = \sum_{j=1}^k \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j \beta_h}$$

$$\begin{aligned}
&= \sum_{j=1}^k \sum_{h \in I_j} \frac{\left(W_h^2 S_h^2 / \beta_h \right) \sum_{j=1}^k \sqrt{\left(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h \right) \left(\sum_{h \in I_j} c_h \beta_h \right)}}{C_o \sum_{j=1}^k \sqrt{\left(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h \right) \left(\sum_{h \in I_j} c_h \beta_h \right)}} \\
&= \frac{\left(\sum_{j=1}^k \sqrt{\left(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h \right) \left(\sum_{h \in I_j} c_h \beta_h \right)} \right)^2}{C_o} \tag{3.42}
\end{aligned}$$

where the symbol ‘(mixed)’ corresponds to the mixed allocation.

The NLPP2 after ignoring restrictions in (3.31) may be described as:

“Find $\alpha_j; j = 1, 2, \dots, k$ that minimize $F_2(\alpha_j)$ given by (3.33) subject to be constraint (3.34)”.

To solve this problem, define the Lagrangian function $\phi_2(\alpha_j, \lambda_2)$ as

$$\phi_2(\alpha_j, \lambda_2) = \sum_{j=1}^k \sum_{h \in I_j} \alpha_j c_h \beta_h + \lambda_2 \left(\sum_{j=1}^k \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j \beta_h} - v \right) \tag{3.43}$$

where λ_2 is the Lagrange multiplier.

As before we have the (k+1) equations as

$$\frac{\partial \phi_2}{\partial \alpha_j} = \sum_{h \in I_j} c_h \beta_h - \lambda_2 \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j^2 \beta_h} = 0; j = 1, 2, \dots, k \tag{3.44}$$

$$\text{and } \frac{\partial \phi}{\partial \lambda_2} = \sum_{j=1}^k \sum_{h \in I_j} \frac{W_h^2 S_h^2}{\alpha_j \beta_h} - v = 0; j = 1, 2, \dots, k \tag{3.45}$$

From (3.44)

$$\alpha_j = \sqrt{\lambda_2} \sqrt{\frac{\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h}{\sum_{h \in I_j} c_h \beta_h}}; j = 1, 2, \dots, k \quad (3.46)$$

Substituting the value of α_j from (3.46) in (3.45) we get

$$\sqrt{\lambda_2} = \frac{\sum_{j=1}^k \sqrt{\frac{(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h)(\sum_{h \in I_j} c_h \beta_h)}{v}}}{v} \quad (3.47)$$

From (3.46) and (3.47)

$$\alpha_j = \frac{1}{v} \left(\sqrt{\frac{(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h)(\sum_{h \in I_j} c_h \beta_h)}{v}} \right) \left(\sum_{j=1}^k \sqrt{\frac{(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h)(\sum_{h \in I_j} c_h \beta_h)}{v}} \right), \quad (3.48)$$

$j = 1, 2, \dots, k$

The values of the sample sizes n_h for the strata belonging to a particular subgroup say G_q , that is for $h \in I_q$ can be obtained by substituting the value of α_j given by (3.48) for $j=p$ in (3.17), where $q \in \{1, 2, \dots, k\}$

The resulting cost is

$$\begin{aligned} C_{(mixed)} &= \sum_{j=1}^k \sum_{h \in I_j} \alpha_j c_h \beta_h \\ &= \frac{1}{v} \sum_{j=1}^k \left[\sum_{h \in I_j} c_h \beta_h \left(\sqrt{\frac{(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h)(\sum_{h \in I_j} c_h \beta_h)}{v}} \right) \right. \\ &\quad \left. \times \left(\sqrt{\frac{(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h)(\sum_{h \in I_j} c_h \beta_h)}{v}} \right) \right] \end{aligned}$$

$$= \frac{\left[\sum_{j=1}^k \sqrt{\left(\sum_{h \in I_j} W_h^2 S_h^2 / \beta_h \right) \left(\sum_{h \in I_j} c_h \beta_h \right)} \right]^2}{v} \quad (3.49)$$

The results obtained in this section can also be obtained by using Cauchy- Schwarz inequality.

3.4 THE INEFFICIENCY OF THE MIXED ALLOCATION

It is well known that the optimum allocation given by (3.5) is the most efficient allocation for fixed cost. But there are certain limitations to the use of optimum allocation in practice. The most severe of all the limitations is the absence of the knowledge of strata variances S_h^2 . In such situations in the formula (3.5) S_h^2 may be replaced by its sample estimate

$$s_h^2 = \frac{1}{n_h - 1} \sum_{i=1}^{n_h} (y_{hi} - \bar{y}_h)^2$$

The values of the sample allocations in this case will be

$$\hat{n}_h = \frac{C_o (W_h s_h / \sqrt{c_h})}{\sum_{h=1}^L W_h s_h \sqrt{c_h}}; h = 1, 2, \dots, L$$

where \hat{n}_h are called the modified optimum allocation. Unfortunately, in general there is no guarantee that this modified optimum allocation is really optimum. At times it proved to be less efficient than a proportional allocation. So that even if an estimate of S_h^2 is available it is not always advisable to use the

modified optimum allocation (see Sukhatme et al (1984)). As an alternative the use of the proposed mixed allocation is advised.

The relative efficiency (R.E.) of the optimum allocation for fixed cost as compared to the corresponding mixed allocation is given by

$$(R.E.)_{opt} = \frac{V_{mixed} - V_{opt}}{V_{opt}} \quad (3.50)$$

where $(R.E.)_{opt}$, stands for the relative efficiency of optimum allocation as compared to the mixed allocation.

The quantity on the R.H.S. of (3.50) can also be called the inefficiency of the mixed allocation as compared to the optimum allocation.

$$\text{Thus } (R.I.E.)_{mixed} = \frac{V_{mixed} - V_{opt}}{V_{opt}} \quad (3.51)$$

where $(R.I.E.)_{mixed}$, stands for the relative inefficiency of the mixed allocation as compared to the optimum allocation.

In the expression (3.50) and (3.51) V_{mixed} is given by (3.42) and V_{opt} (ignoring fpc) is given by

$$V_{opt} = \frac{(\sum_{h=1}^L W_h S_h \sqrt{c_h})^2}{C_o} \quad (3.52)$$

The contribution towards the total relative inefficiency of a particular allocation applied to the subgroup G_p can be assessed by the term

$$\frac{\left(\sum_{h \in I_p} \frac{W_h^2 S_h^2}{\alpha_j \beta_h} - \sum_{h \in I_p} \frac{W_h^2 S_h^2}{n_{h(opt)}} \right)}{V_{opt}} \quad (3.53)$$

in the RHS of (3.51), where $n_{h(opt)}$ denote the sample sizes under optimum allocation for $h \in I_p$.

The expression (3.53) will help in deciding whether to use any particular allocation in a particular subgroup or not. If a particular allocation results in a large contribution towards the total relative inefficiency when applied to a particular, subgroup of strata, then the reasons for applying it may be reviewed.

3.5 A NUMERICAL ILLUSTRATION

In stratification with seven strata the values of N_h, s_h and c_h are given in Table 3.1. It assumed that the total available budget of the survey in $C = 4500$ units which includes an overhead cost $c_0 = 500$ units. The data are artificially constructed to illustrate the use mixed allocation.

We have $C_0 = C - c_0 = 4500 - 500 = 4000$ units of cost available for measurements.

In Table 3.1 the strata are so arranged that

- (i) Strata 1,2 & 3 constitute sub group G_1 in which equal allocation is to be used.

- (ii) Strata 4 & 5 constitute subgroup G_2 in which proportional allocation is to be used.
- (iii) Strata 6 & 7 constitute subgroup G_3 in which optimum allocation is to be used.

Table 3.1

Values of N_h, s_h and c_h for seven strata

Stratum No h	Stratum size N_h	Stratum S.D. s_h	Per units of measurement c_h
1	472	5.237	6
2	559	5.821	8
3	425	5.238	7
4	218	25.528	12
5	233	22.232	11
6	328	15.129	10
7	265	40.125	15

Thus, $I_1 = \{1,2,3\}$

$I_2 = \{4,5\}$

$I_3 = \{6,7\}$

It can be seen that $I_j; j = 1,2,3$ are mutually exclusive and exhaustive as

$$I_1 \cap I_2 = I_1 \cap I_3 = I_2 \cap I_3 = \phi \text{ and } \bigcup_{j=1}^3 I_j = \{1,2,3,4,5,6,7\}$$

The reason for using equal allocation to strata 1,2 and 3 is that these strata are relatively more homogeneous as compared to other strata since their corresponding estimated strata S.D.(s_h) are small. Proportional allocation is used in strata 4 and 5 because they have relatively smaller size (N_h) among the remaining four strata. The above set up will help in reducing the variance $V(\bar{y}_{st})$ under mixed allocation.

Table 3.2

Sample sizes under over all optimum allocation

h	W_h	s_h	c_h	$W_h s_h$	$W_h s_h / \sqrt{c_h}$	$W_h s_h \sqrt{c_h}$	$n_{h(opt)}$ (rounded)
1	0.189	5.237	6	0.990	0.404	2.425	35
2	0.224	5.821	8	1.304	0.461	3.688	40
3	0.170	5.238	7	0.890	0.336	2.355	29
4	0.087	25.528	12	2.221	0.641	7.694	56
5	0.093	22.232	11	2.067	0.623	6.855	55
6	0.131	15.129	10	1.982	0.627	6.268	55
7	0.106	40.125	15	4.253	1.098	16.472	96
Σ						45.757	

Table 3.2 gives the sample sizes when as gives overall optimum allocation is used.

The estimated variance $v(\bar{y}_{st})$ under optimum allocation ignoring fpc is

$$\begin{aligned}
 v_{opt} &= \frac{\left(\sum_{h=1}^L W_h s_h \sqrt{c_h} \right)^2}{C_o} \\
 &= \frac{(45.757)^2}{4000} \\
 &= 0.5234
 \end{aligned} \tag{3.54}$$

The application of the mixed allocation to the various subgroups of the strata according to the given scheme may be characterized by letting

$$n_{h(m)} = \alpha_j \beta_h; j = 1, 2, 3; h \in I_j \tag{3.55}$$

where $n_{h(m)}$; $h = 1, 2, \dots, L$ denote the sample sizes under mixed allocation and

β_h are defined as below.

In subgroup G_1 for applying equal allocation

$$\beta_h = 1 \text{ for } h \in I_1 = \{1, 2, 3\}$$

In subgroup G_2 for applying proportional allocation

$$\beta_h = 1 \text{ for } h \in I_2 = \{4, 5\}$$

In subgroup G_3 for applying optimum allocation

$$\beta_h = (W_h s_h / \sqrt{c_h}) \text{ for } h \in I_3 = \{6, 7\}$$

In Table 3.3 the values of $W_h^2 s_h^2 / \beta_h$ and $c_h \beta_h$ are tabulated. These values are to be used in the calculation of α_j ; $j = 1, 2, 3$.

Table 3.3Values of $W_h^2 s_h^2 / \beta_h$ and $c_h \beta_h$

h	$W_h s_h$	$W_h^2 s_h^2$	c_h	β_h	$W_h^2 s_h^2 / \beta_h$	$c_h \beta_h$
1	0.990	0.980	6	1	0.980	6
2	1.304	1.700	8	1	1.700	8
3	0.890	0.792	7	1	0.792	7
Subtotal for $h \in I_1$					3.472	21.000
4	2.221	4.933	12	0.087	56.701	1.044
5	2.067	4.272	11	0.093	45.935	1.023
Sub total for $h \in I_2$					102.636	2.067
6	1.982	3.928	10	0.627	6.270	6.270
7	4.253	18.088	15	1.098	16.470	16.470
Subtotal for $h \in I_3$					22.740	22.740

Table 3.4 gives the values of α_j ; $j = 1, 2, 3$

Table 3.4

Calculation of α_j

Subgroup No. J	(A) $\sum_{h \in I_j} W_h^2 s_h^2 / \beta_h$	(B) $\sum_{h \in I_j} c_h \beta_h$	(C) $\sqrt{(A)/(B)}$	(D) $\sqrt{(A)(B)}$	α_j $= 4000 \times \frac{(C)}{\sum(D)}$
1	3.472	21.000	0.407	8.539	35.512
2	102.636	2.067	7.047	14.565	614.868
3	22.740	22.740	1	22.740	87.252
$\sum(D)$				45.844	

The values of β_h from Table 3.3 and values of α_j from Table 3.4 when substituted in the formula (3.55) gives the mixed allocation $n_{h(m)}; h = 1, 2, \dots, 7$ as

For $J = 1$, that is $h \in I_1 = \{1, 2, 3\}$

$$n_{1(m)} = \alpha_1 \beta_1 = 35.512 \times 1 = 35.512 \cong 35$$

$$n_{2(m)} = \alpha_1 \beta_2 = 35.512 \times 1 = 35.512 \cong 35 \quad 35$$

$$n_{3(m)} = \alpha_1 \beta_3 = 35.512 \times 1 = 35.512 \cong 35 \quad 35$$

For $J = 2$, that is $h \in I_2 = \{4, 5\}$

$$n_{4(m)} = \alpha_2 \beta_4 = 614.868 \times 0.087 = 53.493 \cong 54$$

$$n_{5(m)} = \alpha_2 \beta_5 = 614.868 \times 0.093 = 57.183 \cong 57$$

For $J = 3$, that in $h \in I_3 = \{6, 7\}$

$$n_{6(m)} = \alpha_3 \beta_6 = 87.252 \times 0.627 = 54.707 \cong 55$$

$$n_{7(m)} = \alpha_3 \beta_7 = 87.252 \times 1.098 = 95.803 \cong 96$$

The estimated variance $v(\bar{y}_{st})$ (ignoring fpc) under mixed allocation is given as

$$v_{mixed} = \sum_{h=1}^7 \frac{W_h^2 s_h^2}{n_{h(m)}} = 0.5356$$

3.6 CONCLUSION

The estimated relative inefficiency of the mixed allocation as compared to the overall optimum allocation is given by (3.51) is

$$\begin{aligned} (R.I.E.)_{mixed} &= \frac{v_{mixed} - v_{opt}}{v_{opt}} \times 100\% \\ &= \frac{0.5253 - 0.5234}{0.5234} \times 100\% \\ &= 0.363\% \end{aligned}$$

The estimated relative inefficiency of equal and proportional allocations are given as

$$\begin{aligned} (R.I.E.)_{equal} &= \frac{\sum_{h \in I_1} \frac{W_h^2 s_h^2}{n_{h(m)}} - \sum_{h \in I_1} \frac{W_h^2 s_h^2}{n_{h(m)}}}{v_{opt}} \times 100\% \\ &= \frac{0.0992 - 0.0978}{0.5234} \times 100\% \\ &= 0.267\% \end{aligned}$$

$$\begin{aligned}
(R.I.E.)_{prop} &= \frac{\sum_{h \in I_1} \frac{W_h^2 s_h^2}{n_h(m)} - \sum_{h \in I_1} \frac{W_h^2 s_h^2}{n_h(m)}}{v_{opt}} \times 100\% \\
&= \frac{0.1663 - 0.1658}{0.5234} \times 100\% \\
&= 0.096\% \quad \text{respectively.}
\end{aligned}$$

It can be seen that

$$(R.I.E.)_{mixed} = (R.I.E.)_{equal} + (R.I.E.)_{prop}.$$

Since the *R.I.E.* in using mixed allocation is only 0.363% we conclude that we used mixed allocation safely instead of overall optimum allocation.

CHAPTER-IV

DOUBLE SAMPLING FOR STRATIFICATION WITH SUBSAMPLING THE NONRESPONDENTS: A DYNAMIC PROGRAMMING APPROACH

CHAPTER IV

DOUBLE SAMPLING FOR STRATIFICATION WITH SUBSAMPLING THE NONRESPONDENTS: A DYNAMIC PROGRAMMING APPROACH

4.1 INTRODUCTION

In stratified sampling, the population is divided into L strata which are homogeneous within themselves and whose means are widely different. The stratum weights are used in estimating unbiasedly the mean or the total of the character under study.

If these weights are not known, the technique of double sampling can be used, which consists of selecting a preliminary sample of size n' by simple random sampling, without replacement (SRSWOR), to estimate the stratum weights and then selecting the subsample of n units with n_h units from the h -th stratum, to collect information on the characteristic under study, such

that $\sum_{h=1}^L n_h = n$

Rao (1973) proposed the method of double sampling for stratification (DSS) for the estimation of the population mean \bar{Y} , of the variate y , using the values of the auxiliary variate collected at the first phase for stratification only.

Ige and Tripathi (1987) used the information collected at the first phase for stratification as well as in constructing ratio and difference estimators of the population mean \bar{Y}

One of the sources of error in surveys is non-contact or refusals. In a household survey the selected family may not be available at home when the interviewer calls. The selected person may refuse to cooperate, saying that he has not time to answer question or that he consider the purpose of the survey to be senseless. Persuasion and further recalls are therefore necessary for achieving complete coverage of the sample. But it is expensive to call and call again. At the same time we cannot afford to neglect the non-response. Results based on response alone will not apply to the entire population from which the sample was selected. Experience from different surveys show that non-response generally differs from the response in several respects and neglecting them will introduce a bias in the results. Under these circumstances, one solution is to take a small subsample of the non-respondents and use all the persuasion, ingenuity and other resources at our command to get a response from them. The two samples can then be combined suitably to get a better estimate of the population parameter.

Hansen and Hurwitz (1946) discussed a method of tackling total non-response in mail interviews. Rao (1986) applied this method of subsampling

the non-respondents for the ratio estimation of the mean when the population mean of the auxiliary character is known

Using an auxiliary variable Okafor (1994) derived the DSS estimator based on the subsampling of the non-respondents, when there is total response on the auxiliary variable and incomplete response on the main character

For practical application of any allocation integer values of the sample sizes are required. This could be done by simply rounding off noninteger sample sizes to their nearest integral values. When the sample sizes are large enough and (or) the measurement cost in various strata are not too high, the rounded off sample allocation may work well.

However in situations other than described above the rounded off sample allocations may become infeasible and (or) non optimal. This means that the rounded off values may violate some of the constraint of the problem and (or) there may exist other sets of integer sample allocations with a better value of the objective function of the formulated NLPP. In such situations we have to use some integer programming technique to obtain an optimum integer solution. In this chapter the problem of obtaining an optimum allocation in DSS, when there is incomplete response on the main character and total response on the auxiliary character, is considered as an all integer nonlinear programming problem (AINLPP). A solution

procedure is developed using the dynamic programming technique. A numerical example is also presented to illustrate the computational details.

4.2 THE PROBLEM

From a population of N units a large sample of size n' is selected by simple random sampling without replacement (SRSWOR). Information on the auxiliary variable x is collected with which an unbiased estimate $w_h = n'_h / n'$ of the true stratum weight $W_h = N_h / N$ is computed.

where n'_h is the number of units in the initial sample that falls in stratum h ,

$$(h=1,2,\dots,L), \text{ with } \sum_{h=1}^L n'_h = n'$$

In each stratum a subsample of size $n_h = v_h n'_h$, ($0 < v_h < 1$), v_h is prefixed, is selected from n'_h by SRSWOR. The main character y is then observed on these n_h units, $h=1,2,\dots,L$.

The DSS estimator of the population mean for the total response is

$$\bar{y}_{ds} = \sum_{h=1}^L w_h \bar{y}_h \tag{4.1}$$

where $\bar{y}_h = \frac{1}{n_h} \sum_{i=1}^{n_h} y_{hi}$, sample mean

The variance of \bar{y}_{ds} is

$$V(\bar{y}_{ds}) = \left(\frac{1}{n'} - \frac{1}{N} \right) S_y^2 + \frac{1}{n'} \sum_{h=1}^L W_h \left(\frac{1}{v_h} - 1 \right) S_y^2 \quad (4.2)$$

where $S_y^2 = \frac{1}{N-1} \sum_{h=1}^L \sum_{i=1}^{N_h} (y_{hi} - \bar{Y})^2$

and $S_{y_h}^2 = \frac{1}{N-1} \sum_{i=1}^{N_h} (y_{hi} - \bar{Y}_h)^2$, variance of y in h-th stratum.

Let

n_{1h} : unit respond at the first call from the n_h units selected in stratum h .

n_{2h} : units do not respond.

Thus the subsample of size n_h is again subdivided into respondent and non-respondent subsamples of sizes n_{1h} and n_{2h} respectively, where $n_{1h} + n_{2h} = n_h$. A subsample of size m_{2h} out of the n_{2h} non-respondents of h-th stratum is selected and interviewed with improved methods, where $m_{2h} = k_h^* n_{2h}$ ($0 < k_h^* < 1$), k_h^* is prefixed.

An unbiased, estimator \bar{y}_{ds}^* for \bar{Y} based on the sample means from respondents and non-respondents (in second attempt) is given as

$$\bar{y}_{ds}^* = \sum_{h=1}^L w_h \bar{y}_h^*, \text{ where } \bar{y}_h^* = \frac{n_{1h} \bar{y}_{1h} + n_{2h} \bar{y}_{m_{2h}}}{n_h} \quad (4.3)$$

\bar{y}_{1h} = sample mean for respondents based on n_{1h} units

$\bar{y}_{m_{2h}}$ = sample mean for the non-respondents based on m_{2h} units

The variance of \bar{y}_{ds}^* is

$$V(\bar{y}_{ds}^*) = V(\bar{y}_{ds}) + \frac{1}{n'} \sum_{h=1}^L W_{2h} \frac{1 - k_h^*}{k_h^* v_h} S_{2yh}^2 \quad (4.4)$$

$W_{2h} = N_{2h} / N$, population proportion of the non-respondents in stratum h .

S_{2yh}^2 , is the population variance among the non-respondents in stratum h .

(see Hansen and Hurwitz (1946) and Rao (1986)).

The problem now is to find the optimum sizes of the subsamples m_{2h} , $h=1, 2, \dots, L$ for which $V(\bar{y}_{ds}^*)$ given by (4.4) is minimum for a fixed cost. This problem may be divided into two phases.

Phase I: In this phase the optimum values of n_h , $h=1, 2, \dots, L$ are obtained for which $V(\bar{y}_{ds}^*)$ is minimum for a fixed sample size $n = \sum_{h=1}^L n_h$.

Phase II: In this phase the optimum values of m_{2h} ; $h=1, 2, \dots, L$ are obtained for a fixed total cost of the survey.

4.3 FORMULATION OF THE PHASE-I PROBLEM

Using (4.2) and (4.4) the problem of first phase can be formulated as

$$\text{Minimize } V(\bar{y}_{ds}^*) = \left(\frac{1}{n'} - \frac{1}{N} \right) S_y^2 + \frac{1}{n'} \sum_{h=1}^L W_h \left(\frac{1}{v_h} - 1 \right) S_{yh}^2$$

$$+ \frac{1}{n'} \sum_{h=1}^L W_{2h} \left(\frac{1-k_h^*}{k_h^* v_h} \right) S_{2yh}^2 \quad (4.5)$$

$$\text{subject to } \sum_{h=1}^L n_h = n \quad (4.6)$$

$$1 \leq n_h \leq N_h \quad (4.7)$$

$$\text{and } n_h \text{ are integers; } h = 1, 2, \dots, L \quad (4.8)$$

Ignoring the terms independent of n_h the objective function in (4.5) can be expressed as

$$\begin{aligned} Z(n_1, n_2, \dots, n_L) &= \frac{1}{n'} \sum_{h=1}^L \left(\frac{W_h n'_h S_{yh}^2 + W_{2h} \{(1-k_h^*)/k_h^*\} n'_h S_{2yh}^2}{n_h} \right) \\ &= \sum_{h=1}^L \frac{a_h}{n_h} \end{aligned}$$

$$\text{where } a_h = \left[\frac{W_h n'_h S_{yh}^2 + W_{2h} \{(1-k_h^*)/k_h^*\} n'_h S_{2yh}^2}{n'} \right] \quad (4.9)$$

The problem (4.5)-(4.8) may be simplified as

$$\text{Minimize } Z(n_1, n_2, \dots, n_L) = \sum_{h=1}^L \frac{a_h}{n_h} \quad (4.10)$$

$$\text{subject to } \sum_{h=1}^L n_h = n \quad (4.11)$$

$$1 \leq n_h \leq N_h \quad (4.12)$$

$$\text{and } n_h \text{ are integers; } h=1,2,\dots,L \quad (4.13)$$

The restriction (4.12) are imposed to avoid over sampling, that is, the situation where $n_h > N_h$ and to have the representation of every stratum in the sample.

4.4 SOLUTION OF THE PHASE-I PROBLEM

Ignoring restrictions in (4.12) and (4.13) and using Lagrangians multipliers technique, the optimum value of n_h that minimize (4.10) subject to (4.11) may be obtained as given below.

$$\phi(n_h, \lambda) = \sum_{h=1}^L \left(\frac{a_h}{n_h} \right) + \lambda \left(\sum_{h=1}^L n_h - n \right)$$

differentiating ϕ partially w.r.t. n_h and equating to zero we get

$$\frac{\partial \phi}{\partial n_h} = -\frac{a_h}{n_h^2} + \lambda = 0; \quad h=1,2,\dots,L$$

$$\text{or } a_h = \lambda n_h^2; \quad h=1,2,\dots,L$$

$$n_h = \frac{\sqrt{a_h}}{\sqrt{\lambda}}; \quad h=1,2,\dots,L$$

Taking summation on both the sides we get

$$\sum_{h=1}^L n_h = \frac{1}{\sqrt{\lambda}} \sum_{h=1}^L \sqrt{a_h} \quad \text{or } n = \frac{1}{\sqrt{\lambda}} \sum_{h=1}^L \sqrt{a_h}$$

$$\text{or } \frac{1}{\sqrt{\lambda}} = \frac{n}{\sum_{h=1}^L \sqrt{a_h}}$$

which gives

$$n_h = \frac{n\sqrt{a_h}}{\sum_{h=1}^L \sqrt{a_h}}; \quad h = 1, 2, \dots, L \quad (4.14)$$

If the above values of n_h satisfies (4.12) also the non-linear programming problem (NLPP) (4.10)-(4.12) is solved.

In case either some or all of the n_h given by (4.14) violates (4.12) or to get an integer solution restricted by (4.13) the Lagrange multipliers technique could not provide the solution and some other constrained optimization technique is to be used. In the next section a computational procedure to obtain integer values of n_h is developed using dynamic programming technique.

4.5 SOLUTION OF THE PHASE-I PROBLEM USING THE DYNAMIC PROGRAMMING TECHNIQUE

The problem (4.10)-(4.13) can be restated as

$$\text{Minimize } Z(n_1, n_2, \dots, n_L) = \frac{a_1}{n_1} + \frac{a_2}{n_2} + \dots + \frac{a_L}{n_L} \quad (4.15)$$

$$\text{subject to } n_1 + n_2 + \dots + n_L = n \quad (4.16)$$

$$1 \leq n_1 \leq N_1, \dots, 1 \leq n_L \leq N_L \quad (4.17)$$

and n_h are integers; $h=1,2,\dots,L$

(4.18) The objective function and the constraints of the AINLPP

(4.15)-(4.18) are the sum of independent functions of n_h , $h=1,2,\dots,L$

The AINLPP, which is an L -stage decision problem, can be decomposed into L -stage single variable decision problems.

In the following a solution procedure for solving the formulated AINLPP using dynamic programming technique is developed.

Consider the sub-problem called the k -th sub-problem, involving the first ($k < L$) strata and let $f(k, r)$ be the minimum value of the objective

function for the first k strata with total sample size r , i.e.

$$f(k, r) = \min \sum_{h=1}^k \frac{a_h}{n_h} \quad (4.19)$$

$$\text{subject to } \sum_{h=1}^k n_h = r \quad (4.20)$$

$$1 \leq n_h \leq N_h \quad (4.21)$$

$$\text{and } n_h \text{ are integers, } h=1,2,\dots,k \quad (4.22)$$

Thus the problem (4.15)-(4.18) is equivalent to the problem of finding $f(L, n)$. $f(L, n)$ is found recursively by finding $f(k, r)$ for $k=1,2,\dots,L$ and $r=0,1,\dots,n$

$$\text{Now } f(k, r) = \min \left(\frac{a_k}{n_k} + \sum_{h=1}^{k-1} \frac{a_h}{n_h} \right)$$

$$\text{subject to } \sum_{h=1}^k n_h = r - n_k$$

$$1 \leq n_h \leq N_h$$

and n_h are integers, $h = 1, 2, \dots, k$

For fixed integer value of n_k , $1 \leq n_k \leq \min[r, N_k]$, $f(k, r)$ is given by

$$f(k, r) = \frac{a_k}{n_k} + \left\{ \min \sum_{h=1}^{k-1} \frac{a_h}{n_h} \mid \sum_{h=1}^{k-1} n_h = r - n_k, \right.$$

$$\left. 1 \leq n_h \leq N_h; n_h \text{ are integers, where } h = 1, 2, \dots, k-1 \right\}$$

But by definition the terms in $\{ \}$ above is equal to $f(k-1, r - n_k)$.

Suppose we assume that for a given k , $f(k-1, r)$ is known for all possible $r = 0, 1, \dots, n$. Then

$$f(k, r) = \min_{n_k=1, 2, \dots, n} \left[\frac{a_k}{n_k} + f(k-1, r - n_k) \right] \quad (4.23)$$

This is the required dynamic programming recursive formula. Using the relation (4.23) for each $k = 1, 2, \dots, L$ and $r = 0, 1, \dots, n$, $f(L, n)$ can be calculated.

Initially we set $f(k, r) = \infty$, if $r < k$ since we wish to have $n_h \geq 1$, for each $h = 1, 2, \dots, k$, r must be at least equal to k .

Also $f(1, r) = \min[a_1 / n_1, \text{ subject to } n_1 = r, 1 \leq n_1 \leq N_1]$

$$\text{Thus } f(1, r) = \begin{cases} \infty & \text{for } r > N_1 \text{ or } r < 1 \\ a_1 / r & \text{for } 1 \leq r \leq N_1 \end{cases}$$

We tabulate the value of $f(k, r)$ and the optimal n_k , for each k , systematically. Then from $f(L, n)$, optimal n_L can be found; from $f(L-1, n-n_L)$ optimal n_{L-1} can be found and so on until finally we find optimal n_1 . (see Arthenari and Dodge (1981)).

4.6 FORMULATION OF THE PHASE-II PROBLEM

For the second phase of the solution consider the variance function given in (4.5)

$$\begin{aligned} V(\bar{y}_{ds}^*) &= \left(\frac{1}{n'} - \frac{1}{N} \right) S_y^2 + \frac{1}{n'} \sum_{h=1}^L W_h \left(\frac{1}{v_h} - 1 \right) S_{yh}^2 \\ &\quad + \frac{1}{n'} \sum_{h=1}^L W_{2h} \left(\frac{1-k_h^*}{k_h^* v_h} \right) S_{2yh}^2 \end{aligned} \quad (4.24)$$

Assuming the cost function [see Okafor (1994)]

$$C = C_1 n' + \sum_h C_{2h} n_h + \sum_k C_{21h} n_{1h} + \sum_h C_{22h} n_{2h} k_h^*$$

where

C_1 : cost of getting information on the first phase sample.

C_{2h} : cost of first attempt on the main character in stratum h .

C_{21h} : cost of processing the results on the main character from the respondent at the first attempt sample in the stratum h .

C_{22h} : cost of getting and processing results on the main character from the sub sample of the non-respondents at the second phase sample in stratum h .

We also must have $1 \leq m_{2h} \leq n_{2h}$

Ignoring the terms independent of m_{2h} in the R.H.S. of (4.24) and putting $k_h^* = m_{2h} / n_{2h}$ and $v_h = n_h / n'_h$.

The problem becomes

$$\text{Minimize } Z(m_{21}, m_{22}, \dots, m_{2L}) = \frac{1}{n'} \sum_{h=1}^L W_{2h} \left(\frac{n_{2h}}{m_{2h}} \right) \frac{n'_h}{n_h} S_{2yh}^2 \quad (4.25)$$

$$\text{subject to } \sum_{h=1}^L C_{22h} m_{2h} \leq C_0 \quad (4.26)$$

$$\text{and } 1 \leq m_{2h} \leq n_{2h} \quad (4.27)$$

m_{2h} are integers, $h = 1, 2, \dots, L$.

$$\text{And } C_0 = C_1 n' + \sum_h C_{2h} n_h + \sum_h C_{21h} n_{1h}$$

Let

$$b_h = \frac{1}{n'} W_{2h} n_{2h} \frac{n'_h}{n_h} S_{2yh}^2 \quad (4.28)$$

The AINLPP (4.25)-(4.27) may be restated as

$$\text{Minimize } Z(m_{21}, m_{22}, \dots, m_{2L}) = \sum_{h=1}^L \frac{b_h}{m_{2h}}$$

(4.29)

$$\text{subject to } \sum_{h=1}^L C_{22h} m_{2h} \leq C_0 \quad (4.30)$$

$$\text{and } 1 \leq m_{2h} \leq n_{2h} \quad (4.31)$$

$$\text{where } m_{2h} \text{ are integers, } h=1, 2, \dots, L \quad (4.32)$$

4.7 SOLUTION OF THE PHASE-II PROBLEM

Like phase-I applying Lagrangian multipliers technique, with equality in (4.30) and ignoring (4.31) and (4.32) we get,

$$\phi(m_{2h}, \lambda) = \sum_{h=1}^L \frac{b_h}{m_{2h}} + \lambda \left(\sum_{h=1}^L C_{22h} m_{2h} - C_0 \right)$$

Differentiating ϕ with respect to m_{2h} and λ and equating to zero we get

$$\frac{\partial \phi}{\partial m_{2h}} = -\frac{b_h}{m_{2h}^2} + \lambda C_{22h} = 0$$

$$\frac{\partial \phi}{\partial \lambda} = \sum_h C_{22h} m_{2h} - C_0 = 0$$

Solving the above equations we get the optimum value of m_{2h} as

$$m_{2h} = C_0 \frac{b_h / C_{22h}}{\sum_{h=1}^L \sqrt{b_h} C_{22h}} \quad (4.33)$$

4.8 SOLUTION OF THE PHASE-II PROBLEM USING THE DYNAMIC PROGRAMMING TECHNIQUE

Let $f(k, r)$ be the minimum value of the objective function of the problem (4.29)-(4.32), the first k strata with $C_0 = r$ i.e.

$$f(k, r) = \left\{ \min \sum_{h=1}^k \frac{b_h}{m_{2h}} \left| \sum_{h=1}^k C_{22h} m_{2h} \leq r \right. \right. \\ \left. \left. 1 \leq m_{2h} \leq n_{2h}, m_{2h} \text{ are integer, } h = 1, 2, \dots, k \right\} \quad (4.34)$$

with the above definition of $f(k, r)$ the problem is equivalent to the problem of finding $f(L, C_0)$. $f(L, C_0)$ is found recursively by using (4.34) for $k = 1, 2, \dots, L$ and $r = 0, 1, \dots, C_0$.

$$\text{Now } f(k, r) = \min \left(\frac{b_k}{m_{2k}} + \sum_{h=1}^{k-1} \frac{b_h}{m_{2h}} \right)$$

$$\text{subject to } \sum_{h=1}^{k-1} C_{22h} m_{2h} = r - C_{22k} m_{2k}$$

$$\text{and } 1 \leq m_{2h} \leq n_{2h},$$

where m_{2h} are integers, $h = 1, 2, \dots, k-1$.

$$f(k, r) = \left\{ \begin{array}{l} \min \left(\frac{b_k}{m_{2k}} + \sum_{h=1}^{k-1} \frac{b_h}{m_{2h}} \right) \mid \sum_{h=1}^{k-1} C_{22} m_{2h} = r - C_{22k} m_{2k} \\ \text{and.. } 1 \leq m_{2h} \leq n_{2h}, m_{2h} \text{ are integers } h=1, 2, \dots, k-1 \end{array} \right\}$$

and $1 \leq m_{2h} \leq n_{2h}, m_{2h}$ are integers $h=1, 2, \dots, k-1$

For a fixed integer value of $m_{2k}, 1 \leq m_{2k} \leq [r, n_{2k}], f(k, r)$, is given by

$$\text{or } \left\{ \min \left(\frac{b_k}{m_{2k}} + \sum_{h=1}^{k-1} \frac{b_h}{m_{2h}} \right) \mid \sum_{h=1}^{k-1} C_{22} m_{2h} = r - C_{22k} m_{2k} \right. \\ \left. 1 \leq m_{2h} \leq n_{2h}, \text{ and } m_{2h} \text{ are integers } h=1, 2, \dots, k-1 \right\} \quad (4.35)$$

By definition the terms in the braces is equivalent to $f(k-1, r)$ is known for all possible $r = 0, 1, \dots, C_0$. Then

$$f(k, r) = \min_{m_{2k}=1, 2, \dots, C_0} \left[\frac{b_k}{m_{2k}} + f(k-1, r - C_{22k} m_{2k}) \right] \quad (4.36)$$

Using the relation (4.36) for each $k=1, 2, \dots, L$ and $r=0, 1, \dots, C_0$, $f(L, C_0)$ can be calculated. Initially we set $f(k, r) = \infty$, if $r \leq k$. Since we wish to have $m_{2h} \geq 1$ for each $h=1, 2, \dots, k$; r must be at least equal to k .

Also $f(1, r) = \min[b_1 / m_{21}]$ Subject to $m_{21} = r, 1 \leq m_{21} \leq n_{21}$

$$\text{Thus } f(1, r) = \begin{cases} \infty & \text{for } r > n_{21} \text{ or } r < 1 \\ b_1 / r & \text{for } 1 \leq r \leq n_{21} \end{cases}$$

We tabulate the value of $f(k, r)$ and the optimal m_{2k} , for each k , systematically. Then from $f(L, C_0)$ optimal m_{2L} can be found from

$f(L-1, C_0 - m_{2L})$. Optimal m_{2L-1} can be found and so on, until finally we find optimal m_{21} .

4.9 NUMERICAL EXAMPLE

The following numerical example demonstrates the use of the solution procedure. The data used in this example is from Murthy (1967). Here DSS is used to estimate the mean area under cultivation. The area of each village and the area cultivated in the village are converted to hectares and grouped into three strata. Within each stratum, the population was again subdivided into respondent and non-respondent groups. Villages with larger area considered in non-respondent group.

Table 4.1 and 4.2 gives the population parameters obtained from the data as given in Okafor (1994).

Table 4.1

Overall stratum population parameters

Stratum	W_h	S_{yh}^2	v_h	k_h^*
0-930	0.336	39974.81	0.4	0.5
931-1700	0.352	61455.48	0.5	0.6
1701-4300	0.313	172425.05	0.6	0.7

It is assumed that $N = 200$, $n' = 100$, $n = 50$

Using proportional allocation n'_h may be obtained as

$$n'_1 = 33.6 \cong 34, n'_2 = 35.2 \cong 35, \text{ and } n'_3 = 31.3 \cong 31$$

Table 4.2

Class stratum population parameters

Stratum	Class	S_{yh}^2	W_h
0-930	Respondent	7162.51	0.188
	Non Respondent	14549.99	0.148
931-1700	Respondent	19564.45	0.219
	Non Respondent	17386.54	0.133
1701-4300	Respondent	5042.50	0.188
	Non Respondent	71175.11	0.125

For L=3 the Phase-I problem (4.15)-(4.18) can be expressed as

$$\text{Minimize } Z = \frac{a_1}{n_1} + \frac{a_2}{n_2} + \frac{a_3}{n_3} \quad (4.37)$$

$$\text{subject to } n_1 + n_2 + n_3 = 50 \quad (4.38)$$

$$\left. \begin{array}{l} 1 \leq n_1 \leq 34 \\ 1 \leq n_2 \leq 35 \\ 1 \leq n_3 \leq 31 \end{array} \right\} \quad (4.39)$$

$$\text{where } n_h \text{ are integers; } h=1,2,3 \quad (4.40)$$

Table (4.3) gives the optimum values of n_h using formula (4.14). These values of n_h satisfy (4.39) also, hence they will solve NLPP (4.37)-(4.40) completely.

Table 4.3

Calculation of n_h using formula (4.14)

h	a_h	$\sqrt{a_h}$	$n\sqrt{a_h}$	n_h
1	5236 5381	72 363928	3618 1964	12 176312 \cong 12
2	8157 2253	90 317359	4515 868	15 197245 \cong 15
3	18085 764	134 48332	6724 1661	22 627748 \cong 23
$\sum \sqrt{a_h}$		= 297 16461		

The optimal value of the objective function is $Z^* = 1766.09$

For the sake of illustration, the dynamic programming approach to find the integer optimum allocation in Phase-I is also applied to the same problem. Execution of the computer program (in C language, given in Appendix-I) for the procedure given in Section 4.5 for solving the AINLPP (4.19)-(4.22) gives the following solution to the Phase I problem

$$n_1 = 12, \quad n_2 = 15, \quad n_3 = 23$$

The corresponding value of the objective function is $Z^* = 1766.5308$

It can be seen that this solution is same as given in Table 4.3 except for a negligible change in the value of the objective function

For formulating the Phase-II problem, let $C_{22h} = 10, 12, 8$ for $h = 1, 2, 3$ respectively and $C_0 = 100$

Since W_{1h} and W_{2h} are known for $h = 1, 2, 3$ they are used to work out the expected values of n_{2h} , $h = 1, 2, 3$ as $n_{2h} = n_h W_{2h} / (W_{1h} + W_{2h})$. These values are tabulated in Table 4.4

Table 4.4
Calculation of n_{2h}

h	W_{1h}	W_{2h}	n'_h	n_h	S_{2yh}^2	C_{22h}	n_{2h}
1	0.188	0.148	33.60	12	14549.99	10	$5.2857 \cong 5$
2	0.219	0.133	35.20	15	17386.54	12	$5.6676 \cong 6$
3	0.188	0.125	31.20	23	71175.11	8	$9.1853 \cong 9$

For $L = 3$, the Phase-II problem as given in (4.29) to (4.32) is

$$\text{Minimize } Z = \frac{b_1}{m_{21}} + \frac{b_2}{m_{22}} + \frac{b_3}{m_{23}} \quad (4.41)$$

$$\text{subject to } C_{221}m_{21} + C_{222}m_{22} + C_{223}m_{23} \leq C_0 \quad (4.42)$$

$$\left. \begin{array}{l} 1 \leq m_{21} \leq n_{21} \\ 1 \leq m_{22} \leq n_{22} \\ 1 \leq m_{23} \leq n_{23} \end{array} \right\} \quad (4.43)$$

where m_{2h} are integers, $h = 1, 2, 3$ (4 44)

Table 4 5 gives the optimum values of m_{2h} using formula (4 35)

These values of m_{2h} are infeasible, since they violate the restriction

$$\sum_{h=1}^3 C_{22h} m_{2h} \leq C_0 \text{ in (4 37), hence as an alternative, the dynamic}$$

programming approach given in Section 4 8 may be used

Table 4.5

Calculation of m_{2h} using formula (4 35)

h	b_h	$\sqrt{b_h}$	$\sqrt{b_h / C_{22h}}$	$\sqrt{b_h C_{22h}}$	m_{2h}
1	318 70212	17 85223	5 6453708	56 453708	2 6707656 \cong 3
2	307 54977	17 537097	5 0625238	60 750286	2 3950268 \cong 2
3	1108 5576	33 295009	11 771563	94 172506	5 5690027 \cong 6
$\sum \sqrt{b_h C_{22h}}$				= 211 3765	

Execution of the computer program (in C language, given in Appendix-II of this chapter) for the procedure developed in Section 4 8 for solving the AINLPP (4 29)-(4 32) gives the following results

$$m_{21} = 3, \quad m_{22} = 2, \quad m_{23} = 5$$

The optimum value of the objective function (4 41) is $Z^* = 481 72045$

APPENDIX-I

```
#include<stdio.h>
#define K_MAX 3
#define R_MAX 50
#define INF 9999999.0

main( )
{
    int l,n[4][51],k,r,i,m,nk;
    float f[4][51],min;
    float a[4]={1,5236.5381,8157.2253,18085.764};
    float Nk[4]={1,34,35,31};
    FILE *op;
    op=fopen("result1.dat","w+");
    f[0][0]=0.0;
    f[1][0]=INF;
    f[2][0]=INF;
    f[3][0]=INF;
    l=0;
    /*Initialization of zero point functions */
    for(i=1;i<=50;i++)
        f[1][i]=INF;
    /*Starting with k */
    for(k=1;k<=K_MAX;k++)
    {
        /*Starting with r */
        for(r=1;r<=R_MAX;r++)
        {
            if(r<k)
                f[k][r]=INF;
            min=INF;
            for(nk=1;nk<=r;nk++)
            {
                if(nk>=1 && nk<=Nk[k])
                /* Implementing the recursion function */
                f[k][r]=a[k]/nk+f[k-1][r-nk];
                if(f[k][r]<min)
                {
                    min=f[k][r];
                    n[k][r]=nk;
                } /* End of if */
            } /* End of nk loop */
            f[k][r]=min;
        } /* End of r loop */
    }
}
```

```

    } /* End of K loop */
/* Saving Output in a file */
fprintf(op,"|-----|
-----|\n");
fprintf(op," r f[1,r]      n1      f[2,r]      n2
      f[3,r]      n3\n");
fprintf(op,"|-----|
-----|
|\n");
for(r=1;r<=R_MAX;r++)
  for(k=1;k<=K_MAX;k++)
  {
    if(f[k][r]==INF)
    {
      f[k][r]=0;
      n[k][r]=0;
    }
    if(k==1)
fprintf(op," %d      %10.4f
%d\t",r,f[k][r],n[k][r]);
    if(k>1)
fprintf(op,"      %10.4f      %d\t",f[k][r],n[k][r]);
    if(k==3)
      fprintf(op," \n");
  }
/* Appending the result to the output file */
m=R_MAX;
for(k=K_MAX;k>=1;k--)
{
fprintf(op," \nThe value of n[%d]=%d",k,n[k][m]);
  m=m-n[k][m];
}
fprintf(op," \n|-----END-----|
|\n");
getch( );
return;
}
/* End of Program */

```

APPENDIX-II

```

#include<stdio.h>
#define K_MAX 3
#define R_MAX 100
#define INF 9999999.0

    main( )
    {
        int l,m2[4][101],k,j,r,i,m,m2k;
        double f[4][101],min;
        double
b[4]={1,318.70212,307.54977,1108.5576};
        double n2k[4]={1,5,6,9};
        double c22k[4]={1,10,12,8};
        FILE *op;
        op=fopen("result4.dat","w+");
        f[0][0]=0.0;
                f[1][0]=INF;
                f[2][0]=INF;
                f[3][0]=INF;
                l=0;
        for(i=1;i<=100;i++)
            f[1][i]=0.0;
        for(k=1;k<=K_MAX;k++)
            {
                for(r=1;r<=R_MAX;r++)
                    {
                        if(r<k)

f[k][r]=INF;

                                                min=INF;
                for(m2k=1;m2k<=r;m2k++)
                    {
                        if(r<c22k[k]*m2k)
                            f[k-1][r-c22k[k]*m2k]=INF;
                            if(m2k>=1
&& m2k<=n2k[k])
                                f[k][r]=b[k]/m2k+f[k-1][r-c22k[k]*m2k];

if(f[k][r]<min)

min=f[k][r];
                m2[k][r]=m2k;

```

```

} /* End
of if */
} /* End of m2k loop
*/
} /* End of r loop
*/
} /* End of K loop */
/* Saving Output in a file */
fprintf(op,"|-----|
-----|\n");
fprintf(op," r f[1,r] m21 f[2,r]
m22 f[3,r] m23\n");
fprintf(op,"|-----|
-----|\n");
for(r=1;r<=R_MAX;r++)
for(k=1;k<=K_MAX;k++)
{
if(f[k][r]==INF)
{
f[k][r]=0;
m2[k][r]=0;
}
if(k==1)
fprintf(op," %d %10.4f
%d\t",r,f[k][r],m2[k][r]);
if(k>1)
fprintf(op,"%10.4f %d\t",f[k][r],m2[k][r]);
if(k==3)
fprintf(op,"\n");
}
/* Appending the result to the output
file */
m=R_MAX;
for(k=K_MAX;k>=1;k--)
{
fprintf(op,"\n The value of
m2[%d]=%d",k,m2[k][m]);
m=m-m2[k][m];
}
fprintf(op,"-----END-----|
|\n");
getch();
return;
}

/* END OF PROGRAM */

```

CHAPTER-V

THE PROBLEM OF OPTIMUM STRATIFICATION UNDER NEYMAN ALLOCATION: A MATHEMATICAL PROGRAMMING APPROACH

CHAPTER-V

THE POBLEM OF OPTIMUM STRATIFICATION UNDER NEYMAN ALLOCATION: A MATHEMATICAL PROGRAMMING APPROACH

5.1 INTRODUCTION

As given in chapter I, Section IA. 5, the use of stratified sampling involves the solution of four carefully formulated optimization problems according to the objective and available resources to the sample survey. These four optimization problems are related to the optimum choice of the

- (i) Stratification variable
- (ii) Number of strata
- (iii) Stratum boundaries
- (iv) Sample size allocations

In this chapter the problem of selecting the optimum strata boundaries is discussed as anMPP and a solution procedure is proposed that uses dynamic programming technique. This chapter is based on my research paper entitled “Optimum Stratification for exponential study variable under Neyman Allocation” accepted for presentation in the 5th International Symposium on Optimization and Statistics, to be held in this department during December 28-30, 2002.

The basic consideration involving in the formation of strata is that the strata should be internally as homogenous as possible, that is stratum variances S_h^2 are as small as possible. If the distribution of the study variable is available the strata would be created by cutting this distribution at suitable points.

Given the number of strata, Dalenius and Gruney (1951) suggested that the strata boundaries be so determined that $W_h S_h$ remain constant.

Mahalanobis (1952) and Hansen, Hurwitz and Madow (1953) have suggested that strata boundaries be so determined that $W_h \bar{Y}_h$ remain constraint. Dalenius and Hodges (1959) have supported the work of Dalenius and Gruney (1951).

Dalenius (1957) has worked out the best stratum boundaries under proportional and Neyman allocation. Ekman (1959) has suggested approximation to complicated theoretical solutions. Cochran (1961) has examined the applications of these approximations through the empirical studies. Sethi (1963) has showed that the above suggestions fail to provide optimum strata boundaries for certain types of populations. He derived the solutions for optimum stratification points for certain populations. Hess, Sethi and Balakrishnan (1966) have applied these solutions to some empirical studies and made a comparison of various approximations. Singh and Sukhatme (1969) have suggested several approximate methods to obtain optimal points of stratification. Singh & Sukhatme (1973) have suggested certain rules for

obtaining optimal stratification points based on auxiliary information. Some others who worked on this problem are Singh (1977), Unnithan (1978), Yadav and Singh (1984) etc.

Khan et al (2002b) have formulated the problem of optimum stratification as a mathematical programming problem and developed a solution procedure using dynamic programming technique. They have applied their procedure to work out optimum strata boundaries to populations having uniform and right triangular distributions.

Most of the authors who worked on this problem obtained minimal equations for optimum strata boundaries. Unfortunately these equations are difficult to solve for exact solutions. So that only approximate solution can be obtained. Some authors suggested iterative procedures that are very slow even to obtain a local minimum of the objective function. Moreover, the iterative procedures may oscillate and there is no guarantee that they will provide us with the approximate global minimum.

In this chapter the approach of Khan et al (2002b) is extended to work out optimum strata boundaries for an exponential population under Neyman allocation.

5.2 THE PROBLEM

Let the population under study is to be stratified into L strata and the estimation of the population mean is of interest. Let x_0 and x_L be the smallest

and largest values of the study variable x in the population. The problem of optimum stratification can be described as to find the intermediate stratum boundaries x_1, x_2, \dots, x_{L-1} such that the variance of the stratified sample mean \bar{x}_{st} under Neyman (1934) allocation is minimum.

The variance of the stratified sample mean

$$\bar{x}_{st} = \sum_{h=1}^L W_h \bar{x}_h \quad (5.1)$$

under Neyman allocation is given as

$$V(\bar{x}_{st}) = \frac{\left(\sum_{h=1}^L W_h S_h \right)^2}{n} - \frac{\sum_{h=1}^L W_h S_h^2}{N} \quad (5.2)$$

where the symbols have the same meaning as described in Section 1A.4 of Chapter I except that the study variable is denoted by x .

If the finite population correction is ignored, minimizing expression on the right hand side of (5.2) is equivalent to minimizing

$$\sum_{h=1}^L W_h S_h \quad (5.3)$$

Let $f(x)$ denotes frequency function of the study variable x , $x_0 \leq x \leq x_L$. The problem of determining the strata boundaries is equivalent to cut up the range

$$x_L - x_0 = d \text{ (say)} \quad (5.4)$$

at points $x_1 \leq x_2 \leq \dots, \leq x_{L-1}$ such that (5.3) is minimum. Where the values of

W_h and S_h are obtained by

$$W_h = \int_{x_{h-1}}^{x_h} f(x)dx \quad (5.5)$$

$$S_h^2 = \frac{1}{W} \int_{x_{h-1}}^{x_h} f(x)dx - \mu_h^2 \quad (5.6)$$

$$\text{where } \mu_h = \frac{1}{W_h} \int_{x_{h-1}}^{x_h} xf(x)dx \quad (5.7)$$

and (x_{h-1}, x_h) are the boundaries of hth the stratum.

When the frequency function $f(x)$ is known, using (5.5), (5.6) and (5.7),

$W_h S_h$ could be expressed as a function of x_h and x_{h-1} only.

Let $f_h(x_h, x_{h-1}) = W_h S_h$

Then the problem of determining the optimum strata boundaries (OSB) can be expressed as:

“Find x_1, x_2, \dots, x_{L-1} that minimize $\sum_{h=1}^L f_h(x_h, x_{h-1})$, subject to the constraints

$$x_0 \leq x_1 \leq x_2 \leq \dots, \leq x_{L-1} \leq x_L”$$

Define

$$y_h = x_h - x_{h-1} ; \quad h = 1, 2, \dots, L$$

where $y_h \geq 0$ denotes the width of the hth stratum.

With the above definition of y_h (5.4) can be expressed as

$$\sum_{h=1}^L y_h = \sum_{h=1}^L (x_h - x_{h-1}) = x_L - x_0 = d$$

The k th stratification point x_k ; $k=1,2,\dots,L-1$ can then be given as

$$x_k = x_0 + y_1 + y_2 + \dots + y_k$$

Then the problem of determining optimum strata boundaries can be considered as the problem of determining optimum strata widths and can be expressed as the following Mathematical Programming Problem (MPP):

$$\left. \begin{array}{l} \text{Minimize } \sum_{h=1}^L f_h(y_h, x_{h-1}) \\ \text{subject to } \sum_{h=1}^L y_h = d \\ \text{and } y_h \geq 0; h = 1, 2, \dots, L \end{array} \right\} \quad (5.8)$$

For $h=1$ the term $f_1(y_1, x_0)$ in the objective function of (5.8) is a function of y_1 alone, as x_0 is known for $h=2$ the term $f_2(y_2, x_1) = f_2(y_2, x_0 + y_1)$ will become a function of y_2 alone once y_1 is known. Thus, we may rewrite the MPP (5.8) expressing the objective function a function of y_h alone as:

$$\left. \begin{array}{l} \text{Minimize } \sum_{h=1}^L f_h(y_h) \\ \text{subject to } \sum_{h=1}^L y_h = d \\ \text{and } y_h \geq 0; h = 1, 2, \dots, L \end{array} \right\} \quad (5.9)$$

Let x follows an exponential frequency function:

$$f(x) = e^{-x}, x \geq 0 \quad (5.10)$$

$$= 0; \text{ otherwise}$$

In practice the actual populations are often finite, assuming the largest value of x in the population as D , (5.10) can be rewritten as

$$f(x) = e^{-x}; 0 \leq x \leq D$$

$$= 0; \text{ otherwise}$$

$$\Rightarrow x_o = 0 \text{ and } x_L = D$$

From (5.5)

$$W_h = \int_{x_{h-1}}^{x_h} f(x) dx$$

$$= \int_{x_{h-1}}^{y_h + x_{h-1}} f(x) dx \quad (\text{because } x_h = y_h + x_{h-1})$$

$$= \int_{x_{h-1}}^{y_h + x_{h-1}} e^{-x} dx$$

$$= \left[-e^{-x} \right]_{x_{h-1}}^{y_h + x_{h-1}}$$

$$= e^{-x_{h-1}} - e^{-(y_h + x_{h-1})}$$

$$\text{or } W_h = e^{-x_{h-1}} (1 - e^{-y_h}) \quad (5.11)$$

From (5.7)

$$\mu_h = \frac{1}{W_h} \int_{x_h}^{y_h + x_{h-1}} x f(x) dx$$

$$= \frac{1}{W_h} \int_{x_{h-1}}^{y_h + x_{h-1}} x e^{-x} dx$$

$$\begin{aligned}
&= \frac{1}{W_h} \left[-xe^{-x} + \int e^{-x} dx \right]_{x_{h-1}}^{y_h+x_{h-1}} \\
&= \frac{1}{W_h} \left[-xe^{-x} - e^{-x} \right]_{x_{h-1}}^{y_h+x_{h-1}} \\
&= \frac{1}{W_h} \left[-e^{-x}(1+x) \right]_{x_{h-1}}^{y_h+x_{h-1}} \\
&= \frac{1}{W_h} \left[e^{-x_{h-1}}(1+x_{h-1}) - e^{-(y_h+x_{h-1})}(1+y_h+x_{h-1}) \right] \\
&= \frac{1}{W_h} e^{-x_{h-1}} \left[(1+x_{h-1}) - e^{-y_h}(1+y_h+x_{h-1}) \right] \\
&= \frac{e^{-x_{h-1}} \left[(1+x_{h-1})(1-e^{-y_h}) - y_h e^{-y_h} \right]}{W_h}
\end{aligned}$$

$$\text{Therefore } \mu_h = \frac{e^{-x_{h-1}} \left[(1+x_{h-1})(1-e^{-y_h}) - y_h e^{-y_h} \right]}{e^{-x_{h-1}}(1-e^{-y_h})}$$

$$\text{or } \mu_h = \frac{\left[(1+x_{h-1})(1-e^{-y_h}) - y_h e^{-y_h} \right]}{1-e^{-y_h}} \quad (5.12)$$

$$\text{From (5.6) } S_h^2 = \frac{1}{W_h} \int_{x_{h-1}}^{y_h+x_{h-1}} x^2 f(x) dx - \mu_h^2$$

$$= \frac{1}{W_h} \int_{x_{h-1}}^{y_h+x_{h-1}} x^2 e^{-x} dx - \mu_h^2$$

$$= \frac{1}{W_h} \left[-x^2 e^{-x} + \int 2x e^{-x} dx \right]_{x_{h-1}}^{y_h+x_{h-1}} - \mu_h^2$$

$$= \frac{1}{W_h} \left[-x^2 e^{-x} - 2x e^{-x} - 2e^{-x} - 2e^{-x} \right]_{x_{h-1}}^{y_h+x_{h-1}} - \mu_h^2$$

$$\begin{aligned}
&= \frac{1}{W_h} \left[-e^{-x} (x^2 + 2x + 2) \right]_{x_{h-1}}^{y_h + x_{h-1}} - \mu_h^2 \\
&= \frac{1}{W_h} \left[e^{-x_{h-1}} (x_{h-1}^2 + 2x_{h-1} + 2) - e^{-(y_h + x_{h-1})} ((y_h + x_{h-1})^2 \right. \\
&\quad \left. + 2(y_h + x_{h-1}) + 2) - \mu_h^2 \right] \\
&= \frac{e^{-x_{h-1}} \left[(x_{h-1}^2 + 2x_{h-1} + 2) - e^{-y_h} (y_h^2 + 2y_h x_{h-1} + x_{h-1}^2 + 2y_h + 2x_{h-1} + 2) \right]}{W_h} \\
&\quad - \mu_h^2 \\
&= \frac{\left[(x_{h-1}^2 + 2x_{h-1} + 2) - e^{-y_h} (x_{h-1}^2 + 2x_{h-1} + 2) - y_h e^{-y_h} (y_h + 2x_{h-1} + 2) \right]}{(1 - e^{-y_h})} \\
&\quad - \mu_h^2 \\
&= \frac{(x_{h-1}^2 + 2x_{h-1} + 2) \left(1 - e^{-y_h} \right) - y_h e^{-y_h} (y_h + 2x_{h-1} + 2)}{(1 - e^{-y_h})} \\
&\quad - \left[\frac{(1 - x_{h-1})(1 - e^{-y_h}) - y_h e^{-y_h}}{1 - e^{-y_h}} \right]^2 \quad \text{(using (5.12))}
\end{aligned}$$

Putting $a_h = (1 - e^{-y_h})$ we get

$$\begin{aligned}
S_h^2 &= \frac{(x_{h-1}^2 + 2x_{h-1} + 2)a_h - y_h e^{-y_h} (y_h + 2x_{h-1} + 2)}{a_h} \\
&\quad - \frac{(1 + x_{h-1})^2 a_h^2 - y_h^2 e^{-2y_h}}{a_h^2}
\end{aligned}$$

$$\begin{aligned}
&= \frac{a_h^2 - y_h^2 e^{-y_h} a_h - y_h^2 e^{-2y_h}}{a_h^2} \\
&= \frac{a_h^2 - y_h^2 e^{-y_h} (a_h + e^{-y_h})}{a_h^2} \\
&= \frac{a_h^2 - y_h^2 e^{-y_h}}{a_h^2} \\
\text{or } S_h^2 &= \frac{(1 - e^{-y_h})^2 - y_h^2 e^{-y_h}}{(1 - e^{-y_h})^2}
\end{aligned}$$

Which gives

$$S_h = \frac{\left[(1 - e^{-y_h})^2 - y_h^2 e^{-y_h} \right]^{1/2}}{(1 - e^{-y_h})}$$

Using (5.11) we get

$$W_h S_h = e^{-x_{h-1}} \left[(1 - e^{-y_h})^2 - y_h^2 e^{-y_h} \right]^{1/2} \quad (5.13)$$

Using (5.11), (5.12) and (5.13) the MPP (5.9), can be restated as:

$$\left. \begin{aligned}
&\text{Minimize } \sum_{h=1}^L e^{-x_{h-1}} \left[(1 - e^{-y_h})^2 - y_h^2 e^{-y_h} \right]^{1/2} \\
&\text{subject to } \sum_{h=1}^L y_h = d \\
&\text{and } y_h \geq 0; h = 1, 2, \dots, L
\end{aligned} \right\} \quad (5.14)$$

where x follows the exponential distribution as defined in (5.10).

5.3 THE SOLUTION

Consider the following subproblems of (5.9) for first k ($<L$) strata

$$\left. \begin{array}{l} \text{Minimize } \sum_{h=1}^L f_h(y_h) \\ \text{subject to } \sum_{h=1}^L y_h = d_k \\ \text{and } y_h \geq 0; h = 1, 2, \dots, k \end{array} \right\} \quad (5.15)$$

where $d_k < d$ is the total width available for division into k strata.

Note that $d_k = d$ for $k = L$.

We then have

$$d_k = y_1 + y_2 + \dots + y_k$$

$$d_{k-1} = y_1 + y_2 + \dots + y_{k-1} = d_k - y_k$$

$$d_{k-2} = y_1 + y_2 + \dots + y_{k-2} = d_{k-1} - y_{k-1}$$

\vdots

$$d_2 = y_1 + y_2 = d_3 - y_3$$

and $d_1 = d_2 - y_2$

Let $f(k, d_k)$ denotes the minimum value of objective function of (5.15), that

$$\text{is, } f(k, d_k) = \min \left[\sum_{h=1}^k f_h(y_h) \mid \sum_{h=1}^k y_h = d_k \text{ and } y_h \geq 0; h = 1, 2, \dots, k \right]$$

The recurrence relation of the Dynamic Programming thus be given as

$$f(k, d_k) = \min_{0 \leq y_k \leq d_k} [f_k(y_k) + f(k-1, d_k - y_k)], k \geq 2 \quad (5.16)$$

For the first stage (i.e. k=1):

$$f(1, d_1) = f_1(d_1) \Rightarrow y^* = d_1 \quad (5.17)$$

From $f(L, d)$ the optimum width of Lth stratum, y_L^* , is obtained ;

from $f(L-1, d - y_L^*)$ the optimum width of (L-1) th stratum y_{L-1}^* , is obtained

and so on until y_1^* is obtained.

Using (5.16) and (5.17) we get for first stage (k=1)

$$f(1, d_1) = \left[\left(1 - e^{-d_1}\right)^2 - d_1^2 e^{-d_1} \right]^{1/2} \text{ at } y_1^* = d_1 \quad (5.18)$$

because $x_{k-1} = x_0 = 0$ when $K = 1$

For the stage $k \geq 2$

$$f(k, d_k) = \min_{0 \leq y_k \leq d_k} \left[e^{-(d_k - y_k)} \left[\left(1 - e^{-y_k}\right)^2 - y_k^2 e^{-y_k} \right]^{1/2} + f(k-1, d_k - y_k) \right] \quad (5.19)$$

where $x_{k-1} = x_0 + y_1 + y_2 + \dots + y_{k-1} = d_k - y_k$

5.4 A NUMERICAL EXAMPLE

Relation (5.18) and (5.19) are the required relations of the dynamic programming. Execution of the computer program in 'Java SDK 2', given in

Appendix, of this chapter, gives the optimum stratum boundaries for the exponential study variable with density function

$$f(x) = \begin{cases} e^{-x}, & x > 0 \\ 0, & \text{elsewhere} \end{cases} \quad (5.20)$$

for 2, 3, 4 and 5 strata. The results are presented in a tabular form in Table 5.1

along with the values of $\sum_{h=1}^L W_h S_h$.

Table 5.1

Optimum strata boundaries for 2, 3, 4 and 5 strata

No. of strata L	Strata widths y_h^*	Strata boundary points $x_h^* = x_{h-1}^* + y_h^*$	Optimum value of the objective function $\sum_{h=1}^L f_h(y_h) = \sum_{h=1}^L W_h S_h$
2	$y_1^* = 1.2610$ $y_2^* = 18.7390$	$x_1^* = x_0 + y_1^* = 1.2610$	0.5341
3	$y_1^* = 0.7678$ $y_2^* = 1.2501$ $y_3^* = 17.9821$	$x_1^* = x_0 + y_1^* = 0.7678$ $x_2^* = x_1^* + y_2^* = 2.0179$	0.3648
4	$y_1^* = 0.5509$ $y_2^* = 0.7638$ $y_3^* = 1.2513$ $y_4^* = 17.4340$	$x_1^* = x_0 + y_1^* = 0.5509$ $x_2^* = x_1^* + y_2^* = 1.3147$ $x_3^* = x_2^* + y_3^* = 2.5650$	0.2770
5	$y_1^* = 0.4393$ $y_2^* = 0.5610$ $y_3^* = 0.7569$ $y_4^* = 1.2688$ $y_5^* = 16.9740$	$x_1^* = x_0 + y_1^* = 0.4393$ $x_2^* = x_1^* + y_2^* = 1.0003$ $x_3^* = x_2^* + y_3^* = 1.7572$ $x_4^* = x_3^* + y_4^* = 2.0260$	0.2233

The total width available for cutting stratum boundaries is taken as 20 units, that is $x_L = D = 20$, because the area above $x = 20$ for exponential distribution given in (5.20) is almost zero.

5.5 CONCLUSION

Unnithan (1978) showed that the iterative procedure by Dalenius and Hodge (1959) is slow even to obtain a local minimum; also it does not suggest any stopping rule and may oscillate. He also suggested an iterative solution procedure using modified Newton's method. Both these procedures require initial approximate solutions. Also there is no guarantee that these procedures will provide a global optimum. The advantage of the proposed solution procedure is that it provides a global minimum.

APPENDIX

```
import java.io.*;
import java.util.*;

public class OptimumNew
{
    private RandomAccessFile randReader[] = null;
    private double e=2.718281828;
    private double increment = 0.10;
    private int intPreci = 1;
    private int intStage = 1;
    private int Dk = 999;
    DataOutputStream outputStream[];
    double storedFk[];

    public static void main(String args[])
    {
        new OptimumNew();
    }

    public OptimumNew()
    {
        System.out.println("enter the Stage value (1 to 9 only):");
        String str = Readline.readLine();
        intStage = Integer.parseInt(str);

        System.out.println("enter the summation Yk ( Dk ) value (integer
only):");
        str = Readline.readLine();
        Dk = Integer.parseInt(str);
        System.out.println("enter the desired precesion 1- 9 (integer
only):");
        str = Readline.readLine();
        intPreci = Integer.parseInt(str);

        try
        {
            randReader = new RandomAccessFile[intStage];
```

```

        for(int i =0; i < intStage; i++)
        {
            File file = new File("./Stage"+(i+1)+".txt");
            randReader[i] = new RandomAccessFile(file, "r");
        }
        FileOutputStream fos[] = new
FileOutputStream[intStage];
        outputStream = new DataOutputStream[intStage];
        for(int i =0; i < intStage; i++)
        {
            File file = new File("./Stage"+(i+1)+".txt");
            fos[i] = new FileOutputStream(file);
            outputStream[i] = new DataOutputStream(fos[i]);
        }
        funF1D1() ;
        for(int i = 1; i < intStage; i++)
            funFkDk(i) ;
        backWardCalculation();

    }
    catch(Exception ex)
    {
        ex.printStackTrace();
    }
}

```

```

void funF1D1()
{
    storedFk = new double[(int)(Dk*Math.pow(10,
intPreci)+1)];

    double Y1=0;
    double dblTmp1 = 0;
    double fx= 0;
    double d1 = 0;
    long d1Count=0;
    int count = 0;
    String strD1 = "", strFx="", strY1="";
    increment = Math.pow(10, -intPreci);
    //System.out.println(increment);
    while(d1 <= Dk)
    {

```

```

        Y1 = d1;
        fx = (1 - Math.pow(e, -Y1))*(1 - Math.pow(e, -
Y1)) - Y1*Y1*Math.pow(e, -Y1);
        if(fx<0.0)
        {
            System.out.println("SQRT OF THE -VE
QUANTITY in funFkDk_");
            System.out.println("d1="+d1+",
Y1="+Y1+" , fx= "+fx+"\n");
            System.exit(0);
        }
        else
            dblTmp1=Math.sqrt(fx);

        fx=Math.pow(e, -(d1-Y1))*dblTmp1;
        storedFk[count] = fx;
        count++;
        strFx = Double.toString(fx);
        while(strFx.length() < 25)
        {
            strFx = "0" + strFx;
        }
        strY1 = Double.toString(Y1);
        while(strY1.length() < 25)
        {
            strY1 = "0" + strY1;
        }
        strD1 = Double.toString(d1);
        while(strD1.length() < 25)
        {
            strD1 = "0" + strD1;
        }
        try
        {
            outputStream[0].writeBytes(strD1+" " +
strY1+" " + strFx+"\n");
        }
        catch(Exception ex)
        {
            ex.printStackTrace();
        }
        //d1 += increment;
        d1Count++;

```

```

        d1 = d1Count*Math.pow(10, -
intPrci);//increment;
    }
}

double readFkDk1(int K, double Dk)
{
    double tmpDk = Dk*Math.pow(10, intPrci);
    long lDk = (long)Dk;
    long n1 = (long)tmpDk;
    //Math.round(Dk*Math.pow(10, intPrci));
    String str= "";
    double ret=0, data1 =0, data2 =0;
    n1 = n1*78;
    try
    {
        if(n1 < 0 || n1 > randReader[K].length()) return 0;
        randReader[K].seek(n1);
        str = randReader[K].readLine();
        data1 = Double.parseDouble(str.substring(51));
        if(str != null && str.length() >= 75)
        {
            data2 =
Double.parseDouble(str.substring(26, 51));
        }
        else
            data2 = data1;
        ret = data1+ (data2-data1)*(Dk*100 -
lDk*100)/100;
        //System.out.println( "fkdk- Dk passed =" + Dk + ",
line = "+n1/78 +", Fx=" + ret );
    }
    catch(Exception ex)
    {
        System.out.println( "fkdk- Dk passed =" + Dk + ",
line = "+n1/78 +", str=" + str );
        ex.printStackTrace();
        System.exit(0);
    }
    return ret;
}
}

```



```

double readFkDk(int K, double Dk)
{
    double tmpDk = Dk*Math.pow(10, intPreci);
    long lDk = (long)Dk;
    int n1 = (int)tmpDk; //Math.round(Dk*Math.pow(10,
intPreci));

    String str= "";
    double ret=0, data1 =0, data2 =0;
    try
    {
        data1 = storedFk[n1];
        if(n1 < storedFk.length-1)
        {
            data2 = storedFk[n1+1];
        }
        else
            data2 = data1;
        ret = data1+ (data2-data1)*(Dk*100 -
lDk*100)/100;
        //System.out.println( "fkdk- Dk passed =" + Dk + ",
line = "+n1/78 +", Fx=" + ret );
    }
    catch(Exception ex)
    {
        System.out.println( "readFkDk- Dk passed =" + Dk
+ ", line = "+n1/78 +", str=" + str );
        ex.printStackTrace();
        System.exit(0);
    }
    return ret;
}

void funFkDk(int K)
{
    if(K > 1)
        readStoredFk(K);
    double Yk=0;
    double dblTmp1 = 0;
    long multi = (long)Math.pow(10, intPreci+1);
    double fx= 0;

```



```

System.out.println("Sqrt of the
-ve quantity in fun2_");

System.out.println("Yk="+Yk+" ,increTmp="+increTmp+" ,dk="+dk+"
,dblTmp1="+dblTmp1+" \n"+Math.pow(e, -Yk));
System.exit(0);
}
else
dblTmp1 =
Math.sqrt(dblTmp1);

fx = Math.pow(e,-(dk-Yk)) *
dblTmp1 + readFkDk(K-1, dk-Yk);
if(minFx > fx)
{
minFx = fx;
minYk = Yk;
}
Yk += increTmp;
}
lowLimit = minYk-increTmp;
upperLimit = minYk + increTmp;
if(upperLimit > dk ) upperLimit = dk;
if(lowLimit < 0 ) lowLimit = 0;
increTmp = increTmp/10;
}

strFx = Double.toString(minFx);
while(strFx.length() < 25)
{
strFx = "0" + strFx;
}
strY1 = Double.toString(minYk);
while(strY1.length() < 25)
{
strY1 = "0" + strY1;
}
strD1 = Double.toString(dk);
while(strD1.length() < 25)
{
strD1 = "0" + strD1;
}
try

```

```

        {
            outputStream[K].writeBytes(strD1+" " +
strY1+" " + strFx+"\n");
        }
        catch(Exception ex)
        {
            ex.printStackTrace();
        }
        Yk = dk;
        dkCount++;
        dk = dkCount*Math.pow(10, -
intPreci);//increment;
    }
    try
    {
        System.out.println(K+" file-
"+randReader[K].length());
    }
    catch(Exception ex)
    {
        ex.printStackTrace();
    }
}

void backWardCalculation()
{
    try
    {
        File tmpFile = new File("./resultNew.txt");
        RandomAccessFile rand = new
RandomAccessFile(tmpFile, "rw");
        rand.seek(rand.length());
        double fxx[] = new double[intStage];
        double fyy[] = new double[intStage];
        double fdd[] = new double[intStage];
        int kk = intStage-1;
        fxx[kk] = readFkDk1(kk, Dk);
        fyy[kk] = readYk(kk, Dk);
        fdd[kk]= Dk;
        rand.writeBytes("\n Date: " + new Date() + "\nNumber of
stage = "+ intStage +" , Dk = " + Dk + " , Precision = "+ intPreci
);
        //System.out.println( "Yk- Dk =" + Dk );
    }
}

```

```

        for( int i =kk-1; i >= 0 ; i--)
        {
            fdd[i]= fdd[i+1] - fyy[i+1];
            fxx[i] = readFkDk1(i, fdd[i+1]-fyy[i+1]);
            fyy[i] = readYk(i, fdd[i+1]-fyy[i+1]);
            //System.out.println(fdd[i+1] + ", fdd=" + fdd[i] );
        }

        for( int i =0; i <= kk ; i++)
        {
            rand.writeBytes("\nY" + (i+1) + " = " + fyy[i] + ",
D" + (i+1) + " = " + fdd[i]);
        }
        rand.writeBytes("\nfx" + (kk+1) + " = " + fxx[kk]+
"\n\n\n");
    }
    catch(Exception ex)
    {
        ex.printStackTrace();
    }
}
}

```

```

        double readYk(int K, double Dk)
        {
            double tmpDk = Dk*Math.pow(10,
intPreci);

            long lDk = (long)Dk;
            long n1 = (long)tmpDk;
            //Math.round(Dk*Math.pow(10, intPreci));
            double ret=0, data1 =0, data2 =0;
            String str= "";
            n1 = n1*78;
            try
            {
                if(n1 < 0 || n1 >
randReader[K].length()) return 0;

                randReader[K].seek(n1);
                str= randReader[K].readLine();
                data1 =
Double.parseDouble(str.substring(26,51));
                str= randReader[K].readLine();
                if(str != null && str.length() >= 75)
                {

```

```

        data2 =
Double.parseDouble(str.substring(26, 51));
    }
    else
        data2 = data1;
        ret = data1+ (data2-data1)*(Dk*100
-1Dk*100)/100;
        //System.out.println( "Dk passed =" +
Dk + ", line = "+n1/78 +", Fx=" + ret );
    }
    catch(Exception ex)
    {
        System.out.println( K+",Dk passed
=" + Dk + ", line = "+n1/78 +", str=" + str );
        ex.printStackTrace();
    }
    }
    return ret;
}
}

```

```

void readStoredFk(int k)
{
    k--;
    try
    {
        File file = new File("./Stage"+(k+1)+".txt");
        RandomAccessFile randTmp = new
RandomAccessFile(file, "r");
        //
        randReader[k].seek(0);
        System.out.println( "filelength read= "
+randReader[k].length() );
        String str = null;
        int line = 0;
        System.out.println( "filelength= "
+randTmp.length() + ", array=" + storedFk.length);
        while((str = randTmp.readLine()) != null
&& line < storedFk.length)
        {
            storedFk[line] =
Double.parseDouble(str.substring(51));
            line++;
        }
    }
}

```

```

        }
        System.out.println( "k= " +k + ", line=" +
line);
    }
    catch(Exception ex)
    {
        ex.printStackTrace();
    }
}

```

```

//*****
*****

```

```

    static class Readline
    {

        public static void main(String args[])
        {
            try{
                // 1. Create an InputStreamReader using the
standard input stream
                InputStreamReader isr = new InputStreamReader(
System.in );

                // 2. Create a BufferedReader using the
InputStreamReader created.
                BufferedReader stdin = new BufferedReader( isr );

                // 3. Don't forget to prompt the user
                System.out.print( "Type some data for the program:
" );

                // 4. Use the BufferedReader to read a line of text
from the user.
                String input = stdin.readLine();

                // 5. Now, you can do anything with the input string
that you need to.
                // Like, output it to the user.
                System.out.println( "input = " + input );
            }catch(Exception ex){ex.printStackTrace();}
        }
    }

```

```

public static String readLine()
{
    String input = "0";
    try{
        // 1. Create an InputStreamReader using the
standard input stream
        InputStreamReader isr = new InputStreamReader(
System.in );

        // 2. Create a BufferedReader using the
InputStreamReader created.
        BufferedReader stdin = new BufferedReader( isr );

        // 4. Use the BufferedReader to read a line of text
from the user.
        input = stdin.readLine();

    }catch(Exception
ex){ex.printStackTrace();System.exit(0);}
    finally
    {
    }
    return input;
}
}
}
}

```


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REFERENCES

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