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Integrated testing strategies can be optimal for chemical risk classification $\stackrel{\Leftrightarrow}{\approx}$

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Abstract

There is an urgent need to refine strategies for testing the safety of chemical compounds. This need arises both from the financial and ethical costs of animal tests, but also from the opportunities presented by new in-vitro and in-silico alternatives. Here we explore the mathematical theory underpinning the formulation of optimal testing strategies in toxicology. We show how the costs and imprecisions of the various tests, and the variability in exposures and responses of individuals, can be assembled rationally to form a Markov Decision Problem. We compute the corresponding optimal policies using well developed theory based on Dynamic Programming, thereby identifying and overcoming some methodological and logical inconsistencies which may exist in the current toxicological testing. By illustrating our methods for two simple but readily gen-

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eralisable examples we show how so-called integrated testing strategies, where information of different precisions from different sources is combined and where different initial test outcomes lead to different sets of future tests, can arise naturally as optimal policies.

Keywords: Optimal Integrated Testing Strategies, Dynamic Programming

1. Introduction

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Notions of what society deems to be an acceptable testing regime for new chemicals are in a constant state of flux. Until 1999 it was acceptable in the EU to perform tests on guinea pigs in order to determine whether certain cosmetic

- ⁵ products were hazardous for human skin (Program (1999)). After 1999 this was replaced by the mouse LLNA (Local Lymph Node Assay), another animal-based method. More recently the EU ethical climate has changed again; by 2018 no new chemical to be used in the cosmetics industry can be tested on animals. Instead, chemicals need to be classified reliably using information from emerg-
- ¹⁰ ing in-vitro and in-silico assays, supplemented where possible by mathematical models. These new methods are likely to be less accurate than in-vivo tests, but are generally cheaper and less ethically problematic to implement. This presents a problem common across toxicology in general: can we make good predictions about the risks associated with new chemicals without using ani-
- ¹⁵ mals at all? In other words, how best can we assemble uncertain information based on non-animal assays, so as to arrive at optimal ethical testing regimes?

Many important papers have emerged on this topic (Gabbert and Weikard (2010),Gabbert and van Ierland (2010),Gabbert and Weikard (2013),Jaworska J (2010),Jaworska et al. (2013),Jaworska et al. (2015),Norlen H (2014)).

Indeed (Gabbert and Weikard (2010)) develops a theory that determines the optimal exposure level of any particular member of the population to the chemical and uses this theory to solve a decision problem of how to pick which chemical to test for hazard first from some finite set of possible chemicals. (Gabbert and

- van Ierland (2010)) develops a framework allowing one to compute the optimal battery of tests to assess a generic toxicological endpoint by means of a cost effectiveness analysis (CEA). By contrast, (Gabbert and Weikard (2013)) develops a framework in which adaptive cost sensitive Integrated Testing Strategies can be derived by means of a Value of Information technique (VOI). The au-
- thors there distinguish between decision problems for competitive businesses and regulators. Furthermore, (Jaworska et al. (2015)) begins by improving and generalising previous work (Jaworska J (2010), Jaworska et al. (2013)) by developing more accurate potency class predictions of skin sensitisation potential of chemicals via theory of Bayesian Networks and then uses these results together
- ³⁵ with VOI framework to derive Optimal Integrated Testing Strategies for the assessment of chemical hazard of chemicals. Finally, similar to (Gabbert and van Ierland (2010)) and (Gabbert and Weikard (2013)), (Norlen H (2014)) uses CEA in the context of performing a cost effectiveness analysis in the special case of acute oral toxicity.

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However, none of these explicitly accounts for the individual differences between humans both in the exposure (i.e. environmental variability) and in the toxicity corresponding caused by that exposure (i.e. individual variability). To be more specific, only Gabbert and Weikard (2010) introduces a concept of toxicity for-

⁴⁵ mally but treats it as constant for all members of the population. Moreover, none of these papers combines these with the financial costs of chemical risk classification in a mathematically rigorous fashion. Any new testing strategy must be able to deal rationally with contradictory evidence. For example, one in-vitro assay may predict that certain chemical is

- a skin sensitiser, while another in-silico assay may predict that the same chemical is actually safe. The classification part of the argument in van der Veen et al. (2014) deals with this problem using a combination of majority voting and Bayesian Statistics. Jaworska J (2010) proposes assembling a Bayesian Network and combines this with the Weight-of-Evidence approach to overcome this is-
- sue. Thomas et al. (2012) proposes a strategy of "averaging probabilities", using empirical estimates of precision of each assay and then averaging these out in one "meta-assay". Each of these solutions may be pragmatic and defendable within the authors' given problem, but an over-arching logical framework would be a helpful step in confirming the value and risk associated with the removal of animal tests.
- 60 of animal tests.

In what follows we shall propose a mathematical framework which seeks to simultaneously overcome the shortcomings mentioned above. The issues of imprecision, and of environment- and individual-level variability, fall naturally ⁶⁵ within theories developed for evolutionary ecology (Currey et al. (2007)). The efficient assimilation of evidence can then be handled by well-developed theories of Markov Decision Processes (Bellman (1957)).

We finish this chapter by surveying previous work in Toxicology and Medicine that is based on this mathematical theory. To the best of our knowledge very little work in Toxicology uses Markov Decision Processes, to be more precise, these are the works of Chang (2010) and Korthikanti et al. (2010). Chang (2010) is a rich summary of techniques used in contemporary in house pharmacological research and decision making illustrated with numerous examples. Among a

vast range of mathematical and computational techniques used are the Markov Decision Processes as applied to the optimal decision making of a pharmaceutical business on whether to proceed from earlier (Phase 1 clinical trial) to later (Phases 2 and 3 of clinical trials) stages. Our work generalises this work in a number of directions. Firstly, the models in Chang (2010) are developed for

- the sake of a making a commercial business more profitable and do not take into account the regulatory aspect of Toxicology, i.e. the fact that the company may actually incur fines from regulatory bodies and lawsuits from individual consumers in case they exhibit adverse outcomes as a consequence of using the drug. The Markov Decision Process model in this paper takes this into account
- via the mechanisms of misclassification costs: in case the company declares an unsafe chemical as safe there will be serious consequences. Equally, if the company actually declares a safe chemical as unsafe it will lose money by not selling the safe product in the market for which it possibly had an advantage over its competitors. Thus our work bridges the two worlds: it allows the company to
- maximise its profits while simultaneously acts in the best interest of the general public. Secondly, Chang (2010) does not take neither variable exposure to chemical among different members of the target population nor precision of measurements used to test safety in the account. Instead, it relies on toxicity levels observed in recently tested chemicals to draw conclusions on the new
- ⁹⁵ chemical of interest while the transition probabilities of the Markov Decision Model model are estimated from historical data and power calculations which, by nature, cannot guarantee precisions of estimates in advance in the case of unknown moments. Another problem with this approach is that there is no guarantee the new chemical will share toxicity thresholds with the previously
- ¹⁰⁰ used chemicals. These issues motivate the truly novel part of the methodological work presented in this paper which is applicable in a general setting not necessarily restricted to Toxicology and Medicine. Namely, the transition probabilities between states of the Markov Decision Process in this paper are based on the idea of integrating evidence from different measurements of the same quantity
- in a non-contradictory way by using the information on the precision from the instruments/assays used in the process. Knowing the value obtained in a less accurate measurement and its precision we can get a probability distribution on the values more accurate measurement of the same quantity can possibly take. This simple observation has far reaching consequences; namely since the states

- of the Markov Decision Process in our model are the collection of measurements the above allows for a logically sound way of defining transition probabilities of the model by first performing cheapest possible tests for each parameter of interest. This procedure is justified by the existence of devices of varying precision in many fields of human work. As far as Toxicology is concerned, this corre-
- ¹¹⁵ sponds to in-vivo, in-vitro and in-silico tests. By choosing to start our analysis with a cheap in-silico test we remedy the problem outlined above: indeed we get the transition probabilities of the model without having to resort to further unknown characteristics of the chemical therefore bypassing a potentially circular argument. Another interesting work involving Markov Decision Process
- ¹²⁰ in Toxicology is a theoretical paper of Korthikanti et al. (2010). The authors approach the problem of model checking for a Markov Decision Process from the Computer Science point of view using the language of Mathematical Logic; actually as it turns out, the authors do not work with a conventional probabilstic definition of a Markov Decision Process but define their own in another
- set-up. Although motivated by an example of Insulin compartment model the paper soon drifts into proving results in Mathematical Logic and holds little practical value.

When it comes to applications in Medicine the literature is much larger. We
discuss in detail a variety of different applications (Fakih (2006), Kurt et al. (2011), Shechter et al. (2008), Alagoz et al. (2007), Alterovitz et al. (2008), Nunes et al. (2017) and Sloan (2007)). As mentioned in the above, the main novelty of this paper, logically consistent aggregation of evidence from different measurements in a non-contradictory way and without a need to resort to
empirical estimates still stands. Fakih (2006) develops a general framework for learning efficient approaches to medical diagnosis. It resembles this manuscript and Bayer-Zubek (2004) in the sense that the states are cumulative history of observations, which in turn, guarantees the Markovian nature of the process but it resorts to the empirical estimates of transition probabilities between states
based on observed frequencies. Kurt et al. (2011) develops a discrete time,

but unlike this paper, an infinite-horizon Markov Decision Process to maximise the patient's quality-adjusted life years prior to them having either a stroke or developing a Coronary Heart Disease. The infinite horizon is justified by the large number of visits to the doctor by Type 2 Diabetes patients. The model

- resembles the one of ours in that it has a terminal state which in turn insures convergence. Finally, transition probabilities are computed as a combination of equations based on medical knowledge and empirical observations. Shechter et al. (2008) develops a Markov Decision Process that aims to maximise the expected lifetime or quality-adjusted life years. Similarly to Kurt et al. (2011) this
- ¹⁵⁰ model contains an absorbing state which can be reached from any other state and is an infinite-horizon problem for the same reason as Kurt et al. (2011). Furthermore, exactly as Kurt et al. (2011), the rewards in the problem are not measured only in monetary units as in our case but instead in the units of Health Economics, namely quality-adjusted life years, while the transition
- probabilities of the model are estimated empirically from data. Alagoz et al. (2007) develops a Markov Decision Process for optimal choice of when to go for a liver transplant and then, furthermore, should one accept the part of a liver of a living-donor or the entire cadaveric liver. This is again a discrete-time infinite-horizon problem for the same reasons as in the previous papers. Again,
- the model contains a naturally emerging absorbing state while the cost is measured in Health Economics units. Transition probabilities are estimated from clinical data. Furthermore, the authors derive in closed form the sufficient conditions for the existence of simple threshold-based optimal policies. Alterovitz et al. (2008) develops a Markov Decision Process to optimise the probability of
- the steering needle reaching a target through the noisy environment generated by the uncertainties emerging from the needle- soft tissue interactions and other obstacles that are too small to detect with great but still limited resolution of medical images. The problem does not have costs in the sense of the papers above. They formulate the problem as a discrete time acyclic Markov Decision
- ¹⁷⁰ Process and solve it by value iteration. Finally, the transition probabilities do not emerge from the imprecision of measurements of needle's position, turning

angle and velocity but instead are a consequence of the stochastic environment in which the needle operates. Nunes et al. (2017) develops a Markov Decision Problem for the optimal control of hospital admissions for non-emergency pa-

- tients. Transition probabilities are computed by means of an interplay between a complicated model based on deep understanding of the specific process and empirical considerations. This is a discrete state model considered where the limit of expected average reward is the object to maximise while the costs are expressed in monetary units. Sloan (2007) develops a Markov Decision Process
- to determine the optimal replacement of broken medical equipment with either new equipment or already used (but sterilised and refurbished) equipment. As in this paper, cost and penalties are expressed in monetary terms. This is a discrete time Markov Decision Process. Transition probabilities between states are made on the authors' guesses and serve to demonstrate how the optimal policies change subject to variation of these. This is similar to our paper with the difference that we study the sensitivity of optimal policies to monetary costs instead.
- In summary, we develop a Markov Decision Process for the assessment of hazard of new chemicals while taking into account the exposure, inter-person variability in toxicity and misclassification costs. The result is optimal cost-effective sequential testing strategies for assessment of risk of chemicals. These so-called "optimal policies" rationally integrate outputs from tests with different costs and precisions. The integrated testing frameworks which emerge achieve the keys aims outlined in Pastoor et al. (2014),Embry et al. (2014).

2. Theory

2.1. Individual- and Population-level Models

Much existing literature in quantitative toxicology seems to ignore, or implicitly to average out, the large differences that may exist between the exposure of individuals to potentially toxic chemicals, and the variable effects of a fixed exposure to these chemicals on these individuals. When the purpose of testing is to identify the risk of rare toxicologically harmful outcomes, it seems mathematically expedient to account for this variability explicitly.

We start by providing an alternative view on exposure. Since we cannot know the exact exposure of every single human to the chemical of interest, the natural way to model it is probabilistically. We shall therefore denote exposure experienced by an individual by a random variable X having some distribution with density function f, say, where f would typically correspond to one of many classical distributions (normal, log-normal, gamma). For illustrative purposes, here we will assume f is characterised by a single unknown parameter a^* , say. This could correspond, for example, to the variance of the exposure distribution.

We now turn to modelling the variability in toxicity caused by a given exposure. A typical dose-response curve is sigmoidal: Toxicity is typically small up to some threshold level of exposure, beyond which it increases rapidly with exposure before eventually saturating. This behaviour is efficiently captured by a relationship of the form

$$g(x) = Ax^p/(1+x^p),$$

where A is the level at which toxicity saturates, x is the exposure and p is a ²¹⁵ model parameter to be inferred from data. However, due to individual-level variability, it makes sense to model toxicity at a given level of exposure as a random variable, having mean specified by the function g together with some unknown variance to be inferred from data. Rigorously, we shall denote toxicity at a level of exposure X = x by a random variable T^x . This random variable will have mean

$$\mathbb{E}(T^x) := g(x)$$

Similarly to the case of exposure (above), we denote the probability density function h^x for the random variable T^x as being characterised by a single parameter b^* .

- For the sake of transparency, in what follows assume A and p are known, and that the unknown parameters are a^* (relating to exposure variability) and b^* (relating to individual variability in response to exposure). The purpose of any optimal testing regime will be, broadly, to find the most efficient way to estimate ranges for a^* and b^* so as to determine whether or not a chemical is safe. This
- ²³⁰ purpose will be made more precise below, while the generalisation to wider sets of unknown parameters is left to the Discussion.

2.2. Definition of safety

From the point of view of a pharmaceutical company it is impossible to predict whether a particular individual who buys their product will be adversely affected

- by it. Again, only probabilistic language makes sense: for example a criterion that no more than 1 in 1000000 customers is affected at a level exceeding some defined threshold. Suppose γ represents this threshold of toxicity and let Tbe the random variable indicating the toxicity experienced by an individual sampled randomly from the population. Rigorously, using the logic and notation
- 240 introduced for exposure and toxicity above, we will declare our chemical of interest to be safe for human use if and only if

$$P(T \ge \gamma) \le 10^{-6}.\tag{1}$$

By conditioning on the exposure level and using the generalised law of total probability we obtain:

$$P(T \ge \gamma) = \int_{0}^{\infty} P(T^{x} \ge \gamma) f(x) dx$$
(2)

$$= \int_0^\infty \int_\gamma^\infty h^x(t) \mathrm{d}t f(x) \mathrm{d}x. \tag{3}$$

In other words, the chemical is safe for human use if and only if:

$$\int_0^\infty \int_\gamma^\infty h^x(t) f(x) \mathrm{d}x \mathrm{d}t \le 10^{-6}.$$
 (4)

²⁴⁵ The important mathematical step here it that calculating the probability a chemical is safe is reduced to finding the value of a well defined (but unknown)

double integral i.e. to evaluating a (possibly complicated) function of a^* and b^* alone. Thus (4) can be, in general, written as

$$F(a^*, b^*) \le 10^{-6} \tag{5}$$

for some well defined function F. In other words, all we need to deduce whether our chemical is safe or not is knowledge of the parameters a^* and b^* . Some values of a^* and b^* correspond to safe chemicals, and some to unsafe chemicals. It is the job the any testing strategy to identify the real values of a^* and b^* with sufficient accuracy to determine in which category to place the chemical.

2.3. Precision of Measurements

- Biological practice shows that one hardly ever obtains the same numerical value upon repeating the same test in the laboratory. Rather, the values observed are random variables. Suppose we have several in-vitro and in-silico tests at our disposal for the assessment of our chemical of interest. Some tests will relate to exposure (i.e. tests attempting to estimate a^*) and some will relate to toxicity
- (to estimate b^*). All tests will be inaccurate, and we assume that in-vivo tests are the most accurate, followed by in-vitro and in-silico tests respectively.

For mathematical convenience we assume that test outcomes can only take a finite number of values (i.e. that the data are discrete rather than continuous).

²⁶⁵ This simply allows integrals such as equation 4 to expressed as finite sums, which does not lose generality because the scale of discretisation can be arbitrarily fine.

Specifically, suppose that a series of n independent in-vitro tests for a^* has outcomes (random variables) $a_1^v, ...a_n^v$, and that m independent in-silico tests for a^* have outcomes $a_1^s, ...a_m^s$. Similarly let $b_1^v, ...b_p^v$ and $b_1^s, ...b_q^s$ be the outcomes of pindependent in-vitro, and q independent in-silico, tests for b^* .

Because in-vitro tests are likely to be more accurate than in-silico tests, it is reasonable to assume that the random variables a_i^v have a smaller variance than

- the a_i^s (and similarly for b). Explicitly, we define constants l_i and r_j (for i = 1...nand j = 1...m) and δ_i and ε_j (for i = 1...q and j = 1...q) which capture the accuracy of these tests, such that each $|a_i^v - a^*| \leq l_i$ and each $|a_i^s - a^*| \leq r_i$, and similarly each $|b_j^v - b^*| \leq \delta_j$ and each $|b_j^s - b^*| \leq \varepsilon_j$. Note that the variances of the l_i and δ_j are smaller than those of the r_i and ε_j , respectively.
- 280 2.4. The cost structure

There are several types of costs associated with toxicological testing. First of all, carrying out each assay is costly; although hopefully much cheaper than in vivo studies, even in silico approaches require computing infrastructure and human resources. There are also costs associated to the misclassification of chemicals.

For example, a "false negative" classifying an unsafe chemical as safe (see (32) for the definition of safety) will involve costs from lawsuits and negative publicity. On the other hand, a "false positive" classifying a safe chemical as unsafe would yield costs in terms of wasted R&D and the loss of income which could be generated from selling the product. In what follows we shall be using the following notation:

 $Ca_i^v = \text{cost of in vitro assay } i$ related to estimating a^* $Ca_i^s = \text{cost of in silico assay } i$ related to estimating a^* $Cb_j^v = \text{cost of in vitro assay } j$ related to estimating b^* $Cb_j^s = \text{cost of in vitro assay } \#\text{i related to estimating } b^*$ $MC_{falseneg} = \text{cost of classifying an unsafe chemical as safe}$

 $MC_{falsepos} =$ Cost of classifying a safe chemical as unsafe

Within the framework developed here, we will show how these costs can, and indeed should, become a part of the development of optimal testing strategies.

2.5. Markov Decision Problems

Our ultimate goal is to find optimal policies for sequential decision making under uncertainty. This is likely to involve the decision about which tests are necessary to depend on the outcomes of earlier tests on the same, or a similar,

chemical (Bayer-Zubek (2004)). The optimal testing policy will be the one that minimises the expected total cost, including the costs of misclassification. Due to limited resources it is clear that the decision of whether to declare the chemical safe or unsafe has to be reached after some finite number of tests have been performed. The theory of (finite horizon) Markov Decision Problems provides

the necessary mathematical machinery, by describing the interaction between a decision maker and the environment.

The general set up for Markov Decision Problems is as follows (see Bellman (1957) for a more detailed treatment). At each (discrete) time step, the testing

- process is in some state s, say, and the decision maker chooses one of several available actions, a, say. This action moves the system to some new state s'and gives the decision maker reward R(s, a, s'). In our case in costs which can simply be thought of as negative rewards. The Markov part of the name comes from the demand that the probability of transition from state s to state s' upon choosing action a depends only on the state s, and not on the previous history of the process. This function is denoted by P(s'|s, a). The ultimate goal of the decision maker is to maximise the expected reward (minimise the expected loss)
 - by choosing her actions optimally.
- Formally speaking, a policy π is simply a function that maps states to actions. The value function of a state s under some policy π , denoted $V^{\pi}(s)$ is formally defined by:

$$V^{\pi}(s) = E[\sum R(s, a, s') | s_0 = s].$$
(6)

In other words, the values function is the expected sum of future costs/rewards on the assumption that we start at state s. Moreover, we introduce the special state $s_{terminal}$ which is the absorbing state, meaning that when the process reaches this state it remains there forever. We demand that

$$V^{\pi}(s_{terminal}) = 0 \tag{7}$$

This relation must hold for all policies π simultaneously.

The optimal value function V^* of the Markov Decision Process specified above 335 is:

$$V^*(s) = \min_{\pi} V^{\pi}(s), \text{ for all states } s \tag{8}$$

It can be demonstrated that V^* satisfies the following equation (Bayer-Zubek (2004)):

$$V^*(s) = \min_{all \ actions} \ \sum_{all \ states \ s} P(s'|s;a)[R(s,a,s') + V^*(s')] \tag{9}$$

Again, this relation has to be specified for all states simultaneously. The optimal policy π^* is defined state-by-state according to the following relation:

$$\pi^*(s) = \arg\min_{all \ states \ a} \sum_{all \ states \ s'} P(s'|s,a) [R(s,a,s') + V^*(s')]$$
(10)

2.6. Chemical Risk Classification Problem as a Markov Decision Problem

Recall that, in our general framework, the chemical is safe if and only if:

$$F(a^*, b^*) \le 10^{-6} \tag{11}$$

where a^* and b^* are the parameters we seek to estimate using various tests. Recall also that we have n in vitro and m in silico tests for estimating a^* . Similarly, we have p in vitro and q in silico tests for estimating b^* . The accuracies and costs of these tests are as defined in sections 2.3 and 2.5.

Essentially, our goal is to come up with a sequential testing strategy which tells ³⁵⁰ us whether (12) holds or not while incurring the smallest expected cost. In what follows we shall describe how to write this as a Markov Decision Problem. To this end we follow the reasoning from (Bayer-Zubek (2004)) very closely indeed. We elaborated on the analogy between problem solved there and our problem in the Introduction. The only fundamental difference is that we have to propose

the way of defining the transition probabilities whereas in (Bayer-Zubek (2004)) the authors estimate these are estimated empirically from data.

We have seen in the previous section that a Markov Decision Problem is fully specified by its actions, states and transition probabilities. We first describe the actions of our Markov Decision Problem. These comprise the n + m + p + qavailable tests and two classification actions ("classify as safe" and "classify as unsafe"). States will represent the complete history of previously observed values upon testing, namely the results of performing various in vitro and/or in silico tests. Indeed, let s stand for the current state of the system. Moreover, assume that the next test to be done is G, say, where G can be any of the tests available to us. Then the system will move to the state $s' = s \cup \{G = g\}$ where g is one of the possible values of G. The probability of this transition will therefore depend on all the values previously obtained by measurements which is the information kept in s. In other words:

$$\mathbb{P}(s'|s,G) = \mathbb{P}(G=g|s) \tag{12}$$

Observe that this ensures Markovian nature of the problem. There will be an initial state which will correspond to the observed values obtained by performing two cheapest in silico tests (or alternatively, expert knowledge), one for each of the parameters. This provides us with a (potentially very wide due to imprecision) initial range of values which a^* and b^* can take, at a lowest cost possible.

We note that, because the same test cannot usefully be performed indefinitely, our Markov Decision Problem is acyclic (Bayer-Zubek (2004)), and that this implies important computational consequences (see sections 3.1 and 3.2). Finally, we introduce a special state called the terminal state, denoted by $s_{terminal}$. This

- has the property that every classification action (independently of the state at which it has been performed) forces transition to $s_{terminal}$ with probability one. In other words, this state is absorbing, indicating that an evidence-based classification decision has been reached.
- All that remains is to introduce the costs of our available classification actions. We introduce the quantity $MC(\omega, \zeta)$ to stand for the misclassification cost for classifying a chemical as ω once its true class is ζ ; $\omega, \zeta \in \{safe, unsafe\}$. Since we do not know the true risk class of the chemical at the time of testing (hence the testing), a natural way of thinking of misclassification cost is to treat it as a
- random variable. We will say that the cost of classification action ω performed at some state *s* takes value $MC(\omega, \zeta)$ with probability $P(\zeta|s)$, where this last quantity stands for the chance that the true risk class of the chemical is ζ given that the system is at state *s*. The expected cost of the action is then:

$$R(s,\omega) = \sum_{\zeta} P(\zeta|s) MC(\omega,\zeta)$$
(13)

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Putting all this together, and following Bayer-Zubek (2004), we can deduce that the equation (9) for the optimal value function collapses to

$$V^*(s) = \min_{a \in all \ actions \ available \ at \ state \ s} [R(s,a) + \sum_{s'} P(s'|s,a)V^*(s')]$$
(14)

3. Calculation

The preceding work was general and technical, but showed that the problem of finding optimal testing policies for toxicological testing can be mapped to the existing mathematical framework of Markov Decision Processes. To show the value of this result, and to explore its general consequences, it is expedient to proceed via examples.

- In these examples it is assumed that the outcome of each test, and the values of the unknown parameters a^* and b^* , are all integers. This is not a practical limitation, because these integers may refer to measurements at an arbitrarily fine scale; this assumption simply allows methods from discrete (as opposed to continuous) mathematics to be applied and exact results to be computed. As
- described above, we start by performing the cheapest in silico tests for a^* and b^* (one for each). For sake of simplicity and illustration, we assume the outcomes of these tests are

$$a_1^s = 6, \ b_2^s = 6.$$
 (15)

and that the accuracy of these tests is such that $|a_1^s - a^*| \leq 3$ and $|b_1^s - b^*| \leq 3$. This immediately implies that $a^* \in \{3, 4, 5, 6, 7, 8, 9\}$ and $b^* \in \{3, 4, 5, 6, 7, 8, 9\}$.

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Moreover, assume relation (10) defines the following points in the (a^*, b^*) -plane:

$$K := \{ (a^*, b^*) \in \mathbb{R}^2 : F(a^*, b^*) \le 10^{-6} \} = \{ (3, 6), (4, 5), (4, 6), (4, 7), (5, 4), (5, 5), (5, 6), (5, 7), (5, 8), (6, 3), (6, 4), (6, 5), (6, 6), (6, 7), (6, 8), (6, 9), (7, 4), (7, 5), (7, 6), (7, 7), (7, 8), (8, 4), (8, 6), (8, 7), (9.6) \}$$

25 "safe points" of region K are indicated in the Figure 1 below via red points. Remaining 24 points on the grid represent the "unsafe points".



Figure 1: region K

In this set up, we know from our first tests that $(a^*, b^*) \in \{3, 4, 5, 6, 7, 8, 9\} \times \{3, 4, 5, 6, 7, 8, 9\}$ (where \times denotes Cartesian product). The aim of the testing strategy is to simply to determine, with maximum efficiency, whether or not (a^*, b^*) is an element of K.

3.1. Example 1: The case of independent evidence

Consider a simple situation where, on top of the two initial in silico tests, we have one more in vitro test available for each of our parameters a^* and b^* . These will be denoted by a^v and b^v , respectively. We also have two possible classification actions, namely to classify our chemical as either safe or unsafe. Recall that the goal is to come up with the optimal sequential testing strategy, that is the one that minimises the expected cost. Evidence is independent in

the sense that a^* -related tests tell us nothing on b^* and vice versa (we shall deal with the case of dependent evidence on the next example). In the spirit of Section 2.6 we introduce states, actions and transition probabilities needed for specifying the corresponding Markov Decision Problem. There will be the initial state corresponding to the observation (15). We will denote this state by

435 {}. Moreover, we will assume in vitro tests are more precise than their in silico counterparts, i.e. that they have a lower absolute measurement error. In order to capture this idea we will assume that:

$$|a^{v} - a^{*}| \le 2 \text{ and } |b^{v} - b^{*}| \le 2$$
(16)

Using the above one deduces that

$$a^{v} \in \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11\}, b^{v} \in \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11\}$$
 (17)

This together with the idea the fact that a^* and b^* related tests are independent, 440 gives rise to 11 + 11 + 121 = 143 more states of the system, all listed below:

$$\{a^{v} = 1\} \dots \{a^{v} = 11\}$$
$$\{b^{v} = 1\}, \dots \{b^{v} = 11\}$$
$$\{a^{v} = 1, b^{v} = 1\} \dots \{a^{v} = 1, b^{v} = 11\}$$
$$\vdots$$
$$\{a^{v} = 11, b^{v} = 1\}, \dots \{a^{v} = 11, b^{v} = 11\}$$

Finally, we have the terminal state, $s_{terminal}$, defined exactly as before.

We now move on to describing the available actions and corresponding transition probabilities. At the initial state $\{\}$, there are four actions available: classify as safe, classify as unsafe, do the a^v test or do the b^v test. This means we must specify the following quantities: $P(is \ safe|\{\}, any \ classification \ action), P(a^v =$

 $i|\{\}, do a^{v} test\}, P(b^{v} = i|\{\}, do b^{v} test); where i \in \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11\}.$

We start with the first quantity in the above. We have seen that we must have: (a^*, b^*) $\in \{3, 4, 5, 6, 7, 8, 9\} \times \{3, 4, 5, 6, 7, 8, 9\}$, so that there are 49 values the ordered pair (a^*, b^*) can possibly take, only 25 of which belong to K. Thus, in the absence of any further knowledge, it makes sense to define:

$$P(chemical \ is \ safe|\{\}, \ any \ classification \ action) = 25/49$$
 (18)

We now present an argument for defining the quantities of the form $P(a^v = i|\{\}, do \ a^v \ test), i \in \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11\}$. Notice that, due to symmetry, the quantities $P(b^v = i|\{\}, do \ b^v \ test), i \in \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11\}$, can be defined in an analogous way.

For $j \in \{3, 4, 5, 6, 7, 8, 9\}$ define sets T_j by

$$T_j := \{j - 2, j - 1, j, j + 1, j + 2\}$$
(19)

Notice than that a^v must belong to at least one of the T'_js since the absolute 460 error of the in vitro measurement is 2. With this in mind we define our quantities of interest as follows:

$$P(a^{v} = i | \{\}, \ do \ a^{v} \ test) := \frac{\sum_{j=3}^{9} I\{i \in T_{j}\}}{\sum_{k=1}^{11} \sum_{j=3}^{9} I\{k \in T_{j}\}}$$
(20)

The idea behind (20) is simple. Observe that if we knew the true value of a^* were j, say, (16) would immediately tell us that $a^v \in T_j$ (since the absolute error of this measurement is 2). The problem, clearly, is that we do not know the exact value of a^* , only that a^* is an element of $\{3, 4, 5, 6, 7, 8, 9\}$. First, we count all those intervals T_j that contain i. Secondly, we scale appropriately to get the normalising constant of the probability density. We then define the probability of interest as a function of the two. For example, notice there is only one suchlike interval that contains value 1, two of them that contain value 2, three of them that contain value 3 and so on. Moreover, let us note that this is only one of the possible ways to define the transition probabilities in a manner consistent with precision of measurements and non-contradictory nature of evidence and is by no means the only one. Thus, the random variable $a^v |\{\}$ has the following

distribution:

$$a^{v} \mid_{\{\}, \ do \ a^{v} \ test} = \begin{cases} 1 & \text{with prob } 1/35 \\ 2 & \text{with prob } 2/35 \\ 3 & \text{with prob } 3/35 \\ 4 & \text{with prob } 3/35 \\ 5 & \text{with prob } 1/7 \\ 6 & \text{with prob } 1/7 \\ 7 & \text{with prob } 1/7 \\ 8 & \text{with prob } 1/7 \\ 8 & \text{with prob } 4/35 \\ 9 & \text{with prob } 3/35 \\ 10 & \text{with prob } 2/35 \\ 11 & \text{with prob } 1/35 \end{cases}$$

As promised before, the corresponding counterpart for b^v test is defined by:

$$b^{v} \mid_{\{\}, \ do \ b^{v} \ test} = \begin{cases} 1 & \text{with prob 1/35} \\ 2 & \text{with prob 2/35} \\ 3 & \text{with prob 3/35} \\ 4 & \text{with prob 4/35} \\ 5 & \text{with prob 1/7} \\ 6 & \text{with prob 1/7} \\ 7 & \text{with prob 1/7} \\ 8 & \text{with prob 1/7} \\ 8 & \text{with prob 4/35} \\ 9 & \text{with prob 3/35} \\ 10 & \text{with prob 2/35} \\ 11 & \text{with prob 1/35} \end{cases}$$
(21)

Now, at the states of the form $\{a^v = i\}, i \in \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11\}$ we have only 3 actions available, namely do b^v test or perform one of the two

classification actions available. This, due to the independence of measurements, simply means that one must define

$$P(chemical \ is \ safe | a^v = i) \ for \ i \in \{1, \dots 11\}.$$

Let $A_0 = B_0 = \{3, 4, 5, 6, 7, 8, 9\}$. Thus A_0 and B_0 can be associated with our knowledge of a^* and b^* at state $\{\}$, respectively. However, we now have new information, namely that $a^v = i$. This, by using the principle of noncontradictory nature of evidence, means that $a^* \in A_0 \cap T_i$. Since we have not done any further b^* -related tests at this stage it makes sense to define (although this can be done in a variety of different ways we believe the one below is the simplest) the following quantity:

$$P(chemical \ is \ safe|a^v = i) := \frac{\#(((A_0 \cap T_i) \times B_0) \cap K))}{\#((A_0 \cap T_i) \times B_0)}$$
(22)

⁴⁷⁰ Here #D stands for the number of elements in the set D. Indeed, numerator in (22) counts only those ordered pairs (a^{*}, b^{*}) that are consistent with the new information and correspond to the safe chemical, whereas the denominator corresponds to all ordered pairs that are consistent with the new information. Yet again, using symmetry and independence, actions in the states of the form
⁴⁷⁵ {b^v = i}, i ∈ {1,...11} are fully specified upon defining:

$$P(chemical \ is \ safe|b^v = j) := \frac{\#((A_0 \times (T_j \cap B_0)) \cap K)}{\#(A_0 \times (T_j \cap B_0))}$$
(23)

Finally, it remains to specify the actions and transition probabilities for the states of the form $\{a^v = i, b^v = j\}$ $i, j \in \{1, ..., 11\}$.

Clearly, there are only two actions available in such states, and these are to classify as safe or unsafe. As before, this boils down to computing

 $P(safe|a^v = i, b^v = j)$. Using the same notation and ideas as before we define these as follows:

$$P(safe|a^{v} = i, b^{v} = j) := \frac{\#(((A_{0} \cap T_{i}) \times (B_{0} \cap T_{j})) \cap K)}{\#((A_{0} \cap T_{i}) \times (B_{0} \cap T_{j}))}$$
(24)

We are now ready to compute the corresponding optimal policies. Because our Markov Decision Problem is acyclic, equation (14) for the optimal value func-

tion can be solved in a single sweep through the states. This means we start by computing the optimal value function at the so called "leaf states", that is the states of the form $\{a^v = i, b^v = j\}$ and then work our way upwards to the states of the form $\{a^v = i\}, \{b^v = j\}$ (notice we need the information above to compute these). Finally, we finish with the computation of $V^*(\{\})$ having computed all other necessary quantities along the way. The optimal policy is then read from top down. The actual computations were implemented using MATLAB, code

is available upon request.

By manipulating the attribute and misclassification costs we were able to obtain several classes of fundamentally different optimal policies. This clearly demonstrates that, even in this apparently simple case, the optimal policy emerges from a complex interplay between the geometry of the feasible region K and the costs present in the problem. Moreover, this shows that simple non-adaptive policies are unlikely to be optimal, that is, having changed your population model and

- the cost structure you have no reason whatsoever to believe the optimal policy will not change as well. Indeed, one would have to repeat all the steps in the above and compute the optimal policy corresponding to one's favourite interplay between the ingredients. Finally, let us note that this gives the problem a strongly Bayesian flavour. The policies thus obtained, and some motivation
- for what they emerge, are summarized below. For convenience we introduce the "cost vector" Ω to represent all the costs associated with the problem. Mathematically:

$$\Omega := (MC_{falsepos}, MC_{falseneg}, Ca^{v}, Cb^{v})$$
(25)

First we look at the following situation:

$$\Omega = (200, 100, 80, 50) \tag{26}$$

- The optimal policy is $\pi^*(\{\}) = "classify as unsafe"$. In plain English means that it is optimal to classify the chemical as unsafe without doing any further test at all. The cost of testing, and of misclassification, mean that the chemical's development should be abandoned.
- ⁵¹⁵ Secondly, we look at the following cost structure:

$$\Omega = (100, 200, 80, 50) \tag{27}$$

In this case we obtain $\pi^*(\{\}) = "classify as safe"$. In contrast to Case 1, the change in the balance between the consequences of false negatives versus false positives means that we should immediately classify our chemical as safe.

⁵²⁰ In reality, the cost of testing is likely to be dwarfed by the cost of misclassification, with false negatives being particularly costly. The motivates the following example captured via respective cost structure:

$$\Omega = (10000, 20000, 80, 50) \tag{28}$$

In this case we obtain the following:

 $\begin{aligned} & \pi^*(\{\}) = \text{"do } b^v \text{ test"} \\ & \pi^*(\{b^v \in \{4, 5, 6, 7, 8\}\}) = \text{"do } a^v \text{ test"} \\ & \pi^*(\{b^v \in \{1, 2, 3, 9, 10, 11\}\}) = \text{"classify as unsafe"} \\ & \pi^*(\{b^v = 4, a^v = 6\}) = \text{"classify as safe"} \\ & \pi^*(\{b^v = 4, a^v = i\}) = \text{"classify as unsafe" for all } i \in \{1, 2, 3, 4, 5, 7, 8, 9, 10, 11\} \\ & \text{530} \quad \pi^*(\{b^v = 5, a^v \in \{5, 6, 7\}\}) = \text{"classify as safe"} \\ & \pi^*(\{b^v = 5, a^v \in \{1, 2, 3, 4, 8, 9, 10, 11\}\} = \text{"classify as unsafe"} \\ & \pi^*(\{b^v = 6, a^v \in \{4, 5, 6, 7, 8\}\}) = \text{"classify as safe"} \end{aligned}$

 $\begin{aligned} \pi^*(\{b^v=6, a^v \in \{1, 2, 3, 9, 10, 11\}\} &= \text{``classify as unsafe''} \\ \pi^*(\{b^v=7, a^v \in \{5, 6, 7\}\}) &= \text{``classify as safe''} \\ \text{535} \quad \pi^*(\{b^v=7, a^v \in \{1, 2, 3, 4, 8, 9, 10, 11\}\} &= \text{``classify as unsafe''} \\ \pi^*(\{b^v=8, a^v=6\}) &= \text{``classify as safe''} \\ \pi^*(\{b^v=8, a^v=i\}) &= \text{``classify as unsafe'' for all } i \in \{1, 2, 3, 4, 5, 7, 8, 9, 10, 11\} \end{aligned}$

The above is a mathematical description of the optimal sequential testing policy. In simple English, the first four lines in the above scheme can be interpreted as follows: Start by doing the b^v test. If this test gives you values 1,2,3,9,10 or 11 then classify your chemical as unsafe, otherwise do the a^v test. If the b^v test gave value 4, then if the a^v test gives value 6 then classify your chemical as safe, but otherwise classify it as unsafe. The remaining statements in the statements in the statement in t

The details of this example are not as important as the general picture which emerges: even in this very simple scenario, the optimal policy is a set of adaptive tests where the outcome of the earlier tests dictates what subsequent tests (if

any) are needed. For an adaptive testing policy to emerge from such a limited range of possibilities, and for it to have such clear dependence on relative costs, is evidence that adaptive testing policies may be expected to be optimal in more realistic scenarios.

3.2. Example # 2: The case of dependent evidence

As in the previous example, we first carry out the two cheapest (and probably therefore inaccurate) in silico tests, one for each parameter. Now suppose we have no more b^* - related tests available, but have two more a^* - related in vitro tests at our disposal, denoted a^{v1} and a^{v2} , respectively. We will assume these two tests have different precision, and that both are more precise than their in silico counterpart. Mathematically, we capture this by imposing that:

$$|a^{v1} - a^*| \le 2 \text{ and } |a^{v2} - a^*| \le 1$$
(29)

Using the ideas from Section 3.1 we deduce that the corresponding Markov Decision Problem will have a total of 79 states. Detailed description of these and the corresponding transition probabilities can be found in the Appendix.

As in the previous example, it is possible to manipulate attribute and misclassification costs to obtain a range of optimal policies. Examples of these are summarized below via corresponding Ω vectors. We begin by looking at the cost structure given by:

$$\Omega = (200, 100, 80, 50) \tag{30}$$

570 In this case the optimal policy is to classify the chemical as unsafe without doing any test.

Secondly, we consider the following costs:

$$\Omega = (100, 200, 80, 50) \tag{31}$$

Then the optimal policy is to immediately classify the chemical as safe without 575 doing any test.

Finally we look at the case capture via the following more realistic cost structure:

$$\Omega = (20000, 10000, 80, 50) \tag{32}$$

The optimal policy turns out to be a complex mixture, interested reader can find the details in the Appendix.

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We finish this section by illustrating another useful feature of our model. Observe that if $a^{v1} = 1$ then (24) tells us that a^* is an element of $\{-1, 0, 1, 2, 3\}$. However, we actually know for sure that a^* is an element of $\{3, 4, 5, 6, 7, 8, 9\}$. Putting these two pieces of information together we deduce that $a^* = 3$, i.e. we have thus identified a^* with absolute certainty. Therefore, common sense tells us that it should not be optimal to do any further a^* -related tests after that since we would be simply spending money for learning nothing we already know. Indeed, the above tells us that $\pi^*(\{a^{vt_1} = 1\}) =$ "classify as unsafe". This (and other similar situations) are not coincidental. Interested reader can find the formal proof of this fact in the Appendix.

Discussion and Conclusion

This paper provides a mathematically rigorous framework for the computation of optimal sequential testing strategies for chemical risk classification. Its novelty lies in the fact that it does so by taking into account various important ⁵⁹⁵ and realistic ingredients. These include population model, errors in laboratory measurements and the corresponding cost structure. Mathematically, the heart of the argument comes from Machine Learning, more specifically the rich theory of Markov Decision Problems. Optimal Sequential Policies were computed in two simple but highly illustrative and easily generalisable examples, namely tests that are independent and those whose outcomes are deeply correlated by the common truth they are revealing with possibly different precision. Indeed, in case of more parameters the corresponding Markov Decision Problem shall

The key take-home message for toxicological testing is that it is possible, and indeed necessary, to search for optimal testing strategies. Markov Decision Problems provide a mathematical framework with which to achieve this. We show that the strategies which emerge must explicitly take into account the costs both of testing and of misclassification. Moreover, the optimal testing

simply have more states, all other aspects will be identical.

⁶¹⁰ policies are typically adaptive, where the outcome of any given test influences the decision as to which test (or classification) to apply next. This may be interpreted as a mathematical avocation of the recent interest in adaptive testing strategies from both industry and regulatory bodies.

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