1	Accepted version: Biological Conservation
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3	Quantifying imperfect detection in an invasive pest fish and the implications for
4	conservation management
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15	Running title: Imperfect detection of an invasive fish
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29 In managing non-native species, surveillance programmes aim to minimise the 30 opportunity for invasions to develop from initial introductions through early detection. 31 However, this is dependent on surveillance methods being able to detect species at low levels of abundance to avoid false-negative recordings through imperfect 32 33 detection. We investigated through field experimentation the ability to detect 34 Pseudorasbora parva, a highly invasive pest fish in Europe, in relation to their known density and sampling method. Secure pond mesocosms of area 100  $m^2$  contained P. 35 parva densities from 0.02 to 5.0  $\text{m}^{-2}$ ; each density was in triplicate. These were 36 37 searched using point sampling electric fishing and deployment of fish traps (nonbaited and baited). No fish were captured at densities  $< 0.5 \text{ m}^{-2}$  using any method and 38 39 this was considered their detection threshold. Point sample electric fishing was the 40 least effective detection method, producing high proportions of false-negative data 41 even at high fish densities. Baited traps were the most effective detection method. Probability of detection of *P. parva* was 1.0 for baited traps at all densities  $> 0.5 \text{ m}^{-2}$ , 42 whereas for electric fishing it only exceeded 0.95 at 5.0  $\text{m}^{-2}$  using high searching 43 44 effort. These data reveal that small pest fishes such as *Pseudorasbora parva* may be 45 prone to imperfect detection when at low densities and this is consistent with a 46 number of other invasive species. This indicates the importance of designing 47 surveillance programmes using methods of known statistical power to optimise 48 conservation resource expenditure and enhance management outcomes.

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50 Keywords: Non-native species; detection threshold; electric fishing; false-negative;
51 probability of detection.

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54 Biological invasions are a conservation issue that have the potential to negatively 55 impact biodiversity and raise global concern over biotic homogenisation (McKinney 56 and Lockwood, 1999). Consequently, a common goal of the conservation 57 management of biological invasions is removing invasive 'pest' species and then keeping areas pest-free (Moore et al., 2010). To be effective, adequate resources 58 59 require to be apportioned to surveillance in order to increase the opportunity for new 60 introductions to be detected (Moore et al., 2010; Britton et al., 2010a). Early detection 61 then enables management actions to be taken in the incipient phases of invasion that 62 inhibit establishment and minimise dispersal rates that should impede or even prevent 63 invasion (Hulme, 2006; Christy et al., 2010; Willson et al., 2010). However, this is 64 reliant on the surveillance methods being capable of capturing the species when they 65 are in low abundance, with 'imperfect detection' referring to situations when 66 introduced individuals have not been able to be detected (Rout et al., 2009a,b). Imperfect detection also inhibits the evaluation of eradication operations; in these 67 68 situations, 'false-negative' data must be minimised in order to reduce the chance of 69 errors occurring in its evaluation (Simberloff, 2003; Rout 2009; Rout et al., 2009a,b; 70 Delaney and Leung, 2010).

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Knowledge on the chance of imperfect detection occurring during either surveillance or eradication evaluation is enhanced when the detection probability of the species and sampling methods are quantified. Low probability of detection of invasive species has been recognised across a number of taxonomic groups, including mustelids (King et al., 2009), snakes (Willson et al., 2010) and plants (Rout, 2009),

and has enabled probability models to be developed (Harvey et al., 2009; Rout et al.,
2009a,b; Christy et al., 2010). The utility of these models is they enable the
development and optimisation of search strategies (Cacho et al., 2006), reduce the
chance of false-negative data being collected (Rout et al., 2009a), and enable better
prioritisation of conservation resources within management programmes (Delaney
and Leung, 2010).

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84 Imperfect detection can be a general issue when surveying fish populations due to, 85 for example, inefficiencies that arise from issues of fish size and water depth that 86 inhibit capture (Janac and Jurajda, 2005; Cowx et al., 2001; Copp et al., 2010). 87 Quantification of imperfect detection has not, however, been applied to the 88 conservation management of non-native fishes, despite the potential of their invasions 89 to cause substantial negative impacts in the environment (Gozlan et al., 2010a). These 90 invasion issues are well demonstrated by the Asian cyprinid fish topmouth gudgeon 91 Pseudorasbora parva (Temminck and Schlegel). Highly invasive in Europe since the 92 1960s, they are now present in at least 32 countries (Gozlan et al., 2010b). 93 Introductions generally occur through inter-country movements of fish in the 94 aquaculture trade, with their release into the wild via small, outdoor aquaculture 95 ponds that are connected to open waters (Britton et al., 2007, 2008, 2010b). In 96 colonised waters, they are generally considered pests due to their numerical 97 dominance and small body sizes (25 to 90 mm fork lengths; Pinder et al., 2005; 98 Britton et al., 2007; Gozlan et al., 2010b). Ecological impacts include their sharing of 99 trophic space with native fishes resulting in decreased growth (Britton et al., 2010c) 100 and the transmission of a novel pathogen (Gozlan et al., 2009). Whilst considerable 101 effort and resource has been exerted in some European countries to control their invasion (Britton et al., 2008; Britton et al., 2010b), the evaluation of eradication
operations is inhibited by a paucity of knowledge on their probability of detection (i.e.
the opportunity for collecting false-negative data). Moreover, there are no known
active surveillance programmes for this fish at the present time, despite extensive
work on their dispersal and associated impacts (Gozlan et al., 2010b).

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108 Consequently, the aim of this work was to quantify the imperfect detection of 109 introduced P. parva in experimental pond systems in relation to known population 110 density, sampling method and search effort. This was achieved through comparing the 111 efficacy of different sampling approaches in detecting their presence in order to 112 produce an unbiased estimate of their probability of detection. This was completed in 113 replicated mesocosm systems through experimentally manipulating the sampling 114 method, search effort and fish densities. A model was then produced to estimate the 115 probability of detecting an individual P. parva at different densities and levels of 116 sampling effort. Given the paucity of data on the probability of detection of invasive 117 pest fish generally then these outputs will have management applications to other pest 118 fishes. They should also provide an important conservation case study in the utility of 119 applying imperfect detection to improve management programmes of invasive 120 species.

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# **2. Materials and Methods**

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124 2.1 Experimental design and search methodology

125 A series of secure, replicated semi-natural mesocosm ponds of  $100 \text{ m}^2$  and maximum 126 depth 1.5 m were set up in June 2010. These were designed to broadly represent the 127 outdoor aquaculture ponds that P. parva are typically introduced to when transferred in the aquaculture industry (cf. Section 1). These mesocosms were located on a bio-128 129 secure aquaculture site where *P. parva* were already present and established. In each mesocosm, the initial step was to place a 4  $m^2$  keep cage containing 5 live *P. parva* 130 131 for 24 hours. As all fish survived, these cages were removed and P. parva of fork lengths 60 to 80 mm introduced into the mesocosms at densities of 0.02, 0.1, 0.5, 1.0, 132 1.7. 2.7 and 5.0  $m^{-2}$ . Each density was replicated 3 times. The mesocosms were then 133 134 left for 14 days to allow the *P. parva* to acclimatise to the conditions before searches 135 commenced.

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137 Surveillance of the mesocosms for completing the searches of P. parva was 138 completed in July and August 2010. The searches commenced using non-quantitative 139 point sampling electric fishing (Copp, 2010) using a Smith Root LR24 backpack 140 fisher operating at approximately 0.5-A pulsed DC, where the conductivity of each 141 pond was in the region of 350 to 400 µs. Electric fishing was preferred to the use of a 142 micromesh seine net (Cowx et al. 2001) as submerged vegetation would have 143 prevented the net's effective use. Point sampling was preferred to continuous electric 144 fishing as this enabled less-disruptive sampling within the different habitats of the mesocosms (e.g., in- and outside of macrophyte cover) and enabled strong 145 146 quantification of fishing effort. Moreover, electric fishing in this manner has been 147 deployed to detect *P. parva* in the evaluation of their eradication operations (Britton et 148 al., 2010b). Each mesocosm was sampled once per week over a 3 week period. On 149 each occasion, the mesocosms were all sampled for the detection/ non-detection of P. 150 parva through the electric fishing of 30 randomly selected point samples where each 151 point was fished for a standard period of 10 s (Copp, 2010). Detection/ non-detection

of *P. parva* was recorded according to whether any individuals were observed within the electric field. At the conclusion of sampling each mesocosm on each occasion, turbidity was assessed using a Secchi disk (nearest cm), surface weed cover estimated (to the nearest 5 %) and the time of day recorded. Throughout the sampling period, the water temperature in each mesocosm was recorded every 30 minutes to the nearest 0.1 °C using a temperature logger.

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159 Following completion of the electric fishing, each mesocosm was sampled using a 160 rectangular fish trap with a circle alloy frame of length 107 cm, width and height 27.5 161 cm, mesh diameter 2 mm and with funnel shaped holes of 6.5 cm diameter at either 162 end to allow fish entry and hence their capture. The traps were set either as non-163 baited, i.e. no attractant was added to the trap, or baited, where 5 fishmeal pellets of 164 21 mm diameter were placed in the trap as an attractant (Dynamite Baits 2010). The 165 traps were fished on 6 occasions, with 3 days between each fishing occasion, with 1 166 trap set in each mesocosm for 1 hour. The sampling schedule of the mesocosms was 167 designed randomly to minimise bias, but was set up to ensure each mesocosm was 168 sampled on 6 occasions covering sampling using 3 non-baited and 3 baited traps. 169 Recording of the detection/ non-detection of P. parva in each mesocosm was 170 determined by their presence/ absence in the trap at the conclusion of the hour. To 171 minimise handling of the fish, they were able to be released without the traps being 172 removed fully from the water. The exception was on the final sampling occasion 173 when the traps were removed from the water and the captured fish counted to enable 174 trap catch per unit effort to be expressed as the number of captured fish per trap hour  $(n \operatorname{trap}^{1} h^{1}).$ 175

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# 177 2.2 Statistical testing and probability of detection

178 The initial step in testing the search data was to determine the detection threshold for 179 each sampling method, expressed as the minimum density at which at least 1 P. parva 180 was detected by the method on at least 1 occasion. Comparisons were then made to 181 identify the effectiveness of each method to provide a presence/ absence search 182 methodology for *P. parva* and then their ability to provide measures of relative 183 abundance. This latter comparison was completed by comparing the number of 184 sampled fish with their known density and testing using regression methods. Multiple 185 regression was then used to identify the relative effects of fish density and the other 186 measured variables (weed cover (%), turbidity, time of day and water temperature) on 187 catch per unit effort. Their effects were compared using their standardized beta 188 coefficients ( $\beta$ ) and their significance; those variables with the largest  $\beta$  values made 189 the strongest singular contribution to explaining the relative abundance (the dependent 190 variable) when all the other model variables were controlled.

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192 As per Delaney and Leung (2010), logistic regression was used to test for a 193 relationship between the probability of detection of at least one individual in a 194 mesocosm its POD measured as the binary yes (detection) or no (non-detection), 195 against the number of point samples (up to the maximum of 30) or trap hours required 196 to detect that fish (S) and the density of P. parva (n), where a, b and c were the 197 regression coefficients, and where the compliment of POD is the probability of a false-negative: POD =  $e^{(a+bn+cS)} / 1 + e^{(a+bn+cS)}$  (Equation 1; Delaney and Leung, 2010). 198 199 From this model, the sampling intensity required to detect defined *P. parva* densities 200 at given PODs was determined and displayed using a contour plot.

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The detection threshold for *P. parva* for all sampling methods was  $0.5 \text{ m}^{-2}$ , i.e. no fish were detected at densities of 0.02 and 0.1 m<sup>-2</sup> and so were considered as falsenegative data (Fig. 1). Electric fishing and non-baited traps continued to produce a proportion of samples that were false-negative even at densities of 5 m<sup>-2</sup> whereas baited traps did not produce any false-negative data at densities > 0.5 m<sup>-2</sup> (Fig. 1).

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At densities above 0.5  $\text{m}^{-2}$ , electric fishing was only able to estimate *P. parva* 210 211 apparent presence/ absence; even in mesocosms of higher *P. parva* density, detection 212 generally involved observing a single individual in the electric field. Overall, catches 213 of P. parva were significantly lower in non-baited traps than baited traps at all densities where fish were captured (ANOVA  $F_{1,5} = 298.1$ ; P < 0.01; Fig. 2). Whilst 214 215 the non-baited fish traps did provide a significant relationship between P. parva 216 density and relative abundance (Fig. 2), this relationship for the baited fish traps was highly significant (Fig. 2). The calibration equation for determining Pseudorasbora 217 *parva* density (*n*) from catch per unit effort values (c) in the baited traps was  $n = (c \times a)^{n}$ 218 219 (0.027) + (0.0106). In the non-baited and baited traps, multiple regression analysis 220 revealed that *P. parva* density was the only significant variable in explaining the 221 variation in the catch per unit effort data (Table 1). Turbidity (Secchi disk depth range 222 0.3 to 0.6 m), weed cover (range 25 to 55 %), water temperature (range 17.1 to 17.6 223 <sup>o</sup>C) and the time of day the samples were taken (range 08.30 to 16.30) had no 224 significant effects (Table 1). For baited traps, the variables in the model explained most of the variation in the catch per unit effort data ( $R^2 = 0.89$ ; Table 1); whilst this 225

was reduced in the non-baited traps ( $R^2 = 0.64$ ), this may be related to their weaker relationship between fish density and catch per unit effort (Fig. 2).

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229 The multiple logistic regressions revealed that both P. parva density (n) and the 230 number of electric fishing point samples (S) had a significant effect on POD; this was 231 also similar for the non-baited fish traps (Table 2). For the baited fish traps, the trap 232 hours required for detection (S) was not significant (P > 0.05; Table 2) as at densities 233  $> 0.5 \text{ m}^{-2}$ , the setting of 1 baited trap for 1 hour resulted in positive detection. Use of 234 the regression coefficients in Equation 1 enabled the probability of detection to be 235 determined as a function of sampling effort (number of point samples/ trap hours 236 required for detection) and P. parva density for each sampling method. For baited traps, the probability of detection using a single trap for 1 hour at all densities > 0.5237  $m^{-2}$  (their detection threshold) was 1.0. For non-baited traps, the probability of 238 detecting *P. parva* at 0.5 m<sup>-2</sup> was 0.78 in 1 trap hour, 0.94 in 2 trap hours and 0.99 in 3 239 trap hours. At densities > 0.5 m<sup>-2</sup>, POD was > 0.95 in 1 non-baited trap hour. The 240 241 POD of point sampling electric fishing was more complex, as revealed by the contour 242 plot produced from the model coefficients (Table 2) that plotted POD as a function of 243 point sample number and *P. parva* density (Fig. 3). Probability of detection only exceeded 0.95 at 5.0 m<sup>-2</sup> when at least 30 point samples were taken. At a density of 244  $0.5 \text{ m}^{-2}$ , POD only exceeded 0.95 when 54 point samples were taken. 245

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## 247 **4. Discussion**

It has been argued that there is a local and global requirement to address biological invasions and this will be assisted by enhanced surveillance methods that specifically

251 target the detection of newly introduced species (McKinney and Lockwood, 1999; 252 Hulme, 2006). Indeed, efforts to limit incursions by newly introduced species tend to 253 be preferable to conducting eradication operations over large spatial areas through 254 reduced expense and increased opportunity for success (Jarrad et al., 2010). Thus, 255 quantifying the probability of detection according to methodology, search effort and 256 the abundance of the target species is a fundamental step in identifying the optimal 257 design of a surveillance programme. Moreover, the increasing importance of 258 designing surveillance methodologies that incorporate known statistical power is 259 demonstrated in Australia where the Western Australian Government has imposed a 260 condition that surveillance programmes must include a specified statistical power in 261 order to detect newly introduced species (80 %; Jarrad et al., 2010). Consequently, the 262 conservation relevance of the outputs of this study are three-fold: (i) it should enable 263 the optimal design of surveillance and eradication evaluation surveys for invasive P. 264 parva and similar small bodied, invasive fishes using methodologies of quantified 265 statistical power; (ii) it has experimentally quantified the statistical power of detection 266 for a highly invasive species when other studies have had to rely on either simulations 267 (e.g. Tyre et al., 2003) or field observations (e.g. King et al., 2009); and (iii) it reveals 268 in a different taxonomic group to similar studies (e.g. Harvey et al., 2009; King et al., 269 2009; Rout, 2009; Willson et al., 2010) that the issue of imperfect detection remains 270 inherent within invasive species management and so has to be firmly embedded 271 within surveillance and eradication programmes.

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In designing surveillance strategies for new and extant populations of invasive species, the collection of presence/ absence data only (rather than abundance) is increasing (Delaney and Leung, 2010). This is because it potentially provides greater

276 spatial coverage for the amount of resource applied and statistical approaches for 277 testing these data are becoming increasingly powerful (Tyre et al., 2003; Wintle et al., 278 2005; Delaney and Leung, 2010). It is, however, only relatively recently that 279 searchers have started account for the possibility of imperfect detection in surveillance 280 (Harvey et al., 2009), despite their serious implications in the (mis)interpretation of 281 search data (Rout et al., 2009a,b; Delaney and Leung, 2010). In the mesocosms used 282 here, false-negative data were always collected when fish density was below  $0.5 \text{ m}^{-2}$ . 283 This may have been a symptom of low random encounter probabilities due to the 284 limited number of replicated mesocosms and the low number of fish therein. However, subsequent work in other mesocosms of  $< 0.5 \text{ m}^{-2}$  using baited traps 285 286 deployed for 24 hours also produced false-negative data, despite studies in other 287 animals showing the duration of trap deployment increases the opportunity of capture (e.g. Gust and Inglis, 2006). Consequently, these data do suggest that newly 288 289 introduced *P. parva* are unlikely to be detected in their incipient phases and may only 290 be detected at higher densities when natural dispersal may have already occurred and 291 the opportunity for taking effective management is constrained (Simberloff, 2003; 292 Hulme, 2006; Gozlan et al., 2010b). This does represent a serious management issue 293 and highlights the requirement for increased quarantine procedures that aim to prevent 294 introductions rather than for a new introduction to be detected. Indeed, such data 295 should be used to assist the determination of the optimal allocation of resources 296 between quarantine and surveillance through providing data on search efficiency and 297 its subsequent cost (Moore et al., 2010).

298

The choice of surveillance method for detecting new introductions remains a key component of determining whether the species will be captured or if a false-negative 301 result is achieved. Whilst we demonstrated that baited traps were the most effective 302 method to detect *P. parva* in ponds, their utility in rivers of moderate flow or above 303 may be lower and electric fishing may be more appropriate, particularly across large 304 spatial areas. Given the issues of reduced probability of detection of *P. parva* when 305 using point abundance electric fishing, then alternative methods may require 306 exploring that were not covered by this study, such as use of more continuous electric 307 fishing over more extended periods, or use of micro-mesh seine nets in areas of 308 preferred P. parva habitat (cf. Beyer et al., 2007). Indeed, whilst the search strategy 309 deployed in this study was geared around identifying the most appropriate rapid 310 detection tool for lentic *P. parva* populations, search effort remains a key component 311 within surveillance and increased effort by electric fishing (point-sampling or 312 continuous) may have provided increased probability of detection. Notwithstanding, 313 inefficiency of capture remains commonplace for many invasive animals through 314 issues such as trap avoidance, low detection opportunity due to low numbers of 315 individuals across large spatial areas and use of cryptic habitats inhibiting capture (Dorcas and Willson, 2009; King et al., 2009; Willson et al., 2010). Consequently, 316 317 this has to be recognised in their management and accounted for whenever possible.

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In summary, this study quantified the imperfect detection of *P. parva* in small pond systems, similar to those ponds used in aquaculture where initial introductions of the species into new spatial areas often occur. It revealed that management programmes that rely heavily on detecting invasive species at low densities may be inhibited by imperfect detection. It suggests that the issue of collecting false-negative data in surveillance and eradication evaluation programmes is as apparent in a small invasive pest fish as with many other invasive taxa, demonstrating that quantifying the

statistical power of detection methods should substantially improve the design ofsearch strategies and so enhance long-term conservation outcomes.

328

#### 329 Acknowledgements

330

The work was supported by the Natural Environment Research Council (NERC
research grant reference number NE/H000429/1).

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Table 1. Multiple regression output of catch per unit effort of (a) non-baited and (b) baited traps versus the variables of *Pseudorasbora parva* density, weed cover (expressed as % of surface area), turbidity (measured as Secchi disk depth), time of day and water temperature that the sample was taken. See Section 3 for the data range of each variable between the mesoscosms.

(a) Non-baited traps

Overall model: $R^2 = 0.64$ ; $F_{5,30} = 10.47$ ; $p < 0.01$					
Variable	$\beta$ (standardised)	Р			
Fish density	0.83	< 0.01			
Weed cover	0.56	0.21			
Turbidity (as Secchi disk depth, cm)	0.35	0.43			
Time of day	0.34	0.54			
Water temperature	0.01	0.74			

(b) Baited traps

Overall model:  $R^2 = 0.89$ ;  $F_{5,30} = 24.35$ ; p < 0.01

Variable	β (standardised)	Р
Fish density	0.69	< 0.01
Weed cover	-0.41	0.24
Turbidity (as Secchi disk depth, cm)	0.15	0.64
Time of day	0.13	0.14
Water temperature	0.11	0.18

Table 2. Multiple logistic regression coefficients used in equation 1, and their statistical significance, for the sampling approaches of point abundance electric fishing, non-baited fish traps and baited fish traps.

Sampling method	Parameter	Symbol in	Coefficient	Standard	Р
		equation 1		error	
Point sampling	Constant	a	3.19	0.85	< 0.03
electric fishing	Fish density	b	0.56	0.15	< 0.02
	Point samples	с	0.11	0.03	< 0.01
Non-baited fish	Constant	a	21.25	7.54	< 0.05
traps	Fish density	b	42.12	19.59	< 0.01
	Trap number	с	1.54	0.54	< 0.04
Baited fish traps	Constant	a	15.49	4.21	< 0.03
	Fish density	b	55.12	22.14	< 0.01
	Trap number	с	2.10	1.92	> 0.05

# **Figure captions**

Figure 1. Relationship of detection success (proportion of sampling occasions that resulted in the detection of *Pseudorasbora parva*) and *Pseudorasbora parva* density by sampling methods, where:  $\times$  point sample electric fishing, • non-baited fish traps, and o baited fish traps. Error bars represent 95 % confidence limits.

Figure 2. Relationship of mean catch per unit effort of non-baited (•) and baited (o) fish traps for *Pseudorasbora parva* at known densities in the 100 m<sup>2</sup> mesocosms. Error bars represent 95 % confidence limits. The solid line represents the significant linear relationship between fish density and catch per unit effort for the baited traps ( $R^2 = 0.99$ ;  $F_{1,60} = 274.4$ , P < 0.001) and the dashed line for the significant relationship between fish density and catch per unit effort for the non-baited traps ( $R^2 = 0.74$ ;  $F_{1,60} = 174.4$ , P < 0.01)

Figure 3. Contour plot of predicted probability of detection (POD) versus known densities of *Pseudorasbora parva* in 100 m<sup>2</sup> mesocosms and the number of electric fishing points required for detection. Key: i: < 0.19; ii: 0.20 - 0.39; iii: 0.40 - 0.59; iv: 0.60 - 0.79; and v) 0.80 - 0.95.



Figure 1.



