

Fear of Appreciation in East and Southeast Asia: The Role of the Chinese Renminbi

Victor Pontines and Reza Y. Siregar

The Flinders University, South Australia, The South East Asian Central Banks (SEACEN) Research and Training Centre, Kuala Lumpur, Malaysia

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Fear of Appreciation in East and Southeast Asia: The Role of the Chinese Renminbi*

Victor Pontines^(a) and Reza Y. Siregar^(b)

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Abstract:

Our study brings into light the evidence of a fundamental role of the Chinese renminbi in shaping the exchange rate behavior of other major Asian currencies. We obtain results suggesting that there is an additional dimension to the 'fear of appreciation' or 'fear of floatingin-reverse' behavior, initially coined by Levy-yeyati and Sturzenegger (2007)with regards to the cases of these major Asian currencies. In particular, we find that there is a greater degree of aversion to appreciation of these same Asian currencies, specifically, the Philippine peso and the Thailand baht against the renminbi than against the US dollar. This heightened fear of appreciation against the Chinese currency confirms that trade competition indeed matters in this part of the world and that the fear continues to play a central role in the exchange rate management behavior of major Asian currencies. This is understandable as the rising role of China as a major trading hub in the region as well as globally, must then render the renminbi to exert a significant influence to the rest of the currencies in the region.

Key Words:Asymmetrical Exchange Rate Regime; East and Southeast Asia; RegimeSwitching Models.

JEL Classifications: E58; F31; F41

^(a)/ Victor Pontines (corresponding author) is with the Flinders Business School, Australia. E-mail: victor.pontines@flinders.edu.au

^(b)/ Reza Y. Siregar is with the South East Asian Central Banks (SEACEN) Research and Training Centre in Kuala Lumpur, Malaysia. E-mail: <u>reza@seacen.org</u> and <u>rezasiregar@yahoo.com</u>.

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1. Introduction

After three decades of spectacular growth, China surpassed Japan in the second quarter of 2010 to become the world's second-largest economy behind the United States. According to official figures from both Japan and China, the Japanese economy was valued at US\$ 1.28 trillion in the second quarter of 2010, slightly below China's US\$1.33 trillion economy (Barboza (2010)).Notwithstanding that growth has begun to moderate a bit, China's economy is still officially forecasted to expand by about 10 percent ---continuing a remarkable three-decade streak of double-digit growth.

In their study, Branson and Healy (2005) demonstrate that the structure of China's exports, both in terms of market and commodity circulation, is parallel to that of a number of Asian economies. They argue, therefore, that Asian economies should have the incentive to keep a close watch on renminbi fluctuations in managing their currencies in order to avoid any potential loss in competitiveness. Similarly, Shu, Chow and Chan (2007) claim that the renminbi has indeed started to figure more prominently in the currency baskets of the Asian economies, primarily due to competitive relationships between Mainland China and these economies in the global export markets. One such cursory evidence is the significant rise in the bivariate correlation between renminbi and individual Asian currencies (Table 1).

The influences of the renminbi on the Asian currencies have long been debated amongst policy makers and academics, but only a few recent works have explicitly demonstrated those impacts. One study carried out by Colavecchio and Funke (2007) investigate the spill-over effects from the renminbi to Asian currencies in the onshore and offshore forward markets. Similarly, Ho et al. (2005) noted that the renminbi non-deliverable forward rates had some bearing on the spot rates of the Asian currencies during the period of rising speculation on renminbi appreciation in 2003- 2004. Applying the Frankel and Wei (1994) framework, Shu, Chow and Chan (2007) claim that since July 2005, the fluctuation of renminbi that are

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independent of those of the US dollar tend to lead to movements in other Asian currencies in the same direction.

Against this backdrop, our study examines the currencies of four East and Southeast Asian economies, namely the Indonesian rupiah, the Korean won, the Philippines peso and the Thailand baht. Immediately after the 1997 East Asian crisis, the monetary authorities of these economies announced their commitments to abandon their rigid soft-US dollar policies and let their currencies to float more flexibly. Furthermore, to regain confidence on their domestic monetary policies, the central banks of these economies have officially adopted inflation targeting policy regimes. Intrinsically, the exchange rate is expected to move more flexibly under an inflation targeting policy.

The first primary contention to be dealt with in this paper is that while the adoption of the inflation targeting policy in these four Asian economies may lead to more flexible exchange rate movements, under various reasons, it is however possible that the degree of flexibility is significantly higher on one side of the market.¹ In other words, the monetary authorities in these four Asian economies can asymmetrically manage their exchange rates wherein they can allow for some currency depreciation while substantially limiting the extent of currency appreciation. For lack of better alternatives, this exchange rate intervention behavior has been coined by Levy-yeyati and Sturzenegger (2007) as 'fear of floating in reverse' or 'fear of appreciation.²

Past studies have looked into this issue of asymmetrical exchange rate behavior across a number of Asian currencies. Ramachandran and Srinivasan (2007) for instance employ a simple dummy variable to account for depreciations and appreciations. On the other hand, Srinivasan, Mahambare and Ramachandran (2008) explore a cubic loss function to account for possible asymmetric behavior. The study of Stigler, Patnaik and Shah (2009) apply time series

¹See, for instance, Stigler et al (2009) for this various set of plausible reasons.

²Such asymmetry in exchange rate behavior have been pointed out much earlier by Almekinders and Eijffinger (1996) when they find that during the post-Louvre period, US and German central banks have tried to counteract appreciations of their currency more strongly than depreciations.

techniques in the context of structural change of the exchange rate regime. Furthermore, Pontines and Rajan (2010) similarly employs the cubic loss function of Srinivasan et al (2008) and tests the model across a broader set of emerging Asian economies that are known to operate a variety of managed floats viz., India, Indonesia, Korea, Philippines, Singapore, and Thailand.

However, these previous studies have all emphasized in the context of an asymmetrical management of their domestic currencies with respect to the US dollar. Hardly any of these earlier studies have considered the possible role of a 'third currency' that may exists in accounting for asymmetrical exchange rate intervention behavior. In particular, none has considered the plausible role of the Chinese renminbi in shaping exchange rate management of these Asian currencies, especially post 2005 exchange rate policy reform in China.³This then leads us to the second and related key policy concern to be assessed in this paper. In view of the alleged bias on one side of the market -- some domestic currency depreciation is allowed while substantially restraining the extent of domestic currency appreciation explicitly against the US dollar -- does this aversion to appreciation also extends and applies to their domestic currencies with respect to the Chinese renminbi? Given the nuances and the intricacies that it entails of the emerging dominant role of the Chinese economy in Asia, in particular, and in the world, in general, this is a logical and interesting research question that is worth pursuing.

In our empirical work, the well-known class of regime switching model known as the Markov regime-switching will be employed. Under this framework we treat asymmetry in exchange rate behavior as being equivalent to evidence of persistence. Testing for the

³In July 2005, the People's Bank of China (PBC) announced its policy intention to set the value of renminbi with reference to a basket of foreign currencies, and to allow the currency to move more flexibly. The outcome was arguably a more measured and deliberate approach akin to a managed crawl whereby the currency was allowed to appreciate by around 20 percent against the dollar from 2005 to late 2008.

presence of persistence is then achieved via the comparison of transition probabilities across two different regimes. To the best of our knowledge, our paper is the only one thus far that employs such technique in the context of flexible exchange rate policies under inflation targeting regimes in East Asia.⁴

The rest of the paper is set out as follows. The empirical section is presented next, and it includes a discussion of the model specifications and econometric testing employed. The key empirical findings are analysed and highlighted in Section 3. In the following section, we present and assess bilateral trade patterns of the four Asian economies. The main objective of this section is to evaluate the test results presented in Section 3 and to further substantiate the analyses. A brief concluding section ends the paper.

2. Empirical Approach

We now turn to the empirical approach adopted in this paper. One widely used approach in estimating the weight or influence of a currency on another is the following Frankel and Wei (1994) framework formally expressed as:

$$\Delta e_{j} = \alpha_{0} + \sum_{i=1}^{k} \alpha_{i} \Delta e_{i}$$
(1)

In equation (1), (e_j) and (e_i) s are, respectively, exchange rate of the currency under study and those currencies that might influence it. A significant (α_i) would suggest that currency (i) has an impact on (e_j) , and (α_i) is interpreted as the weight of currency (i) in the currency basket. In their work, Shu, Chow and Chan (2007) applies this approach and claims that there is a clear

⁴The paper that comes very close to our work is by Leon and Najarian (2005) which adopts similar set of comprehensive tests but undertakes it in the context of deviations from PPP. In addition, Stigler et al (2009) also adopts threshold modeling procedures but uses the self-exciting threshold autoregressive (SETAR) process and the momentum threshold autoregressive model (MTAR) in approaching the issue of asymmetric exchange rate behavior.

rise in the importance of the renminbi in influencing other Asian currencies during the post July 2005.⁵

Whilst the Frankel and Wei framework would be useful in revealing the weight or importance of the renminbi, the framework is not explicitly designed to convey any information on the degree of aversion of the exchange rate management, either to currency appreciation or depreciation. For instance, the estimates of the (*e_i*)s in equation (1) are time-invariant coefficients and as such cannot account for asymmetrical behavior. Thus, this conventional approach will not address our primary question on the possibility of fear of appreciation in exchange rate behavior of our four Asian currencies. The empirical approach adopted in this paper then takes the view that in order to test for the existence of fear or aversion to appreciations, one has to resort to a time-varying approach and one such suggested strategy is via a simple comparison of the estimated probabilities of remaining in the state or regime of depreciation as against that of the state or regime of appreciation.

In this regard, in order to circumvent the limitation of the Frankel and Wei framework, our study instead employs the Markov-regime switching approach. Earlier introduced to the time-series econometric literature by Hamilton (1988, 1989), the general idea behind the Markov-regime switching is that the parameters of the autoregressive process may not be time-invariant, as assumed by linear models. More precisely, the parameters may be time-invariant as long as a particular regime prevails, but change once the regime changes.

Denoting $\Delta \ln exr$ as the first difference of the log of the nominal exchange rate (defined as either the domestic currency per US dollar or the domestic currency per Chinese renminbi), we estimate a univariate regime switching model which can be written in the form:

$$\Delta \ln exr_t = \alpha(s_t) + \sum_{k=1}^p \alpha_k(s_t) \Delta \ln exr_{t-k} + \sum (s_t) \varepsilon_{\Delta \ln exr,t} \qquad \varepsilon_{\Delta \ln exr,t} \sim N(0,1)$$
(2)

⁵The other Asian currencies in their study include Hong Kong dollar, Indian rupee, Indonesian rupiah, Korean won, Malaysian ringgit, Philippines peso, Singapore dollar, New Taiwan dollar and Thai baht.

Where s_t is the unobservable regime, and ε_t is the innovation process with a regime-dependent variance-covariance matrix $\Sigma(s_t)$. Maximum likelihood estimation of (2) yields the 'smoothed' probabilities, representing the ex-post inference about the system being in regime *i* at year *t*. It is assumed that s_t follows an irreducible ergodic two-regime Markov process with the transition matrix **P**:

$$\mathbf{P} = \begin{bmatrix} p_{11} & p_{12} \\ p_{21} & p_{22} \end{bmatrix}$$
(3)

Constant transition probabilities p_{ij} are assumed:

$$p_{ij} = \Pr(s_t = j | s_{t-1} = i), \qquad \sum_{j=1}^{2} p_{ij} = 1, \qquad \forall i, j \in \{1, 2\}$$
 (4)

The Markov-regime switching model employed here is well-suited to testing whether East Asian currencies exhibit aversion to appreciations because it yields these transition probabilities. Focusing on p_{11} (the probability of remaining in the lower regime, i.e., regime 1, given that the process was in the lower regime at time (t-1)) and p_{22} (the probability of remaining in the higher regime, i.e., regime 2, given that the process was in the higher regime at time (t-1)), these transition probabilities can be utilized to test for the degree of asymmetrical exchange rate policy. If the probability of remaining in the upper regime is greater than that of remaining in the lower regime, it suggests an aversion of monetary authorities to currency appreciations. Or, more formally:

$$\Pr(s_t = 2 | s_{t-1} = 2) > \Pr(s_t = 1 | s_{t-1} = 1)$$
(5)

3. Empirical Results

Our estimation is based on monthly data for the period 2000:m1 -2006:m12 for four inflation-targeting emerging Asian economies, viz., Indonesia, Korea, Philippines and Thailand.

The sample period was chosen in view of the fact that the bulk of the reserves accumulation for these economies occurred during this period (i.e. post Asian crisis of 1997-98) as well as to steer clear of the substantial exchange rate instabilities experienced by these economies during the global financial crisis of 2007-2009. All nominal exchange rate (domestic currency per US\$ and the domestic currency per Chinese renminbi) data are obtained from the Pacific Exchange Rate Service (http://fx.sauder.ubc.ca).

The starting point in the empirical analysis is to formally test the null hypothesis of no regime switching in the mean, the variance-covariance matrix, and the autoregressive coefficients against the alternative of regime switching. The null hypothesis is equivalent to a constant mean and autoregressive parameters, homoscedastic linear AR model. The Davies (1987) approach is employed for testing this hypothesis. As shown in the last row of Table 2, the null of no regime switching is rejected for all four East Asian currencies. On this basis, the estimated results of the nonlinear regime switching AR model wherein the first difference of the log of the nominal exchange rate defined as the domestic currency per US dollar are reported in Table 2. The results yields plausible estimates of the coefficients in both regimes 1 and 2 and for the most part, the coefficients are statistically significant at conventional levels.

The estimated results of the nonlinear regime switching AR model of the first difference of the log of the nominal exchange rate defined this time as the domestic currency per Chinese renminbi are reported in Table 3. Again, the regime switching AR is strongly preferred to the linear AR specification as presented in the last row of Table 3. Furthermore, the estimated results in Table 3 offer similar observations as compared to the estimated results of the nonlinear regime switching AR model using the domestic currency per US dollar as the definition of the nominal exchange rate (table 1) – plausible estimated coefficients and for the most part statistically significant at conventional levels.

More importantly, the analysis of the standardized residuals emanating from the estimated nonlinear regime switching AR models as previously presented in Tables 2 and 3

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provide evidence of no residual serial correlation in the residual series (Ljung-Box Q test) as well as there is no evidence of conditional heteroscedasticity (ARCH-LM test) in any of the residual series (see Table 4).⁶ This provides further evidence that the regime-switching dynamics adopted in this paper captures fairly well the nonlinearity and conditional heteroscedasticity in the respective nominal exchange rate series.

We now turn to the core results of this paper and examine the estimated transition probabilities as reported in Table 5. It is remarkably clear that for all of the four East Asian currencies considered, the probability of remaining in the upper regime (p_{22}) is greater than that of remaining in the lower regime (p_{11}), suggesting an aversion to appreciation. This finding is true not just in the case of the nominal exchange rate being defined as the domestic currency per US dollar, but also applies as such in the case of the nominal rate being the domestic currency per Chinese renminbi.

More to the point, it is striking from the results that while the extent of this observed aversion to appreciation varies by degree amongst the four East Asian currencies per nominal exchange rate definitions, the extent of this observed bias on one side of the market also differs in terms of the two comparator base currencies. For instance, the degree of the observed aversion to appreciation in the case of the Indonesian rupiah and the Korean won are both indistinguishable irrespective of whether one looks at the US dollar as the base currency or the Chinese renminbi. In contrast, the cases of the Philippine peso and the Thailand baht exemplify a much greater aversion to appreciations when the Chinese renminbi is used as the comparator base currency.

4. Recent Trends in Bilateral Trade Patterns and the Dominant Role of Mainland China

⁶A slight exception to this observation is the case of the nonlinear regime switching AR model estimated in the case of the Philippine peso using the domestic currency per Chinese renminbi as the definition of the nominal exchange rate wherein the null of no residual serial correlation is rejected at the 10% level of significance.

Our empirical testing brings to light the evidence of a fundamental role of the Chinese renminbi in shaping the exchange rate management of other major Asian economies. The greater degree of aversion to appreciation of the Philippine peso and the Thailand baht against the renminbi than against the US dollar appears to suggest that there is an even more heightened fear of appreciation against the renminbi that is exemplified by the policy makers in these two countries during the period of observation. In view of this, we are led to conclude that the fear of appreciation against the Chinese renminbi, rather than the US dollar, contributed more to the asymmetrical exchange rate behavior on the part of monetary authorities in the Philippines and Thailand. Likewise, given relatively identical degrees of aversion to appreciation of the Indonesian rupiah and the Korean won against the US dollar and the renminbi during the observed period, we also cannot rule out right-away the seeming underlying and extensive influence of the renminbi to an asymmetrical exchange rate behavior on the part of monetary authorities authorities in Indonesia and Korea.

How can we possibly explain such an influential role that the renminbi has on the Asian currencies? One of the frequently advanced factors behind the increasing influence of the Chinese renminbi is the rise of China as a major trading partner of countries around the globe, including that of the four Asian economies in our study. A simple calculation of trade ratios for all four Asian economies in terms of their bilateral trade patterns with China, US and Japan consistently reveal the fast and emerging dominant role of China in the region (Figures 1-4).⁷The steep increase in total trade with China started to be noticeable in the immediate aftermath of the 1997 East Asian crisis in all four Asian economies.

⁷The calculation of the trade ratio proceeds as follows. We take the ratio of total trade (exports plus imports) with China and total trade with that of the US and Japan, respectively. So for the case of Indonesia, for instance, the ratio of total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with the US is calculated. We also do this for the ratio of total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports of Indonesia with China against the total bilateral exports and imports o

In addition, we also find that these four Asian economies experience distinct overall trade balance positions against China (Figures 1b-4b). On one hand, Indonesia and Korea have persistently been successful in maintaining healthy trade surpluses in their overall trade with China since the mid-1990s. In stark contrast, Thailand, throughout the 1990s and until the period of observation has persistently suffered a trade deficit with China, and its bilateral trade deficit seems to be widening beginning in 2003 and onwards. Meanwhile, with the recent steep rise in its exports of goods to China, the Philippines only started to record a trade surplus with China in 2003 and from 2005 onwards, after over a decade of sustained bilateral trade deficits with China. In light of these presented stylized facts, we can reasonably surmised that out of the four Asian economies examined, a country that either suffers a persistently bilateral trade deficit with China (Thailand) or a country that experiences less sustained bilateral trade surpluses with China (Philippines) is more likely to exhibit or display a greater aversion to appreciation against that of the Chinese renminbi. This assertion is simply in conjunction with the argument that trade competition matters in this part of the world and definitely continues to play a central role in the exchange rate management behavior of the four major Asian economies examined in this paper.

5. Conclusion

Despite the full adoption of inflation targeting policy as their monetary policy framework, we assert that there are still significant concerns in the four Asian economies examined in this paper, namely, Indonesia, Korea, Philippines and Thailand, of allowing market forces to completely determine their exchange rates. To be more specific, these same Asian economies continue to be characterized by a greater aversion toward appreciation, or more commonly known as 'fear of appreciation'. This finding confirms the recent evidence obtained by the related work of Pontines and Rajan (2010) for these same set of Asian economies. Our study, however, goes one step further by bringing into light the vital role of the Chinese renminbi in

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shaping the exchange rate behavior of these same four major Asian currencies. Whereas, some of the previous studies continue to underscore the return of the soft US dollar pegging in Asia post-crisis,⁸ our findings indicate that that there is an even more greater dimension to this story – the heightened fear of appreciation against the renminbi than against the US dollar, and much more so in the cases of the Philippine peso and the Thailand baht. This is understandable as the rising role of China as a major trading hub in the region and globally has clearly contributed to the significant influence of the renminbi to the rest of the currencies in the region.

⁸ See, for instance, McKinnon and Schnabl (2004).

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Table 1:

	Before reform	After reform	
	(Jan 5, 1999 – Jul 20, 2005)	(Jul 21, 2005 – Apr 2, 2007)	
Indonesian rupiah	-0.436	0.625	
Korean won	-0.040	0.764	
Malaysian ringgit	-0.186	0.927	
Philippine peso	-0.730	0.890	
Singapore dollar	-0.083	0.893	
Thai baht	-0.391	0.955	

Correlation between the renminbi and selected East and Southeast Asian currencies

Source: Shu, Chow and Chan (2007).

	Indonesian	Korean	Philippine	Theiland
	runiah	WOn		i nallano Raht
	ταριατί	WOIT	1 030	Ban
		Regime 1		
intercept	-0.000 (0.015)	-0.002 (0.002)	-0.003 (0.004)	-0.002 (0.001)***
(∆ln <i>exr</i>) _{t-1}	4.339 (0.521)***	-0.333 (0.169)**	-0.408 (0.193)**	0.167 (0.064)***
(∆ln <i>exr</i>) _{t-2}	0.569 (0.198)***		0.733 (0.239)***	-0.529 (0.068)***
(∆ln <i>exr</i>) _{t-3}			-0.520 (0.223)**	0.495 (0.068)***
(∆ln <i>exr</i>) _{t-4}			-0.277 (0.142)**	
$(\Delta \ln exr)_{t=5}$			-0.426 (0.240)*	
$(\Delta \ln exr)_{t=6}$			0.078 (0.213)	
(),			· · · · ·	
		Regime 2		
intercept	0.002 (0.003)	-0.002 (0.002)	-0.000 (0.003)	-0.002 (0.001)***
$(\Delta \ln exr)_{t-1}$	0.035 (0.092)	0.631 (0.209)***	0.562 (0.139)***	0.442 (0.147)***
$(\Delta \ln exr)_{t-2}$	0.038 (0.145)		-0.233 (0.142)	0.097 (0.156)
$(\Delta \ln exr)_{t-3}$			0.360 (0.095)***	-0.047 (0.154)
$(\Delta \ln exr)_{t-4}$			0.117 (0.265)	
$(\Delta \ln exr)_{t=5}$			0.008 (0.121)	
$(\Delta \ln exr)_{t=6}$			0.090 (0.114)	
, ,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,			. ,	
Linearity test	[0.00]	[0.00]	[0.00]	[0.00]
Numbers in parentheses and square brackets are standard errors and <i>p</i> -values,				

 Table 2

 Estimation results for regime switching models of East Asian currencies per US\$

respectively. Significancelevels: *10%, **5%, ***1%.

	Indonesian	Korean	Philippine	Thailand
	rupiah	won	peso	Baht
		Regime 1		
intercept	-0.000 (0.007)	-0.000 (0.000)	-0.009 (0.003)	0.000 (0.000)
$(\Delta \ln exr)_{t-1}$	4.346 (0.535)***	-0.153 (0.131)	-0.514 (0.184)**	0.094 (0.038)**
$(\Delta \ln exr)_{t-2}$	0.568 (0.198)***	-0.244 (0.140)*		-0.142 (0.044)***
		Regime 2		
intercept	0.001 (0.002)	-0.000 (0.000)	0.002 (0.001)**	0.000 (0.00)
$(\Delta \ln exr)_{t-1}$	0.055 (0.191)	0.601 (0.174)***	0.467 (0.112)***	0.513 (0.142)***
(∆ln <i>exr</i>) _{t-2}	-0.002 (0.457)	0.161 (0.171)		0.045 (0.129)
Linearity test	[0.00]	[0.01]	[0.00]	[0.00]
Numbers in parentheses and square brackets are standard errors and <i>p</i> -values, respectively. Significancelevels: *10%, **5%, ***1%.				

Table 3Estimation results for regime switching models of East Asian currencies per remninbi

	Indonesian rupiah	Korean won	Philippine peso	Thailand baht		
	Per USD					
Ljung-Box <i>Q</i> ARCH-LM(8)	[0.63] [0.69]	[0.51] [0.21]	[0.75] [0.73]	[0.35] [0.64]		
per remninbi						
Ljung-Box Q ARCH-LM(8)	[0.88] [0.52]	[0.55] [0.22]	[0.06] [0.11]	[0.25] [0.95]		

Table 4Diagnostic test results

Numbers in square brackets are *p*-values.

Table 5Transition Probabilities

	per USD		per remninbi	
	<i>p</i> ₁₁	<i>p</i> ₂₂	p_{11}	<i>p</i> ₂₂
Indonesian rupiah Korean won	0.0245	0.9268 0.5765	0.0194 0.3730 0.2570	0.9322 0.5763
Thailand baht	0.6014	0.8524	0.2370	0.9308





Note: a). Indo-Jap: total exports and imports between Indonesia and Japan.

b). Indo-USA: total exports and imports between Indonesia and USA.

c). Indo-Chn: total exports and imports between Indonesia and the People Republic of China.

Source: authors' computations and IMF Direction of Trade Statistics

Figure 1b: Export and Import of Indonesia with China

(in million US dollars)



Source: IMF Direction of Trade Statistics.

Figure 2: Trade Ratio for Korea



Note: a). Kor-Japan: total exports and imports between Korea and Japan.

b). Kor-USA: total exports and imports between Korea and USA.

c). Kor-Chn: total exports and imports between Korea and the People Republic of China.

Source: authors' computations and IMF Direction of Trade Statistics

Figure 2b: Export and Import of Korea with China

(in million US dollars)



Source: IMF Direction of Trade Statistics.

Figure 3: Trade Ratio for the Philippines



Note: a). Phi-Japan: total exports and imports between Philippines and Japan.

b). Phi-USA: total exports and imports between Philippines and USA.

c). Phi-Chn: total exports and imports between Philippines and the People Republic of China.

Source: authors' computations and IMF Direction of Trade Statistics

Figure 3b: Export and Import of Philippines with China



(in million US dollars)

Source: IMF Direction of Trade Statistics.



Note: a). Thai-Japan: total exports and imports between Thailand and Japan.

b). Thai-USA: total exports and imports between Thailand and USA.

c). Thai-Chn: total exports and imports between Thailand and the People Republic of China.

Source: authors' computations and IMF Direction of Trade Statistics

Figure 4b: Export and Import of Thailand with China



(in million US dollars)

Source: IMF Direction of Trade Statistics.