

DRAFT VERSION SEPTEMBER 7, 2010
Preprint typeset using L^AT_EX style emulateapj v. 11/10/09

THE ATACAMA COSMOLOGY TELESCOPE: A MEASUREMENT OF THE COSMIC MICROWAVE BACKGROUND POWER SPECTRUM AT 148 AND 218 GHz FROM THE 2008 SOUTHERN SURVEY

SUDEEP DAS^{1,2,3}, TOBIAS A. MARRIAGE^{3,4}, PETER A. R. ADE⁵, PAULA AGUIRRE⁶, MANDANA AMIRI⁷, JOHN W. APPEL², L. FELIPE BARRIENTOS⁶, ELIA S. BATTISTELLI^{8,7}, J. RICHARD BOND⁹, BEN BROWN¹⁰, BRYCE BURGER⁷, JAY CHERVENAK¹¹, MARK J. DEVLIN¹², SIMON R. DICKER¹², W. BERTRAND DORIESE¹³, JOANNA DUNKLEY^{14,2,3}, ROLANDO DÜNNER⁶, THOMAS ESSINGER-HILEMAN², RYAN P. FISHER², JOSEPH W. FOWLER^{13,2}, AMIR HAJIAN^{9,3,2}, MARK HALPERN⁷, MATTHEW HASSELFIELD⁷, CARLOS HERNÁNDEZ-MONTEAGUDO¹⁵, GENE C. HILTON¹³, MATT HILTON^{16,17}, ADAM D. HINCKS², RENÉE HLOZEK¹⁴, KEVIN M. HUFFENBERGER¹⁸, DAVID H. HUGHES¹⁹, JOHN P. HUGHES²⁰, LEOPOLDO INFANTE⁶, KENT D. IRWIN¹³, JEAN BAPTISTE JUIN⁶, MADHURI KAUL¹², JEFF KLEIN¹², ARTHUR KOŠOWSKY¹⁰, JUDY M LAU^{21,22,2}, MICHELE LIMON^{23,12,2}, YEN-TING LIN^{24,3,6}, ROBERT H. LUPTON³, DANICA MARSDEN¹², KRISTA MARTOCCHI^{25,2}, PHIL MAUSKOPF⁵, FELIPE MENANTEAU²⁰, KAVILAN MOODLEY^{16,17}, HARVEY MOSELEY¹¹, CALVIN B. NETTERFIELD²⁶, MICHAEL D. NIEMACK^{13,2}, MICHAEL R. NOLTA⁹, LYMAN A. PAGE², LUCAS PARKER², BRUCE PARTRIDGE²⁷, BETH REID^{28,2}, NEELIMA SEHGAL²¹, BLAKE D. SHERWIN², JON SIEVERS⁹, DAVID N. SPERGEL³, SUZANNE T. STAGGS², DANIEL S. SWETZ^{12,13}, ERIC R. SWITZER^{25,2}, ROBERT THORNTON^{12,29}, HY TRAC^{30,31}, CAROLE TUCKER⁵, RYAN WARNE¹⁶, ED WOLLACK¹¹, YUE ZHAO²

Draft version September 7, 2010

ABSTRACT

We present measurements of the cosmic microwave background (CMB) power spectrum made by the Atacama Cosmology Telescope at 148 GHz and 218 GHz, as well as the cross-frequency spectrum between the two channels. Our results clearly show the second through the seventh acoustic peaks in the CMB power spectrum. The measurements of these higher-order peaks provide an additional test of the Λ CDM cosmological model. At $\ell > 3000$, we detect power in excess of the primary anisotropy spectrum of the CMB. At lower multipoles $500 < \ell < 3000$, we find evidence for gravitational lensing of the CMB in the power spectrum at the 2.8σ level. We also detect a low level of Galactic dust in our maps, which demonstrates that we can recover known faint, diffuse signals.

Subject headings: cosmology: cosmic microwave background, cosmology: observations

¹ Berkeley Center for Cosmological Physics, LBL and Department of Physics, University of California, Berkeley, CA, USA 94720

² Joseph Henry Laboratories of Physics, Jadwin Hall, Princeton University, Princeton, NJ, USA 08544

³ Department of Astrophysical Sciences, Peyton Hall, Princeton University, Princeton, NJ USA 08544

⁴ Current address: Dept. of Physics and Astronomy, The Johns Hopkins University, 3400 N. Charles St., Baltimore, MD 21218-2686

⁵ School of Physics and Astronomy, Cardiff University, The Parade, Cardiff, Wales, UK CF24 3AA

⁶ Departamento de Astronomía y Astrofísica, Facultad de Física, Pontificia Universidad Católica, Casilla 306, Santiago 22, Chile

⁷ Department of Physics and Astronomy, University of British Columbia, Vancouver, BC, Canada V6T 1Z4

⁸ Department of Physics, University of Rome “La Sapienza”, Piazzale Aldo Moro 5, I-00185 Rome, Italy

⁹ Canadian Institute for Theoretical Astrophysics, University of Toronto, Toronto, ON, Canada M5S 3H8

¹⁰ Department of Physics and Astronomy, University of Pittsburgh, Pittsburgh, PA, USA 15260

¹¹ Code 553/665, NASA/Goddard Space Flight Center, Greenbelt, MD, USA 20771

¹² Department of Physics and Astronomy, University of Pennsylvania, 209 South 33rd Street, Philadelphia, PA, USA 19104

¹³ NIST Quantum Devices Group, 325 Broadway Mailcode 817.03, Boulder, CO, USA 80305

¹⁴ Department of Astrophysics, Oxford University, Oxford, UK OX1 3RH

¹⁵ Max Planck Institut für Astrophysik, Postfach 1317, D-85741 Garching bei München, Germany

¹⁶ Astrophysics and Cosmology Research Unit, School of Mathematical Sciences, University of KwaZulu-Natal, Durban, 4041, South Africa

¹⁷ Centre for High Performance Computing, CSIR Campus, 15 Lower Hope St., Rosebank, Cape Town, South Africa

¹⁸ Department of Physics, University of Miami, Coral Gables, FL, USA 33124

¹⁹ Instituto Nacional de Astrofísica, Óptica y Electrónica (INAOE), Tonantzintla, Puebla, Mexico

²⁰ Department of Physics and Astronomy, Rutgers, The State University of New Jersey, Piscataway, NJ USA 08854-8019

²¹ Kavli Institute for Particle Astrophysics and Cosmology, Stanford University, Stanford, CA, USA 94305-4085

²² Department of Physics, Stanford University, Stanford, CA, USA 94305-4085

²³ Columbia Astrophysics Laboratory, 550 W. 120th St. Mail Code 5247, New York, NY USA 10027

²⁴ Institute for the Physics and Mathematics of the Universe, The University of Tokyo, Kashiwa, Chiba 277-8568, Japan

²⁵ Kavli Institute for Cosmological Physics, 5620 South Ellis Ave., Chicago, IL, USA 60637

²⁶ Department of Physics, University of Toronto, 60 St. George Street, Toronto, ON, Canada M5S 1A7

²⁷ Department of Physics and Astronomy, Haverford College, Haverford, PA, USA 19041

²⁸ Institut de Ciències del Cosmos (ICC), University of Barcelona, Barcelona 08028, Spain

²⁹ Department of Physics, West Chester University of Pennsylvania, West Chester, PA, USA 19383

³⁰ Department of Physics, Carnegie Mellon University, Pittsburgh, PA 15213

³¹ Harvard-Smithsonian Center for Astrophysics, Harvard University, Cambridge, MA, USA 02138

1. INTRODUCTION

Accurate measurements of the arcminute scale temperature anisotropies in the mm-wave sky are uncovering a complex, yet revealing picture (Lueker et al. 2010; Fowler et al. 2010, hereafter F10). On intermediate scales ($500 \lesssim \ell \lesssim 3000$), the primordial acoustic features imprinted on the cosmic microwave background (CMB) at last scattering ($z \simeq 1100$) dominate, with subtle distortions expected from gravitational lensing by intervening large-scale structure. This intermediate range of multipoles is often called the *damping tail* of the CMB, as the acoustic oscillations are exponentially damped due to photon diffusion (Silk 1968; Bond & Efstathiou 1984). On the smallest scales ($\ell \gtrsim 3000$), emission from radio and dust-enshrouded star-forming galaxies, together with the Sunyaev-Zel'dovich (SZ) effect (Zeldovich & Sunyaev 1969) dominates over primary CMB fluctuations.

The damping tail measurements are an additional test of the predictions of the Λ CDM cosmological model, a model that is an excellent fit to current CMB data (e.g., Larson et al. 2010; Reichardt et al. 2009b; Brown et al. 2009), large-scale structure measurements (e.g., Reid et al. 2010; Percival et al. 2010), supernova observations (e.g., Kessler et al. 2009; Amanullah et al. 2010), and a host of other astronomical observations (see, e.g., Spergel et al. 2007 and Komatsu et al. 2010 for reviews). The amplitude of the fluctuations in the damping tail is a sensitive probe of matter fluctuations at $k \simeq 0.1 - 0.25 \text{ Mpc}^{-1}$ — thus, precision measurements constrain the spectral index of the primordial curvature perturbations, n_s , and its variation with scale. Because the positions of the high order acoustic peaks are sensitive to the evolution of the sound speed of the universe and its composition, these measurements constrain the primordial helium fraction and the number of relativistic species including neutrinos (see e.g. White 2001; Bashinsky & Seljak 2004; Trotta & Hansen 2004; Ichikawa et al. 2008a,b; Komatsu et al. 2010). While the damping tail gives us leverage on early universe cosmology, the *composite tail* ($\ell \gtrsim 3000$) is sensitive to a variety of astrophysical phenomena, including the properties of the intracluster medium, the redshift distribution and clustering of dusty sub-mm galaxies (Hall et al. 2010), and the physics of reionization (Huffenberger & Seljak 2005; Sehgal et al. 2007). Sehgal et al. (2010b) describe theoretical expectations for these small-scale measurements.

In this paper, we present a measurement of the CMB power spectrum over the range of multipoles $500 < \ell < 10000$ from the Atacama Cosmology Telescope (ACT) using the 148 GHz and 218 GHz channel data from the 2008 observing season. ACT is a mm-wave, arcminute-resolution telescope (Fowler et al. 2007; Swetz et al. 2010) custom built to make precise observations of the microwave sky over the damping and composite tail regimes. In recent years, several groups have reported rapidly improving measurements of fluctuations over various portions of this multipole range: Bolocam (Sayers et al. 2009), QUaD (Brown et al. 2009; Friedman et al. 2009), APEX-SZ (Reichardt et al. 2009a), ACBAR (Reichardt et al. 2009b), SZA (Sharp et al. 2010), BIMA (Dawson et al.

2006), CBI (Sievers et al. 2009), SPT (Lueker et al. 2010; Hall et al. 2010) and ACT (F10). The current paper enhances the results of F10 in a few ways. First, we augment our 148 GHz data with results from our 218 GHz channel. Second, the power spectrum estimation methods have been revised to provide enough angular frequency resolution to detect the acoustic features on the damping tail. Third, the area used for the power spectrum analysis has been increased from $\simeq 220 \text{ deg}^2$ to $\simeq 300 \text{ deg}^2$.

This paper is one in a set of papers describing ACT and its 2008 Southern survey. The ACT instrument is described in Swetz et al. (2010), a 148 GHz point source catalog is presented in Marriage et al. (2010b) (hereafter M10), and a 148 GHz SZ cluster catalog is presented in Marriage et al. (2010a). Menanteau et al. (2010) discuss the multi-wavelength followup of ACT clusters, while Sehgal et al. (2010a) present σ_8 constraints from SZ cluster detections. On the power spectrum side, Hajian et al. (2010) report on the calibration of ACT maps using cross-correlations with WMAP seven-year maps, and Dunkley et al. (2010) present the constraints on cosmological parameters derived from the power spectrum presented here.

The paper is organized as follows. In Section 2, we briefly review the ACT instrument and observations made so far, touching upon mapmaking and beam determination techniques, and the calibration of our data. In Section 3, we describe the method used in power spectrum estimation. Simulations used to test our pipeline are discussed in Section 4. We present our results in Section 5 and discuss tests performed to validate our results. The main sources of foreground contamination and the methods used to treat them are discussed in Section 6. We discuss the lensing contribution to the power spectrum in Section 7, before concluding in Section 8.

2. OBSERVATIONS AND MAPMAKING

ACT is a 6-meter off-axis Gregorian telescope (Fowler et al. 2007) situated at an elevation of 5190 meters on Cerro Toco in the Atacama desert in northern Chile. ACT has three frequency bands centered at 148 GHz (2.0 mm), 218 GHz (1.4 mm) and 277 GHz (1.1 mm) with angular resolutions of roughly $1'.4$, $1'.0$ and $0'.9$, respectively. The high altitude site in the arid desert is excellent for microwave observations due to low precipitable water vapor and stability of the atmosphere. The tropical location of ACT permits observations on both the northern and southern celestial hemispheres. Further details on the instrument are presented in Swetz et al. (2010), F10 and references therein¹. As of this writing, ACT has completed three observing seasons (2007, 2008, 2009) surveying two stripes on the sky: a 9° -wide stripe centered on declination $-53^\circ 5'$ and a 5° -wide stripe centered on the celestial equator. The power spectrum presented here is derived from the southern stripe data from the two lower frequency channels in 2008. An important aspect of ACT's scanning strategy is the cross-linking of observations. Every point in the survey area is scanned along two different directions during the night. We use constant elevation scans centered on two different

¹ ACT Collaboration papers are archived at <http://www.physics.princeton.edu/act/>

azimuths — once when the survey area is rising on the sky, and again when it is setting toward the end of the observing night. In principle, this cross-linking allows an unbiased reconstruction of *all* modes in the map, and permits recovery of the power spectrum for multipoles as low as a few hundred with errors completely dominated by cosmic variance. A more detailed account of the observing strategy of ACT can be found in F10 and M10.

2.1. Mapmaking

For the 148 GHz data, we use the maps from F10, in which details of the data reduction and mapping can be found. We model the data as $\mathbf{d} = \mathbf{P}\mathbf{m} + \mathbf{n}$, where \mathbf{P} projects the map \mathbf{m} into the time stream, and \mathbf{n} is the noise, with covariance matrix \mathbf{N} . There are no constraints on the contents of \mathbf{P} and \mathbf{m} , which can contain multiple components. For ACT, in addition to solving for the map of the sky, we solve simultaneously for noise correlated amongst the detectors. Under this model, the data can be described by

$$\mathbf{d} = \mathbf{P}\mathbf{m} + \mathbf{A}\mathbf{c} + \mathbf{n} \quad (1)$$

where \mathbf{A} are (assumed constant) patterns of correlation across the array (such as a common mode) and \mathbf{c} are the timestreams associated with each pattern in \mathbf{A} . This is mathematically equivalent to having separate blocks in a generalized projection matrix and a generalized map solution. The maximum likelihood solution, $\tilde{\mathbf{m}}$, is then given by solving the standard mapping equation:

$$\mathbf{P}^T\mathbf{N}^{-1}\mathbf{P}\tilde{\mathbf{m}} = \mathbf{P}^T\mathbf{N}^{-1}\mathbf{d}. \quad (2)$$

We solve for $\tilde{\mathbf{m}}$ iteratively using a preconditioned conjugate gradient (PCG) method (Press et al. 2007; Hinshaw et al. 2007).

The higher atmospheric noise in the 218 GHz data required some changes to the mapping pipeline. For the 148 GHz data, we found that taking the array patterns \mathbf{A} to be the eigenvectors corresponding to the 10 largest eigenvalues of the data covariance matrix for each 15-minute chunk of time-ordered data (TOD) worked well for correlated noise rejection. For the 218 GHz data, substantial atmospheric power remains with this prescription, and so we altered it as follows: first, for each TOD we take the band-limited data between 0.25 and 4 Hz, find the eigenvalues and eigenvectors of the corresponding covariance matrix, and keep all modes with eigenvalues larger than 3.5^2 times the median eigenvalue (equivalently modes with timestream amplitudes larger than 3.5 times the median). We typically find 30–50 modes. Then, we create the covariance matrix from the data high-pass filtered above 4 Hz, project out the modes already found in the 0.25–4 Hz band, and keep all remaining modes with eigenvalue larger than 2.5^2 times the median. We typically find 1–2 additional modes in this step. Of the several different mode removal schemes we tried, we found this fairly aggressive one gave the best signal-to-noise on intermediate and small angular scales, where the 218 GHz data are most valuable, at the price of worse signal-to-noise and slower convergence of the mapper on large scales. Since the method estimates both the correlated modes and the map of the sky simultaneously, mode removal does not bias the maps, although it makes some sky map modes noisier.

The second change in the mapping is the use of a preconditioner with the conjugate gradient technique. In general, a preconditioner is an approximation to $(\mathbf{P}^T\mathbf{N}^{-1}\mathbf{P})^{-1}$ that is used to speed the convergence of the conjugate gradient solution. In general, we would use a two-piece preconditioner, with separate parts for the sky solution and the correlated noise component. However, if the columns of \mathbf{A} are diagonal (which they are by construction), the natural preconditioner for the correlated part of the noise is simply the identity matrix and so we neglect it. The sky preconditioner is a diagonal matrix with elements equal to the inverse of the “ n_{obs} ” map — a map of the number of observations in each pixel. We find the use of this preconditioner greatly aids in the convergence of the conjugate gradient solution for 218 GHz data, unlike for the 148 GHz data where it makes little difference.

2.2. Beam Measurements

The window function due to the beam shape is obtained in a method similar to that followed in Hincks et al. 2009, hereafter H09. We use 22 (26) maps of Saturn from throughout the 2008 season to estimate the beam shapes for 148 GHz (218 GHz). The maps are made with an independent pipeline from that used in H09, but produce consistent results. Each map is binned into a symmetrized radial beam profile, with $0'.17$ bins, out to a radius, θ_{max}^b , of $8'(6')$ for 148 GHz (218 GHz). The dominant noise source in the maps comes from the slow variation of atmospheric brightness, and since this projects into relatively large angular scales it causes significant noise correlation between radial bins in each profile. The individual profiles are thus used to estimate the mean profile, and the full covariance matrix between bins is estimated from the scatter between them. This is in contrast to H09, where the mean profile was computed from a stacked map, and only diagonal errors were considered. The full covariance formulation serves to propagate the large scale map noise into low- ℓ uncertainty in the window function. The zero level of the radial profile is poorly constrained in the maps due to atmospheric contamination at large radii. Because the sharp circular edge of the cold Lyot stop produces an Airy pattern at large scales, the radial profile is expected to fall as $1/\theta^3$ (H09), so a fit to this form beyond θ_{max}^b estimates the zero level and its covariance with the binned profile. The resulting function is also used to extrapolate the beam when computing the beam transform.

Following the same prescription as H09, a small set of basis functions is fitted to the inner part of the beam profile. A natural choice of basis functions in harmonic space are Zernike polynomials which are compact on the unit disk, or their transforms in angular space which are Bessel functions divided by a linear function of radius. For 148 GHz (218 GHz) only 25 (18) basis functions are required to approximate the beam profile at the 1% (2%) level within $1'$ and below the 15% level within $8'(6')$. Using a small set of basis functions simplifies calculation of the covariance matrix of the window function (Page et al. 2003) from the covariance of the coefficients of the basis functions and the wing fit. For the purpose of calibration against WMAP maps (Hajian et al. 2010), the normalization of the window function is fixed

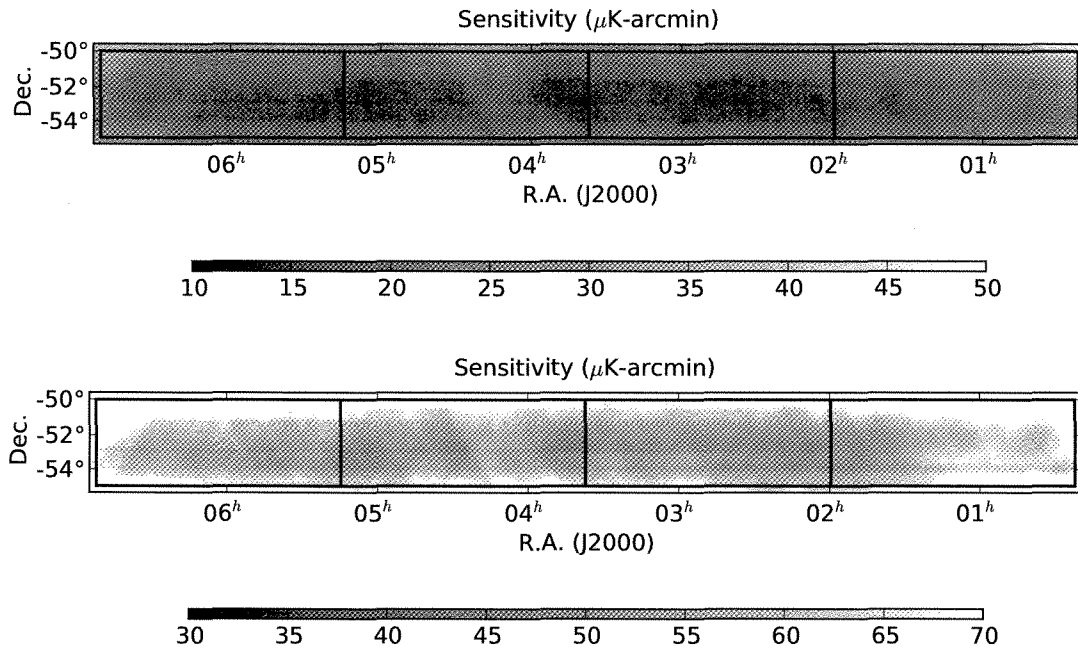


FIG. 1.— The rms temperature uncertainty for one-arcminute pixels for the 148 GHz maps (top) and 218 GHz map (bottom). Also shown in bold lines are the boundaries of the four rectangular patches used for spectral analysis.

to unity at $\ell = 700$, and the calibration error is factored out of the covariance matrix (described in H09) so that the window function has zero error at $\ell = 700$. Between $\ell = 1000$ and 10000, uncertainties in the window function at 148 GHz are between 0.7 and 0.4%, and at 218 GHz between 1.5 and 0.7%. The uncertainty in the window function is significantly higher for $\ell < 700$, where atmospheric contamination makes measurements very difficult. In this paper, we use the beam window function only. The full covariance matrix of the window function is used in Dunkley et al. (2010).

2.3. Calibration

ACT scans every point in the survey area both when it is rising and setting. This cross-linking along with the unbiased map-making method described in Section 2.1 allows the reconstruction of all modes in the maps, without biasing the large-angle modes. As described in Hajian et al. (2010), we calibrate the 148 GHz ACT maps directly to WMAP 7-year 94 GHz maps (Jarosik et al. 2010) of the identical regions. Thus, the cosmological analysis is done with the same data used for the calibration. By matching the ACT-WMAP cross-spectrum to the ACT power spectrum and the WMAP 7-year power spectrum (Larson et al. 2010) in the range $400 < \ell < 1000$, Hajian et al. (2010) calibrate the 148 GHz ACT spectrum to 2% fractional temperature uncertainty. Similar methods are applied to 218 GHz maps, but their larger map noise levels on large angular scales mean that the result does not currently improve on the results of an independent method of calibration, based on observations of Uranus², described below.

² It may be noted that although we use observations of Saturn for evaluating beam shape, we do not use them for calibration. Re-

ACT made approximately thirty usable observations of the planet Uranus during its 2008 season. We generated a map per observation after calibrating the time-ordered data to detector power units, and determined the peak response of the planet with corrections for temperature dilution due to the finite instrumental beam size. The result is then compared to the Uranus temperature in CMB differential units at the effective band center to obtain a calibration factor (see Hajian et al. 2010, for more details). We take the brightness temperature of Uranus to be 107 ± 6 K and 96 ± 6 K for the 148 GHz and 218 GHz bands respectively. These temperatures are based on a reprocessing of the data presented by Griffin & Orton (1993), in combination with WMAP 7-year measurements of Mars and Uranus brightnesses (Weiland et al. 2010). We find an uncertainty of 7% in the Uranus-based calibration that is dominated by the 6% uncertainty in the planet’s temperature. The absolute calibration is consistent with the WMAP-based calibration described above. For 218 GHz, we adopt the Uranus-based calibration.

3. POWER SPECTRUM METHOD

Das et al. (2009) describe the basis of our power spectrum method. This paper improves on the analyses of Fowler et al. (2010) by optimizing the angular frequency resolution of the spectrum over the damping tail to resolve the acoustic features, as well as by extending the multipole range of the power spectrum to $\ell = 10000$.

3.1. Fields used for Spectra

liable modeling of the brightness temperature of Saturn is complicated by the effects of its rings. The effective temperature changes with the ring inclination over the season. This is not the case for Uranus, whose brightness temperature is fairly stable.

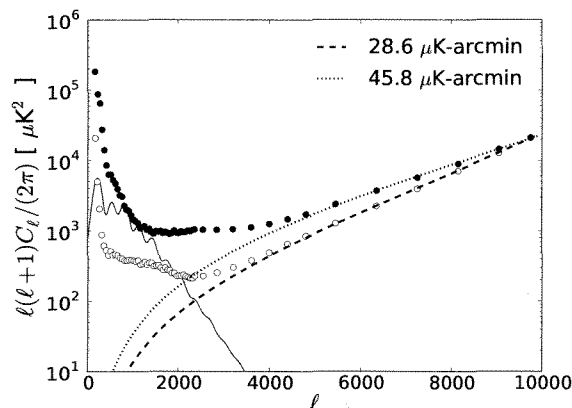


FIG. 2.— Beam deconvolved noise power spectra of the 148 GHz (open circles) and the 218 GHz (filled circles) maps shown against a theoretical lensed CMB spectrum (solid line). The theoretical white noise levels are shown using dashed (for 148 GHz) and dotted (for 218 GHz) lines.

The power spectrum is computed from independent analysis of four contiguous patches in the ACT southern stripe (see Fig. 1) that together cover a rectangular area of 296 deg^2 from $\alpha = 00^{\text{h}}22^{\text{m}}$ to $06^{\text{h}}52^{\text{m}}$ (5.5° to 103°) in right ascension and from $\delta = -55^\circ$ to -50° in declination. Each patch is $5^\circ \times 14.8$ in size. The full dataset is also split into four subsets by distributing data from each of four successive nights into different subsets, thereby producing four independent maps of identical coverage and roughly equal depth. We refer to each subset as a *split*. Therefore, for each of the two frequency bands, we have four patches with four splits in each patch. Each split is properly cross-linked, i.e., it represents the maximum likelihood solution of a dataset containing both rising and setting observations. A typical white noise level in the 148 GHz season map (all four splits taken together) range from 25–50 $\mu\text{K-arcmin}$, while for the 218 GHz map it varies between 40 and 70 $\mu\text{K-arcmin}$, with the noise being higher toward the edges of the maps (Fig. 1). The noise power spectra of the 148 GHz and 218 GHz maps are shown in Fig. 2. At small scales, the noise is typically white, while at larger scales there is a plateau of noise with a red tail towards low multipoles, which is primarily attributable to atmospheric contamination.

3.2. Preprocessing of Maps

Before patches are cut from the splits, each split map is high-pass filtered in Fourier-space. The high-pass filter is designed to remove all modes below $\ell = 100$, and suppress the modes between $\ell = 100$ and 500 with a sine-squared function, which rises from zero to unity within that range. The very lowest ($\ell < 100$) multipoles are dominated by extremely large atmospheric noise, and rolling off the filter to $\ell = 500$ prevents these modes from leaking into and unnecessarily biasing the higher multipole portion of the spectrum.

Next, we *prewhiten* the maps. Prewhitening is a local, real-space operation on the map designed to reduce the dynamic range of its Fourier components (Das et al. 2009). Prewhitening is particularly important for the current spectrum, as it is designed to specifically target

the damping tail of the CMB ($1000 < \ell < 3000$). This region of the spectrum has a steep slope ($C_\ell \sim \ell^{-4}$). In the absence of prewhitening, finite boundary effects and application of the point source masks cause a large amount of power to be aliased from low to high multipoles contaminating the spectrum at large multipoles. Although an unbiased estimate of the spectrum is obtained by deconvolving the mode-coupling window from the spectrum (See Section 3.5 and Das et al. 2009.), this contamination adds to the uncertainty on the estimate, resulting in unnecessarily large error bars. Prewhitening is performed by taking the difference of two versions of the same map, one convolved with a disc of radius $1'$ and the other with a disc of $3'$. To simulate discs with a pixelated kernel, we use the cloud-in-cell approach where we assign appropriate weights for pixels under the disc according to its area of overlap with the disc. This operation approximately amounts to taking the Laplacian of the map, which effectively multiplies the map Fourier transform by $\sim (\ell R)^2$ and the spectrum by $\sim (\ell R)^4$ in the multipole range $900 < \ell < 3000$. Here $R = 1'$ is the radius of the smaller disc. This flattens out the damping tail but makes the less colored lower multipole portion of the spectrum steeply rising. This transition occurs around $\ell \simeq 900$, where $\ell^2 R^2 \approx 0.02$. Therefore, we add 2% of the original map back to the resulting map and this flattens out the $\ell < 900$ portion of the spectrum. Note that the prewhitening kernel is designed to flatten out by $\ell \sim 3500$ and therefore leaves the high multipole composite tail (which is already white) largely unaffected.

We perform these actions in real space, but calculate the resulting prewhitening transfer function in Fourier-space so we can undo the prewhitening in the final map estimate. Prewhitening reduces the scatter in the $\ell = 2000$ –4000 region of the power spectrum, which provides significant statistical weight for extracting information on the SZ and correlated point source signals.

3.3. The Data Window

After the maps are prewhitened, four patches are cut from each split. Each split patch is then multiplied with a window. The window is a product of three components — a point source mask, the n_{obs} map, and a tapering function. The details of the point source mask are given in Section 6. To simplify the application, we generate a single n_{obs} map per patch for each frequency by summing over the four splits for that patch. For the cross-frequency spectra we again generate a single n_{obs} map by summing the n_{obs} maps from the two frequencies. Multiplication by the n_{obs} map essentially downweights the relatively scantily observed and poorly cross-linked pixels toward the top and bottom edges of the map. Finally, to avoid ringing from the sharp edges of a patch, each patch is multiplied by a taper that gently rolls to zero at the edge of the map. We use a simple taper generated by taking a tophat function which is unity in the center and zero over ten pixels at the edges, and convolving this with a Gaussian kernel of FWHM $\simeq 2.5'$.

We will denote the two frequency channels as A and B , and use i, j, k, l to denote the sub-season splits. The patches will be denoted by Greek indices. We will distinguish the data windows only by their patch indices, and denote them by $W_\alpha(\hat{\mathbf{n}})$, suppressing the channel labels. Thus, split i of the patch α of frequency A , after

windowing, becomes³:

$$\tilde{T}_\alpha^{iA}(\hat{\mathbf{n}}) \equiv W_\alpha(\hat{\mathbf{n}})T_{F;\alpha}^{iA}(\hat{\mathbf{n}}) \quad (3)$$

where $T_{F;\alpha}^{iA}(\hat{\mathbf{n}})$ denotes the map after filtering and prewhitening, and is related to the original map $T_\alpha^{iA}(\hat{\mathbf{n}})$ through the Fourier-space relation:

$$T_{F;\alpha}^{iA}(\ell) = F_\ell T_\alpha^{iA}(\ell), \quad (4)$$

where F_ℓ is a product of the prewhitening and beam transfer functions, the pixel window function and the high-pass sine-squared filter. Here, and in what follows, the boldface ℓ is meant to represent a 2D wave-vector in Fourier-space, while we use regular ℓ to represent its the absolute value. Symbols subscripted by ℓ represent quantities that are isotropic, while ones shown with an argument (ℓ) or subscripted by ℓ , stand for explicitly 2D quantities in Fourier-space.

3.4. Azimuthal Fourier-mode Weights and Binning

After windowing the data, a 2D pseudo-spectrum is computed as:

$$\tilde{C}_\ell^{iA \times jB} = \text{Re} \left[\tilde{T}_\ell^{*iA} \tilde{T}_\ell^{jB} \right], \quad (5)$$

where the patch index has been suppressed to simplify notation. At this stage, we crop out and retain a rectangular area of the 2D spectrum defined by $-10000 < \ell_x < 10000$ and $-10000 < \ell_y < 10000$, throwing out all information at $\ell_x, \ell_y > 10000$. This trimming reduces the number n_F of Fourier-space pixels by a factor of ~ 4 , and is crucial for several subsequent operations, especially the direct computation of the mode-coupling matrix (see Section 3.5), which involves a computational step that scales as n_F^2 . The 1D binned spectrum \tilde{C}_b results from averaging the 2D spectrum in annular bins:

$$\tilde{C}_b^{iA \times jB} = \sum_\ell P_{b\ell} \tilde{C}_\ell^{iA \times jB}, \quad (6)$$

where $P_{b\ell}$ is the binning matrix, whose value is zero when ℓ is outside the annulus defined by bin index b . Because the noise properties are not isotropic in Fourier space, we employ a 2D weighted average, with the weights determined by the noise power spectrum. We briefly describe here how the weights are computed. For a given frequency, say, A , we compute the mean 2D auto spectrum from the four split patches and subtract the mean 2D cross-power spectrum from it. This gives an estimate of the 2D noise power spectrum for the split patch N_ℓ^{AA} . We assume that each split has roughly equal noise, and approximate the season noise power spectrum as $N_\ell^{AA}/4$. We estimate the 2D variance of the spectral estimator $C_\ell^{iA \times jB}$ as

$$\sigma^2(C_\ell^{iA \times jB}) \simeq (C_\ell^{\text{th}} b_\ell^{A^2} + N_\ell^{AA}/4)(C_\ell^{\text{th}} b_\ell^{B^2} + N_\ell^{BB}/4), \quad (7)$$

³ Throughout the paper, we use $X(\hat{\mathbf{n}})$ to denote a real space quantity, and $X(\ell)$ or X_ℓ to denote its Fourier transform. We use the tilde to denote windowing (not Fourier transformation), so that $\tilde{X}(\hat{\mathbf{n}})$ is $W(\hat{\mathbf{n}})X(\hat{\mathbf{n}})$ after multiplication by a window. All calculation is done under the flat-sky approximation.

where C_ℓ^{th} is an azimuthally symmetric theoretical CMB spectrum, and b_ℓ represents the beam transfer function. The first estimate of the azimuthal weights is chosen to be the inverse of the variance (Eq. [7]):

$$w_\ell^{iA \times jB} = \frac{1}{\sigma^2(C_\ell^{iA \times jB})}. \quad (8)$$

The resulting 2D *Fourier-space map*⁴ of weights is typically noisy, and before using it to bin the power spectrum, we treat it in the following manner. In every annulus, defined by the bins b , we identify outlier pixels whose values are greater than a threshold α_b (typically chosen to be 10 times the median in that annulus) and set their values to α_b . We then divide the pixels in each annulus by their mean, and smooth the resulting 2D Fourier-space map with a Gaussian with FWHM of three pixels. These operations are designed to produce well-behaved Fourier-space weight maps, with emphasis on bringing out the overall azimuthal variation in the weights by removing outliers and radial dependence. The Fourier-space weight map used in the 148 GHz \times 148 GHz spectrum is shown in Fig. 3. A comparison of this Fourier-space weight map with the signal and noise levels in Fig. 2 shows the expected behavior of the weights being isotropic in the signal-dominated regime at low ℓ , with the anisotropy becoming more pronounced as the noise begins to dominate. The main sources of anisotropy in the weights are the rays of excess noise along and perpendicular to the scan directions, leading to two slightly rotated (because the rising and setting scans are not exactly orthogonal) ‘‘X’’ shaped patterns of low weight regions. Another significant noise term is attributable to scan synchronous noise, likely caused by instabilities induced by acceleration at scan turn-arounds. The noise leads to horizontal striping in the maps and manifests itself as an excess of power in a vertical strip $-90 < \ell < 90$ in the Fourier space. We set the pixels inside this strip in the 2D Fourier-space weight map to zero. In terms of the 2D weights, the binning matrix can be expressed as,

$$P_{b\ell} = \frac{w_\ell}{\sum_{\ell \in b} w_\ell}. \quad (9)$$

It is noteworthy that both the n_{obs} weighting in real space and the azimuthal weighting in Fourier space lead to a 5-10% percent lowering of the uncertainties in the final spectral estimates.

3.5. Mode-coupling

From Equations (3) to (6), we can express the binned pseudo-spectrum \tilde{C}_b in terms of the underlying spectrum C_ℓ :

$$\tilde{C}_b^{iA \times jB} = \sum_{\ell, \ell'} P_{b\ell} |W(\ell - \ell')|^2 F_\ell^2 C_{\ell'}^{iA \times jB}. \quad (10)$$

This quantity can be related to a binned version of the the true spectrum C_b through an inverse binning opera-

⁴ The 2D Fourier transform $X(\ell)$ of a real-space map $X(\hat{\mathbf{n}})$ is defined on a 2D pixelated grid in ℓ -space. We use the term ‘‘Fourier-space map’’ to denote any 2D quantity defined on the same ℓ -space grid. Such a quantity can be solely defined in ℓ -space and is not necessarily derived from a Fourier transform of some known real-space quantity.

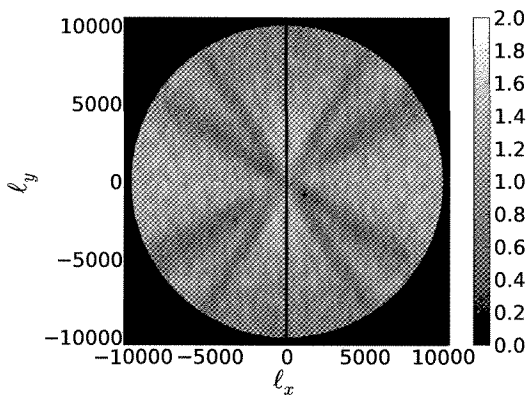


FIG. 3.— Azimuthal weights used for binning the 148×148 GHz spectrum. The horizontal stripe is an artifact of the symmetry of the Fourier space, and does not affect our calculations. The vertical stripe of zero weight is added in by hand, as described in the text. Due to the symmetry of Fourier transforms, only the upper half plane of this weight map is independent. The small grainy disk at the center corresponds to $\ell < 500$ where we do not perform any smoothing — these modes are discarded from the final power spectrum.

tor $Q_{\ell b}$, which is unity when $\ell \in b$ and zero otherwise:

$$\begin{aligned} \tilde{C}_b^{iA \times jB} &= \sum_{\ell, \ell', b'} P_{b\ell} |W(\ell - \ell')|^2 F_{\ell'}^2 Q_{\ell' b'} C_{b'}^{iA \times jB} \\ &\equiv \sum_{b'} M_{bb'} C_{b'}^{iA \times jB}, \end{aligned} \quad (11)$$

where $M_{bb'}$ is the mode-coupling matrix. We compute the $M_{bb'}$ exactly without resorting to any of the one-dimensional approximations often used in the flat-sky case (Hivon et al. 2002; Das et al. 2009; Lueker et al. 2010). The mode-coupling matrices are well behaved and stable to inversion. We define the unbiased, decorrelated estimator of the power spectrum (indicated with the circumflex):

$$\hat{C}_b^{iA \times jB} = \sum_{b'} M_{bb'}^{-1} \tilde{C}_{b'}^{iA \times jB}. \quad (12)$$

For a single frequency spectrum, we compute the cross-power spectrum in each patch as the mean of the six cross-power spectra (from four splits). We also compute the auto-spectrum as the mean of four auto-spectra from the splits. The variance for a given patch is determined by combining the noise estimate (obtained as one-fourth the difference of the mean auto and the mean cross-spectrum) in that patch with a theoretical signal power spectrum. We combine the four mean cross-power spectra from the patches with inverse variance weighting to obtain the final spectrum.

For the cross-frequency power spectra, we compute twelve cross-power spectra in each patch by crossing each split for a frequency with the three other splits from the other frequency. We do not cross the same splits (consisting of data from the same nights) between the two frequencies because of the possible contamination from covariant atmospheric noise among the two channels. We proceed by averaging those twelve per patch, and then combining the cross-spectra from the patches with inverse weighting as above. In this case, we estimate the

variance in a patch as the square root of the product of the single-frequency variances in that patch.

For parameter estimation (see Dunkley et al. 2010), we wish to compare our spectral estimates \hat{C}_b to theoretical power spectra C_ℓ^{th} . To do so we find the bandpower window function, $B_{b\ell}$ which converts C_ℓ^{th} to binned theoretical spectra C_b^{th} . (We suppress the superscripts on the C s for clarity.) We begin with Equation (10) which relates the binned spectrum to the underlying 2D spectrum, and introduce a matrix $I_{\ell\ell'}$ that describes the mapping of a theoretical spectrum defined at integer multipoles onto our discrete Fourier pixels. The binned theoretical spectrum follows from Equation (12);

$$\begin{aligned} C_b^{\text{th}} &= \sum_{b', \ell', \ell} M_{bb'}^{-1} P_{b'\ell'} |W(\ell'' - \ell')|^2 F_{\ell'}^2 I_{\ell'\ell} C_\ell^{\text{th}} \\ &\equiv B_{b\ell} C_\ell^{\text{th}}. \end{aligned} \quad (13)$$

Since each rectangular Fourier-space pixel spans a range of integral multipoles, it gets contributions from several multipoles in the theoretical spectrum. We choose $I_{\ell\ell'}$ as a Gaussian with its peak at the pixel center (ℓ_x, ℓ_y) , and width corresponding to the pixel size, so as to give more weight toward the center of the pixel. The bandpower window functions are insensitive to the exact form of this response function. As with the mode-coupling matrix, we compute $B_{b\ell}$ exactly.

3.6. Summary of power spectrum method

The angular power spectrum is conceptually simple, but measuring it over a wide range of angular scales with sufficient angular resolution to see acoustic peaks from maps over a small portion of sky with significant variation in sensitivity over the map, is technically challenging. As a quick reference, we summarize here our multiple-step procedure to accomplish this task:

- Divide of the data into independent “splits” each comprising a quarter of the total observation nights, and construct independent maps for each split;
- High-pass filter the maps, to eliminate the large angular scale modes which are poorly constrained and can bleed power into the smaller-scale modes;
- Prewhiten the maps through real-space convolutions so that the power spectrum of the maps is flattened, which reduces aliasing of power from large scales to small scales given the steep power spectrum of the maps;
- Divide each split map into four independent patches and construct a window for each patch which accounts for varying sensitivity over the patch, perform masking of point sources, and edge tapering to avoid spurious power from a sharp cut-off;
- Calculate 2D binned pseudo-spectra from the Fourier transformation of the windowed, prewhitened, filtered maps, for each pair of frequency, split, and patch values;

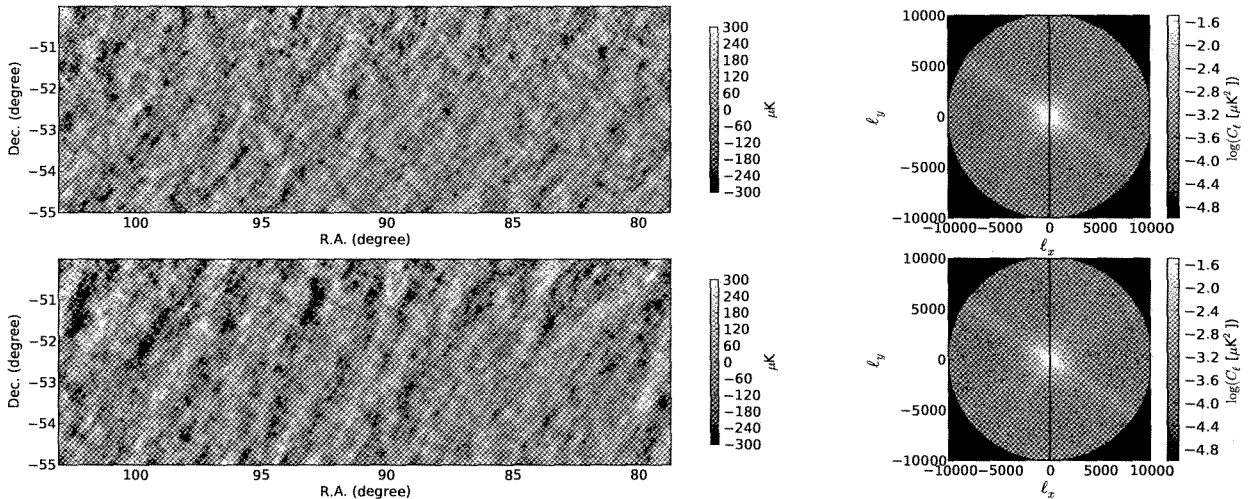


FIG. 4.— Anisotropic noise in data and simulated patches. *Top*: Difference map of two splits in an 148 GHz data patch (*left*), and its 2D power spectrum (*right*). *Bottom*: Same as above, except for a simulated patch. This random realization was seeded by the noise power spectrum of data patch shown in the top panel, as described in the text.

- Construct a one-dimensional binned pseudo-spectrum via azimuthal averaging of the two-dimensional spectrum over bins in multipole ℓ ;
- Estimate the true binned power spectrum from the binned one-dimensional pseudo-spectrum by inverting a mode-coupling matrix which accounts for the effects of beam profile, prewhitening, filtering, pixelization, and windowing.

We combine the spectra from individual patches with inverse variance weighting to obtain the final spectrum. We also construct a bandpower window function which converts a theoretical power spectrum into the corresponding binned power spectrum which we estimate from the maps. We verify these techniques on sky simulations, described below.

4. SIMULATIONS

Simulations play an important role both in verifying that our mapper is working correctly and in understanding our spectrum. Specifically, we use simulations to test for map convergence, estimate bandpower covariances, and confirm analytic estimates of the power spectrum uncertainties.

To test for convergence, we perform a set of end-to-end simulations of the 148 GHz and 218 GHz maps. We simulate signal-only time streams by mock observations of input celestial maps (from Sehgal et al. 2010b) and generate noise time-streams which capture the main features of the noise in the data. To test for convergence, we simulate a signal-plus-noise map, subtract from it a noise-only map at the same iteration and compare the power spectrum of the resulting difference map with that from a signal-only map. For 148 GHz, the maps are well converged to multipoles as low as $\ell \gtrsim 200$ by iteration 1000, and we keep $\ell \geq 500$ in our final spectrum. For the 218 GHz maps, where higher atmospheric noise warranted a slightly different mapmaking approach (see Section 2.1), convergence at large scales is slow. After

1000 iterations, the 218 GHz maps are converged down to $\ell = 1000$, but not at smaller ℓ . Moreover, the 218 GHz spectrum is highly noise-dominated at low multipoles. Thus, we limit spectra involving 218 GHz to $\ell \geq 1500$ in our analyses.

In order to investigate the bandpower covariance and to validate the analytic prescription for uncertainties in the spectrum, we ran a large set of Monte Carlo simulations that contain key aspects of the noise properties manifested in the data. Due to the iterative nature of our

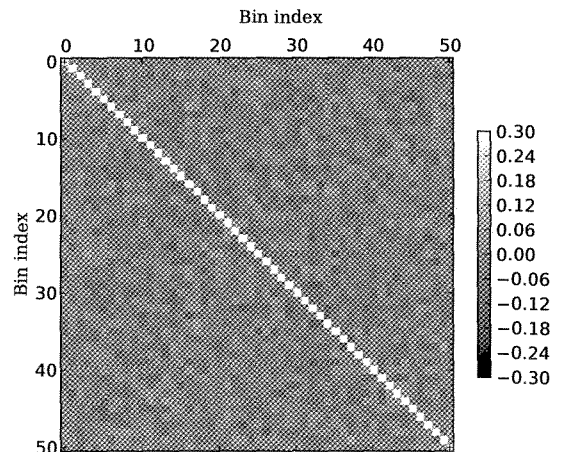


FIG. 5.— Normalized covariance matrix for decorrelated bandpowers for 148 GHz \times 148 GHz spectrum, based on 960 signal+noise Monte Carlo simulations. The bins are defined in Table 1. All values on the diagonal are unity by definition. The colorbar has been stretched to reveal the variations in the off-diagonal elements.

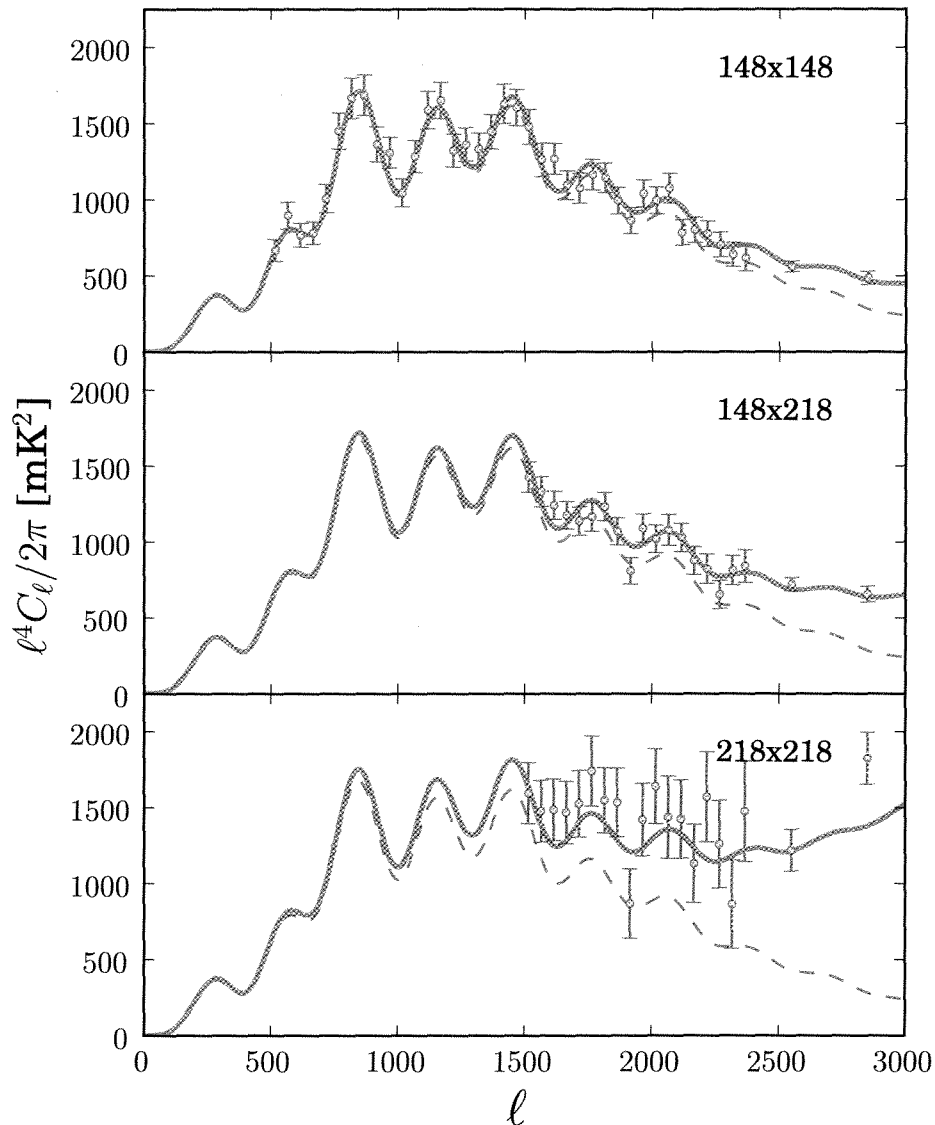


FIG. 6.— Intermediate multipole $500 < \ell < 3000$ portion of the power spectra plotted with an ℓ^4 scaling to emphasize the acoustic peaks. The thick orange curve shows the best-fit model including the CMB secondaries and point source contribution taken from our companion paper Dunkley et al. (2010). The model depends on the frequency combination. The thin dashed teal line shows the best-fit lensed CMB-only theory, which is frequency independent. From top to bottom, the three panels show the 148 GHz, the 148×218 GHz and the 218 GHz spectra. For this plot, data spectra and errors from Table 1 have been scaled by best-fit calibration factors from Dunkley et al. (2010) of 1.02^2 , 1.02×1.09 and 1.09^2 for the 148 GHz, the 148×218 GHz and the 218 GHz spectra, respectively.

map-making pipeline, it is prohibitively expensive to run a large number of end-to-end simulations starting from simulated timestreams. Therefore, we adopt an intermediate approach, in which we find a prescription for the data for generating realistic realizations of noise in map space, to which we add signal realizations. We start with two quarter-season split maps, $T_0(\hat{\mathbf{n}})$ and $T_1(\hat{\mathbf{n}})$, with corresponding n_{obs} maps labeled $n_0(\hat{\mathbf{n}})$ and $n_1(\hat{\mathbf{n}})$. We use the difference of these maps to estimate a one-hit noise map $D_{01}(\hat{\mathbf{n}})$ (i.e., a noise map representative of the

variance in each pixel had it been observed only once),

$$D_{01}(\hat{\mathbf{n}}) = \frac{T_0(\hat{\mathbf{n}}) - T_1(\hat{\mathbf{n}})}{\sqrt{n_0^{-1}(\hat{\mathbf{n}}) + n_1^{-1}(\hat{\mathbf{n}})}}. \quad (14)$$

With the other two splits, we compute another one-hit noise map $D_{23}(\hat{\mathbf{n}})$. We compute the average of the 2D power spectra of D_{01} and D_{23} , and use its square root as the amplitude for a Gaussian random field, and generate

random one-hit noise maps by randomizing the phases and converting back to real space. Finally, we divide one such map by $\sqrt{n_i(\hat{\mathbf{n}})}$ to generate a realization of the noise for split i . This procedure captures the non-white and anisotropic aspects of map noise, as shown by Fig. 4.

We also simulate a random residual point source component in our Monte Carlo realizations. The point source model for 148 GHz consists of two components - a radio source population and a dusty sub-mm population, both based on Toffolatti et al. (1998). Source maps are generated assuming a Poisson distribution. For each flux bin S , spanning the range 0.01 – 20 mJy, we generate the number of point sources in that bin as a Poisson realization from the model flux distribution dN/dS (the upper limit is chosen to be the approximate 5σ detection threshold at 148 GHz). Once the number of sources in each bin is generated, their individual fluxes are dithered to redistribute the fluxes within each bin, and the sources are laid down at random positions in the map. For 218 GHz we scale the radio and IR components by their appropriate spectral indices, based on the findings of Dunkley et al. (2010). With these settings, we closely reproduce the level of the Poisson point source signal seen in the high multipole regime of the data power spectra. We estimate the Poisson contribution to the error bars on bandpowers from the scatter of 500 simulations of point source-only maps, and add that as a correction to the (signal+noise)-only error-budget obtained from the signal. At the current level of noise in our maps, these corrections are largely negligible, and we ignore their contribution in the analytic error bar estimation and band power covariance.

We estimate 148 GHz and 218 GHz spectra for 960 signal+noise realizations. Each realization involves simulating four patches and four splits in each patch for each frequency, and estimating the spectra with exactly the same methods as used on the data. The resulting spectra are used for quantifying the covariance between band powers. Fig. 5 displays the normalized covariance matrix, $\text{Cov}(i, j) / \sqrt{\text{Cov}(i, i)\text{Cov}(j, j)}$, where i, j denote bin indices, showing that correlations between adjacent bins are insignificant at the 10% level.

The uncertainties in the band powers are evaluated using the analytic prescription described in the Appendix. First, a large set of isotropic white noise simulations are generated to validate the analytic formula for the single frequency and the cross-frequency error bars. These simulated error bars are found to be in agreement with analytic prescription to less than a percent. The next set of simulations are run with anisotropic noise, Poisson point source realizations, and with real-space and Fourier-space weighting as described in Section 3, and small corrections to analytic prescription due to the anisotropic nature of the noise and weighting are evaluated against the Monte Carlo simulations.

5. POWER SPECTRUM

Applying the methods described in Section 3, we compute the decorrelated bandpowers for the 148, 218 and 148×218 GHz spectra from the unbiased map solutions. The bandpowers are presented in Table 1 and are displayed in Figures 6 and 7. In Fig. 6 we have plotted the lower multipole portion of the power spectrum $500 < \ell < 3000$ with an ℓ^4 scaling to emphasize the

higher-order acoustic peaks in the primary CMB spectrum. In the 148 GHz spectrum we clearly resolve the second through the seventh peaks of the CMB. As discussed in Dunkley et al. (2010), the higher-order peaks provide leverage on cosmological parameters such as the spectral index and its running, the primordial helium fraction and the number of relativistic species. Fig. 7, on the other hand, emphasizes the high multipole tails of the spectra, where the signal is dominated by emission from dusty star-forming galaxies and unresolved radio sources. The intermediate range $2000 < \ell < 4000$ gets a significant contribution from the thermal Sunyaev-Zeldovich effect, and lets us constrain the amplitude of the SZ spectrum.

The data are checked for consistency by performing various null tests, in which difference maps are constructed to cancel true sky signals, and their power spectra examined. We describe the suite of tests here. As described in Section 3.1, the TODs are split into four subsets, generating four maps with independent detector noise. We verify that these maps are consistent with each other by generating difference maps from each pair, and computing the two-way cross spectra from the three independent pairs of difference maps:

$$C_{\ell}^{(i-j) \times (k-l)} = \left\langle \Delta T^{(i-j)*}(\ell) \Delta T^{(k-l)}(\ell) \right\rangle \quad (15)$$

where $\Delta T^{(i-j)} \equiv T^i - T^j$ and the indices i, j, k, l range from 0-3. The same point source mask used to mask the full dataset is applied to the difference maps before calculating the spectra; all other settings also remain the same. The difference maps, $\Delta T^{(i-j)}$, are downweighted by the same n_{obs} maps summed over the four splits used in the windowing of the CMB data (see Section 3.3). Similarly, when binning the power spectra, we use the same azimuthal weights in each patch described in Section 3.3. The three spectra are shown in the top two panels of Fig. 8 for the 148 (top left) and 218 (top right) GHz maps. We compute the bandpowers in the range $500 < \ell < 10000$ ($1500 < \ell < 10000$) for 148 (218) GHz. The mean of the three spectra is consistent with null with $\chi^2 = 42$ (14) for 148 (218) GHz, with 51 (31) degrees of freedom.

Another null test probes the consistency of data taken with the telescope accelerating as it reverses direction at the ends of the scan (turnarounds). Note that for the standard maps, no turnaround cut is applied. Four new split maps are made cutting data near the turnarounds, amounting to losing $\simeq 24\%$ of the total data. Two difference maps are made by pairing one split of the standard map with a different split of the new maps (we avoid differencing the same splits as they have very similar noise structure), and a two-way cross-power spectrum is produced. Any artifact due to the turnaround would be left in these difference maps and might produce excess power. However, we find the resulting spectra, shown in the second row of Fig. 8, to be consistent with null with a $\chi^2 = 62$ and 18 respectively for the 148, 218 maps (again for 51 and 31 degrees of freedom).

The last two null tests probe possible systematics associated with the focal plane and from diurnal effects. The first compares maps made from only the top half of the detector array with those from the bottom half. The sec-

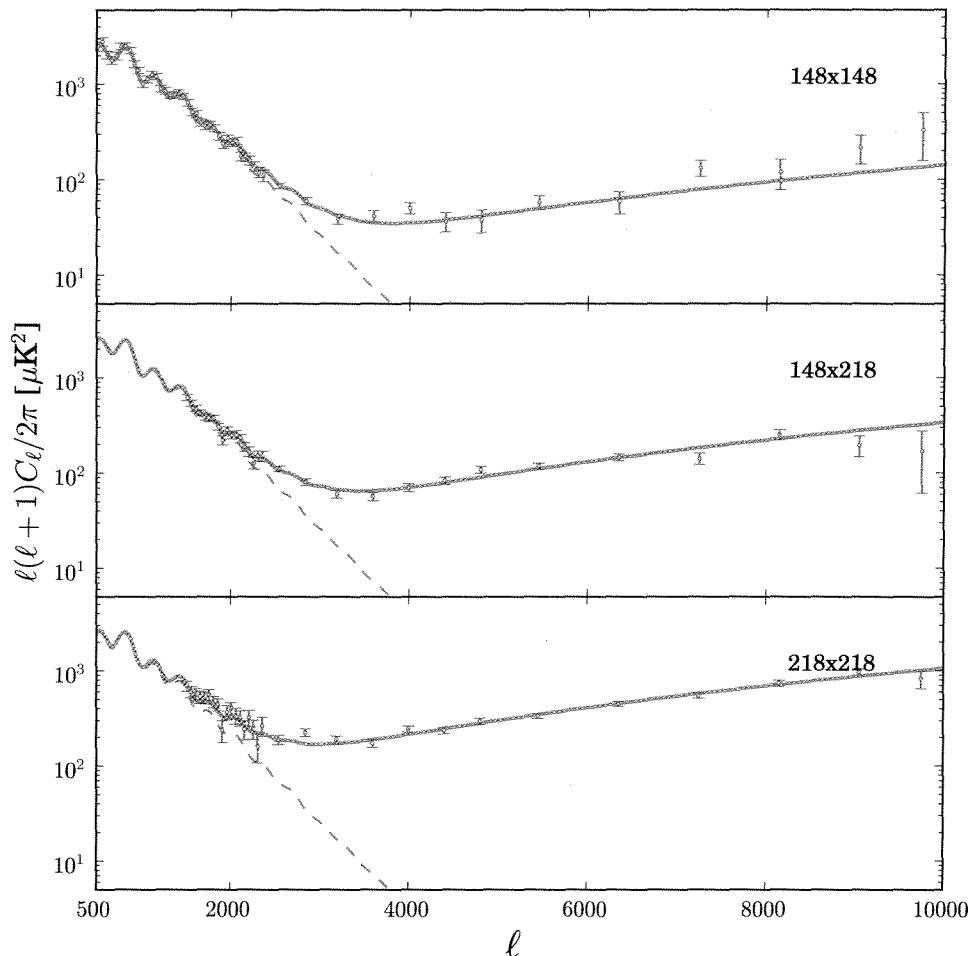


FIG. 7.— Single and cross-frequency spectra plotted with the conventional $\ell(\ell + 1)$ scaling, highlighting the behavior at large multipoles. The thick orange curve shows the best-fit model including the CMB secondaries and point source contribution taken from the companion paper Dunkley et al. (2010). The model depends on the frequency combination. The thin dashed teal line shows the best-fit lensed CMB-only theory, which is frequency independent. From top to bottom, the three panels show the 148 GHz, the 148×218 GHz and the 218 GHz spectra. For this plot, data spectra and errors from Table 1 have been scaled by best-fit calibration factors from Dunkley et al. (2010) of 1.02^2 , 1.02×1.09 and 1.09^2 for the 148 GHz, the 148×218 GHz and the 218 GHz spectra, respectively.

ond compares maps processed with data only from the middle half of the night to the standard maps. The resulting spectra are shown in the final two panels of Fig. 8, and are also consistent with null.

6. FOREGROUNDS

The main foregrounds in the 148 GHz and 218 GHz bands at the angular scales considered here are point sources. We mask the detected ones as described below. Another foreground is the diffuse Galactic dust, which is discussed below in some detail. In the companion paper, Dunkley et al. (2010), we also consider the contributions to the power spectra from the thermal and kinetic SZ effects and the clustering of infrared point sources.

The point sources include radio sources and dusty star-forming galaxies (SFG). The radio sources are typically active galactic nuclei exhibiting synchrotron-dominated

spectra, with emission extending down to lower radio frequencies. The dusty SFG are characterized by absorption of ultraviolet photons from star-forming regions by dust. The dust reemits into graybody radiation that rises with increasing frequency into sub-millimeter bands.

For the purpose of extracting the underlying CMB signal, it is important to identify and mask bright point sources in the maps before the power spectrum is computed. A detailed study of the point source population detected in the ACT 148 GHz 2008 survey is presented in M10. Point sources are detected with the matched filter method (Tegmark & de Oliveira-Costa 1998; Wright et al. 2009; Vieira et al. 2010). To construct the point source masks, we consider detections with $\text{SNR} > 5$ in either frequency. Over the 296 deg^2 area used for power spectrum estimation, we find a total

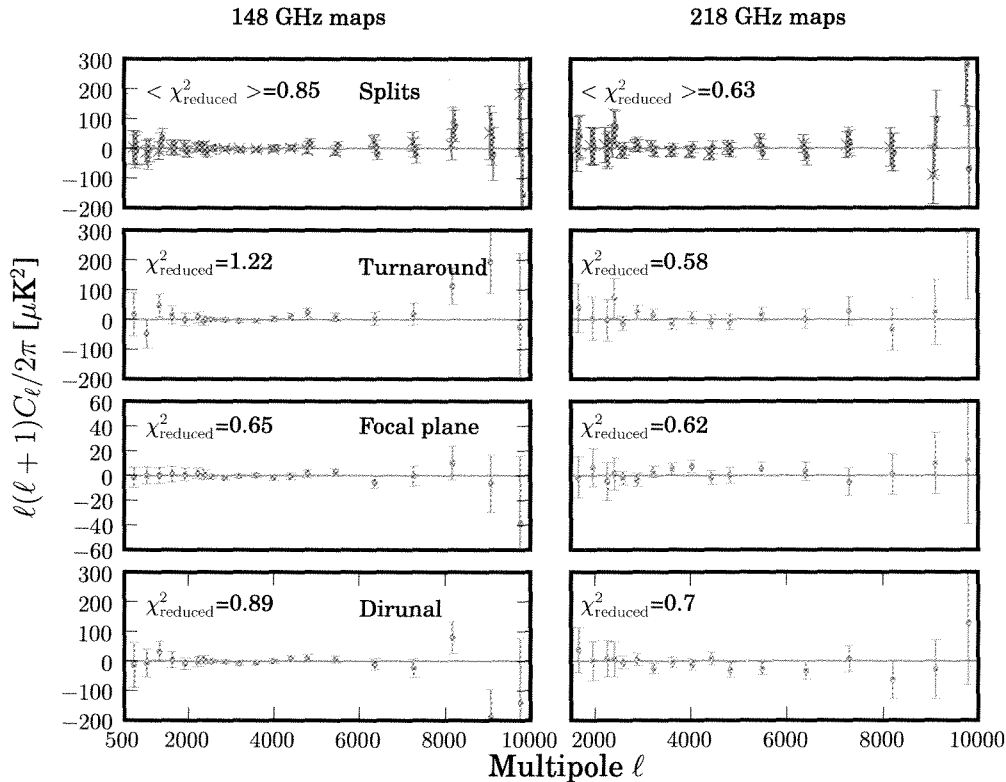


FIG. 8.— Null tests of the ACT 148 (left) and 218 (right) GHz maps. The top row plots the null spectra from difference maps made from the four different time splits described in Section 3.1. The second row illustrates the systematic check on whether data near the telescope turnarounds are contaminated, while the bottom two rows check for a systematic gain difference across halves of the focal plane and for diurnal effects, respectively. The difference spectra and the chi-squares are computed with the same bins as the standard power spectra. For display purposes only, we have re-binned the spectra in this plot with larger bins. The spectra are consistent with null in all cases.

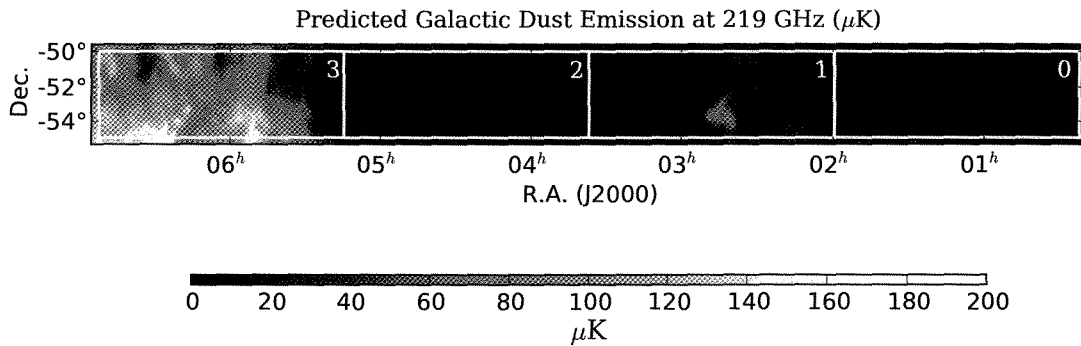


FIG. 9.— Galactic thermal dust emission template at 219 GHz based on Finkbeiner et al. (1999) “model 8”. We cross correlate the template patches (numbered here) with ACT patches to look for a correlated dust signal in ACT maps. Patches 0-3 are located off the Galactic plane with central Galactic latitudes of -64° , -56° , -43° , and -28° , respectively. For the dust emission the effective central frequency in the ACT 218 GHz band is 220 GHz (Swetz et al. 2010).

of 164 point sources, taking the union of the detections at the two frequencies. Before computing the spectrum, we mask out a $10'$ diameter region around each of the 164 point sources. This amounts to masking about 1% of the total area. As discussed in Section 3.2, the prewhitening operation effectively guards against the leakage of power due to the application of this mask.

We estimate the Galactic dust contribution to our

power spectra by cross-correlating ACT maps with the predictions for infrared cirrus emission at CMB frequencies based on the multicomponent dust model “8” of Finkbeiner et al. (1999) (hereafter FDS). The FDS-based dust maps at 148 GHz and 219 GHz are available as a part of the Sehgal et al. (2010b) simulations⁵. We resample these maps into ACT pixelization, and cut out

⁵ These simulations are available at

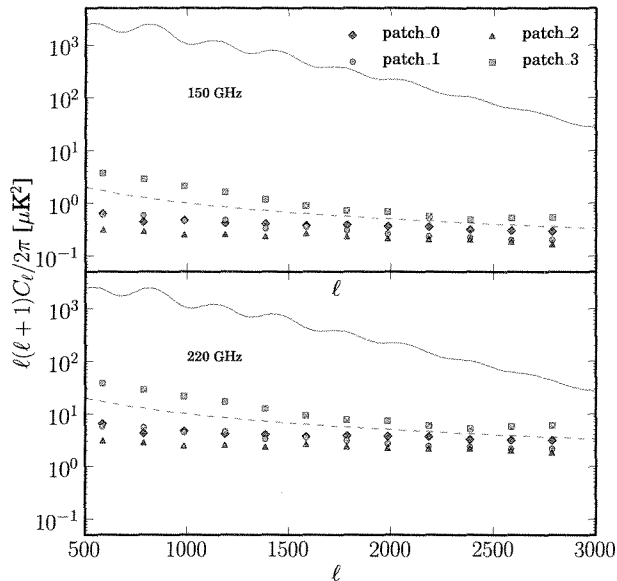


FIG. 10.— Power spectra of dust in the FDS template in the patches shown in Fig. 9 at 150 GHz and 220 GHz in CMB differential temperature units. A theoretical lensed CMB power spectrum is shown for comparison (continuous curve). Also shown is a $C_\ell \propto \ell^{-3}$ curve (dashed) normalized to $\ell^2 C_\ell/2\pi = 2$ (20) μK^2 at $\ell = 500$ for 150 GHz (220 GHz). Note that a temperature power law index of $\beta = 1.7$ between 150 GHz and 220 GHz corresponds to a factor of 10 in the power spectra.

the ACT patch regions. The 219 GHz FDS map is shown in Fig. 9. Patches 0 through 3 labeled in the figure are situated off the Galactic plane with central Galactic latitudes of -64° , -56° , -43° and -28° , respectively. The power spectra of the predicted dust signal in each patch are shown in Fig. 10.

We proceed with the cross-correlation as follows. We express the ACT map as a sum of the cosmic and galactic dust components: $T_{\text{ACT}} = T_{\text{CMB}} + \xi T_{\text{FDS}}$, where ξ is the dust amplitude, predicted to be unity by FDS. We take two approaches to estimate ξ : one based on real-space operations, and the other based on power spectra.

In the real space approach, we prewhiten both the ACT and FDS maps, convolve the ACT maps with an 6'1 FWHM Gaussian beam appropriate to the FDS resolution, apply the ACT point source mask described above, filter out modes below $\ell = 500$ (1500) for 150 GHz (220 GHz) and produce the maps $\mathcal{T}_{\text{ACT}}^i(\hat{\mathbf{n}})$ and $\mathcal{T}_{\text{FDS}}^i(\hat{\mathbf{n}})$ in patch i . Then we estimate ξ in patch i as

$$\xi_{\text{real}}^i = \frac{\int d^2\hat{\mathbf{n}} \mathcal{T}_{\text{ACT}}^i(\hat{\mathbf{n}})\mathcal{T}_{\text{FDS}}^i(\hat{\mathbf{n}})}{\int d^2\hat{\mathbf{n}} \mathcal{T}_{\text{FDS}}^i(\hat{\mathbf{n}})\mathcal{T}_{\text{FDS}}^i(\hat{\mathbf{n}})}. \quad (16)$$

The error on the above estimate is obtained by cross-correlating 960 random CMB plus noise simulations (see Section 4) with the dust maps. In the spectrum-based approach, we estimate ξ in each multipole bin b as the ratio of the ACT-FDS cross-spectrum to the auto-spectrum

http://lambda.gsfc.nasa.gov/toolbox/tb_cmbsim_ov.cfm

To be precise, we use the actual frequencies for which the FDS maps were made, and the effective dust frequency for ACT, but in practice, the differences are negligible.

of the FDS map for the same patch:

$$\xi_b^i = \frac{C_b^{\text{ACT}\times\text{FDS};i}}{C_b^{\text{FDS}\times\text{FDS};i}}. \quad (17)$$

The ACT-FDS cross-spectrum is calculated using both pixel-space weighting and Fourier-space azimuthal weighting, as with the ACT power spectrum, with the FDS and ACT beams deconvolved. The final estimate, ξ_{spec}^i is obtained as the mean across the bins over the range $500 < b < 3000$ ($1500 < b < 3000$) for 150 GHz (220 GHz) where the upper limit is dictated by the FDS resolution. The uncertainty on ξ_{spec}^i is estimated as the scatter in the same quantity computed from the cross-spectra between the FDS maps and 960 random CMB plus noise realizations. As a check for systematics, we circularly shift the FDS patches by one patch, such that patch_0 of ACT is crossed with patch_3 of FDS, patch_1 of ACT with patch_0 of FDS etc., and recompute the real and Fourier space estimates of ξ . The results are shown in Fig. 11. The errors shown arise from treating the FDS template as deterministic.

This figure shows that a dust-correlated signal is observed. More work will be needed to understand how well the FDS template represents faint dust well off the Galactic plane (see, e.g., Veneziani et al. 2010). The fact that such a small signal ($\ell^2 C_\ell/2\pi \sim 0.1 - 1.0 \mu\text{K}^2$ at $\ell = 3000$ and 148 GHz) can be recovered with our map-based technique demonstrates the power of our maximum-likelihood map estimation. Recently Hall et al. (2010) have also reported a measurement of infrared cirrus in the SPT maps.

Subtracting the FDS template from our maps has negligible effect on the power spectra, consistent with the minuteness of the dust signal.

7. LENSING OF THE CMB

The CMB photons are deflected by large-scale structure potentials along their path from the last scattering surface at $z \simeq 1100$ to us (see Lewis & Challinor 2006, for a review). The typical (rms) deflection in the Λ CDM model is about 2'7 and the deflection pattern is coherent over degree scales, comparable to the size of the acoustic features on the primary CMB. These coherent deflections produce distortions of the hot and cold spots on the CMB, leading to a broadening of their size distribution. In the power spectrum, this effect manifests itself as the smoothing of the acoustic peaks, which can be used as a signal to look for lensing. The first attempt at a detection of lensing in the power spectrum was made in Reichardt et al. (2009b) using the data from the ACBAR experiment. They quantified the effect of including lensing in their analysis through the log-ratio of the lensed to unlensed Bayesian evidence ($\Delta \ln \mathcal{E}$), and by combining WMAP5 and ACBAR datasets found $\Delta \ln \mathcal{E} = 2.63$. Calabrese et al. (2008) analyzed the ACBAR data with a different approach, where they introduced a scaling of the power spectrum of the lensing potential ($C_\ell^\Psi \rightarrow A_L C_\ell^\Psi$, with $A_L = 0$ corresponding to no lensing and $A_L = 1$ to the standard Λ CDM expectation). With WMAP5+ACBAR they found $A_L = 3.0^{+0.9}_{-0.9}$ (68% confidence level). According to footnote 17 of Reichardt et al. (2009b), a reanalysis of the ACBAR data using this pa-

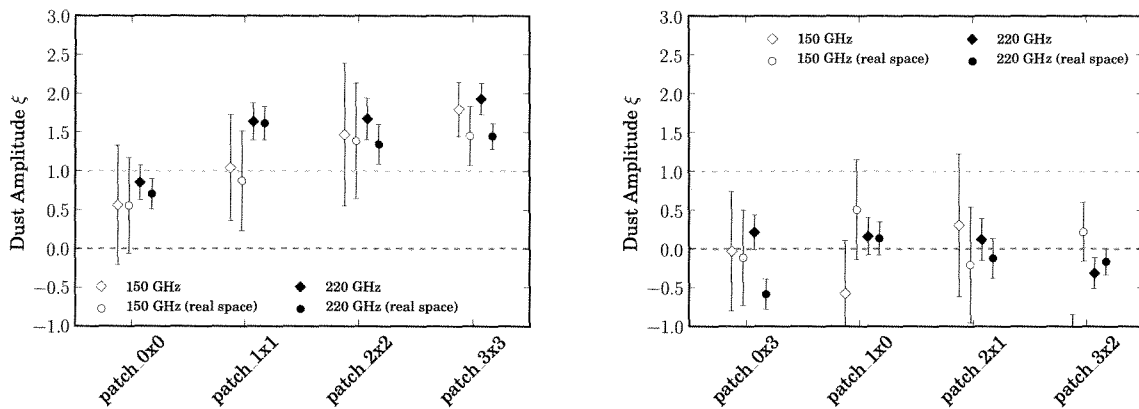


FIG. 11.— *Left*: Amplitude, ξ , of dust emission in ACT patches relative to FDS predictions for 150 GHz (open symbols) and 220 GHz (closed) symbols. Results obtained through the real-space and Fourier-space techniques are shown with circles and diamonds, respectively. *Right*: Results of a null test performed by shuffling around the FDS patches, so that patch 0 of ACT is crossed with patch 3 of FDS etc.

parameterization yields $A_L = 1.60^{+0.55}_{-0.26}$.

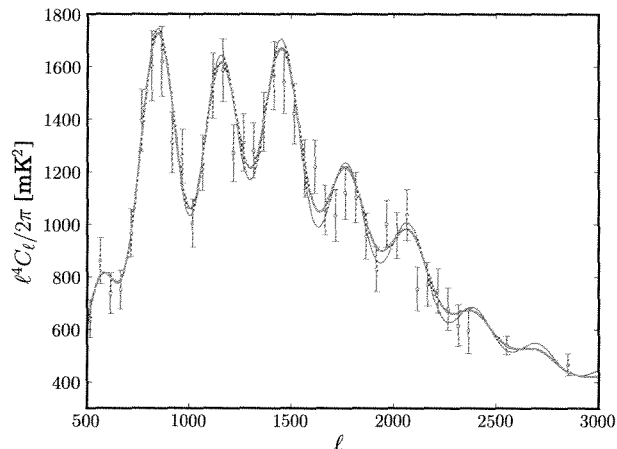


FIG. 12.— Lensing of the CMB smooths out the acoustic peaks in the CMB power spectrum. The best fit model with lensed CMB, secondaries, and point sources is shown as the thick orange curve, while the same with no lensing is shown with the thin green curve.

Our 148 GHz power spectrum is shown against lensed and unlensed models in Fig. 12. We use the parametrization of Calabrese et al. (2008), and using the parameter estimation methodology described in the companion paper Dunkley et al. (2010), we constrain the lensing parameter A_L based on WMAP7 and ACT power spectra. Fig. 13 shows the marginalized 1D likelihood for A_L using WMAP7+ACT. We find $A_L = 1.3^{+0.5(+1.2)}_{-0.5(-1.0)}$ at 68% (95%) confidence, with the best-fit lensed CMB spectrum with $A_L = 1$ having an improved goodness-of-fit to the WMAP7+ACT data of $\Delta\chi^2 = 8$ less than the unlensed model.

We check for systematics that might have given rise to a spurious lensing signal. The projection scheme (cylindrical-equal-area) used for the ACT maps is not particularly optimized for lensing studies — so we test whether this projection could introduce a lensing-like signal. We simulate a low-noise unlensed CMB sig-

nal and run it through the mapmaking pipeline, and try to reconstruct a “lensing convergence” in the resulting map using standard quadratic estimator techniques (Hu & Okamoto 2002). We find the reconstructed convergence power spectrum to be consistent with null (there is a small known bias at high multipoles that we have entirely traced to mode-coupling due to the finiteness of the patch), showing that the projection does not introduce any significant lensing-like signal in our patches. To further test if any other step in our pipeline could produce spurious peak smearing in the power spectrum, we generate an end-to-end simulation of a noisy map exactly as described in Section 4, only this time replacing the lensed CMB signal time-stream with its unlensed version. The resulting maps are then processed through the power spectrum pipeline, and the power spectrum is then analyzed with the parameter estimation method described in Dunkley et al. (2010). We find the lensing amplitude parameter, A_L , described above to be consistent with zero (see Fig. 13).

Here we have used the smearing of the acoustic features to look for the lensing signal. More promising ways of extracting this signal involve using information beyond the two-point correlation (optimal quadratic estimators) and cross-correlation with tracers of large-scale structure. Early efforts cross-correlating the lensing reconstruction in WMAP 3-year maps with luminous red galaxies, radio sources and quasars (Smith et al. 2007; Hirata et al. 2008) yielded $\sim 3\sigma$ evidence for lensing. Data from WMAP are not the best suited for this kind of study because of its limited angular resolution. With arcminute resolution CMB data, such as from ACT, there is a much higher potential for a detection. Such efforts are underway with the ACT maps.

8. DISCUSSION

We have presented a measurement of the power spectrum of the CMB sky observed with the Atacama Cosmology Telescope at 148 GHz and 218 GHz. The 148 GHz spectrum spans a large dynamic range from $\ell = 500$ to $\ell = 10000$, covering the damping tail, where the primary anisotropies with the higher-order acoustic peaks dominate, to the composite high multipole tail of the CMB where emission from dusty galaxies and radio sources and the SZ effect contribute. The second through the seventh

TABLE 1
SINGLE FREQUENCY BANDPOWERS
 $\mathcal{B}_b = \ell_b(\ell_b + 1)C_b/2\pi$ (μK^2)

ℓ range	central ℓ_b	148 GHz		148 GHz \times 218 GHz		218 GHz	
		\mathcal{B}_b	$\sigma(\mathcal{B}_b)$	\mathcal{B}_b	$\sigma(\mathcal{B}_b)$	\mathcal{B}_b	$\sigma(\mathcal{B}_b)$
491 - 540	515	2423.0	270.8	-	-	-	-
541 - 590	565	2701.4	274.8	-	-	-	-
591 - 640	615	1952.1	205.4	-	-	-	-
641 - 690	665	1701.9	167.7	-	-	-	-
691 - 740	715	1895.7	179.5	-	-	-	-
741 - 790	765	2386.5	202.0	-	-	-	-
791 - 840	815	2415.5	200.5	-	-	-	-
841 - 890	865	2168.0	178.2	-	-	-	-
891 - 940	915	1567.5	138.3	-	-	-	-
941 - 990	965	1353.6	109.9	-	-	-	-
991 - 1040	1015	976.0	89.2	-	-	-	-
1041 - 1090	1065	1089.5	92.9	-	-	-	-
1091 - 1140	1115	1229.6	100.2	-	-	-	-
1141 - 1190	1165	1169.7	88.2	-	-	-	-
1191 - 1240	1215	861.9	73.4	-	-	-	-
1241 - 1290	1265	819.3	68.7	-	-	-	-
1291 - 1340	1315	740.8	60.6	-	-	-	-
1341 - 1390	1365	746.8	60.5	-	-	-	-
1391 - 1440	1415	782.4	64.4	-	-	-	-
1441 - 1490	1465	718.8	55.8	-	-	-	-
1491 - 1540	1515	619.3	50.1	559.6	43.6	585.9	87.6
1541 - 1590	1565	495.6	44.5	489.0	40.2	508.0	83.0
1591 - 1640	1615	467.4	38.8	428.3	35.3	480.1	77.9
1641 - 1690	1665	380.4	34.0	382.0	32.7	446.5	74.0
1691 - 1740	1715	351.6	33.6	347.8	32.1	436.8	75.0
1741 - 1790	1765	359.6	32.3	336.7	30.7	470.6	74.1
1791 - 1840	1815	334.4	29.4	336.9	28.3	396.1	65.3
1841 - 1890	1865	274.8	25.3	276.6	25.8	371.5	65.0
1891 - 1940	1915	226.3	23.3	198.8	23.7	199.7	62.1
1941 - 1990	1965	259.7	23.1	254.6	23.9	310.2	61.9
1991 - 2040	2015	235.8	21.6	226.0	22.5	340.2	60.6
2041 - 2090	2065	242.8	22.8	227.7	23.9	283.9	63.5
2091 - 2140	2115	168.8	18.6	206.9	21.2	268.3	57.8
2141 - 2190	2165	164.7	17.8	168.4	19.7	203.4	54.9
2191 - 2240	2215	152.2	17.2	151.0	19.7	269.7	60.5
2241 - 2290	2265	132.2	15.8	114.6	18.1	206.7	56.3
2291 - 2340	2315	114.9	14.7	136.5	18.1	135.8	54.2
2341 - 2390	2365	106.3	15.4	135.3	19.1	222.3	59.2
2391 - 2700	2550	82.9	5.4	99.7	6.7	157.9	21.2
2701 - 3000	2850	57.5	5.3	72.6	6.5	189.1	20.8
3001 - 3400	3200	37.2	4.8	54.0	5.5	158.3	16.9
3401 - 3800	3600	39.8	5.6	50.9	5.9	145.1	16.6
3801 - 4200	4000	48.1	6.8	62.9	6.7	204.5	17.9
4201 - 4600	4400	34.9	8.3	74.6	7.8	198.7	19.7
4601 - 5000	4800	36.4	10.2	98.0	9.3	248.7	22.2
5001 - 5900	5450	55.8	9.6	106.3	8.3	281.7	18.9
5901 - 6800	6350	56.6	15.4	131.7	12.5	378.3	26.8
6801 - 7700	7250	127.8	25.6	127.7	19.1	463.7	38.1
7701 - 8600	8150	115.3	42.4	228.0	30.0	620.6	55.6
8601 - 9500	9050	208.3	73.1	175.9	48.6	812.8	85.0
9501 - 9900	9750	316.9	172.4	152.4	108.1	695.0	176.1

acoustic peaks of the CMB are clearly visible in this spectrum. For the 148×218 GHz and 218×218 GHz spectra, we present measurements from $1500 < \ell < 10000$. The high multipole information from these spectra help us constrain the nature of the point source populations that contribute to this range, and in turn, constrain the SZ contribution to the 148 GHz spectrum. We measure the presence of a very faint dust signal at high Galactic latitudes. Recovering such a faint signal gives us confidence in the fidelity of the maps. We find evidence of gravitational lensing of the CMB in the power spectrum at the 2.8σ level. Constraints on cosmological parameters and the interpretation of high multipole spectra in terms of point source populations and the SZ effect are presented in the companion paper by Dunkley et al. (2010).

ACT is on the Chajnantor Science preserve, which was made possible by the Chilean Comisión Nacional de Investigación Científica y Tecnológica. We are grateful for the assistance we received at various times from the ALMA, APEX, ASTE, CBI/QUIET, and NANTEN2 groups. The PWV data come from the public APEX weather website. Field operations were based at the Don Esteban facility run by Astro-Norte. Reed Plimpton and David Jacobson worked at the telescope during the 2008 season. We thank Norm Jarosik for support throughout the project. This work was supported by the U.S. National Science Foundation through awards AST-0408698 for the ACT project, and PHY-0355328, AST-0707731 and PIRE-0507768. Funding was also provided by Princeton University and the University of Pennsyl-

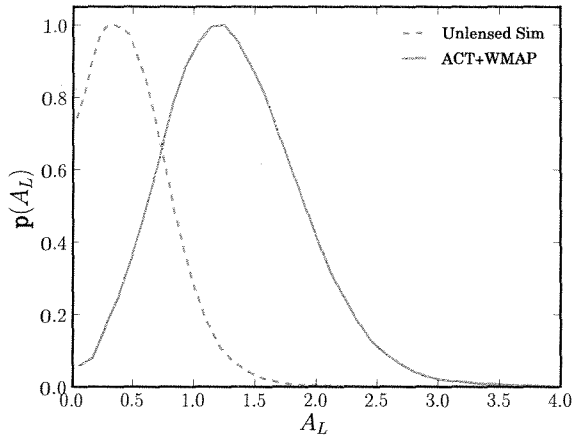


FIG. 13.— One dimensional marginalized distribution (solid line) for the lensing parameter A_L , that scales the expected lensing potential such that $C_\phi \rightarrow A_L C_\phi$ (Calabrese et al. 2008). An unlensed CMB spectrum would have $A_L = 0$, and the standard lensing case has $A_L = 1$. With ACT+WMAP7 we find $A_L = 1.3^{+0.5+1.2}_{-0.5-1.0}$ (68% and 95% limits), a 2.8σ detection of lensing. The dashed line shows the marginalized distribution for A_L , obtained from an unlensed simulation.

vania. The PIRE program made possible exchanges between Chile, South Africa, Spain and the US that enabled this research program.

SD acknowledges support from the Berkeley Center for Cosmological Physics Fellowship. SD would like to thank Christian Reichardt and Oliver Zahn for useful discussions. We thank Bruce Bassett for suggestions on testing lensing in the power spectrum. SD, AH, and TM

were supported through NASA grant NNX08AH30G. JD acknowledges support from an RCUK Fellowship. RH received funding from the Rhodes Trust. ADH received additional support from a Natural Science and Engineering Research Council of Canada (NSERC) PGSD scholarship. AK and BP were partially supported through NSF AST-0546035 and AST-0606975, respectively, for work on ACT. LI acknowledges partial support from FONDAPE Centro de Astrofísica. RD was supported by CONICYT, MECESUP, and Fundación Andes. ES acknowledges support by NSF Physics Frontier Center grant PHY-0114422 to the Kavli Institute of Cosmological Physics. KM, M Hilton, and RW received financial support from the South African National Research Foundation (NRF), the Meraka Institute via funding for the South African Centre for High Performance Computing (CHPC), and the South African Square Kilometer Array (SKA) Project. YTL acknowledges support from the World Premier International Research Center Initiative, MEXT, Japan. NS is supported by the U.S. Department of Energy contract to SLAC no. DE-AC3-76SF00515. We acknowledge the use of the Legacy Archive for Microwave Background Data Analysis (LAMBDA). Support for LAMBDA is provided by the NASA Office of Space Science. The data will be made public through LAMBDA (<http://lambda.gsfc.nasa.gov/>), and the ACT website (<http://www.physics.princeton.edu/act/>).

APPENDIX

CROSS-FREQUENCY CROSS-SPECTRUM ERRORBARS

Here we derive an analytic expression for the expected error bars on the cross-frequency cross-power spectrum, assuming the signal and the noise can be approximated as Gaussian random fields. We do not treat the Poisson component from point sources here, as the Poisson noise correction is negligible at our current noise level. As before, we denote the two frequency channels as A and B and use lowercase romans i, j, k, l , etc. to denote the sub-season data splits.

The estimator for the cross-frequency cross-power spectrum in bin b is given by:

$$\hat{C}_b^{A \times B} = \frac{1}{n_d(n_d - 1)} \sum_{i,j; i \neq j} \hat{C}_b^{(iA \times jB)}, \quad (\text{A1})$$

where n_d is the number of sub-season data splits (four in our case), and

$$\hat{C}_b^{(iA \times jB)} = \frac{1}{\nu_b} \sum_{\ell \in b} T_\ell^{*iA} T_\ell^{jB}, \quad (\text{A2})$$

where the sum is over pixels in the annulus in Fourier-space contained in the bin b , and ν_b is the effective number of pixels in the annulus over which the sum is taken. Evidently,

$$\langle \hat{C}_b^{(iA \times jB)} \rangle = C_b, \quad (\text{A3})$$

$$\langle \hat{C}_b^{A \times B} \rangle = C_b. \quad (\text{A4})$$

First, we will compute the expected covariance between a pair of estimators:

$$\Theta_{bb}^{(iA \times jB); (kA \times lB)} = \left\langle \left(\hat{C}_b^{(iA \times jB)} - \langle \hat{C}_b^{(iA \times jB)} \rangle \right) \left(\hat{C}_b^{(kA \times lB)} - \langle \hat{C}_b^{(kA \times lB)} \rangle \right) \right\rangle$$

where $i \neq j$ and $k \neq l$. From the definition (A2), we can write

$$\Theta_{bb}^{(iA \times jB);(kA \times lB)} = \frac{1}{\nu_b^2} \sum_{\ell \in b} \sum_{\ell' \in b} \left[\langle T_{\ell}^{*iA} T_{\ell}^{jB} T_{\ell'}^{*kA} T_{\ell'}^{lB} \rangle \right] - \langle C_b^{(iA \times jB)} \rangle \langle C_b^{(kA \times lB)} \rangle. \quad (\text{A5})$$

Applying Wick's Theorem to the above formula, we get,

$$\begin{aligned} \Theta_{bb}^{(iA \times jB);(kA \times lB)} &= \frac{1}{\nu_b^2} \left[\nu_b^2 \langle C_b^{(iA \times jB)} \rangle \langle C_b^{(kA \times lB)} \rangle + \nu_b \langle C_b^{(iA \times kA)} \rangle \langle C_b^{(jB \times lB)} \rangle + \nu_b \langle C_b^{(iA \times lB)} \rangle \langle C_b^{(kA \times jB)} \rangle \right] \\ &\quad - \langle C_b^{(iA \times jB)} \rangle \langle C_b^{(kA \times lB)} \rangle \end{aligned} \quad (\text{A6})$$

$$= \frac{1}{\nu_b} \left[\langle C_b^{(iA \times kA)} \rangle \langle C_b^{(jB \times lB)} \rangle + \langle C_b^{(iA \times lB)} \rangle \langle C_b^{(kA \times jB)} \rangle \right]. \quad (\text{A7})$$

Now we will examine various cases of the above variance.

Type 1: All four splits different, $i \neq j \neq k \neq l$. In this case, all four ensemble averages are cross-power spectra and evaluate to the underlying spectrum C_b , giving

$$\Theta_{bb}^{(iA \times jB);(kA \times lB); \text{Type 1}} = \frac{2}{\nu_b} C_b^2. \quad (\text{A8})$$

Type 2a: One split in common, of channel A, ($i = k \neq l \neq j$). Here, the first ensemble average in (A7) is an auto-frequency auto-power spectrum which evaluates out to $(C_b + N_b^{AA})$, where N_b^{AA} is the noise in the channel A data split. This gives

$$\Theta_{bb}^{(iA \times jB);(kA \times lB); \text{Type 2a}} = \frac{1}{\nu_b} (2C_b^2 + C_b N_b^{AA}). \quad (\text{A9})$$

Type 2b: One split in common, of channel B, ($l = j \neq k \neq i$).

$$\Theta_{bb}^{(iA \times jB);(kA \times lB); \text{Type 2b}} = \frac{1}{\nu_b} (2C_b^2 + C_b N_b^{BB}). \quad (\text{A10})$$

Type 2c: One split in common, one of channel A, the other of channel B ($i = l$ or $k = j$).

$$\Theta_{bb}^{(iA \times jB);(kA \times lB); \text{Type 2c}} = \frac{1}{\nu_b} (2C_b^2 + C_b N_b^{AB}), \quad (\text{A11})$$

where N_b^{AB} represents any noise that is correlated across the two frequency channels.

Type 3a: Two splits in common ($i = k, j = l$). In this case, we have

$$\Theta_{bb}^{(iA \times jB);(kA \times lB); \text{Type 3a}} = \frac{1}{\nu_b} [2C_b^2 + C_b(N_b^{AA} + N_b^{BB}) + N_b^{AA} N_b^{BB}]. \quad (\text{A12})$$

Type 3b: Two splits in common ($i = l, j = k$). This evaluates to

$$\Theta_{bb}^{(iA \times jB);(kA \times lB); \text{Type 3b}} = \frac{1}{\nu_b} [2C_b^2 + 2C_b N_b^{AB} + N_b^{AB} N_b^{AB}]. \quad (\text{A13})$$

Next, we turn our attention to the variance of the mean spectrum estimator, $\hat{C}_b^{A \times B}$, defined in (A1):

$$\Theta_{\text{mean}; \text{mean}} = \left\langle \left(\hat{C}_b^{A \times B} - \langle \hat{C}_b^{A \times B} \rangle \right) \left(\hat{C}_b^{A \times B} - \langle \hat{C}_b^{A \times B} \rangle \right) \right\rangle. \quad (\text{A14})$$

Expanding one of the terms out, we have

$$\Theta_{\text{mean}; \text{mean}} = \frac{1}{n_d(n_d - 1)} \sum_{i,j; i \neq j} \left\langle \left(\hat{C}_b^{(iA \times jB)} - \langle \hat{C}_b^{(iA \times jB)} \rangle \right) \left(\hat{C}_b^{A \times B} - \langle \hat{C}_b^{A \times B} \rangle \right) \right\rangle. \quad (\text{A15})$$

Each of the $n_d(n_d - 1)$ terms in the sum is statistically equivalent to the other and should evaluate to be the same. Hence,

$$\Theta_{\text{mean}; \text{mean}} = \Theta^{(iA \times jB); \text{mean}}, \quad (\text{A16})$$

which can be expanded out as

$$\Theta_{\text{mean}; \text{mean}} = \frac{1}{n_d(n_d - 1)} \sum_{k,l; k \neq l} \Theta^{(iA \times jB);(kA \times lB)}. \quad (\text{A17})$$

The last piece of the calculation is to figure out how many times each of the types defined above appear in this sum. Since i and j are fixed, it is apparent that for Type 1, k can be chosen in $(n_d - 2)$ ways, and l in $(n_d - 3)$ ways. So, Type 1 appears $(n_d - 2)(n_d - 3)$ times. For Type 2a k is fixed by equality with i , so l can be chosen only in $(n_d - 2)$ ways. Therefore, Type 2a appears $(n_d - 2)$ times. From symmetry, Type 2b also appears $(n_d - 2)$ times. For Type 2c, either $i = l$ or $k = j$ which gives $2(n_d - 2)$ terms. Finally, for each of Type 3a and 3b, all four indices are completely fixed, so they appear only once. Putting these together, we have:

$$\begin{aligned} \Theta^{\text{mean;mean}} &= \frac{1}{n_d(n_d - 1)} \left[(n_d - 2)(n_d - 3) \frac{2}{\nu_b} C_b^2 \right. \\ &\quad + (n_d - 2) \frac{1}{\nu_b} (2C_b^2 + C_b N_b^{AA}) \\ &\quad + (n_d - 2) \frac{1}{\nu_b} (2C_b^2 + C_b N_b^{BB}) \\ &\quad + 2(n_d - 2) \frac{1}{\nu_b} (2C_b^2 + C_b N_b^{AB}) \\ &\quad + \frac{1}{\nu_b} \{ 2C_b^2 + C_b(N_b^{AA} + N_b^{BB}) + N_b^{AA} N_b^{BB} \} \\ &\quad \left. + \frac{1}{\nu_b} \{ 2C_b^2 + 2C_b N_b^{AB} + N_b^{AB} N_b^{AB} \} \right] \\ &= \frac{1}{\nu_b} \frac{1}{n_d(n_d - 1)} \left[n_d(n_d - 1) 2C_b^2 + (n_d - 1) C_b (N_b^{AA} + N_b^{BB}) \right. \\ &\quad \left. + 2(n_d - 1) C_b N_b^{AB} + (N_b^{AA} N_b^{BB} + N_b^{AB} N_b^{AB}) \right] \end{aligned}$$

which simplifies to

$$\Theta^{\text{mean;mean}} = \frac{1}{\nu_b} \left[2C_b^2 + \frac{C_b}{n_d} (N_b^{AA} + N_b^{BB}) + \frac{2}{n_d} C_b N_b^{AB} + \frac{(N_b^{AA} N_b^{BB} + N_b^{AB} N_b^{AB})}{n_d(n_d - 1)} \right]. \quad (\text{A18})$$

Note that if there is only one channel $A \equiv B$, this reduces to the familiar form (see Eq. (9) in Fowler et al. 2010):

$$\Theta^{\text{mean;mean}} = \frac{1}{\nu_b} \left[2C_b^2 + 4 \frac{C_b}{n_d} N_b^{AA} + 2 \frac{(N_b^{AA})^2}{n_d(n_d - 1)} \right]. \quad (\text{A19})$$

Finally, if we have n_p patches with equal noise, the above variance should be divided by n_p . In practice, a weighted combination of the individual patch variances is used. Under the assumption of isotropic noise and filtering, the effect of the data window, W , can be taken into account by correcting the number of modes as $\nu_b \rightarrow \nu_b w_2^2 / w_4$, where w_n represents the n -th moment of the window. There is a small additional correction due the anisotropic nature of the noise and filtering, which we calibrate against Monte Carlo simulations.

For estimating parameters from the three spectra, $C_b^{148 \times 148}$, $C_b^{148 \times 218}$ and $C_b^{218 \times 218}$ a joint likelihood function is written as

$$-2 \ln L = \begin{pmatrix} C_b^{148 \times 148} - C_b^{148 \times 148, \text{theory}} \\ C_b^{148 \times 218} - C_b^{148 \times 218, \text{theory}} \\ C_b^{218 \times 218} - C_b^{218 \times 218, \text{theory}} \end{pmatrix}^T \Sigma^{-1} \begin{pmatrix} C_b^{148 \times 148} - C_b^{148 \times 148, \text{theory}} \\ C_b^{148 \times 218} - C_b^{148 \times 218, \text{theory}} \\ C_b^{218 \times 218} - C_b^{218 \times 218, \text{theory}} \end{pmatrix}, \quad (\text{A20})$$

where

$$\Sigma = \begin{pmatrix} \Sigma_b^{148-148, 148-148} & \Sigma_b^{148-148, 148-218} & \Sigma_b^{148-148, 218-218} \\ \Sigma_b^{148-218, 148-148} & \Sigma_b^{148-218, 148-218} & \Sigma_b^{148-218, 218-218} \\ \Sigma_b^{218-218, 148-148} & \Sigma_b^{218-218, 148-218} & \Sigma_b^{218-218, 218-218} \end{pmatrix} \quad (\text{A21})$$

with

$$\Sigma_b^{A-B, C-D} = \left\langle \left(\widehat{C}_b^{A \times B} - \langle \widehat{C}_b^{A \times B} \rangle \right) \left(\widehat{C}_b^{C \times D} - \langle \widehat{C}_b^{C \times D} \rangle \right) \right\rangle, \quad (\text{A22})$$

and can be computed in a similar fashion as above.

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