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Optimal Hedging of Basket Barrier Options with Additive Models and its Application to Equity Value Separation Problem*

Yuji Yamada[†]

Abstract

At the heart of optimal hedging with additive models is to replicate the payoff of European basket options using separate options as close as possible. In this paper, we extend those technique for the case of path-dependent barrier options, where the mean square error of the payoffs between the basket barrier option and the sum of options on the individual assets is minimized over any smooth payoff functions. To this end, we propose to represent the underlying assets using the Brownian bridge decomposition and show that computations involving conditional expectations of basket barrier options boil down to those of unconditional expectations. This procedure enables us to provide an algorithm to compute the necessary and sufficient condition for the optimal hedging problem based on the Monte Carlo method. Then, we consider to apply our methodology to the Black-Cox type first passage time structural model, where a defaultable company possesses/runs multiple assets/projects and the default may occur the first time the asset value hits a certain lower threshold before the maturity. We formulate the equity value separation problem using additive models, in which individual equity values are introduced so that their sum approximates the total equity value as close as possible. It is also shown that any portion of total equity value may be assigned as an initial value of each individual equity when using the optimal smooth functions. Finally, we examine the contributions of individual equity values to default or survival by applying a certain normalization for conditional expectations via numerical experiments to illustrate our proposed methodology.

Keywords: Basket barrier options, Optimal hedging, Additive models, Smooth functions, First passage time structural models

1 Introduction

Options theory has played an important role not only for pricing applications but also for default risk modeling known as the so-called structural models. In this context, Merton (1974) provided a pioneering work which indicates that the default is assumed to occur when the firm value is below the face value (or the settlement value) of the debt at the maturity. Thus, the equity holders are considered to possess European call options whose values may be computed based on the Black-Scholes-Merton formula (Black and Scholes (1973), Merton (1973)) under the standard assumptions. On the other hand, Black and Cox (1976) extends the original model for the case where the default may occur the first time the asset value hits a certain lower threshold before the

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maturity, providing the equity holders with the down-and-out knock-out options. In either cases, developments of options theory lead to useful interpretations of the relations between default risk and equity values.

In this paper, we consider to generalize the optimal hedging technique with additive models in Yamada (2010–2012) to address the case of basket barrier options and discuss its application to the first passage time structural model. The optimal hedging problem is formulated as follows: Minimize the mean square error between the terminal payoff of basket options and the sum of smooth functions on individual underlying assets over any smooth payoff functions. Since each payoff may be replicated using standard European options as shown in Carr and Madan (2001), the resulting optimal hedging technique may be interpreted to provide a static hedging strategy for basket options using individual European options. It should be mentioned that our approach is related to the multivariate generalization of static hedging problem in Carr et. al (1998), in which the barrier option payoff on the single underlying asset is replicated based on the standard European options. Also note that, for hedging European basket options using options on individual assets, a super-hedging strategy consisting of the weighted sum of options with the same types (i.e., calls or puts) may be available as described in Hobson et al. (2005) and Su (2008), where the super-hedging strategy is to find a portfolio whose terminal value is always larger than that of the multivariate option. Moreover, there is another approach for hedging basket options using dynamic trading strategy, where a semi-definite programming based receding horizon control approach has been developed in Primbs (2009). To the best of our knowledge, hedging path-dependent basket options in terms of individual options has been uncovered in spite of the problem importance.

At the heart of optimal hedging with additive models in Yamada (2010–2012) is to solve the necessary and sufficient condition expressed in terms of a system of linear equations of conditional expectations, where the computations involving conditional expectations for the multivariate derivatives is shown to boil down to those of unconditional expectations based on the Independence Lemma (see e.g., S.E. Shreve (2004)) by assuming that the underlying assets follow multivariate geometric Brownian motions. On the other hand, one cannot apply the Independence Lemma directly to the path-dependent options due to the autocorrelations of Brownian motions. In this paper, we propose to represent the underlying assets using the Brownian bridge processes and prove that computations involving conditional expectations of basket barrier options may be reduced to those of unconditional expectations based on the Independence Lemma. This procedure enables us to provide an algorithm to solve the necessary and sufficient condition for the optimal hedging problem based on the Monte Carlo method.

Then, we consider to apply our methodology to the Black-Cox type first passage time structural model, where a defaultable company possesses/runs multiple assets/projects and the default may occur the first time the asset value hits a certain lower threshold before the maturity. We formulate the equity value separation problem using additive models, in which individual equity values are introduced so that their sum approximates the total equity value as close as possible under the risk neutral probability measure. It is also shown that any portion of total equity value may be assigned as an initial value of each individual equity when using the optimal smooth functions. Finally, we examine the contributions of individual equity values to default or survival by applying a certain normalization for conditional expectations based on the numerical experiments to illustrate our proposed methodology.

The rest of this paper is organized as follows: In Section 2, we formulate the optimal hedging problem with additive models, discuss some properties of the problem, and introduce the necessary and sufficient condition with respect to the optimal smooth functions. In Section 3, we show a key result that enables us to provide a computational procedure of optimal smooth functions for barrier options. In Section 4, we discuss how we can apply the optimal hedging technique with additive models for the equity value separation problem. Numerical

experiments are also provided to illustrate our proposed methodology in Section 5. Section 6 offers some concluding remarks.

2 Problem formulation

2.1 Optimal hedging problem with additive models

Let $S_{i,t}$, $i = 1, \dots, m$ be the values of m underlying assets at time $t \in [0, \infty)$ under a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ and filtration $\{\mathcal{F}_t\}_{t \in [0, \infty)}$. Also, let A_t and τ denote the value of basket portfolio and the first passage time with a lower threshold L given as

$$A_t := S_{1,t} + \dots + S_{m,t}, \quad t \in [0, \infty).$$

and

$$\tau := \arg \min \{t > 0 | A_t = L\}, \quad (2.1)$$

respectively. Then, the optimal hedging problem with additive models is defined as follows:

$$\min_{f_i \in \mathcal{S}} \mathbb{E} \left[\left\{ G_T - \sum_{i=1}^m f_i(S_{i,T}) \right\}^2 \right] \quad (2.2)$$

where G_T is the terminal value of basket barrier option at the expiration $T > 0$, e.g.,

$$G_T = I_{\{\tau > T\}} \times (A_T - K)^+, \quad (2.3)$$

in the case of a call with a strike price K , and \mathcal{S} a set of smooth functions. Note that the problem formulation in (2.2) clearly addresses the case of European options in Yamada (2010–2012) when $L = 0$ and that we will consider the case where the payoff is of the form of call options in (2.3) for simplicity of the notation.

There are some useful interpretations of the problem (2.2). First, the problem (2.2) may be considered as a multivariate generalization of static hedging problem for the single variate barrier option in Carr et. al (1998) using European options. Since each payoffs defined by optimal smooth functions may be replicated based on the standard European options as shown in Carr and Madan (2001), our problem aims to find a static hedging strategy using European options on individual underlying assets that approximates the basket barrier option payoff as close as possible in the minimum mean square error sense. Second, our problem may provide a separation problem of the equity value using individual smooth functions in the sense of first passage time structural models, in which the equity value of a defaultable company with multiple assets may be given by the value of basket barrier option in structural models. We will discuss this equity value separation problem with additive models later in this paper.

2.2 Necessary and sufficient condition

Similar to the European options case in Yamada (2012), we use the following lemma, which is introduced in Chapter 5 of Hastie and Tibshirani (1990):

Lemma 1 *Smooth functions f_1^*, \dots, f_m^* provide minimizers of problem (2.2), if and only if the following conditions are satisfied:*

$$\sum_{j=1}^m \mathbb{E} [f_j^*(S_{j,T}) | S_{i,T}] = \mathbb{E} [G_T | S_{i,T}], \quad i = 1, \dots, m \quad (2.4)$$

By taking unconditional expectation for both sides of equation (2.4), we have

$$\mathbb{E}[G_T] = \sum_{i=1}^m \mathbb{E}[f_i^*(S_{i,T})]. \quad (2.5)$$

Therefore, it holds that

$$\text{Var} \left[G_T - \sum_{i=1}^m f_i^*(S_{i,T}) \right] = \mathbb{E} \left[\left(G_T - \sum_{i=1}^m f_i^*(S_{i,T}) \right)^2 \right]. \quad (2.6)$$

Conditions (2.5) and (2.6) suggest that minimizing the mean square error corresponds to the variance minimization with zero mean constraint, i.e., the hedge is considered mean-variance optimal.

As in Yamada (2012), we assume that the underlying asset values are modeled via correlated geometric Brownian motions,

$$dS_{i,t} = \mu_i S_{i,t} dt + \sigma_i S_{i,t} dW_{i,t}, \quad i = 1, \dots, m, \quad (2.7)$$

where $W_{1,t}, \dots, W_{m,t}$ are (correlated) Brownian motions with $dW_{i,t} dW_{j,t} = \rho_{ij} dt$, $i, j = 1, \dots, m$, $i \neq j$ on $(\Omega, \mathcal{F}, \mathbb{P})$, and μ_i , σ_i and ρ_{ij} are given mean rate of return, volatility, and correlation coefficients parameters. One of the advantages for considering (2.7) is that there exists a dynamic trading strategy to replicate the terminal payoff $f_i^*(S_{i,T})$ once the optimal smooth functions are specified, similar to the Black-Scholes-Merton dynamic hedging strategy (Black and Scholes (1973), Merton (1973)).

Since the sigma-algebra generated by $W_{i,T}$ contains the same information as that by $S_{i,T}$ under the price dynamics in (2.7), condition (2.4) in Lemma 1 may be rewritten as

$$\sum_{j=1}^m \mathbb{E}[f_j^*(S_{j,T}) | W_{i,T}] = \mathbb{E}[G_T | W_{i,T}], \quad i = 1, \dots, m. \quad (2.8)$$

Noting that G_T is nonnegative in general, there exists a unique function \hat{g}_i satisfying $\mathbb{E}[G_T | W_{i,T}] = \hat{g}_i(W_{i,T})$ for each $i = 1, \dots, m$ ¹, i.e., condition (2.8) may further be written as

$$\sum_{j=1}^m \mathbb{E}[f_j^*(S_{j,T}) | W_{i,T}] = \hat{g}_i(W_{i,T}), \quad i = 1, \dots, m. \quad (2.9)$$

Let $p_{j|i}(w_j | w_i)$ denote conditional probability density functions (PDFs) of $W_{j,T}$ given $W_{i,T} = w_i \in \mathfrak{R}$, i.e.,

$$p_{j|i}(w_j | w_i) := \frac{1}{\sqrt{2\pi(1-\rho_{ij}^2)T}} \exp \left\{ -\frac{(w_j - \rho_{ij}w_i)^2}{2(1-\rho_{ij}^2)T} \right\}. \quad (2.10)$$

Then, the problem boils down to searching over a set of real functions, f_i^* , $i = 1, \dots, m$, satisfying the following condition for any $w_i \in \mathfrak{R}$:

$$f_i^*(S_{i,0}e^{\nu_i T + \sigma_i w_i}) + \sum_{j \neq i} \int f_j^*(S_{j,0}e^{\nu_j T + \sigma_j w_j}) p_{j|i}(w_j | w_i) dw_j = \hat{g}_i(w_i), \quad i = 1, \dots, m. \quad (2.11)$$

If we have a methodology to compute $\hat{g}_i(w_i)$ efficiently for any $w_i \in \mathfrak{R}$, then one can find f_i^* , $i = 1, \dots, m$ by constructing a set of linear equations with suitable discretization of (2.11) as shown in Yamada (2012).

¹See pp. 81 in Shreve (2004).

3 Computational method

In this section, we show a key result that enables us to express $\hat{g}(w_i)$ using unconditional expectation and provide the computational procedure.

3.1 Representation of underlying assets using Brownian bridge decomposition

To this end, we use an equivalent representation for the solution of SDE (2.7) based on the following Cholesky decomposition: Assume that $\Sigma dt \in \mathfrak{R}^{m \times m}$ denotes the covariance matrix of $dS_{j,t}/S_{j,t}$, $j = 1, \dots, m$ being decomposed as $\Sigma = \Lambda_1 \Lambda_1^\top dt \in \mathfrak{R}^{m \times m}$ by a lower triangular matrix $\Lambda_1 \in \mathfrak{R}^{m \times m}$, where

$$\Lambda_1 := \begin{bmatrix} \sigma_{11} & 0 & \cdots & 0 \\ \sigma_{21} & \sigma_{22} & \ddots & \vdots \\ \vdots & \vdots & \ddots & 0 \\ \sigma_{m1} & \sigma_{m2} & \cdots & \sigma_{mm} \end{bmatrix} \in \mathfrak{R}^{m \times m}, \quad \sigma_{11} = \sigma_1.$$

Then, there exist independent Brownian motions, $B_{j,t}$, $j = 1, \dots, m$, on $(\Omega, \mathcal{F}, \mathbb{P})$ such that

$$S_{j,t} = S_{j,0} \exp \left(\nu_j t + \sum_{k=1}^j \sigma_{jk} B_{k,t} \right), \quad \nu_j := \mu_j - \frac{\sigma_j^2}{2}, \quad j = 1, \dots, m, \quad (3.1)$$

Note that $S_{1,t}$ depends on $B_{1,t}$ only and therefore, $W_{1,t} \equiv B_{1,t}$ holds.

Although the lower triangular matrix is unique in the Cholesky decomposition, we have m different ways of representations by sorting $S_{j,t}$ as $j \in \{1, \dots, m\}$, $\{2, \dots, m, 1\}$, $\{3, \dots, m, 1, 2\}$, \dots , or $\{m, 1, \dots, m-1\}$. Let Λ_i be the lower triangle matrix of the Cholesky decomposition when $S_{j,t}$ is sorted as $j \in \{i, \dots, m, 1, \dots, i-1\}$, i.e., $\Lambda_i \Lambda_i^\top dt$ is the covariance matrix of $dS_{j,t}/S_{j,t}$, $j = i, \dots, m, 1, \dots, i-1$. Similar to the case $j \in \{1, \dots, m\}$, $S_{j,t}$ may be represented as in (3.1) by replacing $j \in \{1, \dots, m\}$ with $j \in \{i, \dots, m, 1, \dots, i-1\}$, and in this case, we see that $S_{i,t}$ depends on $B_{i,t}$ only and that $W_{i,t} \equiv B_{i,t}$ holds.

In the case of European options, i.e., $L = 0$, there exists an m -variate function h_i such that

$$G_T = (A_T - K)^+ = h_i(W_{i,T}, B_{i+1,T}, \dots, B_{m,T}, B_{1,T}, \dots, B_{i-1,T}). \quad (3.2)$$

Noting that $W_{i,t}$ is independent of the other Brownian motions, we can apply the Independence Lemma² that a function \hat{h}_i of a dummy variable $w_i \in \mathfrak{R}$,

$$\hat{h}_i(w_i) := \mathbb{E}[h_i(w_i, B_{i+1,T}, \dots, B_{m,T}, B_{1,T}, \dots, B_{i-1,T})], \quad (3.3)$$

satisfies the following condition:

$$\begin{aligned} \hat{h}_i(W_{i,T}) &= \mathbb{E}[G_T | W_{i,T}] \\ &= \mathbb{E}[h_i(W_{i,T}, B_{i+1,T}, \dots, B_{m,T}, B_{1,T}, \dots, B_{i-1,T}) | W_{i,T}]. \end{aligned} \quad (3.4)$$

Clearly, conditions (3.3) and (3.4) indicate that $\hat{g}_i \equiv \hat{h}_i$, $i = 1, \dots, m$ holds.

On the other hand, G_T depends on the entire paths of multivariate underlying assets, $S_{i,t}$, $i = 1, \dots, m$, in the case of barrier options, i.e., $L > 0$ in (2.3), and one cannot apply the Independence Lemma directly to derive equivalent conditions to (3.3) due to the autocorrelations of Brownian motions,

$$\mathbb{E}[W_{i,t} W_{i,s}] = s \wedge t, \quad i = 1, \dots, m, \quad s, t \in [0, T].$$

²See pp. 73 in Shreve (2004).

To identify the independence condition, here we propose to apply the Brownian bridge decomposition below.

Let $X_{i,t}$, $t \in [0, T]$ be a Brownian bridge process such that

$$X_{i,t} := B_{i,t} - \frac{t}{T}B_{i,T}, \quad i = 1, \dots, m. \quad (3.5)$$

Then, the independent Brownian motions, $B_{i,t}$, $i = 1, \dots, m$, may be expressed as

$$B_{i,t} = X_{i,t} + \frac{t}{T}B_{i,T}, \quad i = 1, \dots, m. \quad (3.6)$$

Condition (3.6) implies that each Brownian motion may be described as a function of t , Brownian bridge process $X_{i,t}$, and the terminal condition $B_{i,T}$. It can easily be shown that $X_{i,t}$ is a Gaussian process and that $B_{i,T}$ is independent of $X_{i,t}$, $i = 1, \dots, m$.

Consider the case that $S_{i,t}$ is sorted as $i \in \{1, \dots, m\}$. Using the Brownian bridge, $S_{i,t}$, $t \in [0, T]$ is rewritten as

$$S_{i,t} = S_{i,0} \exp \left[\nu_i t + \sum_{j=1}^i \sigma_{ij} \left(X_{j,t} + \frac{t}{T}B_{j,T} \right) \right], \quad i = 1, \dots, m. \quad (3.7)$$

Since each $S_{i,t}$ is a function of $(t, B_{1,T}, \dots, B_{m,T}, X_{1,t}, \dots, X_{m,t})$ with $B_{1,t} \equiv W_{1,t}$, the barrier option payoff G_T is a function of the following arguments:

$$W_{1,T}, B_{2,T}, \dots, B_{m,T}, \{t, X_{1,t}, \dots, X_{m,t}\}_{t \in [0, T]}, \quad (3.8)$$

i.e., there is a function ϕ_1 satisfying

$$G_T = \phi_1 \left(W_{1,T}, B_{2,T}, \dots, B_{m,T}, \{t, X_{1,t}, \dots, X_{m,t}\}_{t \in [0, T]} \right). \quad (3.9)$$

Noting that $W_{1,T}$ is independent of $B_{2,T}, \dots, B_{m,T}$ and $X_{1,t}, \dots, X_{m,t}$ for any $t \in [0, T]$, we can now apply the Independence Lemma as follows: The function \hat{g} which gives the conditional expectation,

$$\hat{g}_1(W_{1,T}) = \mathbb{E}[G_T | W_{1,T}] = \mathbb{E} \left[\phi_1 \left(W_{1,T}, B_{2,T}, \dots, B_{m,T}, \{t, X_{1,t}, \dots, X_{m,t}\}_{t \in [0, T]} \right) \middle| W_{1,T} \right] \quad (3.10)$$

satisfies the following condition for a nonrandom dummy variable $w_1 \in \mathfrak{R}$:

$$\hat{g}_1(w_1) = \mathbb{E} \left[\phi_1 \left(w_1, B_{2,T}, \dots, B_{m,T}, \{t, X_{1,t}, \dots, X_{m,t}\}_{t \in [0, T]} \right) \right]. \quad (3.11)$$

Similarly, we can construct a function ϕ_k , $k = 2, \dots, m$ by sorting $S_{i,t}$ as $i \in \{2, \dots, m, 1\}$, $\{3, \dots, m, 1, 2\}$, \dots , and $\{m, 1, \dots, m-1\}$ and applying the same procedure. As a result, we obtain the following theorem:

Theorem 1 *Assume that G_T is given by (2.3). Then, for each $i \in \{1, 2, \dots, m\}$ and a (nonrandom) dummy variable $w_i \in \mathfrak{R}$, there exists a function ϕ_i satisfying*

$$\hat{g}_i(w_i) = \mathbb{E} \left[\phi_i \left(w_i, \{B_{j,T}\}_{j=1, \dots, m, j \neq i}, \{t, X_{1,t}, \dots, X_{m,t}\}_{t \in [0, T]} \right) \right]. \quad (3.12)$$

3.2 Computational procedure for barrier options

We see that, for any given real number $w_i \in \mathfrak{R}$, $i = 1, \dots, m$, $\hat{g}_i(w_i)$ is computed using a Monte Carlo method for unconditional expectation (3.12). In this case, we need to generate Brownian bridge processes, $X_{i,t}$, $i = 1, \dots, m$ based on the discrete observations, but in fact, we can generate the Brownian bridge sample paths by using the Brownian motion sample paths. The total procedure is summarized as follows:

Step 0: Let N be appropriate number of observations and $\delta_T := T/N$. For given $i \in \{1, \dots, m\}$ and w_i , repeat Steps 1–3 below.

Step 1: For the time periods, $t_n := n \cdot \delta_T$, $n = 0, 1, \dots, N$, generate sample paths of m -dimensional (independent) Brownian motions, B_{i,t_n} , $i = 1, \dots, m$.

Step 2: For each B_{i,t_n} , compute the Brownian bridge processes, as

$$X_{i,t_n} := B_{i,t_n} - \frac{t_n}{t_N} B_{i,t_N}, \quad i = 1, \dots, m, \quad n = 0, 1, \dots, N. \quad (3.13)$$

Step 3: Substitute X_{i,t_n} and B_{i,t_N} into ϕ_i in (3.12) for $i = 1, \dots, m$ as

$$\phi_i \left(w_i, \{B_{j,t_N}\}_{j=1, \dots, m, j \neq i}, \{t_n, X_{1,t_n}, \dots, X_{m,t_n}\}_{n=0, 1, \dots, N} \right). \quad (3.14)$$

Step 4: After repeating Steps 1–3 sufficiently many times, compute the average of equation (3.14) to obtain an estimated value of $\hat{g}_i(w_i)$, $i = 1, \dots, m$.

4 Equity value separation using additive models

In this section, we demonstrate how we can apply the optimal hedging technique with additive models for the equity value separation problem. To this end, we consider a defaultable firm owing a debt paying no coupon until the maturity $T > 0$. Also, assume that the firm consists of m individual assets, $S_{i,t}$, $i = 1, \dots, m$, where the firm's asset value is denoted by

$$A_t := S_{1,t} + \dots + S_{m,t}, \quad t \in [0, \infty). \quad (4.1)$$

4.1 Individual equity values

Let E_t , $t \in [0, T]$ denote the equity shareholders' value of the firm whose terminal value value, E_T , depends on whether or not the firm is default. For example, in the case of Merton's structural model, the default is defined when the terminal asset value A_T is below D and the terminal equity value may be given as

$$E_T = (A_T - D)^+, \quad (4.2)$$

where $D > 0$ is the total amount of face value. On the other hand, in the case of first passage time structural model, the default is defined by the first time τ the asset value hits a certain lower threshold $L > 0$ before the maturity $T > 0$, i.e., $\tau < T$, and E_T is given as

$$E_T = I_{\{\tau > T\}} \times (A_T - D)^+. \quad (4.3)$$

As in (4.1), the asset values are considered to be additive, i.e., the total asset value of the firm corresponds to the sum of individual asset values. Then, how about equity values? Are equity values considered to be additive? Or in the similar content, is the equity value, denoted by E_t , splittable into individual equity values related to each assets? One candidate to define such equity values is to apply the Merton's structural model for individual asset values as well as the total asset value of (4.2). However, the usage of individual equity values defined by the Merton's structural model may always lead to an overestimation of the total equity value from

the following discussion: Let $\tilde{\mathbb{E}}$ be conditional expectation under risk neutral probability measure $\tilde{\mathbb{P}}$, and define individual equity values, $E_{i,T}$, $i = 1, \dots, m$, by

$$E_{i,t} = e^{-r(T-t)} \tilde{\mathbb{E}} [E_{i,T} | \mathcal{F}_t], \quad E_{i,T} = (S_{i,T} - D_i)^+, \quad (4.4)$$

where D_i , $i = 1, \dots, m$ denote face values of the debts assigned for each assets such that $D_1 + \dots + D_m = D$. Then it holds that

$$\sum_{i=1}^m E_{i,T} = \sum_{i=1}^m (S_{i,T} - D_i)^+ \geq \left(\sum_{i=1}^m (S_{i,T} - D_i) \right)^+ = (A_T - D)^+ = E_T. \quad (4.5)$$

Condition (4.5) implies that the sum of individual equity values (under the Merton's structural model) is always greater than or equal to the total equity value.

In stead of applying the Merton's structural model for individual assets directly, here we consider to introduce individual equity values so that their sum approximates the total equity value as close as possible. A formal definition may be described as follows:

Definition 1 Let $\mathcal{F}_{i,t}$, $t \in [0, T]$, $i = 1, \dots, m$ be filtration generated by each individual asset process, $S_{i,t}$. Then, the individual equity value, denoted by $E_{i,t}$, is defined as a stochastic process adapted to $\mathcal{F}_{i,t}$ whose sum approximates the total equity value E_t as close as possible, i.e.,

$$E_t \simeq E_{1,t} + \dots + E_{m,t}, \quad t \in [0, T]. \quad (4.6)$$

We discuss that the individual equity values satisfying Definition 1 may be obtained using a solution of the optimal hedging problem with additive models. To this end, we define the following optimal hedging problem with additive models under the risk neutral measure $\tilde{\mathbb{P}}$:

$$\min_{f_i \in \mathcal{S}} \tilde{\mathbb{E}} \left[\left\{ E_T - \sum_{i=1}^m f_i(S_{i,T}) \right\}^2 \right], \quad (4.7)$$

Also, let f_i^* , $i = 1, \dots, m$ denote optimal smooth functions for the optimal hedging problem (4.7).

4.2 Optimal hedging problem under risk neutral measure

At first, we assume that $L = 0$ and consider optimal smooth functions, f_i^* , $i = 1, \dots, m$, with E_T defined by (4.2). For these optimal smooth functions, there exists a function $h_{i,t}^*$ such that

$$h_{i,t}^*(S_{i,t}) := \tilde{\mathbb{E}} [f_i^*(S_{i,T}) | S_{i,t}], \quad t \in [0, T], \quad i = 1, \dots, m. \quad (4.8)$$

from the Markov property (see, e.g., Shreve (2004)), where $h_{i,T}^* = f_i^*$. Then, an interesting question is to ask if $h_{i,t}^*$, $i = 1, \dots, m$ provide optimal smooth functions of the problem,

$$\min_{h_{i,t} \in \mathcal{S}} \tilde{\mathbb{E}} \left[\left\{ E_t - \sum_{i=1}^m h_{i,t}(S_{i,t}) \right\}^2 \right], \quad (4.9)$$

for any given $t \in [0, T]$. This statement holds true from the following theorem:

Theorem 2 (Yamada 2012) For any given $t \in [0, T]$, the smooth functions $h_{i,t}^*$, $i = 1, \dots, m$ of (4.8) provide minimizers for the problem (4.9).

Clearly, $h_{i,t}^*(S_{i,t})$ is adapted to $\mathcal{F}_{i,t}$ for $t \in [0, T]$ and the sum of $h_{i,t}^*(S_{i,t})$ approximates the total equity value $E_{i,t}$ in the mean square optimal sense from Theorem 2. Therefore, we conclude that $h_{i,t}^*(S_{i,t})$, $i = 1, \dots, m$ provides a candidate of the individual equity value in Definition 1.

In the case of $L > 0$ where the terminal equity value is defined by (4.3), one cannot obtain the property corresponding to Theorem 2 due to the path-dependent features in τ . Although the further investigation may be our future work, here we show that the optimal smooth functions, f_i^* , $i = 1, \dots, m$, of the problem (4.7) may provide individual equities for the total equity value being approximated by a multivariate European option. Let $\psi^*(S_{1,T}, \dots, S_{m,T})$ be the optimal projection of E_T onto the sigma algebra generated by the terminal values of individual assets, $\sigma(S_{1,T}, \dots, S_{m,T})$, under the risk neutral probability measure, $\tilde{\mathbb{P}}$, i.e., ψ^* is a solution of the following problem:

$$\min_{\psi_i \in \mathcal{S}^m} \tilde{\mathbb{E}} \left[\{E_T - \psi(S_{1,T}, \dots, S_{m,T})\}^2 \right] \quad (4.10)$$

where \mathcal{S}^m is a set of m -dimensional smooth functions. Then, it can readily be confirmed that ψ^* provides conditional expectation of E_T given $\sigma(S_{1,T}, \dots, S_{m,T})$, i.e.,

$$\psi^*(S_{1,T}, \dots, S_{m,T}) = \tilde{\mathbb{E}}[E_T | S_{1,T}, \dots, S_{m,T}]. \quad (4.11)$$

Using ψ^* in (4.11), we demonstrate that the optimal hedging problem (4.7) boils down to the following problem:

$$\min_{f_i \in \mathcal{S}} \tilde{\mathbb{E}} \left[\left\{ \psi^*(S_{1,T}, \dots, S_{m,T}) - \sum_{i=1}^m f_i(S_{i,T}) \right\}^2 \right]. \quad (4.12)$$

Since $f_{i,t}^*$ are optimal smooth functions for the problem (4.7), the following conditions are satisfied:

$$\sum_{j=1}^m \tilde{\mathbb{E}} [f_j^*(S_{j,T}) | S_{i,T}] = \tilde{\mathbb{E}} [E_T | S_{i,T}], \quad i = 1, \dots, m \quad (4.13)$$

Based on the tower property for conditional expectations, the right hand side of (4.13) may be written as

$$\begin{aligned} \tilde{\mathbb{E}} [E_T | S_{i,T}] &= \tilde{\mathbb{E}} \left[\tilde{\mathbb{E}} [E_T | S_{1,T}, \dots, S_{m,T}] | S_{i,T} \right] \\ &= \tilde{\mathbb{E}} [\psi^*(S_{1,T}, \dots, S_{m,T}) | S_{i,T}]. \end{aligned} \quad (4.14)$$

Therefore, it holds that

$$\sum_{j=1}^m \tilde{\mathbb{E}} [f_j^*(S_{j,T}) | S_{i,T}] = \tilde{\mathbb{E}} [\psi^*(S_{1,T}, \dots, S_{m,T}) | S_{i,T}], \quad i = 1, \dots, m, \quad (4.15)$$

and hence, we conclude that $f_{i,t}^*$ provides optimal smooth functions for the problem (4.12) as well as those for the problem (4.7).

The above argument (together with Theorem 2) implies that the functions, \hat{h}_i , $i = 1, \dots, m$, satisfying

$$\tilde{\mathbb{E}} [f_i^*(S_{i,T}) | \mathcal{F}_t] = \tilde{\mathbb{E}} [f_i^*(S_{i,T}) | S_{i,t}] = \hat{h}_i(S_{i,t}), \quad i = 1, \dots, m, \quad (4.16)$$

provide the optimizers of

$$\min_{h_{i,t} \in \mathcal{S}} \tilde{\mathbb{E}} \left[\left\{ \hat{E}_t - \sum_{i=1}^m h_{i,t}(S_{i,t}) \right\}^2 \right], \quad (4.17)$$

where \hat{E}_t is defined by

$$\hat{E}_t := \tilde{\mathbb{E}}[\psi^*(S_{1,T}, \dots, S_{m,T}) | \mathcal{F}_t] \quad (4.18)$$

and $\psi^*(S_{1,T}, \dots, S_{m,T})$ is the mean square optimal approximation of E_T using conditional expectation (4.11). Since $\hat{E}_t \neq E_t$ in general, f_i^* , $i = 1, \dots, m$ do not necessarily provide desired solutions of individual equity values with respect to the total equity value E_t . As mentioned earlier in this subsection, this is mainly due to the path-dependence of barrier options and the further investigation may be required in our future work. In this paper, we regard $\hat{h}_i(S_{i,t})$ in (4.16) as proxies of individual equity values and illustrate the relations between the individual equity values and the equity debt ratio via numerical experiments in the next section.

4.3 Initial value assignment

Before providing the numerical experiments, we show that any portion of total equity value may be assigned as an initial value of each individual equity when using the optimal smooth functions of (4.7). Let \bar{f}_i be another smooth function defined by

$$\bar{f}_i(S_{i,T}) := f_i^*(S_{i,T}) + \eta_i \mathbb{E}[E_T] - \mathbb{E}[f_i^*(S_{i,T})], \quad i = 1, \dots, m \quad (4.19)$$

where η_i is any parameter satisfying $\eta_1 + \dots + \eta_m = 1$. Then, the following condition holds from (2.5):

$$\sum_{i=1}^m \bar{f}_i(S_{i,T}) = \sum_{i=1}^m f_i^*(S_{i,T}). \quad (4.20)$$

Condition (4.20) suggests that \bar{f}_i , $i = 1, \dots, m$ defined in (4.19) are optimizers of the problem (4.7) as well.

Instead of using f_i^* , we define $E_{i,t}$ as

$$E_{i,t} := e^{-r(T-t)} \mathbb{E}[\bar{f}_i(S_{i,T}) | \mathcal{F}_t], \quad i = 1, \dots, m, \quad (4.21)$$

Then, at the maturity $t = T$, the sum of $E_{i,T}$, $i = 1, \dots, m$ provides an optimal approximation of the terminal equity value E_T , i.e., $E_T \simeq \sum_{i=1}^m E_{i,T}$, in the sense of minimum mean square errors³. On the other hand, the initial value may be computed as

$$E_{i,0} = e^{-rT} \mathbb{E}[\bar{f}_i(S_{i,T})] = e^{-rT} \mathbb{E}[f_i^*(S_{i,T})] + \eta_i e^{-rT} \mathbb{E}[E_T] - e^{-rT} \mathbb{E}[f_i^*(S_{i,T})] = \eta_i E_0, \quad (4.22)$$

which indicates that individual equity values may be constructed so that their initial values are assigned with arbitrarily portions of the total equity value at time $t = 0$. This result may be useful for the situation that the initial values of individual equities are given a priori. For example, in the case of merger of two companies, their individual equity values may be observable until the merger is carried out and one can set the initial values of individual equities using the ones right before the merger. Even though individual equity values become unobservable after the merger, it may be possible to use the individual equity values estimated from the optimal hedging technique with additive models provided in this section.

5 Numerical experiment

Here we illustrate our optimal hedging technique for the equity value separation problem via numerical experiments. Assume that a firm possesses five individual assets, $S_{i,t}$, $i = 1, \dots, 5$, which follow SDEs in (2.7) with

³For the Merton's structural model of $E_T = (A_T - D)^+$, $E_t \simeq \sum_{i=1}^m E_{i,t}$ holds for any $t \in [0, T]$ in the sense of minimum mean square errors.

volatility and correlation parameters of Table 5.1, where we note that the higher the asset number the smaller the volatility. We also suppose that $\mu_i = 0$, $i = 1, \dots, 5$ and $r = 0$ to investigate the effect of volatility of each asset on default or survival. In this case, the physical probability measure provides the risk neutral probability measure. We set $T = 3$ (years), $S_{i,0} = 20$, $i = 1, \dots, m$, and $A_0 = \sum_{i=1}^5 S_{i,0} = 100$, whereas the face value D of the zero-coupon debt (maturing at T) is varied according to the debt/equity ratio, denoted by D/E, as D/E = 0.5, 1, 2, 4, e.g., $D = A_0 \times 0.5 / (1 + 0.5)$ when D/E = 0.5. The default threshold L is given by $L = D$ for the first time structural model (Barrier option) and $L = 0$ for the Merton's structural model (European option).

Table 5.1: Volatilities and correlations

	S_1	S_2	S_3	S_4	S_5
S_1	1				
S_2	0.476	1			
S_3	0.291	0.341	1		
S_4	0.315	0.287	0.389	1	
S_5	0.298	0.346	0.406	0.457	1
Volatility	0.549	0.421	0.307	0.232	0.157

Table 5.2: Default probabilities

D/E	0.5	1	2	4
European option (%)	0.29	5.57	21.6	38.1
Barrier option (%)	0.46	9.47	35.8	63.2

For finding optimal smooth functions of the problem (4.7), we need to compute $\hat{g}(w_i)$ by applying Steps 1)–4) in Section 3.2 for each $i \in \{1, \dots, 5\}$ and a dummy variable w_i . Since we can use the same set of random variables to compute the average of ϕ_i in Step 4 for any $w_i \in \mathfrak{R}$, the number of random variables required for the total computation is the same as that of Monte Carlo simulation for independent Brownian motion sample paths. We generate 30,000 sample paths of five dimensional independent Brownian motions, where the number of observations is assumed to be $N = 150$. In this case, the basic time period is given by $\delta_T = T/N = 1/50$ ($\simeq 1$ week) and the total number of independent random variables generated for this simulation is $5 \times 30,000 \times 150 = 225 \times 10^5$. Then, we compute the Brownian bridge processes as in (3.6). Note that, when $L = 0$ for the Merton's structural model, we need the terminal values of five dimensional Brownian motions only. Table 5.2 summarizes the default probabilities with respect to D/E = 0.5, 1, 2, 4, i.e., $\mathbb{P}(A_T \leq D)$ for the Merton's structural model entitled (European option) and $\mathbb{P}(\tau \leq T)$ for the first passage time structural model (Barrier option). We see that the default probability with D/E = 4 for the first passage time structural model is quite high, whereas those with D/E = 0.5 are almost zero for both cases. Also, the default probabilities for the first passage time structural model are more than 50% higher than those for the Merton's structural model, i.e., more than 1/3 of sample paths related to default hit the threshold $L = D$ before the maturity in this example.

Table 5.3: Squared correlation coefficient

D/E	0.5	1	2	4
European option (%)	100.0	99.9	99.5	98.7
Barrier option (%)	100.0	99.7	96.6	84.6

Table 5.4: Minimum mean square errors

D/E	0.5	1	2	4
European option	0.103	1.86	10.5	21.6
Barrier option	0.161	7.30	68.3	254

Tables 5.3 and 5.4 illustrate the squared correlation coefficients corresponding to the coefficients of deter-

mination and the minimum mean square errors, respectively. The two rows in both figures are those between E_T and the sum of optimal smooth functions for the Merton's and the first passage time structural models. (European and Barrier options). From these results, although the higher the default probabilities the worse the sizes of hedge errors for both the squared correlation coefficients and the minimum mean square errors, we can get a reasonably good hedge performance even with $D/E = 4$ for the first passage time structural model, e.g., the squared correlation coefficient is about 85% indicating that 85% of total fluctuation may be explained by the sum of optimal smooth functions.

Finally, we investigate the effect of default or survival on individual equity values estimated from the optimal hedging technique with additive models. Here we compute normalized conditional expectations given survival or default related to the first passage time τ , where the normalized conditional expectation of $E_{i,T}$ given survival, $\tau > T$, is defined as a quantity being proportional to the following value:

$$\mathbb{E} [I_{\{\tau > T\}} E_{i,T}] \quad (5.1)$$

Similarly, the one given default is proportional to

$$\mathbb{E} [I_{\{\tau \leq T\}} E_{i,T}] \quad (5.2)$$

Since the larger (5.1) or (5.2) the higher correlation between the individual equity value and the default or the survival event, conditions (5.1) and (5.2) with a certain normalization may provide contribution rates of the total equity value given survival or default. Let $V_i^{(s)}$ and $V_i^{(d)}$, $i = 1, \dots, m$ denote such normalized conditional expectations given by

$$\begin{aligned} V_i^{(s)} &= a \cdot \mathbb{E} [I_{\{\tau > T\}} E_{i,T}] + b \\ V_i^{(d)} &= a \cdot \mathbb{E} [I_{\{\tau \leq T\}} E_{i,T}] + b \end{aligned}$$

where a and b are constant parameters such that

$$\sum_{i=1}^m V_i^{(s)} = 1, \quad \sum_{i=1}^m V_i^{(d)} = 0.$$

Note that conditions (5.1) and (5.2) may be estimated once smooth functions, f_i^* , $i = 1, \dots, m$, are specified using the same set of sample paths generated for the Monte Carlo simulation.

Fig. 5.1 shows the relation between the asset number vs. normalized conditional expectations given survival or default for European options. The value of each bin in the left hand side denotes the normalized conditional expectation $V_i^{(s)}$ given survival (i.e., $A_T > 0$) for $D/E = 0.5, 1, 2, 4$ with respect to the asset number $i = 1, \dots, 5$. Since the sum of each bin with the same D/E is 1, the vertical axis may be interpreted to provide a contribution rate of asset i to the equity value of the firm conditioned on survival. From the figure, we see that each equity value is indifferent on average for $D/E = 0.5$ in which default probability is very small, whereas the larger the volatility the higher contribution to the equity value and the difference becomes more significant with the larger D/E . This indicates that, in the case where the firm is still survival under default risk, the source of excess profit is mainly brought from a high volatility asset (or project) and the contribution of low volatility asset (or project) is not so significant. The right hand side of Fig 5.1 shows the normalized conditional expectations, $V_i^{(d)}$, $i = 1, \dots, 5$, given the firm is default for $D/E = 0.5, 1, 2, 4$ and $i = 1, \dots, 5$. Note that $V_i^{(d)}$ for each asset can be negative even though the asset value is always positive. From the figure, we see that a smaller (and negative) value is obtained when volatility is higher and that the lower the volatility the higher the contribution to excess profit when the firm is default.

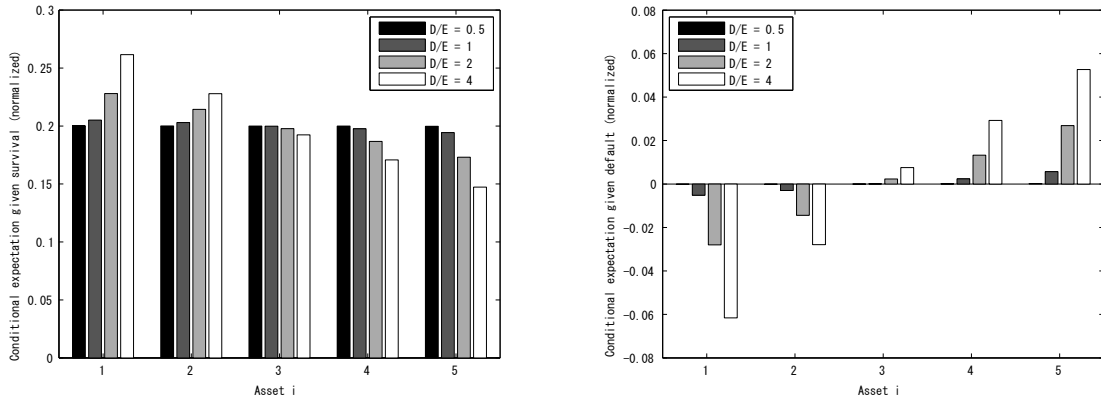


Fig. 5.1: Normalized conditional expectations given survival (Left; $V_s^{(i)}$, $i = 1, \dots, 5$) or default (Right; $V_d^{(i)}$, $i = 1, \dots, 5$) for European options

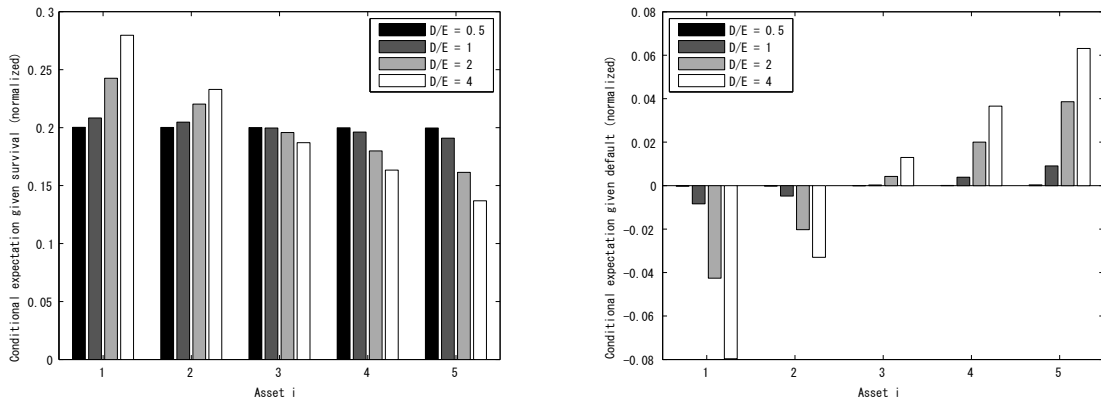


Fig. 5.2: Normalized conditional expectations given survival (Left; $V_i^{(s)}$, $i = 1, \dots, 5$) or default (Right; $V_i^{(d)}$, $i = 1, \dots, 5$) for Barrier options

Fig. 5.2 shows normalized conditional expectations for Barrier options, where the left and right hand plots compare $V_i^{(s)}$ given survival and $V_i^{(d)}$ given default, respectively, for $D/E = 0.5, 1, 2, 4$ and $i = 1, \dots, 5$. Although the default probability for Barrier options in Table 5.2 is much higher than that for European options for every D/E , one can observe a similar tendency for the normalized conditional expectations given survival as shown in the left hand side of Fig. 5.2, i.e., the conditional expectations given survival depends on D/E and a higher contribution to the excess profit is observed when volatility is higher. On the other hand, for the conditional expectations given default, the difference is a little emphasized compared to those for European options, in particular for $D/E = 4$, and the higher the volatility the higher impact on the default.

6 Conclusion

In this paper, we first generalized the optimal hedging technique with additive models in Yamada (2010–2012) for the case of path-dependent barrier options, where the problem is to find optimal payoff functions on individual options to replicate the payoff of path-dependent basket option as close as possible. To solve the necessary and sufficient condition for the optimal hedging problem, we proposed to represent the underlying assets based on the Brownian bridge decomposition and showed that computations involving conditional expectations of basket barrier options may be reduced to those of unconditional expectations. We also provided an algorithm to compute the unconditional expectations using a Monte Carlo method. We then applied our methodology to the Black-Cox type first passage time structural model, where a defaultable company is assumed to possess/run multiple assets/projects and the default may occur the first time the asset value hits a certain lower threshold before the maturity. We formulated the equity value separation problem using additive models and introduced individual equity values so that their sum approximates the total equity value as close as possible. It was also shown that any portion of total equity value may be assigned as an initial value of each individual equity when using the optimal smooth functions. Numerical experiments were also included to illustrate our methodology, in which we estimated contributions of individual equity values to default or survival by applying a certain normalization for conditional expectations.

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