# Genome-wide association studies for feedlot and growth traits in $cattle^1$

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**ABSTRACT:** A genome wide-association study for production traits in cattle was carried out using genotype data from the 10K Affymetrix (Santa Clara, CA) and the 50K Illumina (San Diego, CA) SNP chips. The results for residual feed intake (RFI), BW, and hip height in 3 beef breed types (Bos indicus, Bos taurus, and B. indicus  $\times$  B. taurus), and for stature in dairy cattle, are presented. The aims were to discover SNP associated with all traits studied, but especially RFI, and further to test the consistency of SNP effects across different cattle populations and breed types. The data were analyzed within data sets and within breed types by using a mixed model and fitting 1 SNP at a time. In each case, the number of significant SNP was more than expected by chance alone. A total of 75 SNP from the reference population with 50K chip data were significant (P < 0.001) for RFI, with a false discovery rate

of 68%. These 75 SNP were mapped on 24 different BTA. Of the 75 SNP, the 9 most significant SNP were detected on BTA 3, 5, 7, and 8, with  $P < 6.0 \times 10^{-5}$ . In a population of Angus cattle divergently selected for high and low RFI and 10K chip data, 111 SNP were significantly (P < 0.001) associated with RFI, with a false discovery rate of 7%. Approximately 103 of these SNP were therefore likely to represent true positives. Because of the small number of SNP common to both the 10K and 50K SNP chips, only 27 SNP were significantly (P < 0.05) associated with RFI in the 2 populations. However, other chromosome regions were found that contained SNP significantly associated with RFI in both data sets, although no SNP within the region showed a consistent effect on RFI. The SNP effects were consistent between data sets only when estimated within the same breed type.

Key words: beef and dairy cattle, body weight, feed intake, height, residual feed intake, single nucleotide polymorphism

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#### **INTRODUCTION**

In genomic selection, the estimation of breeding values is based on genetic markers. This is particularly

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useful for traits that are very expensive to measure, such as residual feed intake (**RFI**). Genomic selection relies on linkage disequilibrium (**LD**) between genetic markers, such as SNP, and QTL that affect the trait. This LD generates an association between some markers and the trait. In beef cattle, some studies (Barendse et al., 2007; Nkrumah et al., 2007; Sherman et al., 2009) have reported associations between markers and RFI. For instance, Barendse et al. (2007), using a commercial SNP chip containing approximately 10,000 (**10K**) SNP, analyzed 8,786 polymorphic SNP in 189 Australian beef cattle, chosen on the basis of being phenotypically high and low for RFI, and they detected 161 SNP associated with RFI at P < 0.01. Development of a commercial 50,000 (**50K**) SNP chip provided

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the opportunity to conduct a more powerful genomewide association study (**GWAS**) for RFI.

Whether markers associated with a trait are to be used for genomic selection or for mapping the QTL to a chromosomal region, it is necessary to confirm in independent populations the associations that have been discovered in 1 population. Often such attempts at confirmation have been unsuccessful (e.g., Pryce et al., 2010b). Failure to confirm associations could be due to 3 reasons: 1) the original discovery was a false positive, 2) the association is specific to that breed either because the QTL does not segregate in another breed or because the phase or strength of LD differs between breeds, or 3) there is a lack of statistical power in either the discovery or validation population, or in both populations. In this paper, the importance of these 3 reasons for failure to validate associations is examined.

Ideally, an SNP allele that is associated with an increase in a trait, such as RFI, in 1 breed will also be associated with an increase in other breeds. However, cattle breeds differ in the LD phase between markers in the 50K SNP chip (de Roos et al., 2008, 2009), so it is expected that they will differ in the LD phase between SNP and QTL. This might mean that the association between SNP and QTL is still significant in the second breed but reversed in sign. More likely, the association is simply weak and not significant. If the QTL is segregating in the second breed, it is likely that different SNP, close to the QTL, will now show a significant association with the QTL. Therefore, to confirm the detection of a QTL, 3 types of evidence are investigated: Is the SNP, discovered to be associated with the trait in the discovery population, significantly associated with the trait in the second (validation) population; is the direction of the association the same; and is there another SNP in the same vicinity that shows a significant association with the trait under investigation?

In addition to RFI, we present data on GWAS for BW and hip height, which are model quantitative traits and for which data are more widely available than for RFI. To do this, GWAS using data from 3 breed types of beef cattle [*Bos taurus* (**Bt**), *Bos indicus* (**Bi**), and their crosses] and 2 SNP chips (10K and 50K) were conducted. A GWAS for stature or height in dairy cattle using the same 50K SNP chip was also carried out to determine if associations could be confirmed across dairy and beef breeds.

Genetic correlations between traits could be due to QTL that have pleiotropic effects on multiple traits or could be due to closely linked QTL, each affecting different traits. Nkrumah et al. (2007) found QTL affecting DMI, feed conversion ratio, and ADG together in similar locations on the bovine genome map. Indeed, Moore et al. (2009) pointed out that it is important to investigate the effects of QTL on other traits when studying the molecular basis of RFI to avoid unfavorably correlated responses when selecting for RFI. Therefore, we investigated if SNP associated with RFI were also associated with feed intake or growth rate. If a trait has no phenotypic correlation with RFI, then SNP should rarely be associated with both traits entirely because of false discovery. Therefore, if SNP have a significant effect on 2 uncorrelated traits more often than expected by chance, this is evidence that the association is real.

The objectives of this study were to detect SNP associated with RFI, growth, and height in 3 breed types of beef cattle (Bt, Bi, and their crosses) and in dairy cattle, and to validate SNP effects across different data sets and breed types.

#### MATERIALS AND METHODS

Animal Care and Use Committee approval was not obtained for this study because no new animals were handled in this experiment.

#### SNP Data

The SNP marker data used in this study were obtained from 2 different sources: one used the BovineSN-P50K BeadChip (Illumina, San Diego, CA) and the other one used the Parallele SNP10K chip (Affymetrix, Santa Clara, CA). The 50K SNP were at random positions with approximately equal spacing (median interval of 37 kbp) along the bovine genome (Matukumalli et al., 2009). The 10K SNP were with mean intermarker spacing of 258 kbp (Fidanza et al., 2001). The SNP were ordered by chromosome position using Bovine Genome Build 4.0 (http://www.ncbi.nlm.nih. gov/projects/genome/guide/cow/).

Beef and Dairy Data. A total of 53,798 SNP were genotyped using the 50K chip. Preliminary investigations of the genotype data set showed that all genotypes had more than 95% quality scores and the proportion of missing genotypes was less than 2.1%. A minor allele frequency of < 0.05 was found for 16,008 SNP, and 8,469 SNP deviated from Hardy-Weinberg equilibrium (P < 0.0001). However, these were not removed from further analyses. Out of the initial 53,798 SNP, 50,650 were polymorphic and included in the GWAS. Additionally, 8,201 SNP, which were genotyped using the 10K chip, were evaluated for their effects on RFI. There were 2,390 SNP in common on both the 10K and 50K SNP chips. For dairy data, a preedited genotype data set consisting of 39,048 SNP loci (Hayes et al., 2009) was used for the association analyses.

#### Animals and Populations

**Beef Cattle.** Phenotype and genotype data held in 3 cattle databases were used (Table 1). From the Beef Cooperative Research Centre phase I (**CRCI**) records, phenotypic records on RFI and growth traits on 852 steers with 50K SNP genotype data were obtained (Table 2). These steers were from 7 different pure breeds of 3 breed types. The 4 breeds (Angus, Murray Grey, Shorthorn, and Hereford) were Bt, 1 breed

		Beef data se	t	Dairy	data set
Breed type and trait	CRCI	CRCII	Trangie	Holstein, reference	Holstein, validation
Breed type					
Bt	×		×	×	×
Bi	×	×			
$Bt \times Bi$	×	×			
SNP chip					
50K	×	×		×	×
10K			×		
Trait					
RFI	×				
DFI	×				
ADG	×				
mMWT	×				
w1LWT		×			
pwHH	×				
w1HH		×			
Stature				×	×

**Table 1.** Breed type and trait description<sup>1</sup>

<sup>1</sup>CRCI = Cooperative Research Centre phase I; CRCII = Cooperative Research Centre phase II; Trangie = Angus Trangie selection line; Bt = *Bos taurus*; Bi = *Bos indicus*; Bt × Bi = cross between *B. taurus* and *B. indicus*; 50K = commercial SNP chip containing approximately 50,000 SNP (Illumina, San Diego, CA); 10K = commercial SNP chip containing approximately 10,000 SNP (Affymetrix, Santa Clara, CA); RFI = residual feed intake, kg/d; DFI = daily feed intake, kg/d; ADG, kg/d; mMWT = metabolic midweight, kg<sup>0.75</sup>; w1LWT = end of wet-season 1 BW, kg; pwHH = postweaning hip height, cm; w1HH = end of wet-season 1 hip height, cm.

(Brahman) was Bi, and 2 breeds (Santa Gertrudis and Belmont Red) were Bt  $\times$  Bi synthetic breeds (Johnston et al., 2003). From the Beef Cooperative Research Centre phase II (**CRCII**) data set, records were obtained for 1,456 cows with 50K SNP chip data plus BW and height data. These cows were from 2 breed types: Bi (Brahman) and Bt  $\times$  Bi crosses (Tropical Composites; Barwick et al., 2009). Third, records for 379 Angus (Bt) cattle that had been genotyped using the 10K SNP chip were obtained. These cattle were from the divergent RFI selection lines based at the Trangie Agricultural Research Centre, New South Wales, Australia (Arthur et al., 2001a). Although Trangie selection line, CRCI, and CRCII animals are different, they could be related because of common ancestors.

**Dairy Cattle.** Data for bulls with 50K SNP chip data were extracted from the Australian Dairy Herd Improvement Scheme database. There were 588 Holstein bulls that received EBV based on the phenotypic records of their daughters before 2005 (reference data set) and 117 Holstein bulls proven between 2005 and 2007 (validation data set; Table 1).

**Table 2.** Number of records (N), mean and SD, and estimates of heritability  $(h^2)$  and associated SE for all traits studied<sup>1</sup>

							А	11	
Trait	CRC phase	Ν	Bt	Bi	Bt $\times$ Bi	Mean	SD	$h^2$	SE
Beef cattle									
RFI, 10K		379	379			-0.20	1.3	0.89	0.09
RFI	Ι	852	486	78	288	-0.04	1.2	0.18	0.13
DFI	Ι	852	486	78	288	12.3	2.1	0.16	0.13
ADG	Ι	852	486	78	288	1.40	0.4	0.24	0.14
mMWT	Ι	852	486	78	288	93.8	11.4	0.31	0.15
w1LWT	II	1,456		590	866	301.3	44.3	0.61	0.11
pwHH	Ι	812	466	65	281	116.4	6.5	0.25	0.18
w1HH	II	1,224		360	864	126.0	5.8	0.60	0.12
Dairy cattle									
Stature, reference		588				0.54	0.72	0.78	0.12
Stature, validation	_	117				0.55	0.80	0.71	0.13

<sup>1</sup>A dash (—) indicates data were not available. CRC = Cooperative Research Centre; Bt = Bos taurus; Bi = Bos indicus; Bt × Bi = cross between *B. taurus* and *B. indicus*; RFI = residual feed intake, kg/d; 10K = commercial SNP chip containing approximately 10,000 SNP (Af-fymetrix, Santa Clara, CA); DFI = daily feed intake, kg/d; ADG, kg/d; mMWT = metabolic midweight, kg<sup>0.75</sup>; w1LWT = end of wet-season 1 BW, kg; pwHH = postweaning hip height, cm; w1HH = end of wet-season 1 hip height, cm; stature = height, cm; reference = reference data set; validation = validation data set.

#### Traits Studied

**Beef Cattle.** The CRCI steers were approximately 1 yr old before being recorded in a research feedlot for 4 traits: RFI, ADG, daily feed intake (**DFI**), and metabolic midweight (**mMWT**), and before the feedlot period for postweaning hip height (**pwHH**), following standard procedures described by Johnston et al. (2003) and Robinson and Oddy (2004). Residual feed intake is a measure of feed efficiency and is calculated as the difference in feed intake above or below that expected or predicted on the basis of metabolic BW and growth rate (Arthur et al., 2001b). The CRCII heifers were recorded for first postweaning wet-season BW and hip height (Barwick et al., 2009). The Angus cattle with 10K SNP data were bulls and heifers and were measured for the feedlot traits at a younger age than the CRCI steers (Arthur et al., 2001b).

**Dairy Cattle.** Deregressed EBV for stature in the Holstein reference and validation data sets were used for GWAS. Further details of how deregressed EBV were calculated are given in Pryce et al. (2010a). Stature EBV for dairy bulls are calculated from phenotypes of 2-yr-old daughters measured at the highest point of the sacrum at the hip bones and converted to a scale of 1 to 9, where 1 is approximately 130 cm and 9 is 150 cm.

#### Estimate of LD in the Populations Studied

The LD between pairs of SNP markers  $(r^2)$  was used to estimate the extent of LD in the populations studied. The average pair-wise  $r^2$  for each population was calculated using the LDMAX procedure in GOLD (Abecasis and Cookson, 2000), using the conventional measure of  $r^2$  (Hill and Robertson, 1968; Devlin and Risch, 1995). The average of the LD estimates  $(\bar{r}^2)$  for each population was then calculated in every 10-kbp interval, and it was corrected for sample size (=  $\bar{r}^2 - 1/N$ , where N is sample size).

#### Statistical Analyses

**Phenotype and Model Used.** The association between each SNP and each of the traits was assessed by a regression analysis using ASReml software (Gilmour et al., 2002). The mixed model applied was as follows: trait ~ mean + fixed effects + SNP<sub>i</sub> + animal + error, with animal and error fitted as random effects. The *i*th SNP (SNP<sub>i</sub>) was fitted as a covariate effect. Fixed effects were different for the CRCI and CRCII data sets. For the CRCI data set, breed, herd of origin, sex, year of measurement, season, market BW destination, and nutritional treatment were fitted as class variables, and age deviation from the group mean was fitted as a covariate, whereas for CRCII data the effects of breed, herd of origin, sire group, cohort, calving month, and their first-degree interactions were fitted as fixed effects (Barwick et al., 2009). The fixed effects used for the Angus Trangie selection line data set were contemporary group and linear covariate for age (Arthur et al., 2001b). When the same model was used without fitting  $\text{SNP}_i$ , estimates of heritability in beef cattle were calculated based on the genotyped animals and their 5-generation ancestors (Table 2).

In the dairy cattle, the estimates of heritabilities of deregressed EBV for stature were calculated. The SNP were evaluated for their effects on stature by using a mixed model fitting the mean, SNP as a fixed effect, and animal as a random effect.

Significance of SNP. The SNP were tested for a significant association with particular traits at different probability thresholds (Table 3). In a GWAS, many thousands of significance tests are performed. Therefore, the number of SNP that were significant to the number expected by chance was compared by using a false discovery rate (**FDR**) as

$$FDR = \frac{P(1-s)}{s(1-P)},$$

where P is a defined probability threshold and s is a proportion of SNP that are nominally significant at the defined threshold (= number of significant SNP divided by number of total SNP). This is equivalent to the FDR formula of Storey (2002). The number of true positive SNP then equals to (1 - FDR) multiplied by the number of significant SNP at a particular probability threshold.

The correlations of SNP effects between RFI, ADG, and DFI were estimated. The effects of SNP with large SE are sometimes large but the effects are poorly estimated. Therefore, the SNP effects were divided by their SE before correlations of the SNP effects were calculated.

Validation of SNP. The SNP that were significantly associated with RFI in the 50K SNP data were tested for an association in the 10K SNP data. The 2 data sets had only 2,390 SNP in common. Therefore, we also tested whether significant associations with RFI were found within the same 1-Mbp intervals in both data sets by using  $\chi^2$  tests. If a 1-Mbp interval did not contain any SNP in one of the 50K (reference) or 10K (validation) data sets or in both data sets, then the particular 1-Mbp intervals were removed. After removing those 1-Mbp intervals, each 1-Mbp interval was scored as containing or not containing 1 or more significant SNP from the 10K data set and from the 50K data set. The data were then in the form of a  $2 \times 2$  table, in which 1-Mbp regions were classified as significant or not significant in 2 different experiments. We tested the significance of the agreement between experiments using a  $\chi^2$  test. However, some 1-Mbp regions contained more SNP than others and so might be more likely to contain a significant SNP. This would bias the  $\chi^2$ test. Therefore, we carried out the permutation tests

#### Bolormaa et al.

		No. of SNP			FDR, $\%$	
Trait	P < 0.0001	P < 0.001	P < 0.01	P < 0.0001	P < 0.001	P < 0.01
Beef cattle						
RFI, 10K	36	111	468	2	7	17
RFI	11	75	615	46	67	82
DFI	8	76	624	63	67	81
ADG	11	83	698	46	61	72
mMWT	6	78	694	84	65	73
w1LWT	29	156	935	17	32	54
pwHH	13	75	632	39	67	80
w1HH	26	134	833	19	38	60
Dairy cattle						
Stature, reference	26	173	912	15	22	42
Stature, validation	9	70	589	43	56	66

**Table 3.** Number of significant SNP and false discovery rate (FDR) at different thresholds (P < 0.01) for all traits studied in beef and dairy cattle<sup>1</sup>

 $^{1}$ RFI = residual feed intake; 10K = using a commercial SNP chip containing approximately 10,000 SNP (Affymetrix, Santa Clara, CA); DFI = daily feed intake; mMWT = metabolic midweight; w1LWT = end of wet-season 1 BW; pwHH = postweaning hip height; w1HH = end of wet-season 1 hip height; stature = height; reference = reference data set; validation = validation data set.

to establish an appropriate significance threshold for the  $\chi^2$  statistic. We did permutation tests with 10,000 repetitions to derive the distribution of the test statistic under the null hypothesis to calculate the significance of the association in 1-Mbp intervals between the 2 data sets. The permutation test was performed using the reference data set (e.g., 50K data) with real effects and the validation data set (e.g., 10K data), but with the significance status of SNP permutated across the genome. The number of SNP considered as being significant in the validation data was the same as the number of significant SNP (P < 0.05) with real effects, but for the permutation test, significant SNP were chosen at random. A  $\chi^2$  test for each of 10,000 permutations was calculated, and this empirical distribution of  $\chi^2$  statistics under the null hypothesis was used to test the significance of the association in the same 1 Mbp.

For BW and height traits, a validation test of SNP associations was carried out using the results from the analyses of the same 50K SNP chip data in 2 different populations (CRCI and CRCII cattle). Additionally, the 3 breed types within the CRCI data (Bt, Bi, and  $Bt \times Bi$ ) were also analyzed separately as well as in a joint analysis. Similarly, the 2 breed types (Bi and Bt  $\times$  Bi) represented in the CRCII data set were analyzed separately as well as jointly. The number of records in each breed line is given in Table 2. The number of SNP that were significant in both CRC data sets was counted, and for these SNP, 2 parameters to assess the agreement between the results were calculated: the correlation between SNP effects in the 2 data sets, and the proportion of SNP in which the effects were in the same direction; that is, the proportion in which the same SNP allele increased the trait.

Similarly, a validation of SNP for stature in the dairy reference and validation populations was carried out by examining the proportion of SNP effects with the same direction in the 2 data sets. Finally, the SNP significantly associated with height in the beef and dairy cattle were compared. This was done at the same SNP position and within 1-Mbp regions.

Information about particular genes, located near SNP significantly associated with RFI, was extracted from online sources (http://www.ensembl.org/index. html, http://www.genecards.org/cgi-bin/cardsearch. pl#top, and http://www.uniprot.org).

#### RESULTS

#### Summary Statistics

Raw means, SD, and heritability estimates are given in Table 2. Heritability estimates are based on small sample sizes and so are subject to large SE. In the Trangie animals, the estimate of heritability for RFI is biased upward because animals with extreme phenotypes for RFI were chosen for the experiment. The estimates of heritability using the full RFI data set (n =1,177) was 0.39 (Arthur et al., 2001a). The proportion of genetic variance relative to the total variance was used to calculate "heritability" for deregressed EBV of stature in both the Holstein reference and validation populations (Table 2). This measures the reliability of the progeny test rather than the heritability of the raw trait.

#### LD in the Populations Studied

The average  $r^2(\bar{r}^2)$  declined as a function of distance between markers. Linkage disequilibrium in all populations decreased rapidly over short distances (Figure 1) but remained slightly greater than zero over long distances. The dairy cattle population had the greatest LD. Linkage disequilibrium was greatest in Bt, followed by Bt × Bi, and then by Bi.



Figure 1. Relationship between genetic distances and values of linkage disequilibrium (mean  $r^2$  corrected for sample size) between SNP markers in different breed types [beef *Bos taurus* (Bt), *Bos indicus* (Bi), and crosses of Bt × Bi, and dairy Bt (HFall, all Holstein-Friesian bulls)].

#### RFI

A total of 75 SNP from the 50K chip data for the CRCI steers were significant (P < 0.001) for RFI, with an FDR of 67% (Table 3). These 75 SNP were mapped on 24 different BTA. Of the 75 SNP, the 9 most significant SNP were detected on BTA 3, 5, 7, and 8, with  $P \leq 6.0 \times 10^{-5}$ . A broad peak including the 3 most significant SNP (of these 9 SNP) was detected between 86 and 94 Mbp of BTA 8 (Figure 2). Of these 3 SNP on BTA 8, two were in high LD ( $r^2 = 0.58$ ) with each other, and the other SNP was in less LD ( $r^2 < 0.16$ ). In the Trangie population (10K chip data), 111 SNP were significantly associated with RFI, with an FDR of 7% (Table 3). Approximately, 103 of these SNP were therefore likely to represent true positives.

No separate data set for RFI based on the 50K chip could be used for validation. Therefore, the results from the 10K SNP chip were used to validate those from the 50K SNP chip. Of the 2,390 SNP in common between the 2 data sets (50K and 10K SNP data), 27 of them were significant at P < 0.05 in both data sets (Table 4). This is not convincingly more than expected by chance. Every 1-Mbp interval was classified as significant or not significant according to whether it contained or did not contain 1 or more significant SNP in both experiments. Out of 2,131 intervals, 406 intervals included at least 1 (sometimes up to 11) significant SNP (P < 0.05) in both the 50K and 10K data sets ( $\chi^2 = 9.25$ ; Table 5), and this  $\chi^2$  was significant based on the 10,000 permutation tests. As an example of this tendency to find QTL for RFI in the same region in both data sets, the 10K data also showed a high and broad peak for RFI on BTA 8 at a position of 82 to 94 Mbp (Figure 2).

### Pleiotropy of SNP Affecting RFI

Residual feed intake was moderately correlated with DFI  $(r_P = 0.56)$  and weakly correlated with ADG  $(r_P$ = 0.12), although this latter correlation was expected to be zero. Average daily gain and DFI had a high, positive correlation ( $r_P = 0.63$ ). A similar number of significant SNP (P < 0.001) were detected for ADG and DFI (Table 3). The correlation between SNP effects estimated for RFI and DFI was moderately positive (0.58), whereas the correlation for ADG and DFI was high (0.71; Table 6). Residual feed intake is a measurement that is corrected phenotypically for ADG, and the correlations of SNP effects reflected this. The numbers of SNP that were significant (P <(0.05) for both RFI and DFI and for both ADG and DFI were 651 and 973, respectively. The proportion of these significant SNP effects in the same direction was 100% (Table 6).

Because RFI and ADG are almost uncorrelated, SNP are not expected to be associated with both traits unless there are QTL with a pleiotropic effect on both traits. In total, 162 SNP (P < 0.05) had significant effects for both RFI and ADG, which is no more than expected by chance, and the proportion of effects of these SNP in the same direction was 40% (Tables 5 and 6). When 1-Mbp intervals were considered instead of individual SNP, the number of intervals that contained 1 or more significant SNP (P < 0.05) for both RFI and ADG was 845 out of 2,530 1-Mbp intervals (Table 5). This was not more than expected by chance ( $\chi^2 =$ 2.22). In addition, 576 intervals (out of 2,530) containing SNP were significant (P < 0.05) for all of RFI, ADG, and mMWT (e.g., Table 7).



Figure 2. Significant SNP (P < 0.05) for residual feed intake (RFI), ADG, and daily feed intake (DFI) on BTA 8. 10K indicates using a commercial SNP chip data set containing approximately 10,000 SNP (Affymetrix, Santa Clara, CA), whereas 50K indicates using a commercial SNP chip data set containing approximately 50,000 SNP (Illumina, San Diego, CA).

#### Growth Traits Using the 50K Chip

**Beef Cattle.** The number of significant SNP at a threshold of P < 0.001 (Table 3) was 78 for mMWT in the feedlot and 75 for postweaning height in the CRCI population but was 156 for end of wet-season BW and was 134 for height in the larger CRCII data. Consequently, the FDR were less in the CRCII data (32 and 38%) compared with the CRCI data (65 and 67%).

Table 8 shows the number of SNP that were significant for BW or for height in one of the CRCI populations and in one of the CRCII populations. Table 8 also gives the correlations of SNP effects for BW and height between the CRCI and CRCII data sets across breed types, as well as the proportion of SNP whose effects were in the same direction. The number of SNP significant in both the CRCI and CRCII data sets was no more than expected by chance and, in most cases, the proportion of effects in the same direction did not depart significantly from the 50% expected by chance. The exception was for the  $Bt \times Bi$  data sets, for which 72% of the SNP effects were in the same direction for BW and 66% were in the same direction for height (Table 8). When 1-Mbp intervals were considered, the number of regions that contained a significant SNP (P< 0.05) in both data sets was 959, which is more than the number expected by chance  $(\chi^2 = 6.27; \text{ Table 5}),$ but was not significant by the permutation test.

**Dairy Cattle.** The proportion of SNP that were significant in the Holstein reference data set was greater than in the smaller validation data set and was as high as in the larger CRCII data set (Table 3). This resulted in an FDR for the dairy reference data set of

22% at P < 0.001 but a greater FDR (56%) for the dairy validation set. When the same threshold of P < 0.05 was set in both the reference and validation data sets, the number of SNP that were significant in both data sets was 215, which is not greater than expected by chance (Table 5), but the proportion of significant SNP effects in the same direction was 68%. By considering 1-Mbp intervals, the number of intervals containing at least 1 significant SNP (P < 0.05) in both the reference and validation data sets was 640 ( $\chi^2 = 2.12$ , which is not significant).

Validation of SNP Across Beef vs. Dairy Cattle Breeds. The number of SNP that were significant (P < 0.05) for both dairy stature and beef hip height was 202, which is somewhat more than expected by chance (Table 5). When the presence of significant SNP within 1-Mbp intervals along each chromosome was examined, the  $\chi^2$  test was also not significant ( $\chi^2$ = 2.91; Table 5). As shown in Table 5, there were several 1-Mbp intervals containing SNP significant (P <(0.05) for both the dairy and beef height traits. For dairy stature and beef first postweaning wet-season hip height, there were 857 1-Mbp intervals containing SNP significant for these 2 traits. For dairy stature and beef postweaning hip height, there were 821 1-Mbp intervals containing SNP significant for these 2 traits. This suggests that some significant SNP in the dairy reference data set were near significant SNP in the beef data sets (in this case, within the same 1-Mbp intervals), but this trend was not significant by the permutation test. In general, regions of the genome contain many significant SNP for height and BW in different populations, such as the broad peak detected on BTA 3 from 102.159 to

Table 4. Significant SNP (P < 0.05) for residual feed intake (RFI) at particular positions and within 1-Mbp intervals in both the 10K and 50K data sets<sup>1</sup>

SNP		Position,	50K,	10K,		$Location,^2$	50K,	10K,	50K,	10K,
name	BTA	Mbp	P-value	<i>P</i> -value	BTA	Mbp	No. of $SNP^3$	No. of SNP	${P_{\min}}^4$	$P_{\min}$
352323	1	103459113	0.0265	0.0015	2	22 to 23	1	2	0.0053	0.0009
347872	1	140599889	0.0313	0.0081	2	24 to 25	2	3	0.0003	0.0101
352046	1	35152843	0.0415	0.0326	2	63 to $64$	2	1	0.0002	0.0338
345175	2	113984723	0.0143	0.0186	3	105 to 106	7	2	0.0000	0.0032
348132	2	133058384	0.0064	0.0095	4	41  to  42	3	2	0.0001	0.0211
353948	2	83913947	0.0336	0.0154	4	91 to 92	2	6	0.0002	0.0010
350236	3	24374862	0.0147	0.0028	5	51  to  52	2	6	0.0021	0.0001
342691	5	116152845	0.0310	0.0324	5	75 to 76	3	4	0.0004	0.0073
349813	5	90670437	0.0034	0.0468	5	85 to 86	2	5	0.0010	0.0246
347570	8	21215935	0.0033	0.0438	5	110 to 111	2	2	0.0000	0.0084
347480	8	93873871	0.0001	0.0264	7	102  to  103	3	4	0.0000	0.0011
349887	9	36408577	0.0255	0.0216	8	2  to  3	2	3	0.0025	0.0004
352056	11	16684267	0.0489	0.0210	8	86 to 87	6	3	0.0000	0.0031
344077	11	51903129	0.0344	0.0264	8	90 to 91	7	2	0.0078	0.0001
352038	12	81260515	0.0306	0.0104	8	93 to 94	6	3	0.0001	0.0199
342868	14	59149744	0.0101	0.0399	8	104  to  105	4	1	0.0009	0.0392
343856	16	16731684	0.0170	0.0381	9	14 to 15	1	1	0.0086	0.0009
346839	16	33946102	0.0234	0.0270	9	60 to 61	3	1	0.0044	0.0001
345143	16	46045401	0.0167	0.0200	10	18 to 19	2	1	0.0006	0.0242
349182	18	24355937	0.0407	0.0009	11	1 to 2	2	3	0.0040	0.0008
352299	18	43829131	0.0161	0.0036	12	55 to 56	2	4	0.0064	0.0000
345848	18	45787269	0.0214	0.0010	17	10 to 11	3	2	0.0008	0.0317
353716	19	17750262	0.0239	0.0017	17	43 to $44$	4	1	0.0045	0.0001
353167	19	19698761	0.0317	0.0487	17	57 to 58	5	4	0.0024	0.0002
353494	20	51402608	0.0315	0.0294	18	3  to  4	4	2	0.0041	0.0008
354432	26	2527236	0.0188	0.0061	20	33 to $34$	5	1	0.0006	0.0135
348792	26	32256982	0.0396	0.0064	24	10 to 11	2	1	0.0003	0.0336
					25	12  to  13	1	4	0.0039	0.0005
					27	21 to 22	2	1	0.0084	0.0004
					28	36 to 37	4	2	0.0009	0.0115

 $^{1}10K$  = commercial SNP chip containing approximately 10,000 SNP (Affymetrix, Santa Clara, CA); 50K = commercial SNP chip containing approximately 50,000 SNP (Illumina, San Diego, CA).

<sup>2</sup>Indicates 1-Mbp range.

<sup>3</sup>No. of SNP = number of significant SNP (P < 0.05) for both the 50K and 10K data sets within 1-Mbp intervals; P = F-probabilities.

 ${}^{4}P_{\min} = P$ -value of SNP found to be the least among significant SNP (P < 0.05) across 2 data sets at particular locations of the 1-Mbp interval.

109,411 Mbp (Figure 3), which contained significant SNP in both the beef and dairy data sets.

#### DISCUSSION

This paper discussed the results of a GWAS for traits related to BW, height, and feed intake in beef and dairy cattle genotyped using 50K and 10K SNP chips. Although more significant associations were found between SNP genotype and trait than expected by chance, many of the FDR were disappointingly large. Some of the factors affecting FDR are apparent from the results. The larger sample size (in the CRCII than in the CRCI, and in the dairy reference set than in the dairy validation set) was associated with decreased FDR. The reduced FDR in the 10K data set was most likely due to the use of lines of Angus cattle selected for high and low RFI. This increased the range of breeding values for RFI and so increased the power of the analysis. Part of the variation in the Trangie selection line data could be due to genetic drift, but we have attempted to correct for this by fitting an animal model in the analysis. The greater "heritability" of progeny means compared with single-animal phenotypes was associated with reduced FDR in the dairy cattle compared with the beef cattle GWAS using the same number of cattle. In addition, the use of a single breed in the dairy cattle experiment compared with 7 beef breeds across Bi and Bt would also have contributed to the greater FDR in the beef cattle experiment, as explained below.

At the density of SNP used here, the phase of LD would not be expected to be consistent across cattle breeds (deRoos et al., 2008). Consequently, when multiple breeds are combined, the associations between an SNP and a QTL are likely to be in different directions in different breeds and hence partially cancel out. Fitting a model with an effect of SNP nested within breed (results not included) lacked the power to determine the phase of LD accurately because of the small number of animals within each breed.

The FDR was reduced slightly by using a more stringent *P*-value in the significance test but at the cost of reducing the number of true associations detected. Therefore, rather than rely on a very stringent significance test, confirmation of associations discovered be-

Trait and data set	SNP or interval	Total No.	No. significant in data set 1	No. significant in data set 2	No. significant in data sets 1 and 2	$\times^2$	P-value <sup>2</sup>
RFI-based 50K and 10K chips in beef cattle							
$50 \mathrm{K}^3$ and $10 \mathrm{K}^4$	Individual SNP	2,390	130	369	27	2.99	0.084
$50 \mathrm{K}^3$ and $10 \mathrm{K}^4$	1-Mbp interval	2,131	1,207	660	406	9.25	0.012
RFI- and ADG-based 50K chip in beef cattle	4						
$RFI$ , $all^3$ and $ADG$ , $all^4$	Individual SNP	50,633	2,826	2,995	162	0.18	0.672
$RFI$ , $all^3$ and $ADG$ , $all^4$	1-Mbp interval	2,530	1,417	1,476	845	2.22	NS
LWT <sup>5</sup> -based 50K chip in beef cattle							
CRCI <sup>3</sup> and CRCII <sup>4</sup>	Individual SNP	49,957	2,699	3,050	176	0.58	0.446
CRCI <sup>3</sup> and CRCII <sup>4</sup>	1-Mbp interval	2,532	1,585	1,484	959	6.27	NS
$HH^{3}$ -based 50K chip in beef cattle	r.						
CRCI <sup>3</sup> and CRCII <sup>4</sup>	Individual SNP	49,926	2,846	3,037	198	4.04	0.044
CRCI <sup>3</sup> and CRCII <sup>4</sup>	1-Mbp interval	2,532	1,543	1,492	938	5.68	NS
Stature-based 50K chip in dairy cattle							
Holstein, reference <sup>3</sup> and Holstein, validation <sup>4</sup>	Individual SNP	39,040	3,215	2,462	215	0.86	0.354
Holstein, reference <sup>3</sup> and Holstein, validation <sup>4</sup>	1-Mbp interval	2,527	1,370	1,147	640	2.12	NS
Growth-based 50K chip in beef and dairy cattle							
pwHH, all <sup>3</sup> and Holstein, reference <sup>4</sup>	Individual SNP	38,367	2,191	3,157	202	3.02	0.074
pwHH, all <sup>3</sup> and Holstein, reference <sup>4</sup>	1-Mbp interval	2,530	1,492	1,371	821	1.03	NS
w1HH, all <sup>3</sup> and Holstein, reference <sup>4</sup>	Individual SNP	38,521	2,614	3,176	227	0.71	0.399
w1HH, $all^3$ and Holstein, reference <sup>4</sup>	1-Mbp interval	2,530	1,543	1,371	857	2.91	NS
$^{1}$ RFI = residual feed intake, kg/d; 50K = commerci	ial SNP chip containing appro	ximately 50,000 SN	IP (Illumina, San Die	go, CA); $10K = com$	mercial SNP chip contain	ing approxin	nately 10,000
SNP (Affymetrix, Santa Clara, CA); ADG, kg/d; LW	T = BW, kg; mMWT = meta	bolic mid-weight, k	$g^{0.75}$ ; w1LWT = end $e^{0.75}$	of wet-season 1 BW,	kg; $CRCI = Cooperative$	Research C	entre phase I;
CRCII = Cooperative Research Centre phase II; HH =	= hip height, cm; pwHH $=$ po	stweaning hip heigh	it, cm; w1HH = end $c$	of wet-season 1 hip h	eight, $cm$ ; stature = heigh	t, cm.	

**Table 5.** Validation of individual SNP or 1-Mbp chromosome regions affecting traits in 2 data sets<sup>1</sup>

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<sup>2</sup>*P*-value = probability of  $\chi^2$  using permutation tests with 10,000 repetitions; NS = not significant. <sup>3</sup>Data set 1. <sup>4</sup>Data set 2. <sup>5</sup>Bt × Bi<sub>CRCI</sub> and Bt × Bi<sub>CRCI</sub>, where Bt is *Bos taurus*, Bi is *Bos indicus*, and Bt × Bi is a cross between *B. taurus* and *B. indicus*.

## 1692

Bolormaa et al.

Table 6. Pleiotropic effects of SNP across feedlot traits in the 50K CRCI data<sup>1</sup>

Item	RFI	DFI	ADG	mMWT
RFI	2,826	0.58	-0.05	-0.06
DFI	651 (100%)	2,733	0.71	0.68
ADG	162 (40%)	973 (100%)	2,995	0.60
mMWT	182 (31%)	961 (100%)	786 (100%)	3,141

<sup>1</sup>On diagonal = number of significant SNP (P < 0.05) for each trait; above diagonal = correlations of SNP effects between traits; below diagonal = number of significant SNP in both traits (in parentheses, proportion of significant SNP effects to be in the same direction for both traits). 50K = commercial SNP chip containing approximately 50,000 SNP (Illumina, San Diego, CA); CRCI = Cooperative Research Centre phase I; RFI = residual feed intake, kg/d; DFI = daily feed intake, kg/d; ADG, kg/d; mMWT = metabolic midweight, kg<sup>0.75</sup>.

tween SNP and traits was sought by confirming them in an independent population of cattle. The most obvious confirmation would be to find the same SNP as significant in both data sets. However, this was rarely the case because of the lack of power in both the discovery and validation data sets and because of the use of different breeds for validation and discovery. Variation in the LD phase between breeds means that an SNP that is significant in one breed may not be significant in another breed, even if the same QTL is segregating in both breeds.

The most powerful confirmation test appears to be finding that SNP that are significant in both populations have effects in the same direction more often than expected by chance. For instance, the number of SNP that were significant in the Bt  $\times$  Bi populations from both CRCI and CRCII was no more than expected by chance. However, among those SNP that were significant in both populations, 72% had an effect in the same direction for height and 66% had an effect in the same direction for BW. Similar results are shown for the reference and validation Holstein populations. The negative correlations between the CRCI Bi and CRCII Bi populations could be due to the very small number of animals (n = 78) in the CRCI Bi data set. This test is powerful because it tests a very specific null hypothesis, that is, that 50% of SNP will have effects in opposite directions. Unfortunately, this null hypothesis appears to be true unless the 2 populations are from the same or closely related breeds.

Table 7. Significant SNP (P < 0.05) within 1-Mbp intervals for all of RFI, ADG, and mMWT in CRCI animals<sup>1</sup>

BTA	Position	RFI, No. of SNP	ADG, No. of SNP	mMWT, No. of SNP	$\substack{\text{RFI,}\\ {P_{\min}}^2}$	$ADG, P_{\min}$	$\substack{\text{mMWT,}\\ P_{\min}}$
2	106 to 107	2	3	1	0.0001	0.0019	0.0009
3	51  to  52	6	1	4	0.0004	0.0199	0.0124
3	84 to 85	4	2	3	0.0002	0.0075	0.0356
3	105 to 106	7	1	1	0.0000	0.0445	0.0131
4	46 to 47	1	2	2	0.0258	0.0000	0.0004
4	91 to 92	2	4	1	0.0002	0.0114	0.0239
6	41  to  42	3	2	5	0.0002	0.0132	0.0067
6	111 to 112	3	3	3	0.0007	0.0095	0.0066
8	86 to 87	6	2	2	0.0000	0.0047	0.0155
8	87 to 88	5	2	3	0.0001	0.0140	0.0121
8	88 to 89	7	1	2	0.0009	0.0382	0.0301
8	89 to 90	7	3	2	0.0006	0.0104	0.0053
8	104  to  105	4	2	2	0.0009	0.0253	0.0187
9	78 to 79	4	3	1	0.0003	0.0037	0.0450
10	18 to 19	2	1	2	0.0006	0.0432	0.0103
11	46 to 47	1	2	10	0.0483	0.0002	0.0002
14	17 to 18	6	5	5	0.0005	0.0008	0.0081
16	25 to 26	1	4	1	0.0417	0.0006	0.0005
17	10 to 11	3	3	2	0.0008	0.0358	0.0194
17	37 to 38	2	1	2	0.0005	0.0492	0.0051
19	38 to 39	2	1	2	0.0003	0.0417	0.0346
20	30 to 31	1	2	2	0.0001	0.0085	0.0047
22	45 to 46	1	1	2	0.0003	0.0356	0.0243
23	18 to 19	1	1	1	0.0432	0.0001	0.0009
23	49 to 50	5	3	1	0.0002	0.0274	0.0448

<sup>1</sup>RFI = residual feed intake, kg/d; ADG, kg/d; mMWT = metabolic midweight, kg<sup>0.75</sup>; CRCI = Cooperative Research Centre phase I; No. of SNP = number of significant SNP (P < 0.05) for all 3 traits found within 1-Mbp intervals.

 $^{2}P_{\min} = P$ -value of SNP found to be the least among significant SNP (P < 0.05) across all 3 traits at particular locations of the 1-Mbp interval.

Trait	$\mathrm{All}_{\mathrm{CRCI}}$ : $\mathrm{All}_{\mathrm{CRCII}}$	${\operatorname{Bt}}_{\operatorname{CRCI}}:\operatorname{Bi}_{\operatorname{CRCII}}$	$Bt_{CRCI}:Bt \times Bi_{CRCII}$	Bicrci:Bicrcii	$Bi_{CRCI}$ :Bt × $Bi_{CRCII}$	$Bt \times Bi_{CRCI}:Bi_{CRCII}$	Bt $\times$ Bi <sub>CRCI</sub> :Bt $\times$ Bi <sub>CRCI</sub>
BW							
No. of SNP	242	206	209	245	203	159	176
Correlation	0.26	-0.10	-0.04	-0.21	0.03	0.07	0.43
% of same direction	61	46	48	39	50	53	72
Hip height							
No. of SNP	208	161	149	213	226	197	198
Correlation	0.12	0.00	0.02	0.09	-0.17	0.10	0.36
% of same direction	53	47	51	55	43	54	66

Further evidence that the associations found were real is provided by finding SNP in the same 1-Mbp region significantly associated with RFI in the 2 independent data sets (50K and 10K). Even though we found associations in the 1-Mbp interval for the BW and hip height data sets within beef breeds ( $\chi^2 = 6.27$  and 5.68, respectively), the permutation tests with 10,000 repetitions showed they were not significant. This might suggest that those significant associations in 1-Mbp intervals for BW and hip height were due to the unequal distribution of SNP across the genome. If the number of SNP in each 1-Mbp interval is unequal, the  $\chi^2$  test is inappropriate. Although the permutation test is better, it is less powerful. Therefore, some important findings could be among the results that were not significant by the permutation tests.

The low power of GWAS with <1,000 animals is indirect evidence that few QTL are affecting these traits with large effects, and most QTL have small effects. However, the evidence showed that some of the associations are real, and, in particular, those found in more than 1 data set are unlikely to be false discoveries.

We have identified several chromosome regions that appear to contain polymorphisms or QTL affecting RFI. For example, a region on BTA 8 from 86 to 94 Mbp contains several SNP that were significant for RFI in the 50K or 10K experiments as well as 1 SNP that was significant in both. There were also SNP significantly associated with ADG and mMWT in this region, and there were several significant associations with RFI in both the 10K and 50K data sets at 51.05 to 51.77 Mbp on BTA 5. These results could be due to a single QTL that is in linkage disequilibrium with SNP located in millions of base pairs away or they could reflect more than 1 QTL in these regions. Within this region is a gene encoding hydroxysteroid  $(17-\beta)$  dehydrogenase 3, which is important for steroid metabolism, and another gene encoding SHC (Src homology 2 domain containing) transforming protein 3 (SHC3). The SHC3 is a signal transduction protein involved in recognition of phosphorylated tyrosine. In humans, SRC homology 2 domain-containing-transforming protein C3 plays a role as a signaling adaptor that couples activated growth factor receptors to signaling pathway in neurons and is also involved in the signal transduction pathways of neurotrophin-activated Trk receptors in cortical neurons.

Genetic correlations between traits imply that QTL have pleiotropic effects on multiple traits. High correlations of SNP effects were found between RFI and DFI as well as between ADG and DFI. However, when correlated traits are analyzed, the sampling errors tend to be correlated, so they do not represent independent evidence for the existence of a QTL. Therefore, uncorrelated traits such as RFI and ADG are useful to investigate the pleiotropic effects of QTL. [Nkrumah et al. (2007) found similar results.] As an example, 3 SNP on BTA 2 situated near 109,093,402 bp (within 42 kbp on both sides) had significant effects (with prob-



Figure 3. Significant SNP (P < 0.05) for growth [metabolic midweight (mMWT), end of wet-season 1 BW (w1LWT), postweaning hip height (pwHH), end of wet-season 1 hip height (w1HH), and stature (height)] in all beef Cooperative Research Centre cattle and in Holstein-Friesian (HF) bulls on BTA 3. Circled = a high, broad peak that contains significant SNP for growth in both beef and dairy cattle. ref = reference data set; val = validation data set.

ability thresholds between  $8.0 \times 10^{-4}$  and  $2.0 \times 10^{-2}$ ) for the feedlot RFI, ADG, and DFI traits. The gene for IGFBP2 is located on BTA 2 near 109 Mbp.

On the other hand, it is also possible that the association between an SNP and more than 1 trait reflects the effect of multiple QTL, each affecting a single trait, rather than 1 QTL affecting multiple traits (pleiotropy). With the current density of SNP marker panels, it is difficult to distinguish multiple QTL in close proximity. It may be possible to distinguish between these QTL by using denser SNP panels.

Other researchers have reported QTL for RFI in cattle (Barendse et al., 2007; Nkrumah et al., 2007; Sherman et al., 2009). Perhaps because of the lack of power and high FDR, there is not a close agreement between the studies, despite some overlap between the cattle used by Barendse et al. (2007) and the 50K SNP database used in this report. However, there are some chromosomal regions where significant associations with RFI are in common with the other studies (Table 9). Barendse et al. (2007) also detected RFI QTL (P = 0.006) on BTA 8 situated at 21.2 cM. This SNP was significant with the Angus (10K SNP) and CRCI (50K SNP) data (P = 0.044 and P = 0.003, respectively). In addition, Barendse et al. (2007) identified 2 SNP at stringent thresholds (P < 0.0009) on BTA 1 and 20, but these 2 SNP (at the same position) had no effect on RFI in the Angus and CRCI data sets used here. However, there were significant neighboring SNP on BTA 1 and on BTA 20 within 1-Mbp intervals: 3 RFI 50K SNP (P < 0.032) and one 10K SNP (P =0.035), respectively (Table 9). Sherman et al. (2009),

in Canadian cattle (Angus, Charolais, and composites), mapped 2 QTL for RFI ( $P = 1.2 \times 10^{-5}$  and 7.6  $\times$  $10^{-5}$ ) on BTA 1 and 3, respectively. The putative positions of their RFI QTL on these 2 chromosomes were in approximately the same positions as 2 significant SNP in this study ( $P < 2.7 \times 10^{-3}$  at 6 Mbp on BTA 1 and  $P < 3.5 \times 10^{-4}$  at 82 Mbp on BTA 3; Table 9). Other significant SNP from the present GWAS found on BTA 8, 11, 17, 18, 21, 22, 24, 25, and 26 were also near to those reported by Sherman et al. (2009) and Nkrumah et al. (2007; Table 9). In the report by Sherman et al. (2009), the most significant QTL for DFI ( $P = 1.38 \times$  $10^{-10}$ ) was found on BTA 7 at 54 cM. The closest significant SNP to this DFI QTL in the present study was observed at 55.4 Mbp, with a threshold of P = 0.004. Nkrumah et al. (2007) also reported the association of SNP with ADG. The 7 QTL for ADG from their study were near SNP (P < 0.008) affecting ADG on BTA 7, 11, 14, 17, 18, 20, and 28 in this study.

A group of SNP for stature in the dairy population and for height and BW for beef cattle were found to be significant (P < 0.001) on BTA 5, situated at the region of 120.9 to 121.5 Mbp. Schrooten et al. (2000), testing German Holstein Friesian cattle using 277 microsatellite markers, found an indication of suggestive QTL for stature, chest width, and birth weight on BTA 5 (at 122 cM). Furthermore, Hiendleder et al. (2003) reported significant QTL for stature on BTA 6 (at 66 cM) in the German Holstein breed by using microsatellite markers. In the present study, 3 SNP positioned at 66 Mbp on BTA 6 were found to be significant for hip height in the CRCI and CRCII beef cattle data sets.

Table 9. Significant SNP for residual feed intake (RFI) and ADG across 3 data sets<sup>1</sup>

$\frac{Position}{Mbp}^2$	50K RFI	50K ADG	10K RFI	$Publication^{3}$
6	P = 0.0027			3
10	P = 0.0007			3
26	P = 0.0253	P = 0.0071		1
82	P = 0.0003	<b>D</b>		3
83		P = 0.0007		2
60	P = 0.0092	P = 0.0313	P = 0.0306	1
82	P = 0.0003			1
50	P = 0.0184	P = 0.0211		1
21	P = 0.0438		P = 0.0033	1
21	P = 0.0012	P = 0.0073	P = 0.0438	1
80	P = 0.0060			2
19		P = 0.0079		2
30	P = 0.0072			3
55	P = 0.0064	P = 0.0193	P = 0.0000	1
14		P = 0.0154	P = 0.0121	1
74		P = 0.0016		2
9		P = 0.0042		2
18	P = 0.0045			2
56	P = 0.0026			3
28	P = 0.0020		P = 0.0185	1, 3
47		P = 0.0049		2
64	P = 0.0048			2
2			P = 0.0347	1
65		P = 0.0029		2
4	P = 0.0079			3
26	P = 0.0044			3
51	P = 0.0079	P = 0.0247	P = 0.0484	1
4	P = 0.0076			3
14	P = 0.0039			3
23		P = 0.0003		2
	$\begin{array}{r} \text{Position,}^2 \\ \text{Mbp} \end{array} \\ \hline 6 \\ 10 \\ 26 \\ 82 \\ 83 \\ 60 \\ 82 \\ 50 \\ 21 \\ 21 \\ 21 \\ 80 \\ 19 \\ 30 \\ 55 \\ 14 \\ 74 \\ 9 \\ 18 \\ 56 \\ 28 \\ 47 \\ 64 \\ 2 \\ 65 \\ 4 \\ 26 \\ 51 \\ 4 \\ 14 \\ 23 \\ \end{array}$	$\begin{array}{c c} \mbox{Position},^2 \\ \mbox{Mbp} & 50K \mbox{ RFI} \\ \hline 6 & P = 0.0027 \\ 10 & P = 0.0007 \\ 26 & P = 0.0253 \\ 82 & P = 0.0003 \\ 83 \\ \hline 60 & P = 0.0092 \\ 82 & P = 0.0003 \\ 50 & P = 0.0184 \\ 21 & P = 0.0438 \\ 21 & P = 0.0012 \\ 80 & P = 0.0060 \\ 19 \\ 30 & P = 0.0060 \\ 19 \\ 30 & P = 0.0072 \\ 55 & P = 0.0064 \\ 14 \\ 74 \\ 9 \\ 18 & P = 0.0072 \\ 55 & P = 0.0026 \\ 28 & P = 0.0026 \\ 28 & P = 0.0020 \\ 47 \\ 64 & P = 0.0026 \\ 28 & P = 0.0026 \\ 47 \\ 64 & P = 0.0048 \\ 2 \\ 65 \\ 4 & P = 0.0079 \\ 26 & P = 0.0044 \\ 51 & P = 0.0079 \\ 4 & P = 0.0079 \\ 4 & P = 0.0076 \\ 14 & P = 0.0039 \\ 23 \end{array}$	$\begin{array}{c cccc} \mbox{Position,}^2 & & & & & & \\ \mbox{Mbp} & 50K \mbox{ RFI} & 50K \mbox{ ADG} \\ \hline 6 & P = 0.0027 & & & \\ 10 & P = 0.0007 & & \\ 26 & P = 0.0253 & P = 0.0071 & \\ 82 & P = 0.0003 & & \\ 83 & & P = 0.0007 & \\ 60 & P = 0.0092 & P = 0.0313 & \\ 82 & P = 0.0003 & & \\ 50 & P = 0.0184 & P = 0.0211 & \\ 21 & P = 0.0438 & & \\ 21 & P = 0.0012 & P = 0.0073 & \\ 80 & P = 0.0012 & P = 0.0073 & \\ 80 & P = 0.0060 & & \\ 19 & & P = 0.0079 & \\ 30 & P = 0.0060 & & \\ 19 & & P = 0.0079 & \\ 30 & P = 0.0064 & P = 0.0193 & \\ 14 & & P = 0.0154 & \\ 74 & & P = 0.0016 & \\ 9 & & P = 0.0042 & \\ 18 & P = 0.0026 & & \\ 28 & P = 0.0026 & & \\ 28 & P = 0.0020 & & \\ 47 & & P = 0.0020 & \\ 47 & & P = 0.0020 & \\ 47 & & P = 0.0029 & \\ 4 & P = 0.0048 & & \\ 2 & & \\ 65 & & P = 0.0029 & \\ 4 & P = 0.0079 & & \\ 26 & P = 0.0044 & & \\ 51 & P = 0.0079 & P = 0.0247 & \\ 4 & P = 0.0076 & & \\ 14 & P = 0.0039 & & \\ 23 & & P = 0.0003 & \\ \end{array}$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$

 $^{1}$ BTA = chromosome number; 50K = commercial SNP chip containing approximately 50,000 SNP (Illumina, San Diego, CA); 10K = commercial SNP chip containing approximately 10,000 SNP (Affymetrix, Santa Clara, CA).

<sup>2</sup>http://www.ncbi.nlm.nih.gov/projects/genome/guide/cow/.

 $^{3}1$  = Barendse et al. (2007); 2 = Nkrumah et al. (2007; total number of markers associated with RFI was 8 at a chromosome-wise threshold of P < 0.05); 3 = Sherman et al. (2009; total number of markers associated with RFI was 19 at a chromosome-wise threshold of P < 0.05). All publications referenced the estimated QTL position in centimorgans. We assumed that 1 cM approximately equaled 1 Mbp to compare the results.

The most notable high and narrow peak for the size of effect was observed on BTA 2 near 20 Mbp in the Holstein validation set. Dozens of SNP on BTA 3, which were located from 102.159 to 109,411 Mbp, were associated with different growth traits in the beef data sets as well as stature in the dairy data set.

In conclusion, the GWAS reported here revealed more significant associations between SNP and traits than expected by chance. The FDR was less in the analyses with a larger number of animals, with a more stringent significance test using 1 breed only, with a more highly heritable measurement, and in a population with large genetic variance for RFI because of divergent selection. The direction of the effect of an SNP on traits such as BW and height was consistent only within a breed, probably because of the inconsistency of the LD phase between breeds. This implies that the power to detect SNP when all breed types are analyzed together is reduced because the association between the SNP and the trait is not consistent across breeds. The ability to confirm an association in an independent data set is greatest if the confirmation is carried out in the same breed as used for the discovery of the association. In this case, the most powerful confirmation test is that the direction of the association between an SNP and a trait is the same in the discovery and validation data sets. Although most effects of QTL on RFI appear to be small, associations have been found in more than one data set between RFI and SNP located on BTA 5 and 8.

#### LITERATURE CITED

- Abecasis, G. R., and W. O. Cookson. 2000. GOLD—Graphical overview of linkage disequilibrium. Bioinformatics 16:182–183.
- Arthur, P. F., J. A. Archer, R. M. Herd, and G. J. Melville. 2001a. Response to selection for net feed intake in beef cattle. Pages 135–138 in Proc. Assoc. Adv. Anim. Breed. Genet., Queenstown, New Zealand. Assoc. Adv. Anim. Breed. Genet., Queenstown, New Zealand.
- Arthur, P. F., J. A. Archer, D. J. Johnston, R. M. Herd, E. C. Richardson, and P. F. Parnell. 2001b. Genetic and phenotypic variance and covariance components for feed intake, feed effi-

ciency, and other postweaning traits in Angus cattle. J. Anim. Sci. 79:2805–2811.

- Barendse, W., A. Reverter, R. J. Bunch, B. E. Harrison, W. Barris, and M. B. Thomas. 2007. A validated whole-genome association study of efficient food conversion in cattle. Genetics 176:1893–1905.
- Barwick, S. A., M. L. Wolcott, D. J. Johnston, H. M. Burrow, and M. T. Sullivan. 2009. Genetics of heifer performance in 'wet' and 'dry' seasons and their relationships with steer performance in two tropical beef genotypes. Anim. Prod. Sci. 49:367–382.
- de Roos, A. P. W., B. J. Hayes, R. Spelman, and M. E. Goddard. 2008. Linkage disequilibrium and persistence of phase in Holstein Friesian, Jersey and Angus cattle. Genetics 179:1503– 1512.
- de Roos, A. P. W., B. J. Hayes, R. Spelman, and M. E. Goddard. 2009. Reliability of genomic predictions across multiple populations. Genetics 183:1545–1553.
- Devlin, B., and N. Risch. 1995. A comparison of linkage disequilibrium measures for fine-scale mapping. Genomics 29:311–322.
- Fidanza, J., M. Glazer, D. Mutnick, G. McGall, and C. Frank. 2001. High capacity substrates as a platform for a DNA probe array genotyping assay. Nucleosides Nucleotides Nucleic Acids 20:533–538.
- Gilmour, A. R., B. J. Gogel, B. R. Gullis, S. J. Welham, and R. Thompson. 2002. ASReml User Guide Release 1.0. VSN Int. Ltd., Hemel Hempstead, UK.
- Hayes, B. J., P. J. Bowman, A. J. Chamberlain, K. Savin, C. P. van Tassell, T. S. Sonstegard, and M. E. Goddard. 2009. A validated genome wide association study to breed cattle adapted to an environment altered by climate change. PLoS ONE 4:e6676 doi:10.1371/journal.pone.0006676.
- Hiendleder, S., H. Thomsen, N. Reinsch, J. Bennewitz, B. Leyhe-Horn, C. Looft, N. Xu, I. Medjugorac, I. Russ, C. Kühn, G. A. Brockmann, J. Blümel, B. Brenig, F. Reinhardt, R. Reents, G. Averdunk, M. Schwerin, M. Förster, E. Kalm, and G. Erhardt. 2003. Mapping of QTL for body conformation and behavior in cattle. J. Hered. 94:496–506.
- Hill, W. G., and A. Robertson. 1968. Linkage disequilibrium in finite populations. Theor. Appl. Genet. 38:226–231.
- Johnston, D. J., A. Reverter, H. M. Burrow, V. H. Oddy, and D. L. Robinson. 2003. Genetic and phenotypic characterisation of

animal, carcass, and meat quality traits from temperate and tropically adapted beef breeds. 1. Animal measures. Aust. J. Agric. Res. 54:107–118.

- Matukumalli, L. K., C. T. Lawley, R. D. Schnabel, J. F. Taylor, M. F. Allan, M. P. Heaton, J. O'Connell, S. S. Moore, T. P. Smith, T. S. Sonstegard, and C. P. Van Tassell. 2009. Development and characterization of a high density SNP genotyping assay for cattle. PLoS ONE 4:e5350.
- Moore, S. S., F. D. Mujibi, and E. L. Sherman. 2009. Molecular basis for residual feed intake in beef cattle. J. Anim. Sci. 87(E. Suppl.):E41–E47. doi:10.2527/jas.2008-1418.
- Nkrumah, J. D., E. L. Sherman, C. Li, E. Marques, D. H. Crews Jr., R. Bartusiak, B. Murdoch, Z. Wang, J. A. Basarab, and S. S. Moore. 2007. Primary genome scan to identify putative quantitative trait loci for feedlot growth rate, feed intake, and feed efficiency of beef cattle. J. Anim. Sci. 85:3170–3181.
- Pryce, J. E., S. Bolormaa, A. J. Chamberlain, P. J. Bowman, K. Savin, M. E. Goddard, and B. J. Hayes. 2010a. A validated genome-wide association study in 2 dairy cattle breeds for milk production and fertility traits using variable length haplotypes. J. Dairy Sci. 93: 3331–3345.
- Pryce, J. E., M. Haile-Mariam, K. Verbyla, P. J. Bowman, M. E. Goddard, and B. J. Hayes. 2010b. Genetic markers for lactation persistency in primiparous Australian dairy cows. J. Dairy Sci. 93:2202–2214.
- Robinson, D.L., and V.H. Oddy. 2004. Genetic parameters for feed efficiency, fatness, muscle area and feeding behaviour of feedlot finished beef cattle. Livest. Prod. Sci. 90:255–270. doi:10.1016/j.livprodsci.2004.06.011.
- Schrooten, C., H. Bovenhuis, W. Coppieters, and J. A. M. Van Arendonk. 2000. Whole genome scan to detect quantitative trait loci for conformation and functional traits in dairy cattle . J. Dairy Sci. 83:795–806.
- Sherman, E. L., J. D. Nkrumah, C. L. R. Bartusiak, B. Murdoch, and S. S. Moore. 2009. Fine mapping quantitative trait loci (QTL) for feed intake and efficiency in beef cattle. J. Anim. Sci. 87:37–45.
- Storey, J. D. 2002. A direct approach to false discovery rates. J. R. Stat. Soc., B 64:479–498.