

1 **Seasonal climate prediction: a new source of information for the**
2 **management of wind energy resources**

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ABSTRACT

18 Climate predictions tailored to the wind energy sector represent an innovation in the use
19 of climate information to better manage the future variability of wind energy resources. Tra-
20 ditionally, wind energy users employed a simple approach based on an estimate of a retro-
21 spective climatology. Instead, climate predictions can better support the balance between
22 energy demand and supply, as well as decisions relative to the scheduling of maintenance
23 work. One limitation for the use of the climate predictions is the bias, which has until now
24 prevented their incorporation in wind energy models because they require variables with
25 similar statistical properties to those observed. To overcome this problem, two techniques of
26 probabilistic climate forecast bias adjustment are considered here: a simple bias correction
27 and a calibration method. Both approaches assume the seasonal distributions are Gaussian.
28 These methods are linear and robust, and neither requires parameter estimation; essential
29 features for the small sample sizes of current climate forecast systems. This paper is the
30 first to explore the impact of the necessary bias adjustment on the forecast quality of an op-
31 erational seasonal forecast system, using the European Centre for Medium-Range Weather
32 Forecasts seasonal predictions of near-surface wind speed to produce useful information for
33 wind energy users. The results reveal to what measure the bias adjustment techniques are
34 indispensable to produce statistically consistent and reliable predictions, particularly the cal-
35 ibration method. The forecast quality assessment shows that calibration is a fundamental
36 requirement for a high-quality climate service.

37 **1. Introduction**

38 The demand for renewable energy sources as an alternative to fossil-fuel sources has increased
39 due to reasons such as the need to mitigate the climate change resulting from anthropogenic green-
40 house gas emissions, the interest in the creation of new economic opportunities and the provision
41 of energy access to people living in areas without access to other sources of energy (Renewable En-
42 ergy Policy Network for the 21st Century 2015; Solomon 2007). Furthermore, the 21st Congress
43 of the Parties for the United Nations Framework Convention on Climate Change (COP21) agree-
44 ment has recently proposed several policies to promote the energy efficiency and replace the fossil
45 fuels by the use of renewable energies (Lane 2016). Wind energy is the cheapest option for the new
46 sources of power generating capacity and the second leading renewable energy source worldwide,
47 only exceeded by hydropower in terms of installed capacity (Pryor and Barthelmie 2010; Santos
48 et al. 2015). In recent years, wind power installed capacity has experienced a rapid growth, with
49 a total of 370 GW installed worldwide in 2014. As a consequence, wind energy has become a
50 key element of the electricity supply in many parts of the world (World Wind Energy Association
51 2015).

52 Operational and economic issues related to wind energy, such as the need to match supply with
53 demand at all times under the intermittent nature of wind, require the modeling and forecasting
54 of wind power generation processes at a range of temporal and spatial scales (Pinson 2013). Pre-
55 diction of the variability of wind energy resources, which has been identified as a challenge to the
56 grid integration of wind energy systems (Najafi et al. 2016; Füss et al. 2013), is a key piece of
57 the decision-making processes because it allows end users to take informed, precautionary action
58 with potential cost savings to their operations. Hence, more efficient energy management strongly
59 depends on having accurate resource forecasts. Wind energy forecasting options have been tra-

ditionally limited to short (from hours to a few days) time scales because near-surface winds and thus wind energy production, strongly depend on the meso- and synoptic-scale variability (Graff et al. 2014; Pryor and Barthelmie 2010). At longer time scales, the assessment of the economic feasibility of future wind farms is a function of, among other things, the expected energy yield and the maintenance requirements over their life span of periods from a month to several decades. However this information is not readily available to the relevant users, who have to rely on past information based on observations, and this is often only available as short time series. The need of climate information representative of the next few decades has raised the interest of the wind industry in climate projections, which are increasingly being used in long-term resource evaluation (Hueging et al. 2013; Reyers et al. 2015; Vautard et al. 2014).

Focusing on time scales from one month to a decade into the future, current energy practices use an approach based on the future climate being a repetition of an estimate of the climatology (Garcia-Morales and Dubus 2007). However, advances in climate prediction science that cover the climate information gap between weather forecasting and climate change projections can be considered as an alternative to the state-of-the-art by providing predictive information that helps users to take more informed decisions and move beyond using only climatological information.

It has been shown recently that climate predictions are capable to provide additional value for wind energy applications, especially for the management of power production plants (Clark et al. 2017; García-Bustamante et al. 2009; Lynch et al. 2014; Troccoli 2010). For instance, climate predictions could allow electricity system operators to estimate the future production generated by wind farms and use it as input for load-balance models. Should this potential of climate prediction materialize, the matching of supply and demand could be optimized and significant cost savings be made with a better anticipation of market changes. This framework will favor greater penetration of the renewable electricity into the markets.

84 While the scenario described is of great interest to the renewable energy community, little
85 progress had been made in practice. However in recent years the skill of the climate predictions has
86 significantly improved (Doblas-Reyes et al. 2013). For instance, seasonal prediction systems (i.e.
87 those providing information for periods ranging from one month to slightly longer than one year
88 into the future) are now providing skillful forecasts for extra-tropical regions where no substantial
89 skill was found before (Clark et al. 2017; Dunstone et al. 2016; Scaife et al. 2014). This will
90 promote their application wind energy decision making as illustrated for different energy sources
91 (De Felice et al. 2015; Garcia-Morales and Dubus 2007). Currently however there are very few
92 instances of the application of seasonal predictions in the wind energy industry. Improved climate
93 information which includes seasonal forecasts may change this, for example by allowing innova-
94 tive wind energy insurance and helping to cover high risk periods associated with persistent lower
95 than expected wind resource.

96 Seasonal predictions will be beneficial if they are skillful enough, but also if they must be tai-
97 lored to the potential users in a decision-making context. In particular, seasonal predictions have
98 systematic errors that make them unusable unless they are post-processed to have similar statis-
99 tical features as the observational reference employed. This problem has been recognized by the
100 climate science community as one of the main challenges for moving to a better use of climate
101 predictions (Buontempo et al. 2014; Coelho and Costa 2010). The recent FP7 European projects
102 on climate services EUPORIAS¹ and SPECS² have tried to address these challenges and support
103 the development of sectorial climate services in Europe through the involvement of stakeholders
104 in the definition of effective ways to develop climate information.

¹<http://www.euporias.eu/>

²<http://www.specs-fp7.eu/>

105 This paper raises the limits associated with current seasonal prediction systems for their use in
106 wind energy applications. It focuses on the description of appropriate bias adjustment techniques
107 to overcome some of these limits and promote the use of the climate prediction information in
108 those occasions in which it can provide greater accuracy than current approaches. The methodol-
109 ogy described recognizes that end users must be provided with information about the prediction
110 uncertainty (Alessandrini et al. 2013), so that a probabilistic approach is adopted because it is
111 more valuable in user-specific loss functions (Pinson and Tastu 2013).

112 An overview of the necessary steps to provide climate predictions to the wind energy sector is
113 provided in Fig. 1, which summarizes the main challenges addressed in this paper. Section 2 of
114 the paper introduces the data sets and describes one of the most widely used seasonal prediction
115 systems and its limitations. Section 3 describes appropriate bias adjustment techniques and intro-
116 duces forecast quality assessment measures and explains their relevance in a user context. Section
117 4 presents the impact of the bias adjustments over the wind speed seasonal forecasts including
118 an analysis of the changes in the statistical properties of the post-processed predictions. Finally,
119 Section 5 reports the concluding remarks and provides a wider context for future work in the
120 dissemination of climate predictions in user-relevant formats.

121 **2. Data**

122 In this study we use the 10-m wind speed forecasts from the European Centre for Medium-
123 Range Weather Forecasts (ECMWF) System 4 (System 4) operational seasonal prediction system
124 (Molteni et al. 2011), which is based on a global climate model, with coupled atmospheric and
125 oceanic components. System 4 comprises of the ECMWF atmospheric model, the Integrated
126 Forecast System (IFS) CY36R4 with a T255 spectral truncation (horizontal resolution of approxi-
127 mately 80 km) and 91 vertical levels reaching up to 0.1 hPa, which is coupled to the ocean model

128 NEMO (Nucleus for European Modeling of the Ocean) version 3.0. The ocean model uses a grid
129 with horizontal resolution of around 1° in the extratropics with equatorial refinement and 42 levels
130 in the vertical. The atmosphere and ocean are coupled using a version of the OASIS3 (Ocean
131 Atmosphere Sea Ice Soil) coupler developed at the CERFACS (Centre Européen de Recherche et
132 de Formation Avancée en Calcul Scientifique).

133 System 4 is run in ensemble prediction mode. Ensemble predictions are a way to deal with
134 uncertainties in the climate system, in particular those associated with the imperfections of the
135 initial conditions and in the model formulation (Slingo and Palmer 2011). For this reason, the
136 operational System 4 forecasts are produced at the beginning of each month with 51-member
137 ensembles. Each member of the ensemble uses slightly different initial conditions and different
138 realizations of stochastic representations of sub-grid physical processes in the atmosphere. This
139 allows the prediction of the forecast uncertainty (measured by the ensemble dispersion), along
140 with the prediction itself. The simulations are performed for up to seven months into the future.

141 Traditionally seasonal prediction systems do not produce operational forecasts of wind speeds at
142 turbine height levels. Instead, wind speeds are made available at 10- or at different pressure levels.
143 It is difficult to interpolate directly to hub height as the physical height of pressure levels is not
144 constant over time. For that reason 10-m wind speeds have been selected for this analysis. Should
145 the renewable energy community show an interest in seasonal prediction systems to deliver wind
146 speed at hub-height, this might be possible by the forecast systems.

147 The analysis in this paper focuses on the boreal winter as the winter season has larger wind speed
148 variability in the Northern Hemisphere (Archer and Jacobson 2013).. In addition, the analysis of
149 the seasonal predictions of wind speed in winter can be relevant due to the higher variability
150 of wind power supply in that particular season (Bett and Thornton 2015). This illustrates the
151 potential of seasonal predictions for end users as they potentially have more impact where the

152 inter-annual variability is the largest, although other seasons have also been analyzed (Fig. S1)
153 and the conclusions apply equally. The predictions considered here are those issued on the 1st of
154 November, for which three-month statistics for the December-January-February (DJF, also known
155 as one-month lead seasonal forecast) period are made. Predictions over the period 1981–2013 have
156 been used in the study. The prediction for DJF in 2013 has been used as an operational forecast
157 and the predictions over 1981-2012 have been used as the retrospective predictions (hindcasts) to
158 be used in the validation process. This consideration aims to emulate true operational prediction
159 conditions when no observed information about the future is available.

160 To evaluate the System 4 prediction quality, we compare the predicted 10-m wind speed with
161 the corresponding variable of the ERA-Interim reanalysis (Dee et al. 2011). This reanalysis uses
162 the ECMWF Integrated Forecast System (IFS) atmospheric model to assimilate observational data
163 of many types, including in-situ observations and satellite retrievals, to produce a spatially and
164 temporally complete 'best-guess' gridded observational data set. ERA-Interim has the same reso-
165 lution as System 4. This resolution is fairly coarse, but this product offers uniform global coverage
166 in exchange. Given the sparsity of global wind observations reanalyses have demonstrated their
167 potential usefulness for large-scale wind energy applications (Cannon et al. 2015). The problems
168 related with the lack of long enough historical data needed have also promoted the use of reanaly-
169 ses by the wind industry (Rose and Apt 2015).

170 For this reason, and being aware that reanalysis estimates could often be far from point observed
171 values, the reanalysis has been used as the best available estimate of wind speed. The choice
172 of reanalysis is arbitrary and the conclusions are equally valid when using other reanalysis, both
173 global or regional. Further work is needed to assess the seasonal predictions for specific wind farm
174 locations.

175 **3. Methodology**

176 *a. Data Processing*

177 The wind speed forecasts are affected by biases resulting from the inability numerically repro-
178 duce all the relevant processes responsible of climate variability (Doblas-Reyes et al. 2013). Apart
179 from biases in the mean and other characteristics of the distribution of the simulated variables, for
180 probabilistic forecasts additional difficulties appear such as the lack of forecast reliability (Pinson
181 2012), which quantifies the agreement between the predicted probabilities and observed relative
182 frequencies of a particular event. This is important from a wind energy point of view since reliable
183 probabilities are expected to be included in decision-making processes. Hence, climate predictions
184 require a bias adjustment stage to statistically resemble the observational reference, minimize fore-
185 cast errors and formulate reliable probabilities. The bias adjustment of the wind speed has been
186 identified as a requirement of the wind energy sector to fulfill acceptable reliability requirements
187 to be used in their decision-making processes (Alessandrini et al. 2013).

188 This paper illustrates the relative merits of different techniques for the statistical bias adjustment
189 of ensemble forecasts to address different aspects of the forecast error. Two approaches, a simple
190 bias correction and a calibration method, have been selected.

191 1) SIMPLE BIAS CORRECTION

192 The simple bias correction is based on the assumption that both the reference and predicted dis-
193 tributions of seasonal wind speed, are well approximated by a Gaussian (normal) distribution. The
194 adjustment creates predictions with the same mean and standard deviation as the reference data
195 set. This is a zero-order approach for the correction of the systematic mean error that has been
196 previously applied to correct temperature and precipitation (Leung et al. 1999). The Gaussian as-

197 sumption is a limitation of the approach because the monthly and seasonal wind speed distribution
198 can be, at times, slightly non-gaussian.

199 The bias correction scheme can be summarized in this way:

$$y_{ij} = (x_{ij} - \bar{x}) \frac{\sigma_{ref}}{\sigma_e} + \bar{o}. \quad (1)$$

200 Seasonal mean anomalies are calculated by subtracting the ensemble mean of the seasonal av-
201 erages (\bar{x}) from the seasonal average of each forecast (x_{ij}) for each year i and for each member j .
202 A new seasonal mean (y_{ij}) is calculated by multiplying the seasonal mean anomaly by the ratio of
203 the standard deviation of the reference data set (σ_{ref}) to the interannual standard deviation of the
204 ensemble members (σ_e), and adding the climatology of the reference data set (\bar{o}). This is done for
205 each grid cell separately, resulting in a new wind speed forecast ensemble, with the same ensemble
206 mean and standard deviation as the reference.

207 2) CALIBRATION METHOD

208 The calibration can be considered as a way of obtaining predictions with interannual variance
209 equivalent to that of a reference data set in a similar way to the bias correction method, but at
210 the same time ensuring an increased reliability of the probability predictions. Here we apply the
211 variance inflation technique (Von Storch and Zwiers 2001). This calibration strategy has been
212 selected because an inflation of the ensemble spread is required to obtain reliable probabilities and
213 it is applied as in (Doblas-Reyes et al. 2005).

214 If x_i is the ensemble-mean prediction for any grid point at year i and z_{ij} is the difference of
215 ensemble member j with the ensemble mean, then the calibrated estimate of the ensemble member
216 j can be expressed as

$$y_{ij} = \alpha x_i + \beta z_{ij}. \quad (2)$$

217 The coefficients α and β are defined as follows:

$$\alpha = \text{abs}(\rho) \frac{\sigma_{ref}}{\sigma_{em}}, \quad (3)$$

$$\beta = \sqrt{1 - \rho^2} \frac{\sigma_{ref}}{\sigma_e}. \quad (4)$$

219 The σ_{em} is the standard deviation of the ensemble mean (the time series of xi), σ_e is the standard
220 deviation of the ensemble, σ_{ref} is the standard deviation of the reference and ρ is the correlation
221 between the ensemble mean of the retrospective forecasts and the reference data set. The α and
222 β coefficients are found under two constraints. The former is that the standard deviation of the
223 inflated prediction is the same as that for the reference and the latter is that the predictable signal
224 after the inflation is made equal to the correlation of the ensemble mean with the reference data
225 set.

226 *b. Forecast quality assessment*

227 Seasonal forecast systems, as in any other forecasting process, have to be systematically com-
228 pared to a reference, preferably observations, to assess their overall quality in a multifaceted pro-
229 cess known as forecast quality assessment (Mason and Baddour 2008). This is a fundamental
230 step to the prediction problem because a prediction has no value without an estimate of its quality
231 based on past performance (Doblas-Reyes et al. 2013). Moreover the quantification of the uncer-
232 tainty is one of the most crucial aspects for the successful development of wind industry and the
233 minimization of the financial risk.

234 Three sources of uncertainty in common scoring metrics of probabilistic forecasts should be
235 considered: improper estimates of probabilities from small-sized ensembles, insufficient number
236 of forecast cases, and imperfect reference values due to observation errors. A way to alleviate
237 these problems is to use several scoring measures to offer a comprehensive picture of the forecast

238 quality of the system (Jolliffe and Stephenson 2012) and to apply statistical inference as often as
239 required.

240 The reader should note that these sources of uncertainty are independent of the uncertainty of
241 the individual forecasts: the user should consider and be provided with, both types of uncertainty
242 when making decisions where this information is included.

243 Several scoring measures are used in this paper, including skill and reliability measures such
244 as the reliability diagram and the rank histogram. Forecast quality has been used to evaluate the
245 performance of the seasonal predictions system as well as the impact of the two bias adjustment
246 techniques over the forecast quality. The goal is to offer the most general and, a priori, relevant
247 information for a user in the wind energy sector instead of the traditional view offered by climate
248 scientists where the information provided to the users is mainly based on correlation, which is very
249 useful, but gives only a small part of the information user requires.

250 1) SKILL SCORES

251 The skill estimates based on the performance of the system in the past, may guide users about
252 the expected performance of the future forecasts (Weisheimer and Palmer 2014), always with the
253 caveat that the predictability of the climate system might change over time. Skill scores are a tool
254 for end users to develop alternative strategies to their baseline information to minimize the risk
255 and to perform an optimal management (Pinson et al. 2009). Skill scores for both deterministic
256 (ensemble mean) and probabilistic predictions are considered.

257 The Pearson correlation coefficient between the ensemble mean and the reference data set has
258 been used as a measure of the linear correspondence between the forecasts and the reference. This
259 deterministic skill measure is invariant to changes in scale, hence the bias correction and calibra-
260 tion of the forecasts do not change the correlation of the ensemble mean with the observations.

261 However the bias adjustment techniques that are illustrated in this paper (defined in section 3.a)
262 have been applied in leave-one-out cross-validation to mimic as closely as possible an operational
263 context in which new coefficients might be estimated to predict each year. In cross-validation mode
264 the prediction to be adjusted is removed from the sample used to estimate the coefficients. As a
265 result the correlation of the post-processed forecast changes relative to the correlation computed
266 directly with the uncorrected forecasts.

267 A comprehensive measure of the predictive skill for the probabilistic seasonal predictions of
268 categorical events is the ranked probability skill score (RPSS) (Epstein 1969; Wilks 2011). This is
269 a squared distance between the cumulative probabilities of the categorical forecast and reference
270 vectors relative to a naive forecast strategy, that in our case has been taken as the climatology (made
271 of all the possible events recorded in the past) because this is the preferred current choice of the
272 users targeted by this analysis. The RPSS is based on the rank probability score (RPS), a measure
273 of the squared distance between the forecast and the reference cumulative probabilities. In the
274 present case the RPSS has been computed based on categorical forecasts for terciles. Three equi-
275 probable events associated with the two terciles of the climatological distribution of the reference:
276 wind speed exceeding the upper tercile (above normal category), not exceeding the lower tercile
277 (below normal category) and values between the two terciles (normal category). The probabilities
278 have been computed as the fraction of ensemble members in the corresponding category. This is
279 only one example, other categories could be defined if they better represent the decisions involved
280 in precautionary climate action. The individual values of the reference data set in the verification
281 time series can fall in any of the three categories with probability determined by the probability
282 density function (PDF) for the target season.

283 The continuous ranked probability skill score (CRPSS) is a commonly used probabilistic skill
284 score (Jolliffe and Stephenson 2012) that has been used to evaluate the predictive skill of the full

285 probability distribution. It is based on the continuous ranked probability score (CRPS), a score
286 that reduces to the mean absolute error if a deterministic forecast is used. The CRPS measures the
287 difference between the predicted and observed cumulative distributions and it can be converted
288 into a skill score, measuring the performance of a forecast relative to the climatology.

289 The RPSS and CRPSS range between 1 to $-\infty$. Skill scores below 0 are defined as unskillful,
290 those equal to 0 are equal to the climatology forecast, and anything above 0 is an improvement
291 upon climatology, up to 1, which indicates a ‘perfect’ forecast.

292 Fair scores to ensemble forecasts have been recently introduced (Fricker et al. 2013; Ferro 2014).
293 A skill score is fair when it favours predictions with ensemble members that perform as if they
294 have been sampled from the same distribution than the reference dataset. The fair version of the
295 RPSS and CRPSS have been used in order to give an estimate of what the skill is when an infinite
296 ensemble size is used (a measure of potential skill). The differences between the results of the fair
297 and the basic scores are small as has been shown for the RPSS in the supplementary material (Fig.
298 S2).

299 2) RELIABILITY

300 Reliability analysis of prediction systems remains as a prime concern for the wind energy sec-
301 tor, as for any user of probability predictions, due to the risks and uncertainties involved in the
302 forecasting of wind resources (Chaudhry and Hughes 2012).

303 Rank histograms are a simple tool to evaluate the reliability of ensemble forecasting systems
304 (Elmore 2005). They are generated by dividing the observations among a limited number of bins,
305 thereby defining a set of exhaustive and mutually exclusive events. Then the observed frequencies
306 for these bins are compared with the corresponding forecast probabilities. Rank histograms help
307 to know if the forecast is assumed to be reliable and then it is expected to be flat. However,

308 some deviations from uniformity can appear for reliable forecasts due to randomness. The rank
309 histograms have been displayed on probability paper (Bröcker 2008). In the y-axis rank histograms
310 display cumulative probabilities instead of the traditional observed frequency which indicate how
311 probable that observed frequency would be if the prediction was reliable. This information is
312 useful to identify if the deviations from a reliable behavior are systematic or merely random. In
313 addition the readability of the rank histogram is further improved by scaling the ordinate by a logit-
314 transformation, that has the effect of displaying both small and large probabilities equidistantly.
315 On the right the 90, 95, and 99 percent simultaneous confidence intervals have been represented.

316 Rank histograms illustrate if the ensemble members and the verifying observation come from
317 the same probability distribution, in which case the forecasts are statistically consistent then no
318 calibration of the ensemble is needed. This happens when the rank histogram is flat (as if coming
319 from a uniform distribution). However, because of sampling variations the histograms are almost
320 never flat. To assess if the deviations from flatness are attributed to chance or deficiencies in the
321 forecasts, goodness-of-fit test statistics are computed: Pearson χ^2 , the Jolliffe-Primo test statistic
322 for slope (JP slope) and the Jolliffe-Primo test statistic for convexity (JP convex) (Jolliffe and
323 Primo 2008). The Jolliffe-Primo statistics are obtained from the decomposition of the Pearson
324 χ^2 in components that allow the identification of bias (slope) or under/over-dispersion (convexity)
325 in the forecast ensemble. The detailed mathematical definition of this goodness-of-fit test can be
326 found in the appendix of Jolliffe and Primo 2008.

327 Reliability diagrams are a common diagnostic of probabilistic predictions that assess both reli-
328 ability and skill. They consist of a plot of the observed relative frequency against the predicted
329 probability of a dichotomous event, providing a quick visual assessment of the impact of tuning
330 probabilistic forecast systems. A perfectly reliable system should draw a line as closely as possible
331 to the diagonal, within a certain measure of uncertainty.

332 The information provided by the reliability diagram should be interpreted with care because even
333 a perfectly reliable forecast system is not expected to have an exactly diagonal reliability diagram
334 due to the limited samples typical of seasonal forecast systems (Jolliffe and Stephenson 2012).
335 To deal with this problem we have included consistency bars (Bröcker and Smith 2007) in these
336 diagrams. They indicate how likely the observed relative frequencies are, under the assumption
337 that predicted probabilities are accurate.

338 To draw a reliability diagram, discretization and grouping into probability bins (ten in this pa-
339 per) of the probability forecasts have to be done. A reliability diagram also includes the frequency
340 of the forecast probabilities included in each bin, which is known as sharpness diagram. Sharp-
341 ness gives an indication of the variation in forecast probabilities issued by the prediction system,
342 independently of the observations.

343 The rank histogram and the reliability diagram are complementary tools to assess the reliability
344 of the system. The former assesses the full forecast ensemble and does not require the formulation
345 of forecast probabilities, an aspect that is necessary in the case of the reliability diagram, where
346 one assesses the features of both the forecast system and the statistical model that transforms the
347 ensemble into probabilities.

348 **4. Results**

349 Total wind power installed indicates the wind power capacity available in each wind farm. It has
350 been represented in Fig. 2 to identify which are the most important locations from a wind energy
351 user point of view. To illustrate the performance of the seasonal predictions two key regions for the
352 wind energy sector because wind farms are located there have been selected. For the selection of
353 the regions we have also taken into account the potential skill available in such regions (Fig. S1).
354 The first region is in Canada [longitude: 112.5°-113.2°W and latitude: 50.3°-51.0°N]. This country

355 is an important player in terms of energy resources (Vaillancourt et al. 2014) and a global leader
356 in the sustainable development of wind energy. This region had an exceptional year in 2014 for
357 wind energy development, ranking seventh globally in terms of new installed capacity (Canadian
358 Wind Energy Association 2015) that year. The North Sea region [longitude: 9.8°-10.6°E and
359 latitude: 58.0°-58.7°] is the second region considered. It is the most important region for offshore
360 energy activities in Europe due to the large and consistent wind resource, the relatively shallow
361 water that minimizes the cost of the wind farms and the proximity to developed electricity markets
362 (Schillings et al. 2012).

363 Fig. 3 displays the predictions for the uncorrected, bias corrected and calibrated sets for these
364 two regions. The effect of the bias adjustment over the predictions is that when the corrections are
365 applied, the hindcasts (grey dots) show similar mean and variance to the reference data set (black
366 dots). After the bias adjustment the probabilities in each category differ as a result of the changes
367 in the ensemble distribution. The skill changes accordingly with the bias adjustment, showing a
368 decrease in the correlation and an increase in the probabilistic skill scores. The decrease of the
369 correlation is due to the cross-validation, which leads to an implicit leakage of information and
370 a degeneracy in this measure of potential skill (Barnston and van den Dool 1993; Barnston et al.
371 2012). The improvement of the fair RPSS and CRPSS are associated with the reduction of the
372 systematic errors. Contrary to the correlation, the RPSS and the CRPSS are both sensitive to the
373 systematic differences in the statistical properties (mean, variance) of the predicted variables with
374 respect to those in the observations as well as to the inadequacy of the ensemble dispersion to
375 act as a prediction of the forecast error (the lack of reliability). This is a useful example of the
376 importance of using more than one forecast quality measure, in particular when dealing with user
377 relevant variables.

378 The information provided by global forecast systems is relatively coarse. In a global context,
379 the sizes of the two selected regions are small. Besides, for a small region the skill is expected
380 to be noisier and less robust than for a larger one. In order to explore how the size of the region
381 affects the forecast quality we have estimated the forecast quality for larger regions (Fig. S2).
382 The comparison shows that the skill differences are small when a larger region is considered.
383 Future work will focus on the formulation of predictions for specific sites. This is a non-trivial
384 task because the bias adjustment techniques necessary in seasonal forecasting require long-enough
385 observational references that are not readily available.

386 The forecast system considered allows estimating the global forecast quality of the different
387 sets of predictions. The fair RPSS maps for the uncorrected, bias corrected and calibrated wind
388 speed are shown in Fig. 4. The uncorrected predictions (Fig. 4 (a)) display very low scores all
389 around the world. The highest values are found in tropical regions, in particular in some regions
390 of North East of South America and North Western Africa. This maximum can be explained
391 because the largest predictability at seasonal timescales is attributed to anomalies in the tropical
392 sea surface temperatures (SST) resulting from coupled ocean-atmosphere phenomena, in particular
393 those related to El Niño-Southern Oscillation events (Kirtman and Pirani 2009) that affect mainly
394 the regions mentioned above.

395 Fig. 4 (b) and (c) show that the fair RPSS increases globally when bias adjustment is applied.
396 This kind of assessments are widely available for variables like temperature and precipitation,
397 but are not available for wind speed. The skill improvement has been quantified in the Fig. 4
398 (d) and (e), which indicate that the skill scores for the bias adjusted predictions increase more
399 than 1 relative to the uncorrected ones. The fair RPSS maps (Fig. 4 (b) and (c)) for the post-
400 processed predictions have their maximum values in the tropics. Although the skill is relatively
401 low at extratropical latitudes, some positive skill is found in those regions. For instance, some

402 regions in Europe as the North Sea or Scandinavia display positive values. Wind speed predictions
403 show the highest skill in Northern Europe, while in Southern Europe negative RPSS values is
404 found. This is in agreement with previous work (e.g. Weisheimer et al. 2011) indicating that
405 seasonal dynamical predictions have limited forecast quality over Europe.

406 The skill improvement is also present in South-eastern Asia, central United States or North-
407 eastern South America where positive values appear when bias correction and calibration tech-
408 niques are applied. The bias adjustment allows the skill in those regions associated with ENSO
409 teleconnections (Hamlington et al. 2015; Quan et al. 2006), as well as with other sources of sea-
410 sonal to interannual predictability, such as the persistence of the North Pacific decadal oscillation
411 (Gershunov and Cayan 2003) to emerge. Wind speed with positive skill in North American regions
412 has important implications for the wind energy sector in this economically active region.

413 The differences between the correlation and CRPSS before and after the bias adjustment of the
414 wind speed forecasts have been included in the Fig. S3 and S4. The correlation of the uncorrected
415 forecasts is always higher due to the cross-validation leakage mentioned above. It is noticeable
416 that the correlation spatial distribution in the calibrated hindcasts is noisier than the two other types
417 of forecasts considered. This is due to the coefficients estimated in the calibration having a smaller
418 spatial decorrelation length and being less robust than the mean and variance used in the simple
419 bias correction.

420 For the uncorrected predictions (Fig. 5 (a) and (b)), the overpopulated lower ranks and the neg-
421 ative slope in the rank histogram illustrate that a positive unconditional bias is present in the data.
422 These biases appear for the predictions of both regions, although the effect of this deficiency seems
423 more important in Canada (Figure 5 (a)) where all the observations are exceeded by the majority
424 of the ensemble members, leaving the highest rank categories almost empty. The bias corrected
425 and calibrated forecasts show more homogeneously populated ranks indicating that the reliability

426 of the ensemble improves when the bias adjustment is applied. However, the deviation of the flat-
427 ness of these rank histograms could be the result of some forecast deficiencies still remaining after
428 the bias adjustment. For instance, for the calibrated forecasts in Canada (Fig. 5 (e)), the rank 50
429 shows a very large value that might indicate that the ensemble overestimates the true uncertainty
430 range.

431 To assess if the deviations from flatness of the rank histograms are attributed to either chance or
432 deficiencies in the forecasts, goodness-of-fit test statistics, with the null hypothesis being that the
433 rank histogram is uniform, are computed and included in Table 1. The three statistical tests, the
434 Pearson χ^2 , the JP slope and JP convex, allow us to identify if the forecasts are biased or whether
435 the ensemble has over or under-dispersion.

436 Table 1 shows that departures from flatness exist for the uncorrected forecasts, especially in
437 Canada, where the tests take very high values, showing that the ensembles are under-dispersive, as
438 evidenced by the high JP convex test. The high values of the JP slope show that the forecasts are
439 also affected by biases. The uncorrected forecasts in the North Sea have also biases and are under-
440 dispersive, although the statistical tests have smaller values than those in Canada. The results are
441 statistically significant, with the p-values being virtually zero.

442 The tests applied to the simple bias corrected and the calibrated forecasts indicate that the de-
443 viation from flatness is minimised when the bias adjustment is applied. The Pearson χ^2 for the
444 calibrated data in the Canada region has higher values than the bias corrected ones (p-value 0.01),
445 while the JP tests provide no evidence of departures from flatness with p-values higher than 0.01.
446 Consequently this result shows that the biases and the under-dispersion in the raw ensemble are
447 corrected, and the deviations from uniformity are independent of these specific problems. Making
448 sure that the ensemble is well calibrated, which is a critical aspect of the forecast for the user,
449 because it suggests that the ensemble predictions represent the forecast error, within statistical

450 sampling, and can be trusted in specific applications that have been developed using meteorologi-
451 cal observational references.

452 To further analyse the impact of the bias adjustment on reliability, reliability diagrams (Fig. 6)
453 allow the comparison between the observed frequencies with forecast probabilities (obtained from
454 the ensemble forecasts) for binary events. The events are defined by the thresholds of the lower and
455 upper terciles, as for the RPSS but in a dichotomous way. If the prediction system is reliable, then
456 a good agreement should exist between forecast probabilities and observed relative frequencies
457 and the graph should be close to the diagonal.

458 The slope of the reliability diagrams is positive. This shows that as the forecast probability of
459 the event occurring increases, so does the verified chance of observing the event and therefore the
460 forecasts have some reliability. The reliability curves for the three events have a steeper slope than
461 the diagonal in both regions suggesting that the probability forecasts are overconfident. For the
462 uncorrected forecasts in Canada (Fig. 6 (a)), the curve for the below-normal category (blue line)
463 flattens when the forecast probability is above 0.45. This means that when the forecast probability
464 is higher than 0.45 there is no relationship between the forecast probabilities and the frequency of
465 the observed below-normal wind speeds. The reliability diagram for the uncorrected predictions
466 in the North Sea (Fig. 6 (b)) shows only a narrow set of probabilities issued, with values ranging
467 from 0.1 to 0.5 for the above (red line) and below normal (blue line) categories and from 0.4 to
468 0.7 for the normal category (orange line). In addition the above-normal category is so steep that
469 falls outside the consistency bars. This illustrates the poor reliability for that event in the North
470 Sea when the predictions are uncorrected.

471 The reliability curves of the bias corrected predictions (Fig. 6 (c,d)) show similar features to the
472 uncalibrated ones. One should bear in mind that, apart from correcting the mean and standard
473 deviation of the forecast distribution, the simple bias correction does not have any additional im-

474 pact on the predictions and, hence, no substantial changes beyond the effect of the cross-validation
475 should be expected in the reliability diagram.

476 The calibrated predictions for the above-normal and below-normal events (Fig. 6 (e,f)) have re-
477 liability diagrams with their points lying closer to the diagonal than found for the uncorrected and
478 bias corrected predictions. This corresponds to a better agreement between the forecast proba-
479 bilities and the probability of the observed event than in the other two cases suggesting that the
480 overconfidence has been corrected. In the North Sea (Fig. 6 (f)) the slope of the curve for the
481 normal category (orange line) (Fig. 6 (f)) becomes horizontal suggesting that the system can not
482 discriminate between predictable and unpredictable normal wind speeds in this region, which is
483 not surprising because normal events might not have strong signals, which are those associated
484 with the predictability of the system.

485 In addition, for the predictions of below-normal and above-normal wind speeds after calibration
486 the sharpness diagrams (Fig. 6 (e,f)) show more homogeneously populated bins for both regions.
487 This means that the forecast system is able to predict those events with a larger range of forecast
488 probability values. Conversely, the uncorrected and simple bias corrected predictions display their
489 frequency peaks near the climatological frequency, so that they predict often the event with a
490 climatological probability. These results show the improvement in the reliability of the predictions
491 obtained when calibration is applied, improvements that are particularly relevant to the users.

492 **5. Conclusions**

493 Seasonal predictions have not yet been widely taken into account by the wind-energy sector.
494 However, some applications in the energy sector of this type of forecasts have been recently iden-
495 tified. They illustrate that predictions at seasonal time scales can be used as input by the industry

496 in decision-making processes to replace the current naive climatological information. In this paper
497 we illustrate a strategy for the use of wind speed seasonal predictions by the wind energy sector.

498 After describing one of the most popular operational seasonal forecast systems, ECMWF's Sys-
499 tem 4, and its forecast quality characteristics, two different bias adjustment techniques to correct
500 the typical deficiencies of the predictions of global forecast systems are described. It is shown
501 that bias adjustment is indispensable for the predictions to be usable. The System 4 predictions
502 have skill in predicting wind speed at seasonal time scales, especially in the tropics, but also in
503 extratropical regions of relevance to the wind-energy sector. This is an encouraging result that has
504 not been documented elsewhere. However, dynamical seasonal predictions suffer from a number
505 of important systematic errors that also affect wind speed predictions. bias adjustment methods
506 are required for the predictions to have the same statistical properties of the observational refer-
507 ence and hence to be applicable by the users. Concerning the bias adjustment, the simple bias
508 correction and the calibration methods produce predictions with statistical properties that allow
509 their actual application. The most important gain in forecast quality for the seasonal predictions
510 comes through the increase in their skill and reliability, the latter a critical aspect of the forecasts
511 from the user perspective. These gains in forecast quality cannot be evidenced using correlation,
512 which suggests that more than one forecast quality measure is needed even in a user context.

513 The predictions and the impact of the bias adjustment are illustrated on two skillful regions that
514 are crucial for the wind energy sector, the North Sea and central Canada. A further analysis of the
515 predictions reveals that both the bias correction and calibration methods produce an improvement
516 in the consistency of the ensemble. Besides, the reliability diagrams demonstrate that the calibra-
517 tion method, which also corrects the deficiencies in the ensemble spread, provides more reliable
518 predictions than the simple bias correction technique. Improvements in reliability are fundamental
519 from a user perspective because it guarantees the trustworthiness of the predictions.

520 Our work demonstrates that calibration is necessary because it produces an improvement in both
521 skill and reliability, making this technique essential for the seasonal predictions to be usable. The
522 development of these strategies is part of a recent initiative undertaken by the climate community
523 where climate services are developed to provide more relevant, reliable and action-oriented climate
524 information (Buontempo et al. 2014). This paper illustrates the fact that seasonal predictions of
525 near-surface wind speed have skill in several regions where there is substantial installed power,
526 and that after bias adjustment the predictions are reliable for their use.

527 Future improvements include the combination of seasonal predictions from different sources,
528 based on both dynamical and empirical-statistical forecast systems. The global and illustrative
529 character of this paper requires the use of a reanalysis as reference data. The verification against
530 other reanalyses and regional observed wind speed data might offer slightly different results be-
531 cause of the observational uncertainty, which is an additional factor that will be taken into account
532 in future analyses, but the need of a bias adjustment process will be unavoidable. Finally, there are
533 simple ways to convert the wind speed into energy density that will be explored from the seasonal
534 prediction point of view, while the use of empirical downscaling could offer additional benefits
535 when considering seasonal predictions for specific power plants.

536 The work described here opens the field to the next step in the development of a climate service:
537 the creation of tailored products that facilitate the widespread use of climate predictions by the
538 wind-energy sector (Step 4 in Fig.1). The release of climate services can range from knowledge
539 transfer (informing, documenting and providing training in the best bias adjustment techniques)
540 to the creation of operational online interactive interfaces to allow wind industry user easily ex-
541 plore probabilistic predictions. An example of a prototype of interactive platform that incorporates
542 bias-adjusted predictions can be found at Project Ukko³ interface designed in the framework of

³<http://www.project-ukko.net>

543 the EUPORIAS project. In addition, the New European Wind Atlas (NEWA⁴), which is currently
544 in development will provide access to skill evaluations of climate predictions. Further interactions
545 between the climate science community and renewable energy community are also indispensable
546 to quantify the actual economic value of climate predictions and evaluate the predictions perfor-
547 mance in the past. This is a necessary step to demonstrate to energy stakeholders the saliency of
548 climate predictions outcomes.

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714 by Jolliffe and Primo 2008. They have been computed from the rank histograms
715 (Fig. 4) of 10-m wind speed forecasts from ECMWF System 4 in winter (DJF)
716 for the period 1981-2012. 36

	Uncorrected		Simple bias corrected		Calibrated	
	Canada	North Sea	Canada	North Sea	Canada	North Sea
Pearson χ^2	2080.37	148.25	36.12	57.25	77.56	68.62
p value	0.00	2.01×10^{-11}	0.94	0.25	0.01	0.05
JP slope	291.69	38.46	0.17	0.1	0.03	0.12
p value	2.12×10^{-65}	0.56×10^{-11}	0.68	0.75	0.87	0.72
JP convex	319.69	11.32	1.15	1.77	0.73	3.07
p value	1.69×10^{-71}	0.76×10^{-04}	0.28	0.18	0.39	0.06

717 **Table 1.** Goodness-of-fit tests: Pearson χ^2 , JP-slope and JP-convex statistics formulated by Jolliffe and Primo
718 2008. They have been computed from the rank histograms (Fig. 4) of 10-m wind speed forecasts from ECMWF
719 System 4 in winter (DJF) for the period 1981-2012.

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721 **Fig. 1.** Main steps for the development of a climate service for the wind energy sector based on
 722 seasonal climate predictions. Steps 2 and 3 in the diagram outline the main challenges
 723 addressed in this paper. 38

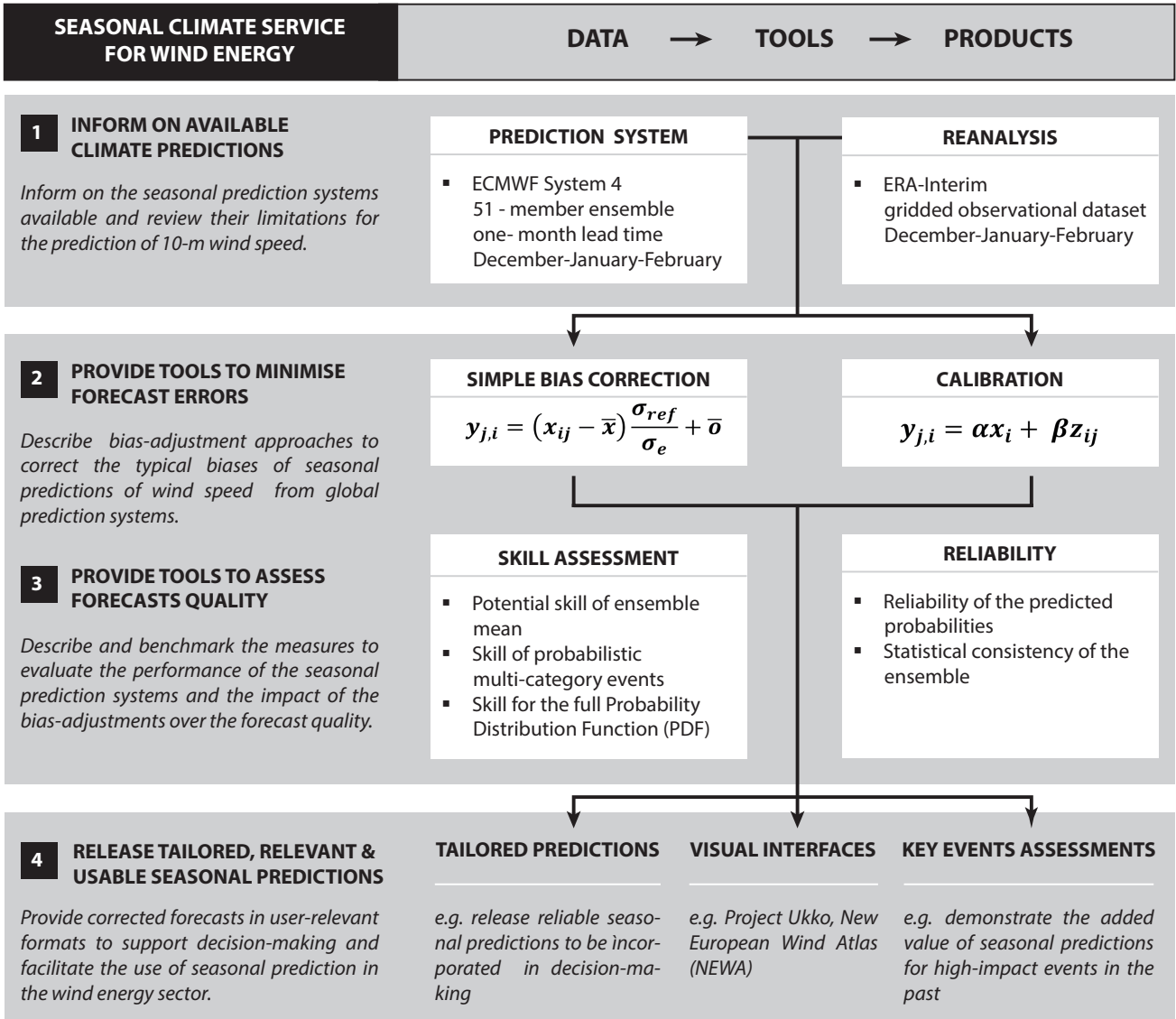
724 **Fig. 2.** Total installed wind power capacity for each individual wind farm (operational and under
 725 construction have been included) in 2015 (Source:www.thewindpower.net). 39

726 **Fig. 3.** Time series of 10-m wind speed from ECMWF System 4 and ERA-Interim reanalysis in
 727 winter (DJF). These predictions have been initialized on the first of November for the period
 728 of 1981-2013. The ensemble members of the hindcasts are represented as small grey dots
 729 and the ensemble mean is represented with a large grey dot for each start date. The grey
 730 horizontal line shows the mean of the hindcast in whole period (1981-2012) and the blue and
 731 red horizontal lines show its lower and upper terciles, respectively. The ensemble members
 732 of the forecast year (2013) are represented as red dots. The percentages indicate the fraction
 733 of members in each category, which are limited by the terciles. The black dots represent
 734 the 10-m wind speed values of ERA-Interim. The black horizontal line shows the mean of
 735 the ERA-Interim in the 1981-2012 period. Correlation, RPSS and CRPSS are shown in the
 736 upper part of each panel. 40

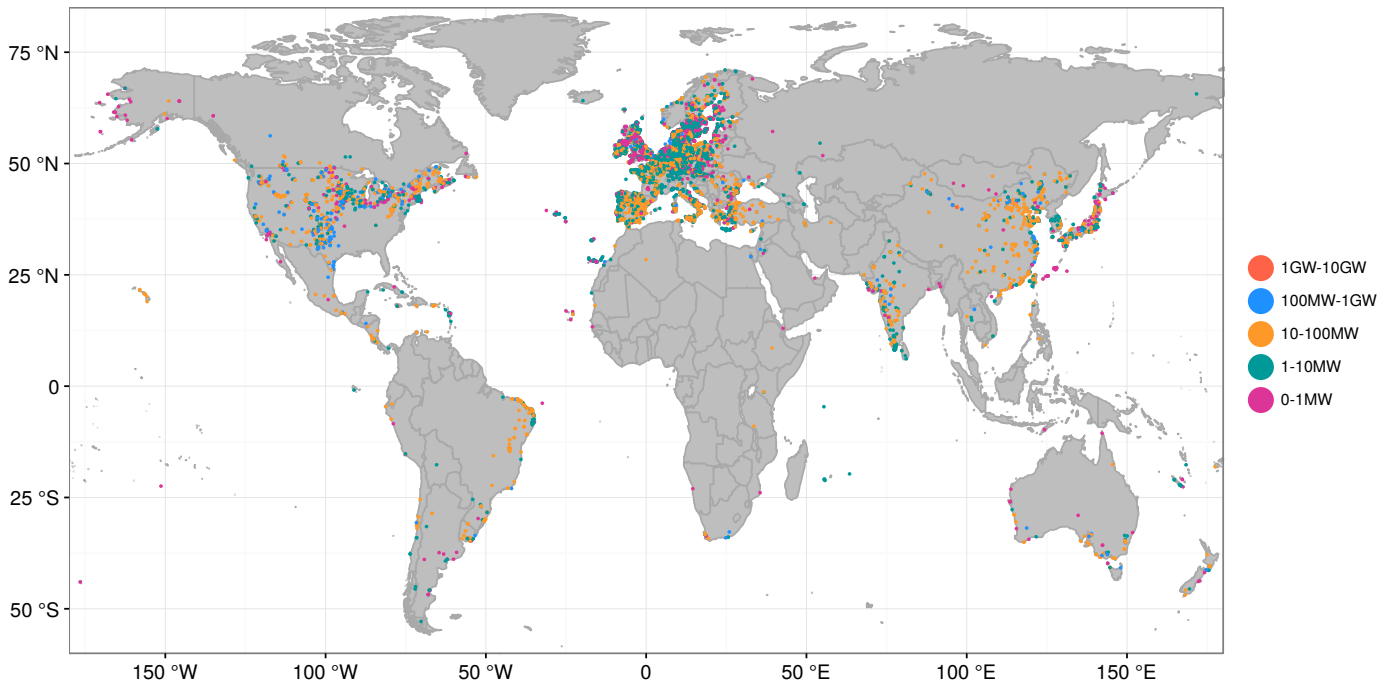
737 **Fig. 4.** Fair Ranked Probability Skill Score (RPSS) for tercile events of 10-m wind speed forecasts
 738 from ECMWF System 4 and ERA-Interim reanalysis in winter (DJF). These predictions
 739 have been initialized on the first of November for the period of 1981-2012. 41

740 **Fig. 5.** Rank histograms of 10-m wind speeds forecasts from ECMWF System 4 and ERA-Interim
 741 reanalysis in winter (DJF). These predictions have been initialized on the first of November
 742 for the period of 1981-2012. These rank histograms have been represented on probability
 743 paper to show if the deviations from a reliable behavior are systematic or random. The x-axis
 744 represents the ranks. The probabilities of the cumulative observed frequency on a log-it scale
 745 are shown in the y-axis. On the right 90, 95 and 99 percent simultaneous confidence intervals
 746 are indicated. If all ranks were equally likely on average, approximately 90 percent of all
 747 rank histogram would be contained in the 90 percent confidence interval and approximately
 748 10 percent of all rank histograms would have at least one bar that falls outside this interval. . . . 42

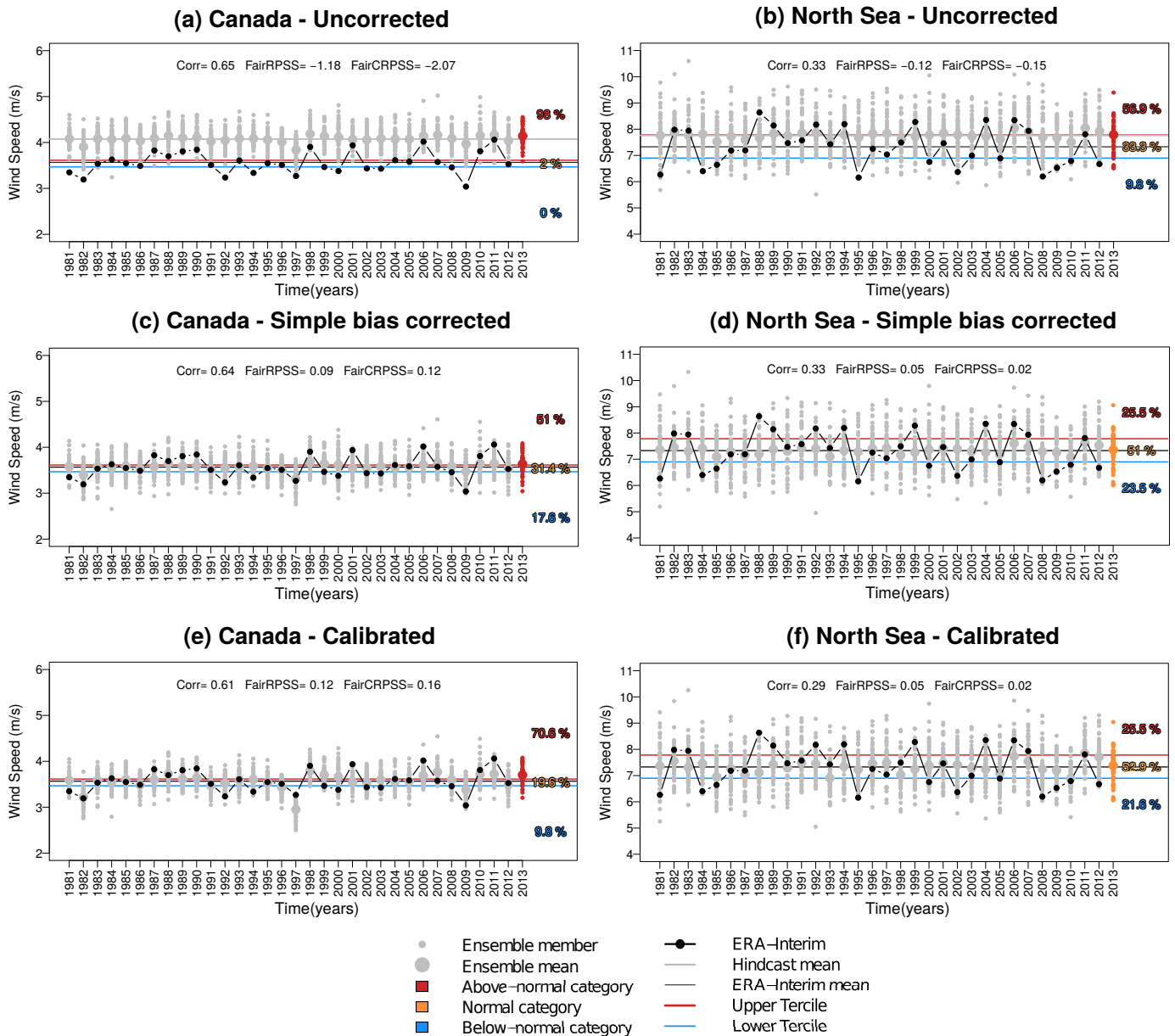
749 **Fig. 6.** Reliability diagrams of 10-m wind speeds forecasts from ECMWF System 4 and ERA-
 750 Interim reanalysis in winter (DJF). These predictions have been initialized on the first of
 751 November for the period of 1981-2012. Three events are represented: above-normal wind
 752 speeds (red line), normal wind speeds (orange) and below-normal wind speeds (blue). Right
 753 panels show the sharpened diagrams with the distribution of samples for each bin and each
 754 event. The consistency bars have been represented as vertical lines to illustrate how likely
 755 the observed relative frequencies are under the assumption that predicted probabilities are
 756 reliable. 43



757 FIG. 1. Main steps for the development of a climate service for the wind energy sector based on seasonal
758 climate predictions. Steps 2 and 3 in the diagram outline the main challenges addressed in this paper.

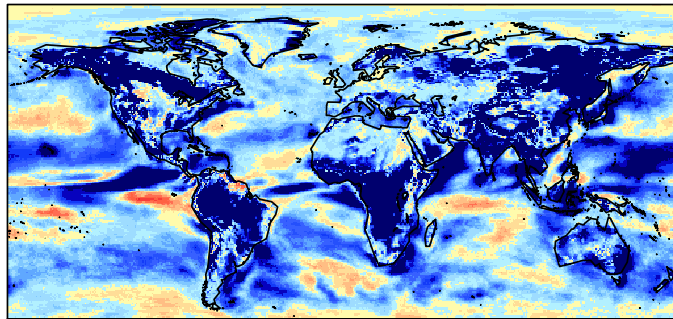


759 FIG. 2. Total installed wind power capacity for each individual wind farm (operational and under construction
760 have been included) in 2015 (Source:www.thewindpower.net).

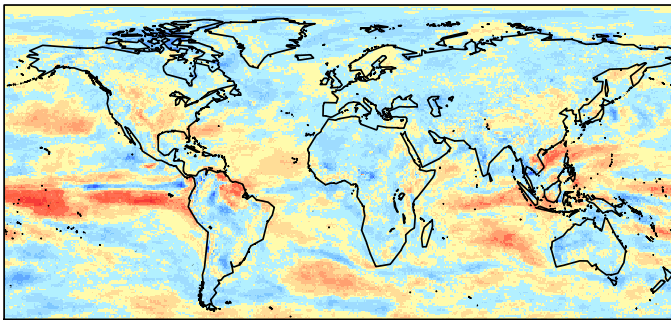


761 FIG. 3. Time series of 10-m wind speed from ECMWF System 4 and ERA-Interim reanalysis in winter (DJF).
 762 These predictions have been initialized on the first of November for the period of 1981-2013. The ensemble
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 764 grey dot for each start date. The grey horizontal line shows the mean of the hindcast in whole period (1981-2012)
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 766 the forecast year (2013) are represented as red dots. The percentages indicate the fraction of members in each
 767 category, which are limited by the terciles. The black dots represent the 10-m wind speed values of ERA-Interim.
 768 The black horizontal line shows the mean of the ERA-Interim in the 1981-2012 period. Correlation, RPSS and
 769 CRPSS are shown in the upper part of each panel.

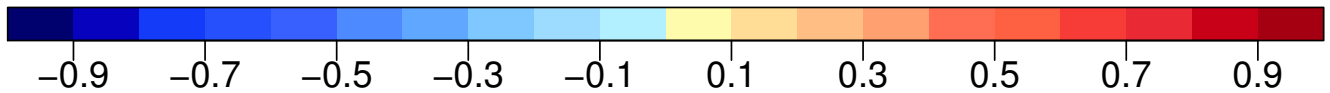
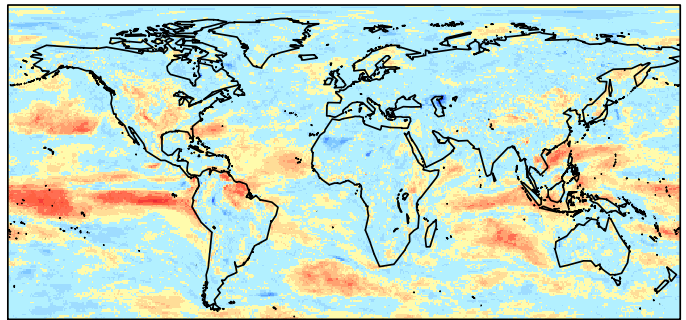
(a) Uncorrected



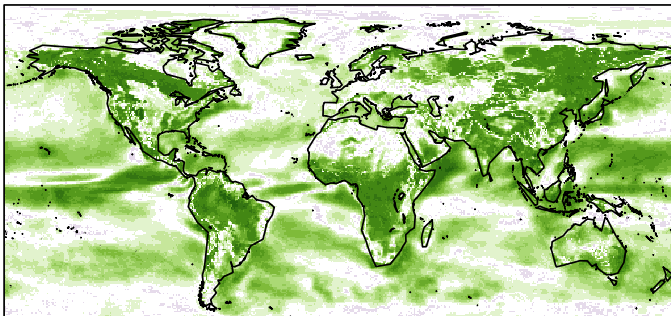
(b) Simple bias corrected



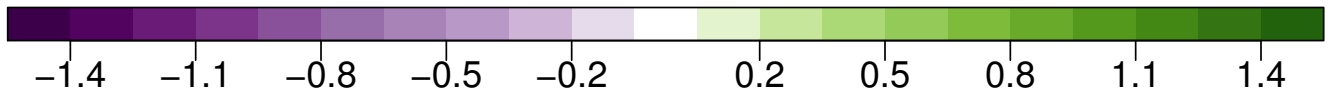
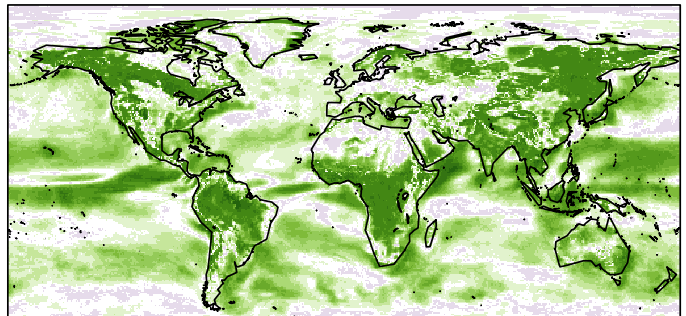
(c) Calibrated



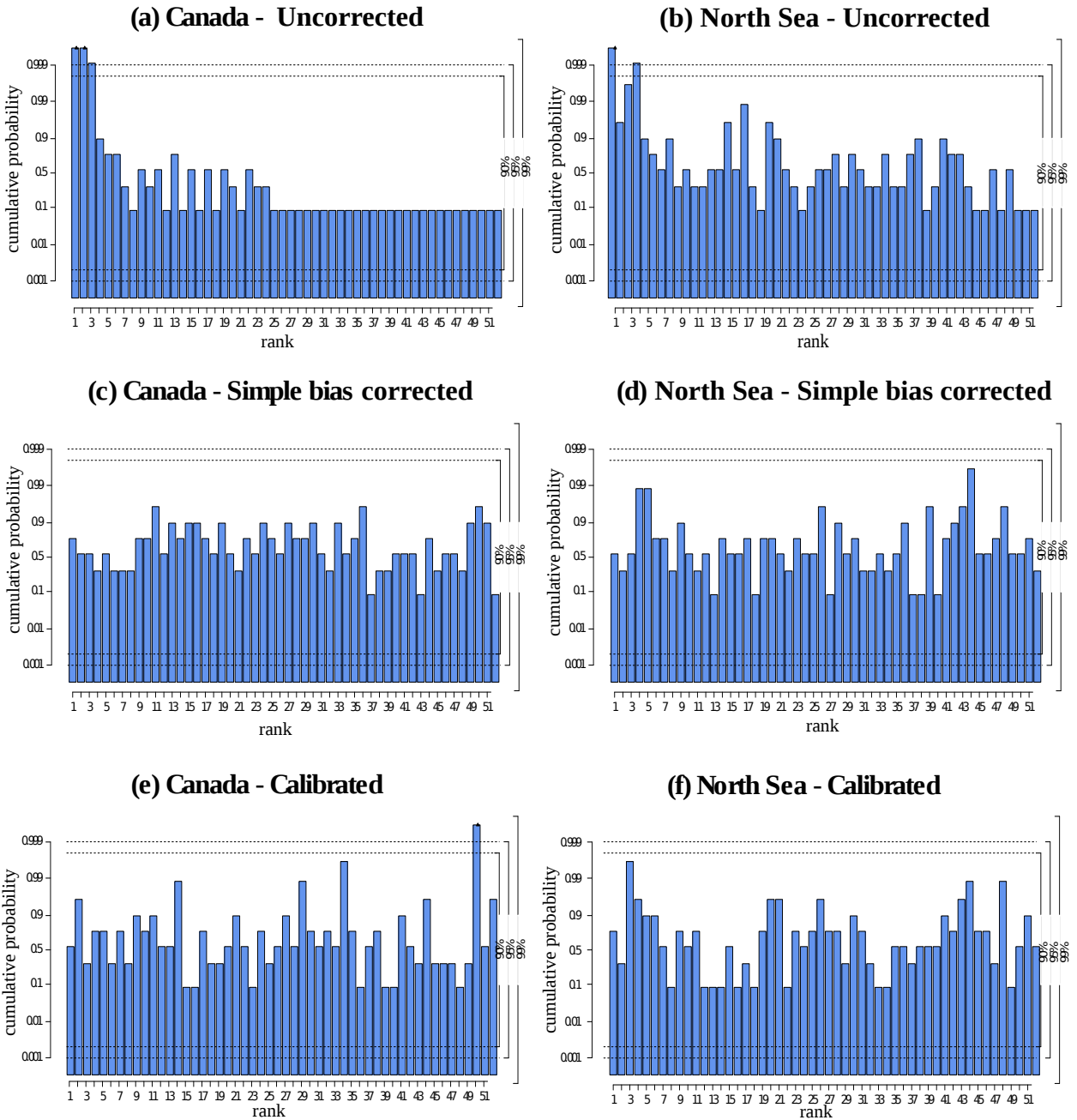
(d) Simple bias corrected – Uncorrected



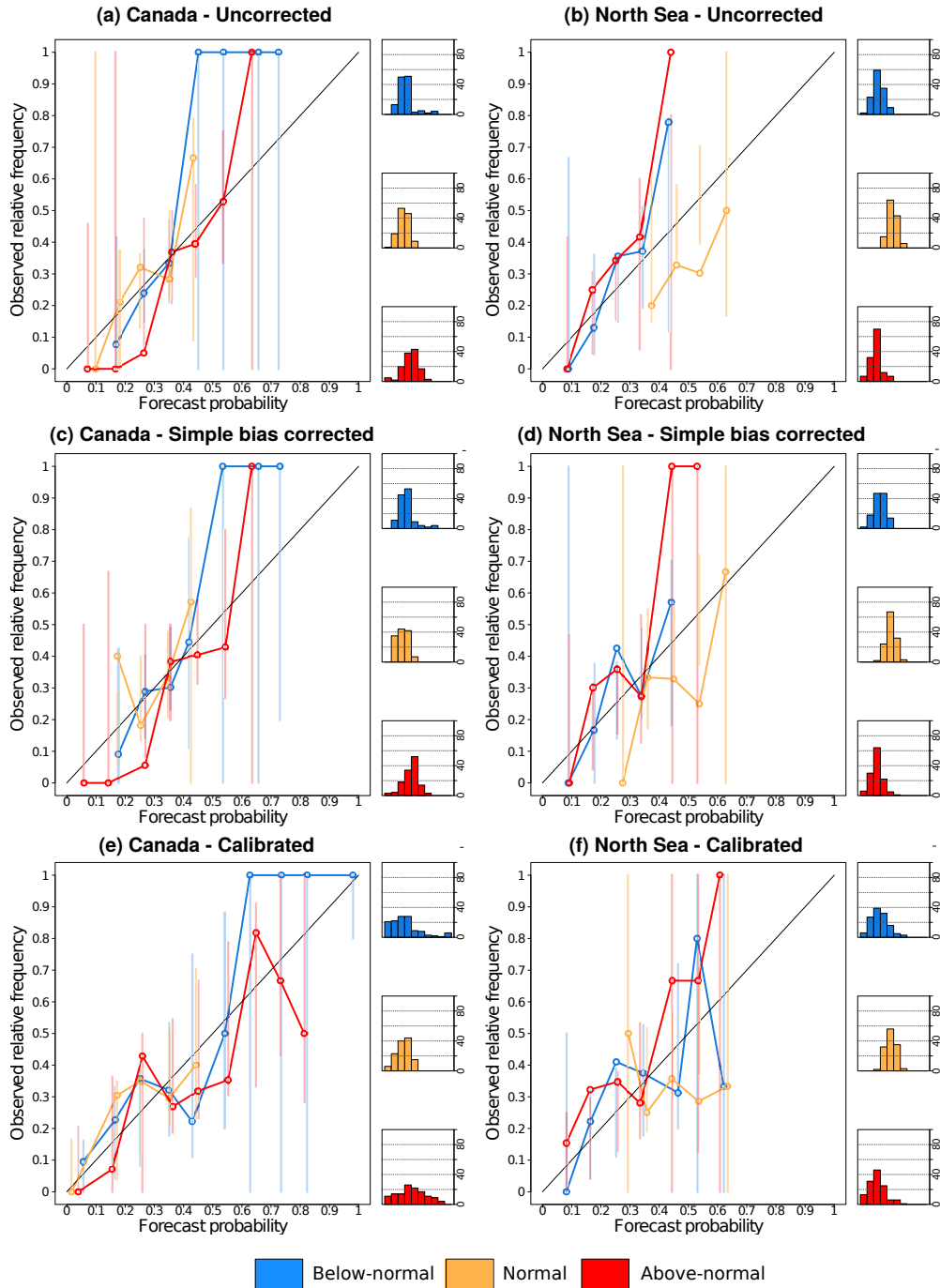
(e) Calibrated – Uncorrected



770 FIG. 4. Fair Ranked Probability Skill Score (RPSS) for tercile events of 10-m wind speed forecasts from
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 780 would have at least one bar that falls outside this interval.



781 FIG. 6. Reliability diagrams of 10-m wind speeds forecasts from ECMWF System 4 and ERA-Interim re-
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